

# **Integrated Science Assessment for Carbon Monoxide – First External Review Draft**

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# Acronyms and Abbreviations

$\alpha$	alpha, ambient exposure factor (varies between 0 and 1)
a	air exchange rate of the microenvironment
AA	abdominal aorta(s)
ABR	auditory brainstem response
ACS	American Cancer Society
ACS-CPS-II	ACS Cancer Prevention Study II
ADP	adenosine diphosphate
AEFV	area under the expiratory flow-volume curve
AGL	above ground level
Akt	Akt cell signaling pathway - protein family (of protein kinases B [PKB])
AM, a.m.	morning hours
AMI	acute myocardial infarction
AMP	adenosine monophosphate
ANOVA	analysis of variance
APO E	apolipoprotein E
APO E <sup>-/-</sup>	mouse strain used as a model of atherosclerosis
ARI	acute respiratory infection
AP	action potential
APD	action potential duration
APEX	Air Pollution Exposure (population-based model)
APHEA	Air Pollution and Health: A European Approach (multi-cities analyses)
APHEA2	extended analysis of APHEA
APTT	activated partial thromboplastin time, (blood coagulability endpoint)
AQ	air quality
AQCD	Air Quality Criteria Document
AQS	(EPA) Air Quality System (database)
AR	gastronomy reared with artificial feeding system (milk substitutes using gastronomy-feeding)
ARCO	gastronomy reared + CO exposure
ARIC	Atherosclerosis Risk in Communities (study)
ARID	gastronomy reared with iron deficient diet
ARIDCO	gastronomy reared with iron deficient diet + CO exposure
ATP	adenosine triphosphate
ATS	American Thoracic Society
avg	average
AVP	aortic valve prosthesis (surgically implantable biological prosthesis)

$\beta$	beta (beta coefficient, slope)
B lymphocytes	bursa-dependent lymphocytes
BALF	bronchoalveolar lavage fluid
BC	black carbon
BEAS-2B	human bronchial epithelial cell line
BEIS	Biogenic Emissions Inventory System
BELD	Biogenic Emissions Landcover Database
BHR	bronchial hyper-responsiveness
BK <sub>Ca</sub>	voltage and Ca <sup>2+</sup> -activated K <sup>+</sup> channel(s)
BP	blood pressure
BQ-123	a selective endothelin A (ET <sub>A</sub> ) receptor antagonist
BS	black smoke
BSP	black smoke particles
C	carbon
C <sub>a</sub>	ambient concentration
CA	cardiac arrhythmia
Ca <sup>2+</sup>	calcium ion
CAA	Clean Air Act
CAD	coronary artery disease
CALINE	California Line Source Dispersion Model
CALPUFF	Gaussian puff modeling system to simulate air quality dispersion and accesses long range transport of pollutants. Distributed by TRC Solutions.
CAMP	Childhood Asthma Management Program (study)
cAMP	cyclic AMP
CAP(s)	compound action potential(s)
CASAC	Clean Air Scientific Advisory Committee
CASN	Cooperative Air Sampling Network
CAth	cardiac atherosclerosis
CBSA	Core-Based Statistical Area (containing at least one urban area of 10,000 people [as determined by the 2000 census]; replaces the older MSA definition from the 1990 census).
CCGG	Carbon Cycle Greenhouse Gases Group (within ESRL)
CD	cardiac dysrhythmias
Cd	cadmium
CD-1	mouse strain
CDC	Centers for Disease Control and Prevention
CdCl <sub>2</sub>	cadmium chloride
CFK	Coburn-Forster-Kane (equation or modeling)
CFR	Code of Federal Regulations
cGMP	cyclic GMP
CH <sub>2</sub> O	formaldehyde
CH <sub>2</sub> O <sub>2</sub>	formic acid

CH <sub>3</sub>	methyl groups
CH <sub>3</sub> CHO	acetaldehyde
CH <sub>3</sub> CO	acetyl radical(s)
CH <sub>3</sub> CO <sub>3</sub> NO <sub>2</sub>	PAN, peroxyacetyl nitrate
CH <sub>3</sub> O <sub>2</sub>	methyl peroxy radical
CH <sub>3</sub> OOH	methyl hydroperoxide
CH <sub>4</sub>	methane
ChAT	choline acetyl-transferase
CHD	coronary heart disease
CHF	congestive heart failure
CI	confidence interval(s)
CIS	cerebral ischemic stroke
C <sub>j</sub>	airborne concentration at location j
CL/P	cleft lip with or without palate
CNS	central nervous system
CO	carbon monoxide
CO <sub>2</sub>	carbon dioxide
COD	coefficient of divergence
CoH, COH	coefficient of haze (a measurement of visibility interference in the atmosphere, as the quantity of dust and smoke in a theoretical 1,000 linear feet of air).
COHb	carboxyhemoglobin (% concentration measured in (mL CO/mL blood))
[COHb] <sub>0</sub>	carboxyhemoglobin concentration at time zero (0)
[COHb] <sub>t</sub>	carboxyhemoglobin concentration at time t
COMb	carboxymyoglobin
Complex IV	mitochondrial cytochrome <i>c</i> oxidase complex assembly, respiratory chain complex IV in the mitochondrial inner membrane.
Cong	U.S. Congress
CONUS	contiguous U.S. (United States of America)
COPD	chronic obstructive pulmonary disease
CPS II	Cancer Prevention Study II
C-R	concentration-response (relationship)
CRC	Coordinating Research Council
CrMP	collapsin response mediator protein
CRP	C-reactive protein
CSA	Combined Statistical Area (an aggregate of adjacent CBSAs tied by specific commuting behaviors), as defined by the 2000 census
Cu	copper (element)
CVD	cardiovascular disease
Δ	delta, change, difference
d	straight-line distance between monitor pairs
Dahl/Rapp	salt-sensitive model of hypertension (substrain of Sprague Dawley rat)
D.C.	District of Columbia



df	degrees of freedom
df/yr	degrees of freedom per year
$D_L$	lung diffusing capacity
$D_{LCO}$	lung diffusing capacity of CO ( in mL/min/mmHg)
$D_mCO$	capacity for diffusion of CO into the muscle
DMT-1	divalent metal transporter-1 protein (transport and detoxification of metals)
DMV	dorsal motor nucleus of the vagus nerve
DNA	deoxyribonucleic acid
DOCA	Deoxycorticosterone acetate
$dP/dt_{LV}$	left ventricular maximal and minimal first derived pressure ( $+dP/dt_{LV}$ , $-dP/dt_{LV}$ )
$dP/dt_{RV}$	right ventricular maximal and minimal first derived pressure ( $+dP/dt_{RV}$ , $-dP/dt_{RV}$ )
DSA	deletion/substitution/addition
dt	time spent in each location
$\Sigma$	sigma, sum of a series
E	exposure over some duration
$E_a$	exposure to pollutant of ambient origin
E/A	mitral ratio of peak early to late diastolic filling velocity (E/A).
EC	elemental carbon
ED	emergency department
EKG, ECG	electrocardiogram; electrical activity of the heart over time, measured by an electrocardiograph
$E_{na}$	exposure to pollutant of non-ambient origin
eNOS	endothelial nitric oxide synthase
EPA	U.S. Environmental Protection Agency
EPO	erythropoietin (stimulates production of new red blood cells)
EPR	Electron Paramagnetic Resonance (spectroscopy), aka Electron Spin Resonance (spectroscopy)
EPRI	Electric Power Research Institute
ERK1/2	ERK-1 (MAPK p42) and ERK-2 (MAPK p44) (extracellular signal-regulated kinases [in cell signaling pathway])
ESRL	Earth System Research Laboratory (within GMD of NOAA)
ET-1	endothelin-1; vasoconstrictor that mediates regulation of vascular tone
$ET_A$	endothelin A ( $ET_A$ ) receptor subtype
ETS	environmental tobacco smoke
EXPOLIS	six-city European air pollution study
Factor VII	FVII, a vitamin K-dependent serine protease glycoprotein (also known as stable factor or proconvertin), initiates the process of coagulation.
Factor VIIIIC	blood-coagulation factor VIII. (antihemophilic factor, part of factor VIII/von Willebrand factor complex; acts in the intrinsic pathway of blood coagulation
FAS	apoptosis stimulating fragment
FC	interference filter
Fe	iron

Fe <sup>2+</sup>	ferrous iron
Fe <sup>3+</sup>	ferric iron
FEF	forced expiratory flow (Liters/second)
FEF <sub>25-75</sub>	determined from the time in seconds at which 25% and 75% of the vital capacity is reached. The volume of air exhaled between these two times is the FEF <sub>25-75</sub> . This value reflects the status of medium and small sized airways.
FEM	Federal equivalent method
FEV <sub>1</sub>	forced expiratory volume in 1 second (volume of air exhaled into the spirometer mouthpiece in one second, using maximum effort)
f <sub>i</sub>	fraction of time spent indoors
F <sub>1</sub> CO	fractional concentration of CO in ambient air (in ppm)
F <sub>inf</sub>	infiltration factor
f <sub>o</sub>	fraction of time spent outdoors
FR	Federal Register
FGR	fetal growth restriction(s)
FRM	Federal reference method
FSH	follicle stimulating hormone
FVC	forced vital capacity (volume of air that is expelled into the spirometer following a maximum inhalation effort)
FVII	Factor VII, a vitamin K-dependent serine protease glycoprotein (also known as stable factor or proconvertin), initiates the process of coagulation
FW	fresh weight
G0/G1 phase	Phase of Cell Cycle in which cell commits to mitosis (division)
g, mg, µg	gram(s), milligram(s), microgram(s)
g/mole	unit of molar mass (grams/mole)
GAM	generalized additive model(s)
GD	gestational day
GEE	generalized estimating equations
GEM	gas extraction monitor
GFAP	glial fibrillary acidic protein
GFC	gas filter correlation
GLM	generalized linear models
GLMM	generalized linear mixed models
GMD	Global Monitoring Division (of NOAA)
GMP	guanosine monophosphate
GSH	glutathione
GSSG	oxidized glutathione
GTP	guanosine triphosphate
GWP(s)	global warming potential(s)
H	atomic hydrogen, hydrogen radical, height
h	hour
H <sub>2</sub>	molecular hydrogen

H <sub>2</sub> O	water
H <sub>2</sub> O <sub>2</sub>	hydrogen peroxide
H9c2	rat embryonic cardiomyocytes (cell line)
Hb	hemoglobin
HC(s)	hydrocarbon(s)
HCFC(s)	hydrochlorofluorocarbon(s)
HCO	formyl radical
HEAPSS	Health Effects of Air Pollution among Susceptible Subpopulations (study)
hectoPascal	atmospheric level measured by atmospheric pressure (e.g., 700 hectoPascals ≈10,000 feet above sea level)
HEK293	human embryonic kidney cells (experimentally transformed cell line)
Hem	hemorrhagic (stroke)
Hep3B	Human hepatocarcinoma cell line
HF	heart failure
HF	high energy frequency (HRV parameter)
HFLFR	high energy frequency versus low energy frequency ratio (HRV parameter)
HH	hypobaric hypoxia (measured in torr)
HIF-1 $\alpha$	oxygen-responsive component of the hypoxia-inducible factor (HIF) 1 complex
his-Fe <sup>2+</sup> -his	iron binding scheme in neuroglobin (Nb); hexacoordinated deoxy ferrous (Fe <sup>2+</sup> ) form of Nb, bound to histidine
HO	heme oxygenase (microsomal enzyme that degrades heme protein to form endogenous CO and iron (Fe <sup>2+</sup> ))
HO <sub>2</sub>	hydroperoxy radical, hydrogen dioxide
HO-1	inducible isoform of heme oxygenase (induced by stressors)
HO-2	constitutively expressed isoform of heme-oxygenase
HO/CO	heme oxygenase/carbon monoxide system (signaling pathway)
HR	heart rate (beats per minute), hazard ratio
H/R	hypoxia followed by reoxygenation
HRV	heart rate variability (beat-to-beat alterations in the heart determined by analyses of time and frequency domains in ECG(s)).
HUVEC(s)	human umbilical vein endothelial cell(s)
h $\nu$	photon
H/W	height to width ratio
I	electrical current
IARC	International Agency for Research on Cancer
IC	inferior colliculus (an auditory integrative section of the midbrain)
I <sub>ca,L</sub>	electrical current (I) through the L-type Ca <sup>2+</sup> channel
ICAM-1	intercellular adhesion molecule
ICD	implantable cardioverter defibrillator(s)
ICR	substrain of CD1 mouse, Institute for Cancer Research
IDW	inverse-distance-weighted
IHD	ischemic heart disease

IL-6	interleukin-6
IL-8	interleukin-8
INDAIR	Indoor Air Model
I/O	indoor to outdoor concentration ratio
IOM	Institute of Medicine
i.p.	intraperitoneal injection
IQR	interquartile range
IR	immunoreactivity
IS	ischemic stroke
ISA	Integrated Science Assessment
Isch	ischemic (stroke)
ITA	internal thoracic artery of the heart (coronary artery bypass surgery graft)
$I_{to}$	transient outward current
i.v.	intravenous (injection)
IUGR	intrauterine growth restriction
$K^+$	potassium ion
k	dissociation rate, rate of CO loss from the microenvironment
$k_{CO}$	dissociation rate of carbon monoxide from hemoglobin
kHz	kilohertz (1,000 cycles per second; frequency of an electrical signal)
$K_m$	Michaelis Constant; the substrate concentration which gives $\frac{1}{2} V_{max}$ (in the Michaelis-Menten equation of enzyme kinetics)
km	kilometer(s)
$K_m/V_{max}$	slope of Michaelis-Menten equation; measures the efficiency of an enzyme
$k_{O_2}$	Dissociation rate of oxygen from hemoglobin
L, dL, mL, $\mu$ L	Liter, deciLiter, milliLiter, microLiter
lag 0	same day as the hospital, ED, clinic, physician visit
lag 0-3	the three previous days before the hospital, ED, clinic, physician visit
LBW	low birth weight (<2,500 grams, ( $\approx$ 5lbs, 8 oz))
LCA+	leucocyte common antigen cells
LD	lactational day
LDH	lactate dehydrogenase (release of LDH is an indicator of cell membrane damage)
LDL	low-density lipoprotein
leeward	downwind
LF	low energy frequency (HRV parameter)
LF/HF	low energy frequency to high energy frequency ratio (HRV parameter)
LH	lutenizing hormone
LOAEL	lowest observed adverse effect level
LOD	limit of detection
LOESS	locally weighted scatterplot smoothing (modeling method); aka LOWESS
LPS	lipopolysaccharide
LTP	long-term potentiation (hippocampal)

LUR	land use regression (model[s])
LV	left ventricle (of heart)
LV+S	left ventricular plus septum
LVDP	left ventricular developed pressure
LVESP	left ventricular end diastolic pressure
LVSF	left ventricular shortening fraction
LVW	left ventricular work
m, cm, $\mu$ m, nm	meter(s), centimeter, micrometer(s) = micron(s), nanometer(s)
M, mM, $\mu$ M, nM, pM	Molar, milliMolar, microMolar, nanoMolar, picoMolar
M	Haldane coefficient representing the CO chemical affinity for Hb [or Mb]), Reaction mediator.
ma	moving average
MAPK	mitogen-activated protein kinase(s), MAP kinase(s)
MAO-A	monoamine oxidase A (mitochondrial enzyme; deaminates amine-neurotransmitters)
max	maximum
Mb	myoglobin
MC	ultrafine particle mass concentration
METs	metabolic equivalent unit(s)
$\text{mg}/\text{m}^3$	unit of chemical concentration in air (milligrams per cubic meter)
MHC	major histocompatibility complex
MI	myocardial infarction, “heart attack”
Miller PEE	Miller equation prediction equation estimates
min	minute(s)
MIP-2	macrophage inflammatory protein-2
mitral E to A ratio	mitral ratio of peak early to late diastolic filling velocity (E/A).
MMEF	maximal midexpiratory flow
mmHg	millimeters of mercury (unit of pressure)
MMP	matrix metalloproteinase
MOA(s)	mode(s) of Action
MOBILE6	(EPA) Mobile source emission factor model
MODIS	Moderate Resolution Imaging Spectroradiometer
mol, nmol	mole, nanomole
MONICA	Monitoring of Trends and Determinants in Cardiovascular Disease (study)
MOPITT	Measurement of Pollution in the Troposphere (instrument)
MPO	myeloperoxidase (indicative of neutrophil degranulation)
MPT	mitochondrial permeability transition
MR	maternally reared; rodent pups reared with their respective dams
mRNA	messenger RNA
ms	millisecond(s)
MSA	Metropolitan Statistical Area
MSNA	muscle sympathetic nerve activity

MT	million tons
MV <sub>O2</sub>	myocardial oxygen consumption
n, N	sample size
N	number of monitor days
N <sub>2</sub>	nitrogen gas
NAAQS	National Ambient Air Quality Standards
NADPH	nicotinamide adenine dinucleotide phosphate
NADH-TR	nicotinamide adenine dinucleotide - tetrazolium reductase; histochemical stain of muscle tissue - the space between the myofibrils.
NAPAP	National Acid Precipitation Assessment Program
NARSTO	(formerly) North American Research Strategy for Tropospheric Ozone
NAS	National Academy of Sciences
NASA	National Aeronautics and Space Administration
Nb	neuroglobin
NC	ultrafine particle number concentration
NDIR	nondispersive infrared (detection)
NE	norepinephrine
NEI	(EPA) National Emissions Inventory
NF-κB	nuclear factor kappa B
Ni	nickel (element)
NIHL	noise-induced hearing loss
NMDA	N-methyl-D-aspartate
NMHC(s)	nonmethane hydrocarbon(s)
NMMAPS	National Morbidity, Mortality, and Air Pollution Study
nmol	nanomole
NN	normal-to-normal (NN or RR) time interval between each QRS complex in the EKG
nNOS	neuronal nitric oxide synthase (NOS)
NO	nitric oxide
NO•	nitric oxide free radicals, free radical species of nitric oxide
NO <sub>2</sub>	nitrogen dioxide
NOAA	National Oceanic and Atmospheric Administration
NOAEL	no observed adverse effect level
NO•-Hb	nitrosyl bound Hb
NO•-Mb	nitrosyl bound Mb
NO <sub>x</sub>	nitrogen oxides, oxides of nitrogen
NRC	National Research Council
NTS	nucleus of the solitary tract (in brainstem)
O <sub>2</sub>	oxygen
O <sub>3</sub>	ozone
O <sub>2</sub> Hb	oxyhemoglobin (% concentration in mL O <sub>2</sub> / mL blood)
O <sub>2</sub> Mb	oxymyoglobin

OAE	otoacoustic emissions (testing)
OAQPS	Office of Air Quality Planning and Standards
OC	organic carbon
OH, OH OH•	hydroxyl group, hydroxyl radical(s)
OR	odds ratio(s)
OS	occlusive stroke
OSPM	Operational Street Pollution Model
P	penetration (of pollutant)
P, p	probability
P	number of paired hourly observations
P <sub>1</sub> CO	CO partial pressure in inhaled air (in mmHg)
P13K	phosphoinositide 3-kinase
p21 <sup>Waf1/Cip1</sup>	inhibitor of CDK(s) (cyclin-dependent kinases) - regulates cell cycle arrest following DNA damage.
p38	p38 MAP kinase
P90	90th percentile of the absolute difference in concentrations
P450	cytochrome P450
P <sub>A</sub>	alveolar pressure
PA	pulmonary artery (myocytes)
PACF	partial auto-correlation functions
P <sub>A</sub> CO	alveolar pressure for carbon monoxide
PAF	platelet activating factor
PAH	polycyclic aromatic hydrocarbon
PAHT	pulmonary artery hypertension
PAN	peroxyacetyl nitrate (CH <sub>3</sub> CO <sub>3</sub> NO <sub>2</sub> )
P <sub>A</sub> O <sub>2</sub>	alveolar pressure for oxygen
P <sub>a</sub> O <sub>2</sub>	arterial oxygen pressure
PARP	poly(ADP-ribose) polymerase; (involved in DNA repair; cleaved by caspases during early apoptosis)
Pb	lead
PB	barometric pressure (in mmHg)
PBN	the free radical inhibitors PBN (a spin trap) or N-tert-butyl-alpha-phenylnitron (free radical inhibitor) an organic spin trap agent designed specifically to form "stable" adducts with free radicals in electron spin resonance studies.
$P_{\bar{C}}$	average partial pressure in lung capillaries
pCO	partial pressure of CO in lung capillaries (in mmHg)
P <sub>CO</sub>	partial pressure of CO
$P_{\bar{C}}O_2$	average partial pressure of O <sub>2</sub> in lung capillaries (mmHg)
PDGF	platelet derived growth factor
PEE	prediction equation estimates
PEF	peak expiratory flow (L/min)
PEFD(s)	Personal Exposures Frequency Distributions

PEM(s)	personal exposure monitor(s)
%	percent
PHD	pulmonary heart disease
$P_1$	partial pressure of inhaled air
Pi	inorganic phosphate
PIH	primary intracerebral hemorrhage
PKB	protein kinases B
PM	particulate matter
p.m.	afternoon/evening hours
PM <sub>2.5</sub>	fine particulate matter, particles with a nominal mean aerodynamic diameter less than or equal to 2.5 $\mu\text{m}$
PM <sub>10</sub>	particles with a mean aerodynamic diameter less than or equal to 10 $\mu\text{m}$
PM <sub>10-2.5</sub>	coarse particulate matter, coarse fraction of PM <sub>10</sub> (referred to as thoracic coarse particles or coarse-fraction particles; generally including particles with a nominal mean aerodynamic diameter greater than 2.5 $\mu\text{m}$ and less than or equal to 10 $\mu\text{m}$ ).
PMN	polymorphonuclear leukocytes
pmol	picomole
PNC	particle number concentration / count
PND	post natal day
pNEM/CO	probabilistic NAAQS Exposure Model for CO (preceded APEX)
PNN	proportion of interval differences of successive normal-beat intervals in EKG
PNN <sub>50</sub>	proportion of interval differences of successive normal-beat intervals greater than 50 ms in EKG
PNS	peripheral nervous system
pO <sub>2</sub>	partial pressure of oxygen in lung capillaries (in mmHg)
ppb	parts per billion
ppm	parts per million
PRB	policy-relevant background
PT	prothrombin time (blood coagulability endpoint)
PTB	preterm birth (birth after the 20th week, but before the 38th week of human pregnancy)
PVCD	peripheral vascular and cerebrovascular disease
PvO <sub>2</sub>	venous oxygen tension
$\dot{P}V\text{O}_2$	peak oxygen consumption
$\dot{Q}$	cardiac output
QCP	Quantitative Circulatory Physiology (model)
$\dot{Q}_M$	blood flow to muscle
$\dot{Q}_{OT}$	blood flow to other tissues
r	correlation coefficient
R <sup>2</sup>	coefficient of determination
R State	R state of hemoglobin; structural shape of protein when binding oxygen (oxy state)
RA	radial artery of the heart
RAW 264.7	mouse macrophage cell line



RBC	red blood cell
redox status	ratio of interconvertible reduced/oxidized forms of a molecule
rho(0)	rho(0) cells (cells lacking mitochondrial DNA)
rMSSD	mean squared differences of successive difference normal-beat to normal-beat (NN or RR) time intervals between each QRS complex in the EKG
RNA	ribonucleic acid
ROE	Report on the Environment
ROFA	residual oil fly ash (particles)
ROS	reactive oxygen species
RR	normal-to-normal (NN or RR) time interval between each QRS complex in the EKG
RR	risk ratio(s)
RUPERT	Reducing Urban Pollution Exposure from Road Transport (study)
RV	right ventricle (of heart)
RVEDP	right ventricular end diastolic pressure
RVESP	right ventricular end-systolic pressure
RVSF	right ventricular shortening fraction
RVW	right ventricular work
S	second(s)
SA	sphinganine
SAA	serum amyloid A
SAB	(EPA) Science Advisory Board
SBP	systolic blood pressure
SDNN	standard deviation normal-to-normal (NN or RR) time interval between each QRS complex in the EKG
Se	selenium
SEM	standard error of mean
sEng	soluble endoglin
SES	socioeconomic status
Sess	Session (of Congress)
SF <sub>6</sub>	sulfur hexafluoride (tracer gas)
sFlt	soluble Fms-like tyrosine kinase-1
SGA	small for gestational age
sGC	soluble guanylate cyclase
SHEDS	(EPA) Stochastic Human Exposure and Dose Simulation (model)
SHR	Spontaneously hypertensive rat strain
SIDS	sudden infant death syndrome
SIPs	State Implementation Plan(s)
siRNA	small inhibitory RNA (silencing RNA)
SLAMS	State and Local Air Monitoring Stations
SMC	smooth muscle cell(s)
Sn	tin

snMP	tin-(IV)-mesoporphyrin (HO inhibitor)
SNP	single-nucleotide polymorphism
SnPP-IX	tin protoporphyrin IX, (HO inhibitor)
SO	sphingosine
SO <sub>2</sub>	sulfur dioxide
SO <sub>4</sub> <sup>2-</sup>	sulfate
SOD	superoxide dismutase
SOD1	cytoplasmic superoxide dismutase
SOD2	mitochondrial superoxide dismutase
SOPHIA	Study of Particles and Health in Atlanta
ST-segment	segment of EKG between QRS complex, and T wave. ST-segment elevation may indicate myocardial infarction
STAT1 / STAT3	signal transducers and activators of transcription (transcription factors involved in cell signaling)
STEMS	Space-Time Exposure Modeling System
STN	(EPA) Speciation Trends Network
STPD	standard temperature and pressure, dry
SV	stroke volume
SVEB	supraventricular (atrium or atrioventricular node) ectopic beats
τ	photochemical lifetime
T	trial
T lymphocytes	thymus-dependent lymphocytes
T State	T state of hemoglobin; structural shape of protein when not binding oxygen (deoxy state)
TBARS	thiobarbituric acid reactive substances
TC	total carbon
TFAM	mitochondrial transcription factor A
Tg	teragram(s)
TH	tyrosine hydroxylase
TH <sub>b</sub>	total blood concentration of hemoglobin (in g Hb/mL blood)
THP-1	human monocyte-derived cell line, (can differentiate into macrophages)
TIA	transient ischemic attack (mini-stroke)
TNF-α	tissue necrosis factor alpha (WBC protein boosts immune system, too much can cause inflammation)
torr	unit of pressure, equal to 0.001316 (1/760) atmosphere
TPM	total particulate matter
TSP	total suspended particles
UFP	ultrafine particle(s)
U.K.	United Kingdom
ULTRA	(Exposure and Risk Assessment for Fine and )Ultrafine Particles in Ambient Air (Study)
URI	upper respiratory infection
U.S.	United States of America
USC	U.S. Code

UV	ultraviolet
V	vanadium
$\dot{V}$	minute ventilation
$\dot{V}_A$	alveolar ventilation (in mL/min at STPD)
V <sub>A</sub>	alveolar volume, a mesurement of lung size
V <sub>b</sub>	blood volume ( in mL)
$\dot{V}_{CO}$	CO uptake rate (the product of D <sub>L</sub> CO and the mean P <sub>A</sub> CO).
$\dot{V}_{CO}$	endogenous CO production rate (in mL/min at STPD)
V <sub>D</sub>	Dead space volume
V <sub>E</sub>	ventilation rate
VEGF	vascular endothelial growth factor
VLf	very low energy frequency (HRV parameter)
V <sub>max</sub>	maximum velocity (catalyzed by a fixed enzyme concentration, in the Michaelis-Menten equation of enzyme kinetics)
VO <sub>2</sub> max	maximum volume per time, of oxygen (maximal oxygen consumption, maximal oxygen uptake or aerobic capacity)
VOC(s)	volatile organic compound(s)
vol	volume
VPB	ventricular premature beat
V/Q	ventilation-perfusion ratio
vWF	von Willebrand factor (part of factor VIII/von Willebrand factor complex; acts in the intrinsic pathway of blood coagulation)
W	width
WBC	white blood cell
WHI	Women's Health Initiative (Study)
windward	upwind
WKY	Wistar-Kyoto rat strain
ww	wet weight
X <sub>ij</sub> , X <sub>ik</sub>	observed hourly concentrations for time period <i>i</i> at sites <i>j</i> and <i>k</i>
yr	year
Z/H	elevation of the measurement (Z) scaled by height (H)
ZIP (code)	Zone Improvement Plan (system of postal codes used in the U.S.)
Zn	zinc
ZnPP IX	Zn protoporphyrin IX, HO inhibitor

# Chapter 1. Introduction

1           The Integrated Science Assessment (ISA) is a concise evaluation and synthesis of the most policy-  
2 relevant science for reviewing the national ambient air quality standards (NAAQS). Because the ISA  
3 communicates critical science judgments relevant to the NAAQS review, it forms the scientific  
4 foundation for the review of the NAAQS for carbon monoxide (CO). The existing primary CO standards  
5 include a 1-hour (h) average (avg) standard set at 35 parts per million (ppm), and an 8-h avg standard set  
6 at 9 ppm, neither to be exceeded more than once per year. There is currently no secondary standard for  
7 CO.

8           The ISA accurately reflects “the latest scientific knowledge useful in indicating the kind and extent  
9 of identifiable effects on public health which may be expected from the presence of [a] pollutant in  
10 ambient air” (U.S. Code, 1970). Key information and judgments formerly contained in the Air Quality  
11 Criteria Document (AQCD) for CO are incorporated in this assessment. Additional details of the pertinent  
12 scientific literature published since the last review, as well as selected older studies of particular interest,  
13 are included in a series of annexes. This first external draft ISA thus serves to update and revise the  
14 information available at the time of the previous AQCD for CO in 2000.

15           The Integrated Plan for Review of the NAAQS for CO (U.S. EPA, 2008b) identifies key policy-  
16 relevant questions that provide a framework for this review of the scientific evidence. These questions  
17 frame the entire review of the NAAQS and thus are informed by both science and policy considerations.  
18 The ISA organizes and presents the scientific evidence such that it, when considered along with findings  
19 from risk analyses and policy considerations, will help the U.S. Environmental Protection Agency (EPA)  
20 address these questions during the NAAQS review:

- 21           ▪ Has new information altered the scientific support for the occurrence of health effects  
22           following short- and/or long-term exposure to levels of CO found in the ambient air?
- 23           ▪ To what extent is key evidence becoming available that could inform our understanding of  
24           human subpopulations that are particularly sensitive to CO exposures? Is there new or  
25           emerging evidence on health effects beyond cardiovascular and respiratory endpoints  
26           (e.g., systemic effects, developmental effects, birth outcomes) that suggest additional  
27           sensitive subpopulations should be given increased focus in this review (e.g., neonates)?
- 28           ▪ What do recent studies focused on the near-roadway environment, including bus stops and  
29           intersections, tell us about high-exposure human subpopulations and the health effects of CO?  
30           What information is available on elevated exposures due to other transportation sources, such

- 1 as shipping, port operations, and recreational vehicles? What is the effect of altitude on CO  
2 sources and health effects?
- 3 ■ At what levels of CO exposure do health effects of concern occur?
  - 4 ■ To what extent is key scientific evidence becoming available to improve our understanding of  
5 the health effects associated with various time periods of CO exposures, including not only  
6 daily, but also chronic (months to years) exposures? To what extent is critical research  
7 becoming available that could improve our understanding of the relationship between various  
8 health endpoints and different lag periods (e.g., single day, multi-day distributed lags)?
  - 9 ■ To what extent does the evidence suggest that alternate dose indicators other than  
10 carboxyhemoglobin (COHb) levels (e.g., tissue oxygenation) should be evaluated to  
11 characterize the biological effect?
  - 12 ■ Has new information altered conclusions from previous reviews regarding the plausibility of  
13 adverse health effects caused by CO exposure?
  - 14 ■ To what extent have important uncertainties identified in the last review been reduced and/or  
15 have new uncertainties emerged?
  - 16 ■ Have new information or scientific insights altered the scientific conclusions regarding the  
17 occurrence of direct (or indirect) welfare effects associated with levels of CO found in the  
18 ambient air?

## 1.1. Legislative Requirements

19 Two sections of the Clean Air Act (CAA) govern the establishment and revision of the NAAQS.  
20 Section 108 (42 U.S.C. 7408) directs the Administrator to identify and list “air pollutants” that “in [her]  
21 judgment, may reasonably be anticipated to endanger public health and welfare” and whose “presence ...  
22 in the ambient air results from numerous or diverse mobile or stationary sources” and to issue air quality  
23 criteria for those that are listed (42 U.S.C. 7408). Air quality criteria are intended to “accurately reflect the  
24 latest scientific knowledge useful in indicating the kind and extent of identifiable effects on public health  
25 or welfare which may be expected from the presence of [a] pollutant in ambient air...” 42 U.S.C. 7408(b).

26 Section 109 (42 U.S.C. 7409) of the Clean Air Act directs the EPA Administrator to propose and  
27 promulgate “primary” and “secondary” National Ambient Air Quality Standards (NAAQS) for pollutants  
28 listed under section 108. Section 109(b)(1) defines a primary standard as one “the attainment and  
29 maintenance of which in the judgment of the Administrator, based on such criteria and allowing an

1 adequate margin of safety, are requisite to protect the public health.”<sup>1</sup> A secondary standard, as defined in  
2 section 109(b)(2), must “specify a level of air quality the attainment and maintenance of which, in the  
3 judgment of the U.S. EPA Administrator, based on such criteria, is required to protect the public welfare  
4 from any known or anticipated adverse effects associated with the presence of [the] pollutant in the  
5 ambient air.”<sup>2</sup> The requirement that primary standards include an adequate margin of safety was intended  
6 to address uncertainties associated with inconclusive scientific and technical information available at the  
7 time of standard setting. It was also intended to provide a reasonable degree of protection against hazards  
8 that research has not yet identified. See *Lead Industries Association v. EPA*, 647 F.2d 1130, 1154 (D.C.  
9 Cir 1980), cert. denied, 449 U.S. 1042 (1980); *American Petroleum Institute v. Costle*, 665 F.2d 1176,  
10 1186 (D.C. Cir. 1981) cert. denied, 455 U.S. 1034 (1982). The aforementioned uncertainties are  
11 components of the risk associated with pollution at levels below those at which human health effects can  
12 be said to occur with reasonable scientific certainty. Thus, in selecting primary standards that include an  
13 adequate margin of safety, the Administrator is seeking not only to prevent pollution levels that have been  
14 demonstrated to be harmful but also to prevent lower pollutant levels that may pose an unacceptable risk  
15 of harm, even if the risk is not precisely identified as to nature or degree.

16 In selecting a margin of safety, the EPA considers such factors as the nature and severity of the  
17 health effects involved, the size of sensitive population(s) at risk, and the kind and degree of the  
18 uncertainties that must be addressed. The selection of any particular approach to providing an adequate  
19 margin of safety is a policy choice left specifically to the Administrator’s judgment. See *Lead Industries*  
20 *Association v. EPA*, supra, 647 F.2d at 1161-62.

21 In setting standards that are “requisite” to protect public health and welfare, as provided in Section  
22 109(b), EPA’s task is to establish standards that are neither more nor less stringent than necessary for  
23 these purposes. In so doing, EPA may not consider the costs of implementing the standards. See *Whitman*  
24 *v. American Trucking Associations*, 531 U.S. 457, 465-472, 475-76 (D.C. Cir. 2001).

25 Section 109(d)(1) requires that “not later than December 31, 1980, and at 5-year intervals  
26 thereafter, the Administrator shall complete a thorough review of the criteria published under section 108  
27 and the national ambient air quality standards...and shall make such revisions in such criteria and  
28 standards and promulgate such new standards as may be appropriate...” Section 109(d)(2) requires that an  
29 independent scientific review committee “shall complete a review of the criteria...and the national  
30 primary and secondary ambient air quality standards...and shall recommend to the Administrator any

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<sup>1</sup> The legislative history of section 109 of the Clean Air Act indicates that a primary standard is to be set at “the maximum permissible ambient air level . . . which will protect the health of any [sensitive] group of the population,” and that for this purpose “reference should be made to a representative sample of persons comprising the sensitive group rather than to a single person in such a group” [S. Rep. No. 91-1196, 91st Cong., 2d Sess. 10 (1970)].

<sup>2</sup> Welfare effects as defined in section 302(h) [42 U.S.C. 7602(h)] include, but are not limited to, “effects on soils, water, crops, vegetation, man-made materials, animals, wildlife, weather, visibility and climate, damage to and deterioration of property, and hazards to transportation, as well as effects on economic values and on personal comfort and well-being.”

1 new...standards and revisions of existing criteria and standards as may be appropriate...” Since the early  
2 1980s, this independent review function has been performed by the Clean Air Scientific Advisory  
3 Committee (CASAC) of EPA’s Science Advisory Board (SAB).

## 1.2. History of the NAAQS for CO

4 On April 30, 1971, EPA promulgated identical primary and secondary NAAQS for CO, under  
5 section 109 of the Clean Air Act, set at 9 ppm, 8-h avg and 35 ppm, 1-h avg, neither to be exceeded more  
6 than once per year (36 FR 8186). In 1979, EPA published the *Air Quality Criteria Document for Carbon*  
7 *Monoxide* (U.S. EPA, 1979a), which updated the scientific criteria upon which the initial CO standards  
8 were based. A Staff Paper (U.S. EPA, 1979b) was prepared and, along with the AQCD, served as the basis  
9 for development of proposed rulemaking (45 FR 55066) published on August 18, 1980. Delays due to  
10 uncertainties regarding the scientific basis for the final decision resulted in EPA announcing a second  
11 public comment period (47 FR 26407). Following substantial reexamination of the scientific data, EPA  
12 prepared an Addendum to the 1979 AQCD (U.S. EPA, 1984b) and an updated Staff Paper (U.S. EPA,  
13 1984a). Following review by CASAC, EPA announced its final decision (50 FR 37484) not to revise the  
14 existing primary standard and to revoke the secondary standard for CO on September 13, 1985, due to a  
15 lack of evidence of direct effects on public welfare at ambient concentrations.

16 In 1987, EPA initiated action to revise the criteria for CO and released a revised AQCD for CASAC  
17 and public review. In a “closure letter” (McClellan, 1991) sent to the Administrator, the CASAC  
18 concluded that the AQCD (U.S. EPA, 2000) “. . . provides a scientifically balanced and defensible  
19 summary of current knowledge of the effects of this pollutant and provides an adequate basis for the EPA  
20 to make a decision as to the appropriate primary NAAQS for CO.” A revised Staff Paper subsequently  
21 was reviewed by CASAC and the public, and in a “closure letter” (McClellan, 1992) sent to the  
22 Administrator, CASAC stated “. . . that a standard of the present form and with a numerical value similar  
23 to that of the present standard would be supported by the present scientific data on health effects of  
24 exposure to carbon monoxide.” Based on the revised AQCD (U.S. EPA, 2000) and staff conclusions and  
25 recommendations contained in the revised Staff Paper (U.S. EPA, 1992), the Administrator announced the  
26 final decision (59 FR 38906) on August 1, 1994, that revision of the primary NAAQS for CO was not  
27 appropriate at that time.

28 In 1997, revisions to the AQCD were initiated. A workshop was held in September 1998 to review  
29 and discuss material contained in the revised AQCD. On June 9, 1999, CASAC held a public meeting to  
30 review the draft AQCD and a draft exposure analysis methodology document. Comments from CASAC  
31 and the public were considered in a second draft AQCD, which was reviewed at a CASAC meeting, held  
32 on November 18, 1999. After revision of the second draft AQCD, the final AQCD (U.S. EPA, 2000) was

1 released in August 2000. EPA put the review on hold when Congress called on the National Research  
2 Council (NRC) to conduct a review of the impact of meteorology and topography on ambient CO  
3 concentrations in high altitude and extreme cold regions of the U.S. In response, the NRC convened the  
4 committee on Carbon Monoxide Episodes in Meteorological and Topographical Problem Areas, which  
5 focused on Fairbanks, Alaska as a case study in an interim report, which was completed in 2002. A final  
6 report, *Managing Carbon Monoxide Pollution in Meteorological and Topographical Problem Areas*, was  
7 published in 2003 (NRC, 2003) and offered a wide range of recommendations on management of CO air  
8 pollution, cold start emissions standards, oxygenated fuels, and CO monitoring. EPA did not complete the  
9 review which started in 1997.

### 1.3. Document Development

10 EPA initiated the current review of the NAAQS for CO on September 13, 2007 with a call for  
11 information from the public (72 FR 52369). In addition to the call for information, publications were  
12 identified through an ongoing literature search process that includes extensive computer database mining  
13 on specific topics. Additional publications were identified by EPA scientists in a variety of disciplines by  
14 combing through peer-reviewed scientific literature obtained through these ongoing literature searches,  
15 reviewing previous EPA reports, and a review of reference lists from important publications. All relevant  
16 epidemiologic, human clinical, and animal toxicological studies, including those related to exposure-  
17 response relationships, mode(s) of action (MOA), or susceptible subpopulations published since the last  
18 review were considered. Added to the body of research were EPA's analyses of air quality and emissions  
19 data, studies on atmospheric chemistry, transport, and fate of these emissions, as well as issues related to  
20 exposure to CO. A literature search conducted on the ecological effects of ambient CO did not identify  
21 any relevant information. Further information was acquired from consultation with scientific experts and  
22 the public.

### 1.4. Document Organization

23 The ISA is composed of five chapters. This introductory chapter presents background information,  
24 and provides an overview of EPA's framework for making causal judgments. Chapter 2 is an integrated  
25 summary of key findings and conclusions regarding the source to dose paradigm, MOA, and important  
26 health effects of CO, including cardiovascular, nervous system, perinatal/developmental, respiratory, and  
27 mortality outcomes. Chapter 3 highlights key concepts and evidence relevant to understanding the  
28 sources, ambient concentrations, atmospheric behavior, and exposure to ambient CO. Chapter 4 describes



1 the dosimetry and pharmacokinetics of CO, including formation and fate of carboxyhemoglobin (COHb).  
2 Chapter 5 presents a discussion of the MOA of CO and evaluates and integrates epidemiologic, human  
3 clinical, and animal toxicological information on the health effects of CO, including cardiovascular and  
4 systemic effects, central nervous system (CNS) effects, birth outcomes and developmental effects,  
5 respiratory effects, and mortality.

6 A series of annexes supplement this ISA. The annexes provide tables summarizing additional  
7 details of the pertinent literature published since the last review, as well as selected older studies of  
8 particular interest. These annexes contain information on:

- 9       ▪ atmospheric chemistry of CO, sampling and analytic methods for measurement of CO  
10       concentrations, emissions, sources and human exposure to CO (Annex A)
- 11       ▪ studies on the dosimetry and pharmacokinetics of CO (Annex B)
- 12       ▪ epidemiologic studies of health effects from short- and long-term exposure to CO (Annex C)
- 13       ▪ controlled human exposure studies of health effects related to exposure to CO (Annex D);  
14       and
- 15       ▪ toxicological studies of health effects in laboratory animals (Annex E)

16 Within the Annexes, detailed information about methods and results of health studies is  
17 summarized in tabular format, and generally includes information about concentrations of CO and  
18 averaging times, study methods employed, results and comments, and quantitative results for  
19 relationships between effects and exposure to CO.

## 1.5. Document Scope

20 For the current review of the primary CO standards, relevant scientific information on human  
21 exposures and health effects associated with exposure to ambient CO has been assessed. Health effects  
22 resulting from accidental exposures to very high concentrations of non-ambient CO (i.e., carbon  
23 monoxide poisoning) are not directly relevant to ambient exposures, and as such, a discussion of these  
24 effects has deliberately been excluded from this document. For a detailed review of the effects of high  
25 level exposures to CO, the reader is referred to the extensive body of literature related to carbon  
26 monoxide poisoning (Ernst and Zibrak, 1998; Penney, 2007; Raub et al., 2000). The possible influence of  
27 other atmospheric pollutants on the interpretation of the role of CO in health effects studies is considered.  
28 This includes other pollutants with the potential to co-occur in the environment (e.g., nitrogen dioxide  
29 [NO<sub>2</sub>], sulfur dioxide [SO<sub>2</sub>], ozone [O<sub>3</sub>], and particulate matter [PM]). The review also assesses relevant  
30 scientific information associated with known or anticipated public welfare effects that may be identified.

1 As discussed in Section 1.3, a literature review of the ecological effects of ambient CO identified no  
2 relevant information, and thus these effects are not assessed. With regard to climate effects this review  
3 includes updates and additional data available since the 2000 CO AQCD on the interaction of largely  
4 urban CO emissions and hydroxyl radical concentrations, and on background concentrations in the U.S.

## 1.6. EPA Framework for Causal Determination

5 The EPA has developed a consistent and transparent basis to evaluate the causal nature of air  
6 pollution-induced health or environmental effects. The framework described below establishes uniform  
7 language concerning causality and brings more specificity to the findings. This standardized language was  
8 drawn from across the federal government and wider scientific community, especially from the recent  
9 National Academy of Sciences (NAS) Institute of Medicine (IOM) document, *Improving the Presumptive*  
10 *Disability Decision-Making Process for Veterans* (IOM, 2008), the most recent comprehensive work on  
11 evaluating causality.

12 This introductory section focuses on the evaluation of health effects evidence. While focusing on  
13 human health outcomes, the concepts are also generally relevant to causality determination for welfare  
14 effects. This section:

- 15       ▪ describes the kinds of scientific evidence used in establishing a general causal relationship  
16       between exposure and health effects;
- 17       ▪ defines cause, in contrast to statistical association;
- 18       ▪ discusses the sources of evidence necessary to reach a conclusion about the existence of a  
19       causal relationship;
- 20       ▪ highlights the issue of multifactorial causation;
- 21       ▪ identifies issues and approaches related to uncertainty; and
- 22       ▪ provides a framework for classifying and characterizing the weight of evidence in support of  
23       a general causal relationship.

24 Approaches to assessing the separate and combined lines of evidence (e.g., epidemiologic, human  
25 clinical, and animal toxicological studies) have been formulated by a number of regulatory and science  
26 agencies, including the IOM of the NAS (IOM, 2008), International Agency for Research on Cancer  
27 (IARC, 2006), *EPA Guidelines for Carcinogen Risk Assessment* (U.S. EPA, 2005), Centers for Disease  
28 Control and Prevention (CDC, 2004), and National Acid Precipitation Assessment Program (NAPAP,  
29 1991). These formalized approaches offer guidance for assessing causality. The frameworks are similar in

1 nature, although adapted to different purposes, and have proven effective in providing a uniform structure  
2 and language for causal determinations. Moreover, these frameworks have supported decision-making  
3 under conditions of uncertainty.

### **1.6.1. Scientific Evidence Used in Establishing Causality**

4 Causality determinations are based on the evaluation and synthesis of evidence from across  
5 scientific disciplines; the type of evidence that is most important for such determinations will vary by  
6 pollutant or assessment. The most compelling evidence of a causal relationship between pollutant  
7 exposures and human health effects comes from human clinical studies. This type of study experimentally  
8 evaluates the health effects of administered exposures in human volunteers under highly-controlled  
9 laboratory conditions.

10 In epidemiologic or observational studies of humans, the investigator does not control exposures or  
11 intervene with the study population. Broadly, observational studies can describe associations between  
12 exposures and effects. These studies fall into several categories: cross-sectional, prospective cohort, and  
13 time-series studies. “Natural experiments” offer the opportunity to investigate changes in health with a  
14 change in exposure; these include comparisons of health effects before and after a change in population  
15 exposures, such as closure of a pollution source.

16 Experimental animal data complement the clinical and observational data; these studies can help  
17 characterize effects of concern, exposure-response relationships, susceptible subpopulations and MOAs.  
18 In the absence of clinical or epidemiologic data, animal data alone may be sufficient to support a likely  
19 causal determination, assuming that humans respond similarly to the experimental species.

### **1.6.2. Association and Causation**

20 “Cause” is a significant, effectual relationship between an agent and an effect on health or public  
21 welfare. “Association” is the statistical dependence among events, characteristics, or other variables. An  
22 association is prima facie evidence for causation; alone, however, it is insufficient proof of a causal  
23 relationship between exposure and disease. Unlike an association, a causal claim supports the creation of  
24 counterfactual claims; that is, a claim about what the world would have been like under different or  
25 changed circumstances (IOM, 2008). Much of the newly available health information evaluated in this  
26 ISA comes from epidemiologic studies that report a statistical association between ambient exposure and  
27 health outcome.

28 Many of the health and environmental outcomes reported in these studies have complex etiologies.  
29 Diseases such as asthma, coronary heart disease (CHD) or cancer are typically initiated by multiple

1 agents. Outcomes depend on a variety of factors, such as age, genetic susceptibility, nutritional status,  
2 immune competence, and social factors (Gee and Payne-Sturges, 2004; IOM, 2008). Effects on  
3 ecosystems are often also multifactorial with a complex web of causation. Further, exposure to a  
4 combination of agents could cause synergistic or antagonistic effects. Thus, the observed risk represents  
5 the net effect of many actions and counteractions.

### 1.6.3. Evaluating Evidence for Inferring Causation

6 Moving from association to causation involves elimination of alternative explanations for the  
7 association. In estimating the causal influence of an exposure on health or environmental effects, it is  
8 recognized that scientific findings incorporate uncertainty. Uncertainty can be defined as a state of having  
9 limited knowledge where it is impossible to exactly describe an existing state or future outcome; e.g., the  
10 lack of knowledge about the correct value for a specific measure or estimate. Uncertainty characterization  
11 and uncertainty assessment are two activities that lead to different degrees of sophistication in describing  
12 uncertainty. Uncertainty characterization generally involves a qualitative discussion of the thought  
13 processes that lead to the selection and rejection of specific data, estimates, scenarios, etc. The uncertainty  
14 assessment is more quantitative. The process begins with simpler measures (e.g., ranges) and simpler  
15 analytical techniques and progresses, to the extent needed to support the decision for which the  
16 assessment is conducted, to more complex measures and techniques. Data will not be available for all  
17 aspects of an assessment and those data that are available may be of questionable or unknown quality. In  
18 these situations, evaluation of uncertainty can include professional judgment or inferences based on  
19 analogy with similar situations. The net result is that the assessments will be based on a number of  
20 assumptions with varying degrees of uncertainty.

21 Uncertainties commonly encountered in evaluating health evidence for the criteria air pollutants are  
22 outlined below for epidemiologic and experimental studies. Various approaches to characterizing  
23 uncertainty include classical statistical methods, sensitivity analysis, or probabilistic uncertainty analysis,  
24 in order of increasing complexity and data requirements. The ISA generally evaluates uncertainties  
25 qualitatively in assessing the evidence from across studies; in some situations quantitative analysis  
26 approaches, such as meta-regression may be used.

27 Controlled human exposure studies evaluate the effects of exposures to a variety of pollutants in a  
28 highly controlled laboratory setting. Also referred to as human clinical studies, these experiments allow  
29 investigators to expose subjects to known concentrations of air pollutants under carefully regulated  
30 environmental conditions and activity levels. In some instances, controlled human exposure studies can  
31 also be used to characterize concentration-response relationships at pollutant concentrations relevant to  
32 ambient conditions. Controlled human exposures are typically conducted using a randomized crossover

1 design with subjects exposed both to CO and a clean air control. In this way, subjects serve as their own  
2 controls, effectively controlling for many potential confounders. However, human clinical studies are  
3 limited by a number of factors including a small sample size and short exposure time. The repetitive  
4 nature of ambient CO exposures at levels that can vary widely may lead to cumulative health effects, but  
5 this type of exposure is not practical to replicate in a laboratory setting. In addition, although subjects do  
6 serve as their own controls, personal exposure to pollutants in the hours and days preceding the controlled  
7 exposures may vary significantly between and within individuals. Finally, human clinical studies require  
8 investigators to adhere to stringent health criteria for a subject to be included in the study, and therefore  
9 the results cannot necessarily be generalized to an entire population. Although some human clinical  
10 studies have included health-compromised individuals such as asthmatics or individuals with chronic  
11 obstructive pulmonary disease (COPD) or coronary artery disease (CAD), these individuals must also be  
12 relatively healthy and do not represent the most sensitive individuals in the population. Thus, a lack of  
13 observation of effects from human clinical studies does not necessarily mean that a causal relationship  
14 does not exist. While human clinical studies provide important information on the biological plausibility  
15 of associations observed between air pollutant exposure and health outcomes in epidemiologic studies,  
16 observed effects in these studies may underestimate the response in certain subpopulations.

17 Epidemiologic studies provide important information on the associations between health effects  
18 and exposure of human populations to ambient air pollution. These studies also help to identify  
19 susceptible or vulnerable subgroups and associated risk factors. There are important methodological  
20 issues to be considered in evaluating results from air pollution epidemiologic studies, especially the  
21 potential for confounding and/or effect modification; and exposure measurement error.

22 Scientific judgment is needed regarding sources and magnitude of potential confounding by  
23 covariates, together with judgment about how well the existing constellation of study designs, results, and  
24 analyses address this potential threat to inferential validity. One key consideration is evaluation of the  
25 potential contribution of CO to health effects when it is a component of a complex air pollutant mixture.  
26 There are multiple ways by which CO might cause or be associated with adverse health effects. First, the  
27 reported CO effect estimates in epidemiologic studies may reflect independent CO effects on health.  
28 Second, ambient CO may be serving as an indicator of complex ambient air pollution mixtures that share  
29 the same source as CO (e.g., motor vehicle emissions). Finally, copollutants may mediate the effects of  
30 CO, or CO may influence the toxicity of copollutants.

31 Epidemiologists use the term “interaction” or “effect modification” to denote the departure from  
32 the observed joint risk from what might be expected based on the separate effects of the factors. In  
33 addition, confounding can result in the production of an association between adverse health effects and  
34 CO that is actually attributable to another factor that is associated with CO in a particular study.  
35 Multivariate models are the most widely used strategy to address confounding in epidemiologic studies,

1 but such models are not always easily interpreted when assessing effects of covarying pollutants such as  
2 fine particulate matter (PM<sub>2.5</sub>) and NO<sub>2</sub>.

3         Inferring causation requires consideration of potential confounders. In confounding, the apparent  
4 effect of the exposure of interest is distorted because the effect of an extraneous factor is mistaken for or  
5 mixed with the actual exposure effect, which may be null. When associations are found in epidemiologic  
6 studies, one approach to remove spurious associations from possible confounders is to control for  
7 characteristics that may differ between exposed and unexposed persons; this is frequently termed  
8 “adjustment.” Multivariable regression models constitute one tool for estimating the association between  
9 exposure and outcome after adjusting for characteristics of participants that might confound the results.  
10 The use of multipollutant regression models has been the prevailing approach for controlling potential  
11 confounding by copollutants in air pollution health effects studies. Finding the likely causal pollutant  
12 from multipollutant regression models is made difficult by the possibility that one or more air pollutants  
13 may be acting as a surrogate for an unmeasured or poorly-measured pollutant or for a particular mixture  
14 of pollutants. Further, the correlation between the air pollutant of interest and various copollutants may  
15 show temporal and spatial incongruities that can influence exposures and health effects. Thus, results of  
16 models that attempt to distinguish gaseous and particle effects must be interpreted with caution. Despite  
17 these limitations, the use of multipollutant models is still the prevailing approach employed in most air  
18 pollution epidemiologic studies and may provide some insight into the potential for confounding or  
19 interaction among pollutants.

20         Another way to adjust for potential confounding is through stratified analysis, i.e., examining the  
21 association within homogeneous groups with respect to the confounding variable. Stratified analysis can  
22 also be used to examine potential effect modification. The use of stratified analyses has an additional  
23 benefit: it allows examination of effect modification through comparison of the effect estimates across  
24 different groups. If investigators successfully measured characteristics that distort the results, adjustment  
25 of these factors help separate a spurious from a true causal association. Appropriate statistical adjustment  
26 for confounders requires identifying and measuring all reasonably expected confounders. Deciding which  
27 variables to control for in a statistical analysis of the association between exposure and disease depends  
28 on knowledge about possible mechanisms and the distributions of these factors in the population under  
29 study. Identifying these mechanisms makes it possible to control for potential sources that may result in a  
30 spurious association.

31         Measurement error is another problem encountered when adjusting for spurious associations.  
32 Controlling for confounders, whether by adjustment or stratification, is only successful when the  
33 confounder is well-measured. Considered together, the effects of a well-measured covariate may be  
34 overestimated in contrast to a covariate measured with greater error. There are several components that  
35 contribute to exposure measurement error in these studies, including the difference between true and  
36 measured ambient concentrations, the difference between average personal exposure to ambient pollutants

1 and ambient concentrations at central monitoring sites, and the use of average population exposure rather  
2 than individual exposure estimates. Previous assessments have examined the role of measurement error in  
3 time-series epidemiologic studies using simulated data and mathematical analyses and suggested that  
4 “transfer of effects” would only occur under unusual circumstances (e.g., “true” predictors having high  
5 positive or negative correlation; substantial measurement error; or extremely negatively correlated  
6 measurement errors) (U.S. EPA, 2004).

7 Confidence that unmeasured confounders are not producing the findings is increased when multiple  
8 studies are conducted in various settings using different subjects or exposures; each of which might  
9 eliminate another source of confounding from consideration. Thus, multi-city studies which use a  
10 consistent method to analyze data from across locations with different levels of covariates can provide  
11 insight on potential confounding in associations. The number and degree of diversity of covariates, as  
12 well as their relevance to the potential confounders, remain matters of scientific judgment. Intervention  
13 studies, because of their experimental nature, can be particularly useful in characterizing causation.

14 In addition to clinical and epidemiologic studies, the tools of experimental biology have been  
15 valuable for developing insights into human physiology and pathology. Laboratory tools have been  
16 extended to explore the effects of putative toxicants on human health, especially through the study of  
17 model systems in other species. Background knowledge of the biological mechanisms by which an  
18 exposure might or might not cause disease can prove crucial in establishing, or negating, a causal claim.  
19 Testable hypotheses about the causal nature of proposed mechanisms or modes of action are central to  
20 utilizing experimental data in causal determinations.

21 Interpretations of experimental studies of air pollution effects in animals, as in the case of  
22 environmental comparative toxicology studies, are affected by limitations associated with extrapolation  
23 models. The differences between humans and rodents with regard to pollutant absorption and distribution  
24 profiles based on metabolism, hormonal regulation, breathing pattern, exposure dose, and differences in  
25 lung structure and anatomy all have to be taken into consideration. Also, in spite of a high degree of  
26 homology and the existence of a high percentage of orthologous genes across humans and rodents  
27 (particularly mice), extrapolation of molecular alterations at the gene level is complicated by species-  
28 specific differences in transcriptional regulation. Given these molecular differences, there are  
29 uncertainties associated with quantitative extrapolations at this time between laboratory animals and  
30 humans of observed pollutant-induced pathophysiological alterations under the control of widely varying  
31 biochemical, endocrine, and neuronal factors.

## 1.6.4. Application of Framework for Causal Determination

1 EPA uses a two-step approach to evaluate the scientific evidence on health or environmental effects  
2 of criteria pollutants. The first step determines the weight of evidence in support of causation and  
3 characterizes the strength of any resulting causal classification. The second step includes further  
4 evaluation of the quantitative evidence regarding the concentration-response relationships and the loads or  
5 levels, duration and pattern of exposures at which effects are observed.

6 To aid judgment, various “aspects”<sup>1</sup> of causality have been discussed by many philosophers and  
7 scientists. The most widely cited aspects of causality in epidemiology, and public health, in general, were  
8 articulated by Sir Austin Bradford Hill in 1965 and have been widely used (CDC, 2004; IARC, 2006;  
9 IOM, 2008; U.S. EPA, 2005, 2008d). Several adaptations of the Hill aspects have been used in aiding  
10 causality judgments in the ecological sciences (Adams, 2003; Buck et al., 2000; Collier, 2003; Fox, 1991;  
11 Gerritsen et al., 1998). These aspects (Hill, 1965) have been modified (Table 1-1) for use in causal  
12 determinations specific to health and welfare effects or pollutant exposures (U.S. EPA, 2008f).<sup>2</sup> Some  
13 aspects are more likely than others to be relevant for evaluating evidence on the health or environmental  
14 effects of criteria air pollutants. For example, the analogy aspect does not always apply, especially for the  
15 gaseous criteria pollutants, and specificity would not be expected for multi-etiological health outcomes,  
16 such as asthma or cardiovascular disease, or ecological effects related to acidification. Aspects that  
17 usually play a larger role in determination of causality are consistency of results across studies, coherence  
18 of effects observed in different study types or disciplines, biological plausibility, exposure-response  
19 relationship, and evidence from “natural” experiments.

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**Table 1-1. Aspects to aid in judging causality.**

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<b>Consistency of the observed association</b>	An inference of causality is strengthened when a pattern of elevated risks is observed across several independent studies. The reproducibility of findings constitutes one of the strongest arguments for causality. If there are discordant results among investigations, possible reasons such as differences in exposure, confounding factors, and the power of the study are considered.
<b>Coherence</b>	An inference of causality from epidemiologic associations may be strengthened by other lines of evidence (e.g., clinical and animal studies) that support a cause-and-effect interpretation of the association. Evidence on ecological or welfare effects may be drawn from a variety of experimental approaches (e.g., greenhouse, laboratory, and field) and subdisciplines of ecology (e.g., community ecology, biogeochemistry and paleological/historical reconstructions). The coherence of evidence from various fields greatly adds to the strength of an inference of causality. The absence of other lines of evidence, however, is not a reason to reject causality.
<b>Biological plausibility.</b>	An inference of causality tends to be strengthened by consistency with data from experimental studies or other sources demonstrating plausible biological mechanisms. A proposed mechanistic linking between an effect and exposure to the agent is an important source of support for causality, especially when data establishing the existence and functioning of those mechanistic links are available. A lack of biologic understanding, however, is not a reason to reject causality.

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<sup>1</sup> The “aspects” described by Hill (1965) have become, in the subsequent literature, more commonly described as “criteria.” The original term “aspects” is used here to avoid confusion with ‘criteria’ as it is used, with different meaning, in the Clean Air Act.

<sup>2</sup> The Hill aspects were developed for interpretation of epidemiologic results. They have been modified here for use with a broader array of data, i.e., epidemiologic, controlled human exposure, and animal toxicological studies, as well as in vitro data, and to be more consistent with EPA’s Guidelines for Carcinogen Risk Assessment.



<b>Strength of the observed association</b>	The finding of large, precise risks increases confidence that the association is not likely due to chance, bias, or other factors. However, given a truly causal agent, a small magnitude in the effect could follow from a lower level of exposure, a lower potency, or the prevalence of other agents causing similar effects. While large effects support causality, modest effects therefore do not preclude it.
<b>Biological gradient (exposure-response relationship)</b>	A clear exposure-response relationship (e.g., increasing effects associated with greater exposure) strongly suggests cause and effect, especially when such relationships are also observed for duration of exposure (e.g., increasing effects observed following longer exposure times). There are, however, many possible reasons that a study may fail to detect an exposure-response relationship. Thus, although the presence of a biologic gradient may support causality, the absence of an exposure-response relationship does not exclude a causal relationship.
<b>Experimental evidence.</b>	The strongest evidence for causality can be provided when a change in exposure brings about a change in occurrence or frequency of health or welfare effects.
<b>Temporal relationship of the observed association</b>	Evidence of a temporal sequence between the introduction of an agent, and appearance of the effect, constitutes another argument in favor of causality.
<b>Analogy</b>	Structure activity relationships and information on the agent's structural analogs can provide insight into whether an association is causal. Similarly, information on mode of action for a chemical, as one of many structural analogs, can inform decisions regarding likely causality.

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1           While these aspects provide a framework for assessing the evidence, they do not lend themselves to  
2 being considered in terms of simple formulas or fixed rules of evidence leading to conclusions about  
3 causality (Hill, 1965). For example, one cannot simply count the number of studies reporting statistically  
4 significant results or statistically nonsignificant results and reach credible conclusions about the relative  
5 weight of the evidence and the likelihood of causality. Rather, these important considerations are taken  
6 into account with the goal of producing an objective appraisal of the evidence, informed by peer and  
7 public comment and advice, which includes weighing alternative views on controversial issues.  
8 Additionally, it is important to note that the aspects in Table 1-1 cannot be used as a strict checklist, but  
9 rather to determine the weight of the evidence for inferring causality. In particular, not meeting one or  
10 more of the principles does not automatically preclude a determination of causality (e.g., see discussion in  
11 (CDC, 2004).

## 1.6.5. Determination of Causality

12           In the ISA, EPA assesses the results of recent relevant publications, building upon evidence  
13 available during the previous NAAQS review, to draw conclusions on the causal relationships between  
14 relevant pollutant exposures and health or environmental effects. This ISA uses a five-level hierarchy that  
15 classifies the weight of evidence for causation, not just association<sup>1</sup>; that is, whether the weight of  
16 scientific evidence makes causation at least as likely as not, in the judgment of the reviewing group. In  
17 developing this hierarchy, EPA has drawn on the work of previous evaluations, most prominently the  
18 IOM's *Improving the Presumptive Disability Decision-Making Process for Veterans* (IOM, 2008), EPA's  
19 Guidelines for Carcinogen Risk Assessment (U.S. EPA, 2005), and the U.S. Surgeon General's smoking  
20 reports (CDC, 2004). In the ISA, EPA uses a series of five descriptors to characterize the weight of

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<sup>1</sup> It should be noted that the CDC and IOM frameworks use a four-category hierarchy for the strength of the evidence. A five-level hierarchy is used here to be consistent with the EPA Guidelines for Carcinogen Risk Assessment and to provide a more nuanced set of categories.

1 evidence for causality (U.S. EPA, 2008d, f). This weight of evidence evaluation is based on various lines  
 2 of evidence from across the health and environmental effects disciplines. These separate judgments are  
 3 integrated into a qualitative statement about the overall weight of the evidence and causality. The five  
 4 descriptors for causal determination are described in Table 1-2.

**Table 1-2. Weight of evidence for causal determination.**

	<b>Health Effects</b>	<b>Ecological and Welfare Effects</b>
<b>Causal relationship</b>	Evidence is sufficient to conclude that there is a causal relationship between relevant pollutant exposures and the health outcome. That is, a positive association has been observed between the pollutant and the outcome in studies in which chance, bias, and confounding could be ruled out with reasonable confidence. Evidence includes, for example, controlled human exposure studies; or observational studies that cannot be explained by plausible alternatives or are supported by other lines of evidence (e.g., animal studies or mode of action information). Evidence includes replicated and consistent high-quality studies by multiple investigators.	Evidence is sufficient to conclude that there is a causal relationship between relevant pollutant exposure and the outcome. Causality is supported when an association has been observed between the pollutant and the outcome in studies in which chance, bias, and confounding could be ruled out with reasonable confidence. Controlled exposure (laboratory or small- to medium-scale field studies) provides the strongest evidence for causality, but the scope of inference may be limited. Generally, determination is based on multiple studies conducted by multiple research groups, and evidence that is considered sufficient to infer a causal relationship is usually obtained from the joint consideration of many lines of evidence that reinforce each other.
<b>Likely to be a causal relationship</b>	Evidence is sufficient to conclude that a causal relationship is likely to exist between relevant pollutant exposures and health outcome but important uncertainties remain. That is, a positive association has been observed between the pollutant and the outcome in studies in which chance and bias can be ruled out with reasonable confidence but potential issues remain. For example: a) observational studies show positive associations but copollutant exposures are difficult to address and/or other lines of evidence (controlled human exposure, animal, or mode of action information) are limited or inconsistent; or b) animal evidence from multiple studies, sex, or species is positive but limited or no human data are available. Evidence generally includes replicated and high-quality studies by multiple investigators.	Evidence is sufficient to conclude that there is a likely causal association between relevant pollutant exposures and the outcome. That is, an association has been observed between the pollutant and the outcome in studies in which chance, bias and confounding are minimized, but uncertainties remain. For example, field studies show a relationship, but suspected interacting factors cannot be controlled, and other lines of evidence are limited or inconsistent. Generally, determination is based on multiple studies in multiple research groups.
<b>Suggestive of a causal relationship</b>	Evidence is suggestive of a causal relationship between relevant pollutant exposures and the health outcome, but is limited because chance, bias and confounding cannot be ruled out. For example, at least one high-quality study shows a positive association but the results of other studies are inconsistent.	Evidence is suggestive of an association between relevant pollutant exposures and the outcome, but chance, bias and confounding cannot be ruled out. For example, at least one high-quality study shows an association, but the results of other studies are inconsistent.
<b>Inadequate to infer a causal relationship</b>	Evidence is inadequate to determine that a causal relationship exists between relevant pollutant exposures and health outcome. The available studies are of insufficient quantity, quality, consistency or statistical power to permit a conclusion regarding the presence or absence of an association between relevant pollutant exposure and the outcome.	The available studies are of insufficient quality, consistency or statistical power to permit a conclusion regarding the presence or absence of an association between relevant pollutant exposure and the outcome.
<b>Suggestive of no causal relationship</b>	Evidence is suggestive of no causal relationship between relevant pollutant exposures and health outcome. Several adequate studies, covering the full range of levels of exposure that human beings are known to encounter and considering susceptible or vulnerable subpopulations, are mutually consistent in not showing a positive association between exposure and the outcome at any level of exposure.	Several adequate studies, examining relationships between relevant exposures and outcomes, are consistent in failing to show an association between exposure and the outcome at any level of exposure.

Source: U.S. EPA (2008f)

5 For the CO ISA, determination of causality involved the evaluation of evidence for different types  
 6 of health effects associated with short- and long-term exposure periods. In making determinations of  
 7 causality for CO, evidence was evaluated for health outcome categories, such as cardiovascular effects,  
 8 and then conclusions were drawn based upon the integration of evidence from across disciplines  
 9 (e.g., epidemiology, clinical studies and toxicology) and also across the suite of related individual health  
 10 outcomes. To accomplish this integration, evidence from multiple and various types of studies was

1 considered. Response was evaluated over a range of observations which was determined by the type of  
2 study and methods of exposure or dose and response measurements. Results from different protocols were  
3 compared and contrasted.

4 In drawing judgments regarding causality for the criteria air pollutants, EPA focuses on evidence of  
5 effects at relevant pollutant exposures. To best inform reviews of the NAAQS, these evaluations go  
6 beyond a determination of causality at any dose or concentration to emphasize the relationship apparent at  
7 relevant pollutant exposures. Concentrations generally within an order of magnitude or two of ambient  
8 pollutant measurements are considered to be relevant for this determination. Building upon the  
9 determination of causality are questions relevant to quantifying health or environmental risks based on  
10 our understanding of the quantitative relationships between pollutant exposures and health or welfare  
11 effects. While the causality determination is based primarily on evaluation of health or environmental  
12 effects evidence, EPA also evaluates evidence related to the doses or levels at which effects are observed.  
13 Considerations relevant to evaluation of quantitative relationships for health and environmental effects are  
14 summarized below.

## **Effects on Human Populations**

15 Important questions regarding quantitative relationships include:

- 16 ▪ What is the concentration-response or dose-response relationship in the human population?
- 17 ▪ What is the interrelationship between incidence and severity of effect?
- 18 ▪ What exposure conditions (dose or exposure, duration and pattern) are important?
- 19 ▪ What subpopulations appear to be differentially affected i.e., more susceptible or vulnerable  
20 to effects?

21 Addressing these questions requires evaluating the entirety of policy-relevant quantitative evidence  
22 regarding the concentration-response relationships including levels of pollutant and exposure durations at  
23 which effects were observed, and subpopulations that differ in susceptibility or vulnerability from the  
24 general population. This integration of evidence resulted in identification of a study or set of studies that  
25 best approximated the concentration response relationship for the U.S. population, given the current state  
26 of knowledge and the uncertainties that surrounded these estimates.

27 An important consideration in characterizing the public health impacts associated with exposure to  
28 a pollutant is whether the concentration-response relationship is linear across the full concentration range  
29 encountered, or if nonlinear relationships exist along any part of this range. Of particular interest is the  
30 shape of the concentration-response curve at and below the level of the current standards. The shape of  
31 the concentration-response curve varies, depending on the type of health outcome, underlying biological  
32 mechanisms and dose. At the human population level, however, various sources of variability and

1 uncertainty (such as the low data density in the lower concentration range, possible influence of  
2 measurement error, and individual differences in susceptibility to air pollution health effects) tend to  
3 smooth and “linearize” the concentration-response function. Additionally, many chemicals and agents  
4 may act by perturbing naturally occurring background processes that lead to disease, which also linearizes  
5 population concentration-response relationships (Clewell and Crump, 2005; Crump et al., 1976; Hoel,  
6 1980). These attributes of population dose-response may explain why the available human data at ambient  
7 concentrations for some environmental pollutants (e.g., PM, O<sub>3</sub>, lead [Pb], environmental tobacco smoke  
8 [ETS], radiation) do not exhibit evident thresholds for cancer or noncancer health effects, even though  
9 likely mechanisms include nonlinear processes for some key events. These attributes of human population  
10 dose-response relationships have been extensively discussed in the broader epidemiologic literature  
11 (Rothman and Greenland, 1998).

### **Effects on Ecosystems or Public Welfare**

12 Key questions for understanding the quantitative relationships between exposure (or concentration  
13 or deposition) to a pollutant and risk to ecosystems or the public welfare include:

- 14       ▪ What elements of the ecosystem (e.g., types, regions, taxonomic groups, populations,  
15       functions, etc.) appear to be affected, or are more sensitive to effects?
- 16       ▪ Under what exposure conditions (amount deposited or concentration, duration and pattern)  
17       are effects seen?
- 18       ▪ What is the shape of the concentration-response or exposure-response relationship?

19 Evaluations of causality typically characterize how the probability of ecological and welfare effects  
20 change in response to exposure. A challenge to the quantification of exposure-response relationships for  
21 ecological effects is the variability across ecosystems. Ecological responses are evaluated within the range  
22 of observations, so a quantitative relationship may be determined for a given ecological system and scale.  
23 There is great regional and local variability in ecosystems, thus an exposure-response relationship  
24 generally cannot be determined at the larger national or even regional scale. Quantitative relationships  
25 therefore are available site by site. For example, an ecological response to deposition of a given pollutant  
26 can differ greatly between ecosystems. Where results from greenhouse or animal ecotoxicological studies  
27 are available, they may be used to aid in characterizing exposure-response relations, particularly relative  
28 to mechanisms of action, and characteristics of sensitive biota.

## 1.6.6. Concepts in Evaluating Adversity of Health Effects

1 In evaluating the health evidence, a number of factors can be considered in determining the extent  
2 to which health effects are “adverse” for health outcomes such as changes in lung function or in  
3 cardiovascular health measures. Some health outcome events, such as hospitalization for respiratory or  
4 cardiovascular diseases, are clearly considered adverse; what is more difficult is determining the extent of  
5 change in the more subtle health measures that is adverse. What constitutes an adverse health effect may  
6 vary between populations. Some changes in healthy individuals may not be considered adverse while  
7 those of a similar type and magnitude are potentially adverse in more susceptible individuals.

8 For example, the extent to which changes in lung function are adverse has been discussed by the  
9 American Thoracic Society (ATS) in an official statement titled *What Constitutes an Adverse Health*  
10 *Effect of Air Pollution?* (American Thoracic Society, 2000). This statement updated the guidance for  
11 defining adverse respiratory health effects that had been published 15 years earlier  
12 (American Thoracic Society, 1985), taking into account new investigative approaches used to identify the  
13 effects of air pollution and reflecting concern for impacts of air pollution on specific susceptible groups.  
14 In the 2000 update, there was an increased focus on quality of life measures as indicators of adversity and  
15 a more specific consideration of population risk. Exposure to air pollution that increases the risk of an  
16 adverse effect to the entire population is viewed as adverse, even though it may not increase the risk of  
17 any identifiable individual to an unacceptable level. For example, a population of asthmatics could have a  
18 distribution of lung function such that no identifiable individual has a level associated with significant  
19 impairment. Exposure to air pollution could shift the distribution such that no identifiable individual  
20 experiences clinically relevant effects; this shift toward decreased lung function, however, would be  
21 considered adverse because individuals within the population would have diminished reserve function  
22 and, therefore, would be at increased risk to further environmental insult.

23 It is important to recognize that the more subtle health outcomes may be linked to health events  
24 that are clearly adverse. For example, air pollution has been shown to affect markers of transient  
25 myocardial ischemia such as ST-segment abnormalities and onset of exertional angina. In some cases,  
26 these effects are silent yet may still increase the risk of a number of cardiac events, including myocardial  
27 infarction and sudden death.

## 1.7. Summary

28 This first external review draft ISA is a concise evaluation and synthesis of the most  
29 policy-relevant science for reviewing the NAAQS, and it communicates critical science judgments  
30 relevant to the NAAQS review. It reviews the most policy-relevant evidence from health and

1 environmental effects studies, including mechanistic evidence from basic biological science. Annexes to  
2 the ISA provide additional details of the literature published since the last review. A framework for  
3 making critical judgments concerning causality was presented in this chapter. It relies on a widely  
4 accepted set of principles and standardized language to express evaluation of the evidence. This approach  
5 can bring rigor and clarity to the current and future assessments. This ISA should assist EPA and others,  
6 now and in the future, to accurately represent what is presently known—and what remains unknown—  
7 concerning the effects of CO on human health and public welfare.

# Chapter 2. Integrative Health Effects Overview

1 The subsequent chapters of this ISA present the most policy-relevant information related to the  
2 review of the NAAQS for CO, including a synthesis of the evidence presented in the 2000 CO AQCD  
3 with the results of more recent studies. This chapter integrates important findings from the disciplines  
4 evaluated in the assessment of the CO scientific literature, which include atmospheric sciences, ambient  
5 air data analyses, exposure assessment, dosimetry, and health effects research (animal toxicology,  
6 controlled human exposure, and epidemiologic studies). The EPA framework for causal determinations  
7 described in Chapter 1 has been applied to the body of evidence evaluated in this assessment in order to  
8 characterize the relationship between exposure to CO at relevant pollutant exposures and health effects.  
9 The EPA framework applied here employs a five-level hierarchy for causal determination:

- 10       ▪ Causal relationship
- 11       ▪ Likely to be a causal relationship
- 12       ▪ Suggestive of a causal relationship
- 13       ▪ Inadequate to infer a causal relationship
- 14       ▪ Suggestive of no causal relationship

15 This evaluation led to causal determinations for a range of health outcomes, and characterization of  
16 the magnitude of these responses, including responses in susceptible or vulnerable subpopulations, over a  
17 range of relevant concentrations. This integration of evidence also provides a basis for characterizing the  
18 concentration-response relationships of CO and adverse health outcomes for the U.S. population, given  
19 the current state of knowledge.

20 This chapter summarizes and integrates the recently available scientific evidence along with key  
21 findings and conclusions from the 2000 CO AQCD that best informs consideration of the policy-relevant  
22 questions that frame this assessment as presented in Chapter 1. Section 2.1 discusses the trends in ambient  
23 concentrations and sources of CO and provides a brief summary of factors influencing personal exposure  
24 to ambient CO. Section 2.2 summarizes CO dosimetry and pharmacokinetics and describes what is  
25 known regarding the hypoxic and non-hypoxic modes of action of CO. Section 2.3 integrates the evidence  
26 for studies that examined health effects related to short- and long-term exposure to CO and discusses  
27 important uncertainties identified in the interpretation of the scientific evidence. Finally, Section 2.4

1 presents the public health impacts associated with exposure to CO, and includes evidence of effects in  
2 potentially susceptible and vulnerable populations from CO exposure.

## 2.1. Ambient Concentrations, Sources, and Exposure

3 CO is formed by incomplete combustion of carbon-containing fuels and by photochemical  
4 reactions in the atmosphere. Nationally, on-road mobile sources constituted more than half of total CO  
5 emissions in 2002, or ~63 of ~109 million tons (MT) of total CO emissions, based on the most recent  
6 publicly available data from EPA's National Emissions Inventory (NEI). On-road mobile source emissions  
7 decreased by 16% since 1997, which is the most recent year covered in the 2000 CO AQCD. In  
8 metropolitan areas in the U.S., as much as 70-75% of all CO emissions result from on-road vehicle  
9 exhaust. The majority of these on-road CO emissions are derived from gasoline-powered vehicles since  
10 diesel vehicles emit relatively little CO. When emissions from incomplete combustion of fuels powering  
11 non-road mobile sources, such as farm and construction equipment, lawnmowers, boats, ships,  
12 snowmobiles, and aircraft are included, all mobile sources accounted for ~80% of total CO emissions in  
13 the U.S. in 2002. Other sources of CO include wildfires, controlled vegetation burning, residential  
14 biomass combustion, and industrial processes. While CO emissions from non-road mobile sources, fire,  
15 and industry have remained fairly constant, on-road mobile source CO emissions have decreased by  
16 roughly 5% per year since the early 1990s.

17 Significant reductions in ambient CO concentrations and in the number of NAAQS exceedances  
18 have been observed over the past 25 years, a continuation of trends documented in the 2000 CO AQCD.  
19 Nationwide ambient CO data from the EPA Air Quality System (AQS), for the years 2005-2007, show  
20 that the median 1-h daily maximum (max) concentration across the U.S. was 0.7 ppm; the mean was  
21 0.9 ppm; the 95th percentile was 2.4 ppm; and the 99th percentile was 3.8 ppm. The median 8-h daily  
22 max ambient CO concentration for the years 2005-2007 was 0.5 ppm; the mean was 0.7 ppm; the 95th  
23 percentile was 1.7 ppm; and the 99th percentile was 2.6 ppm. The current CO NAAQS are 35 ppm (1-h  
24 avg) and 9 ppm (8-h avg), not to be exceeded more than once per year (yr). During the years 2005-2007,  
25 1-h and 8-h CO concentrations did not exceed the NAAQS level more than once per year at any  
26 monitoring site. Moreover, in these 3 years, a 1-h avg concentration in excess of 35 ppm was reported  
27 only once, in 2007, in Ogden, UT (39 ppm), and there were only 7 reported 8-h avg values nationwide in  
28 excess of 9 ppm in all 3 years. Seasonally divided box plots of data from 2005-2007 compiled for  
29 spatially diverse urban metropolitan areas illustrate the tendency for higher median CO concentrations  
30 and wider variations in concentrations in the winter and fall compared with the spring and summer.

31 Policy-relevant background (PRB) concentrations of CO were computed for this assessment using  
32 data for the years 2005-2007 collected at 12 remote sites in the U.S. which are part of the National



1 Oceanic and Atmospheric Administration's (NOAA) Global Monitoring Division (GMD) and are not part  
2 of the EPA national regulatory network. The 3-year avg CO PRB averaged ~0.13 ppm in Alaska,  
3 ~0.10 ppm in Hawaii, and ~0.13 ppm over the contiguous U.S. (CONUS). (Note that the analysis for  
4 North American PRB in this assessment was made by segregating the three Alaska sites based on their  
5 high latitude and the two Hawaii sites based on their distance from the continent and then treating the  
6 remaining seven sites as representative of the CONUS PRB.)

7 A person's total personal exposure to CO is a combination of ambient and non-ambient exposures.  
8 Ambient exposure can be further broken down into direct exposure while outdoors and exposure to  
9 ambient CO which has infiltrated into buildings and vehicles. Several studies have shown that in the  
10 absence of indoor sources such as ETS or gas stoves, CO concentrations are highest in automobiles and  
11 near highly traveled roadways. Specific concentrations vary as a function of vehicle ventilation rate and  
12 CO emissions source strength and are related to, for example, on-road vehicle numbers and speed. In-  
13 vehicle CO concentrations are typically reported to be between 2 and 5 times higher than ambient  
14 concentrations measured at the roadside, but have been reported to be as much as 25 times higher. A  
15 portion of these high concentrations comes from the vehicle's own engine emissions entrained with the  
16 ambient concentrations from the roadway, although the exact fraction of this self-pollution in the in-  
17 vehicle total varies with body integrity of the vehicle and is not known with certainty. Concentrations of  
18 CO, like those of other on-road vehicle pollutants, often display an inverse relationship with distance  
19 from the roadway. For CO, this fall-off with distance can be very sharp up to ~70 meters (m), beyond  
20 which the concentration increment from roadway emissions is largely indistinguishable from local  
21 ambient concentrations. The siting and location instructions for federal regulatory CO monitors explicitly  
22 recognize the need to measure at distances close to mobile sources with a requirement for monitored areas  
23 that at least one monitor be sited to measure max concentration. This is often accomplished with a  
24 monitor situated at the CFR-defined microscale of 2-10 m from the roadway; these microscale monitors  
25 also have sample inlets mounted at  $3 \pm 0.5$  m above ground level (AGL), unlike those at larger scale  
26 distances whose inlet heights can vary between 2 and 15 m AGL. In 2007, there were at least 70 CO  
27 monitors described as microscale reporting to AQS.

28 Although the correlations across CO monitors sited to sample at different scales can be greater than  
29 0.8 in some areas, they also can vary widely from within and between cities across the U.S. as a function  
30 of natural and urban topography, meteorology, and strength and proximity to sources. At the same time,  
31 personal exposure monitoring captures both ambient and non-ambient CO concentrations. Because  
32 non-ambient CO exposure is not expected to be correlated with ambient CO concentrations at the  
33 monitor, the non-ambient contribution to total personal CO exposure complicates interpretation of health  
34 effects relationships observed in epidemiologic studies. For the general U.S. population, exposure error  
35 analysis for epidemiologic studies, as summarized in Chapter 3, indicates that fixed-site measured  
36 ambient CO concentration is generally a good indicator of ambient exposure to CO.

## 2.2. Dosimetry, Pharmacokinetics, and Mode of Action

### 2.2.1. Dosimetry and Pharmacokinetics

1 CO elicits various health effects by binding to and altering the function of a number of heme-  
2 containing molecules, mainly hemoglobin (Hb). The formation of COHb reduces the oxygen  
3 (O<sub>2</sub>)-carrying capacity of blood and impairs the release of O<sub>2</sub> from oxyhemoglobin (O<sub>2</sub>Hb) to the tissues.  
4 The 2000 CO AQCD has a detailed description of the well-established Coburn-Forster-Kane (CFK)  
5 equation, which has been used for many years to model COHb formation. More recent work since then  
6 has developed models that include myoglobin (Mb) and extravascular storage compartments, as well as  
7 other dynamics of physiology relevant to CO uptake and elimination. These models have indicated that  
8 CO has a biphasic elimination curve, due to initial washout from the blood followed by a slower flux  
9 from the tissues. The flow of CO between the blood and alveolar air or tissues is controlled by diffusion  
10 down the CO concentration gradient. The uptake of CO is governed not only by this CO pressure  
11 differential, but also by physiological parameters, such as minute ventilation and lung diffusing capacity,  
12 that can, in turn, be affected by factors such as exercise, age, and medical conditions (e.g., obstructive  
13 lung disease). Susceptible populations, including health compromised individuals and developing fetuses,  
14 are at a greater risk from COHb induced health effects due to altered CO kinetics, compromised  
15 cardiopulmonary processes, and increased baseline hypoxia levels. Altitude also may have a substantial  
16 effect on the kinetics of COHb formation, especially for visitors to high altitude areas. Compensatory  
17 mechanisms, such as increased cardiac output, combat the decrease in barometric pressure. Altitude also  
18 increases the endogenous production of CO through upregulation of heme oxygenase (HO). CO is  
19 considered a second messenger and is endogenously produced from the catabolism of heme proteins by  
20 enzymes such as HO-1 (the inducible form of heme oxygenase) and through endogenous lipid  
21 peroxidation. Finally, CO is removed from the body by expiration and oxidation to CO<sub>2</sub>.

### 2.2.2. Mode of Action

22 The diverse effects of CO are dependent upon dose, duration of exposure, and the cell types and  
23 tissues involved. Responses to CO are not necessarily due to a single process and may instead be  
24 mediated by a combination of effects including COHb-mediated hypoxic stress and mechanisms unrelated  
25 to tissue hypoxia including free radical production and the initiation of cell signaling. However binding of  
26 CO to reduced iron in heme proteins with subsequent alteration of heme protein function is the underlying  
27 mechanism for both the hypoxic and non-hypoxic biological responses to CO.

1 As discussed in the 2000 CO AQCD, the most well-known pathophysiological effect of CO is  
2 tissue hypoxia caused by binding of CO to Hb. Not only does the formation of COHb reduce the  
3 O<sub>2</sub>-carrying capacity of blood, but it also impairs the release of O<sub>2</sub> from O<sub>2</sub>Hb. Compensatory alterations  
4 in hemodynamics, such as vasodilation and increased cardiac output, protect from tissue hypoxia. At  
5 ambient CO concentrations, these compensatory changes are slight and likely tolerable in people with a  
6 healthy cardiovascular system. However, people with cardiovascular detriments may be unable to endure  
7 these small changes in hemodynamics which may lead to the presentation of health effects as described in  
8 Sections 5.2.1 and 5.2.2. Binding of CO to Mb, as discussed in the 2000 CO AQCD and in Section  
9 4.3.2.3, can also impair the delivery of O<sub>2</sub> to tissues. Mb has a high affinity for CO; however,  
10 physiological effects are seen only after high dose exposures to CO, resulting in COMb concentrations far  
11 above baseline levels.

12 Non-hypoxic mechanisms underlying the biological effects of CO have been the subject of a  
13 substantial amount of recent research since the 2000 CO AQCD. Most of these mechanisms are related to  
14 CO's ability to bind heme-containing proteins other than Hb and Mb. These mechanisms, which may be  
15 interrelated, include alteration in nitric oxide (NO) signaling, inhibition of cytochrome *c* oxidase, heme  
16 loss from proteins, disruption of iron homeostasis, and alteration in cellular redox status. CO is a  
17 ubiquitous cell signaling molecule and the physiological functions of HO-derived CO are numerous. The  
18 endogenous generation and release of CO from HO-1 and HO-2 is tightly controlled, as is any  
19 homeostatic process. Thus, exogenously-applied CO has the capacity to disrupt myriad heme-based  
20 signaling pathways due to its nonspecific nature. Only a limited amount of information is available  
21 regarding the impact of exogenous CO on tissue and cellular levels of CO. However recent animal studies  
22 demonstrated increased tissue CO levels and biological responses following exposure to 50 ppm CO.  
23 Whether or not environmentally relevant exposures to CO can affect endogenous CO signaling pathways  
24 and lead to adverse health effects is an open question for which there are no definitive answers at this  
25 time.

## 2.3. Health Effects

**Table 2-1. Causal determinations for health effects outcomes.**

Outcome Category	Exposure Period	Causality Determination
Cardiovascular morbidity	Short-term	Likely to be a causal relationship
Central nervous system effects	Short- and long-term	Suggestive of a causal relationship
Birth outcomes and Developmental effects	Long-term	Suggestive of a causal relationship
Respiratory morbidity	Short-term	Suggestive of a causal relationship
	Long-term	Inadequate to infer a causal relationship
Mortality	Short-term	Suggestive of a causal relationship
	Long-term	Suggestive of no causal relationship

### 2.3.1. Cardiovascular Morbidity

1           The most compelling evidence of a CO-induced effect on the cardiovascular system at COHb  
2 levels relevant to the current NAAQS comes from a series of controlled human exposure studies among  
3 individuals with CAD (see Section 5.2). These studies, described in the 1991 and 2000 CO AQCDs,  
4 demonstrate consistent decreases in the time to onset of exercise-induced angina and ST-segment changes  
5 following CO exposures resulting in COHb levels of 3-6%, with one multicenter study reporting similar  
6 effects at COHb levels as low as 2.4%. No human clinical studies have evaluated the effect of controlled  
7 exposures to CO resulting in COHb levels lower than 2.4%. Human clinical studies published since the  
8 2000 CO AQCD have reported no association between CO and ST-segment changes or arrhythmia;  
9 however, none of these studies included individuals with diagnoses of heart disease.

10           While the exact physiological significance of the observed ST-segment changes among individuals  
11 with CAD is unclear, ST-segment depression is a known indicator of myocardial ischemia. It is also  
12 important to note that the individuals with CAD who participated in these controlled exposure studies  
13 were not representative of the most sensitive individuals in the population. In fact, the most sensitive  
14 individuals may respond to levels of COHb lower than 2.4%. Variability in activity patterns and severity  
15 of disease among individuals with CAD is likely to influence the critical level of COHb which leads to  
16 adverse cardiovascular effects.

17           The degree of ambient CO exposure which leads to attainment of critical levels of COHb will also  
18 vary between individuals. First of all, endogenous CO production varies as described in Section 4.5, but  
19 generally results in less than 1% COHb. Secondly, nonambient exposures to CO, such as exposure to  
20 ETS, can increase COHb above baseline levels. Ambient exposures will result in an additive increase in

1 COHb. Using mathematical modeling to predict changes in COHb in healthy inactive adults (Quantitative  
2 Circulatory Physiology [QCP] model, Section 4.2.3), it can be estimated that exposure to 35 ppm CO for  
3 1 h results in an increase of 0.6% COHb over baseline and exposure to 9 ppm CO for 8 h results in an  
4 increase of 0.8% COHb over baseline. Furthermore, 24-h exposure to 3 ppm CO results in an increase of  
5 0.4% COHb above baseline which can also be obtained following 1-h exposure to 30 ppm CO.  
6 Consequently, exposure to CO at concentrations relevant to the NAAQS has the potential to increase  
7 COHb to levels associated with adverse cardiovascular health effects in some individuals.

8 Findings of controlled human exposure studies are coherent with findings of recent epidemiologic  
9 studies conducted since the 2000 CO AQCD, which observed associations between ambient CO  
10 concentration and emergency department (ED) visits and hospital admissions for ischemic heart disease  
11 (IHD), congestive heart failure (CHF) and all-cause cardiovascular disease (CVD). All but one of these  
12 epidemiologic studies were conducted in locations where the entire distribution of CO concentrations  
13 were at or below the level of the current NAAQS, with mean 24-h avg concentrations ranging from  
14 0.5 ppm (Montreal, Canada) to 9.4 ppm (Tehran, Iran) (Table 5-7). A single study reported a negative  
15 association between CO concentration and hospital admissions and ED visits for IHD among all ages; all  
16 other associations were positive, with increases in hospital admissions and ED visits for IHD between  
17 0.2% and 19.8% per standardized increase in CO concentration (Figure 5-1). These recent studies build  
18 upon the conclusions of the 2000 CO AQCD that short-term variations in ambient CO concentrations are  
19 associated with daily hospital admissions for heart disease.

20 These health outcomes are consistent with a role for CO in limiting O<sub>2</sub> availability (i.e., hypoxic  
21 mechanisms) in individuals with CAD. However, recent toxicological studies suggested that CO may also  
22 act through non-hypoxic mechanisms by disrupting cellular signaling. Studies in healthy animals  
23 demonstrated oxidative injury and inflammation in response to 50-100 ppm CO while studies in disease  
24 models demonstrate effects on heart rhythm and exacerbation of cardiomyopathy and vascular remodeling  
25 in response to 35-50 ppm CO. Furthermore, in utero exposure to 150 ppm CO alters postnatal  
26 electrophysiological maturation in rat cardiomyocytes. Further investigations will be useful in determining  
27 the importance of non-hypoxic mechanisms following environmentally-relevant CO exposures. Taken  
28 together, the evidence from epidemiologic, human clinical, and toxicological studies is **sufficient to**  
29 **conclude that a causal relationship is likely to exist between relevant short-term CO exposures and**  
30 **cardiovascular morbidity.**

### 2.3.2. Central Nervous System Effects

31 Exposure to high levels of CO has long been known to adversely affect CNS function, with  
32 symptoms following acute CO poisoning including headache, dizziness, cognitive difficulties,

1 disorientation, and coma. However, the relationship between ambient levels of CO and neurological  
2 function is less clear and has not been evaluated in epidemiologic studies. Studies of controlled human  
3 exposures to CO discussed in the 2000 CO AQCD reported inconsistent neural and behavioral effects  
4 following exposures resulting in COHb levels of 5-20%. No new human clinical studies have evaluated  
5 central nervous system or behavioral effects of exposure to CO. At ambient-level exposures, healthy  
6 adults may be protected against CO-induced neurological impairment owing to compensatory responses  
7 including increased cardiac output and cerebral blood flow. However, these compensatory mechanisms  
8 are likely impaired among certain potentially susceptible groups including individuals with reduced  
9 cardiovascular function.

10 Toxicological studies that were not discussed in the 2000 CO AQCD employed rodent models to  
11 show that low level CO exposure during the in utero period can adversely affect adult outcomes including  
12 behavior, neuronal myelination, neurotransmitter levels or function, and the auditory system (discussed in  
13 Section 5.3). In utero CO exposure, including both intermittent and continuous exposure, has been shown  
14 to impair multiple behavioral outcomes in offspring including active avoidance behavior (150 ppm CO),  
15 non-spatial memory (75 and 150 ppm CO), spatial learning (endogenous CO inhibition), homing behavior  
16 (150 ppm CO), locomotor movement (150 ppm CO), and negative geotaxis (125 and 150 ppm). In two  
17 separate studies, in utero CO exposure (75 and 150 ppm) was associated with significant myelination  
18 decrements without associated changes in motor activity in adult animals. Multiple studies demonstrated  
19 that in utero CO exposure affected glutamatergic, cholinergic, catecholaminergic, and dopaminergic  
20 neurotransmitter levels or transmission in exposed male rodents. Possible or demonstrated adverse  
21 outcomes from the CO-mediated aberrant neurotransmitter levels or transmission include respiratory  
22 dysfunction (200 ppm CO), impaired sexual behavior (150 ppm CO), and an adverse response to  
23 hyperthermic insults resulting in neuronal damage (200 ppm). Finally, perinatal CO exposure has been  
24 shown to affect the developing auditory system of rodents, inducing permanent changes into adulthood.  
25 This is manifested with atrophy of cochlear cells innervating the inner hair cells (25 ppm CO), decreased  
26 immunostaining associated with impaired neuronal activation (12.5 ppm CO), impaired myelination of  
27 auditory associated nerves (25 ppm CO), decreased energy production in the sensory cell organ of the  
28 inner ear or the organ of corti (25 ppm CO), some of which is mechanistically proposed to be mediated by  
29 reactive oxygen species (ROS) (25 ppm CO). Functional tests of the auditory system of neonatally, low  
30 level CO-exposed rodents, using otoacoustic emissions (OAE) testing (50 ppm CO) and amplitude  
31 measurements of the 8th cranial nerve action potential (12, 25, 50, 100 ppm CO), revealed decrements in  
32 auditory function at PND22 and permanent changes into adulthood using AP testing (50 ppm CO).  
33 Together, these animal studies demonstrate that in utero exposure to CO can adversely affect adult  
34 behavior, neuronal myelination, neurotransmission, and the auditory system in adult male rodents.  
35 Considering the combined evidence from controlled human exposure and toxicological studies, the

1 evidence is **suggestive of a causal relationship between relevant short- and long-term CO exposures**  
2 **and central nervous system effects.**

### 2.3.3. Birth Outcomes and Developmental Effects

3 The most compelling evidence for a CO-induced effect on birth and developmental outcomes is for  
4 preterm birth (PTB) and cardiac birth defects. These outcomes were not addressed in the 2000 CO  
5 AQCD, which included only two studies that examined the effect of ambient CO on low birth weight  
6 (LBW). Since then, a number of studies have been conducted looking at varied outcomes, including PTB,  
7 birth defects, fetal growth (including LBW), and infant mortality.

8 There is limited epidemiologic evidence that CO during early pregnancy (e.g., first month and first  
9 trimester) is associated with an increased risk of PTB. The only U.S. studies to investigate the PTB  
10 outcome were conducted in California, and these reported consistent results whereby all studies reported a  
11 significant association with CO exposure during early pregnancy, and exposures were assigned from  
12 monitors within close proximity of the mother's residential address. Additional studies conducted outside  
13 of the U.S. provide supportive, though less consistent, evidence of an association between CO  
14 concentration and PTB.

15 Very few epidemiologic studies have examined the effects of CO on birth defects. Two of these  
16 studies found maternal exposure to CO to be associated with an increased risk of cardiac birth defects.  
17 This insult to the heart is coherent with results of human clinical studies demonstrating the heart as a  
18 target for CO effects (Section 5.2). Animal toxicological studies provide additional evidence for such an  
19 insult to the heart, and reported transient cardiomegaly at birth after continuous in utero CO exposure (60,  
20 125, 250 and 500 ppm CO), delayed myocardial electrophysiological maturation (150 ppm CO), or  
21 systemic splenic immunocompromise (75 or 150 ppm CO). Toxicological studies have also shown that  
22 exogenous continuous in utero CO exposure (250 ppm) induced teratogenicity in rodent offspring in a  
23 dose-dependent manner that was further exacerbated by dietary protein restriction (65 ppm CO) or zinc  
24 depletion (500 ppm CO). Toxicological studies of exogenous CO exposure over the duration of gestation  
25 have shown skeletal alterations (7 h/day, CO 250 ppm) or limb deformities (24 h/day, CO 180 ppm) in  
26 prenatally exposed offspring.

27 There is evidence of ambient CO exposure during pregnancy having a negative effect on fetal  
28 growth in epidemiologic studies. In general, the reviewed studies, summarized in Figures 5-7 through 5-9,  
29 reported small reductions in birth weight (ranging ~5-20 g). Several studies examined various  
30 combinations of birth weight, LBW, and small for gestational age (SGA)/intrauterine growth restriction  
31 (IUGR) and inconsistent results are reported across these metrics. It should be noted that having a  
32 measurable, even if small, change in a population is different than having an effect on a subset of

1 susceptible births and increasing the risk of IUGR/LBW/SGA. It is difficult to conclude if CO is related  
2 to a small change in birth weight in all births across the population, or a marked effect in some subset of  
3 births.

4 In general, there is limited epidemiologic evidence that CO is associated with an increased risk of  
5 infant mortality during the neonatal or post-neonatal periods. In support of this limited evidence, animal  
6 toxicological studies do provide some evidence that exogenous CO exposure to pups in utero significantly  
7 increased postnatal mortality (7 h/day and 24 h/day, 250 ppm CO; 24 h/day, 90 or 180 ppm CO) and  
8 prenatal mortality (7 h/day, 250 ppm CO).

9 Evidence exists for additional developmental outcomes which have been examined in toxicological  
10 studies, but not epidemiologic or human clinical studies, including behavioral abnormalities, learning and  
11 memory deficits, locomotor effects, neurotransmitter changes, and changes in the auditory system.  
12 Structural aberrations of the cochlea involving neuronal activation (12.5, 25 and 50 ppm CO) and  
13 auditory related nerves (25 ppm CO) were seen in pups after in neonatal CO exposure. Auditory  
14 functional testing using OAE (50 ppm CO) and ABR (12, 25, 50, 100 ppm CO) on rodents exposed  
15 perinatally to CO showed that CO-exposed neonates had auditory decrements at PND22 (OAE and ABR)  
16 and permanent changes into adulthood with ABR (50 ppm CO).

17 Overall, there is limited, though positive, epidemiologic evidence for a CO-induced effect on PTB  
18 and birth defects, and weak evidence for a decrease in birth weight, other measures of fetal growth, and  
19 infant mortality. Animal toxicological studies provide support and coherence for these effects. Both  
20 hypoxic and non-hypoxic mechanisms that could lead to such effects have been proposed in the  
21 toxicological literature (Section 5.1), though a clear understanding of the mechanisms underlying  
22 reproductive and developmental effects is still lacking. Taking into consideration the positive evidence for  
23 some birth and developmental outcomes from epidemiologic studies and the resulting coherence for these  
24 associations in animal toxicological studies, the evidence is **suggestive of a causal relationship**  
25 **between long-term exposures to relevant CO concentrations and developmental effects and birth**  
26 **outcomes.**

### 2.3.4. Respiratory Morbidity

27 New epidemiologic studies, supported by the body of literature summarized in the 2000 CO  
28 AQCD, provide evidence of positive associations between short-term exposure to CO and respiratory-  
29 related outcomes including pulmonary function, respiratory symptoms, medication use, hospital  
30 admissions, and ED visits (discussed in Section 5.5). However, the interpretation of the results from  
31 epidemiologic studies is difficult due to the lack of extensive copollutant analyses along with the  
32 moderate to high correlation between CO and other combustion/traffic generated pollutants. To date the



1 majority of the literature has not extensively examined the association between CO and respiratory  
2 morbidity due to studies focusing primarily on effects associated with exposure to other criteria  
3 pollutants, namely PM and O<sub>3</sub>. This has contributed to the inability to disentangle the effects attributed to  
4 CO from the larger complex air pollution mix. In addition, uncertainty as to a biological mechanism to  
5 explain the respiratory-related effects observed in the epidemiologic literature further complicates the  
6 interpretation of these results, especially considering the low ambient CO concentrations reported (24-h  
7 avg: 0.35–2.1 ppm). However, animal toxicological studies do provide some evidence that short-term  
8 exposure to CO (50-100 ppm) can cause oxidative injury and inflammation and alter pulmonary vascular  
9 remodeling. Human clinical studies have not extensively examined the effect of short-term exposure to  
10 CO on respiratory morbidity, specifically pulmonary function. The limited number of clinical studies that  
11 have been conducted prior to and since the 2000 CO AQCD provide very little evidence of any adverse  
12 effect of CO on the respiratory system at COHb levels <10%. Although human clinical studies have not  
13 provided evidence to support CO-related respiratory health effects, the epidemiologic studies that  
14 examined the effects of short-term exposure to CO and lung-related outcomes show positive associations  
15 and animal toxicological studies demonstrate the potential for an underlying biological mechanism, which  
16 together provide evidence that is **suggestive of a causal relationship between short-term exposure to**  
17 **relevant CO concentrations and respiratory morbidity.**

18 Currently, only a few studies have been conducted that examine the association between long-term  
19 exposure to CO and respiratory morbidity. Although some studies did observe associations between long-  
20 term exposure to CO and respiratory health outcomes key uncertainties still exist. These uncertainties  
21 include: the lack of replication and validation studies to evaluate new methodologies (i.e.,  
22 Deletion/Substitution/Addition (DSA) algorithm) that have been used to examine the association between  
23 long-term exposure to CO and respiratory health effects; whether the respiratory health effects observed  
24 in response to long-term exposure to CO can be explained by the proposed biological mechanisms; and  
25 the lack of co-pollutant analyses to disentangle the respiratory effects associated with CO due to its high  
26 correlation with NO<sub>2</sub> and other combustion-related pollutants. Overall, the evidence available is  
27 **inadequate to conclude that a causal relationship exists between long-term exposure to relevant**  
28 **CO concentrations and respiratory morbidity.**

### 2.3.5. Mortality

29 Among the gaseous pollutants examined in time-series mortality studies, CO is the least frequently  
30 studied criteria air pollutant. Because CO was mostly treated as a potential confounder in these studies,  
31 the information available regarding the nature of the association between short-term exposure to CO and  
32 mortality is limited compared to the other pollutants. However, the recently available multi-city studies,

1 which consist of larger sample sizes, and single-city studies generally confirmed the findings reported in  
2 the 2000 CO AQCD (see Section 5.6).

3 The multi-city studies report comparable CO mortality risk estimates for total (non-accidental)  
4 mortality with the APHEA2 European multi-city study showing slightly higher estimates for  
5 cardiovascular mortality in single-pollutant models. However, when examining potential confounding by  
6 copollutants these studies consistently showed that CO mortality risk estimates were reduced when NO<sub>2</sub>  
7 was included in the model, but this observation may not be “confounding” in the usual sense in that NO<sub>2</sub>  
8 may also be a surrogate marker of other pollutants or pollution sources (i.e., traffic).

9 Only one of the multi-city studies focused specifically on the CO-mortality association (the  
10 APHEA study), examining: (1) model sensitivity; (2) the CO-mortality concentration-response (C-R)  
11 relationship; and (3) potential effect modifiers of CO mortality risk estimates. The APHEA2 study  
12 performed a sensitivity analysis, which indicated an approximate 50 - 80% difference in CO risk  
13 estimates from a reasonable range of alternative models. In addition, the study examined the CO-mortality  
14 C-R relationship through a grid search of varying threshold points, and found only weak evidence of a CO  
15 threshold at 0.5 mg/m<sup>3</sup> (0.43 ppm), but this result is complicated by the lowest 10% of the CO distribution  
16 for seven of the 19 cities examined being at or above 2 mg/m<sup>3</sup> (1.74 ppm). The examination of a variety  
17 of city-specific variables to identify potential effect modifiers of the CO-mortality relationship found that  
18 geographic region explained most of the heterogeneity in CO mortality risk estimates with the  
19 CO-mortality associations being stronger in western and southern European cities than eastern cities. A  
20 similar pattern has been reported for black smoke (BS) and SO<sub>2</sub> in previous APHEA studies, but the  
21 geographic variability observed does not provide specific information which could be used to evaluate the  
22 CO-mortality association.

23 The results from the single-city studies evaluated are generally consistent with the multi-city  
24 studies in that some evidence of a positive association was found for mortality upon short-term exposure  
25 to CO. However, the CO-mortality associations were often, but not always, attenuated when other  
26 copollutants were included in the regression models. In addition, limited evidence was available to  
27 identify cause-specific mortality outcomes (e.g., cardiovascular causes of death) associated with short-  
28 term exposure to CO.

29 The new multi- and single-city studies evaluated provide evidence that an association between  
30 short-term exposure to CO and mortality exists, but limited evidence is available to evaluate cause-  
31 specific mortality outcomes associated with CO exposure and it is unclear if CO is acting alone or as a  
32 surrogate for other combustion-related pollutants. In addition, the results also underscore the limitation of  
33 current analytical methods to disentangle the health effects associated with one pollutant in the complex  
34 air pollution mixture. Overall, the epidemiologic evidence is **suggestive of a causal relationship**  
35 **between short-term exposure to relevant CO concentrations and mortality.**

1           The evaluation of new epidemiologic studies conducted since the 2000 CO AQCD that investigated  
2 the association between long-term exposure to CO and mortality consistently found null or negative  
3 mortality risk estimates. No such studies were discussed in the 2000 CO AQCD. The re-analysis of the  
4 American Cancer Society (ACS) data by Jerrett et al. (2003) found no association between long-term  
5 exposure to CO and mortality. Similar results were obtained in an updated analysis conducted by the  
6 original ACS investigators when using earlier (1980) CO data, but negative associations were found when  
7 using more recent (1982-1998) data. The Women’s Health Initiative (WHI) Study also found no  
8 association between CO and CVD events (including mortality) using the data from recent years  
9 (1994-1998), while the series of Veterans Cohort studies found no association or a negative association  
10 between mean annual 95th percentile of hourly CO values and mortality. In addition, a cross-sectional  
11 study of U.S. counties reported results that are generally consistent with the cohort studies: positive  
12 associations between long-term exposure to PM<sub>2.5</sub> and sulfate (SO<sub>4</sub><sup>2-</sup>) and mortality, and generally  
13 negative associations with CO. Overall, the consistent null and negative associations observed across  
14 epidemiologic studies which included cohort populations encompassing potentially susceptible  
15 subpopulations (i.e., post-menopausal women and hypertensive men) combined with (1) the lack of  
16 evidence for respiratory and cardiovascular morbidity outcomes following long-term exposure to CO; and  
17 (2) the absence of a proposed mechanism to explain the progression to mortality following long-term  
18 exposure to CO provide supportive evidence that is **suggestive of no causal relationship between long-**  
19 **term exposure to relevant CO concentrations and mortality.**

## 2.4. Public Health Impacts

### 2.4.1. Concentration-Response Relationship

20           Currently, very limited information is available in the human clinical and epidemiologic literature  
21 regarding the CO C-R relationship and the potential presence of a CO threshold. Two human clinical  
22 studies described in the 2000 CO AQCD have evaluated the C-R relationship between CO and onset of  
23 exercise-induced angina among individuals with CAD, but at the high end of CO concentrations (i.e., CO  
24 levels above the current NAAQS). Anderson et al. (1973) exposed 10 adult men with stable angina for 4 h  
25 to CO concentrations of 50 and 100 ppm, which resulted in average COHb levels of 2.9% and 4.5%,  
26 respectively. Both exposures significantly decreased the time to onset of exercise-induced angina relative  
27 to room air control (1.6% COHb). However, there was no difference in response between the two  
28 exposure concentrations of CO. In a much larger study, 63 adults with stable angina were exposed for 1 h  
29 to two concentrations of CO (average exposure concentrations of 117 and 253 ppm) resulting in average

1 pre-exercise COHb levels of 2.4% and 4.7% (Allred et al., 1989). Relative to control (average COHb  
2 0.7%), COHb levels of 2.4% and 4.7% were observed to decrease the time to onset of angina by 4.2% (p  
3 = 0.054) and 7.1% (p = 0.004), respectively. In addition, these investigators reported a statistically  
4 significant decrease in the time to exercise-induced ST-segment depression with increasing COHb levels.  
5 These findings provide some evidence of a significant C-R relationship at COHb concentrations between  
6 2.4 and 4.7%. However, the human clinical literature has yet to evaluate the C-R relationship at lower CO  
7 concentrations or COHb levels.

8 One study in the epidemiologic literature attempted to examine the C-R relationship at the low end  
9 of CO concentrations through a threshold analysis. Samoli et al. (2007) in their examination of the  
10 association between short-term exposure to CO and mortality conducted an ancillary analysis to examine  
11 the potential presence of a CO threshold. In this analysis the authors compared city-specific models to the  
12 threshold model, which consisted of thresholds at 0.5 mg/m<sup>3</sup> (0.43 ppm) increments. Samoli et al. (2007)  
13 then computed the deviance between the two models and summed the deviances for a given threshold  
14 over all cities. While the minimum deviance suggested a potential threshold of 0.43 ppm (the lowest  
15 threshold examined), the comparison with the linear no-threshold model indicated weak evidence (p-  
16 value >0.9) for a threshold. However, determining the presence of a threshold at the very low range of CO  
17 concentrations (i.e., at 0.43 ppm) in this data set is challenging, because, in seven of the 19 European  
18 cities examined, the lowest 10% of the CO distribution was at or above 2 mg/m<sup>3</sup> (1.74 ppm). By only  
19 using the 12 cities in the analysis that had minimum CO concentrations approaching 0.5 mg/m<sup>3</sup>  
20 (0.43 ppm), a limited number of observations were examined around the threshold of interest, which  
21 subsequently contributed to the inability to draw conclusions regarding the potential presence of a  
22 threshold with any certainty.

## 2.4.2. Potentially Susceptible and Vulnerable Subpopulations

23 The examination of both susceptible and vulnerable subpopulations to CO exposure allows for the  
24 NAAQS to provide an adequate margin of safety for both the general population and sensitive  
25 subpopulations (see Section 5.7 for a more detailed discussion). During the evaluation of the CO  
26 literature, numerous studies were identified that examined whether underlying factors increased the  
27 susceptibility or vulnerability of an individual to CO-related health effects. In this ISA, a susceptible  
28 subpopulation is defined as those individuals with intrinsic biological characteristics that might exhibit an  
29 adverse health effect to a pollutant at concentrations lower than those needed to elicit the same response  
30 in the general population or those individuals that might elicit a more adverse health effect at the same  
31 concentration. A vulnerable subpopulation is defined as those individuals with external, nonbiological  
32 factors that increase the risk of adverse health effects, such as differential exposure or living at altitude.

1           The most important susceptibility characteristic for increased risk due to CO exposure is CAD.  
2 Individuals with heart disease may be at a greater risk from CO exposure since they may already have  
3 compromised O<sub>2</sub> delivery. CO is notable among air pollutants because it is especially harmful in  
4 individuals with impaired cardiovascular systems. Persons with a normal cardiovascular system can  
5 tolerate substantial concentrations of CO, if they vasodilate in response to the hypoxemia produced by  
6 CO. In contrast, individuals unable to vasodilate in response to CO exposure may show evidence of  
7 ischemia at low concentrations of COHb. Many of the controlled human exposure studies have focused  
8 on individuals with IHD. Other medical conditions that confer increased susceptibility include obstructive  
9 lung disease, which impairs airflow to the lungs, and conditions such as anemia that alter the blood O<sub>2</sub>  
10 carrying capacity or content and result in a greater risk from COHb induced hypoxia and decreased tissue  
11 O<sub>2</sub> delivery. Age is important for susceptibility in older adults, who have an increased elimination time for  
12 COHb. Newborns can also be considered a susceptible subpopulation due to critical phases of  
13 development and differences between fetal and maternal CO pharmacokinetics in utero.

14           Subpopulations considered vulnerable to CO-related health effects include people who spend a  
15 considerable amount of time in or near traffic, live at high altitude, exercise and/or use medications which  
16 may increase endogenous CO production. Individuals that spend a substantial amount of time on or near  
17 heavily traveled roadways, such as commuters and those living or working near freeways, are likely to  
18 experience elevated CO concentrations and therefore constitute a potentially vulnerable subpopulation  
19 due to differential exposure. Vulnerable subpopulations also include individuals at high altitude, who  
20 undergo physiologic changes that favor increased CO uptake and COHb formation. Exercising  
21 individuals are also considered vulnerable because exercise facilitates CO uptake and transport by  
22 increasing gas exchange efficiency and the COHb elimination rate decreases with physical activity.  
23 Individuals who use certain medications which may increase endogenous CO production and the baseline  
24 level of COHb are also considered vulnerable to CO-related health effects.

# Chapter 3. Source to Exposure

## 3.1. Introduction

1           This chapter contains basic information about concepts and findings in atmospheric sciences and  
2 exposure assessment pertaining to CO levels relevant for establishing a foundation for the detailed  
3 discussions of health effects data in subsequent chapters. Section 3.2 provides an overview of the sources  
4 of CO and examples of their spatial distribution. Atmospheric chemistry involved in the production and  
5 removal of CO by oxidation processes is discussed in Section 3.3. Descriptions of CO measurement  
6 methods and monitor siting requirements and locations are presented in Section 3.4. Data for ambient CO  
7 concentrations are characterized in Section 3.5. Policy relevant background concentrations of CO,  
8 i.e., those concentrations defined to result from uncontrollable emissions, are also presented in Section  
9 3.5. Finally, factors related to human exposure to CO are discussed in Section 3.6.

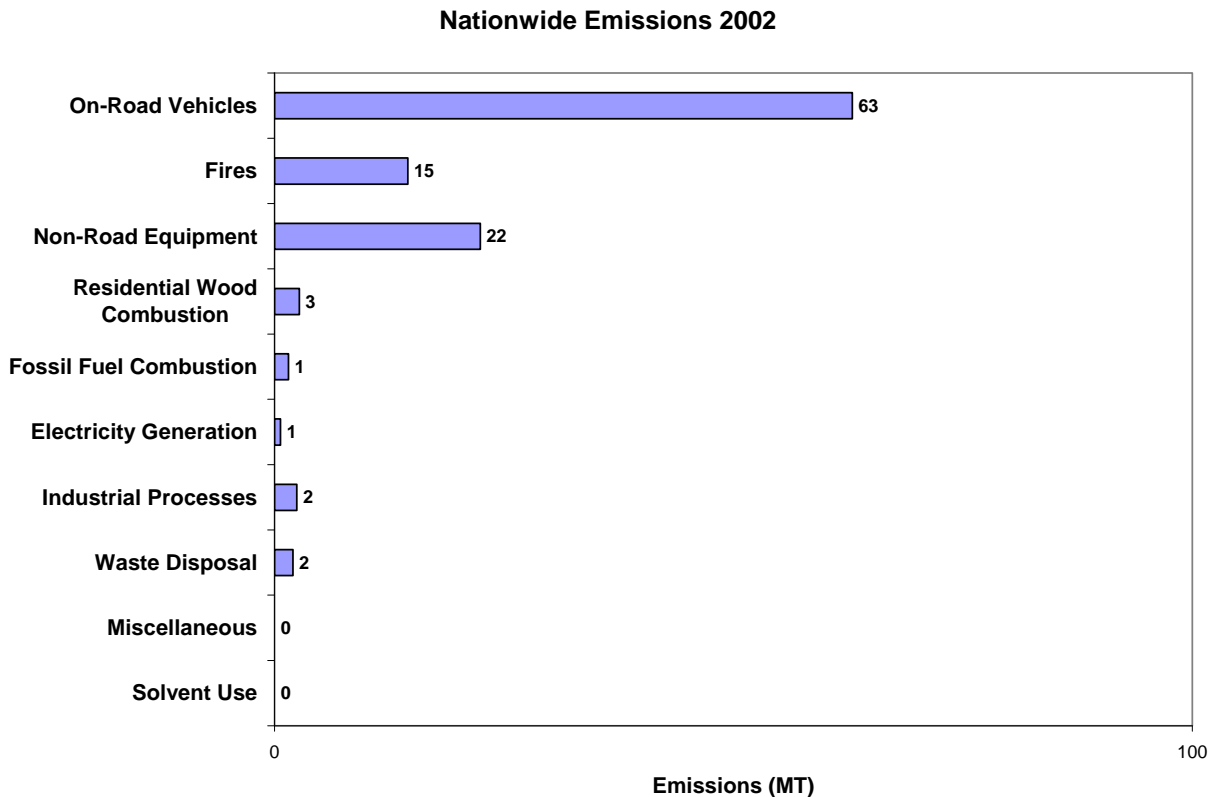
## 3.2. Sources and Emissions of CO

10           CO is a colorless, odorless, tasteless gas consisting of one carbon (C) atom covalently bonded to  
11 one O<sub>2</sub> atom; its molar mass is 28.0101 g/mol. CO is formed primarily by incomplete combustion of  
12 carbon-containing fuels and photochemical reactions in the atmosphere. In general, any increase in fuel  
13 O<sub>2</sub> content, burn temperature, or mixing time in the combustion zone will tend to decrease production of  
14 CO relative to CO<sub>2</sub>; hence CO emissions from large fossil-fueled power plants are typically very low  
15 since the boilers at these plants are tuned for highly efficient combustion with the lowest possible fuel  
16 consumption. Internal combustion engines used in mobile sources, by contrast, have widely varying  
17 operating conditions and, thus, inherently higher and varying CO formation.

18           Figure 3-1 lists CO emissions totals in tons segregated by individual source sectors in the U.S. for  
19 2002, which is the most recent publicly available data. In the U.S., CO emissions data are tracked in the  
20 National Emissions Inventory (U.S. EPA, 2006a), a composite of data from various sources including  
21 industries and state, tribal, and local air agencies. NEI data are collected for all states, the District of  
22 Columbia, the U.S. territories of Puerto Rico and Virgin Islands, and some of the territories of federally  
23 recognized American Indian nations. Different data sources use different data collection methods, most of  
24 which are based on engineering calculations and estimates rather than measurements. Most fuel  
25 combustion and industrial sources, for example, estimate their CO emissions using EPA-approved  
26 emission factors, as do on-road and non-road mobile source emitters (U.S. EPA, 2007). Although these

1 estimates are generated using well-established approaches, uncertainties are inherent in the emission  
2 factors and models used to represent sources for which emissions have not been directly measured.

3       Nationally, on-road mobile sources in the NEI constituted more than half of total CO emissions in  
4 2002, or ~63 MT of ~109 MT total. For this reason, high concentrations of CO can often occur in areas of  
5 heavy traffic. In metropolitan areas in the U.S., for example, as much as 70-75% of all CO emissions  
6 came from on-road vehicle exhaust in the 2002 NEI (U.S. EPA, 2006a). The majority of these on-road CO  
7 emissions derive from gasoline-powered vehicles since the O<sub>2</sub> content, pressure, and temperature required  
8 for diesel fuel ignition result in much less CO production. When the emissions from incomplete  
9 combustion of fuels powering non-road mobile sources were included, all mobile sources accounted for  
10 ~80% of total CO emissions in the U.S. in 2002; see Figure 3-1.

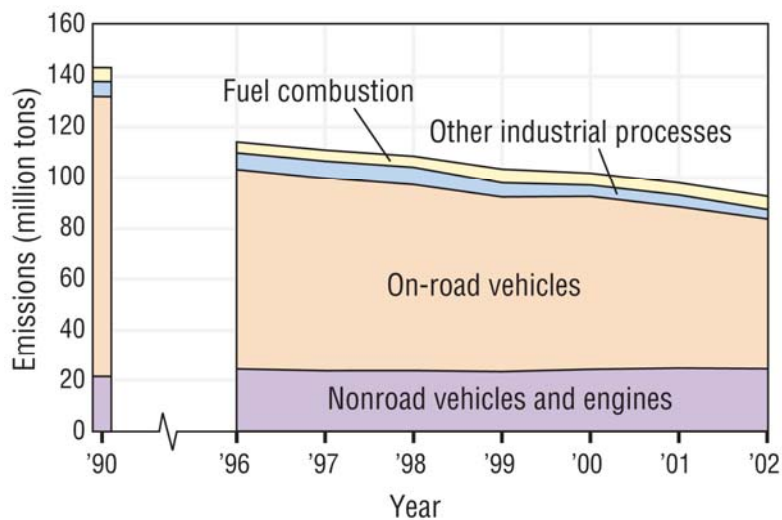


Source: U.S. EPA (2006a)

**Figure 3-1. CO emissions (MT) in the U.S. by source sector in 2002.**

11       Figure 3-2 shows present and historical CO emissions from the traditionally inventoried  
12 anthropogenic source categories: (1) fuel combustion, which includes emissions from coal-, gas-, and oil-  
13 fired power plants and industrial, commercial, and institutional sources, as well as residential heaters  
14 (e.g., wood-burning stoves) and boilers; (2) industrial processes, which includes chemical production,

1 petroleum refining, metals production, and industrial processes other than fuel combustion; (3) on-road  
 2 vehicles, which includes cars, trucks, buses, and motorcycles; and (4) non-road vehicles and engines, such  
 3 as farm and construction equipment, lawnmowers, chainsaws, boats, ships, snowmobiles, aircraft,  
 4 locomotive, and others. Using these NEI data, trends in the national CO emissions can be computed and  
 5 compared over time. So, for example, the national-scale estimated anthropogenic CO emissions decreased  
 6 35% between 1990 and 2002; see Figure 3-2. The trend plot in Figure 3-2 demonstrates that controls in  
 7 the on-road vehicle sector have produced nearly all the national-level CO reductions since 1990. Data are  
 8 presented here for 1990 and from 1996 to 2002 because only the 1990 data has been updated to be  
 9 comparable to the more recent inventories made since 1996.



Source: U.S. EPA (2006a, 2008c)

**Figure 3-2. Trends in anthropogenic CO emissions (MT) in the U.S. by source category for 1990 and 1996-2002.**

10 With the exception of this downward trend resulting from emissions controls, anthropogenic CO  
 11 emissions demonstrate less interannual variability than biogenic emissions (Bergamaschi et al., 2000).  
 12 Several recent reports using both ambient concentrations and fuel-based emissions estimates have  
 13 explored this annual-to-decadal emissions decrease in anthropogenic CO in finer detail. They included  
 14 Harley et al. (2001; 2005), Parrish et al. (2002; 2006), Pollack et al. (2004), and Mobley et al. (2005). The  
 15 consistent conclusion from those investigations has been that annual average U.S. on-road vehicle CO  
 16 emissions decreased at a rate of ~5% per year since the early 1990s. Additional analyses by Harley et al.  
 17 (2005), Parrish et al. (2002) and Parrish (2006) were also consistent with the suggestion in Pollack et al.  
 18 (2004) that the EPA MOBILE6 vehicle emissions model (<http://www.epa.gov/otaq/m6.htm>) now  
 19 overestimates vehicle CO emissions by a factor of ~2.



1 Biogenic emissions can vary widely from year to year (Bergamaschi et al., 2000). CO isotopes,  
2 most notably  $^{14}\text{C}$ , have been especially useful in helping refine emissions estimates and partition them  
3 between combustion and other sources. Weinstock and Niki (1972) analyzed measurements of the  $^{14}\text{C}$   
4 content of CO in Tonawanda, NY (a suburb of Buffalo) obtained by MacKay et al. (1963) during the  
5 winter of 1960 and found at that time that a large proportion of the CO was produced by sources  
6 containing modern carbon and not fossil carbon. The main source of this modern carbon CO was  
7 conjectured by Weinstock and Niki (1972) to have been from photochemical oxidation of  $\text{CH}_4$ ; however,  
8  $\text{CH}_4$  has a long atmospheric life time (about 8 years) but does oxidize to compounds like methyl radical  
9 and also is a contributor to  $\text{O}_3$  formation. Other work suggested that CO sources such as isoprene  
10 oxidation and biomass burning were also important. For example, Conny (1998) reviewed studies using  
11 measurements of  $^{14}\text{C}$  to characterize winter sources of CO in Las Vegas, NV, and Albuquerque, NM.  
12 These studies concluded that the contribution from residential wood burning could have been as high as  
13 30% in the samples collected.

14 Estimates of non-anthropogenic CO emissions are made using the Biogenic Emissions Inventory  
15 System (BEIS) model with data from the Biogenic Emissions Landcover Database (BELD) and annual  
16 meteorological data; see <http://www.epa.gov/ttnchie1/emch/biogenic>. National biogenic emissions,  
17 excluding fires, were estimated to contribute ~5% of total CO emissions from all sources in 2002; fires in  
18 2002 added another 13%, or ~14.5 MT, to the national CO emissions total. Geogenic emissions of CO  
19 also included in this inventory, include volcanic gases released from molten rock in the earth's mantle.  
20 Mixing ratios of dissolved CO in this rock vary in a range from 0.01 to 2% as a function of the rock  
21 stratum surrounding the volcano and other geologic conditions. This high variability and infrequent  
22 though often violent release mean geogenic CO measurements are very difficult to make with precision.  
23 On a global scale, however, the magnitude of their contribution is small relative to all anthropogenic  
24 sources. Photodecomposition of organic matter in oceans, rivers, lakes, and other surface waters, and  
25 from soil surfaces also releases CO (Goldstein and Galbally, 2007). However, soils can act as a CO source  
26 or a sink depending on soil moisture, UV flux reaching the soil surface, and soil temperature (Conrad and  
27 Seiler, 1985). Soil uptake of CO is driven by anaerobic bacteria (Inman et al., 1971). Emissions of CO  
28 from soils appear to occur by abiotic processes, such as thermodecomposition or photodecomposition of  
29 organic matter. In general, warm and moist conditions found in most soils favor CO uptake, whereas hot  
30 and dry conditions found in deserts and some savannas favor the release of CO (King, 1999).

31 Biomass burning consists of wildfires and the intentional burning of vegetation to clear new land  
32 for agriculture and population resettlement; to control the growth of unwanted plants on pasture land; to  
33 manage forest resources with prescribed burning; to dispose of agricultural and domestic waste; and as  
34 fuel for cooking, heating, and water sterilization. Globally, most wildfires may be ignited directly as the  
35 result of human activities leaving only 10 to 30% initiated by lightning (Andreae, 1991). However,  
36 because fire management practices suppress natural wildfires, the buildup of fire fuels increases the

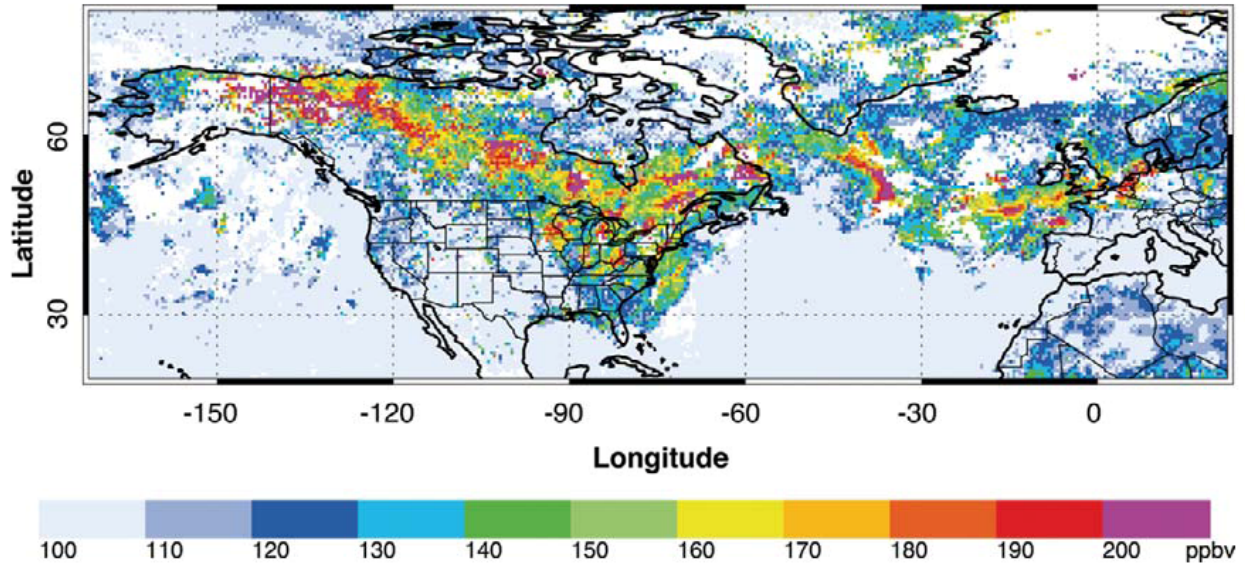
1 susceptibility of forests to more severe but less frequent fires in the future. There is considerable  
2 uncertainty in attributing the fraction of wildfire emissions to human activities because the emissions  
3 from naturally occurring fires that would have been present in the absence of fire suppression practices  
4 are not known.

5 Biomass burning also exhibits strong seasonality and interannual variability (van der Werf et al.,  
6 2006), with most biomass burned during the local dry season. This is true for both prescribed burns and  
7 wildfire. The unusually warm and dry weather in central Alaska and western Yukon in the summer of  
8 2004, for example, contributed to the burning of 11 million acres there. These fires, the largest on record  
9 for this region, produced CO emissions easily tracked by the Measurement of Pollution in the  
10 Troposphere (MOPITT) instrument on NASA's Terra satellite; see Figure 3-3. The high CO concentration  
11 measured by MOPITT coincided with the surface location of fires using aerosol plumes identified by the  
12 Moderate Resolution Imaging Spectroradiometer (MODIS) also on Terra. Subsequent modeling by Pfister  
13 et al. (2005) showed that the CO contribution from these fires in July 2004 was 30 ( $\pm 5$ ) teragrams (Tg)  
14 that summer, or in the range of the total U.S. anthropogenic CO emissions during the same time.

15 The smoldering phase of combustion yields higher CO emissions than the flaming phase. Using  
16 controlled combustion chamber experiments Lobert et al. (1991) found that with a wide variety of  
17 vegetation types, on average, 84% of the CO from biomass fires was produced during the smoldering  
18 phase and 16% during the flaming phase of combustion.

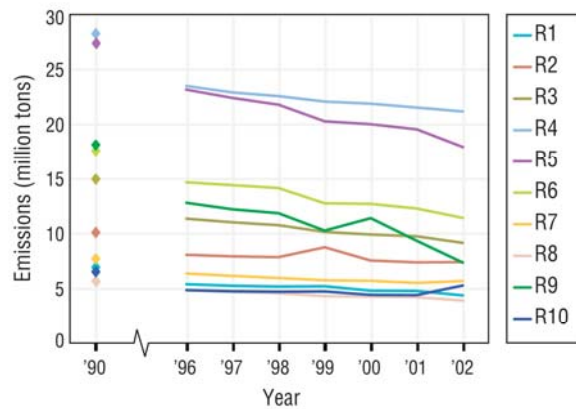
19 CO emissions data for EPA's ten administrative Regions in the U.S. depicted in Figure 3-4 show a  
20 more nuanced view of the national concentrations and trends described just above. Net anthropogenic CO  
21 emissions were estimated to have declined in all EPA Regions between 1990 and 2002 with the largest  
22 decrease (10.8 MT) occurring in Region 9.

23 On still finer scales, CO emissions from on-road mobile sources or from fires can dominate in  
24 different places across the U.S. Figure 3-5 illustrates this variability with CO state-level emissions total  
25 and selected county totals in 2002 for Colorado. (Annex A includes analogous data for Alaska, Utah,  
26 Massachusetts, Georgia, California, and Alabama.) In Colorado, emissions from fires and on-road  
27 vehicles were nearly equal: ~0.9 MT from fires and ~1.1 MT from on-road vehicles, though emissions  
28 varied strongly across counties with urban Denver County dominated by on-road vehicle emissions at  
29 71% and rural Garfield County dominated by fire emissions at 67%.



Source: Fishman et al. (2008)

**Figure 3-3.** CO concentrations measured by satellite at the 700 hectoPascal level (~10,000 feet above sea level) from MOPITT for the period 15-23 July 2004 during intense wildfires in Alaska and Yukon.



Data are presented for 1990 and 1996-2002, as datasets from these inventory years are all fully up to date. Data are available for inventory years 1991-1995, but these data have not been updated to allow comparison with data from 1990 and 1996-2002.



Source: U.S. EPA (2006a, 2008c)

**Figure 3-4.** Trends in sub-national CO emissions in the 10 U.S. EPA Regions for 1990 and 1996 to 2002.

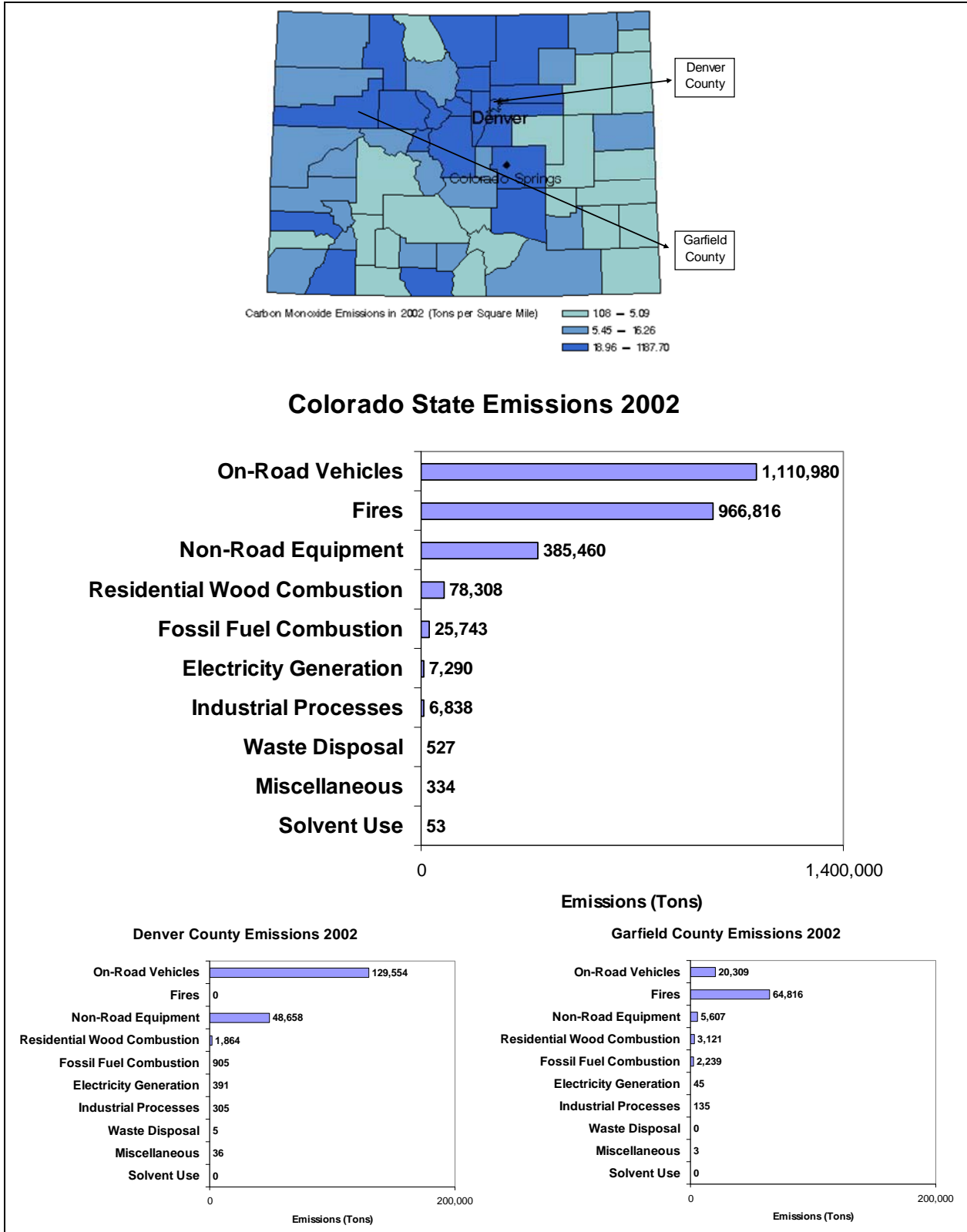
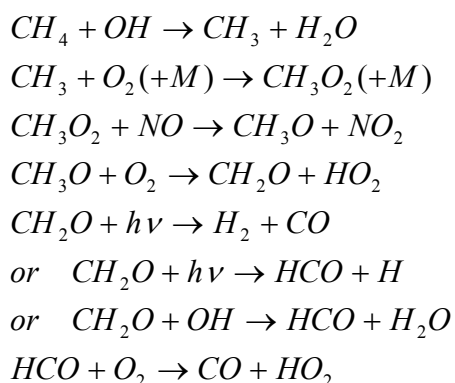


Figure 3-5. CO emissions density map and distribution for the state of Colorado, and for selected counties in Colorado.

### 3.3. Physics and Chemistry of Atmospheric CO

1 In addition to being emitted directly by combustion sources, CO is produced by photooxidation of  
2 methane (CH<sub>4</sub>) and other VOCs including nonmethane hydrocarbons (NMHCs) in the atmosphere, and of  
3 organic molecules in surface waters and soils. CH<sub>4</sub> oxidation is summarized in the following reaction  
4 sequence:



5 where M is a reaction mediator providing collisional energy but is neither created nor destroyed.

6 Photolysis of formaldehyde (CH<sub>2</sub>O) proceeds by two pathways. The first produces molecular  
7 hydrogen (H<sub>2</sub>) and CO with a reaction yield of 55% in conditions of clear skies and low zenith angles; the  
8 second yields a hydrogen radical (H) and the formyl radical (HCO). HCO then reacts with O<sub>2</sub> to form  
9 hydroperoxy radical (HO<sub>2</sub>) and CO. Reaction of methyl peroxy radical (CH<sub>3</sub>O<sub>2</sub>) with HO<sub>2</sub> radicals  
10 (reaction not shown) to form methyl hydroperoxide (CH<sub>3</sub>OOH) is also operative, especially in low oxides  
11 of nitrogen (NO<sub>x</sub>) conditions. Heterogeneous removal of the water-soluble intermediate products  
12 CH<sub>3</sub>OOH, CH<sub>2</sub>O, and radicals will decrease CO yields from CH<sub>4</sub> oxidation.

13 While oxidation of CH<sub>2</sub>O nearly always produces CO and some small quantities of formic acid  
14 (CH<sub>2</sub>O<sub>2</sub>) in the reaction of CH<sub>2</sub>O with HO<sub>2</sub> (not shown here), oxidation of acetaldehyde (CH<sub>3</sub>CHO) does  
15 not always yield two CO molecules. Reaction of CH<sub>3</sub>CHO with OH can yield acetyl radicals (CH<sub>3</sub>CO)  
16 which then will participate with O<sub>2</sub> in a termolecular recombination reaction to form peroxyacetyl radicals,  
17 which then can react with nitric oxide (NO) to form CH<sub>3</sub> and CO<sub>2</sub>; or the peroxyacetyl radicals can react  
18 with NO<sub>2</sub> to form peroxyacetyl nitrate (PAN), CH<sub>3</sub>CO<sub>3</sub>NO<sub>2</sub>. In this way, one carbon atom is oxidized  
19 directly to CO<sub>2</sub> without passing through CO. The yield of CO from these pathways depends on the OH  
20 concentration and the photolysis rate of CH<sub>3</sub>CHO, as well as on the abundance of NO, since peroxyacetyl  
21 radicals also will react with other odd hydrogen radicals like OH.

22 Estimating the CO yield from oxidation of hydrocarbons (HCs) larger than CH<sub>4</sub> requires computing  
23 the yields of CH<sub>2</sub>O, CH<sub>3</sub>CHO, CH<sub>3</sub>CO, and analogous radicals from oxidation of the parent molecules.  
24 Moreover, the extent of heterogeneous removal of soluble intermediate products also affects oxidation of  
25 more complex HCs. However, the detailed gas-phase kinetics for many HCs with more than a few

1 carbons is still unknown, and this is especially the case for several important classes of VOCs including  
2 the aromatics, biogenic HCs including isoprene, and their intermediate oxidation products like epoxides,  
3 nitrates, and carbonyls. It has long been known that as much as 30% of the carbon in HCs in many urban  
4 areas is in the form of aromatics largely from mobile sources since gasoline contains significant quantities  
5 of aromatics (Grosjean and Fung, 1984; Seila et al., 1989). Yet mass balance analyses performed on  
6 irradiated smog chamber mixtures of aromatic HCs indicate that only about one-half of the carbon is in  
7 the form of compounds that can be identified. In addition, reactions like the oxidation of terpenes that  
8 produce condensable products are also significant because these reactions produce secondary organic  
9 aerosols, thereby reducing the potential yield of CO. The CO yield from oxidation of CH<sub>4</sub>, for example, is  
10 ~0.9 on a per carbon basis (Kanakidou and Crutzen, 1999). Yields from other compounds range from less  
11 than 0.1 for anthropogenic alkanes (Altshuller, 1991) to ~0.7 for some other non-CH<sub>4</sub> HCs; see  
12 Kanakidou and Crutzen (1999) for CO yields from other HCs.

13 The major pathway for removal of CO from the atmosphere is reaction with OH to produce CO<sub>2</sub>  
14 and H radicals that rapidly combine with O<sub>2</sub> to form HO<sub>2</sub> radicals with a rate constant at 1 atm in air of  
15 ~2.4 × 10<sup>-13</sup> cm<sup>3</sup>/molecule/s (Finlayson-Pitts and Pitts Jr). The mean tropospheric photochemical lifetime  
16 (τ) of CO in the northern hemisphere is ~57 days (Khalil and Rasmussen, 1990; Thompson and Cicerone,  
17 1986). Owing to variation in atmospheric water vapor, OH concentration, and insolation, shorter τ are  
18 found nearer the tropics and longer ones at higher latitudes. During winter at high latitudes CO has nearly  
19 no photochemical reactivity on urban and regional scales. The CO τ is shorter than the characteristic time  
20 scale for mixing between the hemispheres of ~1 year; hence a large gradient in concentrations can exist  
21 between the hemispheres. In addition, the CO τ at high latitudes is long enough to result in much smaller  
22 gradients between 30° latitude and the pole of either hemisphere. The typical residence times of CO in  
23 urban areas when assuming a diel-average OH concentration of 3 × 10<sup>6</sup>/cm<sup>3</sup> in urban areas is ~16 days, so  
24 CO will not be destroyed in urban areas where it is emitted and will likely be mixed on continental and  
25 larger scales. OH concentrations are orders of magnitude lower in indoor environments and so CO will  
26 generally not be destroyed by indoor air reactions.

27 Recent data do not alter the current well-established understanding of the role of urban and regional  
28 CO in continental and global-scale chemistry outlined in the 2000 CO AQCD (U.S. EPA, 2000), and  
29 subsequently confirmed in the global assessments of climate change by the Intergovernmental Panel on  
30 Climate Change (IPCC, 2001, 2007). CO is a weak contributor to greenhouse warming because its  
31 fundamental absorption band near 4.63 μm is far from the spectral maximum of earth's longwave  
32 radiation at ~10 μm. However, because reaction with CO is also the major sink for OH, changes in CO  
33 concentrations can lead to changes in the concentrations of other trace gases whose loss processes involve  
34 OH attack. Some of these trace gases, CH<sub>4</sub>, for example, absorb infrared radiation from the earth's surface  
35 and contribute to the greenhouse effect directly; others, including the hydrochlorofluorocarbons (HCFCs)

1 and methyl chloride and methyl bromide, can deplete stratospheric O<sub>3</sub>, increasing the surface-incident UV  
2 flux.

3 This indirect effect of CO on stratospheric O<sub>3</sub> concentrations is opposite in sign to the effect of CO  
4 on O<sub>3</sub> in the troposphere where CO reacts in a manner similar to other VOCs in the presence of NO<sub>x</sub> and  
5 UV to create O<sub>3</sub>. (See the detailed description of O<sub>3</sub> formation from VOCs and NO<sub>x</sub> in the 2008 NO<sub>x</sub> ISA  
6 (U.S. EPA, 2008e). The long chemical lifetime and one-to-one stoichiometry of CO oxidation (whereby  
7 one molecule of CO converts only one molecule of NO to NO<sub>2</sub>) means that CO has significantly lower O<sub>3</sub>  
8 forming potential than other VOCs in the troposphere. Carter (1998) computed a maximum incremental  
9 reactivity for CO of 0.07 g O<sub>3</sub> for 1 g CO, as compared to reactivities of total on-road vehicle exhaust  
10 emissions in the range of 3 to 4. However, because the total mass of CO emissions is substantially greater  
11 than those of the other VOCs with higher carbon numbers and faster reactivities, CO can contribute  
12 greatly to O<sub>3</sub> formation even though its photochemical processing is slow. Using data from instrumented  
13 models including that of Jeffries (1995), NRC (1999) estimated, for example, that CO can contribute 15-  
14 25% of the total O<sub>3</sub> forming potential of gasoline exhaust emissions; this figure varies from region to  
15 region. The contribution of CO to urban and regional O<sub>3</sub> concentration is often less than 10% owing to its  
16 very slow reactivity on these scales and to locally variable radical concentration ratios.

17 Because the greenhouse warming effects from CO are nearly completely indirect, and because CO  
18 concentrations are spatially heterogeneous, neither the IPCC nor EPA computes global warming potentials  
19 (GWPs) for CO, just as they do not for tropospheric O<sub>3</sub>, NO, NO<sub>2</sub>, or VOCs (U.S. EPA, 2008a).  
20 Additionally, urban and regional-scale oxidation of CO to CO<sub>2</sub> under current atmospheric conditions  
21 proceeds very slowly and IPCC considers production of CO<sub>2</sub> through this pathway to be double counting  
22 of CO effects (IPCC, 2007).

## 3.4. Ambient Measurements

### 3.4.1. Ambient Measurement Instruments

23 To promote uniform enforcement of the air quality standards set forth under the Clean Air Act, EPA  
24 has established provisions in the Code of Federal Regulations (CFR) under which analytical methods can  
25 be designated as federal reference methods or federal equivalent methods (FRM or FEM, respectively).  
26 Measurements for determinations of NAAQS compliance must be made with FRMs or FEMs. As of  
27 December 2008, 19 automated FRMs and no FEMs had been approved for CO  
28 (<http://www.epa.gov/ttn/amtic/criteria.html>).

1 All EPA FRMs for CO operate on the principle of nondispersive infrared (NDIR) detection and can  
2 include the gas filter correlation (GFC) methodology. NDIR is an automated and continuous method  
3 based on the specific absorption of infrared radiation by the CO molecule. Most commercially available  
4 analyzers incorporate a gas filter to minimize interferences from other gases and operate near atmospheric  
5 pressure. The most sensitive trace-level versions of these instruments can detect minimum CO  
6 concentrations of ~0.04 ppm; the required lower detection limit for FRMs in the EPA network is 1.0 ppm  
7 (40 CFR 53.20 Table B-1).

8 NDIR is based on the physics of CO's characteristic infrared absorption near 4.63  $\mu\text{m}$ . NDIR  
9 methods have several practical advantages over other techniques for CO detection in that they are not  
10 sensitive to flow rate changes, require no wet chemicals, are reasonably independent of ambient air  
11 temperature changes, are sensitive over wide concentration ranges, and have fast response times. Earlier  
12 concerns over zero-drift and nonlinear span responses have been addressed. An extensive and  
13 comprehensive review of NDIR, GFC, and alternative, non-FRM techniques for CO detection including  
14 tunable diode laser spectroscopy, gas chromatography, mercury liberation, and resonance fluorescence  
15 was made for the 2000 CO AQCD (U.S. EPA, 2000), and the reader is directed there for additional  
16 information. The description here is limited to a brief outline of the FRM NDIR and GFC techniques.

17 GFC spectroscopy analyzers are used most frequently now in documenting compliance with  
18 ambient air standards. A GFC monitor has all of the advantages of an NDIR instrument and the additional  
19 advantages of smaller size, no interference from CO<sub>2</sub>, and very small interference from water vapor.  
20 During operation, air flows continuously through a sample cell. Radiation from the infrared source is  
21 directed by optical transfer elements through two main optical subsystems: (1) the rotating gas filter and  
22 (2) the optical multipass (sample) cell. The beam exits the sample cell through an interference filter (FC),  
23 which limits the spectral passband to a few of the strongest CO absorption lines. Detection of the  
24 transmitted radiation occurs at the infrared detector. The gas correlation cell is constructed with two  
25 compartments, one filled with 0.5 atm CO, and a second with pure N<sub>2</sub>. Radiation transmitted through the  
26 CO is completely attenuated at the wavelengths where CO absorbs strongly. The radiation transmitted  
27 through the nitrogen gas (N<sub>2</sub>) is reduced by coating the exit window of the cell with a neutral attenuator  
28 so that the amounts of radiation transmitted by the two cells are made approximately equal in the  
29 passband that reaches the detector. In operation, radiation passes alternately through the two cells as they  
30 are rotated to establish a signal modulation frequency. If CO is present in the sample, the radiation  
31 transmitted through the CO is not appreciably changed, whereas that through the N<sub>2</sub> cell is changed. This  
32 imbalance is linearly related to CO concentrations in ambient air.



## 3.4.2. Ambient Sampling Network Design

### 3.4.2.1. Monitor Siting Requirements

1 CO monitoring is included at all active NCore sites and State and Local Air Monitoring Stations  
2 (SLAMS) where continued measurements of CO using FRM are required until discontinuation is  
3 approved by the EPA Regional Administrator. Where SLAMS CO monitoring is required, at least one of  
4 the sites must be a max concentration site for that specific area. In 2007, there were ~385 CO monitors  
5 reporting values to the EPA Air Quality System (AQS) database. Where CO monitoring is ongoing, 40  
6 CFR Part 58 requires at least one CO monitor to capture maximum levels in a given region. This is often  
7 done with a monitor situated at the CFR-defined microscale distance (2-10 m) from the side of a roadway  
8 for CO. Microscale monitor locations also have sample inlets mounted at  $3 \pm 0.5$  m above ground level,  
9 unlike the monitors sampling for larger scales, whose inlet heights can vary between 2 and 15 m. For the  
10 CFR-defined middle (up to 500 m) and neighborhood (~500 m-4 km) scale monitoring, the minimum  
11 monitor distance from a major roadway is inversely related to the average daily traffic counts on that  
12 roadway to ensure that measurements are not substantially influenced by any one roadway. For example,  
13 the minimum distance of a middle scale CO monitor from a roadway with an average daily traffic count  
14 of 50,000 vehicles per day is 135 m. More detail on siting requirements can be found in 40 CFR Part 58  
15 Appendices D through E.

### 3.4.2.2. Spatial and Temporal Coverage

16 Figure 3-6 depicts the distribution of the ~385 regulatory CO monitors operating in the U.S. in  
17 2007. Data from 285 of the ~385 CO monitors operating year-round in the years 2005-2007 met the 75%  
18 data completeness criterion for inclusion in the multi-year ambient data analyses for this assessment. The  
19 greatest density of monitors is in the CSAs for Los Angeles, CA and San Francisco, CA, and along the  
20 Mid-Atlantic sea board. Monitors are also located in regions where biomass burning is more prevalent,  
21 such as Anchorage, AK, but not all of these monitors report values from all seasons of all years.

22 Eleven metropolitan regions were chosen for closer investigation of monitor siting based on their  
23 relevance to the health studies assessed in subsequent chapters of this ISA and to demonstrate specific  
24 points about geospatial distributions of CO emissions and concentrations. These regions were: Anchorage,  
25 AK; Atlanta, GA; Boston, MA; Denver, CO; Houston, TX; Los Angeles, CA; New York City, NY;  
26 Phoenix, AZ; Pittsburgh, PA; Seattle, WA; and St. Louis, MO. Core-Based Statistical Areas (CBSAs) and  
27 Combined Statistical Areas (CSAs), as defined by the U.S. Census Bureau (<http://www.census.gov/>),  
28 were used to determine which counties, and hence which monitors, to include for each metropolitan

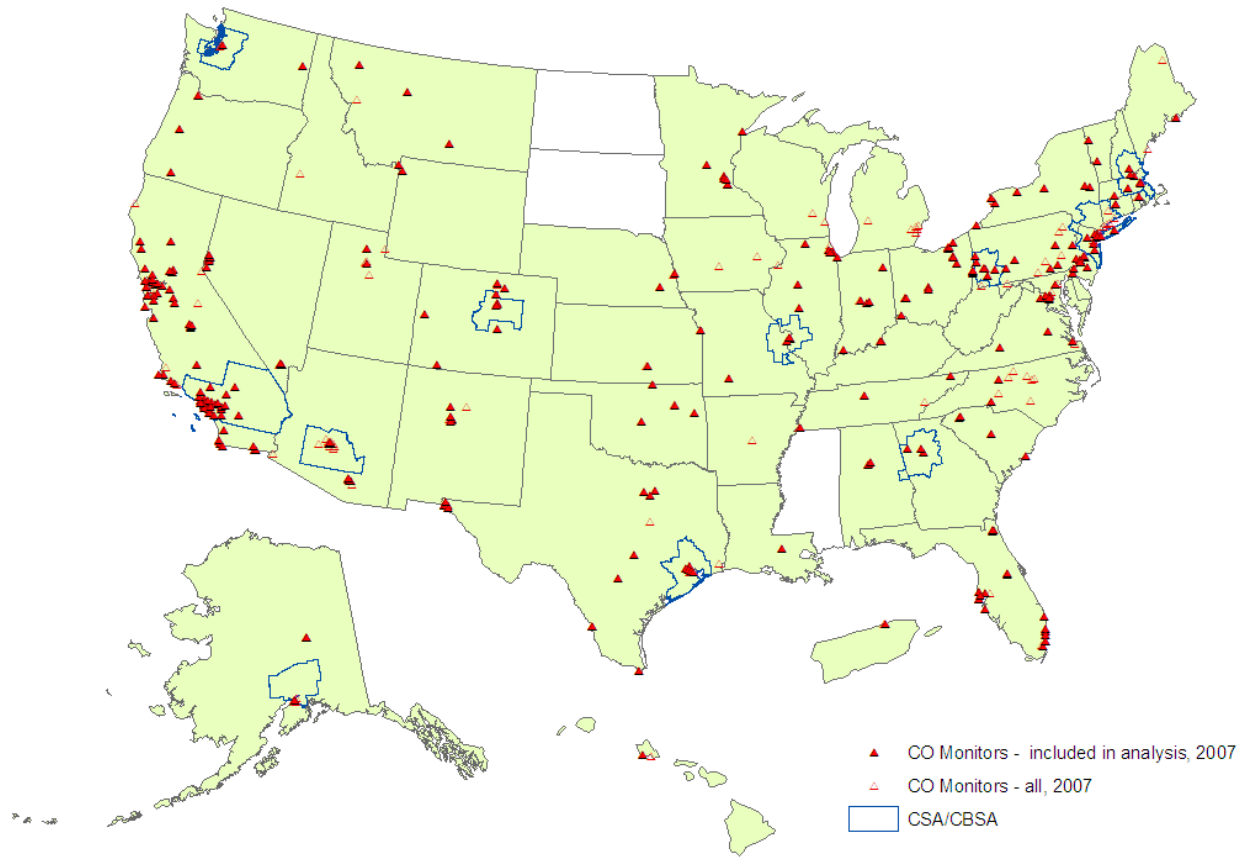
1 region.<sup>1</sup> As an example, Figures 3-7 through 3-10 display CO monitor density with respect to population  
2 density (for total population and elderly adults aged 65 and over) for Phoenix and Pittsburgh. (Annex A  
3 includes analogous plots for the other nine metropolitan regions.)

4 Although ambient monitors for CO and other criteria pollutants are explicitly not located in order  
5 to monitor population exposures, it is instructive to test the utility of the current network for  
6 characterizing this exposure. Tables 3-1 and 3-2 show the population density around CO monitors for the  
7 total population and for elderly adults aged 65 and over for each CSA/CBSA. The percentage of  
8 population within specific radii of the monitors for each city was, for the most part, similar between the  
9 total and elderly populations. In the cases of Anchorage, Denver, Phoenix, and St. Louis however, the  
10 percentage of the elderly population within given radii of the monitors was considerably different  
11 compared with the total population. Between-city disparities in population density were larger. Los  
12 Angeles, with 85%, and Denver, with 68%, had the largest proportion of the total population within 15  
13 km of a monitor. Seattle, with 18%, had the lowest population coverage in large part because ambient CO  
14 concentrations there require only a single CO monitor. As concerns the elderly population, Los Angeles,  
15 at 83%, Anchorage, at 73%, and Denver, at 70%, had the greatest population coverage within 15 km of a  
16 monitor, whereas Seattle, at 18%, again, had the lowest coverage. Proximity to monitoring stations is  
17 considered further in sections 3.5 and 3.6 regarding spatial variability within cities. Figures 3-7 through 3-  
18 10 show that multiple CO monitors in Phoenix and Pittsburgh were in the city center, and that Pittsburgh  
19 also had monitors in outlying areas of moderate to low population density. The CO monitors in Phoenix  
20 appear to provide good coverage of areas of the highest total population density, while areas of high  
21 elderly population density were located at greater distances from the monitors. CO monitors in Pittsburgh  
22 appear to sample areas of high population density well for both the total and elderly populations.

23 Figures 3-13 and 3-15 in section 3.5 and additional figures in Annex A show the locations of CO  
24 monitors for the 11 CSAs/CBSAs in relation to major roadways, including Interstate highways, U.S.  
25 highways, state highways, and other major roadways required for traffic network connectivity. In most  
26 cases, the monitors were concentrated near the center of the CSA/CBSA. Regional background sites were  
27 not included on the maps unless they lay within the CSA/CBSA.

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<sup>1</sup> A CBSA represents a county-based region surrounding an urban center of at least 10,000 people determined using 2000 census data and replaces the older Metropolitan Statistical Area (MSA) definition from 1990. The CSA represents an aggregate of adjacent CBSAs tied by specific commuting behaviors. The broader CSA definition was used when selecting monitors for the cities listed above with the exception of Anchorage and Phoenix, which are not contained within a CSA. Therefore, the smaller CBSA definition was used for these metropolitan areas.



**Figure 3-6. Map of ~385 CO monitor locations in the U.S. in 2007. Locations are indicated with triangles: filled triangles show locations of the 285 monitors used in data analysis for this assessment; open triangles are at locations with monitors which did not meet the data completeness requirements for analysis; blue lines mark the boundaries of the 11 CSAs/CBSAs used in the data analysis for this assessment.**

### Phoenix Core Based Statistical Area

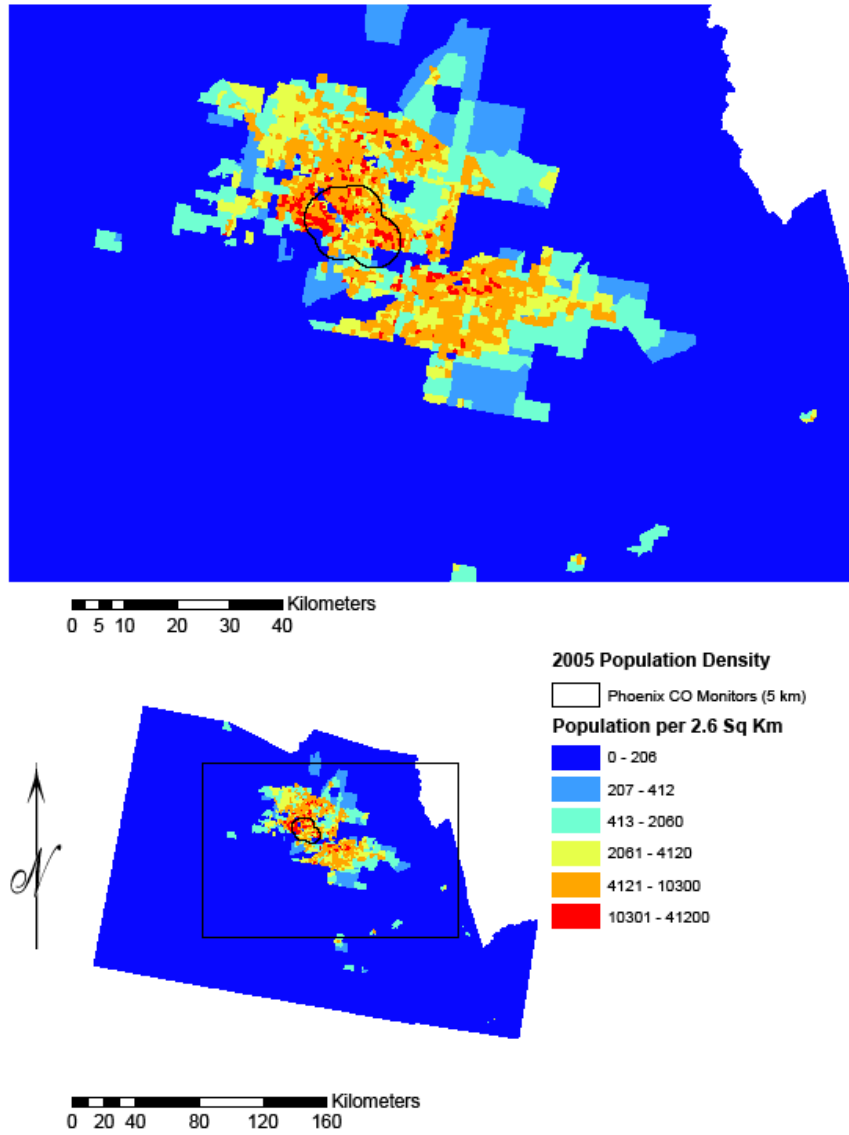


Figure 3-7. Map of CO monitor locations with respect to population density in the Phoenix, AZ CBSA, total population.

### Phoenix Core Based Statistical Area

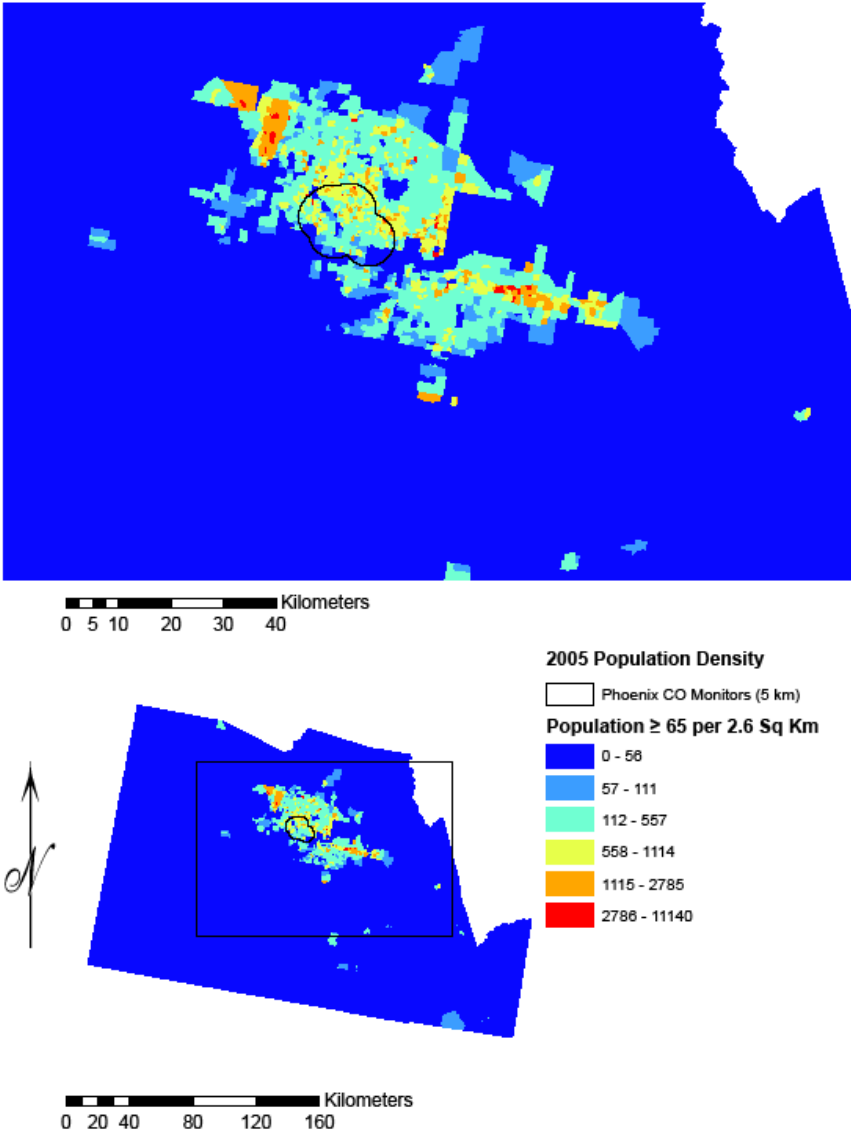


Figure 3-8. Map of CO monitor locations with respect to population density in the Phoenix, AZ CBSA, age 65 and older.

### Pittsburgh Combined Statistical Area

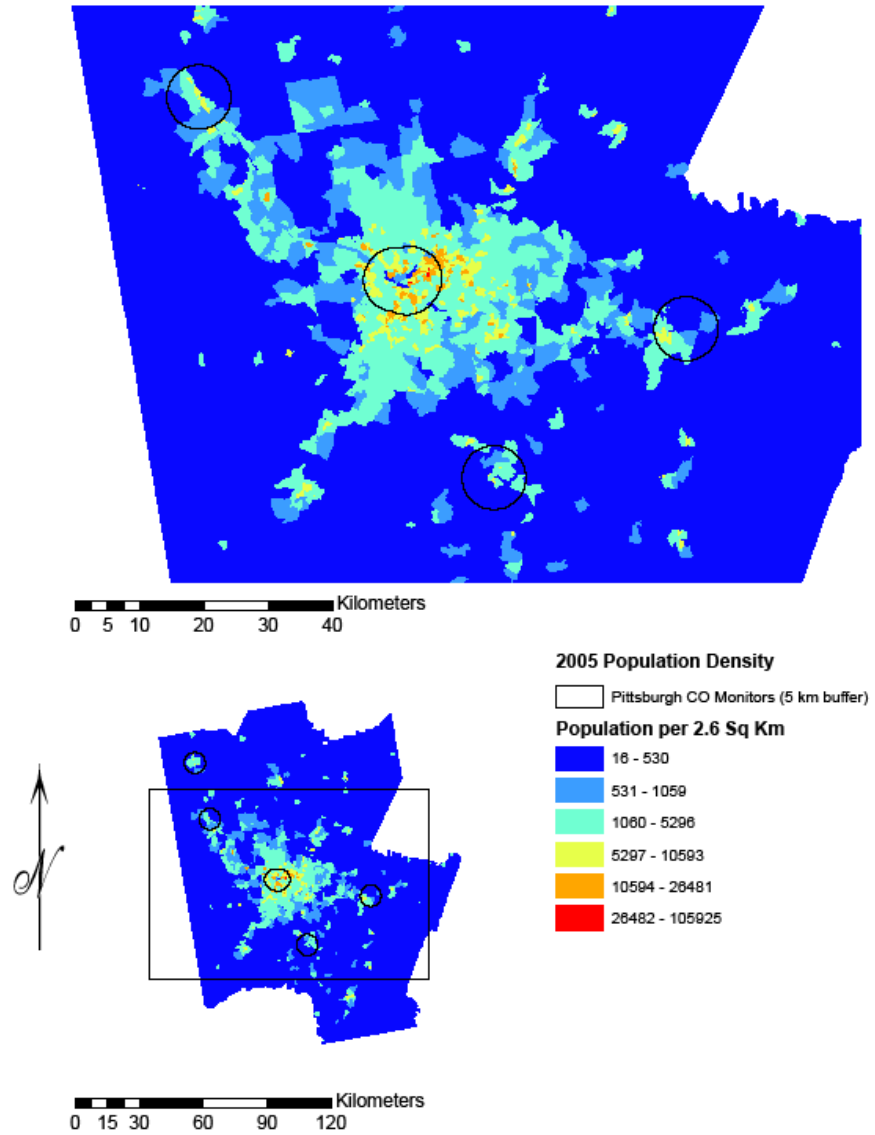


Figure 3-9. Map of CO monitor locations with respect to population density in the Pittsburgh, PA CSA, total population.

### Pittsburgh Combined Statistical Area

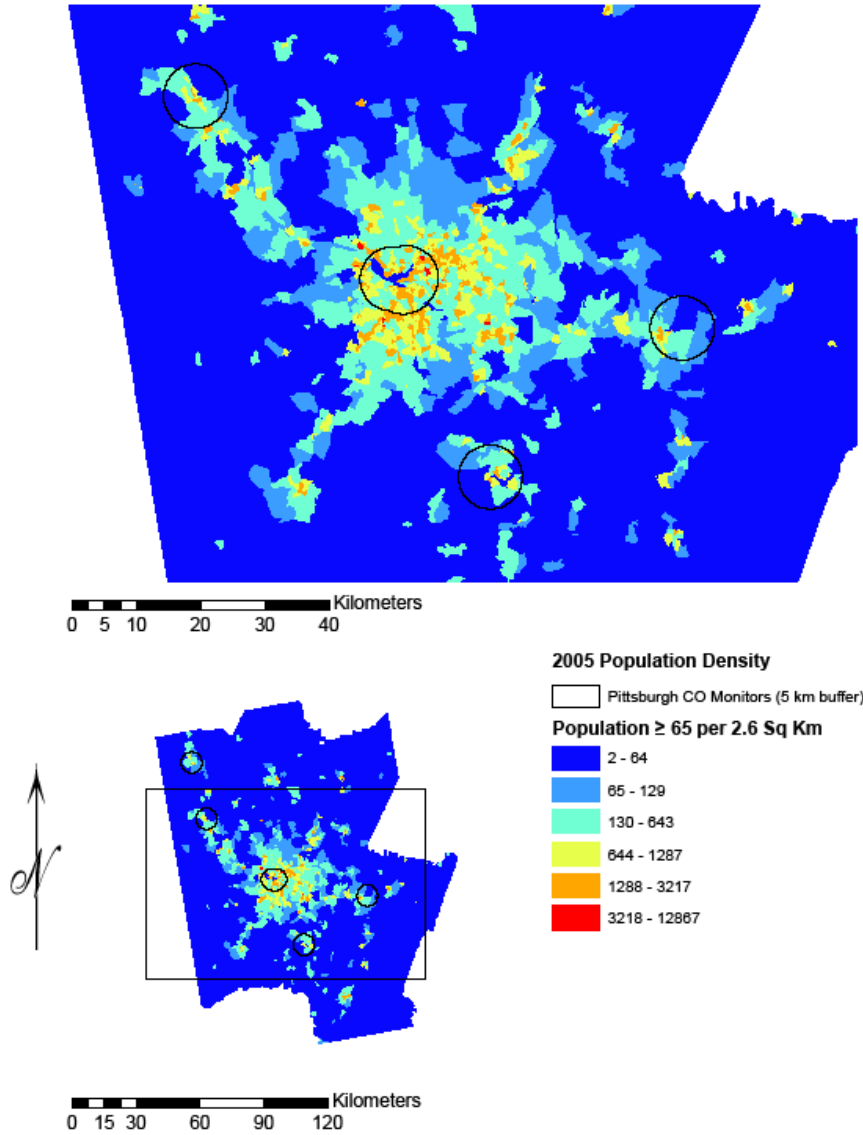


Figure 3-10. Map of CO monitor locations with respect to population density in the Pittsburgh, PA CSA, age 65 and older.

**Table 3-1. Proximity to CO monitors for the total population by city.**

Region	Total CSA/ CBSA	≤ 1 km		≤ 5 km		≤ 10 km		≤ 15 km	
	N	N	%	N	%	N	%	N	%
Anchorage, AK	352,225	5,391	1.53	131,608	37.36	212,834	60.43	239,842	68.09
Atlanta, GA	5,316,742	5,480	0.10	149,772	2.82	672,701	12.65	1,444,986	27.18
Boston, MA	7,502,707	95,732	1.28	1,180,054	15.73	2,432,846	32.43	3,418,353	45.56
Denver, CO	2,952,039	26,096	0.88	497,598	16.86	1,091,444	36.97	1,720,360	58.28
Houston, TX	5,503,320	2,9068	0.53	599,796	10.90	1,669,117	30.33	2,506,830	45.55
Los Angeles, CA	17,655,319	202,340	1.15	4,064,309	23.02	11,928,427	67.56	15,074,972	85.38
New York, NY	22,050,940	201,350	0.91	3,711,369	16.83	8,385,801	38.03	12,454,837	56.48
Phoenix, AZ	3,818,147	47,478	1.24	503,433	13.19	1,033,102	27.06	1,581,887	41.43
Pittsburgh, PA	2,515,383	29,136	1.16	369,965	14.71	895,252	35.59	1,359,596	54.05
Seattle, WA	3,962,434	4,814	0.12	94,649	2.39	279,976	7.07	699,490	17.65
St. Louis, MO	2,869,955	16,638	0.58	255,499	8.90	886,412	30.89	1,303,636	45.42

**Table 3-2. Proximity to CO monitors for adults aged 65 and older by city.**

Region	Total CSA/ CBSA	≤ 1 km		≤ 5 km		≤ 10 km		≤ 15 km	
	N	N	%	N	%	N	%	N	%
Anchorage, AK	17,742	361	2.03	8,986	50.65	12,038	67.85	12,990	73.22
Atlanta, GA	362,201	423	0.12	12,758	3.52	54,148	14.95	111,232	30.71
Boston, MA	945,790	8,272	0.87	131,198	13.87	297,392	31.44	430,502	45.52
Denver, CO	232,974	2,541	1.09	42,760	18.35	102,783	44.12	163,682	70.26
Houston, TX	377,586	1,703	0.45	42,312	11.21	130,567	34.58	182,049	48.21
Los Angeles, CA	1,626,663	17,974	1.10	380,079	23.37	1,069,188	65.73	1,355,461	83.33
New York, NY	2,710,675	29,534	1.09	427,601	15.77	940,121	34.68	1,429,215	52.73
Phoenix, AZ	388,150	2,877	0.74	35,839	9.23	77,244	19.90	125,300	32.28
Pittsburgh, PA	449,544	5,383	1.20	66,967	14.90	166,440	37.02	255,220	56.77
Seattle, WA	390,372	556	0.14	12,142	3.11	3,1036	7.95	69,858	17.90
St. Louis, MO	358,747	3,203	0.89	42,890	11.96	127,274	35.48	184,491	51.43

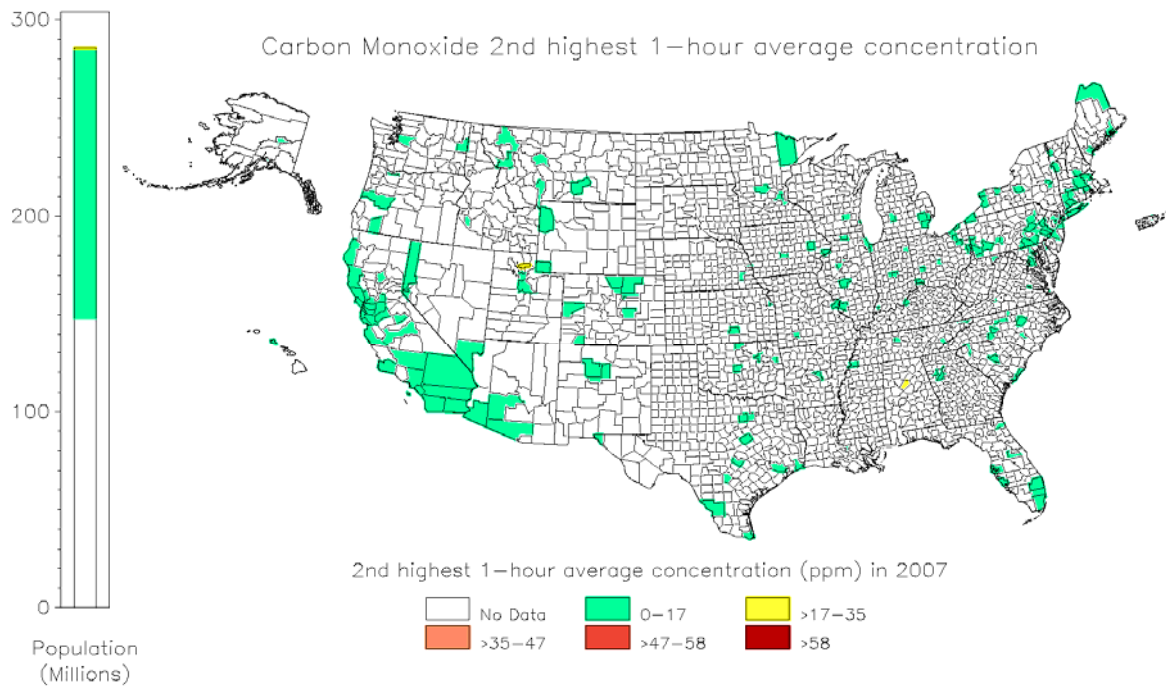


## 3.5. Environmental Concentrations

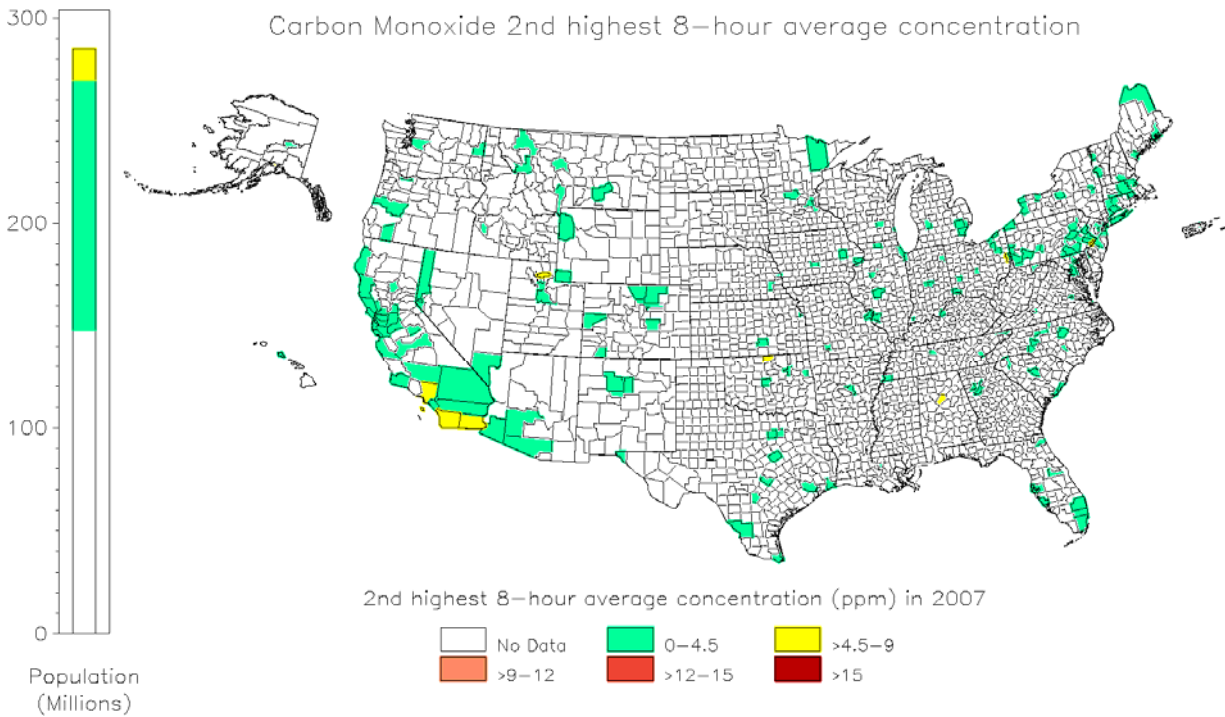
### 3.5.1. Spatial Variability

#### 3.5.1.1. National Scale

1           The current NAAQS designates that the level of the NAAQS can only be exceeded once per year at  
2 a given location. Figures 3-11 and 3-12 show the second-highest 1-h and second-highest 8-h county-  
3 average CO concentrations, respectively, over the U.S. along with estimates of the fraction of U.S. total  
4 population exposed to those concentrations. Although 93% of the U.S. counties are not represented in  
5 AQS reporting, based on their population densities and proximity to sources, those counties are not  
6 expected to have higher concentrations than the ones analyzed here in the absence of extreme events such  
7 as wildfires. Continuous hourly averages are reported from U.S. monitoring stations. 1-h and 8-h CO data  
8 were available for 243 counties and autonomous cities or municipalities (e.g., Anchorage, AK,  
9 Washington, DC) where CO monitors met the 75% data completeness criteria used in this analysis for the  
10 years 2005-2007. In 2007, no monitored location reported a second-highest 1-h CO concentration above  
11 35 ppm, the level of the current 1-h NAAQS; see Figure 3-11. Moreover, only two monitored locations,  
12 one in Weber Co., UT and the other in Jefferson Co., AL (including Birmingham, AL), reported second-  
13 highest 1-h CO concentrations above 18 ppm, or approximately one-half the level of the 1-h standard.  
14 Figure 3-12 shows that no counties reported second-highest 8-h CO concentrations above 9 ppm, the level  
15 of the 8-h NAAQS. Only nine counties reported second-highest 8-h CO concentrations above 4.5 ppm:  
16 Jefferson Co., AL, Kay Co., OK, Imperial Co., CA, Weber Co., UT, Philadelphia Co., PA, Anchorage  
17 Municipality, AK, Los Angeles Co., CA, Jefferson Co., OH, and San Diego Co., CA.



**Figure 3-11.** County-level map of second-highest 1-h avg CO concentrations in the U.S. in 2007. The bar on the left shows the total U.S. population who live in counties with CO concentrations in the range indicated. Note that approximately 150 million people live in counties with no CO monitors.



**Figure 3-12.** County-level map of second-highest 8-h avg CO concentrations in the U.S. in 2007. The bar on the left shows the total U.S. population who live in counties with CO concentrations in the range indicated. Note that approximately 150 million people live in counties with no CO monitors.

**Table 3-3. Distribution of 1-h avg CO concentration (ppm) derived from AQS data.**

	n	Mean	Min	Percentiles									Max
				1	5	10	25	50	75	90	95	99	
<i>NATIONWIDE STATISTICS (N = NUMBER OF OBSERVATIONS)</i>													
2005-2007	7,180,700	0.5	0.0	0.0	0.0	0.1	0.2	0.4	0.6	0.9	1.2	2.1	39.0
2005	2,391,962	0.5	0.0	0.0	0.0	0.1	0.2	0.4	0.6	1.0	1.3	2.3	22.3
2006	2,402,153	0.5	0.0	0.0	0.0	0.1	0.2	0.4	0.6	0.9	1.2	2.1	35.3
2007	2,386,585	0.4	0.0	0.0	0.0	0.1	0.2	0.3	0.5	0.8	1.1	1.9	39.0
Winter (December - February)	1,752,340	0.6	0.0	0.0	0.0	0.1	0.3	0.4	0.7	1.2	1.6	2.7	20.0
Spring (March - May)	1,826,167	0.4	0.0	0.0	0.0	0.1	0.2	0.3	0.5	0.8	1.0	1.7	35.3
Summer (June - August)	1,811,082	0.4	0.0	0.0	0.0	0.0	0.2	0.3	0.5	0.7	0.9	1.5	39.0
Fall (September - November)	1,791,111	0.5	0.0	0.0	0.0	0.1	0.2	0.4	0.6	1.0	1.3	2.2	24.1
<i>NATIONWIDE STATISTICS, POOLED BY SITE (N = NUMBER OF SITES)</i>													
2005-2007	285	0.5	0.0	0.0	0.1	0.2	0.3	0.4	0.6	0.7	0.8	1.0	1.5
2005	285	0.5	0.0	0.0	0.1	0.2	0.4	0.5	0.6	0.8	0.9	1.3	1.6
2006	285	0.5	0.0	0.0	0.1	0.2	0.3	0.4	0.6	0.7	0.8	1.2	1.4
2007	285	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.7	0.7	1.1	1.5
Winter (December - February)	285	0.6	0.0	0.0	0.2	0.2	0.4	0.5	0.7	0.9	1.1	1.5	1.6
Spring (March - May)	285	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.7	0.7	1.0	1.6
Summer (June - August)	285	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.6	0.7	1.1	1.5
Fall (September - November)	285	0.5	0.0	0.0	0.1	0.2	0.4	0.4	0.6	0.8	0.9	1.1	1.5
<i>STATISTICS FOR INDIVIDUAL CSAS/CBSAS (2005-2007) (N = NUMBER OF OBSERVATIONS)</i>													
Anchorage <sup>a</sup>	25,672	1.1	0.0	0.1	0.2	0.3	0.5	0.7	1.3	2.3	3.1	5.0	13.1
Atlanta	76,683	0.5	0.0	0.0	0.2	0.2	0.3	0.4	0.6	0.8	1.1	1.6	10.8
Boston	171,975	0.4	0.0	0.0	0.0	0.1	0.2	0.4	0.5	0.7	0.9	1.4	10.0
Denver	129,038	0.5	0.0	0.0	0.1	0.2	0.3	0.4	0.6	1.0	1.3	2.2	9.3
Houston	123,925	0.3	0.0	0.0	0.0	0.0	0.2	0.3	0.4	0.6	0.8	1.4	4.6
Los Angeles	592,960	0.5	0.0	0.0	0.0	0.1	0.2	0.3	0.6	1.0	1.4	2.3	8.4
New York	226,673	0.5	0.0	0.0	0.1	0.1	0.3	0.5	0.6	0.9	1.1	1.6	5.8
Phoenix	127,477	0.8	0.0	0.0	0.1	0.2	0.3	0.5	1.0	1.9	2.5	3.6	7.8
Pittsburgh	179,758	0.3	0.0	0.0	0.0	0.0	0.1	0.2	0.4	0.6	0.8	1.2	6.7
Seattle	25,818	0.8	0.0	0.1	0.2	0.3	0.4	0.6	0.9	1.3	1.6	2.5	5.9
St. Louis	77,142	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.7	0.9	1.4	5.7
Not in the 11 cities	5,449,251	0.5	0.0	0.0	0.0	0.1	0.2	0.4	0.6	0.9	1.2	2.1	39.0

<sup>a</sup>CO monitoring is only available for quarters 1 and 4; since monitoring data is not available year-round, Anchorage is not included in the nationwide statistics shown in this table.

1 Table 3-3 contains the distribution of hourly CO measurements reported to AQS for 2005-2007. All  
2 monitoring locations meeting the 75% data completeness criteria have been included in this table.  
3 Anchorage did not meet the data completeness criteria since its monitoring sites were required to report  
4 CO measurements during the first and fourth quarters of each year. Anchorage was included in the table,  
5 however, for an approximate comparison with the other CSAs and CBSAs reporting year-round  
6 measurements to AQS. (Anchorage was not, however, included in the national averages shown in the

1 table.) The nationwide mean, median, and interquartile range for 1-h measurements reported for  
2 2005-2007 were 0.5, 0.4 and 0.4 ppm, respectively, and these statistics did not change by more than  
3 0.1 ppm over the 3-year period. The largest recorded second-highest 1-h concentration, 26.3 ppm, for this  
4 period was reported in 2006 in Birmingham, AL (AQS site ID: 0107360004). The absolute highest 1-h  
5 concentration, 39 ppm, between 2005 and 2007, was reported in Ogden, UT (AQS site ID: 490570006).  
6 Concentrations are generally highest in the winter (December-February) and fall (September-November)  
7 and decrease on average during the spring (March-May) and summer (June-August).

8 Nationwide statistics pooled by site are listed in the center of Table 3-3 and illustrate the  
9 distribution of the site average CO concentrations recorded at the 285 monitoring sites for 2005-2007 (see  
10 Figure 3-6 for these sites). The site reporting the highest 3-year pooled 1-h avg CO concentration,  
11 1.5 ppm, was located in San Juan, Puerto Rico (AQS site ID: 721270003). The eleven individual  
12 CSAs/CBSAs discussed earlier are included in the table, none of which reported concentrations above the  
13 value of the 1-h NAAQS. Four of the eleven cities (Boston, Houston, Pittsburgh and St. Louis) had 95th  
14 percentile 1-h CO concentrations below 1 ppm; the 95th percentile concentrations for the remaining cities  
15 were below 3.1 ppm. Lack of year-round monitoring in Anchorage prevented a direct comparison with the  
16 other metropolitan regions. However, Anchorage exhibited a 1-h CO distribution shifted higher in  
17 concentration when compared to the U.S. average during fall or winter.

**Table 3-4. Distribution of 24-h avg CO concentration (ppm) derived from AQS data.**

	n	Mean	Min	Percentiles									Max
				1	5	10	25	50	75	90	95	99	
<i>NATIONWIDE STATISTICS (N = NUMBER OF OBSERVATIONS)</i>													
2005-2007	303,843	0.5	0.0	0.0	0.0	0.1	0.3	0.4	0.6	0.9	1.1	1.7	7.0
2005	101,184	0.5	0.0	0.0	0.0	0.1	0.3	0.4	0.6	0.9	1.1	1.8	5.8
2006	101,652	0.5	0.0	0.0	0.0	0.1	0.3	0.4	0.6	0.9	1.1	1.6	7.0
2007	101,007	0.4	0.0	0.0	0.0	0.1	0.2	0.4	0.5	0.8	1.0	1.6	6.9
Winter (December - February)	74,144	0.6	0.0	0.0	0.1	0.2	0.3	0.5	0.7	1.1	1.3	2.0	7.0
Spring (March - May)	77,317	0.4	0.0	0.0	0.0	0.1	0.2	0.4	0.5	0.7	0.9	1.4	6.4
Summer (June - August)	76,562	0.4	0.0	0.0	0.0	0.1	0.2	0.3	0.5	0.7	0.8	1.3	6.9
Fall (September - November)	75,820	0.5	0.0	0.0	0.0	0.1	0.3	0.4	0.6	0.9	1.1	1.7	5.8
<i>NATIONWIDE STATISTICS, POOLED BY SITE (N = NUMBER OF SITES)</i>													
2005-2007	285	0.5	0.0	0.0	0.1	0.2	0.3	0.4	0.6	0.7	0.8	1.0	1.5
2005	285	0.5	0.0	0.0	0.1	0.2	0.4	0.5	0.6	0.8	0.9	1.3	1.6
2006	285	0.5	0.0	0.0	0.1	0.2	0.3	0.4	0.6	0.7	0.8	1.2	1.4
2007	285	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.7	0.7	1.1	1.5
Winter (December - February)	285	0.6	0.0	0.0	0.2	0.2	0.4	0.5	0.7	0.9	1.1	1.5	1.6
Spring (March - May)	285	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.7	0.7	1.0	1.6
Summer (June - August)	285	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.6	0.7	1.1	1.5
Fall (September - November)	285	0.5	0.0	0.0	0.1	0.2	0.4	0.4	0.6	0.8	0.9	1.1	1.5
<i>STATISTICS FOR INDIVIDUAL CSAS/CBSAS (2005-2007) (N = NUMBER OF OBSERVATIONS)</i>													
Anchorage <sup>a</sup>	1,074	1.1	0.0	0.2	0.2	0.4	0.6	0.9	1.4	1.9	2.4	3.3	4.6
Atlanta	3,229	0.5	0.0	0.1	0.2	0.2	0.3	0.4	0.6	0.8	0.9	1.2	1.6
Boston	7,446	0.4	0.0	0.0	0.1	0.1	0.3	0.4	0.5	0.7	0.8	1.1	2.2
Denver	5,363	0.5	0.0	0.1	0.2	0.2	0.3	0.5	0.6	0.9	1.1	1.5	2.3
Houston	5,188	0.3	0.0	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.6	0.9	1.9
Los Angeles	25,803	0.5	0.0	0.0	0.1	0.1	0.2	0.4	0.6	1.0	1.2	1.7	3.8
New York	9,513	5.1	0.0	0.0	0.1	0.2	0.4	0.5	0.6	0.8	1.0	1.3	2.5
Phoenix	5,348	0.8	0.0	0.1	0.2	0.3	0.4	0.6	1.1	1.6	1.9	2.5	3.4
Pittsburgh	7,497	0.3	0.0	0.0	0.0	0.0	0.1	0.2	0.4	0.6	0.7	1.0	1.9
Seattle	1,079	0.8	0.1	0.2	0.3	0.4	0.5	0.7	0.9	1.2	1.4	1.8	2.4
St. Louis	3,216	0.4	0.0	0.0	0.1	0.2	0.3	0.4	0.5	0.7	0.8	1.0	1.9
Not in the 11 cities	230,161	0.5	0.0	0.0	0.0	0.1	0.2	0.4	0.6	0.8	1.1	1.6	7.0

<sup>a</sup>CO monitoring is only available for quarters 1 and 4; since monitoring data is not available year-round, Anchorage is not included in the nationwide statistics shown in this table.

1 Table 3-4 contains the distribution of 24-h avg CO concentrations derived from the 1-h  
2 concentrations reported to AQS and summarized in Table 3-3. The nationwide mean, median, and  
3 interquartile range for 24-h avg for 2005-2007 were 0.5, 0.4 and 0.3 ppm, respectively. These were  
4 similar to those for the 1-h values. The maximum 24-h avg concentration in these years, 7 ppm, was  
5 reported in Birmingham, AL (AQS site ID: 010736004).

**Table 3-5. Distribution of 1-h daily max CO concentration (ppm) derived from AQS data.**

	n	Mean	Min	Percentiles									
				1	5	10	25	50	75	90	95	99	Max
<i>NATIONWIDE STATISTICS (N = NUMBER OF OBSERVATIONS)</i>													
2005-2007	303,843	0.9	0.0	0.0	0.1	0.3	0.4	0.7	1.2	1.8	2.4	3.8	39.0
2005	101,184	1.0	0.0	0.0	0.2	0.3	0.5	0.8	1.3	2.0	2.6	4.1	22.3
2006	101,652	0.9	0.0	0.0	0.1	0.3	0.4	0.7	1.2	1.9	2.4	3.9	35.3
2007	101,007	0.8	0.0	0.0	0.1	0.2	0.4	0.7	1.1	1.7	2.1	3.4	39.0
Winter (December - February)	74,144	1.2	0.0	0.0	0.2	0.3	0.5	0.9	1.6	2.5	3.1	4.7	20.0
Spring (March - May)	77,317	0.8	0.0	0.0	0.1	0.3	0.4	0.7	1.0	1.6	2.0	3.0	35.3
Summer (June - August)	76,562	0.7	0.0	0.0	0.1	0.2	0.4	0.6	0.9	1.3	1.6	2.5	39.0
Fall (September - November)	75,820	1.0	0.0	0.0	0.2	0.3	0.5	0.8	1.3	2.0	2.5	3.8	24.1
<i>NATIONWIDE STATISTICS, POOLED BY SITE (N = NUMBER OF SITES)</i>													
2005-2007	285	0.9	0.1	0.1	0.3	0.5	0.6	0.8	1.1	1.5	1.7	2.3	3.9
2005	285	1.0	0.1	0.1	0.4	0.5	0.7	0.9	1.2	1.6	2.0	2.5	3.7
2006	285	0.9	0.1	0.1	0.3	0.5	0.6	0.9	1.1	1.6	1.8	2.3	4.8
2007	285	0.8	0.1	0.1	0.3	0.4	0.6	0.8	1.0	1.4	1.6	2.0	3.1
Winter (December - February)	285	1.2	0.0	0.1	0.4	0.6	0.8	1.0	1.5	2.1	2.5	3.4	4.1
Spring (March - May)	285	0.8	0.1	0.1	0.3	0.4	0.6	0.8	1.0	1.3	1.5	2.1	4.0
Summer (June - August)	285	0.7	0.0	0.1	0.2	0.3	0.5	0.6	0.8	1.1	1.3	2.2	3.3
Fall (September - November)	285	1.0	0.1	0.1	0.3	0.5	0.7	0.9	1.2	1.7	2.0	2.4	4.1
<i>STATISTICS FOR INDIVIDUAL CSAS/CBSAS (2005-2007) (N = NUMBER OF OBSERVATIONS)</i>													
Anchorage <sup>a</sup>	1,074	2.6	0.0	0.3	0.6	0.8	1.3	2.2	3.5	5.0	6.1	7.6	13.1
Atlanta	3,229	0.8	0.0	0.2	0.3	0.3	0.4	0.7	1.1	1.4	1.7	2.2	10.8
Boston	7,446	0.7	0.0	0.1	0.2	0.3	0.4	0.6	0.9	1.2	1.6	2.6	10.0
Denver	5,363	1.2	0.1	0.2	0.4	0.5	0.7	1.0	1.5	2.2	2.7	3.9	9.3
Houston	5,188	0.7	0.0	0.0	0.1	0.2	0.4	0.6	0.8	1.3	1.7	2.6	4.6
Los Angeles	25,803	1.0	0.0	0.1	0.2	0.3	0.5	0.8	1.3	2.0	2.6	4.0	8.4
New York	9,513	0.9	0.0	0.1	0.2	0.4	0.6	0.8	1.1	1.5	1.8	2.5	5.8
Phoenix	5,348	1.9	0.0	0.3	0.5	0.6	0.9	1.6	2.5	3.5	4.1	5.3	7.8
Pittsburgh	7,497	0.6	0.0	0.0	0.0	0.1	0.2	0.5	0.8	1.1	1.4	2.0	6.7
Seattle	1,079	1.5	0.2	0.4	0.5	0.7	0.9	1.3	1.8	2.4	2.9	4.3	5.9
St. Louis	3,216	0.8	0.0	0.1	0.3	0.4	0.5	0.6	0.9	1.3	1.7	2.7	5.7
Not in the 11 cities	230,161	0.9	0.0	0.0	0.1	0.2	0.4	0.7	1.2	1.8	2.4	3.8	39.0

<sup>a</sup>CO monitoring is only available for quarters 1 and 4; since monitoring data is not available year-round, Anchorage is not included in the nationwide statistics shown in this table.

1 Table 3-5 contains the distribution of 1-h daily max CO concentrations derived from 1-h values  
2 reported to AQS for all monitors meeting the inclusion criteria described earlier. The nationwide mean,  
3 median, and interquartile range for 1-h daily max concentrations reported for 2005-2007 were 0.9, 0.7 and  
4 0.8 ppm, respectively.

**Table 3-6. Distribution of 8-h daily max CO concentration (ppm) derived from AQS data.**

	n	Mean	Min	Percentiles									
				1	5	10	25	50	75	90	95	99	Max
<i>NATIONWIDE STATISTICS (N = NUMBER OF OBSERVATIONS)</i>													
2005-2007	303,843	0.7	0.0	0.3	0.3	0.3	0.3	0.5	0.8	1.3	1.7	2.6	10.9
2005	101,184	0.7	0.0	0.3	0.3	0.3	0.3	0.6	0.9	1.4	1.8	2.8	9.7
2006	101,652	0.7	0.0	0.3	0.3	0.3	0.3	0.5	0.8	1.3	1.7	2.6	9.8
2007	101,007	0.6	0.0	0.3	0.3	0.3	0.3	0.5	0.8	1.2	1.5	2.3	10.9
Winter (December - February)	74,144	0.9	0.0	0.3	0.3	0.3	0.4	0.7	1.1	1.7	2.1	3.2	9.8
Spring (March - May)	77,317	0.6	0.0	0.3	0.3	0.3	0.3	0.5	0.7	1.1	1.3	2.0	9.6
Summer (June - August)	76,562	0.5	0.0	0.3	0.3	0.3	0.3	0.4	0.6	0.9	1.1	1.7	10.9
Fall (September - November)	75,820	0.7	0.0	0.3	0.3	0.3	0.3	0.6	0.9	1.4	1.8	2.7	9.0
<i>NATIONWIDE STATISTICS, POOLED BY SITE (N = NUMBER OF SITES)</i>													
2005-2007	285	0.7	0.2	0.3	0.3	0.4	0.5	0.6	0.8	1.0	1.2	1.7	2.1
2005	285	0.7	0.3	0.3	0.3	0.4	0.5	0.6	0.9	1.1	1.4	1.9	2.2
2006	285	0.7	0.2	0.3	0.3	0.4	0.5	0.6	0.8	1.1	1.2	1.8	2.4
2007	285	0.6	0.2	0.3	0.3	0.4	0.5	0.6	0.7	1.0	1.1	1.6	2.0
Winter (December - February)	285	0.9	0.2	0.3	0.4	0.4	0.6	0.8	1.1	1.4	1.7	2.4	2.6
Spring (March - May)	285	0.6	0.2	0.3	0.3	0.4	0.4	0.5	0.7	0.9	1.1	1.6	2.2
Summer (June - August)	285	0.5	0.2	0.3	0.3	0.3	0.4	0.5	0.6	0.8	0.9	1.5	2.0
Fall (September - November)	285	0.7	0.2	0.3	0.3	0.4	0.5	0.6	0.9	1.2	1.3	1.8	2.2
<i>STATISTICS FOR INDIVIDUAL CSAS/CBSAS (2005-2007) (N = NUMBER OF OBSERVATIONS)</i>													
Anchorage <sup>a</sup>	1,074	1.7	0.3	0.3	0.4	0.6	0.9	1.5	2.3	3.3	3.9	5.0	6.5
Atlanta	3,229	0.6	0.0	0.2	0.2	0.3	0.4	0.5	0.8	1.1	1.3	1.7	2.5
Boston	7,446	0.6	0.3	0.3	0.3	0.3	0.3	0.5	0.7	0.9	1.1	1.8	5.8
Denver	5,363	0.8	0.3	0.3	0.3	0.3	0.5	0.7	1.0	1.4	1.8	2.4	3.4
Houston	5,188	0.5	0.3	0.3	0.3	0.3	0.3	0.4	0.6	0.9	1.1	1.7	3.3
Los Angeles	25,803	0.7	0.3	0.3	0.3	0.3	0.3	0.6	0.9	1.5	1.8	2.7	6.2
New York	9,513	0.7	0.3	0.3	0.3	0.3	0.4	0.6	0.9	1.2	1.4	1.8	3.0
Phoenix	5,348	1.3	0.3	0.3	0.3	0.4	0.6	1.0	1.8	2.5	3.0	3.8	5.8
Pittsburgh	7,497	0.5	0.3	0.3	0.3	0.3	0.3	0.3	0.6	0.9	1.0	1.5	3.7
Seattle	1,079	1.1	0.3	0.3	0.4	0.5	0.7	1.0	1.4	1.8	2.2	3.2	4.0
St. Louis	3,216	0.6	0.3	0.3	0.3	0.3	0.3	0.5	0.7	0.9	1.2	1.9	4.2
Not in the 11 cities	230,161	0.7	0.0	0.3	0.3	0.3	0.3	0.5	0.8	1.3	1.6	2.5	10.9

<sup>a</sup>CO monitoring is only available for quarters 1 and 4; since monitoring data is not available year-round, Anchorage is not included in the nationwide statistics shown in this table.

1 Table 3-6 contains the distribution of 8-h daily max concentrations derived from the 1-h CO  
2 concentrations reported to AQS. This was done by first calculating the average concentration for each  
3 successive 8-h period, thereby producing 24 8-h avg per day. The maximum of these values for a given  
4 monitor within a given day (midnight-to-midnight) was used as the 8-h daily max statistic for that  
5 monitor and day. The nationwide mean, median, and interquartile range for 8-h daily max concentrations



1 reported for 2005-2007 were 0.7, 0.5, and 0.5 ppm, respectively. The highest 8-h daily max concentration,  
2 10.9 ppm, was recorded at a monitor located five miles north of Newkirk, OK (AQS site ID: 400719010).

### 3.5.1.2. Urban Scale

3 This section covers urban variability in CO concentrations reported to AQS at the individual  
4 CSA/CBSA level. Phoenix, AZ and Pittsburgh, PA were selected for this assessment to illustrate the  
5 variability in CO concentrations measured across contrasting metropolitan regions. Information on the  
6 other nine cities evaluated for this assessment is included in Appendix A. Maps of the Phoenix CBSA and  
7 Pittsburgh CSA shown in Figures 3-13 and 3-15, respectively, illustrate the location of all CO monitors  
8 included in the analyses described above. Letters on the maps identify the individual monitor locations  
9 and correspond with the letters provided in the accompanying box plots (Figures 3-14 and 3-16) and pair-  
10 wise comparison tables (Tables 3-7 and 3-8). The box plots for each monitor include the hourly CO  
11 concentration median and interquartile range with whiskers extending from the 5th to the 95th percentile.  
12 Data from 2005-2007 were used to generate the box plots which are stratified by season as follows:  
13 1 = winter (December-February), 2 = spring (March-May), 3 = summer (June-August), and 4 = fall  
14 (September-November). The comparison tables include the Pearson correlation coefficient ( $r$ ), the 90th  
15 percentile of the absolute difference in concentrations (P90) in ppm, the coefficient of divergence (COD)  
16 and the straight-line distance between monitor pairs ( $d$ ) in km. The COD provides an indication of the  
17 variability across the monitoring sites within each CSA/CBSA and is defined as follows:

$$COD_{jk} = \sqrt{\frac{1}{p} \sum_{i=1}^p \left( \frac{X_{ij} - X_{ik}}{X_{ij} + X_{ik}} \right)^2}$$

Equation 3-1

18 where  $X_{ij}$  and  $X_{ik}$  represent the observed hourly concentrations for time period  $i$  at sites  $j$  and  $k$ , and  $p$  is  
19 the number of paired hourly observations. A COD of 0 indicates there are no differences between  
20 concentrations at paired sites (spatial homogeneity), while a COD approaching 1 indicates extreme spatial  
21 heterogeneity. Similar maps, box plots and comparison tables for the nine remaining CSAs/CBSAs  
22 described earlier in this chapter along with Phoenix and Pittsburgh are included in Annex A.

# Phoenix Core Based Statistical Area

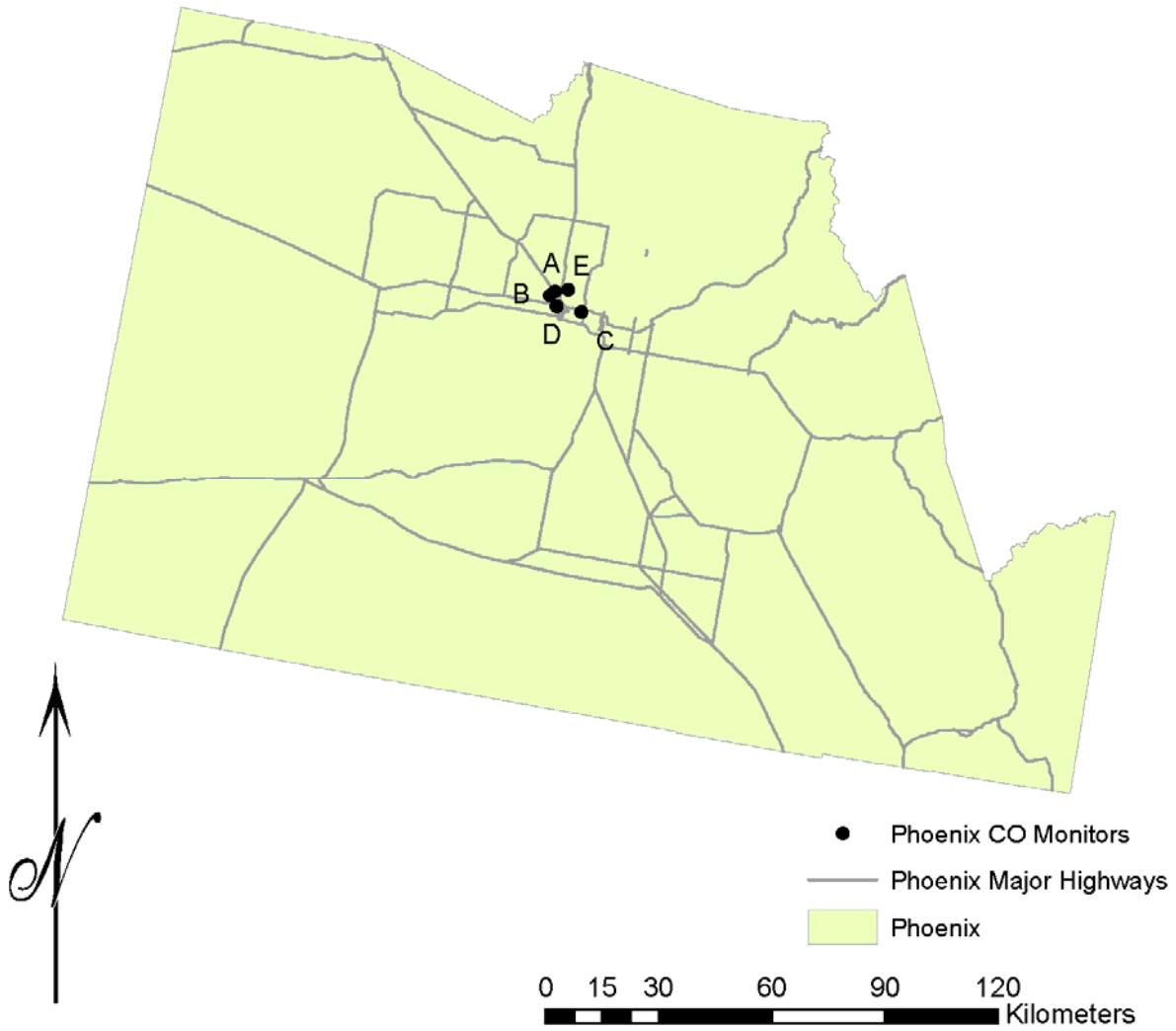
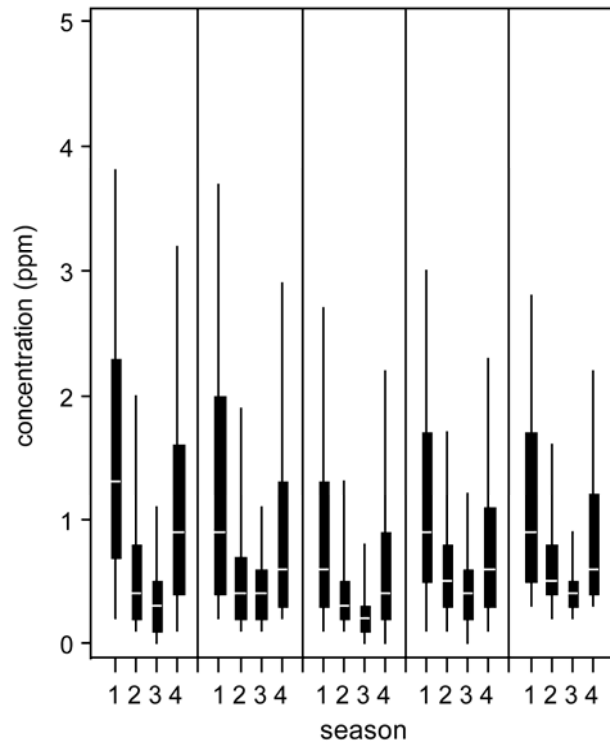


Figure 3-13. Map of CO monitor locations with AQS Site IDs for Phoenix, AZ.

	A	B	C	D	E
Mean	0.93	0.84	0.58	0.76	0.79
Obs	25382	25589	25657	25414	25435
SD	0.95	0.88	0.64	0.72	0.64



**Figure 3-14.** Box plots illustrating the seasonal distribution of 2005-2007 hourly CO concentrations in Phoenix, AZ. Note: 1 = winter, 2 = spring, 3 = summer, and 4 = fall on the x-axis. The box plots for each monitor include the hourly CO concentration median and interquartile range with whiskers extending from the 5th to the 95th percentile.

**Table 3-7. Table of inter-sampler comparison statistics, as defined in the text, including Pearson r, P90 (ppm), COD and d (km) for each pair of hourly CO monitors reporting to AQS for 2005-2007 in Phoenix, AZ.**

Phoenix					
	A	B	C	D	E
A	1.00	0.89	0.80	0.86	0.84
	0.0	0.7	1.1	0.8	0.9
	0.0	0.37	0.43	0.39	0.37
	0.0	1.6	8.9	3.9	3.5
B		1.00	0.81	0.88	0.89
		0.0	0.9	0.6	0.7
		0.0	0.38	0.34	0.24
		0.0	9.4	3.4	4.9
C			1.00	0.81	0.85
			0.0	0.7	0.6
			0.0	0.41	0.36
			0.0	6.6	6.8
D				1.00	0.83
				0.0	0.6
				0.0	0.33
				0.0	5.2
E					1.00
					0.0
					0.0
					0.0

Legend	
r	1.00
P90	0.0
COD	0.0
d	0.0

# Pittsburgh Combined Statistical Area

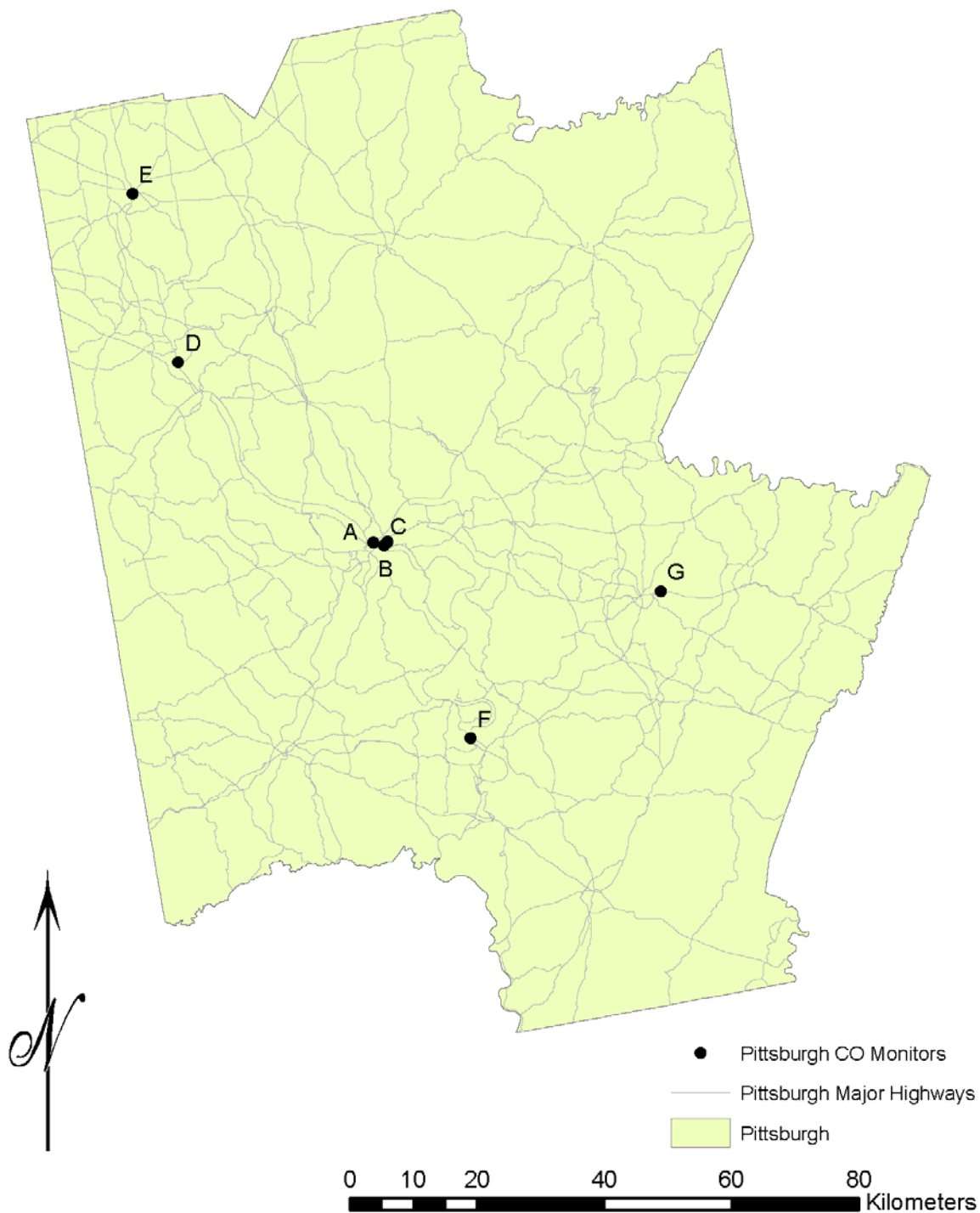
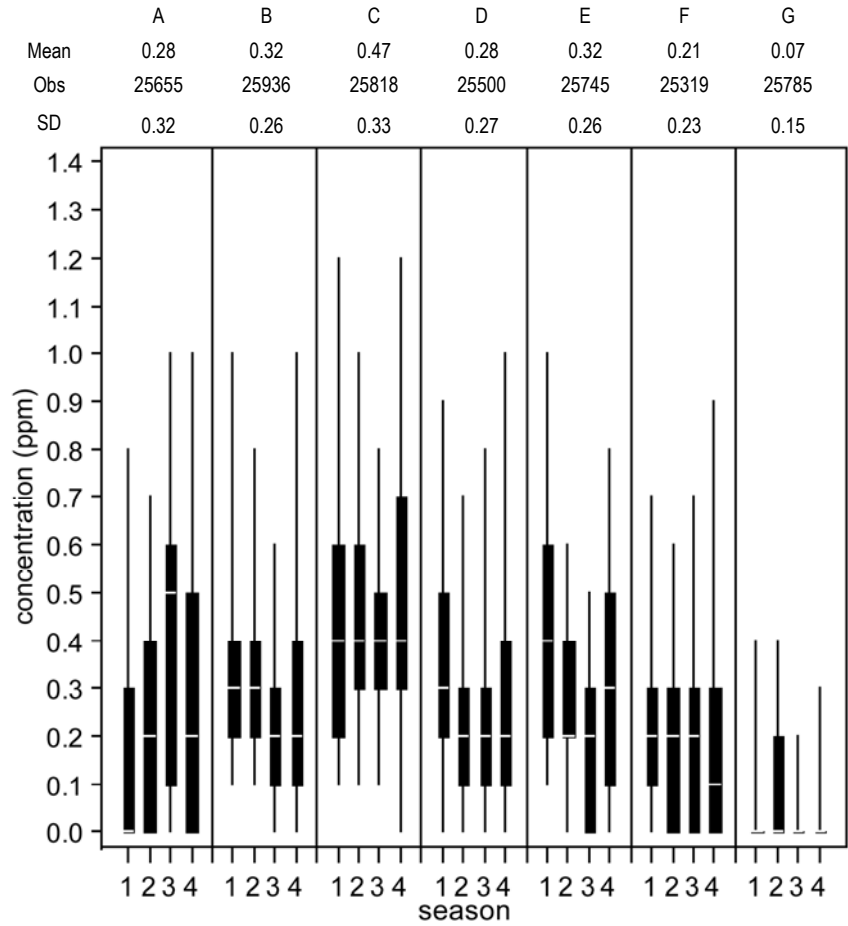


Figure 3-15. Map of CO monitor locations with AQS Site IDs for Pittsburgh, PA.



**Figure 3-16. Box plots illustrating the seasonal distribution of 2005-2007 hourly CO concentrations in Pittsburgh, PA. Note: 1 = winter, 2 = spring, 3 = summer, and 4 = fall on the x-axis. Median interquartile range is shown with whiskers extending from the 5th to the 95th percentile.**

**Table 3-8. Table of inter-sampler comparison statistics, as defined in the text, including Pearson r, P90 (ppm), COD and d (km) for each pair of hourly CO monitors reporting to AQS for 2005-2007 in Pittsburgh, PA.**

Pittsburgh							
	A	B	C	D	E	F	G
A	1.00	0.52	0.43	0.05	0.11	0.42	0.18
	0.0	0.5	0.6	0.7	0.6	0.5	0.7
	0.0	0.69	0.68	0.74	0.72	0.73	0.88
	0.0	2.2	1.8	41.9	66.8	34.4	45.8
B		1.00	0.73	0.30	0.42	0.29	0.35
		0.00	0.40	0.50	0.40	0.50	0.50
		0.00	0.39	0.54	0.51	0.62	0.86
		0.00	0.70	43.40	68.00	33.60	43.70
C			1.00	0.20	0.39	0.25	0.30
			0.0	0.7	0.6	0.7	0.8
			0.0	0.56	0.51	0.65	0.88
			0.0	43.4	68.2	33.3	44.1
D				1.00	0.16	0.09	0.02
				0.0	0.6	0.6	0.7
				0.0	0.57	0.69	0.87
				0.0	27.5	75.0	84.1
E					1.00	0.26	0.21
					0.0	0.5	0.6
					0.0	0.68	0.87
					0.0	101.0	104.1
F						1.00	0.09
						0.0	0.5
						0.0	0.90
						0.0	37.8
G	<b>Legend</b>						
	r						1.00
	P90						0.0
	COD						0.0
	d						0.0

1  
2 The Phoenix CBSA in Figure 3-13 incorporates an area of 37,786 km<sup>2</sup> with the CO monitors  
3 densely packed near the urban center. The maximum straight-line distance between CO monitors in the  
4 Phoenix CBSA is 9.4 km. By contrast, the Pittsburgh CSA in Figure 3-15 is less than half the size, having  
5 an area of 14,627 km<sup>2</sup>, but the monitors are farther apart with a maximum straight-line distance between  
6 them of 104.1 km. Of the eleven CSAs/CBSAs investigated, Phoenix has among the highest correlations  
7 across monitors, while Pittsburgh has among the lowest. This discrepancy is partly due to the close  
8 proximity of the monitors as noted here, but horizontal distance alone does not explain the range of  
9 correlations observed between monitors. For example, the three monitors located in the downtown urban  
10 core of Pittsburgh (sites A, B and C; AQS site IDs 420030010, 420030031 and 420030038; maximum

1 distance = 2.2 km) exhibited considerably lower correlations ( $0.43 \leq r \leq 0.73$ ) than any of the monitor  
2 pairs in Phoenix ( $0.80 \leq r \leq 0.89$ ), located up to 9.4 km apart.

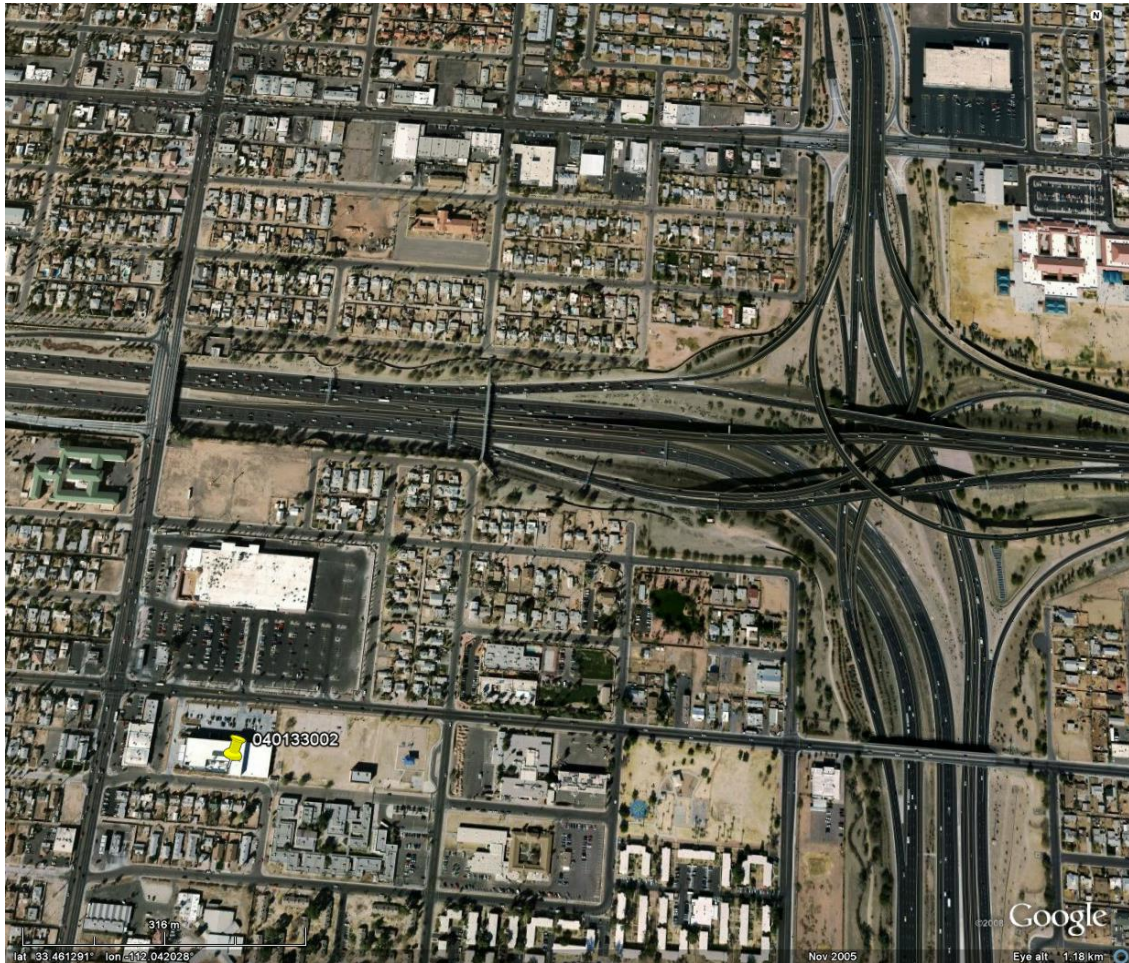
3 The correlation structures for measurements at the monitors in each of the eleven CSAs/CBSAs  
4 included in this analysis reveal a wide range of response between monitors in each city and among the  
5 cities judged against each other. While this wide range is produced by the interactions of many physical  
6 and chemical elements, the location of each monitor and the uniqueness of its immediate surroundings  
7 can often explain much of the agreement or lack thereof. Compare, for example, Figures 3-17 and 3-18,  
8 the aerial views of monitors C (AQS site ID 040133002) and E (AQS site ID 040139997), respectively, in  
9 Phoenix. Monitor C has a correlation value with monitor E of 0.85 even though C is less than 1 km from  
10 the intersection of two major highways and E is surrounded by neighborhood residences. Because all  
11 other correlations of monitors in Phoenix are also  $>0.8$ , the network of CO monitors there appears to have  
12 captured the behavior of emitted and transported CO with high fidelity. In contrast, monitors in Pittsburgh  
13 show a low degree of correlation, with only two values  $>0.5$  and eight values  $\leq 0.2$ . The topography of  
14 Pittsburgh is mountainous and therefore more complex than that of Phoenix, and this could account for  
15 some of the disparity across space. But in addition, several CO monitors in Pittsburgh are sited at  
16 distances far from the urban core, making them less likely to correlate with those monitors in the high-  
17 density urban core. Figure 3-19 depicts the three monitors in the Pittsburgh urban core (monitor A: AQS  
18 site ID 420030010, monitor B: AQS site ID 420030031, and monitor C: AQS site ID 420030038), which  
19 demonstrated the highest correlation values in this CSA for 2005-2007. Figure 3-20 shows the location of  
20 a more distant monitor, monitor F (AQS site ID 421250005), which does not correlate well with the three  
21 monitors located downtown, very likely because it is located at a great distance from any strong CO  
22 sources in the downtown urban core. This analysis demonstrates that agreement between monitors on an  
23 urban scale is a complex function of monitor siting, location to sources, geography, and  
24 micrometeorology.

25 In addition to high correlations, Phoenix exhibits similar average concentration levels across  
26 monitors with  $0.24 \leq \text{COD} \leq 0.43$ ; see Table 3-7. The seasonal patterns shown in Figure 3-14 also  
27 demonstrate homogeneity across monitors for Phoenix. By contrast, Pittsburgh had more variability  
28 between monitored values at the sites shown in Figure 3-16 with the range in COD shifted upward  
29 ( $0.39 \leq \text{COD} \leq 0.90$ ); see Table 3-8. For the three sites located in the downtown urban core of Pittsburgh  
30 (sites A, B and C), the COD is as high as 0.69, indicating greater spatial variability in CO concentrations  
31 in Pittsburgh, even for these more proximal monitors.

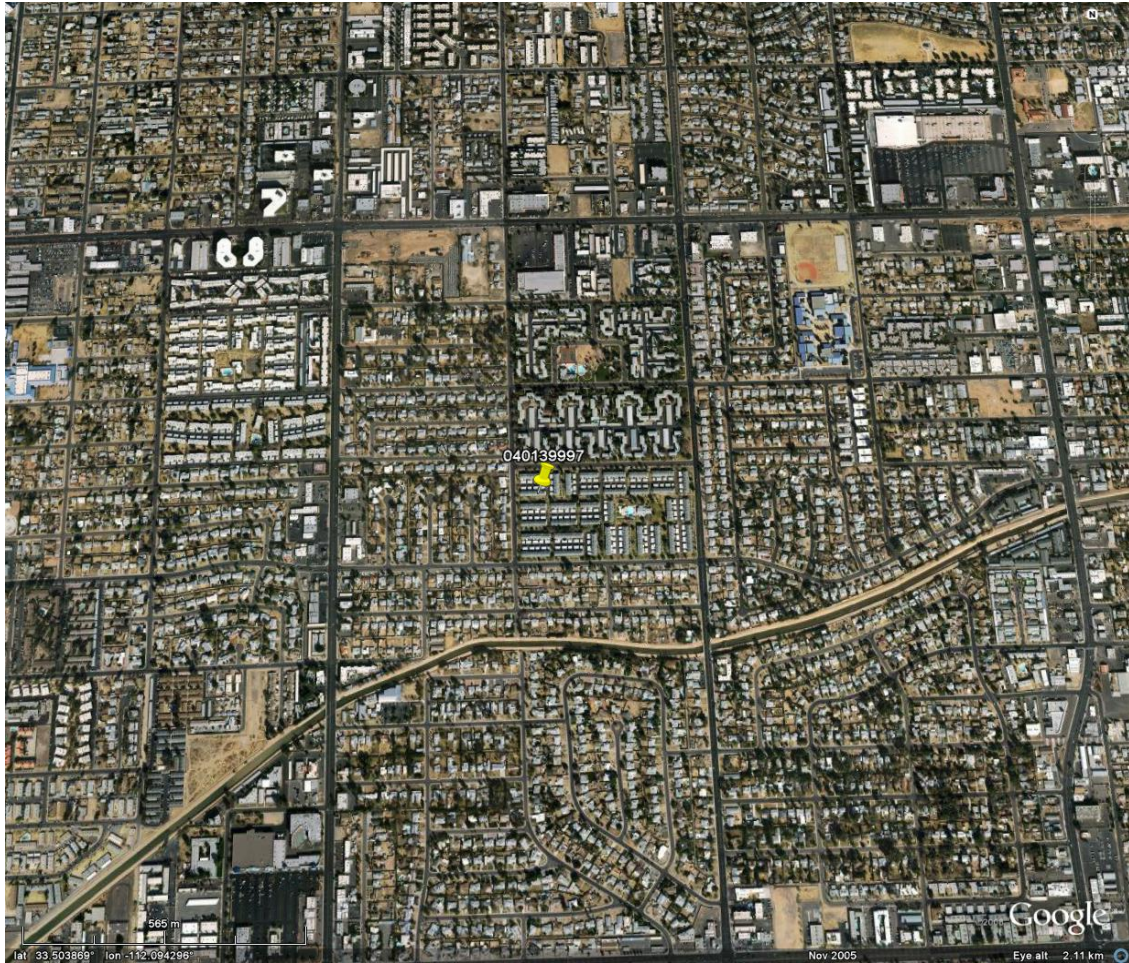
32 The information in Table 3-8 should be used with some caution since many of the reported  
33 concentrations for the years 2005-2007 are very near or even well below the monitors' stated lowest  
34 detection limits. Because ambient concentrations are now in large part very near to the FRM detection  
35 limit of 1.0 ppm and the coarsely reported measurement resolution is 0.1 ppm, the comparison statistics  
36 shown in these tables might be biased to show a specious heterogeneity in the box plots. These cautions



1 are especially important for interpreting data from site G in Pittsburgh and selected other sites in other  
2 metropolitan regions where one or more monitors consistently reported values below the stated detection  
3 limit. Such monitors include Houston site D, AQS site ID 482011035, and New York site I, AQS site ID  
4 361030009, included in Annex A.



**Figure 3-17. Aerial view of the location of CO monitor C. AQS ID 040133002 (marked by the yellow pin) in Phoenix, AZ, depicting its proximity to major roadways and neighborhood residences.**



**Figure 3-18.** Aerial view of the location of CO monitor E. AQS ID 040139997 (marked by the yellow pin) in Phoenix, AZ, depicting its proximity to neighborhood residences and secondary surface roads.



**Figure 3-19.** Aerial view of the location of CO monitors A, B, and C. AQS IDs 420030010, 420030031, and 420030038 (marked by yellow pins) in Pittsburgh, PA, depicting their proximity to major roadways and areas of high commercial density.



**Figure 3-20. Aerial view of the location of CO monitor F. AQS ID 421250005 (marked by the yellow pin) in the Pittsburgh, PA CSA, depicting its proximity to secondary surface roads and water ways.**

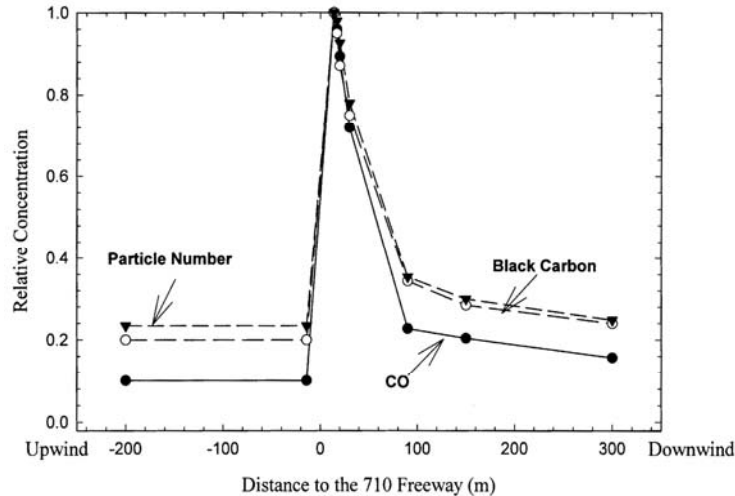
### **3.5.1.3. Neighborhood Scale**

1            Roadway density is an important determinant of the spatial distribution of CO concentrations in  
2 urban areas. Mobile sources are the largest single source of CO, and their abundance and density affect  
3 the magnitude of CO production. Urban topography around roadways also affects CO transport and  
4 dispersion (e.g., Mfula et al., 2005; Rigby and Toumi, 2008). CO concentration gradients have been  
5 observed in near-road environments (Baldauf et al., 2008; Pirjola et al., 2006; Zhu et al., 2002) such that  
6 CO concentrations are often higher on or near roadways compared with upwind sites, and can decrease  
7 with increasing distance from the road. For example, Zhu et al. (2002) found that on-road CO  
8 concentrations were approximately 10 times higher than at an upwind monitoring site, as shown in Figure  
9 3-21. At 30 m downwind from the road, relative CO concentrations were decreased to 7 times above  
10 upwind levels, and were approximately 2.5 times the upwind levels at 100 m. Concentrations continued to

1 decrease and were still somewhat higher than upwind levels at the final monitoring site 300 m away.  
2 Other traffic-related pollutants (BC, particle number) followed a similar pattern, although with different  
3 downwind:upwind ratios. This indicates that near-road and on-road exposures are important in CO  
4 exposure assessment.

5 Field measurements, computational modeling, and wind tunnel experiments have shown that  
6 roadway design and roadside features can affect CO and other pollutant concentrations near roadways.  
7 Field measurements reported by Baldauf et al. (2008) indicated that noise barriers could reduce near-road  
8 pollutant concentrations by as much as 50 percent, although this effect was highly dependent on  
9 meteorological conditions. This study also showed that the presence of mature vegetation, where the noise  
10 barrier further reduced concentrations and flattened the concentration gradient away from the road. Urban  
11 dispersion and wind-field modeling by Bowker et al. (2007) also demonstrated the influence of noise  
12 barriers and vegetation on the concentrations and spatial variability of inert pollutants emitted from traffic  
13 sources. Wind tunnel work reported by Baldauf et al. (In press) demonstrated the effects of noise barriers  
14 as well as roadway design characteristics, such as the presence of cut and elevated roadway segments.  
15 These results indicated that cut sections reduced concentrations and altered the gradient away from the  
16 road for inert gases emitted from traffic sources such as CO. These results showed similar concentrations  
17 as Zhu et al. (2002) for roadway segments at-grade with no obstructions to air flow as well as elevated  
18 roadway segments with fill conditions.

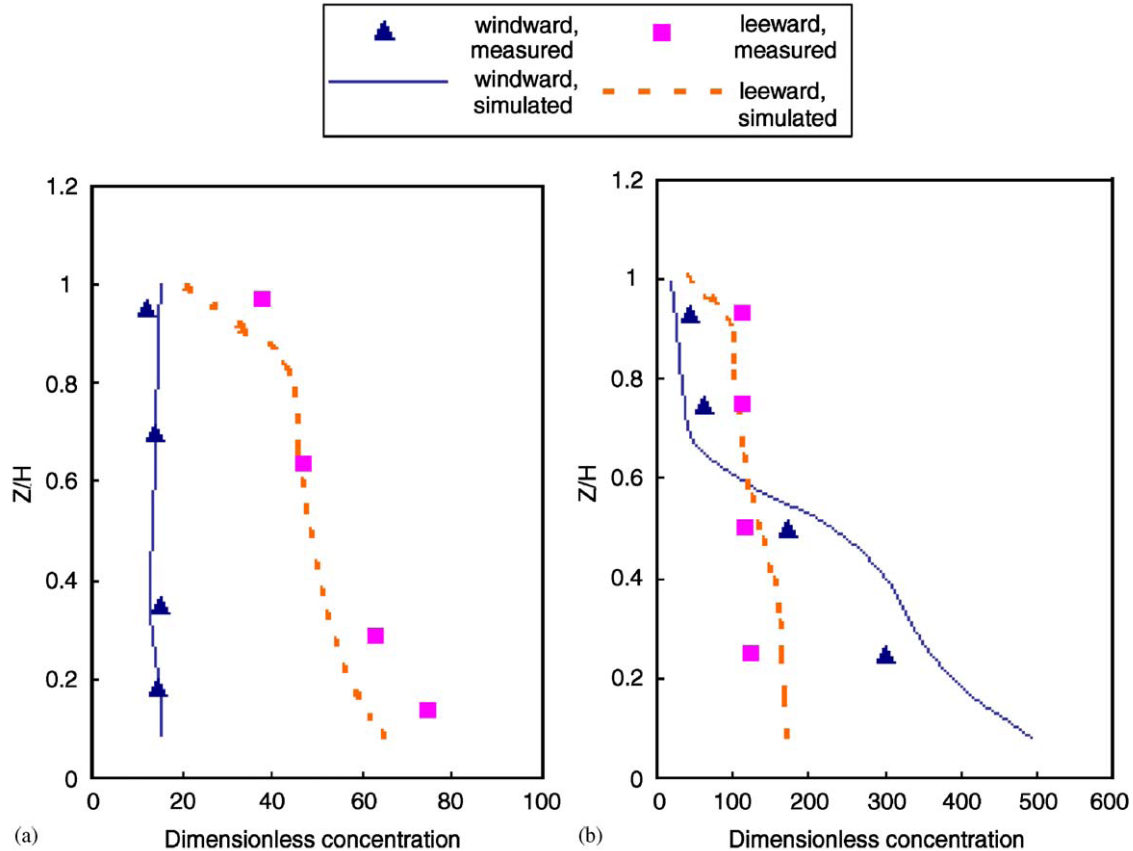
19 The geometry of urban street canyons has a profound effect on the distribution of CO  
20 concentrations on a micro-scale. A number of studies have performed computational and wind tunnel  
21 modeling of street canyons using nonreactive tracers and demonstrated the potential variability in  
22 concentration within a canyon (e.g., Borrego et al., 2006; Chang and Meroney, 2003; Kastner-Klein and  
23 Plate, 1999; So et al., 2005; Xiaomin et al., 2006). Because CO is a pollutant with very low reactivity on  
24 urban and regional scales, results from these models are directly relevant to CO concentration  
25 distributions in street canyons. Influential parameters include canyon height to width ratio (H/W), source  
26 positioning, wind speed and direction, building shape, and upstream configuration of buildings.  
27 Figure 3-22 shows dimensionless concentrations obtained from wind tunnel and computational fluid  
28 dynamics simulations of tracer gas transport and dispersion in an infinitely long street canyon with a line  
29 source centered at the bottom of the canyon (Xiaomin et al., 2006). When the canyon height was equal to  
30 the street width (typical of moderate density suburban or urban fringe residential neighborhoods) and  
31 lower background wind speed existed, concentrations on the leeward (downwind) canyon wall were four  
32 times those of the windward (upwind) wall near ground level. When the canyon height was twice the  
33 street width (typical of higher-density cities) and background winds were somewhat higher, near ground-  
34 level concentrations on the windward canyon wall were roughly three times higher than those measured at  
35 the leeward wall. These results suggest that the magnitude of microscale CO concentrations may vary by  
36 factors of three or four times at different locations within a street canyon.



Source: Zhu et al. (2002)

**Figure 3-21. Relative concentrations of CO and copollutants at various distances from the 710 freeway in Los Angeles.**

1 In a multi-site survey of curbside CO concentration in London, U.K., Croxford and Penn (1998)  
 2 observed differences in concentration related to the side of the street on which the monitor was positioned  
 3 relative to the wind direction. Bogo et al. (2001) measured CO concentrations in a street canyon with  
 4 building height-street width ratio of 1 in Buenos Aires, Argentina using a continuous CO monitor. Similar  
 5 to the Xiaomin et al. (2006) simulation results for a height-width ratio of 1, Bogo et al. (2001) observed  
 6 slightly higher leeward concentrations than windward concentrations within the canyon, where  
 7 recirculating airflow inside the canyon causes pollutants to collect in higher concentration on one side.  
 8 For part of this study, they estimated aggregated emissions from several fixed sources, including thermal  
 9 power stations and home heating, and from traffic. Mobile source emissions were estimated to be more  
 10 than 800 times higher than fixed source emissions given assumptions regarding vehicle fleet. Moreover,  
 11 they estimated that up to 30% of the measured CO concentrations came from within the local street  
 12 canyon in which measurements were made.

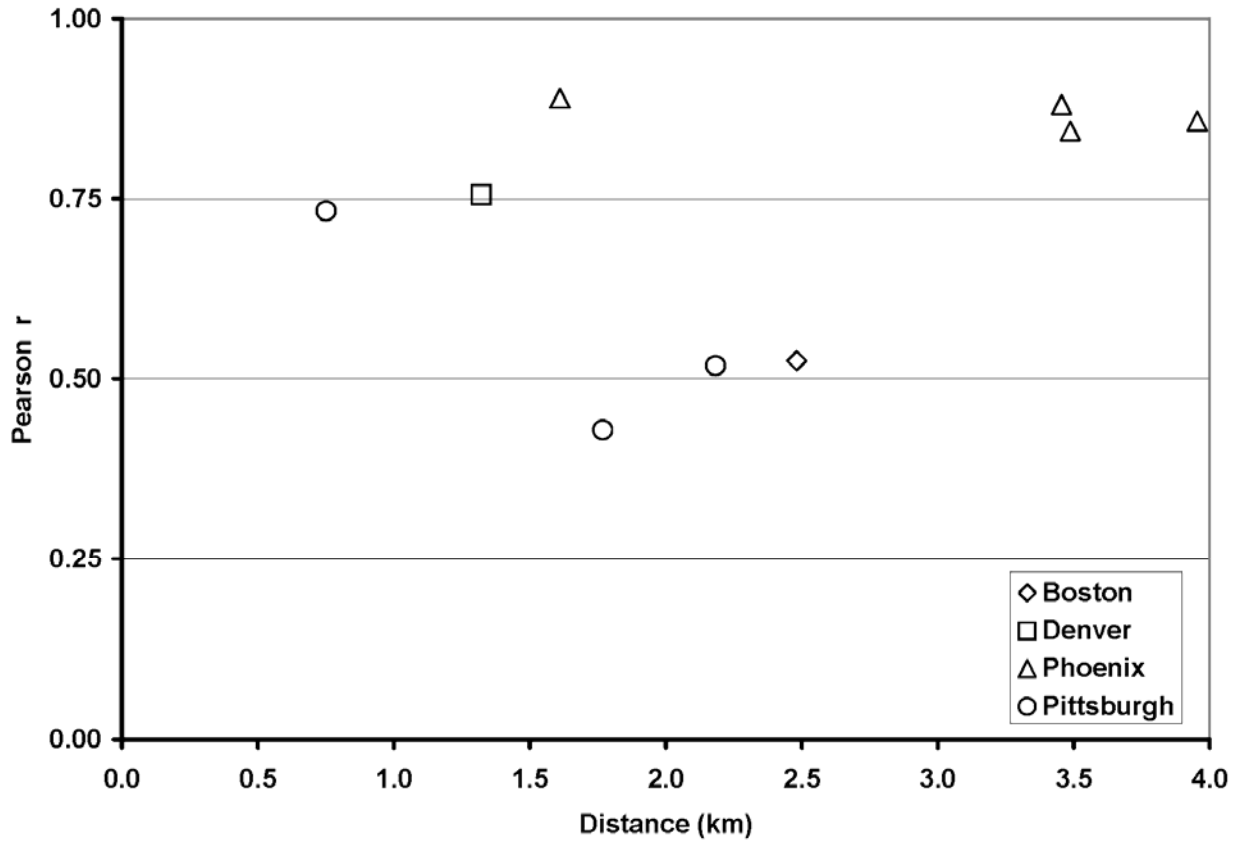


Source: Xiaomin et al. (2006)

**Figure 3-22.** Dimensionless tracer gas concentration as a function of elevation at windward and leeward locations and street canyon aspect ratios ( $H/W$ ).  $Z/H$  is the elevation of the measurement ( $Z$ ) scaled by building height ( $H$ ). (a) Dimensionless concentration on the windward and leeward sides of the canyon when  $H/W = 1$  and wind speed = 3 m/s. (b) Dimensionless concentration on the windward and leeward sides of the canyon when  $H/W = 2$  and wind speed = 5 m/s. Computational fluid dynamics modeling was performed, and measurements were obtained in wind tunnel simulations.

1 Identification of neighborhood-scale variability is important for interpreting data from ambient CO  
 2 monitors. Figure 3-23 shows a scatter plot of CO correlation as a function of distance (Pearson  $r$ ) for  
 3 monitors sited within 4 km of each other in four of the eleven CSAs and CBSAs. (The other seven CSAs  
 4 or CBSAs were excluded from this analysis because their monitors were spaced farther than 4 km from  
 5 each other.) The purpose of this plot is not to show a simple relationship between inter-sampler  
 6 correlation and distance but to illustrate how this relationship varies between cities at the neighborhood  
 7 scale. High correlation is seen for the monitors in the Phoenix, AZ CBSA, while those in the Pittsburgh,  
 8 PA, CSA have much lower inter-sampler correlations with steeper slope ( $\Delta r/\Delta d = -0.04$  for Phoenix and  
 9  $\Delta r/\Delta d = -0.27$  for Pittsburgh). The inter-sampler correlation observed for the Boston, MA CSA is also

- 1 substantially lower than those observed for Phoenix. As discussed in section 3.5.1.2, these differences
- 2 could be due to a variety of factors, including but not limited to natural and urban topography, different
- 3 traffic conditions on roads located near monitors, and proximity of the monitors to the roadway.



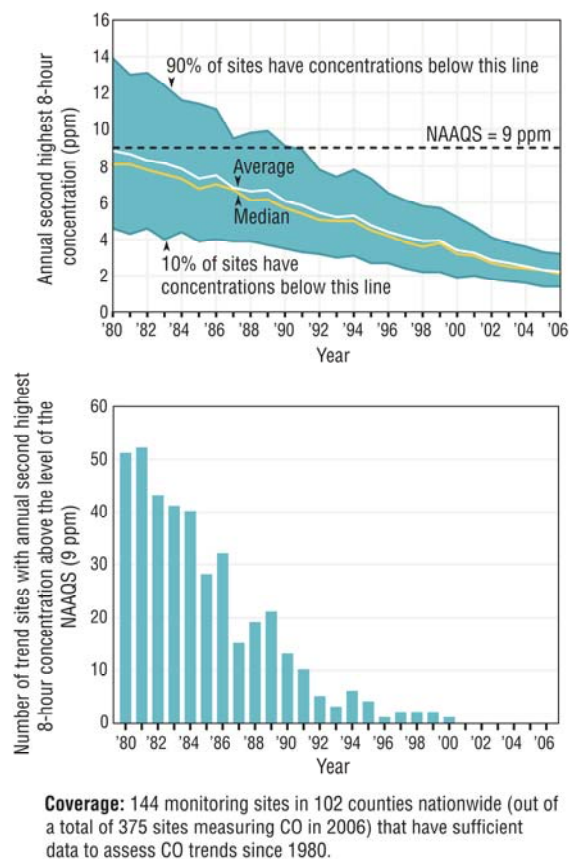
**Figure 3-23. Inter-sampler correlations as a function of distance between CO monitors for samplers located within 4 km (neighborhood scale) for Boston, Denver, Phoenix, and Pittsburgh.**



## 3.5.2. Temporal Variability

### 3.5.2.1. Multi-year Trends

1            Figures 3-24 (top) shows ambient CO concentrations in ppm from 1980 to 2006 based on  
2 continuous measurements averaged over 8-h time segments. The 8-h NAAQS is indicative of exposures  
3 occurring over a sustained period of time, an outdoor worker's exposure over the course of a day, for  
4 example. Figure 3-24 (bottom) depicts trends in the annual second-highest 8-h CO concentrations for 144  
5 sites in 102 counties nationwide having data either in the State and Local Air Monitoring Stations  
6 (SLAMS) network or from other special purpose monitors.

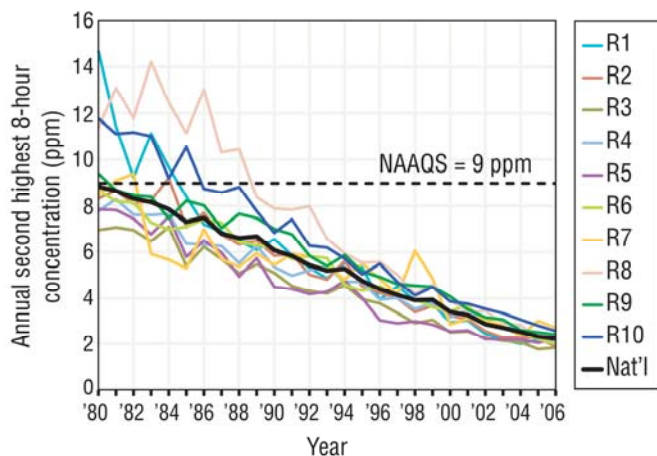


Source: U.S. EPA (2008c)

**Figure 3-24.** (Top) Trends in ambient CO in the U.S., 1980-2006, reported as the annual second highest 8-h concentrations (ppm) for the mean, median, 10% and 90% values. (Bottom) Trends in ambient CO in the U.S., 1980-2006, reported as the number of trend sites (y-axis) with annual second-highest 8-h concentrations above the level of the NAAQS (9 ppm).

1           The 2006 annual second highest 8-h CO concentration averaged across 144 monitoring sites  
2 nationwide was 75% below that for 1980, and is the lowest concentration recorded during the past 27  
3 years; see Figure 3-24 (top). Since 1992, more than 90% of all these sites have reported CO  
4 concentrations below the 8-h NAAQS of 9 ppm; see Figure 3-24. The mean annual second highest 8-h  
5 ambient CO concentration has been below 5 ppm since 2004. The downward trend in CO concentrations  
6 in the 1990s parallels the downward trend observed in CO emissions, attributed largely to decreased  
7 mobile source emissions; see Figure 3-2. In addition, of the 144 sites used to determine this trend, from a  
8 total of 375 monitoring sites operating in 2006, the number reporting second-highest CO concentrations  
9 above the level of the CO NAAQS declined to zero over the same period; see Figure 3-24 (bottom).

10           Consistent with the nationwide trends in emissions and concentrations, CO concentrations in all ten  
11 EPA Regions have steadily decreased since 1980, with reductions over this period ranging from 68% in  
12 Region 7 to 85% in Region 1; see Figure 3-25. This is also consistent with declining emissions seen in  
13 many regions of the U.S., shown in Figure 3-4. Reductions in anthropogenic CO emissions in the U.S.  
14 and elsewhere in the extra-tropical northern hemisphere are a major cause for the observed decrease in  
15 hemispheric and global-average CO concentrations observed since 1991 (Bakwin et al., 1994), although  
16 ~30% of the decline between 1991 and 2001 was attributed to decreases in CO following the eruption of  
17 Mt. Pinatubo (Novelli et al., 2003).



**Coverage:** 141 monitoring sites in the EPA Regions (out of a total of 375 sites measuring CO in 2006) that have sufficient data to assess CO trends since 1980.



Source: U.S. EPA (2008c)

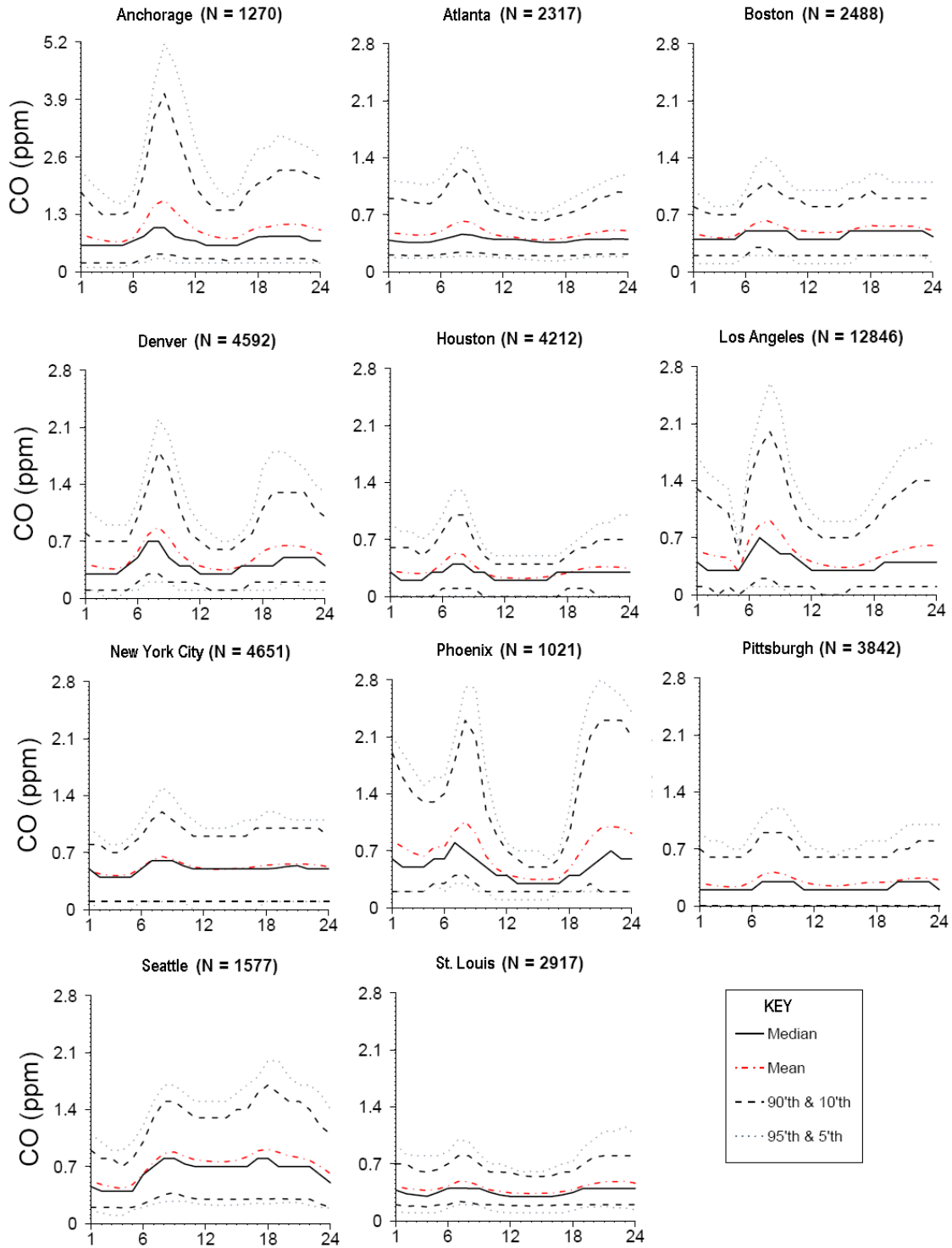
**Figure 3-25. Trends in ambient CO in the U.S., 1980-2005, reported as the annual second highest 8-h concentrations (ppm) for the EPA Regions 1 through 10, along with a depiction of the geographic extent of those Regions**

### 3.5.2.2. Hourly Variation

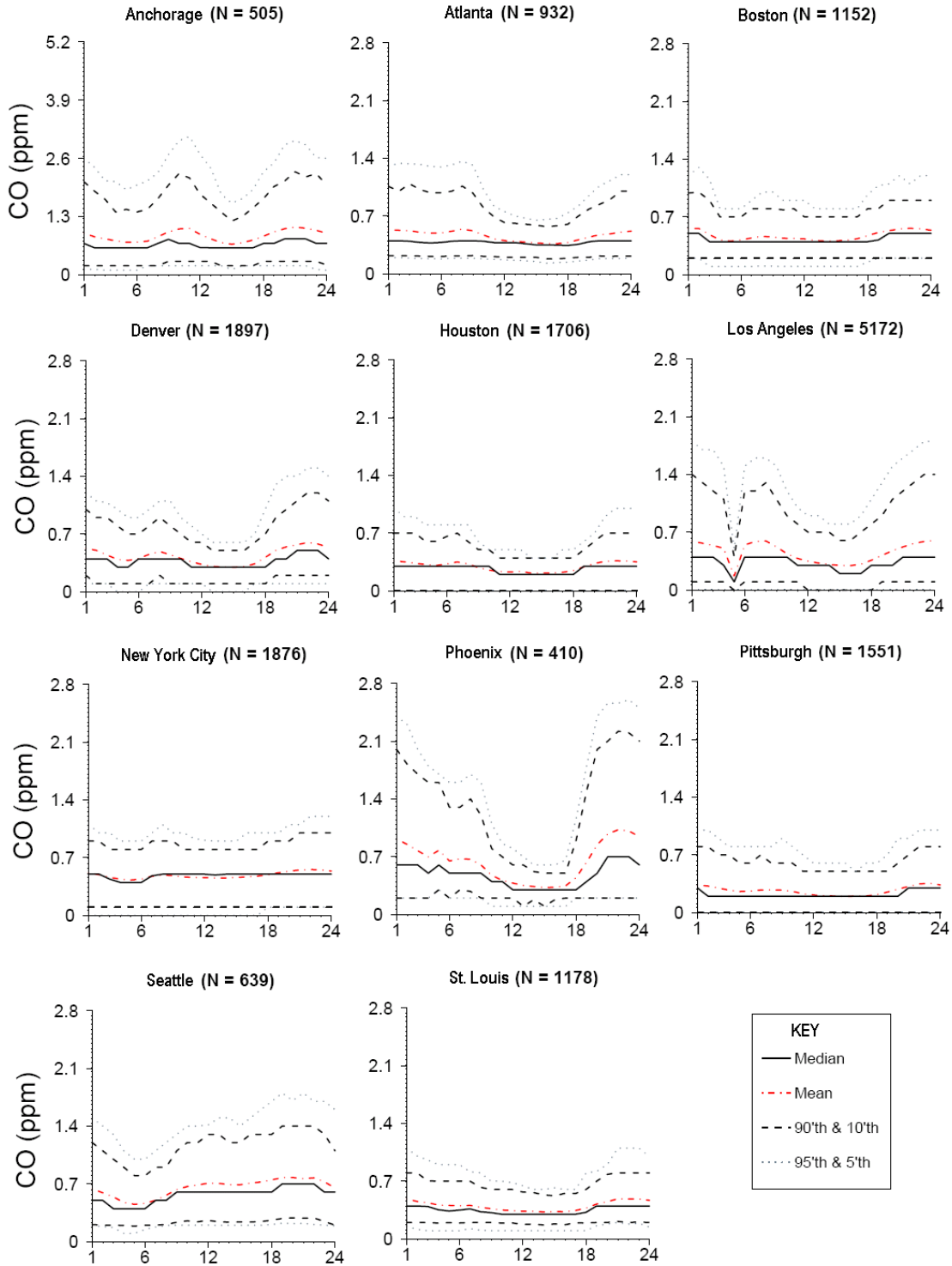
1 Weekday and weekend diel variation for the mean, median, 5th, 10th, 90th, and 95th percentiles of  
 2 hourly CO concentration over 2005-2007 are shown in Figures 3-26 and 3-27, respectively among the  
 3 eleven CSAs and CBSAs examined in this assessment. The weekday data showed that the Anchorage  
 4 mean, median, 5th and 10th percentile CO concentration curves exhibit more pronounced morning and  
 5 evening rush hour peak CO levels. Boston, Denver, Houston, Los Angeles, Phoenix, Pittsburgh, and St.  
 6 Louis all exhibited similar trends, although the magnitude of the concentrations shown was roughly twice  
 7 as high for Anchorage as the other cities. The curves had less overall variability for Boston, Pittsburgh,  
 8 and St. Louis. The Atlanta plot shows that the median concentration was fairly constant throughout the  
 9 24-h period, with a slightly elevated mean during the morning hours. The 90th and 95th percentile curves  
 10 exhibit stronger morning and evening CO concentration peaks. New York City shows fairly constant CO  
 11 mean and median concentration throughout the day with slight elevations throughout the morning rush  
 12 hour and a slight trough between 1:00 and 5:00 AM. The Seattle plot shows a daytime plateau beginning  
 13 around 5:00 AM and lasting until roughly 10:00 PM, with higher concentrations during morning and  
 14 afternoon rush hour. Differences in hourly variation among the eleven CSAs and CBSAs reflect city-to-

1 city variation in source characteristics and meteorology. For instance, the rush hour peaks in many cities  
2 likely correspond to increased mobile source emissions during those periods. Local meteorology and  
3 topography, which influence mixing heights, can also affect hourly variation in CO concentration.

4 Figure 3-27 illustrates weekend diel trends for the eleven CSAs and CBSAs considered in this  
5 assessment. For Anchorage during the period 2005-2007, the mean and median concentration curves  
6 peaked during the morning and evening hours. A daytime concentration trough is evident. The 90th and  
7 95th percentiles of concentration were similar but more pronounced. The shape of this plot is also  
8 characteristic of Atlanta, Boston, Denver, Houston, Los Angeles, Phoenix, Pittsburgh, Seattle, and St.  
9 Louis, although the Anchorage CO concentrations are nearly 100% higher than concentrations in the other  
10 cities. The weekend diel plot for New York shows that the mean and median CO concentrations remain  
11 fairly constant throughout the day, with a slight reduction between 2:00 and 7:00 AM. The 90th and 95th  
12 percentile curves illustrate more diel variation.



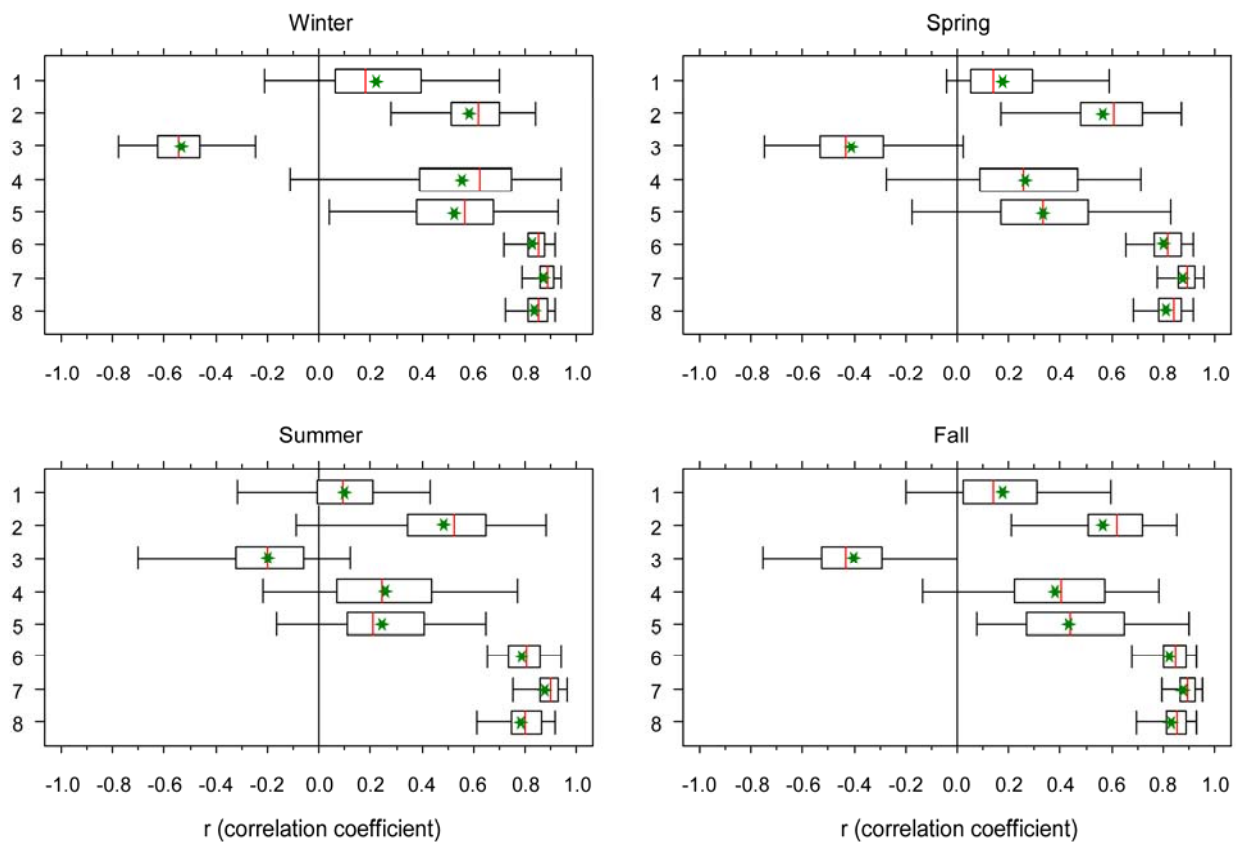
**Figure 3-26.** Diel plot generated from weekday hourly CO data (ppm) for the eleven CSAs and CBSAs 2005-2007. Included are the number of monitor days (N) and the median, mean, 5th, 10th, 90th and 95th percentiles of CO concentration. Note that the y-axis of the Anchorage CBSA plot is scaled to 5.2 ppm while the other plots are scaled to 2.8 ppm.



**Figure 3-27.** Diel plot generated from weekend hourly CO data (ppm) for the eleven CSAs and CBSAs 2005-2007. Included are the number of monitor days (N) and the median, mean, 5th, 10th, 90th and 95th percentiles of CO concentration. Note that the y-axis of the Anchorage CBSA plot is scaled to 5.2 ppm while the other plots are scaled to 2.8 ppm.

### 3.5.3. Associations with Copollutants

1 Associations between hourly CO and other copollutants, including SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>, PM<sub>10</sub>, and PM<sub>2.5</sub>  
 2 are provided in box plots for all U.S. monitors in Figure 3-28. The figure also shows the correlation of the  
 3 24-h avg CO concentration with the daily max 1-h and daily max 8-h CO concentrations as well as the  
 4 correlation between the daily max 1-h and daily max 8-h concentrations. AQS data were obtained from all  
 5 available co-located monitors across the U.S. after application of 75% completeness criteria. Pearson  
 6 correlation coefficients (r) were calculated using 2005-2007 data. Correlation plots analogous to Figure 3-  
 7 28 for select individual cities are provided in Annex A.



**Figure 3-28. Seasonal plots of nationwide correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations. Also shown are correlations between 24-h avg CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (The numbers in this caption refer to those on the y-axis of each seasonal plot.) Red bars denote the median, and green stars denote the arithmetic mean.**

1 The nationwide mean and median of the correlation with CO were positive for NO<sub>2</sub>, PM<sub>10</sub>, and  
2 PM<sub>2.5</sub>; near zero for SO<sub>2</sub>; and negative for O<sub>3</sub>. These findings might reflect common combustion sources  
3 for CO, NO<sub>2</sub>, and PM<sub>2.5</sub>. In all cases, a wide range of correlation values were reported. There were  
4 significant positive correlations for all CO-CO (24-h avg, daily max 1-h, daily max 8-h) comparisons  
5 rendered. Among those copollutants with positive associations, NO<sub>2</sub> had the highest mean and median  
6 correlations, followed by PM<sub>2.5</sub> and PM<sub>10</sub> (correlations vary by season). The correlations of CO with SO<sub>2</sub>  
7 and PM<sub>10</sub> were not significantly different from zero for any season; SO<sub>2</sub> would not be expected to  
8 correlate well with CO because SO<sub>2</sub> emanates primarily from industrial sources. Correlations between CO  
9 and NO<sub>2</sub> were significant and positive for winter, spring, and fall. Correlations between CO and PM<sub>2.5</sub>  
10 were significant and positively correlated for winter and fall. Correlations between CO and O<sub>3</sub> were  
11 significant and negative for winter only, when CO emissions tend to be high and O<sub>3</sub> formation is low. The  
12 copollutant correlation plots for individual cities shown in Annex A illustrate higher and more statistically  
13 significant correlations of CO with both PM<sub>10</sub> and PM<sub>2.5</sub> in several but not all of the select cities  
14 displayed. It is widely believed that sources of CO and PM<sub>2.5</sub> are highly correlated because they are both  
15 emitted directly during incomplete combustion and because secondary nitrate PM comes from NO<sub>x</sub>,  
16 which is largely produced from mobile sources. The wide confidence intervals displayed in the  
17 nationwide plots reflect the large pool of data in addition to the micrometeorological factors in each city.  
18 Additionally, CO monitors tend to be located apart from other pollutant monitors based on the criteria  
19 from the CFR of siting at least one monitor near the highest source of CO. Lack of co-location can affect  
20 the correlation of reported CO with PM<sub>10</sub> and PM<sub>2.5</sub> concentrations; this limitation also would inhibit the  
21 significance of correlations between CO and PM<sub>10-2.5</sub>.

22 Several studies reported correlations between ambient CO and copollutants. Reported relationships  
23 were generally consistent with the correlation data reported by the AQS. Sarnat et al. (2001) reported  
24 significant positive Spearman's correlations of CO with NO<sub>2</sub> ( $r = 0.76$ ) and PM<sub>2.5</sub> ( $r = 0.69$ ) and  
25 significant negative correlations of CO with O<sub>3</sub> ( $r = -0.67$ ) in Baltimore. Correlation of CO with SO<sub>2</sub> was  
26 insignificant ( $r = -0.12$ ). The Sarnat et al. (2001) study focused on correlations of ambient and personal  
27 PM<sub>2.5</sub> with gaseous copollutants, so seasonal information is only available for the correlation between  
28 PM<sub>2.5</sub> and CO. High correlation of ambient CO with NO<sub>2</sub> is expected given that both are closely related to  
29 mobile source combustion emissions. Sarnat et al (2005) also reported significant year-round association  
30 between CO and PM<sub>2.5</sub> and significant associations between CO and SO<sub>4</sub><sup>2-</sup> aerosols. Tolbert et al. (2007)  
31 reported correlations between multiple pollutants in Atlanta and also showed the highest Spearman's  
32 correlation for CO with NO<sub>2</sub> ( $r = 0.70$ ). CO was also reported to have fairly high correlation with PM<sub>2.5</sub>  
33 elemental carbon (EC) ( $r = 0.66$ ), PM<sub>2.5</sub> organic carbon (OC) ( $r = 0.59$ ), and PM<sub>2.5</sub> total carbon (TC)  
34 ( $r = 0.63$ ). Correlations were reported to be much lower for CO with O<sub>3</sub> ( $r = 0.27$ ) and PM<sub>2.5</sub> SO<sub>4</sub><sup>2-</sup>  
35 ( $r = 0.14$ ). Kim et al. (2006) measured CO, NO<sub>2</sub>, and PM<sub>2.5</sub> at ambient fixed sites in Toronto, Canada and

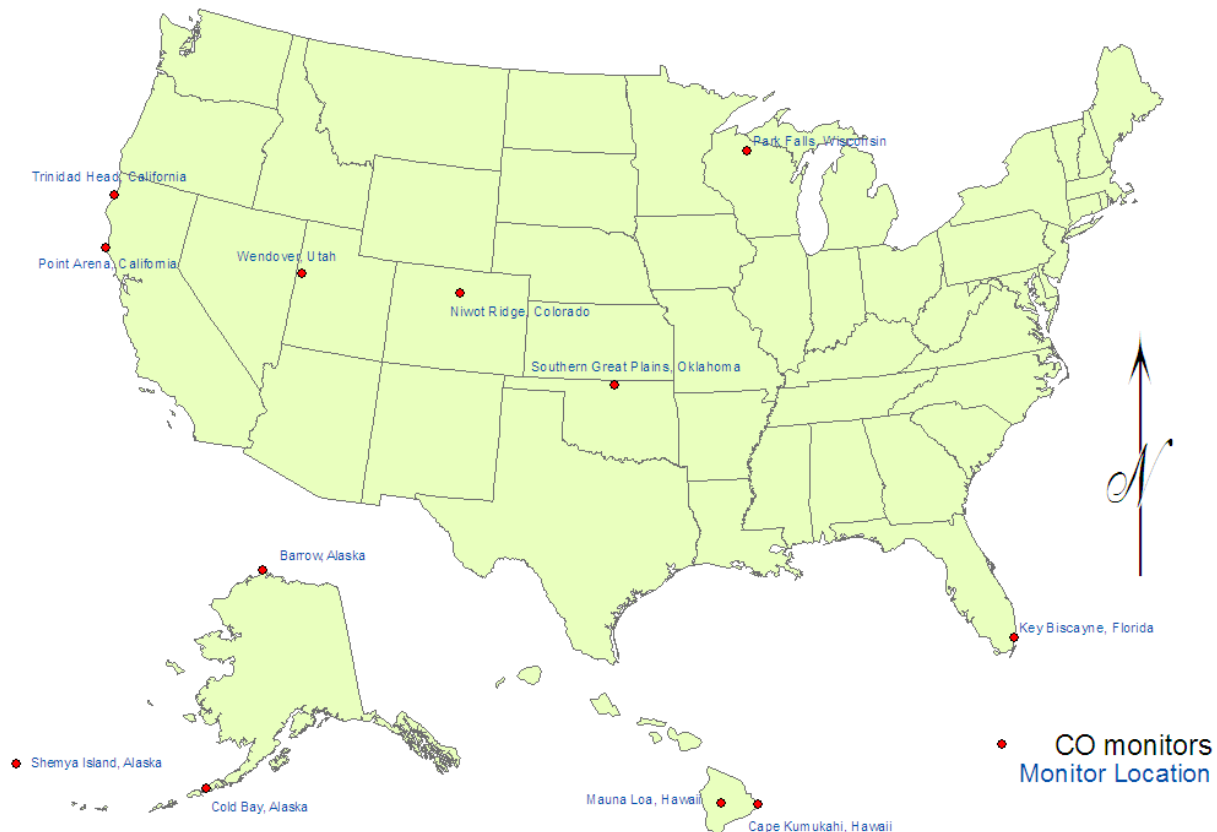


1 found associations, averaged over monitoring stations, of CO with PM<sub>2.5</sub> (Spearman's  $r = 0.38$ , non-  
2 significant) and of CO with NO<sub>2</sub> ( $r = 0.72$ , significant).

### 3.5.4. Policy-Relevant Background

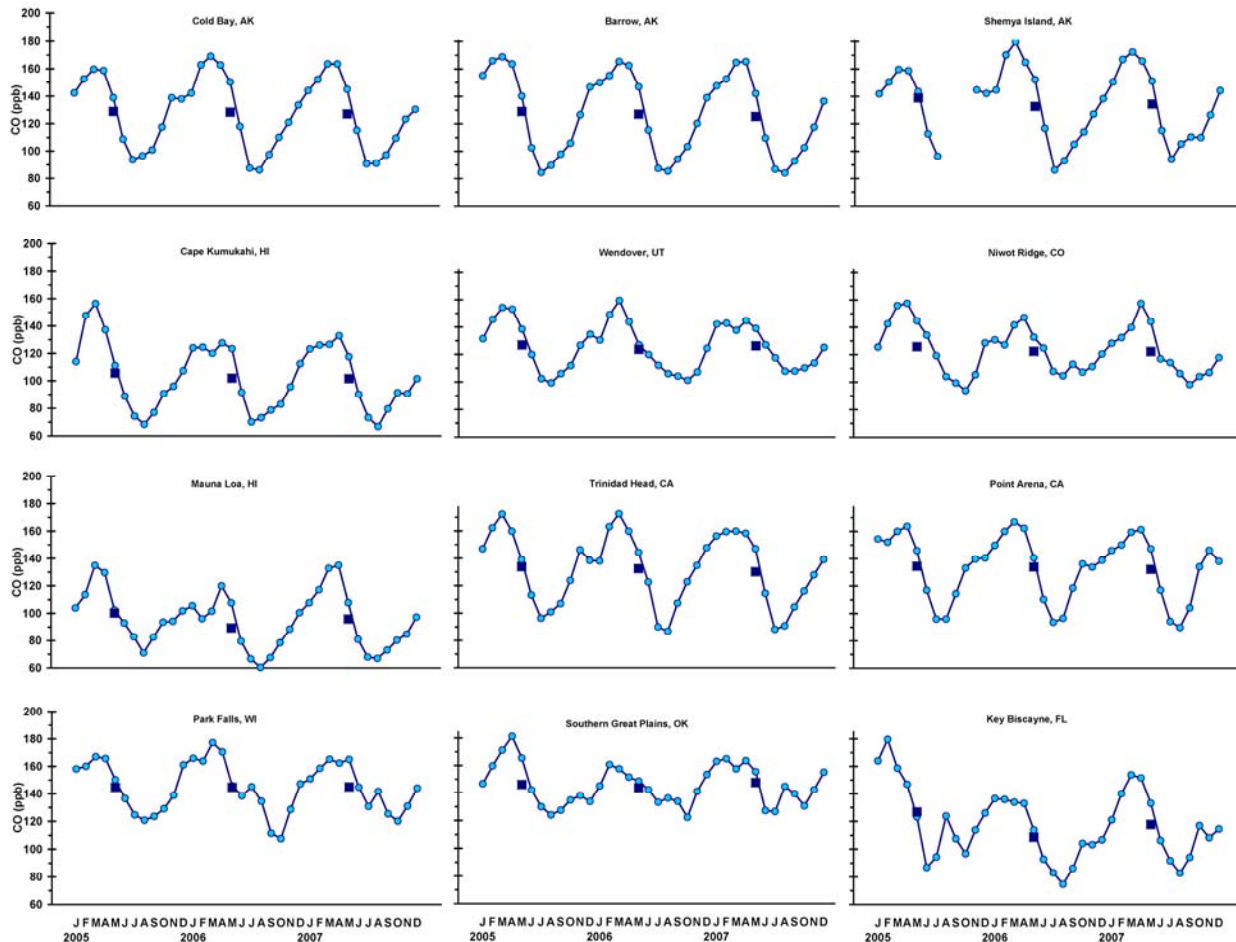
3 Background concentrations of pollutants used for informing policy decisions about national  
4 standards in the U.S. are commonly referred to as policy-relevant background (PRB) concentrations. PRB  
5 concentrations are those that are exclusive of anthropogenic emissions in the U.S., Canada and Mexico  
6 (North America), and consist of world-wide biogenic emissions (including North America) and  
7 anthropogenic emissions elsewhere in the world.

8 PRB concentrations of CO can best be determined from the extensive and long-running network  
9 of remote-site baseline CO measurements conducted by NOAA's Earth System Research Laboratory  
10 (ESRL), Global Monitoring Division (GMD), as part of their Carbon Cycle Greenhouse Gases Group  
11 (CCGG) Cooperative Air Sampling Network (CASN) (<http://www.esrl.noaa.gov/gmd/ccgg/iadv>). CO data  
12 through December 2007 are available with extensive quality assurance and control information from a  
13 worldwide network of 72 nodes active in December 2008. ESRL GMD uses the highly sensitive gas  
14 chromatography-mercury liberation photometric detection technique with precision to 1 parts per billion  
15 (ppb) in 50 ppb (2 ppb in 200 ppb) and accuracy to 1.5 ppb in 500 ppb (2 ppb in 200 ppb) (Novelli et al.,  
16 2008).



**Figure 3-29. Map of the baseline monitor sites used in this assessment to compute policy-relevant background concentrations.**

1 This assessment used data from 2005 through 2007 at 12 remote sites in the U.S. to determine PRB  
 2 (see the map in Figure 3-29): Cold Bay, AK; Barrow, AK; Shemya Island, AK; Cape Kumukahi, HI;  
 3 Mauna Loa, HI; Trinidad Head, CA; Point Arena, CA; Wendover, UT; Niwot Ridge, CO; Park Falls, WI;  
 4 Southern Great Plains, OK; and Key Biscayne, FL. Average concentrations for each month and for each  
 5 year, 2005-2007 are shown for each site in Figure 3- 30. All sites demonstrate the well-known seasonality  
 6 in background CO with minima in the summer and fall and maxima in the winter and spring. Summer-  
 7 time minima may be related to photochemical reaction of CO with OH, as described in section 3.3.  
 8 Analysis for North American PRB is made here by segregating the three Alaska sites (based on their high  
 9 latitude) and the two Hawaii sites (based on their distance from the continent) and treating the remaining  
 10 seven sites as representative of the CONUS background. The 3-year avg CO PRB in Alaska ranged from  
 11 127 to 135 ppb with an average of 130 ppb; in Hawaii from 95.3 to 103.1 ppb with an average of 99.2  
 12 ppb; and over the CONUS from 118 to 146 ppb with an average of 132 ppb.



Figures 3-30. Monthly (circles) and annual (squares) average CO concentrations (ppb), 2005-2007 . Cold Bay, AK; Barrow, AK; Shemya Island, AK; Cape Kumukahi, HI; Wendover, UT; Niwot Ridge, CO; Mauna Loa, HI; Trinidad Head, CA; Point Arena, CA; Park Falls, WI; Southern Great Plains, OK; and Key Biscayne, FL.

## 3.6. Issues in Exposure Assessment

### 3.6.1. Summary of Findings from 2000 CO AQCD

1           The 2000 CO AQCD (U.S. EPA, 2000) describes the results of studies completed prior to 1999 on  
 2 personal exposures and microenvironmental concentrations of CO. Although these studies may no longer  
 3 be representative of current exposure levels due to declining ambient CO concentrations, the personal-  
 4 microenvironmental-ambient relationships are still instructive. Time spent commuting, particularly in  
 5 cars, was a major contributor to personal CO exposures. Many studies measured in-vehicle concentrations  
 6 of CO and found elevated concentrations compared to fixed-site monitors. Roadside CO monitors were

1 elevated compared to ambient levels, and equal to or lower than in-vehicle levels (e.g., Ott, 1994; Rodes  
2 et al., 1999). A small portion of the CO concentrations inside a vehicle cabin comes from the vehicle  
3 itself, while a substantial fraction comes from roadway traffic emissions entering the cabin via air  
4 exchange. Studies summarized in the 2000 CO AQCD found that in-vehicle CO concentrations were  
5 generally two to five times higher than ambient CO concentrations. High traffic volumes contributed to  
6 increased in-vehicle concentrations.

7       Prior to the 2000 CO AQCD, it was well-known that CO levels in residences may be elevated  
8 above ambient due to non-ambient indoor sources, such as cooking, space heating, and smoking.  
9 Separation of indoor CO into ambient and non-ambient components is important for determining the  
10 effect of ambient CO concentrations, although this has not been done successfully in previous studies.  
11 Two large studies done in Denver, CO and Washington, DC in the early 1980s found that fixed-site  
12 monitor concentrations were higher than personal exposures for those with low-level exposures, while  
13 fixed site monitor concentrations were lower than exposures for those with high-level exposures (Akland,  
14 1985; Johnson, 1984). Non-ambient sources contributing to high total exposures likely obscured this  
15 relationship. In Denver, gas stove operation, passive smoking, and attached garages increased residential  
16 indoor exposure by 2.6, 1.6, and 0.4 ppm respectively compared to individuals without those sources  
17 present. Categorical analyses found significantly higher personal exposures on high ambient  
18 concentration days than on low ambient concentration days, suggesting that personal exposures are  
19 related to ambient levels. Non-ambient exposures tend to obscure the relationship between ambient CO  
20 concentrations and personal exposure.

### 3.6.2. General Exposure Concepts

21       A general framework for human exposure modeling was described in the 2008 Draft PM ISA, 2008  
22 NO<sub>x</sub> ISA, and 2008 SO<sub>x</sub> ISA (U.S. EPA, 2008e, f). A brief conceptual model of human exposure is  
23 provided here with respect to CO for the readers' convenience. An individual's daily exposure to CO can  
24 be described based on a compartmentalization of the person's activities over a time period of interest:

$$E = \int C_j dt$$

*Equation 3-2*

25 where  $E$  = exposure over some duration,  $C_j$  = CO concentration at location  $j$ , and  $dt$  = time spent in  
26 location  $j$ . This basic equation can be broken down into a microenvironmental model that accounts for  
27 exposure to CO of ambient ( $E_a$ ) and non-ambient ( $E_{na}$ ) origin of the form:

$$E = E_a + E_{na}$$

Equation 3-3

1 This assessment focuses on the ambient component of exposure because the non-ambient portion of  
 2 exposure is subject to individual behavior. Daily  $E_a$  can be expressed in terms of the fraction of time spent  
 3 outdoors and indoors being exposed to ambient CO concentration,  $C_a$  (Wilson et al., 2000):

$$E_a = \left( f_o + \sum f_i F_{inf_i} \right) C_a$$

Equation 3-4

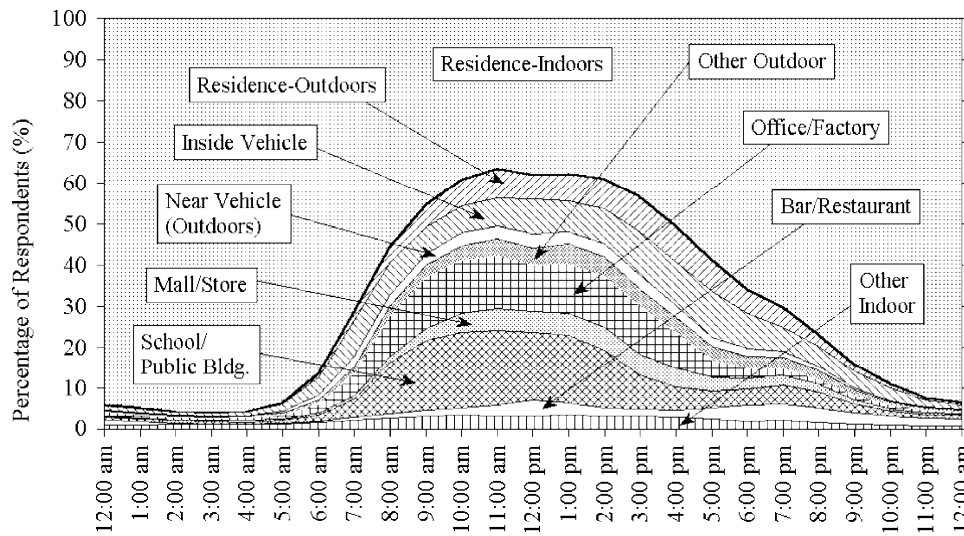
4 where  $f$  = fraction of time, subscript  $o$  = outdoor, subscript  $i$  = indoor, and  $F_{inf}$  = infiltration factor.  $F_{inf}$   
 5 quantifies the equilibrium fraction of the CO concentration outside the microenvironment that penetrates  
 6 inside the microenvironment and remains in mixture. It is a function of the building air exchange  
 7 characteristics and the properties of the gas. Assuming steady state ventilation conditions, the infiltration  
 8 factor is a function of the penetration ( $P$ ) of CO, the air exchange rate ( $a$ ) of the microenvironment, and  
 9 the rate of CO loss ( $k$ ) in the microenvironment;  $F_{inf} = Pa/(a+k)$ . Given that  $k \rightarrow 0$  for CO,  $F_{inf}$  reduces to  
 10  $P$ . Equation 3-4 is subject to the constraint  $f_o + \sum f_i = 1$  to reflect the total exposure over a specified time  
 11 period. The indoor term has a summation because indoor exposure can occur in various  
 12 microenvironments throughout a time period of interest. “Outdoor” exposure can occur in parks, yards,  
 13 sidewalks, and on bicycles or motorcycles. “Indoor” refers to being inside any aspect of the built  
 14 environment, e.g., home, office buildings, enclosed vehicles (automobiles, trains, buses), and/or  
 15 recreational facilities (movies, restaurants, bars). The complex human activity patterns that dictate  
 16 exposure across the population (all ages) are illustrated in Figure 3-31 (Klepeis et al., 2001). This figure  
 17 illustrates the diversity of daily activities among the population as well as the proportion of time spent in  
 18 each microenvironment. Although activities in Klepeis et al. (2001) are presented over a day, information  
 19 from this figure can be extracted to deduce variation over a 1-h or 8-h period.

20 As seen in equation 3-4, ambient exposure is a linear function of the ambient CO concentration.  
 21 The ambient exposure factor,  $\alpha$ , is the ratio between the personal exposure to ambient CO and the ambient  
 22 concentration of CO (or the ambient exposure factor); i.e.,  $\alpha = E_a/C_a$ .  $\alpha$  is therefore the proportionality  
 23 factor in equation 3-4, i.e.,  $\alpha = f_o + \sum f_i F_{inf_i}$  ( $\alpha$  varies between 0 and 1). If a person’s exposure occurs in a  
 24 single microenvironment, the ambient component of a microenvironmental CO concentration can be  
 25 represented as the product of the ambient concentration and  $F_{inf}$ . If there are no significant local outdoor  
 26 sources and sinks of CO, then  $C_a$  can be approximated by the concentration measured by an ambient  
 27 monitor. If local sources and sinks exist and are significant, then the ambient component of outdoor air  
 28 must be estimated using dispersion models, land use regression models, receptor models, fine scale  
 29 chemistry-transport models or some combination of these techniques.

1 A variety of approaches can be used to estimate exposure to ambient CO. In some cases, individual  
2 personal exposures are measured with personal exposure monitors (PEMs), where personal samples are  
3 taken to estimate population exposure. In other cases, ambient concentrations are used as an exposure  
4 indicator. The ambient concentration may be based on measurements made at a single ambient monitor or  
5 as the average of several ambient monitors.

6 Very few recent exposure assessment studies utilized measurements of CO concentration and none  
7 of the recent exposure error studies used measurements of personal and ambient CO exposure data,  
8 although Sheppard et al. (2005) presented a conceptual model for nonreactive pollutants that could  
9 include CO. Many recent studies analyzing exposure error used PM data (e.g., Sheppard et al., 2005;  
10 Wilson and Brauer, 2006; Zeger et al., 2000). The lack of recent CO data contributes to uncertainty  
11 regarding personal-ambient relationships. However, review of exposure error studies for PM, under the  
12 assumption that it is nonreactive, is instructive for discerning the relative influence of ambient and non-  
13 ambient exposures if consideration is made for differences between CO and PM with respect to indoor  
14 source variability and infiltration.

15 Wilson and Brauer (2006) showed significantly stronger associations between health effects and  
16 ambient exposure than between health effects and total personal exposure. The use of personal exposure  
17 in population exposure assessment studies could cause various errors in the health effect estimate because  
18 the correlation between personal exposure and ambient concentration may be different for each subject in  
19 a panel and may not be statistically significant (Wilson et al., 2007). The correlation between ambient  
20 concentration and personal exposure may be high for some subjects, in which case the exposure error  
21 caused by using ambient concentration instead of personal exposure may be small. In other subjects, the  
22 correlation may be low or negative (and not statistically significant). In this case, the exposure error will  
23 be high and may obscure relationships between ambient exposure and health effects. Differences in  
24 correlation observed in these studies occurred in part because the ambient exposure factor,  $\alpha$ , was  
25 different for each subject. The relationships between ambient concentration and the ambient component  
26 of personal exposure were statistically significant for all subjects. The analyses of Wilson and Brauer  
27 (2006) and Wilson et al. (2007) focused on PM, which is subject to indoor-outdoor differences in size  
28 distribution and chemical composition related to differential infiltration. However, this issue is still  
29 relevant for CO because non-ambient CO exposure is uncorrelated with ambient CO exposure and  
30 therefore could obscure health effects relationships observed in epidemiologic studies. Therefore, ambient  
31 concentration is better than total personal exposure as an indicator of ambient CO exposure.



Source: Klepeis et al. (2001)

**Figure 3-31. Distribution of time that a sample population spends in various environments, from the National Human Activity Pattern Survey.**

### 3.6.3. Monitoring Issues Associated with Exposure Assessment

#### 3.6.3.1. Exposure Assessment using Community-Based Ambient Monitors

1 Instrumentation and associated monitoring errors are described in section 3.4; exposure error  
 2 related to instrumental measurement error in ambient monitors is described here. Because there will likely  
 3 be some random component to instrumental measurement error, the correlation of the measured CO  
 4 concentration with the true CO concentration will likely be less than 1. Sheppard et al. (2005) indicated  
 5 that instrument error in the individual or daily average concentrations have “the effect of attenuating the  
 6 estimate of  $\alpha$ .” However, Zeger et al. (2000) stated that the “instrument measurement error in the ambient  
 7 levels...is close to the Berkson type” and in order for this error to cause substantial bias in later estimation  
 8 of the health outcome, the error term (the difference between the true concentrations and the measured  
 9 concentrations) must be strongly correlated with the measured concentrations. Zeger et al. (2000)  
 10 suggested that, “Further investigations of this correlation in cities with many monitors are warranted.”

#### 3.6.3.2. Personal Exposure Monitors

11 Portable monitors for measuring personal CO exposure include the Langan and Draeger monitors,  
 12 both of which use electrochemical oxidation-reduction techniques (Langan, 1992). These monitors  
 13 continuously log CO concentrations, making them suitable for use in personal monitoring studies. More  
 14 detail on personal CO monitoring is provided in the 2000 CO AQCD (U.S. EPA, 2000).

## Measurement Error in Personal Exposure Monitors

1 Personal electrochemical CO monitors are subject to interference and drift, and have a relatively  
2 high detection limit (approximately 1 ppm) relative to current ambient concentrations. Previous studies in  
3 the 1980's and 1990's, when ambient levels were higher, were able to successfully deploy these monitors,  
4 but more recent exposure studies have avoided personal CO measurements due to the high percentage of  
5 non-detects. The lack of a suitable personal monitor for measuring low-level exposures (<1 ppm) has  
6 hampered field studies assessing personal exposure to ambient CO. Chang et al. (2001) evaluated the  
7 Langan CO monitor as part of an air quality sampling manifold. At high (0.4-3.0 ppm) CO  
8 concentrations, the instrument correlated well ( $R^2 = 0.93$ ) with a reference CO monitor with the Langan  
9 underestimating the CO concentration by 41%. When ambient levels fell consistently below that level,  
10 coefficient of determination ( $R^2$ ) between the Langan and reference monitor fell to  $R^2 = 0.4$  in summer  
11 and  $R^2 = 0.59$  in winter with the arithmetic average concentration underestimated by 47% in summer and  
12 by 63% in winter. Chang et al. (2001) pointed out the need for frequent instrument zeroing to minimize  
13 instrument drift. Abi Esber et al. (2007a) evaluated a similar personal electrochemical CO sensor, the  
14 GEM™ 2000, by comparing measured concentrations with those obtained through co-located grab bag  
15 sampling in a vehicle cabin. Differences between the GEM™ 2000 and the reference samples were fairly  
16 low during weekday driving (differences = 2.1-10.6%). Differences on Sundays, when traffic was  
17 significantly lower than during weekdays, were dependent on vehicle ventilation conditions, with better  
18 agreement when vehicle ventilation allowed for higher cabin CO concentrations (differences = 3.4-5.6%),  
19 but the electrochemical sensor did not compare well with reference values when concentrations were low  
20 (differences = 20-71%). In general, it is difficult to separate the large instrumental measurement error  
21 seen at concentrations below instrument LOD from variation in non-ambient exposures. This large  
22 variation in personal measurements can result in high levels of classical measurement error (Sheppard et  
23 al., 2005).

### 3.6.4. Indoor/Outdoor Relationships and Infiltration

24 CO is a relatively inert gas, making the indoor decay rate negligible compared to typical air  
25 exchange rates (~1/h). In the absence of indoor sources, this would lead to an indoor-outdoor  
26 concentration ratio (I/O) of approximately 1. For this reason, few studies have measured I/O for CO.  
27 Polidori et al. (2007) measured I/O of 0.94-1.21 for two retirement communities in the Los Angeles area.  
28 The authors suggested that similarity between I/O for CO and NO<sub>x</sub> can be attributed to lack of indoor  
29 sources of either gas. Chaloulakou and Mavroidis (2002) reported I/O for CO measurements in the  
30 absence of indoor sources in a school building in Athens, Greece and found that I/O varies with season.  
31 During the summer, median I/O was reported to be 0.57 on weekdays, 0.91 on Saturdays, and 0.81 on



1 Sundays. In winter, median I/O was reported to be 0.82 during weekdays, 0.90 on Saturdays, and 0.74 on  
2 Sundays. The authors attributed the lower weekday I/O values to traffic-based peaks in outdoor CO  
3 concentrations that were not translated to indoors. In a related work, Chaloulakou et al. (2003) reported  
4 the median I/O over all days as 0.8 for the same school and 0.9 for an Athens office building with no ETS  
5 (the presence of other sources was not clearly stated but assumed zero). However, observed indoor values  
6 are often greater than outdoor concentrations in the presence of indoor sources. A recent study in the U.K.  
7 reported I/O of 3.9-4.3 in homes with gas cookers (Dimitroulopoulou et al., 2006), which is consistent  
8 with previous studies. A multipollutant study conducted in 2000-2001 attempted to measure I/O for CO  
9 and calculated residential infiltration factors, but low CO concentrations resulted in a large number of  
10 non-detects (Williams et al., 2003).

### 3.6.5. Personal/Ambient Relationships

#### 3.6.5.1. Panel and Population Exposure Studies

11 Although several multi-pollutant exposure studies have been conducted recently in the U.S., (e.g.,  
12 Sarnat, 2006), most have not included CO in the suite of pollutants, possibly due to limitations in personal  
13 monitoring techniques. A few studies conducted in Europe and Canada measured personal-ambient  
14 relationships for CO.

15 The EXPOLIS study (Georgoulis et al., 2002) found that 48-h personal exposures were  
16 significantly correlated with ambient concentrations in each of five European cities (Athens, Basel,  
17 Helsinki, Milan, and Prague). Controlling for source terms, including ETS, traffic, and natural gas  
18 appliances, regression coefficients between personal exposure and ambient concentration ranged from  
19 0.28 in Milan to 1.99 in Helsinki. Regression coefficients greater than 1 may indicate location of a fixed  
20 site monitor away from local sources (e.g., roadways), resulting in a lower ambient value than  
21 experienced in the urban core. As part of this study, personal CO exposure was measured for a panel of 50  
22 office workers in Milan (Bruinen de Bruin et al., 2004). Average measured 1-h personal exposures were  
23 7.3 ppm in comparison with 5.0 ppm for fixed site 1-h measurements. Average 8-h (3.3 ppm) and 24-h  
24 (2.1 ppm) CO concentrations were the same for personal and fixed site measurements. Percentage of time  
25 exposed, exposures, and percentage of exposure from the Bruinen de Bruin et al. (2004) study, in the  
26 absence of non-ambient CO from ETS and gas cooking, are shown in Table 3-9. The largest percentage of  
27 CO exposure was attributed to home indoor exposure in the absence of indoor sources, while the highest  
28 exposure levels were observed during transit. Scotto di Marco et al. (2005) found similar results. Bruinen  
29 de Bruin et al. (2004) and Scotto di Marco et al. (2005) found that mobile source emissions were  
30 important contributors to personal exposure, as described in the following subsection.

**Table 3-9. Percentage of time exposed to ambient CO (adjusted to reflect the absence of non-ambient CO from ETS and gas cooking), average CO exposures, and percentage of exposure estimated for the population.**

	Percent of time exposed (%)	Exposure (ppm)	Percent of exposure (%)
<b>INDOORS</b>	<b>89.6</b>		<b>81.1</b>
Home	56.5	1.8	49.4
Work	29.1	1.9	26.8
Other	4.1	2.5	4.9
<b>OUTDOORS</b>	<b>1.8</b>		<b>2.1</b>
Home	0.2	2.3	0.2
Work	0.6	2.1	0.6
Other	1.0	2.6	1.2
<b>IN-TRANSIT</b>	<b>8.5</b>		<b>16.8</b>
Walking	3.0	3.0	4.4
Train/metro	0.7	3.0	1.0
Bus/tram	2.0	3.8	3.7
Motorbike	0.2	4.5	0.4
Car/taxi	2.6	5.7	7.2

Source: Bruinen de Bruin et al. (2004).

1           EXPOLIS also looked at the special case of children’s exposure to CO because children generally  
2 do not produce CO in their daily activities and have no occupational exposures. Alm et al. (2000; 2001)  
3 reported higher personal exposures than ambient concentrations for children aged 3-6 years old in  
4 Helsinki. Their mean daily max 1-h exposure was 5.2 ppm, compared to 1.4 ppm measured at a fixed-site  
5 monitor. For daily max 8-h and 24-h avg concentrations, the corresponding values were 2.9 ppm and  
6 2.1 ppm for personal exposure and 0.8 and 0.6 ppm, respectively, for fixed site measurements. The  
7 Spearman rank correlation, although statistically significant, was relatively low ( $r = 0.15$ ) between  
8 individual 24-h avg exposure and the ambient monitor. The correlation improved when the average  
9 exposure of children measured on the same day ( $r = 0.33$ , 3-6 children) or the same week ( $r = 0.55$ , 10-23  
10 children) was compared to the monitor data. A regression model using questionnaire data found that  
11 parental smoking status, parental education, and presence of a gas stove explained 12% of the variability  
12 in the 8-h max exposures, indicating that other factors, such as time spent outdoors and proximity to  
13 roadways are likely to be important in determining personal exposure.

14           Kim et al. (2006) reported mean CO concentrations of 1.4 ppm for a panel of 28 cardiac-  
15 compromised individuals in Toronto, Canada. Corresponding fixed-site monitor mean concentrations  
16 ranged from 0.5-1.4 ppm, with an overall mean of 1.0 ppm. The observed higher personal exposures may  
17 have been due to both indoor sources and proximity to roadways when outdoors. Personal-ambient  
18 Spearman correlations ranged from -0.65 to 0.93, with a median of  $r = 0.31$ , indicating that while

1 moderate correlations are observed overall, inter-individual differences based on time spent in different  
2 microenvironments have a strong influence on the observed correlation.

3 Lai et al. (2004) measured relationships between personal CO exposure and microenvironmental  
4 (home indoor, home outdoor, and work indoor) concentrations in Oxford, U.K.. The highest personal  
5 exposures were associated with smoking, cooking, and transportation while low correlations were  
6 observed between personal and indoor residential concentrations, further indicating the importance of  
7 indoor sources and the need to separate ambient contributions to personal exposure from total personal  
8 exposure.

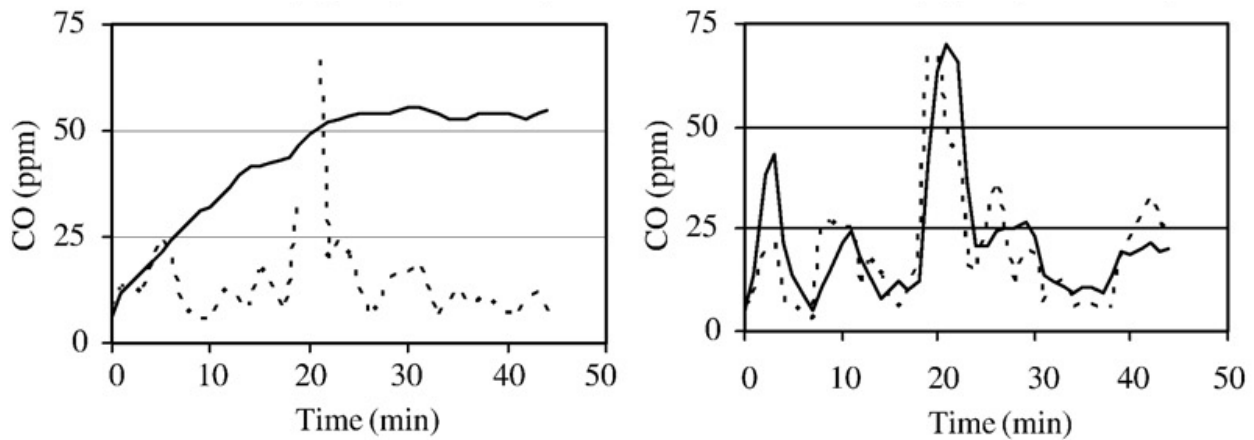
### 3.6.5.2. Commuting Time CO Exposure Studies

9 A number of studies focused on transit-time CO exposure, which can occur in a vehicle or while  
10 walking or cycling. Kaur (2005) found that transit time exposures in London, U.K. were significantly  
11 higher than measurements made at a fixed site background monitor away from traffic ( $0.3 \pm 0.1$  ppm) for  
12 car riders ( $1.3 \pm 0.2$  ppm), taxi riders ( $1.1 \pm 0.1$  ppm), bicyclers ( $1.1 \pm 0.2$  ppm), walkers ( $0.9 \pm 0.2$  ppm),  
13 and bus riders ( $0.8 \pm 0.1$  ppm). Curbside measurements ( $1.5 \pm 0.7$  ppm) in this study were slightly higher  
14 than car riders' exposures. Duci et al. (2003) found that average in-transit concentrations were highest for  
15 cars (winter:  $21.4 \pm 4$  ppm), followed by pedestrians (winter:  $11.5 \pm 2.6$  ppm; summer:  $10.1 \pm 1.7$  ppm),  
16 buses (winter:  $10.4 \pm 2.9$  ppm; summer:  $9.4 \pm 3.6$  ppm), trolley (winter:  $9.6 \pm 1.9$  ppm; summer:  $8.2 \pm$   
17  $3$  ppm), and rail (winter:  $4 \pm 0.6$  ppm; summer:  $3.4 \pm 0.7$  ppm). Duci et al. (2003) did not provide fixed  
18 site CO concentrations but stated that in-transit exposures were higher in each case. Additional analyses  
19 from the EXPOLIS study indicated that on-road mobile source emissions were the most important source  
20 of CO exposure for non-ETS-exposed subjects (Bruinen de Bruin et al., 2004; Scotto di Marco et al.,  
21 2005). Scotto di Marco et al. (2005) found that, for a panel of 201 adult Helsinki residents (aged 25-55  
22 years), subjects spent 8.1% of their time in transport, which accounted for 12.6% of their total exposure  
23 (range of means =  $0.96$  ppm on a train –  $2.8$  ppm in a car). Similarly, in a panel study of 50 office  
24 workers, Bruinen de Bruin et al. (2004) found that, in the absence of non-ambient sources, the subjects  
25 spent 8.5% of their time in transit, which accounted for 16.8% of their total exposure, with 2.6% of time  
26 spent in a car or taxi accounting for 7.2% of exposure (mean =  $5.7$  ppm). Commuting time was an  
27 important predictor of exposure, such that subjects living in low CO concentration suburban areas and  
28 commuting to work experienced higher levels than urban residents with short commute times.

29 Gómez-Perales et al. (2004; 2007) measured CO exposures on buses, mini-buses, and metro cars in  
30 Mexico City, Mexico to be 12 ppm, 15 ppm, and 7 ppm, respectively. These values are higher than  
31 CONUS measurements and those presented by Kaur et al. (2005), but the relative difference between the  
32 minibus and bus exposures in the Gómez-Perales et al. (2004; 2007) study are similar to those seen for the  
33 taxi-to-bus or car-to-bus comparisons in Kaur et al. (2005). Alm et al. (1999) reported in-vehicle CO

1 concentrations of 5.7 ppm in the morning and 3.1 ppm in the afternoon commute for Kuopio, Finland.  
2 Abi Esber et al. (2007b) report results from CO concentration measurements taken within an automobile  
3 in Beirut, Lebanon during the morning commute period of 7:30 to 9:30 a.m. Weekday trip CO levels  
4 ranged from 10.8 ppm with the windows open and vents closed to 37.4 ppm when driving with windows  
5 and vents closed. Mean and standard deviation for ambient CO concentrations, obtained using a roadside  
6 monitor in Beirut during the periods September-December 2003, August-September 2004, and  
7 May-August 2005 were  $1.4 \pm 0.7$ ,  $1.6 \pm 0.4$ , and  $1.1 \pm 0.7$  ppm, respectively.

8 Abi Esber and El-Fadel (2008) compared the amount of CO produced by an automobile, driving  
9 the same route of Beirut described in Abi Esber et al. (Abi Esber et al., 2007b) above, by sampling CO  
10 directly from the engine of the vehicle and separately from the cabin of the car under three different  
11 ventilation conditions. For the case when one window was half-open and vents were closed, engine CO  
12 concentrations averaged 12.6 ppm while in-vehicle concentrations averaged 17.7 ppm, which was a  
13 40.5% increase. With windows closed and the air conditioner operating on “recirculating air” mode, CO  
14 concentrations averaged 13 ppm from the engine and 30.2 ppm in the vehicle cabin, a 132% increase.  
15 With windows closed and the air conditioner on “fresh air” mode, engine CO concentrations averaged  
16 18.3 ppm while in-vehicle concentrations were 20.5 ppm, which was only a 12% increase. Figure 3-32  
17 shows that the time series for the cabin and engine CO samples are very similar for the fresh air scenario,  
18 but for the recirculating air ventilation the concentration increases as a logarithmic-type function as CO  
19 builds up in the cabin of the vehicle.



Source: Abi Esber and El Fadel (2008)

**Figure 3-32. Comparison of in-vehicle (solid line) and engine (dotted line) results for (left) driving with windows closed and air conditioner in recirculating air mode, and (right) driving with windows closed and air conditioner in fresh air mode.**

1 Riedeker et al. (2003) measured CO concentrations inside patrol cars during shifts. Troopers  
2 recorded in a time-activity diary the ventilation settings of their cars and exit/entry from the vehicle, and  
3 the air conditioning was typically set to recirculation mode during the shifts. Riedeker et al. (2003) found  
4 that CO concentrations (mean, SD:  $2.6 \pm 1.1$  ppm) were higher than ambient monitor concentrations  
5 ( $0.8 \pm 0.3$  ppm). They were also higher than roadside CO concentrations ( $1.1 \pm 0.3$  ppm), indicating that  
6 either the vehicle itself contributes to in-cabin CO, or on-road concentrations are higher than roadside  
7 concentrations, or both. Riedeker et al. (2003) noted that within-shift variability was higher than between-  
8 shift variability, which underscores the variability in police officers' activities during a given shift. No  
9 data was segregated by ventilation settings. Chang et al. (2000) measured CO concentrations during a  
10 scripted activity study in Baltimore, MD in 1998 and 1999. Mean 1-h CO concentrations were near the  
11 1 ppm detection limit of the Langan CO monitor. Microenvironmental CO concentrations were  
12 significantly correlated with concentrations measured at a fixed-site ambient monitor for residential, other  
13 indoor, in-vehicle, and outdoor near-road microenvironments during the winter. Significant correlations  
14 were observed only for residential microenvironments during the summer. The location of the ambient  
15 monitor near a roadway may have contributed to the lack of correlation with concentrations measured at  
16 outdoor locations away from roadways. Microenvironmental concentrations inside vehicles were  
17 significantly higher than those for other microenvironments.

18 Several factors may affect commuters' exposures. Van Wijnen et al. (1995) observed that while  
19 bicycling, cyclists inhale 2.3 times as much air as pedestrians and drivers so that this group may, in fact,  
20 inhale the most CO. Likewise, in their review of roadway exposures to CO and PM, Kaur et al. (2007)  
21 listed a number of uncontrolled factors that may influence exposure. Vertical CO concentration gradients  
22 have been documented with concentrations decreasing with height; lower breathing zone height among  
23 children may make them more likely to be exposed to higher CO concentrations. With respect to  
24 transportation, Kaur et al. (2007) suggested that vehicle ventilation, speed, position in traffic, and  
25 start/stop activity influence in-vehicle exposures. Gómez-Perales et al. (2007) also noted that meteorology  
26 can impact in-vehicle exposures, with evening increases in wind speed causing a 50% reduction in CO  
27 exposures among bus and minibus commuters. Alm et al. (1999) made a similar observation in a study of  
28 urban commuters' exposure within a vehicle.

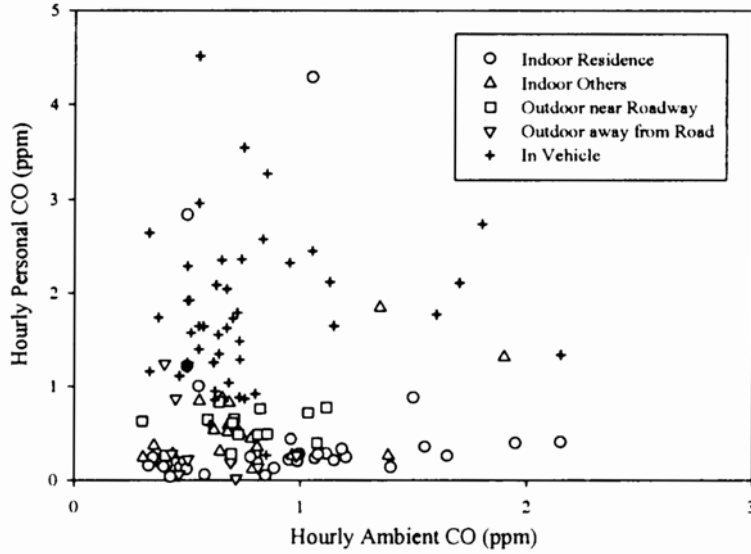
29 Studies of vehicle self-pollution are instructive for considering potential for ambient CO infiltration  
30 in vehicles. Behrentz et al. (2004) used sulfur hexafluoride ( $\text{SF}_6$ ) tracer gas emitted from school bus  
31 engines to determine the proportion of in-vehicle pollution related to self-pollution. Based on the  $\text{SF}_6$   
32 concentration, they calculated that 0.04-0.29% of the bus cabin air contained exhaust for high emitting  
33 diesel engines, 0.01-0.03% for "regular" diesel buses, 0.02-0.04% for buses fitted with a particle trap, and  
34 0.03-0.04% for buses running on compressed natural gas.  $\text{SF}_6$  concentrations were higher when bus  
35 windows were closed.

### 3.6.5.3. CO Exposure Assessment Variability and Error

1 In the context of determining the effects of ambient pollutants on human health, the association  
2 between the ambient component of personal exposures and ambient concentrations is more relevant than  
3 the association between total personal exposures (ambient component + non-ambient component) and  
4 ambient concentrations. If there are no non-ambient sources of a pollutant, the total personal exposure is  
5 equal to the ambient personal exposure. However, non-ambient sources could significantly affect personal  
6 exposures to CO. Unlike PM, CO has no chemical identifiers that can be used to apportion ambient and  
7 non-ambient sources of CO (other than radioactive isotopes, whose detection is rare). For the general U.S.  
8 population, exposure error analysis for epidemiologic studies indicates that fixed-site measured ambient  
9 CO concentration is generally a good indicator of ambient exposure to CO, as discussed in more detail  
10 below.

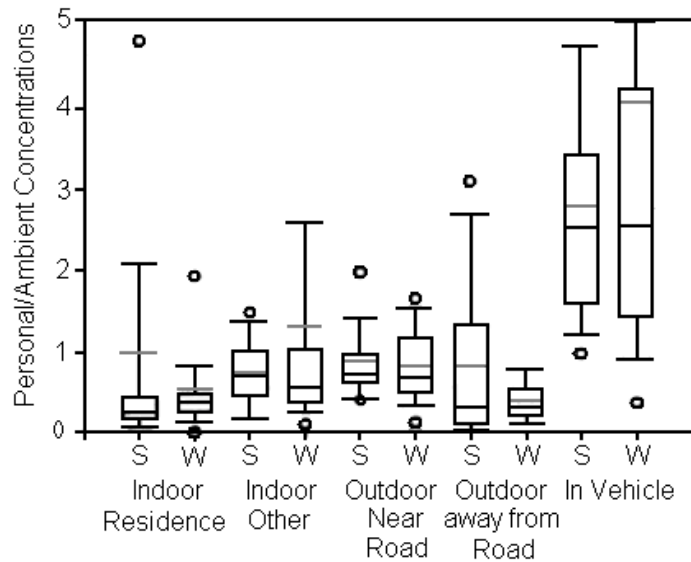
11 Figure 3-33 shows hourly vs. personal CO concentration data obtained by Chang et al. (2000) for a  
12 1998-1999 multi-pollutant sampling campaign in Baltimore, MD. Personal exposures were obtained in  
13 five separate microenvironments in this study. A high degree of scatter is evident in this figure, which  
14 suggests that these personal exposures are influenced by both ambient and non-ambient sources of CO.  
15 Figure 3-34 is a box plot of the personal-to-ambient CO concentration ratio for the same five  
16 microenvironments. Figure 3-34 shows that personal exposures in vehicles were on average 2.8 times  
17 higher than ambient during the summer and 4.1 times higher than ambient in the winter. For the other four  
18 microenvironments tested, the average ratio was around 1. Wide variability is seen in these plots,  
19 particularly during the summer. Much of that variability could be due to the influence of non-ambient  
20 sources, which would then result in poor correlation between total personal exposure and ambient  
21 concentration.

22 Zeger et al. (2000) stated that there are three error terms in the estimate of a person's exposure: 1)  
23 error from using pooled population data in lieu of individual data, 2) error between total exposure and  
24 ambient concentration, and 3) error between the actual and measured ambient concentration. Zeger et al.  
25 (2000) also described that errors related to individual variability and measurement (as stated above for the  
26 latter) are of the Berkson type and therefore not expected to bias estimates of exposure or health effects.  
27 Moreover, Sheppard et al. (2005) simulated ambient and non-ambient exposures to a non-reactive  
28 pollutant and observed that non-ambient CO exposure has no effect on the association between ambient  
29 CO exposure and health outcomes for the case where ambient and non-ambient sources were  
30 independent.



Source: Chang et al. (2000)

**Figure 3-33.** Hourly personal vs. ambient CO concentrations obtained in Baltimore, MD. Summer of 1998 in five settings: indoor residence, indoor other, outdoor near road, outdoor away from road, in vehicle.



Source: Adapted from Chang et al. (2000)

**Figure 3-34.** Box plots of the ratio of personal to ambient concentrations obtained in Baltimore, MD. Summer of 1998 and winter of 1999 in five settings: indoor residence, indoor other, outdoor near road, outdoor away from road, in vehicle. The dotted line shows the mean, and the solid line shows the median. S = summer; W = winter.

1 For community time-series epidemiology, the community averaged concentration, not the  
2 concentration at each fixed monitoring site, is the concentration variable of concern (Zeger et al., 2000).  
3 The correlation between the concentration at a central community ambient monitor and the community  
4 averaged concentration depends on at least the following three factors. First, the distribution of monitors  
5 across space: CO emissions from traffic might show spatial heterogeneity near roadway sources but have  
6 a more homogeneous distribution farther downwind. It follows that use of a large number of samplers will  
7 dampen inter-sampler variability. Second, the relationship between the measurement at the ambient  
8 monitoring site and the community average concentration: if the site is selected to measure a “hot spot” or  
9 pollution from a nearby source, estimates of community exposure could be skewed upwards. Third,  
10 terrain features or source location: different terrain features or source locations across sub-communities  
11 may differ in the temporal pattern of pollution. Intra-urban spatial heterogeneity was discussed in detail in  
12 Section 3.5.1.2. Community exposure may not be well-represented when monitors cover large areas with  
13 several sub-communities having different sources and topographies. For example, with the exception of  
14 two closely located samplers, there is poor inter-sampler correlation among the Pittsburgh, PA CO  
15 samplers. This variability reflects the wide topographical differences throughout the Pittsburgh CSA as  
16 well as variation in traffic usage and distance to near-road sources at each point in contrast to the Phoenix  
17 CBSA, where inter-sampler correlation was high.

18 Epidemiologic studies of long-term exposure rely on differences among communities in long-term  
19 average ambient concentrations. If exposure errors are different in the different communities, the  
20 differences in long-term ambient concentrations among communities may not represent the differences in  
21 long-term average exposures. For example, there may be community to community differences in  
22 measurement error, in the average ambient exposure factor ( $\alpha$ ) or the average non-ambient exposure. This  
23 could happen if exposure to newly formed pollutants generated by vehicular traffic or pollutants from  
24 other localized sources differed among the spatial areas. Thus, in a regression of health effects against  
25 average concentration (as an indicator for average exposure) there could be a different amount of error  
26 (either positive or negative) in the exposure indicator for each spatial area. This could add error and bias  
27 the slope up or down. Geostatistical tools enable the use of concentration fields that include spatial  
28 variations in concentration. However, it has not yet been possible to include individual or small-area  
29 variations in the personal exposure to ambient concentrations or variations in personal exposures to  
30 indoor-generated pollutants in long-term studies of the associations of pollutants with health effects.

### 3.6.6. Multi-pollutant Exposures

31 Since incomplete combustion is the primary source of ambient CO, exposure to ambient CO is  
32 accompanied by exposure to other combustion-related pollutants, such as NO<sub>x</sub> and PM. Thus, ambient



1 CO is often considered a surrogate for exposure to traffic-generated pollutants. However, the specific mix  
2 of CO with NO<sub>x</sub> and PM depends on the source; for example, the mixture generated by gasoline engines  
3 differs from that produced by gas combustion. Correlations between ambient CO and ambient PM<sub>2.5</sub>,  
4 PM<sub>10</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and O<sub>3</sub> from AQS data and the peer-reviewed literature were presented in section 3.5.3.  
5 Nationwide ambient CO was most highly correlated with ambient NO<sub>2</sub> followed by PM<sub>2.5</sub> and PM<sub>10</sub>.  
6 Correlations between CO and PM<sub>2.5</sub> were not always statistically significant on a national basis;  
7 correlations were insignificant for ambient CO with ambient SO<sub>2</sub> and ambient PM<sub>10</sub>, and ambient CO was  
8 negatively correlated with ambient O<sub>3</sub>. Correlations of ambient CO with PM<sub>2.5</sub> and PM<sub>10</sub>, were much  
9 higher and more statistically significant at some of the select cities, as shown in Annex A.

10 Sarnat et al. (2001) analyzed the relationship between personal exposure to PM<sub>2.5</sub> and ambient CO  
11 concentrations and found significant associations during winter of ambient CO with personal exposure to  
12 total PM<sub>2.5</sub> (slope = 3.99) and to ambient PM<sub>2.5</sub> (slope = 6.30). Ambient CO was not significantly  
13 associated with personal PM<sub>2.5</sub> exposure during summer. Sarnat et al. (2005) also reported significant  
14 association between ambient CO concentrations and personal PM<sub>2.5</sub> exposure for winter (three 12-day  
15 sessions in 2000, dates not specified) only and significant year-round associations between ambient CO  
16 concentrations and personal exposure to SO<sub>4</sub><sup>2-</sup> aerosols.

17 Relationships between personal CO exposures and copollutants were reported less frequently in the  
18 literature, but results from these studies were consistent with the findings cited above. In a study of  
19 personal exposures to CO, PM<sub>2.5</sub>, and ultrafine PM in a street canyon, Kaur et al. (2005) found low  
20 Pearson's correlation of personal CO exposure with personal PM<sub>2.5</sub> exposure (r = 0.23). Personal CO  
21 exposure had much better correlation with personal ultrafine PM exposure (r = 0.68). Chang et al. (2000)  
22 reported correlations of personal CO exposure with personal PM<sub>2.5</sub>, personal toluene, and personal  
23 benzene exposures in Baltimore, MD at five locations, labeled indoor residential, indoor nonresidential,  
24 outdoors near roadway, outdoors away from road, and in vehicle. Much variability was observed in the  
25 correlations for different locations and seasons (winter vs. summer). In general, the correlations tended to  
26 be stronger and more significant in the winter. Chang et al. (2000) pointed out that indoor air exchange  
27 rates tend to be lower in winter, which could cause these correlations to be more sensitive to non-ambient  
28 sources. Significant associations of CO with benzene and toluene were observed in-vehicle. Because CO  
29 exposures most often occur together with exposure to other combustion-related pollutants, interpretation  
30 of health studies using ambient CO data can be a challenge, as discussed further in Chapter 5.

### 3.6.7. Exposure Modeling

31 A number of modeling techniques to describe air pollutant exposures at the individual and  
32 population level have been published since 2002. Ambient CO exposure models often assume that

1 vehicular traffic is the sole source of emissions given that roughly 90% of urban CO is estimated to come  
2 from traffic (Gulliver and Briggs, 2005). Individual-level models can predict personal exposure based on  
3 a person's path of travel and on conditions of the surrounding environment using traffic and meteorology  
4 data. One such example is the Space-Time Exposure Modeling System (STEMS) (Gulliver and Briggs,  
5 2005). Population-based methods, such as the Air Pollution Exposure (APEX) model  
6 ([www.epa.gov/ttn/fera/human\\_apex.html](http://www.epa.gov/ttn/fera/human_apex.html)), Stochastic Human Exposure and Dose Simulation (SHEDS)  
7 (Burke et al., 2001), and EXPOLIS models (Kruize et al., 2003), involve stochastic treatment of the  
8 model input factors. Another approach is to predict location-based exposures using a deterministic model  
9 such as the CMAQ model, California Line Source Dispersion Model (CALINE), CALPUFF (long-range  
10 plume transport model), or Operational Street Pollution Model (OSPM) for determination of street-level  
11 pollution coupled with infiltration models to represent indoor exposure to ambient levels (Appel et al.,  
12 2008; Gilliam et al.; Hering, 2007; Mensink and Cosemans, 2008; Wilson and Zawar-Reza, 2006).  
13 Stochastic and deterministic methods are often combined, as described below. Land use regression (LUR)  
14 models have also been developed to describe pollution levels as a function of geographic source behavior  
15 (Gilliland et al., 2005; Ryan and LeMasters, 2007; Veen et al., 1997). Similarly, other spatial interpolation  
16 techniques have been used to determine geographic distributions of exposure (Marshall et al., 2008).  
17 Compartmental models, such as the Indoor Air Model (INDAIR), can be used to assess exposure to  
18 infiltrated ambient air pollutants in a deterministic or probabilistic framework (Dimitroulopoulou et al.,  
19 2001; 2006). A detailed description of various modeling techniques is provided in the 2008 NO<sub>x</sub> ISA and  
20 2008 SO<sub>x</sub> ISA annexes (2008f, g), and an explanation of associated modeling errors is provided in the  
21 2008 Draft PM ISA (U.S. EPA, 2008f). Applications of these models to assessment of CO exposure are  
22 described in the following paragraphs.

23 The STEMS model maps exposures based on inputs for traffic levels, atmospheric dispersion,  
24 background concentrations, and geography. Gulliver and Briggs (2005) tested the STEMS model for CO  
25 and observed some correlation between modeled and measured CO concentrations ( $R^2 = 0.41$ ), which was  
26 consistent with results for PM<sub>10</sub> and NO<sub>x</sub>. Exposures were estimated from the predicted ambient CO  
27 concentration using a term similar to  $\alpha$  that varied depending on whether the individual was walking or in  
28 a vehicle. Gulliver and Briggs (2005) noted that a limitation to modeling CO is the scarcity of background  
29 CO data obtained at rural sites. For this reason, they assumed a constant value obtained from estimates  
30 made over the North Atlantic Ocean. Although the authors only presented detailed results for a model of  
31 PM<sub>10</sub> based on traffic and meteorology in Northampton, U.K., they found that the majority of variation on  
32 a given day in modeled exposure among school children was due to differences in travel routes. Variation  
33 across days was also influenced by background and meteorological conditions. Similar results can be  
34 expected for CO based on the tendency for variation of the CO concentration profile on the neighborhood  
35 and micro-scales (Jerrett et al., 2004). Flachsbart (Flachsbart) tested numerous meteorological, traffic, and  
36 background CO input variables in a regression approach to predicting CO exposure among individuals

1 while traveling in a vehicle. This work showed travel time and average speed of on-road vehicles to be  
2 important determinants of CO exposure in a vehicle. Results from individual models of this nature can be  
3 pooled to develop a distribution for examination of population effects or for comparison with population  
4 exposure models.

5 Bruinen de Bruin et al. (2004) utilized the EXPOLIS model to predict CO population exposures in  
6 Milan, Italy. The simulation results showed that the U.S. 8-h NAAQS level was exceeded in one case out  
7 of 1,000. The model also showed that exposures exceeded 20 ppm in one case out of 100,000. The results  
8 were not shown to be very sensitive to the number of microenvironments (e.g., outdoors, indoors, in  
9 vehicle) included in the model. The probabilistic NAAQS Exposure Model for CO (pNEM/CO) was used  
10 in the 1992 and 2000 CO AQCDs (U.S. EPA, 1992, 2000) to estimate personal CO exposures among  
11 adult populations. Given reductions in ambient CO levels, details, and results from those efforts are  
12 provided here. The current version of this model, now known as APEX, was used for exposure  
13 assessment in the O<sub>3</sub> NAAQS review (U.S. EPA, 2006b).

14 To examine indoor concentrations of ambient CO, Dimitroulopoulou et al. (2006) used the  
15 probabilistic formulation of the INDAIR model to examine indoor exposure to ambient CO, along with  
16 NO<sub>x</sub> and PM for a given distribution of background CO levels, meteorology, residential air exchange rate,  
17 and residential room dimensions. They found that 24-h avg CO concentration increased from 1.86 ppm  
18 outdoors to 1.90-1.93 ppm indoors in the absence of non-ambient sources, and that indoor 24-h avg CO  
19 concentration could increase to 1.93-2.00 ppm in the presence of smoking and to 1.98-2.32 ppm in the  
20 presence of gas cooking. Similarity between the outdoor and non-source indoor concentrations was  
21 attributed to the lack of CO loss mechanisms. With maximum CO concentrations modeled at 3.3 ppm in  
22 the absence of sources, the probability density function of indoor exposure to ambient fell well below the  
23 8-h NAAQS. In the Reducing Urban Pollution Exposure from Road Transport (RUPERT) study, Bell et  
24 al. (2004) presented methodology to use the probabilistic form of INDAIR for development of personal  
25 exposure frequency distributions of CO, NO<sub>x</sub>, and PM based on time spent in residential, transportation,  
26 school, office, and recreational environments with inputs from transportation source categories (Chen et  
27 al., 2008).

28 Another set of approaches to improve exposure estimates in urban areas involves construction of a  
29 concentration surface over the geographic area. This does not estimate exposure directly because it does  
30 not account for activity patterns or concentrations in different microenvironments. It provides an  
31 improved estimate of the expected local outdoor concentration near residences, schools or workplaces,  
32 and roadways across the area. There are two main types of approaches: spatial interpolation of measured  
33 concentrations, and regression models using land use, roadway characteristics, and other variables to  
34 predict concentrations at receptors in the domain. Rigorous first-principles models, such as dispersion  
35 models and chemical transport models, can also be used for this type of application, but are less suitable  
36 because they have intensive resource requirements and are typically applied over larger domains.

1 Marshall et al. (2008) compared four spatial interpolation techniques for estimation of CO  
2 concentrations in Vancouver, BC. The investigators assigned a daily average CO concentration to each of  
3 the 51,560 postal code centroids using one of the following techniques: (1) the concentration from the  
4 nearest monitor within 10 km, (2) the average of all monitors within 10 km, (3) the inverse-distance-  
5 weighted (IDW) average of all monitors in the area, and (4) the IDW average of the three closest monitors  
6 within 50 km. Method 1 (the nearest-monitor approach) and Method 4 (IDW-50 km) had similar mean  
7 and median estimated annual average concentrations, although the 10th-90th percentile range was smaller  
8 for IDW-50. This is consistent with the averaging of extreme values inherent in IDW methods. The  
9 Pearson correlation coefficient between the two methods was 0.88. Methods 2 and 3 were considered sub-  
10 optimal and were excluded from further analysis. In the case of method 2, a single downtown high-  
11 concentration monitor skewed the results in the vicinity, partially as a result of the asymmetric layout of  
12 the coastal city of Vancouver. Method 3 was too spatially homogenous, because it assigned most locations  
13 a concentration near the regional average, except for locations immediately adjacent to a monitoring site.  
14 LUR results were also reported in this study for NO and NO<sub>2</sub>, and indicated that LUR's higher spatial  
15 precision reflects neighborhood-scale effects from nearby land use, but may not account for urban-scale  
16 variation. These results highlight the variation in local concentration estimates with choice of estimation  
17 technique.

## 3.7. Summary and Conclusions

### 3.7.1. Sources of CO

18 In the U.S., on-road mobile sources constituted more than half of total CO emissions in 2002, or  
19 ~63 MT of ~109 MT total. In metropolitan areas in the U.S., for example, as much as 70-75% of all CO  
20 emissions can come from on-road vehicle exhaust (U.S. EPA, 2006a). The majority of these on-road CO  
21 emissions derive from gasoline-powered vehicles since the O<sub>2</sub> content, pressure, and temperature  
22 required for diesel fuel ignition do not produce large quantities of CO. Anthropogenic CO emissions are  
23 estimated to have decreased 35% between 1990 and 2002. On-road vehicle sector emissions controls have  
24 produced nearly all these national-level CO reductions. Nationally, biogenic emissions, excluding fires,  
25 were estimated to contribute ~5% of total CO emissions from all sources in 2002, and fires in 2002 added  
26 another 13%, or ~14.5 MT, to the national CO emissions total.

## 3.7.2. Physics and Chemistry of Atmospheric CO

1 CO is produced by photooxidation of CH<sub>4</sub> and other VOCs in the atmosphere. Estimating the CO  
2 yield from oxidation of HCs larger than CH<sub>4</sub> requires computing the yields of several intermediate  
3 products and reactants from oxidation of the parent molecules. The major pathway for removal of CO  
4 from the atmosphere is reaction with OH to produce CO<sub>2</sub> and HO<sub>2</sub>. The mean photochemical lifetime ( $\tau$ )  
5 of CO in the northern hemisphere is ~57 days. During winter at high latitudes, CO has nearly no  
6 photochemical reactivity on urban and regional scales.

## 3.7.3. Ambient CO Measurements

7 As of March 2008, 19 automated FRMs and no FEMs had been approved for CO. All EPA FRMs  
8 for CO operate on the principle of nondispersive infrared (NDIR) detection and can include gas filter  
9 correlation (GFC). The required lower detection limit for FRMs in the EPA network is 1.0 ppm. In 2007,  
10 there were 285 CO monitors meeting the 75% completeness requirements and reporting values year-  
11 round to the AQS in the 50 states, plus the District of Columbia, Puerto Rico, and the Virgin Islands. At  
12 least 70 monitors across the U.S. have been positioned at microscale within 10 m of a road to capture  
13 near-road concentrations. At larger scales, monitor distance from a road is inversely related to the road's  
14 average daily traffic count to capture community averages.

## 3.7.4. Environmental CO Concentrations

15 CO concentration data for 1-h and 8-h intervals were available for 243 counties and autonomous  
16 cities or municipalities that maintained active CO monitoring stations meeting the 75% completeness  
17 criteria for the years 2005-2007. There were no violations of the 1-h or 8-h NAAQS in those years. The  
18 nationwide mean, median, and interquartile range for 1-h measurements reported between 2005-2007  
19 were 0.5, 0.4, and 0.4 ppm, respectively, and these statistics did not change by more than 0.1 ppm for  
20 each year of the 3-year period. The nationwide mean, median, and interquartile range for 8-h daily max  
21 concentrations reported between 2005-2007 were 0.7, 0.5, and 0.5 ppm, respectively. The 2006 annual  
22 second highest 8-h CO concentration averaged across 144 monitoring sites nationwide was 75% below  
23 that for 1980 and is the lowest concentration recorded during the past 27 years. The mean annual second  
24 highest 8-h ambient CO concentration has been below 5 ppm since 2004. The downward trend in CO  
25 concentrations in the 1990s parallels the downward trend observed in CO emissions and can be attributed  
26 largely to decreased mobile source emissions.

27 The correlation structures for measurements at the monitors in each of the 11 CSAs/CBSAs  
28 examined for this assessment reveal a wide range of response between monitors in each city and among

1 the cities judged against each other. While this wide range is produced by the interactions of many  
2 physical and chemical elements, the location of each monitor and the uniqueness of its immediate  
3 surroundings can often explain much of the agreement or lack thereof. CO concentrations can be elevated  
4 near roadways and decrease with increasing distance from the road. Likewise, micro- and neighborhood-  
5 scale variation related to urban topography or microenvironmental source distribution may have a  
6 significant effect on ambient CO concentrations with respect to street canyon concentrations. Anchorage,  
7 AK had concentrations roughly twice those of the other metropolitan areas. Most of the CSAs/CBSAs  
8 examined here had diel concentration curves with pronounced morning and evening rush hour peak CO  
9 levels, although diel CO concentrations had less variability for New York City, Atlanta, and Seattle than  
10 for the other eight cities. For most metropolitan areas examined here, concentrations were generally  
11 highest in the winter (December-February) and fall (September-November) and decreased, on average,  
12 during the spring (March-May) and summer (June-August). Measurements near or below the required  
13 FRM lower detection limit of 1.0 ppm coupled with the coarsely reported measurement resolution  
14 (0.1 ppm) can artificially influence the comparison statistics shown in the tables and result in apparent  
15 heterogeneity in the box plots (Figures 3-13 through Figures 3-16).

16 For nationwide copollutant correlation data, ambient CO had the highest mean and median  
17 correlations with ambient NO<sub>2</sub>, followed by PM<sub>2.5</sub> (correlations vary by season). The SO<sub>2</sub> and PM<sub>10</sub>  
18 correlations did not appear to be statistically significantly different from zero for any season, and  
19 statistical significance was only observed in winter and fall for PM<sub>2.5</sub>. Correlations between CO and NO<sub>2</sub>  
20 were significant for all seasons but summer. Correlations between CO and O<sub>3</sub> were significant and  
21 negatively correlated for winter only. The copollutant correlation plots shown in Annex A illustrate higher  
22 and more statistically significant correlations between CO and PM in several but not all of the select cities  
23 displayed.

24 This assessment has used data from 2005-2007 at 12 remote sites as part of the international CCGG  
25 CASN in the CONUS, Alaska, and Hawaii to determine PRB. All sites demonstrate the well-known  
26 seasonality in background CO with minima in the summer and fall and maxima in the winter and spring.  
27 The 3-year avg CO PRB in Alaska was 130 ppb; in Hawaii it was 99 ppb; and over the CONUS it was  
28 132 ppb.

### 3.7.5. Exposure Assessment and Implications for Epidemiology

29 Very few recent exposure assessment studies involve CO concentration data. The studies of  
30 personal exposure to CO presented here generally found that the largest fraction of an individual's  
31 exposure to ambient CO occurs indoors but the highest CO exposure levels occur in transit. Among  
32 commuters, exposures were highest for those traveling in automobiles in comparison with those traveling

1 in buses and motorbikes and with those cycling or walking. A portion of that exposure is thought to come  
2 from the vehicle in which the exposed person travels. Commuting time was an important predictor of  
3 exposure, such that subjects living in low-CO concentration suburban areas and commuting to work  
4 experienced higher CO levels than residents of higher-CO concentration urban areas with short commute  
5 times. Additional analyses indicated that on-road mobile emissions were the most important source of CO  
6 exposure for non-ETS-exposed subjects. Results have also indicated that time spent outdoors and  
7 proximity to roadways may have been important in determining personal exposure in children.  
8 Uncontrolled factors that may influence exposure include vertical CO concentration gradients; vehicle  
9 ventilation, speed, position in traffic, and start/stop activity; and meteorology.

10 For the general U.S. population, exposure error analysis for epidemiologic studies indicates that  
11 fixed-site measured ambient CO concentration is generally a good indicator of ambient exposure to CO.  
12 Errors associated with exposure to non-ambient CO sources, for example smoking or gas cooking, are of  
13 the Berkson type and therefore not expected to bias estimates of exposure or health effects. Simulations of  
14 ambient and non-ambient exposures to a non-reactive generic pollutant demonstrated that non-ambient  
15 exposure had no effect on health effects outcomes for the case where ambient and non-ambient sources  
16 were independent. Likewise for community time-series epidemiology, selected monitoring sites are  
17 thought to be representative of community averages. Measurement at a “hot spot” could skew community  
18 exposure estimates upward. Topographical or source variability could produce differences in the temporal  
19 pattern of pollution. Differences such as those produced by variation in local sources among communities  
20 in long-term exposure studies may also produce error in estimation of the ambient exposure factor. Such  
21 differences could add error, and therefore influence the slope of health effects estimate regressions up or  
22 down.

23 Several studies have examined multi-pollutant exposure. Since incomplete combustion is the  
24 primary source of ambient CO, exposure to ambient CO is accompanied by exposure to other  
25 combustion-related pollutants, such as NO<sub>x</sub> and PM. High correlations of ambient CO with NO<sub>2</sub> and  
26 PM<sub>2.5</sub> have been observed in the peer-reviewed literature and AQS data. Thus, ambient CO is often  
27 considered a surrogate for exposure to traffic-generated pollutants. Because CO exposures most often  
28 occur together with exposure to other combustion-related pollutants, interpretation of health studies using  
29 ambient CO data can be a challenge, as discussed further in Chapter 5.

# Chapter 4. Dosimetry and Pharmacokinetics of Carbon Monoxide

## 4.1. Introduction

1           Inhaled ambient CO elicits various health effects by binding with and altering the function of a  
2 number of heme-containing molecules, mainly Hb. Traditional concepts for CO pathophysiology have  
3 been based on the high affinity of CO for deoxyhemoglobin, resulting in COHb formation and consequent  
4 reduction in O<sub>2</sub>-carrying capacity of blood and impaired O<sub>2</sub> delivery to tissues. Research on the basics of  
5 CO pharmacokinetics date back to the 1890s, but since the late 1970s has become limited. Current  
6 literature primarily focuses on endogenous CO produced by the metabolic degradation of heme by heme  
7 oxygenase (HO) and its role as a gaseous messenger. This chapter reviews the physiology and  
8 pharmacokinetics of CO. The chapter draws heavily from Chapter 5 of the previous AQCD (U.S. EPA,  
9 2000). Relevant new data are included when available. New explanations of recent models of Hb binding  
10 are included, as well as discussion on tissue CO concentrations using new methods of extraction.

11           CO binds with a number of heme-containing molecules including Mb and cytochromes, but none  
12 have been studied as extensively as Hb. The primary focus of this chapter is placed on the models and  
13 kinetics of such binding and the factors influencing this event. The chapter discusses effects at ambient or  
14 near ambient levels of CO leading to low COHb levels ( $\leq 5\%$ ); however few studies are available at  
15 ambient CO concentrations. Both human and animal studies using higher CO exposure concentrations,  
16 resulting in moderate to high COHb levels (<20%), are discussed where needed to understand CO  
17 kinetics, pathophysiologic processes, and mechanisms of cytotoxicity. Where human studies could not  
18 experimentally test certain hypotheses or were unavailable, animal experiments were used as surrogates.  
19 CO uptake and elimination has been shown to be inversely proportional to body mass over  
20 environmentally relevant exposure levels, meaning the smaller the animal, the faster the rate of absorption  
21 and elimination (Klimisch et al., 1975; Tyuma et al., 1981). However, the basic mechanisms of CO  
22 toxicity between experimental animals and humans are similar and are thus extrapolated from animals to  
23 humans in this chapter, keeping in mind a number of interspecies differences.



## 4.2. Carboxyhemoglobin Formation

### 4.2.1. The Coburn-Forster-Kane and Other Models

1           Investigators have modeled the effect of CO binding to Hb in a number of ways. Empirical and  
2 mechanistic models are two distinct approaches that have been taken to model in vivo COHb formation  
3 after CO exposure. First, empirical models were used to predict COHb by regressing concentration and  
4 duration of exogenous CO exposure with COHb. These methods were reviewed in depth in the previous  
5 AQCD (U.S. EPA, 2000). These models are limited to estimating COHb in the exact conditions on which  
6 the models were based. These simple models include those by Peterson and Stewart (1970) and Ott and  
7 Mage (1978), as well as various others (Selvakumar et al., 1992; Singh et al., 1991). Using these simple  
8 models, it was shown that the presence of brief ambient CO concentration spikes averaged over hourly  
9 intervals may lead to underestimating the COHb concentration by as much as 21% of the true value. To  
10 avoid this problem, it was suggested that ambient CO measurements be monitored and averaged over 10–  
11 15 minute periods (Ott and Mage, 1978).

12           Secondly, mechanistic models use physical and physiological processes and an understanding of  
13 biological processes to predict COHb production. The most commonly used mechanistic method for  
14 predicting levels of blood COHb after CO inhalation is the Coburn-Forster-Kane equation or CFK model  
15 developed in 1965 (Coburn et al., 1965). This differential equation was developed to examine endogenous  
16 CO production, using the major physiological and physical variables influencing this value. Since then, it  
17 has been shown to provide a good approximation to the COHb level at a steady level of inhaled  
18 exogenous CO (Peterson and Stewart, 1975; Stewart et al., 1973b). The CFK model describes a four-  
19 element, physical system containing an exogenous CO source, a transfer interface, an endogenous CO  
20 source, and a storage compartment. The linear CFK model assumes O<sub>2</sub>Hb concentration is constant and is  
21 as follows in Equation 4-1:

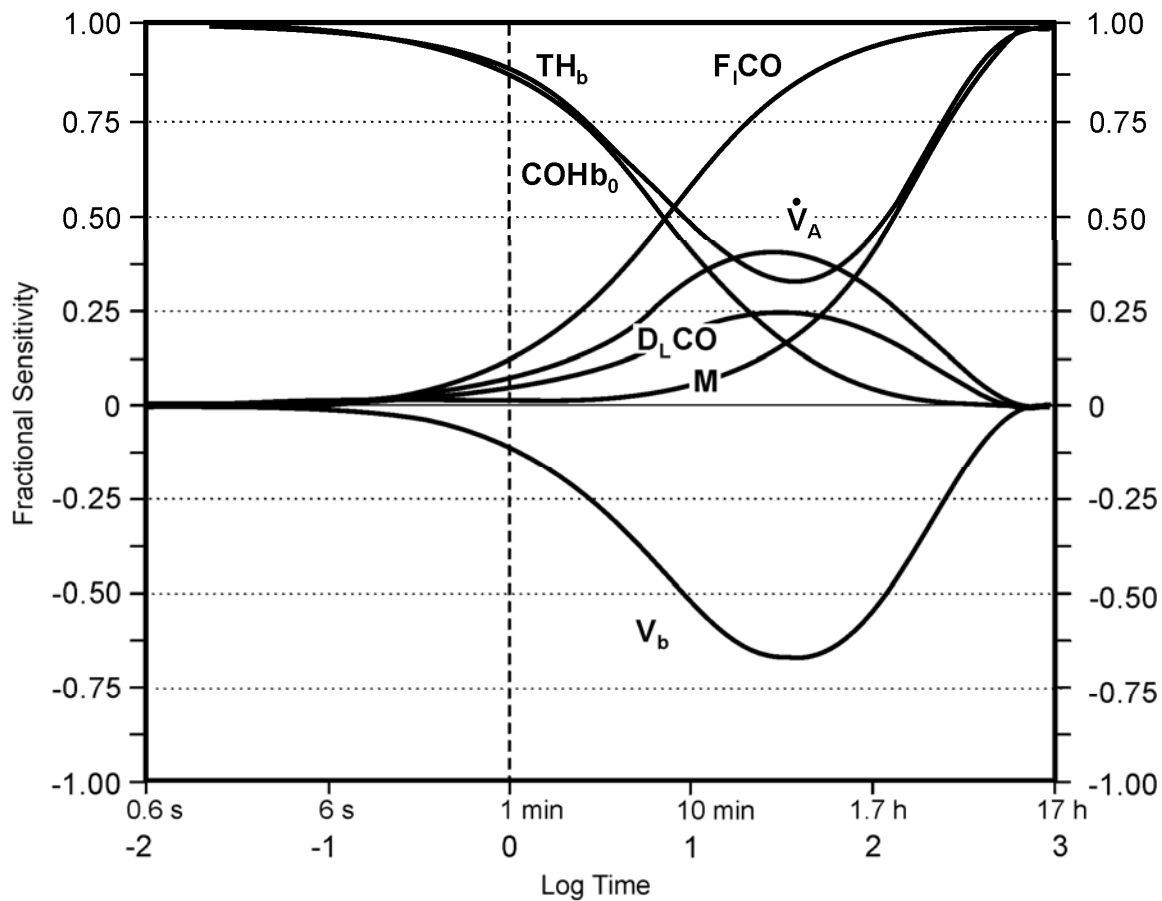
$$V_b \frac{d[\text{COHb}]_t}{dt} = \dot{V}_{\text{CO}} - \frac{[\text{COHb}]_0 P_{\text{C}} \text{O}_2}{[\text{O}_2\text{Hb}]M} \left( \frac{1}{\frac{1}{D_L \text{CO}} + \frac{P_B - 47}{\dot{V}_A}} \right) + \left( \frac{P_I \text{CO}}{\frac{1}{D_L \text{CO}} + \frac{P_B - 47}{\dot{V}_A}} \right)$$

*Equation 4-1*

22           where  $V_b$  is blood volume in milliliters (mL);  $[\text{COHb}]_t$  is the COHb concentration at time  $t$  in mL CO/mL  
23 blood, at standard temperature and pressure, dry (STPD);  $\dot{V}_{\text{CO}}$  is the endogenous CO production rate in  
24 mL/min, STPD;  $[\text{COHb}]_0$  is the COHb concentration at time zero in mL CO/mL blood, STPD;  $[\text{O}_2\text{Hb}]$  is  
25 the O<sub>2</sub>Hb concentration in mL O<sub>2</sub>/mL blood, STPD;  $M$  is the Haldane coefficient representing the CO

1 chemical affinity for Hb;  $P_{\bar{c}O_2}$  is the average partial pressure of  $O_2$  in lung capillaries in mmHg;  $\dot{V}_A$  is  
2 the alveolar ventilation in mL/min, STPD;  $D_LCO$  is the lung diffusing capacity of CO in mL/min/mmHg,  
3 STPD;  $P_B$  is the barometric pressure in mmHg; and  $P_I CO$  is the CO partial pressure in inhaled air in  
4 mmHg.

5 This model assumes instant equilibration of COHb concentration between venous and arterial  
6 blood, gases in the lung, and COHb concentrations between blood and extravascular tissues, which is not  
7 physiologically representative. The nonlinear CFK equation incorporates the interdependence of COHb  
8 and  $O_2Hb$  levels since they are derived from the same pool of blood Hb. The nonlinear equation is more  
9 physiologically accurate; however the linear CFK equation gives a good approximation to the nonlinear  
10 solution over a large range of values during CO uptake and during low levels of CO elimination (Smith,  
11 1990). The linear equation prediction of COHb concentration at or below 6% will only differ  $\pm 0.5\%$  from  
12 the nonlinear equation prediction. Sensitivity analysis of the CFK equations has shown that alterations in  
13 each variable of the equation will affect the outcome variably at different times of exposure (McCartney,  
14 1990). Figure 4-1 illustrates the temporal changes in fractional sensitivities of the principal physiological  
15 determinants of CO uptake for the linear form of the CFK equation, where  $TH_b$  is the total blood  
16 concentration of Hb in g Hb/mL blood and  $F_I CO$  is the fractional concentration of CO in ambient air  
17 in ppm. The fractional sensitivity of unity means that, for example, a 5% error in the selected variable  
18 induces a 5% error in the predicted COHb value by the nonlinear model. As Figure 4-1 demonstrates, a  
19 constant or given percent error in one variable of the model does not generally produce the same error in  
20 the calculated blood COHb, and the error is time dependent. Thus, each variable influencing CO uptake  
21 and elimination will exert its maximal influence at different times of exposure.



Source: modified from McCartney (1990)

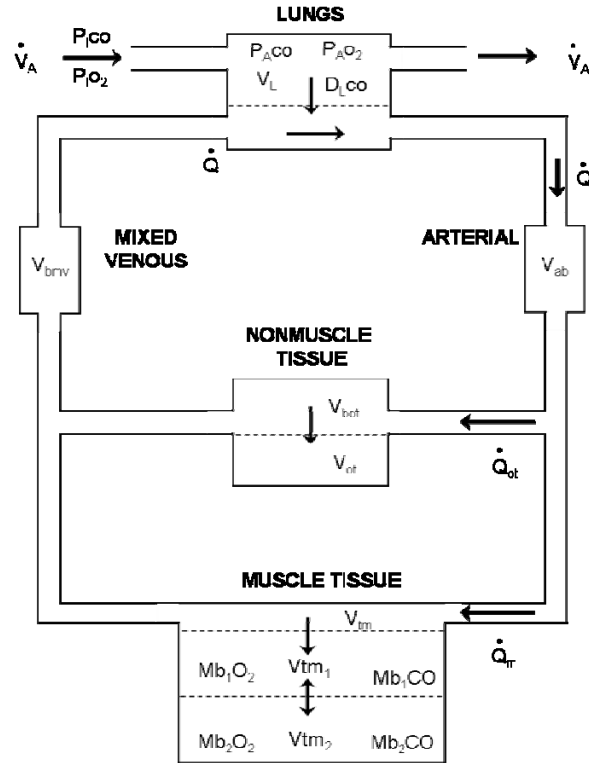
**Figure 4-1. Plot of fractional sensitivities of selected variables versus time of exposure.**

1 The mechanistic CFK model contains a number of assumptions under which the model is solely  
 2 applicable, including 1) ventilation is a continuous process, 2) equilibrium between plasma CO  
 3 concentration and COHb concentration is obtained in the pulmonary system, 3) percent COHb can exceed  
 4 100% saturation in the linear model, and 4) it does not account for the shape of the O<sub>2</sub> or CO saturation  
 5 versus pO<sub>2</sub> or pCO relation (McCartney, 1990). Estimations outside of these assumptions have been  
 6 attempted but with less predictive agreement. For example, transient exposures such as those that would  
 7 simulate everyday conditions would violate the assumption of a single, well-mixed vascular compartment.  
 8 COHb levels during exposure of subjects exposed to frequent but brief high CO exposures (667 to  
 9 7,500 ppm for 75 s to 5 min) were not accurately predicted by CFK modeling (Benignus et al., 1994;  
 10 Tikuisis et al., 1987a; Tikuisis et al., 1987b). Consistently, the predicted COHb value overpredicted  
 11 venous COHb (0.8 to 6%) and underpredicted arterial COHb (1.5 to 6.1%) and this disparity increased  
 12 after exercise. Individual differences between arterial and venous COHb varied from 2.3 to 12.1%  
 13 (Benignus et al., 1994). These inaccuracies between measured and predicted COHb values disappeared  
 14 after simulated mixing of arterial and venous blood and thus are likely due to delays in mixing of arterial

1 and venous blood and differences in cardiac output and lung wash-in. A modified CFK was created to  
2 adjust for these issues and produce a more accurate COHb prediction (Smith et al., 1994).

## 4.2.2. Multicompartment Model

3 In addition to the limitations discussed above, the CFK model does not account for extravascular  
4 storage sites for CO, such as muscle Mb. CO will undergo reversible muscle Mb binding, similar to Hb,  
5 as well as uptake into other extravascular tissues (Vreman et al., 2006). A five compartment model has  
6 been proposed to predict CO uptake and distribution from acute inhalation exposure and contains  
7 components for lung, arterial blood, venous blood, muscle tissue, and nonmuscle tissue (Bruce and Bruce,  
8 2003; Bruce et al., 2008; 2006). This model structure is illustrated in Figure 4-2. This model includes the  
9 dynamics of CO storage in the lung and its dependence on ventilation and CO pressure of mixed venous  
10 blood, relaxes the assumption Hb is saturated by including the role of CO in altering the O<sub>2</sub> dissociation  
11 curve, includes a subcompartmentalized muscle tissue compartment, accounts for dissolved CO in blood  
12 and tissue, and predicts COHb based on age and body dimensions. This multicompartment model is  
13 limited by its exclusion of cellular metabolism or Mb diffusion, simplification of within tissue bed spatial  
14 variability, and assumes ventilation and P<sub>A</sub>O<sub>2</sub> are constant. This model better predicts COHb levels when  
15 inspired CO levels change rapidly or when incomplete blood mixing has occurred, and better predicts the  
16 CO washout time course compared to the CFK equation. Bruce and Bruce (2003) compared the two  
17 models and found similar results for long term exposure settings (1,000 min), however, the  
18 multicompartment model predicted somewhat lower COHb levels compared to the CFK model over  
19 transient CO uptake conditions when using data taken from Peterson and Stewart (1970).



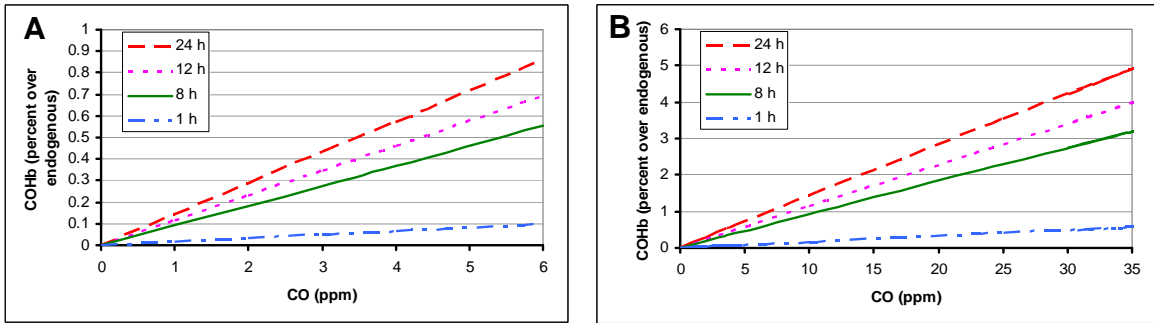
Source: Modified from Bruce et al. (2008)

**Figure 4-2. Overall structure of the multicompartment model of storage and transport of CO. Includes compartments for lung, arterial blood, venous blood, muscle tissue, and nonmuscle tissue. The muscle compartment is divided into two subcompartments for diffusion of gases within the tissue.**

### 4.2.3. Mathematical Model Usage

1 Since measurements of COHb in the population are not readily available, mathematical models are  
 2 used to predict the resulting COHb levels from various CO exposure scenarios. Figure 4-3 illustrates the  
 3 predictions of COHb after 1, 8, 12, or 24 h of CO exposure at a range of concentrations in a healthy,  
 4 inactive, adult human. The Quantitative Circulatory Physiology (QCP) model, which integrates human  
 5 physiology using over 4,000 variables and equations based on published biological interactions, was used  
 6 to predict these values (Abram et al., 2007; Benignus et al., 1987; 2006). This dynamic model uses the  
 7 nonlinear CFK equation with modifications presented in Smith et al. (1994). The data in Figure 4-3 are  
 8 presented as the change in percent COHb from endogenous due to CO exposure to NAAQS relevant  
 9 levels from 0-35 ppm (A) and relevant ambient levels from 0-6 ppm (B). Endogenous CO production  
 10 varies as described in Section 4.5 but generally results in less than 1% COHb, with a QCP modeled value  
 11 of 0.39% at time zero. Figure 4-3 illustrates that 35 ppm CO for 1 h results in an increase of 0.56% COHb  
 12 over endogenous levels and 9 ppm CO for 8 h results in a 0.83% increase over endogenous levels. Also,

1 these graphs show that long term, low concentration CO exposure results in equivalent COHb levels to  
 2 higher concentration, acute exposure. For example, COHb resulting from 35 ppm for 1 h (0.56%) is  
 3 approximately equivalent to 6 ppm for 8 h (0.55%) or 4 ppm for 24 h (0.57%).



**Figure 4-3.** Predicted COHb increments over endogenous levels resulting from 1, 8, 12, and 24 h CO exposures in a modeled healthy human at rest. The QCP model used a dynamic nonlinear CFK with affinity constant  $M = 230$ . The data are presented as the change in percent COHb from endogenous levels due to CO exposure to 0-6 ppm (A) and 0-35 ppm (B). Note: Graph A is a subset of Graph B.

## 4.3. Absorption, Distribution, and Elimination

### 4.3.1. Pulmonary Absorption

4 Pulmonary uptake of CO accounts for all environmental CO absorption and occurs at the  
 5 respiratory bronchioles and alveolar ducts and sacs. CO and O<sub>2</sub> share various physico-chemical  
 6 properties, thus allowing for the extension of the knowledge about O<sub>2</sub> kinetics to those of CO despite the  
 7 differences in the reactivity of the gases. The exchange of CO between the air and the body depends on a  
 8 number of physical (e.g., mass transfer and diffusion), as well as physiological factors (e.g., alveolar  
 9 ventilation and cardiac output), which are controlled by environmental conditions, physical exertion, and  
 10 other processes discussed in Section 4.4. The ability of the lung to take up inhaled CO is measured by  
 11  $D_LCO$ , and CO uptake ( $V_{CO}$ ) representing the product of  $D_LCO$  and the mean alveolar pressure ( $P_ACO$ ).  
 12 The importance of dead space volume, gas mixing and homogeneity, and ventilation/perfusion matching  
 13 were discussed in depth in the 2000 AQCD (U.S. EPA, 2000).

### 4.3.1.1. Mass Transfer of Carbon Monoxide

1 Mass transfer refers to the molecular and convective transport of CO molecules within the body  
 2 stores, driven by random molecule motion from high to low concentrations. CO enters through the airway  
 3 opening (mouth and nose) and transfers in a gas phase to the alveoli. CO transport is due to convective  
 4 flow, the mechanical action of the respiratory system, and diffusion in the acinar zone of the lung (Engel,  
 5 1983). Then, CO diffuses in a “liquid” phase across the air-blood interface, binding red blood cell (RBC)  
 6 Hb. At environmental CO levels, CO uptake into RBC is limited by the reaction rate of binding of CO to  
 7 O<sub>2</sub>Hb forming COHb. Pulmonary capillary RBC CO diffusion is rapidly achieved (Chakraborty et al.,  
 8 2004; Gibson and Roughton, 1955; Reeves and Park, 1992; Roughton and Forster, 1957). The rate and  
 9 level of COHb depends upon the pCO, pO<sub>2</sub> in the air, time of exposure, and the ventilation rate (Roughton  
 10 and Forster, 1957). Most of the body CO is bound to Hb; however, 10-15% of the total body CO is  
 11 located in extravascular tissues primarily bound to other heme proteins (Coburn, 1970a). Considerable  
 12 concentrations of CO have been measured in spleen, lung, kidney, liver, muscle, and heart (Vreman et al.,  
 13 2005; Vreman et al., 2006), whereas less CO is localized to fatty tissues, such as adipose and brain (Table  
 14 4-1). The transfer of CO occurs by a partitioning of CO between Hb and tissue. Less than 1% of the total  
 15 body CO stores appear as dissolved in body fluids, due to the insolubility and small tissue partial pressure  
 16 of CO (Coburn, 1970a). Transport pathways and body stores of CO are shown in Figure 4-4.

**Table 4-1a. CO concentration in pmol/100 g ww tissue – human.**

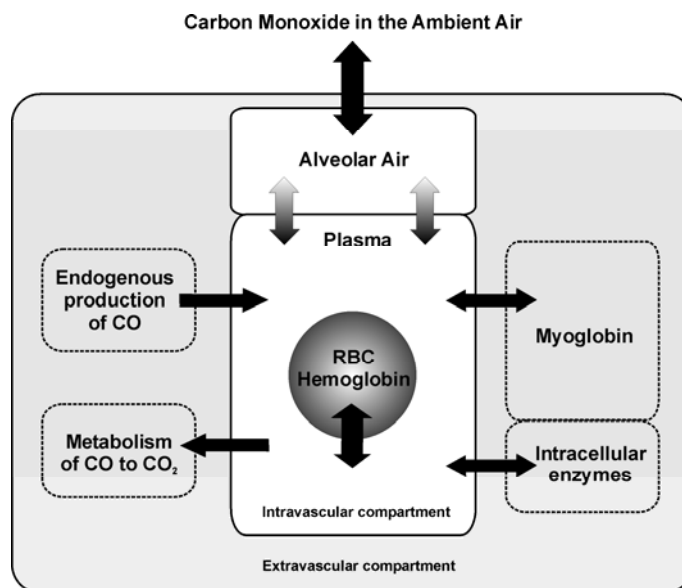
Exposure	Adipose	Brain	Muscle	Heart	Kidney	Lung	Spleen	Blood	% COHb
Background	2 ± 1	3 ± 3	15 ± 9	31 ± 23	23 ± 18	57 ± 59	79 ± 75	165 ± 143	1.5 ± 1.2
Fire	5 ± 4	7 ± 5	24 ± 16	54 ± 33	27 ± 11	131 ± 127	95 ± 69	286 ± 127	3.8 ± 3.2
Fire + CO	18 ± 29	17 ± 14	168 ± 172	128 ± 63	721 ± 427	1097 ± 697	2290 ± 1409	3623 ± 1975	40.7 ± 28.8
CO asphyxiation	25 ± 27	72 ± 38	265 ± 157	527 ± 249	885 ± 271	2694 ± 1730	3455 ± 1347	5196 ± 2625	56.4 ± 28.9

Source: Vreman et al. (2006)

**Table 4-1b. CO concentration in pmol/mg fresh weight – adult mouse.**

Exposure	Brain	Muscle	Heart	Kidney	Lung	Spleen	Liver	Intestine	Testes	Blood	% COHb
Background	2 ± 0	10 ± 1	6 ± 1	7 ± 2	3 ± 1	6 ± 1	5 ± 1	4 ± 2	2 ± 1	45 ± 5	0.5
500 ppm CO	18 ± 4	14 ± 1	100 ± 18	120 ± 12	250 ± 2	229 ± 55	115 ± 31	9 ± 7	6 ± 3	2648 ± 400	28
30 µM heme	2 ± 0	7 ± 1	14 ± 3	7 ± 2	8 ± 3	11 ± 1	8 ± 3	3 ± 1	2 ± 0	88 ± 10	0.9

Source: Vreman et al. (2005)



Source: Adapted from Coburn (1967)

**Figure 4-4. Diagrammatic presentation of CO uptake and elimination pathways and CO body stores.**

#### 4.3.1.2. Lung Diffusion of Carbon Monoxide

1 Lung diffusion of CO is an entirely passive process of gas diffusion across the alveolo-capillary  
 2 membrane, through the plasma, across the RBC membrane and into the RBC stroma, where CO binding  
 3 to Hb rapidly occurs. Membrane and blood phase transfer are governed by physico-chemical laws,  
 4 including Fick's first law of diffusion. The diffusing capacity of the lung for CO, represented as  $D_LCO$ , is  
 5 a measurement of the partial pressure difference between inspired and expired CO. Due to the rapid  
 6 binding of CO to Hb, a high pressure differential between air and blood exists when CO air levels are  
 7 increased. Inhalation of CO-free air reverses the pressure differential (higher CO pressure on the blood  
 8 side than the alveolar side), and then CO is released into the alveolar air. Since CO is also produced  
 9 endogenously, CO release will also be affected by this production pressure. However, the air-blood



1 gradient for CO is usually higher than the blood-air gradient; therefore, CO uptake will be a  
2 proportionately faster process than CO elimination.

3 A number of factors have been found to affect  $D_LCO$  including Hb concentration, cardiac output  
4 ( $\dot{Q}$ ), erythrocyte flow, COHb concentration,  $P_ACO_2$ , body position, exercise, time of day, age, etc.  
5 (Forster, 1966; Hsia, 2002).  $D_LCO$  consistently decreases after intense bouts of exercise, likely due to the  
6 redistribution of blood volume to the periphery (Hanel et al., 1997; Manier et al., 1991). However, in  
7 going from rest to exercise  $D_LCO$  can increase linearly from: lung expansion leading to unfolding and  
8 distension of alveolar septa, opening and/or distension of capillaries as  $\dot{Q}$  increases, increased capillary  
9 hematocrit, and more homogeneous distribution of capillary erythrocytes (Hsia, 2002).  $D_LCO$  is less  
10 dependent upon lung volume at mid-range vital capacity, but at extreme volumes the diffusion rate is  
11 varied, higher than average at total lung capacity and lower at residual volume (McClellan et al., 1981).

## 4.3.2. Tissue Uptake

### 4.3.2.1. The Respiratory Tract

12 The upper respiratory tract contributes little to the overall COHb uptake. The lung has nearly  
13 constant exposure to CO; however, relatively little CO diffuses into the tissue except for at the alveolar  
14 region en route to the circulation. No detectable uptake of CO was observed in the human nasal cavity or  
15 upper airway (Guyatt et al., 1981) or in the monkey oronasal cavity after high CO exposure (Schoenfisch  
16 et al., 1980).

### 4.3.2.2. The Blood

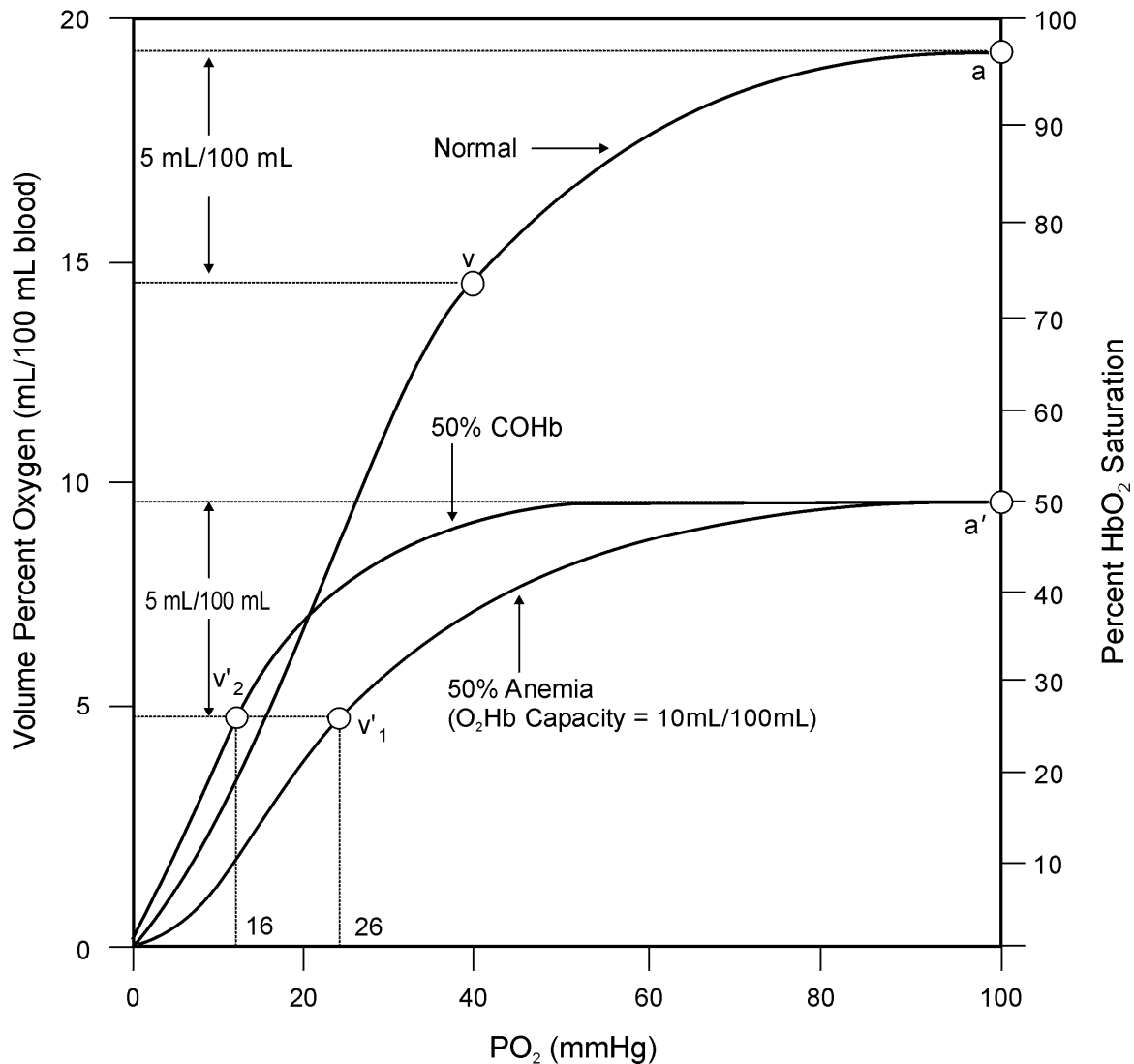
17 The blood is the largest reservoir for CO, where it reversibly binds to Hb. The chemical affinity of  
18 CO for adult human Hb is approximately 218 times greater than that of  $O_2$ , meaning one part CO and 218  
19 (210-250) parts  $O_2$  would form equal parts of  $O_2Hb$  and COHb (Engel et al., 1969; Rodkey et al., 1969;  
20 Roughton, 1970). This would happen when breathing air containing 21%  $O_2$  and 960 ppm CO. This  
21 concept was presented by Haldane and Smith (1898) and later represented as the Haldane constant M  
22 (210-250) in the Haldane equation by Douglas, Haldane, and Haldane (1912). M is relatively unaffected  
23 by changes in physiological pH,  $CO_2$ , temperature, or 2,3-diphosphoglycerate:  
24

$$COHb \div O_2Hb = M \times (pCO \div pO_2)$$

Equation 4-2

1           The Hb association rate for CO is 10% slower than O<sub>2</sub> and occurs in a cooperative manner  
2 (Chakraborty et al., 2004; Sharma et al., 1976). Hb is composed of four globin chains each containing a  
3 heme group capable of binding CO or O<sub>2</sub>. The associative reaction rates become faster with successive  
4 heme binding, attributed to interactions within the protein and to strains imposed on the heme and its  
5 ligands (Alcantara et al., 2007). More simply, the greater the number of heme sites bound to CO, the  
6 greater the affinity of free heme sites for O<sub>2</sub>, thus causing Hb to bind and retain O<sub>2</sub> that would normally be  
7 released to tissues. Cooperativity is greatly reduced in CO dissociation, but the rate of dissociation of CO  
8 from Hb is orders of magnitude slower than O<sub>2</sub> ( $k_{CO} = 4 \times 10^{-4} k_{O_2}$ ), which accounts for the high affinity  
9 values (Chakraborty et al., 2004). The half-time of dissociation reaction is about 11 seconds at 37 °C  
10 (Holland, 1970). In general, CO uptake to COHb equilibrium is slower in humans and large animals,  
11 requiring 8-24 h, than in smaller species such as rats, which will equilibrate in 1-2 h (Penney et al., 1988).

12           CO binding to Hb also has effects on the O<sub>2</sub> dissociation curve of the remaining Hb by shifting the  
13 curve progressively to the left and altering the normal S-shaped curve to become more hyperbolic due to  
14 increased cooperative O<sub>2</sub> binding (Roughton, 1970). This is referred to as the “Haldane effect” and causes  
15 tissues to have more trouble obtaining O<sub>2</sub> from the blood, even compared to the same extent of reduced  
16 Hb resulting from anemia. For example, Figure 4-5 (as explained in the 2000 AQCD) illustrates that an  
17 acute anemia patient (50% of Hb) at a venous pO<sub>2</sub> of 26 mmHg ( $v'_1$ ), 5 vol % of O<sub>2</sub> (50% saturation) was  
18 extracted from the blood. In contrast, a CO poisoned person with 50% COHb, the venous pO<sub>2</sub> will have to  
19 drop to 16 mmHg ( $v'_2$ ) to release the same 5 vol % O<sub>2</sub>. This more severe effect on O<sub>2</sub> pressure may lead  
20 to brain O<sub>2</sub> depletion and loss of consciousness if any higher demand of O<sub>2</sub> is needed (e.g., exercise).



Source: U.S. EPA (1991)

**Figure 4-5. O<sub>2</sub>Hb dissociation curve of normal human blood, of blood containing 50% COHb, and of blood with only 50% Hb because of anemia.**

### 4.3.2.3. Heart and Skeletal Muscle

1 Mb is a globular heme protein that facilitates O<sub>2</sub> diffusion from the muscle sarcoplasm to  
 2 mitochondria, acting as an O<sub>2</sub> supply buffer to maintain adequate pO<sub>2</sub> for mitochondria when the O<sub>2</sub>  
 3 supply changes, as in exercise. O<sub>2</sub> has a greater affinity for Mb than Hb, which allows small changes in  
 4 tissue pO<sub>2</sub> to release large amounts of O<sub>2</sub> from O<sub>2</sub>Mb (Wittenberg et al., 1975). Small reductions in O<sub>2</sub>  
 5 storage capacity of Mb, due to CO binding, may have a profound effect on the supply of O<sub>2</sub> to the tissue.

6 Like Hb, Mb will undergo reversible CO binding, however the affinity constant is approximately  
 7 eight-times lower than Hb (M = 20-40 versus 218, respectively) (Haab, 1990). The association rate

1 constant of CO and Mb is approximately 27 times lower than O<sub>2</sub>, however the dissociation rate constant is  
2 approximately 630 times lower than O<sub>2</sub> (Gibson et al., 1986) causing CO to be retained and possibly  
3 stored in the muscle. CO levels have been measured in human muscle and heart tissues with less than 2%  
4 COHb concentrations at background levels averaging 15 and 31 picomole (pmol) CO/100 grams ww,  
5 respectively (Vreman et al., 2006). Under conditions of CO asphyxiation, tissue concentrations increased  
6 to 265 and 527 pmol CO/100 grams ww muscle and heart tissue, respectively; however, heart tissue  
7 concentrations varied widely between individuals (See Table 4-1a). The capacity for diffusion of CO into  
8 the muscle is represented by the coefficient D<sub>m</sub>CO and is generally larger in males than in females, likely  
9 due to the differences in muscle mass and capillary density (Bruce and Bruce, 2003). COMb  
10 concentrations in the heart and skeletal muscle increase with work load, causing an increase in  
11 COMb/COHb that is not as greatly seen at rest (Sokal et al., 1984). Subjects with 2% COHb, but not  
12 those with 20% COHb levels, showed a significant uptake of CO from the blood to the muscle with  
13 increasing work intensity of the quadriceps muscle (Richardson et al., 2002).

#### 4.3.2.4. Other Tissues

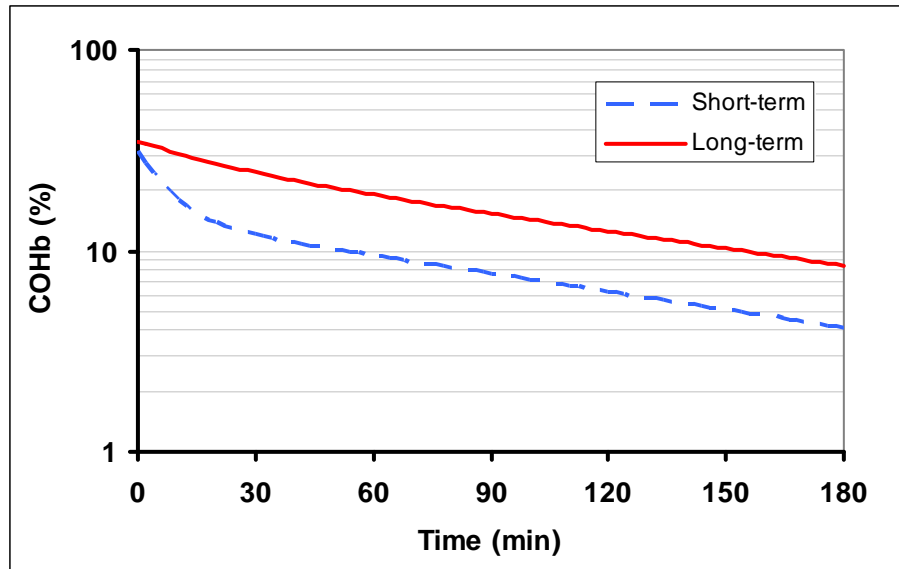
14 CO binds with other hemoproteins, such as cytochrome P450, cytochrome *c* oxidase, catalase, and  
15 peroxidase, but the possibility of this binding influencing CO-O<sub>2</sub> kinetics has not been established. CO  
16 transfers between COHb and tissue, the extent of which varies between organs. Blood to tissue flux  
17 causes less CO to be expired following CO exposure than what is lost from the blood in terms of COHb  
18 (Roughton and Root, 1945). This value is estimated to be 0.3-0.4% min<sup>-1</sup> or 0.24 mL/min (Bruce and  
19 Bruce, 2003; Prommer and Schmidt, 2007). The equilibration rate from blood to tissue is controversial.  
20 Newly modeled CO trafficking kinetics shows that CO continues to be taken up by the muscle and  
21 extravascular tissues well beyond the end of exposure because of a less than instant equilibration (Bruce  
22 and Bruce, 2006). Tables 4-1a and 4-1b contain tissue CO concentrations from human and mouse under  
23 different CO exposure conditions. The distribution of CO between the different organs was shown to  
24 follow the same pattern versus percent of the blood CO concentration, irrespective of the level of blood  
25 CO (Vreman et al., 2006). These results are in conflict with older papers suggesting that negligible  
26 retention of CO occurs in the liver or brain (Sokal et al., 1984; Topping, 1975).

#### 4.3.3. Pulmonary and Tissue Elimination

27 Blood COHb concentrations are generally considered to have a monotonically decreasing, second-  
28 order (logarithmic or exponential) elimination rate from equilibrium. However, more recent reports have  
29 presented evidence for a biphasic washout curve, especially after short-term CO exposure (Figure 4-6)  
30 (Bruce and Bruce, 2006; Shimazu et al., 2000; Wagner et al., 1975). This event is modeled by a two-

1 compartment system where the initial rapid decrease is the washout rate from the blood, followed by a  
2 slower phase due to CO flux from the muscle and extravascular compartments back to the blood. Tissue  
3 elimination rates have been reported as slower than those for blood (Landaw, 1973). The biphasic curve is  
4 more obvious after short-term CO exposure (less than 1 h), whereas long-term CO exposure (5 h or more)  
5 results in a virtually monoexponential elimination, which could account for the historical findings.  
6 However, this elimination curve also follows a biphasic curve with a slightly higher rate of elimination  
7 initially (Shimazu et al., 2000). Differences in elimination kinetics could also be a result of the variation  
8 in CO exposure duration (Weaver et al., 2000).

9       The elimination of COHb is affected by a number of factors, including duration of exposure,  $P_aO_2$ ,  
10 minute ventilation, the time post-exposure for analysis due to extravascular stores, as well as inter-  
11 individual variability (Bruce and Bruce, 2006; Landaw, 1973; Shimazu, 2001). The elimination rate does  
12 not seem to be dependent upon the CO exposure source (e.g., fire, non-fire CO exposure) (Levasseur et  
13 al., 1996). In addition, in a series of poisoned patients, the COHb elimination half-life was not influenced  
14 by gender, age, smoke inhalation, history of loss of consciousness, concurrent tobacco smoking, degree of  
15 initial metabolic acidosis (base excess), or the initial COHb level (Weaver et al., 2000). COHb  
16 elimination half-life falls as the fractional inspired  $O_2$  concentration increases. While breathing air at sea  
17 level pressure, the expected half-life in adult males is approximately 285 minutes, but may be shorter in  
18 adult females. With inhalation of normobaric 40%  $O_2$ , the half-life falls to 75 minutes and further to 21  
19 minutes when breathing 100%  $O_2$  because of greater competition for Hb by  $O_2$  (Landaw, 1973). Another  
20 study reports the half-life falls to 74 minutes (mean) after breathing 100%  $O_2$ , although the range in this  
21 particular study was 26 to 148 minutes (Weaver et al., 2000). In addition, COHb half-life will fall further  
22 after normocapnic hyperoxic hyperpnea (i.e., hyperventilation while maintaining normal  $CO_2$  pressure in  
23 high  $O_2$ ) (Takeuchi et al., 2000).



Adapted from Shimazu et al. (2000)

Figure 4-6. Changes in blood COHb after short-term and long-term exposure to CO, representing the biphasic nature of CO elimination. Note: y-axis is log-scale.

## 4.4. Conditions Affecting Uptake and Elimination

### 4.4.1. Environment and Activity

1 Elevated CO exposure and COHb levels are dependent upon the changes in CO concentration in  
 2 the local environment. Pedestrians are exposed to high levels of CO for short time periods from vehicle  
 3 exhaust at busy intersections (see also Chapter 3.6). Higher exposure can also result from riding in an  
 4 automobile or stopping at busy intersections (Ott, 1994). Indoor exposure occurs from ETS and unvented  
 5 combustion appliances, such as natural gas cooking stoves, attached garages, and gas fireplaces, which  
 6 can accumulate to over 100 ppm (Dutton et al., 2001). Recreational exposure at levels exceeding 200 ppm  
 7 and peaks of 1,600 ppm could occur in indoor ice rinks using fossil fuel powered ice resurfacers and  
 8 coliseums housing malfunctioning equipment or poor ventilation (Levesque et al., 2000; Pelham et al.,  
 9 2002). Certain occupations provide instances and conditions for transient moderate-to-high CO levels,  
 10 including fire fighters and machinery operators. Such transient exposures have the ability to increase  
 11 COHb levels. For example, exposure for 5 minutes or less of a resting individual to 6,600 ppm CO will  
 12 result in up to 20% COHb (Benignus et al., 1994).

13 Exercise is an important determinant of CO kinetics and toxicity due to the extensive increase in  
 14 gas exchange. O<sub>2</sub> consumption can increase more than 10 fold during exercise. Similarly, ventilation,

1 membrane and lung diffusing capacity, pulmonary capillary blood volume, and cardiac output increase  
2 proportional to work load. The majority of these changes facilitate CO uptake and transport, by increasing  
3 gas exchange efficiency.

4 The COHb elimination rate decreases with physical activity (Joumard et al., 1981). Healthy  
5 subjects exposed to CO and achieving COHb levels of approximately 4-5% observed a significant  
6 detriment to exercise duration and maximal effort capability (measured by metabolic equivalent units)  
7 (Adir et al., 1999). It is possible that CO lowers the anaerobic threshold, allowing earlier fatigue of the  
8 skeletal muscles and decreased maximal effort capability.

## 4.4.2. Altitude

9 Increased altitude changes a number of factors that contribute to the uptake and elimination of CO.  
10 The relationship between altitude and CO exposure has been discussed in depth in the 2000 AQCD and  
11 other documents (U.S. EPA, 1978). In an effort to maintain proper O<sub>2</sub> transport and supply, physiological  
12 changes occur as compensatory mechanisms to combat the decreased barometric pressure and resulting  
13 altitude induced hyperbaric hypoxia (HH). HH, unlike CO hypoxia, causes humans to hyperventilate,  
14 which reduces arterial blood CO<sub>2</sub> (hypocapnia) and increases alveolar partial pressure of O<sub>2</sub>. Hypocapnia  
15 will lead to difficulty of O<sub>2</sub> dissociation and decreased blood flow, thus reducing tissue O<sub>2</sub> supply. HH  
16 increases blood pressure (BP) and cardiac output and leads to redistribution of blood from skin to organs  
17 and from blood vessels to extravascular compartments. Generally these changes will favor increased CO  
18 uptake and COHb formation, as well as CO elimination. In hypoxic conditions both CO and O<sub>2</sub> bind  
19 reduced Hb through a competitive-parallel reaction (Chakraborty et al., 2004). Breathing CO (9 ppm) at  
20 rest at altitude produced higher COHb compared to sea level (McGrath et al., 1993), whereas high altitude  
21 exposure with exercise caused a decrease in COHb levels versus similar exposure at sea level (Horvath et  
22 al., 1988). This decrease could be a shift in CO storage or suppression of COHb formation, or both.  
23 Altitude also increases the baseline COHb levels by inducing endogenous CO production. Initial HH  
24 increased lung HO-1 protein and activity, whereas chronic HH induced endogenous CO production in  
25 nonpulmonary sites (see Section 4.5) (Carraway et al., 2000).

26 As the length of stay increases at high altitude, acclimatization occurs, inducing hyperventilation,  
27 polycythemia or increased red blood cell count, and increased tissue capillarity and Mb content in skeletal  
28 muscle, which could also favor increased CO uptake. Most of the early adaptive changes gradually revert  
29 to sea level values. However, differences in people raised at high altitude persist even after  
30 reacclimatization to sea level (Hsia, 2002).

### 4.4.3. Physical Characteristics

1 Certain physical characteristics (e.g., age, sex, pregnancy) can alter the variables that influence the  
2 uptake, distribution, and elimination of CO. Values of CO uptake and elimination change with age. Young  
3 children eliminate COHb more rapidly than adults after CO exposure (Joumard et al., 1981; Klasner et al.,  
4 1998). After infancy, the COHb half-life increases with age, nearly doubling between 2 and 70 years  
5 (Joumard et al., 1981). The rate of this increase in CO elimination is very rapid in the growing years (2 to  
6 16 years of age), but slows beyond adolescence. Alveolar volume and  $D_LCO$  increase with increasing  
7 body length of infants and toddlers (Castillo et al., 2006), suggesting a further degree of lung  
8 development and faster CO uptake. After infancy, increasing age decreases  $D_LCO$  and increases  $V_A/Q$   
9 mismatch, causing it to take longer to both load and eliminate CO from the blood (Neas and Schwartz,  
10 1996).

11 COHb concentrations are generally lower in female subjects than in male subjects (Horvath et al.,  
12 1988) and the COHb half-life may be longer in healthy men than in women of the same age, which may  
13 be partially explained by differences in muscle mass or the slight correlation between COHb half-life and  
14 increased height (Joumard et al., 1981). The rate of decline of  $D_LCO$  with age is lower in middle-aged  
15 women than in men; however, it evens out towards older age (Neas and Schwartz, 1996). Women also  
16 tended to be more resistant to altitude hypoxia (Horvath et al., 1988).

17 Fetal CO pharmacokinetics do not follow the same kinetics as maternal CO exposure, making it  
18 difficult to estimate fetal COHb based on maternal levels. Human fetal Hb has a higher affinity for CO  
19 than adult Hb (Di Cera et al., 1989). Maternal and fetal COHb concentrations have been modeled as a  
20 function of time using a modified CFK equation (Hill et al., 1977). At steady-state conditions, the fetal  
21 COHb is up to 10% higher than the maternal COHb levels, for example, exposure to 30 ppm CO results  
22 in a maternal COHb of 5% and a fetal COHb of 5.5%. The fetal CO uptake lags behind the maternal for  
23 the first few hours but later may overtake the maternal values. Similarly, during washout, the fetal COHb  
24 levels are maintained for longer, with a half-life of around 7.5 hours versus the maternal half-life of  
25 around 4 hours.

26 Ethnicity does alter physiological variables that determine CO uptake and kinetics. Lung volumes  
27 are 10-15% less in both Asian and African-American populations when compared to Caucasians. This  
28 causes a reduced alveolar surface area (20% less than estimated values) for gas exchange, leading to a  
29 13% difference in diffusion capacity,  $D_LCO$  (Pesola et al., 2004; Pesola et al., 2006). Certain factors such  
30 as socioeconomic status (SES) were not controlled for in these studies. SES has been shown to affect  
31 pulmonary function, including decreasing  $D_LCO$  (Hegewald and Crapo, 2007).



#### 4.4.4. Health Status

1 Health status can influence the toxicity involved with CO exposure by influencing the severity of  
2 hypoxia resulting from CO exposure. Any condition that would alter the blood O<sub>2</sub> carrying capacity or  
3 content will result in a greater risk from COHb induced hypoxia and decreased tissue O<sub>2</sub> delivery. The  
4 severity of this effect depends upon the initial level of hypoxia.

5 Anemias are a group of diseases that result in insufficient blood O<sub>2</sub> or hypoxia due to Hb  
6 deficiency through hemolysis, hemorrhage, or reduced hematopoiesis. Anemia may result from pathologic  
7 conditions characterized by chronic inflammation such as malignant tumors or chronic infections  
8 (Cavallin-Stahl et al., 1976a, b). The bodies of people with anemia compensate causing cardiac output to  
9 increase as both heart rate and stroke volume increase. The endogenous production of CO, thus COHb, is  
10 increased in patients with hemolytic anemia due to increased heme catabolism, causing an increased  
11 baseline COHb concentration. One of the most prevalent anemias arises from a single-point mutation of  
12 Hb, causing sickle cell diseases. The Hb affinity for O<sub>2</sub> and O<sub>2</sub> carrying capacity is reduced causing a shift  
13 to the right in the O<sub>2</sub> dissociation curve. It is well documented that African-American populations have a  
14 higher incidence of sickle cell anemia, which may be a risk factor for CO hypoxia.

15 Chronic obstructive pulmonary disease (COPD) is often accompanied by a number of changes in  
16 gas exchange, including increased deadspace volume (V<sub>D</sub>) and ventilation-perfusion ratio (V/Q)  
17 inequality (Marthan et al., 1985), which could slow both CO uptake and elimination. Patients with  
18 pulmonary sarcoidosis may have a decrease in lung volumes, a loss of D<sub>L</sub>CO, and gas exchange  
19 abnormalities during exercise, including decreased arterial oxygen pressure (P<sub>a</sub>O<sub>2</sub>) and increased alveolar-  
20 arterial oxygen pressure difference (Lamberto et al., 2004).

21 Individuals with heart disease may be at a greater risk from CO exposure since they may already  
22 have compromised O<sub>2</sub> delivery. Time to onset of angina was reduced after exposure to 100 ppm carbon  
23 monoxide, compared to clean air (Kleinman et al., 1998). Hyperlipidemic patients may have decreased  
24 CO diffusion capacity, a loss of V/Q gradient, and a decrease in P<sub>a</sub>O<sub>2</sub> (Enzi et al., 1976).

### 4.5. Endogenous CO Production and Metabolism

25 Humans breathing air containing no environmental sources of CO will still have a low measurable  
26 level of circulating COHb. This is due to endogenous CO production from heme protein catabolism  
27 among other sources. In the natural degradation of RBC Hb, the porphyrin ring of heme is broken at the  
28 methene bridge and catabolized in an O<sub>2</sub> dependent manner by HO complexed with NADPH-cytochrome  
29 P450 reductase and biliverdin reductase to biliverdin, Fe, and CO. Biliverdin is then further broken down  
30 into bilirubin, a powerful endogenous antioxidant. Two main HO isoforms exist, HO-1 and HO-2.

1 Expression of HO-1 is inducible, whereas HO-2 is constitutively expressed. The major site of heme  
2 catabolism, thus the major organ of CO production is the liver, followed by the spleen, brain, and  
3 erythropoietic system (Berk et al., 1976). These rates of CO formation may be due to higher levels of HO  
4 activity in these tissues. The whole body production rate of CO is approximately 16.4  $\mu\text{mol/h}$  (0.42 mL/h)  
5 and produces between 400-500  $\mu\text{mol CO}$  per day (Coburn et al., 1964; Coburn et al., 1966; Coburn,  
6 1970b). The endogenous rate of CO formation has been shown to vary little between different tissues,  
7 ranging from 0.029 nmol/mg protein/h in chorionic villi of term human placenta to 0.28 nmol/mg  
8 protein/h in rat olfactory receptor neurons in culture and in rat liver perfusate (Marks et al., 2002),  
9 however these estimations are questionable since CO is quickly scavenged in the cytosol of living cells.  
10 CO is endogenously produced in the nose and paranasal sinus which may contribute to exhaled CO  
11 concentrations (Andersson et al., 2000).

12 HO mediated metabolism functions as the rate-limiting enzyme step in heme degradation and  
13 endogenous CO production (Wu and Wang, 2005). Three isoforms of HO exist, but HO-1 is the only  
14 inducible form (Maines and Kappas, 1974; Maines et al., 1986; McCoubrey et al., 1997). Endogenous CO  
15 production can be increased by the up-regulation of HO-1 expression and activity by inducers such as  
16 oxidative stress, hypoxia, heavy metals, sodium arsenite, heme and heme derivatives, various cytokines,  
17 and also exogenous CO (Wu and Wang, 2005). High levels of CO (2,500 ppm) have been shown to  
18 increase HO-1 activity in the brain of rats, as well as liberate intracellular heme to further stimulate  
19 endogenous CO production (Cronje et al., 2004).

20 Not all endogenous CO production is derived from Hb breakdown. Other hemoproteins, such as  
21 Mb, cytochromes, peroxidases, and catalase, contribute 20-25% to the total amount of endogenous CO  
22 (Berk et al., 1976). All of these sources result in a normal blood COHb concentration between 0.4 and  
23 0.7% (Coburn et al., 1965). This baseline level of endogenous production can be altered by drugs or a  
24 number of physiological conditions that alter RBC destruction, other hemoprotein breakdown, or bilirubin  
25 production. Because of this sensitivity, ranges of endogenous COHb levels in the population are  
26 uncertain. Nicotinic acid (Lundh et al., 1975), allyl-containing compounds (acetamids and barbiturates)  
27 (Mercke et al., 1975b), diphenylhydantoin (Coburn, 1970b), progesterone (Delivoria-Papadopoulos et al.,  
28 1974), and contraceptives (Mercke et al., 1975a) will increase CO production. Compounds such as carbon  
29 disulfide and sulfur-containing chemicals (parathion and phenyltiourea) will increase CO by acting on  
30 P450 system moieties (Landaw et al., 1970). The P450 system may also cause large increases in CO  
31 produced from the metabolic degradation of dihalomethanes leading to very high (>10%) COHb levels  
32 (Manno et al., 1992), which can be further enhanced by prior exposure to HCs or ethanol (Pankow et al.,  
33 1991; Wirkner and Poelchen, 1996). HO can catalyze the release of CO from the auto-oxidation of  
34 phenols, photo-oxidation of organic compounds, and lipid peroxidation of cell membrane lipids (Rodgers  
35 et al., 1994).

1 Any disturbance in RBC hemostasis by acceleration of destruction of hemoproteins will lead to  
2 increased production of CO. Pathologic conditions such as anemias, hematomas, thalassemia, Gilbert's  
3 syndrome with hemolysis, and other hematological diseases and illness will accelerate CO production  
4 (Berk et al., 1974; Hampson, 2007; Meyer et al., 1998; Solanki et al., 1988). Patients with hemolytic  
5 anemia exhibit COHb levels 2- to 3-times higher than healthy individuals and CO production rates 2- to  
6 8-times higher (Coburn et al., 1966). Endogenous CO production rate varied from 0.70 to 3.44 mL/h in  
7 anemic patients (Coburn et al., 1965). Women experience fluctuating COHb levels during pregnancy as  
8 well as through the menstrual cycle when endogenous CO production doubles in the progesterone phase  
9 (Delivoria-Papadopoulos et al., 1974; Mercke and Lundh, 1976).

10 Altitude has been shown to be positively associated with baseline COHb concentrations (McGrath,  
11 1992; 1993). This increase in COHb with altitude induced hypoxia has also been associated with  
12 increases in the mRNA, protein, and activity of HO-1 in rats and cells leading to enhanced endogenous  
13 CO production (Carraway et al., 2002; Lee et al., 1997). Whether other variables such as an accelerated  
14 metabolism or a greater pool of Hb, transient shifts in body stores, or a change in the elimination rate of  
15 CO play a role has not been explored.

16 Endogenous CO is removed from the body mainly by expiration and oxidation. CO will diffuse  
17 across the alveolar-capillary membrane and then is exhaled. This event has been used as a noninvasive  
18 measurement of endogenous CO and CO body load (Stevenson et al., 1979). CO can also be oxidized to  
19 CO<sub>2</sub> by cytochrome *c* oxidase in the mitochondria (Fenn, 1970; Young and Caughey, 1986). However, the  
20 rates of CO metabolism are much slower than the rates of endogenous CO production, with the rate of  
21 consumption representing only 10% of the rate of CO production in dogs (Luomanmaki and Coburn,  
22 1969).

## 4.6. Summary and Conclusions

23 CO elicits various health effects by binding with and altering the function of a number of heme-  
24 containing molecules, mainly Hb. The formation of COHb reduces the O<sub>2</sub>-carrying capacity of blood and  
25 impairs the release of O<sub>2</sub> from O<sub>2</sub>Hb to the tissues. COHb formation has been modeled mainly by the  
26 CFK equation, but more recent models have included Mb and extravascular storage compartments, as  
27 well as other dynamics of CO physiology. These models have indicated that CO has a biphasic  
28 elimination curve, due to initial washout from the blood followed by a slower flux from the tissues. The  
29 flow of CO between the blood and alveolar air or tissues is controlled by diffusion down the CO  
30 concentration gradient. The uptake of CO is governed not only by this CO pressure differential, but also  
31 by physiological factors, such as minute ventilation and lung diffusing capacity, that can, in turn, be  
32 affected by conditions such as exercise, age, and health. Susceptible populations, including health

1 compromised individuals and developing fetuses, are at a greater risk from COHb induced health effects  
2 due to altered CO kinetics, compromised cardiopulmonary processes, and increased baseline hypoxia  
3 levels. Altitude may also significantly affect the kinetics of COHb formation. Compensatory mechanisms,  
4 such as increased cardiac output, combat the decrease in barometric pressure. Altitude also increases the  
5 endogenous production of CO through upregulation of HO-1. CO is considered a second messenger and  
6 is endogenously produced from the catabolism of heme proteins by enzymes such as HO-1. Finally, CO is  
7 removed from the body by expiration and oxidation to CO<sub>2</sub>.

# Chapter 5. Integrated Health Effects

## 5.1. Mode of Action of CO Toxicity

### 5.1.1. Introduction

1           The diverse effects of CO are dependent upon concentration and duration of exposure as well as on  
2 the cell types and tissues involved. Responses to CO are not necessarily due to a single process and may  
3 instead be mediated by a combination of effects including COHb-mediated hypoxic stress and  
4 mechanisms unrelated to tissue hypoxia including free radical production and the initiation of cell  
5 signaling. However, binding of CO to reduced iron in heme proteins with subsequent alteration of heme  
6 protein function is the common mechanism underlying the biological responses to CO.

### 5.1.2. Hypoxic Mechanisms

7           As discussed in the 2000 CO AQCD, the most well-known pathophysiological effect of CO is  
8 tissue hypoxia caused by binding of CO to Hb. Not only does the formation of COHb reduce the O<sub>2</sub>-  
9 carrying capacity of blood, but it also impairs the release of O<sub>2</sub> from O<sub>2</sub>Hb. Compensatory alterations in  
10 hemodynamics, such as vasodilation and increased cardiac output, protect from tissue hypoxia. At  
11 ambient CO concentrations, these compensatory changes are slight and likely tolerable in people with a  
12 healthy cardiovascular system. However, people with cardiovascular detriments may be unable to endure  
13 these small changes in hemodynamics which may lead to the presentation of health effects as described in  
14 Sections 5.2.1 and 5.2.2.

15           The 2000 CO AQCD reported changes in vasodilation due to CO levels between 500-2000 ppm  
16 (Kanten et al., 1983; MacMillan, 1975). Vasodilation can be dependent on or independent of perturbations  
17 in O<sub>2</sub> supply (Koehler et al., 1982). For example, cerebral blood flow elevations that were independent of  
18 O<sub>2</sub> availability were blocked by the inhibition of nitric oxide synthase (NOS) indicating a role for the free  
19 radical species nitric oxide (NO•) in mediating vasodilation (Meilin et al., 1996).

20           Increased cardiac output was also discussed in the 2000 CO AQCD as a compensatory response to  
21 CO-mediated tissue hypoxia. Findings of studies measuring hemodynamic alterations following CO  
22 exposure were equivocal and sometimes contradictory (Penney, 1988). While most studies reported a  
23 positive correlation between COHb and cardiac output at COHb levels above 20%, one study  
24 demonstrated increased cardiac output in humans following acute exposure to 5% CO resulting in COHb  
25 levels around 9% (Ayres et al., 1973). There was no increase in cardiac output following acute exposure

1 to 0.1% CO in this latter study. Increased heart rate and stroke volume (SV) were also observed in  
2 response to CO exposure (Stewart et al., 1973b); however, some experiments found no change in SV in  
3 humans with 18-20% COHb (Vogel and Gleser, 1972) or 12.5% COHb (Klausen et al., 1968). The 2000  
4 CO AQCD reported that BP was generally unchanged in human CO exposure studies, while a number of  
5 animal studies demonstrated CO-induced hypotension (Penney et al., 1988). No changes in forearm blood  
6 flow, BP, or heart rate were reported in humans experiencing approximately 8% COHb (Hausberg and  
7 Somers, 1997). However, high concentration animal exposures (3,000 to 10,000 ppm) showed diminished  
8 organ blood flow (Brown and Piantadosi, 1992). In depth discussion of hemodynamic changes resulting  
9 from CO exposure in recent human clinical studies can be found in Section 5.2.2.

10 Binding of CO to Mb, as discussed in the 2000 CO AQCD and in Section 4.3.2.3, can also impair  
11 the delivery of O<sub>2</sub> to tissues. Mb has a high affinity for CO; however, physiological effects are seen only  
12 after high dose exposures to CO, resulting in COMb concentrations far above baseline levels. High  
13 energy phosphate production in cardiac myocytes was inhibited when COMb concentrations exceeded  
14 40%, corresponding to an estimated COHb level between 20-40% (Wittenberg and Wittenberg, 1993).  
15 Conversely, rat hearts perfused with 6% CO (60,000 ppm) exhibited no change in high energy phosphate  
16 production, respiration rate, or contractile function (Chung et al., 2006; Glabe et al., 1998). These  
17 conflicting studies employed CO levels that are not relevant for ambient exposure to CO.

### 5.1.3. Non-Hypoxic Mechanisms

18 Non-hypoxic mechanisms underlying the biological effects of CO were discussed in the 2000 CO  
19 AQCD and are summarized below. Most of these mechanisms are related to CO's ability to bind heme-  
20 containing proteins other than Hb and Mb (Raub and Benignus, 2002). Since then, additional experiments  
21 have confirmed and extended these findings. While the majority of the older studies utilized  
22 concentrations of CO far higher than ambient levels, many of the newer studies have employed more  
23 environmentally relevant levels of CO.

#### 5.1.3.1. Non-Hypoxic Mechanisms Reviewed in the 2000 CO AQCD

24 Inhibition of heme-containing proteins such as cytochrome *c* oxidase and cytochrome P450  
25 oxidases may alter cellular function. CO interacts with the ferrous heme *a*<sub>3</sub> of the terminal enzyme of the  
26 electron transport chain, cytochrome *c* oxidase (Petersen, 1977). Cytochrome *c* oxidase inhibition not  
27 only interrupts cellular respiration and energy production, but can also enhance reactive oxygen species  
28 (ROS) production. In vivo studies observed CO binding to cytochrome *c* oxidase under conditions where  
29 COHb concentrations are above 50% (Brown and Piantadosi, 1992). It is unlikely that this could arise  
30 under physiological conditions or under conditions relevant to ambient exposures.

1 Studies indicated that CO exposure produces a pro-oxidant cellular environment by a second  
2 mechanism. CO exposure at concentrations of 10-20 ppm and above caused platelets of laboratory rats, as  
3 well as cultured bovine pulmonary endothelial cells, to release NO• (Thom and Ischiropoulos, 1997). An  
4 increase in available NO• was also seen in the lung and brain of CO-exposed rats (Ischiropoulos et al.,  
5 1996; Thom et al., 1999b). NO• combines with superoxide to form the highly active oxidant species,  
6 peroxynitrite (Thom et al., 1997), which can lead to the activation and sequestration of leukocytes in brain  
7 vessels (Thom et al., 2001b) and aorta, oxidation of plasma lipoproteins (Thom et al., 1999a), and the  
8 formation of protein nitrotyrosine (Ischiropoulos et al., 1996; Thom et al., 1999a; Thom et al., 1999b).  
9 The mechanism leading to NO• release is likely the displacement of NO• from nitrosyl bound Mb (NO•-  
10 Mb) or Hb (NO•-Hb) by CO. The rate of this event is slow; however it occurs at environmentally relevant  
11 concentrations of CO (Thom et al., 1997).

12 CO exposure also increases the production of other pro-oxidant species, including hydrogen  
13 peroxide (H<sub>2</sub>O<sub>2</sub>) and hydroxyl radical (OH•). High level CO exposure (2,500 ppm) increases OH• in rat  
14 brain and this response is distinct from tissue hypoxia (Piantadosi et al., 1997). The mechanism for  
15 enhanced H<sub>2</sub>O<sub>2</sub> production is unclear. The release of H<sub>2</sub>O<sub>2</sub> in the lung of CO-exposed rats was dependent  
16 upon the production of NO•, as it was inhibited by the pretreatment with a NOS inhibitor (Thom et al.,  
17 1999b). It is possible that peroxynitrite formed after CO exposure inhibited electron transport at  
18 complexes I through III, or that cytochrome *c* oxidase interference led to mitochondrial dysfunction and  
19 ROS production.

20 It has also been suggested that CO leads to vasorelaxation by three different mechanisms. First, CO  
21 may inhibit the P450-dependent synthesis of vasoconstrictors (Wang, 1998). Vasodilation has been  
22 demonstrated via this P450-mediated mechanism with high concentrations of CO (approximately  
23 90,000 ppm) (Coceani et al., 1988). In the case of cytochrome P450 enzymes, tissue CO levels may have  
24 to be abnormally increased to elicit a physiological response since the in vitro Warburg binding  
25 coefficients, the ratio of CO to O<sub>2</sub> at which half the reactive sites are occupied by CO, for cytochrome  
26 P450s range from 0.1-12 (Piantadosi, 2002). P450 inhibition may also reduce the hypoxia-induced  
27 expression of mitogens such as erythropoietin (EPO), vascular endothelial growth factor (VEGF),  
28 endothelin-1 (ET-1), and platelet derived growth factor (PDGF) which may decrease smooth muscle  
29 proliferation in response to hypoxia (Wang, 1998). CO may also interfere with the metabolism of  
30 barbiturates and other drugs; however, this is probably due to the hypoxic actions of CO rather than to  
31 P450 inhibition (Roth and Rubin, 1976a, b).

32 Secondly, CO has been shown to play a physiological role in vasomotor control and in signal  
33 transduction by activation of soluble guanylate cyclase (sGC), causing a conversion of GTP to cyclic  
34 GMP (cGMP). CO interacts with sGC and forms an unstable complex with the heme core of the protein.  
35 The resulting porphyrin IX intermediate triggers cGMP production (Ndisang et al., 2004). CO causes  
36 vascular relaxation, independent of the endothelium, in human arterial rings (Achouh et al., 2008), rat tail

1 artery (Wang et al., 1997), and rat thoracic aorta (Lin and McGrath, 1988), but not cerebral vessels  
2 (Andresen et al., 2006; Brian et al., 1994). Activation of sGC by CO has been linked to  
3 neurotransmission, vasodilation, bronchodilation, inhibition of platelet aggregation, and inhibition of  
4 smooth muscle proliferation (Brune and Ullrich, 1987; Cardell et al., 1998; Morita et al., 1997; Verma et  
5 al., 1993).

6 CO-mediated vasorelaxation can also be caused by activation of voltage-activated potassium ( $K^+$ )  
7 channels or calcium ( $Ca^{2+}$ )-activated  $K^+$  channels, which leads to membrane hyperpolarization, voltage-  
8 dependent  $Ca^{2+}$  channel closing, reduction of resting  $Ca^{2+}$  concentration, and then vascular tissue  
9 relaxation (Farrugia et al., 1993; Wang et al., 1997). This effect may be linked to sGC activity; however it  
10 has also been reported to occur independently (Dubuis et al., 2003; Naik and Walker, 2003).  
11 Developmental stage and tissue type will determine whether  $K^+$  channels or the sGC/cGMP pathway  
12 plays more of a role in vasorelaxation (Ndisang et al., 2004).

13 Collectively, these older studies demonstrated that exposures to high concentrations of CO resulted  
14 in altered functions of heme proteins other than Hb and Mb. Decreased cellular respiration and energy  
15 production and increased pro-oxidant status following cytochrome *c* oxidase inhibition would likely  
16 predispose towards cellular injury and death. The increase in free  $NO^{\bullet}$  following its release from  
17 sequestered stores could also contribute to the pro-oxidant status. Increased ROS and reactive nitrogen  
18 species are known to promote cell signaling events leading to inflammation and endothelial dysfunction.  
19 An inappropriate increase in vasorelaxation due to inhibition of vasoconstrictor production or to  
20 activation of vasodilatory pathways (sGC and ion channels) could potentially limit compensatory  
21 alterations in hemodynamics. Alternatively, CO-binding to sGC could result in decreased vasorelaxation  
22 by interfering with the binding of  $NO^{\bullet}$  to sGC.  $NO^{\bullet}$  can also activate sGC, and with a 30-fold greater  
23 affinity than CO is one-thousand fold more potent with respect to vasodilation and sGC activation (Stone  
24 and Marletta, 1994). CO could further contribute to endothelial dysfunction by this mechanism. Although  
25 the 2000 CO AQCD made no definitive links between these non-hypoxic mechanisms of CO and  
26 CO-mediated health effects, it did establish the potential for CO to interfere with basic cellular and  
27 molecular processes which could lead to dysfunction and/or disease.

### 5.1.3.2. Recent Studies of Non-Hypoxic Mechanisms

28 Since the 2000 CO AQCD, new studies have provided additional evidence for non-hypoxic  
29 mechanisms of CO which involve the binding of CO to reduced iron in heme proteins. These  
30 mechanisms, which may be inter-related, are described below and include:



- 1           ▪ Alteration in NO• signaling
- 2           ▪ Inhibition of cytochrome *c* oxidase
- 3           ▪ Heme loss from protein
- 4           ▪ Disruption of iron homeostasis
- 5           ▪ Alteration in cellular redox status

6           This assessment evaluates these non-hypoxic mechanisms in terms of their potential to contribute  
7 to health effects associated with environmentally-relevant CO exposures. As discussed above, CO at high  
8 concentrations may promote oxidative stress, cell injury and death, inflammation and endothelial  
9 dysfunction. Whether lower CO concentrations trigger these same processes is of key interest since these  
10 can potentially contribute to adverse health effects.

11           In addition, a large number of studies published since the 2000 CO AQCD has focused on the role  
12 of COs as an endogenous signaling molecule and the potential therapeutic effects of exogenous CO  
13 administered at high concentrations. This assessment addresses these topics only briefly, as they pertain to  
14 the evaluation of health effects associated with environmentally-relevant CO exposures.

### **Alteration in NO• Signaling**

15           Work by Thorup et al. (1999) demonstrated altered NO• signaling in isolated rat renal resistance  
16 arteries. In one set of experiments, rapid release of NO• was observed in response to a supra-physiologic  
17 dose of CO (100 nM). These findings are similar to those of Thom and colleagues (Ischiropoulos et al.,  
18 1996; Thom et al., 1994; Thom and Ischiropoulos, 1997; Thom et al., 1997; Thom et al., 1999a; Thom et  
19 al., 1999b; Thom et al., 2000; Thom et al., 2006) who demonstrated NO• release, presumably from  
20 sequestered stores, in platelets, endothelial cells, aorta and lung in response to CO (see above).  
21 Furthermore in a second set of experiments, Thorup et al. (1999) demonstrated inhibition of NOS in  
22 isolated rat renal resistance arteries. Here rapid NOS-dependent release of NO• following carbachol  
23 stimulation was blocked by pretreatment with 100 nM NO•. Both sets of studies illustrate the potential of  
24 CO to alter processes dependent on endogenous NO•. This could be critical in the case of cGMP-  
25 mediated vasodilation since, as discussed above, NO• activates sGC to a greater extent than CO. Thus in  
26 the presence of excess CO, NO•-dependent vasodilation may be significantly less.

### **Inhibition of Cytochrome *c* Oxidase**

27           High concentrations of CO are known to inhibit cytochrome *c* oxidase, the terminal enzyme in the  
28 mitochondrial electron transport chain, resulting in inhibition of cellular respiration and the formation of  
29 superoxide from mitochondrial substrates. Several recent studies demonstrated CO-mediated decreases in  
30 cytochrome *c* oxidase activity in model systems ranging from isolated mitochondria to whole animals. In

1 a study by Alonso et al. (2003), exposure of isolated mitochondria from human skeletal muscle to 50-  
2 500 ppm CO for 5 min decreased cytochrome *c* oxidase activity. Similarly, exposure of cultured  
3 macrophages to 250 ppm CO for 1 h inhibited cytochrome *c* oxidase (Zuckerbraun et al., 2007). In this  
4 study, increased ROS were observed following exposure to 250 ppm, as well as to CO concentrations as  
5 low as 50 ppm, for 1 h. Animal studies demonstrated that exposure of rats to 250 ppm CO for 90 min  
6 inhibited cytochrome *c* oxidase activity in myocardial fibers (Favory et al., 2006a) and exposure of mice  
7 to 1,000 ppm CO for 3 h decreased cytochrome *c* oxidase activity in heart mitochondria (Iheagwara et al.,  
8 2007). COHb levels were reported to be 61% in the latter study.

### **Heme Content Loss from Proteins**

9 In addition to decreasing the activity of cytochrome *c* oxidase, exposure of mice to 1,000 ppm CO  
10 for 3 h resulted in decreased protein levels and heme content of cytochrome *c* oxidase (Iheagwara et al.,  
11 2007). CO-mediated heme release was also seen in a study by Cronje et al. (2004), and was followed by  
12 increased endogenous CO production through the activation of HO-2 and the induction of HO-1. Loss of  
13 heme from proteins leads to loss of protein function and often to protein degradation.

### **Disruption of Iron Homeostasis**

14 Exposure of rats to 50 ppm CO for 24 h increased levels of iron and ferritin in the bronchoalveolar  
15 lavage fluid (BALF), decreased lung non-heme iron and increased liver non-heme iron (Ghio et al.,  
16 2008). Furthermore in this same study, exposure of respiratory epithelial cells to 10-100 ppm CO for 24 h  
17 caused a dose-dependent decrease in cellular non-heme iron and ferritin. Heme loss observed in other  
18 studies (Cronje et al., 2004; Iheagwara et al., 2007) might be expected to result in disruption of iron  
19 homeostasis. Iron homeostasis is critical for the sequestration of free iron and the prevention of iron-  
20 mediated redox cycling which can lead to ROS generation and lipid peroxidation.

### **Alteration in Cellular Redox Status**

21 Recent studies demonstrated that exposure to low, moderate and high levels of CO increased  
22 cellular oxidant stress in cultured cells (Kim et al., 2008; Zuckerbraun et al., 2007). A dose-dependent  
23 increase in dichlorofluorescein (DCF) fluorescence (an indicator of ROS) occurred following 1-h  
24 exposure to 50-500 ppm CO in macrophages and following 1-h exposure to 250 ppm CO in hepatocytes.  
25 NOS inhibition had no effect on the increase in DCF fluorescence in CO-treated macrophages indicating  
26 that the effects were not due to an interaction of CO and NO<sup>•</sup> (Zuckerbraun et al., 2007). Mitochondria  
27 were identified as the source of the increased ROS since mitochondria-impaired cells (rho zero cells and  
28 treatment with antimycin A) did not respond to CO with an increase in DCF fluorescence. Furthermore,  
29 1-h exposure to 250 ppm CO inhibited mitochondrial cytochrome *c* oxidase enzymatic activity in  
30 macrophages (Zuckerbraun et al., 2007). Recently, inhibition of cytochrome *c* oxidase was demonstrated

1 in HEK-293 cells transfected with HO-1 and in macrophages with induced HO-1, and this effect was  
2 attributed to endogenously produced CO (D'Amico et al., 2006). However inhibition of cytochrome *c*  
3 oxidase at CO concentrations below 250 ppm has not been definitively demonstrated. In hepatocytes,  
4 exposure to 250 ppm CO for 1 h resulted in Akt phosphorylation and nuclear translocation of nuclear  
5 factor kappa B (NF- $\kappa$ B), effects which were blocked by antioxidants (Kim et al., 2008). Significant  
6 increases in apoptosis were also observed in this model. Thus, CO appeared to uncouple mitochondrial  
7 respiration leading to ROS-induced programmed cell death.

8 Further evidence for cellular redox stress is provided by studies in which glutathione stores were  
9 altered following CO exposure in vitro (Kim et al., 2008; Patel et al., 2003). In addition, mitochondrial  
10 redox stress was observed in livers of rats exposed to 50 ppm CO (Piantadosi et al., 2006). Furthermore,  
11 an adaptive increase in intracellular antioxidant defenses (i.e., superoxide dismutase) was observed in  
12 endothelial cells exposed to 10 ppm CO for 40 min (Thom et al., 2000) and mitochondrial biogenesis was  
13 observed in hearts of mice exposed to 50 and 250 ppm CO for 1 h (Suliman et al., 2007).

14 Several mechanisms could contribute to the cellular redox stress elicited by CO exposure. As  
15 discussed above, inhibition of cytochrome *c* oxidase results in the formation of superoxide from  
16 mitochondrial substrates. However, interactions of CO with heme proteins can lead to the release of heme  
17 and free iron which could also lead to the generation of ROS. As mentioned above, increased ROS  
18 generation has been linked to cellular injury and death, inflammation, and endothelial dysfunction.

19 Two of the above-mentioned studies demonstrated that CO-mediated mechanisms were unrelated  
20 to hypoxia by showing that hypoxic conditions failed to mimic the results obtained with CO. Hence the  
21 mitochondrial redox stress and mitochondrial pore transition observed in livers from rats exposed to CO  
22 (Piantadosi et al., 2006) and the cardiac mitochondrial biogenesis observed in mice exposed to CO  
23 (Suliman et al., 2007) could be attributed specifically to non-hypoxic mechanisms of CO.

24 Recent studies also demonstrated non-hypoxic mechanisms of CO which do not directly involve  
25 heme protein interactions. These mechanisms are described below and include:

- 26     ▪ Alteration in ion channel activity
- 27     ▪ Modulation of protein kinase pathways

### **Alteration in Ion Channel Activity**

28 Work by Dubuis et al. (2002) demonstrated increased current through  $\text{Ca}^{2+}$ -activated  $\text{K}^+$  channels in  
29 smooth muscle cells from pulmonary arteries of rats exposed to 530 ppm CO for 3 weeks. These findings  
30 provide further evidence for non cGMP-dependent vasodilatory actions of CO.

## Modulation of Protein Kinase Pathways

1 Endogenously produced CO is a gaseous second messenger molecule in the cell. Work from  
2 numerous laboratories has demonstrated the potential for CO to be used as a therapeutic gas with  
3 numerous possible clinical applications, since it can produce anti-inflammatory, anti-apoptotic, and anti-  
4 proliferative effects (Ryter et al., 2006). These studies generally involved pretreatment with CO followed  
5 by exposure to another agent 18-24 h later. There is extensive literature on this topic as reviewed by Ryter  
6 et al. (2006) and others. A number of these processes are mediated through cGMP while others involve  
7 redox-sensitive kinase pathways, possibly secondary to CO-dependent generation of ROS. For example,  
8 250 ppm CO inhibited growth of airway smooth muscle cells by attenuating the activation of the  
9 extracellular signal-regulated kinase 1/2 (ERK 1/2) pathway, independent of sGC and other MAP kinases  
10 (Song et al., 2002). A second example is provided by the study of Kim et al. (2005) where 250 ppm CO  
11 inhibited PDGF- induced smooth muscle cell proliferation by upregulating p21<sup>Waf1/Cip1</sup> and caveolin-1, and  
12 down-regulating cyclin A expression. These events were dependent upon cGMP and the p38 MAPK  
13 pathway (Kim et al., 2005). Thirdly, rat endothelial cells exposed to 15 ppm CO escaped anoxia-  
14 reoxygenation mediated apoptosis via modulation of the signaling pathways involving phosphoinositide  
15 3-kinase (PI3K), Akt, p38 MAP kinase, Signal Transducers and Activators of Transcription (STAT-1) and  
16 STAT-3 (Zhang et al., 2005). In a fourth study, Akt was found to be responsible for the CO-induced  
17 activation of NF- $\kappa$ B, protecting against hepatocyte cell death (Kim et al., 2008). While research focusing  
18 on therapeutic applications of CO generally involves high level, short-term exposure to CO (250-  
19 1,000 ppm for up to 24 h), some studies found effects below 20 ppm (Zhang et al., 2005). Few if any  
20 studies on the therapeutic effects of CO have explored the dose-response relationship between CO and  
21 pathway activation/deactivation, so it remains unclear how these effects may be related to  
22 environmentally-relevant exposures.

## Concentration-Response Relationships

23 In many cases the concentrations of exogenous CO required for these non-hypoxic mechanisms is  
24 much higher than what would be expected to result from exposures at ambient levels (Alonso et al., 2003;  
25 Favory et al., 2006b; Iheagwara et al., 2007; Thorup et al., 1999). However in some studies the effects are  
26 mimicked by upregulation of HO-1 which would result in increased local production of CO as well as of  
27 iron and biliverdin (D'Amico et al., 2006; Imai et al., 2001; Thorup et al., 1999). For example, HO-1  
28 overexpression attenuated carbachol-mediated NO<sup>•</sup> release and NO<sup>•</sup>-mediated vasodilation, similar to the  
29 effects of exogenous CO in these same models (Imai et al., 2001; Thorup et al., 1999). In the study by  
30 D'Amico et al. (2006), overexpression of HO-1 in cells inhibited cellular respiration by 12% and  
31 decreased cytochrome c oxidase activity by 23%. It is not clear how comparable these conditions

1 involving increased intracellular concentrations of endogenous CO are to increased intracellular  
 2 concentrations of CO resulting from ambient CO exposures.

3 There is a growing appreciation that non-hypoxic mechanisms may contribute to the effects  
 4 associated with CO poisoning (Ischiropoulos et al., 1996; Thom et al., 1994; Weaver et al., 2007). On the  
 5 other hand, recent studies suggest that exogenous CO at low concentrations may have beneficial anti-  
 6 inflammatory, anti-proliferative and cytoprotective effects under certain circumstances (Ryter et al.,  
 7 2006). Since the focus of this assessment is on mechanisms which are relevant to ambient exposures it is  
 8 important to understand which mechanisms may occur at “low” (50 ppm and less) and “moderate”  
 9 (50-500 ppm CO) concentrations of CO. Hence, both recent animal studies and relevant older ones which  
 10 add to the understanding of mechanisms in this range of CO concentrations are briefly summarized in  
 11 Table 5-1. It should be noted that most of the above-mentioned non-hypoxic mechanisms are  
 12 demonstrated at CO concentrations of 50 ppm and less.

**Table 5-1. Responses to low and moderate CO exposures.**

Reference	Model System	CO Exposure	Response	Notes
<i>IN VITRO</i>				
Alonso et al. (2003)	Human muscle mitochondria	50, 100, 500 ppm 5 min	Decreased cytochrome <i>c</i> oxidase activity	
Thom and Ischiropoulos (1997)	Rat platelets	10 ppm	Increased free NO*	
Thom et al. (1997)	Bovine pulmonary artery endothelial cells	20 ppm 30-60 min	Increased free NO* and peroxynitrite	Reported to correspond to 7% COHb
Thom et al. (2000)	Bovine pulmonary artery endothelial cells	10 ppm 40 min	Increased MnSOD and protection against toxic effects of 100 ppm CO	Adaptive responses
Song et al. (2002)	Human aortic smooth muscle cells	50-500 ppm 24 h	Inhibition of cellular proliferation	Blocked activation of ERK1/2 pathway, independent of sGC and other MAP kinases
Kim et al. (2005)	Rat pulmonary artery smooth muscle cells	250 ppm 1 h	Inhibited PDGF- induced smooth muscle cell proliferation	Upregulated p21 <sup>Waf1/Cip1</sup> and caveolin-1, and down-regulated cyclin A expression.
Kim et al. (2008)	Rat hepatocytes	250 ppm 1 h 2x per day 250 ppm 1 h	Blocked spontaneous apoptosis Increased mitochondrial ROS generation, increased mitochondrial glutathione oxidation, and decreased cellular ascorbic acid	CO induced Akt phosphorylation via ROS production CO activated NFκB
Zhang et al. (2005)	Rat pulmonary artery endothelial cells	15 ppm 0.5-24 h	Blocked anoxia-reoxygenation mediated apoptosis	Modulation of PI3K/Akt/p38 MAP kinase and STAT-1 and STAT-3
Zuckerbraun et al. (2007)	Mouse macrophages	50 and 250 ppm 1 h	Increased ROS generation (dose dependent response for 50-500 ppm CO)	Mitochondrial derived ROS and cytochrome <i>c</i> oxidase inhibition demonstrated for 250 ppm

Reference	Model System	CO Exposure	Response	Notes
Ghio et al. (2008)	Human bronchial epithelial cells	10-100 ppm 24 h	Dose-dependent decrease in cellular non-heme iron (effect at 10 ppm was significant, effect at 50 ppm maximal) Dose-dependent decrease in cellular ferritin at 50-100 ppm 50 ppm blocked iron uptake by cells 50 ppm increased iron release from cells	Compare with in vivo experiment in same paper
<i>IN VIVO</i>				
Ghio et al. (2008)	Rats	50 ppm 24 h	Mild neutrophil accumulation in BALF Increased lavage MIP-2, protein, LDH Lavage iron and ferritin were increased by CO Lung non-heme iron was decreased by CO Liver non-heme iron was increased by CO	
Thom et al., (1999b)	Rats	50 ppm 1 h  100 ppm 1 h	Increased nitrotyrosine in aorta Leukocyte sequestration in aorta after 18 h Albumin efflux from skeletal muscle microvasculature 3 h after CO LDL oxidation	Effects blocked by NOS inhibitor
Thom et al., (1999a)	Rats	100 ppm 1 h 50 ppm 1 h	Elevated free NO* during CO exposure (EPR) Elevated nitrotyrosine in lung homogenates Lung capillary leakage 18 h after exposure	Inhibition of NOS abrogated CO effects
Sorhaug et al. (2006)	Rats	200 ppm 72 weeks	No changes in lung morphology No pulmonary hypertension No atherosclerotic lesions in systemic vessels Ventricular hypertrophy	
Loennechen et al. (1999)	Rats	100 and 200 ppm 1-2 weeks	Increased ET-1 mRNA in the heart ventricles, increased right and left ventricular weight	12 and 23% COHb
Favory et al. (2006b)	Rats	250 ppm 90 min	Complex IV inhibition in myocardial fibers Inhibition of vasodilatory response to acetylcholine and SNP, increased coronary perfusion pressure and contractility	11% COHb
Piantadosi et al. (2006)	Rats	50 ppm CO or HH for 1, 3, or 7 days	Liver mitochondrial oxidative and nitrosative stress, altered mitochondrial permeability pore transition sensitivity	CO effects not mimicked by HH
Suliman et al. (2007)	Mice	50 and 250 ppm 1 h	Cardiac mitochondrial biogenesis	Activation of GC involved. No role for NOS. Increased mitochondrial H <sub>2</sub> O <sub>2</sub> and activation of Akt proposed
Wellenius et al. (2004)	Rats Model of MI	35 ppm 1 h	Decreased delayed ventricular beat frequency	Altered arrhythmogenesis
Wellenius et al. (2006a)	Rats Model of MI	35 ppm 1 h	Decreased supraventricular ectopic beats	Altered arrhythmogenesis
Carraway et al. (2002)	Rats Model of hypoxic pulmonary vascular remodeling	Hypobaric hypoxia ± 50 ppm CO 3 weeks	CO promoted remodeling and increased pulmonary vascular resistance	
Gautier et al. (2007)	Rats Model of right ventricle hypertrophy secondary to chronic hypoxia	3 weeks of HH with 50 ppm CO during last week	Rats with pulmonary hypertension were more sensitive to CO which altered the right ventricular adaptive response to pulmonary hypertension leading to ischemic lesions	
Melin et al. (2005)	Rats Model of right ventricle hypertrophy secondary to chronic hypoxia	50 ppm 10 wk	CO increased cardiac dilation and decreased left ventricular function	

Reference	Model System	CO Exposure	Response	Notes
Melin et al. (2002)	Rats Model of right ventricle hypertrophy secondary to chronic hypoxia	50 ppm 10 wk	CO increased right ventricular hypertrophy, decreased right ventricular diastolic function and increased left ventricular weights	

### 5.1.3.3. Implications of Non-Hypoxic Mechanisms

1 A key issue in understanding the biological effects of environmentally-relevant exposures to CO is  
2 whether the resulting partial pressures of CO in cells and tissues can initiate cell signaling or perturb  
3 signaling which normally is mediated by endogenously generated CO.

4 Several aspects need to be considered. One is that, for most cells and tissues, Hb acts as a buffer to  
5 limit the availability of free CO. At the same time, COHb delivers CO to cells and tissues. CO  
6 dissociation from Hb occurs as a function of CO diffusion down a concentration gradient. Hence, greater  
7 release of CO from COHb will occur under conditions of low cell/tissue CO concentration.

8 A second consideration is the role played by O<sub>2</sub> in competing with CO for binding to intracellular  
9 heme protein targets. In general, heme proteins (e.g., cytochrome *c* oxidase) are more sensitive to CO  
10 when O<sub>2</sub> is limited. Hence hypoxic conditions would be expected to enhance the effects of CO. This  
11 concept is demonstrated in the study by D'Amico et al. (2006).

12 A third consideration is whether certain cell types serve as primary targets for the effects of CO.  
13 Besides the RBC, the first cells encountering CO which dissociates from Hb will be the endothelial cells  
14 lining blood vessels. White blood cells in the circulation may also be first-line targets of Hb-dissociated  
15 CO. An exception to this situation is in the lungs where epithelial and inflammatory cells found in  
16 airways and alveoli are exposed to free CO prior to CO binding to Hb. These lung cells may also serve as  
17 unique targets for CO. Processes such as endothelial dysfunction, inflammatory cell activation and  
18 respiratory epithelial injury may ensue as a result of preferential targeting of these cell types.

19 Fourth, it should be considered that adaptation to chronic exogenous CO exposure might occur and  
20 that intermittent exogenous CO exposure might have unexpected effects.

21 Since there is potential for exogenous CO to affect endogenous pools of CO, it is important to  
22 know the concentrations of CO in cells and tissues before and after exogenous exposures. Table 5-2  
23 summarizes findings from 4 recent studies relevant to this issue. It should be noted that exposure to  
24 50 ppm CO resulted in a 3-5 fold increased in tissue CO concentration.

**Table 5-2. Tissue concentration of CO following inhalation exposure.**

Reference	CO Exposure	Tissue CO Concentrations	COHb	Notes
Cronje, et al. (2004)	Rat 2,500 ppm 45 min	Blood: 27,500 (800) pmol/mg Heart: 800 (300) pmol/mg Muscle: 90 (80) pmol/mg Brain: 60 (40) pmol/mg Control levels in parentheses	66-72%	CO concentration increased in the heart but not in brain or skeletal muscle after CO exposure A later report stated that these tissue CO values were too high due to a computational error (Piantadosi et al., 2006)
Vreman, et al. (2005)	Mice 500 ppm 30 min	Blood: 2648 ± 400 (45) pmol/mg Heart: 100 ± 18 (6) pmol/mg Muscle: 14 ± 1 (10) pmol/mg Brain: 18 ± 4 (2) pmol/mg Kidney: 120 ± 12 (7) pmol/mg Spleen: 229 ± 55 (6) pmol/mg Liver: 115 ± 31 (5) pmol/mg Lung: 250 ± 2 (3) pmol/mg Intestine: 9 ± (4) pmol/mg Testes: 6 ± 3 (2) pmol/mg Control levels in parentheses	28%	CO concentration relative to 100% blood: Lung: 9.4% Spleen: 8.6% Kidney: 4.5% Liver: 4.3% Heart: 3.8% Brain: 0.7% Muscle: 0.5% Intestine: 0.3%, Testes: 0.2%
Piantadosi et al. (2006)	Rats 50 ppm 1-7 days	Liver: 30-40 pmol/mg Control liver 10 pmol/mg	4-5% Control 1%	CO concentration plateaued after 1 day
Suliman et al. (2007)	Mice 50, 250 and 1250 ppm 1 h	Heart (left ventricle) 50 ppm: 50 pmol/mg 250 ppm: 95 pmol/mg 1250 ppm: 160 pmol/mg Control heart: 9 pmol/mg		No mention of COHb% but exposures were similar to those in Cronje et al. (2004)

1 Furthermore, endogenous CO production is increased during inflammation, hypoxia, increased  
 2 heme availability and other conditions where HO-1 or HO-2 activity is increased. A few studies reported  
 3 increased COHb levels and/or cell and tissue concentrations of CO resulting from enhanced endogenous  
 4 CO production. Table 5-3 summarizes these findings.

**Table 5-3. Tissue concentration of CO following increased endogenous production.**

Reference	Exposure	Tissue CO	COHb	Notes
Carraway, et al. (2000)	Rats Hypobaric Hypoxia for 21 days		1.5-2.8% Control 0.5%	COHb highest after days 1 and 21 at 3-4 fold higher than controls
Piantadosi et al. (2006)	Rats Hypobaric Hypoxia 1-7 days	Liver: 5-12 pmol/mg Control liver 10 pmol/mg	1-1.25% Control 1%	CO concentration plateaued after 1 day
Vreman et al. (2005)	Mice 30 µM heme	Blood: 88 ± 10 (45) pmol/mg Heart: 14 ± 3 (6) pmol/mg Muscle: 7 ± 1 (10) pmol/mg Brain: 2 ± 0 (2) pmol/mg Kidney: 7 ± 2 (7) pmol/mg Spleen: 11 ± 1 (6) pmol/mg Liver: 8 ± 3 (5) pmol/mg Lung: 8 ± 3 (3) pmol/mg Intestine: 3 ± 1 (4) pmol/mg Testes: 2 ± 0 (2) pmol/mg Control levels in parentheses	0.9%	CO concentration relative to 100% blood: Heart: 16% Spleen: 13% Lung: 9% Liver: 9% Kidney: 8% Muscle: 8% Intestine: 3% Brain: 2% Testes: 2%



1           It should be noted that increased cellular and tissue concentrations of biliverdin and iron  
2 accompany the increased endogenous production of CO by HO-1 and HO-2. Biliverdin and iron have  
3 known biological effects, with biliverdin exhibiting antioxidant properties and iron exhibiting pro-oxidant  
4 properties (Piantadosi, 2008), which could impact interpretation of results from studies in which HO-1  
5 and HO-2 activities are increased. Hence the situations of increased endogenous CO production and that  
6 of exogenous CO exposure are not equivalent. Nonetheless, additional measurements of CO levels in cells  
7 and tissues following increased endogenous production and following inhalation of exogenous CO is  
8 essential for a better understanding of the relationship between the CO tissue dose and the response.

9           In summary, CO is a ubiquitous cell signaling molecule and the physiological functions of HO-  
10 derived CO are numerous. The endogenous generation and release of CO from HO-1 and HO-2 is tightly  
11 controlled, as is any homeostatic process. Thus, exogenously-applied CO has the capacity to disrupt  
12 myriad heme-based signaling pathways due to its nonspecific nature. Only a limited amount of  
13 information is available regarding the impact of exogenous CO on tissue and cellular levels of CO.  
14 However recent animal studies demonstrated increased tissue CO levels and biological responses  
15 following exposure to 50 ppm CO. Whether or not environmentally relevant exposures to CO can affect  
16 endogenous CO signaling pathways and lead to adverse health effects is an open question for which there  
17 are no definitive answers at this time.

## 5.2. Cardiovascular Effects

18           This section characterizes the evidence from epidemiologic, controlled human exposure and animal  
19 toxicological studies on the cardiovascular effects of CO. While epidemiologic studies evaluated the  
20 effects of ambient exposures, experimental studies employed higher than ambient concentrations of CO  
21 but not levels of exposure associated with poisoning.

### 5.2.1. Epidemiologic Studies

22           The 2000 CO AQCD concluded that epidemiologic studies provided evidence that short-term  
23 variations in ambient CO concentrations were associated with daily hospital admissions for heart disease.  
24 The following section reviews the literature since the 2000 CO AQCD, including new studies on  
25 physiological cardiac endpoints and biomarkers and additional studies of daily hospital admissions for  
26 heart disease that support the evidence in the 2000 CO AQCD.

## 5.2.1.1. Epidemiologic Studies with Short-Term Exposure

### Heart Rate and Heart Rate Variability

1 Heart rate variability (HRV) refers to the beat-to-beat alterations in the heart and is generally  
2 determined by analyses of time and frequency domains measured by electrocardiograms (ECG). The time  
3 domains often analyzed are (a) normal-to-normal (NN or RR) time interval between each QRS complex,  
4 (b) standard deviation of the normal-to-normal interval (SDNN), and (c) mean squared differences of  
5 successive difference normal-beat to normal-beat intervals (rMSSD), shorter time domain variables  
6 results in lower HRV. The frequency domains often analyzed are a) the ratio of low energy frequency (LF)  
7 to high energy frequency (HF) and b) the proportion of interval differences of successive normal-beat  
8 intervals greater than 50 ms (PNN<sub>50</sub>), reflecting autonomic balance. Decreased HRV is associated with a  
9 variety of adverse cardiac outcomes such as arrhythmia, myocardial infarction (MI), and heart failure (De  
10 Jong and Randall, 2005; Deedwania et al., 2005; Huikuri et al., 1999; Rajendra Acharya et al., 2006).

11 Two studies investigated the association between ambient air pollution, including CO, and HRV in  
12 Boston, MA and reported inconsistent results. The earlier of these studies recruited twenty-one 53-to  
13 87-year old active residents and performed up to 12 ECG assessments on each subject over a period of 4  
14 months (during summer 1997). Particles (PM<sub>10</sub>, PM<sub>2.5</sub>) and several gaseous pollutants (O<sub>3</sub>, NO<sub>2</sub>, and SO<sub>2</sub>)  
15 were monitored at fixed sites (up to 4.8 miles from the study site) while CO was monitored 0.25 miles  
16 from each participants' residence. Lag periods for the preceding 1-h, 4-h, and 24-h before each subject's  
17 HRV assessment were analyzed and results showed that only PM<sub>2.5</sub> and O<sub>3</sub> were associated with HRV  
18 parameters (Gold et al., 2000).

19 A similar study by the same group of researchers two years later involved 28 older subjects (aged  
20 61-89) who were living at or near an apartment complex located on the same street as the Harvard School  
21 of Public Health. The subjects were seen once a week for up to 12 weeks and HRV parameters (SDNN,  
22 r-MSSD, PNN<sub>50</sub>, LF/HF ratio) were measured for 30 minutes each session. Data for PM<sub>2.5</sub>, black carbon  
23 (BC), and CO were recorded at the Harvard School of Public Health (<1 km from the residence) while  
24 data for NO<sub>2</sub>, O<sub>3</sub>, and SO<sub>2</sub> were collected from government regulatory monitoring sites. There were  
25 moderate correlations between CO and PM<sub>2.5</sub> (r = 0.61) and NO<sub>2</sub> (r = 0.55), but not with SO<sub>2</sub> (r = 0.18) or  
26 O<sub>3</sub> (r = 0.21). Similarly PM<sub>2.5</sub> was associated with HRV, whereas in contrast to the previous study, CO  
27 was associated<sup>1</sup> with a negative change in SDNN (% change: -13 [95% CI: -24.06 to -1.88]), r-MSSD  
28 (% change: -31.88 [95% CI: -38 to -7.5]), and PNN<sub>50</sub> (% change: -46.25 [95% CI -103.95 to -9.38]) per  
29 0.5 ppm increase in 24-h avg CO concentration) (Schwartz et al., 2005).

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<sup>1</sup> The effect estimates from epidemiologic studies have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations throughout this section (text, tables, and figures).

1 A later Boston, MA study examined HRV parameters (SDNN, LF, HF, LF/HF ratio) among 603  
2 persons from the Normative Aging Study, a longitudinal study that originally recruited 2,280 men in the  
3 greater Boston area during 1963. The cohort members were examined (November 2000 to October 2003)  
4 and the ECG data were linked to air pollution data for PM<sub>2.5</sub>, particle number concentration, BC, O<sub>3</sub>, NO<sub>2</sub>,  
5 SO<sub>2</sub>, and CO. Lagged pollutant effects for a 4-h, 24-h, and 48-h moving avg were used. Since previous  
6 studies established variable CO results, the main pollutant effects were with PM<sub>2.5</sub> and O<sub>3</sub> while CO was  
7 not associated with HRV (Park et al., 2005b).

8 A study in Mexico City selected 30 subjects from the outpatient clinic at the National Institute of  
9 Cardiology and followed them for ~10 hours (starting at 0900 hrs) (Riojas-Rodriguez et al., 2006). Each  
10 subject was connected to a Holter ECG monitor (e.g., a portable ECG monitor) and also given personal  
11 PM<sub>2.5</sub> and CO monitors. The subjects went about their usual daily activities and the personal PM<sub>2.5</sub> and  
12 CO data were linked to various ECG parameters (heart rate [HR], R-R, LF, HF) at various lags. In  
13 copollutant models (PM<sub>2.5</sub> and CO) personal CO exposure for the same 5-min period was significantly  
14 associated with a decrease in LF and very low energy frequency (VLF) parameters with coefficients equal  
15 to -0.024 (95% CI: -0.041 to -0.007) and -0.034 (95% CI: -0.061 to -0.007) respectively for a 1 ppm  
16 increase in 1-h CO concentration.

17 In Mexico City, 34 residents from a nursing home underwent HRV analysis every other day for 3  
18 months (Holguin et al., 2003). Exposure assessment for ambient PM<sub>2.5</sub> was based on data recorded at a  
19 monitor on the roof of the nursing home while exposures to ambient O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and CO were derived  
20 from data recorded at a fixed site 3 km from the nursing home. Exposures for the same day and 1-day lags  
21 were analyzed and only O<sub>3</sub> and PM<sub>2.5</sub> were positively associated with HRV.

22 Wheeler et al. (2006) examined 18 individuals with COPD and 12 individuals with recent MI  
23 living in Atlanta, GA. Morning ECG readings were collected by a Holter system by a field technician in  
24 the subjects' homes. Ambient air pollution exposures for PM<sub>2.5</sub>, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub> and CO were derived from  
25 data recorded at fixed sites throughout metropolitan Atlanta. Three exposure periods were analyzed: the  
26 hour of the ECG reading, 4-h mean and 24-h mean before the reading. While positive effects were  
27 reported for NO<sub>2</sub> and PM<sub>2.5</sub>, no quantitative results were reported for CO.

28 After reviewing 2,000 patient charts, Dales (2004) recruited 36 subjects with CAD from the  
29 Toronto Western Hospital's noninvasive cardiac diagnostic unit. HR and HRV (SDNN, N-N, HF, LF,  
30 HF/LH ratio) were assessed 1 day each week for up to 10 weeks by a Holter monitoring system. Personal  
31 air sampling for PM<sub>2.5</sub> and CO was carried out for the same 24-h period whereby subjects went about  
32 their usual daily activities for that period. Stratified results showed that among those not on beta-receptor-  
33 blockers, personal CO exposure was positively associated with SDNN (p = 0.02). However, in the group  
34 taking beta blockers there was a negative association (p = 0.06). Personal exposure to PM<sub>2.5</sub> was not  
35 associated with HRV.

1 HR was examined among a sub-sample of the Monitoring of Trends and Determinants in  
2 Cardiovascular Disease (MONICA) study (n = 2,681) in Augsburg, Germany by Peters and colleagues  
3 (1999a). Total suspended particles (TSP), SO<sub>2</sub>, and CO data were collected from a single monitoring  
4 station located in the center of the city and linked to each subject to estimate exposures on the same day  
5 and 5 days prior. A 0.5 ppm change in 24-h CO concentration was associated with an increase in HR of  
6 approximately 1 beat per minute, whereas CO based on a 5-day exposure had no effect on HR.

7 Liao et al (2004) investigated men and women aged 45-64 years from the Atherosclerosis Risk in  
8 Communities (ARIC) study (Washington County, MD; Forsyth County, NC; and selected suburbs of  
9 Minneapolis, MN). The sample sizes were 4,899, 5,431, 6,232, 4,390 and 6,784 for analyses involving  
10 PM<sub>10</sub>, O<sub>3</sub>, CO, NO<sub>2</sub>, and SO<sub>2</sub> respectively. County level exposure estimates for 24 h CO were calculated  
11 for 1, 2, and 3 days prior to clinical examination. A 0.5 ppm increase in 24-h CO concentration (at lag 1)  
12 was associated with an increase in HR (beats/minute) ( $\beta = 0.357$ ,  $p < 0.05$ ). CO was not significantly  
13 associated with changes in SDNN.

14 The Exposure and Risk Assessment for Fine and Ultrafine Particles in Ambient Air (ULTRA) study  
15 was carried out in three European cities: Amsterdam, the Netherlands, Erfurt, Germany, and Helsinki,  
16 Finland, whereby a panel of subjects with CAD was followed for 6 months with biweekly clinical visits,  
17 which included an ECG reading to assess HRV (Timonen et al., 2006). The time domain measures of  
18 HRV (SDNN and rMSSD) were analyzed along with frequency domain measures, which included power  
19 spectrum densities for LF and HF. Exposures to ambient air pollution (PM<sub>2.5</sub>, PM<sub>10</sub>, NO<sub>2</sub>, CO) were  
20 derived from data recorded at fixed monitoring site networks within each city. Correlation coefficients for  
21 NO<sub>2</sub> and CO ranged from 0.32 to 0.86 in the three cities. CO was moderately correlated with PM<sub>10</sub> in  
22 Helsinki ( $r = 0.40$ ) and with PM<sub>2.5</sub> in Amsterdam ( $r = 0.58$ ) and more highly correlated with PM<sub>10</sub> in  
23 Erfurt ( $r = 0.77$ ). Various lag periods were examined including lag 0 (24 h prior to the clinical visit)  
24 through a 0-2-day avg lag and a 0-4-day avg lag. In total there were 1,266 ECG recordings used in the  
25 final analyses. In the pooled analyses (e.g., across cities) a 0.5 ppm increase in 24-h CO concentration  
26 was associated with a decrease in LF/HF ratio at lag 1-day ( $\beta -16.4$  [95% CI: -29.9 to -0.3]), and a  
27 decrease in SDNN and HF at lag 2-day ( $\beta -3.4$  [95% CI: -6.1 to -0.4];  $\beta = -17.6$  [95% CI: -34.4 to -0.9],  
28 respectively). However, the same study reported no effect for CO on BP and HR (Ibald-Mulli et al.,  
29 2004).

30 A small panel study in Kuopio, Finland, which was designed as the pilot study for the ULTRA  
31 study examined simultaneous ambulatory ECG and personally monitored CO readings among 6 male  
32 patients with CAD (Tarkiainen et al., 2003). The patients were asked to follow their usual daily activities,  
33 but data were recorded only three times with 1-week intervals. The CO exposures were divided into low  
34 ( $\leq 2.7$  ppm) and high ( $>2.7$  ppm) and during the high CO exposure r-MSSD increased on average by 2.4  
35 ms. However, there was no effect on RR or SDNN.

1 A study in Taiwan recruited 83 patients (aged 40-75) from the National Taiwan University  
 2 Hospital, Taipei and conducted ambulatory ECG readings using a Holter system (Chan et al., 2005).  
 3 Ambient air pollution exposures for PM<sub>10</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and CO were derived from 12 fixed site monitoring  
 4 stations across Taipei. Lag periods of 1 h to 8 h prior to the ECG reading were analyzed and only NO<sub>2</sub>  
 5 was associated with HRV parameters (SDNN and LF). CO was not associated with HRV.

6 The ST-segment of an ECG represents the period of slow repolarization of the ventricles and  
 7 ST-segment depression can be associated with adverse cardiac outcomes. Gold et al. (2005) recruited a  
 8 panel of 28 older adults living at or near an apartment complex located within 1 km of a monitoring site in  
 9 Boston, MA. Each subject underwent weekly ECGs for 12 weeks in summer 1999 with the main outcome  
 10 of interest being the ST-segment. Air pollution data in the form of PM<sub>2.5</sub>, black carbon (BC), and CO were  
 11 collected from a central site within 0.5 km of the residences of the subjects and averaged over various lag  
 12 periods (1-24 h, 12 h and 24 h moving average [ma] ) before the ECG. The final analyses included 24  
 13 subjects with 269 observations and results showed consistent negative associations of ST-segment level  
 14 with increased BC with the strongest association with the 5-h lag. CO during the same lag period also  
 15 showed a negative association with ST-segment depression, however only BC remained significant in  
 16 multipollutant models.

17 In summary, few studies have examined the effect of CO on HR and while two of the three studies  
 18 reported a positive association, further research is warranted to corroborate the current results. Similarly,  
 19 while a larger number of studies have examined the effect of CO on various HRV parameters, mixed  
 20 results have been reported throughout these studies. Furthermore, with several HRV parameters often  
 21 examined, there are mixed results across the studies as to the HRV parameters that are positively  
 22 associated with CO exposure. Table 5-4 shows a summary of the reviewed studies.

**Table 5-4. Summary of studies investigating the effect of CO exposure on HRV parameters.**

Study	Location (Sample Size)	Cardiac Endpoint	Exposure Assessment	Mean CO Level (ppm)	Copollutants
Gold et al. (2000)	Boston, MA (n = 21)	HR, SDNN, r-MSSD	Ambient	Mean: 0.47(24 h) Range: 0.12-0.82	PM <sub>10</sub> , PM <sub>2.5</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Schwartz et al. (2005)	Boston, MA (n = 28)	SDNN, r-MSSD, PNN, LF/HF	Ambient	25th, 50th, 75th percentiles: 0.38, 0.45, 0.54	PM <sub>2.5</sub> , BC, NO <sub>2</sub> , O <sub>3</sub>
Park et al. (2005b)	Boston, MA (n = 4 97)	SDNN, LF, HF, LF/HF	Ambient	Mean: 0.50 (24 h) Range: 0.13-1.8	PM <sub>2.5</sub> , BC, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Riojas-Rodriguez et al. (2006)	Mexico City, Mexico (n = 30)	HF, LF, VLF, HR, R-R	Personal	Mean: 2.9 (11 h) Range: 0.1-18	PM <sub>2.5</sub>
Holguin et al. (2003)	Mexico City, Mexico (n = 34)	HF, Double check LF, LF/HF	Ambient	Mean: 3.3(24 h) Range: 1.8-4.8	PM <sub>2.5</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>

Study	Location (Sample Size)	Cardiac Endpoint	Exposure Assessment	Mean CO Level (ppm)	Copollutants
Wheeler et al. (2006)	Atlanta, GA (n = 30)	SDNN, r-MSSD, PNN, LF, HF, LF/HF	Ambient	Mean: 362 ppb (4h) 25th, 50th, 75th percentiles: 221.5, 304.3, 398.1	PM <sub>2.5</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Dales et al. (2004)	Toronto, Canada (n = 36)	SDNN, HF, LF, LF/HF, N-N	Personal	Mean: 2.4* Range: 0.4-16.5	PM <sub>2.5</sub>
Peters et al. (1999a)	Augsburg, Germany (n = 2681)	HR	Ambient	Mean: 3.6 Range: 1.5-7.1	TSP, SO <sub>2</sub>
Liao et al. (2004)	Maryland, North Carolina, Minnesota, (n = 4899-6784)	HR, SDNN, LF, HF	Ambient	Mean: 0.65 (24 h)	PM <sub>10</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Timonen et al. (2006)	Amsterdam, the Netherlands; Erfurt, Germany; Helsinki, Finland (n = 131)	SDNN, HF, LF/HF	Ambient	Mean: 0.35-0.52 Range: 0.09-2.17	PM <sub>2.5</sub> , PM <sub>10</sub> , NO <sub>2</sub>
Ibald-Mulli et al. (2004)	Amsterdam, the Netherlands; Erfurt, Germany; Helsinki, Finland (n = 131)	BP, HR	Ambient	Mean: 0.35-0.52 Range: 0.09-2.17	UFP, PM <sub>10</sub> , PM <sub>2.5</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Tarkiainen et al. (2003)	Kuopio, Finland (n = 6)	PNN, SDNN, r-MSSD	Personal	Mean: 4.6 Range: 0.5-27.4	None
Chan et al. (2005)	Taipei, Taiwan (n = 83)	SDNN, r-MSSD, LF	Ambient	Mean: 1.1 Range: 0.1-7.7	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Gold et al. (2005)	Boston, MA (n = 28)	ST-segment	Ambient	Median: 0.56 (12 h) Maximum: 1.04	PM <sub>2.5</sub> , BC

\*95th percentile of 24-h levels

## Arrhythmia

1 Cardiac arrhythmia refers to a broad group of conditions where there is irregular electrical activity  
2 in the heart. The main types of arrhythmias are fibrillation, tachycardia, and bradycardia, all of which can  
3 be associated with the upper (atria) and lower (ventricle) chambers of the heart. Briefly, fibrillation refers  
4 to when a chamber of the heart quivers chaotically rather than pumps in an orderly fashion, tachycardia  
5 refers to a rapid heart beat (e.g., >100 beats/minute) while bradycardia refers to a slow heart beat  
6 (e.g., <60 beats/minute). A few air pollution panel studies have examined the occurrence of cardiac  
7 arrhythmias by analyzing data recorded by implantable cardioverter defibrillators (ICDs) among cardiac  
8 patients. The majority of these studies were conducted in North America with the main outcome  
9 investigated being tachycardia. Results of these studies provide little evidence for an association between  
10 cardiac arrhythmia and ambient CO.

11 For example, Dockery and colleagues (2005) analyzed the relationship between ambient air  
12 pollution and the daily incidence of ventricular tachyarrhythmia among 203 patients with ICDs in Boston,  
13 MA. An hourly city average for the Boston metropolitan area was calculated for CO, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, SO<sub>4</sub><sup>2-</sup>,  
14 BC, and PM<sub>2.5</sub>. Although positive associations between ventricular arrhythmic episode days were found  
15 for all mean pollutant levels on the same day and previous days, none of these associations approached  
16 statistical significance. However, when the analyses were stratified by patients who had a previous

1 incidence of ventricular arrhythmia within 3 days, or greater than 3 days to the day of interest, a 0.5 ppm  
2 increase in 24-h CO concentration was positively associated with incidence of ventricular arrhythmia  
3 (OR: 1.68 [95% CI: 1.18-2.41]) among those who had a ventricular arrhythmia within the last 3 days.

4 A similar study in eastern Massachusetts examined cardiac arrhythmia by analyzing defibrillator  
5 discharges precipitated by either ventricular tachycardia or fibrillation among 100 cardiac patients (Peters  
6 et al., 2000b). Exposure to ambient CO was estimated for the same day, 1-day, 2-day, 3-day, and a 5-day  
7 mean lag period. Co was moderately correlated with PM<sub>10</sub> (r = 0.51) and PM<sub>2.5</sub> (r = 0.56) and more highly  
8 correlated with NO<sub>2</sub> (r = 0.71). When analyzing patients who had at least one defibrillator discharge  
9 (n = 33) there was no association with CO. However, when analyzing patients who had at least 10  
10 discharges (n = 6), a 0.5 ppm increase in 24-h CO concentration (lag 0-4) was associated with an  
11 increased odds of a defibrillator discharge (OR: 1.66 [95% CI: 1.01-2.76]).

12 In contrast, other air pollution panel studies conducted in St Louis, MO (among 56 subjects) (Rich  
13 et al., 2006b), Atlanta, GA (among 518 subjects) (Metzger et al., 2007), Boston, MA (among  
14 203 subjects) (Rich et al., 2005), and Vancouver, Canada (Rich et al., 2004; Vedal et al., 2004) (among 34  
15 and 50 subjects respectively) did not find an association between short term changes in ambient CO and  
16 occurrence of cardiac arrhythmia in patients with implantable defibrillators. The study in Boston also  
17 examined atrial fibrillation episodes among the same group of subjects and also did not find an  
18 association with ambient CO (Rich et al., 2005).

19 An alternative method used to assess the relationship between cardiac arrhythmia and ambient air  
20 pollution is to analyze cardiac data recorded via ECG. Two studies have employed this method and  
21 reported inconsistent results. A study in Steubenville, OH, which is located in an industrial area, examined  
22 weekly ECG data among 32 non-smoking older adults for 24 weeks during summer and fall (Sarnat et al.,  
23 2006). Ambient exposures for up to 5 days prior to the health assessment (based on a 5-day moving  
24 average) were calculated for PM<sub>2.5</sub>, SO<sub>4</sub><sup>2-</sup>, elemental carbon (EC), O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and CO from data  
25 recorded at one central monitoring site. Increases in ambient CO were not associated with increased odds  
26 of having at least one arrhythmia during the study period.

27 In contrast, a study in Germany examined the relationship between ambient air pollution and the  
28 occurrence of supraventricular (atria) and ventricular tachycardia recorded via monthly 24-h ECGs among  
29 57 subjects over a 6 month period (Berger et al., 2006). Exposure estimates were calculated for ambient  
30 ultrafine particles, PM<sub>2.5</sub>, CO, NO, NO<sub>2</sub>, and SO<sub>2</sub> for various lag periods (0-23 h, 24-47 h, 48-71 h,  
31 72-95 h, and 5-day avg) prior to the ECG. Results showed that a 0.5 ppm increase in ambient 24-h CO  
32 concentration (lag 0-4 days prior to ECG) was positively associated with the occurrence of  
33 supraventricular tachycardia (OR: 1.36 [95% CI: 1.08-1.74]). However, ambient CO was not associated  
34 with ventricular tachycardia.

35 In summary, few studies have examined associations between CO and the occurrence of cardiac  
36 arrhythmias, and these studies provided little evidence of a CO effect on cardiac arrhythmias. While more

1 studies analyzed data from ICDs, very few reported significant associations. This was similar for the  
 2 mixed results from the two studies that analyzed ECG data to evaluate cardiac arrhythmias in association  
 3 with CO exposures. Table 5.5 summarizes the reviewed studies.

**Table 5.5. Summary of studies investigating the effect of CO exposure on cardiac arrhythmias.**

Study	Location, Sample Size	Cardiac Endpoint	Exposure Assessment	Mean CO Level (ppm)	Copollutants
<i>ARRHYTHMIAS (AMONG PATIENTS WITH ICDs)</i>					
Dockery et al. (2005)	Boston, MA (n = 203)	Ventricular Tachycardia	Ambient	25th, 50th, 75th, 95th, percentiles: 0.53, 0.80, 1.02, 1.37 (2-day)	PM <sub>2.5</sub> , BC, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub> , SO <sub>4</sub> <sup>2-</sup>
Peters et al. (2000b)	Massachusetts, (n = 100)	Ventricular fibrillation or tachycardia	Ambient	Mean: 0.58 (24 h) Max: 1.66	PM <sub>2.5</sub> , PM <sub>10</sub> , BC, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub> , SO <sub>4</sub> <sup>2-</sup>
Rich et al. (2006b)	Boston, MA (n = 56)	Ventricular arrhythmia	Ambient	25th, 50th, 75th percentiles: 0.4, 0.5, 0.6 (24 h)	PM <sub>2.5</sub> , EC, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Metzger et al. (2007)	Atlanta, GA (n = 518)	Ventricular Tachycardia	Ambient	Mean: 1.7 (1 h) Range: 0.1-7.7	PM <sub>10</sub> , PM <sub>2.5</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Rich et al. (2005)	Boston, MA (n = 203)	Atrial fibrillation	Ambient	25th, 50th, 75th, 95th, percentiles: 0.53, 0.80, 1.02, 1.37 (2-day)	PM <sub>2.5</sub> , BC, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Rich et al. (2004)	Vancouver, Canada (n = 34)	ICD discharge due to arrhythmia	Ambient	Mean: 0.55 (24 h) IQR: 0.16	PM <sub>2.5</sub> , PM <sub>10</sub> , EC, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub> , SO <sub>4</sub>
Vedal et al. (2004)	Vancouver, Canada (n = 50)	ICD discharge due to arrhythmia	Ambient	Mean: 0.6 (24 h) Range: 0.3-1.6	PM <sub>10</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
<i>ARRHYTHMIAS (VIA ECG)</i>					
Sarnat et al. (2006)	Steubenville, OH (n = 32)	Atrial or ventricular tachycardia	Ambient	Mean: 0.2 (24 h) Range: 0.1, 1.5	PM <sub>2.5</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub> , SO <sub>4</sub> , EC
Berger et al. (2006)	Erfurt, Germany (n = 57)	Atrial or ventricular tachycardia	Ambient	Mean: 0.45 (24 h) Min, Med, Max 0.10, 0.38, 1.68	PM <sub>10</sub> , PM <sub>2.5</sub> , NO <sub>2</sub> , NO, SO <sub>2</sub> , UF

## Cardiac Arrest

4 Cardiac arrest refers to the abrupt loss of heart function due to failure of the heart to contract  
 5 effectively during systole, which can lead to sudden cardiac death if not treated immediately. Very few  
 6 studies have investigated the association between ambient CO exposure and the risk of cardiac arrest and  
 7 none reported a significant link between increased CO exposure and the occurrence of cardiac arrest.



1 Two similar studies were conducted in Seattle, WA, and both did not report an association between  
2 ambient CO and cardiac arrest. Both studies employed a case-crossover design and examined air pollution  
3 exposures for black smoke particles (BSP), PM<sub>10</sub>, SO<sub>2</sub>, and CO. The correlation coefficient for PM<sub>10</sub> and  
4 CO was 0.81. The first of these studies examined paramedic-attended out-of-hospital primary cardiac  
5 arrests among 362 cases (between 1998-1994) in Seattle and King County, WA whereby lags of 0-5 days  
6 were analyzed (Levy et al., 2001). The second of these studies examined out-of-hospital primary cardiac  
7 arrest for a ten year period (1985-1994) among subjects within a health organization database (the Group  
8 Health Cooperative of Puget Sound) whereby 0-day through 2-day lags were analyzed (Sullivan et al.,  
9 2003).

### **Myocardial Infarction**

10 As previously stated, MI is commonly referred to as ‘heart attack’ and is another cardiac outcome  
11 that has received limited attention within the area of air pollution research. Only one study has  
12 investigated the association between short-term changes in ambient CO and the onset of MI. Peters and  
13 colleagues (2001) employed a case-crossover study design to analyze short term exposures (0-5 h and 0-  
14 5 days before the onset of MI) to particles (PM<sub>10</sub>, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, BC) and gases (CO, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>)  
15 among 772 patients with MI in the greater Boston area. While all pollutants showed positive associations  
16 with the onset of MI, only PM<sub>2.5</sub> reached statistical significance with the main exposure period being 2 h  
17 before the onset.

### **Blood Pressure**

18 Only two studies have investigated whether short-term ambient CO influences BP. The earlier of  
19 these two studies examined BP among 2607 men and women aged 25-64 who participated in the  
20 Augsburg, Germany MONICA study (Ibald-Mulli et al., 2001). Exposures to ambient TSP, SO<sub>2</sub> and CO  
21 (from one monitor in the center of the city) during the same day as the BP reading and an average over the  
22 5 days prior were examined. Results showed that ambient CO had no association with BP.

23 Similarly, the second of these studies extracted baseline and repeated-measures of cardiac  
24 rehabilitation data from a Boston, MA hospital for 62 subjects with 631 visits and analyzed ambient air  
25 pollution exposures (with particular focus on PM<sub>2.5</sub>) averaged over various periods up to 5 days before the  
26 visit (Zanobetti et al., 2004b). While results showed significant associations between increased BP and  
27 ambient PM<sub>2.5</sub>, SO<sub>2</sub>, O<sub>3</sub>, and BC, there was no significant effect for CO.

### **Blood Markers of Coagulation and Inflammation**

28 Several studies have investigated the association between ambient CO and various blood markers  
29 related to coagulation and inflammation. The main endpoints analyzed have been plasma fibrinogen,

1 Factor VII, C-reactive protein (CRP), prothrombin, intercellular adhesion molecule (ICAM-1), and white  
2 blood cell count (WBC).

3 Pekkanen et al. (2000) examined the association between daily concentrations of air pollution and  
4 concentrations of plasma fibrinogen measured among 4982 male and 2223 female office workers in  
5 Whitehall, London, U.K. between September 1991 and May 1993. Plasma fibrinogen data were linked to  
6 ambient exposure to BS, PM<sub>10</sub>, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and CO, where the exposures were derived from data  
7 recorded at 5 fixed sites across London. There was a high correlation between levels of CO and NO<sub>2</sub>  
8 ( $r = 0.81$ ) and more moderate correlations with PM<sub>10</sub> ( $r = 0.57$ ) and SO<sub>2</sub> ( $r = 0.61$ ). The pollution data on  
9 the same day when the blood sampling was done (lag 0) and on the 3 previous days (lags 1-3) were  
10 analyzed. Results showed that ambient CO at all lags was significantly associated with an increase in  
11 plasma fibrinogen. Results were similar for NO<sub>2</sub> while all other pollutants were not associated with an  
12 increase in plasma fibrinogen.

13 Liao et al. (2005) examined associations between various air pollutants and hemostatic and  
14 inflammatory markers (fibrinogen, factor VIII-C, von Willebrand factor, serum albumin, WBC) among  
15 10,208 middle-aged males and females from the ARIC study. Exposure estimates for ambient PM<sub>10</sub>, NO<sub>2</sub>,  
16 SO<sub>2</sub>, O<sub>3</sub> and CO were calculated for days 1-3 prior to the blood sampling. A 0.5 ppm increment in 24-h  
17 CO concentration was significantly associated with 0.015 g/dL decrease in serum albumin among persons  
18 with a history of CVD. CO was not associated with other hemostatic or inflammatory factors.

19 In Israel, Steinvil et al. (2008) examined WBC, fibrinogen, and CRP among 3,659 study subjects  
20 enrolled in the Tel-Aviv Sourasky Medical Center inflammation survey, in which subjects lived <11km  
21 from an ambient air pollution monitor. Air pollution data in the form of PM<sub>10</sub>, NO<sub>2</sub>, SO<sub>2</sub>, O<sub>3</sub>, and CO were  
22 derived from data recorded at fixed sites. The correlations coefficients were high between CO and NO<sub>2</sub>  
23 ( $r = 0.86$ ) and PM<sub>10</sub> ( $r = 0.75$ ). Exposures for lag days 1-7 were analyzed and ambient CO had a  
24 significant negative effect on fibrinogen only among males. Significant associations were reported for lag  
25 0 (e.g., same day) and lags 2-5 with the decrease in fibrinogen ranging from -5.5 mg/dL to -9.8 mg/dL per  
26 0.5 ppm increase in 24-h CO concentration. A similar negative effect for CO was observed on WBC  
27 among males only. The average CO exposure over the week prior to the sampling yielded the largest  
28 reduction in WBC (-263 cells/ $\mu$ L).

29 In a German study, R ckerl and colleagues (2006) recruited 57 non-smoking male patients with  
30 CHD who were scheduled for 12 subsequent clinical visits where samples of blood were collected. The  
31 authors tested the primary hypothesis that CRP would increase in association with a rise in air pollution  
32 levels. CRP is an acute phase protein that increases during inflammatory processes in the body. Other  
33 markers of inflammation (serum amyloid A [SAA]), cell adhesion (E-selectin, von Willebrand factor  
34 antigen [vWF], ICAM-1), and coagulation (fibrinogen, factor VII [FVII], prothrombin fragment 1+2)  
35 were also examined. Ambient air pollution in the form of PM<sub>10</sub>, ultrafine particles (UFP), EC, NO<sub>2</sub>, and  
36 CO was monitored at one central site and a 24-h avg immediately preceding the clinic visit (lag 0) and up

1 to 5 days (lags 1-4) was calculated for each patient. For CRP, the odds of observing CRP concentrations  
2 above the 90th percentile were 2.41 (95% CI: 1.23-5.02) in association with a 0.5 ppm increase in 24-h  
3 CO concentration (lag 2). CO concentration during lags 1 and 2 was associated with observing ICAM-1  
4 concentrations above the 90th percentile (OR: 2.41 [95% CI: 1.49-4.04]; OR: 3.17 [95% CI: 1.77-6.11],  
5 respectively). CO concentration during lags 0-3 was associated with a decrease in FVII.

6 A similar study by Ruckerl and colleagues (2007) was conducted among 1,003 MI survivors across  
7 6 European cities (Athens, Greece; Augsburg, Germany; Barcelona, Spain; Helsinki, Finland; Rome,  
8 Italy; Stockholm, Sweden). The study compared repeated measurements of interleukin-6 (IL-6), CRP and  
9 fibrinogen with concurrent ambient levels of air pollution (particle number count [PNC], PM<sub>10</sub>, PM<sub>2.5</sub>,  
10 NO, NO<sub>2</sub>, O<sub>3</sub>, SO<sub>2</sub>, CO) from fixed sites across each city. Lags 0-1 and the 5-day mean prior to the blood  
11 sampling were analyzed and ambient CO was not associated with any of the inflammatory endpoints.

12 Baccarelli et al. (2007) recruited 1,218 healthy individuals from the Lombardia region in Italy and  
13 assessed whether blood coagulability is associated with ambient air pollution. The main blood  
14 coagulability endpoints of interest were prothrombin time (PT) and activated partial thromboplastin time  
15 (APTT), which are measures of the quality of the coagulation pathways, assuming that, if shortened these  
16 measures would reflect hypercoagulability. Air pollution data (PM<sub>10</sub>, O<sub>3</sub>, NO<sub>2</sub>, and CO) were obtained  
17 from 53 fixed stations across the Lombardia region, which was divided into nine different study areas and  
18 a network average for each pollutant was calculated across the available monitors within each of the nine  
19 study areas. The analyses examined air pollution at the time of the blood sampling as well as averages for  
20 the 7 days prior and 30 days prior. Results showed that ambient CO at the time of blood sampling was  
21 associated with a decrease in PT (coefficient = -0.11 [95% CI: -0.18 to -0.05, p <0.001), indicating  
22 hypercoagulability. However, PM<sub>10</sub> and NO<sub>2</sub> at the time of blood sampling were also associated with a  
23 decrease in PT and results from multipollutant models were not reported. Acute phase reactants such as  
24 fibrinogen, and naturally occurring anticoagulants such as antithrombin, protein C and protein S were  
25 examined and none were associated with ambient air pollution.

26 In summary, despite the small number of studies, there was some evidence of a significant link  
27 between CO exposure and blood markers of coagulation and inflammation. Table 5.6 summarizes the  
28 reviewed studies.

**Table 5-6. Summary of studies investigating the effect of CO exposure on blood markers of coagulation and inflammation.**

Study	Location, Sample Size	Cardiac Endpoint	Exposure Assessment	Mean CO Level (ppm)	Copollutants
Pekkanen et al (2000)	London, U.K. (n = 7205)	Plasma fibrinogen	Ambient	Mean: 1.22 (24 h) 10th, 50th, 90th, Max: 0.61, 1.04, 2.0, 8.61	PM <sub>10</sub> , BS, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Liao et al (2005)	USA (n = 10,208)	Fibrinogen, VII-C, WBC, albumin, vWF	Ambient	Mean: 1.4 (24 h)	PM <sub>10</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Steinvil et al (2008)	Tel-Aviv, Israel (n = 3659)	CRP, fibrinogen, WBC	Ambient	Mean: 0.8 25th, 50th, 75th percentiles: 0.7, 0.8, 1.0	PM <sub>10</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Ruckerl et al (2006)	Erfurt, Germany (n = 57)	CRP, SAA, cell adhesions and coagulation	Ambient	Mean: 0.45 (24 h) Range: 0.10, 1.68	PM <sub>10</sub> , PM <sub>2.5</sub> , UFP, EC, NO <sub>2</sub>
Ruckerl et al (2007)	Six European cities (n = 1003)	IL-6, CRP, fibrinogen	Ambient	Mean (24 h): 0.29-1.48	PM <sub>10</sub> , PM <sub>2.5</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>
Baccarelli et al (2007)	Lombardia region, Italy (n = 1218)	PT, APTT, fibrinogen, anticoagulants	Ambient	Mean: 1.14-3.11 Max: 5.52-11.43	PM <sub>10</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub>

## Hospital Admissions and Emergency Department Visits

1 Since the 2000 CO AQCD there have been a number of studies investigating the effect of ambient  
 2 CO on hospital admissions and ED visits for cardiovascular diseases. Some of these studies have focused  
 3 solely on one specific CVD outcome, and these studies are discussed first. This is followed by a  
 4 discussion of studies that investigated admissions for all CVD outcomes (e.g., non-specific) or a variety  
 5 of specific CVD outcomes.

### *Ischemic Heart Disease, Myocardial Infarction, and Angina Pectoris*

6 Ischemic heart disease (IHD), also known as coronary heart disease (CHD), is caused by  
 7 inadequate circulation of the blood to the heart muscle, which is a result of the heart arteries being  
 8 blocked by cholesterol deposits. IHD can lead to sudden episodes such as MI (“heart attack”) or death, as  
 9 well as chronic conditions such as angina pectoris (chest pain).

10 **Ischemic Heart Disease.** Very few studies have focused directly on hospitalizations for IHD.  
 11 There is a lot of variation among these studies with regard to methods employed and results reported. It  
 12 should be noted that within these studies IHD included MI and angina pectoris (ICD-9 codes 410-414;  
 13 ICD-10 codes I20, I21-I23, I24). Mann and colleagues (2002) investigated the modifying effect of  
 14 secondary diagnosis of arrhythmia and congestive heart failure (CHF) on the relationship between  
 15 hospital admissions for IHD (ICD-9: 410-414) and ambient air pollutants for the period of 1988 to 1995  
 16 in southern California. There were 54,863 visits analyzed and a 0.75 ppm increase in 8-h max CO

1 concentration was associated with a 2.69% (95% CI: 1.21-4.19) increase in same-day IHD admissions  
2 among persons with a secondary diagnosis of CHF, a 2.23% (95% CI: 1.35-3.13) increase among persons  
3 with a secondary diagnosis of arrhythmia, and a 1.21% (95% CI: 0.49-1.94) among persons without either  
4 secondary diagnosis. Of all pollutants examined (PM<sub>10</sub>, NO<sub>2</sub>, O<sub>3</sub>, CO), only NO<sub>2</sub> showed similar positive  
5 effects to CO and no multipollutant models were analyzed. The correlation coefficients between CO and  
6 NO<sub>2</sub> ranged from 0.64 to 0.86 across the seven regions. This study indicated that people with IHD and  
7 accompanying CHF and /or arrhythmia are a sensitive group in relation to the effects of ambient air  
8 pollution.

9 By using a time-series approach, ED visits for IHD (ICD-9: 410-414) in Montreal, Canada  
10 (1997-2002) were examined in relation to ambient CO concentrations (lags 0 and 1) (Szyszkowicz, 2007).  
11 A total of 4,979 visits were analyzed and results showed significant positive effects with a 0.5 ppm  
12 increase in 24-h CO concentration (lag 0) attributing to a 14.1% (95% CI: 5.8-20.6) increase in daily ED  
13 visits among all patients. Stratified analyses showed that this effect was mostly among male patients  
14 (19.8% [95% CI: 9.2-31.6]). NO<sub>2</sub> was the only other pollutant examined, and it too showed significant  
15 positive associations with ED visits for IHD for same-day exposure; however, no multipollutant models  
16 were examined.

17 Lee and colleagues (2003b) examined daily counts of hospital admissions for IHD in Seoul, Korea  
18 for the period from December 1997 to December 1999. Single-day lags 0 to 5 were analyzed and the lag  
19 period with the strongest association for each pollutant was chosen. For CO, lag 5 showed the strongest  
20 effect with a 1 ppm increase in 1-h maximum (max) CO concentration associated with a daily increase in  
21 the number of hospital admissions for IHD; however, this was only among patients 64+ years of age (RR:  
22 1.07 [95% CI: 1.01-1.13]). All other pollutants (PM<sub>10</sub>, O<sub>3</sub>, NO<sub>2</sub>) except SO<sub>2</sub> showed similar significant  
23 effects and in a two-pollutant model with PM<sub>10</sub> the CO effect attenuated toward the null.

24 Other studies have examined hospital admissions for IHD while investigating a broad group of  
25 CVD outcomes. A study was conducted in Atlanta, GA, where over 4 million ED visits from 31 hospitals  
26 for the period 1993 to 2000 were analyzed (Study of Particles and Health in Atlanta [SOPHIA]). Several  
27 articles have been published from this research with two examining cardiovascular admissions in relation  
28 to CO concentrations. The first of these (Metzger et al., 2004a) used a time-series design and analyzed a  
29 3-day moving average over single-day lags 0-2 as the a priori lag structure. Although of borderline  
30 statistical significance, CO was positively associated with an increase in ED visits for IHD (RR 1.016  
31 [95% CI: 0.999-1.034] per 1 ppm increase in 1-h max CO concentration).

32 The second of these reports (Peel et al., 2007) examined the association of ambient air pollution  
33 levels and cardiovascular morbidity in visits with and without specific secondary conditions (e.g., co-  
34 morbidity). Within a time-stratified case-crossover design using the same lag structure already mentioned,  
35 the main results showed that a 1 ppm increase in 1-h max CO concentration was associated with an

1 increase in IHD among those without diabetes (OR: 1.023 [95% CI: 1.004-1.042]), and without CHF  
2 (OR: 1.024 [95% CI: 1.006-1.042]).

3 Two Australian studies have also examined associations between ambient CO concentrations and  
4 increased hospital admissions for various CVD outcomes. The first of these studies (Barnett et al., 2006)  
5 analyzed data from 5 of the largest cities in Australia (Brisbane, Canberra, Melbourne, Perth, Sydney) and  
6 two New Zealand cities (Auckland, Christchurch) for the period 1998-2001. A time-stratified case-  
7 crossover design was employed and the age groups of 15-64 years and  $\geq 65$  years were analyzed for the  
8 0-1 lag period. The pooled estimates across all cities showed that a 0.75 ppm increase in 8-h max CO  
9 concentration was associated with a 1.9% (95% CI: 0.7-3.2) increase in admissions for IHD, but only  
10 among the elderly group ( $\geq 65$  years).

11 The second of the Australian studies (Jalaludin et al., 2006) examined ED visits for CVD outcomes  
12 in the elderly (65+ years) in Sydney for the period 1997 to 2001. Using a time-series approach, single-day  
13 lags of 0, 1, 2, 3 and an average over lags 0 and 1 were examined. A 0.75 ppm increase in 8-h max CO  
14 concentration (lag 0) was associated with increases in IHD emergency department visits of 3.1%  
15 (95% CI: 1.3-4.9).

16 **Myocardial Infarction.** Linn et al. (2000) examined the association between ambient air pollution  
17 and hospital admissions for cardiopulmonary illnesses in metropolitan Los Angeles for the years  
18 1992-1995. Using a time-series approach, a 0.5 ppm increase in same-day 24-h avg CO concentration was  
19 associated with a 2.0% increase in MI hospital admissions among people aged  $>30$  years. When the  
20 analyses were stratified by season, no significant effects were observed (No quantitative seasonal effects  
21 reported). A time-series study in Denver, Colorado, investigated daily hospital admissions for various  
22 CVD outcomes among older adults ( $>65$  years) across 11 hospitals (Koken et al., 2003). Data between  
23 July and August for the period 1993-1997 were analyzed. Single-day lags 0 to 4 were examined and CO  
24 showed no association with hospital admissions for MI (quantitative results were not reported).

25 As part of the HEAPSS (Health Effects of Air Pollution among Susceptible Subpopulations) study,  
26 Lanki et al. (2006) investigated the association between traffic-related exposure to air pollutants and  
27 hospitalization for first acute myocardial infarction (AMI). Data were collected from five European cities  
28 with either AMI registers (Augsburg, Barcelona), or hospital discharge registers (Helsinki, Rome,  
29 Stockholm). Correlation coefficients between CO and NO<sub>2</sub> ranged from 0.43 to 0.75 across the five cities,  
30 and for PM<sub>10</sub> the range was 0.21 to 0.56. A total of 26,854 hospitalizations were analyzed and pooled  
31 estimates from all 5 cities showed that there was a weak positive association with AMI hospitalizations  
32 and 24-h avg CO concentrations at lag 0 (RR: 1.014 [95% CI: 1.000-1.029] per 0.5 ppm increase), but  
33 more so when only using data from the three cities (Helsinki, Rome, Stockholm) with hospital discharge  
34 registers (RR: 1.020 [95% CI: 1.003-1.035] per 0.5 ppm increase). When analyses were stratified by  
35 fatality and age, results showed that the CO effect was significantly associated with fatal AMI among the

1 <75-year age group (RR: 1.080 [95% CI: 1.017-1.144), and with non-fatal AMI in the  $\geq$  75-year age  
2 group (RR: 1.044 [95% CI: 1.011-1.076).

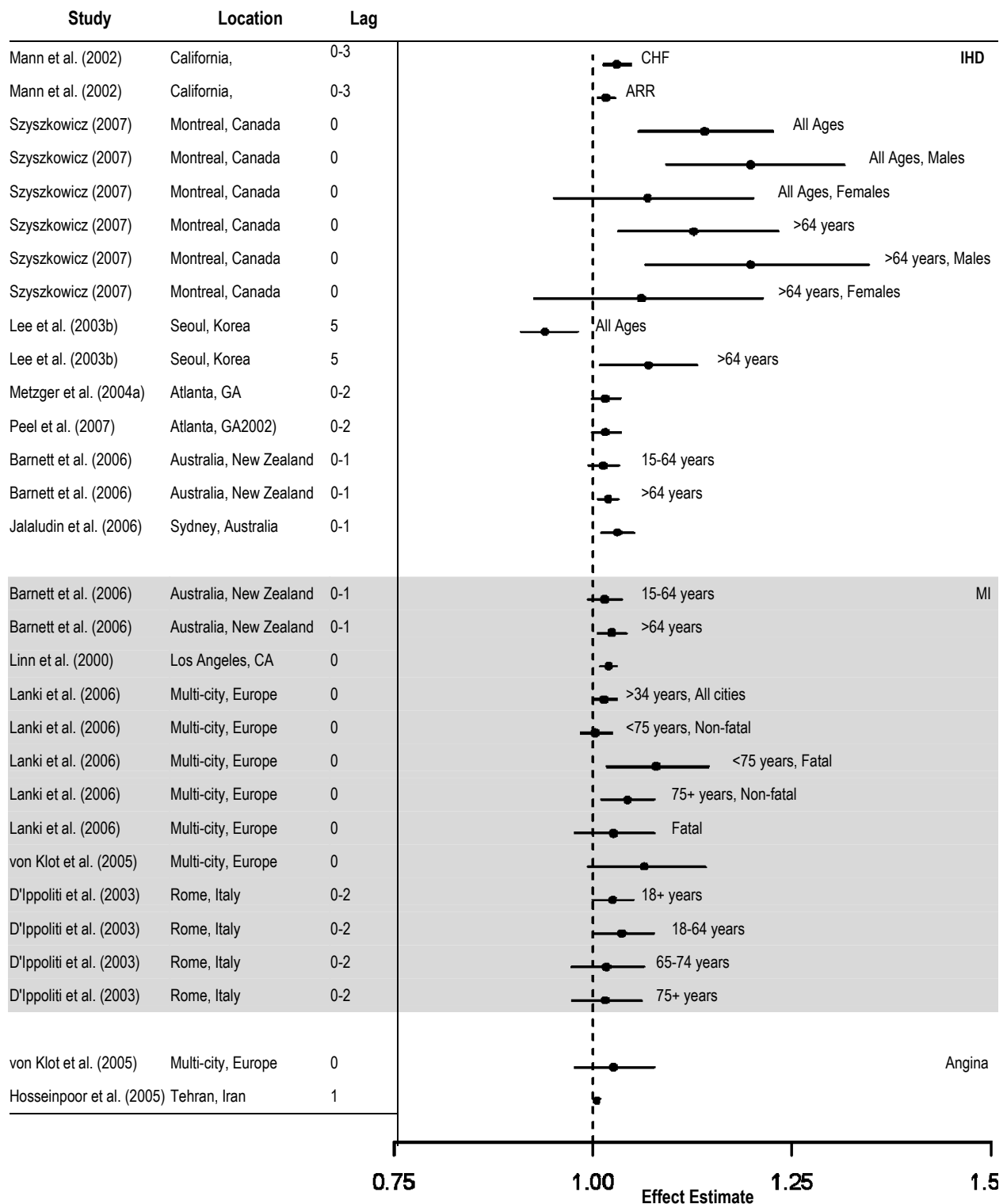
3 Further analyses within the HEAPSS cohort was conducted where the event of cardiac readmission  
4 among the first MI survivors (n = 22, 006) was linked to ambient air pollution (von Klot et al., 2005). The  
5 readmissions of interest were those with primary diagnosis of AMI, angina pectoris, dysrhythmia, and  
6 heart failure that occurred at least 29 days after the index event. Single-day lags 0 to 3 were examined and  
7 pooled estimates from all 5 cities showed that a 0.5 ppm increase in same-day (lag 0) CO was associated  
8 with an increase in cardiac (e.g., any of the diagnoses) readmissions (RR: 1.041 [95% CI: 1.003-1.076])  
9 and this persisted in two-pollutant models that included either PM<sub>10</sub> or O<sub>3</sub>. Correlation coefficients with  
10 CO ranged from 0.21 to 0.57 for PM<sub>10</sub> and 0.44 to 0.75 for NO<sub>2</sub>.

11 A study in Rome, Italy, also found an association between ambient CO and hospitalizations for first  
12 episode MI among 6,531 subjects (January 1995-June 1997) (D'Ippoliti et al., 2003). A case-crossover  
13 design with stratification of time into separate months was used to select referent days as the days falling  
14 on the same day of the week within the same month as the index day. CO concentration was positively  
15 associated for lag 2 (OR: 1.019 [95% CI: 1.001-1.037]). The other pollutants analyzed were NO<sub>2</sub> and TSP,  
16 both of which exhibited a significant positive effect at lag 0. TSP also showed a significant positive effect  
17 at lag 0-2 and when entered into a model with CO, the CO effect did not persist.

18 The previously mentioned Australian and New Zealand study that analyzed data from seven cities  
19 (Brisbane, Canberra, Melbourne, Perth, Sydney, Auckland, and Christchurch) for the period 1998-2001  
20 also reported an association between CO and MI hospitalization (Barnett et al., 2006). The pooled  
21 estimates across all cities showed that a 0.75 ppm increase in 8-h max CO concentration was associated  
22 with a 2.4% (95% CI: 0.6-4.1) increase in admissions for MI, but only among older adults ( $\geq$  65 years).

23 **Angina Pectoris.** In the current literature, only one study was identified that focused solely on  
24 angina pectoris as an endpoint. Admissions data for angina pectoris were collected from 25 academic  
25 hospitals in Tehran, Iran, and linked to ambient air pollution for the period of 1996 to 2001 (Hosseinpoor  
26 et al., 2005). Using a time-series approach, single-day lags of 0 to 3 were analyzed and a 0.5 ppm increase  
27 in 24-h avg CO concentration at lag 1 was associated with increased hospital admissions for angina (OR:  
28 1.005 [95% CI: 1.003-1.007]). This result persisted in a multipollutant model that also included NO<sub>2</sub>,  
29 PM<sub>10</sub>, and O<sub>3</sub> with CO being the only significant pollutant (OR: 1.005 [95% CI: 1.001-1.008]). Figure 5-1  
30 shows the effect estimates associated with daily admissions for various forms of IHD from selected  
31 studies. Table 5.7 shows a summary of the IHD hospital admission studies that examined CO exposures.

32 In summary, the majority of studies reported significant increases in the daily number of  
33 admissions for IHD and MI in relation to CO exposures. In studies that stratified by age groups and/or  
34 sex, the effects were larger among the elderly and males. Among the different lag periods being  
35 examined, the associations were more commonly observed with same day CO (lag 0) or an average over  
36 the same day and previous day (lag 0-1).



**Figure 5-1.** Summary of effect estimates (95% confidence intervals) associated with hospital admissions for various forms of IHD. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.



**Table 5-7. Summary of IHD hospital admission studies.<sup>1</sup>**

Study	Location	Endpoints Examined	Copollutants	Lags Examined	CO Levels (ppm)
<i>STUDIES THAT FOCUSED SOLELY ON IHD, MI, OR ANGINA</i>					
Mann et al. (2002)	Southern California (1988-1995)	IHD	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	0,1,2, 2-4ma	Mean: 2.07 (8h)
Szyszkowicz (2007)	Montreal, Canada (1997-2002)	IHD	NO <sub>2</sub>	0,1	Mean: 0.5 (24 h)
Lee et al. (2003b)	Seoul, Korea (1997-1999)	IHD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3,4,5	Mean: 1.8
Lanki et al. (2006) <sup>2</sup>	5 European cities (1992-2000)	MI (first acute)	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub> , PNC	0,1,2,3	Highest city was Rome. 25th = 1.5 75th = 2.9 mg/m <sup>3</sup>
von Klot et al. (2005) <sup>2</sup>	5 European cities (1992-2001)	MI, Angina, Cardiac*	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub> , PNC	0,1,2,3	Mean: highest city was Rome: 1.9 (24 h)
D'Ippoliti et al. (2003) <sup>2</sup>	Rome, Italy (1995-1997)	MI	TSP, NO <sub>2</sub> , SO <sub>2</sub>	0,1,2,3,4, 0-2	Mean: 3.8 (24 h)
Hosseinpoor et al. (2005) <sup>2</sup>	Tehran, Iran (1996-2001)	Angina	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3	Mean: 9.4 (24 h)
<i>STUDIES THAT EXAMINED IHD, MI, OR ANGINA AMONG OTHER CVDS</i>					
Metzger et al. (2004a)	Atlanta, GA (1993-2000)	IHD, All CVD, CD, CHF, PVCD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2ma	Mean: 1.5 (1 h)
Peel et al. (2007)	Atlanta, GA (1993-2000)	IHD, All CVD, CD, CHF, PVCD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2ma	Mean: 1.5 (1 h)
Barnett et al. (2006)	Australia and New Zealand (1998-2001)	IHD, MI, All CVD, CA, Stroke	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	Lag 0-1	Mean: (8 h) 0.5- 2.1
Jalaludin et al. (2006)	Sydney, Australia (1997-2001)	IHD, All CVD, Stroke, Cardiac	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3, 0-1	Mean: 0.82 (8 h)
Linn et al. (2000)	Los Angeles, CA (1992-1995)	MI, All CVD, CHF, CA, OS	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	0	Mean: (24 h) Winter 1.7, Spring 1.0, Summer 1.2, Fall 2.1
Koken et al. (2003)	Denver, CO (1993-1997)	MI, CAth, PHD, CD, CHF	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3,4	Mean: 0.9 ppm (24 h)

<sup>1</sup>Cardiac = AMI, angina, dysrhythmia, or HF; CA = Cardiac arrhythmia; CAth = Cardiac atherosclerosis; CD = cardiac dysrhythmias; CHF = Congestive heart failure; PHD = Pulmonary heart disease; OS = Occlusive stroke; PVCD = peripheral vascular and cerebrovascular disease, ma = moving average.

<sup>2</sup>These studies presented CO concentrations in the units mg/m<sup>3</sup>. The concentrations were converted to ppm using the conversion factor 1 ppm = 1.15 mg/m<sup>3</sup>, which assumes standard atmosphere and temperature.

## Stroke

1 A stroke is the result of either the blood supply to the brain being blocked (e.g., embolism), which  
2 refers to an ischemic stroke (80% of strokes), or the occurrence of a burst blood vessel or hemorrhaging,  
3 referred to as a hemorrhagic stroke. Hemorrhagic stroke has two main groupings; intracerebral  
4 hemorrhagic stroke (10% of strokes), which is when a blood vessel in the brain leaks, and subarachnoid  
5 hemorrhage (3% of strokes), which is bleeding under the outer membranes of the brain. The third type of  
6 stroke is a transient ischemic attack (TIA), or mini-stroke, which has the same early symptoms as a  
7 normal stroke but the symptoms disappear within 24 hours, leaving no apparent deficits.

8 A small number of air pollution studies have investigated hospital admissions for the three main  
9 forms of stroke with the majority reporting positive associations with ambient CO and lag periods  
10 between 0 and 3 days.

11 A U.S. study across 9 cities investigated hospital admissions for ischemic and hemorrhagic stroke  
12 among Medicare beneficiaries aged 65+ years of age (155,503 ischemic and 19,314 hemorrhagic  
13 admissions from the ED) (Wellenius et al., 2005a). Single-day lags 0 to 2 were examined and based on a  
14 pooled estimate, same-day CO (lag 0) was associated with an increase in admissions of 1.98%  
15 (95% CI: 0.86-3.12) per 0.5 ppm increase in 24-h CO concentration) for ischemic stroke admissions but  
16 not hemorrhagic stroke admissions (-1.14%, 95% CI: -3.40 to 1.18). All other pollutants examined (PM<sub>10</sub>,  
17 NO<sub>2</sub>, SO<sub>2</sub>) were associated with an increase in ischemic stroke admissions, but not hemorrhagic stroke  
18 admissions.

19 Villeneuve and colleagues (2006a) studied ED visits for hemorrhagic strokes, acute ischemic  
20 strokes and transient ischemic attacks among individuals 65+ years of age at 5 hospitals within the  
21 Edmonton area in Canada between April 1992 and March 2002 (12,422 visits). Within a time-stratified  
22 case-crossover design the analyses were stratified by two seasonal groups (October-March and April-  
23 September) and CO only had an effect on ischemic stroke during April-September. A 0.5 ppm increase the  
24 CO concentration for a 3-day avg across lags 0-2 was associated with a 32% increase in risk (OR: 1.32  
25 [95% CI 1.09-1.60]). CO had no effect on any other stroke subtype. In two-pollutant models the CO effect  
26 on ischemic stroke persisted after controlling for PM<sub>10</sub>, PM<sub>2.5</sub>, SO<sub>2</sub>, and O<sub>3</sub>. When all seasons and all  
27 strokes were combined there was no statistically significant association between all the pollutants  
28 examined and increased admissions for stroke.

29 In Kaohsiung City, Taiwan, CO averaged over lags 0-2 was associated with increased admissions  
30 for stroke across 63 hospitals (Tsai et al., 2003b). From 1997 to 2000 a total of 23,179 admissions were  
31 analyzed and on warm days ( $\geq 20^{\circ}\text{C}$ ) the odds ratios for primary intracerebral hemorrhage and ischemic  
32 stroke were 1.39 (95% CI: 1.16-1.66) and 1.39 (95% CI: 1.25-1.53) respectively for a 0.5 ppm increase in  
33 24-h CO concentration. For the same increase in CO on cool days ( $<20^{\circ}\text{C}$ ) the odds ratios were 1.33  
34 (95% CI: 0.38-2.55) for intracerebral hemorrhage and 2.68 (95% CI: 1.59-4.49) for ischemic stroke.

1 These results persisted in two-pollutant models that included PM<sub>10</sub>, SO<sub>2</sub>, and O<sub>3</sub>, but did not persist when  
2 controlling for NO<sub>2</sub>.

3 Earlier research conducted in metropolitan Los Angeles examined hospital admissions for  
4 cardiopulmonary illnesses from 1992-1995 (Linn et al., 2000). Using a time-series approach, a 0.5 ppm  
5 increase in 24-h CO concentration (lag 0) was associated with a 2.18% (95% CI: 1.73-2.62) increase in  
6 occlusive (ischemic) stroke hospital admissions among people aged >30 years. When the analyses were  
7 stratified by season there was a 1.8% increase during winter, a 4.55% increase during summer, and a 1.6%  
8 increase during fall (results for spring were not reported).

9 A study in Taipei, Taiwan analyzed 8,582 emergency admissions for cerebrovascular diseases,  
10 hemorrhagic stroke, ischemic stroke, and all strokes during 1997 to 2002 (Chan et al., 2006). Single-day  
11 lags 0 to 3 were analyzed and a 0.75 ppm increase in 8-h max CO concentration (lag 2) was associated  
12 with an increase in cerebrovascular diseases (OR: 1.03 [95% CI: 1.01-1.05]) and all strokes (OR: 1.03  
13 [95% CI: 1.01-1.05]). These results persisted in two- and three-pollutant models that included O<sub>3</sub> and  
14 PM<sub>10</sub>. There was no association with individual ischemic or hemorrhagic stroke. CO was moderately  
15 correlated with PM<sub>10</sub> (r = 0.47) and PM<sub>2.5</sub> (r = 0.44), and the correlation was higher with NO<sub>2</sub> (r = 0.77).

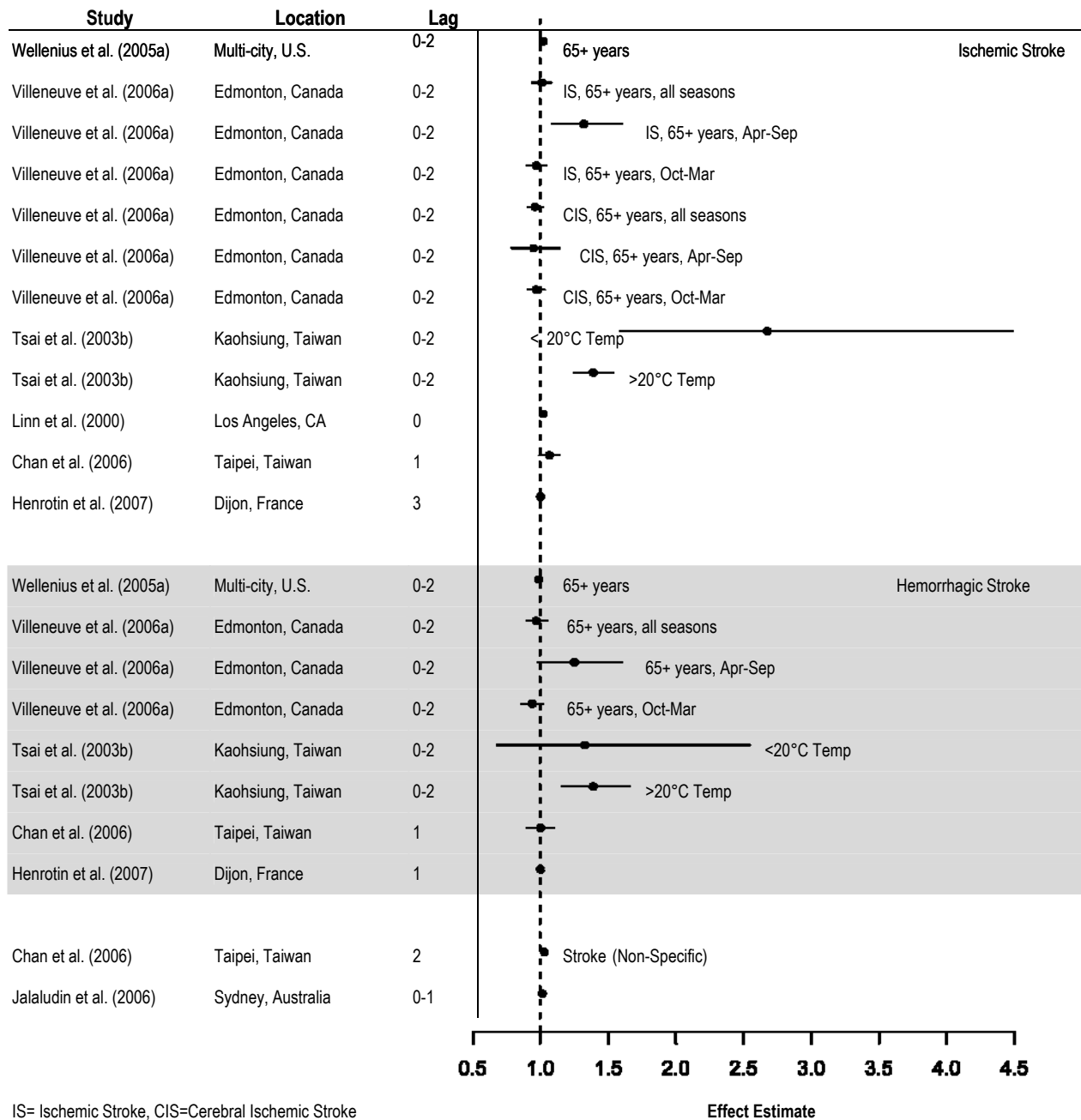
16 The only time-series study that focused specifically on stroke hospital admissions that did not  
17 report a significant association with ambient CO was conducted in Dijon, France (Henrotin et al., 2007).  
18 Hospital admissions for different types of first-ever stroke (e.g., ischemic, hemorrhagic) among subjects  
19 over 40 years of age were analyzed for the period of 1994 to 2004. A bi-directional case-crossover study  
20 design was employed where single-day lags of 0 to 3 were examined and CO had no significant  
21 association across all lags. This was also the case when the analyses were stratified by gender and types  
22 of ischemic stroke (large arteries, lacunar, cardioembolic, transient). Of all pollutants examined (PM<sub>10</sub>,  
23 NO<sub>x</sub>, O<sub>3</sub>, SO<sub>2</sub>, CO) only O<sub>3</sub> showed a significant effect.

24 Two Australian studies examined associations between ambient CO and hospital admissions for  
25 various CVDs. The first of these studies analyzed data from five of the largest cities in Australia  
26 (Brisbane, Canberra, Melbourne, Perth, Sydney) and two New Zealand cities (Auckland, Christchurch)  
27 for the period 1998-2001 (Barnett et al., 2006). A time-stratified case-crossover design was employed and  
28 the age groups of 15-64 years and ≥ 65 years were analyzed for the 0-1 lag period (average over lag 0 and  
29 1). The pooled estimates across all cities showed that CO had no effect on stroke admissions (quantitative  
30 results not reported).

31 The second of the Australian studies examined ED visits for CVDs in older adults (65+ years) in  
32 Sydney for the period from 1997 to 2001 (Jalaludin et al., 2006). Using a time-series approach, single-day  
33 lags of 0 to 3 and an average over lags 0 and 1 (e.g., lag 0-1) were examined and CO showed no effect on  
34 stroke ED visits. When the analyses were stratified by cool and warm periods a 0.75 ppm increase in 8-h  
35 max CO concentration during the cool period was associated with a 3.8% (95% CI: 0.76-6.94) increase in  
36 stroke ED visits.

1           Figure 5-2 shows the effect estimates associated with daily admissions for stroke from selected  
2 studies. Table 5-8 shows a summary of the stroke hospital admission studies that examined CO exposures.

3           In summary, there was some evidence that increased ambient CO concentrations were associated  
4 with an increase in the number of hospital admissions for stroke. The largest positive effects came from  
5 the Taiwan study in Kaohsiung (Tsai et al., 2003b) with slightly larger effects during the warmer period  
6 ( $>20^{\circ}\text{C}$ ). Similarly, in the Canadian study by Villeneuve and colleagues (2006a) there was a stronger  
7 effect during the warmer period (April-September).



**Figure 5-2. Summary of effect estimates (95% confidence intervals) associated with ED visits and hospital admissions for stroke. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.**

**Table 5-8. Summary of stroke hospital admission studies.<sup>1</sup>**

Study	Location	Type Of Stroke Examined	Copollutants	Lags Examined	CO Levels (ppm)
<i>STUDIES THAT FOCUSED SOLELY ON STROKE</i>					
Wellenius et al. (2005a)	9 cities, USA (1993-1999)	Isch, Hem	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub>	0,1, 2	25th, 50th, 75th percentiles: 0.73, 1.02, 1.44
Villeneuve et al. (2006a)	Edmonton, Canada (1992-2002)	Isch, Hem, TIA	NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1, 0-2	Mean: 0.8 (24 h)
Tsai et al. (2003b)	Kaohsiung, Taiwan (1997-2000)	Isch, Hem	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2	Mean: 0.79 (24 h)
Chan et al. (2006)	Taipei, Taiwan (1997-2002)	All, Isch, Hem	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3	Mean: 1.7 (8h)
Henrotin et al. (2007) <sup>2</sup>	Dijon, France (1994-2004)	Isch, Hem	PM <sub>10</sub> , NO <sub>x</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3	Mean: 0.59 (24 h)
<i>STUDIES THAT EXAMINED STROKE AMONG OTHER CVDS</i>					
Linn et al. (2000)	Los Angeles, CA (1992-1995)	Isch	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	Lag 0	Mean: (24 h) Winter 1.7, Spring 1.0, Summer 1.2, Fall 2.1
Barnett et al. (2006)	Australia and New Zealand (1998-2001)	All	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	Lag 0-1	Mean: (8h) 0.5-2.1
Jalaludin et al. (2006)	Sydney, Australia (1997-2001)	All	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3, 0-1	Mean: 0.82 (8h)

<sup>1</sup>Isch = Ischemic; Hem = Hemorrhagic; TIA = transient ischemic attack

<sup>2</sup>These studies presented CO concentrations in the units mg/m<sup>3</sup>. The concentrations were converted to ppm using the conversion factor 1 ppm = 1.15 mg/m<sup>3</sup>, which assumes standard atmosphere and temperature.

## Congestive Heart Failure

1 Heart failure (HF) is a condition in which the heart is unable to adequately pump blood to the rest  
 2 of the body. It does not refer to the cessation of the heart, but more to the inability of the heart to operate  
 3 at an optimal capacity. HF is often called congestive heart failure (CHF), which refers to when the  
 4 inadequate pumping leads to a buildup of fluid in the tissues. The underlying causes of CHF are  
 5 hypertension, CAD, MI, and diabetes.

6 Wellenius and colleagues (2005b) examined the rate of hospitalization for CHF among 55,019  
 7 Medicare recipients (aged ≥ 65 years) residing in Allegheny County, PA, during 1987-1999. A time-  
 8 stratified case-crossover design was employed and single-day lags of 0 to 3 were analyzed and a 1 ppm  
 9 increase in 1-h max CO concentration on the same-day (lag 0) was associated with a 9.31%  
 10 (95% CI: 6.77-11.92) increase in the rate of hospitalization for CHF. This result persisted in two-pollutant  
 11 models that included PM<sub>10</sub>, NO<sub>2</sub>, O<sub>3</sub>, and SO<sub>2</sub>. CO was moderately correlated with SO<sub>2</sub> (r = 0.54) and  
 12 PM<sub>10</sub> (r = 0.57) and more highly correlated with NO<sub>2</sub> (r = 0.70).

13 Another U.S. study recruited 125 patients diagnosed with CHF who were admitted to Johns  
 14 Hopkins Bayview Medical Center in Baltimore, MD (Symons et al., 2006). The patients were interviewed  
 15 after admission through the ED during their stays in overnight wards. The interview was designed to

1 collect information about symptom onset, health conditions, and factors related to air pollution exposure.  
2 Various lag periods (single day and cumulative days 0 to 3) prior to the onset of symptoms were analyzed  
3 and although the focus of this study was exposure to PM<sub>2.5</sub>, of all the pollutants examined (PM<sub>2.5</sub>, CO,  
4 NO<sub>2</sub>, O<sub>3</sub>) only 8-h max CO concentration at lag 2 was significantly associated with the onset of CHF  
5 symptoms (OR: 1.68 [95% CI: 1.28- 2.80]).

6 Earlier research conducted in metropolitan Los Angeles, CA examined hospital admissions for  
7 cardiopulmonary illnesses 1992-1995 (2000). Using a time-series approach, a 0.5 ppm increase in same-  
8 day 24-h avg CO concentration was associated with a 1.25% increase in CHF hospital admissions among  
9 people aged >30 years. When the analyses were stratified by seasons only summer showed a significant  
10 increase (3.7%); however, the study did not report the results for the other seasons.

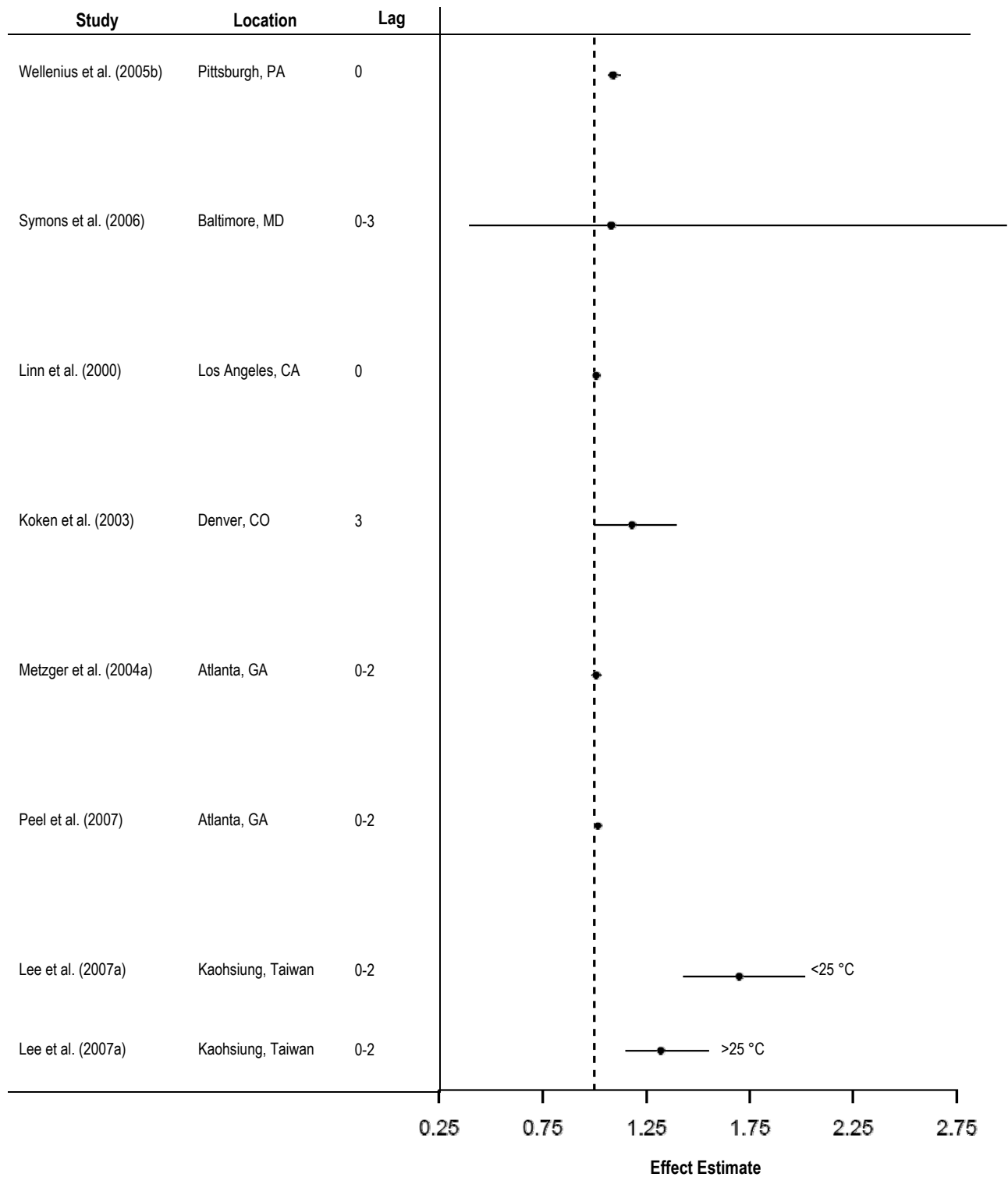
11 A time-series study in Denver, Colorado, investigated daily admissions for various CVDs among  
12 older adults (>65 years) across 11 hospitals (Koken et al., 2003). Single-day lags 0 to 4 were examined  
13 and an increase of 0.5 ppm in 24-h avg CO concentration for lag 3 was associated with an 18%  
14 (95% CI: 0.2-39.3) increase in risk of hospitalization for CHF.

15 As stated earlier, a study was conducted in Atlanta, GA, where over 4 million ED visits from  
16 31 hospitals for the period 1993 to 2000 were analyzed (Metzger et al., 2004a). A time-series design was  
17 used and a 3-day moving average over single-day lags 0-2 as the a priori lag structure was analyzed.  
18 Results showed that 1-h max CO concentration was not associated with an increase in ED visits for CHF  
19 (RR: 1.010 [95% CI: 0.988-1.032] per 1 ppm increase). When the analyses examined the same CVDs  
20 among those with and without specific secondary conditions (e.g., co-morbidity) 1-h max CO  
21 concentration was associated with an increase in ED visits for CHF only among those with COPD (OR:  
22 1.058 [95% CI: 1.003-1.115] per 1 ppm increase) (Peel et al., 2007).

23 In Kaohsiung city, Taiwan, a study analyzed 13,475 admissions for CHF across 63 hospitals for the  
24 period 1996 through 2004 (Lee et al., 2007a). A 0.5 ppm increase in 24-h avg CO concentration averaged  
25 over lag days 0-2 was positively associated with CHF hospital admissions on cool days (<25°C) (OR:  
26 1.70 [95% CI: 1.43-2.01) with a slightly weaker effect on warm days (>25°C) (OR: 1.32 [95% CI: 1.15-  
27 1.55]). These results persisted in two-pollutant models that included PM<sub>10</sub>, SO<sub>2</sub>, O<sub>3</sub>, and models with NO<sub>2</sub>  
28 only on warmer days, not with NO<sub>2</sub> on cooler days.

29 Figure 5-3 shows the effect estimates for associations between CO and daily admissions for HF  
30 from selected studies. Table 5-9 summarizes the HF hospital admission studies that examined CO  
31 exposures.

32 In summary, many of the studies that examined associations between ambient CO concentrations  
33 and daily hospital admissions for CHF reported significant associations at lags of 0 to 3 days.



**Figure 5.3.** Summary of effect estimates (95% confidence intervals) associated with hospital admissions for CHF. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.



**Table 5-9. Summary of HF hospital admission studies.**

Study	Location	Endpoints Examined	Copollutants	Lags Examined	CO Levels (ppm)
<i>STUDIES THAT FOCUSED SOLELY ON HF</i>					
Wellenius et al. (2005b)	Pittsburgh, PA (1987-1999)	CHF	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3	Mean: 1.03 (24 h)
Symons et al. (2006)	Baltimore, MD (2002)	CHF	PM <sub>2.5</sub> , NO <sub>2</sub> , O <sub>3</sub>	0,1,2,3	Mean: 0.4 (24 h)
Lee et al. (2007a)	Kaohsiung, Taiwan (1996-2004)	CHF	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2	Mean: 0.76 (24 h)
<i>STUDIES THAT EXAMINED HF AMONG OTHER CVDS</i>					
Linn et al. (2000)	Los Angeles, CA (1992-1995)	CHF, MI, All CVD, CA, OS	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	0	Mean: (24 h) Winter 1.7; Spring 1.0 Summer 1.2; Fall 2.1
Koken et al. (2003)	Denver, CO (1993-1997)	CHF, MI, CAth, PHD, CD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3	Mean: 0.9 (24 h)
Metzger et al. (2004a)	Atlanta, GA (1993-2000)	CHF, IHD, All CVD, CD, PVCD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2ma	Mean 1.5 (1 h)
Peel et al. (2007)	Atlanta, GA (1993-2000)	CHF, IHD, All CVD, CD, PVCD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2ma	Mean 1.5 (1 h)

\*Cardiac = AMI, angina, dysrhythmia, or HF; CA = Cardiac arrhythmia; CAth = Cardiac atherosclerosis; CD = cardiac dysrhythmias; CHF = Congestive heart failure; PHD = Pulmonary heart disease; OS = Occlusive stroke; PVCD = peripheral vascular and cerebrovascular disease, ma = moving average.

## Cardiovascular Diseases

1 The following section reviews studies that have investigated the effect of CO on ED visits and  
 2 hospital admissions for all CVD outcomes (e.g., non-specific). Several of these studies also examined  
 3 specific CVDs and were briefly discussed in previous sections.

4 As discussed earlier, a study was conducted in Atlanta, GA where over 4 million ED visits from  
 5 31 hospitals for the period 1993 to 2000 were analyzed (SOPHIA). Several articles have been published  
 6 from this research with three examining cardiovascular admissions in relation to CO exposures. The first  
 7 of these used a time-series design and analyzed a 3-day moving average over single-day lags 0-2 as the a  
 8 priori lag structure (Metzger et al., 2004a). Results showed that a 1 ppm increase in 1-h max CO  
 9 concentration was associated with an increase in daily ED visits for all CVDs (RR: 1.017  
 10 [95% CI: 1.008-1.027]). This persisted in two-pollutant models that included NO<sub>2</sub> and PM<sub>2.5</sub>.

11 The second of these publications examined the association of ambient air pollution levels and  
 12 cardiovascular morbidity in visits with and without specific secondary conditions (Peel et al., 2007).  
 13 Within a time-stratified case-crossover design, a 3-day moving average over single-day lags 0-2 was used  
 14 as the a priori lag structure. Results from the case-crossover analyses on all cardiovascular and peripheral  
 15 vascular and cerebrovascular disease were similar to the time-series results presented earlier. Results from  
 16 the various co-morbidity analyses are presented in Table 5-10. Similar to the results from the earlier  
 17 publication, CO was mostly associated with peripheral vascular and cerebrovascular disease (PVCD)  
 18 among those with and without the co-morbidities, except among those with CHF. Overall, there is limited,  
 19 if any, evidence of susceptibility to the effects of CO concentration for those with comorbid conditions.

**Table 5-10. Association of ambient air pollution levels and cardiovascular morbidity in visits with and without specific secondary conditions.**

Co-morbidity	IHD	Dysrhythmias	PVCD	CHF
<i>HYPERTENSION</i>				
- With	1.007 (0.978-1.037)	1.065 (1.015-1.118)	1.038 (1.004-1.074)	1.037 (0.997-1.079)
- Without	1.022 (1.000-1.043)	1.008 (0.988-1.029)	1.027 (1.002-1.054)	1.010 (0.985-1.037)
<i>DIABETES</i>				
- With	0.985 (0.945-1.027)	1.058 (0.976-1.146)	1.065 (1.012-1.121)	1.020 (0.975-1.067)
- Without	1.023 (1.004-1.042)	1.014 (0.995-1.034)	1.025 (1.003-1.048)	1.018 (0.993-1.044)
<i>COPD</i>				
- With	0.996 (0.938-1.057)	0.972 (0.878-1.077)	1.113 (1.027-1.205)	1.058 (1.003-1.115)
- Without	1.018 (1.000-1.036)	1.018 (0.999-1.038)	1.026 (1.004-1.047)	1.011 (0.987-1.036)
<i>CHF</i>				
- With	0.956 (0.907-1.007)	1.065 (0.968-1.173)	1.072 (0.981-1.172)	-
- Without	1.024 (1.006-1.042)	1.015 (0.996-1.034)	1.029 (1.008-1.051)	-
<i>DYSRHYTHMIAS</i>				
- With	1.028 (0.985-1.072)	-	1.072 (1.011-1.138)	1.004 (0.960-1.051)
- Without	1.014 (0.995-1.033)	-	1.026 (1.004-1.048)	1.023 (0.998-1.049)

PVCD - peripheral vascular and cerebrovascular disease, IHD = ischemic heart disease, CHF = congestive heart failure.

Source: Peel et al. (2007)

1 The third study utilizing the SOPHIA data extended the time period to include 1993 through 2004  
2 (Tolbert et al., 2007) and focused on two large outcome groups: a respiratory diseases group and a  
3 cardiovascular diseases group. The combined cardiovascular case group included the following groups of  
4 primary ICD-9 diagnostic codes: IHD (410-414), cardiac dysrhythmias (427), CHF (428), and peripheral  
5 vascular and cerebrovascular disease (433-437, 440, 443-445, 451-453). Results showed that a 1 ppm  
6 increase in 1-h max CO concentration was associated with an increase in daily ED visits for all CVDs  
7 (RR: 1.016 [95% CI: 1.008-1.024]). CO was the strongest predictor of CVD effects in models with two-  
8 pollutant combinations of NO<sub>2</sub>, CO and TC, as well as in a model including all three pollutants.

9 Earlier research conducted in Los Angeles, CA, showed that a 0.5 ppm increase in same-day 24-h  
10 avg CO concentration was associated with a 1.6% increase in CVD hospital admissions among people  
11 aged >30 years (2000). When the analyses were stratified by season the significant CO effect was  
12 strongest during winter (1.9% increase) followed by summer (1.8%) and fall (1.4%) with no effect in  
13 spring.

14 In contrast to other North American studies, a study in Spokane, WA, did not find an association  
15 between CO (lags of 1 to 3 days) and an increase in the number of daily cardiac hospital admissions  
16 (quantitative results not reported) (Slaughter et al., 2005). Similarly, a time-series study in Windsor,  
17 Ontario, did not find an association between ambient CO and daily hospital admissions for CVDs  
18 (defined as HF, IHD, or dysrhythmias) (Fung et al., 2005). A total of 11,632 cardiac admissions were

1 analyzed for the period of 1995 to 2000. The lag periods analyzed in this study were lag 0 (same-day), a  
2 2-day avg (lag 0-1), and a 3-day avg (lag 0-2). For a 1 ppm increase in 1-h max CO concentration the  
3 mean percent change in daily admissions for the <65 age group (lag 0) was -2.6 (95% CI: -6.2 to 3.3); and  
4 for the 65+ age group, 0.4 (95% CI: -1.9 to 2.7). The authors reported moderate to low correlations with  
5 NO<sub>2</sub> (r = 0.38), PM<sub>10</sub> (r = 0.21) and SO<sub>2</sub> (r = 0.16).

6 Two case-crossover studies in Taiwan reported an association between ambient CO and hospital  
7 admissions for CVDs. In Taipei, a total of 74,509 CVD admissions from 47 hospitals for the period of  
8 1997-2001 were analyzed (Chang et al., 2005). An increase of 0.5 ppm in 24-h avg CO concentration  
9 (average over lags 0-2) during warmer periods ( $\geq 20^{\circ}\text{C}$ ) was associated with an increase in daily hospital  
10 admissions (OR: 1.09 [95% CI: 1.065-1.121] but not cooler periods ( $<20^{\circ}\text{C}$ ) (OR: 0.98  
11 [95% CI: 0.93-1.004]). These results persisted after controlling for PM<sub>10</sub>, SO<sub>2</sub>, or O<sub>3</sub> in two-pollutant  
12 models. An identical study in Kaohsiung analyzed 29,661 CVD admissions for the period 1997-2000  
13 (Yang et al., 2004b). Results showed that a 0.5 ppm increase in 24-h avg CO concentration was associated  
14 with an increase in CVD hospital admissions during both the warmer periods (OR: 1.50  
15 [95% CI: 1.38-1.63] and cooler periods (OR: 1.89 [95% CI: 1.69-2.12])).

16 Similarly, two Australian studies also reported associations between ambient CO concentrations  
17 and increased hospital admissions among older adults. The first of these studies analyzed data from five  
18 of the largest cities in Australia (Brisbane, Canberra, Melbourne, Perth, Sydney) and two New Zealand  
19 cities (Auckland, Christchurch) for the period 1998-2001 (Barnett et al., 2006). The combined estimates  
20 showed that an increase of 0.75 ppm in the average 8-h max CO concentration over the current and  
21 previous day (lag 0-1) was associated with a 1.8% (95% CI: 0.7-2.8) increase in all CVD admissions  
22 among those aged 65+ years. Among those aged 15-64 years there was a smaller increase in CVD  
23 admissions (1.0% [95% CI: 0.2-1.7]). The second of the Australian studies examined ED visits for CVDs  
24 in older adults (65+ years) in Sydney for the period 1997 to 2001 (Jalaludin et al., 2006). A 0.75 ppm  
25 increase in 8-h max CO concentration for single-day lags 0 and 1 was associated with increases in  
26 admissions of 2.5% (95% CI: 1.6-3.5) and 1.4% (95% CI: 0.5-2.4) respectively. Based on an average over  
27 lags 0 and 1 (e.g., lag 0-1) there was an increase of 2.6% (95% CI: 1.5-3.6). There were positive increases  
28 of approximately 3% in CVD ED visits during the cool (May-October) period, but not the warm period  
29 (November-April).

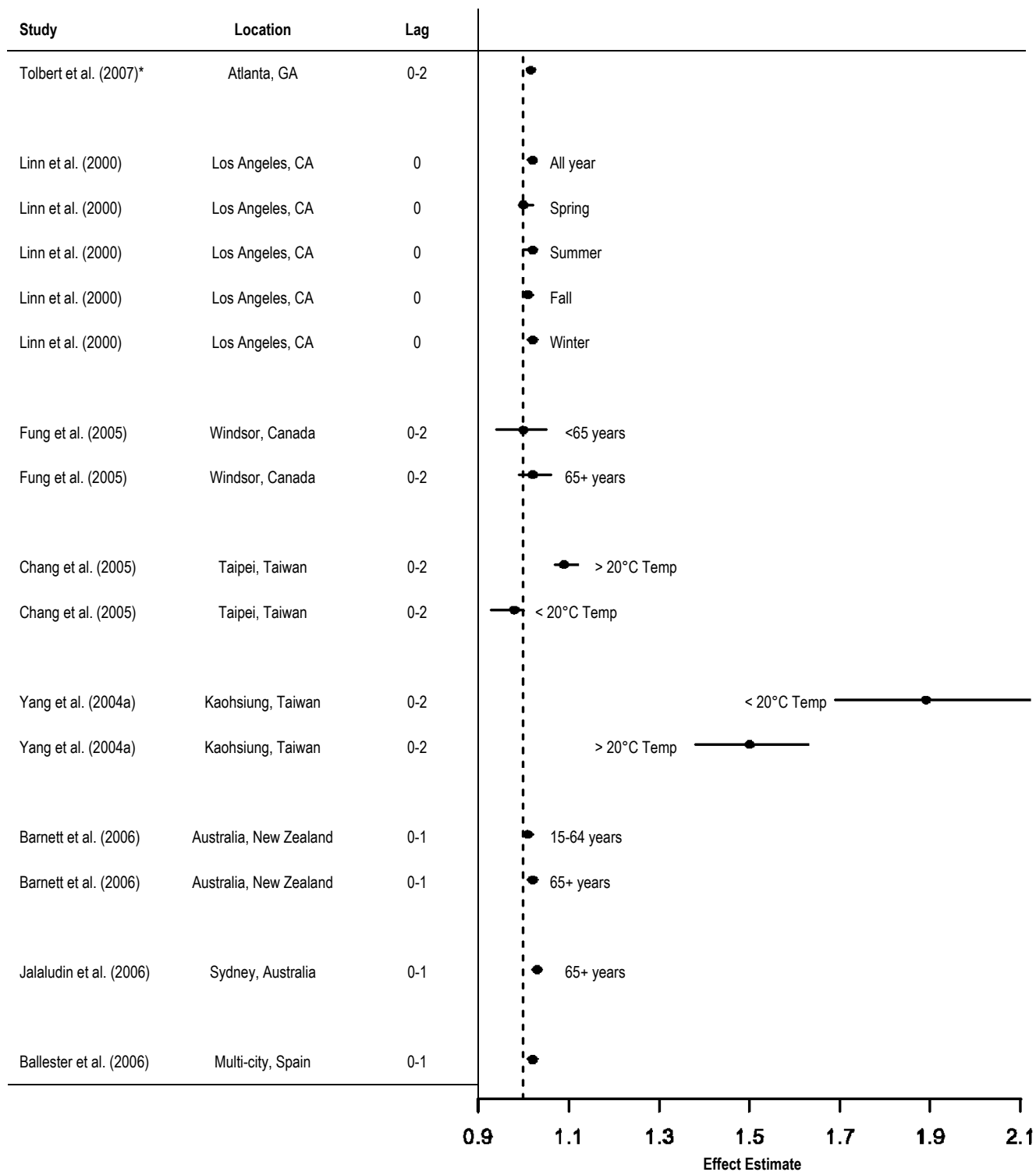
30 Very few studies investigating the association between CO and cardiovascular hospital admissions  
31 have been conducted in European cities. Ballester et al. (2001) analyzed emergency hospital admissions  
32 in Valencia, Spain for the period 1994 to 1996. The mean daily number of CVD admissions was 7 and  
33 when using a time-series approach there was no association between CO and admissions for all CVDs  
34 (RR: 1.009 [95% CI: 0.99-1.016] per 1 ppm increase in 1-h max CO concentration), heart diseases (RR:  
35 1.010 [95% CI: 0.993-1.028] per 1 ppm increase), and cerebrovascular diseases (RR: 0.985  
36 [95% CI: 0.959-1.012] per 1 ppm increase). When the analyses were stratified by hot and cold seasons,

1 only CO concentrations during the hot season were associated with an increase in all cardiovascular  
2 admissions (RR: 1.033 [95% CI: 1.006-1.064] per 1 ppm increase), heart disease admissions (RR: 1.033  
3 [95% CI: 1.000-1.067] per 1 ppm increase), and cerebrovascular admissions (RR: 1.074  
4 [95% CI: 1.007-1.113] per 1 ppm increase).

5 Ballester et al. (2006) extended this research to include data from 14 Spanish cities for the period  
6 of 1995 to 1999. An average exposure period over lags 0-1 was analyzed and for the combined estimates  
7 a 0.75 ppm increase in 8-h max CO concentration was associated with a 1.77% (95% CI: 0.56-2.99)  
8 increase in all cardiovascular emergency hospital admissions and a larger increase of 3.57%  
9 (95% CI: 1.12-6.08) for heart disease admissions. These results persisted in two-pollutant models that  
10 included NO<sub>2</sub>, O<sub>3</sub> and SO<sub>2</sub>.

11 Table 5.11 summarizes the non-specific CVD hospital admission studies that examined CO  
12 exposures. Figure 5-4 shows the effect estimates associated with daily admissions for non-specific CVD  
13 hospital admissions from selected studies.

14 In summary, many of the studies that examined associations between ambient CO concentrations  
15 and ED visits and daily hospital admissions for CVD reported significant associations at short (0-1 day)  
16 lags. Among studies that conducted stratified analyses, there were slightly stronger effects among older  
17 adults and possibly during warmer periods.



\*Represents the results reported by Metzger et al. (2004b) and Peel et al. (2007)

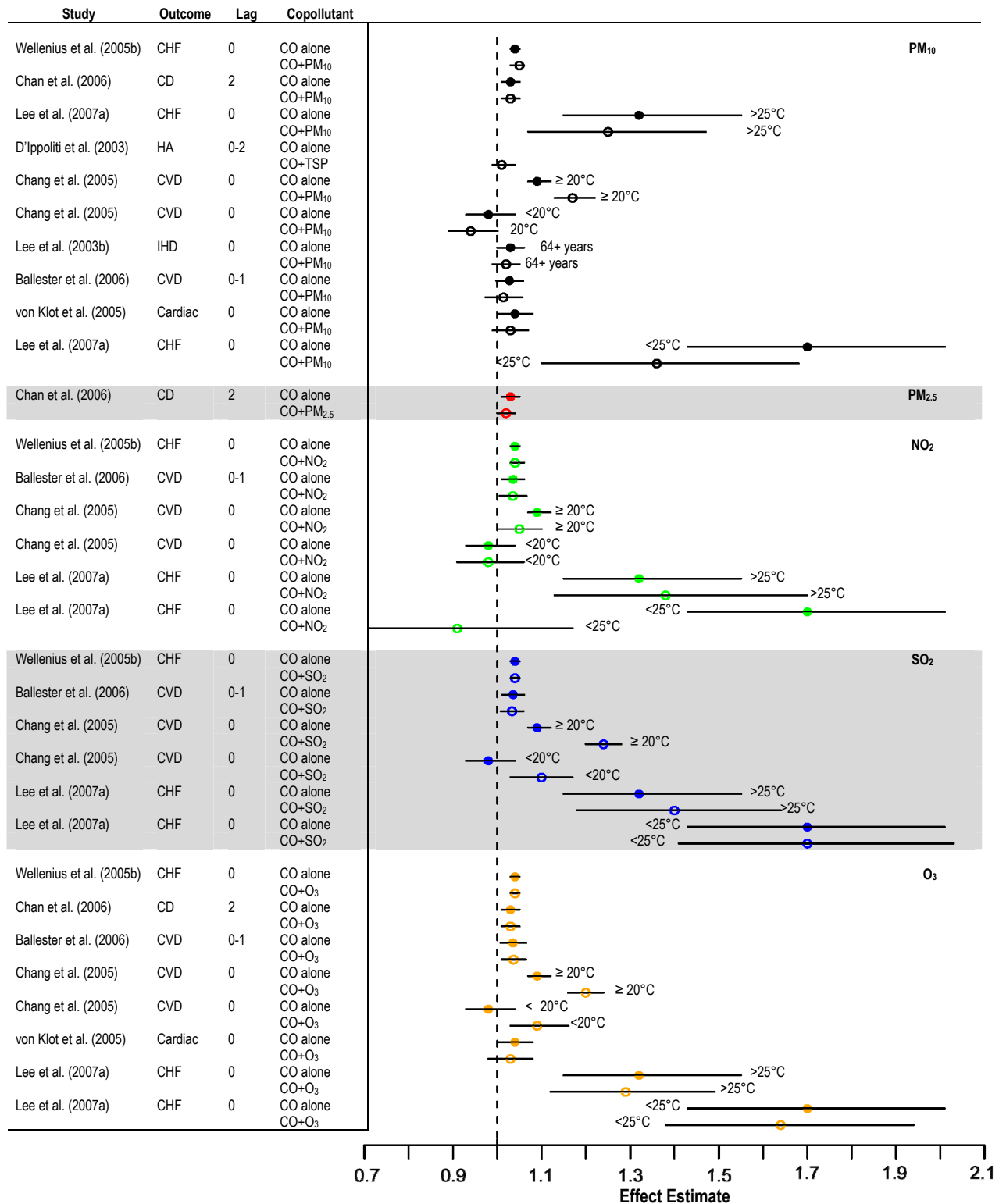
**Figure 5-4. Summary of effect estimates (95% confidence intervals) associated with hospital admissions for CVD. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.**

**Table 5-11. Summary of non-specific CVD hospital admission studies.**

Study	Location	CVD Codes	Copollutants	Lags Examined	CO Levels (ppm)
Metzger et al. (2004a)	Atlanta, GA (1993-2000)	All CVD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2ma	Mean: 1.5 (1 h)
Peel et al. (2007)	Atlanta, GA (1993-2000)	All CVD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2ma	Mean 1.5 (1 h)
Tolbert et al. (2007)	Atlanta, GA (1993-2004)	All CVD	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2ma	Mean 1.6 (1 h)
Linn et al. (2000)	Los Angeles, CA (1992-1995)	All CVD	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	0	Mean: (24 h) Winter 1.7; Spring 1.0; Summer 1.2; Fall 2.1
Slaughter et al. (2005)	Spokane, WA (1995-2001)	All CVD (ICD9: 390-459)	PM <sub>10</sub> , PM <sub>2.5</sub> , CO	1,2,3	Mean: range across 5 monitors 0.42-1.82 (24 h)
Fung et al. (2005)	Windsor, Canada (1995-2000)	All CVD (HF, IHF, or Dysrhythmia)	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0, 0-1, 0-2	Mean: 1.3 (24 h)
Chang et al. (2005)	Taipei, Taiwan (1997-2001)	All CVD (ICD9: 410-429)	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2	Mean: 1.37 (24 h)
Yang et al. (2004b)	Kaohsiung, Taiwan (1997-2000)	All CVD (ICD9: 410-429)	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-2	Mean: 0.79 (24 h)
Barnett et al. (2006)	Australia and New Zealand (1998-2001)	All CVD (ICD9: 390-459)	PM <sub>10</sub> , NO <sub>2</sub> , O <sub>3</sub>	0-1	Mean: (8h) 0.5-2.1
Jalaludin et al. (2006)	Sydney, Australia (1997-2001)	All CVD (ICD9: 390-459)	PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0,1,2,3, 0-1	Mean: 0.82 (8h)
Ballester et al. (2001) <sup>1</sup>	Valencia, Spain (1994-1996)	All CVD (ICD9: 390-459)	BS, NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	1,2,3,4,5	Mean: 0.54 (24 h)
Ballester et al. (2006) <sup>1</sup>	Multi-city, Spain (1995-1999)	All CVD (ICD9: 390-459)	BS, PM <sub>10</sub> , TSP, NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub>	0-1	Mean: range across 14 cities 0.12-0.24 (8h)

<sup>1</sup> These studies presented CO concentrations in the units mg/m<sup>3</sup>. The concentrations were converted to ppm using the conversion factor 1 ppm = 1.15 mg/m<sup>3</sup>, which assumes standard atmosphere and temperature.

1 Figure 5-5 summarizes the effects of CO concentration on ED visits and hospital admissions for all  
2 CVD outcomes other than stroke from studies that presented the results from two-pollutant models.  
3 Generally, the CO effect estimates from these studies are robust to the inclusion of copollutants, including  
4 PM<sub>10</sub>, PM<sub>2.5</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and O<sub>3</sub>. In all but one instance (Lee et al., 2007a) (<25°C adjusted for NO<sub>2</sub>) when  
5 the single pollutant effect estimate was positive for CO, it remained positive after the addition of any of  
6 the copollutants investigated.



**Figure 5-5. Effect estimates from studies of ED visits and hospital admissions for CVD outcomes other than stroke from single pollutant (CO only, closed circles) and copollutant (CO plus PM, NO<sub>2</sub>, SO<sub>2</sub> and O<sub>3</sub>, open circles) models. Effect estimates are standardized to 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO levels, and 0.5 ppm for 24-h avg CO concentrations.**

### 5.2.1.2. Epidemiologic Studies with Long-Term Exposure

1           Only two studies examined CVD outcomes in association with long-term exposure to CO.  
2 Rosenlund et al. (2006) investigated long-term exposure (30 years) to urban air pollution and the risk of  
3 MI in Sweden. The study included 2,246 cases and 3,206 controls aged between 45 to 70 years and  
4 residing in Stockholm County during 1992 to 1993. A detailed postal questionnaire was completed by  
5 4067 subjects and all addresses inhabited during more than 2 years since 1960 were geocoded. The  
6 exposures were then derived from dispersion calculations based on emissions data for each decade since  
7 1960. These calculations were estimates of annual mean levels of traffic-generated NO<sub>x</sub>, NO<sub>2</sub>, CO, PM<sub>10</sub>,  
8 and PM<sub>2.5</sub>, with the addition of SO<sub>2</sub> from heating sources. The analyses were stratified by all cases,  
9 nonfatal cases, fatal cases, in-hospital death, and out-of-hospital death. Based on a 30 year avg exposure  
10 all pollutants were not associated with overall MI incidence. However, increased CO was associated with  
11 out-of-hospital death from MI (OR: 1.81 [95% CI: 1.02-3.23] per 0.5 ppm increase in 30-year avg CO  
12 concentration). Similar results were reported for NO<sub>2</sub>. The correlation between the 30-year NO<sub>2</sub> and CO  
13 exposures was reasonably strong (r = 0.74) and multipollutant models with both these pollutants included  
14 (NO<sub>2</sub>, CO) were not examined. No other pollutants were significantly associated with all other MI  
15 outcomes.

16           A small-area ecologic study analyzed mortality and hospital admissions for stroke across 1,030  
17 census districts in Sheffield, U.K. (Maheswaran et al., 2005b). Stroke counts within each census district  
18 were linked to modeled air pollution data which was then grouped into quintiles of exposure. For stroke  
19 hospital admissions, when the analyses were adjusted for only sex and age demographics there was an  
20 exposure-response pattern exhibited across the quintiles of CO exposure with all levels reaching  
21 significance (RR: 1.37 [95% CI: 1.24-1.52] for the highest exposure group compared to the lowest  
22 group). However, this result did not persist when also adjusting for a deprivation index and smoking rates  
23 across the districts (RR: 1.11 [95% CI: 0.99-1.25]).

### 5.2.1.3. Summary of Epidemiologic Studies of Exposure to CO and Cardiovascular Effects

24           A substantial number of epidemiologic studies have examined the potential association between  
25 exposure to CO and various physiological cardiac endpoints or biomarkers. Overall, despite some mixed  
26 results reported among panel and retrospective cohort studies, there was evidence that exposure to CO has  
27 an effect on HR, various HRV parameters, and blood markers of coagulation and inflammation.  
28 Conversely, based on results from panel studies there was little evidence of a link between CO and  
29 cardiac arrhythmia, cardiac arrest, the occurrence of myocardial infarction, and increased BP.



1 Studies of ED visits and hospital admissions provide evidence that CO is associated with various  
2 forms of CVD with lag periods ranging from 0 to 3 days. There is little evidence that ambient CO is  
3 associated with an increase in hospital admissions for ischemic stroke. Studies of hospital admissions and  
4 ED visits for IHD and CHF provide the strongest evidence of ambient CO being associated with adverse  
5 CVD outcomes. It is difficult to determine from this group of studies the extent to which CO is  
6 independently associated with CVD outcomes or if CO is a marker for the effects of another traffic-  
7 related pollutant or mix of pollutants. On-road vehicle exhaust emissions are a nearly ubiquitous source of  
8 combustion pollutant mixtures that include CO and can be an important contributor to CO in near-road  
9 locations. Although this complicates the efforts to disentangle specific CO-related health effects, the  
10 evidence indicates that CO associations generally remain robust in copollutant models, are coherent with  
11 the effects demonstrated by controlled human exposure and animal toxicological studies, and supports a  
12 direct effect of short-term CO exposure on CVD morbidity at ambient concentrations below the current  
13 NAAQS level. Such direct effects are plausible considering that long-term, low concentration CO  
14 exposure (See Section 4.2.3) could result in a COHB level approaching those used in controlled human  
15 exposure studies (See Section 5.2.2).

## 5.2.2. Controlled Human Exposure Studies

16 Controlled human exposure studies provide valuable information related to the health effects of  
17 exposure to air pollutants following short term exposures. These types of experiments are often referred to  
18 as human clinical studies and are conducted in a laboratory setting under carefully regulated exposure  
19 concentrations, environmental conditions, and subject activity levels. Human clinical studies are typically  
20 conducted using a randomized crossover study design with subjects exposed to both the pollutant(s) of  
21 interest and a clean air control. Results of controlled human exposure studies can be used to provide  
22 coherence with the evidence from epidemiologic studies by expanding the understanding of potential  
23 mechanisms for the observed health outcomes. However, they may also provide information that can be  
24 used directly in quantitatively characterizing the exposure concentration-health response relationships at  
25 ambient or near-ambient concentrations. Human clinical studies are limited by a number of factors  
26 including a small sample size and relatively short exposure time. In addition, although health-  
27 compromised individuals have been included in human clinical studies, all subjects participating must be  
28 relatively healthy and do not represent the most sensitive individuals in the population.

29 Several human clinical studies cited in the 2000 CO AQCD observed changes in measures of  
30 cardiovascular function among individuals with CAD following short term exposures to CO. In a multi-  
31 laboratory study of men with stable angina, Allred et al. (1989, 1991) evaluated the effect of CO exposure  
32 on exercise-induced angina and ST-segment changes indicative of myocardial ischemia. Relative to clean

1 air exposure (COHb  $\approx$  0.6%), exposures to CO resulting in COHb concentrations of 2.4% and 4.7% were  
2 shown to decrease the time required to induce ST-segment changes by 5.1% ( $p = 0.02$ ) and 12.1%  
3 ( $p < 0.001$ ), respectively. These changes were well correlated with the onset of exercise-induced angina. A  
4 number of other studies involving individuals with stable angina have also demonstrated a CO-induced  
5 decrease in time to onset of angina as well as reduction in duration of exercise at COHb concentrations  
6 between 3 and 6% (Adams et al., 1988; Anderson et al., 1973) (Kleinman et al., 1989; Kleinman et al.,  
7 1998). However, Sheps et al. (1987) observed no change in time to onset of angina or maximal exercise  
8 time following a 1-h exposure to 100 ppm CO (targeted COHb of 4%) among a group of 30 patients with  
9 CAD. In a subsequent study conducted by the same laboratory, a significant increase in number of  
10 ventricular arrhythmias during exercise was observed relative to room air among individuals with CAD  
11 following a 1-h exposure to 200 ppm CO (targeted COHb of 6%), but not following a 1-h exposure to  
12 100 ppm CO (targeted COHb of 4%) (Sheps et al., 1990). It should be noted that although the subjects  
13 evaluated in the studies described above are not necessarily representative of the most sensitive  
14 population, the level of disease in these individuals was relatively severe, with the majority either having  
15 a history of MI or having  $\geq 70\%$  occlusion of one or more of the coronary arteries.

16 The 2000 CO AQCD presented very little evidence of CO-induced changes in cardiovascular  
17 function in healthy adults. Davies and Smith (1980) exposed healthy young adults continuously for 7 days  
18 to CO concentrations of 0, 15, or 50 ppm. In this study, a marked ST-segment depression was  
19 demonstrated in only 1 out of 16 subjects following exposure to 15 ppm CO (2.4% COHb) or 50 ppm CO  
20 (7.2% COHb). Since the publication of the 2000 CO AQCD, no new human clinical studies have been  
21 published involving controlled CO exposures among subjects with CAD. However, a number of new  
22 studies have evaluated changes in various measures of cardiovascular and systemic responses following  
23 controlled exposures to CO in healthy adults. Adir et al. (1999) exposed 15 young healthy adult males to  
24 room air or CO for approximately 4 min, using a CO exposure concentration which had been shown to  
25 produce the targeted COHb level of 4-6%. Following each exposure, subjects performed an exercise  
26 treadmill test at their maximal capacity. Exposure to CO was not observed to cause arrhythmias,  
27 ST-segment changes, or changes in myocardial perfusion (thallium scintigraphy) during post-exposure  
28 exercise. However, CO was demonstrated to decrease the post-exposure duration of exercise by  
29 approximately 10% ( $p = 0.0012$ ). In addition, the authors reported significant CO-induced decreases in  
30 metabolic equivalent units ( $p < 0.001$ ), which is a relative measure of  $O_2$  consumption. These results  
31 support the findings of several studies cited in the 2000 CO AQCD which observed decreases in exercise  
32 duration and maximal aerobic capacity among healthy adults at COHb levels  $\geq 3\%$  (Drinkwater et al.,  
33 1974; Ekblom and Huot, 1972; Horvath et al., 1975; Raven et al., 1974). While these decreases in  
34 exercise duration were relatively small and only likely to be noticed by competing athletes, the findings  
35 are nonetheless important in providing coherence with the observed effects of CO on exercise-induced  
36 myocardial ischemia among patients with CAD.

1 Kizakevich et al. (2000) evaluated the cardiovascular effects of increasing CO concentration in  
2 healthy adults engaged in upper and lower body exercise. Subjects were initially exposed for 4-6 min to  
3 CO concentrations between 1,000 and 3,000 ppm, followed by continued exposure to 27, 55, 83, and  
4 100 ppm to maintain COHb levels of 5, 10, 15, and 20%, respectively. Relative to room air control, CO  
5 exposure was not observed to cause ST-segment changes or affect cardiac rhythm at any concentration  
6 during either upper or lower body exercise. Compensation mechanisms for reduced O<sub>2</sub> carrying capacity  
7 during CO exposure were demonstrated, with statistically significant increases in heart rate occurring at  
8 COHb levels  $\geq$  5%, and statistically significant increases in cardiac output and cardiac contractility  
9 observed at COHb levels  $\geq$  10%. In a human clinical study designed to evaluate the contribution of CO to  
10 cardiovascular morbidity associated with cigarette smoking, Zevin et al. (2001) exposed 12 healthy male  
11 smokers for 7 consecutive days to clean air, CO, or cigarette smoke, with each subject serving as his own  
12 control. The COHb levels were similar between the exposures to cigarette smoke and CO, with average  
13 concentrations of 6% and 5%, respectively. Cigarette smoke, but not CO, was observed to significantly  
14 increase plasma levels of CRP and plasma platelet factor 4 relative to the air control arm of the study.  
15 Neither cigarette smoke nor CO was shown to affect BP. Hanada et al. (2003) observed an increase in leg  
16 muscle sympathetic nerve activity (MSNA) following controlled exposures to CO (COHb  $\approx$ 20%) under  
17 normoxic or hyperoxic conditions. Although an increase in the magnitude of sympathetic activation is  
18 typically associated with regional vasoconstriction, no CO-induced changes in femoral venous blood flow  
19 were observed in this study. These findings are in agreement with those of Hausberg et al. (1997) who  
20 observed no change in forearm blood flow or BP in a study of 10 healthy men and women following a  
21 controlled exposure to CO (COHb  $\approx$  8%). Interestingly, one recent study did observe an increase in retinal  
22 blood flow, retinal vessel diameter, and choroidal blood flow following controlled exposures to CO at a  
23 concentration of 500 ppm (Resch et al., 2005). This protocol resulted in COHb concentrations of 5.6%  
24 and 9.4% following exposures of 30 and 60 min, respectively, with statistically significant increases in  
25 retinal and choroidal blood flow observed at both time points relative to synthetic air control. This  
26 CO-induced change in ocular hemodynamics may have been due to local tissue hypoxia; however, the  
27 clinical significance of this finding is unclear. Exposures to CO have also been shown to affect skeletal  
28 muscle function, with one recent human clinical study reporting a decrease in muscle fatigue resistance in  
29 healthy adult males using both voluntary and electrically-induced contraction protocols following  
30 controlled exposures to CO resulting in an average COHb level of 6% (Morse et al., 2008).

31 In summary, controlled human exposures to CO among individuals with CAD have been shown to  
32 consistently increase markers of myocardial ischemia at COHb levels between 3 and 6%, with one study  
33 reporting similar effects following CO exposures resulting in COHb concentrations of 2.4%. No such  
34 effects have been observed in healthy adults following controlled exposures to CO. Although some  
35 studies have reported CO-induced hemodynamic changes among healthy adults at COHb concentrations  
36 of as low as 5%, this effect has not been consistent across studies.

### 5.2.3. Toxicological Studies

1 While novel toxicological research on environmental levels of CO was limited in the 2000 CO  
2 AQCD, several sections reported potential relationships between CO exposure and cardiovascular effects.  
3 Conflicting experimental data relating to the role of CO in promoting atherosclerotic vessel disease was  
4 discussed. While some animal studies have linked chronic CO exposure with atherosclerosis development  
5 resulting from increased fatty streaking and cellular lipid loading (Davies et al., 1976; Thomsen, 1974;  
6 Turner et al., 1979), other studies have failed to see this association (Penn et al., 1992; Stupfel and  
7 Bouley, 1970). Ventricular hypertrophy has also been shown after chronic CO exposure (Penney et al.,  
8 1984; Penney et al., 1988).

9 The following sections describe recent studies dealing with toxicity of relatively low levels of CO.  
10 There has been little new research with the overt purpose of examining environmentally-relevant levels of  
11 CO. For the most part, studies were designed to mimic exposures related to cigarette smoke, either side-  
12 stream or mainstream, accidental CO poisoning, or for the purposes of therapeutic application. Thus, few  
13 studies examined levels of CO within the current 1 h (35 ppm) or 8 h (9 ppm) NAAQS levels, and fewer  
14 still examined concentration response curves to delineate no effects levels. However, it is apparent that  
15 CO, at low to moderate levels (35-250 ppm), has pathophysiological effects on the cardiovascular system  
16 and on relatively ubiquitous cellular pathways.

17 CO exposure at environmentally-relevant levels is unlikely to overwhelm normal physiology in a  
18 healthy cell; however, susceptibility may be rendered by disease or early development. A common theme  
19 appears to be the vulnerability of vascular cells, especially the endothelium, which could be considered  
20 the first organ of contact once taken up into the circulation. While relatively little research has been  
21 conducted since the 2000 CO AQCD, several key studies conducted at near-environmental CO levels  
22 provide important clues to the potential public health implications of ambient CO exposure.

#### 5.2.3.1. Endothelial Dysfunction

23 While the preferential binding to heme and effective displacement of O<sub>2</sub> by CO has been well  
24 established for over a century, new information from various fields of study are beginning to elucidate  
25 non-hypoxic mechanisms that may lead to cardiovascular abnormalities associated with CO exposure.  
26 Research by Thom, Ischiropoulos, and colleagues (Ischiropoulos et al., 1996; Thom et al., 1994; Thom  
27 and Ischiropoulos, 1997; Thom et al., 1997; Thom et al., 1999a; Thom et al., 1999b; Thom et al., 2000;  
28 Thom et al., 2006) has focused on CO-mediated displacement of NO• from heme-binding sites. Some of  
29 this work demonstrates a specific pathway by which severe CO poisoning can lead to the release NO•  
30 from platelets with subsequent neutrophil activation and vascular injury (Ischiropoulos et al., 1996; Thom  
31 et al., 2006). The steps include (1) peroxyntirite generation from the reaction of NO• from platelets with

1 neutrophil-derived superoxide followed by (2) stimulation of intravascular neutrophil degranulation that  
2 can result in (3) myeloperoxidase deposition along the vascular lining. Products from myeloperoxidase-  
3 mediated reactions can cause endothelial cell activation (Thom et al., 2006) and can lead to endothelial  
4 dysfunction. The concentrations used in these studies are greatly in excess of the NAAQS levels, but  
5 certainly within the range of accidental or occupational exposures. Research by these same investigators  
6 at more environmentally-relevant CO levels was partially reviewed in the 2000 CO AQCD. The release of  
7 free NO• was noted in isolated rat platelets exposed to 10-20 ppm CO (Thom and Ischiropoulos, 1997).  
8 Increased nitrotyrosine content of the aorta was observed in rats exposed to 50 ppm CO for 1 h (1999a;  
9 Thom et al., 1999b). Furthermore in this same study, a 1-h exposure to 100 ppm CO led to albumin efflux  
10 from skeletal muscle microvasculature at 3 h and leukocyte sequestration in the aorta at 18 h. LDL  
11 oxidation was also reported. These effects were dependent on NOS but not on neutrophils or platelets. A  
12 second study demonstrated NO-dependent effects of 50-100 ppm CO in lungs and is described in Section  
13 5.5.4 (Thom et al., 1999b). Studies in cultured endothelial cells were also conducted using buffer  
14 saturated with 10-100 ppm CO (Thom et al., 1997). These experiments were designed to mimic  
15 conditions where blood COHb levels were between 3.8 and 28% resulting in exposure of endothelial cells  
16 to 11-110 nM CO. CO-stimulated release of NO• from endothelial cells along with peroxynitrite  
17 formation and delayed cell death was observed at CO concentrations of 22 nM and higher (Thom et al.,  
18 1997). A more recent study demonstrated adaptive responses in endothelial cells exposed to this same  
19 range of CO concentrations (Thom et al., 2000). Specifically, 1-h exposure to 11 nM CO resulted in  
20 MnSOD and HO-1 induction and resistance to the apoptotic effects of 110 nM CO. These protective  
21 effects of CO were mediated by NO•, as demonstrated using an inhibitor of NOS and a scavenger of  
22 peroxynitrite. Collectively, these experiments demonstrated altered oxidative stress, the initiation of  
23 inflammation, increased microvascular permeability and altered cell signaling in animals and isolated  
24 cells following exposure to 10-100 ppm CO.

25 CO is an endogenous regulator of vasomotor tone through vasodilatory effects mediated by  
26 activation of soluble guanylate cyclase and activation of large conductance Ca<sup>2+</sup> activated K<sup>+</sup> channels.  
27 However, CO does not cause vasodilation in every vascular bed. For example, 5, 100, 500 and 2,500 ppm  
28 CO administered by inhalation to near-term fetal lambs did not induce pulmonary vasodilation and the  
29 HO inhibitor zinc protoporphyrin IX failed to affect baseline vascular tone (Grover et al., 2000). In some  
30 cases CO promotes vasoconstriction, which is thought to be mediated by inhibition of endothelial NOS  
31 (Johnson et al., 2003; Thorup et al., 1999) or decreased NO• bioavailability. An interesting series of  
32 studies has also suggested that endogenous CO derived from HO-1 which is induced in a variety of  
33 disease models (salt-sensitive forms of hypertension, metabolic syndrome in obese rats) is responsible for  
34 skeletal muscle arterial endothelial dysfunction (Johnson et al., 2003, 2004; Teran et al., 2005). Additional  
35 studies will be useful in determining whether environmentally-relevant concentrations of CO have  
36 detrimental effect on pre-existing conditions such as hypertension, metabolic syndrome or pregnancy.

1 Several recent animal studies examined the vascular effects of controlled exposures to complex  
2 combustion mixtures containing CO. Vascular dilatation was decreased following exposure to diesel (4 h  
3 at 4 ppm) (Knuckles et al., 2008) and gasoline engine emissions (6 h/day x 1, 3, and 7 day at 80 ppm)  
4 (Lund et al., 2009). Furthermore, evidence of vascular ROS following gasoline emissions has been shown  
5 in certain animal models (6 h/day x 50 day at 8-80 ppm) (Lund et al., 2007). While none of these studies  
6 examined the potential role of CO, alone or in combination, it is clearly a common factor in the various  
7 combustion atmospheres, and future work will be needed to reveal its importance on vascular health.

### 5.2.3.2. Cardiac Remodeling Effects

8 Cardiomyopathy, or abnormal growth of the cardiac muscle, can manifest in different ways,  
9 depending on the nature of the insult. The adverse effects of cardiac hypertrophy are due to reduction of  
10 ventricular chamber volume and a diminishing efficiency of the heart. Such concentric hypertrophy  
11 typically occurs in response to chronic increases in load, as occurs with hypertension. Ischemia of the  
12 cardiac tissue can also lead to cardiac remodeling and myopathy. During and after an acute infarction or  
13 obstruction of major coronary vessels, downstream tissues can suffer severe regional ischemia that leads  
14 to significant necrosis. Such regions will lose the ability to contract, and surrounding tissue will show  
15 deficits in contractility. Decreased contractility is often a result of structural thinning of the ventricular  
16 wall, as well as metabolic impairments. Chronic ischemia, such as may result from CAD, may similarly  
17 impair cardiomyocyte function and cause decreased contractility and remodeling. However, ultimately  
18 cardiomyopathies are of a complex origin involving mismanagement of fluid balance, abnormal hormonal  
19 influences (epinephrine, angiotensin), and insufficient perfusion/nutrition. Assessing the role of  
20 exogenous CO in altering pathways leading to cardiomyopathy is a relatively new endeavor and several  
21 new findings are of great interest.

22 The heart is a known target for CO toxicity, potentially due to its high rate of O<sub>2</sub> consumption.  
23 Direct effects of CO on the healthy heart have only been observed at relatively high concentrations. For  
24 example, a recent study by Sorhaug et al. (2006) demonstrated cardiac hypertrophy in rats exposed for 72  
25 weeks to 200 ppm CO. COHb levels were reported to be 14.7%. Neither structural signs of hypertension  
26 in the pulmonary arteries or atherosclerotic lesions in the systemic arteries were observed. Cardiac  
27 hypertrophy was also demonstrated in rats exposed to 100-200 ppm CO for 1-2 weeks (Loennechen et al.,  
28 1999). This response was accompanied by an increase in endothelin-1 expression. COHb levels were  
29 reported to be 12-23% in this study.

30 Direct effects of CO on the healthy heart have also been demonstrated following short-term  
31 exposures. In a study by Favory et al. (2006b), rats were exposed to 90 min of 250 ppm CO, which led to  
32 peak COHb values of roughly 11%; recovery of 96 h was needed for COHb levels to return to baseline.  
33 The authors noted that within the first 24 hours of recovery, while COHb values decreased from 11% to

1 5%, the coronary vascular perfusion pressure and the left ventricular developed pressure were  
2 significantly increased compared to baseline. Concomitantly, the ratio of cGMP to cAMP decreased and  
3 the sensitivity of the coronary vascular bed to both acetylcholine and a NO• donor were reduced by CO  
4 exposure. The authors concluded that the discordant alterations in contractility (increased) and perfusion  
5 (decreased) may place the heart at risk of O<sub>2</sub> limitations following this exposure to CO.

6 Several studies examined the impact of lower levels (50 ppm) on pre-existing or concurrent cardiac  
7 pathologies. In one such study, CO exacerbated the effects of a hypoxia-based model of right ventricular  
8 remodeling and failure (Gautier et al., 2007). In controlled laboratory settings, chronic hypobaric hypoxia  
9 (HH) caused right ventricular hypertrophy as a result of pulmonary arterial vasoconstriction and increased  
10 pulmonary resistance. Using such a model (Wistar rats exposed for 3 weeks to hypoxia), CO (50 ppm  
11 during the last week of hypoxia, continuous) only increased COHb from 0.5% to 2.4% in the hypoxia  
12 model, yet had significant effects on blocking compensatory functional responses to hypoxia, such as  
13 increased fractional shortening and contractility. Also, while right ventricular weight was increased by  
14 hypoxia alone, significant pathology related to necrosis was observed in the hypoxia + CO-exposed rats.  
15 The reduced coronary perfusion of the right ventricle in hypoxia + CO-exposed rats may help explain the  
16 histopathological findings. The authors cited previous work demonstrating that exogenous CO can inhibit  
17 NOS (Thorup et al., 1999), which is essential for coronary dilation and angiogenesis. Thus, this study  
18 provided evidence that exogenous CO may interrupt or downregulate pathways that endogenous CO may  
19 activate.

20 In two studies by Melin et al. (2002; 2005), Dark Agouti rats were exposed for 10 weeks to either  
21 HH, 50 ppm CO or HH plus 50 ppm CO. CO exposure amplified the right ventricular cardiac hypertrophy  
22 and decreased the right ventricular diastolic function which occurred in response to HH. In addition, the  
23 combined exposure led to effects on left ventricular morphology and function which were not seen with  
24 either exposure alone. Changes in HRV were also reported. Results from both of these studies combined  
25 with results of Gautier and colleagues (2007) indicated that CO may interfere with normal homeostatic  
26 responses to hypoxia. This could occur by blocking HIF-1 $\alpha$ -responsive elements (vascular endothelial  
27 growth factor, erythropoietin) or other cell signaling pathways.

28 In a similar study, Carraway et al. (2002) exposed rats to HH (380 torr) with or without co-  
29 exposure to CO (50 ppm). These exposures were continuous for up to 21 days and focused on pulmonary  
30 vascular remodeling. While the addition of CO to HH did not alter the thickness or diameter of vessels in  
31 the lung, there was a significant increase in the number of small (<50  $\mu$ m) diameter vessels compared to  
32 control, HH only, and CO-only exposures. Despite the greater number of vessels, the overall pulmonary  
33 vascular resistance was increased in the combined CO + hypoxic exposure, which the authors attributed  
34 to enhancement of muscular arterioles and  $\beta$ -actin. Results of this study taken together with results from  
35 the Gautier et al. (2007) and Melin studies (2002; 2005) suggested that the combined effect of low levels  
36 of CO with hypoxia is an enhanced right ventricle workload and an exacerbated cardiomyopathy related

1 to pulmonary hypertension. The population at risk of primary pulmonary hypertension is low, but  
2 secondary pulmonary hypertension is a frequent complication of COPD and certain forms of heart failure.

### 5.2.3.3. Electrocardiographic Effects

3 In two related studies, Wellenius et al. (2004; 2006b) examined the effect of CO on a rat model of  
4 ischemia-related arrhythmia that was previously shown to produce significant results with exposures to  
5 PM (Wellenius et al., 2002). ECG changes were observed during exposure to residual oil fly ash (ROFA)  
6 particles in a rat model of MI. Thus, using an anesthetized model of post-infarction myocardial sensitivity,  
7 Wellenius and colleagues tested the effects of 35 ppm CO (1-h exposure) on the induction of spontaneous  
8 arrhythmias in Sprague Dawley rats (Wellenius et al., 2004). CO exposure caused a statistically  
9 significant decrease (60.4%) in ventricular premature beat (VPB) frequency during the exposure period in  
10 rats with a high number of pre-exposure VPB. No interaction was observed with co-exposure to carbon  
11 concentrated particles, which independently reduced VPB frequency during the post-exposure period  
12 when administered alone. In a follow-up publication, results from the analysis of supraventricular ectopic  
13 beats (SVEB) were provided (Wellenius et al., 2006b). A decrease in the number of SVEB was observed  
14 with CO (average concentration 37.9 ppm) compared to filtered air. While the authors concluded that CO  
15 exposure did not increase risk of SVEB in this particular rodent model of coronary occlusion, the fact that  
16 cardiac electrophysiological dynamics are significantly altered by short-term exposure to low level CO  
17 may be of concern for other models of susceptibility.

### 5.2.3.4. Summary of Cardiovascular Toxicology

18 Recent studies demonstrated that short-term exposure to 50-100 ppm CO resulted in aortic injury  
19 as measured by increased nitrotyrosine and the sequestration of activated leukocytes in healthy rats. In  
20 addition, skeletal muscle microvascular permeability was increased. Short term-exposure to 35 ppm CO  
21 altered cardiac electrophysiology in a rat model of myocardial ischemia. Furthermore short-term exposure  
22 to 50 ppm CO exacerbated cardiac pathology and impaired function in animal models of hypertrophic  
23 cardiomyopathy and/or pulmonary hypertension. Ventricular hypertrophy was observed in healthy rats in  
24 response to chronic exposures of 100-200 ppm CO. These studies provide a strong basis for the  
25 development of adverse health effects resulting from exposures to CO at environmentally-relevant  
26 concentrations.



## 5.2.4. Summary of Cardiovascular Effects

1           The most compelling evidence of a CO-induced effect on the cardiovascular system at COHb  
2 levels relevant to the current NAAQS comes from a series of controlled human exposure studies among  
3 individuals with CAD (see Section 5.2). These studies, described in the 1991 and 2000 CO AQCDs,  
4 demonstrate consistent decreases in the time to onset of exercise-induced angina and ST-segment changes  
5 following CO exposures resulting in COHb levels of 3-6%, with one multicenter study reporting similar  
6 effects at COHb levels as low as 2.4%. No human clinical studies have evaluated the effect of controlled  
7 exposures to CO resulting in COHb levels lower than 2.4%. Human clinical studies published since the  
8 2000 CO AQCD have reported no association between CO and ST-segment changes or arrhythmia;  
9 however, none of these studies included individuals with diagnoses of heart disease.

10           While the exact physiological significance of the observed ST-segment changes among individuals  
11 with CAD is unclear, ST-segment depression is a known indicator of myocardial ischemia. It is also  
12 important to note that the individuals with CAD who participated in these controlled exposure studies  
13 were not representative of the most sensitive individuals in the population. In fact, the most sensitive  
14 individuals may respond to levels of COHb lower than 2.4%. Variability in activity patterns and severity  
15 of disease among individuals with CAD is likely to influence the critical level of COHb which leads to  
16 adverse cardiovascular effects.

17           The degree of ambient CO exposure which leads to attainment of critical levels of COHb will also  
18 vary between individuals. First of all, endogenous CO production varies as described in Section 4.5, but  
19 generally results in less than 1% COHb. Secondly, nonambient exposures to CO, such as exposure to  
20 ETS, can increase COHb above baseline levels. Ambient exposures will result in an additive increase in  
21 COHb. Using mathematical modeling to predict changes in COHb in healthy inactive adults (Quantitative  
22 Circulatory Physiology [QCP] model, Section 4.2.3), it can be estimated that exposure to 35 ppm CO for  
23 1 h results in an increase of 0.6% COHb over baseline and exposure to 9 ppm CO for 8 h results in an  
24 increase of 0.8% COHb over baseline. Furthermore, 24 h exposure to 3 ppm CO results in an increase of  
25 0.4% COHb above baseline which can also be obtained following 1-h exposure to 30 CO ppm.  
26 Consequently, exposure to CO at concentrations relevant to the NAAQS has the potential to increase  
27 COHb to levels associated with adverse cardiovascular health effects in some individuals.

28           Findings of controlled human exposure studies are coherent with findings of recent epidemiologic  
29 studies conducted since the 2000 CO AQCD, which observed associations between ambient CO  
30 concentration and ED visits and hospital admissions for IHD, CHF and all-cause cardiovascular disease.  
31 All but one of these epidemiologic studies were conducted in locations where the entire distribution of  
32 CO concentrations were at or below the level of the current NAAQS, with mean 24-h avg concentrations  
33 ranging from 0.5 ppm (Montreal, Canada) to 9.4 ppm (Tehran, Iran) (Table 5-7). A single study reported a  
34 negative association between CO concentration and hospital admissions and ED visits for IHD among all

1 ages; all other associations were positive, with increases in hospital admissions and ED visits for IHD  
2 between 0.2% and 19.8% per standardized increase in CO concentration (Figure 5-1). These recent  
3 studies build upon the conclusions of the 2000 CO AQCD that short-term variations in ambient CO  
4 concentrations are associated with daily hospital admissions for heart disease.

5 These health outcomes are consistent with a role for CO in limiting O<sub>2</sub> availability (i.e., hypoxic  
6 mechanisms) in individuals with CAD. However, recent toxicological studies suggested that CO may also  
7 act through non-hypoxic mechanisms by disrupting cellular signaling. Studies in healthy animals  
8 demonstrated oxidative injury and inflammation in response to 50-100 ppm CO while studies in disease  
9 models demonstrate effects on heart rhythm and exacerbation of cardiomyopathy and vascular remodeling  
10 in response to 35-50 ppm CO. Furthermore, in utero exposure to 150 ppm CO alters postnatal  
11 electrophysiological maturation in rat cardiomyocytes. Further investigations will be useful in determining  
12 the importance of non-hypoxic mechanisms following environmentally-relevant CO exposures. Taken  
13 together, the evidence from epidemiologic, human clinical and toxicological studies is sufficient to  
14 conclude that **a causal relationship is likely to exist between relevant short-term CO exposures and**  
15 **cardiovascular morbidity.**

## 5.3. Central Nervous System Effects

### 5.3.1. Controlled Human Exposure Studies

16 The behavioral effects of controlled human exposures to CO have been examined by several  
17 laboratories, and these studies were summarized in the 2000 CO AQCD. Briefly, decreases in visual  
18 tracking as well as visual and auditory vigilance were observed following exposures to CO resulting in  
19 COHb levels between 5% and 20% (Benignus et al., 1987; Fodor and Winneke, 1972; Horvath et al.,  
20 1971; Putz et al., 1979). One study reported similar behavioral effects (time discrimination) among a  
21 group of healthy volunteers with COHb levels <3% (Beard and Wertheim, 1967); however, subsequent  
22 studies were unable to replicate these findings at such low exposure concentrations (Otto et al., 1979;  
23 Stewart et al., 1973a). These outcomes represent a potentially important adverse effect of CO exposure  
24 resulting in COHb levels ≥ 5%, although it is important to note that these findings have not been  
25 consistent across studies. Similarly, some studies demonstrated decreases in reaction time as well as  
26 decrements in cognitive function and fine motor skills following controlled exposures to CO; however,  
27 these studies were not typically conducted using double-blind procedures, which may significantly affect  
28 the outcome of behavioral studies (Benignus, 1993). It should be noted that all behavioral studies of

1 controlled CO exposure were conducted in normal, healthy adults. No new human clinical studies have  
2 evaluated CNS or behavioral effects of exposure to CO.

### 5.3.2. Toxicological Studies

3 The evidence for toxicological effects of CO exposure in laboratory animal models comes from in  
4 utero or perinatal exposure to relatively low levels of CO (25 to 750 ppm). Affected endpoints from this  
5 early, developmental CO exposure include behavior, memory, learning, locomotor ability, peripheral  
6 nervous system myelination, auditory decrements, and neurotransmitter changes. These data are  
7 addressed in detail in the birth outcomes section of the ISA (Section 5.4.2).

### 5.3.3. Summary of Central Nervous System Effects

8 Exposure to high levels of CO has long been known to adversely affect CNS function, with  
9 symptoms following acute CO poisoning including headache, dizziness, cognitive difficulties,  
10 disorientation, and coma. However, the relationship between ambient levels of CO and neurological  
11 function is less clear and has not been evaluated in epidemiologic studies. Studies of controlled human  
12 exposures to CO discussed in the 2000 CO AQCD reported inconsistent neural and behavioral effects  
13 following exposures resulting in COHb levels of 5-20%. No new human clinical studies have evaluated  
14 central nervous system or behavioral effects of exposure to CO. At ambient-level exposures, healthy  
15 adults may be protected against CO-induced neurological impairment owing to compensatory responses  
16 including increased cardiac output and cerebral blood flow. However, these compensatory mechanisms  
17 are likely impaired among certain potentially susceptible groups, including individuals with reduced  
18 cardiovascular function.

19 Toxicological studies that were not discussed in the 2000 CO AQCD employed rodent models to  
20 show that low level CO exposure during the in utero or perinatal period can adversely affect adult  
21 outcomes including behavior, neuronal myelination, neurotransmitter levels or function, and the auditory  
22 system (discussed in Section 5.3). In utero CO exposure, including both intermittent and continuous  
23 exposure, has been shown to impair multiple behavioral outcomes in offspring including active avoidance  
24 behavior (150 ppm CO), non-spatial memory (75 and 150 ppm CO), spatial learning (endogenous CO  
25 inhibition), homing behavior (150 ppm CO), locomotor movement (150 ppm CO), and negative geotaxis  
26 (125 and 150 ppm). In two separate studies, in utero CO exposure (75 and 150 ppm) was associated with  
27 significant myelination decrements without associated changes in motor activity in adult animals.  
28 Multiple studies demonstrated that in utero CO exposure affected glutamatergic, cholinergic,  
29 catecholaminergic, and dopaminergic neurotransmitter levels or transmission in exposed male rodents.

1 Possible or demonstrated adverse outcomes from the CO-mediated aberrant neurotransmitter levels or  
2 transmission include respiratory dysfunction (200 ppm CO), impaired sexual behavior (150 ppm CO), and  
3 an adverse response to hyperthermic insults resulting in neuronal damage (200 ppm). Finally, perinatal  
4 CO exposure has been shown to affect the developing auditory system of rodents, inducing permanent  
5 changes into adulthood. This is manifested with atrophy of cochlear cells innervating the inner hair cells  
6 (25 ppm CO), decreased immunostaining associated with impaired neuronal activation (12.5 ppm CO),  
7 impaired myelination of auditory associated nerves (25 ppm CO), decreased energy production in the  
8 sensory cell organ of the inner ear or the organ of corti (25 ppm CO), some of which is mechanistically  
9 proposed to be mediated by ROS (25 ppm CO). Functional tests of the auditory system of neonatally, low  
10 level CO-exposed rodents, using OAE testing (50 ppm) and amplitude measurements of the 8th cranial  
11 nerve action potential (12, 25, 50, 100 ppm), revealed decrements in auditory function at PND22 and  
12 permanent changes into adulthood using action potential (AP) testing (50 ppm). Together, these animal  
13 studies demonstrated that in utero or perinatal exposure to CO can adversely affect adult behavior,  
14 neuronal myelination, neurotransmission, and the auditory system in adult male rodents. Considering the  
15 combined evidence from controlled human exposure and toxicological studies, **the evidence is**  
16 **suggestive of a causal relationship between relevant short- and long-term CO exposures and**  
17 **central nervous system effects.**

## 5.4. Birth Outcomes and Developmental Effects

### 5.4.1. Epidemiologic Studies

18 Although the body of literature is growing, the research focusing on adverse birth outcomes is  
19 limited when compared to the numerous studies that have examined the more well-established health  
20 effects of air pollution. Various dichotomized measures of birth weight, such as LBW, SGA, and IUGR,  
21 have been the most examined outcomes in air pollution research while preterm birth (PTB), congenital  
22 malformations, and infant mortality are less studied.

23 In the 2000 CO AQCD only two studies were cited that examined the effect of ambient air  
24 pollution on adverse birth outcomes and both of these studies investigated LBW as an endpoint  
25 (Alderman et al., 1987; Ritz and Yu, 1999). At that time this area of research was in its infancy and since  
26 then there has been increasing interest.

### 5.4.1.1. Preterm Birth

1 A small number of air pollution-birth outcome studies have investigated the possible association  
2 between PTB and maternal exposure to CO with the majority of U.S. studies conducted in southern  
3 California. The earliest of these studies examined exposures to ambient CO during the first month of  
4 pregnancy and the last 6 weeks prior to birth among a cohort of 97,158 births in southern California  
5 between 1989 and 1993 (Ritz et al., 2000). The exposure assessment within this study was based on data  
6 from fixed site monitors that fell within a 2-mile radius of the mother's ZIP code area. The crude relative  
7 risks for PTB associated with a 1 ppm increase in 3-h avg CO concentration (6:00 to 9:00 a.m.) during the  
8 last 6 weeks prior to birth and the first month of pregnancy were 1.04 (95% CI: 1.03-1.5) and 1.01  
9 (95% CI: 1.00-1.03) respectively. However, when the authors controlled for other risk factors, only the  
10 effect associated with CO during the last 6 weeks prior to birth persisted (RR: 1.02 [95% CI: 1.01-1.03]).  
11 Furthermore, when the analyses included variables for either season or other pollutants the CO effect  
12 estimates generally were reduced.

13 Expanding on this research, Wilhelm and Ritz (2005) examined PTB among a cohort of 106,483  
14 births in Los Angeles County, CA between 1994 and 2000. Based on data recorded at monitoring stations  
15 of varying proximities to the mother's residence, the main exposure windows examined were the first  
16 trimester and the last 6 weeks prior to birth. Among women living within a 1-mile radius of a CO  
17 monitoring station, a 0.5 ppm increase in 24-h avg CO concentration during the first trimester was  
18 associated with a 3% (RR: 1.03 [95% CI: 1.00-1.06]) increased risk of PTB. This result persisted after  
19 simultaneously adjusting for NO<sub>2</sub> and O<sub>3</sub> (RR: 1.05 [95% CI: 1.00-1.10]), but not with the inclusion of  
20 PM<sub>10</sub> into the regression model (RR: 0.99 [95% CI: 0.91-1.09]). The result from the single pollutant  
21 model for CO exposures averaged over the 6 weeks prior to birth was similar in magnitude but failed to  
22 reach statistical significance (RR: 1.02 [95% CI: 0.99-1.04]).

23 A limitation of many air pollution-birth outcome studies is the limited availability of detailed  
24 information on maternal lifestyle factors and time-activity patterns during pregnancy. To assess possible  
25 residual confounding due to these factors, Ritz and colleagues (Ritz et al., 2007) were able to analyze  
26 detailed maternal information from a survey of 2,543 of 6,374 women sampled from a cohort of 58,316  
27 eligible births in 2003 in Los Angeles County. Based on data from the closest monitor to the mother's ZIP  
28 code area, exposures to CO, NO<sub>2</sub>, O<sub>3</sub>, and PM<sub>2.5</sub> during the first trimester and last 6 weeks prior to  
29 delivery were examined and results from the overall cohort (n = 58,316) with limited maternal  
30 information were compared to the more detailed nested case-control cohort (n = 2,543). Within the overall  
31 cohort, CO during the first trimester was associated with an increased risk of 25% (OR: 1.25  
32 [95% CI: 1.12-1.38]; highest exposure group >1.25 ppm vs. lowest ≤ 0.58 ppm). This result persisted  
33 within the nested case-control cohort (OR: 1.21 [95% CI: 0.88-1.65]) where factors such as passive  
34 smoking and alcohol use during pregnancy were included in the model; however, the confidence intervals

1 were wider due to the smaller sample. Any possible association between CO and PTB was less evident  
2 during the last 6 weeks prior to birth. A strength of this study was that it also highlighted how there was  
3 little change in the air pollution effect estimates when controlling for more detailed maternal information  
4 (e.g., smoking, alcohol use), as opposed to only controlling for more limited maternal information that is  
5 routinely collected on birth registry forms.

6 In contrast to the Los Angeles studies, a case-control study of PTB across California for the period  
7 1999 through 2000 found a positive, though not statistically significant association, with 24-h CO  
8 concentration during the entire pregnancy (OR: 1.03 [95% CI: 0.98-1.09] per 0.5 ppm increase), the first  
9 month of gestation (OR: 1.05 [95% CI: 0.99-1.10] per 0.5 ppm increase), and the last 2 weeks of gestation  
10 (OR: 1.00 [95% CI: 0.96-1.04] per 0.5 ppm increase) (Huynh et al., 2006). Although there was an  
11 indication of an effect during early pregnancy, the small sample size (when compared to other studies)  
12 may have influenced the lack of statistical significance. Furthermore, exposures within this study were  
13 assigned based on a county-level average which may explain the lack of effect, given the poor level of  
14 exposure assessment.

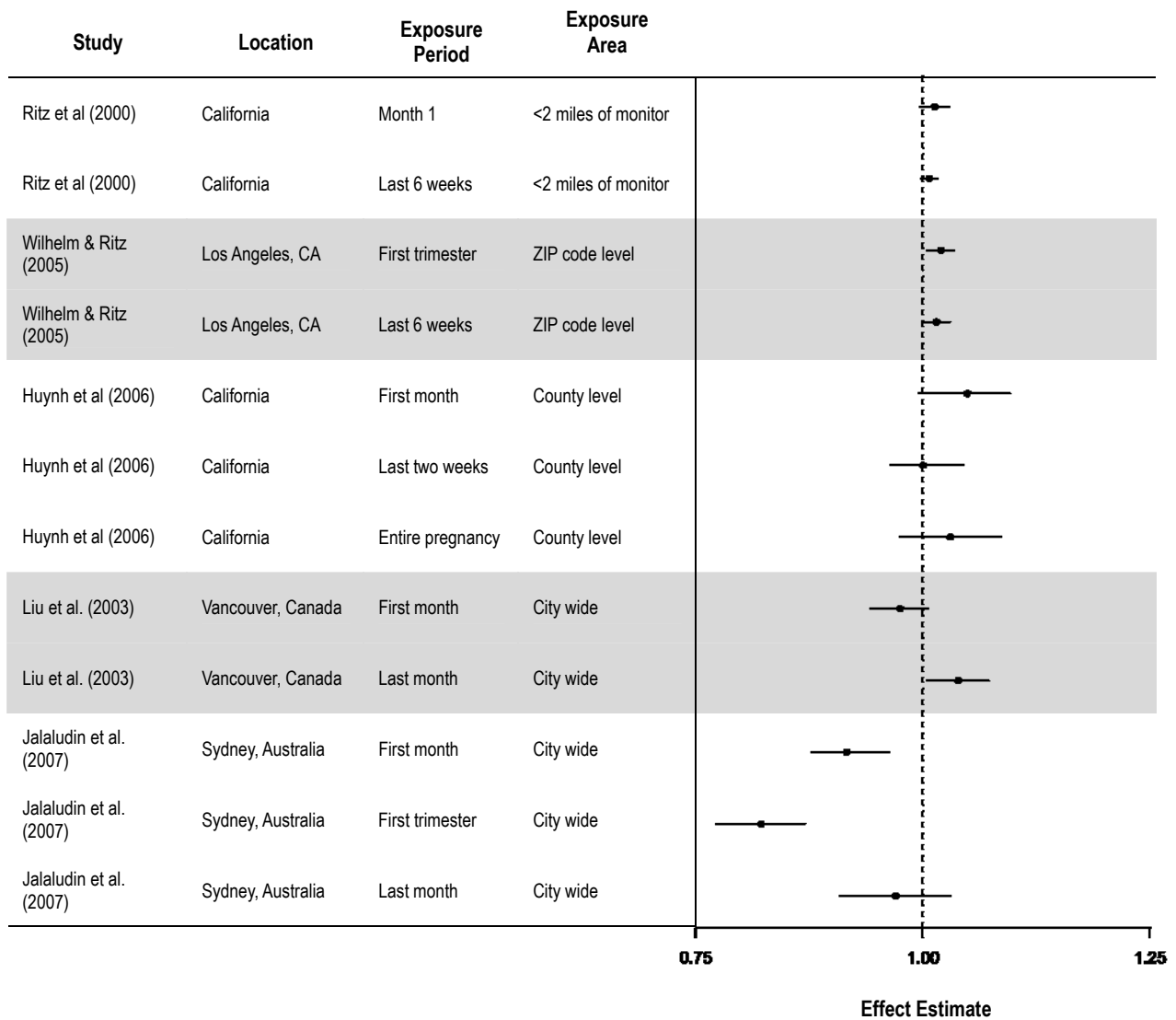
15 Studies outside of the U.S. have been conducted in Canada, Australia, and Korea with mixed  
16 results reported. In Vancouver, Canada, based on a city-wide average across available monitoring sites,  
17 24-h avg CO concentration during the last month of pregnancy was associated with a 4% (OR: 1.04  
18 [95% CI: 1.00-1.07]) increased risk of PTB per 0.5 ppm increase while there was no association found  
19 during the first month of pregnancy (OR: 0.98 [95% CI: 0.94-1.00]) (Liu et al., 2003). This study  
20 investigated maternal exposures to ambient gaseous pollutants (CO, NO<sub>2</sub>, SO<sub>2</sub>, O<sub>3</sub>) averaged over the first  
21 and last month of pregnancy among a cohort of 229,085 births between 1985 and 1998.

22 In a cohort of 52,113 births in Incheon, Korea between 2001-2002, CO concentrations during the  
23 first trimester was associated with a 26% (RR: 1.26 [95% CI: 1.11-1.44]) increased risk of PTB for the  
24 highest quartile of exposure when compared to the lowest quartile (Leem et al., 2006). There was also a  
25 strong significant trend exhibited across the quartiles. A similar result was found for 24-h avg CO  
26 concentration during the last trimester although the effect was less pronounced (RR: 1.16  
27 [95% CI: 1.01-1.24]). To assign the maternal exposures to CO, this study used a kriging technique, which  
28 is a statistical mapping technique that allows the prediction of an average concentration over a spatial  
29 region from data collected at specific points. The spatial average CO concentrations were then linked to  
30 each study subject's residential address.

31 Conversely, a study in Sydney, Australia, examined maternal exposure to ambient air pollution  
32 during the first and last month, and the first and last trimester of pregnancy among a cohort of 123,840  
33 births between 1998-2000 and found no association between PTB and CO (Jalaludin et al., 2007).  
34 Maternal exposure estimates in this study were based on a city-wide average of available monitoring sites  
35 and also based on data from fixed sites within 5 km of the mother's postcode area. The odds ratios for  
36 PTB associated with 8-h avg CO concentrations during the first trimester and last three months of

1 gestation were 1.18 (95% CI: 0.85-1.63) and 1.08 (95% CI: 0.95-1.22), respectively, when including  
2 births within 5 km of a monitor. Interestingly, when all births were included in the analyses and the  
3 exposure was based on a city-wide average, these effects had become protective for the first trimester  
4 (OR: 0.82 [95% CI: 0.77-0.87]) and last three months of gestation (OR: 0.99 [95% CI: 0.92-1.07]). This  
5 suggests that exposures based on data from fixed sites closer to the mother's address are more likely to  
6 detect an effect than a city-wide average.

7 Figure 5-6 shows the risk ratios for the risk of delivering a preterm infant from the reviewed  
8 studies. Table 5-12 provides a brief overview of the PTB studies. In summary there are mixed results  
9 across the studies. Although these studies are difficult to compare directly due to the different exposure  
10 assessment methods employed, there is some evidence that CO during early pregnancy (e.g., first month  
11 and trimester) is associated with an increased risk of PTB. The most consistency is exhibited within the  
12 studies conducted around Los Angeles, CA and surrounding areas whereby all studies reported a  
13 significant association with CO exposure during early pregnancy, and exposures were assigned from  
14 monitors within close proximity of the mother's residential address (Ritz et al., 2000; Ritz et al., 2007;  
15 Wilhelm and Ritz, 2005). It should also be noted that the mixed results when analyzing different cohorts  
16 that resided within varying proximities to a monitor may be attributable to analyzing different  
17 populations.



**Figure 5-6. Summary of effect estimates (95% confidence intervals) for PTB associated with maternal exposure to ambient CO. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.**



**Table 5-12. Brief summary of PTB studies.**

Study	Location (Sample Size)	Mean CO (ppm)	Exposure Assessment	Exposure Window
Ritz et al. (2000)	California (n = 97,158)	2.7 (6-9 a.m.)	<2 miles of monitor	Month 1 Last 6 weeks
Wilhelm and Ritz (2005)	Los Angeles, CA (n = 106,483)	1.4 (24 h)	Varying distances to monitor	Month 1 Last 6 weeks
Ritz et al. (2007)	Los Angeles, CA (n = 58,316)	0.87 (24 h)	Nearest monitor to ZIP code	Entire pregnancy Trimester 1 Last 6 weeks
Huynh et al. (2006)	California (n = 42,692)	0.8 (24 h)	County level	Entire pregnancy Month 1 Last 2 weeks
Liu et al. (2003)	Vancouver, Canada (n = 229,085)	1.0 (24 h)	City wide average	Month 1 Last month
Leem et al. (2006)	Incheon, Korea (n = 52,113)	0.9 (24 h)	Residential address within Dong-based on Kriging	First trimester Last trimester
Jalaludin et al. (2007)	Sydney, Australia (n = 123,840)	0.9 (8 h)	City wide average and <5 km from monitor	First month First trimester Last trimester Last month

### 5.4.1.2. Birth Weight, Low Birth Weight, and Intrauterine Growth Restriction/Small for Gestational Age

1 With birth weight routinely collected in vital statistics and a powerful predictor of infant mortality,  
 2 it is the most studied outcome within air pollution-birth outcome research. Air pollution researchers have  
 3 analyzed birth weight as a continuous variable, and/or as a dichotomized variable in the forms of low  
 4 birth weight (LBW) (<2,500g [5 lbs, 8 oz]) and SGA.

5 It should be noted that the terms small for gestational age (SGA), which is defined as a birth weight  
 6 <10th percentile for gestational age (and often sex), and intrauterine growth restriction (IUGR) are used  
 7 interchangeably. However, this definition of SGA does have limitations. For example, using this  
 8 definition of IUGR may overestimate the percentage of ‘growth-restricted’ neonates as it is unlikely that  
 9 10% of neonates have growth restriction (Wollmann, 1998). On the other hand, when the 10th percentile  
 10 is based on the distribution of live births at a population level the percentage of SGA among preterm  
 11 births is most likely underestimated (Hutcheon and Platt, 2008).

12 Nevertheless, the terms SGA and IUGR are often used interchangeably and it therefore should be  
 13 noted that SGA represents a statistical description of a small neonate, whereas the term IUGR is reserved  
 14 for those with clinical evidence of abnormal growth. Meaning that all IUGR neonates will be SGA, but

1 not all SGA neonates will be IUGR (Wollmann, 1998). In the following sections the terms SGA and  
2 IUGR are referred to as each cited study used the terms.

3 Over the past decade a number of studies examined various metrics of birth weight in relation to  
4 maternal exposure to CO with the majority conducted in the U.S. Given that most studies examined  
5 multiple birth weight metrics, in order to avoid overlap of the studies the following section focuses on  
6 each study only once and presents results for each metric within that study.

7 Most of the U.S. studies have been conducted in southern California with inconsistent results  
8 reported with regard to gestational timing of the CO effects. The first of these studies was reviewed in the  
9 2000 CO AQCD and is briefly summarized here. Ritz and Yu (1999) examined the effect of ambient CO  
10 during the last trimester on LBW among 125,573 births in Los Angeles between 1989 and 1993. When  
11 compared to neonates born to women in the lowest CO exposure group (<2.2 ppm), neonates born to  
12 women in the highest exposure group (5.5 ppm - 95th percentile) had a 22% (OR: 1.22 [95% CI: 1.03-  
13 1.44]) increased risk of being born as LBW.

14 Building upon this research, Wilhelm and Ritz (2005) reported similar results when extending this  
15 study to include 136,134 births for the period of 1994–2000. Exposure to ambient CO during each  
16 trimester was based on data recorded at monitoring stations of varying proximities to the mother's  
17 residence. For women residing within 1 mile of a station, there was 36% (OR: 1.36 [95% CI: 1.04-1.76])  
18 increased risk of having a term LBW baby for women in the highest CO exposure category  
19 (1.84 ppm-75th percentile) for the third-trimester. There was also an increased risk of term LBW (OR:  
20 1.28 [95% CI: 1.12-1.47]) among women in the highest exposure group when the analyses included  
21 women within a 5 mile radius of a station. However, when the analyses included women within a 1-2  
22 mile, or 2-4 mile radius of a station, the CO effects failed to reach statistical significance and there was no  
23 evidence of an exposure-response pattern exhibited across the varying distances to a station. Furthermore,  
24 none of the significant CO results persisted after controlling for other pollutants. Although standard errors  
25 were certainly increased after controlling for the other pollutants leading to non-significant results, some  
26 of the effect sizes were similar, providing some consistency. It is interesting to note, however, that  
27 maternal exposure to CO during trimesters one and two was not associated with LBW (results not  
28 reported).

29 Further validation in association with exposure times was observed in an analysis using a subset of  
30 participants in the Children's Health Study. Salam and colleagues (2005) found that CO only during the  
31 first trimester was associated with reduced fetal growth. Their research examined birth weight, LBW, and  
32 IUGR among a subset of participants in the Children's Health Study (Peters et al., 1999b) who were born  
33 in California between 1975-1987 (n = 3901). The study examined term births with a gestational age  
34 between 37-44 weeks. Exposures in this study were based on CO data from up to the three nearest  
35 monitoring sites within 50 km of the centroid of the mother's ZIP code. Exposures for the entire  
36 pregnancy and each trimester were analyzed and a 0.5 ppm increase in 24-h CO concentration during the

1 first trimester was associated with a 7.8 g (95% CI: 15.1-0.4) decrease in birth weight, which also  
2 translated to a 6.7% (OR: 1.07 [95% CI: 1.00-1.13]) increased risk of IUGR; however, there was no  
3 significant association with LBW (OR: 1.00 [95% CI: 0.88-1.16]).

4 In contrast to the previous studies, another California study of 18,247 singleton births born at 40  
5 weeks gestation during 2000 found no association between ambient 24-h CO concentration and reduced  
6 birth weight or SGA where the highest quartile of exposure was 0.98 ppm. Based on data from fixed sites  
7 within 5 miles of the mother's residence, exposures to CO and PM<sub>2.5</sub> during the entire pregnancy and each  
8 trimester were analyzed. Although CO during the entire pregnancy was associated with a 20 g  
9 (95% CI: 40.1-0.8) reduction in birth weight, this did not persist after controlling for PM<sub>2.5</sub>. PM<sub>2.5</sub> was  
10 found to have a strong effect on birth weight within each trimester (Parker et al., 2005).

11 Two similar studies were conducted in the Northeast of the U.S. with inconsistent results. A study  
12 of 89,557 singleton term births in Boston, MA, Hartford, CT, Philadelphia, PA, Pittsburgh, PA, and  
13 Washington, DC between 1994-1996 found that exposure to ambient CO during the third trimester was  
14 associated with an increased risk of LBW (OR: 1.14 [95% CI: 1.03-1.27] per 0.5 ppm increase) (Maisonet  
15 et al., 2001). When stratified by race this effect was only significant among African Americans for the  
16 first and third trimesters (First OR: 1.32 [95% CI: 1.22-1.43]; Third OR: 1.20 [95% CI: 1.09-1.32]).  
17 Exposures to PM<sub>10</sub> and SO<sub>2</sub> were examined and there was no strong evidence that these pollutants were  
18 associated with LBW. Exposures for this study were based on a city-wide average of monitors within the  
19 mother's city of residence. The second study examined 358,504 births at 32-44 weeks gestation between  
20 1999-2002 in Connecticut and Massachusetts (Bell et al., 2007). 24-h CO exposures were estimated from  
21 fixed sites within each mother's county of residence (e.g., county level). CO averaged over the entire  
22 pregnancy was associated with a reduction in birth weight of 27.0 g (95% CI: 21.0-32.8). This result  
23 persisted after controlling for each additional pollutant (PM<sub>10</sub>, PM<sub>2.5</sub>, NO<sub>2</sub>, and SO<sub>2</sub>) in two-pollutant  
24 models. However, this reduction in birth weight did not translate to an increased risk of LBW (OR: 1.05  
25 [95% CI: 0.97-1.12]). When controlling for exposure during each trimester, the reduction in birth weight  
26 associated with a 0.5 ppm increase in 24-h CO concentration during the first trimester ranged from 18.8 to  
27 16.5 g while the reductions associated with third trimester exposure ranged between 23.3 and 27.2 g. It is  
28 interesting to note that, although the exposures were based on data averaged at the county level, CO was  
29 associated with a reduction in birth weight. Whereas, in a previously cited California study by Huynh and  
30 colleagues (Huynh et al., 2006) exposures were also at the county level yet there was no association with  
31 PTB. This difference may be due to the counties being smaller in New England than in California,  
32 resulting in more precise exposure estimates.

33 Two studies in Canada investigated the effects of ambient air pollution on fetal growth with  
34 exposures derived from a city-wide average across the available monitoring sites. The first of these  
35 studies was among a cohort of 229,085 singleton term births (37-42 weeks gestation) in Vancouver, BC  
36 with monthly and trimester exposures to CO investigated in relation to LBW and IUGR (Liu et al., 2003).

1 For a 0.5 ppm increase in 24-h CO concentration during the first month of pregnancy there was an  
2 increased risk of IUGR (OR: 1.03 [95% CI: 1.00-1.05]) and this was of borderline significance when CO  
3 was averaged over the first trimester (OR: 1.02 [95% CI: 1.00-1.05]). This result persisted after  
4 controlling for other gaseous pollutants. Conversely, maternal exposure to CO was not associated with  
5 LBW. The more recent of these 2 studies examined 386,202 singleton term births (37-42 weeks gestation)  
6 in Calgary, Edmonton and Montreal between 1986 and 2000 (Liu et al., 2007). The study examined  
7 monthly and trimester exposures to CO with IUGR being the only endpoint. A 0.5 ppm increase in 24-h  
8 CO concentration was associated with an increased risk of IUGR in the first (OR: 1.09  
9 [95% CI: 1.07-1.11]), second (OR: 1.07 [95% CI: 1.05-1.09]), and third trimesters (OR: 1.09  
10 [95% CI: 1.07-1.11]) of pregnancy. This result translated to CO exposure having a positive effect on  
11 IUGR within each individual month of pregnancy with the highest effect during the first and last months.  
12 This result persisted after controlling for concurrent NO<sub>2</sub> and PM<sub>2.5</sub>.

13 Two studies in Sao Paulo, Brazil, a city with notably high levels of air pollution (mean CO  
14 3.7 ppm) investigated associations between maternal exposures to CO in relation to reduced birth weight  
15 and LBW within two consecutive time periods and found similar results. In both studies the exposures  
16 were derived from a city-wide average across the available monitoring sites. The first study examined  
17 179,460 singleton term births during 1997 and found that a 0.75 ppm increase in 8-h CO concentration  
18 averaged over the first trimester was associated with a 17.3 g (95% CI: 31.0-3.7) reduction in birth weight  
19 (Gouveia et al., 2004). The second of these studies examined 311,735 singleton births (37-41 weeks  
20 gestation) between 1998 and 2000 and reported a 6.0 g reduction in birth weight associated with a  
21 0.5 ppm increase in 24-h CO concentration averaged over the first trimester (Medeiros and Gouveia,  
22 2005). It is important to note that neither of these studies found an association between CO exposure and  
23 an increased risk of LBW. Therefore, despite CO during the first trimester being associated with reduced  
24 birth weight, it was not associated with LBW.

25 Similar to the two studies in Sao Paulo, Brazil, researchers in Seoul, South Korea conducted two  
26 studies using data from two consecutive time periods. Both of these studies based the exposure estimates  
27 on a city-wide average from all available fixed sites and as would be expected, the results pertaining to  
28 CO were similar for both studies. For example, Ha and colleagues (2001) examined maternal exposures to  
29 CO during the first and third trimesters among 276,763 singleton term births in Seoul between 1996 and  
30 1997. Exposure to CO during the first trimester was associated with a decrease in birth weight of 13.3 g,  
31 which also translated into an increased risk of LBW (RR: 1.10 [95% CI: 1.05-1.14] per 0.5 ppm increase  
32 in 24-h CO concentration). When Lee and colleagues (2003a) extended this study to include singleton  
33 term births (37-44 weeks gestation) for the period of 1996 to 1998 with 24-h CO concentrations averaged  
34 over each month of pregnancy and trimester, CO exposure during the first trimester was associated with  
35 an increased risk of LBW (OR: 1.04 [95% CI: 1.01-1.07] per 0.5 ppm increase). No associations were

1 found in the third trimester for any of the pollutants. Monthly-specific exposures showed that the risk of  
2 LBW tended to increase with CO exposure between months 2-5 of pregnancy.

3 In contrast to other studies reporting that early and late pregnancy are the critical periods for CO  
4 exposure, a Sydney, Australia study of 138,056 singleton births between 1998-2000 reported a reduction  
5 in birth weight of 21.7 g (95% CI: 38.2-5.1) and 17.2 g (95% CI: 33.4-0.9) associated with a 0.75 ppm  
6 increase in maternal exposure to 8-h CO averaged over the second and third trimesters respectively  
7 (Mannes et al., 2005). However, this result did not persist after controlling for other pollutants (PM<sub>10</sub>,  
8 NO<sub>2</sub>) and was only significant when including births where the mother resided within 5 km of a monitor.  
9 Furthermore, this result did not translate to an increased risk of SGA, which was defined as a birth weight  
10 two standard deviations below the mean. The odds ratios for SGA for CO exposures during the first,  
11 second and third trimesters were 0.96 (95% CI: 0.91-1.03), 0.99 (95% CI: 0.92-1.07), and 1.01  
12 (95% CI: 0.93-1.08) respectively. While the majority of studies restrict the analyses to term births as a  
13 method of controlling for gestational age, it is important to note that the Sydney study used all births and  
14 controlled for gestational age in the birth weight analyses and SGA was derived from each gestational age  
15 group.

16 Of all studies reviewed, only two failed to find an association between maternal exposure to CO  
17 and adverse birth outcomes. In northern Nevada, Chen and colleagues (2002) examined CO, PM<sub>10</sub>, and O<sub>3</sub>  
18 exposures among a cohort of 39,338 term births (37-44 weeks gestation) between 1991 and 1999 and  
19 found no association between CO exposure during the entire pregnancy (and each trimester) and a  
20 reduction in birth weigh or an increased risk of LBW. For a 0.75 ppm increase in 8-h CO concentration  
21 averaged over the entire pregnancy there was a reduction in birth weight of 6 g, however it failed to reach  
22 statistical significance. Exposures for this study were based on data from all monitoring sites across  
23 Washoe County, Nevada.

24 In a retrospective cohort study among 92,288 singleton term births (37-44 weeks gestation) in  
25 Taipei and Kaoshiung, Taiwan between 1995-1997, maternal exposures to CO, SO<sub>2</sub>, O<sub>3</sub>, NO<sub>2</sub>, and PM<sub>10</sub> in  
26 each trimester of pregnancy were examined and only SO<sub>2</sub> during the third trimester showed evidence of  
27 attributing to LBW. Exposure assessment was based on data from the monitor closest to the centroid of  
28 the mother's residential district and the final analyses only included mothers whose district centroid was  
29 within 3 km of a monitor. CO exposures were grouped into low (~1.1 ppm), medium (~1.2-15.0 ppm),  
30 and high (>15.0 ppm) and when compared to the lowest exposure group, the odds ratio for LBW in the  
31 highest exposure group was 0.90 (95% CI: 0.75-1.09) for the first trimester, 1.00 (95% CI: 0.82-1.22) for  
32 the second trimester, and 0.86 (95% CI: 0.71-1.03) for the third trimester (Lin et al., 2004a).

33 Table 5-13 provides a brief overview of the birth weight studies. In summary, there is evidence of  
34 ambient CO during pregnancy having a negative effect on fetal growth. From the reviewed studies  
35 Figure 5-7 shows the change in birth weight (grams), Figure 5-8 shows the effect estimates for LBW, and  
36 Figure 5-9 shows the effect estimates for SGA. In general the reported reductions in birth weight are

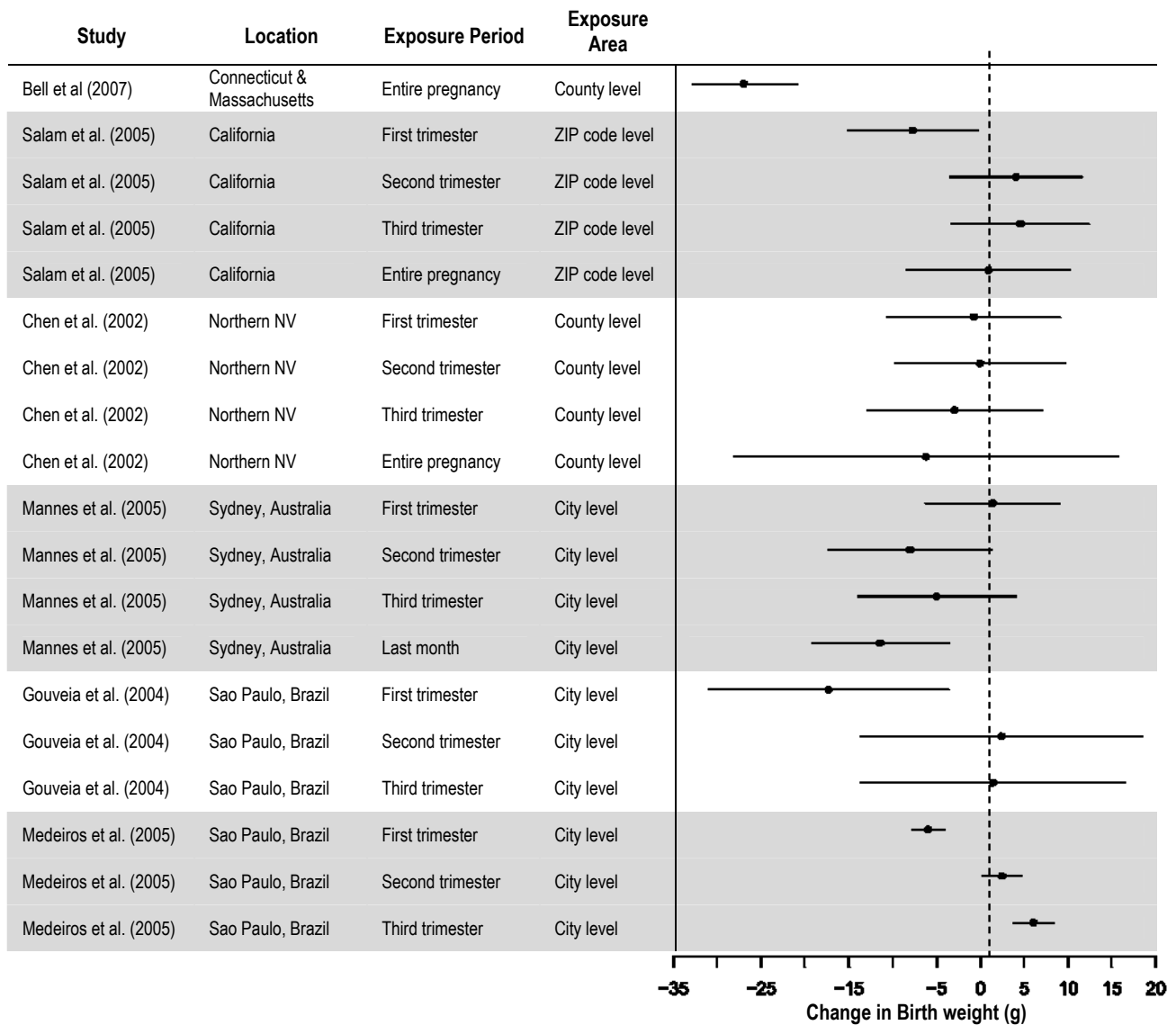
1 small (ranging ~10-20g). It is difficult to conclude whether CO is related to a small change in birth weight  
2 in all births across the population, or a marked effect in some subset of births.

3 Furthermore, there is a large degree of inconsistency across these studies. This may be due to  
4 several factors such as inconsistent exposure assessment and statistical methods employed, different CO  
5 concentrations, and/or different demographics of the birth cohorts analyzed. The main inconsistency  
6 among these findings is the gestational timing of the CO effect. Although the majority of studies reported  
7 significant effects during either the first or third trimester, other studies failed to find a significant effect  
8 during these periods. Several studies found an association with exposure during the entire pregnancy,  
9 providing evidence for a possible accumulative effect; however, these results are inconclusive and this  
10 may be the result of correlated exposure periods.

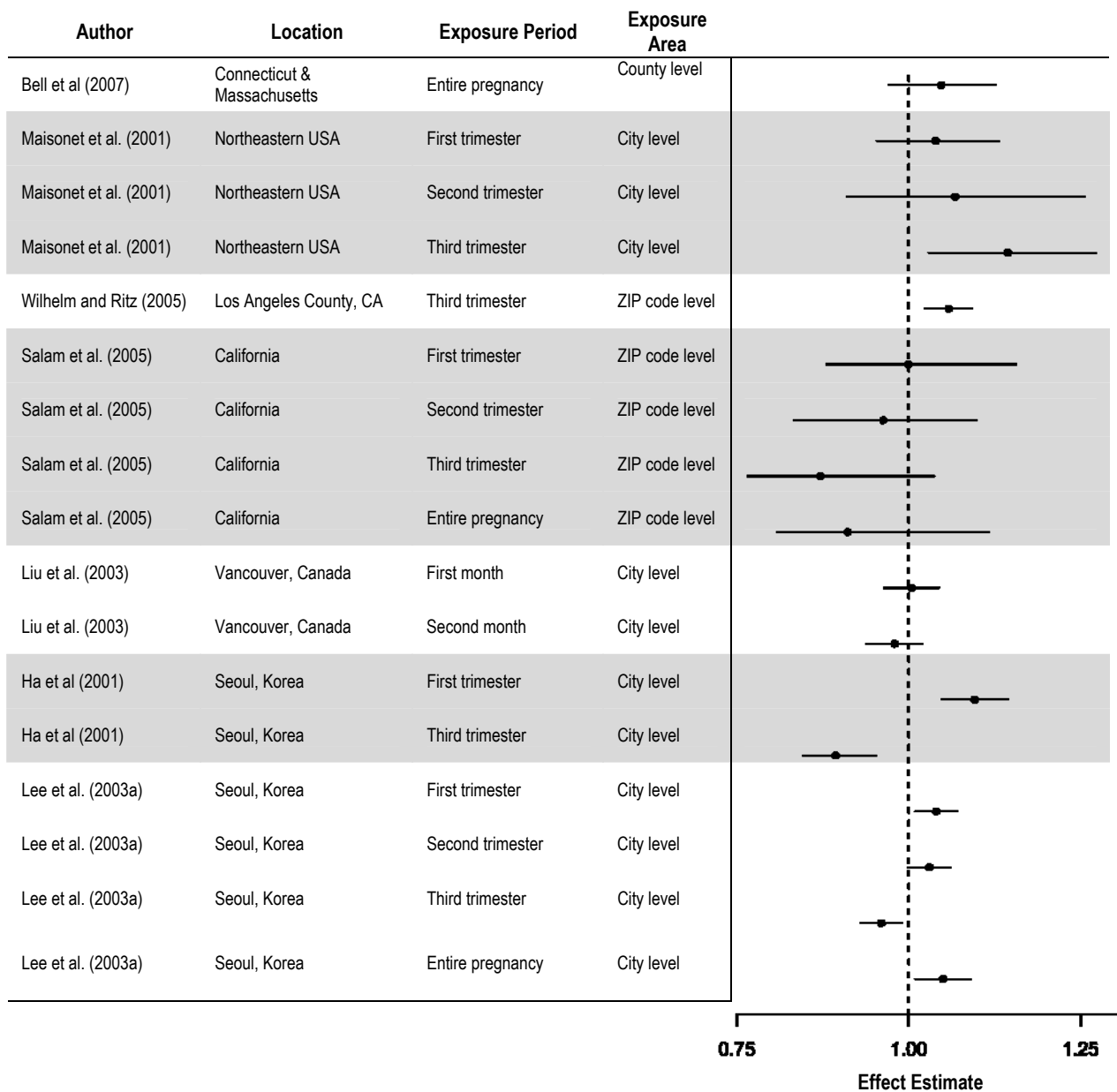
11 Several studies examined various combinations of birth weight, LBW, and SGA/IUGR and  
12 inconsistent results are reported across these metrics. For example, several studies reported an association  
13 between maternal exposure to CO and decreased birth weight yet the decrease in birth weight did not  
14 translate to an increased risk of LBW or SGA. However, it needs to be noted that a measureable change,  
15 even if only a small one, on a population is different than an effect on a subset of susceptible births which  
16 may increase the risk of IUGR/LBW/SGA.

17 The possibility exists that the small reductions in birth weight associated with maternal CO  
18 exposures are the result of residual confounding associated with other factors (e.g., other pollutants,  
19 temperature, and spatial/temporal variation in maternal factors) or other correlated pollutants. For  
20 example, in some studies the CO effect did not persist after controlling for other pollutants (Mannes et al.,  
21 2005; Parker et al., 2005; Wilhelm and Ritz, 2005) while in some studies it did persist (Bell et al., 2007;  
22 Gouveia et al., 2004; Liu et al., 2003, 2007), and other studies did not report results from multipollutant  
23 models (Ha et al., 2001; Lee et al., 2003a; Maisonet et al., 2001; Medeiros and Gouveia, 2005). In  
24 addition, various methods have been employed to control for seasonality and trends (e.g., month of birth,  
25 season of birth, year of birth, smoothed function of time), which may explain some of the mixed results.

26 The two U.S. studies conducted in the Northeast compared results from analyses stratified by race.  
27 The earlier of these studies found an association between CO and LBW among African Americans but not  
28 among whites and hispanics (Maisonet et al., 2001). In contrast, despite reporting an 11g reduction in  
29 birth weight among African-Americans and a 17 g reduction among whites, the more recent of the two  
30 studies found no significant difference between these reductions by race (Bell et al., 2007). Parker and  
31 colleagues (2005) also tested for interactions between race and found no significant association.

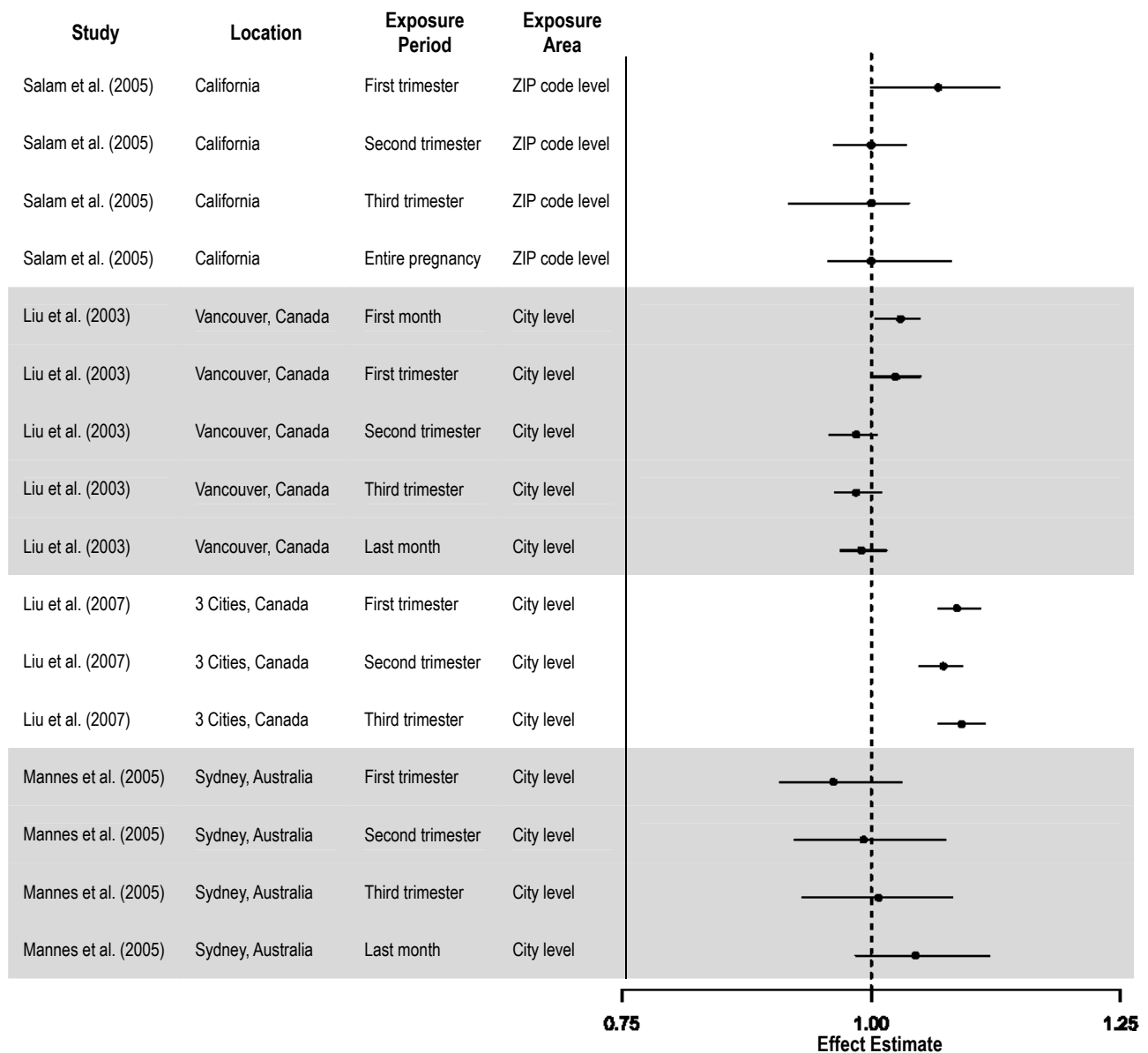


**Figure 5-7. Summary of change in birth weight (95% confidence intervals) associated with maternal exposure to ambient CO. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.**



**Figure 5-8. Summary of effect estimates (95% confidence intervals) for LBW associated with maternal exposure to ambient CO. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.**





**Figure 5-9. Summary of effect estimates (95% confidence intervals) for SGA associated with maternal exposure to ambient CO. Effect estimates have been standardized to a 1 ppm increase in ambient CO for 1-h max CO concentrations, 0.75 ppm for 8-h max CO concentrations, and 0.5 ppm for 24-h avg CO concentrations.**

**Table 5-13. Brief summary of birth weight studies.**

Study	Outcomes Examined	Location (Sample Size)	Mean CO (ppm)	Exposure Assessment	Exposure Windows
<i>UNITED STATES</i>					
Ritz and Yu (1999)	LBW	Los Angeles, CA (n = 125, 573)	2.6 (6-9 a.m.)	<2 miles of monitor	Trimester 3
Wilhelm and Ritz (2005)	LBW	Los Angeles County, CA (n = 136, 134)	1.4 (24 h)	Varying distances from monitor	Trimesters 1, 2, 3
Salam et al. (2005)	Birth weight LBW IUGR	California (n = 3901)	1.8 (24 h)	ZIP code level	Entire pregnancy Trimesters 1, 2, 3
Parker et al. (2005)	Birth weight SGA	California (n = 18,247)	0.75 (8 h)	<5 miles from monitor	Entire pregnancy Trimesters 1, 2, 3
Maisonet et al. (2001)	LBW	Boston, MA; Hartford, CT; Philadelphia & Pittsburg, PA; Washington DC (n = 103,465)	1.1 (24 h)	City wide average	Trimesters 1, 2, 3
Bell et al. (2007)	Birth weight LBW	Connecticut and Massachusetts, (n = 358,504)	0.6 (24 h)	County level average	Entire pregnancy Trimesters 1, 3
Chen et al. (2002)	Birth weight LBW	Northern Nevada, (n = 36,305)	0.9 (8 h)	County level	Trimesters 1, 2, 3
<i>CANADA</i>					
Liu et al. (2003)	LBW IUGR	Vancouver, Canada (n = 229,085)	1.0 (24 h)	City wide average	Trimester 1
Liu et al. (2007)	IUGR	Calgary, Edmonton, Montreal, Canada (n = 386,202)	1.1 (24 h)	City wide average	Trimesters 1, 2, 3
<i>SOUTH AMERICA</i>					
Gouveia et al. (2004)	Birth weight LBW	Sao Paulo, Brazil (n = 179,460)	3.7 (8 h)	City wide average	Trimesters 1, 2, 3
Medeiros et al. (2005)	Birth weight LBW	Sao Paulo, Brazil (n = 311,735)	3.0 (24 h) (Presented in graph)	City wide average	Trimesters 1, 2, 3
<i>AUSTRALASIA</i>					
Ha et al. (2001)	Birth weight LBW	Seoul, Korea (n = 276,763)	1.2 (24 h)	City wide average	Trimesters 1 and 3
Lee et al. (2003a)	LBW	Seoul, Korea (n = ?)	1.2 (24 h)	City wide average	Entire pregnancy Trimesters 1, 2, 3
Mannes et al. (2005)	Birth weight SGA	Sydney, Australia (n = 138,056)	0.8 (8 h)	City wide average and <5 km from monitor	Trimesters 1, 2, 3 Last 30 days
Lin et al. (2004b)	LBW	Taipei, Kaoshiung, Taiwan (n = 92,288)	Taipei 1.1, Kaoshiung 8.1	<3 km of monitor	Entire pregnancy Trimesters 1, 2, 3

### 5.4.1.3. Congenital Anomalies

1           Despite the growing evidence of an association between ambient air pollution and various adverse  
2 birth outcomes, only three studies investigated the effect of temporal variations in ambient air pollution  
3 on congenital anomalies. Given the higher prevalence and associated mortality, heart defects have been  
4 the main focus of two of these three recent air pollution studies. The other study's focus was cleft  
5 lip/palate.

6           The first of these studies was conducted in southern California (Ritz et al., 2002). Exposure to  
7 ambient CO, NO<sub>2</sub>, O<sub>3</sub> and PM<sub>10</sub> during each of the first three months of pregnancy was examined among  
8 births during 1987-1993. Maternal exposure estimates were based on data from the fixed site closest to  
9 the mother's ZIP code area and when using a case-control design where cases were matched to 10  
10 randomly selected controls, results showed that CO during the second month of pregnancy was associated  
11 with cardiac ventricular septal defects. The CO exposures were grouped by quartiles (25th = 1.14,  
12 50<sup>th</sup> = 1.57, 75th = 2.39 ppm) and when compared to those in the lowest quartile exposure group  
13 (<1.14 ppm), the odds ratios for ventricular septal defects across the 3 exposure groups were 1.62  
14 (95% CI: 1.05-2.48), 2.09 (95% CI: 1.19-3.67), and 2.95 (95% CI: 1.44-6.05) respectively. In a  
15 multipollutant model a similar exposure-response pattern was exhibited across the quartiles with the  
16 highest quartile of exposure reaching statistical significance (OR: 2.84 [95% CI: 1.15-6.99]). The only  
17 other pollutant associated with a defect was O<sub>3</sub> during the second month of pregnancy, which was  
18 associated with aortic artery and valve defects.

19           The second study was conducted in Texas (Gilboa et al., 2005), where exposure to ambient CO,  
20 NO<sub>2</sub>, SO<sub>2</sub>, O<sub>3</sub> and PM<sub>10</sub> during the 3rd to 8th week of gestation was examined among births between  
21 1997-2000. Maternal exposure estimates were calculated by assigning the data from the closest monitor to  
22 the mother's residential address. If data were missing on a particular day then data from the next closest  
23 site were used. The median distances from a monitor ranged from 8.6-14.2 km with maximum distances  
24 ranging from 35.5-54.5 km. The main results showed that CO was associated with multiple conotruncal  
25 defects and Tetralogy of Fallot. CO exposures were grouped into quartiles of much lower concentrations  
26 (25th = 0.4, 50th = 0.5, 75th = 0.7 ppm) than the California study and when compared to the lowest  
27 quartile, the odds ratios for conotruncal defects across the 3 CO exposure groups were 1.38  
28 (95% CI: 0.97-1.97), 1.17 (95% CI: 0.81-1.70), and 1.46 (1.03-2.08) respectively without a significant  
29 test for trend (p for trend = 0.0870). Whereas, a strong exposure-response pattern was exhibited across  
30 the quartiles of CO exposure for Tetralogy of Fallot (25th OR: 0.82 [95% CI: 0.52-1.62]; 50th OR: 1.27  
31 [95% CI: 0.75-2.14]; 75th OR: 2.04 [95% CI: 1.26-3.29]; p for trend = 0.0017). The only significant  
32 associations found with other pollutants were between PM<sub>10</sub> and isolated atrial septal defects, and SO<sub>2</sub>  
33 and ventricular septal defects.

1           The last of these three studies was a case-control study that examined maternal exposure to various  
2 air pollutants during the first three months of pregnancy and the risk of delivering an infant with an oral  
3 cleft, namely cleft lip with or without palate (CL/P). Birth data from the Taiwanese birth registry from  
4 2001 to 2003 was linked to air pollutant data that were spatially interpolated from all fixed monitoring  
5 sites across Taiwan. Based on data at the center of the townships or districts, exposure estimates for PM<sub>10</sub>,  
6 SO<sub>2</sub>, NO<sub>x</sub>, O<sub>3</sub>, and CO were averaged over each of the first three months of pregnancy. Interestingly, of  
7 all the pollutants examined, only O<sub>3</sub> during the first two months of pregnancy was significantly associated  
8 with an increased risk of CL/P. In multipollutant models CO was not associated with CL/P (Hwang and  
9 Jaakkola, 2008).

10           The main results from the southern California study showed that CO was associated with an  
11 increased risk of ventricular septal defects and this was exhibited by an exposure-response pattern across  
12 the quartiles of exposure, yet there was no indication that ambient CO concentration in Texas was  
13 associated with ventricular septal defects. Conversely, ambient CO concentration in Texas was associated  
14 with an increased risk of conotruncal defects, yet there was no indication that CO in southern California  
15 was associated with conotruncal defects, and on the contrary, reported results of a protective effect.

16           Interestingly, similar inconsistencies were also found for PM<sub>10</sub> between these studies. For example,  
17 PM<sub>10</sub> in Texas was associated with an increased risk of atrial septal defects, yet there was no indication of  
18 such an effect in southern California where PM<sub>10</sub> concentrations were markedly higher.

19           The authors of the Texas study (Gilboa et al., 2005) provide little discussion toward the  
20 inconsistent results with the southern California study. One suggestion is the different CO concentrations  
21 across the studies with the 75th quartile in southern California being 2.39 ppm while in Texas it was much  
22 lower at 0.7 ppm. However, this suggests that different defects are associated with different  
23 concentrations of CO, yet it still does not explain why particular associations were reported in Texas and  
24 not southern California where concentrations were higher. Similarly, the authors of the Texas study  
25 (Gilboa et al., 2005) also suggested the inconsistency was due to different exposure periods. In Texas the  
26 exposures were averaged over the 3rd to 8th week while in southern California the exposures were  
27 averaged over the second month of pregnancy. However, there was no reason provided as to why this  
28 small difference in the examined exposure period would explain the inconsistent results.

29           Overall, there is little evidence that maternal exposure to CO is associated with an increased risk of  
30 congenital anomalies, namely heart defects and cleft lip and palate. Further research is required to  
31 corroborate these findings.

#### 5.4.1.4. Neonatal and Post-Neonatal Mortality

1 A handful of studies examined the effect of ambient air pollution on neonatal and post-neonatal  
2 mortality with the former the least studied. These studies varied somewhat with regard to the outcomes  
3 and exposure periods examined, and study designs employed.

##### Neonatal

4 In Sao Paulo, Brazil, a time-series study examined daily counts of neonatal (up to 28 days after  
5 birth) deaths for the period of 1998-2000 in association with concurrent day exposure to SO<sub>2</sub>, CO, O<sub>3</sub>,  
6 PM<sub>10</sub>. Moving averages from 2 to 7 days were examined. The mean city-wide CO concentration was  
7 2.8 ppm and there was no association between daily ambient CO and neonatal deaths. Despite CO being  
8 correlated with PM<sub>10</sub> ( $r = 0.71$ ) and SO<sub>2</sub> ( $r = 0.55$ ), only PM<sub>10</sub> and SO<sub>2</sub> were associated with an increase in  
9 the daily rate of neonatal deaths (Lin et al., 2004a).

##### Post-Neonatal

10 Two studies in the U.S. examined the potential association between ambient CO and post-neonatal  
11 (from 28 days to 1 year after birth) mortality and inconsistent results were reported. These studies,  
12 however, varied somewhat in study design.

13 The first of these studies employed a case-control design and examined all infant deaths during the  
14 first year of life among infants born alive during 1989-2000 within 16 km from a monitoring site within  
15 the South Coast Air Basin of California. Exposures for 2-week, 1-month, 2-month, and 6-month periods  
16 before death were linked to each individual death. Extensive analyses were conducted for all-cause infant  
17 deaths, respiratory causes of death, and sudden infant death syndrome (SIDS). Given the long time period  
18 of the data analyzed, in order to alleviate the confounding trends in infant mortality and CO levels this  
19 study was able to match by year (Ritz et al., 2006). Ambient 1-h max CO concentrations averaged over  
20 the 2 months before death were associated with an 11% (OR: 1.11 [95% CI: 1.06-1.16]) increase in risk of  
21 all-cause post-neonatal death (per 1 ppm increase) and a 19% (OR: 1.19 [95% CI: 1.10-1.28]) increase in  
22 risk of SIDS. In the multipollutant models (including PM<sub>10</sub>, NO<sub>2</sub>, O<sub>3</sub>) the positive CO mortality effect  
23 decreased by around 50% and failed to reach statistical significance. Based on exposure from 2 weeks  
24 before death, CO was associated with an increased risk of respiratory related post-neonatal deaths  
25 occurring 28 days to 1 year after birth (OR: 1.14 [95% CI: 1.03-1.25] per 1 ppm increase) and 28 days to  
26 3 months after birth (OR: 1.20 [95% CI: 1.02-1.40]), but no effect was observed for respiratory related  
27 deaths occurring 4-12 months after birth. These results persisted in the multipollutant models and  
28 exposure-response patterns were exhibited across the exposures groupings of 1.02 to <2.08, and  
29  $\geq 2.08$  ppm. To control for gestational age and birth weight the analyses were stratified by 'term/normal-  
30 weight infants' and 'preterm and/or LBW infants.' When these two strata were analyzed, CO was

1 associated with an increased risk of all-cause death and SIDS within both strata (ORs ranged from 1.12 to  
2 1.46). However, these effects did not persist in multipollutant models (Ritz et al., 2006).

3 The second of these 2 studies examined 3,583,495 births, including 6,639 post-neonatal deaths  
4 occurring in 96 counties throughout the U.S. (in counties with more than 250,000 residents) between  
5 1989 and 2000 (Woodruff et al., 2008). Only exposure during the first two months of life was examined  
6 and these were based on an average of CO concentrations recorded across all available monitors within  
7 the mother's county of residence. In contrast to the other postnatal mortality study in California, CO  
8 averaged over the first two months of life was not associated with all-cause death (OR: 1.01  
9 [95% CI: 0.94-1.09] per 0.5 ppm increase in 24-h CO concentration), or with respiratory related deaths  
10 (OR: 1.08 [95% CI: 0.91-1.54] per 0.5 ppm increase in 24-h CO concentration), SIDS (OR 0.85  
11 [95% CI: 0.70-1.04] per 0.5 ppm increase in 24-h CO concentration), or other causes of post-neonatal  
12 mortality (OR: 1.03 [95% CI: 0.96-1.09] per 0.5 ppm increase in 24-h CO concentration). These null  
13 findings may be due to higher error of the exposure assessment at the county-level as opposed to using  
14 data from monitors within close proximity to the residence.

15 The only other postnatal mortality studies have been conducted throughout Asia. Two identical  
16 studies in Taiwan failed to find an association between daily counts of post-neonatal deaths and ambient  
17 air pollutants, including CO. The data analyzed were from the cities of Taipei (Yang et al., 2006) and  
18 Kaohsiung (Tsai et al., 2006b) with ambient CO concentrations being 1.6 ppm and 0.8 ppm respectively.  
19 Both studies examined deaths for the period of 1994-2000 and employed a case-crossover design that  
20 compared air pollution levels 1 week before and after each infant's death.

21 Similarly, another study in South Korea examined post-neonatal mortality for the period of 1995-  
22 1999 using a time-series design. Same-day CO was not associated with all-cause death (RR: 1.02  
23 [95% CI: 0.97-1.06] per 0.5 ppm increase). However, same-day CO was associated with post-neonatal  
24 mortality when the analyses were restricted to respiratory mortality (RR: 1.33 [95% CI: 1.01-1.76] per  
25 0.5 ppm increase) (Ha et al., 2003).

26 In general, the inconsistent exposure periods examined among these studies allows for limited  
27 direct comparison and interpretation. Nevertheless, there is limited evidence that CO is associated with an  
28 increased risk of infant mortality during the post-neonatal period. The exposure periods examined varied  
29 from the same-day CO to lag periods up to a 6 month period prior to birth with one study alternatively  
30 exploring exposures averaged over the first two months of life. Furthermore, given that birth weight and  
31 gestational age are strong predictors of infant mortality, in all of the reviewed studies these factors have  
32 not been considered at either the design or analysis stage. Hence, the link between fetal exposures,  
33 neonatal exposures, and post-neonatal exposures, and the possible interaction that birth weight and  
34 gestational age may have on the results yielded from these examined exposure periods, needs further  
35 attention within this field of research.

### **5.4.1.5. Summary of Epidemiologic Studies of Birth Outcomes and Developmental Effects**

1           There is some evidence that CO during early pregnancy (e.g., first month and first trimester) is  
2 associated with an increased risk of PTB. Only two studies examined the effects of CO on birth defects.  
3 Both studies found maternal exposure to CO to be associated with an increased risk of cardiac birth  
4 defects. This insult to the heart is coherent with the CO effects on the heart characterized in Section 5.2.  
5 In general, there is limited evidence that CO is associated with an increased risk of infant mortality during  
6 the post-neonatal period.

7           There is evidence of ambient CO during pregnancy having a negative effect on fetal growth. In  
8 general, the reviewed studies (Figures 5.7 through 5.9) reported small reductions in birth weight (ranging  
9 ~10-20 g). Although the majority of studies reported significant effects during either the first or third  
10 trimester, other studies failed to find a significant effect during these periods. Several studies examined  
11 various combinations of birth weight, LBW, and SGA/IUGR and inconsistent results are reported across  
12 these metrics. For example, six studies reported an association between maternal exposure to CO and  
13 decreased birth weight yet the decrease in birth weight did not translate to an increased risk of LBW or  
14 SGA. It should be noted that having a measurable, even if small, change in a population is different than  
15 having an effect on a subset of susceptible births, which may increase the risk of IUGR/LBW/SGA. It is  
16 difficult to conclude if CO is related to a small change in birth weight in all births across the population,  
17 or a marked effect in some subset of births.

## **5.4.2. Toxicological Studies of Birth Outcomes and Developmental Effects**

### **5.4.2.1. Birth Outcomes**

#### **Decreased Birth Weight**

18           Multiple reports have been published associating low level CO exposure in laboratory animals and  
19 decrements in birth weight; some of these studies noted reduced growth rates persisting in the neonatal  
20 period. Prigge et al. (1977) saw significant decreases in near-term fetal body weight (GD21) after 21 days  
21 of continuous exposure of pregnant Wistar rats to CO (125, 250, or 500 ppm). Fechter and Annau (1977)  
22 exposed pregnant rats to 150 ppm CO continuously during gestation via inhalation and found 5%  
23 significantly decreased birth weights at PND1 in gestationally exposed pups versus control animals with  
24 weight decrements persisting to weaning; lactational cross fostering did not ameliorate the reduced  
25 growth rates. Dams exposed to CO during gestation had COHb over gestation of 15% with control dams

1 having less than 1%. Decreased birth weight and pre-weaning weight were seen in CO-exposed pups  
2 despite a lack of weight decrement in CO-exposed dams versus air-exposed control dams.

3 Penney et al. (1983) exposed pregnant rats to CO (157, 166, and 200 ppm) over GD6-GD19 and  
4 found significant decreases in near term fetal rat weight at GD20; gestation in rats is ~ 22 days. Carmines  
5 et al. (2008) exposed Sprague-Dawley rats to ~600 ppm CO via nose-only inhalation (levels similar to  
6 those seen in cigarette smoke) during GD6-GD19 of gestation for 2 h/day and found significant decreases  
7 in birth weight (0.5 g or 13%) of exposed pups versus controls. Dam COHb was 30% immediately after  
8 exposure. The half life in rats at this exposure level is ~ 42 minutes. Maternal body weight was  
9 unchanged during gestation, but corrected terminal body weight (body weight minus uterine weight) was  
10 significantly elevated in CO-exposed dams at term. There were no external malformations (teratogenicity)  
11 seen. Singh et al. (1992; 2006) found significant decreases in birth weight in gestationally CO-exposed  
12 mouse pups (65, 125, 250 or 500 ppm) in two studies. In the 2006 study, mice were exposed to CO for 6  
13 h/day for the first 2 weeks of pregnancy, and in the 1992 study, dams were exposed to CO (GD0-GD18)  
14 longer with fetuses collected at GD18. Astrup et al. (1972) found significant decreases (11 and 20%,  
15 respectively) in birth weight of rabbits exposed to either 90 or 180 ppm CO continuously over the  
16 duration of gestation. Fechter and Annau (1977) reported decreased birth weight, albeit not significant, of  
17 Long-Evans rats exposed in utero to 150 ppm CO continuously throughout gestation (dam COHb 15%),  
18 but by PND4 and through pre-weaning to PND21, exposed pups weighed significantly less than controls.  
19 Tolcos et al. (2000a) found significant decreases in body weight and crown to rump length in guinea pig  
20 fetuses after being exposed to 200 ppm CO for 10h/day from GD23-GD25 until GD61-GD63, at which  
21 time the fetuses were collected (term ranges from GD68 to GD72). However, in other studies, Tolcos  
22 found no significant differences in birth weight of guinea pig pups after a similar exposure (GD23-GD25  
23 to term). During pregnancy, fetal and maternal COHb levels were 13% and 8.5%, respectively.

### **Pregnancy Loss and Perinatal Death**

24 Schwetz et al. (1979) exposed CF-1 mice or New Zealand rabbits to 250 ppm CO for either 7 h/day  
25 or 24 h/day over GD6-GD15 (mice) or GD6-GD18 (rabbits), yielding 4 exposure paradigms. The fetuses  
26 were then collected by C-section at the termination of exposure, which was near term. Maternal COHb in  
27 the 7 h/day exposure groups was approximately 10-15% COHb in rabbits and mice; COHb was not  
28 followed in the 24 h exposure groups. The mice exposed to 7 h/day CO had a significant increase in the  
29 number of resorbed pups (not seen in 24 h/day CO exposure). Fetal mouse weight was significantly  
30 greater than control in the 7 h exposures and significantly less in the 24 h exposure groups with  
31 corresponding significant differences in crown to rump length in the two groups. Rabbits seemed to be  
32 less affected by CO exposure manifesting with no significant perinatal death or pregnancy loss.

33 Astrup et al. (1972) studied the effect of CO on fetal development after continuous CO exposure  
34 (90 or 180 ppm CO) over the duration of gestation in rabbits. COHb was 16-18% and 8-9% in the 180



1 and 90 ppm exposure groups, respectively. In the immediate neonatal period, 24 h postpartum, 35%  
2 (180 ppm) and 9.9% (90 ppm) of CO-exposed animals died. In the postpartum period after the first 24  
3 hours and extending out to PND21, 90 ppm CO-exposed pups experienced 25% mortality versus 13% in  
4 controls; there was no difference from control at the 180 ppm CO exposure level. Gestation length was  
5 unchanged with CO exposure.

6 Fechter and Annau (1977) exposed Long-Evans rats in utero to 150 ppm CO continuously through  
7 gestation (dam COHb 15%) and saw no effects of CO on litter mortality or pup number at PND1.

## **Maternal Diet**

### *Maternal Protein Intake and Neonatal Mouse Mortality*

8 Pregnant CD-1 mice were exposed intermittently (6 h/day for first two weeks of pregnancy) to CO  
9 (0, 65, or 125 ppm) in combination with protein modified diets [27% (supplemental protein), 16%  
10 (control), 8% (low), or 4% (very low protein)] to assess the role of CO exposure coupled with low-protein  
11 diet on neonatal mortality at 1 week of age (Singh, 2006). Litter size was not affected by CO exposure.  
12 Pup weight was inversely related to CO exposure and directly related to dam diet protein content during  
13 pregnancy. Pup mortality at birth was directly related to CO exposure in certain protein groups  
14 (supplemental, and 4% protein) and inversely related to the dam's dietary protein content. At 1 week of  
15 age, pup mortality was significantly increased by CO-exposure. Dietary protein restriction also induced  
16 pup mortality at 1 week of age; all pups in the 4% protein diet died by 1 week of age. CO exposure  
17 (65 ppm only) combined with a normal protein diet significantly increased pup mortality at 1 week.  
18 Animals receiving supplemental protein diets with CO exposure (65 and 125 ppm) had pups that had  
19 significant increases in mortality at 1 week versus control air pups. Control protein diet pups had  
20 significantly increased pup mortality at 1 week with CO exposure (65 ppm only) versus control air  
21 animals (0 ppm CO). Contrary to other findings, low protein diet (8%) combined with CO (125 ppm) led  
22 to a slight yet significant decrease in pup mortality at 1 week of age versus control (0 ppm CO). In  
23 summary, these data show that in utero CO exposure induced increased neonatal mouse deaths at 1 week  
24 in supplemental protein and normal protein diet exposure groups.

### *Maternal Low Protein Diet and CO-Dependent Teratogenicity*

25 The role of diet as a contributor to teratogenicity of CO (0, 65, 125, or 250 ppm CO) in CD-1 mice  
26 given a various protein diets (27%, 16%, 8% or 4% protein) during pregnancy was explored by Singh et  
27 al. (1992). Timed pregnant CD-1 mice were exposed continuously to CO from GD8-GD18 of pregnancy  
28 at which point animals were sacrificed and fetuses collected. Subsequent work by this group has shown  
29 that low protein diets plus CO exposure act in an additive fashion to increase placental COHb in mice  
30 (Singh et al., 1992; 2003). As expected, all levels of CO and the lowest protein diet (8 or 4% protein)

1 given to the dams during gestation resulted in significantly decreased fetal weight of normal fetuses at  
2 GD18. CO exposure did not produce maternal toxicity except for a significant decrease in maternal  
3 weight at GD18 with 4 and 8% protein diets versus control diet in non-CO-exposed animals. Dam dietary  
4 protein levels were inversely related to gross malformations including jaw changes. All concentrations of  
5 CO exposure significantly increased the percentage of litters with malformations. There was a CO dose-  
6 dependent increase in percent of litters with malformations within each maternal dietary protein level.  
7 Skeletal malformations were present in offspring with the percent of litters affected inversely related to  
8 dietary protein levels. CO exposure concomitant with a low protein diet exacerbated the percent of  
9 skeletal malformations in offspring. The percent of dead, resorbed, or grossly malformed fetuses was  
10 directly related to CO concentration and inversely related to maternal dietary protein levels. CO and  
11 maternal dietary protein restriction had a synergistic effect on mouse offspring mortality and an additive  
12 effect on malformations.

### *Zinc-Supplemented Protein-Deficient Maternal Diet and Mortality and Teratogenicity*

13 Further studies by Neggers and Singh (2006) explored how teratogenicity and fetal mortality were  
14 affected by zinc (Zn) modulation in CO-exposed pregnant dams (CD-1 mouse) given protein insufficient  
15 diets. Developmental toxicity of CO was attenuated by protein supplementation, i.e., protein  
16 supplemented animals (27%) showed a significantly lower incidence of fetal mortality versus 8% and  
17 16% protein groups. Further, dietary restriction of both protein and Zn with co-exposure to CO during  
18 gestation increased the incidence of pup mortality and malformations including gastroschisis.

19 Earlier studies by Schwetz et al. (1979) found fetal skeletal alterations in lumbar ribs or spurs were  
20 significantly increased in the 7 h/day and 24 h/day gestationally CO-exposed fetal CF-1 mice collected  
21 near term after CO exposure (250 ppm), over GD6-GD15 (dam gestational COHb 10-15% for 7 h/day  
22 exposure, 24 h/day dam COHb not measured); these changes were not seen in similarly exposed fetal  
23 rabbits (1979).

24 Astrup et al. (1972) studied the effect of CO exposure on fetal rabbit development via continuous  
25 CO exposure (90 or 180 ppm with gestational dam COHb of 9% and 17%, respectively) over the duration  
26 of gestation. Three pups (n = 123) in the 180 ppm CO group had deformities in their extremities at birth,  
27 whereas no control and no 90 ppm CO-exposed animals manifested with this malformation.

### *Skeletal Abnormalities*

28 Earlier studies by Schwetz et al. (1979) found fetal skeletal alterations in lumbar ribs or spurs were  
29 significantly increased in the 7 h/day and 24 h/day gestationally CO-exposed fetal CF-1 mice collected  
30 near term after CO exposure (250 ppm), over GD6-GD15 (dam gestational COHb 10-15% for 7h/day  
31 exposure, 24 h/day dam COHb not measured); these changes were not seen in similarly exposed fetal  
32 rabbits.

1 Astrup et al. (1972) studied the effect of CO exposure on fetal rabbit development via continuous  
2 CO exposure (90 or 180 ppm with gestational dam COHb of 9 and 17%, respectively) over the duration  
3 of gestation. Three pups (n = 123) in the 180 ppm CO group had deformities in their extremities at birth,  
4 whereas no control and no 90 ppm CO-exposed animals manifested with this malformation.

## **Endogenous CO**

### *Spontaneous Abortions*

5 Idiopathic spontaneous abortions are more frequent in women with HO-1 polymorphisms in their  
6 genome (Denschlag et al., 2004). To evaluate the role of HO-1 in spontaneous abortion, a mouse model  
7 that spontaneously undergoes abortion (CBA/J x DBA/2J mice) (Zenclussen et al., 2006) was used with  
8 and without HO adenovirus treatment to see if pregnancy outcome could be modulated by changing HO  
9 concentration. HO-1 is known to protect organs from rejection (Kotsch et al., 2006) and thus, HO activity  
10 may protect the developing fetus from rejection by the non-self maternal immune system. Pregnancy  
11 outcome was significantly better (abortion rate significantly decreased) in mice overexpressing HO due to  
12 adenovirus transfer. Thus, in this model, it appears that upregulation of the HO/CO system is able to  
13 protect the developing fetus from spontaneous abortion.

### *Depressed Vascular Reactivity during Pregnancy and CO*

14 CO through the production of soluble guanylate cyclase is able to stimulate the relaxation of  
15 vascular smooth muscle (Villamor et al., 2000). Further, the role of HO expression is important in the  
16 maintenance and outcome of pregnancy and lactation. During pregnancy, there is increased blood volume  
17 without a concurrent increase in systemic BP; this is accomplished by a decrease in total peripheral  
18 vascular resistance, to which CO contributes (Zhao et al., 2008). In humans, genetic polymorphisms in  
19 HO-1 (microsatellite polymorphisms associated with altered HO-1 transcription) are linked to idiopathic  
20 recurrent miscarriages (Denschlag et al., 2004) and administering HO-inhibitors to pregnant rodents  
21 induced total litter loss (Alexandreaanu and Lawson, 2002). Various pathologies of pregnancy including  
22 IUGR and pre-eclampsia are associated with significant decreases in placental HO activity (Denschlag et  
23 al., 2004; McLaughlin et al., 2003). Thus, the HO/CO system appears to be crucial in maintaining  
24 pregnancy.

### *CO and Vasular Relaxation during Pregnancy*

25 Isolated rat aortic rings and tail artery rings from pregnant dams (Sprague-Dawley rats) can be  
26 relaxed by submersion in exogenous CO solutions (Longo et al., 1999). Further, the administration of the  
27 HO inhibitor SnMP induced increased maternal BP (systolic, diastolic, and mean arterial pressure) during  
28 pregnancy in FVB mice (Zhao et al., 2008). Zhao also showed pregnancy induced increased total body  
29 CO excretion as measured in inhalation chambers, and that this increased CO production could be

1 significantly decreased by SnMP administration. Pregnant dam abdominal aortas (AA) are significantly  
2 dilated with pregnancy and SnMP treatment leads to AA vasoconstriction to levels similar to non-  
3 pregnant mice. Zhao looked at blood flow changes using Doppler technology and observed a significant  
4 increase in uterine artery blood flow velocity following SnMP administration. Thus, it appears that the  
5 increased CO production during pregnancy may partially account for the decreased peripheral vascular  
6 resistance seen in pregnancy that prevents the increased blood volume of pregnancy from affecting BP.

### *Placenta*

7 Women living at high altitude (chronic hypoxia exposure), women with pre-eclampsia, or women  
8 who had pregnancies with fetal growth restrictions (FGR) produced term placenta with significant  
9 decreases in HO-2 versus women living at lower altitude with uncomplicated pregnancies (Barber et al.,  
10 2001; Lyall et al., 2000). FGR and pre-eclampsia HO-2 changes were detected in endothelial cells; HO-1  
11 immunostaining was very low in the placenta. Women living at high altitude have an increased risk of  
12 adverse pregnancy outcomes versus women living at lower altitudes (Zamudio et al., 1995). Oxygenation  
13 is important in early pregnancy and triggers trophoblast invasion of the spiral arteries (Kingdom and  
14 Kaufmann, 1997). Isolated human placenta exposed to solutions containing CO demonstrated a  
15 concentration-dependent decrease in perfusion pressure (Bainbridge et al., 2002) further demonstrating  
16 the role of CO in maintaining basal vasculature tone.

17 The endogenous generation of CO in the placenta has been demonstrated in chorionic villi of term  
18 placenta (McLaughlin et al., 2001) with HO activity highest in the placenta near term (McLean et al.,  
19 2000). CO can be generated from multiple endogenous sources; the source of CO in term human placental  
20 chorionic villi was found to be the catalysis of heme by HO and not endogenous lipid peroxidation  
21 (Ahmed et al., 2005). Term human placental cell types including syncytiotrophoblasts and  
22 cytotrophoblasts were grown in cell culture under basal and hypoxic conditions to explore changes in HO  
23 expression (Newby et al., 2005). HO-1 was expressed at significantly lower amounts in  
24 syncytiotrophoblasts versus cytotrophoblasts at normoxic conditions. HO-1 was unchanged in  
25 cytotrophoblasts under hypoxia, but HO-1 was significantly decreased in hypoxic syncytiotrophoblasts.  
26 HO-2 was unchanged in either cell type with hypoxia. These cell culture data can give insight into what  
27 cell types might be responsive to hypoxia through the HO/CO system in the human placenta. Nonetheless,  
28 these data should be interpreted remembering that isolated cell culture models lack the interaction with  
29 the intact placenta and may lead to different physiological outcomes.

### *Uterus at Parturition*

30 The addition of exogenous CO to isolated human and rat uterine tissue during pregnancy failed to  
31 induce relaxation of uterine tissue (Longo et al., 1999). Thus, exogenous CO failed to quiet the  
32 spontaneous contractility of rat or human myometrium (uterine smooth muscle). CO is not able to relax

1 all types of vascular smooth muscle (Brian et al., 1994), and pregnancy appears to modulate the response  
2 of tissues to CO (Katoue et al., 2005).

3 HO-1 induction and HO concentration have been shown to be regulated by estrogen in the rat  
4 (Sprague Dawley) uterus (Cella et al., 2006) during pregnancy and in non-gravid rats. This agrees with  
5 work by Tschugguel et al. (2001) in which CO was generated by primary endothelial cells from human  
6 umbilical veins and uterine arteries after exogenous 17- $\beta$  estradiol administration.

### *Ovaries*

7 **Ovarian Follicular Atresia.** As a part of normal follicular maturation in the ovaries, the majority of  
8 follicles undergo atresia via apoptosis prior to ovulation. Harada et al. (2004) harvested porcine  
9 granulosa cells from ovaries and found that cells naturally undergoing atresia or cell death more strongly  
10 expressed HO-1 than did successful follicles. Addition of the HO substrate hemin or the HO inhibitor Zn  
11 protoporphyrin IX (ZnPP IX) significantly induced or inhibited granulosa cell apoptosis, respectively. In  
12 this porcine model, HO was able to augment granulosa cell apoptosis allowing for proper follicular  
13 maturation.

14 **Ovarian Steroidogenesis.** HO-1 and HO-2 are localized in the ovaries in rats (Sprague Dawley)  
15 and treatment of rat ovaries in vitro with CrMP, an inhibitor of HO, or with hemin, a substrate for HO  
16 induced steroidogenic changes in the ovaries (Alexandreaunu and Lawson, 2002). CrMP significantly  
17 decreased ovarian production of gonadotrophin-induced androstenedione and progesterone without  
18 affecting estradiol levels. Hemin treatment caused androstenedione and estradiol production from rat  
19 ovaries in vitro. Thus, the HO/CO pathway may play an important role in rat ovarian steroidogenesis.

### *Anterior Pituitary and Heme Oxygenase*

20 HO-1 and HO-2 are expressed in rat (Sprague Dawley) anterior pituitary and the secretion of  
21 gonadotropins and prolactin is affected by HO inhibitors and HO substrates (Alexandreaunu and Lawson,  
22 2003a). The estrogen-induced afternoon surge of LH was advanced forward in time by chronic  
23 administration of the HO inhibitor CrMP and this advance could be reversed by concomitant  
24 administration of hemin, a HO substrate. The serum FSH surge was unaffected by CrMP or hemin but in  
25 vitro treatment of GnRH-stimulated pituitaries with hemin led to a significant increase in FSH release.  
26 The estrogen-dependent afternoon prolactin surge was inhibited or delayed by CrMP and CrMP+hemin  
27 and CrMP significantly decreased prolactin release. In vitro studies using pituitary explants showed that  
28 LH release was significantly increased by CrMP administration and unaffected by hemin. Modulation of  
29 the HO/CO system in the anterior pituitary of the female rat led to altered secretion of gonadotropins and  
30 prolactin.

### *Lactation and Pregnancy and HO Modulation*

1 The role of the HO/CO system in estrous cyclicity, pregnancy, and lactation were monitored using  
2 a Sprague Dawley rat model (Alexandreaanu and Lawson, 2002). Pregnancy outcomes used dams that  
3 were injected daily (GD5-GD16) with hemin, a HO substrate, or CrMP, a HO inhibitor that were followed  
4 until 2-days post-term at which point they were euthanized and uteri examined for implantation sites and  
5 fetuses. Another group of female rats were monitored for estrous cycle changes after receiving daily  
6 hemin (10 or 11 days of injections) or CrMP (11 or 14 days of injections). A third group of animals was  
7 monitored during lactation after receiving either CrMP or hemin exclusively during lactation (lactation  
8 days LD5-LD15); the lactational effect was monitored by weighing pups out to weaning (LD21). Hemin  
9 had no effect on any of the outcomes in this study, except a slight effect on litter weight gain during  
10 lactational exposure. CrMP decreased time in estrous in a dose-dependent manner. CrMP exposure  
11 significantly increased uterine fetal resorptions and led to no live births, possibly due to vasoconstriction  
12 and associated ischemia of the placental vascular bed. CrMP induced decreased litter weight gain during  
13 lactation, which the authors attributed to decreased maternal milk production or ejection problems as  
14 cross-fostered pups regained weight lost during nursing on CrMP dams. The lactational effects seen in  
15 this model may be explained by changes in prolactin as previously reported (Alexandreaanu and Lawson,  
16 2002). These data indicate that the HO/CO pathway is important in estrous cyclicity, parturition, and  
17 lactation. Nonetheless, CrMP may also inhibit NO• production, a mechanism that is distinct from  
18 CO-dependent effects.

### **Summary of Toxicological Studies on Birth Outcomes**

19 Endogenous CO production by various organs and the vascular system during pregnancy is  
20 important for maintaining pregnancy and for the proper development of the fetus and/or offspring  
21 postnatally. Exogenous CO addition has been shown to alter perinatal effects and birth outcomes in  
22 various toxicological studies. Endogenous CO production is required for maintenance of a pregnancy as  
23 has been shown by use of HO inhibitors; further, genetic polymorphisms in HO-1, which are associated  
24 with increased HO expression upon stress, are associated with idiopathic recurrent miscarriages in a study  
25 of a Caucasian population. A rodent study of spontaneous abortion demonstrated better pregnancy  
26 outcome in animals in which HO-1 levels were increased by adenovirus transfer.

27 Rodent studies have explored the role of exogenous CO exposure on birth outcomes. Exogenous  
28 CO exposure to rodent fetuses in utero significantly increased postnatal mortality. CO exposure in utero  
29 induced teratogenicity in offspring in a dose-dependent manner that was further exacerbated by dietary  
30 protein restriction or zinc depletion. Two studies noted skeletal abnormalities following in utero CO  
31 exposure (180-250 ppm).

32 CO induces cGMP production by vascular smooth muscle beds, which can regulate vascular tone.  
33 Thus, endogenous CO production is associated with maintenance of decreased peripheral vascular

1 resistance during pregnancy that prevents the increased blood volume of pregnancy from affecting BP.  
2 This has been demonstrated in the human placenta and cells from the placenta with exogenous CO  
3 addition further decreasing perfusion pressure. However, exogenous CO addition to human and rat uterine  
4 tissue during pregnancy was unable to relax spontaneous myometrial contraction.

5 Estrogens have been shown to upregulate HO-1 levels in the uterus in non-gravid and pregnant rat  
6 uteri; other studies have shown isolated cell culture systems administered 17-beta estradiol can generate  
7 CO. In the female reproductive tract, the endogenous HO/CO system is involved in proper follicular  
8 maturation, ovarian steroidogenesis, the secretion of gonadotropin and prolactin by the anterior pituitary,  
9 in maintenance of lactation, and estrous cyclicity in rodent studies.

## 5.4.2.2. Developmental Effects

### CNS Developmental Effects

#### *Behavioral*

10 **Active Avoidance Behavior.** To assess behavioral changes after in utero exposure, pregnant Wistar  
11 rats were exposed to CO (0, 75 or 150 ppm) continuously over GD0-GD20 (De Salvia et al., 1995). Male  
12 pups from exposed dams were evaluated for active avoidance behavior (mild shock avoidance) during  
13 acquisition and reacquisition. This work was designed to expand on the studies of Mactutus and Fechter  
14 (1985), who showed delayed acquisition (120 days of age) of an active avoidance task and disruption of  
15 retention at a later test date (360 days) after continuous in utero CO exposure (150 ppm CO, dam COHb  
16 concentrations of  $15.6 \pm 1.1\%$ ), and to determine if these behavioral changes were permanent. De Salvia  
17 et al. (1995) found there were no significant behavioral impairments in the low dose animals (75 ppm).  
18 Animals exposed to the higher dose of CO (150 ppm) in utero had significantly impaired acquisition (at  
19 3 and 18 months of age) and reacquisition (at 18 months of age) of conditioned avoidance behavior. The  
20 authors speculated that this CO-dependent behavioral change may be mediated through neurotransmitter  
21 signaling, specifically changes in dopamine in the neostriatum or nucleus accumbens. These studies  
22 demonstrate that low level CO exposure in utero can lead to permanent behavioral changes in male  
23 offspring.

24 Mactutus and Fechter (1984) also found that acquisition in a two-way conditioned avoidance  
25 (flashing light warnings followed by mild footshock) test failed to improve with age of in utero  
26 CO-exposed (150 ppm, dam COHb 15%) Long-Evans rats (male and female offspring) in contrast to  
27 air-exposed controls who improved with age/maturation, indicating a failure in the associative process of  
28 learning. They also found impairments in reacquisition performance, an index of retention, in PND31 rats  
29 that had received continuous in utero CO exposure. Prenatal CO exposure induced learning and memory  
30 deficits in male and female offspring.

1           **Habituation and Non-Spatial Working Memory.** Investigators have used animal models to study  
2 how low dose CO exposure (75 or 150 ppm) during gestation can affect behavioral outcomes  
3 (habituation, exploration of novel objects, and motor activity) in offspring as adults. Giustino et al. (1999)  
4 exposed primiparous pregnant Wistar rats to CO (0, 75 or 150 ppm) by inhalation from GD0-GD20.  
5 Blood COHb concentrations (mean %  $\pm$  SEM) on GD20 have been reported (0 ppm:  $1.6 \pm 0.1$ ; CO  
6 75 ppm:  $7.36 \pm 0.2$ ; CO 150 ppm:  $16.1 \pm 0.9$ ). At age 40 weeks, male offspring were given two trials. In  
7 the first trial (T1), two similar objects were presented. In the second trial (T2), one object from the first  
8 trial was presented as well as one novel object. Exploration time was defined as time exploring objects  
9 during each trial. Global habituation was quantified as a comparison of the time spent exploring the two  
10 objects in T1 to time spent exploring objects in T2. Discrimination between new and familiar objects was  
11 measured in T2 by contrasting the time spent exploring the familiar object to time spent exploring the  
12 new object. This object recognition tests for the preference that rats have for investigating novel objects in  
13 lieu of familiar objects and is a measurement of non-spatial working memory. The results of this study  
14 showed gestationally CO (75 and 150 ppm) exposed animals at 40 days of age have a significantly  
15 decreased T1 during exploration of novel objects. Global habituation of control and 75 ppm CO rats  
16 showed exploratory times with  $T2 < T1$ . However, T1 and T2 were not significantly different in the  
17 150 ppm CO group, showing a lack of habituation after the second exposure to a previously viewed  
18 object. Also, the decreased time spent with a familiar object by control rats was not seen in CO-exposed  
19 animals (75 or 150 ppm). Further, the authors speculated that the mesolimbic dopaminergic system may  
20 be responsible for these changes, possibly involving the nucleus accumbens. The human literature shows  
21 a possible connection with these CO-dependent rodent effects; infants whose mothers smoked during  
22 pregnancy manifest with habituation defects. Nonetheless, CO is just one of many constituents of  
23 cigarette smoke. The results from these animal toxicology studies showed that in utero exposure to CO  
24 affects non-spatial working memory in young adult male rats.

25           **Long-Term Potentiation and Spatial Learning.** The above study agreed with earlier work showing  
26 that hippocampal long-term potentiation (LTP) (Zhuo et al., 1993) of brain sections was significantly  
27 affected by CO exposure with ZnPPIX, a HO inhibitor, blocking hippocampal LTP, but was in contrast to  
28 Bing et al. (1995) who showed enhanced spatial learning in the Morris water maze in rodents exposed to  
29 the HO inhibitor tin protoporphyrin (Sn-PP). Thus, distinct types of learning may be differentially  
30 regulated by CO exposure, and endogenous CO, as modulated with the above inhibitors, may manifest  
31 with different outcomes versus exogenous CO.

32           **Homing Behavior.** Development of homing behavior, orientation by the rat toward its home cage,  
33 was significantly delayed in rats prenataly exposed continuously to 150 ppm CO (150 ppm) (Fechter and  
34 Annau, 1980).

35           **Locomotor Effects.** Fechter and Annau (1977) exposed Long-Evans rats in utero to 150 ppm CO  
36 continuously through gestation (dam COHb 15%) and saw significant effects of CO on pup locomotor



1 activity at PND1 and PND4 (both after subcutaneous L-DOPA administration to induce movement) and at  
2 PND14, but not at PND21. CO-exposed pups showed significantly less activity than air-exposed controls  
3 through the pre-weaning window.

4 Fechter and Annau (1980) exposed rats to CO (150 ppm) continuously throughout gestation and  
5 found that exposed offspring manifested with poorer than normal performance on the negative geotaxis, a  
6 reflexive response that results in a directional movement with or against gravity. In these studies, negative  
7 geotaxis was defined as performing a 180° turn to face the top of an inclined plane. Continuous prenatal  
8 CO exposure (125 ppm, GD7-GD18) in CD-1 mice impaired negative geotaxis at PND10 in a study by  
9 Singh (1986). The standardization and use of geotaxis as a vestibular, motor or postural metric in infant  
10 rodents has been debated in the literature (Kreider and Blumberg, 2005).

11 Prenatal exposure to CO (125 ppm, GD7-GD18) significantly affected the righting reflex (the  
12 turning of an animal from its supine position to its feet) in exposed CD-1 mice on PND1. Also, the aerial  
13 righting score, or turning 180° and landing on the feet when dropped from the supine position at a height,  
14 was significantly decreased in pups exposed to CO in utero (65 and 125 ppm) (Singh, 1986) at PND14.  
15 These behavioral tests indicated neuromuscular, vestibular or postural effects in the CO-exposed neonate.  
16 Earlier studies by Fechter and Annau (1977) identified an early window of sensitivity for CO-dependent  
17 motor activity deficits of PND1-PND4, with recovery by PND21. Carratu's (2000) PND40 and PND90  
18 monitoring of motor activity of in utero CO-exposed animals may be too late to detect CO-dependent  
19 changes.

### *Neuronal*

20 **Sciatic Nerve Myelination.** In utero exposure (GD0-GD20) to low levels of CO (0, 75 or 100 ppm)  
21 and its effect on sciatic nerve myelination in male offspring was studied in Wistar rats (Carratù et al.,  
22 2000). The dam CO blood concentration expressed as %COHb was determined for 0 ppm (GD10:  $0.97 \pm$   
23  $0.02$ ; GD20:  $1.62 \pm 0.1$ ), 75 ppm (GD10:  $7.20 \pm 0.12$ ; GD20:  $7.43 \pm 0.62$ ), 150 ppm (GD10:  $14.42 \pm$   
24  $0.52$ ; GD20:  $16.08 \pm 0.88$ ). There were significant increases in %COHb in all CO-exposed animals. The  
25 myelin sheath thickness of the nerve fibers was significantly decreased in CO-exposed animals (75 and  
26 150 ppm). Axon diameter was not affected by CO exposure. Even though CO affected myelination, it did  
27 not significantly affect motor activity of CO-exposed mice at 40 and 90 days. In conclusion, in utero  
28 exposure of male rats to CO induces decreased myelination of the sciatic nerve without concomitant  
29 decreases in motor activity.

30 **Brain and Peripheral Nervous System Myelination.** The effect of in utero exposure (GD0-GD20)  
31 of Wistar rats to CO (0 or 150 ppm) on pup neuronal myelination was studied in the peripheral nervous  
32 system and the brain of male offspring at 90 days of age (Carratù et al., 2000). CO exposure during  
33 development induces hypoxia, and hypoxia can induce sphingomyelin changes which could lead to  
34 impaired myelination and motor activity decrements. This study reported maternal COHb (mean %  $\pm$

1 SEM) as  $1.9 \pm 0.04$  and  $16.02 \pm 0.98$  in control and 150 ppm CO-exposed animals, respectively. Prenatal  
2 CO exposure had no effect on brain sphinganine (SA) or sphingosine (SO) levels in male offspring at  
3 90 days of age. However, the sciatic nerve had significant increases in SO after CO exposure, no changes  
4 in SA and a significant decrease in the SA/SO ratio at 90 days of age. Motor activity, which could be  
5 affected by changes in myelination, showed no differences between CO and control animals at 90 days of  
6 age. SO is an intermediate in sphingolipid turnover and SA is an intermediate of de novo sphingolipid  
7 biosynthesis. These results demonstrate that sphingolipid homeostasis in the PNS but not CNS is  
8 interrupted in offspring exposed in utero to CO without a manifestation of changes in motor activity.

### *Neurotransmitter Changes*

#### 9 **Medullar Cholinergic and Catecholaminergic Changes after In Utero CO Exposure and SIDS.**

10 One theory related to Sudden Infant Death Syndrome (SIDS) relates to aberrant development of brain  
11 stem nuclei controlling respiratory, cardiovascular, and arousal activity. To address these changes in an  
12 animal model, Tolcos et al. (2000a) exposed pregnant guinea pigs to CO (0 or 200 ppm) for 10 hours per  
13 day over the last 60% of gestation leading to fetal and maternal COHb levels of 13% (versus control of  
14 0.25%) and 8.5% (versus control of 1.6%), respectively. Guinea pigs are a good model because they  
15 display a similar time course of CNS development to humans with the majority of development occurring  
16 in utero. Pups and sows were collected near term and CO-exposed pups were found to have significant  
17 decrements in body, brain, and liver weights as well as decreased crown to rump length when compared  
18 to control pups. Medullar volume was also significantly decreased in CO-exposed pups. Neurotransmitter  
19 systems were affected after CO exposure. Specifically, the catecholaminergic system of the brainstem  
20 displayed significant decreases in immunoreactivity for tyrosine hydroxylase (TH), which is likely due to  
21 decreased cell number in specific medullar regions. This was consistent with earlier work showing  
22 aberrant respiratory responses (to asphyxia and CO<sub>2</sub>) of offspring with prenatal CO exposure (McGregor  
23 et al., 1998). The cholinergic system is also affected by prenatal CO exposure with significant increases in  
24 choline acetyl-transferase (ChAT) immunoreactivity of the medulla and no changes in muscarinic  
25 acetylcholine receptor. This is in contrast to human infants with SIDS that show decreased brainstem  
26 muscarinic receptor binding (Kinney et al., 1995). ChAT changes in this study (Tolcos et al., 2000a) were  
27 from areas of the medulla associated with tongue innervation, which is crucial to swallowing and which  
28 may be impaired in SIDS infants.

29 **Mesolimbic Dopaminergic Effects and Impaired Sexual Behavior.** Exposure (GD0-GD20) of  
30 pregnant Wistar rats to CO (0, 75 or 150 ppm) and its effects on adult (5 and 10 months of age) male  
31 offspring sexual behavior and mesolimbic dopaminergic function was accessed by Cagiano et al. (1998).  
32 Maternal COHb at GD10 was 1, 7, and 15% (0, 75 and 150 pm CO, respectively) and at GD20 was 1.5, 7,  
33 and 16% (0, 75, 150 ppm CO, respectively). Using the aforementioned parameters, only animals at the  
34 150 ppm CO exposure were significantly affected by CO exposure. At 5 months of age, CO-exposed male

1 offspring showed decrements in sexual behavior including an increase in mount to intromission latency, a  
2 decrease in mount to intromission frequency, and a decrease in ejaculation frequency. Further,  
3 administration of amphetamine, which stimulates copulatory activity, did not change CO-induced  
4 increases in mount to intromission latency or decreases in mount to intromission frequency. Basal  
5 extracellular dopamine concentration in the nucleus accumbens was unchanged after CO exposure.  
6 However, when stimulated with amphetamine administration, control rats had increased release of  
7 dopamine that was absent with CO-exposed rats. This demasculination of the CO-exposed offspring  
8 paralleled earlier studies of mice exposed gestationally to hypoxia (Hermans et al., 1993). When rats were  
9 followed at ten months of age, there were no significant changes in copulatory activity or neurochemical  
10 parameters of the CO-exposed male offspring showing a recovery from earlier decrements. In summary,  
11 in utero exposure to CO induced impairment of copulatory sexual behavior in male offspring with  
12 accompanying changes in the mesolimbic dopaminergic system at five months of age. By ten months of  
13 age, these changes were no longer detectable in CO-exposed males.

14 **Dopamine in the Neostriatum Is Affected by In Utero Plus Perinatal CO Exposure.** Exposure of  
15 Long Evans rat dams and pups continuously to CO (75, 150, or 300 ppm with maternal COHb of 11, 19,  
16 and 27%, respectively) from conception to PND10 induced significant elevations in dopamine in the  
17 striatum at PND21 in CO-exposed offspring versus air exposed controls (Fechter et al., 1987).

18 **Cerebellum Weight and Neurotransmitter Changes after In Utero and Perinatal CO Exposure.**  
19 Long Evans dams and pups exposed to CO (75, 150, or 300 ppm with COHb 9%, 15% and 24%) over the  
20 duration of gestation and out to PND10, during the period of cerebellular development in the rat which is  
21 equivalent to the in utero cerebellar development in the human, yielded a dose-dependent reduction in  
22 cerebellum wet weight (significant at 150 and 300 ppm) at PND21 with increases in norepinephrine (NE)  
23 concentration in the extrinsic noradrenergic system (a system that terminates in the cerebellum) in  
24 CO-exposed animals monitored from PND14 to PND42, which is in contrast to a lack of NE change seen  
25 in the cerebral cortex with in utero only CO exposure (Storm and Fechter, 1985a, b). With the same  
26 exposures as Storm and Fechter, CO-exposed (300 ppm only) pups at PND21 had significant decreases in  
27 cerebellar GABA content, decreased uptake of exogenous radio-labeled GABA, decreased fissures in the  
28 cerebellum and decreased cerebellum size (Storm and Fechter, 1985b). Thus, in these studies, it appeared  
29 that the cerebellum and neurotransmitters of the cerebellum were significantly affected by CO exposure  
30 during cerebellar development.

31 **Neonatal Hyperthermia Effects on Neurotransmitters and Neuroglia.** To explore the interaction  
32 of hyperthermia and hypoxia (CO-induced), two risk factors for SIDS, pregnant guinea pigs were exposed  
33 to CO (0 or 200 ppm) for 10 h/day for the last 60% of gestation. At PND4 male pups were exposed to  
34 hyperthermia or ambient temperature as a control. Brains were then collected at 1 and 8 weeks of age. CO  
35 exposure had no effect on litter size or litter birth weight. In utero CO exposure sensitized some areas of  
36 the brain to future hyperthermic insults. Specifically, CO plus hyperthermia induced significant increases

1 in serotonin in multiple brain regions (NTS, DMV, and hypoglossal nucleus) at 1 week of age; this change  
2 was no longer evident at 8 weeks of age. Hyperthermia exposure alone induced decreased met-enkephalin  
3 neurotransmitter immunoreactivity at 1 week of age that was absent at 8 weeks and absent in CO plus  
4 hyperthermia exposed animals. Brain stem neurotransmitter (met-enkephalin, serotonin, TH, substance P)  
5 immunohistochemical differences were not apparent with CO treatment alone. At 8 weeks of age, CO  
6 plus hyperthermia exposure induced glial aggregations and gliosis surrounding infarct or necrotic areas in  
7 the brain and the medulla lesions stained positive for glial fibrillary acidic protein (GFAP). GFAP  
8 upregulation is classically seen with neuronal diseases or following neurodegeneration. Gross structural  
9 observations revealed no differences in the medulla or cerebellum following in utero CO exposure alone.  
10 Together, these data showed that CO exposure in utero sensitizes the brain to future hyperthermic insults  
11 leading to generation of necrotic lesions in the brain and changes in neurotransmitter levels.

12 **Glutamatergic Transmission Impairment by In Utero CO Exposure.** Pregnant Wistar rats were  
13 exposed continuously to CO (75 ppm) during gestation (GD5-GD20) (Antonelli et al., 2006). Primary cell  
14 cultures obtained from the cerebral cortex of exposed offspring (PND1) had decreased extracellular  
15 glutamate (basal and K<sup>+</sup>-evoked) levels versus air exposed controls, which may functionally impair  
16 cortical glutamatergic transmission in CO-exposed offspring, possibly affecting learning and memory.

### *Electrophysiological Changes*

17 Gestational exposure of pregnant Wistar rats to continuous CO (0, 75 or 150 ppm with dam COHb  
18 for 150 ppm CO of 15%) for the entirety of gestation yielded electrophysiological changes in the  
19 peripheral nervous system with reversible changes (present at PND40 and absent at PND270) in sodium  
20 channel inactivation kinetics and irreversible changes in the sodium equilibrium potential (Carratu et al.,  
21 1993) in CO-exposed pups (75 and 150 ppm) versus control pups. These voltage clamp studies showed  
22 that in utero CO exposure had both reversible and irreversible effects on sodium channels, which are  
23 essential for proper electrophysiological function of the PNS.

### *The Developing Auditory System*

24 The developing auditory system of rodents has recently been recognized to be a target of CO  
25 exposure. Low concentrations of CO exposure via inhalation (0, 12.5, 25, or 50 ppm) over PND8-PND22  
26 during the period of extensive auditory development and synaptogenesis was studied by Webber , who  
27 looked specifically at the inferior colliculus (IC), an auditory integrative section of the midbrain. Rodent  
28 pups were allowed to be maternally reared (MR) or were removed from their respective dams and  
29 nutritionally supported with gastronomy-reared nutrition (AR), an artificial feeding system or received  
30 AR plus CO exposure (ARCO). AR reared pups were fed a milk substitute comparable to natural rat milk  
31 via intragastric cannulation. AR allowed nursing pups to be exposed to CO without possible  
32 CO-dependent developmental changes being confounded by maternal CO co-exposure or lactational

1 exposures. Half of the pups were collected at PND27 and half were collected at PND75-PND77. Brains  
2 were sectioned and stained for c-Fos, a marker of neuronal activation in the nervous system. AR and MR  
3 exposed animals were found to have no significant differences in c-FOS staining. Further, c-Fos  
4 immunoreactivity in the central IC was significantly decreased in the ARCO-exposed animals at both  
5 PND27 and PND75-PND77 over all dose groups (12.5, 25, or 50 ppm CO); immunostaining of other  
6 subregions of the IC were not affected by CO. These studies showed exposure to low concentrations of  
7 CO during development can lead to permanent changes in the auditory system of mice that persist into  
8 adulthood.

9 Others have shown that mild CO exposure early in development impacts cochlear development  
10 (Lopez et al., 2003). Sprague Dawley rats were reared from PND6 to weaning (PND19-PND20) on milk  
11 substitutes using gastronomy-feeding (AR) or maternally reared (MR). These pups were exposed to low  
12 levels of CO (12 or 25 ppm, ARCO) via inhalation from PND6-PND27. At PND27, the animals were  
13 sacrificed and examinations of the cochlea were performed to determine the effect of mild CO exposure.  
14 In ARCO animals, there was no evidence of damage to the inner or outer hair cells. After CO exposure  
15 (25 ppm), there was atrophy or vacuolization of the nerve cells that innervate the inner (not outer) hair  
16 cells. Also, fibers of the 8th cranial nerve at the level of the internal auditory canal of the ARCO animals  
17 exposed to 25 ppm CO had distorted myelination and vacuolization of the axoplasm. Energy production  
18 markers in the organ of corti and spiral ganglion neurons including cytochrome oxidase and NADH-TR  
19 were significantly decreased in 25 but not 12 ppm CO exposure groups versus control (AR and MR).  
20 Expression of the calcium-mediated myosin ATPase in the organ of corti and spiral ganglion neurons was  
21 significantly decreased in the 25 ppm CO exposure group when compared to controls. Together, these  
22 findings show that specific areas of the cochlea were affected after low level developmental CO exposure.

23 Two non-invasive measurements of auditory function, otoacoustic emissions testing (OAE) and  
24 auditory brainstem response (ABR) are routinely performed on the majority of human newborns in  
25 hospitals throughout the U.S. to detect early hearing loss. OAE requires the insertion of a microphone and  
26 earphone into a sleeping newborn's ear; when a sound is generated, an echo is recorded on the  
27 microphone in a functional ear and absent or reduced in an affected ear. The ABR measures brainwaves  
28 generated after exposure to a sound. Otherwise healthy, term human infants born to smoking mothers  
29 have impaired cochlear development, albeit mild, with decreased amplitudes of transient evoked  
30 otoacoustic emissions versus newborns born to non-smokers (Korres et al., 2007); CO is one of many  
31 potential affective components of cigarette smoke.

32 Stockard-Sullivan et al. (2003) used these two functional tests in Sprague-Dawley rat pups  
33 receiving ARCO (12, 25, or 50 ppm CO) to determine how continuous (22 h/day) perinatal CO exposure  
34 (PND6-PND22) functionally affected hearing in the developing rat. OAEs were measured in two ARCO  
35 groups (25 and 50 ppm) at PND22. At 50 ppm, significant reductions in OAE amplitude were detected at  
36 specific frequencies (7.13 and 8.01 kHz). ABR conduction time was not affected in CO-exposed animals

1 (12, 25, 50, 100 ppm). Using another otoacoustic test revealed significant attenuation of the action  
2 potential of the 8th cranial nerve with ARCO exposure (12, 25, and 50 ppm CO) versus AR controls at  
3 PND22, but this is complicated by the finding that AR control animals had significant attenuation of the  
4 8th cranial nerve AP versus MR control animals, implying that artificial diet contributes to AP changes  
5 before CO was introduced. Nonetheless, the ARCO-dependent attenuation of the 8th cranial nerve AP  
6 (versus AR control) was permanent, persisting until adulthood in the 50 ppm CO exposure group (the  
7 only CO group monitored); OAE was not measured in adult animals. These functional tests reported that  
8 neonatal exposure to low concentrations of CO can induce auditory functional changes in rodents.

9 Multiple high dose CO studies have shown a role for ROS in CO-dependent ototoxicity. Studies  
10 using high dose CO (i.p. injection of pure CO [35 mL/kg] inducing 40% COHb) have shown  
11 CO-dependent ototoxicity in adult guinea pigs as measured by loss of threshold of cochlear compound  
12 action potentials (CAP) could be attenuated using the free radical inhibitors PBN (a spin trap) or  
13 allopurinol (a superoxide scavenger), implicating ROS in high dose CO dependent cochlear damage  
14 (Fechter et al., 1997). Fechter et al. (2002) found that noise-induced hearing loss (NIHL) was potentiated  
15 by CO (500 ppm) coexposure. In the same manuscript, Fechter et al. (2002) showed potentiation of NIHL  
16 post-CO exposure (1,200 ppm for 30 minutes in adult rodents) with significant elevations seen in free  
17 radical production above control animals only in the CO+NIHL group; an interesting finding of this study  
18 was that the acute high dose CO exposure did not induce significant increases in free radical generation  
19 with this high dose acute CO exposure alone. A possible mechanism for this high-dose CO cochlear  
20 damage is via glutamate release. Use of an NMDA inhibitor attenuated acute CO-dependent (1,200 ppm  
21 or i.p. injection of pure CO [35 mL/kg]) CAP threshold disruption at 15 minutes post CO-exposure.  
22 Glutamate is a known ROS generator (Liu and Fechter, 1995) . The glutamate receptor blockade did not  
23 afford long-term protection (1 h or 4 weeks post-CO exposure) from high doses of CO (1,200 ppm or i.p.  
24 injection of 35ml/kg pure CO) (Chen et al., 2001; Liu and Fechter, 1995)

25 Cognizant of the ROS-dependent contribution to the auditory changes seen with high dose  
26 CO-exposure in animal toxicology models, Webber et al. investigated whether limiting iron availability  
27 through dietary iron restriction could confer protection against low dose perinatal CO-dependent auditory  
28 decrements. Briefly, rat pups were reared from PND6 through the duration of lactation with gastronomy-  
29 reared nutrition with adequate iron (AR) or AR with iron-deficient diets (ARID) or pups were reared with  
30 their respective dams (MR). Animals exposed to CO were denoted as ARCO and ARIDCO and received  
31 either 25 or 100 ppm CO from PND9 to PND24 with all animals collected on PND27. ARID, ARIDCO,  
32 and ARCO groups were compared to controls (MR and AR) to determine differences in the auditory  
33 system after low iron exposure and/or CO exposure. Coincidentally, ARIDCO mice became anemic but  
34 ARID mice were not. Cochlea were stained for transferrin, an iron-transport protein, and low iron  
35 exposure (ARID and ARIDCO) induced increased transferrin. Neurofilament loss from the spiral  
36 ganglionic neurons and somas after ARCO treatment was rescued (no detectable neurofilament loss) with

1 low iron (ARIDCO); ARID treatment induced no change in neurofilaments. CuZn superoxide dismutase  
2 (SOD1) was significantly increased with CO exposure (ARCO) and rescued in ARIDCO animals; SOD1  
3 was unchanged in low iron only animals (ARID). Low iron treatment or CO exposure alone led to  
4 significant decreases in c-fos positive cell numbers of the central IC, but c-fos levels were unchanged  
5 after low iron diet concomitant with CO exposure (ARIDCO). These authors postulated that ROS  
6 generated via the interaction of peroxide and iron (via the Fenton reaction or Haber Weiss chemistry) can  
7 lead to an oxidative stress which leads to decreased c-fos expression in the IC and decreased numbers of  
8 neurofilaments; decreasing the available iron to the developing animal decreases the total pool available  
9 for ROS generation. Further, the attenuation of the elevated SOD levels with iron restriction post  
10 CO-exposure gives credence to this model. The changes in c-fos are not completely explained as ARID  
11 mice also show significant decreases in c-fos IC staining.

### *Summary of Toxicological Studies on Developmental Central Nervous System Effects*

12 Toxicological studies employing rodent models have shown that low level CO exposure during the  
13 in utero period can adversely affect adult outcomes including behavior, neuronal myelination,  
14 neurotransmitter levels or function, and the auditory system. In utero CO exposure has been shown to  
15 impair active avoidance behavior in male offspring after in utero CO exposure. Other studies have shown  
16 that after in utero CO exposure, adult male offspring manifest with altered behavior including a lack of  
17 habituation and an absence of familiarity with previously viewed objects (non-spatial memory). In two  
18 separate studies, in utero CO exposure (75 and 150 ppm) was associated with significant PNS  
19 myelination decrements without associated changes in motor activity in adult animals. With in utero CO  
20 exposure peripheral myelination was decreased and the sphingolipid homeostasis (150 ppm CO) was  
21 disrupted; CNS myelination was not affected. Multiple studies demonstrated that in utero CO exposure  
22 affected glutamatergic (75 ppm), cholinergic (200 ppm), catecholaminergic (200 ppm), and dopaminergic  
23 (75 ppm) neurotransmitter levels or transmission in exposed male rodents. Possible or demonstrated  
24 adverse outcomes from the CO-mediated aberrant neurotransmitter levels or transmission include  
25 respiratory dysfunction (150 ppm), impaired sexual behavior (150 ppm), and an adverse response to  
26 hyperthermic insults resulting in neuronal damage (200 ppm). Finally, in utero CO exposure has been  
27 shown to affect the developing auditory system of rodents, inducing permanent changes into adulthood.  
28 The IC showed decreased c-Fos staining indicating decreased neuronal activation post-perinatal CO  
29 exposure (12 and 25 ppm CO). The cochlea also showed adverse effects after in utero CO exposure with  
30 atrophy of nerve cells innervating the inner hair cells (25 ppm CO), impaired myelination of the 8th  
31 cranial nerve (25 ppm CO), and decreased energy production in the organ of corti (25 ppm CO). Auditory  
32 functional testing using OAE (50 ppm CO) and attenuation of the 8th cranial nerve AP (12, 25, 50,  
33 100 ppm CO) on rodents exposed perinatally to CO showed that CO-exposed neonates had auditory  
34 decrements at PND22 (OAE and AP) and into adulthood with AP decrements. Further studies limiting

1 iron consumption in the diet protected CO-exposed mice from cochlear damage at the spiral ganglion,  
2 possibly through a diminuation of ROS. Together, these animal studies demonstrate that in utero or  
3 perinatal exposure to CO can adversely affect adult behavior, neuronal myelination, neurotransmission,  
4 and the auditory system in rodents.

## **Cardiovascular and Systemic Developmental Effects**

### *Myocardial Electrophysiological Maturation*

5 A rat model of in utero exposure was employed to study CO effects on the development of cardiac  
6 myocytes. Results demonstrated that in utero CO exposure (150 ppm) alters postnatal cellular  
7 electrophysiological maturation in the rat heart (Sartiani et al., 2004). Specifically, at 4 weeks of age, the  
8 action potential duration (APD) of isolated cardiac myocytes from CO-exposed animals failed to shorten  
9 or mature as did the APD of control animals. Further, the two ion conduction channels  $I_{to}$  (transient  
10 outward current,  $K^+$ -mediated) and  $I_{Ca,L}$  (L-type  $Ca^{2+}$  current), which largely control the rat APD, were  
11 significantly different from control animals after in utero CO exposure, at 4 weeks of age. These  
12 CO-dependent changes resolved by 8 weeks of age, reflecting a delayed maturation. The authors noted  
13 that human SIDS is increased in infants whose mothers are smokers and cigarette smoke is composed of  
14 approximately 4% CO. Further, these authors postulated that a CO-dependent delay in  
15 electrophysiological maturation of the cardiac myocyte (lack of APD shortening) could lead to  
16 arrhythmias and thus could be associated with SIDS deaths. However, no SIDS-like cardiac aberrations  
17 were followed in intact Holter-monitored rats in this study.

### *Heart Morphological changes after In Utero or Perinatal CO Exposure*

18 Multiple authors have reported cardiomegaly following in utero low level CO exposure. Prigge and  
19 Hochrainer (1977) reported increased fetal Wistar rat heart wet weight or cardiomegaly following  
20 continuous in utero CO (60, 125, 250, and 500 ppm) exposure with no decreases in near term fetal  
21 hematocrit or Hb levels seen at exposures below 250 ppm. Fechter et al (1980) found that prenatal  
22 exposure to CO affected cardiac development in exposed offspring. Long Evans rats that were exposed to  
23 CO continuously (150 ppm) during gestation manifested with significant elevations in wet heart weight at  
24 PND1 as well as heart weight in relation to body weight; body weight was significantly decreased in  
25 CO-exposed pups. At no other age measured (PND4, PND14, or PND21) did the CO-exposed pups show  
26 increased heart weight (absolute or relative wet or dry weight) versus control. At PND14, control pups  
27 that had significantly increased body weight, also had significantly increased absolute heart weights (dry  
28 and wet) versus CO-exposed offspring. Dry to wet weight ratios revealed that the increased heart weight  
29 of CO-exposed pups at birth was due to edema or water content. Penney et al. (1982) studied  
30 CO-dependent (500 ppm) cardiomegaly in neonates (continuous CO exposure for 32 days starting at



1 PND1). Other studies of adult male Charles River derived rats exposed to CO for 6 weeks (at 400 or  
2 500 ppm CO) as adults only developed CO-dependent cardiomegaly during exposure that significantly  
3 regressed within a couple of months after termination of CO exposure (Styka and Penney, 1978).

### *Systemic Immune Toxicology after In Utero CO Exposure*

4 In utero exposure (GD0-GD20) of male Wistar rats to relatively low CO (0, 75, or 150 ppm)  
5 concentrations induced reversible changes in macrophage function (Giustino et al., 1993). The killing of  
6 *Candida albicans* (yeast) by splenic macrophages was significantly decreased at PND15 in in utero  
7 CO-exposed male offspring (75 and 150 ppm) but recovered and maintained function by PND21 and  
8 PND60. Macrophage phagocytosis of *Candida albicans* was significantly reduced at PND15 and PND21  
9 in CO-exposed males (150 ppm only) and recovery was seen at PND60. Superoxide production by the  
10 splenic macrophage respiratory burst was significantly decreased at PND15 and PND21 after in utero CO  
11 exposure (150 ppm only) with recovery to control levels at PND60. In summary, CO exposure in utero  
12 leads to a reversible and dose dependent loss of function of splenic macrophages with decreased killing  
13 ability, decreased phagocytosis, and decreased ROS production during the macrophage respiratory burst.

14 Further studies by the same laboratory showed that in utero exposure of male Wistar rats to CO  
15 exposure induced changes in the frequency of splenic immunocompetent cells (Giustino et al., 1994).  
16 Specifically, there was a significant decrease in the number of leucocyte common antigen cells (LCA+)  
17 cells in PND21 male rats exposed during gestation to 150 ppm CO; other cell subpopulations including  
18 macrophages, Major Histocompatibility (MHC) II cells, T and B lymphocytes did not show significant  
19 decreases in cell numbers with with CO exposure.

### *Summary of Toxicological Studies of Cardiovascular and Systemic Development*

20 In utero CO exposure is associated with various adverse, albeit non-persistent, cardiac aberrations.  
21 Exposure to 150 ppm induced a delayed maturation of the cardiac action potential in CO-exposed  
22 offspring. Specifically, APD failed to shorten at 4 weeks of age but shortened to control duration by 8  
23 weeks of age. Also, ion channels in CO-exposed animals, which are related to APD, were significantly  
24 different from controls at 4 weeks but matured to control levels by 8 weeks of age. In other studies,  
25 continuous in utero CO exposure (60, 125, 250, and 500 ppm) induced cardiomegaly at PND1 which was  
26 transient and regressed by PND4. Systemic immunocompromise was displayed in two studies focusing on  
27 splenic cells. In the first study, a reversible and dose dependent loss of function of splenic macrophages  
28 with decreased killing ability, decreased phagocytosis, and decreased ROS production during the  
29 macrophage respiratory burst was observed (150 ppm but not 75 ppm CO). In the second study, the  
30 distribution of splenic immunocompetent cells was skewed because of a decrease in the number of LCA+  
31 cells in PND21 male rats exposed during gestation to 150 ppm CO. In conclusion, in utero exposure to  
32 low doses of CO (60, 125, 150, 250, or 500 ppm) is able to induce transient changes in cardiac

1 morphology, cardiac action potentials, and systemic immunity that may make the immediate neonatal  
2 period a time for a CO-exposed animal more susceptible to other outside stressors.

### 5.4.3. Summary of Birth Outcomes and Developmental Effects

3 The most compelling evidence for a CO-induced effect on birth and developmental outcomes is for  
4 PTB and cardiac birth defects. These outcomes were not addressed in the 2000 CO AQCD, which  
5 included only two studies that examined the effect of ambient CO on LBW. Since then, a number of  
6 studies have been conducted looking at varied outcomes, including PTB, birth defects, fetal growth  
7 (including LBW), and infant mortality.

8 There is limited epidemiologic evidence that CO during early pregnancy (e.g., first month and first  
9 trimester) is associated with an increased risk of PTB. The only U.S. studies to investigate the PTB  
10 outcome were conducted in California, and these reported consistent results whereby all studies reported a  
11 significant association with CO exposure during early pregnancy, and exposures were assigned from  
12 monitors within close proximity of the mother's residential address. Additional studies conducted outside  
13 of the U.S. provide supportive, though less consistent, evidence of an association between CO  
14 concentration and PTB.

15 Very few epidemiologic studies have examined the effects of CO on birth defects. Two of these  
16 studies found maternal exposure to CO to be associated with an increased risk of cardiac birth defects.  
17 This insult to the heart is coherent with results of human clinical studies demonstrating the heart as a  
18 target for CO effects (Section 5.2). Animal toxicological studies provide additional evidence for such an  
19 insult to the heart, and reported transient cardiomegaly at birth after continuous in utero CO exposure (60,  
20 125, 250 and 500 ppm CO), delayed myocardial electrophysiological maturation (150 ppm CO), or  
21 systemic splenic immunocompromise (75 or 150 ppm CO). Toxicological studies have also shown that  
22 exogenous continuous in utero CO exposure (250 ppm) induced teratogenicity in rodent offspring in a  
23 dose-dependent manner that was further exacerbated by dietary protein restriction (65 ppm CO) or zinc  
24 depletion (500 ppm CO). Toxicological studies of exogenous CO exposure over the duration of gestation  
25 have shown skeletal alterations (7 h/day, CO 250 ppm) or limb deformities (24 h/day, CO 180 ppm) in  
26 prenatally exposed offspring.

27 There is evidence of ambient CO exposure during pregnancy having a negative effect on fetal  
28 growth in epidemiologic studies. In general, the reviewed studies, summarized in Figures 5-7 through 5-9,  
29 reported small reductions in birth weight (ranging ~5-20 g). Several studies examined various  
30 combinations of birth weight, LBW, and SGA/IUGR and inconsistent results are reported across these  
31 metrics. It should be noted that having a measurable, even if small, change in a population is different  
32 than having an effect on a subset of susceptible births and increasing the risk of IUGR/LBW/SGA. It is

1 difficult to conclude if CO is related to a small change in birth weight in all births across the population,  
2 or a marked effect in some subset of births.

3 In general, there is limited epidemiologic evidence that CO is associated with an increased risk of  
4 infant mortality during the neonatal or post-neonatal periods. In support of this limited evidence, animal  
5 toxicological studies do provide some evidence that exogenous CO exposure to pups in utero significantly  
6 increased postnatal mortality (7 h/day and 24 h/day, 250 ppm CO; 24 h/day, 90 or 180 ppm CO) and  
7 prenatal mortality (7 h/day, 250 ppm CO).

8 Evidence exists for additional developmental outcomes which have been examined in toxicological  
9 studies, but not epidemiologic or human clinical studies, including behavioral abnormalities, learning and  
10 memory deficits, locomotor effects, neurotransmitter changes, and changes in the auditory system.  
11 Structural aberrations of the cochlea involving neuronal activation (12.5, 25 and 50 ppm CO) and  
12 auditory related nerves (25 ppm CO) were seen in pups after neonatal CO exposure. Auditory functional  
13 testing using OAE (50 ppm CO) and 8th cranial nerve AP amplitude measurements (12, 25, 50, 100 ppm  
14 CO) on rodents exposed perinatally to CO showed that CO-exposed neonates had auditory decrements at  
15 PND22 (OAE and AP) and permanent changes into adulthood with AP (50 ppm CO).

16 Overall, there is limited, though positive, epidemiologic evidence for a CO-induced effect on PTB  
17 and birth defects, and weak evidence for a decrease in birth weight, other measures of fetal growth, and  
18 infant mortality. Animal toxicological studies provide support and coherence for these effects. Both  
19 hypoxic and non-hypoxic mechanisms that could lead to such effects have been proposed in the  
20 toxicological literature (Section 5.1), though a clear understanding of the mechanisms underlying  
21 reproductive and developmental effects is still lacking. Taking into consideration the positive evidence for  
22 some birth and developmental outcomes from epidemiologic studies and the resulting coherence for these  
23 associations in animal toxicological studies, **the evidence is suggestive of a causal relationship**  
24 **between long-term exposures to relevant CO concentrations and developmental effects and birth**  
25 **outcomes.**

## 5.5. Respiratory Effects

### 5.5.1. Epidemiologic Studies with Short-Term Exposure

#### 5.5.1.1. Pulmonary Function, Respiratory Symptoms, and Medication Use

26 The 2000 CO AQCD briefly discussed the potential acute respiratory health effects associated with  
27 short-term exposure to CO. An evaluation of the epidemiologic literature at the time did not find any  
28 evidence of an association between short-term exposure to CO and lung function, respiratory symptoms,

1 or respiratory disease. As a result, the 2000 CO AQCD did not conclude that a causal association exists  
 2 between short-term exposure to CO and respiratory health effects. The following section evaluates the  
 3 current literature that examines the potential association between short-term exposure to CO and  
 4 respiratory health effects. Table 5-14 lists the studies evaluated in this section along with the respiratory  
 5 health outcomes examined and CO concentrations reported.

**Table 5-14. Range of CO concentrations reported in key respiratory morbidity studies that examined effects associated with short-term exposure to CO.**

Author	Location	Years	Health Outcome	Metric	Mean Concentration (ppm)	Middle/Upper Percentile Concentrations (ppm)
Rabinovitch et al. (2004)	Denver, CO Year 1: n = 41 Year 2: n = 63 Year 3: n = 43	11/1999-3/2000; 11/2000-3/2001; 11/2001-3/2002	Pulmonary function; Medication use	24-h avg	1.0	50th: 0.9 75th: 1.2 Maximum: 3.5
Silkoff et al. (2005)	Denver, CO Year 1: n = 16 Year 2: n = 18	1999-2000 (winter); 2000-2001 (winter)	Pulmonary function; Medication use	24-h avg	1999-2000: 1.2 2000-2001: 1.1	1999-2000 50th: 1.10 75th: 1.43 Maximum: 3.79 2000-2001 50th: 0.975 75th: 1.34 Maximum: 2.81
Fischer et al. (2002) <sup>1</sup>	The Netherlands n = 68	March - April <sup>2</sup>	Pulmonary function	24-h avg	0.80	Max: 1.34
Ranzi et al. (2004) <sup>1</sup>	Emilia-Romagna Region, Italy n = 120	2/1999-5/1999	Pulmonary function; Respiratory symptoms; Medication use	24-h avg	Urban: 1.34 Rural: 1.06	NR
Lagorio et al. (2006) <sup>1</sup>	Rome, Italy (n = 29)	5/1999-6/1999; 11/1999-12/1999	Pulmonary Function	24-h avg	Spring: 1.83 Winter: 10.7 Overall: 6.4	Overall Max: 25.1
Penttinen et al. (2001) <sup>1</sup>	Helsinki, Finland n = 57	11/1996-4/1997	Pulmonary function	24-h avg	NR	50th: 0.35 75th: 0.43 Maximum: 0.96
Timonen et al. (2002) <sup>1</sup>	Kuopio, Finland n = 33	2/1994-4/1994	Pulmonary function	24-h avg	0.52	Maximum: 2.43
Chen et al. (1999)	Taiwan n = 941	5/1995-1/1996	Pulmonary function	1-h max; 24-h avg	NR	1-h max Maximum: 3.6
Delfino et al. (2003)	Los Angeles, CA n = 22	11/1999-1/2000	Asthma symptoms	1-h max; 8-h max	1-h max: 7.7 8-h max: 5.0	1-h max 90th: 12.0 Maximum: 17 8-h max 90th: 7.9 Maximum: 10
Slaughter et al. (2003)	Seattle, WA n = 133	12/1993-8/1995	Asthma symptoms; Medication use	24-h avg	NR	50th: 1.47 75th: 1.87
Yu et al. (2000)	Seattle, WA n = 133	11/1993-8/1995	Asthma symptoms	24-h avg	1.6	50th: 1.47 Maximum: 4.18
Schildcrout et al. (2006)	8 North American cities n = 990	11/1993-9/1995	Asthma symptoms; Medication use	24-h avg	NR	50th: 0.63-1.49 75th: 0.77-1.90 90th: 0.95-2.40

Author	Location	Years	Health Outcome	Metric	Mean Concentration (ppm)	Middle/Upper Percentile Concentrations (ppm)
von Klot et al. (2002) <sup>1</sup>	Erfurt, Germany n = 53	10/1996-3/1997	Asthma symptoms; Medication use	24-h avg	0.78	50th: 0.70 75th: 1.04 Maximum: 2.60
Park et al. (2005a)	Incheon, Korea n = 64	3/2002-6/2002	Asthma symptoms; Medication use	24-h avg	Control days: 0.64 Dust days: 0.65	NR
Rodriguez et al. (2007)	Perth, Australia n = 263	6/1996-7/1998	Symptoms associated with respiratory illness	8-h max	1.41	Maximum: 8.03
de Hartog et al. (2003) <sup>1</sup>	Amsterdam, the Netherlands n = 37 Erfurt, Germany n = 47 Helsinki, Finland n = 47	1998-1999 (winter)	Respiratory symptoms	24-h avg	Amsterdam: 0.52 Erfurt: 0.35 Helsinki: 0.35	Maximum: Amsterdam: 1.39 Erfurt: 2.17 Helsinki: 0.87

<sup>1</sup> These studies presented CO concentrations in the units mg/m<sup>3</sup>. The concentrations were converted to ppm using the conversion factor 1 ppm = 1.15 mg/m<sup>3</sup>, which assumes standard atmosphere and ambient temperature.

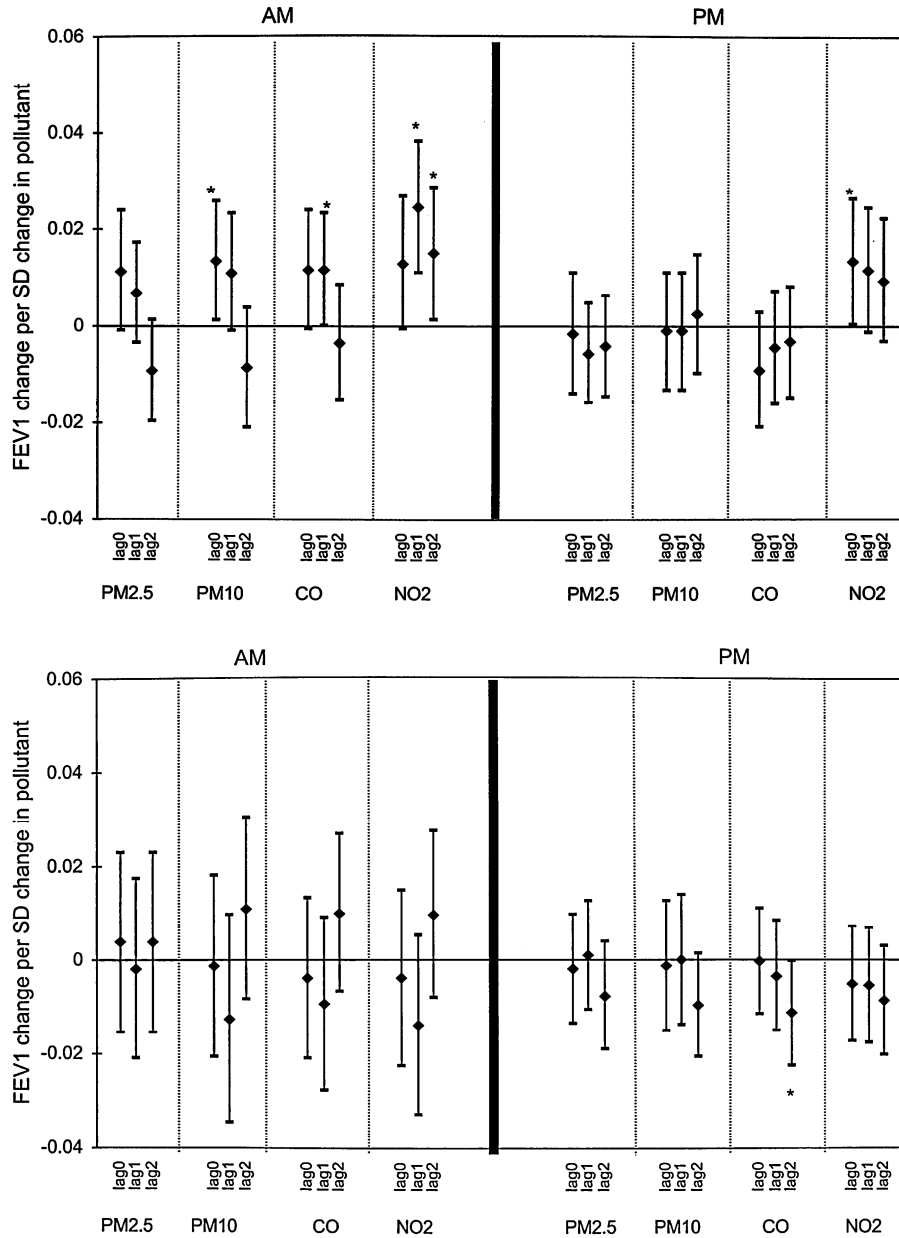
<sup>2</sup> These studies did not provide the year(s) in which air quality data was collected.

## Pulmonary Function

1 Rabinovitch et al. (2004) examined the association between exposure to ambient air pollutants and  
2 asthma exacerbation in a panel of urban minority children, 6-12 years old, with moderate to severe asthma  
3 over three winters in Denver, CO. The investigators examined pulmonary function by measuring forced  
4 expiratory volume in 1 second (FEV<sub>1</sub>) and peak expiratory flow (PEF) in the morning on school days, and  
5 also at night on weekends or other nonschool days. Using a 3-day moving average (lag 0-2) for all  
6 pollutants, Rabinovitch et al. (2004) did not find an association between CO and either lung function  
7 parameter during the morning or at night.

8 Silkoff et al. (2005) also examined lung function in Denver during the winter months, but in a  
9 panel of former smokers that were at least 40 years old and had been diagnosed with COPD. In this study  
10 CO concentrations were similar to those reported in Rabinovitch et al. (2004). The authors examined the  
11 association between exposure to air pollutants and lung function (i.e., FEV<sub>1</sub> and PEF) in both the morning  
12 and the evening. Silkoff et al. (2005) found contradictory results when examining the effects of CO for  
13 each of the winter periods separately, 1999-2000 and 2000-2001. During the analysis of the first winter  
14 (i.e., 1999-2000) CO was not found to be associated with lung function decrements at any lag, but an  
15 increase in FEV<sub>1</sub> during the morning was observed at lag 1. For the second winter (i.e., 2000-2001) the  
16 authors found a significant negative association between CO exposure and FEV<sub>1</sub> in the evening at lag 2  
17 (Figure 5-10). CO was not found to be associated with PEF at any lag during either winter period. Silkoff  
18 et al. (2005) postulated that the difference in results for the two study periods could be due to higher  
19 pollution concentrations along with somewhat lower temperatures and higher humidity in 2000-2001.  
20 However, mean CO levels remained relatively constant between the first and second winters, whereas,  
21 PM<sub>10</sub>, PM<sub>2.5</sub>, and NO<sub>2</sub> concentrations all increased. The decrements in FEV<sub>1</sub> observed in the second

- 1 winter, therefore, may have been due to the slightly worse, although not significantly different, baseline
- 2 lung function of the panel of subjects used during the second winter (Silkoff et al., 2005).



Source: Silkoff et al. (2005)

**Figure 5-10. Estimates for FEV<sub>1</sub> change expressed per SD change of the individual pollutants PM<sub>2.5</sub>, PM<sub>10</sub>, CO, and NO<sub>2</sub> for the 1999-2000 (top) and 2000-2001 (bottom) winters at lags 0, 1, and 2. The asterisk indicates a significant association (p < 0.05).**

1 In the recent literature, the majority of studies that examined the association between short-term  
2 exposure to CO and lung function have been conducted in Europe. The results from these studies  
3 contradict those reported in the U.S.-based studies previously discussed. Negative associations between  
4 short-term exposure to CO and lung function were observed primarily in individuals with underlying  
5 respiratory conditions; however, some evidence also exists for effects in children that live in urban  
6 environments. Penttinen et al. (2001) examined the association between CO and lung function in a panel  
7 consisting of 57 non-smoking adult asthmatics during the winter and spring in Helsinki, Finland. The  
8 authors observed negative associations with PEF (L/min) for a 0.5 ppm increase in 24-h avg CO  
9 concentrations in the morning at lag 1 ( $\beta = -0.54$ , SE = 0.084), and in the afternoon ( $\beta = -1.52$ , SE = 0.29)  
10 and evening ( $\beta = -1.81$ , SE = 0.27) for a 5-day average. In two-pollutant models with daily mean particle  
11 number concentration (PNC), CO effects on PEF in the morning were attenuated at lag 1, but remained  
12 negative. In addition, negative associations with PEF persisted in the afternoon and evening in a two-  
13 pollutant model at lag 0. In this study, high correlations between UFP and other traffic generated  
14 pollutants (e.g., CO and NO<sub>x</sub>) make it difficult to attribute the observed respiratory effects to a specific  
15 pollutant.

16 Lagorio et al. (2006) also conducted a study that examined the association between CO and lung  
17 function in adults. In this study, 3 panels of subjects that resided in Rome, Italy, were selected with  
18 underlying asthma, COPD, or IHD. The ages of the subjects varied depending on the panel, but overall  
19 the subjects ranged from 18-80 years old. In single-pollutant models CO was found to be negatively  
20 associated with both FVC (forced vital capacity) and FEV<sub>1</sub> at most of the lags examined (i.e., 0, 0-1, and  
21 0-2) for both the COPD and asthma panels. No association was observed between CO and FVC or FEV<sub>1</sub>  
22 in the IHD panel. Lagorio et al. (2006) did observe high correlations between CO and PM<sub>2.5</sub>, but not NO<sub>2</sub>.  
23 Unfortunately, copollutant models were not conducted in this analysis to identify whether the CO  
24 associations observed are confounded by other pollutants.

25 Timonen et al. (2002) examined the effect of CO on bronchial responsiveness and pulmonary  
26 function (i.e., FVC, FEV<sub>1</sub>, MMEF, and AEFV) in a panel of children 7-12 years old with chronic  
27 respiratory symptoms during the winter in Kuopio, Finland. The authors found that CO was significantly  
28 associated with decrements in baseline lung function for FVC (mL) at lags 2 (-17.5 mL), 3 (-24.8 mL),  
29 and 4-day avg (-52.5 mL), and for FEV<sub>1</sub> (mL) at lag 3 (-20.9 mL) for a 0.5 ppm increase in 24-h avg CO  
30 concentration. However, CO was not associated with exercise induced changes in lung function. Overall,  
31 Timonen et al. (2002) found that high concentrations of combustion byproducts (i.e., BS, PM<sub>10</sub>, particle  
32 numbers, NO<sub>2</sub>, and CO) were associated with impairment in baseline lung function. These associations,  
33 along with the high correlation between pollutants, contributed to the inability of the authors to conclude  
34 that the lung function effects observed were due to biological changes in lung pathology specific to CO  
35 exposure.

1           Chen et al. (1999) examined the effect of CO on lung function in a panel of 8-13 year old asthmatic  
2 children in Taiwan. The authors observed an association between short-term exposure to CO and  
3 decrements in FVC (mL) at a 2-day lag when using daytime average CO concentrations (from 0800 to  
4 1800) in a single-pollutant model. In addition, the authors found a high correlation between CO and NO<sub>2</sub>  
5 concentrations ( $r = 0.86-0.98$ ), but did not conduct multipollutant analyses to examine the effect of each  
6 pollutant.

7           One additional study, Fischer et al. (2002), examined the association between CO and respiratory  
8 health, specifically lung function in a non-selected cohort study of 68 children ages 10-11 that live in an  
9 urban environment (i.e., Utrecht, the Netherlands). In this study, the authors examined whether eNO was  
10 a more sensitive measure of lung damage than the traditional pulmonary function measurements  
11 (i.e., FVC, FEV<sub>1</sub>, PEF, and MMEF). Fischer et al. (2002) found negative associations between CO and  
12 FEV<sub>1</sub>, PEF, and MMEF at both lags 1 and 2. Additionally, the authors found an association between CO  
13 and an increase in eNO at lag 1. However, the study did not present the correlations between pollutants or  
14 examine copollutant models.

## **Respiratory Symptoms in Asthmatic Individuals**

15           Upon evaluating the literature that examined the association between short-term exposure to CO  
16 and respiratory symptoms in asthmatic individuals, consistent positive results were observed across  
17 studies. Studies consisting of children enrolled in the Childhood Asthma Management Program (CAMP)  
18 study found that CO was positively associated with asthma symptoms. Yu et al. (2000) found a 14%  
19 increase in asthma symptoms ([95% CI: 5-23] per 0.5 ppm increase in 24-h avg CO concentrations at lag  
20 1) in a population of 5-13 year old children ( $n = 133$ ) with asthma in Seattle, WA. These effects persisted  
21 when controlling for previous day's asthma symptoms (12% [95% CI: 5-19] at lag 1). Using the same  
22 population of children, Slaughter et al. (2003) found a significant association between short-term  
23 exposure to CO at lag 1 and asthma severity both with and without controlling for the previous day's  
24 asthma severity, (RR = 1.04 [95% CI: 1.01-1.08]) and (RR = 1.03 [95% CI: 1.00-1.05]), respectively.  
25 Schildcrout et al. (2006) examined the association between air pollutants and asthma symptoms in 990  
26 children ages 5-12 in 8 North American cities. The authors found a positive association between short-  
27 term exposure to CO and asthma symptoms at lag 0 (OR = 1.04 [95% CI: 1.00-1.07] per 0.5 ppm increase  
28 in 24-h avg CO), but similar effects were also observed at lag 1, 2, and the 3-day moving sum. The CO  
29 effects observed persisted when NO<sub>2</sub>, PM<sub>10</sub>, and SO<sub>2</sub> were included in joint pollutant models.

30           Two additional U.S. studies also found positive associations between CO and asthma symptoms,  
31 Rabinovitch et al. (2004) and Delfino et al. (2003). Rabinovitch et al. (2004) found a positive association  
32 between 24-h avg CO concentrations for a 3-day moving average (lag 0-2) and asthma exacerbations (OR  
33 = 1.02 [95% CI: 0.89-1.16] per 0.5 ppm increase in 24-h avg CO) in a population of urban poor children  
34 with moderate to severe asthma in Denver, CO. Delfino et al. (2003) also observed positive associations

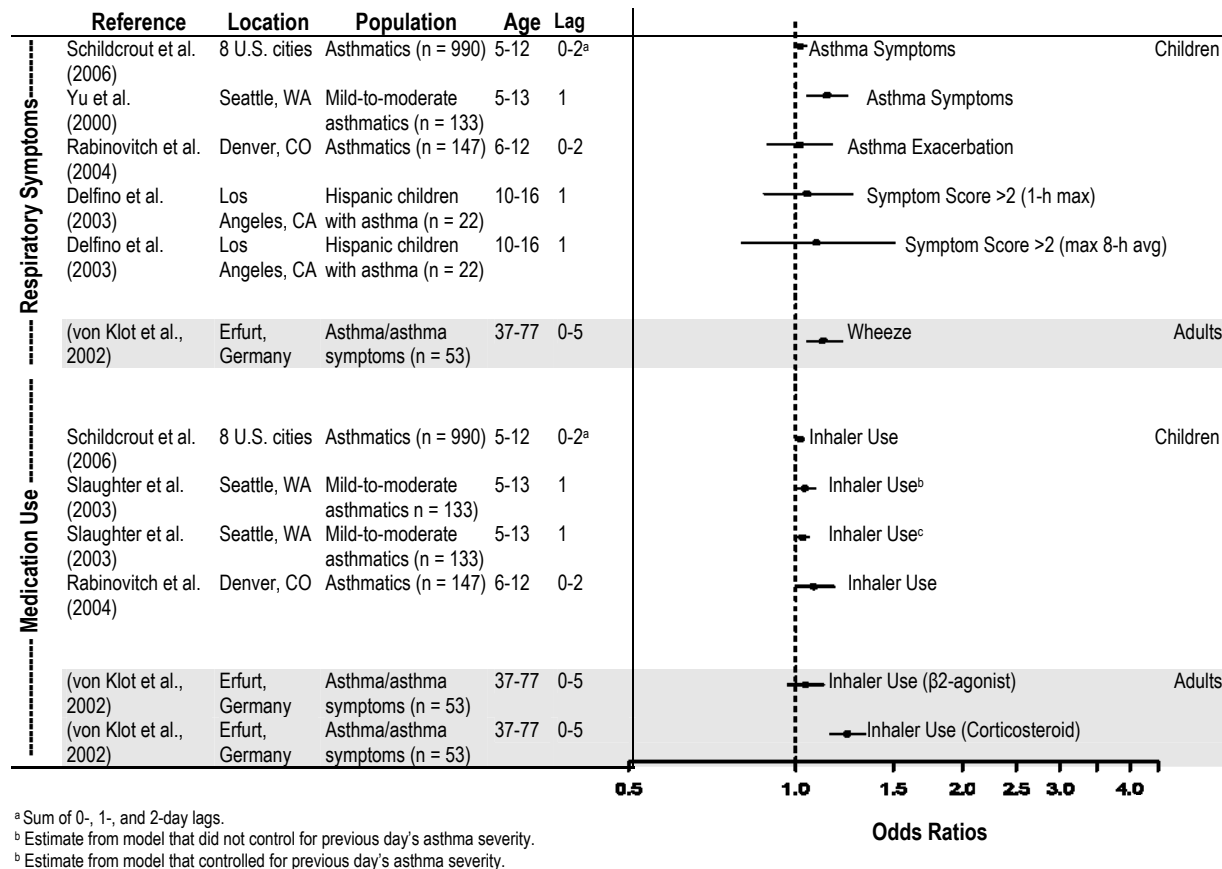


1 between CO and asthma symptoms in a population of Hispanic children with asthma in a Los Angeles,  
2 CA, community. However, positive associations were found only when using the previous day's  
3 maximum 8-h avg CO concentration as the exposure metric. These results are in contrast to those studies  
4 reported above which found positive associations when using 24-h avg CO concentrations as the exposure  
5 metric. It should be noted that in comparison to Rabinovitch et al. (2004) and the other studies discussed  
6 above, the mean ambient concentrations for 1-h maximum and maximum 8-h avg reported by Delfino et  
7 al. (2003) were 7.7 ppm and 5.0 ppm, respectively, both of which are approximately 3.5 times higher than  
8 the corresponding 24-h avg concentrations reported in the other studies.

9 In a panel study consisting of 53 adults with asthma or asthma symptoms in Germany, von Klot  
10 et al. (2002) observed a marginal association between CO concentration and the prevalence of wheezing  
11 at lag 0 (OR = 1.03 [95% CI: 0.97-1.08] per 0.5 ppm increase in 24-h avg CO), and a positive association  
12 for a 5-day mean concentration (OR = 1.12 [95% CI: 1.05-1.21] per 0.5 ppm increase in 24-h avg CO).  
13 However, the authors found CO to be highly correlated with UFP, making it unclear if the effect observed  
14 was due to CO alone. Additionally, Park et al. (2005a) in a panel study of individuals 16-75 years old in  
15 Incheon, Korea with bronchial asthma did not find an association between CO and nighttime asthma  
16 symptoms or cough. Figure 5-11 summarizes the results from studies that provided usable quantitative  
17 results and examined the association between short-term exposure to CO and asthma or respiratory  
18 symptoms in asthmatic individuals.

19 To further examine the effect of CO on asthma and asthma symptoms some studies also analyzed  
20 medication use in asthmatic individuals in response to an increase in air pollution levels. The majority of  
21 U.S.-based studies (i.e., Rabinovitch (2004), Slaughter et al. (2003), and Schildcrout et al. (2006))  
22 focused on rescue inhaler use in children with ages ranging from 5-13 years old. Rabinovitch et al. (2004)  
23 found a weak association (OR = 1.08 [95% CI: 1.00-1.17] per 0.5 ppm increase in 24-h avg CO) between  
24 rescue inhaler use in a population of 6-12 year old urban minority children with moderate to severe  
25 asthma in the winter in Denver, CO, which is evident by the large confidence intervals surrounding the  
26 estimate. Slaughter et al. (2003) in a population of 5-12 year old children with asthma in Seattle, WA  
27 found a positive and significant association with rescue inhaler use both with and without taking into  
28 consideration the previous day's asthma severity, (RR: 1.04 [95% CI: 1.01-1.08] per 0.5 ppm increase in  
29 24-h avg CO) and (RR: 1.03 [95% CI: 1.00-1.05] per 0.5 ppm increase in 24-h avg CO), respectively.  
30 Similar results were observed in a multi-city study conducted by Schildcrout et al. (2006), which analyzed  
31 rescue inhaler use in 990 children ages 5-13 with asthma in eight North American cities. Schildcrout et al.  
32 (2006) found that short-term exposure to CO was positively associated with rescue inhaler use at lags of  
33 0, 2, and a 3-day moving sum, and that the association was fairly robust to an increase in other pollutants  
34 (i.e., NO<sub>2</sub>, PM<sub>10</sub>, and SO<sub>2</sub>) when included in joint models with CO. However, both Slaughter et al. (2003)  
35 and Schildcrout et al. (2006) attributed the associations observed to other combustion byproducts.  
36 Additional studies (Park et al., 2005a; Silkoff et al., 2005; von Klot et al., 2002) conducted in Denver,

1 CO; Erfurt, Germany; and Incheon, Korea, respectively, found results that are consistent with those  
 2 previously reported, but in populations with combined ages ranging from 16-77. Figure 5-11 presents the  
 3 risk estimates from studies that examined the association between short-term exposure to CO and  
 4 medication use in asthmatic individuals.



**Figure 5-11. Asthma symptoms, respiratory symptoms and medication use in asthmatic individuals associated with short-term exposure to CO<sup>1</sup>. Effect estimates were standardized depending on the averaging time used in the study: 0.5 ppm for 24-h avg, 0.75 ppm for max 8-h avg, and 1.0 ppm for 1-h max.**

### Respiratory Symptoms in Non-Asthmatic Individuals

5 In addition to examining the association between short-term exposure to CO and respiratory  
 6 symptoms (e.g., cough, wheeze, shortness of breath, etc.) in asthmatic populations some studies examined  
 7 these respiratory effects in individuals classified as non-asthmatics. Rodriguez et al. (2007) and de Hartog

<sup>1</sup> Effect estimates from Park et al. (2005a) were not included in this figure because the study did not provide the increment at which the effect estimates were calculated.

1 et al. (2003) examined the effect of CO on respiratory symptoms in a panel of 263 children 0-5 years old  
2 at high risk for developing asthma in Perth, Australia, and a 3-city panel of individuals  $\geq 50$  years of age  
3 with CHD during the winter in Amsterdam, the Netherlands, Erfurt, Germany, and Helsinki, Finland,  
4 respectively. Rodriguez et al., (2007) found CO concentrations to be positively associated with  
5 wheeze/rattle chest and runny/blocked nose at both a 5-day lag and a 0-5-day lag in Perth, Australia. It is  
6 unclear which pollutant is driving the effect observed by Rodriguez et al. (2007) because multipollutant  
7 models were not examined and additional analyses were not conducted to further characterize the  
8 associations observed.

9 In a panel of elderly individuals with CHF in three European locations, de Hartog et al. (2003)  
10 observed some marginal associations, specifically between CO concentration and the prevalence of the  
11 respiratory symptom shortness of breath, at lag 3 and 0-5-days. Although a marginal association was  
12 observed, the authors found that the associations between air pollution exposure and respiratory  
13 symptoms were stronger for PM<sub>2.5</sub> than for gaseous air pollutants.

### **Summary of Associations between Short-Term Exposure to CO and Pulmonary Function, Respiratory Symptoms, and Medication Use**

14 A limited body of evidence is available that examined the effect of short-term exposure to CO on  
15 various respiratory health outcomes. Among asthmatics, the studies reviewed generally find positive  
16 associations between short-term exposure to CO and respiratory-related health effects (i.e., decrements in  
17 lung function/lung function growth, respiratory symptoms, and medication use). It is difficult to  
18 determine from this group of studies if CO is independently associated with respiratory outcomes or if  
19 CO is an indicator for other traffic-related pollutants. On-road vehicle exhaust emissions are a nearly  
20 ubiquitous source of combustion pollutant mixtures that include CO and can be an important contributor  
21 to CO-related health effects in near-road locations. A lack of copollutant analyses among this group of  
22 studies complicates the efforts to disentangle the health effects attributed to CO from the larger traffic-  
23 related pollutant mix. Additional uncertainty exists as to a biologically plausible mechanism which could  
24 explain the effect of CO on respiratory health.

#### **5.5.1.2. Respiratory Hospital Admissions, ED Visits and Physician Visits**

25 The 2000 CO AQCD (U.S. EPA, 2000) evaluated a limited amount of literature that examined the  
26 association between short-term exposure to CO and respiratory hospital admissions, ED visits, and  
27 physician visits in the U.S. (i.e., Seattle, WA, Reno, NV, and Anchorage, AK) and Europe (i.e., Barcelona,  
28 Spain). From these studies, the 2000 CO AQCD concluded that positive associations were observed for  
29 short-term exposure to CO with several respiratory outcomes, including asthma and COPD. However, the

1 lack of a biologically plausible mechanism for CO-induced respiratory morbidity at that time brought into  
 2 question the validity of the results observed.

3 Since the 2000 CO AQCD, the number of studies that examined the association between short-term  
 4 exposure to CO and respiratory morbidity has increased; however, the total number of studies published is  
 5 still considerably less than the number that examine the health effects associated with exposure to other  
 6 criteria air pollutants. This section focuses primarily on those studies conducted in the U.S. and Canada  
 7 which examined the potential respiratory health effects associated with CO at concentrations at or similar  
 8 to those found in the U.S. Unlike previous sections, which also evaluated studies conducted outside of the  
 9 U.S., the expansive U.S.-based respiratory hospital admission and ED visits literature provides adequate  
 10 evidence to examine the association between short-term exposure to CO and respiratory HA and ED  
 11 visits. Collectively, the studies conducted outside of the U.S. observed associations that are consistent  
 12 with those observed in the U.S.-based studies evaluated below (see Annex C).

13 Overall, this section focuses on respiratory-related hospital admissions because the majority of the  
 14 literature examines hospital admissions as opposed to ED visits or physician visits (Table 5-15 presents  
 15 the studies evaluated in this section along with the range of CO concentrations measured in each study). It  
 16 must be noted that when examining the association between short-term exposure to CO and health  
 17 outcomes that require medical attention, it is important to distinguish between hospital admissions, ED  
 18 visits, and physician visits for respiratory outcomes (more so than for cardiovascular outcomes). This is  
 19 because it is likely that a small percentage of respiratory ED visits will be admitted to the hospital and,  
 20 therefore, may represent potentially less serious, but more common outcomes. To adequately distinguish  
 21 between the results presented in hospital admission, ED visit, and physician visit studies, each outcome is  
 22 evaluated in individual sections. In addition, each section presents results separately for respiratory health  
 23 outcomes which includes all respiratory diagnoses (ICD-9: 460-519) or selected diseases (e.g., asthma,  
 24 COPD, pneumonia and other respiratory infections) in order to evaluate the potential effect of short-term  
 25 exposure to CO on each outcome.

**Table 5-15. Range of CO concentrations reported in key respiratory hospital admission and ED visit studies that examine effects associated with short-term exposure to CO.**

Author	Location	Type of Visit (ICD9)	Metric	Mean Concentration (ppm)	Middle/Upper Percentile Concentrations (ppm)
Cakmak et al. (2006a)	10 Canadian cities	Hospital Admissions: Respiratory disease (i.e., Acute bronchitis and bronchiolitis; Pneumonia; Chronic and unspecific bronchitis; Emphysema; Asthma; Bronchiectasis; Chronic airway obstruction)	24-h avg	0.8	Maximum: 6.5
Linn et al. (2000)	Los Angeles, CA	Hospital Admissions: Pulmonary; Asthma; COPD	24-h avg	Winter: 1.7; Spring: 1.0; Summer: 1.2; Fall: 2.1	Maximum: Winter: 5.3; Spring: 2.2; Summer: 2.7; Fall: 4.3;

Author	Location	Type of Visit (ICD9)	Metric	Mean Concentration (ppm)	Middle/Upper Percentile Concentrations (ppm)
Slaughter et al. (2005)	Spokane, WA	ED Visits and Hospital Admissions: Respiratory; Asthma; COPD; Pneumonia; Acute Respiratory Infection	24-h avg	Hamilton St.: 1.73 Backdoor Tavern: 1.29 Spokane Club: 1.41 Third and Washington: 1.82 Rockwood: 0.42	95th: 3.05
Burnett et al. (2001)	Toronto, ON, Canada	Hospital Admissions: Respiratory disease (i.e., Asthma; Acute bronchitis/bronchiolitis; Croup; Pneumonia)	1-h max	1.9	50th: 1.8; 75th: 2.3; 95th: 3.3; 99th: 4.0 Maximum: 6.0
Yang et al. (2003)	Vancouver, BC, Canada	Hospital Admissions: Respiratory diseases	24-h avg	0.98	50th: 0.82; 75th: 1.16 Maximum: 4.90
Lin et al. (2003)	Toronto, ON, Canada	Hospital Admissions: Asthma	24-h avg	1.18	50th: 1.10; 75th: 1.40 Maximum: 6.10
Lin et al. (2004c)	Vancouver, BC, Canada	Hospital Admissions: Asthma	24-h avg	0.96	50th: 0.80; 75th: 1.12 Maximum: 4.90
Moolgavkar (2003a) (re-analysis of Moolgavkar 2000a)	Cook County, IL; Los Angeles County, CA	Hospital Admissions: COPD	24-h avg	NR	<b>Cook:</b> 50th: .99; 75th: 1.25 Maximum: 3.91 <b>Los Angeles:</b> 50th: 1.35; 75th: 2.16 Maximum: 5.96
Yang et al. (2005)	Vancouver, BC, Canada	Hospital Admissions: COPD	24-h avg	0.71	50th: 0.64 Maximum: 2.48
Karr et al. (2006)	South Coast Air Basin, CA	Hospital Admissions: Acute bronchiolitis	24-h avg	<b>Lag 1:</b> Index: 1.730 Referrent: 1.750 <b>Lag 4:</b> Index: 1.760 Referrent: 1.790	<b>Lag 1:</b> Index: 50th: 1.52; 75th: 2.26; 90th: 3.16 Maximum: 9.60 <b>Referrent:</b> 50th: 1.51; 75th: 2.29; 90th: 3.23 Maximum: 9.60 <b>Lag 4:</b> Index: 50th: 1.54; 75th: 2.31; 90th: 3.23 Maximum: 8.71 <b>Referrent:</b> 50th: 1.55; 75th: 2.35; 90th: 3.30 Maximum: 9.60
Karr et al. (2007)	South Coast Air Basin, CA	Hospital Admissions: Acute bronchiolitis	24-h avg; Monthly avg	24-h avg: 1.720 Monthly: 1.770	<b>24-h avg:</b> 50th: 1.61; 75th: 2.08; 90th: 2.75 Maximum: 5.07 <b>Monthly avg:</b> 50th: 1.63; 75th: 2.13; 90th: 2.88 Maximum: 8.30
Zanobetti and Schwartz (2006)	Boston, MA	Hospital Admissions: Pneumonia	24-h avg	NR	50th: 0.48; 75th: 0.60; 95th: 0.88
Lin et al. (2005)	Toronto, ON, Canada	Hospital Admissions: Respiratory infections	24-h avg	1.16	50th: 1.05; 75th: 1.37 Maximum: 2.45

Author	Location	Type of Visit (ICD9)	Metric	Mean Concentration (ppm)	Middle/Upper Percentile Concentrations (ppm)
Peel et al. (2005)	Atlanta, GA	ED Visits: All respiratory; Asthma; COPD; URI; Pneumonia	1-h max	1.8	90th: 3.4
Tolbert et al. (2007)	Atlanta, GA	ED Visits: Respiratory diseases (i.e., Asthma; COPD; URI; Pneumonia; Bronchiolitis)	1-h max	1.6	50th: 1.3; 75th: 2.0; 90th: 3.0 Maximum: 7.7
Ito et al. (2007)	New York, NY	ED Visits: Asthma	8-h max	1.31	50th: 1.23; 75th: 1.52; 95th: 2.11
Villeneuve et al. (2006b)	Toronto, ON, Canada	Physicians Visits: Allergic rhinitis	24-h avg	1.1	Maximum: 2.2

## Hospital Admissions

### *Respiratory Disease*

1 All but two of the published hospital admission studies that examined the association between  
2 short-term exposure to CO and all respiratory diseases from North America were conducted in Canada,  
3 and only one study presented results from a combined analysis of multiple cities (Cakmak et al., 2006b).  
4 In a study of 10 of the largest Canadian cities, Cakmak et al. (2006b) examined respiratory hospital  
5 admissions (ICD-9: 466, 480-486, 490-494, 496) in relation to ambient gaseous pollutant concentrations  
6 for the time period 1993-2000. This study reported a 0.37% (95% CI: 0.12-0.50) increase in respiratory  
7 hospital admissions for all ages for a 0.5 ppm increase in 24-h avg CO (lag 2.8 days averaged over the 10  
8 cities<sup>1</sup>). U.S.-based studies (Los Angeles and Spokane) reported similarly weak or null associations for  
9 respiratory disease hospital admissions (Linn et al., 2000; Slaughter et al., 2005). In a study conducted in  
10 Toronto, Canada for the time period 1980-1994, Burnett et al. (2001) reported a relatively strong  
11 association between 1-h max CO and respiratory disease hospital admissions in children less than two  
12 years of age, for the diagnoses of asthma (493), acute bronchitis/bronchiolitis (466), croup (464.4), and  
13 pneumonia (480-486). The authors found a 9.7% (95% CI: 4.1-15.5) increase in hospital admissions for a  
14 2-day avg (lag 0-1) per 1 ppm increase in 1-h max CO. In the two-pollutant model analysis, the estimates  
15 for both CO and O<sub>3</sub> remained elevated. Yang et al. (2003) reported similar results (OR = 1.04  
16 [95% CI: 1.01-1.06] at lag 1 per 0.5 ppm increase in 24-h avg CO) for pediatric (<3 years of age)  
17 respiratory disease (ICD-9: 460-519) admissions in Vancouver for the time period 1986-1998. Yang et al.  
18 (2003) also reported elevated associations with 24-h avg CO and respiratory hospital admissions (ICD-9:  
19 codes 460-519) for ages 65 and over in Vancouver, Canada (OR = 1.02 [95% CI: 1.00-1.04]) at lag 1 for a  
20 0.5 ppm increase in 24-h avg CO. The authors found that the risk estimate remained elevated when O<sub>3</sub>  
21 was included in the model.

<sup>1</sup> To determine the lag for the combined estimate across all 10 cities, Cakmak et al. averaged the strongest associations from lags 0-5 days from each city.

## *Asthma*

1           Some studies that examined the effect of short-term exposure to CO on asthma hospital admissions  
2 conducted all age and age-stratified analyses, specifically to examine effects in children. In two of these  
3 hospital admission studies conducted in Canada, evidence was observed for increased pediatric (ages  
4 6-12) asthma hospital admissions (ICD-9: 493) in boys, but not girls (Lin et al., 2003; 2004c); however, a  
5 biological explanation was not provided which could explain this difference. Lin et al. (2003) used a bi-  
6 directional case-crossover analysis in Toronto, Canada for the years 1981–1993 reported an OR of 1.05  
7 (95% CI: 1.00-1.11) per 0.5 ppm increase in 24-h avg CO for a 1-day lag for boys with similar results  
8 being reported when averaging CO concentrations up to 7 days prior to hospitalization. Risk estimates for  
9 girls did not provide evidence of an association using the same lag structure that was used in the boys’  
10 analysis (OR = 1.00 [95% CI: 0.93-1.06]); lag 1). In a copollutant analysis, the estimates for boys were  
11 essentially unchanged when adjusting for all PM indices (Lin et al., 2003). It should be noted that this  
12 study used a bi-directional case-crossover analysis, which may be biased (Levy et al., 2001). Studies that  
13 examined the various referent selection strategies for the case-crossover study design have concluded that  
14 the preferred control selection strategy is the time-stratified framework (Levy et al., 2001). In an  
15 additional analysis conducted by Lin et al. (2004c), the authors found less consistent evidence for a  
16 greater effect in boys versus girls in Vancouver during the years 1987-1998 using a time-series study  
17 design that stratified results by socioeconomic status (SES). In one additional study that examined asthma  
18 hospital admissions for all ages and genders combined, Slaughter et al. (2005) observed some evidence of  
19 an increase in asthma hospital admissions (ICD-9 493) in Spokane (1995-2000) for CO at lag 2 (RR =  
20 1.03 [ 95% CI: 0.98-1.08]) for a 0.5 ppm increase in 24--h avg, but not for the other two lags examined  
21 (lag 1 and lag 3).

## *Chronic Obstructive Pulmonary Disease*

22           A few of the studies examined the effect of short-term exposure to CO on COPD, or obstructive  
23 lung disease, and hospital admissions. Moolgavkar (2003a) (a reanalysis of (Moolgavkar, 2000a)  
24 examined hospital admissions for COPD plus “allied diseases” (ICD-9 490-496) in two U.S. counties  
25 (Cook County, IL and Los Angeles County, CA) for the years 1987-1995 using Poisson generalized linear  
26 models (GLMs) or generalized additive models (GAM) with the more stringent convergence criteria.  
27 Overall, the results from both models were similar. Using the GAM models the study reported percent  
28 increases of 0.53-1.20% for all ages in Los Angeles County, and 0.17-1.41% for ages 65 and older in  
29 Cook County, for a 0.5 ppm increase in 24-h avg CO and lags ranging from 0 to 5 days. Yang et al. (2005)  
30 reported similar results for COPD hospital admissions (ICD-9 490-492, 494, 496) in Vancouver for ages  
31 65 and older for the years 1994-1998 for a moving average of 0-6 day lags (RR = 1.14 [95% CI: 1.03-  
32 1.23] per 0.5 ppm increase in 24-h avg CO). However, Slaughter et al. (2005) found no association  
33 between short-term exposure to CO and COPD hospital admissions (ICD-9 491, 492, 494, 496) in

1 Spokane, WA at lag 1-day (RR = 0.97 [95% CI: 0.93-1.01] per 0.5 ppm increase in 24-h avg CO) with  
2 similar results being reported for 2- and 3-day lags.

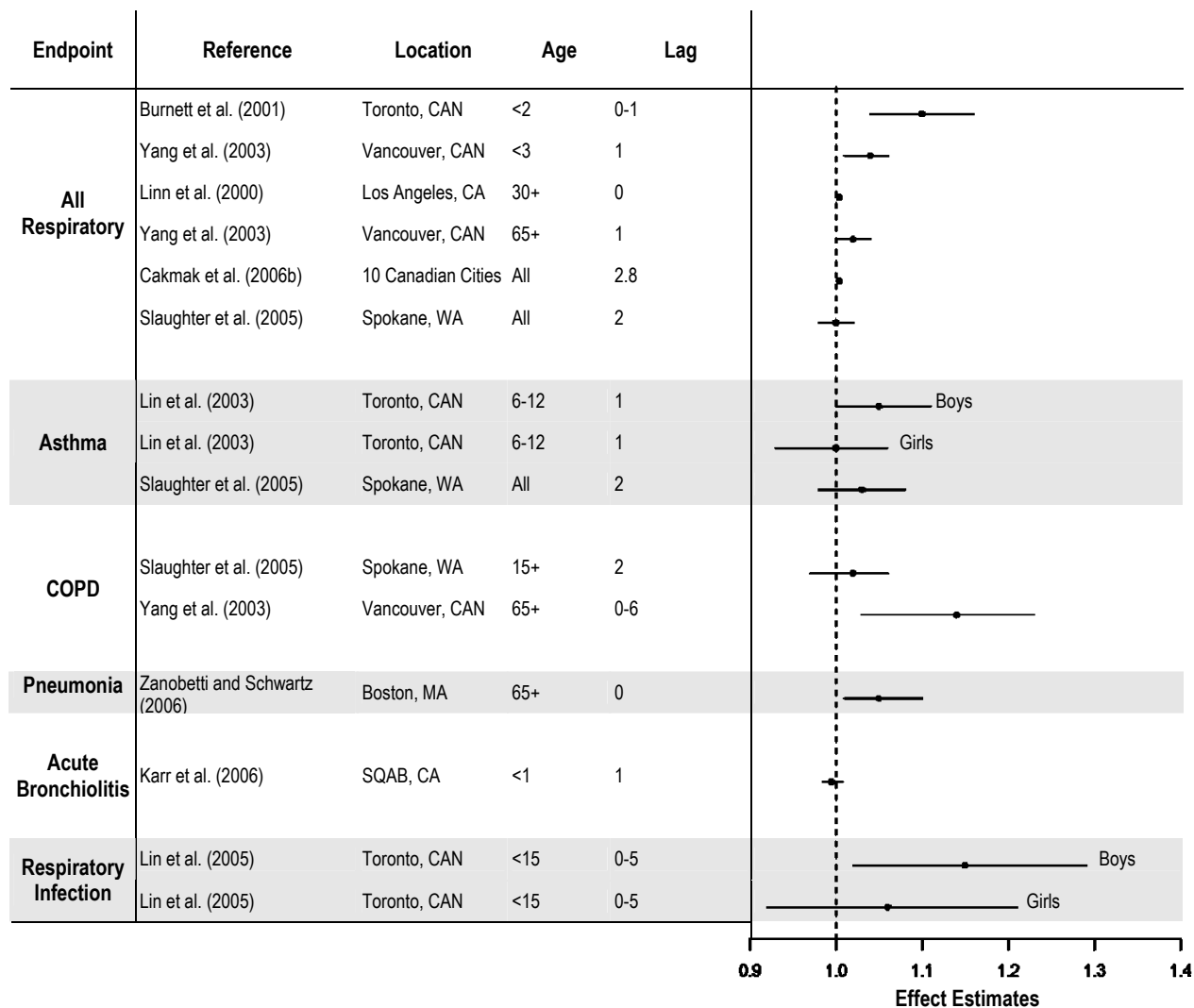
### *Acute Bronchiolitis in Infants*

3 Two studies (Karr et al., 2006; 2007) from the South Coast Air Basin in California examined both  
4 short term (lag 0 or 1) and longer term levels of CO in relation to acute bronchiolitis (ICD-9: 466)  
5 hospital admissions during the first year of life from 1995-2000. Karr et al. (2006) found no evidence of a  
6 short-term association between ambient CO concentrations and hospital admissions for acute bronchiolitis  
7 at lag 1 day (OR= 0.99 [95%CI: 0.98-1.01] per 0.5 ppm increase in 24-h avg CO). In addition, Karr et al.  
8 (2007), which examined longer term exposures (average in the month prior to hospitalization and lifetime  
9 average) in a matched case-control study, also did not provide any evidence of an association with CO.

### *Pneumonia and Other Respiratory Infections*

10 In addition to examining the effect of short-term exposure to CO on health outcomes that can limit  
11 the function of the respiratory system, some studies examined the effect of CO on individuals with  
12 pneumonia (ICD-9: 480-486) separately or in combination with other respiratory infections. Zanobetti  
13 and Schwartz (2006) examined pneumonia hospital admissions (ICD-9 480-487) in Boston, MA, for the  
14 years 1995-1999 for ages 65 and older using a time-stratified case-crossover analysis. The authors  
15 reported an increase in pneumonia hospital admissions at lag 0 of 5.4% (95% CI: 1.2-10.0) per 0.5 ppm  
16 increase in 24-h avg CO. While Zanobetti and Schwartz (2006) did not report multipollutant results, they  
17 suggested that CO was most likely acting as a marker for traffic-related pollutants because CO was highly  
18 correlated with both BC ( $r = 0.80$ ) and  $\text{NO}_2$  ( $r = 0.67$ ), and moderately correlated with  $\text{PM}_{2.5}$  ( $r = 0.52$ ).  
19 Instead of examining the effect of CO on pneumonia hospital admissions separately as was done by  
20 Zanobetti and Schwartz (2006), Lin et al. (2005) presented results for the overall effect of CO on  
21 respiratory infection hospital admissions (ICD-9: 464, 466, 480-487). In this analysis, Lin et al. (2005)  
22 examined the potential increase in respiratory hospital admissions in children less than 15 years of age in  
23 Toronto, Canada for 1998-2001 using a bi-directional case-crossover approach. The authors reported  
24 elevated estimates for boys (OR = 1.15 [95% CI: 1.02-1.29] per 0.5 ppm increase in 24-h avg CO for a 6-  
25 day ma) while the estimate for girls was weaker and with wider confidence intervals (OR = 1.06 [95%CI:  
26 0.92-1.21]). Lin et al. (2005) did not provide an explanation as to why the estimates are stronger for boys  
27 than for girls. It should be noted that this study used a bi-directional case-crossover analysis, which may  
28 be biased (Levy et al., 2001). Studies that have examined the various referent selection strategies for the  
29 case-crossover study design have concluded that the preferred control selection strategy is the time-  
30 stratified framework (Levy et al., 2001).





**Figure 5-12. Summary of associations between short-term exposure to CO and respiratory hospital admissions.<sup>1,2</sup> Effect estimates were standardized depending on the averaging time used in the study: 0.5 ppm for 24-h avg, 0.75 ppm for max 8-h avg, and 1.0 ppm for 1-h max.**

## Emergency Department Visits

### *Respiratory Disease*

- 1 Peel et al. (2005) conducted a large single-city ED study in Atlanta, GA, which included data from
- 2 31 hospitals for the time period 1993–2000. In this study, results were reported for various respiratory-

<sup>1</sup> Risk estimates from Moolgavkar (2003a) were not included in this figure because the study presented a range of effect estimates using different statistical models. The results from this study were more adequately highlighted in the evaluation of the study in the COPD section.

<sup>2</sup> Risk estimates from Lin et al. (2004c) were not included in the figure because the results were stratified by SES and therefore could not be readily compared to effect estimates from Lin et al. (2003).

1 related visits (ICD-9 460-466, 477, 480-486, 491-493, 496, 786.09). In an all ages analysis, the authors  
2 found a RR of 1.011 (95% CI: 1.004-1.019) for a 3-day avg (lag 0-2) per 1 ppm increase in 1-h max CO  
3 concentration for all respiratory diseases. Tolbert et al. (2007) expanded the time period used in the Peel  
4 et al. (2005) study to include ED visits through 2004, and reported similar results for respiratory ED visits  
5 (RR: 1.013 [95% CI: 1.007-1.018] per 1 ppm increase in 1-h max CO). The CO risk estimates from the  
6 Atlanta, GA, ED visits studies were attenuated when O<sub>3</sub>, NO<sub>2</sub>, or PM were added to the model (). In  
7 addition, Tolbert et al. (2007) reported high correlations between CO and NO<sub>2</sub> (r = 0.70) and EC  
8 (r = 0.66); and a moderate correlation with PM<sub>2.5</sub> (r = 0.51). One additional ED visits study that also  
9 examined respiratory disease (Slaughter et al., 2005) presented essentially null results at lag 1 and 2, but  
10 found similar results to Peel et al. (2005) and Tolbert et al. (2007) at lag 3 (RR = 1.02 [95% CI: 1.00-1.03]  
11 per 0.5 ppm increase in 24-h avg CO).

### *Asthma*

12 The association between short-term exposure to CO and asthma ED visits (ICD-9 493, 786.09) was  
13 also examined in Atlanta, GA by Peel et al. (2005). In this study the authors reported results from  
14 distributed lag models including lags 0-13 in addition to a moving average of lags 0, 1, and 2 (lag 0-2) for  
15 specific respiratory outcomes (e.g., asthma). Effect estimates from the distributed lag models were  
16 stronger than those produced from models that used 3-day moving average CO concentrations (RR =  
17 1.010 [95% CI: 0.999-1.022] for lags 0-2 compared to RR = 1.076 [95% CI: 1.047-1.105] for an  
18 unconstrained distributed lag of 0-13 for a 1 ppm increase in 1-h max CO). These results demonstrated  
19 the potential effect of CO exposures up to 13 days prior to an asthma ED visit. Estimates were stronger  
20 for pediatric ED visits (ages 2-18 years) (RR = 1.019 [95% CI: 1.004-1.035] per 1 ppm increase in 1-h  
21 max CO) for a 3-day avg (lag 0-2) compared to all ages (Peel et al., 2005). Slaughter et al. (2005), which  
22 also examined ED visits for Spokane (1995-2001), reported an increase in asthma ED visits for all ages  
23 for CO at lag 3 (RR = 1.03 [95% CI: 1.00-1.05] per 0.5 ppm increase in 24-h avg CO), but not for the  
24 other two lags examined (lags 1 and 2). The results from Ito et al. (2007) also provide evidence of  
25 increased ED visits for asthma (ICD-9 493) for all ages in New York City for 1999-2002, but quantitative  
26 results were not provided. In addition, Ito et al. (2007) stated that increased CO effect estimates were  
27 attenuated when NO<sub>2</sub> was included in the model, but effect estimates remained elevated in two-pollutant  
28 models with either PM<sub>2.5</sub> or O<sub>3</sub>.

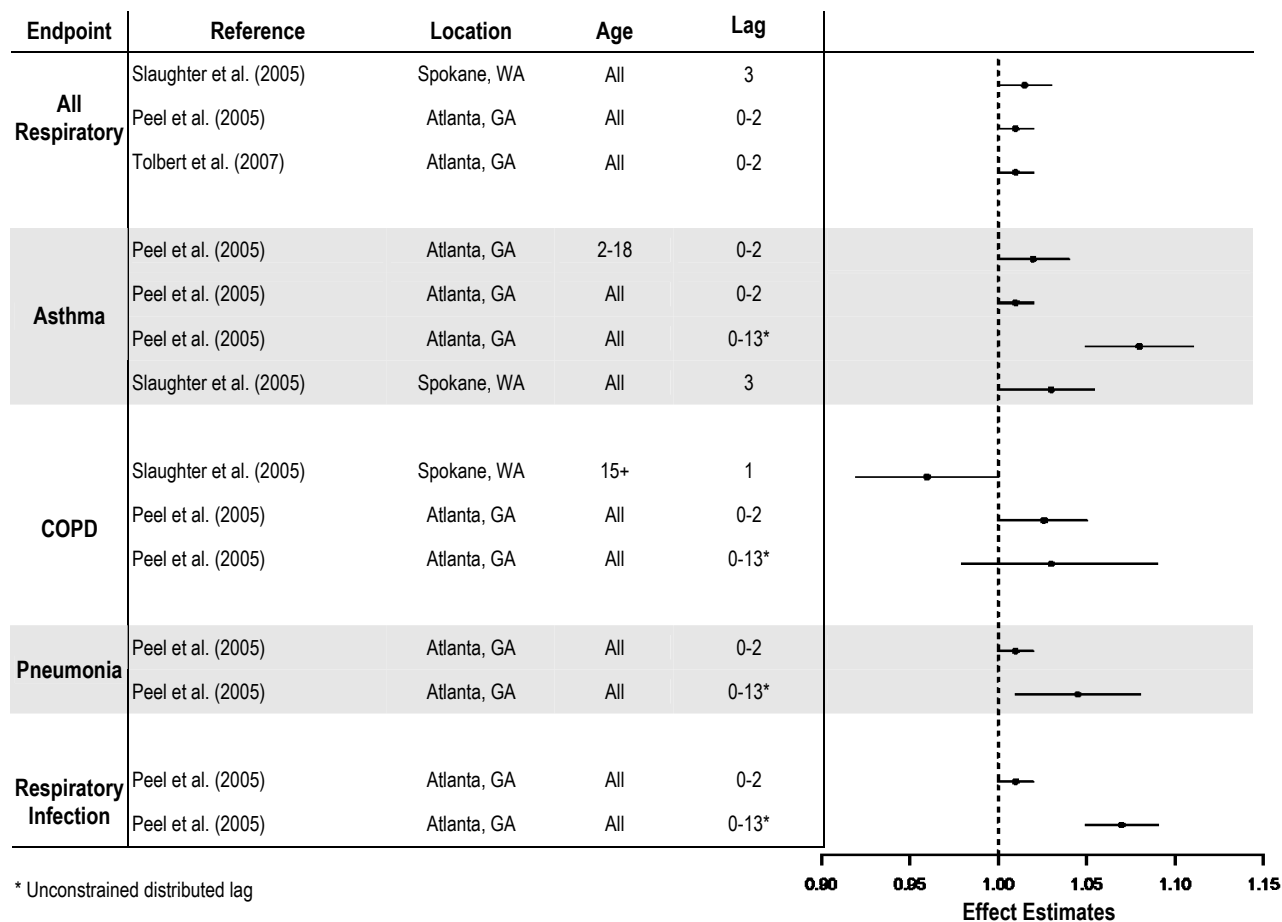
### *Chronic Obstructive Pulmonary Disease*

29 In the examination of the effect of short-term exposure to CO on COPD ED visits (ICD-9 491, 492,  
30 496), Peel et al. (2005) reported elevated estimates for Atlanta, GA for 1993-2000 (RR = 1.03 [95%CI:  
31 1.00-1.05] per 1 ppm increase in 1-h max CO for a moving average of lag 0-2) with similar results for the  
32 distributed lag model (RR = 1.03 [95% CI: 0.98-1.09). However, results from Slaughter et al. (2005) from

1 Spokane were consistent with a null or slightly protective association at lag 1 (RR = 0.96  
2 [95% CI: 0.92-1.00] per 0.5 ppm increase in 24-h avg CO at lag 1) with similar results for lags 2 and 3.

### *Pneumonia and Other Respiratory Infections*

3 Similar to the hospital admission analysis conducted by Zanobetti and Schwartz (2006) discussed  
4 above, Peel et al. (2005) examined the effect of CO on pneumonia separately (ICD-9: 480-486), but also  
5 included an analysis of upper respiratory infection (ICD-9: 460-466, 477) ED visits for all ages in Atlanta,  
6 GA during the years 1993-2000. The authors reported a weak estimate for pneumonia for the three-day  
7 moving average (lag 0-2) (RR = 1.01 [95% CI: 0.996-1.021] per 1 ppm increase in 1-h max CO).  
8 However, when using an unconstrained distributed lag (days 0-13), Peel et al. (2005) observed evidence  
9 of an association (RR = 1.045 [95% CI: 1.01-1.08]). An examination of URI ED visits, the largest of the  
10 respiratory ED groups, found slightly increased risk estimates for both the three-day moving average  
11 (lag 0-2) (RR = 1.01 [95% CI: 1.00-1.02]) and the unconstrained distributed lag for days 0-13 (RR = 1.07  
12 [95% CI: 1.05-1.09]) per 1 ppm increase in 1-h max CO. In copollutant models, CO risk estimates were  
13 largely attenuated when PM<sub>10</sub>, O<sub>3</sub>, or NO<sub>2</sub> were included in the model. Upon conducting an age-stratified  
14 analysis, Peel et al. (2005) also found that infant (less than one year of age) and pediatric (ages 2-18) URI  
15 ED visit CO risk estimates were substantially stronger than all age risk estimates.



**Figure 5-13. Summary of associations between short-term exposure to CO and respiratory ED visits. Effect estimates were standardized depending on the averaging time used in the study: 0.5 ppm for 24-h avg, 0.75 ppm for max 8-h avg, and 1.0 ppm for 1-h max.**

## Physician Visits

1 Although hospital admissions and ED visits are the two most well studied measures of morbidity, a  
 2 few studies also examined the effect of CO on unscheduled physician visits. In a time-series study,  
 3 Villeneuve et al. (2006b) examined the effect of CO on physician visits for allergic rhinitis in individuals  
 4 65 and older in Toronto, Canada. Although quantitative results were only presented in figures, upon  
 5 observation it was evident that estimates were consistent with a null association for lags 0-6 (Villeneuve  
 6 et al., 2006b). In an additional study, Sinclair et al. (2004) reported results for urgent care visits for asthma  
 7 and respiratory infections in a health maintenance organization in Atlanta, GA; however, the study only  
 8 reported statistically significant results, of which none were for CO.

## Summary of Associations between Short-Term Exposure to CO and Respiratory Hospital Admissions, ED Visits, and Physicians Visits

1           Relatively few studies evaluated the association between short-term exposure to ambient CO and  
2 hospital admissions and ED visits for various respiratory outcomes compared to other criteria air  
3 pollutants (e.g., O<sub>3</sub> and PM). Although evidence for consistent positive associations (See Figures 5-12 and  
4 5-13) has been found across these studies, various issues surrounding the association between short-term  
5 exposure to CO and respiratory-related health effects have not been addressed due to: the lack of studies  
6 that examined potential confounders of the CO-respiratory hospital admission and ED visits relationship;  
7 and uncertainty as to a biologically plausible mechanism which could explain the association between CO  
8 exposure and respiratory-related health effects. Some of the studies evaluated suggest that CO acts as an  
9 indicator of combustion (e.g., traffic), which is supported by the moderate to high correlation between CO  
10 and other traffic-related pollutants such as NO<sub>2</sub>, PM<sub>2.5</sub>, EC, or BC and in addition complicates the results  
11 presented. Only two studies examined potential confounding of CO risk estimates by other pollutants  
12 through copollutant models, and found that CO risk estimates were robust or attenuated but remained  
13 positive in two-pollutant models with O<sub>3</sub>, NO<sub>2</sub>, or PM indices.

### 5.5.2. Epidemiologic Studies with Long-Term Exposure

14           The 2000 CO AQCD did not evaluate any studies that examined the effect of long-term exposure to  
15 CO on respiratory health. The following section discusses those studies that analyze the effect of long-  
16 term exposure to CO on pulmonary function, asthma/asthma symptoms, and allergic rhinitis.

#### 5.5.2.1. Pulmonary Function

17           Mortimer et al. (2008) examined the effect of prenatal and lifetime exposures to air pollutants on  
18 pulmonary function in 232 asthmatic children that resided in the San Joaquin Valley of California. The  
19 strong temporal correlation between pollutants and pollutant metrics for different time periods in the  
20 study area contributed to the inability to draw conclusions about the effect of individual pollutant metrics  
21 on pulmonary function (Mortimer et al., 2008). In an attempt to remedy this problem the authors used a  
22 newly developed Deletion/Substitution/Addition (DSA) algorithm “to identify which pollutant metrics  
23 were most predictive of pulmonary function” (Mortimer et al., 2008). This methodology uses an  
24 exploratory process to identify the best predictive model for each outcome of interest. Using this  
25 approach, Mortimer et al. (2008) found that exposure to CO early in life (ages 0-3) was negatively  
26 associated with FEV<sub>1</sub>/FVC and FEF<sub>25-75</sub>/FVC resulting in an effect size of -2.5% and -4.8%, respectively,

1 per IQR increase in CO.<sup>1</sup> Additional negative associations were observed between exposure to CO during  
2 the first 6 years of life and FEF<sub>25</sub> (-6.7%), and prenatal, not trimester-specific exposure, and FEF<sub>25-75</sub> (%  
3 reduction not reported). Overall, Mortimer et al. (2008) found that the effects were limited to subgroups,  
4 including African Americans, individuals diagnosed with asthma before the age of 2 years, and  
5 individuals exposed to maternal smoking during pregnancy. It must be noted that research still needs to be  
6 conducted to validate the aforementioned results obtained using the DSA algorithm and the subsequent  
7 calculation of effect estimates using GEE because the current model could underestimate the uncertainty  
8 surrounding the associations reported (Mortimer et al., 2008). Although the authors did find associations  
9 between long-term exposure to CO and decrements in pulmonary function, they also observed high  
10 correlations between CO and NO<sub>2</sub>, which together are markers for pollutants generated by urban  
11 combustion sources (e.g., mobile sources) (Mortimer et al., 2008).

### 5.5.2.2. Asthma and Asthma Symptoms

12 All of the studies that examined the association between long-term exposure to CO and asthma  
13 and/or asthma symptoms presented consistent, positive results, which is similar to what was observed  
14 when evaluating the effects attributed to short-term exposure to CO. Two U.S.-based studies, Goss et al.  
15 (2004) and Meng et al. (2007), were evaluated which examined the effect of long-term exposure to CO  
16 on: (1) pulmonary exacerbations in a cohort of individuals with cystic fibrosis >6 years of age, and (2)  
17 poorly controlled asthma in a population of asthmatics ≥ 18 years old that resided close to cross-street  
18 intersections in Los Angeles County and San Diego County, respectively. Of these two studies only Goss  
19 et al. (2004) observed a positive association between a respiratory outcome (i.e., pulmonary  
20 exacerbations) and long-term exposure to CO. However, it is unclear if the effects observed are due to CO  
21 alone because the authors did not conduct co-pollutant analyses.

22 When evaluating studies conducted in other countries, Hirsch et al. (1999), in a study conducted in  
23 Germany, and Guo et al. (1999), Wang et al. (1999), and Hwang et al. (2005), in studies conducted in  
24 Taiwan, all found positive associations between long-term exposure to CO and asthma or asthma  
25 symptoms in populations ranging from 6–16 years old. In these studies, the authors addressed the  
26 observed associations differently. Guo et al. (1999) and Hwang et al. (2005) both concluded that it is  
27 unlikely CO directly affects the respiratory system; Hirsch et al. (1999) attributed the increase in the  
28 prevalence of cough and bronchitis to exposure to traffic-related air pollutants (i.e., NO<sub>2</sub>, CO, and  
29 benzene); and Wang et al. (1999) did not interpret the association between long-term exposure to CO and  
30 adolescent asthma. Only Hwang et al. (2005) conducted a co-pollutant analysis and found that the asthma  
31 effects observed were robust to the inclusion of PM<sub>10</sub>, SO<sub>2</sub> and O<sub>3</sub> in the model. This is of note because

---

<sup>1</sup> The study did not present the IQR for CO; therefore, the effect estimates presented were not standardized using the approach mentioned previously in this ISA.

1 inclusion of the traffic-related pollutants NO<sub>x</sub> or NO<sub>2</sub> in some of the aforementioned studies complicated  
2 the overall results because NO<sub>x</sub> and NO<sub>2</sub> are highly correlated with CO, which makes it difficult to  
3 separate the effects attributed to each pollutant.

### 5.5.2.3. Allergic Rhinitis

4 Hwang et al. (2006) and Lee et al. (2003c) both examined the effect of long-term exposure to air  
5 pollutants on the prevalence of allergic rhinitis in a population of schoolchildren in Taiwan. Both studies  
6 found an association between allergic rhinitis prevalence and CO, but they also observed an association  
7 with NO<sub>x</sub>. As a result, although Hwang et al. (2006) and Lee et al. (2003c) observed an increase in the  
8 prevalence of allergic rhinitis in response to an increase in long-term CO levels, they concluded that the  
9 combination of an association being observed for both CO and NO<sub>x</sub> can be attributed to the complex  
10 mixture of traffic-related pollutants and not necessarily CO alone.

### 5.5.2.4. Summary of Associations between Long-Term Exposure to CO and Respiratory Morbidity

11 To date, a limited number of studies have examined the potential association between long-term  
12 exposure to CO and respiratory morbidity. Although studies have reported positive associations for  
13 various respiratory outcomes, the limited evidence available, the new analytical methods employed, and  
14 the lack of studies that examined potential confounders of the CO-respiratory morbidity relationship,  
15 especially due to the high correlation between CO and other traffic-related pollutants, brings into question  
16 the validity of the associations observed.

## 5.5.3. Controlled Human Exposure Studies

17 Human clinical studies provide very little and inconsistent evidence of changes in pulmonary  
18 function following exposure to CO. In one older study, Chevalier et al. (1966) observed a significant  
19 decrease in total lung capacity following a short term exposure to 5,000 ppm resulting in a COHb level of  
20 4%. However, a similar study conducted at a higher CO concentration resulting in COHb levels of  
21 17-19% found no CO-induced changes in lung volume or mechanics (Fisher et al., 1969). The 2000 CO  
22 AQCD reported no evidence of CO-induced changes in exercise ventilation at COHb levels <15% during  
23 submaximal exercise (Koike et al., 1991). In two recent human clinical studies, exposure to CO (COHb ≈  
24 10%) was not found to significantly affect resting pulmonary ventilation compared with exposure to clean  
25 air under either hypoxic or hyperoxic exposure conditions (Ren et al., 2001; Vesely et al., 2004). The  
26 results of these studies demonstrate that the hypoxia- and CO<sub>2</sub>-induced increases in pulmonary ventilation

1 are not affected by CO. One recent study evaluated the potential anti-inflammatory effects of controlled  
2 exposures to CO in the airways of 19 individuals with COPD (Bathoorn et al., 2007). Subjects were  
3 exposed to both CO at concentrations of 100-125 ppm as well as room air for 2 h on each of four  
4 consecutive days. The authors reported a small decrease in sputum eosinophils, as well as a slight increase  
5 in the provocative concentration of methacholine required to cause a 20% reduction in FEV<sub>1</sub> following  
6 exposure to CO. Although this study appears to demonstrate some evidence of an anti-inflammatory  
7 effect of CO among subjects with COPD, it must be noted that two of these patients experienced  
8 exacerbations of COPD during or following CO exposure. A similar study found no evidence of systemic  
9 anti-inflammatory effects following exposure to higher CO concentrations (500 ppm for 1 h) in a group of  
10 healthy adults (Mayr et al., 2005).

#### 5.5.4. Toxicological Studies

11 As discussed in Section 5.2.3., the work of Thom, Ischiropoulos and colleagues (Ischiropoulos et  
12 al., 1996; Thom and Ischiropoulos, 1997; Thom et al., 1997; Thom et al., 1999a; Thom et al., 1999b)  
13 focused on CO-mediated displacement of NO• from heme-binding sites. Although the concentrations of  
14 CO used in many of their studies were far higher than ambient levels, some of this research involved  
15 more environmentally-relevant CO levels. In one study, 1-h exposure of rats to 50 ppm CO resulted in  
16 increased lung capillary leakage 18 h later (Thom et al., 1999a). Increased NO• was observed in the lungs  
17 by electron paramagnetic resonance during 1-h exposure to 100 ppm CO and was accompanied by  
18 increases in H<sub>2</sub>O<sub>2</sub> and nitrotyrosine. All of these effects were blocked by inhibition of NOS. These results,  
19 which were partially discussed in the 2000 CO AQCD, demonstrate the potential for exogenous CO to  
20 interact with NO•-mediated pathways and to lead to pathophysiological effects in the lung.

21 Recent work by Ghio et al. (2008) showed a disruption of cellular iron homeostasis following  
22 exposure to a low level of CO (50 ppm x 24 h) in rats. In lungs of inhalation-exposed rats, non-heme iron  
23 was significantly reduced, while lavagable iron was increased dramatically, suggesting an active removal  
24 of cellular iron. Lavagable ferritin was also increased following the CO exposure. Concurrently, liver iron  
25 levels increased, implying that the anatomical distribution of iron stores may significantly shift  
26 during/after CO exposures. These investigators were able to replicate the effect of loss of cellular iron in  
27 an in vitro model of cultured BEAS-2B cells and reported statistically significant effects at 10 ppm CO  
28 and an apparent maximal effect at 50 ppm CO (concentrations up to 500 ppm did not significantly  
29 enhance the iron loss beyond 50 ppm). Similar responses were observed for cellular ferritin. Both  
30 enhancement of iron removal and diminished iron uptake were noted in CO-exposed cells. Furthermore,  
31 decreased oxidative stress, mediator release and proliferation were noted in respiratory cells. These effects  
32 were reversible with a recovery period in fresh air. Interestingly, the in vivo exposure to CO induced mild,



1 but significant neutrophilia in the lungs compared to air-exposed rats. This finding is contrary to the  
2 concept that CO acts as an anti-inflammatory agent; however, with alterations in iron handling several  
3 potential pathways could be initiated to recruit inflammatory cells into airways. The authors pointed out  
4 that while CO derived from HO activity may have an important role in iron regulation, the non-specific  
5 application of exogenous CO will have little capacity to discriminate between excessive and/or  
6 inappropriate iron which catalyzes oxidative stress and iron which may be required for normal  
7 homeostasis.

8 A chronic inhalation study by Sorhaug et al. (2006) demonstrated no alterations in lung  
9 morphology in Wistar rats exposed to 200 ppm CO for 72 weeks. COHb levels were reported to be 14.7%  
10 and morphological changes were noted in the heart as described in Section 5.2.3.

11 A recent study by Carraway et al. (2002) involved continuous exposure of rats to HH (380 torr)  
12 with or without co-exposure to CO (50 ppm) for up to 21 days. The focus of this study was on remodeling  
13 of the pulmonary vasculature. While the addition of CO to HH did not alter the thickness or diameter of  
14 vessels in the lung, there was a significant increase in the number of small (<50 µm) diameter vessels  
15 compared to control, HH only, and CO-only exposures. Despite the greater number of vessels, the overall  
16 pulmonary vascular resistance was increased in the combined CO + HH exposure, which the authors  
17 attribute to enhancement of muscular arterioles and β-actin.

18 In summary, one older study (Thom et al., 1999a) and two new studies (Carraway et al., 2002; Ghio  
19 et al., 2008) demonstrated effects of 50-100 ppm CO on the lung. Responses included an increase in  
20 alveolar capillary permeability, disrupted iron homeostasis, mild pulmonary inflammation and an  
21 exacerbation of pulmonary vascular remodeling elicited by HH. These results should be considered in  
22 view of the potential for inhaled CO to interact directly with lung epithelial cells and resident  
23 macrophages. However, a chronic study involving 200 ppm CO demonstrated no changes in pulmonary  
24 morphology (Sorhaug et al., 2006).

## 5.5.5. Summary of Respiratory Health Effects

### 5.5.5.1. Short-Term Exposure to CO

25 New epidemiologic studies, supported by the body of literature summarized in the 2000 CO  
26 AQCD, provide evidence of positive associations between short-term exposure to CO and respiratory-  
27 related outcomes including pulmonary function, respiratory symptoms, medication use, hospital  
28 admissions, and ED visits. However, the interpretation of the results from epidemiologic studies is  
29 difficult due to the lack of extensive copollutant analyses along with the moderate to high correlation  
30 between CO and other combustion/traffic generated pollutants. To date the majority of the literature has  
31 not extensively examined the association between CO and respiratory morbidity due to studies focusing

1 primarily on effects associated with exposure to other criteria pollutants, namely PM and O<sub>3</sub>. This has  
2 contributed to the inability to disentangle the effects attributed to CO from the larger complex air  
3 pollution mix (particularly motor vehicle emissions). In addition, uncertainty as to a biological  
4 mechanism to explain the respiratory-related effects observed in the epidemiologic literature further  
5 complicates the interpretation of these results, especially considering the low ambient CO concentrations  
6 reported (24-h avg: 0.35-2.1 ppm). However, animal toxicological studies do provide some evidence that  
7 short-term exposure to CO (50-100 ppm) can cause oxidative injury and inflammation and alter  
8 pulmonary vascular remodeling. Human clinical studies have not extensively examined the effect of  
9 short-term exposure on respiratory morbidity, specifically pulmonary function. The limited number of  
10 human clinical studies that have been conducted prior to and since the 2000 CO AQCD provide very little  
11 evidence of any adverse effect of CO on the respiratory system at COHb levels <10%. Although human  
12 clinical studies have not provided evidence to support CO-related respiratory health effects, the  
13 epidemiologic studies that examined the effects of short-term exposure to CO and lung-related outcomes  
14 show positive associations and animal toxicological studies demonstrate the potential for an underlying  
15 biological mechanism, which together provide **evidence that is suggestive of a causal relationship**  
16 **between short-term exposure to relevant CO concentrations and respiratory morbidity.**

#### 5.5.5.2. Long-Term Exposure to CO

17 Currently, only a few studies have been conducted that examine the association between long-term  
18 exposure to CO and respiratory morbidity. Although some studies did observe associations between long-  
19 term exposure to CO and respiratory health outcomes key uncertainties still exist. These uncertainties  
20 include: the lack of replication and validation studies to evaluate new methodologies (i.e.,  
21 Deletion/Substitution/Addition (DSA) algorithm) that have been used to examine the association between  
22 long-term exposure to CO and respiratory health effects; whether the respiratory health effects observed  
23 in response to long-term exposure to CO can be explained by the proposed biological mechanisms; and  
24 the lack of co-pollutant analyses to disentangle the respiratory effects associated with CO due to its high  
25 correlation with NO<sub>2</sub> and other combustion-related pollutants. Overall, **the evidence available is**  
26 **inadequate to conclude that a causal relationship exists between long-term exposure to relevant**  
27 **CO concentrations and respiratory morbidity.**

## 5.6. Mortality

### 5.6.1. Epidemiologic Studies with Short-Term Exposure to CO

1           The relationship between short-term exposure to CO and mortality has not been extensively  
2 examined over the years due to the majority of epidemiologic studies focusing on mortality effects  
3 associated with exposure to PM and O<sub>3</sub>. As a result, a clear understanding of the association between  
4 short-term exposure to CO and mortality has yet to be developed. This section summarizes the main  
5 findings of the 2000 CO AQCD, and evaluates the newly available information on the relationship  
6 between short-term exposure to CO and daily mortality in an effort to disentangle the CO-mortality effect  
7 from those effects attributed to other criteria air pollutants.

#### 5.6.1.1. Summary of Findings from 2000 CO AQCD

8           The 2000 CO AQCD examined the association between short-term exposure to CO and mortality  
9 through the analysis of twelve single-city time-series studies, and one multi-city study, which included 11  
10 Canadian cities. While the results presented by these studies did provide suggestive evidence that an  
11 association exists between CO and mortality the AQCD concluded that inadequate evidence existed to  
12 infer a causal association between mortality and short-term exposure to ambient concentrations of CO.  
13 The reasons for this conclusion, which are shared with those studies that examined the effect of short-  
14 term exposure to CO on morbidity, were due to: internal inconsistencies and lack of coherence of the  
15 reported results within and across studies; the representativeness of the average ambient CO levels of  
16 spatially heterogeneous ambient CO values derived from fixed monitoring sites or of personal exposures  
17 that often include nonambient CO; and the lack of biological plausibility for any harmful effects  
18 occurring with the very small changes in COHb levels (from near 0 up to 1.0%) over typical baseline  
19 levels (about 0.5%) that would be expected with the low average ambient CO levels (<5.0 ppm, 1-h daily  
20 max) reported in the epidemiologic studies (U.S. EPA, 2000). Additionally, some epidemiologic studies  
21 have also suggested that CO is acting as an indicator for other combustion-related pollutants, which has  
22 led to investigators questioning the CO-mortality relationship even when associations have been  
23 observed.

24           To date the aforementioned issues have not been addressed primarily due to the majority of the  
25 recent time-series mortality studies focusing on the effects of only PM and O<sub>3</sub>. As such, CO has usually  
26 been considered one of the potential confounding copollutants in air pollution epidemiologic studies. As a  
27 result, the available CO information from these PM and O<sub>3</sub> studies most frequently consists of risk  
28 estimates from single- and multipollutant models. Given the limitation that most of these studies were not

1 conducted to examine CO, the goal of this review is to evaluate the CO-mortality association, and  
 2 specifically the: magnitude of associations; evidence of confounding; and evidence of effect modification.

### 5.6.1.2. Multi-City Studies

3 The following sections evaluate the recent literature that examined the association between short-  
 4 term exposure to CO and mortality, and in addition discuss newly available information with regard to the  
 5 issues specific to CO mentioned above. This evaluation focuses primarily on multi-city studies because  
 6 they provide: a more representative sample of potential CO-related mortality effects; and especially useful  
 7 information by analyzing data from multiple cities using a consistent method, and thus avoiding potential  
 8 publication bias.<sup>1</sup> Table 5-16 lists the multi-city studies evaluated along with the mean CO concentrations  
 9 reported in each study.

**Table 5-16. Range of CO concentrations reported in multi-city studies that examine mortality effects associated with short-term exposure to CO.**

Author	Location	Years	Averaging Time	Mean Concentration (ppm)	Range of Mean Concentrations Across Cities (ppm)
Dominici et al. (2003b; 2005a) (reanalysis of Samet et al., 2000b)	82 U.S. cities <sup>1</sup> (NMMAPS)	1987-1994	24-h avg	1.02	Baton Rouge = 0.43 Spokane = 2.19
Burnett et al. (2004)	12 Canadian cities	1981-1999	24-h avg	1.02	Winnipeg = 0.58 Toronto = 1.31
Samoli et al. (2007) <sup>2</sup>	19 European cities (APHEA2)	1990-1997 <sup>3</sup>	8-h max	2.12	Basel = 0.52 Athens = 5.3

<sup>1</sup> The study actually consisted of 90 U.S. cities, but only 82 had CO data.

<sup>2</sup> This study presented CO concentrations in the units mg/m<sup>3</sup>. The concentrations were converted to ppm using the conversion factor 1 ppm = 1.15 mg/m<sup>3</sup>, which assumes standard atmosphere and temperature.

<sup>3</sup> The study period varied from city to city. These years represent the total years in which data was collected across all cities.

### National Morbidity, Mortality, and Air Pollution Study of 90 U.S. Cities

10 The time-series analysis of the largest 90 U.S. cities (82 cities for CO) in the National Morbidity,  
 11 Mortality, and Air Pollution Study (NMMAPS) Dominici et al. (2003b; 2005a) (reanalysis of Samet et al.,  
 12 2000b) is by far the largest multi-city study conducted to date to investigate the mortality effects of air  
 13 pollution, but the study primarily focused on PM<sub>10</sub>. The range in 24-h avg CO concentrations in the  
 14 largest 20 cities (by population size) was 0.66 ppm (Detroit, MI) to 2.04 ppm (New York City). The  
 15 analysis in the original report used GAM with default convergence criteria. In response to the bias

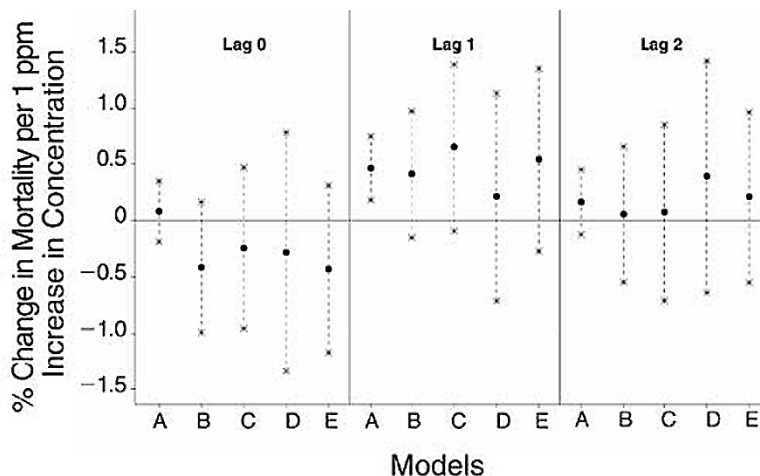
<sup>1</sup> To compare studies in this section that used different averaging times, effects estimates were standardized to the following: 0.5 ppm for studies that used 24-h avg concentrations and 0.75 ppm for studies that used max 8-h avg concentrations. These standardized values represent the range of current mean ambient concentrations across the U.S.

1 observed in the estimates generated using GAM models with default convergence criteria (Dominici et  
2 al., 2002), Dominici et al. (2003b; 2005a) (reanalysis of Samet et al., 2000b) used the data using GAM  
3 with stringent convergence criteria as well as GLM.

4 Focusing on the results obtained using GLM, PM<sub>10</sub> and O<sub>3</sub> (in summer) appeared to be more  
5 strongly associated with mortality than the other gaseous pollutants. The authors stated that the results did  
6 not indicate associations of CO, SO<sub>2</sub>, or NO<sub>2</sub>, with total (non-accidental) mortality. However, as with  
7 PM<sub>10</sub>, the gaseous pollutants CO, SO<sub>2</sub>, and NO<sub>2</sub> each showed the strongest association at a 1-day lag (for  
8 O<sub>3</sub>, a 0-day lag). Figure 5-14 presents the total mortality risk estimates for CO from Dominici et al.  
9 (2003b). The authors found a mortality risk estimate of 0.23% (95% PI: 0.09, 0.36) per 0.5 ppm increase  
10 in 24-h avg CO for a 1-day lag in a single-pollutant model. The inclusion of PM<sub>10</sub> or PM<sub>10</sub> and O<sub>3</sub> in the  
11 model did not reduce CO risk estimates. However, the confidence bands were wider in the multipollutant  
12 models, but this could be attributed to: (1) PM<sub>10</sub> data in many of the cities being collected every 6th day,  
13 as opposed to daily data for gaseous pollutants; and (2) O<sub>3</sub> being collected in some cities only during  
14 warm months. The addition of NO<sub>2</sub> (along with PM<sub>10</sub>) to the model resulted in a reduced CO risk  
15 estimate. Some caution is required when interpreting this apparent reduction because a smaller number of  
16 cities (57 cities<sup>1</sup>) were available for the CO multipollutant analysis with PM<sub>10</sub> and NO<sub>2</sub> compared to the  
17 single-pollutant CO analysis (82 cities). However, most of the cities that did not have NO<sub>2</sub> data (26 out of  
18 32), and subsequently were not included in the multipollutant analysis, were some of the least populated  
19 cities. Thus, the difference in the number of cities in the multi- and single-pollutant analyses is unlikely to  
20 be the underlying cause for the reduction in the CO risk estimate in the CO multipollutant analysis with  
21 PM<sub>10</sub> and NO<sub>2</sub>. In comparison to the PM<sub>10</sub> risk estimates, which were not reduced in multipollutant  
22 models, the CO risk estimates from multipollutant models indicate less consistent associations with  
23 mortality.

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<sup>1</sup> One city was excluded from the multipollutant analysis because it contained NO<sub>2</sub> data, but did not contain CO data.



Source: Dominici et al. (2003b)

**Figure 5-14. Posterior means and 95% posterior intervals of national average estimates for CO effects on total (non-accidental) mortality at lags 0, 1, and 2 within sets of the 90 U.S. cities with available pollutant data. Models A = CO alone; B = CO + PM<sub>10</sub>; C = CO + PM<sub>10</sub> + O<sub>3</sub>; D = CO + PM<sub>10</sub> + NO<sub>2</sub>; E = CO + PM<sub>10</sub> + SO<sub>2</sub>.**

## Canadian Multi-City Studies

1 Since the 2000 CO AQCD two Canadian multi-city studies have been published that examined the  
 2 association between mortality and short-term exposure to air pollutants: (1) an analysis of PM<sub>10</sub>, PM<sub>2.5</sub>,  
 3 PM<sub>10-2.5</sub>, and gaseous pollutants in 8 cities from 1986 to 1996 (Burnett et al., 2000); and (2) an analysis of  
 4 PM<sub>10</sub>, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, and gaseous pollutants in 12 cities from 1981 to 1999 (Burnett et al., 2004). The  
 5 2000 study utilized GAM with default convergence criteria, and upon reanalysis only examined PM  
 6 indices (Burnett and Goldberg, 2003).

7 Burnett et al. (2004) is the most extensive Canadian multi-city study conducted to date, both in  
 8 terms of the length of the study and the number of cities covered. Although the study focused on NO<sub>2</sub>  
 9 because it was the best predictor of short-term mortality fluctuations among the pollutants examined  
 10 (NO<sub>2</sub>, O<sub>3</sub>, SO<sub>2</sub>, CO, PM<sub>2.5</sub>, and PM<sub>10-2.5</sub>), it did present single- and copollutant results for all pollutants  
 11 included in the analysis. The mean CO concentrations reported by Burnett et al. (2004) are similar to  
 12 those reported in NMMAPS (see Table 5-14).

13 Burnett et al. (2004) examined the effect of short-term exposure to CO on total (non-accidental)  
 14 mortality. The authors found the strongest mortality association at lag 1-day for CO, SO<sub>2</sub>, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>,  
 15 PM<sub>10</sub> (arithmetic addition of PM<sub>2.5</sub> and PM<sub>10-2.5</sub>), and CoH, whereas for NO<sub>2</sub>, it was the 3-day moving  
 16 average (i.e., average of 0-, 1-, and 2-day lags), and for O<sub>3</sub>, it was the 2-day moving average. In this study,  
 17 Burnett et al. (2004) used 24-h avg pollutant concentrations because these values showed stronger  
 18 associations than the daily 1-h max values for all of the gaseous pollutants and CoH, but not for O<sub>3</sub>. In a  
 19 single-pollutant model the CO risk estimate for total mortality was 0.33% (95% CI: 0.12-0.54) per

1 0.5 ppm increase in 24-h avg CO with a 1-day lag. After adjusting for NO<sub>2</sub>, the CO risk estimate was  
2 reduced to 0.04% (95% CI: -0.19 to 0.26), while the NO<sub>2</sub> risk estimate was only slightly affected  
3 (increased from 2.25% to 2.35%) when including CO in the model. In this analysis, a copollutant model  
4 including both CO and PM was not presented. The similarity between the results presented in this  
5 Canadian multi-city and NMMAPS is that, in both analyses, CO risk estimates appeared to be sensitive to  
6 the addition of NO<sub>2</sub> in the regression model. However, interpretation of these results requires some  
7 caution because: (1) NO<sub>2</sub> tends to have a more spatially uniform distribution within a city compared to  
8 CO; (2) CO and NO<sub>2</sub> share common sources (e.g., traffic); and (3) CO and NO<sub>2</sub> are often moderately to  
9 highly correlated.

### **Air Pollution and Health: A European Approach**

10 Most of the Air Pollution and Health: A European Approach (APHEA) analyses have focused on  
11 the mortality effects of PM (PM<sub>10</sub> and BS), SO<sub>2</sub>, NO<sub>2</sub>, and O<sub>3</sub>, but not CO. In addition, some of the  
12 analyses have not even considered CO as a potential confounder, such as the extended analysis  
13 (APHEA2) of PM (Katsouyanni et al., 2001), and NO<sub>2</sub>. One study, Gryparis et al. (Gryparis et al., 2004)  
14 did consider CO as a potential confounder in an analysis of O<sub>3</sub>, and found that the addition of CO  
15 increased O<sub>3</sub> mortality risk estimates both in the summer and winter although the number of cities  
16 included in the copollutant model were reduced from 21 to 19. However, the study did not present CO  
17 risk estimates. Unlike other APHEA studies (or the NMMAPS and Canadian multi-city studies), the  
18 Samoli et al. (2007) analysis focused specifically on CO.

19 Samoli et al. (2007) investigated the effect of short-term exposure to CO on total (non-accidental)  
20 and cardiovascular mortality in 19 European cities participating in the APHEA2 project by using a two-  
21 stage analysis to examine city-specific effects and potential sources of heterogeneity in CO-mortality risk  
22 estimates. The mean levels of the max 8-h avg CO concentration in this study ranged from 0.52 ppm  
23 (Basel, Switzerland, and the Netherlands) to 5.3 ppm (Athens, Greece). The max 8-h avg CO  
24 concentration for the APHEA2 study of 2.12 ppm is higher than the estimated max 8-h avg CO  
25 concentrations reported in U.S. (Dominici et al., 2003b, 2005b) and Canadian (Burnett et al., 2004) cities  
26 of 1.53 ppm.<sup>1</sup> In APHEA cities, the correlation between CO and BS ( $r = 0.67-0.82$ ) was higher than the  
27 correlation between CO and PM<sub>10</sub> ( $r = 0.16-0.70$ ) or the correlation between CO and 1-h max NO<sub>2</sub>  
28 ( $r = 0.03-0.68$ ).

29 To examine the CO-mortality relationship, Samoli et al. (2007) conducted a time-series analysis of  
30 individual cities following the revised APHEA2 protocol.<sup>2</sup> The primary results presented by the authors  
31 are from a sensitivity analysis that used two alternative methods to select the extent of adjustment for

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<sup>1</sup> The max 8-h avg concentration for the Dominici et al. (2003b) and Burnett et al. (2004) studies was calculated using the conversion factor of 2:3 to convert 24-h avg concentrations to max 8-h avg concentrations.

<sup>2</sup> The APHEA2 protocol used a Poisson GAM model with penalized splines as implemented in the statistical package R.

1 temporal confounding. These methods consisted of: (1) confining the extent of smoothing to 8 degrees of  
2 freedom per year (df/yr); and (2) selecting the appropriate extent of smoothing through minimization of  
3 the absolute value of the sum of partial auto-correlation functions (PACF) of the residuals, which resulted  
4 in the analysis using on average 5 df/yr for total mortality and 4 df/yr for cardiovascular mortality. The  
5 authors also conducted copollutant analyses using PM<sub>10</sub>, BS, SO<sub>2</sub>, NO<sub>2</sub>, or O<sub>3</sub> (1 h). In the second stage  
6 model Samoli et al. (2007) examined heterogeneity in CO risk estimates between cities by regressing risk  
7 estimates from individual cities on potential effect modifiers including: a) the air pollution level and mix  
8 in each city (i.e., mean levels of pollutants, ratio PM<sub>10</sub>/NO<sub>2</sub>); b) the exposure (number of CO monitors,  
9 correlation between monitors' measurements); c) variables describing the health status of the population  
10 (e.g., crude mortality rate); d) the geographic area (northern, western, and central-eastern European  
11 cities); and e) the climatic conditions (mean temperature and relative humidity levels).

12 Samoli et al. (2007) found that CO was associated with total and cardiovascular mortality. The  
13 primary results represent the combined random effects estimate for a 0.75 ppm increase in max 8-h avg  
14 CO concentrations for the average of 0- and 1-day lag for total mortality (1.03% [95% CI: 0.55-1.53]) and  
15 for cardiovascular mortality (1.08% [95% CI: 0.25-1.90]). These results were obtained using PACF to  
16 choose the extent of adjustment for temporal trends. Although the results obtained using PACF are  
17 insightful, the use of 8 df/yr would have been more consistent with the NMMAPS model (7 df/yr), and  
18 would have allowed for a more accurate comparison of the results between APHEA2 and NMMAPS. The  
19 corresponding risk estimates obtained using the 8 df/yr model are: 0.57% (95% CI: 0.23-0.91) for total  
20 mortality and 0.70% (95% CI: 0.31-1.09) for cardiovascular mortality. In the sensitivity analysis, Samoli  
21 et al. (2007) used 8 or 12 df/yr to adjust for temporal confounding. Both approaches led to similar risk  
22 estimates, but using PACF to choose the extent of smoothing generally resulted in larger CO risk  
23 estimates (by ~50 to 80%).

24 During the examination of results obtained from the copollutant models, the authors noted that  
25 there was indication of confounding of CO risk estimates by BS and NO<sub>2</sub>, but not PM<sub>10</sub>. These results are  
26 consistent with CO, BS, and NO<sub>2</sub> being part of the traffic pollution mixture and PM<sub>10</sub> likely including  
27 secondary aerosols that do not correlate well with traffic-derived pollution. The risk estimates from the  
28 model using 8 df/yr that included NO<sub>2</sub> were: 0.26% (-0.09 to 0.61) for total mortality and 0.37% (-0.05 to  
29 0.80) for cardiovascular mortality. Thus, the inclusion of NO<sub>2</sub> in the model nearly halved the CO risk  
30 estimates (whereas the NO<sub>2</sub> risk estimate was not sensitive to the inclusion of CO in the model). A similar  
31 magnitude of reduction in the CO risk estimates was also observed when including BS in the model.  
32 Overall, the sensitivity of CO risk estimates to the inclusion of NO<sub>2</sub> in the model is consistent with the  
33 results presented in NMMAPS and the Canadian multi-city studies.

34 In the second stage model, Samoli et al. (2007) found that geographic region was the most  
35 significant effect modifier, while the other effect modifiers (mentioned above) did not result in strong  
36 associations. Effects were primarily found in western and southern European cities, and were larger in



1 cities where the standardized mortality rate was lower. Earlier APHEA studies also reported a regional  
2 pattern of air pollution associations for BS and SO<sub>2</sub>, and found that western cities showed stronger  
3 associations than eastern cities. However, the heterogeneity in CO risk estimates by geographic region  
4 does not provide specific information to evaluate the CO-mortality association.

5 An ancillary analysis conducted by Samoli et al. (2007) examined the possible presence of a CO  
6 threshold. The authors compared city-specific models to the threshold model, which consisted of  
7 thresholds at 0.5 mg/m<sup>3</sup> (0.43 ppm) increments. Samoli et al. (2007) then computed the deviance between  
8 the two models and summed the deviances for a given threshold over all cities. While the minimum  
9 deviance suggested a potential threshold of 0.43 ppm (the lowest threshold examined), the comparison  
10 with the linear no-threshold model indicated weak evidence (p-value >0.9) for a threshold. However,  
11 determining the presence of a threshold at the very low range of CO concentrations (i.e., at 0.43 ppm) in  
12 this data set is challenging, because, in seven of the 19 European cities examined, the lowest 10% of the  
13 CO distribution was at or above 2 mg/m<sup>3</sup> (1.74 ppm). Thus, the interpretation of the suggestive indication  
14 of a threshold is limited.

15 In summary, the APHEA2 analysis of CO in 19 cities found an association between CO and total  
16 and cardiovascular mortality in single-pollutant models, but the associations were substantially reduced  
17 when NO<sub>2</sub> or BS was included in copollutant models. The evidence for potential confounding of CO risk  
18 estimates by NO<sub>2</sub> is consistent with the findings from the NMMAPS and Canadian 12 cities studies. In  
19 addition, Samoli et al. (2007) found that geographic region was a potential effect modifier, but such  
20 geographic heterogeneity is not specific to CO, based on previously conducted APHEA studies. Finally,  
21 examination of the CO concentration-response relationship found weak evidence of a CO threshold.

### **Other European Multi-City Studies**

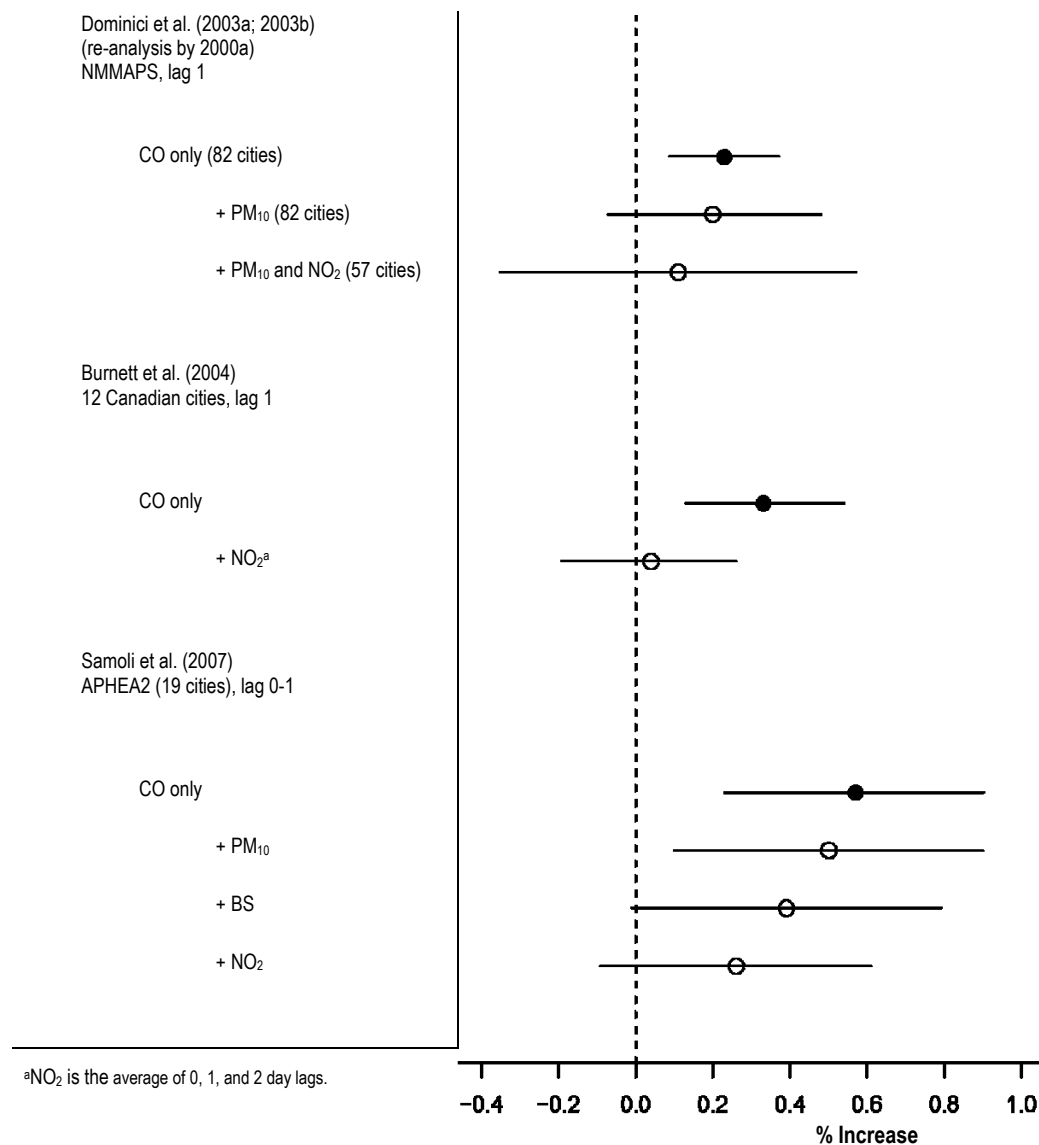
22 An additional European multi-city study was conducted by Biggeri et al. (2005) in eight Italian  
23 cities. The authors examined the effect of short-term exposure to CO on mortality in single-pollutant  
24 models using a time-series approach. In this analysis, because all of the pollutants showed positive  
25 associations with the mortality endpoints examined and the correlations among the pollutants were not  
26 presented, it is unclear if the observed associations are shared or confounded.

### **Summary of Multi-City Studies**

27 In summary, the mortality risk estimates from single-pollutant models are comparable for the  
28 NMMAPS and Canadian 12-city studies, 0.23 and 0.33, respectively, with the estimate from the APHEA2  
29 study being slightly larger (0.57%) (Figure 5-15). In both the NMMAPS and Canadian studies, a 1-day  
30 lag showed the strongest association; but the APHEA2 study used an a priori exposure window, the  
31 average of 0- and 1-day lags, which has been found to be the exposure window most strongly associated  
32 with mortality in PM analyses.

1           The APHEA2 risk estimates presented in Figure 5-15 are from a model that used a fixed amount of  
2 smoothing to adjust for temporal confounding (8 df/yr), which is similar to that used in the NMMAPS  
3 study (7 df/yr). However, the APHEA2 sensitivity analysis suggested an approximate 50 to 80%  
4 difference in CO risk estimates between the models that used 8 or 12 df/yr, and the models that used  
5 minimization of the absolute value of the sum of PACF of the residuals as a criterion to choose the  
6 smoothing parameters. Thus, some model uncertainty likely influences the range of CO risk estimates  
7 obtained from the studies evaluated.

8           The CO risk estimates from the aforementioned studies are also consistently sensitive to the  
9 inclusion of NO<sub>2</sub> in a copollutant model (0.11, 0.03, and 0.26%, for the NMMAPS, Canadian 12-cities  
10 study, and APHEA2, respectively). Thus, these results suggest confounding by NO<sub>2</sub>. However, this  
11 interpretation is further complicated because as with CO, NO<sub>2</sub> itself may be an indicator of combustion  
12 sources, such as traffic. Because CO measurements tend to reflect more local impacts, due to the location  
13 of monitors, than NO<sub>2</sub> (which is a secondary pollutant and therefore more spatially uniform) it is also  
14 possible that CO, the less precisely measured pollutant in terms of spatial distribution, may “lose” in the  
15 multipollutant model. Thus, it may not be accurate to interpret these results as evidence of “confounding  
16 by NO<sub>2</sub>.”



**Figure 5-15. Summary of mortality risk estimates for short-term exposure to CO from multi-city studies. Estimates were standardized to 0.5 ppm and 0.75 ppm for studies that used 24-h avg CO and max 8-h avg CO exposure metrics, respectively.**

### 5.6.1.3. Meta-Analysis of All Criteria Pollutants

1 Stieb et al. (2002) reviewed the time-series mortality studies published between 1985 and 2000,  
 2 and conducted a meta-analysis to estimate combined effects for PM<sub>10</sub>, CO, NO<sub>2</sub>, O<sub>3</sub>, and SO<sub>2</sub>. Because  
 3 many of the studies reviewed in the 2000 analysis used GAM with default convergence criteria,  
 4 Stieb et al. (2003) updated the estimates from the meta-analysis by separating the GAM versus non-GAM  
 5 studies. In this meta-analysis the authors also presented separate combined estimates for single- and  
 6 multipollutant models. Overall, there were more GAM estimates than non-GAM estimates for all of the

1 pollutants except SO<sub>2</sub>. For CO, 4 single-pollutant model risk estimates were identified, resulting in a  
 2 combined estimate of 3.18% (95% CI: 0.76-5.66) per 0.5 ppm increase in 24-h avg CO, and only  
 3 1 multipollutant model risk estimate (0.00% [95% CI: -1.71 to 1.74]) from the non-GAM studies. Thus,  
 4 for CO, this study did not provide useful meta-estimates because the number of studies that contributed to  
 5 the combined estimates for CO was rather small.

#### 5.6.1.4. Single-City Studies

6 In addition to the multi-city studies discussed above, there have also been several single-city U.S.-  
 7 and Canadian-based time-series mortality studies that examined CO. The single-city studies, similar to the  
 8 multi-city studies, often focused on the PM-mortality association, but also provided additional  
 9 information that is not available in the multi-city studies. Because the sample size used in each single-city  
 10 study is small, and subsequently results in wide confidence bands, a quantitative comparison of the results  
 11 from single- and multi-city studies is difficult. In addition, some studies do not present CO results  
 12 quantitatively adding to the inability to adequately compare studies. Table 5-17 lists the single-city studies  
 13 evaluated along with the mean CO concentrations reported in each study.

**Table 5-17. Range of CO concentrations reported in single-city studies that examine mortality effects associated with short-term exposure to CO.**

Study	Location	Years	Averaging Time	Mean Concentration (ppm)	Upper Percentile Concentrations (ppm)
De Leon et al. (2003)	New York, NY	1985-1994	24-h avg	2.45	95th: 4.04
Klemm et al. (2004)	Atlanta, GA	1998-2000	1-h max	1.31	Max: 7.40 75th: 1.66
Vedal et al. (2003)	Vancouver, BC, Canada	1994-1996	24-h avg	0.6	Max: 1.9 90th: 0.9
Villeneuve et al. (2003)	Vancouver, BC, Canada	1986-1999	24-h avg	1.0	Max: 4.9 90th: 1.6
Goldberg et al. (2003)	Montreal, Quebec, Canada	1984-1993	24-h avg	0.8	Max: 5.1 75th: 1.0
Hoek et al. (2000; 2001; Hoek, 2003)	The Netherlands	1986-1994	24-h avg	Entire Country: 0.46 Four Major Cities: 0.59	Max. Entire Country: 2.6 Four Major Cities: 4.6

<sup>1</sup> Study reported median CO concentrations.

#### Single-City Studies Conducted in the United States

14 De Leon et al. (2003) focused on the role of contributing respiratory diseases on the association  
 15 between air pollution (i.e., PM<sub>10</sub>, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and CO) and primary non-respiratory mortality

1 (circulatory and cancer) in New York City, NY during the period 1985-1994. This study only presented  
2 risk estimates graphically for each of the pollutants analyzed, except PM<sub>10</sub>. In single-pollutant models,  
3 PM<sub>10</sub>, CO, SO<sub>2</sub>, and NO<sub>2</sub> all showed the same pattern of association with circulatory mortality for  
4 individuals  $\geq 75$ , indicating a larger risk of death in individuals with contributing respiratory diseases  
5 compared to those without. In two-pollutant models, PM<sub>10</sub> and CO risk estimates were reduced, but each  
6 remained significantly positive.

7 Klemm et al. (2004) analyzed 15 air pollutants for their associations with mortality in Atlanta, GA,  
8 for a two-year period starting in August 1998. These pollutants included PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, UFP surface area  
9 and counts, aerosol acidity, EC, OC, SO<sub>4</sub><sup>2-</sup>, O<sub>3</sub>, CO, SO<sub>2</sub>, and NO<sub>2</sub>. This study presented risk estimates  
10 using three levels of smoothing (quarterly, monthly, and biweekly knots) for temporal trend adjustment,  
11 and suggested that the risk estimates were rather sensitive to the extent of smoothing. It should be noted  
12 that temporal smoothing using biweekly knots is a more aggressive modeling approach than the degrees  
13 of freedom approach used by most studies. In the single-pollutant models for all-cause mortality, the  
14 strongest association, which was also statistically significant, was found for PM<sub>2.5</sub>. CO, SO<sub>4</sub><sup>2-</sup>, and PM<sub>10-2.5</sub>  
15 also showed positive associations with all-cause mortality, but they were not significant (CO: Quarterly  
16 knots and Monthly Knots  $\beta = 0.00002$  [SE = 0.00001]; Biweekly knots  $\beta = 0.00001$  [SE = 0.00002]).  
17 However, CO was significantly associated with circulatory mortality in older adults ( $\geq 65$ ), and these  
18 associations remained when PM<sub>2.5</sub> was included in the model (results were presented graphically).

### Single-City Studies Conducted in Canada

19 Vedal et al. (2003) examined the association between short-term exposure to “low levels” of air  
20 pollution (i.e., PM<sub>10</sub>, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and CO) and daily mortality in Vancouver, British Columbia, Canada  
21 for the years 1994-1996. In this analysis, all of the risk estimates were presented graphically; however, the  
22 results suggested that O<sub>3</sub> in the summer and NO<sub>2</sub> in the winter showed the strongest associations with  
23 mortality. Vedal et al. (2003) found that CO was positively, but not significantly associated with mortality.  
24 Additionally, the association between short-term exposure to NO<sub>2</sub> and mortality was found to be  
25 consistent with the results from the Canadian multi-city study conducted by Burnett et al. (2004).

26 Villeneuve et al. (2003) also conducted an analysis using data from Vancouver, Canada, using a  
27 cohort of 550,000 individuals whose vital status was ascertained between 1986 and 1999. In this study,  
28 PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, TSP, CoH, PM<sub>10</sub>, SO<sub>4</sub><sup>2-</sup>, O<sub>3</sub>, CO, SO<sub>2</sub>, and NO<sub>2</sub> were examined for their associations with  
29 all-cause, cardiovascular, and respiratory mortality. When examining the association between gaseous  
30 pollutants and all-cause mortality in this data set, NO<sub>2</sub> and SO<sub>2</sub> showed the strongest associations, while  
31 the association between CO and all-cause mortality were generally weaker than those for NO<sub>2</sub> and SO<sub>2</sub>.  
32 For cardiovascular mortality, SO<sub>2</sub> risk estimates were smaller than those for NO<sub>2</sub> or CO, while for  
33 respiratory mortality, SO<sub>2</sub> showed the strongest associations. However, the wider confidence bands of

1 these categories and the smaller daily counts make it difficult to assess CO associations with cause-  
2 specific mortality outcomes.

3 Goldberg et al. (2003) contrasted associations between air pollution and mortality in individuals  
4 with underlying CHF vs. mortality in individuals who were identified as having CHF one year prior to  
5 death based on information from the universal health insurance plan in Montreal, Quebec, Canada, during  
6 the period 1984-1993. In this study, Goldberg et al. (2003) examined associations between PM<sub>2.5</sub>, CoH,  
7 SO<sub>4</sub><sup>2-</sup>, O<sub>3</sub>, CO, SO<sub>2</sub>, and NO<sub>2</sub>, and mortality. The authors found no association between any of the air  
8 pollutants and mortality with underlying CHF. However, Goldberg et al. (2003) found positive  
9 associations between air pollution and mortality in individuals diagnosed with CHF one year prior to  
10 death. Of the air pollutants examined, CoH, NO<sub>2</sub>, and SO<sub>2</sub> were most consistently associated with  
11 mortality for ages 65 and older, while CO showed positive but weaker associations compared to these  
12 three pollutants.

### **Single-City Studies Conducted in Other Countries**

13 Of the epidemiologic studies conducted in other countries that examine the association between  
14 short-term exposure to CO and mortality only those studies conducted in European countries that have  
15 CO levels comparable to the U.S. were evaluated. However, because Samoli et al. (2007) conducted a  
16 multi-city study of European cities that focused on short-term exposure to CO, there are only a few  
17 single-city studies that provide additional information, specifically those studies conducted in the  
18 Netherlands. The Netherland studies were evaluated because they provide risk estimates for multiple  
19 pollutants and cause-specific mortality, and consisted of relatively large sample sizes (i.e., the mortality  
20 time-series of the entire country was analyzed).

21 Hoek et al. (2000) re-analyzed by (Hoek, 2003) examined associations between air pollution and  
22 all-cause, cardiovascular, COPD, and pneumonia deaths in the entire Netherlands, the four major cities  
23 combined, and the entire country minus the four major cities for the period 1986 to 1994. The air  
24 pollutants analyzed included BS, PM<sub>10</sub>, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, CO, SO<sub>4</sub><sup>2-</sup> and NO<sub>3</sub><sup>-</sup>. In the single-pollutant  
25 models, all of the pollutants were significantly associated with all-cause mortality at lag 1-day and 0-  
26 6 days when using the entire Netherlands data set. In the two-pollutant model, CO risk estimates were  
27 reduced to null when PM<sub>10</sub>, BS, SO<sub>4</sub><sup>2-</sup> and NO<sub>3</sub><sup>-</sup> were included in the model while the risk estimates for  
28 these copollutants remained significantly positive. BS, CO, and NO<sub>2</sub> were highly correlated (r >0.85) in  
29 this data set, and the authors noted “all these pollutants should be interpreted as indicators for motorized  
30 traffic emissions” (Hoek et al., 2000). The authors found that O<sub>3</sub> showed the most consistent and  
31 independent associations with mortality and that the risk estimates for all of the pollutants were  
32 substantially higher in the summer months than in the winter months. Pneumonia deaths showed the  
33 largest risk estimates for most pollutants including CO. The result from the Hoek et al. (2000) study is  
34 somewhat in contrast to the result from the Samoli et al. (2007) multi-city study in that, in the Hoek et al.

1 (2000) analysis, CO was more sensitive to the addition of PM indices in copollutant models. This may be  
2 due to the high correlation between CO and PM indices in the Netherlands.

3 Hoek et al. (2001) reanalysis by (Hoek, 2003) analyzed the Netherlands data using more specific  
4 cardiovascular causes of death: MI and other IHD, arrhythmia, heart failure, cerebrovascular mortality,  
5 and embolism/thrombosis. In this analysis, the authors analyzed O<sub>3</sub>, BS, PM<sub>10</sub>, CO, SO<sub>2</sub>, and NO<sub>2</sub> in only  
6 single-pollutant models. For all of the pollutants, risk estimates were larger for arrhythmia, heart failure,  
7 and cerebrovascular mortality than for the combined cardiovascular mortality outcome. While the results  
8 suggested larger impacts of air pollution on more specific cardiovascular causes, the authors did not  
9 provide evidence of an association that was specific to any particular pollutant which would aid in  
10 addressing the question of confounding.

### **Summary of Single-City Studies**

11 Overall, it is difficult to identify a clear pattern of CO-mortality associations from the single-city  
12 studies evaluated because of the relatively small sample sizes, expected variation in the air pollution mix  
13 across cities, and differences in the analytical approaches used across studies. However, in all of these  
14 studies, CO showed, at least to some extent, positive associations with all-cause or circulatory/  
15 cardiovascular deaths. In addition the one study that examined specific cardiovascular causes of death  
16 (Hoek et al., 2001)reanalysis by (Hoek, 2003) did not find an association that is specific to CO. Although  
17 the extent of sensitivity of CO-mortality associations varied across studies, it could be attributed to the  
18 likely variation in the correlation between CO and copollutants.

#### **5.6.1.5. Summary of Mortality and Short-Term Exposure to CO**

19 Among the gaseous pollutants examined in time-series mortality studies, CO is the least frequently  
20 studied criteria air pollutant. Because CO was mostly treated as a potential confounder in these studies,  
21 the information available regarding the nature of the association between short-term exposure to CO and  
22 mortality is limited compared to the other pollutants. However, the recently available multi-city studies,  
23 which consist of larger sample sizes, and single-city studies generally confirmed the findings reported in  
24 the 2000 CO AQCD.

25 The multi-city studies which were evaluated, reported comparable CO mortality risk estimates for  
26 total (non-accidental) mortality with the APHEA2 European multi-city study (Samoli et al., 2007)  
27 showing slightly higher estimates for cardiovascular mortality in single-pollutant models. However, when  
28 examining potential confounding by copollutants these studies consistently showed that CO mortality risk  
29 estimates were reduced when NO<sub>2</sub> was included in the model, but this observation may not be  
30 “confounding” in the usual sense in that NO<sub>2</sub> may also be an indicator of other pollutants or pollution  
31 sources (i.e., traffic).

1           Only one of the multi-city studies focused specifically on the CO-mortality association (Samoli et  
2 al., 2007), the APHEA2 study), and in the process examined: (1) model sensitivity; (2) the CO-mortality  
3 C-R relationship; and (3) potential effect modifiers of CO mortality risk estimates. The APHEA study  
4 performed a sensitivity analysis, which indicated an approximate 50 - 80% difference in CO risk  
5 estimates from a reasonable range of alternative models. In addition, the study examined the CO-mortality  
6 concentration-response relationship through a grid search of varying threshold points, and found only  
7 weak evidence of a CO threshold at 0.5 mg/m<sup>3</sup> (0.43 ppm), but this result was complicated by the lowest  
8 10% of the CO distribution for seven of the 19 cities examined being at or above 2 mg/m<sup>3</sup> (1.74 ppm).  
9 The examination of a variety of city-specific variables to identify potential effect modifiers of the  
10 CO-mortality relationship found that geographic region explained most of the heterogeneity in CO  
11 mortality risk estimates with the CO-mortality associations being stronger in western and southern  
12 European cities than eastern cities (Samoli et al., 2007). A similar pattern has been reported for black  
13 smoke (BS) and SO<sub>2</sub> in previous APHEA studies, but the geographic variability observed does not  
14 provide specific information, which could be used to evaluate the CO-mortality association.

15           The results from the single-city studies are generally consistent with the multi-city studies in that  
16 some evidence of a positive association was found for mortality upon short-term exposure to CO.  
17 However, the CO-mortality associations were often, but not always, attenuated when other copollutants  
18 were included in the regression models. In addition, limited evidence was available to identify cause-  
19 specific mortality outcomes (e.g., cardiovascular causes of death) associated with short-term exposure to  
20 CO.

21           The evidence from the recent multi- and single-city studies suggests that an association between  
22 short-term exposure to CO and mortality exists, but limited evidence is available to evaluate cause-  
23 specific mortality outcomes associated with CO exposure; and it is unclear if CO is acting alone or as an  
24 indicator for other combustion-related pollutants. In addition, the results underscore the limitation of  
25 current analytical methods to disentangle the health effects associated with one pollutant in the complex  
26 air pollution mixture. Overall, the epidemiologic evidence **is suggestive of a causal relationship**  
27 **between short-term exposure to relevant CO concentrations and mortality.**

## 5.6.2. Epidemiologic Studies with Long-Term Exposure to CO

28           The 2000 CO AQCD did not evaluate the association between long-term exposure to CO and  
29 mortality because there were no studies at the time that examined this relationship. Since then there have  
30 been several new studies that examined the association between long-term exposure to CO and mortality,  
31 but it should be noted that these studies primarily focused on PM, and CO was only considered in these



1 studies as a potential confounder. Therefore, the information available from these new long-term exposure  
2 studies is somewhat limited, especially in comparison to that for PM.

### 5.6.2.1. U.S. Cohort Studies

#### American Cancer Society Cohort Studies

3 Pope et al. (1995) investigated associations between long-term exposure to PM and mortality  
4 outcomes in the ACS cohort. In this study, ambient air pollution data from 151 U.S. metropolitan areas in  
5 1981 were linked with individual risk factors in 552,138 adults who resided in these areas when enrolled  
6 in the prospective study in 1982. Death outcomes were ascertained through 1989. PM<sub>2.5</sub> and SO<sub>4</sub><sup>2-</sup> were  
7 associated with total (non-accidental), cardiopulmonary, and lung cancer mortality, but not with mortality  
8 for all other causes (i.e., non-accidental minus cardiopulmonary and lung cancer). Gaseous pollutants  
9 were not analyzed in Pope et al. (1995). Jerrett et al. (2003) (using data from (Krewski et al., 2000)  
10 conducted an extensive sensitivity analysis of the Pope et al. (1995). ACS data, augmented with  
11 additional gaseous pollutants data. Due to the smaller number of CO monitors available compared to  
12 SO<sub>4</sub><sup>2-</sup>, the number of metropolitan statistical areas (MSAs) included in the CO analysis were reduced  
13 (from 151 with SO<sub>4</sub><sup>2-</sup>) to 107. The mean annual CO concentrations in these MSAs ranged from 0.19 to  
14 3.95 ppm. CO was weakly negatively correlated with SO<sub>4</sub><sup>2-</sup> (r = -0.07). Among the gaseous pollutants  
15 examined (CO, NO<sub>2</sub>, O<sub>3</sub>, SO<sub>2</sub>), only SO<sub>2</sub> showed positive associations with mortality, and in addition was  
16 the only copollutant that reduced SO<sub>4</sub><sup>2-</sup> risk estimates. For CO, the relative risk estimates for total (non-  
17 accidental) mortality in single and copollutant models with SO<sub>4</sub><sup>2-</sup> was 0.99 (95% CI: 0.96-1.01) and 0.98  
18 (95% CI: 0.96-1.01), respectively, per 0.5 ppm increase in mean annual average CO concentrations.

19 Pope et al. (2002) conducted an extended analysis of the ACS cohort with double the follow-up  
20 time (to 1998) and triple the number of deaths compared to the original Pope et al. (1995) study. In  
21 addition to PM<sub>2.5</sub>, data for all of the gaseous pollutants were retrieved for the extended period and  
22 analyzed for their associations with mortality-specific outcomes. As in the 1995 analysis, the air pollution  
23 exposure estimates were based on the MSA-level averages. The authors found that PM<sub>2.5</sub> and SO<sub>4</sub><sup>2-</sup> were  
24 both associated with total, cardiopulmonary, and lung cancer mortality.<sup>1</sup> In this study, the CO analysis  
25 used two different data sets. The first data set consisted of 1980 data and 113 MSAs; while the second  
26 data set used averages of the years 1982-1998 and 122 MSAs. The authors found, when using the 1980  
27 data, that CO was not associated (risk estimates ~ 1) (See Figure 5-16) with all-cause, cardiopulmonary,  
28 lung cancer, or mortality for all other causes. However, the analysis of the 1982-1998 data found that CO  
29 was negatively (and significantly) associated with all-cause, cardio-pulmonary, and lung cancer mortality.  
30 It is unclear why significant negative associations were observed when analyzing the 1982-1998 data, but

---

<sup>1</sup> These results were presented graphically in Pope et al. (2002) and were estimated for Figure 5-16.

1 evidence from other mortality studies that examined the association between long-term exposure to CO  
2 and mortality does not suggest that CO elicits a protective effect.

### **Women’s Health Initiative Cohort Study**

3 Miller et al. (2007) studied 65,893 postmenopausal women between the ages of 50 and 79 years  
4 without previous CVD in 36 U.S. metropolitan areas from 1994 to 1998. They examined the association  
5 between one or more fatal or nonfatal cardiovascular events and air pollutant concentrations. Exposures to  
6 air pollution were estimated by assigning the annual mean levels of air pollutants measured at the nearest  
7 monitor to the location of residence of each subject on the basis of its five-digit ZIP code centroid, which  
8 allowed estimation of effects due to both within-city and between-city variation of air pollution. The  
9 investigators excluded monitors whose measurement objective focused on a single point source or those  
10 with “small measurement scale (0 to 100 meters).” Thus, presumably these criteria reduced some of the  
11 exposure measurement error associated with monitors that are highly impacted by local sources.

12 During the course of the study, a total of 1,816 women had one or more fatal or nonfatal  
13 cardiovascular events, including 261 cardiovascular-related deaths. Hazard ratios for the initial  
14 cardiovascular event were estimated. The following results are for models that only included subjects  
15 with non-missing exposure data for all pollutants (n = 28,402 subjects, resulting in 879 CVD events). In  
16 the single-pollutant models, PM<sub>2.5</sub> showed the strongest associations with CVD events among all  
17 pollutants (HR = 1.24 [95% CI: 1.04-1.48] per 10-μg/m<sup>3</sup> increase in annual average), followed by SO<sub>2</sub>  
18 (HR = 1.07 [95% CI: 0.95-1.20] per 5-ppb increase in the annual average). For CO the single-pollutant  
19 risk estimate was slightly (but not significantly) negative (HR = 0.96 [95%CI: 0.84-1.10]). In the  
20 multipollutant model, which included all pollutants (i.e., PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, SO<sub>2</sub>, NO<sub>2</sub>, and O<sub>3</sub>), the CO risk  
21 estimate was similar to the one presented in the single-pollutant model (HR = 0.96 [95% CI: 0.82-1.14]).  
22 In addition, CO was not associated with CVD events in a single pollutant model (HR = 1.00 [95%CI:  
23 0.90-1.10] per 0.5 ppm increase in mean annual average CO concentration) that used all available  
24 observations. This study did not examine the correlations among pollutants and, therefore, the extent of  
25 confounding could not be examined, but PM<sub>2.5</sub> was clearly the best predictor of cardiovascular events.

### **The Washington University-EPRI Veterans’ Cohort Mortality Studies**

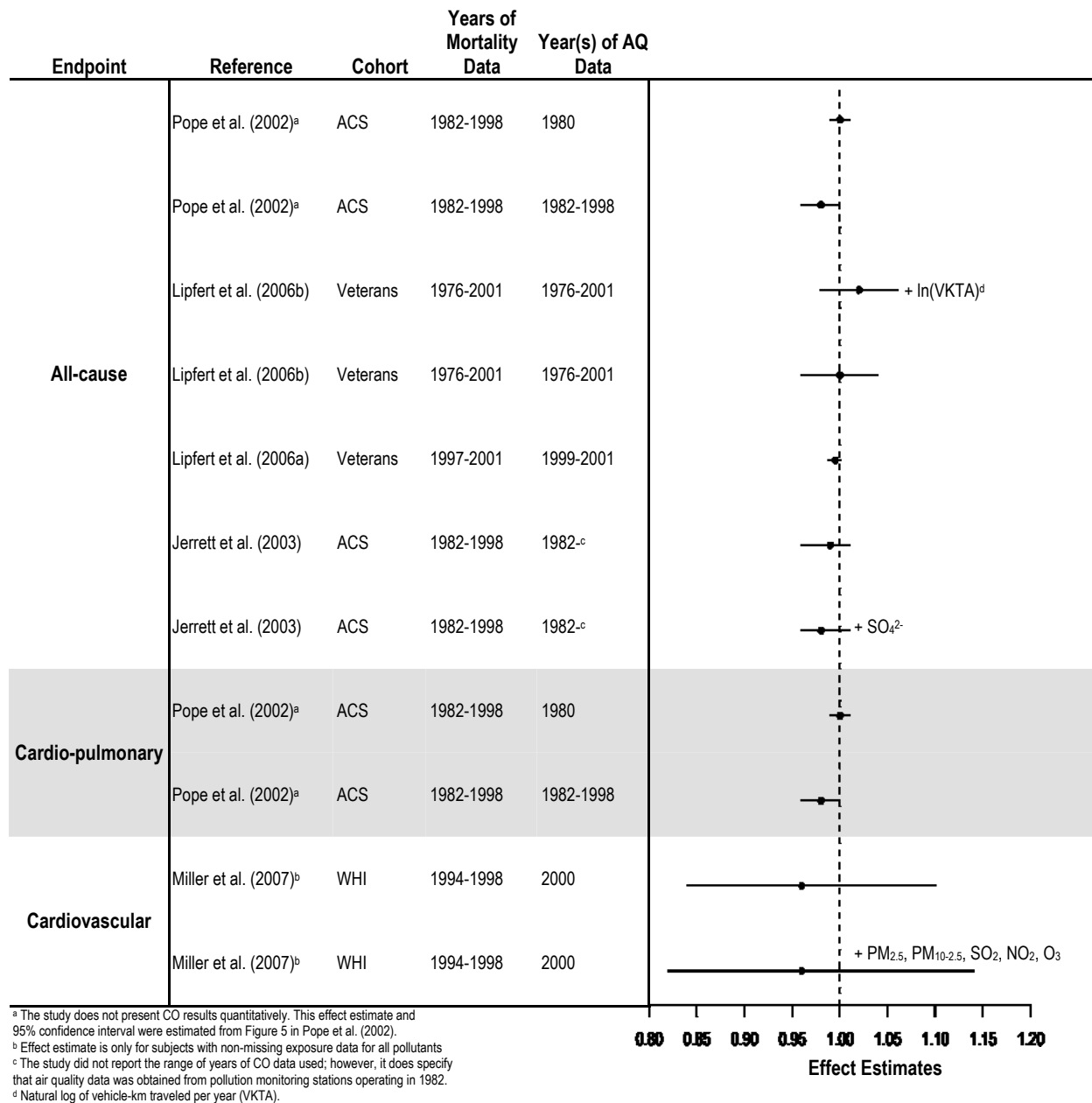
26 Lipfert et al. (2000a) conducted an analysis of a national cohort of ~70,000 male U.S. military  
27 veterans who were diagnosed as hypertensive in the mid 1970s and were followed for approximately 21  
28 years (up to 1996). Demographically, 35% of the cohort consisted of African American men and 57% of  
29 the cohort was defined as current smokers; however, 81% of the cohort had been smokers at one time in  
30 their life. The study examined mortality effects in response to long-term exposure to multiple pollutants  
31 including, PM<sub>2.5</sub>, PM<sub>10</sub>, PM<sub>10-2.5</sub>, TSP, SO<sub>4</sub><sup>2-</sup>, CO, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>, and Pb. Lipfert et al. (2000a) estimated  
32 exposures by indentifying the county of residence at the time of entry to the study. Four exposure periods

1 (1960-1974, 1975-1981, 1982-1988, and 1989-1996) were defined, and deaths during each of the three  
2 most recent exposure periods were considered. The mean annual 95th percentile of hourly CO values  
3 during these periods declined from 10.8 ppm to 2.4 ppm. The authors noted that the pollution risk  
4 estimates were sensitive to the regression model specification, exposure periods, and the inclusion of  
5 ecological and individual variables. Lipfert et al. (2000a) reported that indications of concurrent mortality  
6 risks (i.e., associations between mortality and air quality for the same period) were found for NO<sub>2</sub> and  
7 peak O<sub>3</sub>. The estimated CO mortality risks were all negative, but not significant.

8 Lipfert et al. (2006b) examined associations between traffic density and mortality in the same  
9 Veterans' Cohort, but in this analysis the follow-up period was extended to 2001. As in their 2000 study,  
10 four exposure periods were considered but more recent years were included in the 2006 analysis. The  
11 authors used the mean annual average of the 95th percentile of 24-h avg CO in each of the exposure  
12 periods as the averaging metric. The traffic density variable was the most significant predictor of  
13 mortality in their analysis, remaining so in two- and three pollutant models with other air pollutants  
14 (i.e., CO, NO<sub>2</sub>, O<sub>3</sub>, PM<sub>2.5</sub>, SO<sub>4</sub><sup>2-</sup>, non-SO<sub>4</sub><sup>2-</sup> PM<sub>2.5</sub>, and PM<sub>10-2.5</sub>). In the multipollutant models, mortality  
15 risk estimates were not statistically significant for any of the other pollutants, except O<sub>3</sub>. The natural log  
16 of the traffic density variable (VKTA = vehicle-km traveled per year) was not correlated with CO  
17 ( $r = -0.06$ ), but moderately correlated with PM<sub>2.5</sub> ( $r = 0.50$ ) in this data set. For the 1989-1996 data period,  
18 the estimated mortality relative risk was 1.02 (95% CI: 0.98-1.06) per 1 ppm increase in the mean annual  
19 95th percentile of hourly CO concentration in a single-pollutant model. The two-pollutant model, which  
20 included the traffic density variable, resulted in a relative risk of 1.00 (95% CI: 0.96-1.04). Lipfert et al.  
21 (2006b) note that the low risk estimates for CO in this study were consistent with those observed in other  
22 long-term exposure studies, and may have been due to the localized nature of CO, which can lead to  
23 exposure errors when data from centralized monitors is used to represent an entire county. Interestingly,  
24 as Lipfert et al. (2006b) pointed out, the risk estimates due to traffic density did not vary appreciably  
25 across these four periods even though regulated tailpipe emissions declined during the study period. The  
26 authors speculated that some combination of other environmental factors such as road dust, psychological  
27 stress, and noise (all of which constitute the environmental effects of vehicular traffic) along with spatial  
28 gradients in SES might contribute to the non-negative effects observed.

29 Lipfert et al. (2006a) extended the analysis of the Veterans Cohort data to include the EPA's  
30 Speciation Trends Network (STN) data, which collected chemical components of PM<sub>2.5</sub>. The authors  
31 analyzed the STN data for the year 2002, and again used county-level averages. In addition, they analyzed  
32 PM<sub>2.5</sub> and gaseous pollutants data for 1999 through 2001. As in the other Lipfert et al. (2006b) study,  
33 traffic density was the most important predictor of mortality, but associations were also observed for EC,  
34 vanadium (V), nickel (Ni), and NO<sub>3</sub><sup>-</sup>. Ozone, NO<sub>2</sub>, and PM<sub>10</sub> also showed positive, but weaker  
35 associations. The authors found no association between the mean annual 95th percentile of hourly CO

1 values and mortality (RR = 0.995 [95% CI: 0.988-1.001] per 1 ppm increase in CO concentration) in a  
 2 single-pollutant model. The study did not present multipollutant model results for CO.



**Figure 5-16. Summary of mortality risk estimates for long-term exposure to CO. Estimates were standardized to 0.5 ppm and 1.0 ppm for studies that used mean annual average CO and the mean annual 95th percentile of hourly CO values exposure metrics, respectively.**

### 5.6.2.2. U.S. Cross-Sectional Analysis

1 An ecological cross-sectional analysis involves regressing county- (or city) average health outcome  
2 values on county-average explanatory variables such as air pollution and census statistics. Unlike the  
3 cohort studies described above, to the extent that individual level confounders are not adjusted for, the  
4 cross-sectional study design is considered to be subject to ecologic confounding. However, all of the  
5 cohort studies described above are also semi-ecologic in that the air pollution exposure variables are  
6 ecologic (Kuenzli and Tager, 1997). In this sense, cross-sectional studies may be useful in evaluating the  
7 correlation among exposure variables.

8 Lipfert and Morris (2002) conducted ecological cross-sectional regressions for U.S. counties  
9 (except Alaska) during five periods: 1960-1969, 1970-1974, 1979-1981, 1989-1991, and 1995-1997. They  
10 regressed age-specific (15-44, 45-64, 65-74, 76-84, and 85+) all-cause (excluding AIDS and trauma)  
11 mortality on air pollution, demography, climate, SES, lifestyle, and diet. The authors analyzed TSP, PM<sub>10</sub>,  
12 PM<sub>2.5</sub>, SO<sub>4</sub><sup>2-</sup>, SO<sub>2</sub>, CO, NO<sub>2</sub>, and O<sub>3</sub>. However, air pollution data was only available for limited periods of  
13 time depending on the pollutant: TSP up to 1991; PM<sub>10</sub> between 1995 and 1999; and PM<sub>2.5</sub> between  
14 1979-1984 and 1999. In response to the varying number of counties with valid air pollution data by  
15 pollutant and time, Lipfert and Morris (2002) employed a staged regression approach. In the first stage, a  
16 national model was developed for each dependent variable, excluding air pollution variables. In the  
17 second stage, regressions were performed with the residuals on concurrent and previous periods' air  
18 pollution variables to identify the pollutants of interest. Overall, there were too many results to summarize  
19 because of the large number of age groups, lagged exposure time windows, and mortality study periods  
20 examined in the study, but similar to the ACS cohort studies, PM<sub>2.5</sub> and SO<sub>4</sub><sup>2-</sup> were found to be  
21 consistently and positively associated with mortality. Lipfert and Morris (2002) generally found the  
22 strongest associations in the earlier time periods, and when mortality and air quality were measured in  
23 different periods (e.g., mortality data 1995-1997 and CO data 1970-1974). Also, consistent with the  
24 Lipfert et al. (2000a) and the Pope et al. (2002) cohort studies, CO was frequently negatively (and often  
25 significantly) associated with mortality in older age groups, especially when mortality was matched with  
26 CO levels in more recent time periods. The younger age group (15-44) often showed a positive  
27 association with CO, but considering the small number of deaths attributed to this age group (less than  
28 1% of total deaths), the association was not informative. Overall, this study highlighted that the  
29 CO-mortality associations presented in purely ecologic study designs are generally consistent with those  
30 found in semi-individual cohort studies.

### 5.6.2.3. Summary of Mortality and Long-Term Exposure to CO

1           The evaluation of new epidemiologic studies conducted since the 2000 CO AQCD (2000) that  
2 investigated the association between long-term exposure to CO and mortality consistently found null or  
3 negative mortality risk estimates. No such studies were discussed in the 2000 CO AQCD. The re-analysis  
4 of the ACS data (Pope et al., 1995) by Jerrett et al. (2003) found no association between long-term  
5 exposure to CO and mortality. Similar results were obtained in an updated analysis conducted by Pope  
6 et al. (2002) when using earlier (1980) CO data, but negative associations were found when using more  
7 recent (1982-1998) data. The Women's Health Initiative (WHI) Study also found no association between  
8 CO and CVD events (including mortality) using the data from recent years (1994-1998) (Miller et al.,  
9 2007), while the series of Veterans Cohort studies found no association or a negative association between  
10 mean annual 95th percentile of hourly CO values and mortality (Lipfert et al., 2006a; Lipfert et al.,  
11 2006b). One additional study was identified that used a cross-sectional study design, Lipfert and Morris  
12 (2002), which reported results for a study of U.S. counties that are generally consistent with the cohort  
13 studies: positive associations between long-term exposure to PM<sub>2.5</sub> and SO<sub>4</sub><sup>2-</sup> and mortality, and generally  
14 negative associations with CO. Overall, the consistent null and negative associations observed across  
15 epidemiologic studies which included cohort populations encompassing potentially susceptible  
16 subpopulations (i.e., post-menopausal women and hypertensive men) combined with the lack of evidence  
17 for respiratory and cardiovascular morbidity outcomes following long-term exposure to CO; and the  
18 absence of a proposed mechanism to explain the progression to mortality following long-term exposure to  
19 CO provide supportive evidence that is **suggestive of no causal relationship between long-term**  
20 **exposure to CO and mortality.**

## 5.7. Public Health Impacts

21           This section addresses several issues relating to the broader public health impact from exposure to  
22 ambient CO through a discussion on: (1) the shape of the concentration-response (C-R) relationship for  
23 CO, which is based primarily on controlled human exposure studies; and (2) the identification of  
24 subpopulations which may experience increased risks from CO exposures, through either enhanced  
25 susceptibility (e.g., as a result of pre-existing disease, genetic factors, age) and/or vulnerability associated  
26 with external factors or differential exposure (e.g., altitude, close proximity to sources, activity patterns).

## 5.7.1. Concentration-Response Relationship

1           Currently, very limited information is available in the human clinical and epidemiologic literature  
2 regarding the CO C-R relationship and the potential presence of a CO threshold. Two human clinical  
3 studies described in the 2000 CO AQCD had evaluated the C-R relationship between CO and onset of  
4 exercise-induced angina among individuals with CAD, but at the high end of CO concentrations (i.e., CO  
5 levels above the current NAAQS). Anderson et al. (1973) exposed 10 adult men with stable angina for 4 h  
6 to CO concentrations of 50 and 100 ppm, which resulted in average COHb levels of 2.9% and 4.5%,  
7 respectively. Both exposures significantly decreased the time to onset of exercise-induced angina relative  
8 to room air control (1.6% COHb). However, there was no difference in response between the two  
9 exposure concentrations of CO. In a much larger study, 63 adults with stable angina were exposed for 1 h  
10 to two concentrations of CO (average exposure concentrations of 117 and 253 ppm) resulting in average  
11 pre-exercise COHb levels of 2.4% and 4.7% (Allred et al., 1989). Relative to control (average COHb  
12 0.7%), COHb levels of 2.4% and 4.7% were observed to decrease the time to onset of angina by 4.2%  
13 ( $p = 0.054$ ) and 7.1% ( $p = 0.004$ ), respectively. In addition, these investigators reported a statistically  
14 significant decrease in the time to exercise-induced ST-segment depression with increasing COHb levels.  
15 These findings provide some evidence of a significant C-R relationship at COHb concentrations between  
16 2.4 and 4.7%. However, the human clinical literature has yet to evaluate the C-R relationship at lower CO  
17 concentrations or COHb levels.

18           One study in the epidemiologic literature attempted to examine the C-R relationship at the low end  
19 of CO concentrations through a threshold analysis. Samoli et al. (2007) in their examination of the  
20 association between short-term exposure to CO and mortality conducted an ancillary analysis to examine  
21 the potential presence of a CO threshold. In this analysis the authors compared city-specific models to the  
22 threshold model, which consisted of thresholds at 0.5 mg/m<sup>3</sup> (0.43 ppm) increments. Samoli et al. (2007)  
23 then computed the deviance between the two models and summed the deviances for a given threshold  
24 over all cities. While the minimum deviance suggested a potential threshold of 0.43 ppm (the lowest  
25 threshold examined), the comparison with the linear no-threshold model indicated weak evidence  
26 ( $p$ -value  $>0.9$ ) for a threshold. However, determining the presence of a threshold at the very low range of  
27 CO concentrations (i.e., at 0.43 ppm) in this data set is challenging, because, in seven of the 19 European  
28 cities examined, the lowest 10% of the CO distribution was at or above 2 mg/m<sup>3</sup> (1.74 ppm). By only  
29 using the 12 cities in the analysis that had minimum CO concentrations approaching 0.5 mg/m<sup>3</sup>  
30 (0.43 ppm), a limited number of observations were examined around the threshold of interest, which  
31 subsequently contributed to the inability to draw conclusions regarding the potential presence of a  
32 threshold with any certainty.

## 5.7.2. Potentially Susceptible and Vulnerable Populations

1 Interindividual variation in human responses to air pollutants indicates that not all individuals  
2 exposed to pollutants respond similarly. That is, some subpopulations are at increased risk to the  
3 detrimental effects of pollutant exposure. The NAAQS are intended to provide an adequate margin of  
4 safety for both general populations and susceptible and vulnerable subpopulations, or those subgroups  
5 potentially at increased risk for ambient air pollution health effects. For the purposes of this document, a  
6 susceptible subpopulation is defined as one with an intrinsic biological characteristic (e.g., disease) that  
7 might result in an adverse health effect at pollutant concentrations below those needed to elicit the same  
8 response in the general population, or result in a more adverse effect at the same concentration. A  
9 vulnerable subpopulation is one with a non-biological characteristic (e.g., differential exposure, living at  
10 altitude) that results in increased incidence of health effects from an ambient pollutant relative to the  
11 general population. The previous review of the CO NAAQS identified certain groups within the  
12 population that may be more susceptible to the effects of CO exposure, including individuals (particularly  
13 older adults) with CAD and other vascular diseases, anemia patients, patients with obstructive lung  
14 disease, and infants. Other subgroups considered to be somewhat vulnerable in the previous review  
15 include individuals living at altitude and those using certain medications that can increase endogenous CO  
16 production. Tables 5-18 and 5-19 provide an overview of the characteristics that contribute to susceptible  
17 and vulnerable subpopulations, respectively, which have been observed in the examination of the NAAQS  
18 for all criteria pollutants. Those characteristics of susceptible/vulnerable subpopulations exposed  
19 specifically to CO, as mentioned in the literature that encompasses this ISA, are discussed below.

---

**Table 5-18. Characteristics of susceptible subpopulations.**

---

**Susceptibility Characteristics<sup>1</sup>**

---

Pre-existing disease: Cardiovascular diseases, Anemia, Obesity, Diabetes, Respiratory diseases (e.g., asthma, obstructive lung disease)

---

Age: Children, Older Adults (65+)

---

Infants: Premature, Low Birth Weight

---

Gender

---

Pregnancy

---

Birth Defects

---

Race/Ethnicity

---

Genetic Factors

---

Nutritional status

---

<sup>1</sup> Susceptible (i.e., intrinsic) refers to biological characteristics of an individual, which can include life stage, genetics, and pre-existing disease.



---

**Table 5-19. Characteristics of vulnerable subpopulations.**

---

**Vulnerability Characteristics<sup>1</sup>**

---

Altitude

---

Level of Exercise

---

Proximity to Roadways

---

Medication Use

---

Geographic Location

---

Work Environment (e.g., outdoor workers)

---

Air Conditioning Use

---

SES

---

Education Level

---

Nutritional status

---

<sup>1</sup> Vulnerable (i.e., extrinsic) refers to non-biological variables associated with an individual that can result in a health effect.

### 5.7.2.1. Susceptibility Characteristics

#### Cardiovascular Disease

1           Individuals with heart disease may be at a greater risk from CO exposure since they may already  
2 have compromised O<sub>2</sub> delivery. CO is notable among air pollutants because it is especially harmful in  
3 individuals with impaired cardiovascular systems. Persons with a normal cardiovascular system can  
4 tolerate substantial concentrations of CO, if they vasodilate in response to the hypoxemia produced by  
5 CO. In contrast, individuals unable to vasodilate in response to CO exposure may show evidence of  
6 ischemia at low concentrations of COHb. Many of the controlled human exposure studies have focused  
7 on individuals with IHD.

8           An estimated 81 million American adults (1 in 3) have one or more type of cardiovascular disease  
9 (CVD), with an estimated 47% of these being 60 or more years of age. CVD is the leading cause of death  
10 in the U.S. with nearly 2,400 deaths each day-an average of one death every 37 seconds (Rosamond et al.,  
11 2008). For the major diseases within the category of total CVD, about 73 million Americans have high  
12 BP, 16 million have CHD, 5 million have heart failure, 5 million have stroke, and the estimated  
13 prevalence of congenital cardiovascular defects is estimated to be between 650,000 to 1.3 million  
14 (Rosamond et al., 2008). In the U.S., IHD is the largest major killer, causing 1 in 5 deaths. Because the  
15 numbers of affected people are so high, even relatively small percent increases in cardiovascular mortality  
16 or morbidity in the population could have a large impact on public health.

17           Each year in the U.S. approximately 780,000 people experience a new or recurrent stroke, with the  
18 majority of these being a first stroke (77%). On average this equates to someone in the U.S. having a  
19 stroke every 40 seconds with a death occurring every 3 to 4 minutes. When considered separately from

1 other CVDs, stroke ranks third among all causes of death, behind diseases of the heart and cancer  
2 (Rosamond et al., 2008). While epidemiologic evidence is weak regarding associations between ambient  
3 CO and stroke, diseased individuals could exhibit increased sensitivity to CO exposure.

## **Obstructive Lung Disease**

4 Patients with obstructive lung disease, such as COPD, were identified as a susceptible  
5 subpopulation in the 2000 CO AQCD. COPD is a progressive disease resulting in decreased air flow to  
6 the lungs, and is especially prevalent among smokers. A majority of COPD patients have exercise  
7 limitations as demonstrated by a decrease in O<sub>2</sub> saturation during mild to moderate exercise. This makes  
8 individuals with hypoxia resulting from COPD particularly sensitive to CO during submaximal exercise  
9 typical of normal daily activity. COPD patients who are smokers may have baseline COHb levels of  
10 4-8%, leaving little reserve for increases in COHb due to ambient exposure. COPD is often accompanied  
11 by a number of changes in gas exchange, including increased V<sub>D</sub> and V/Q inequality (Marthan et al.,  
12 1985), which could slow both CO uptake and elimination. Patients with pulmonary sarcoidosis may have  
13 a decrease in lung volumes, a loss of D<sub>L</sub>CO, and gas exchange abnormalities during exercise, including  
14 decreased P<sub>a</sub>O<sub>2</sub> and increased alveolar-arterial oxygen pressure difference (Lamberto et al., 2004). A  
15 controlled human exposure study on 19 individuals with COPD (Bathoorn et al., 2007) found that two of  
16 the patients experienced COPD exacerbation during or following CO exposure at 100-125 ppm for 2 h,  
17 although a slight anti-inflammatory effect was also observed. The few epidemiologic studies that  
18 evaluated the relationship between ambient CO and increased hospital admissions or ED visits for COPD  
19 show weak positive associations. Epidemiologic results were similar for asthmatics, who can also  
20 experience exercise-induced airflow limitation.

## **Anemia**

21 Health status can influence the toxicity involved with CO exposure by influencing the severity of  
22 hypoxia resulting from CO exposure. Any condition that would alter the blood O<sub>2</sub> carrying capacity or  
23 content will result in a greater risk from COHb induced hypoxia and decreased tissue O<sub>2</sub> delivery. The  
24 severity of this effect depends upon the initial level of hypoxia. Anemias are a group of diseases that  
25 result in insufficient blood O<sub>2</sub> or hypoxia due to Hb deficiency through hemolysis, hemorrhage, or  
26 reduced hematopoiesis. Anemia may result from pathologic conditions characterized by chronic  
27 inflammation such as malignant tumors or chronic infections (Cavallin-Stahl et al., 1976a, b). The bodies  
28 of people with anemia compensate causing cardiac output to increase as both heart rate and stroke volume  
29 increase. The endogenous production of CO, thus COHb, is increased in patients with hemolytic anemia  
30 due to increased heme catabolism, causing an increased baseline COHb concentration. One of the most  
31 prevalent anemias arises from a single-point mutation of Hb, causing sickle cell diseases. The Hb affinity  
32 for O<sub>2</sub> and O<sub>2</sub> carrying capacity is reduced causing a shift to the right in the O<sub>2</sub> dissociation curve. It is

1 well documented that African-American populations have a higher incidence of sickle cell anemia, which  
2 may be a risk factor for CO hypoxia.

3 Any disturbance in RBC hemostasis by acceleration of destruction of hemoproteins will lead to  
4 increased production of CO. Pathologic conditions such as anemias, hematomas, thalassemia, Gilbert's  
5 syndrome with hemolysis, and other hematological diseases and illness will accelerate CO production  
6 (Berk et al., 1974; Hampson, 2007; Meyer et al., 1998; Solanki et al., 1988). Patients with hemolytic  
7 anemia exhibit COHb levels 2- to 3-times higher than healthy individuals and CO production rates 2- to  
8 8-times higher (Coburn et al., 1966).

### **Age and Gender**

9 Age and gender alter the variables that influence the uptake, distribution, and elimination of CO.  
10 COHb levels decline more rapidly in young children than adults after CO exposure (Joumard et al., 1981;  
11 Klasner et al., 1998). After infancy, the COHb half-life increases with age, practically doubling between 2  
12 and 70 years (Joumard et al., 1981). The rate of this decrease in CO elimination is very rapid in the  
13 growing years (2 to 16 years of age), but slows beyond adolescence. Alveolar volume and  $D_LCO$   
14 increased with increasing body length of infants and toddlers (Castillo et al., 2006), suggesting a further  
15 degree of lung development and faster CO uptake. After infancy, increasing age decreases  $D_LCO$  and  
16 increases  $V_A/Q$  mismatch, causing it to take longer to both load and eliminate CO from the blood (Neas  
17 and Schwartz, 1996).

18 COHb concentrations are generally lower in female subjects than in male subjects (Horvath et al.,  
19 1988) and the COHb half-life is longer in healthy men than in women of the same age, which may be  
20 partially explained by differences in muscle mass or the slight correlation between COHb half-life and  
21 increased height (Joumard et al., 1981). The rate of decline of  $D_LCO$  with age is lower in middle-aged  
22 women than in men; however, it evens out towards older age (Neas and Schwartz, 1996). Women also  
23 tended to be more resistant to altitude hypoxia (Horvath et al., 1988).

### **Newborns and Young Infants**

24 Fetal CO pharmacokinetics do not follow the same kinetics as maternal CO exposure, making it  
25 difficult to estimate fetal COHb based on maternal levels. Human fetal Hb has a higher affinity for CO  
26 than adult Hb (Di Cera et al., 1989). Maternal and fetal COHb concentrations have been modeled as a  
27 function of time using a modified CFK equation (Hill et al., 1977). At steady-state conditions, the fetal  
28 COHb is up to 10% higher than the maternal COHb levels, for example, exposure to 30 ppm CO results  
29 in a maternal COHb of 5% and a fetal COHb of 5.5%. The fetal CO uptake lags behind the maternal for  
30 the first few hours but later may overtake the maternal values. Similarly, during washout, the fetal COHb  
31 levels are maintained for longer, with a half-life of around 7.5 hours versus the maternal half-life of  
32 around 4 hours. In addition, women experience fluctuating COHb levels during pregnancy as well as

1 through the menstrual cycle when endogenous CO production doubles in the progesterone phase  
2 (Delivoria-Papadopoulos et al., 1974; Mercke and Lundh, 1976).

3 Epidemiologic studies of birth outcomes have examined well-established clinical metrics of infant  
4 health. Preterm birth (PTB, <37 weeks gestation) and low birth weight (LBW, birth weight <2,500 g)  
5 have been established as strong predictors of infant mortality and morbidity (Barker et al., 2002;  
6 Berkowitz and Papiernik, 1993; Li et al., 2003; McIntire et al., 1999). In 2004, 36.5 percent of all infant  
7 deaths in the U.S. were preterm-related (MacDorman et al., 2007). Vital statistics for the year 2005 in the  
8 U.S. showed that the rate for PTB was 12.7%, which has risen 20% since 1990, and the rate for LBW was  
9 8.2%, which has risen 17% since 1990 (Martin et al.).

10 Congenital anomalies remain the leading cause of infant death in the U.S. (Martin et al.). In 2004,  
11 for every 1,000 live births in the U.S. the infant (up to 1 year of age) mortality rate was 6.8 deaths, while  
12 the neonatal (under 28 days of age) mortality rate was 4.5 deaths, and the postneonatal (between 28 days  
13 and 1 year of age) mortality rate was 2.4 deaths (National Center for Health Statistics, 2007). From 1968  
14 to 1995, the proportion of infant mortality attributable to birth defects increased from 14.5% to 22.2 %  
15 (Centers for Disease Control and Prevention, 1998). Limited epidemiologic research has been conducted  
16 on congenital anomalies because of their rare occurrence, which makes it difficult to detect small effects  
17 within defined exposure periods. For example, in 2005 the rates for spina bifida and anencephaly in the  
18 U.S. were 18.0 and 11.3 per 100,000 births respectively (Martin et al., 2007) and as a result of folic acid  
19 fortification the prevalence of these defects declined considerably from 1995 to 2002 (Williams et al.,  
20 2005). Survival rate among infants with spina bifida had also improved due to folic acid fortification (Bol  
21 et al., 2006). The rate of cleft lip/palate is higher at approximately 79.1 per 100,000 births (Martin et al.,  
22 2007).

23 The rate of cardiovascular defects is much higher. Data from the Metropolitan Atlanta Congenital  
24 Defects Program (MACDP), which is one of the most comprehensive birth defect registries in the U.S.,  
25 showed that the prevalence of congenital heart defects had increased between 1968 and 1997. During  
26 1995-1997 the rate was 90.2 per 10,000 births (0.9%) and this had increased from 58.7 per 10,000 births  
27 since 1986-1972 (Botto et al., 2001). Cardiovascular defects are the single largest contributor to infant  
28 mortality attributable to birth defects (Centers for Disease Control and Prevention, 1998). Between 1979  
29 and 1997, 1 in 10 infant deaths (9.8%) was associated with a congenital heart defect, and 1 in 13 infant  
30 deaths (7.4%) was due to a congenital heart defect (Boneva et al., 2001).

### 5.7.2.2. Vulnerability Characteristics

#### Altitude

31 Increased altitude changes a number of factors that contribute to the uptake and elimination of CO.  
32 The relationship between altitude and CO exposure has been discussed in depth in the 2000 CO AQCD

1 and other documents (U.S. EPA, 1978). In an effort to maintain proper O<sub>2</sub> transport and supply,  
2 physiological changes occur as compensatory mechanisms to combat the decreased barometric pressure  
3 and resulting altitude induced HH (HH). HH, unlike CO hypoxia, causes humans to hyperventilate, which  
4 reduces arterial blood CO<sub>2</sub> (hypocapnia) and increases alveolar partial pressure of O<sub>2</sub>. Hypocapnia will  
5 lead to difficulty of O<sub>2</sub> dissociation and decreased blood flow, thus reducing tissue O<sub>2</sub> supply. HH  
6 increases BP and cardiac output and leads to redistribution of blood from skin to organs and from blood to  
7 extravascular compartments. Generally these changes will favor increased CO uptake and COHb  
8 formation, as well as CO elimination. In hypoxic conditions both CO and O<sub>2</sub> bind reduced Hb through a  
9 competitive-parallel reaction (Chakraborty et al., 2004). Breathing CO (9 ppm) at rest at altitude produced  
10 higher COHb compared to sea level (McGrath et al., 1993), whereas high altitude exposure with exercise  
11 caused a decrease in COHb levels versus similar exposure at sea level (Horvath et al., 1988). This  
12 decrease could be a shift in CO storage or suppression of COHb formation, or both. Altitude also  
13 increases the baseline COHb levels by inducing endogenous CO production. Early HH increased lung  
14 HO-1 protein and activity, whereas chronic HH induced endogenous CO production in nonpulmonary  
15 sites (see Section 4.5) (Carraway et al., 2000).

16 As the length of stay increases at high altitude, acclimatization occurs, inducing hyperventilation,  
17 polycythemia or increased red blood cell count, and increased tissue capillarity and Mb content in skeletal  
18 muscle, which could also favor increased CO uptake. Most of the initial adaptive changes gradually revert  
19 to sea level values. However, differences in people raised at high altitude persist even after  
20 reacclimatization to sea level (Hsia, 2002).

21 Altitude has been shown to be positively associated with baseline COHb concentrations (McGrath,  
22 1992; McGrath et al., 1993). This increase in COHb with altitude induced hypoxia has also been  
23 associated with increases in the mRNA, protein, and activity of HO-1 in rats and cells leading to  
24 enhanced endogenous CO production (Carraway et al., 2002; Chin et al., 2007). Whether other variables  
25 (such as an accelerated metabolism or a greater pool of Hb, transient shifts in body stores, or a change in  
26 the elimination rate of CO) play a role has not been explored.

## **Activity Patterns**

27 Exercise is an important determinant of CO kinetics and toxicity due to the extensive increase in  
28 gas exchange. O<sub>2</sub> consumption can increase more than 10 fold during exercise. Similarly, ventilation,  
29 membrane and lung diffusing capacity, pulmonary capillary blood volume, and cardiac output increase  
30 proportional to work load. The majority of these changes facilitate CO uptake and transport, by increasing  
31 gas exchange efficiency.

32 The COHb elimination rate decreases with physical activity (Joumard et al., 1981). Healthy  
33 subjects exposed to CO and achieving COHb levels of approximately 4-5% observed a significant  
34 detriment to exercise duration and maximal effort capability (measured by metabolic equivalent units)

1 (Adir et al., 1999). It is possible that CO lowers the anaerobic threshold, allowing earlier fatigue of the  
2 skeletal muscles and decreased maximal effort capability.

### **Proximity to Roadways**

3 Individuals that spend a substantial amount of time on or near heavily traveled roadways, such as  
4 commuters and those living or working near freeways, are likely to experience elevated CO  
5 concentrations, as discussed in Chapter 3, and therefore constitute a potentially vulnerable subpopulation  
6 due to differential exposure. Studies of commuters have shown that commuting time is an important  
7 predictor of CO exposure for those traveling by car, cycling, and walking, and that on-road CO  
8 concentrations are typically two to five times higher than concentrations measured at the roadside.

### **Medication Use**

9 Not all endogenous CO production is derived from Hb breakdown. Other hemoproteins, such as  
10 Mb, cytochromes, peroxidases, and catalase, contribute 20-25% to the total amount of endogenous CO  
11 (Berk et al., 1976). All of these sources result in a blood COHb concentration between 0.4 and 0.7%  
12 (Coburn et al., 1965). This baseline level of endogenous production can be altered by drugs or a number  
13 of physiological conditions that alter RBC destruction, other hemoprotein breakdown, or bilirubin  
14 production. Nicotinic acid (Lundh et al., 1975), allyl-containing compounds (acetamids and barbiturates)  
15 (Mercke et al., 1975b), diphenylhydantoin (Coburn, 1970b), progesterone (Delivoria-Papadopoulos et al.,  
16 1974), and contraceptives (Mercke et al., 1975a) will increase CO production. Compounds such as carbon  
17 disulfide and sulfur-containing chemicals (parathion and phenyltiourea) will increase CO by acting on  
18 P450 system moieties (Landaw et al., 1970). The P450 system may also cause large increases in CO  
19 produced from the metabolic degradation of dihalomethanes leading to very high (>10%) COHb levels  
20 (Manno et al., 1992), which can be further enhanced by prior exposure to HCs or ethanol (Pankow et al.,  
21 1991; Wirkner and Poelchen, 1996). HO can catalyze the release of CO from the auto-oxidation of  
22 phenols, photo-oxidation of organic compounds, and lipid peroxidation of cell membrane lipids (Rodgers  
23 et al., 1994).

# Annex A. Atmospheric Science

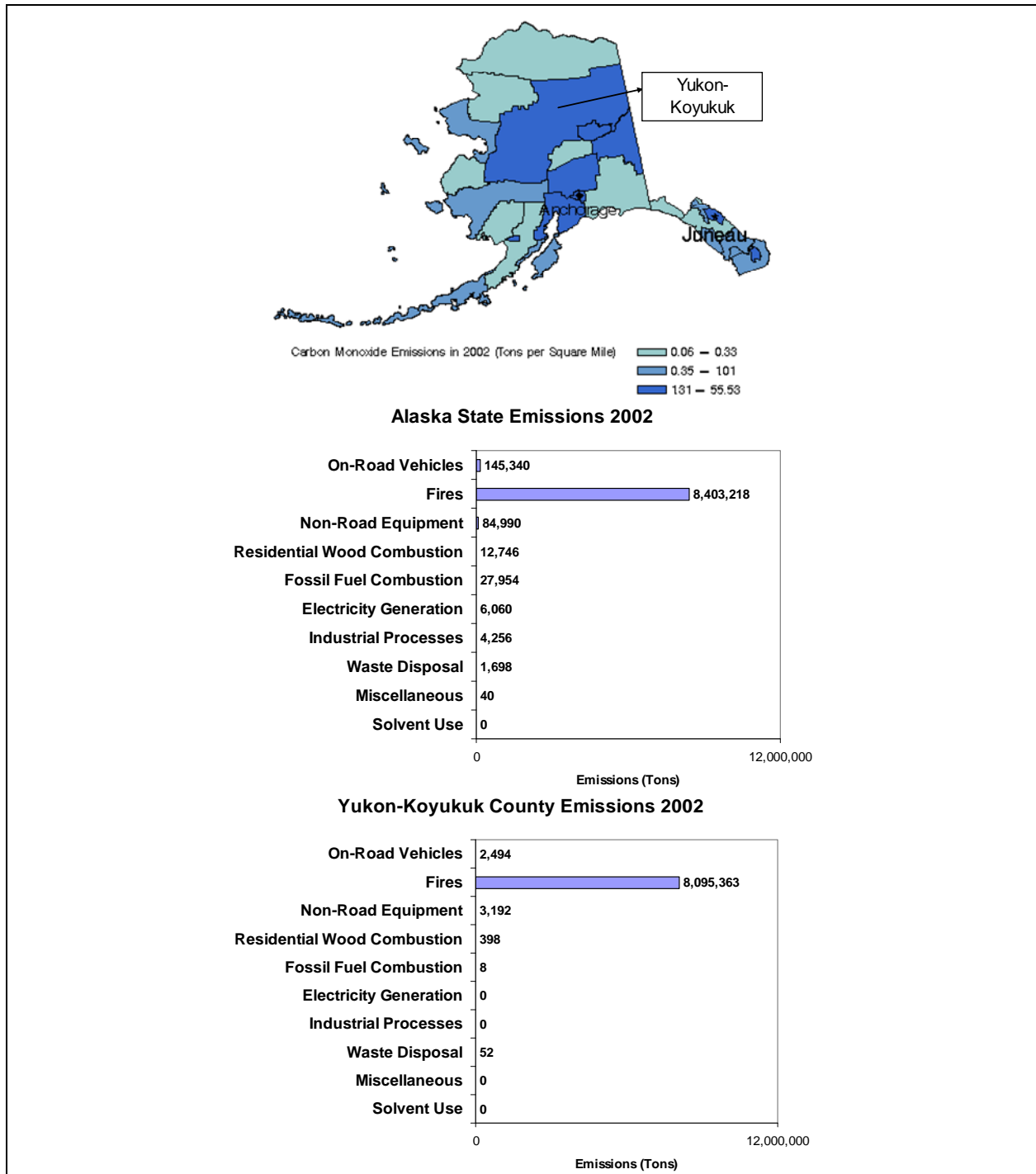
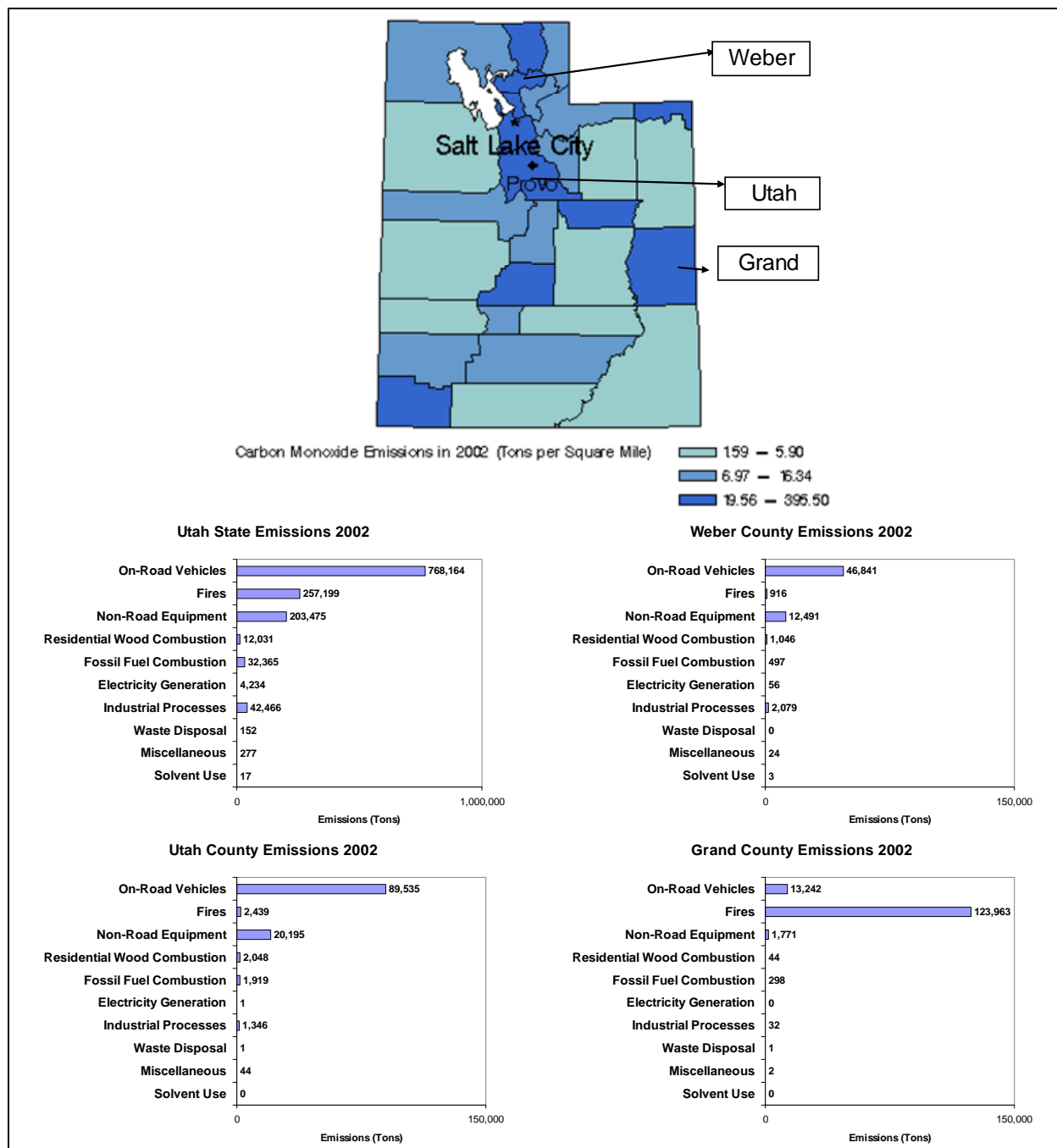
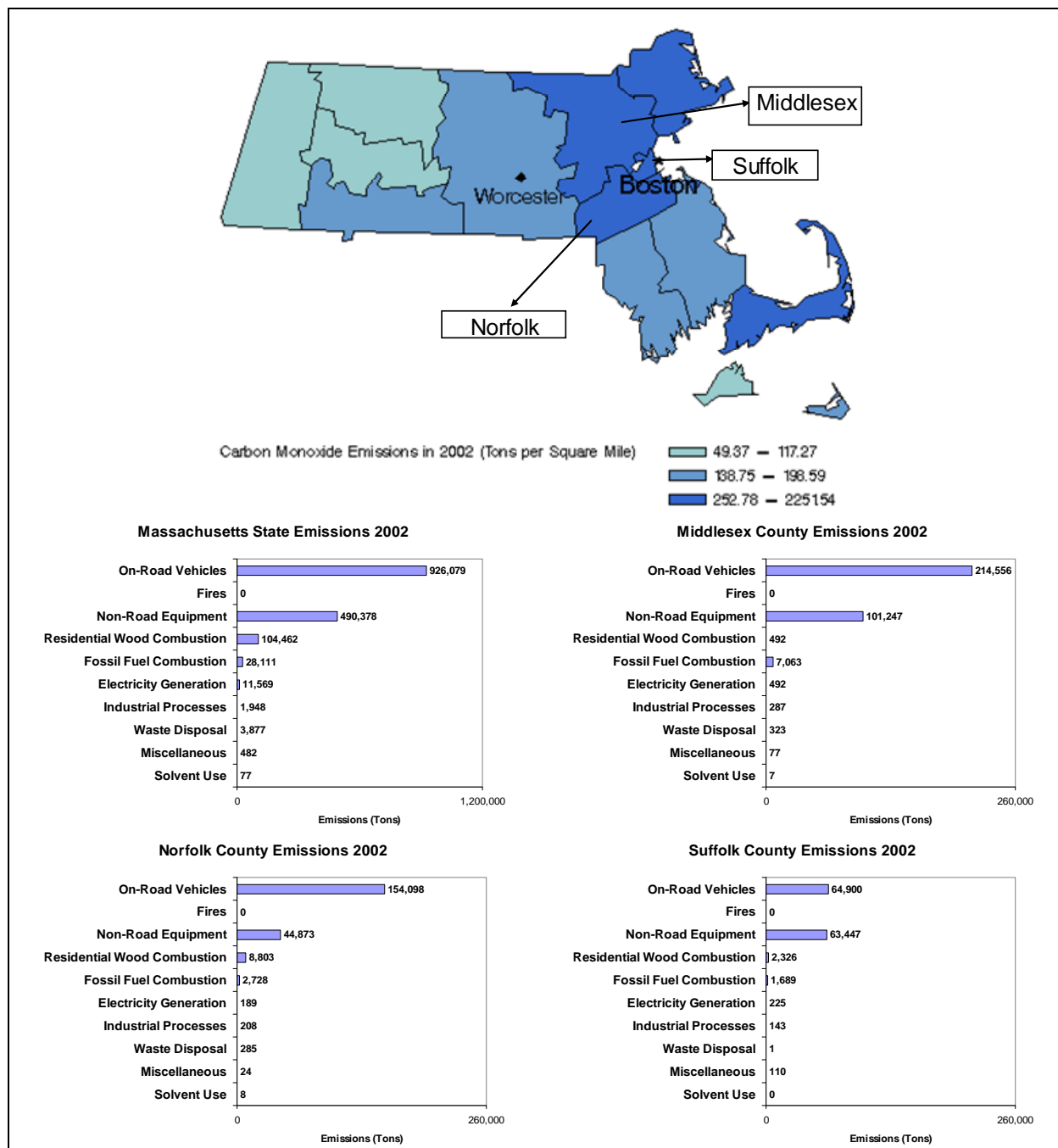


Figure A-1. CO emissions density map and distribution for the state of Alaska and for Yukon-Koyukuk County in Alaska.



**Figure A-2. CO emissions density map and distribution for the state of Utah and for selected counties in Utah.**





**Figure A-3. CO emissions density map and distribution for the state of Massachusetts and for selected counties in Massachusetts.**

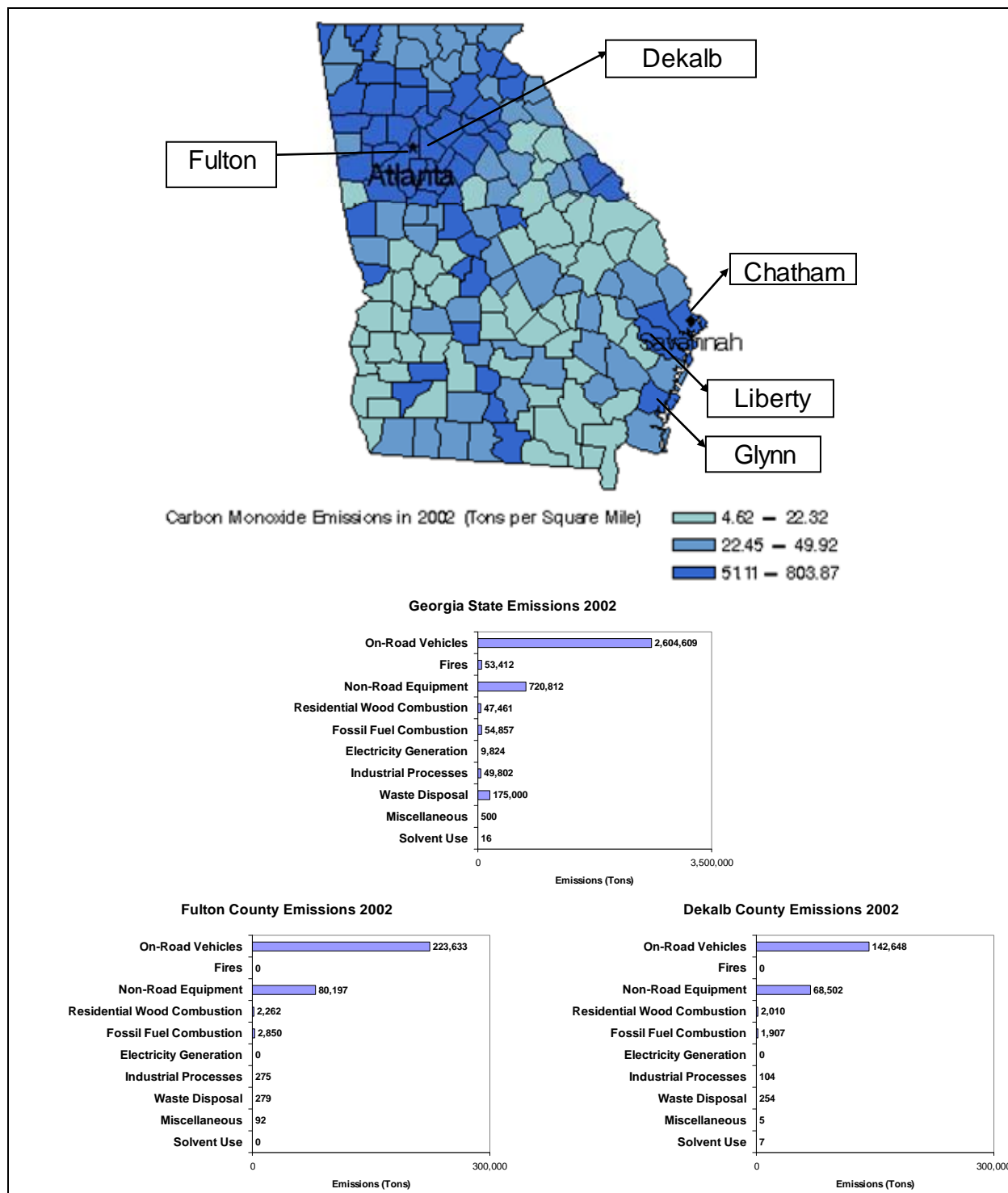


Figure A-4a. CO emissions density map and distribution for the state of Georgia and for selected counties in Georgia (1 of 2).

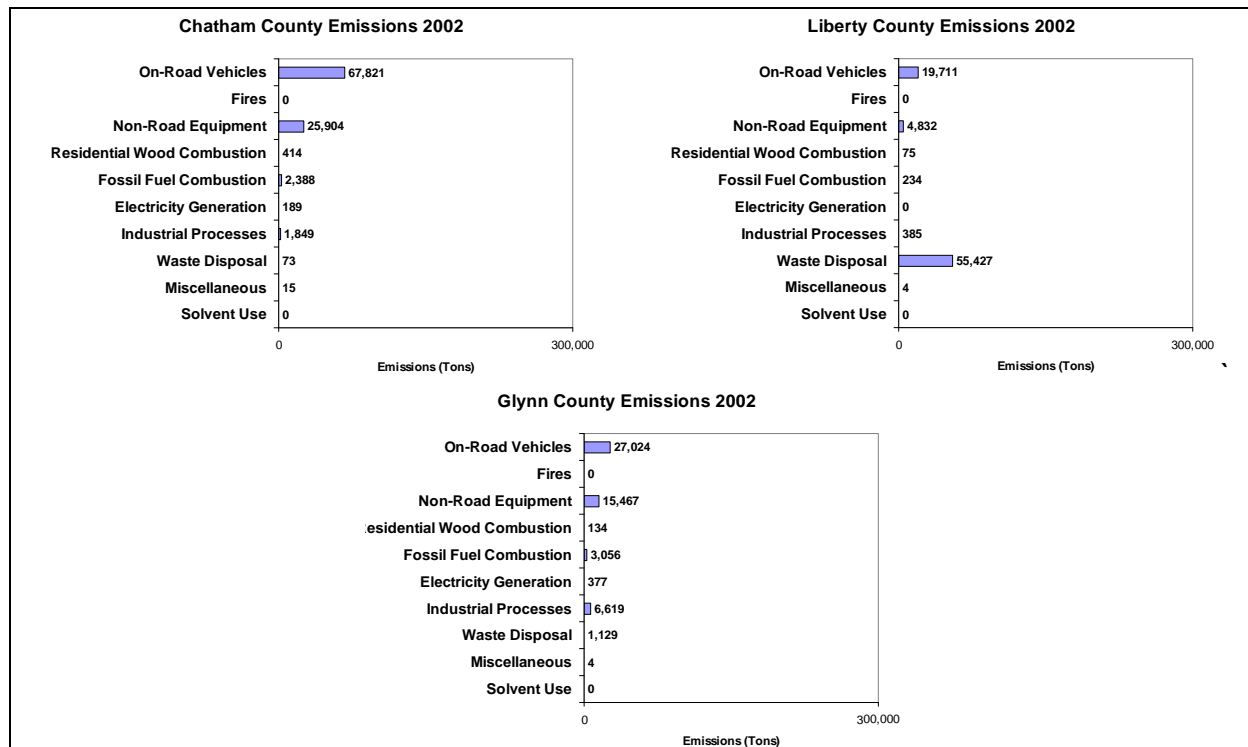
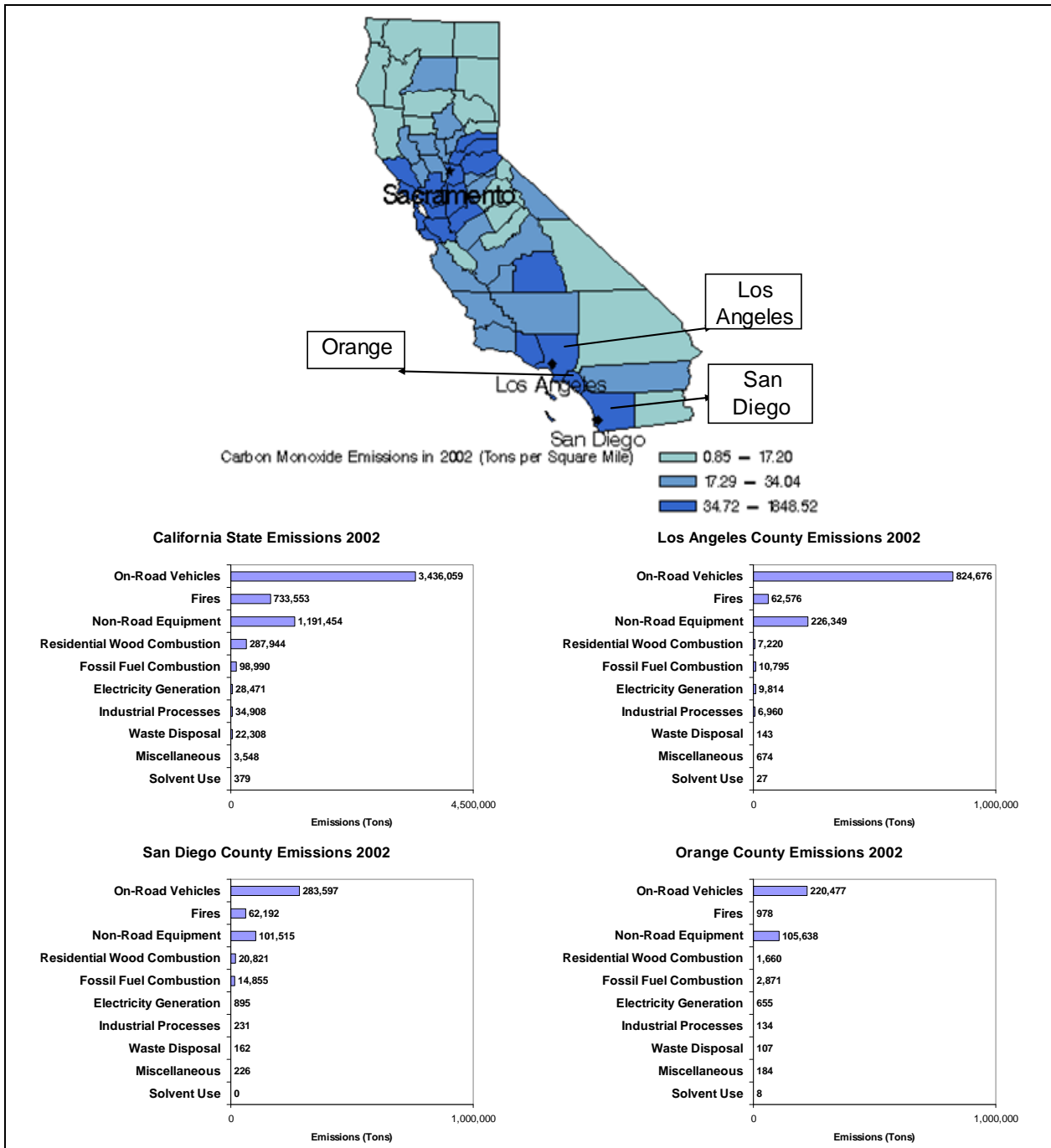
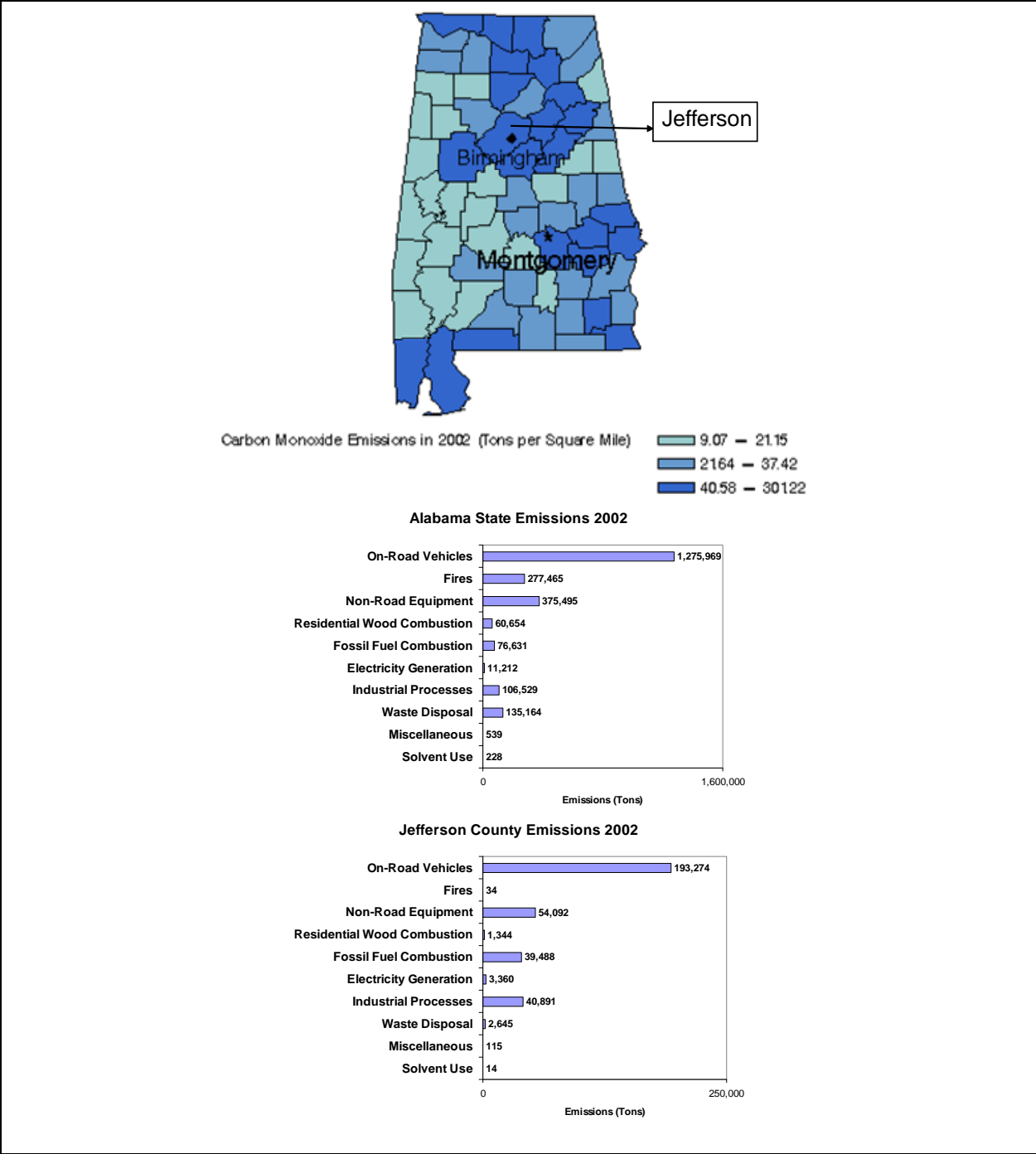


Figure A-4b. CO emissions distribution for selected counties in Georgia (2 of 2).



**Figure A-5. CO emissions density map and distribution for the state of California and for selected counties in California.**



**Figure A-6. CO emissions density map and distribution for the state of Alabama and for Jefferson County in Alabama.**

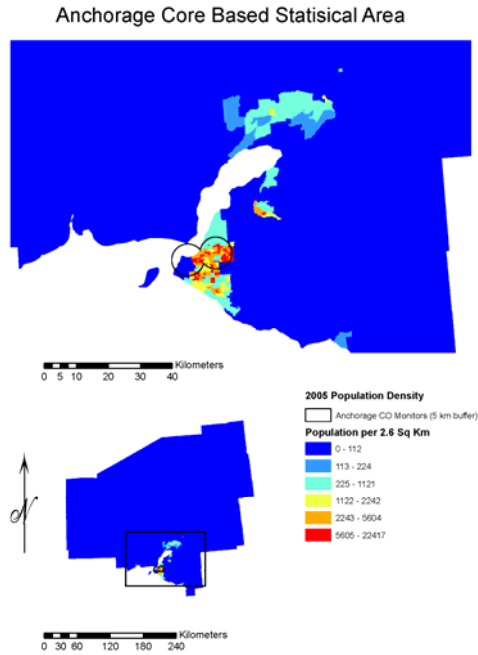


Figure A-7. Map of CO monitor locations with respect to population density in the Anchorage CBSA, total population.

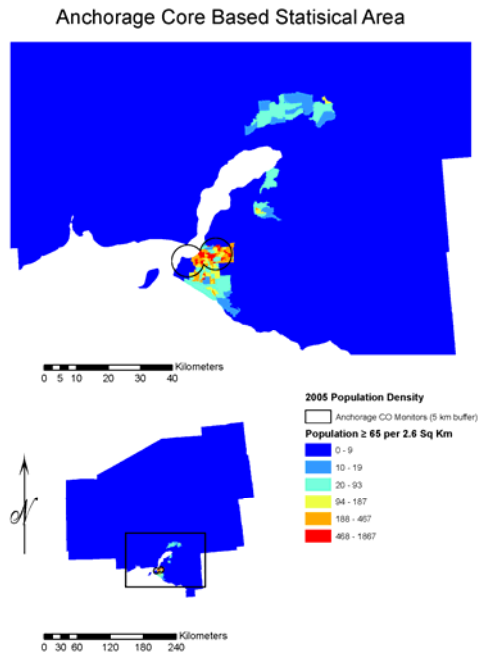


Figure A-8. Map of CO monitor locations with respect to population density in the Anchorage CBSA, ages 65 and older.

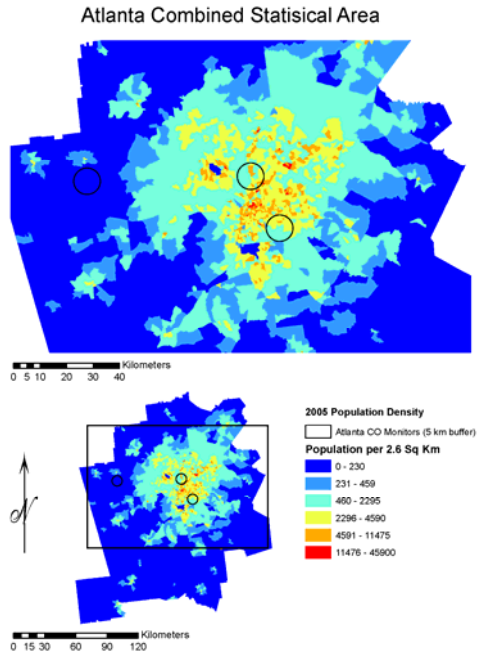


Figure A-9. Map of CO monitor locations with respect to population density in the Atlanta CSA, total population.

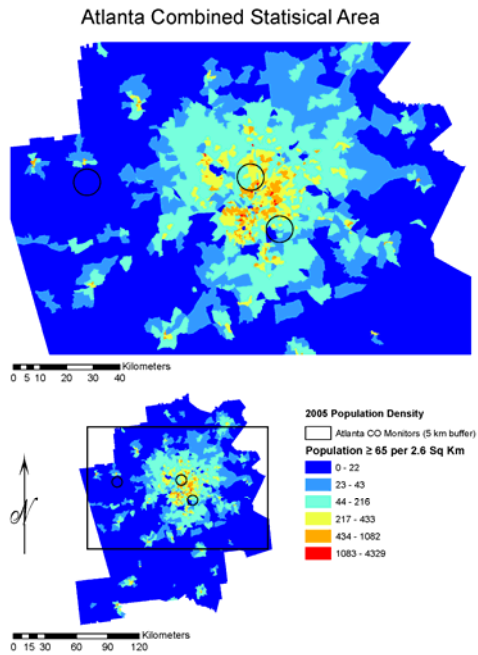


Figure A-10. Map of CO monitor locations with respect to population density in the Atlanta CSA, ages 65 and older.

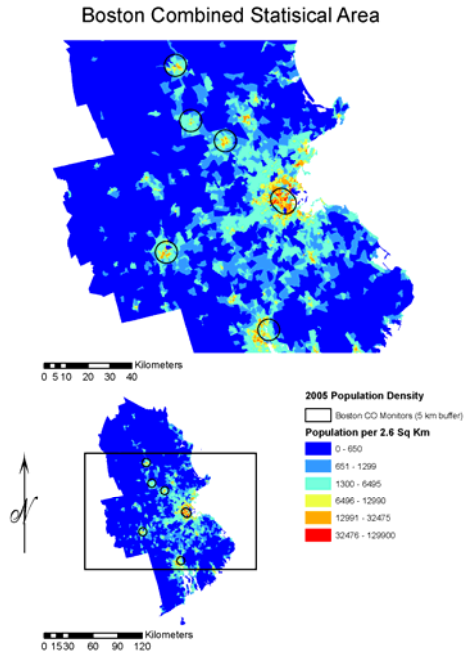


Figure A-11. Map of CO monitor locations with respect to population density in the Boston CSA, total population.

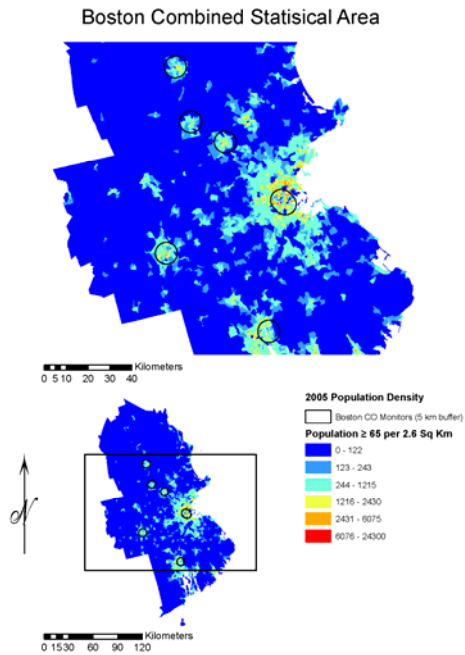


Figure A-12. Map of CO monitor locations with respect to population density in the Boston CSA, ages 65 and older.



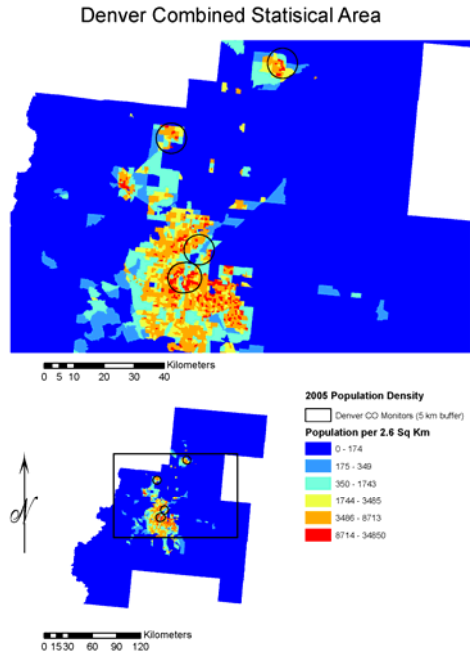


Figure A-13. Map of CO monitor locations with respect to population density in the Denver CSA, total population.

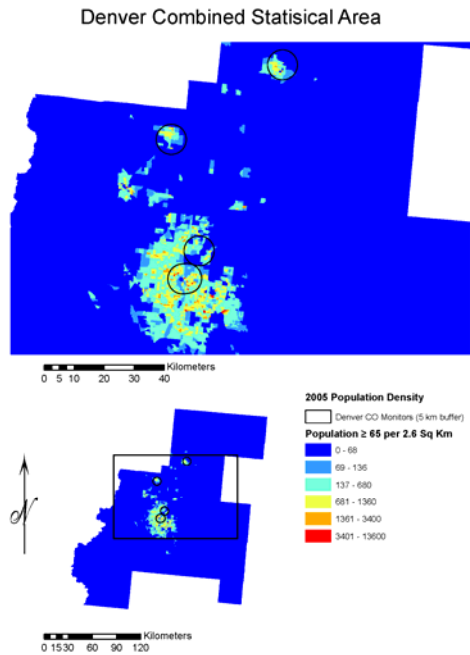


Figure A-14. Map of CO monitor locations with respect to population density in the Denver CSA, ages 65 and older.

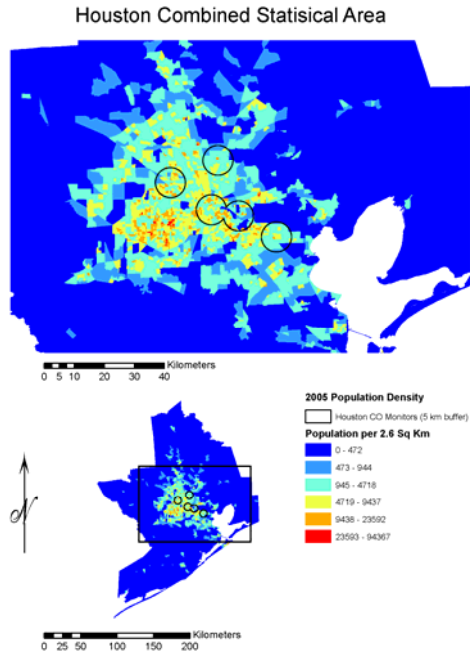


Figure A-15. Map of CO monitor locations with respect to population density in the Houston CSA, total population.

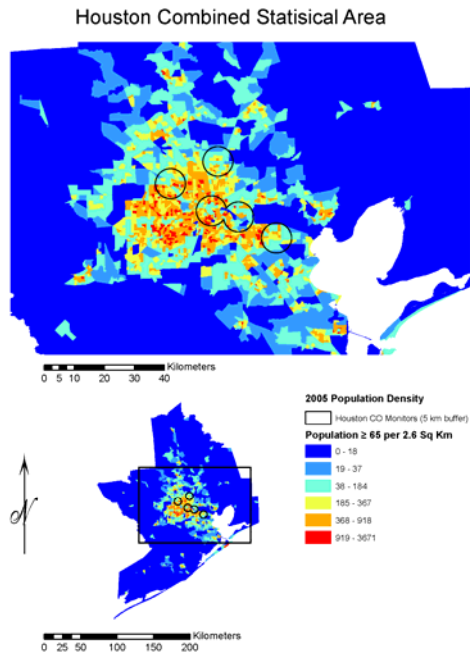


Figure A-16. Map of CO monitor locations with respect to population density in the Houston CSA, ages 65 and older.

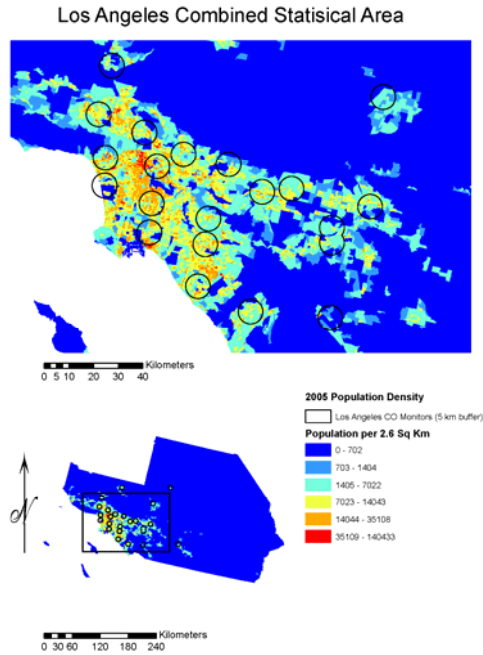


Figure A-17. Map of CO monitor locations with respect to population density in the Los Angeles CSA, total population.

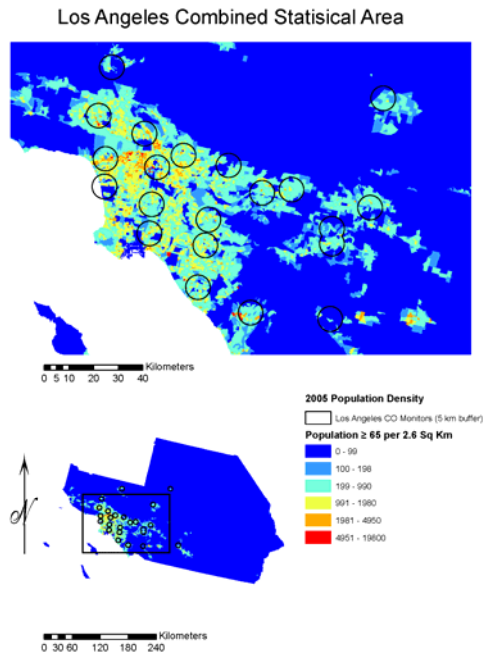


Figure A-18. Map of CO monitor locations with respect to population density in the Los Angeles CSA, ages 65 and older.

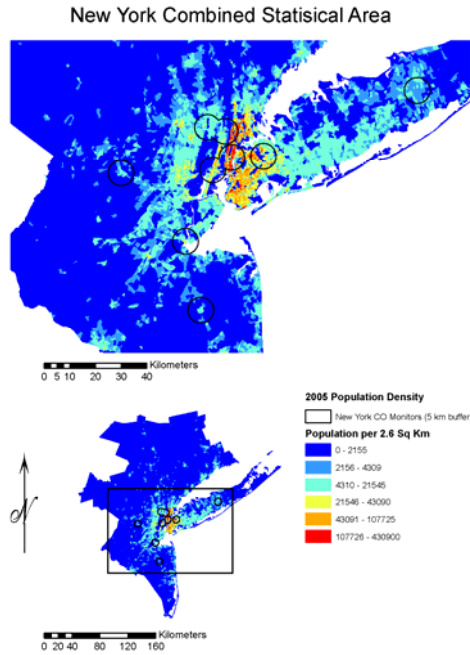


Figure A-19. Map of CO monitor locations with respect to population density in the New York City CSA, total population.

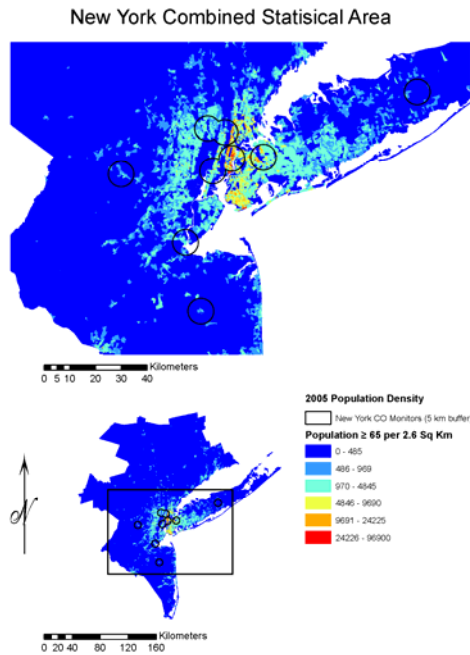


Figure A-20. Map of CO monitor locations with respect to population density in the New York City CSA, ages 65 and older.

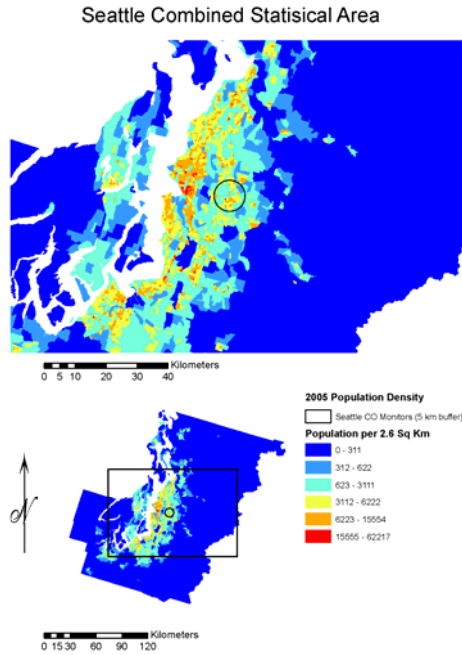


Figure A-21. Map of CO monitor locations with respect to population density in the Seattle CSA, total population.

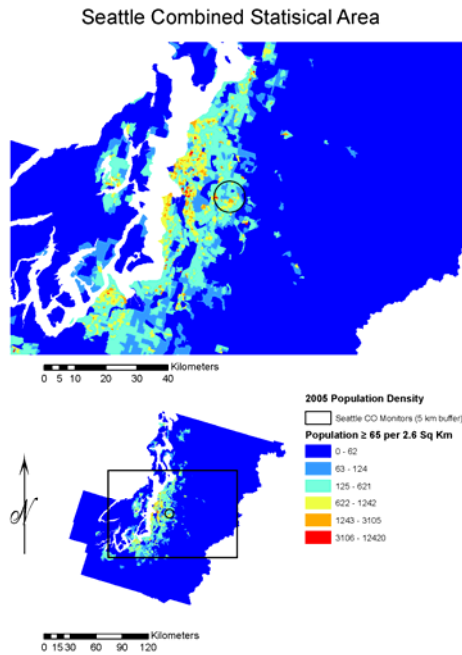


Figure A-22. Map of CO monitor locations with respect to population density in the Seattle CSA, ages 65 and older.

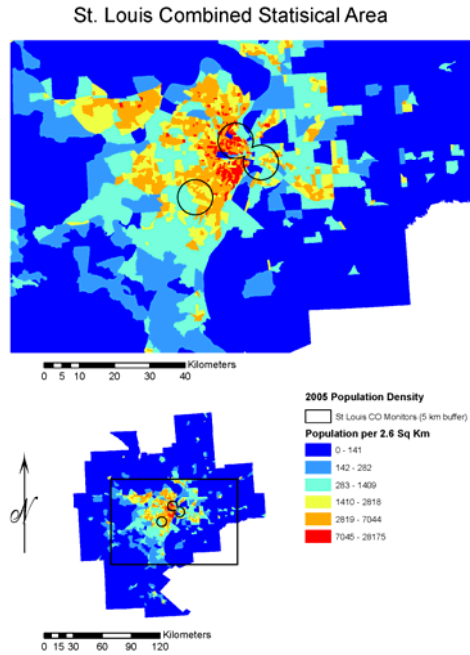


Figure A-23. Map of CO monitor locations with respect to population density in the St. Louis CSA, total population.

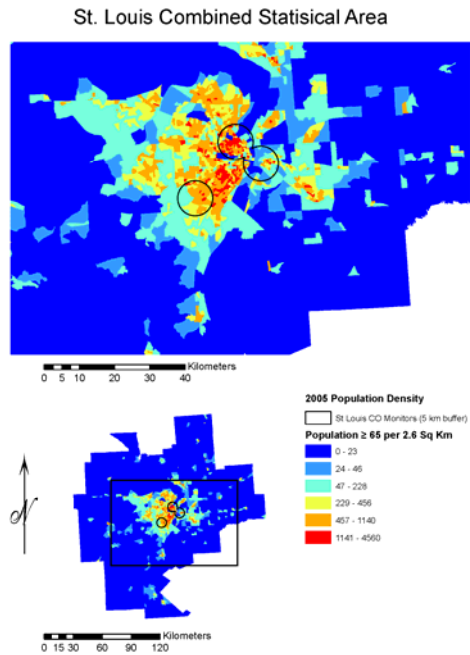


Figure A-24. Map of CO monitor locations with respect to population density in the St. Louis CSA, ages 65 and older.

Anchorage Core Based Statistical Area

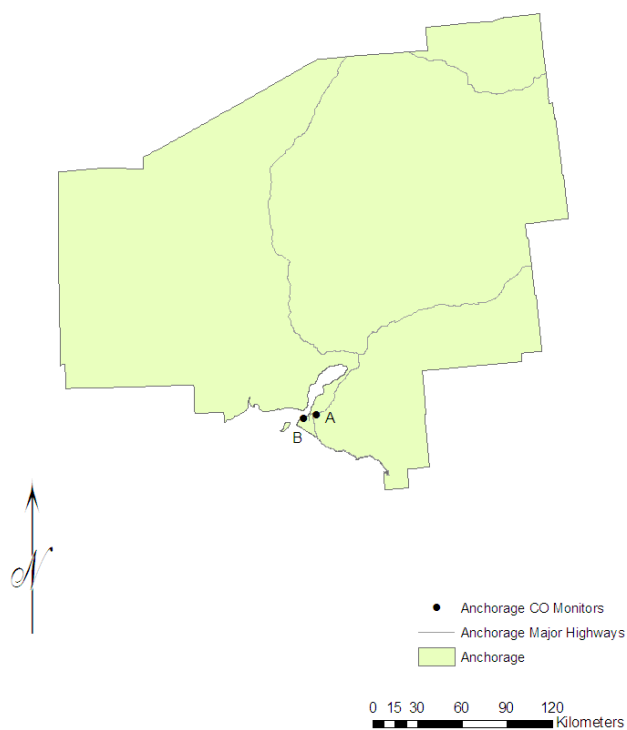
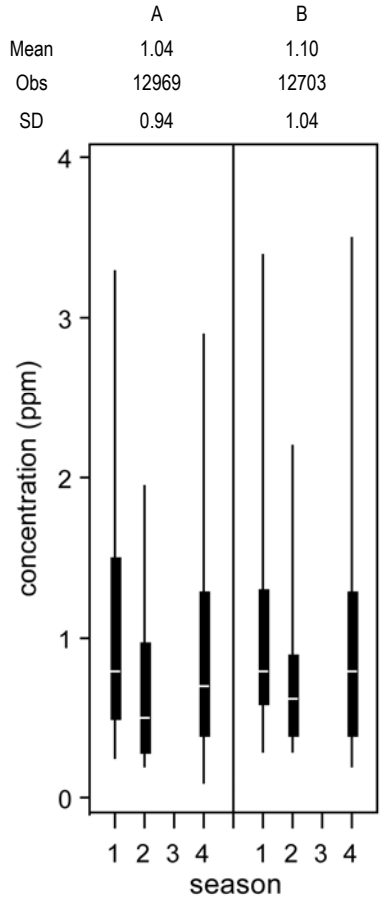


Figure A-25. Map of CO monitor locations with AQS Site IDs for Anchorage, AK.

Table A-1. Table of inter-sampler comparison statistics, including Pearson  $r$ , P90 (ppm), COD, and  $d$  (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in Anchorage, AK.

		Anchorage	
		A	B
A		1.00	0.73
		0.0	1.1
		0.00	0.32
		0.0	9.0
B	Legend		
	$r$		1.00
	P90		0.0
	COD		0.00
	$d$		0.0



**Figure A-26. Box plots illustrating the seasonal distribution of hourly CO concentrations in Anchorage, AK. Note: 1 = winter, 2, = spring, 3 = summer, and 4 = fall on the x-axis.**



Atlanta Combined Statistical Area

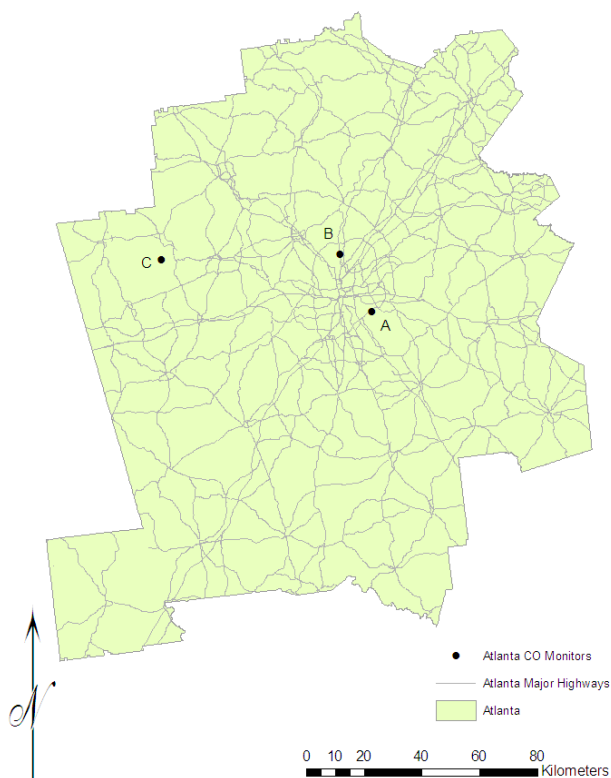
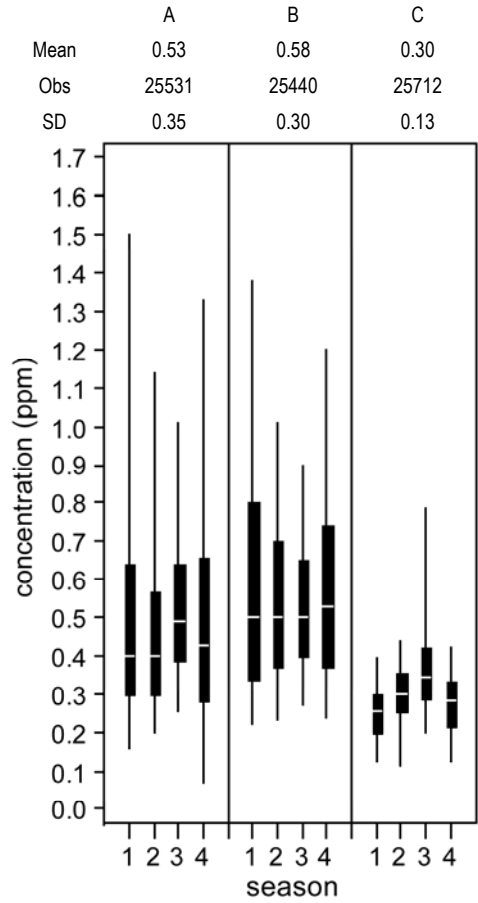


Figure A-27. Map of CO monitor locations with AQS Site IDs for Atlanta, GA.

Table A-2. Table of inter-sampler comparison statistics, including Pearson  $r$ , P90 (ppm), COD, and  $d$  (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in Atlanta, GA.

Atlanta			
	A	B	C
A	1.00	0.60	0.12
	0.0	0.5	0.7
	0.00	0.27	0.37
	0.0	22.5	74.7
B		1.00	0.10
		0.0	0.7
		0.00	0.38
C		0.0	61.7
			1.00
			0.0
			0.00
			0.0



**Figure A-28.** Box plots illustrating the seasonal distribution of hourly CO concentrations in Atlanta, GA. Note: 1 = winter, 2 = spring, 3 = summer, and 4 = fall on the x-axis.

Boston Combined Statistical Area

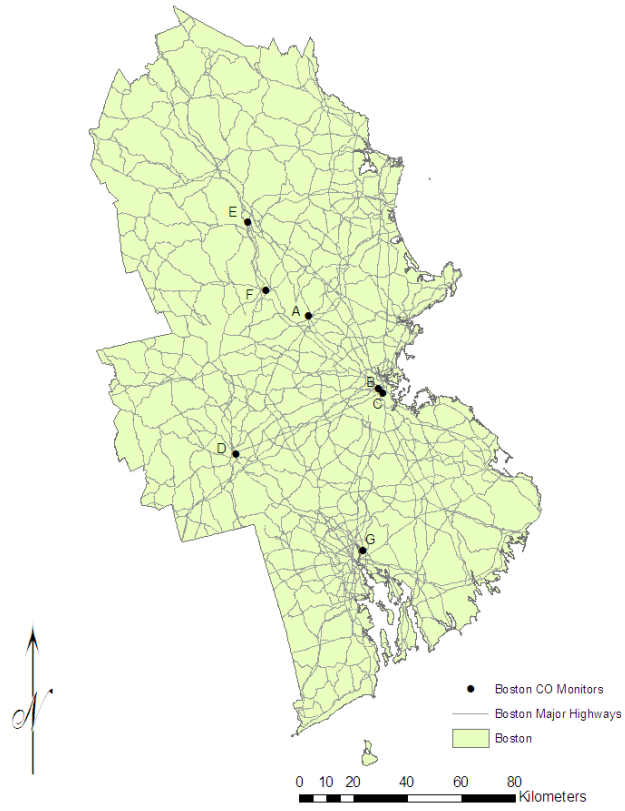


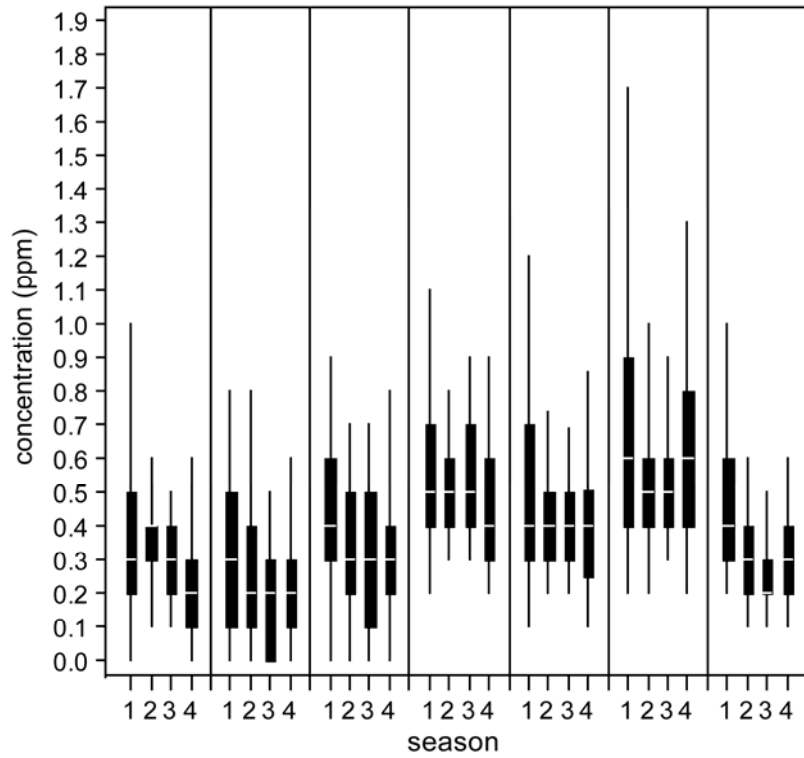
Figure A-29. Map of CO monitor locations with AQS Site IDs for Boston, MA.

**Table A-3. Table of inter-sampler comparison statistics, including Pearson r, P90 (ppm), COD, and d (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in Boston, MA.**

Boston							
	A	B	C	D	E	F	G
A	1.00	0.35	0.50	0.49	0.41	0.50	0.40
	0.0	0.4	0.4	0.5	0.4	0.6	0.4
	0.00	0.58	0.48	0.42	0.41	0.44	0.40
	0.0	37.2	39.7	57.9	41.3	18.3	89.1
B		1.00	0.52	0.34	0.27	0.35	0.34
		0.0	0.4	0.6	0.5	0.7	0.4
		0.00	0.56	0.59	0.58	0.60	0.55
		0.0	2.5	58.0	78.2	55.1	60.2
C			1.00	0.37	0.26	0.38	0.36
			0.0	0.5	0.5	0.6	0.4
			0.00	0.45	0.45	0.46	0.47
			0.0	58.9	80.7	57.5	58.7
D				1.00	0.40	0.46	0.34
				0.0	0.4	0.5	0.5
				0.00	0.28	0.25	0.39
				0.0	85.8	61.5	58.9
E					1.00	0.49	0.29
					0.0	0.5	0.4
					0.00	0.30	0.37
					0.0	26.1	128.6
F						1.00	0.43
						0.0	0.6
						0.00	0.39
						0.0	102.6
G							1.00
							0.0
							0.00
							0.0

Legend	
r	
P90	
COD	
d	

	A	B	C	D	E	F	G
Mean	0.33	0.26	0.36	0.53	0.45	0.60	0.34
Obs	24362	24134	24260	24446	25197	25869	23707
SD	0.22	0.24	0.26	0.23	0.27	0.37	0.22



**Figure A-30.** Box plots illustrating the seasonal distribution of hourly CO concentrations in Boston, MA. Note: 1 = winter, 2, = spring, 3 = summer, and 4 = fall on the x-axis.

Denver Combined Statistical Area

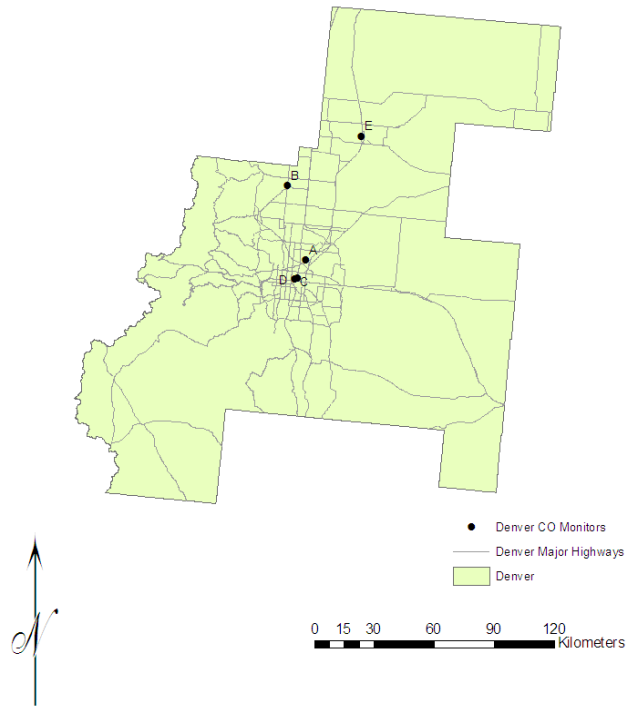


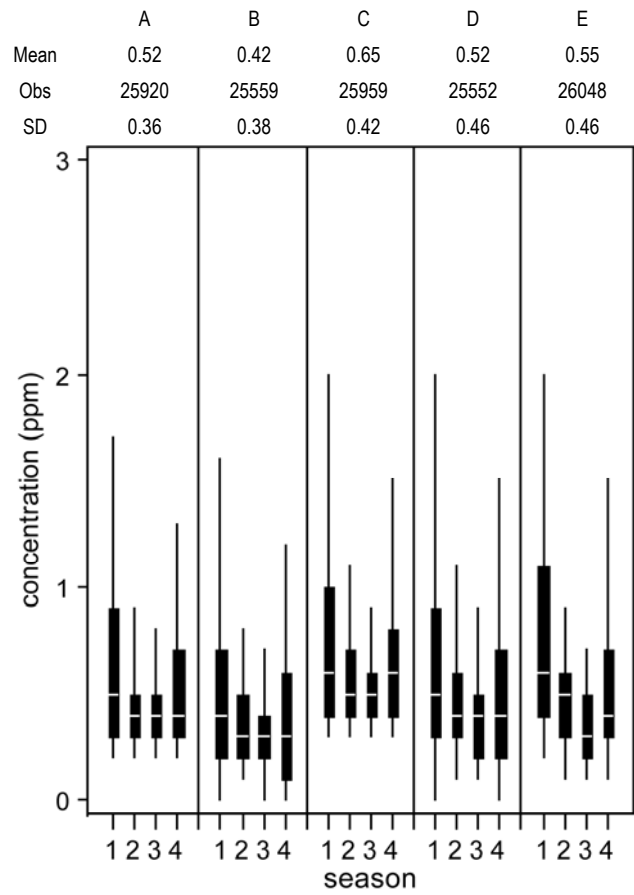
Figure A-31. Map of CO monitor locations with AQS Site IDs for Denver, CO.

**Table A-4. Table of inter-sampler comparison statistics, including Pearson r, P90 (ppm), COD, and d (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in Denver, CO.**

Denver					
	A	B	C	D	E
A	1.00	0.53	0.59	0.64	0.52
	0.0	0.6	0.6	0.5	0.6
	0.00	0.43	0.29	0.37	0.34
	0.0	38.5	10.1	10.9	68.2
B		1.00	0.46	0.49	0.54
		0.0	0.7	0.7	0.6
		0.00	0.44	0.47	0.43
		0.0	46.9	47.0	44.6
C			1.00	0.76	0.45
			0.0	0.5	0.7
			0.00	0.34	0.36
			0.0	1.3	78.3
D				1.00	0.46
				0.0	0.7
				0.00	0.42
				0.0	79.0
E					1.00
					0.0
					0.00
					0.0

Legend	
r	
P90	
COD	
d	



**Figure A-32.** Box plots illustrating the seasonal distribution of hourly CO concentrations in Denver, CO. Note: 1 = winter, 2 = spring, 3 = summer, and 4 = fall on the x-axis.



Houston Combined Statistical Area

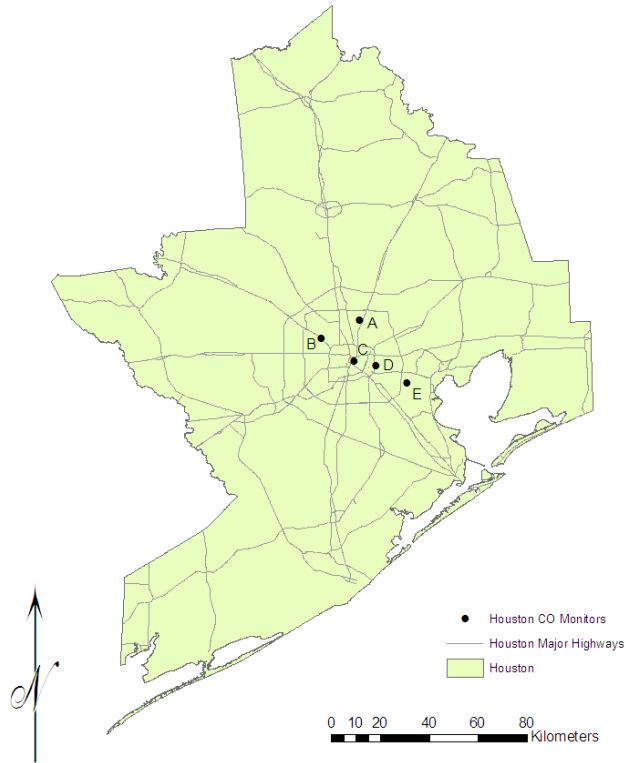


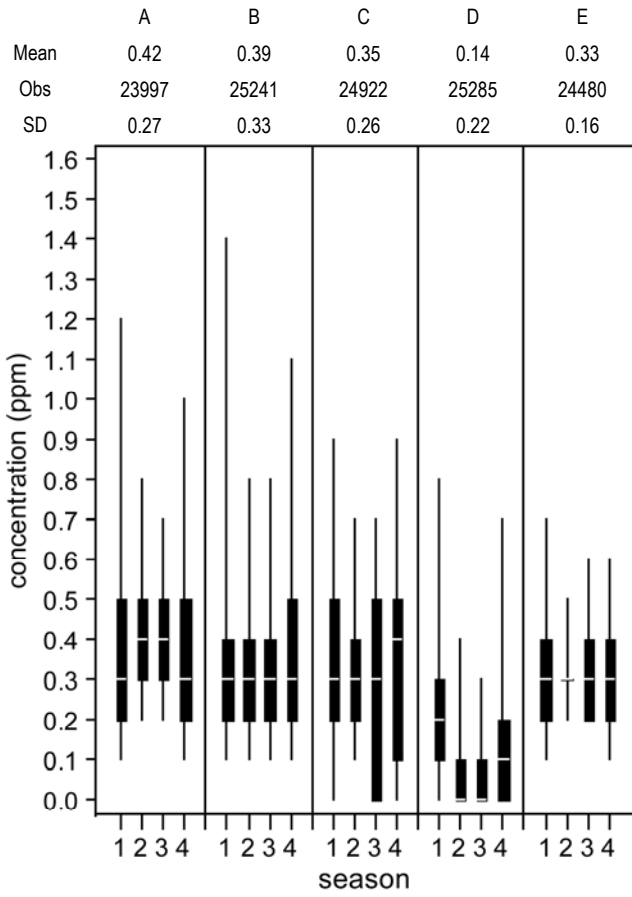
Figure A-33. Map of CO monitor locations with AQS Site IDs for Houston, TX.

**Table A-5. Table of inter-sampler comparison statistics, including Pearson r, P90 (ppm), COD, and d (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in Houston, TX.**

Houston					
	A	B	C	D	E
A	1.00	0.72	0.45	0.56	0.68
	0.0	0.3	0.4	0.5	0.3
	0.00	0.29	0.47	0.73	0.24
	0.0	17.5	16.7	19.8	32.2
B		1.00	0.56	0.65	0.63
		0.0	0.4	0.5	0.4
		0.00	0.47	0.73	0.29
		0.0	16.3	25.2	39.7
C			1.00	0.53	0.43
			0.0	0.5	0.4
			0.00	0.74	0.47
			0.0	9.3	23.5
D				1.00	0.57
				0.0	0.4
				0.00	0.72
				0.0	14.5
E					1.00
					0.0
					0.00
					0.0

Legend	
r	
P90	
COD	
d	



**Figure A-34.** Box plots illustrating the seasonal distribution of hourly CO concentrations in Houston, TX. Note: 1 = winter, 2 = spring, 3 = summer, and 4 = fall on the x-axis.

Los Angeles Combined Statistical Area

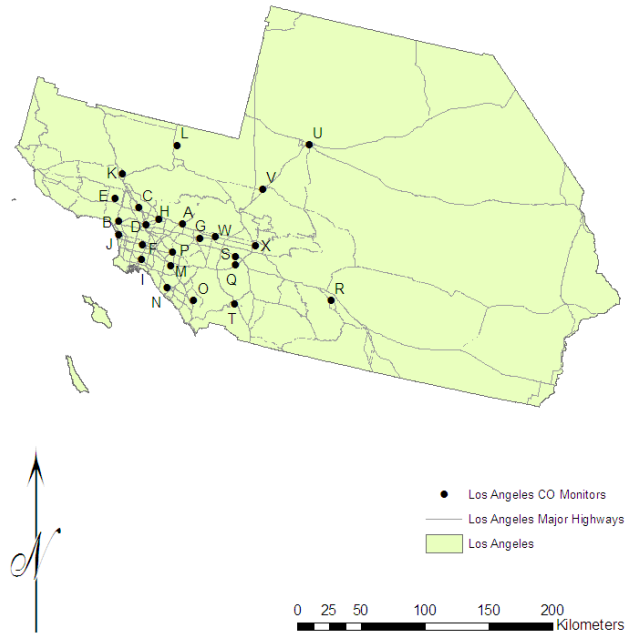
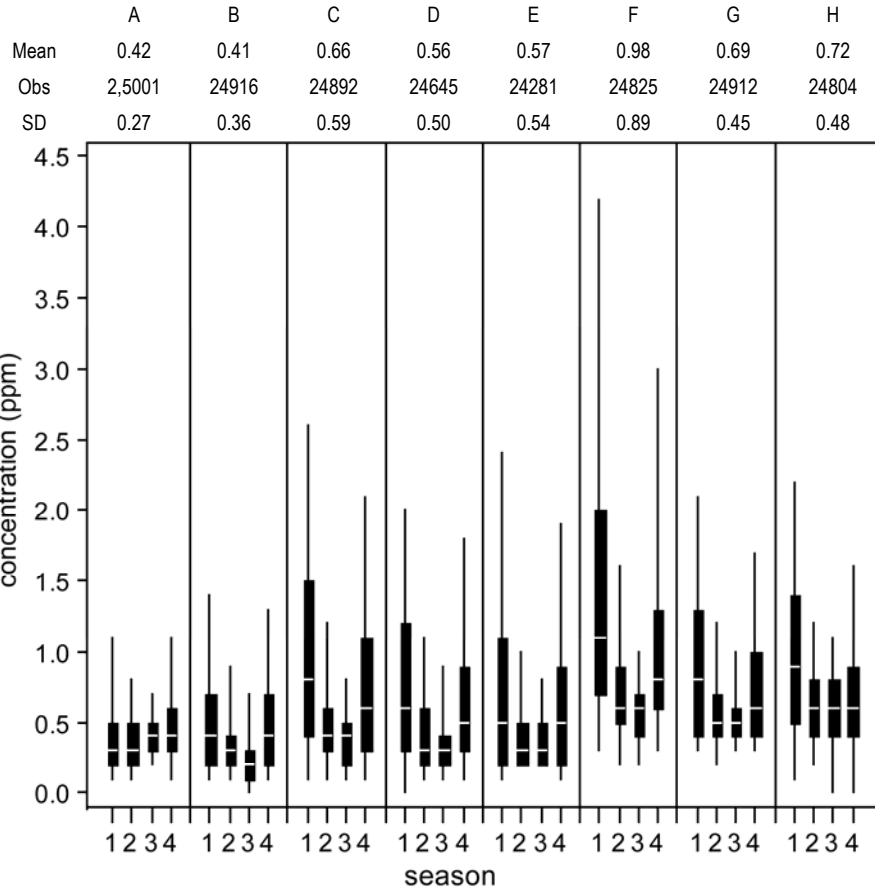


Figure A-35. Map of CO monitor locations with AQS Site IDs for Los Angeles, CA.

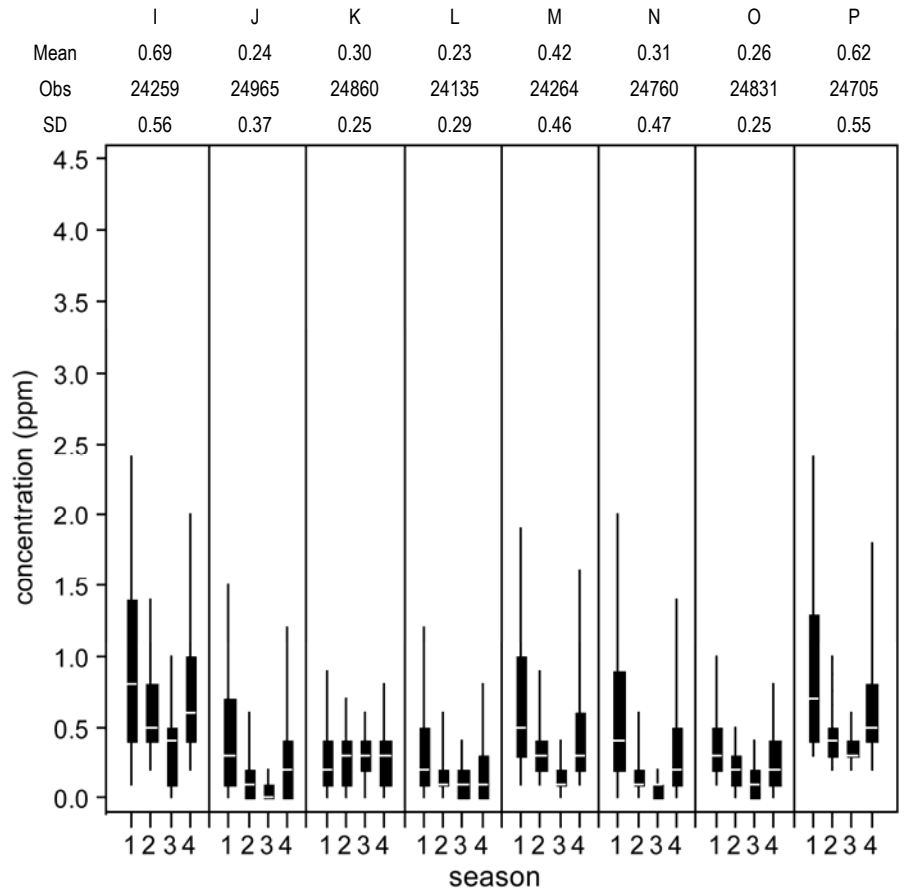
**Table A-6. Table of inter-sampler comparison statistics, including Pearson r, P90 (ppm), COD, and d (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in Los Angeles, CA.**

Los Angeles		A	B	C	D	E	F	G	H	I	J	K	L	M	N	O	P	Q	R	S	T	U	V	W	X
<b>A</b>	1.00	0.45	0.43	0.46	0.44	0.21	0.57	0.41	0.17	0.19	0.43	0.27	0.29	0.14	0.42	0.34	0.54	0.33	0.47	0.56	0.17	0.24	0.61	0.48	
	0.0	0.5	1.0	0.8	0.7	1.5	0.8	0.8	1.1	0.7	0.5	0.6	0.6	0.7	0.5	0.8	0.7	0.5	0.8	0.4	0.6	0.6	0.5	0.6	
	0.00	0.38	0.35	0.38	0.33	0.44	0.33	0.43	0.47	0.72	0.47	0.62	0.43	0.64	0.52	0.33	0.33	0.36	0.43	0.38	0.62	0.50	0.30	0.35	
	0	50.0	36.5	29.0	56.5	35.0	17.7	18.7	42.4	51.0	61.9	62.2	33.9	51.2	60.6	23.6	52.6	131.4	49.2	74.8	117.7	68.4	27.4	59.9	
<b>B</b>	1.00	0.69	0.71	0.64	0.50	0.63	0.49	0.43	0.59	0.36	0.42	0.59	0.47	0.57	0.59	0.57	0.35	0.55	0.55	0.25	0.41	0.50	0.53		
	0.0	0.8	0.6	0.7	1.3	0.7	0.8	0.9	0.6	0.6	0.7	0.5	0.7	0.6	0.6	0.7	0.6	0.7	0.5	0.7	0.6	0.6	0.6		
	0.00	0.38	0.38	0.34	0.50	0.41	0.49	0.48	0.69	0.50	0.60	0.39	0.59	0.50	0.37	0.42	0.38	0.46	0.39	0.62	0.50	0.40	0.41		
	0	18.9	21.2	17.8	26.4	65.0	31.6	35.2	11.4	37.4	74.9	53.7	64.4	85.9	48.5	98.1	178.1	96.1	112.0	161.1	115.3	76.4	109.1		
<b>C</b>	1.00	0.84	0.74	0.70	0.78	0.67	0.62	0.64	0.39	0.59	0.72	0.70	0.69	0.75	0.70	0.40	0.52	0.58	0.35	0.58	0.59	0.65			
	0.0	0.5	0.6	0.9	0.5	0.7	0.8	1.0	1.1	1.1	0.7	0.9	1.0	0.6	0.6	1.2	0.8	1.0	1.2	1.0	0.7	0.7			
	0.00	0.29	0.28	0.35	0.26	0.39	0.41	0.72	0.52	0.65	0.41	0.63	0.56	0.26	0.29	0.46	0.42	0.46	0.67	0.53	0.29	0.35			
	0	14.7	20.0	29.1	53.5	18.1	40.8	27.0	30.1	57.4	51.8	66.3	84.6	43.6	88.2	167.7	85.3	106.5	142.7	97.8	63.8	96.3			
<b>D</b>	1.00	0.64	0.74	0.77	0.67	0.60	0.59	0.34	0.53	0.70	0.58	0.63	0.74	0.62	0.32	0.49	0.53	0.34	0.51	0.57	0.58				
	0.0	0.7	1.0	0.5	0.7	0.8	0.8	0.9	0.9	0.6	0.8	0.8	0.5	0.6	1.0	0.8	0.8	1.0	0.8	0.6	0.6				
	0.00	0.34	0.41	0.33	0.44	0.45	0.73	0.52	0.63	0.40	0.63	0.55	0.31	0.37	0.46	0.45	0.46	0.66	0.53	0.36	0.39				
	0	31.8	15.3	43.8	11.7	27.1	22.7	44.7	67.5	37.3	51.6	70.2	29.8	77.4	157.4	75.1	93.4	143.4	95.9	55.2	87.9				
<b>E</b>	1.00	0.49	0.70	0.54	0.44	0.52	0.53	0.61	0.57	0.55	0.61	0.73	0.42	0.62	0.66	0.31	0.55	0.58	0.71						
	0.0	1.2	0.6	0.8	0.9	0.9	0.8	0.9	0.7	0.8	0.8	0.7	0.6	0.9	0.7	0.8	1.0	0.8	0.6	0.6					
	0.00	0.41	0.30	0.42	0.44	0.72	0.49	0.62	0.41	0.62	0.54	0.29	0.29	0.40	0.40	0.40	0.64	0.50	0.30	0.32					
	0	42.1	73.4	38.0	52.3	29.1	20.4	64.0	68.3	80.7	101.2	61.5	108.0	187.6	105.2	125.1	158.2	115.6	83.8	116.3					
<b>F</b>	1.00	0.65	0.53	0.70	0.63	0.23	0.55	0.71	0.67	0.57	0.78	0.54	0.28	0.35	0.39	0.33	0.51	0.39	0.44						
	0.0	1.0	1.0	1.0	1.5	1.6	1.6	1.1	1.3	1.5	0.9	1.1	1.7	1.2	1.6	1.7	1.5	1.3	1.3						
	0.00	0.29	0.39	0.39	0.76	0.60	0.72	0.51	0.70	0.65	0.30	0.30	0.56	0.45	0.56	0.73	0.61	0.35	0.42						
	0	45.0	23.8	11.8	20.4	58.2	82.5	27.4	38.6	59.5	23.8	74.8	154.5	73.7	86.0	152.6	103.4	57.0	88.6						
<b>G</b>	1.00	0.65	0.52	0.51	0.47	0.61	0.69	0.60	0.70	0.77	0.78	0.47	0.60	0.69	0.40	0.58	0.78	0.68							
	0.0	0.6	0.8	1.0	0.9	0.9	0.7	0.9	0.9	0.5	0.4	1.0	0.7	0.8	1.1	0.9	0.5	0.6							
	0.00	0.34	0.38	0.73	0.53	0.66	0.44	0.66	0.58	0.20	0.19	0.47	0.38	0.46	0.68	0.54	0.18	0.31							
	0	35.4	48.6	63.9	79.6	75.4	31.4	46.3	48.9	24.3	35.0	114.2	31.8	58.1	113.3	62.4	12.0	44.2							
<b>H</b>	1.00	0.41	0.40	0.32	0.45	0.54	0.50	0.55	0.58	0.53	0.39	0.39	0.47	0.31	0.46	0.53	0.52								
	0.0	0.9	1.0	1.0	0.8	0.9	0.9	0.7	0.7	1.0	0.8	0.9	1.1	0.9	0.7	0.7									
	0.00	0.46	0.75	0.58	0.70	0.52	0.68	0.62	0.37	0.35	0.54	0.46	0.53	0.71	0.59	0.35	0.41								
	0	34.7	34.4	46.2	59.6	37.7	54.0	69.5	28.1	70.1	149.6	67.2	89.2	131.7	84.3	46.0	78.6								
<b>I</b>	1.00	0.59	0.15	0.43	0.63	0.64	0.48	0.62	0.46	0.24	0.31	0.36	0.24	0.43	0.28	0.39									
	0.0	1.0	1.2	1.1	0.8	0.9	1.1	0.7	0.8	1.2	1.0	1.1	1.3	1.1	0.9	0.9									
	0.00	0.75	0.60	0.69	0.48	0.64	0.58	0.37	0.39	0.53	0.48	0.54	0.70	0.59	0.41	0.46									
	0	26.4	69.4	94.0	23.2	29.5	52.1	24.6	74.1	152.4	74.0	81.1	159.7	109.6	60.3	90.1									
<b>J</b>	1.00	0.24	0.39	0.58	0.60	0.42	0.59	0.45	0.26	0.43	0.43	0.18	0.41	0.27	0.43										
	0.0	0.6	0.6	0.6	0.5	0.5	0.8	0.9	0.5	0.9	0.5	0.6	0.9	0.8											
	0.00	0.73	0.74	0.71	0.69	0.73	0.72	0.73	0.70	0.74	0.70	0.75	0.72	0.73											
	0	48.7	84.4	47.4	55.8	78.3	44.2	95.0	174.8	93.7	106.1	166.0	118.6	75.8	108.0										
<b>K</b>	1.00	0.40	0.29	0.25	0.28	0.34	0.54	0.40	0.46	0.56	0.25	0.39	0.53	0.50											
	0.0	0.5	0.7	0.7	0.5	0.9	0.8	0.4	0.9	0.3	0.5	0.5	0.6	0.6											
	0.00	0.62	0.53	0.67	0.59	0.52	0.53	0.45	0.56	0.47	0.62	0.54	0.51	0.51											
	0	48.4	81.8	96.2	114.6	73.4	114.5	192.2	110.8	135.3	148.8	110.7	88.3	119.3											
<b>L</b>	1.00	0.56	0.55	0.51	0.63	0.56	0.41	0.41	0.50	0.45	0.67	0.49	0.50												
	0.0	0.7	0.6	0.4	0.9	0.9	0.4	0.9	0.4	0.4	0.4	0.7	0.7												
	0.00	0.60	0.66	0.61	0.64	0.67	0.56	0.67	0.57	0.62	0.57	0.65	0.64												
	0	94.7	112.0	122.7	84.2	104.6	171.9	99.2	132.5	104.0	75.4	77.9	100.4												
<b>M</b>	1.00	0.75	0.67	0.81	0.62	0.37	0.49	0.53	0.31	0.54	0.47	0.55													
	0.0	0.5	0.5	0.5	0.7	0.7	0.8	0.6	0.8	0.6	0.7	0.6													
	0.00	0.55	0.46	0.38	0.46	0.41	0.49	0.42	0.62	0.50	0.45	0.46													
	0	17.3	32.9	10.5	51.3	129.2	51.7	58.7	144.7	93.7	41.6	68.5													

N	1.00	0.62	0.72	0.56	0.34	0.40	0.46	0.28	0.55	0.36	0.50
	0.0	0.5	0.7	0.8	0.6	0.9	0.5	0.7	0.6	0.8	0.7
	0.00	0.56	0.62	0.66	0.59	0.67	0.59	0.69	0.62	0.65	0.64
	0	23.7	27.9	57.1	129.6	59.3	55.1	158.5	107.5	54.9	76.9
O	1.00	0.67	0.61	0.44	0.43	0.51	0.33	0.45	0.57	0.51	
	0.0	0.8	0.8	0.3	0.9	0.3	0.4	0.4	0.7	0.6	
	0.00	0.54	0.58	0.47	0.59	0.50	0.63	0.56	0.56	0.56	
	0	41.5	43.3	107.9	47.5	32.4	152.3	102.5	52.6	64.6	
P	1.00	0.66	0.40	0.50	0.55	0.34	0.58	0.54	0.55		
	0.0	0.6	1.0	0.8	0.8	1.0	0.8	0.6	0.7		
	0.00	0.24	0.42	0.38	0.43	0.66	0.51	0.25	0.33		
	0	51.0	130.6	50.2	63.7	137.1	86.4	35.8	65.7		
Q	1.00	0.48	0.67	0.72	0.35	0.55	0.70	0.73			
	0.0	0.9	0.6	0.7	1.0	0.8	0.5	0.5			
	0.00	0.47	0.37	0.46	0.68	0.54	0.20	0.30			
	0	80.1	6.1	30.5	110.6	62.8	27.4	21.3			
R	1.00	0.40	0.49	0.29	0.38	0.47	0.46				
	0.0	0.9	0.3	0.3	0.4	0.7	0.7				
	0.00	0.50	0.32	0.55	0.44	0.43	0.43				
	0	82.4	75.7	123.6	102.8	104.2	73.3				
S	1.00	0.66	0.19	0.38	0.54	0.64					
	0.0	0.8	1.1	0.9	0.6	0.6					
	0.00	0.49	0.69	0.58	0.37	0.41					
	0	36.64	105.4	57.05	22.78	17.67					
T	1.00	0.30	0.49	0.68	0.69						
	0.0	0.4	0.4	0.6	0.6						
	0.00	0.58	0.47	0.43	0.42						
	0	137.8	92.3	54.8	48.0						
U	1.00	0.43	0.36	0.31							
	0.0	0.4	0.8	0.8							
	0.00	0.59	0.66	0.65							
	0	51.0	103.7	90.2							
V	1.00	0.47	0.51								
	0.0	0.7	0.6								
	0.00	0.52	0.52								
	0	52.7	44.9								
W	1.00	0.64									
	0.0	0.5									
	0.00	0.29									
	0	32.7									
X	r	1.00									
	P90	0.0									
	COD	0.00									
	d	0									

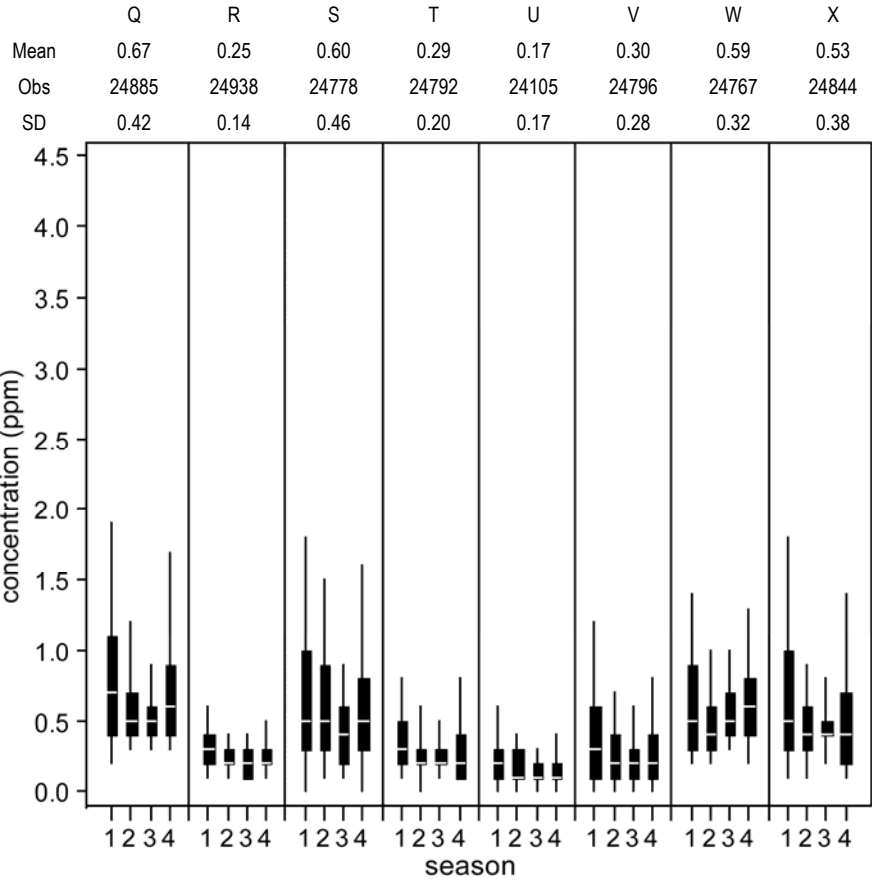


**Figure A-36a.** Box plots illustrating the seasonal distribution of hourly CO concentrations in Los Angeles, CA. Note: 1 = winter, 2, = spring, 3 = summer, and 4 = fall on the x-axis.



**Figure A-36b. Box plots illustrating the seasonal distribution of hourly CO concentrations in Los Angeles, CA. Note: 1 = winter, 2, = spring, 3 = summer, and 4 = fall on the x-axis.**





**Figure A-36c.** Box plots illustrating the seasonal distribution of hourly CO concentrations in Los Angeles, CA. Note: 1 = winter, 2, = spring, 3 = summer, and 4 = fall on the x-axis.

New York Combined Statistical Area

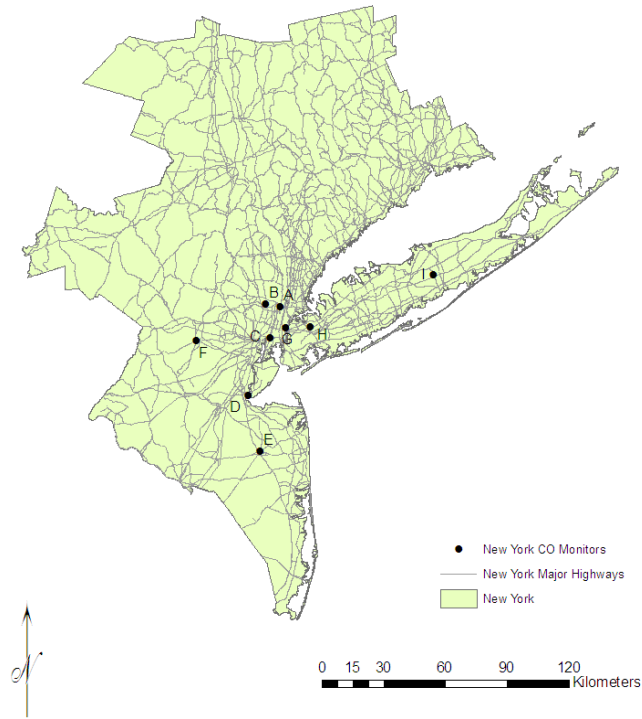


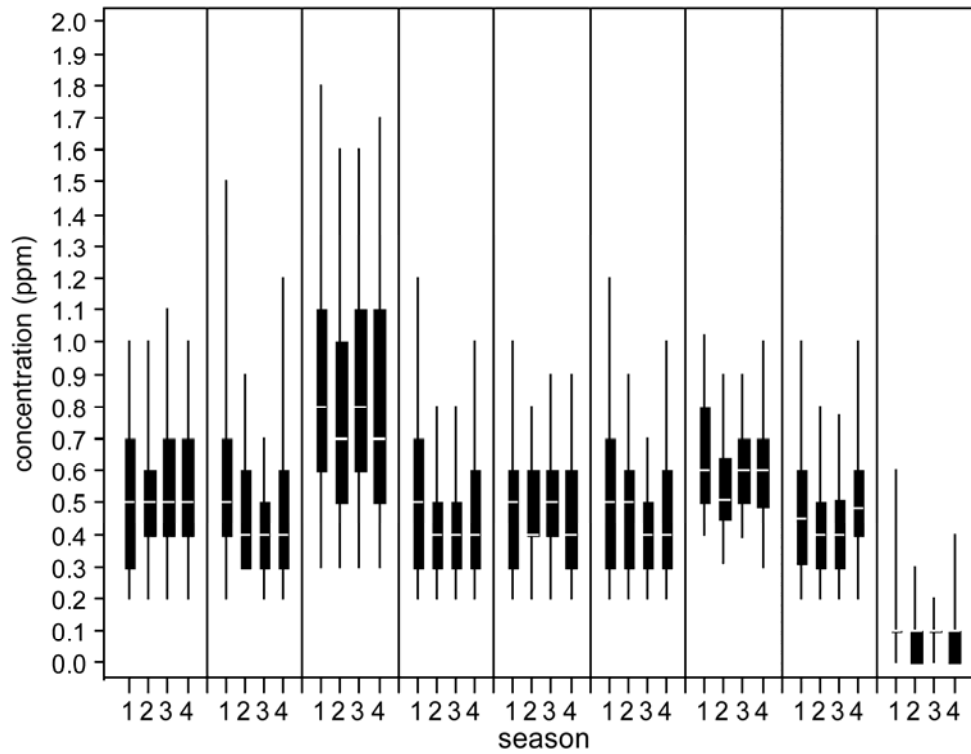
Figure A-37. Map of CO monitor locations with AQS Site IDs for New York City, NY.

**Table A-7. Table of inter-sampler comparison statistics, including Pearson r, P90 (ppm), COD, and d (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in New York City, NY.**

New York									
	A	B	C	D	E	F	G	H	I
A	1.00	0.58	0.65	0.55	0.40	0.56	0.56	0.41	0.30
	0.0	0.4	0.7	0.4	0.4	0.4	0.4	0.5	0.8
	0.00	0.22	0.28	0.25	0.25	0.24	0.23	0.28	0.75
	0.0	7.0	15.9	45.8	70.6	43.7	10.5	17.8	76.5
B		1.00	0.64	0.55	0.35	0.54	0.54	0.59	0.49
		0.0	0.8	0.4	0.5	0.4	0.4	0.4	0.7
		0.00	0.29	0.23	0.26	0.23	0.23	0.23	0.74
		0.0	16.8	45.4	71.5	38.1	15.0	24.5	82.9
C			1.00	0.54	0.32	0.48	0.52	0.43	0.31
			0.0	0.9	0.9	0.9	0.7	0.9	1.3
			0.00	0.35	0.34	0.34	0.24	0.35	0.81
			0.0	29.9	55.0	35.7	8.9	20.5	85.5
D				1.00	0.50	0.57	0.41	0.46	0.33
				0.0	0.4	0.4	0.4	0.4	0.7
				0.00	0.24	0.23	0.28	0.27	0.72
				0.0	27.5	36.7	37.5	45.1	107.8
E					1.00	0.47	0.33	0.33	0.32
					0.0	0.4	0.4	0.4	0.6
					0.00	0.23	0.25	0.27	0.73
					0.0	61.9	61.0	65.0	120.3
F						1.00	0.41	0.34	0.31
						0.0	0.4	0.4	0.7
						0.00	0.26	0.27	0.72
						0.0	43.6	55.8	119.7
G							1.00	0.46	0.29
							0.0	0.4	0.7
							0.00	0.26	0.77
							0.0	12.3	76.8
H								1.00	0.43
								0.0	0.6
								0.00	0.73
								0.0	65.1
I									1.00
									0.0
									0.00
									0.0

Legend	
r	1.00
P90	0.0
COD	0.00
d	0.0

	A	B	C	D	E	F	G	H	I
Mean	0.55	0.52	0.85	0.48	0.50	0.49	0.62	0.47	0.12
Obs	23113	25150	25646	25028	25727	25691	25547	25022	25749
SD	0.27	0.30	0.43	0.27	0.24	0.25	0.21	0.23	0.17



**Figure A-38. Box plots illustrating the seasonal distribution of hourly CO concentrations in New York City, NY. Note: 1 = winter, 2, = spring, 3 = summer, and 4 = fall on the x-axis.**

Seattle Combined Statistical Area

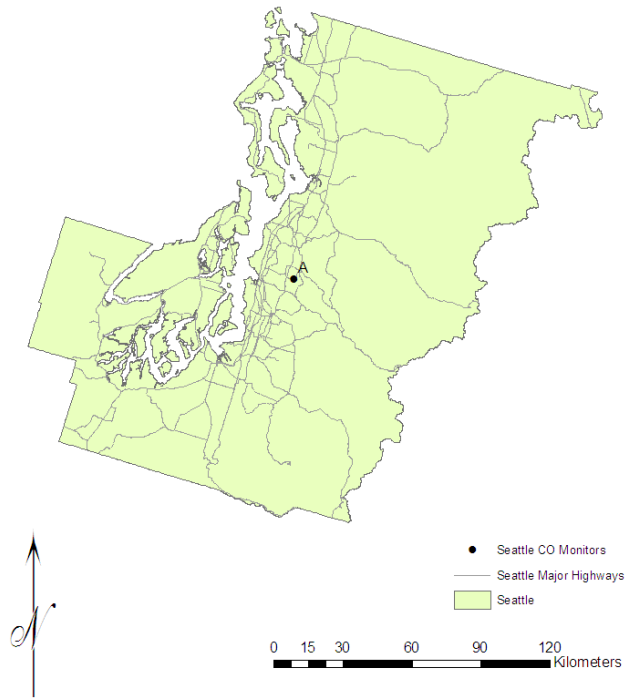
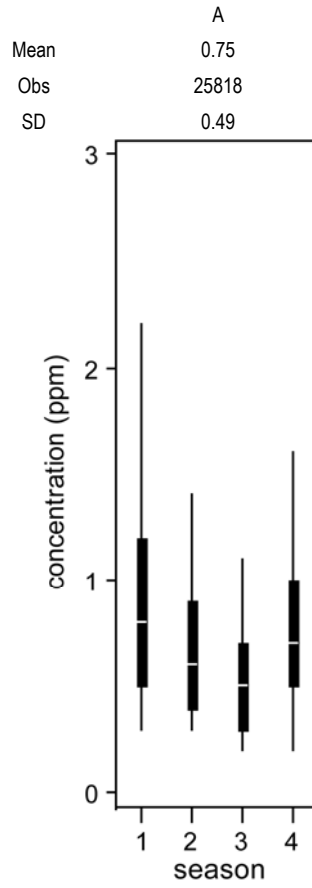


Figure A-39. Map of CO monitor locations with AQS Site IDs for Seattle, WA.



**Figure A-40.** Box plots illustrating the seasonal distribution of hourly CO concentrations in Seattle, WA. Note: 1 = winter, 2 = spring, 3 = summer, and 4 = fall on the x-axis.

St Louis Combined Statistical Area

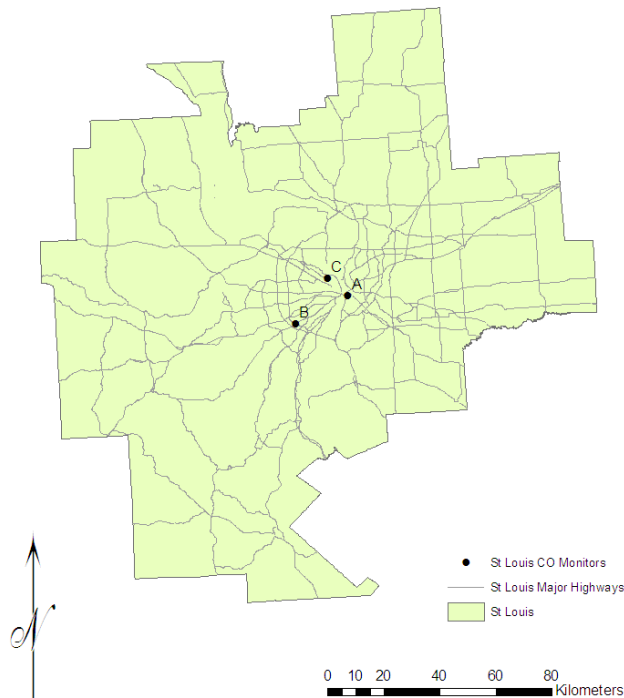


Figure A-41. Map of CO monitor locaitons with AQS Site IDs for St. Louis, MO.

Table A-8. Table of inter-sampler comparison statistics, including Pearson r, P90 (ppm), COD, and d (km), as defined in the text, for each pair of hourly CO monitors reporting to AQS in St. Louis, MO.

St. Louis			
	A	B	C
A	1.00 0.0 0.00 0.0	0.19 0.5 0.40 21.2	0.60 0.3 0.24 9.5
B		1.00 0.0 0.00 0.0	0.19 0.5 0.42 19.8
C			1.00 0.0 0.00 0.0

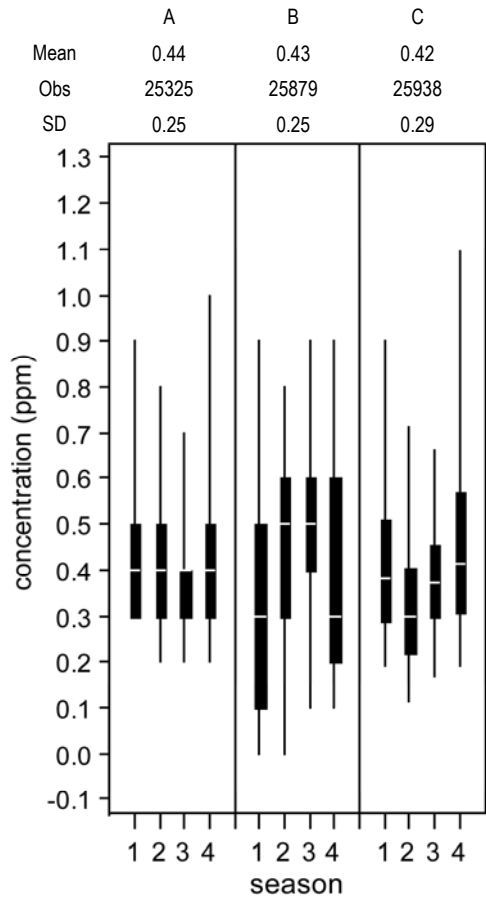
**Legend**

r

P90

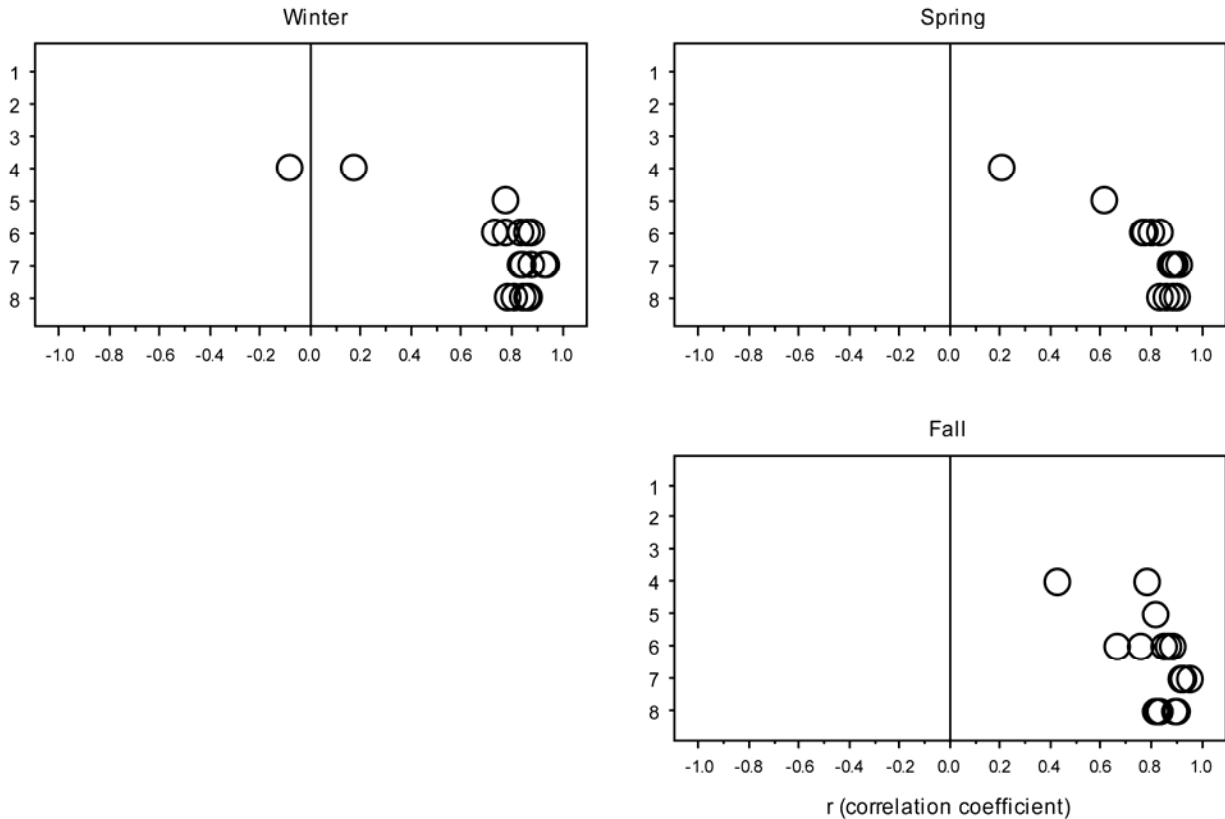
COD

d

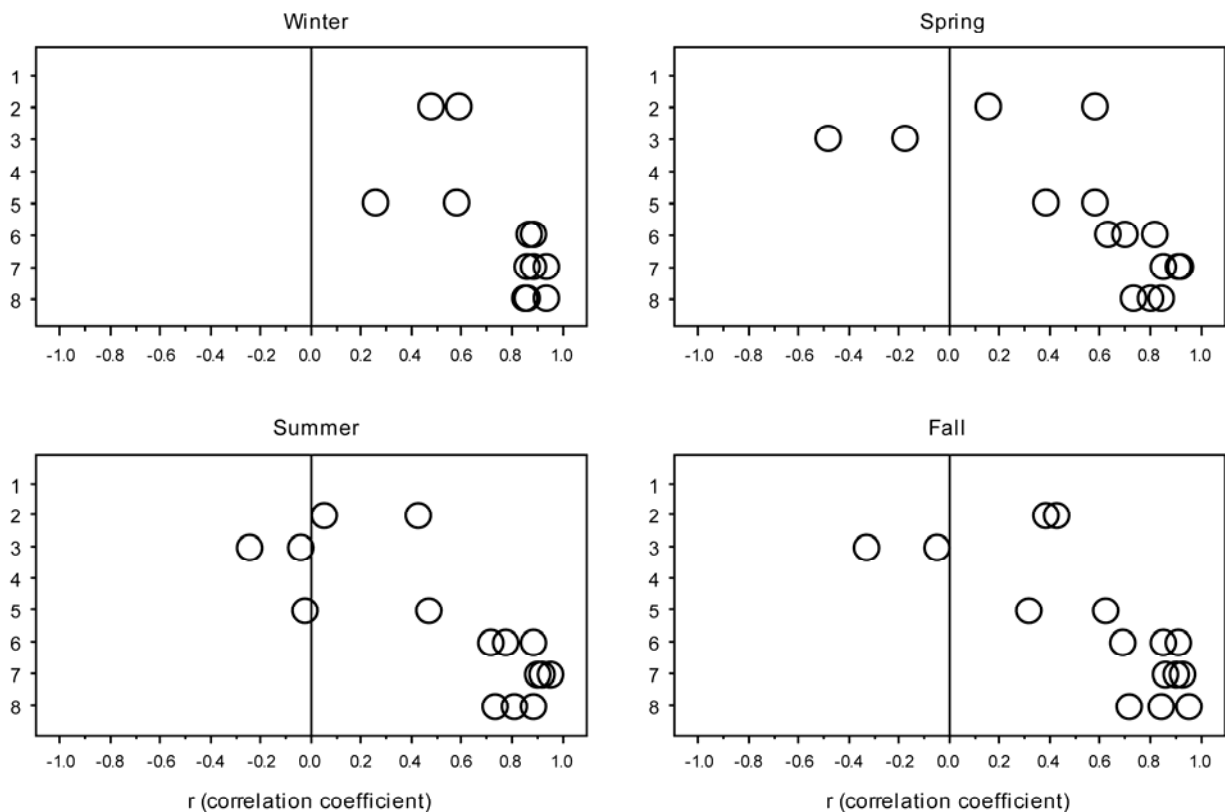


**Figure A-42.** Box plots illustrating the seasonal distribution of hourly CO concentrations in St. Louis, MO. Note: 1 = winter, 2, = spring, 3 = summer, and 4 = fall on the x-axis.

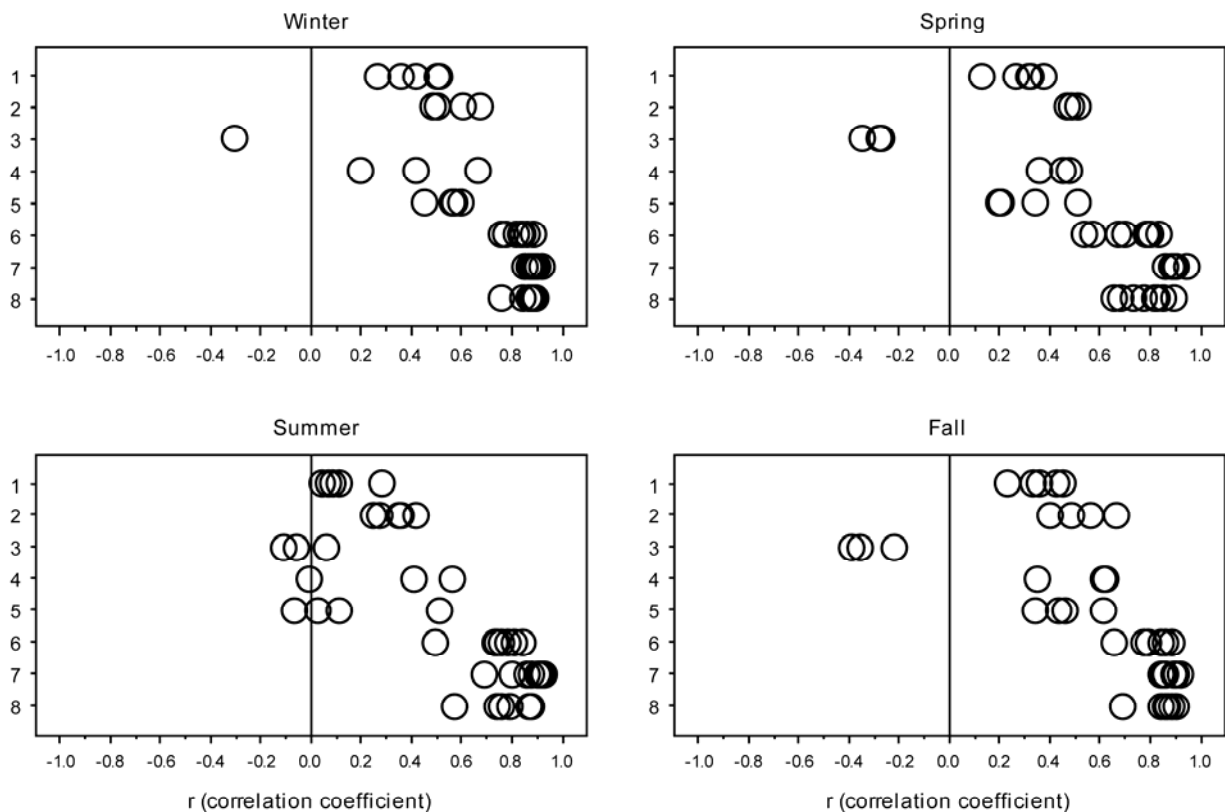




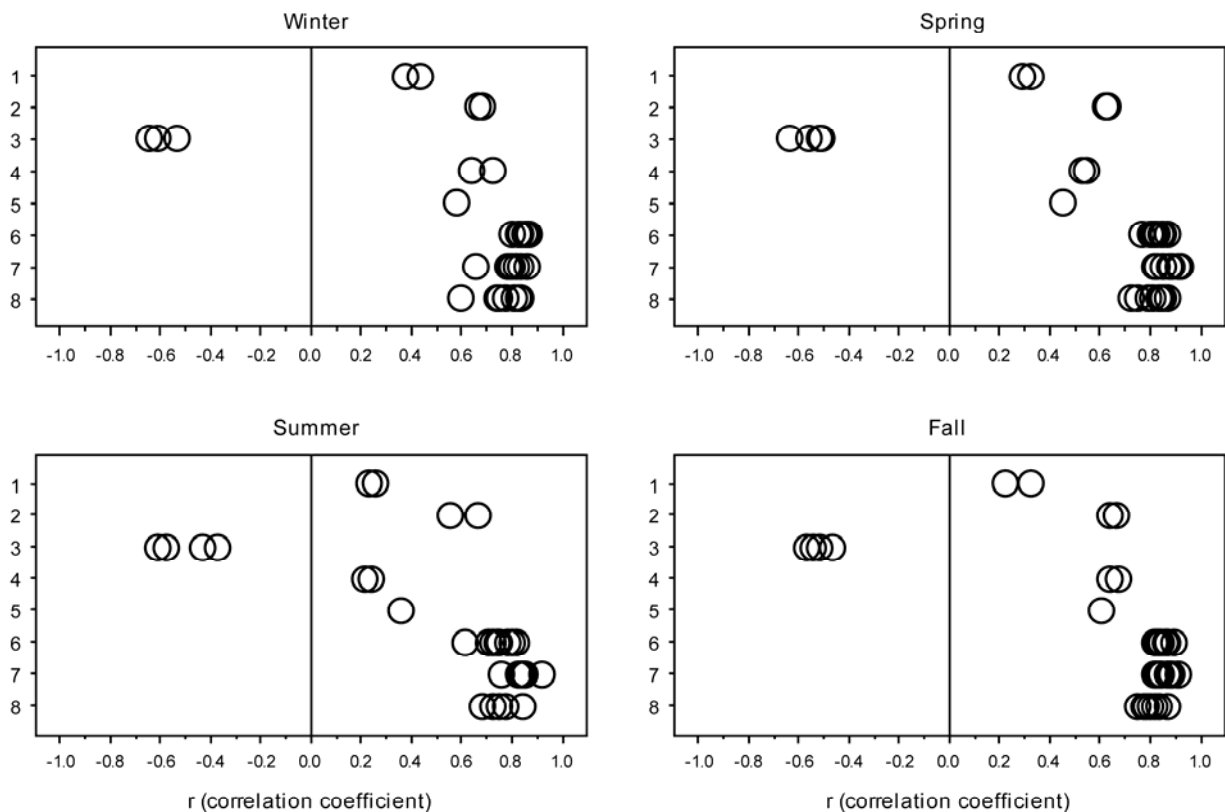
**Figure A-43. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for Anchorage, AK. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.) Note that the data are not obtained for Anchorage during the summer, and so are not presented here.**



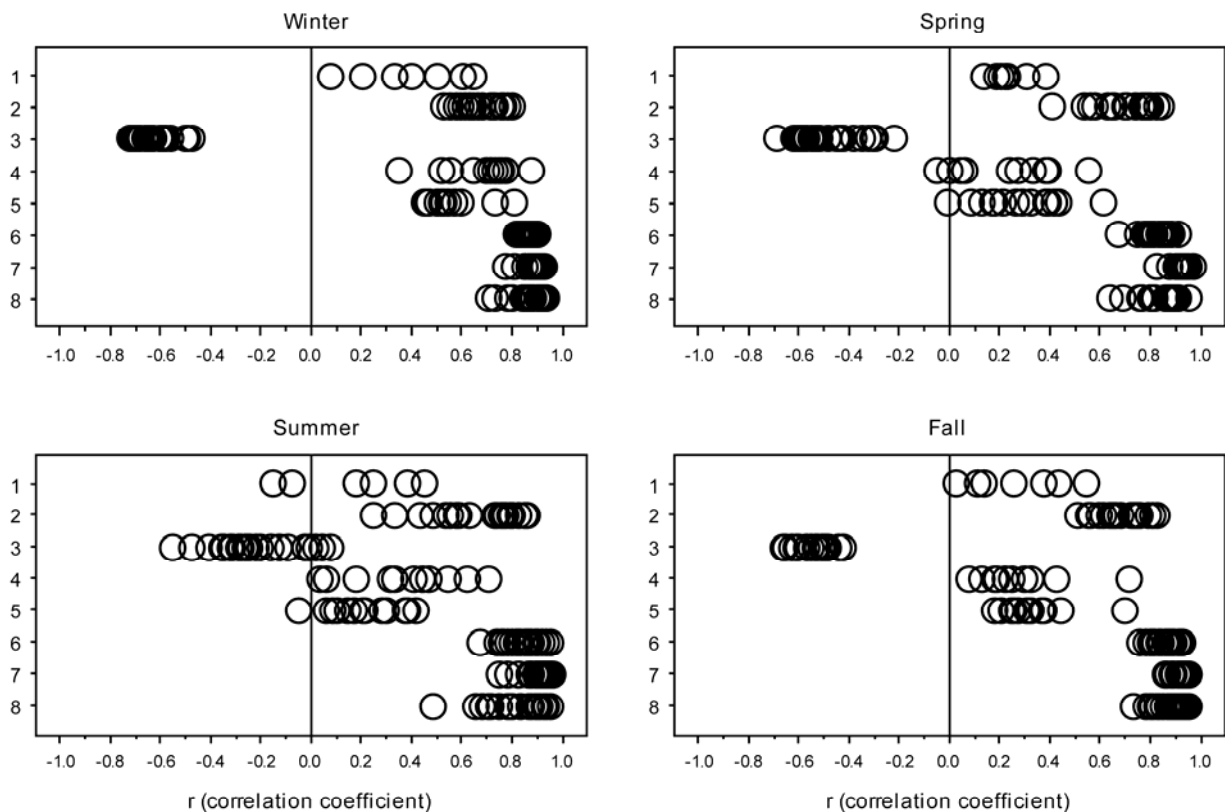
**Figure A-44. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for Atlanta, GA. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.)**



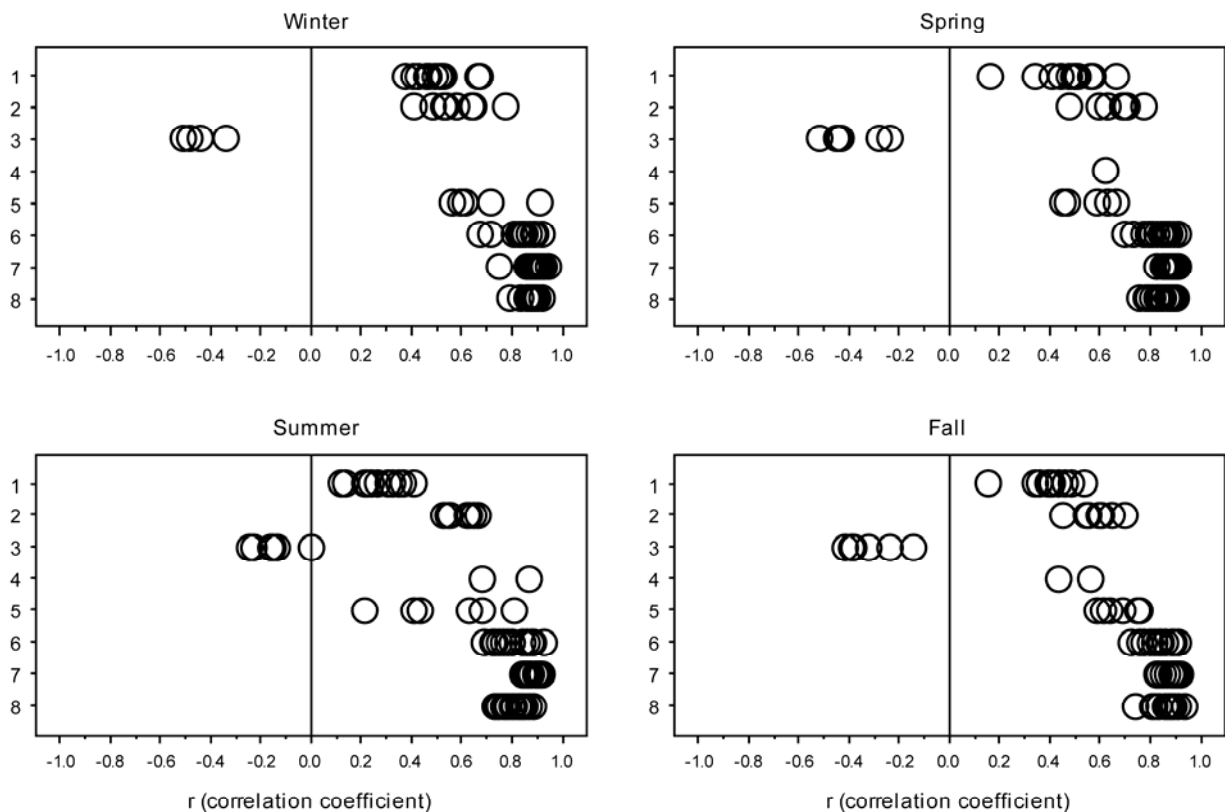
**Figure A-45. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for Boston, MA. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.) Red bars denote the median, and green stars denote the arithmetic mean.**



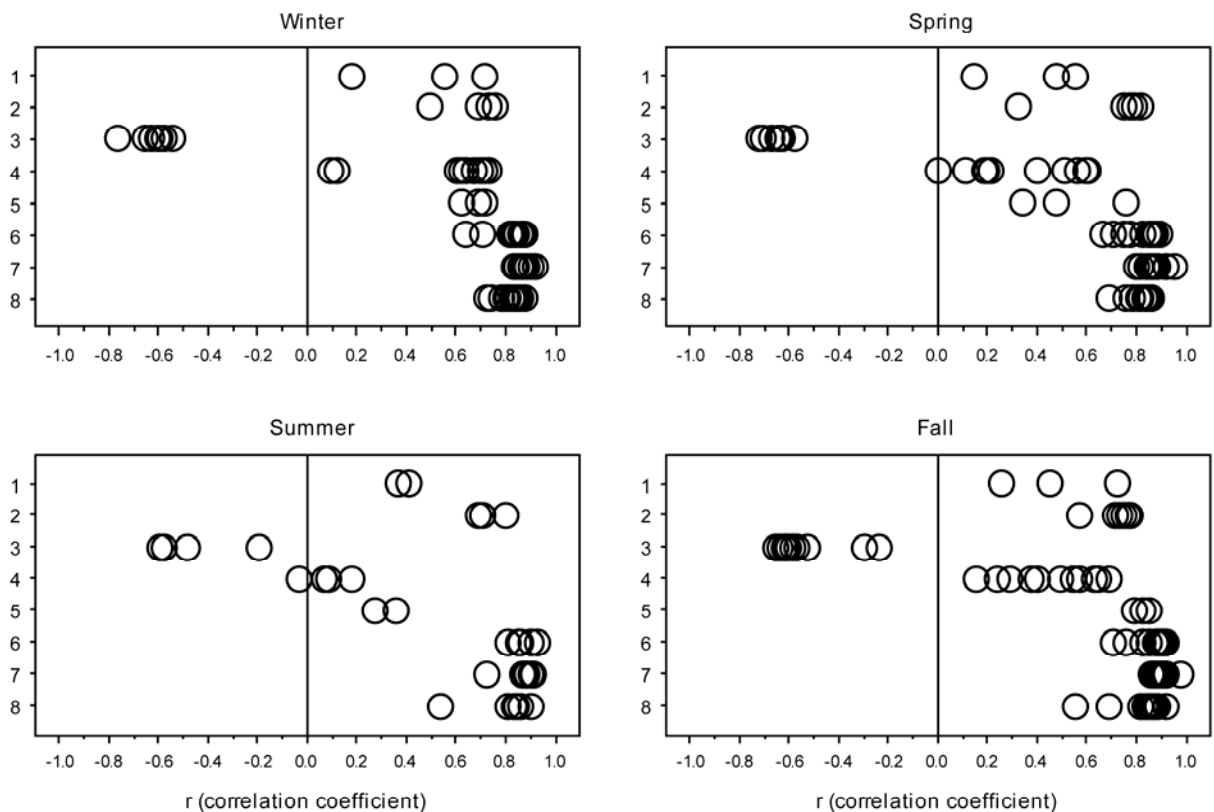
**Figure A-46. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for Denver, CO. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.) Red bars denote the median, and green stars denote the arithmetic mean.**



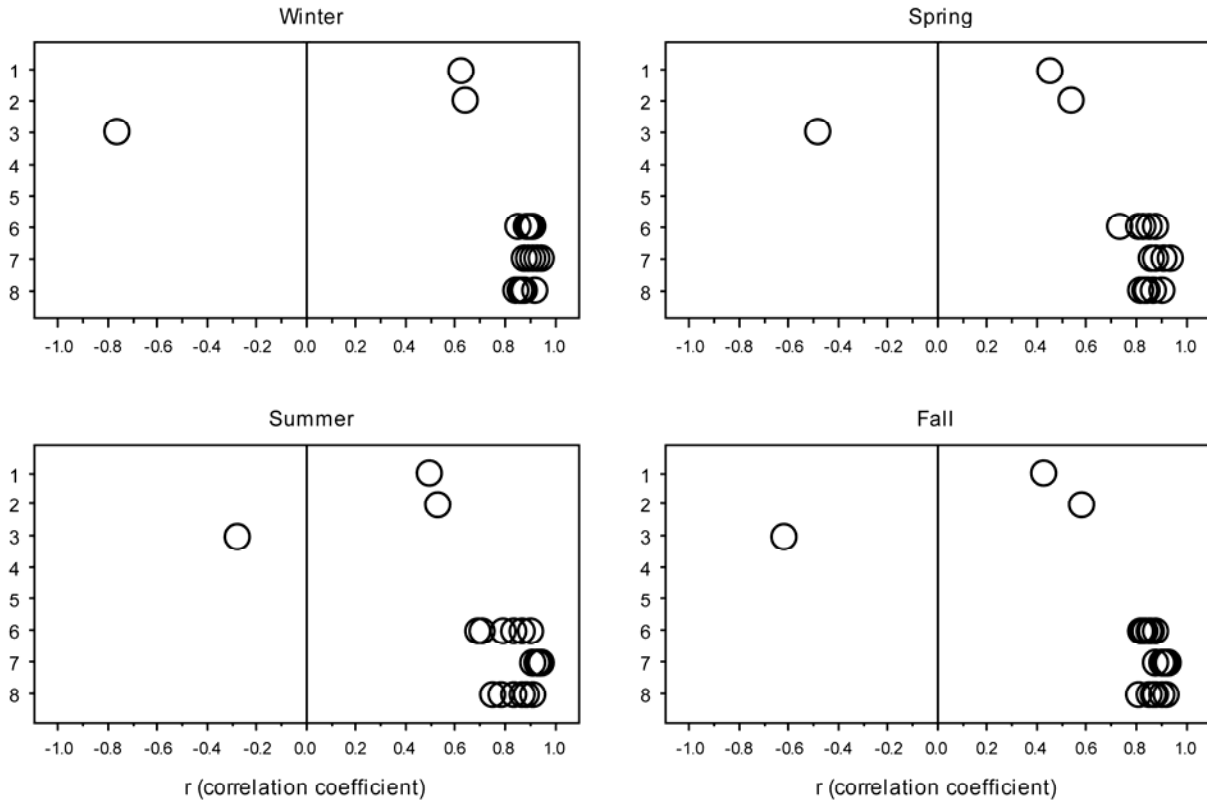
**Figure A-47. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for Los Angeles, CA. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.)**



**Figure A-48. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for New York City, NY. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.)**



**Figure A-49. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for Phoenix, AZ. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.)**



**Figure A-50. Seasonal plots of correlations between hourly CO concentration with hourly (1) SO<sub>2</sub>, (2) NO<sub>2</sub>, (3) O<sub>3</sub>, (4) PM<sub>10</sub>, and (5) PM<sub>2.5</sub> concentrations for Seattle, WA. Also shown are correlations between 24-h average CO concentration with (6) daily max 1-h and (7) daily max 8-h CO concentrations and (8) between daily max 1-h and daily max 8-h CO concentrations. (Refer the numbers in this caption to those on the y-axis of each seasonal plot.)**



# Annex B. Dosimetry Studies

**Table B-1. Recent studies related to CO dosimetry and pharmacokinetics.**

Reference	Purpose	Findings
Alcantara et al. (2007)	To use a quantum mechanics/molecular mechanics approach to understand the cooperativity of Hb ligand binding and differences in energy between T and R Hb functional states.	The ligand binding energies between R and T states differ due to strain induced in the heme and its ligands and in protein contacts in the $\alpha$ and $\beta$ chains.
Adir et al. (1999)	To determine if low concentrations of CO would affect exercise performance and myocardial perfusion in young healthy men.	Men with COHb levels between 4-6% had decreased exercise performance measured by decreased mean duration of exercise (1.52 min) and maximal effort described by metabolic equivalent units (2.04). No changes were seen in lactate/pyruvate ratio, arrhythmias, or myocardial perfusion.
Andersson et al. (2000)	To investigate if CO could be endogenously produced in the nose and paranasal sinuses.	Both nose and paranasal sinuses contained HO-like immunoreactivity, mostly in the respiratory epithelium, indicating local CO production in the upper respiratory airways.
Bruce and Bruce (2003)	To create a mathematical model to predict uptake and distribution of CO in both vascular and tissue compartments during constant or variable inhalation levels of CO.	This model contains 5 compartments: lung, arterial blood, venous blood, muscle tissue, and nonmuscle tissue. It was constructed to include tissue compartment flux and difference between venous and arterial COHb for short exposures which is not possible with the CFK model.
Bruce and Bruce (2006)	To use their mathematical multicompartment model along with experimental data to predict the factors that influence the washout rates of CO, along with predicting the rates of CO uptake, distribution in vascular and extravascular (muscle and non-muscle tissue) compartments, and washout over a range of exposure and conditions.	Rates of CO washout follow a biphasic elimination where washout was faster immediately post exposure. The difference in rates is likely due to slow equilibration between vascular and extravascular compartments. Important factors contributing to washout kinetics include: peak COHb level, exposure duration and concentration, time after exposure samples were obtained, and individual variability.
Bruce et al. (2008)	To develop a mathematical model able to integrate a large body of indirect experimental findings on the uptake and distribution of CO by accounting for arteriole to venule shunting via intra-tissue pathways and diffusion of blood gases into tissues from pre-capillary vessels like arterioles.	The former model of Bruce and Bruce (2006) was altered by adding a mass balance equation for $O_2$ so $pO_2$ is directly calculated in the compartments and the muscle compartment is divided into two sub-compartments of muscle and non-muscle tissue. CO uptake from blood by muscle is much slower than $O_2$ , thus COHb% will fall rapidly while COMb% could remain high.
Carraway et al. (2000)	To test the hypothesis that HO-1 gene expression and protein are upregulated in the lungs of rats during chronic hypoxia.	Rats were exposed to HH (17,000 ft) for 1-21 days. COHb increased after 1 day and progressively after 14 days. HO-1 protein and activity were upregulated during early chronic hypoxia. This HO-1 was localized to inflammatory cells and then to newly muscularized arterioles.
Castillo et al. (2006)	To describe a new method for measurement of CO diffusing capacity ( $D_LCO$ ) and alveolar volume ( $V_A$ ) in sleeping infants (6-22 months old), using a single 4-sec breath-hold technique.	$V_{A30}$ and $D_LCO$ increased with increasing body length and the method could be used as a measurement of lung development and growth.
Chakraborty et al. (2004)	To present an analytical expression for diffusing capacity of CO, NO, $CO_2$ , and $O_2$ to the red blood cell in terms of optimum size and shape of the RBC, thickness of the unstirred plasma layer surrounding the RBC, diffusivities and solubilities of the gas in RBC and boundary layer, hematocrit, and the slope of the dissociation curve.	Results indicate the discoidal shape of the RBC is optimal for $O_2$ uptake and reaction velocity is limited by mass transfer resistance in surrounding stagnant plasma layer. The paper overviews rate constants and reaction kinetics for CO binding to Hb. CO diffusing capacity is shown to be reaction rate limited at low $P_{CO}$ under normoxic and hyperoxic conditions, but diffusion rate limited under hypoxic and high $P_{CO}$ conditions.
Cronje et al. (2004)	To analyze CO uptake and elimination in the brain, muscle, heart, and blood of rats, with the intent of testing the Warburg hypothesis that CO partitioning is directly proportional to the $CO/O_2$ ratio.	Results indicate that tissue and blood [CO] dissociate during CO inhalation, but [CO] does not follow blood [CO] or $1/pO_2$ as in the Warburg theory during intake or elimination. Tissue [CO] increases later during the resolution period and varies significantly among animals and tissues. The deviation from the predicted values in the brain is likely due to the release of heme and increase in NADPH stimulating endogenous CO production by HO.

Reference	Purpose	Findings
Dutton et al. (2001)	To monitor CO, NO <sub>2</sub> , and PAH emissions during the operation of unvented natural gas fireplaces in two residences in Boulder, CO, at various times between 1997 and 2000.	Results showed significant accumulation of CO, NO <sub>2</sub> , and PAH indoors when the fireplaces were used. CO concentrations could exceed 100 ppm. NO <sub>2</sub> concentrations average 0.36 ppm over 4 hours. PAH 4-h time average reached 35 ng/m <sup>3</sup> .
Hampson (2007)	To present a case study of a man with drug-induced hemolytic anemia and hepatic failure.	The man had elevated endogenous CO production resulting in levels of COHb as high as 9.7%.
Hsia (2002)	To review the current concepts and practical relevance of the diffusing capacity/cardiac output interaction, in hopes of aiding in the interpretation of diffusing capacity, membrane diffusing capacity, and capillary blood volume.	This review helped to understand the determinants of changes in diffusing capacity, including hematocrit, erythrocyte distribution, blood volume, lung volume, cardiac output, etc.
Lamberto et al. (2004)	To evaluate which component, alveolar membrane diffusing capacity (D <sub>m</sub> ) and pulmonary capillary blood volume (V <sub>c</sub> ), is responsible for decreased resting D <sub>L</sub> CO in sarcoidosis patients and which component is the best predictor of gas exchange abnormalities.	Patients with pulmonary sarcoidosis had decreased lung volumes, a loss in D <sub>L</sub> CO, and gas exchange abnormalities during exercise including decreased P <sub>a</sub> O <sub>2</sub> and increased alveolar-arterial oxygen pressure difference. D <sub>m</sub> accounted for the majority of the decrease in D <sub>L</sub> CO and was predictive for gas exchange abnormalities.
Levesque et al. (2000)	To describe the results of air quality monitoring in an indoor ice skating rink during Monster Truck and car demolition exhibitions.	Maximum time-weighted average levels of CO were 100 ppm with several peaks exceeding 200 ppm (maximum: 1,600 ppm).
Marks et al. (2002)	To review the analytical methods for measurement of endogenous formation of CO in a variety of tissues.	A variety of methods have been used to measure endogenous CO. The rate of formation varies over a narrow range from 0.029 nmol/mg protein/h to 0.28 nmol/mg protein/h depending on tissue. Brain and liver regions tend to have the highest rates of CO formation likely due to high levels of HO activity in these tissues.
Merx et al. (2001)	To investigate the effect of CO inactivation of Mb in wild-type and myo <sup>-/-</sup> mice on hemodynamics and oxygen dynamics.	Fully oxygenated Mb treated with 20% CO had no change in left ventricular developed pressure or coronary venous PO <sub>2</sub> . Partially O <sub>2</sub> -saturated Mb (87% O <sub>2</sub> Mb) exposed to 20% CO had significantly decreased LVDP (12%) and PvO <sub>2</sub> (30%) in wild-type but not myo <sup>-/-</sup> hearts.
Pelham et al. (2002)	To review the literature on exposure and effects of mainly CO and NO <sub>2</sub> in enclosed ice rinks.	CO levels as high as 300 ppm were recorded after episodes of malfunctioning ice resurfacing equipment or inadequate ventilation.
Pesola et al. (2004)	To determine if healthy African Americans may be misdiagnosed as having respiratory deficient due to comparison using Caucasian-derived prediction equation estimates of D <sub>L</sub> CO.	The lung volume of African American individuals is 10-15% lower than Caucasians. The measured D <sub>L</sub> CO was consistently significantly lower in African Americans than what would be predicted, thus the authors suggest a race correction reduction of the Miller PEE for diffusion of 12%.
Pesola et al. (2006)	To determine if healthy Asians may be misdiagnosed as having respiratory deficient due to comparison using Caucasian-derived prediction equation estimates of D <sub>L</sub> CO.	The lung volume of Asian individuals is 10-15% lower than Caucasians, thus a Chinese derived prediction for D <sub>L</sub> CO should be used.
Prommer and Schmidt (2007)	To determine the error in total Hb mass measurements using the optimized CO-rebreathing method due to loss of CO to Mb	Optimal blood mixing (when venous and arterial blood COHb% are equivalent) was determined to be after 6 min. A small volume of administered CO leaves the vascular space (0.32% per min). 2.3% increase in total Hb mass would be found if CO diffusion was not included.
Richardson et al. (2002)	To combine invasive vascular measures of arterial and venous blood and muscle blood flow with noninvasive magnetic spectroscopy of deoxy-myoglobin and high energy phosphates to determine the effects of mild CO poisoning (20% COHb) in humans during muscular work.	Five humans were analyzed under normoxia, hypoxia, normoxia + CO (20% COHb), and 100% O <sub>2</sub> + CO. Maximum works rates and maximal oxygen uptake were reduced in H, CO <sub>norm</sub> , and CO <sub>hyper</sub> . CO and H caused elevated blood flow. Net muscle CO uptake from blood was less during 20% COHb trials than during normoxia and hypoxia (1-2%) trials.
Shimazu et al. (2000)	To study the effects of short-term (minutes) or long-term (several hours) CO exposure on COHb elimination and developing a mathematical model to simulate this event.	COHb exhibited an initial rapid decrease followed by a slower phase which is compatible with a 2-compartment model and biphasic elimination. Both exposures fit the 2-compartment, single central outlet mathematical model.
Shimazu (2001)	To discuss the findings of Weaver et al. (2000) on COHb t <sub>1/2</sub> .	The authors discuss that CO elimination is biphasic and is heavily affected by duration of exposure which was not taken into account in the Weaver, et al. (2000) paper.
Takeuchi et al. (2000)	To examine the relationship between minute ventilation and rate of COHb reduction during breathing 100% O <sub>2</sub> and during normocapnic hyperoxic hyperpnea.	Patients were exposed to 400 to 1,000 ppm CO, resulting in 10-12% COHb. The half-time of COHb reduction was 78 ± 24 min during 100% O <sub>2</sub> treatment and 31 ± 6 min during normocapnic hyperpnea with O <sub>2</sub> treatment.
Vreman et al. (2005)	To develop a sensitive and reproducible method of CO quantification in rodent (mouse and rat) tissue pre- and post-exposure in hopes of understanding endogenous CO production.	Tissues were sonicated mixed with sulfosalicylic acid for 30 min at 0°C and then liberated CO was analyzed by gas chromatograph. Blood contained the highest CO concentration. Lowest concentrations were found in brain, testes, intestine, and lung (endogenously).

Reference	Purpose	Findings
Vreman et al. (2006)	To test a method of CO quantification in frozen postmortem human tissues from 3 determined categories of fatalities: trauma with no suspected CO exposure (controls), fire-related, and CO asphyxiation.	CO levels were analyzed in adipose, brain, muscle, heart, kidney, lung, spleen, and blood (ordered from approximate low to high tissue concentration). It was suggested that blood, muscle, brain, lung, and kidney are suitable for diagnosing death due to lethal CO exposure due to regression analysis against COHb values.
Weaver et al. (2000)	To determine in COHb half-life is influenced by CO poisoning vs. experimental CO exposure, loss of consciousness, concurrent tobacco smoking, or P <sub>a</sub> O <sub>2</sub> .	COHb t <sub>1/2</sub> determined was 74 ± 25 min with a range from 26 to 148 min by a single exponential decrease function. This is shorter than most clinical studies and was inversely proportionate to P <sub>a</sub> O <sub>2</sub> , however not influenced by age, gender, smoke inhalation, loss of consciousness, tobacco smoking, or method of O <sub>2</sub> treatment.
Wu and Wang (2005)	To review the endogenous production of CO through HO, as well as discuss physiological roles for CO both toxic and therapeutic.	CO is produced endogenously by HO-1 and -2 and acts as a gasotransmitter, inducing cell signaling cascades. The review discusses possible roles for CO in the various organ systems. Also, it discusses the potential pharmacological and therapeutic applications for CO.

# Annex C. Epidemiology Studies

**Table C-1. Studies of CO exposure and cardiovascular morbidity.**

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<i>CHANGES IN HEART RATE AND HEART RATE VARIABILITY</i>			
<b>Author:</b> Chan et al. (2005) <b>Period of Study:</b> December 2001 – February 2002 <b>Location:</b> Taipei, Taiwan	<b>Health Outcome:</b> Various measures of HRV via ambulatory ECG (Holter system) <b>Study Design:</b> Panel <b>Statistical Analyses:</b> Linear regression (mixed effects) <b>Age Groups Analyzed:</b> 40 – 75 years <b>Sample Description:</b> 83 patients from the National Taiwan University Hospital	<b>Averaging Time:</b> 1-h ma <b>Mean (SD) unit:</b> 1.1 ppm <b>Range (Min, Max):</b> 0.1, 7.7 <b>Copollutant:</b> NR	<b>Increment:</b> NR <b>RR Estimate [Lower CI, Upper CI]</b> <b>Lags examined (-h ma):</b> 1, 2, 3, 4, 5, 6, 7, 8 CO had no statistically significant effect on SDNN, rMSSD, LF, HF.
<b>Author:</b> Dales et al. (2004) <b>Period of Study:</b> NR <b>Location:</b> Toronto, Canada	<b>Health Outcome:</b> Various measures of HRV via Holter system <b>Study Design:</b> Panel <b>Statistical Analyses:</b> Linear regression (mixed effects) <b>Age Groups Analyzed:</b> 51 – 88 years (mean 65 years) <b>Sample Description:</b> 36 subjects with pre-existing CAD	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 2.40 ppm (95th percentile) Personal monitoring <b>Range (Min, Max):</b> 0.4, 16.5 <b>Copollutant: correlation</b> PM <sub>2.5</sub> : r = 0.17	<b>Increment:</b> NR <b>Regression co-efficient [Lower CI, Upper CI]</b> Lags examined : NR CO had no statistically significant effect on LF, HF, HFLFR, SDNN among those taking Beta-blockers. Whereas CO had a positive effect on SDNN among those not taking Beta-blockers. Slope = 0.0111 (0.002-0.020, p=0.02)
<b>Author:</b> Gold et al. (2000) <b>Period of Study:</b> June – September 1997 <b>Location:</b> Boston, MA	<b>Health Outcome (ICD9 or ICD10):</b> Heart Rate and various measures of HRV via Holter system <b>Study Design:</b> Panel/Cohort <b>Statistical Analyses:</b> Linear regression (fixed effects/random effects) <b>Age Groups Analyzed:</b> 53 – 87 years <b>Sample Description:</b> 21 active Boston residents observed up to 12 times.	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.47 ppm <b>Range (Min, Max):</b> 0.12, 0.82 <b>Copollutant:</b> NR	<b>Increment:</b> 0.6 ppm <b>% Change [Lower CI, Upper CI]</b> Lags examined : 24-h No significant effect with CO (no results recorded)
<b>Author:</b> Gold et al. (2005) <b>Period of Study:</b> June – September 1999 <b>Location:</b> Boston, MA	<b>Health Outcome:</b> ST- segment. <b>Study Design:</b> Panel <b>Statistical Analyses:</b> Linear regression (mixed models) <b>Age Groups Analyzed:</b> 61 – 88 years <b>Sample Description:</b> 24 Active Boston residents - each observed up to 12 times.	<b>Averaging Time:</b> 1-h, 24-h <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> (ppm) (personal monitoring) 10th = 0.20 90th = 1.08 <b>Copollutant:</b> NR	<b>Increment:</b> NR <b>RR Estimate [Lower CI, Upper CI]</b> Lags examined : 1-24 h Although CO was associated with ST-segment depression in single pollutant models, this result did not persist in multiple pollutant models.
<b>Author:</b> Holguin et al. (2003) <b>Period of Study:</b> February – April 2000 <b>Location:</b> Mexico City, Mexico	<b>Health Outcome:</b> Various measures of HRV via ECG <b>Study Design:</b> Panel <b>Statistical Analyses:</b> GEE <b>Age Groups Analyzed:</b> 60-96 years (mean age 79 years) <b>Sample Description:</b> 34 patients who were permanent residents of a nursing home in the Northeast metropolitan area.	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 3.3 ppm <b>Range (Min, Max):</b> 1.8, 4.8 <b>Copollutant:</b> NR	<b>Increment:</b> 10 ppm <b>Regression Coefficients [Lower CI, Upper CI]</b> Lags examined : 0 Lag 0 : HF : 0.003 (-0.004 to 0.001) LF : 0.001 (-0.006 to 0.008) LF/HF : 0.001 (-0.005 to 0.002)

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Liao et al. (2004) <b>Period of Study:</b> 1996 - 1998 <b>Location:</b> Forsyth County, NC; Selected suburbs of Minneapolis, MN; Jackson, MI	<b>Health Outcome:</b> Heart Rate & various rates of HRV. <b>Study Design:</b> Cohort <b>Statistical Analyses:</b> Linear regression <b>Age Groups Analyzed:</b> 45 – 64 years (mean 62 years) <b>Sample Description:</b> 6784 study subjects from the atherosclerosis risk in communities study	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.65 ppm (0.44) <b>Range (Min, Max):</b> NR <b>Copollutant:</b> NR	<b>Increment:</b> 0.44 ppm <b>Regression coefficients Lags examined : 1</b> Lag 1 : HF (log transformed) : -0.033 LF (log transformed) : 0.006 SDNN : -0.274 Heart Rate (bpm) : 0.404* Confidence Intervals not recorded *p < 0.05
<b>Author:</b> Park et al. (2005b) <b>Period of Study:</b> 2000 - 2003 <b>Location:</b> Boston, MA	<b>Health Outcome:</b> Various measures of HRV via ECG <b>Study Design:</b> Panel/Cohort <b>Statistical Analyses:</b> Linear regression <b>Age Groups Analyzed:</b> 21 – 81 years <b>Sample Description:</b> 497 men from the Normative aging study in Greater Boston	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.50 ppm <b>Range (Min, Max):</b> 0.13, 1.8 <b>Copollutant:</b> NR	<b>Increment:</b> 0.24 ppm <b>% Change in HRV [Lower CI, Upper CI]</b> Lags examined: 4-h ma, 24-h ma, 48-h ma <b>Lag 4-h ma:</b> SDNN (Log10): 2.0 (-2.9 to 7.3) HF (Log10): 8.8 (-4.6 to 24.1) LF(Log10) : 3.2 (-7.0 to 14.6) LF :HF(Log10) : -5.1 (-13.5 to 4.1) <b>Lag 24-h ma:</b> SDNN (Log10): -2.2 (-7.7 to 3.6) HF (Log10): -13.2 (-25.4 to 1.0) LF(Log10) : -0.6 (-11.9 to 12.1) LF :HF(Log10) : 14.5 (2.9-27.5) <b>Lag 48-h ma:</b> SDNN(Log10): -3.4 (-10.2 to 3.9) HF (Log10): -13.8 (-28.9 to 4.4) LF (Log10): -2.4 (-16.2 to 13.6) LF :HF (Log10): 13.2 (-1.1 to 29.6)
<b>Author:</b> Peters et al. (1999a) <b>Period of Study:</b> 1984 - 1985 <b>Location:</b> Augsburg, Germany	<b>Health Outcome:</b> Heart Rate <b>Study Design:</b> Cohort <b>Statistical Analyses:</b> Linear regression (GEE) <b>Age Groups Analyzed:</b> 25 – 64 years <b>Sample Description:</b> 2681 men & women who participated in the MONICA study	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> During air pollution episode: 4.54 mg/m <sup>3</sup> Outside air pollution episode: 4.51 mg/m <sup>3</sup> <b>Range (Min, Max):</b> During air pollution episode: 2.39, 6.85 Outside air pollution episode: 0.91, 11.51 Respectively <b>Copollutant:</b> NR	<b>Increment:</b> 6.6 mg/m <sup>3</sup> <b>Mean Change in Heart Rate (beats/min) [Lower CI, Upper CI]</b> Lags examined: 0, 5-day avg <b>All</b> Lag 0 : 0.97 (0.02-1.91) Lag 5-day avg : 0.70 (-0.09 to 1.48) <b>Men</b> Lag 0 : 0.95 (-0.37 to 2.27) Lag 5-day avg : 0.91 (-0.25 to 2.07) <b>Women</b> Lag 0 : 0.98 (-0.37 to 2.34) Lag 5-day avg : 0.52 (-0.55 to 1.59)
<b>Author:</b> Riojas-Rodriguez et al. (2006) <b>Period of Study:</b> December 2001 – April 2002 <b>Location:</b> Mexico City, Mexico	<b>Health Outcome:</b> Various measures of HRV via Holter system <b>Study Design:</b> Panel <b>Statistical Analyses:</b> Linear regression (mixed effects models) <b>Age Groups Analyzed:</b> 25 – 76 years (mean 55 years) <b>Sample Description:</b> 30 patients from the Outpatient clinic of the National Institute of Cardiology of Mexico	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 2.9 ppm (personal monitor) <b>Range (Min, Max):</b> 0.1, 18.0 <b>Copollutant:</b> NR	<b>Increment:</b> 1 ppm <b>Regression Coefficients [Lower CI, Upper CI]</b> Lags examined (per minutes) : 5, 10 Lag 5 minutes : HF : -0.006 (-0.023 to 0.010) LF : -0.024 (-0.041 to -0.007) VLF : -0.034 (-0.061 to -0.007) Notes: VLF = Very low frequency
<b>Author:</b> Schwartz et al. (2005) <b>Period of Study:</b> 1999 <b>Location:</b> Boston, MA	<b>Health Outcome:</b> Measures of HRV via Holter system <b>Study Design:</b> Panel <b>Statistical Analyses:</b> Linear regression (hierarchical model) <b>Age Groups Analyzed:</b> 61 – 89 years <b>Sample Description:</b> 28 subjects living at or near an apartment complex located on the same street at the Harvard School of Public Health	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> ppm 25th = 0.38; 75th = 0.54 <b>Copollutant: correlation</b> PM <sub>2.5</sub> : r = 0.61 NO <sub>2</sub> : r = 0.55 SO <sub>2</sub> : r = -0.18 O <sub>3</sub> : r = 0.21	<b>Increment:</b> 0.16 ppm <b>% Change in HRV [Lower CI, Upper CI]</b> Lags examined : 24-h, 1-h Lag 1-h: SDNN : -2.6 (-5.6 to 0.5); rMSSD : -3.9 (-10.6 to 3.3); PNN50 : -3.5 (-13.7 to 8.0); LF :HF : 4.5 (-1.2 to 10.5) Lag 24-h: SDNN : -4.2 (-0.6 to -7.7); rMSSD : -10.2 (-2.4 to -17.4); PNN50 : -14.8 (-3.0 to -25.2); LF :HF : 6.2 (-0.6 to 13.4)

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Tarkiainen et al. (2003)</p> <p><b>Period of Study:</b> October 1997 – May 1998</p> <p><b>Location:</b> Kuopio, Finland</p>	<p><b>Health Outcome:</b> Various measures of HRV via Ambulatory ECG (Holter system)</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> ANOVA for repeated errors (GLM)</p> <p><b>Age Groups Analyzed:</b> Age 55 – 68 years</p> <p><b>Sample Description:</b> 6 male patients with angiographically verified CAD</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 4.6 ppm (maximum of CO episode) (personal monitoring)</p> <p><b>Range (Min, Max):</b> 0.5, 27.4 (maximum of CO episode)</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>RR Estimate [Lower CI, Upper CI]</b></p> <p>Lags examined : 5 minute prior to CO episode, 5 minute during CO episode</p> <p>CO had no statically significant effect on NN, SDNN or rMSSD. However, during high CO exposure (&gt;2.7 ppm) CO was associated with an increase in rMSSD of 2.4ms (p=0.034).</p>
<p><b>Author:</b> Timonen et al. (2006)</p> <p><b>Period of Study:</b> 1998 - 1999</p> <p><b>Location:</b> 3 Cities in Europe: Amsterdam, Netherlands; Erfert, Germany; Helsinki, Finland</p>	<p><b>Health Outcome:</b> Stable CAD: Various measures of HRV via ambulatory ECG (Holter system)</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> Linear regression (mixed model)</p> <p><b>Age Groups Analyzed:</b> Mean age across 3 cities; 64 – 71 years.</p> <p><b>Sample Description:</b> 131 subjects with Stable CAD followed for 6 months with bi-weekly clinical visits.</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> Amsterdam: 0.6 mg/m<sup>3</sup> Erfert: 0.4 mg/m<sup>3</sup> Helsinki: 0.4 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> Amsterdam: 0.4, 1.6 Erfert: 0.1, 2.5 Helsinki: 0.1, 1.0</p> <p><b>Copollutant: correlation</b> Amsterdam: PM<sub>2.5</sub>: r = 0.58 NO<sub>2</sub>: r = 0.76 Erfert: PM<sub>10</sub>: r = 0.77 NO<sub>2</sub>: r = 0.86 Helsinki: PM<sub>10</sub>: r = 0.40 NO<sub>2</sub>: r = 0.32</p>	<p><b>Increment:</b> 1 mg/m<sup>3</sup></p> <p><b>Regression co-efficient [Lower CI, Upper CI]</b></p> <p>Lags examined (days): 0, 1, 2, 3, 5-day avg</p> <p><b>SDNN:</b> Lag 0 : -1.21 (-4.44 to 2.03); Lag 1 : -1.71 (-6.05 to 2.63); Lag 2 : -5.69 (-10.7 to -0.72); Lag 3 : 0.66 (-3.83 to 5.15); 5-day avg: -3.60 (-9.88 to 2.68)</p> <p><b>HF:</b> Lag 0 : 5.0 (-15.1 to 25.1); Lag 1 : -2.0 (-37.1 to 33.1); Lag 2 : -30.7 (-59.8 to -1.5); Lag 3 : -9.3 (-35.8 to -17.3); 5-day avg: -15.2 (-53.0 to 22.6)</p> <p><b>LF/HF:</b> Lag 0 : -3.6 (-21.8 to 14.5); Lag 1 : -28.6 (-52.0 to -5.3); Lag 2 : -10.1 (-36.9 to 16.7); Lag 3 : 7.7 (-16.5 to 31.9); 5-day avg: -16.9 (-51.2 to 17.3)</p>
<p><b>Author:</b> Wheeler et al. (2006)</p> <p><b>Period of Study:</b> 1999 - 2000</p> <p><b>Location:</b> Atlanta, GA</p>	<p><b>Health Outcome:</b> Various measures of HRV via Holter system</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> Linear regression (mixed effects models)</p> <p><b>Age Groups Analyzed:</b> Mean 65 years – IQR 55- 73 years.</p> <p><b>Sample Description:</b> 18 subjects with COPD and 12 subjects with recent MI.</p>	<p><b>Averaging Time:</b> 1-h</p> <p><b>Mean (SD) unit:</b> 362.0 ppb</p> <p><b>Range (Min, Max):</b> 25th = 221.5; 75th = 398.1</p> <p><b>Copollutant: correlation</b> PM<sub>2.5</sub>: r = 0.43</p>	<p><b>Increment:</b> NR</p> <p><b>RR Estimate [Lower CI, Upper CI] ; lag :</b></p> <p>Lags examined (h ma): 1, 4, 24</p> <p>No CO results reported.</p>
<b>ONSET OF CARDIAC ARRHYTHMIA</b>			
<p><b>Author:</b> Berger et al. (2006)</p> <p><b>Period of Study:</b> October 2000 – April 2001</p> <p><b>Location:</b> Erfurt, Germany</p>	<p><b>Health Outcome:</b> Runs of supraventricular and ventricular tachycardia recorded via 24-h ECG.</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> Poisson regression (GAM) Linear regression</p> <p><b>Age Groups Analyzed:</b> 52 – 76 years (mean 76years)</p> <p><b>Sample Description:</b> 57 men with CHD</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 0.52 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> 0.11, 1.93</p> <p><b>Copollutant: correlation</b> NR</p>	<p><b>Increment:</b></p> <p>All: 0.27 mg/m<sup>3</sup> 5-day avg : 0.22 mg/m<sup>3</sup></p> <p><b>RR Estimate [Lower CI, Upper CI]</b></p> <p>Lags examined (h): 0, 0-23, 24-47, 48-71, 72-95, 5-day avg</p> <p>Supraventricular extrasystoles: Lag 0 : 1.18 (1.00-1.38) Lag 0-23 : 1.16 (1.02-1.31); Lag 24-47 : 1.13 (1.00-1.28); Lag 48-71 : 1.18 (1.03-1.36); Lag 72-95 : 1.08 (0.98-1.20); 5-day avg: 1.18 (1.04-1.35)</p> <p><b>Mean % Change [Lower CI, Upper CI]</b></p> <p>Hourly Lags examined: 0, 0-23, 24-47, 48-71, 72-95, 5-day avg</p> <p>Ventricular extrasystoles: Lag 0 : 0.0 (-4.1 to 4.4); Lag 0-23 : 1.1 (-3.3 to 5.7); Lag 24-47 : 1.9 (-2.6 to 6.6); Lag 48-71 : 4.2 (-0.3 to 8.9); Lag 72-95 : 2.7 (-1.3 to 6.9); 5-day avg: 3.0 (-1.8 to 8.0)</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Dockery et al. (2005)</p> <p><b>Period of Study:</b> 1995 - 2002</p> <p><b>Location:</b> Boston, MA</p>	<p><b>Health Outcome:</b> Tachyarrhythmias:</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> Logistic regression (GEE)</p> <p><b>Age Groups Analyzed:</b> 19 – 90 years; mean age 64 years</p> <p><b>Sample Description:</b> 203 cardiac patients with ICDs within 40km of air monitoring site at Harvard School of Public Health, Boston</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> 25th = 0.53; 75th = 1.02</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 0.48 ppm</p> <p><b>OR for Ventricular Arrhythmia [Lower CI, Upper CI]</b></p> <p>Lags examined (days): 0, 1, 2, 3</p> <p>Lag 2day ma: 1.14 (0.95-1.29)</p> <p>Among those who had an Arrhythmia – within 3 days : 1.65 (1.17-2.33)</p> <p>later than 3 days : 1.04 (0.83-1.29)</p>
<p><b>Author:</b> Metzger et al. (2007)</p> <p><b>Period of Study:</b> 1993 - 2002</p> <p><b>Location:</b> Atlanta, GA</p>	<p><b>Health Outcome:</b> Cardiac Arrhythmia, ICD, Ventricular tachyarrhythmia</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> Logistic regression (GEE)</p> <p><b>Age Groups Analyzed:</b> 15 – 88 years</p> <p><b>Sample Description:</b> 518 patients with ICDs with at least one ventricular tachyarrhythmic event</p>	<p><b>Averaging Time:</b> 1-h</p> <p><b>Mean (SD) unit:</b> 1.7 ppm</p> <p><b>Range (Min, Max):</b> 0.1, 7.7</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>OR for Tachyarrhythmic event [Lower CI, Upper CI]</b></p> <p>Lags examined (days) : 0</p> <p><b>Results for All events</b></p> <p>Lag 0 : 0.999 (0.970-1.028)</p> <p><b>Events resulting in cardiac pacing or defibrillation</b></p> <p>Lag 0 : 1.008 (0.964-1.054)</p> <p><b>Events resulting defibrillation</b></p> <p>Lag 0 : 1.012 (0.925-1.10.7)</p>
<p><b>Author:</b> Peters et al. (2000b)</p> <p><b>Period of Study:</b> 1995 - 1997</p> <p><b>Location:</b> Eastern Massachusetts</p>	<p><b>Health Outcome:</b> Defibrillated discharges for ventricular tachycardia or fibrillation</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> Mean age of 62 years</p> <p><b>Sample Description:</b> 100 patients with ICDs</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 0.58 ppm</p> <p><b>Range (Min, Max):</b> 25th = 0.43; 75th = 0.66</p> <p><b>Copollutant: correlation</b></p> <p>PM<sub>10</sub>: r = 0.51</p> <p>PM<sub>2.5</sub>: r = 0.56</p> <p>NO<sub>2</sub>: r = 0.71</p> <p>SO<sub>2</sub>: r = 0.41</p> <p>O<sub>3</sub>: r = -0.40</p>	<p><b>Increment:</b> 0.65 ppm (Lags 0, 1, 2, 3); 0.42 ppm (Lag 5-day mean)</p> <p><b>OR for Defibrillated Discharge [Lower CI, Upper CI]</b></p> <p>Lags examined (days): 0, 1, 2, 3, 5-day mean</p> <p><b>At least one discharge:</b></p> <p>Lag 0: 1.07 (0.62-1.86); Lag 1 : 1.06 (0.61-1.85); Lag 2: 1.05 (0.62-1.77); Lag 3 : 0.09 (0.65-1.83); Lag 5-day mean : 1.23 (0.71- 2.12)</p> <p><b>At least 10 discharges:</b></p> <p>Lag 0: 1.12 (0.54-2.32); Lag 1 : 1.13 (0.54-2.33); Lag 2: 1.62 (0.85-3.09); Lag 3 : 1.98 (1.05-3.72); Lag 5-day mean : 1.94 (1.01-.75)</p>
<p><b>Author:</b> Rich et al. (Rich et al.)</p> <p><b>Period of Study:</b> February – December 2000</p> <p><b>Location:</b> Vancouver, Canada</p>	<p><b>Health Outcome:</b> Cardiac arrhythmia via patients ICD</p> <p><b>Study Design:</b> Case-crossover</p> <p><b>Statistical Analyses:</b> Conditional Logistic regression</p> <p><b>Age Groups Analyzed:</b> 15 – 85 years</p> <p><b>Sample Description:</b> 34 patients who experienced at least 1 ICD discharge (8201 person days)</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 553.8 ppb</p> <p><b>Range (Min, Max):</b> IQR: 162.7</p> <p><b>Copollutant: correlation</b></p> <p>PM<sub>10</sub>: r = 0.40</p> <p>SO<sub>2</sub>: r = 0.75</p> <p>NO<sub>2</sub>: r = 0.68</p> <p>O<sub>3</sub>: r = -0.56</p>	<p><b>Increment:</b> NR</p> <p><b>RR Estimate [Lower CI, Upper CI]</b></p> <p>Lags examined (days): 0, 1, 2, 3</p> <p>No significant effect (results not reported in table).</p>
<p><b>Author:</b> Rich et al. (2005)</p> <p><b>Period of Study:</b> 1995 - 1999</p> <p><b>Location:</b> Boston, MA</p>	<p><b>Health Outcome:</b> Ventricular arrhythmias via ICD</p> <p><b>Study Design:</b> Panel/Case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> All</p> <p><b>Sample Description:</b> 203 patients with implanted ICD at the New England Medical Center</p>	<p><b>Averaging Time:</b> 1-h &amp; 24-h</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (percentiles):</b></p> <p>1-h:</p> <p>25th = 0.46</p> <p>75th = 1.04</p> <p>24-h:</p> <p>25th = 0.52</p> <p>75th = 1.03</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 0.56 ppm; 0.54; 0.51; 0.49 respectively for results shown below</p> <p><b>OR Estimate [Lower CI, Upper CI]</b></p> <p>Ventricular Arrhythmia</p> <p>Hours prior to event :</p> <p>0-2 : 1.01 (0.87-1.18)</p> <p>0-6 : 1.00 (0.85-1.17)</p> <p>0-23 : 1.03 (0.84-1.25)</p> <p>0-47 : 1.11 (0.88-1.40)</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Rich et al. (2006b) <b>Period of Study:</b> 2001 & 2002 <b>Location:</b> St. Louis, MO	<b>Health Outcome:</b> Ventricular arrhythmia <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 60 subjects with at least 1 ICD recorded arrhythmia who lived within 40 km of St. Louis – Midwest supersite.	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> 25th = 0.4; 75th = 0.6 <b>Copollutant:</b> NR	<b>Increment:</b> 0.2 ppm <b>OR for Ventricular Arrhythmia [Lower CI, Upper CI]</b> Lags examined : 0-23 h-ma 0-23h-ma : 0.99 (0.80-1.21)
<b>Author:</b> Rich et al. (2006a) <b>Period of Study:</b> 1995 - 1999 <b>Location:</b> Boston, MA	<b>Health Outcome:</b> ICD Episode of Atrial fibrillation <b>Study Design:</b> Panel/case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 203 patients with ICDs at the New England Medical Center	<b>Averaging Time:</b> 1-h & 24-h <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> 1-h: 25th = 0.46; 75th = 1.04 24-h: 25th = 0.52; 75th = 1.03 <b>Copollutant:</b> NR	<b>Increment:</b> Lag (hrs) 0 : 0.58 ppm Lag (hrs) 0-23 : 0.51 ppm <b>OR for Episode of Atrial Fibrillation [Lower CI, Upper CI]</b> Lags (h) : 0, 0-23 Lag 0 : 0.87 (0.56-1.37) Lag 0-23 : 0.71 (0.39-1.28)
<b>Author:</b> Sarnat et al. (2006) <b>Period of Study:</b> 24 weeks during the Summer and Fall of 2000 <b>Location:</b> Steubenville, OH	<b>Health Outcome:</b> Arrhythmia via ECG measurements <b>Study Design:</b> Panel <b>Statistical Analyses:</b> Logistic regression <b>Age Groups Analyzed:</b> 53-90 years (mean age 71) <b>Sample Description:</b> 32 non-smoking older adults	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.02 ppm <b>Range (Min, Max):</b> -0.1, 1.5 <b>Copollutant: correlation</b> PM <sub>2.5</sub> : r = 0.45 SO <sub>2</sub> : r = 0.62 NO <sub>2</sub> : r = 0.66 O <sub>3</sub> : r = -0.37	<b>Increment:</b> 0.2 ppm <b>RR Estimate [Lower CI, Upper CI] ; lag :</b> Lags examined (days): 1, 2, 3, 4, 5, 5-day ma Lag 5-day ma : Supraventricular Ectopy SVE : 0.99 (0.76-1.29) Ventricular Ectopy VE : 1.05 (0.75-1.46)
<b>Author:</b> Vedal et al. (2004) <b>Period of Study:</b> 1997 - 2000 <b>Location:</b> Vancouver, Canada	<b>Health Outcome:</b> Cardiac arrhythmia via patients with ICD <b>Study Design:</b> Panel <b>Statistical Analyses:</b> Logistic regression (GEE) <b>Age Groups Analyzed:</b> Range from 12-77 (mean age 53) <b>Sample Description:</b> 50 patients who experienced 1 or more arrhythmia event days during the four years	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.6 ppm <b>Range (Min, Max):</b> 0.3, 1.6 <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.43 SO <sub>2</sub> : r = 0.62 NO <sub>2</sub> : r = 0.74 O <sub>3</sub> : r = -0.52	<b>Increment:</b> 0.2 ppm <b>RR Estimate [Lower CI, Upper CI]</b> Lags examined (days) : 0, 1, 2, 3 No significant effect for CO (results shown in plots)
<b>CARDIAC ARREST</b>			
<b>Author:</b> Levy et al. (2001) <b>Period of Study:</b> 1988 - 1994 <b>Location:</b> Seattle, WA	<b>Health Outcome:</b> Out of hospital primary cardiac arrest <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> 25-75 years <b>Sample Description:</b> 362 cases	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 1.79 ppm <b>Range (Min, Max):</b> 0.52, 5.92 <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.81 SO <sub>2</sub> : r = 0.29	<b>Increment:</b> NR <b>RR Estimate [Lower CI, Upper CI]</b> Lags examined (days): 0, 1 Lag 1 : 0.99 (0.83-1.18)
<b>Author:</b> Sullivan et al. (2003) <b>Period of Study:</b> 1985 - 1994 <b>Location:</b> Washington State	<b>Health Outcome:</b> Out of Hospital Cardiac Arrest. <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 1,542 members of a large health maintenance organization	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 1.92 ppm <b>Range (Min, Max):</b> 0.52, 7.21 <b>Copollutant:</b> NR	<b>Increment:</b> 1.02 ppm <b>OR Estimate [Lower CI, Upper CI]</b> Lags examined (days) : 0, 1, 2 Lag 0 : 0.95 (0.85-1.05) Lag 1 : 0.97 (0.87-1.08) Lag 2 : 0.99 (0.89-1.11)



Study	Design	Concentrations	CO Effect Estimates (95% CI)
<i>MYOCARDIAL INFARCTION</i>			
<b>Author:</b> Peters et al. (2001) <b>Period of Study:</b> 1995 - 1996 <b>Location:</b> Boston, MA	<b>Health Outcome:</b> Onset of MI: <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 772 participants	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 1.09 <b>Range (percentiles):</b> ppm 5th = 0.49 95th = 1.78 <b>Copollutant:</b> NR	<b>Increment:</b> 2 H – 1 ppm; 24 h – 0.6 ppm <b>OR Estimate [Lower CI, Upper CI]</b> Onset of MI: 2 h prior : 1.22 (0.89-1.67) 24 h prior : 0.98 (0.70-1.36)
<b>Author:</b> Rosenlund et al. (2006) <b>Period of Study:</b> 1992 - 1994 <b>Location:</b> Stockholm, Sweden	<b>Health Outcome:</b> MI <b>Study Design:</b> Case-control <b>Statistical Analyses:</b> Logistic regression <b>Age Groups Analyzed:</b> 45 – 70 years <b>Sample Description:</b> 1,397 cases, 1,870 controls	<b>Averaging Time:</b> <b>Mean (SD) unit:</b> 66.8 µg/m <sup>3</sup> (est. 30yr residential exposure) <b>Range (percentiles):</b> 5th = 13.9; 95th = 295.7 <b>Copollutant:</b> NR	<b>Increment:</b> 300 µg/m <sup>3</sup> <b>OR Estimate [Lower CI, Upper CI] ; lag :</b> Estimated 30 yr avg. exposure All cases : 1.04 (0.89-1.21) Non-fatal cases : 0.98 (0.82-1.16) Fatal cases : 1.22 (0.98-1.52) In-hospital death : 1.16 (0.89-1.51) Out-of-hospital death : 1.36 (1.01-1.84)
<i>CHANGES IN BLOOD PRESSURE</i>			
<b>Author:</b> Ibalde-Mulli et al. (2001) <b>Period of Study:</b> 1984 - 1985 <b>Location:</b> Augsburg, Germany	<b>Health Outcome:</b> BP - SPB <b>Study Design:</b> Cohort <b>Statistical Analyses:</b> Gaussian regression for repeated measures <b>Age Groups Analyzed:</b> 25-64 years <b>Sample Description:</b> 2,607 men & women aged 25-64 years	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 4.1 mg/m <sup>3</sup> <b>Range (Min, Max):</b> 1.7, 8.2 <b>Copollutant:</b> NR	<b>Increment:</b> Lag 0 : 5.6 mg/m <sup>3</sup> 5-day prior avg. <b>Mean Change [Lower CI, Upper CI]</b> SPB mmHg Lag 0 (days): All : 0.53 (-0.66 to 1.72); Men : 0.68 (-0.94 to 2.31); Women : 0.51 (-1.31 to 2.19) 5-day prior avg: All : 1.06 (-0.17 to 2.29); Men : 0.92 (-0.87 to 2.70); Women : 0.91 (-0.87 to 2.70)
<b>Author:</b> Zanobetti et al. (2004a) <b>Period of Study:</b> 1999 - 2001 <b>Location:</b> Boston, MA	<b>Health Outcome:</b> BP <b>Study Design:</b> Cohort/Panel <b>Statistical Analyses:</b> Random effects <b>Age Groups Analyzed:</b> 39 – 90 years <b>Sample Description:</b> 62 subjects with 631 total visits	<b>Averaging Time:</b> 1-h & 120-h avg <b>Mean (SD) unit:</b> Same Hr: 0.81 ppm 120 Hr av: 0.66 ppm <b>Range (Min, Max):</b> Same h: 10th = 0.48; 90th = 1.22 120-h av: 10th = 0.48; 90th = 0.86 <b>Copollutant:</b> NR	<b>Increment:</b> NR <b>RR Estimate [Lower CI, Upper CI]</b> CO had no significant effect on BP
<i>CHANGES IN BLOOD MARKERS OF COAGULATION AND INFLAMMATION</i>			
<b>Author:</b> Baccarelli et al. (2007) <b>Period of Study:</b> 1995 - 2005 <b>Location:</b> Milan, Italy	<b>Health Outcome:</b> Prothrombin time (PT) and Activated partial thromboplastin time (APTT) <b>Study Design:</b> Panel <b>Statistical Analyses:</b> GAMS <b>Age Groups Analyzed:</b> 11-84 years (mean 43years) <b>Sample Description:</b> 1,218 healthy individuals who were partners or friends of patients with thrombosis who attended the thrombosis center of the University of Milan.	<b>Averaging Time:</b> 1-h <b>Mean (SD) unit:</b> NR <b>Range (percentiles):</b> Sept-Nov: 25th = 1.36; 75th = 3.52 Dec-Feb: 25th = 2.00; 75th = 4.31 Mar-May: 25th = 1.03; 75th = 2.14 Jun-Aug: 25th = 0.73; 75th = 1.58 <b>Copollutant:</b> NR	<b>Increment:</b> NR <b>Regression co-efficient [Lower CI, Upper CI]</b> Lags examined (time of blood sampling – avg): 0, 7, 30 <b>PT:</b> Lag 0 : -0.11 (-0.18 to -0.05); Lag 7 : -0.07 (-0.14 to 0.01); Lag 30 : -0.05 (-0.13 to 0.02) <b>APTT:</b> Lag 0 : 0.03 (-0.04 to 0.10); Lag 7 : 0.04 (-0.04 to 0.11); Lag 30 : 0.06 (-0.01 to 0.14) <b>Notes:</b> CO had no effect on fibrinogen, functional antithrombin, functional protein C, protein C antigen, functional protein S, free protein S for all lag periods.
<b>Author:</b> Liao et al. (2005) <b>Period of Study:</b> 1996 to 1998 <b>Location:</b> Forsyth County, NC; Selected suburbs of Minneapolis, MN; Jackson, MI	<b>Health Outcome:</b> Various measures of hemostasis/ inflammation <b>Study Design:</b> Cohort <b>Statistical Analyses:</b> Linear regression <b>Age Groups Analyzed:</b> 45 – 64 years <b>Sample Description:</b> 10208 subjects from the Atherosclerosis Risk in Communities Study	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant:</b> NR	<b>Increment:</b> 0.6 ppm <b>Regression coefficients [SE]</b> Lags examined (days): 1 Lag 1: Fibrinogen (mg/dL) : -0.16 (0.67) Factor VIII –C (%) : 0.45 (0.42) vWF % : -0.29 (0.50) WBC (x 10 <sup>3</sup> /mm <sup>3</sup> ) : 0.003 (0.017) Albumin (g/dL) : -0.018 (0.003)** ** p < 0.01

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Pekkanen et al. (2000)</p> <p><b>Period of Study:</b> 1991 - 1993</p> <p><b>Location:</b> London, England</p>	<p><b>Health Outcome:</b> Fibrinogen</p> <p><b>Study Design:</b> Cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> 35 – 55 years</p> <p><b>Sample Description:</b> 7,205 office workers</p>	<p><b>Averaging Time:</b> 8-h</p> <p><b>Mean (SD) unit:</b> 1.4 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> Min = NR, Max = 9.9</p> <p><b>Copollutant correlation:</b> PM<sub>10</sub>: r = 0.57 NO<sub>2</sub>: r = 0.81 SO<sub>2</sub>: r = 0.61 O<sub>3</sub>: r = -0.45</p>	<p><b>Increment:</b> 1.6 mg/m<sup>3</sup></p> <p><b>% Change in Fibrinogen Concentration [p value] ;</b> Lags examined : 0, 1, 2, 3 Lag 0 : 1.43 (&lt;0.01); Lag 1 : 1.49 (&lt;0.01); Lag 2 : 1.59 (&lt;0.01); Lag 3 : 1.26 (&lt;0.01)</p> <p><b>OR for having Fibrinogen above 3.19 g/l [p value]</b> Lags examined : 0, 1, 2, 3 Lag 0 : 1.17 (0.05); Lag 1 : 1.09 (0.31); Lag 2 : 1.14 (0.11); Lag 3 : 1.22 (&lt;0.01)</p>
<p><b>Author:</b> Ruckerl et al. (2006)</p> <p><b>Period of Study:</b> 2000 - 2001</p> <p><b>Location:</b> Erfert, Germany</p>	<p><b>Health Outcome:</b> Blood markers of inflammation and coagulation</p> <p><b>Study Design:</b> Panel</p> <p><b>Statistical Analyses:</b> Linear and logistic regression (fixed effects)</p> <p><b>Age Groups Analyzed:</b> 51-76 years (mean age 66 years)</p> <p><b>Sample Description:</b> 57 male patients with CHD</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 0.52 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> 0.11, 1.93</p> <p><b>Copollutant correlation:</b> NO<sub>2</sub>: r = 0.82</p>	<p><b>Increment:</b> 0.27 mg/m<sup>3</sup></p> <p><b>OR Estimate for blood marker &gt;90th percentile [Lower CI, Upper CI]</b> Lags examined (h): 0-23, 24-47, 48-71, 5-day avg.</p> <p><b>CRP (C-reactive protein)</b> 0-23 : 0.9 (0.7-1.2); 24-47 : 1.0 (0.7-1.5); 48-71 : 1.5 (1.1-2.1); 5-day avg 1.1 (0.8-1.6)</p> <p><b>ICAM-1 (Intercellular adhesion molecule 1)</b> 0-23 : 0.8 (0.6-1.0); 24-47 : 1.5 (1.2-1.9); 48-71 : 1.7 (1.3-2.3); 5-day avg 1.2 (1.0-1.6)</p> <p><b>% of change from the mean of blood marker</b></p> <p><b>vWF (von Willebrand factor antigen)</b> 0-23 : 4.4 (1.4- 7.5); 24-47 : 2.7 (-0.8 to 6.1); 48-71 : 2.0 (-1.7 to 5.8); 5-day avg : 4.9 (1.0-8.8)</p> <p><b>FVII (Factor VII)</b> 0-23 : -1.4 (-3.8 to 1.1); 24-47 : -2.6 (-4.8 to 0.3); 48-71 : -2.8 (-5.1 to -0.4); 5-day avg : -3.0 (-5.5 to -0.4)</p>
<p><b>Author:</b> Ruckerl et al. (2007)</p> <p><b>Period of Study:</b> May 2003-July 2004</p> <p><b>Location:</b> 6 cities across Europe: Athens, Greece; Augsburg, Germany; Barcelona, Spain; Helsinki, Finland; Rome, Italy; Stockholm, Sweden</p>	<p><b>Health Outcome:</b> Interleukin-6, C-reactive protein, Fibrinogen</p> <p><b>Study Design:</b> Panel/Cohort</p> <p><b>Statistical Analyses:</b> Linear regression (mixed effects)</p> <p><b>Age Groups Analyzed:</b> 37-81 years</p> <p><b>Sample Description:</b> 1,003 MI survivors who had at least 2 valid repeated blood samples</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> Athens: 1.48 mg/m<sup>3</sup> Augsburg: 0.58 mg/m<sup>3</sup> Barcelona: 0.59 mg/m<sup>3</sup> Helsinki: 0.31 mg/m<sup>3</sup> Rome: 1.40 mg/m<sup>3</sup> Stockholm: 0.29 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 0.34 mg/m<sup>3</sup></p> <p><b>% Change in mean [Lower CI, Upper CI]</b> Lags examined: 0, 1, 2, 5-day avg (Pooled estimates)</p> <p><b>Interleukin-6</b> Lag 0 : 0.57 (-0.63 to 1.79); Lag 1 : 0.44 (-0.79 to 1.68); Lag 2 : -2.36 (-4.82 to 0.17); 5-day avg: -0.28 (-2.53 to 2.02)</p> <p><b>C-reactive protein</b> Lag 0 : -0.01 (-1.72 to 1.73); Lag 1 : -1.51 (-3.30 to 0.32); Lag 2 : -2.35 (-6.84 to 2.36); 5-day avg : -0.85 (-5.37 to 3.90)</p> <p><b>Fibrinogen</b> Lag 0 : 0.24 (-0.54 to 0.92); Lag 1 : 0.32 (-0.35 to 1.00); Lag 2 : -0.44 (-1.11 to 0.23); 5-day avg : 0.12 (-0.81 to 1.05)</p>
<p><b>Author:</b> Steinvil et al. (2008)</p> <p><b>Period of Study:</b> 2003 - 2006</p> <p><b>Location:</b> Tel Aviv, Israel</p>	<p><b>Health Outcome:</b> Various measures of inflammation sensitive biomarkers</p> <p><b>Study Design:</b> Cohort</p> <p><b>Statistical Analyses:</b> Linear regression</p> <p><b>Age Groups Analyzed:</b> mean age 46 years</p> <p><b>Sample Description:</b> 3659 subjects living within 11 km of monitoring site</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 0.8 ppm</p> <p><b>Range (percentiles):</b> 25th = 0.7; 75th = 1.0</p> <p><b>Copollutant correlation</b> PM<sub>10</sub>: r = 0.75 NO<sub>2</sub>: r = 0.857 SO<sub>2</sub>: r = 0.671 O<sub>3</sub>: r = -0.656</p>	<p><b>Increment:</b> 0.3 ppm</p> <p><b>Regression co-efficient [Lower CI, Upper CI]</b> Lags examined (days): 0, 1, 2, 3, 4, 5, 6, 7, last week avg</p> <p><b>Fibrinogen – Men</b> Lag 0 : -3.3 (-6.1 to -0.6); Lag 1 : -2.6 (-5.5 to 0.4); Lag 2 : -3.4 (-6.6 to -0.3); Lag 3 : -3.4 (-6.5 to -0.2); Lag 4 : -5.9 (-8.9 to -2.9); Lag 5 : -4.7 (-7.8 to -1.6); Lag 6 : -2.0 (-5.1 to 1.0); Lag 7 : -2.7 (-5.7 to 0.2); Last week avg: -7.7 (-12.1 to -3.3)</p> <p><b>Notes:</b> No effect on fibrinogen among women. CO had no effect on CRP among men and no effect on CRP and WBC among women for all Lag times examined.</p>

**Table C-2. Studies of CO exposure and cardiovascular hospital admissions and ED visits.**

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<i>STROKE</i>			
<p><b>Author:</b> Chan et al. (2006)  <b>Period of Study:</b> 1997 - 2002  <b>Location:</b> Taipei, Taiwan</p>	<p><b>ED Visits</b>  <b>Health Outcome (ICD9):</b> Cerebrovascular disease (430-437); Strokes (430-434); Hemorrhagic stroke (430-432); Ischemic stroke (433-434)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> GAM  <b>Age Groups Analyzed:</b> All  <b>Sample Description:</b> NR</p>	<p><b>Averaging Time:</b> 8-h  <b>Mean (SD) unit:</b> 1.7 ppm  <b>Range (Min, Max):</b> 0.6, 4.4  <b>Copollutant: correlation</b>  O<sub>3</sub>: r = 0.30  SO<sub>2</sub>: r = 0.63  NO<sub>2</sub>: r = 0.77  PM<sub>2.5</sub>: r = 0.44  PM<sub>10</sub>: r = 0.47</p>	<p><b>Increment:</b> 0.8 ppm  <b>OR Estimate [Lower CI, Upper CI]</b>  Lags (days) examined 0, 1, 2, 3  Cerebrovascular disease: Lag 2, 1.03 (1.01, 1.06)  Stroke : Lag 2, 1.03 (1.01, 1.05)  Ischemic and Hemorrhagic stroke : not significant.  Cerebrovascular 2 pollutant model:  CO + O<sub>3</sub> : Lag 2, 1.03 (1.01-1.05)  CO + PM<sub>2.5</sub> : Lag 2, 1.02 (1.00-1.04)  CO + PM<sub>10</sub> : Lag 2, 1.03 (1.01-1.05)</p>
<p><b>Author:</b> Henrotin et al. (2007)  <b>Period of Study:</b> 1994 - 2004  <b>Location:</b> Dijon, France</p>	<p><b>Health Outcome (ICD9 or ICD10):</b> Stroke (Ischemic &amp; Hemorrhagic)  <b>Study Design:</b> Bi-directional Case-crossover  <b>Statistical Analyses:</b> Conditional logistic regression  <b>Age Groups Analyzed:</b> ≥ 40  <b>Sample Description:</b> NR</p>	<p><b>Averaging Time:</b> 24-h  <b>Mean (SD) unit:</b> 683 µg/m<sup>3</sup>  <b>Range (Min, Max):</b> 0, 4014  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 10 µg/m<sup>3</sup>  <b>OR Estimate [Lower CI, Upper CI]</b>  Lags (days) examined: 0, 1, 2, 3.  <b>Ischemic:</b>  Lag 0 : 0.999 (0.997-1.001)  Lag 1 : 0.998 (0.997-1.001)  Lag 2 : 0.999 (0.998-1.001)  Lag 3 : 1.000 (0.998-1.001)  <b>Hemorrhagic:</b>  Lag 0 : 1.000 (0.996-1.004)  Lag 1 : 1.001 (0.997-1.005)  Lag 2 : 0.999 (0.995-1.004)  Lag 3 : 0.998 (0.994-1.002)  Also not significant when stratified by sex.</p>
<p><b>Author:</b> Maheswaran et al. (2005b)  <b>Period of Study:</b> 1994 - 1998  <b>Location:</b> Sheffield, UK</p>	<p><b>Health Outcome (ICD9 or ICD10):</b> Stroke deaths (ICD9: 430-438); Stroke Hospital admissions (ICD10: I60-I69)  <b>Study Design:</b> Ecological  <b>Statistical Analyses:</b> Poisson regression  <b>Age Groups Analyzed:</b> ≥ 45 years  <b>Sample Description:</b> 1,030 census districts</p>	<p><b>Averaging Time:</b> NR  <b>Mean (SD) unit:</b> Quintiles  <b>Range (Min, Max):</b> NR  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR – Quintiles of exposure  <b>RR Estimate [Lower CI, Upper CI]</b>  Adjusted for sex, age, deprivation, smoking.  Quintiles:  2nd : 1.04 (0.94-1.16)  3rd : 1.01 (0.91-1.13)  4th : 1.10 (0.99-1.23)  5th : 1.11 (0.99-1.25)  Adjusted for sex, age:  2nd : 1.11 (1.01-1.22)  3rd : 1.15 (1.04-1.27)  4th : 1.29 (1.17-1.42)  5th : 1.37 (1.24-1.52)</p>
<p><b>Author:</b> Tsai et al. (2003b)  <b>Period of Study:</b> 1997-2000  <b>Location:</b> Kaohsiung, Taiwan</p>	<p><b>Study Design:</b> Case-crossover  <b>Health Outcome (ICD9 or ICD10):</b> Cerebrovascular diseases: ICD9: 430 to 438 (Subarachnoid hemorrhagic stroke 430, Primary intracerebral hemorrhage (PIH): 431-432, Ischemic stroke (IS): 433-435).  <b>Statistical Analyses:</b> NR  <b>Age Groups Analyzed:</b> All  <b>Sample Description:</b> NR</p>	<p><b>Averaging Time:</b> 24-h  <b>Mean (SD) unit:</b> 0.79 ppm  <b>Range (Min, Max):</b> 0.24, 1.72  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 0.8 ppm (IQR)  <b>RR Estimate [Lower CI, Upper CI]</b>  Lag (days): 0-2  &gt;20°C  PIH : OR 1.21 (1.09-1.34)  IS : OR 1.21 (1.14-1.28)  &lt;20°C  PIH : OR 1.18 (0.80-0.72)  IS : OR 1.77 (1.31-2.39)  <b>Notes:</b>  2-pollutant models:  PIH results persisted when adjusting for SO<sub>2</sub> and O<sub>3</sub>  IS results persisted when controlling for PM<sub>10</sub>, SO<sub>2</sub> and O<sub>3</sub></p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Villeneuve et al. (2006a) <b>Period of Study:</b> 1992-2002 <b>Location:</b> Edmonton, Canada	<b>ED Visits</b> (within 5 hospitals) <b>Health Outcome (ICD9):</b> Stroke (430-438); Ischemic (434-436) Hemorrhagic (430-432); Transient Ischemic Attack (435) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> 65+ <b>Sample Description:</b> 12,422 visits	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.8 ppm <b>Range (percentiles):</b> 25th = 0.5; 75th = 1.0 <b>Copollutant correlation :</b> O <sub>3</sub> : r = -0.54 PM <sub>2.5</sub> : r = 0.43 PM <sub>10</sub> : r = 0.30	<b>Increment:</b> 0.5 ppm <b>OR Estimate [Lower CI, Upper CI]</b> Lags (days) examined : 0, 1 & 0-2 Ischemic (April-Sept) Lag 0 : 1.16 (1.00, 1.33) Lag 1 : 1.17 (1.01, 1.36) Lag 0-2 : 1.32 (1.09, 1.60) <b>Notes:</b> - Not significant for all seasons or Oct-Mar. - Hemorrhagic : Not significant for all seasons or Oct-Mar, Apr-Sept. - Transient Ischemic Attack : Not significant for all seasons or Oct-Mar, Apr-Sept.
<b>Author:</b> Wellenius et al. (2005a) <b>Period of Study:</b> NR <b>Location:</b> 9 U.S. cities: Chicago, Detroit, Pittsburgh, Cleveland, Birmingham, New Haven, Seattle, Minneapolis, Salt Lake City	<b>ED Visits</b> <b>Health Outcome:</b> Stroke among Medicare beneficiaries: (Ischemic, hemorrhagic) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Logistic regression <b>Age Groups Analyzed:</b> ≥ 65 years <b>Sample Description:</b> 155,503 visits	<b>Averaging Time:</b> NR <b>Mean (SD) unit:</b> NR <b>Range (percentiles):</b> 25th = 0.73; 50th = 1.02; 75th = 1.44 (ppm) <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.43	<b>Increment:</b> 0.71 ppm <b>% Change [Lower CI, Upper CI]</b> Lag : 0 Ischemic : 2.83 (1.23-4.46) Hemorrhagic : -1.61 (-4.79 to 1.68)
<i>ISCHEMIC HEART DISEASE</i>			
<b>Author:</b> D'Ippoliti et al. (2003) <b>Period of Study:</b> 1995 - 1997 <b>Location:</b> Rome, Italy	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> MI (410) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> 18+ <b>Sample Description:</b> 6,531 patients.	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 4.4 mg/m <sup>3</sup> <b>Range (percentiles):</b> 25th = 2.8; 75th = 4.3 <b>Copollutant: correlation</b> TSP: r = 0.35 SO <sub>2</sub> : r = 0.56 NO <sub>2</sub> : r = 0.31	<b>Increment:</b> 1 mg/m <sup>3</sup> <b>OR Estimate [Lower CI, Upper CI] ; lag :</b> Lags examined (days): 0, 1, 2, 3, 4, 0-2 Acute MI Lag 0 : 1.021 (0.988-1.054) Lag 1 : 1.020 (0.988-1.054) Lag 2 : 1.033 (1.001-1.066) Lag 3 : 1.010 (0.982-1.040) Lag 4 : 1.025 (0.996-1.055) Lag 0-2 : 1.044 (1.000-.089)
<b>Author:</b> Hosseini et al. (2005) <b>Period of Study:</b> 1996 - 2001 <b>Location:</b> Tehran, Iran	<b>Health Outcome:</b> Angina Pectoris (ICD9: 413; ICD10: I20) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 10.8 mg/m <sup>3</sup> <b>Range (Min, Max):</b> 1.6, 57.8 <b>Copollutant:</b> NR	<b>Increment:</b> 1 mg/m <sup>3</sup> <b>RR Estimate [Lower CI, Upper CI]</b> Lags examined (days): 0, 1, 2, 3 Lag 1 : 1.00957 (1.00600-1.01315)

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Lanki et al. (2006)</p> <p><b>Period of Study:</b> 1994 - 2000</p> <p><b>Location:</b> 5 European cities: Augsburg, Germany Barcelona, Spain Helsinki, Finland Rome, Italy Stockholm, Sweden</p>	<p><b>Health Outcome:</b> First AMI (ICD9: 410; ICD10: I21, I22)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson regression (GAM)</p> <p><b>Age Groups Analyzed:</b> 35+ years</p> <p><b>Sample Description:</b> 26,854 Hospital Admissions</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Unit:</b> mg/m<sup>3</sup></p> <p><b>Range (percentiles):</b> Augsburg, Germany 25th = 0.7; 75th = 1.1 Barcelona, Spain 25th = 0.6; 75th = 1.4 Helsinki, Finland 25th = 0.3; 75th = 0.5 Rome, Italy 25th = 1.7; 75th = 2.9 Stockholm, Sweden 25th = 0.3; 75th = 0.5</p> <p><b>Copollutant: correlation</b> PM<sub>10</sub>: r = 0.21 – 0.56 NO<sub>2</sub>: r = 0.43 – 0.75 O<sub>3</sub>: r = -0.023 – 0.020</p>	<p><b>Increment:</b> 0.2 mg/m<sup>3</sup></p> <p><b>RR Estimate [Lower CI, Upper CI] ; lag :</b></p> <p>Lags examined : 0, 1, 2, 3</p> <p>All 5 cities: Lag 0 : 1.005 (1.000-1.010) Lag 1 : 1.002 (0.996-1.007) Lag 2 : 1.002 (0.997-1.007) Lag 3 : 0.998 (0.992-1.003)</p> <p>3 cities with Hospital Discharge Register(HDR): Lag 0 : 1.007 (1.001-1.012) Lag 1 : 1.002 (0.996-1.008) Lag 2 : 1.003 (0.998-1.009) Lag 3 : 1.004 (0.988-1.020)</p> <p>3 cities with HDR – ≤ 75years Fatal: Lag 0 : 1.027 (1.006-1.048) Lag 1 : 1.021 (1.000-1.042) Lag 2 : 1.018 (0.997-1.039) Lag 3 : 1.015 (0.994-1.037)</p> <p>Non-Fatal: Lag 0 : 1.001 (0.995-1.008) Lag 1 : 1.000 (0.994-1.007) Lag 2 : 1.004 (0.998-1.011) Lag 3 : 0.999 (0.992-1.006)</p> <p>3 cities with HDR – ≥ 75years Fatal: Lag 0 : 1.009 (0.992-1.006) Lag 1 : 1.001 (0.985-1.018) Lag 2 : 1.006 (0.990-1.023) Lag 3 : 1.000 (0.983-1.017)</p> <p>Non-Fatal: Lag 0 : 1.015 (1.004-1.086) Lag 1 : 1.006 (0.995-1.017) Lag 2 : 0.995 (0.983-1.006) Lag 3 : 0.998 (0.987-1.009)</p>
<p><b>Author:</b> Lee et al. (2003b)</p> <p><b>Period of Study:</b> 1997 - 1999</p> <p><b>Location:</b> Seoul, Korea</p>	<p><b>Study Design:</b> Time-series</p> <p><b>Health Outcome (ICD9 or ICD10):</b> Angina: ICD10: I20 AMI: ICD10: I21-I23 Other Acute IHDs: ICD10: I24</p> <p><b>Statistical Analyses:</b> Poisson regression, GAM</p> <p><b>Age Groups Analyzed:</b> 64 +</p> <p><b>Sample Description:</b> 822 days</p>	<p><b>Averaging Time:</b> Daily max</p> <p><b>Mean (SD) unit:</b> 1.8 ppm</p> <p><b>Range (percentiles):</b> 25th = 1.2 75th = 2.2</p> <p><b>Copollutant: correlation</b> PM<sub>20</sub>: 0.60 SO<sub>2</sub>: 0.81 NO<sub>2</sub>: 0.79 O<sub>3</sub>: -0.39</p>	<p><b>Increment:</b> 1 ppm (IQR)</p> <p><b>RR Estimate [Lower CI, Upper CI]</b></p> <p>Lags examined (days) : 0, 1, 2, 3, 4, 5, 6</p> <p>All year: Lag 5 : All ages : 0.94 (0.91-0.98) Lag 5 : 64+ age : 1.07 (1.01-1.13)</p> <p>Summer: Lag 5 : All ages : 1.19 (1.02-1.38) Lag 5 : 64+ age : 1.60 (1.27-2.03)</p> <p>2-pollutant model: Lag 5 : 64+ age : CO + PM<sub>10</sub> : 1.04 (0.98-1.11)</p>
<p><b>Author:</b> Maheswaran et al. (2005b)</p> <p><b>Period of Study:</b> 1994 - 1998</p> <p><b>Location:</b> Sheffield, UK</p>	<p><b>Emergency Hospital Admission</b></p> <p><b>Health Outcome (ICD9):</b> CHD (410-414)</p> <p><b>Study Design:</b> Ecological</p> <p><b>Statistical Analyses:</b> Poisson regression</p> <p><b>Age Groups Analyzed:</b> 45+ years</p> <p><b>Sample Description:</b> 11,407 Emergency Hospital Admissions for CHD in patients 45+ years (within 1,030 census districts)</p>	<p><b>Averaging Time:</b> NR</p> <p><b>Mean (SD) unit:</b> Quintiles</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NA</p> <p><b>RR Estimate [Lower CI, Upper CI]</b></p> <p>Lowest quintile reference category</p> <p>Adjusted for sex, age, deprivation, smoking: 2nd : 0.97 (0.89-1.07) 3rd : 0.94 (0.86-1.04) 4th : 0.96 (0.97-1.06) 5th : 0.88 (0.79- 0.98)</p> <p>Adjusted for sex, age: 2nd : 1.09 (1.00-1.19) 3rd : 1.15 (1.05-1.26) 4th : 1.19 (1.09-1.30) 5th : 1.20 (1.09-1.32)</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Mann et al. (2002) <b>Period of Study:</b> 1988 - 1995 <b>Location:</b> Southern California	<b>Health Outcome (ICD9):</b> IHD (IHD) (410-414); Myocardial Infarction (MI) (410) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson regression, GAM <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 54,863 IHD admissions among Southern California Kaiser- Permanente members (within 20km of monitor)	<b>Averaging Time:</b> 8-h <b>Mean (SD) unit:</b> 2.07 ppm <b>Range (Min, Max):</b> 0.30, 11.8 <b>Copollutant: correlation</b> Ranging across 7 regions: NO <sub>2</sub> : r = 0.64, 0.86 O <sub>3</sub> : r = -0.37, 0.28 PM <sub>10</sub> : r = 0.15, 0.40	<b>Increment:</b> 1 ppm <b>% Change [Lower CI, Upper CI]</b> Lags examined (days) : 0, 1, 2, 2 ma, 3 ma, 4 ma <b>With arrhythmia:</b> Lag 0 : 2.99 (1.80-4.99) Lag 1 : 1.51 (0.37-2.66) Lag 2 : 1.26 (0.15-2.38) 2 ma : 2.66 (1.40-3.94) 3 ma : 2.59 (1.27-3.92) 4 ma : 2.25 (0.90-3.63) <b>With CHF:</b> Lag 0 : 3.60 (1.620-5.63) Lag 1 : 3.34 (1.48-5.22) Lag 2 : 1.90 (0.11-3.72) 2 ma : 4.23 (2.13-6.37) 3 ma : 4.14 (1.96-6.37) 4 ma : 4.07 (1.81-6.38) <b>Without secondary diagnosis:</b> Lag 0 : 1.62 (0.65-2.59) Lag 1 : 1.45 (0.54-2.37) Lag 2 : 0.92 (0.04-1.82) 2 ma : 1.83 (0.80-2.86) 3 ma : 1.79 (0.72-2.87) 4 ma : 1.82 (0.71-2.94)
<b>Author:</b> Szyszkwicz (2007) <b>Period of Study:</b> 1997 - 2003 <b>Location:</b> Montreal, Canada	<b>Study Design:</b> Time-series <b>Health Outcome (ICD9 or ICD10):</b> ED Visits. IHD: ICD9: 410-414 <b>Statistical Analyses:</b> Poisson regression (GLMM) <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 4979 ED Visits	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.5 ppm <b>Range (Min, Max):</b> 0.1, 3.1 <b>Copollutant:</b> NR	<b>Increment:</b> 0.2 ppm <b>% Change [Lower CI, Upper CI] ; lag :</b> Lags examined (days): 0, 1 All Patients: Lag 0 : 5.4 (2.3-8.5) Males: Lag 0 : 7.5 (3.6-11.6) Females: Lag 0 : 2.7 (-2.0 to 7.6) Ages ≥ 64 All Patients: Lag 0 : 4.9 (1.3-8.7) Males: Lag 0 : 7.5 (2.6-12.6) Females: Lag 0 : 2.4 (-3.0 to 0) Lag 1 not significant for all results
<b>Author:</b> von Klot et al. (2005) <b>Period of Study:</b> 1992 - 2001 <b>Location:</b> 5 European cities: Augsburg, Germany Barcelona, Spain Helsinki, Finland Rome, Italy Stockholm, Sweden	<b>Health Outcome:</b> Hospital Cardiac (Myocardial Infarction (MI), Angina, Dysrhythmia, Heart Failure) Re-admissions <b>Study Design:</b> Prospective Cohort <b>Statistical Analyses:</b> Poisson regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 22006 survivors of first MI	<b>Averaging Time:</b> 24-h <b>Unit:</b> mg/m <sup>3</sup> <b>Mean (SD) unit:</b> Augsburg, Germany: 0.93 Barcelona, Spain: 1.00 Helsinki, Finland: 0.42 Rome, Italy: 2.21 Stockholm, Sweden: 0.43 <b>Range (Min, Max):</b> NR <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.21 – 0.57 NO <sub>2</sub> : r = 0.44 – 0.75 O <sub>3</sub> : r = -.027 – 0.47	<b>Increment:</b> 0.2 mg/m <sup>3</sup> (0.172 ppm) <b>RR Estimate [Lower CI, Upper CI]</b> Lags examined (days): 0, 1, 2, 3 Lag 0: MI : 1.022 (0.998-.047) Angina : 1.009 (0.992-.02) Cardiac : 1.014 (1.001-.026)
<b>HEART FAILURE</b>			
<b>Author:</b> Lee et al. (2007a) <b>Period of Study:</b> 1996 - 2004 <b>Location:</b> Kaohsiung City, Taiwan	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> CHF (428) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 13475 Hospital Admissions (63 Hospitals)	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.76 ppm <b>Range (Min, Max):</b> 0.14, 1.72 <b>Copollutant:</b> NR	<b>Increment:</b> 0.31 ppm <b>OR Estimate [Lower CI, Upper CI]</b> Lag examined (days): 0-2 ≥ 25°C: 1.19 (1.09-1.31) <25°C: 1.39 (1.24-1.54) <b>Adjusted for PM<sub>10</sub>:</b> ≥ 25°C: 1.15 (1.04-1.27) <25°C: 1.21 (1.206-1.38) <b>Adjusted for SO<sub>2</sub>:</b> ≥ 25°C: 1.23 (1.11-1.36) <25°C: 1.39 (1.24-1.55) <b>Adjusted for NO<sub>2</sub>:</b> ≥ 25°C: 1.22 (1.08-1.39) <25°C: 0.94 (0.81-1.10) <b>Adjusted for O<sub>3</sub>:</b> ≥ 25°C: 1.17 (1.07-1.28) <25°C: 1.36 (1.22-1.51)

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Symons et al. (2006) <b>Period of Study:</b> 2002 (April – November) <b>Location:</b> Johns Hopkins Bayview Medical Center, Baltimore, MD	<b>Hospital Admissions</b> <b>Health Outcome:</b> NR <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 398 Hospital Admissions for CHF	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.4 ppm <b>Range (Min, Max):</b> 0.1, 1.0 <b>Copollutant:</b> NR	<b>Increment:</b> 0.2 ppm <b>OR Estimate [Lower CI, Upper CI]</b> Lags examined (days): 0, 1, 2, 3, cum 1, cum 2, cum 3 Lag 0 : 0.86 (0.67-1.11) Lag 1 : 0.90 (0.70-1.17) Lag 2 : 0.96 (0.73-1.26) Lag 3 : 0.88 (0.67-1.16) Cum. Lag1 : 0.82 (0.60-1.13) Cum. Lag2 : 0.80 (0.54-1.17) Cum. Lag3 : 0.27 (0.46-1.14)
<b>Author:</b> Wellenius et al. (2005b) <b>Period of Study:</b> 1987 - 1999 <b>Location:</b> Pennsylvania, PA	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> CHF (428, 428.1) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> 65+ <b>Sample Description:</b> 54019 Hospital Admissions among Medicare beneficiaries	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 1.03 ppm <b>Range (percentiles):</b> 25th = 0.68; 75th = 1.23 <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.57 NO <sub>2</sub> : r = 0.70 O <sub>3</sub> : r = -0.25 SO <sub>2</sub> : r = 0.54	<b>Increment:</b> 0.55 ppm <b>% Change [Lower CI, Upper CI]</b> Lags examined (days): 0, 1, 2, 3 <b>Lag 0:</b> Single pollutant model: 4.55 (3.33-5.79) Adjusted for PM <sub>10</sub> : 5.18 (3.49-6.89) Adjusted for NO <sub>2</sub> : 4.84 (3.06-6.66) Adjusted for O <sub>3</sub> : 4.35 (3.08-5.64) Adjusted for SO <sub>2</sub> : 4.51 (3.15-5.90)
<b>CARDIOVASCULAR DISEASES – NON-SPECIFIC</b>			
<b>Author:</b> Ballester et al. (2001) <b>Period of Study:</b> 1994 - 1996 <b>Location:</b> Valencia, Spain	<b>ED Visits</b> <b>Health Outcome (ICD9):</b> CVD (390-459); Heart Diseases (410-414, 427, 428); Cerebrovascular Disease (430 – 438) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 6.2 mg/m <sup>3</sup> <b>Range (Min, Max):</b> 0.6, 17.8 <b>Copollutant: correlation</b> BS: r = 0.64 NO <sub>2</sub> : r = 0.03 SO <sub>2</sub> : r = 0.74 O <sub>3</sub> : r = -0.26	<b>Increment:</b> 1 mg/m <sup>3</sup> <b>RR Estimate [Lower CI, Upper CI] ; lag:</b> Lags examined (days): 0, 1, 2, 3, 4, 5 All cardiovascular: Lag 2 : 1.0077 (0.9912-1.0138) Heart Disease: Lag 1 : 1.0092 (0.9945-1.0242) Cerebrovascular Disease: Lag 1 : 0.9874 (0.9646-1.0107)
<b>Author:</b> Ballester et al. (2006) <b>Period of Study:</b> 1995 - 1999 <b>Location:</b> 14 Cities in Spain	<b>Health Outcome (ICD9):</b> All CVD (390-459); Heart Diseases (410-414, 427, 428) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> GAM <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> NR	<b>Averaging Time:</b> 8-h <b>Mean (SD) unit:</b> Range across 14 cities, 1.4-2.8 mg/m <sup>3</sup> <b>Range (percentiles):</b> 10th = 0.4-1.7; 90th = 2.0-3.9 <b>Copollutant:</b> NR	<b>Increment:</b> 1 mg/m <sup>3</sup> <b>% Change [Lower CI, Upper CI]</b> Lags examined (days): 0-1 All CVD: Lag 0-1 : 2.06 (0.65-3.48) Heart Disease: Lag 0-1 : 4.15 (1.31-7.08)
<b>Author:</b> Barnett et al. (2006) <b>Period of Study:</b> 1998 – 2001 <b>Location:</b> Brisbane, Canberra, Melbourne, Perth, Sydney Australia Auckland & Christchurch, New Zealand	<b>Hospital Admissions with Cardiovascular Diseases</b> <b>Health Outcome (ICD9):</b> Arrhythmia (247); Cardiac Disease (390 – 429); Cardiac Failure (428); IHD (410-413); MI (410); Total CVD (390-459) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> 15-64 years & ≥ 65 years <b>Sample Description:</b> NR	<b>Averaging Time:</b> 8-h <b>Mean (SD) unit:</b> ppm Brisbane: 1.7 Canberra: 0.9 Melbourne: 1.0 Perth: 1.0 Sydney: 0.8 Auckland: 2.1 Christchurch: 0.5 <b>Range (Min, Max):</b> ppm Brisbane: 0.0, 7.0 Canberra: 0.0, 5.8 Melbourne: 0.1, 8.0 Perth: 0.1, 4.0 Sydney: 0.0, 4.5 Auckland: 0.2, 7.9 Christchurch: 0.0, 5.4 <b>Copollutant:</b> NR	<b>Increment:</b> 0.9 ppm <b>% Change [Lower CI, Upper CI]</b> Lags examined (days): 0-1 <b>15-64 years</b> Arrhythmia: 2.5 (0.1-4.9) Cardiac: 1.7 (0.5-2.9) Cardiac Failure: 4.2 (0.6-7.8) IHD: 1.6 (-0.6 to 3.9) MI: 1.8 (-0.7 to 4.3) Total CVD: 1.2 (0.3-2.1) <b>≥ 65 years</b> Arrhythmia: 0.1 (-1.8 to 2.1) Cardiac: 2.8 (1.3-4.4) Cardiac Failure: 6.0 (3.5-8.5) IHD: 2.3 (0.9-3.8) MI: 2.9 (0.8-4.9) Total CVD: 2.2 (0.9-3.4)

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Chang et al. (2005) <b>Period of Study:</b> 1997 - 2001 <b>Location:</b> Taipei, Taiwan	<b>Health Outcome (ICD9):</b> CVD Hospital Admissions (410 – 429) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 74,509 CVD hospital admissions (47 Hospitals)	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 1.37 ppm <b>Range (Min, Max):</b> 0.37, 3.66 <b>Copollutant:</b> NR	<b>Increment:</b> 0.49 ppm <b>OR Estimate [Lower CI, Upper CI]</b> Lags examined (days) : 0-2 ≥ 20°C: 1.090 (1.064-1.118) <20°C: 0.984 (0.927-1.044) <b>Adjusted for PM<sub>10</sub>:</b> ≥ 20°C: 1.171 (1.132-1.211) <20°C: 0.946 (0.892-1.003) <b>Adjusted for SO<sub>2</sub>:</b> ≥ 20°C: 1.232 (1.194-1.272) <20°C: 1.098 (1.034-1.165) <b>Adjusted for NO<sub>2</sub>:</b> ≥ 20°C: 1.048 (1.003-1.095) <20°C: 0.983 (0.914-1.058) <b>Adjusted for O<sub>3</sub>:</b> ≥ 20°C: 1.196 (1.161-1.232) <20°C: 1.092 (1.031-1.157)
<b>Author:</b> Fung et al. (2005) <b>Period of Study:</b> 1995 - 2000 <b>Location:</b> Windsor, Ontario, Canada	<b>Hospital Admissions of Cardiovascular Diseases</b> <b>Health Outcome (ICD9):</b> CHF (428); IHD (410 – 414); Dysrhythmias (427) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> GLM <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 11,632 Cardiac hospital admissions	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 1.3 ppm <b>Range (Min, Max):</b> 0.0, 11.8 <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.21 NO <sub>2</sub> : r = 0.38 SO <sub>2</sub> : r = 0.16 O <sub>3</sub> : r = 0.10	<b>Increment:</b> 1.2 ppm <b>% Change [Lower CI, Upper CI]</b> Lags examined (days): 0, 0-1, 0-2 <b>&lt;65 years</b> Lag 0 : -3.1 (-7.4 to 1.4) Lag 0-1 : -2.7 (-8.1 to 3.0) Lag 0-2 : -0.5 (-6.7 to 6.0) <b>≥ 65 years</b> Lag 0 : 0.5 (-2.2 to 3.3) Lag 0-1 : 2.3 (-1.1 to 5.9) Lag 0-2 : 2.8 (-1.1 to 7.0)
<b>Author:</b> Jalaludin et al. (2006) <b>Period of Study:</b> 1997 - 2001 <b>Location:</b> Sydney, Australia	<b>ED Visits</b> <b>Health Outcome (ICD9):</b> All Cardiovascular (390-459); Cardiac Disease (390-429); IHD (410-413); Cerebrovascular or Stroke (430-438) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> GLM & GAM <b>Age Groups Analyzed:</b> 65+ years <b>Sample Description:</b> NR	<b>Averaging Time:</b> 8-h <b>Mean (SD) unit:</b> 0.82 ppm <b>Range (Min, Max):</b> 0.02, 4.63 <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.31 NO <sub>2</sub> : r = 0.71 SO <sub>2</sub> : r = 0.51 O <sub>3</sub> : r = 0.19	<b>Increment:</b> 0.69 ppm <b>% Change [Lower CI, Upper CI]</b> Lags examined (days): 0, 1, 2, 3, 0-1 <b>All Cardiovascular:</b> Lag 0 : 2.32 (1.45-3.19) Lag 1 : 1.33 (0.47-2.20) Lag 0-1 : 2.35 (1.39-3.32) <b>Cardiac Disease:</b> Lag 0 : 2.52 (1.50-3.54) Lag 1 : 1.85 (0.83-2.88) Lag 2 : 1.11 (0.0-2.15) Lag 0-1 : 2.85 (1.71-4.01) <b>IHD:</b> Lag 0 : 2.83 (1.22-4.48) Lag 1 : 1.58 (0.01-3.19) Lag 0-1 : 2.86 (1.07-4.68) <b>Stroke:</b> No results were significant for Stroke. <b>All Cardiovascular Disease:</b> Cool period: Lag 0 : 3.26 (2.00-4.53) <b>Cardiac Disease:</b> Cool period: Lag 0 : 3.43 (1.95-4.93) <b>IHD:</b> Cool period: Lag 0 : 3.64 (1.28-6.06) Warm period: Lag 0 : 2.29 (0.01-4.62) <b>Stroke:</b> Cool period: Lag 0 : 3.54 (0.78-6.37) <b>Notes:</b> Cool : May to October Warm : November to April
<b>Author:</b> Koken et al. (2003) <b>Period of Study:</b> 1993 - 1997 <b>Location:</b> Denver, CO	<b>Hospital Admissions for Cardiovascular Disease</b> <b>Health Outcome (ICD9):</b> MI (410-410.92); Coronary Atherosclerosis (414 – 414.05); Pulmonary Heart Disease (416 - 416.9); Cardiac Dysrhythmia (427 – 427.9); CHF (428) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> GLM <b>Age Groups Analyzed:</b> >65 years <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.9 ppm <b>Range (Min, Max):</b> 0.3, 1.6 <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.25 NO <sub>2</sub> : r = 0.73 SO <sub>2</sub> : r = 0.21 O <sub>3</sub> : r = -0.40	<b>Increment:</b> 0.3 ppm <b>% Change [Lower CI, Upper CI]</b> Lags examined (days): 1, 2, 3, 4 <b>CHF:</b> Lag 3 : 10.5 (0.1-22.0) CO not significantly associated with other Lag periods.



Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Linn et al. (2000)</p> <p><b>Period of Study:</b> 1992 - 1995</p> <p><b>Location:</b> Los Angeles, CA</p>	<p><b>Health Outcome:</b> Hospital Admissions for Cardiovascular, Cerebrovascular, Pulmonary.</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Ordinary least squares regression; Poisson regression</p> <p><b>Age Groups Analyzed:</b> &gt;30 years</p> <p><b>Sample Description:</b> NR</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> Winter: 1.7; Spring: 1.0; Summer: 1.2; Fall: 2.1</p> <p><b>Range (Min, Max):</b> Winter: 0.5; 5.3; Spring: 0.4, 2.2; Summer: 0.3, 2.7; Fall: 0.2, 4.3</p> <p><b>Copollutant: correlation</b>  <b>Winter:</b> PM<sub>10</sub>: r = 0.78; NO<sub>2</sub>: r = 0.89; O<sub>3</sub>: r = -0.43;  <b>Spring:</b> PM<sub>10</sub>: r = 0.54; NO<sub>2</sub>: r = 0.92; O<sub>3</sub>: 0.29  <b>Summer:</b> PM<sub>10</sub>: r = 0.72; NO<sub>2</sub>: r = 0.94; O<sub>3</sub>: 0.03  <b>Fall:</b> PM<sub>10</sub>: r = 0.58; NO<sub>2</sub>: r = 0.84; O<sub>3</sub>: r = -0.36</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>Co-efficient [SE]</b></p> <p>Lags examined (lags) : 0, 1</p> <p><b>Lag 0:</b></p> <p><b>Cardiovascular</b>  All: 0.032 (0.003)* (e.g. 3.2% increase)  Winter: 0.038 (0.006)*  Spring: 0.010 (0.015)  Summer: 0.035 (0.014)*  Fall: 0.027 (0.006)*</p> <p><b>Cerebrovascular</b>  All : 0.009 (0.007)  Winter: -0.008 (0.014)  Spring: 0.107 (0.033)*  Summer: 0.030 (0.033)  Fall: 0.008 (0.012)</p> <p><b>Myocardial Infarction</b>  All : 0.040 (0.009) *</p> <p><b>CHF</b>  All : 0.025 (0.009)*</p> <p><b>Cardiac Arrhythmia</b>  All : 0.023 (0.009)*</p> <p><b>Stroke</b>  All : 0.044 (0.009)*</p> <p><b>Notes:</b>* p &lt; 0.05</p>
<p><b>Author:</b> Metzger et al. (2004a)</p> <p><b>Period of Study:</b> 1993 - 2000</p> <p><b>Location:</b> Atlanta, GA</p>	<p><b>ED Visits</b> (from 31 hospitals)</p> <p><b>Health Outcome (ICD9):</b> Cardiovascular: IHD (410 – 414); Acute MI (410); Dysrhythmia (427); Cardiac Arrest (427.5); CHF (428); Peripheral Vascular &amp; Cerebrovascular Disease (PVCD) (433 – 437, 440, 443, 444, 451-453); Atherosclerosis (440); Stroke (436)</p> <p><b>Study Design:</b> Case-crossover</p> <p><b>Statistical Analyses:</b> Poisson regression (GLM)</p> <p><b>Age Groups Analyzed:</b> All</p> <p><b>Sample Description:</b> 4,407,535 visits</p>	<p><b>Averaging Time:</b> 1-h</p> <p><b>Median (SD) unit:</b> 1.5 ppm</p> <p><b>Range (percentiles):</b> 10th = 0.5; 90th = 3.4</p> <p><b>Copollutant: correlation</b>  PM<sub>10</sub>: r = 0.47  NO<sub>2</sub>: r = 0.68  SO<sub>2</sub>: r = 0.26  O<sub>3</sub>: r = 0.20</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>RR Estimate [Lower CI, Upper CI]</b></p> <p>Lags examined (days): 0-2ma</p> <p>All CVD : 1.017 (1.008-1.027)  Dysrhythmia : 1.012 (0.993-1.031)  CHF : 1.010 (0.988-1.032)  IHD : 1.016 (0.999-1.034)  PVCD : 1.031 (1.010-1.052)</p>
<p><b>Author:</b> Peel et al. (2007)</p> <p><b>Period of Study:</b> 1993 - 2000</p> <p><b>Location:</b> Atlanta, GA</p>	<p><b>ED Visits</b> (from 31 hospitals)</p> <p><b>Health Outcome (ICD9):</b> Cardiovascular: IHD (410 – 414); Dysrhythmia (427); CHF (428); PVCD (433 – 437, 440, 443, 444, 451-453)</p> <p><b>Study Design:</b> Case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> All</p> <p><b>Sample Description:</b> 4,407,535 visits</p>	<p><b>Averaging Time:</b> 1-h</p> <p><b>Mean (SD) unit:</b> 1.8 ppm</p> <p><b>Range (SD):</b> SD: 1.2</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1.2 ppm</p> <p><b>OR Estimate [Lower CI, Upper CI]</b></p> <p>Lags examined (days): 0-2ma</p> <p><b>IHD:</b>  Without Diabetes : 1.023 (1.004-1.420)  Without CHF: 1.024 (1.006-1.042)</p> <p><b>Dysrhythmias:</b>  With Hypertension : 1.065 (1.015-1.118)</p> <p><b>PVCD:</b>  With Hypertension : 1.038 (1.004-1.074)  Without Hypertension: 1.027 (1.002-1.054)  With Diabetes: 1.065 (1.012-1.121)  Without Diabetes: 1.025 (1.003-1.048)  With COPD: 1.113 (1.027-1.205)  Without COPD: 1.026 (1.004-1.047)  Without CHF : 1.029 (1.008-1.051)  With Dysrhythmias: 1.072 (1.011-1.138)  Without Dysrhythmias : 1.026 (1.004-1.048)  <b>CHF:</b> With COPD : 1.058 (1.003-1.115)</p>
<p><b>Author:</b> Slaughter et al. (2005)</p> <p><b>Period of Study:</b> 1995 - 2001</p> <p><b>Location:</b> Spokane, WA</p>	<p><b>Health Outcome (ICD9):</b> Cardiac Hospital Admissions: (390-459)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson regression (GLM &amp; GAM)</p> <p><b>Age Groups Analyzed:</b> All</p> <p><b>Sample Description:</b> NR</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 0.42-1.82</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant correlation :</b>  PM<sub>10</sub>: r = 0.32  PM<sub>2.5</sub>: r = 0.62</p>	<p><b>Increment:</b> NR</p> <p><b>RR Estimate [Lower CI, Upper CI] ; lag :</b></p> <p>Lags examined (days): 1, 2, 3</p> <p>No significant association. Results not reported.</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Yang et al. (2004a) <b>Period of Study:</b> 1997 - 2000 <b>Location:</b> Kaohsiung City, Taiwan	<b>Health Outcome (ICD9):</b> Cardiovascular diseases (410-429) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All <b>Sample Description:</b> 29,661 Cardiovascular hospital admissions (63 hospitals)	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.79 ppm <b>Range (Min, Max):</b> 0.24, 1.72 <b>Copollutant:</b> NR	<b>Increment:</b> 0.28 ppm <b>OR Estimate [Lower CI, Upper CI]</b> Lag examined (days): 0-2 ≥ 25°C: 1.264 (1.205-1.326) <25°C: 1.448 (1.357-1.545) <b>Adjusted for PM<sub>10</sub>:</b> ≥ 25°C: 1.206 (1.146-1.270) <25°C: 1.314 (1.213-1.423) <b>Adjusted for SO<sub>2</sub>:</b> ≥ 25°C: 1.406 (1.327-1.489) <25°C: 1.3450 (1.352-1.555) <b>Adjusted for NO<sub>2</sub>:</b> ≥ 25°C: 1.246 (1.166-1.332) <25°C: 0.905 (0.819-0.999) <b>Adjusted for O<sub>3</sub>:</b> ≥ 25°C: 1.250 (1.191-1.311) <25°C: 1.447 (1.356-1.545)

**Table C-3. Studies of CO exposure and neonatal and postneonatal outcomes.**

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Bell et al. (2007) <b>Period of Study:</b> 1999 - 2002 <b>Location:</b> Connecticut and Massachusetts	<b>Health Outcome:</b> Birth weight and LBW <b>Study Design:</b> Retrospective cohort <b>Statistical Analyses:</b> Linear and logistic regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> 358,504 full term live singleton births (32-44 weeks)	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 0.65 ppm (0.18) <b>Range (Min, Max):</b> NR <b>Copollutant:</b> NR	<b>Increment:</b> Interquartile range – 0.30 ppm <b>Regression co-efficient for birth weight (g) [Lower CI, Upper CI]</b> Entire pregnancy: -16.2 (-19.7 to -12.6) Stratified by race: Black mother : -10.9 (-20.2 to -1.6) White mother : -17.5 (-21.3 to -13.7) <b>OR for LBW [Lower CI, Upper CI]</b> Entire pregnancy : 1.028 (0.983-1.074)
<b>Author:</b> Brauer et al. (2008) <b>Period of Study:</b> 1999 - 2004 <b>Location:</b> Vancouver, Canada	<b>Health Outcome:</b> LBW, PTB and SGA <b>Study Design:</b> Retrospective cohort <b>Statistical Analyses:</b> Logistic regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> 70249 live singleton births	<b>Averaging Time:</b> Land use regression model <b>Mean (SD) unit:</b> 633 µg/m <sup>3</sup> <b>Range (Min, Max):</b> 124, 1409 <b>Copollutant: correlation:</b> PM <sub>10</sub> : r = 0.73 NO <sub>2</sub> : r = 0.75 SO <sub>2</sub> : r = 0.82 O <sub>3</sub> : r = -0.39	<b>Increment:</b> 100 µg/m <sup>3</sup> <b>OR for SGA [Lower CI, Upper CI] ;</b> Entire pregnancy : 1.06 (1.03-1.08) <b>OR for term LBW [Lower CI, Upper CI] ;</b> Entire pregnancy : 1.02 (0.96-1.09) <b>OR PTB [Lower CI, Upper CI] ;</b> Entire pregnancy : 1.16 (1.01-1.33)
<b>Author:</b> Chen et al. (2002) <b>Period of Study:</b> 1991 - 1999 <b>Location:</b> Northern Nevada	<b>Health Outcome:</b> Birth weight & LBW <b>Study Design:</b> Retrospective cohort <b>Statistical Analyses:</b> Linear and logistic regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> 39,338 full term live singleton births (37-44 weeks)	<b>Averaging Time:</b> 8-h <b>Mean (SD) unit:</b> 0.98 ppm <b>Range (Min, Max):</b> 0.25, 4.87 <b>Copollutant:</b> NR	<b>Increment:</b> NR <b>Regression co-efficient for birth weight (g) [SE]</b> Trimesters: First : -1.02 (6.68) Second : -0.07 (6.58) Third : -3.95 (6.76) Entire pregnancy : -8.28 (14.9) <b>Notes:</b> CO not associated with LBW
<b>Author:</b> Conceicao et al. (2001) <b>Period of Study:</b> 1994 - 1997 <b>Location:</b> Sao Paulo, Brazil	<b>Health Outcome:</b> Child mortality, under 5 years of age <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson regression (GAM) <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 4.4 ppm (2.2) <b>Range (Min, Max):</b> NR <b>Copollutant:</b> NR	<b>Increment:</b> NR <b>Regression co-efficient for Child mortality – under 5 years of age [SE] ;</b> Lags examined : 0, 1, 2, 3 Lag 2 : 0.0306 (0.0076) (p < 0.01) Lag chosen for best fitting model

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Gilboa et al. (2005) <b>Period of Study:</b> 1997 - 2000 <b>Location:</b> Texas	<b>Health Outcome:</b> Birth defects (heart defects & orofacial clefts) <b>Study Design:</b> Case-control <b>Statistical Analyses:</b> Conditional Logistic regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> NR	<b>Averaging Time:</b> NR <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant:</b> NR	<b>Increment:</b> Exposure categories (ppm): <0.4; 0.4 – 0.5; 0.5 – 0.7; >0.7 <b>OR for Birth Defects [Lower CI, Upper CI] ;</b> Exposure period : weeks 3 – 8 of pregnancy <b>Conotruncal defects:</b> 1.00; 1.38 (0.97-1.97); 1.17 (0.81-1.70); 1.46 (1.03-2.08) <b>Tetralogy of Fallot:</b> 1.00; 0.92 (0.52-1.62); 1.27 (0.75-2.14); 2.04 (1.26-3.29) <b>Notes:</b> CO was not associated with the following defects: Aortic artery & valve, atrial septal, pulmonary artery & valve, ventricular septal, endocardial cushion & mitral valve , cleft lip, cleft palate, aortic valve stenosis, coarctation of the aorta, ostium secundum.
<b>Author:</b> Gouveia et al. (2004) <b>Period of Study:</b> 1997 <b>Location:</b> Sao Paulo, Brazil	<b>Health Outcome:</b> Birth weight & LBW <b>Study Design:</b> Retrospective cohort <b>Statistical Analyses:</b> Linear and logistic regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> 179,460 live singleton term births (>37 weeks)	<b>Averaging Time:</b> 8-h <b>Mean (SD) unit:</b> 3.7 ppm <b>Range (Min, Max):</b> 1.1, 11.4 <b>Copollutant:</b> NR	<b>Increment:</b> 1 ppm <b>Regression co-efficient for birth weight (g) [Lower CI, Upper CI]</b> Trimesters: First : -23.1 (-41.3 to -4.9); Second : 3.2 (-18.2 to 24.5); Third : 1.9 (-18.2 to 22.0) <b>OR for LBW ) [Lower CI, Upper CI]</b> 4th quartile exposure (compared to lowest quartile): First : 1.02 (0.82-1.27); Second : 1.07 (0.88-1.30); Third : 0.93 (0.76-1.12)
<b>Author:</b> Ha et al. (2001) <b>Period of Study:</b> 1996 - 1997 <b>Location:</b> Seoul, South Korea	<b>Health Outcome:</b> LBW <b>Study Design:</b> Retrospective cohort <b>Statistical Analyses:</b> Logistic regression (GAM) <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> 276 763 full term live singleton births (>37 weeks)	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> Percentiles: 25th: 0.99 ppm 75th: 1.41 ppm <b>Copollutant correlation:</b> TSP: r = 0.73 NO <sub>2</sub> : r = 0.75 SO <sub>2</sub> : r = 0.82 O <sub>3</sub> : r = -0.39	<b>Increment:</b> 0.42 ppm <b>RR for LBW [Lower CI, Upper CI]</b> Trimesters: First : 1.08 (1.04, 1.12) Third : 0.91 (0.87, 0.96)
<b>Author:</b> Ha et al. (2003) <b>Period of Study:</b> 1995 - 1999 <b>Location:</b> Seoul, South Korea	<b>Health Outcome:</b> Post-neonatal mortality (1 month – 1 yr) (also looked at older age groups) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson regression (GAM) <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 1.2 ppm <b>Range (Min, Max):</b> 0.39, 3.38 <b>Copollutant correlation:</b> PM <sub>10</sub> : r = 0.63 NO <sub>2</sub> : r = 0.72 SO <sub>2</sub> : r = 0.75 O <sub>3</sub> : r = -0.46	<b>Increment:</b> 0.57 ppm <b>RR for Post–neonatal mortality (1 month – 1 yr) [Lower CI, Upper CI]</b> Lags examined : 0 Total mortality: Lag 0 : 1.020 (0.976-1.067) Respiratory mortality: Lag 0 : 1.388 (1.009-1.911)
<b>Author:</b> Huynh et al. (2006) <b>Period of Study:</b> 1999 - 2000 <b>Location:</b> California	<b>Health Outcome:</b> PTB (24-36 weeks gestation) <b>Study Design:</b> Case-control <b>Statistical Analyses:</b> Conditional Logistic regression <b>Age Groups Analyzed:</b> Cases = 24-36 weeks gestation; Controls = 39 – 44 weeks <b>Sample Description:</b> 10,673 PTBs (cases); 32,119 term births (controls)	<b>Averaging Time:</b> NR <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant:</b> NR	<b>Increment:</b> 1 ppm Exposure level – Quartiles of exposure for first month and last two weeks of gestation (mg/m <sup>3</sup> ) First : <0.61; Second : 0.61 – 0.82; Third : 0.82 – 1.07; Fourth : >1.07 Quartiles for entire pregnancy and last two weeks of pregnancy were similar. <b>OR for PTB [Lower CI, Upper CI]</b> <b>First month of gestation:</b> Per 1 ppm increase : 1.10 (0.99-1.20) Second quartile : 0.94 (0.88-1.01) Third quartile : 1.04 (0.97-1.11) Fourth quartile : 1.05 (0.96-1.14) <b>Last two weeks of gestation:</b> Per 1 ppm increase : 1.00 (0.93-1.09) Second quartile : 1.03 (0.97-1.10) Third quartile : 1.04 (0.97-1.12) Fourth quartile : 0.99 (0.91-1.08) <b>Entire pregnancy:</b> Per 1 ppm increase : 1.06 (0.95-1.18) Second quartile : 0.97 (0.91-1.04) Third quartile : 0.99 (0.92-1.05) Fourth quartile : 1.02 (0.94-1.09) <b>Lowest quartile used as reference group</b>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Hwang and Jaakkola (2008)</p> <p><b>Period of Study:</b> 2001-2003</p> <p><b>Location:</b> Taiwan</p>	<p><b>Health Outcome:</b> Oral clefts (with or without palate)</p> <p><b>Study Design:</b> Case-control</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 6,530 cases from 721,289 newborns</p>	<p><b>Averaging Time:</b> 8-h</p> <p><b>Mean (SD) unit:</b> 0.69 (0.4)</p> <p><b>Range (Min, Max):</b> 0.25, 2.7</p> <p><b>Copollutant correlation:</b>  PM<sub>10</sub>: r = -0.19  NO<sub>x</sub>: r = 0.82  SO<sub>2</sub>: r = 0.24  O<sub>3</sub>: r = -0.19</p>	<p><b>Increment:</b> 100 ppb</p> <p><b>RR for oral cleft [Lower CI, Upper CI]</b></p> <p>Month 1 : 1.00 (0.96-1.04)</p> <p>Month 2 : 1.00 (0.96-1.03)</p> <p>Month 3 : 1.00 (0.96-1.03)</p>
<p><b>Author:</b> Jalaludin et al. (2007)</p> <p><b>Period of Study:</b> 1998 - 2000</p> <p><b>Location:</b> Sydney, Australia</p>	<p><b>Health Outcome:</b> PTB</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 123,840 full term live singleton births (&lt;42 weeks)</p>	<p><b>Averaging Time:</b> 8-h</p> <p><b>Mean (SD) unit:</b> 0.9 ppm (0.68)</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant correlation:</b>  PM<sub>10</sub>: r = 0.28  NO<sub>2</sub>: r = 0.60  SO<sub>2</sub>: r = 0.24  O<sub>3</sub>: r = -0.21</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>RR for PTB [Lower CI, Upper CI]</b></p> <p><b>First month:</b>  All of Sydney : 0.89 (0.84-0.95)  Within 5km of site : 1.03 (0.68-1.54)</p> <p><b>First trimester:</b>  All of Sydney : 0.77 (0.71-0.83)  Within 5km of site : 1.24 (0.81-1.91)</p> <p><b>1 month prior to birth:</b>  All of Sydney : 0.96 (0.88-1.04)  Within 5km of site : 1.00 (0.86-1.15)</p> <p><b>3 months prior to birth:</b>  All of Sydney : 0.99 (0.90-1.09)  Within 5km of site : 1.11 (0.94-1.31)</p>
<p><b>Author:</b> Lee et al. (2003a)</p> <p><b>Period of Study:</b> 1996 - 1998</p> <p><b>Location:</b> Seoul, South Korea</p>	<p><b>Health Outcome:</b> LBW</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 388,105 full term live singleton births (37-44 weeks)</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 1.2 ppm</p> <p><b>Range (Min, Max):</b> 0.4, 3.4</p> <p><b>Copollutant correlation:</b>  PM<sub>10</sub>: r = 0.47  NO<sub>2</sub>: r = 0.77  SO<sub>2</sub>: r = 0.79</p>	<p><b>Increment:</b> 0.5 ppm</p> <p><b>OR for LBW [Lower CI, Upper CI]</b></p> <p>First : 1.04 (1.01-1.07)</p> <p>Second : 1.03 (1.00-1.06)</p> <p>Third : 0.96 (0.93-0.99)</p> <p>Entire pregnancy : 1.05 (1.01-1.09)</p>
<p><b>Author:</b> Leem et al. (2006)</p> <p><b>Period of Study:</b> 2001 - 2002</p> <p><b>Location:</b> Incheon, Korea</p>	<p><b>Health Outcome:</b> PTB</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 52,113 live singleton births</p>	<p><b>Averaging Time:</b> Kriging was used to estimate exposure</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant correlation:</b>  PM<sub>10</sub>: r = 0.27  NO<sub>2</sub>: r = 0.63  SO<sub>2</sub>: r = 0.31</p>	<p><b>Increment:</b> Exposure level – Quartiles of exposure for first trimester (mg/m<sup>3</sup>)</p> <p>First : 0.47-0.63; Second : 0.6 -0.77;  Third : 0.78-0.90; Fourth : 0.91-1.27</p> <p>- exposure groups for third trimester was similar</p> <p><b>OR for PTB [Lower CI, Upper CI]</b></p> <p><b>First Trimester:</b>  Second quartile : 0.92 (0.81-1.05)  Third quartile : 1.14 (1.01-1.29)  Fourth quartile : 1.26 (1.11-1.44)</p> <p><b>Third Trimester:</b>  Second quartile : 1.07 (0.95-1.21)  Third quartile : 1.07 (0.94-1.22)  Fourth quartile : 1.16 (1.01-1.34)</p> <p><b>Lowest quartile used as reference group.</b></p>
<p><b>Author:</b> Lin et al. (2004a)</p> <p><b>Period of Study:</b> 1998 - 2000</p> <p><b>Location:</b> Sao Paulo, Brazil</p>	<p><b>Health Outcome:</b> Neonatal death (within first 28 days of life)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson regression (GAM)</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> NR</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 2.83 ppm</p> <p><b>Range (Min, Max):</b> 0.54, 10.25</p> <p><b>Copollutant correlation:</b>  PM<sub>10</sub>: r = 0.71  NO<sub>2</sub>: r = 0.67  SO<sub>2</sub>: r = 0.55  O<sub>3</sub>: r = 0.03</p>	<p><b>Increment:</b> NR</p> <p><b>Regression coefficient for neonatal death [SE]</b></p> <p>Lags examined : 0</p> <p>Lag 0 : 0.0061 (0.0110)</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Lin et al. (2004b)</p> <p><b>Period of Study:</b> 1995 - 1997</p> <p><b>Location:</b> Taipei &amp; Kaoshiung, Taiwan</p>	<p><b>Health Outcome:</b> LBW</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 92288 full term live singleton births (&gt;37 weeks) within 3km of monitoring site.</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> Taipei (avg over 5 sites) 0.84 – 1.31 Kaohsiung (avg over 5 sites) 5.56 – 10.05</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> Exposure groups M = Median exposure 1.1-14.2 ppm H = High exposure &gt;14.2 ppm</p> <p><b>OR for LBW [Lower CI, Upper CI]</b></p> <p><b>Trimesters:</b> First : M 1.01 (0.89, 1.16), H 0.90 (0.75, 1.09) Second : M 1.02 (0.90, 1.16), H 1.00 (0.82, 1.22) Third : M 0.88 (0.77, 1.00), H 0.86 (0.71, 1.03) Entire pregnancy : M 0.89 (0.77, 1.01), H 0.77 (0.63, 0.94)</p> <p><b>Notes:</b> Cut off for exposures groups for second and third trimester were similar to those presented above.</p>
<p><b>Author:</b> Liu et al. (2003)</p> <p><b>Period of Study:</b> 1985 - 1998</p> <p><b>Location:</b> Vancouver, BC, Canada</p>	<p><b>Health Outcome:</b> PTB, IUGR, LBW</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 229,085 live singleton births</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 1.0 ppm</p> <p><b>Range (Min, Max):</b> 25th: 0.7; 75th: 1.2</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1.0 ppm</p> <p><b>OR for LBW [Lower CI, Upper CI]</b></p> <p><b>Month of pregnancy:</b> First month: 1.01 (0.93-1.09) Last month: 0.96 (0.88-1.04)</p> <p><b>OR for PTB [Lower CI, Upper CI]</b> First month : 0.95 (0.89-1.01) Last month : 1.08 (1.01-1.15)</p> <p><b>OR for IUGR [Lower CI- Upper CI]</b> First month : 1.06 (1.01-1.10) Last month : 0.98 (0.94-1.03) Trimester 1 : 1.05 (1.00-1.10) Trimester 2 : 0.97 (0.92-1.01) Trimester 3 : 0.97 (0.93-1.02)</p>
<p><b>Author:</b> Liu et al. (2007)</p> <p><b>Period of Study:</b> 1995 - 2000</p> <p><b>Location:</b> Calgary, Edmonton, and Montreal Canada</p>	<p><b>Health Outcome:</b> IUGR</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 386,202 live singleton births</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 1.1 ppm</p> <p><b>Range (Min, Max):</b> 25th: 0.6; 75th: 1.3</p> <p><b>Copollutant correlation:</b> PM<sub>2.5</sub>: r = 0.31 NO<sub>2</sub>: r = 0.71 SO<sub>2</sub>: r = 0.21 O<sub>3</sub>: r = -0.42</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>RR for LBW [Lower CI, Upper CI]</b></p> <p><b>Notes:</b> CO was associated with an increased risk of IUGR of approximately 16% and 23% in the first and nine month of pregnancy. (All results presented in Figures)</p>
<p><b>Author:</b> Maisonet et al. (2001)</p> <p><b>Period of Study:</b> 1994 - 1996</p> <p><b>Location:</b> Northeastern USA</p>	<p><b>Health Outcome:</b> Live birth weight</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 89,557 live singleton term births (37-44 weeks)</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> Percentiles: 25th: 0.93 ppm; 75th: 1.23 ppm</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>OR for LBW [Lower CI, Upper CI]</b></p> <p><b>Trimester:</b> First : 1.08 (0.91-1.28); Second : 1.14 (0.83-1.58); Third : 1.31 (1.06-1.62)</p> <p><b>Stratified results among African-Americans:</b> First : 1.43 (1.18-1.74); Second : 1.27 (0.87-1.86); Third : 1.75 (1.50-2.04)</p> <p><b>Notes:</b> CO had no effect on whites or Hispanics</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Mannes et al. (2005)</p> <p><b>Period of Study:</b> 1998 - 2000</p> <p><b>Location:</b> Sydney, Australia</p>	<p><b>Health Outcome:</b> Birth weight and SGA</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Linear and logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 138,056 full term all singleton births (including stillbirths) (at least 20 weeks gestation)</p>	<p><b>Averaging Time:</b> 8-h</p> <p><b>Mean (SD) unit:</b> 0.8 ppm</p> <p><b>Range (Min, Max):</b> 0.0, 4.6</p> <p><b>Copollutant: correlation</b>  PM<sub>10</sub>: r = 0.26  NO<sub>2</sub>: r = 0.57  O<sub>3</sub>: r = -0.20</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>Regression coefficients for birth weight (g) [Lower CI, Upper CI]</b></p> <p><b>All births:</b>  First trimester : 1.86 (-8.31 to 12.03)  Second trimester : -10.72 (-23.09 to 1.65)  Third trimester : -6.63 (-18.57 to 5.31)  One month prior to birth : -15.28 (-25.59 to -4.97)</p> <p><b>Births within 5 km of monitor:</b>  First trimester : -8.56 (-28.60 to 10.68)  Second trimester : -28.87 (-50.98 to -6.76)  Third trimester : -22.88 (-44.58 to -1.18)  One month prior to birth : -10.41 (-30.03 to 9.21)</p> <p><b>OR for SGA [Lower CI, Upper CI]</b></p> <p><b>All births:</b>  First trimester : 0.95 (0.88-1.04)  Second trimester : 0.99 (0.90-1.10)  Third trimester : 1.01 (0.91-1.11)  One month prior to birth : 1.06 (0.98-1.16)</p> <p><b>Births within 5km of monitor:</b>  First trimester : 0.99 (0.86-1.14)  Second trimester : 1.06 (0.90-1.25)  Third trimester : 1.05 (0.90-1.23)  One month prior to birth : 1.10 (0.96-1.27)</p>
<p><b>Author:</b> Medeiros et al. (2005)</p> <p><b>Period of Study:</b> 1998 - 2000</p> <p><b>Location:</b> Sao Paulo, Brazil</p>	<p><b>Health Outcome:</b> Birth weight and LBW</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Linear and logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 311,735 full term live singleton births (37-41 weeks)</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> Daily mean shown in Figure (see paper)</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>Regression co-efficient for birth weight (g) [Lower CI, Upper CI]</b></p> <p>Trimesters:  First : -11.9 (-15.5 to -8.2); Second : 4.9 (0.5-9.3);  Third : 12.1 (7.6-16.6)</p> <p><b>OR for LBW [Lower CI, Upper CI]</b></p> <p>4th quartile exposure (compared to lowest quartile)  First : 0.98 (0.91-1.06); Second : 0.97 (0.90-1.05);  Third : 1.03 (0.96-1.11)</p>
<p><b>Author:</b> Parker et al. (2005)</p> <p><b>Period of Study:</b> 2000</p> <p><b>Location:</b> California</p>	<p><b>Health Outcome:</b> Birth weight &amp; SGA</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Linear and logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 18,247 full term live singleton births (40 weeks) within 5 miles of a monitor</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 0.78 ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> Quartiles of exposure for first trimester  First : &lt;0.57; Second : 0.57-0.76 ;  Third : 0.76- 0.93; Fourth : &gt;0.93  - exposure groups for other trimesters were similar</p> <p><b>Regression co-efficient for birth weight (g) [Lower CI, Upper CI]</b></p> <p>Trimesters:  4th quartile exposure (compared to lowest quartile)  First : -7.3 (-29.7 to 15.0); Second : 14.2 (-8.9 to 37.3);  Third : -8.4 (-32.2 to 15.3);  Entire pregnancy: -20.5 (-40.1 to -0.8)</p> <p><b>OR for SGA [Lower CI, Upper CI]</b></p> <p>4th quartile exposure (compared to lowest quartile)  First: 0.91 (0.76-1.09); Second: 0.80 (0.66-0.97);  Third: 0.90 (0.75-1.10);  Entire pregnancy: 0.95 (0.81-1.12)</p>
<p><b>Author:</b> Ritz et al. (2000)</p> <p><b>Period of Study:</b> 1989 - 1993</p> <p><b>Location:</b> Southern California</p>	<p><b>Health Outcome:</b> PTB</p> <p><b>Study Design:</b> Retrospective Cohort</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> Eligible study subjects were singletons born at 26-44 weeks gestation</p> <p><b>Sample Description:</b> 97,518 neonates born in Southern California</p>	<p><b>Averaging Time:</b> 6-9 a.m.</p> <p><b>Mean (SD) unit:</b> 2.70 ppm</p> <p><b>Range (Min, Max):</b> 0.36, 9.12</p> <p><b>Copollutant correlation:</b>  PM<sub>10</sub>: r = 0.37  NO<sub>2</sub>: r = 0.60  O<sub>3</sub>: r = -0.44</p>	<p><b>Increment:</b> 3 ppm</p> <p><b>RR for PTB [Lower CI, Upper CI]</b></p> <p><b>Adjusted for various risk factors and season of birth and conception</b>  6 weeks prior to birth : 1.04 (0.99-1.10)  1st month of pregnancy : 1.04 (0.99-1.09)</p> <p><b>Adjusted for various risk factors</b>  6 weeks prior to birth : 1.06 (1.02-1.10)  1st month of pregnancy : 1.01 (0.97-1.04)</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Ritz et al. (2002)</p> <p><b>Period of Study:</b> 1987 - 1993</p> <p><b>Location:</b> Southern California</p>	<p><b>Health Outcome:</b> Birth defects (heart defects &amp; orofacial clefts)</p> <p><b>Study Design:</b> Case-control</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> NR</p>	<p><b>Averaging Time:</b> NR</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> Exposure categories: ppm &lt;1.14; 1.14-1.57; 1.57- 2.39; &gt;2.39</p> <p><b>OR for Birth defects [Lower CI, Upper CI]:</b> Period of exposure – Second month of pregnancy.</p> <p><b>Aortic artery &amp; valve defects:</b> 1.00 (ref group); 1.10 (0.73-1.66); 1.25 (0.74-2.13); 0.93 (0.47-1.85)</p> <p><b>Pulmonary artery &amp; valve anomalies:</b> 1.00 (ref group); 1.09 (0.69-1.73); 0.92 (0.50-1.70); 1.00 (0.46-2.17)</p> <p><b>Ventricular septal defects:</b> 1.00 (ref group); 1.62 (1.05-2.48); 2.09 (1.19-3.67); 2.95 (1.44-6.05)</p> <p><b>Conotruncal defects:</b> 1.00 (ref group); 0.79 (0.47-1.32); 0.73 (0.36-1.47); 0.95 (0.38-2.38)</p> <p><b>Notes:</b> Results also presented for more specific defects, however CO showed no association (see paper Table 3.). CO not associated with orofacial clefts)</p>
<p><b>Author:</b> Ritz et al. (2006)</p> <p><b>Period of Study:</b> 1989 - 2000</p> <p><b>Location:</b> Southern California</p>	<p><b>Health Outcome:</b> Post-neonatal mortality (28 days to 1 yr); All causes; SIDS</p> <p><b>Study Design:</b> Case-control</p> <p><b>Statistical Analyses:</b> Conditional Logistic regression</p> <p><b>Sample Description:</b> Mothers residing within 16 km of monitoring site</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 1.63 ppm</p> <p><b>Range (Min, Max):</b> 0.38, 3.44</p> <p><b>Copollutant: correlation</b>  PM<sub>10</sub>: r = 0.33  NO<sub>2</sub>: r = 0.72  O<sub>3</sub>: r = -0.57</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>OR for Post-neonatal death [Lower CI, Upper CI]</b>  Exposure period : 2 weeks prior to death, 1 month prior to death, 2 months prior to death, 6 months prior to death</p> <p><b>All causes:</b>  2 weeks prior to death : 1.14 (1.03-1.25)  2 months prior to death : 1.11 (1.06-1.16)</p> <p><b>SIDS:</b>  2 months prior to death : 1.19 (1.10-1.28)</p> <p><b>Term/normal weight births 2 months prior to death:</b>  All causes: 1.12 (1.05-1.19)  SIDS : 1.17 (1.07-1.29)  Respiratory : 1.14 (0.95-1.36)</p> <p><b>Preterm &amp;/or LBW births 2 months prior to death:</b>  All causes: 1.12 (1.01-1.25)  SIDS : 1.46 (1.09-1.94)  Respiratory : 1.03 (0.83-1.27)</p> <p><b>Notes:</b> These results did not persist in multipollutant models (CO, NO<sub>2</sub>, PM<sub>10</sub>, O<sub>3</sub>)</p>
<p><b>Author:</b> Ritz et al. (2007)</p> <p><b>Period of Study:</b> January – December 2003</p> <p><b>Location:</b> Los Angeles, CA</p>	<p><b>Health Outcome:</b> PTB</p> <p><b>Study Design:</b> Nested case-control</p> <p><b>Statistical Analyses:</b> Logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> A survey of 2,543 of 6,374 women sampled from a cohort of 58,316 eligible births in Los Angeles county.</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Copollutant correlation:</b>  TSP: r = 0.73  NO<sub>2</sub>: r = 0.75  SO<sub>2</sub>: r = 0.82  O<sub>3</sub>: r = -0.39</p>	<p><b>Increment:</b> Exposure categories (ppm): Less than 0.58: 0.59-0.91; 0.92-1.25; &gt;1.25</p> <p><b>RR for LBW [Lower CI, Upper CI]</b></p> <p><b>First trimester:</b> 1.00 (Ref group); 1.17 (1.08-1.26); 1.15 (1.05-1.26); 1.25 (1.12-1.38)</p> <p><b>6 weeks prior to birth</b> 1.00 (Ref group); 1.00 (0.93-1.08); 1.08 (0.98-1.20); 1.03 (0.93-1.14)</p> <p><b>Entire pregnancy:</b> 1.00 (Ref group); 0.76 (0.70-0.82); 0.84 (0.77-0.91); 1.03 (0.91-1.17)</p>
<p><b>Author:</b> Salam et al. (2005)</p> <p><b>Period of Study:</b> 1975 - 1987</p> <p><b>Location:</b> California</p>	<p><b>Health Outcome:</b> Birth weight, LBW, IUGR</p> <p><b>Study Design:</b> Retrospective cohort</p> <p><b>Statistical Analyses:</b> Linear and logistic regression</p> <p><b>Age Groups Analyzed:</b> NA</p> <p><b>Sample Description:</b> 3,901 infants from the California Children's Health Study</p>	<p><b>Averaging Time:</b> 24-h</p> <p><b>Mean (SD) unit:</b> 1.8 ppm (0.9) (Entire pregnancy)</p> <p><b>Range:</b> NR</p> <p><b>Copollutant: correlation</b>  PM<sub>10</sub>: r = 0.41  NO<sub>2</sub>: r = 0.69  O<sub>3</sub>: r = -0.27</p>	<p><b>Increment:</b> Entire pregnancy 1.2 ppm</p> <p>Trimesters: First : 1.4 ppm; Second : 1.4 ppm; Third : 1.3 ppm</p> <p><b>Regression co-efficient for birth weight (g) [Lower CI, Upper CI]</b>  Trimesters:  First: -21.7 (-42.3 to -1.1); Second: 11.3 (-9.7 to 32.3); Third : 11.8 (-8.4 to 32.1);  Entire pregnancy: 2.2 (-20.1 to 24.4)</p> <p><b>OR for LBW [Lower CI, Upper CI]</b>  Trimesters:  First: 1.0 (0.7-1.5); Second: 0.9 (0.6-1.3); Third: 0.7 (0.5-1.1); Entire pregnancy: 0.8 (0.6-1.3)</p> <p><b>OR for IUGR [Lower CI, Upper CI]</b>  Trimesters:  First: 1.2 (1.0-1.4); Second: 1.0 (0.9-1.1); Third: 1.0 (0.8-1.1); Entire pregnancy: 1.0 (0.9-1.2)</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> Tsai et al. (2006b) <b>Period of Study:</b> 1994 - 2000 <b>Location:</b> Kaoshiung, Taiwan	<b>Health Outcome:</b> Postneonatal death (27 days-1 yr old) <b>Study Design:</b> Case- crossover <b>Statistical Analyses:</b> Poisson regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 8.27 ppm x10 <b>Range (Min, Max):</b> 2.26, 17.7 <b>Copollutant:</b> NR	<b>Increment:</b> Interquartile range : 0.31 ppm <b>OR for Post-neonatal mortality [Lower CI, Upper CI]</b> Lag examined : 0-2 Lag 0-2: 1.051 (0.304-3.630)
<b>Author:</b> Wilhelm et al. (2005) <b>Period of Study:</b> 1994 - 2000 <b>Location:</b> Los Angeles, CA	<b>Health Outcome:</b> Term LBW and PTB <b>Study Design:</b> Retrospective cohort <b>Statistical Analyses:</b> Logistic regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> 518,254 births within 4 miles of a monitoring station. Varied according to analyses.	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> Trimester 1: 1.42 ppm Results for third trimester and 6 weeks prior to birth were similar to first trimester <b>Range (Min, Max):</b> 0.26, 2.82 <b>Copollutant correlation:</b> First Trimester: PM <sub>10</sub> : r = 0.12 PM <sub>2.5</sub> : r = 0.57 NO <sub>2</sub> : r = 0.81 SO <sub>2</sub> : r = -0.31	<b>Increment:</b> 1 ppm <b>RR for PTB [Lower CI, Upper CI]</b> <b>First trimester:</b> Less than 1 mile: 1.06 (1.00-1.12) 1-2 miles: 1.06 (1.03-1.10) 2-4 miles: 1.08 (1.06-1.09) ZIP code level: 1.04 (1.01-1.07) <b>6 weeks prior to birth:</b> Less than 1 mile: 1.04 (0.98-1.09) 1-2 miles: .04 (1.01-1.08) 2-4 miles: 1.01 (0.99-1.02) ZIP code level: 1.03 (1.00-1.06) <b>Notes:</b> All results above did not persist in multipollutant model (CO, NO <sub>2</sub> , O <sub>3</sub> , PM <sub>10</sub> ) <b>OR for term LBW [Lower CI, Upper CI]</b> <b>Third trimester:</b> Less than 1 mile: 1.10 (0.98-1.23) 1-2 miles: 1.05 (0.99-1.13) 2-4 miles: 1.06 (1.02-1.10) ZIP code level: 1.12 (1.05-1.19) <b>Notes:</b> All results above did not persist in multipollutant model (CO, NO <sub>2</sub> , O <sub>3</sub> , PM <sub>10</sub> ) See paper for results based on exposure category groupings.
<b>Author:</b> Woodruff et al. (2008) <b>Period of Study:</b> 1999 - 2002 <b>Location:</b> U.S. counties with >250,000 residents	<b>Health Outcome:</b> Post-neonatal deaths All causes; respiratory; SIDS; ill-defined + SIDS; other causes. <b>Study Design:</b> Retrospective cohort <b>Statistical Analyses:</b> Logistic regression (GEE) <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> All causes: 0.70 ppm <b>Range (Min, Max):</b> Percentiles: 25th: 0.48; 75th: 0.87 <b>Copollutant correlation:</b> PM <sub>10</sub> : r = 0.18 SO <sub>2</sub> : r = 0.27 O <sub>3</sub> : r = -0.46	<b>Increment:</b> 0.39 ppm <b>OR for Post-neonatal mortality [Lower CI, Upper CI]</b> Avg exposure over the first 2 months of life: All causes: 1.01 (0.95-1.07) Respiratory: 1.14 (0.93-1.40) SIDS: 0.88 (0.76-1.03) Ill-defined + SIDS: 0.93 (0.84-1.02) Other causes: 1.02 (0.97-1.07)
<b>Author:</b> Yang et al. (2004b) <b>Period of Study:</b> 1994 - 2000 <b>Location:</b> Taipei, Taiwan	<b>Health Outcome:</b> Post-neonatal mortality (27 days-1 yr old) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Poisson regression <b>Age Groups Analyzed:</b> NA <b>Sample Description:</b> NR	<b>Averaging Time:</b> 24-h <b>Mean (SD) unit:</b> 15.8 ppm x10 <b>Range (Min, Max):</b> 3.20, 48.4 <b>Copollutant:</b> NR	<b>Increment:</b> Interquartile range: 0.56 ppm <b>OR for Post-neonatal mortality [Lower CI, Upper CI]</b> Lag examined : 0-2 Lag 0-2: 1.038 (0.663-1.624)



**Table C-4. Studies of short-term CO exposure and respiratory morbidity**

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Chen et al. (1999)  <b>Period of Study:</b> 5/1995 – 1/1996  <b>Location:</b> 3 Taiwan communities</p>	<p><b>Health Outcome:</b> Lung function (FVC, FEV<sub>1</sub>, FEV<sub>1</sub>/FVC, FEF<sub>25-75</sub>%, PEF)  <b>Study Design:</b> Cross-sectional survey  <b>Statistical Analyses:</b> Multivariate linear model  <b>Population:</b> 941 children (Boys: 453; Girls: 488)  <b>Age Groups Analyzed:</b> 8-13</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 1-h maximum; 24-h avg  <b>Mean (SD) unit:</b> NR  <b>Range (Min, Max):</b> 1-h maximum: (0.4, 3.6)  <b>Copollutant correlation:</b> NO<sub>2</sub>: r = 0.86 – 0.98  <b>Note:</b> To represent the schoolchildren's exposure the daytime avg and peak concentrations were measured from 0800 to 1800.</p>	<p><b>Increment:</b> NR  <b>β Coefficient (SE); lag:</b>  <b>FVC (mL)</b>  <b>24-h avg</b>  -66.6 (40.73); 1  -147.71 (64.48); 2  2.2 (48.13); 7  <b>1-h maximum</b>  -33.25 (20.74); 1  -16.48 (19.67); 2  -5.18 (16.48); 7  <b>FEV<sub>1</sub> (mL)</b>  <b>24-h avg</b>  20.55 (38.24); 1  -82.42 (60.95); 2  48.23 (45.58); 7  <b>1-h maximum</b>  1.2 (19.48); 1  -1.44 (18.57); 2  20.96 (15.67); 7</p>
<p><b>Author:</b> Chen et al. (2000)  <b>Period of Study:</b> 8/1996 – 6/1998  <b>Location:</b> Washoe County, NV</p>	<p><b>Health Outcome:</b> School absenteeism  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> Maximum likelihood  <b>Population:</b> 1st to 6th grade children: 27,793  <b>Age Groups Analyzed:</b> 1st to 6th grade children</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 1-h maximum  <b>Mean (SD) unit:</b> 2.73 (1.154) ppm  <b>Range (Min, Max):</b> (0.65, 2.73)  <b>Copollutant correlation:</b> PM<sub>10</sub>: r = 0.721  O<sub>3</sub>: r = -0.204</p>	<p><b>Increment:</b> 1.0 ppm  <b>% Increase (Lower CI, Upper CI); lag:</b>  3.79% (1.04-6.55); 0</p>
<p><b>Author:</b> de Hartog et al. (2003)  <b>Period of Study:</b> 1998 – 1999  <b>Location:</b> Amsterdam, Netherlands; Erfurt, Germany; Helsinki, Finland</p>	<p><b>Health Outcome:</b> Respiratory symptoms (shortness of breath, being awakened by breathing problems, phlegm, wheezing, tripping heart)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> Logistic regression  <b>Population:</b> Non-smoking individuals with CHD: Amsterdam: 37  Erfurt: 47  Helsinki: 47  <b>Age Groups Analyzed:</b> ≥ 50</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 24-h avg  <b>Mean (SD) unit:</b> Amsterdam: 0.6 mg/m<sup>3</sup>  Erfurt: 0.4 mg/m<sup>3</sup>  Helsinki: 0.4 mg/m<sup>3</sup>  <b>Range (Min, Max):</b> Amsterdam: (0.4, 1.6)  Erfurt: (0.1, 2.5)  Helsinki: (0.1, 1.0)  <b>Copollutant:</b> PM<sub>2.5</sub>; NO<sub>2</sub></p>	<p><b>Increment:</b> 0.25 mg/m<sup>3</sup>  <b>Odds Ratio (Lower CI, Upper CI); lag:</b>  <b>Incidence of symptoms</b>  <b>Shortness of breath</b>  1 (0.92-1.1); 0  0.96 (0.88-1.05); 1  1 (0.92-1.09); 2  1.07 (0.98-1.16); 3  1.03 (0.9-1.18); 0-4  <b>Being awakened by breathing problems</b>  1.02 (0.92-1.14); 1  1.03 (0.93-1.15); 2  1.11 (1-1.22); 3  1.16 (0.98-1.37); 0-4  <b>Phlegm</b>  1.05 (0.93-1.19); 0  1.02 (0.91-1.14); 1  1.08 (0.96-1.22); 2  1.09 (0.97-1.22); 3  1.13 (0.94-1.35); 0-4  <b>Prevalence of symptoms</b>  <b>Shortness of breath</b>  1 (0.94-1.06); 0  0.99 (0.94-1.05); 1  0.99 (0.93-1.05); 2  1.01 (0.95-1.07); 3  0.98 (0.9-1.07); 0-4  <b>Being awakened by breathing problems</b>  1.01 (0.93-1.1); 1  0.99 (0.91-1.08); 2  1.1 (1.02-1.19); 3  1.13 (1-1.29); 0-4</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Delfino et al. (2003)  <b>Period of Study:</b> 11/1999 – 1/2000  <b>Location:</b> Los Angeles, CA</p>	<p><b>Health Outcome:</b> Asthma symptoms (Cough, wheeze, sputum production, shortness of breath, chest tightness) (symptom scores &gt;1, symptoms scores &gt;2); Lung function (PEF)  <b>Study Design:</b> Panel study  <b>Statistical Analyses:</b> Asthma symptoms: GEE  Lung function: Generalized linear mixed model  <b>Population:</b> 22 asthmatic Hispanic children  <b>Age Groups Analyzed:</b> 10-15</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 1-h maximum; 8-h maximum  <b>Mean (SD) unit:</b> 1-h maximum: 7.7 (3.1) ppb  8-h maximum: 5.0 (2.0) ppb  <b>Range (Min, Max):</b> 1-h maximum: (2, 17)  8-h maximum: (1, 10)  <b>Copollutant correlation:</b>  NO<sub>2</sub>: r = 0.65; O<sub>3</sub>: r = -0.17;  Acetaldehyde: r = 0.51;  Acetone: r = 0.28;  Formaldehyde: r = 0.41;  Benzene: r = 0.50;  Ethylbenzene: r = 0.62;  Tetrachloroethylene: r = 0.63;  Toluene: r = 0.71;  m,p - Xylene: r = 0.72;  PM<sub>10</sub>: r = 0.50;  EC: r = 0.60;  OC: r = 0.55;  SO<sub>2</sub>: r = 0.69</p>	<p><b>Increment:</b> 5.0 ppb &amp; 3.0 ppb  <b>Odds Ratio (Lower CI, Upper CI); lag:</b>  <b>1-maximum</b>  <b>Increment:</b> 5.0 ppb  <b>Symptom scores &gt;1</b>  0.95 (0.52-1.75); 0  1.11 (0.75-1.65); 1  <b>Symptom scores &gt;2</b>  0.48 (0.07-3.53); 0  .28 (0.53-3.12); 1  <b>8-h maximum</b>  <b>Increment:</b> 3.0 ppb  <b>Symptom scores &gt;1</b>  0.95 (0.55-1.62); 0  1.2 (0.77-1.86); 1  <b>Symptom scores &gt;2</b>  0.53 (0.10-2.92); 0  1.43 (0.41-5.00); 1</p>
<p><b>Author:</b> Estrella et al. (2005)  <b>Period of Study:</b> 1/2000 – 4/2000  <b>Location:</b> Quito, Ecuador</p>	<p><b>Health Outcome:</b> Acute respiratory infection  <b>Study Design:</b> Prospective study  <b>Statistical Analyses:</b> Logistic regression; Poisson  <b>Population:</b> 960 children  <b>Age Groups Analyzed:</b> 6-11</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> NR  <b>Mean (SD) unit:</b> NR  <b>Range (Min, Max):</b> NR  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR  <b>Odds Ratio (Lower CI, Upper CI); lag:</b>  <b>Acute respiratory infection</b>  ARI in children COHb &gt;2.5% vs. COHb &lt;2.5%:  Adjusted Logistic Regression Model  3.25 (1.65-6.38)  ARI in children COHb &gt;2.5% vs. COHb &lt;2.5%:  Crude Logistic Regression Model  2.06 (1.30-3.20)  Log-Linear Model (Each Percent Increase in COHb above 2.5%)  1.15 (1.03-1.28)</p>
<p><b>Author:</b> Fischer et al. (2002)  <b>Period of Study:</b> NR  <b>Location:</b> Utrecht, Netherlands</p>	<p><b>Health Outcome:</b> Lung function (FVC, FEV<sub>1</sub>, PEF, MMEF)  <b>Study Design:</b> Panel study  <b>Statistical Analyses:</b> Restricted maximum likelihood linear model  <b>Population:</b> 68 children  <b>Age Groups Analyzed:</b> 10-11</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 24-h avg  <b>Mean (SD) unit:</b> 921 µg/m<sup>3</sup>  <b>Range (Min, Max):</b> (319, 1540)  <b>Copollutant:</b> PM<sub>10</sub>; BS; NO<sub>2</sub>; NO</p>	<p><b>Increment:</b> 100 µg/m<sup>3</sup>  <b>mL (SE); lag:</b>  <b>FVC:</b> 0.5 (0.4); 1; 0.1 (0.2); 2  <b>FEV<sub>1</sub>:</b> -0.4 (0.5); 1; -0.2 (0.2); 2  <b>m/s (SE); lag:</b>  <b>PEF:</b> -1.1 (2.8); 1; -0.6 (1.1); 2  <b>MMEF:</b> -0.5 (1.4); 1; -0.3 (0.6); 2</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Lagorio et al. (2006)</p> <p><b>Period of Study:</b> 5/1999 – 6/1999; 11/1999 – 12/1999</p> <p><b>Location:</b> Rome, Italy</p>	<p><b>Health Outcome:</b> Lung function (FVC, FEV<sub>1</sub>)</p> <p><b>Study Design:</b> Time-series panel study</p> <p><b>Statistical Analyses:</b> Generalized estimating equations (GEE)</p> <p><b>Population:</b> COPD panel: 11 Asthma panel: 11 IHD panel: 7</p> <p><b>Age Groups Analyzed:</b> COPD panel: 50-80 Asthma panel: 18-64 IHD panel: 40-64</p> <p><b>Notes:</b> Asthma panel was restricted to never smokers, while COPD and IHD panels include former smokers if smoking cessation occurred at least 1 year prior to enrollment.</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> Overall: 7.4 (6.2) mg/m<sup>3</sup> Spring: 2.1 (0.3) mg/m<sup>3</sup> Winter: 12.3 (4.9) mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> Overall: (1.6, 28.9)</p> <p><b>Copollutant correlation:</b> PM<sub>2.5</sub>: r = 0.67 PM<sub>10-2.5</sub>: r = -0.09 PM<sub>10</sub>: r = 0.55 NO<sub>2</sub>: r = 0.05 O<sub>3</sub>: r = -0.87 SO<sub>2</sub>: r = 0.65</p>	<p><b>Increment:</b> 1 mg/m<sup>3</sup></p> <p><b>β Coefficient (SE); lag:</b></p> <p><b>COPD panel</b> FVC (% of predicted) -0.14 (0.15); 0 -0.13 (0.18); 0-1 0.15 (0.23); 0-2 FEV<sub>1</sub> (% of predicted) -0.05 (0.13); 0 -0.12 (0.16); 0-1 -0.03 (0.2); 0-2 <b>Asthma panel</b> FVC (% predicted) 0.02 (0.12); 0 -0.001 (0.13); 0-1 -0.06 (0.16); 0-2 FEV<sub>1</sub> (% predicted) -0.05 (0.14); 0 -0.16 (0.15); 0-1 -0.28 (0.18); 0-2 <b>IHD panel</b> FVC (% of predicted) 0.176 (0.101); 0 0.132 (0.120); 0-1/ 0.132 (0.165); 0-2 FEV<sub>1</sub> (% of predicted) 0.204 (0.120); 0 0.114 (0.142); 0-1 0.159 (0.194); 0-2</p>
<p><b>Author:</b> Park et al. (2002)</p> <p><b>Period of Study:</b> 3/1996 – 12/1999</p> <p><b>Location:</b> Seoul, Korea</p>	<p><b>Health Outcome:</b> School absenteeism</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Population:</b> ~ 1,264 children (671 Boys, 593 girls)</p> <p><b>Age Groups Analyzed:</b> 1st through 6th grade students</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.11 (0.40) ppm</p> <p><b>Range (Min, Max):</b> (0.39, 2.97)</p> <p><b>Copollutant correlation:</b> PM<sub>10</sub>: r = 0.56; NO<sub>2</sub>: r = 0.70; SO<sub>2</sub>: r = 0.67; O<sub>3</sub>: r = -0.46</p>	<p><b>Increment:</b> 0.52 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>Total Absences:</b> 0.95 (0.94-0.97); 0 <b>Non-Illness Related Absences:</b> 0.99 (0.96-1.02); 0 <b>Illness-Related Absences:</b> 0.96 (0.94-0.98); 0</p>
<p><b>Author:</b> Park et al. (2005a)</p> <p><b>Period of Study:</b> 3/2002 – 6/2002</p> <p><b>Location:</b> Incheon, Korea</p>	<p><b>Health Outcome:</b> Lung function (PEF variability (&gt;20%), Mean PEF); Respiratory symptoms (night respiratory symptoms, cough, inhaler use)</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> GEE; Poisson GAM</p> <p><b>Population:</b> 64 bronchial asthmatics</p> <p><b>Age Groups Analyzed:</b> 16-75</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> Control days: 0.6368 (0.1522) ppm Dust days: 0.6462 (0.0945) ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p>PEF variability (&gt;20%): 1.43 (0.54-3.75) <b>Night respiratory symptoms:</b> 0.98 (0.51-1.86)</p> <p><b>β Coefficient (SE); lag:</b> PEF variability (&gt;20%): 0.9737 (0.3187) Mean PEF (L/min): -10.103 (2.7146) <b>Night respiratory symptoms:</b> -0.018 (0.3654) <b>Cough:</b> 0.0855 (0.1826) <b>Inhaler Use:</b> 0.0796 (0.1733)</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Penttinen et al. (2001)</p> <p><b>Period of Study:</b> 11/1996 – 4/1997</p> <p><b>Location:</b> Helsinki, Finland</p>	<p><b>Health Outcome:</b> Lung function (PEF)</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> First order autoregressive linear model</p> <p><b>Population:</b> 57 non-smoking adult asthmatics</p> <p><b>Age Groups Analyzed:</b> NR</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Median unit:</b> 0.4 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> (0.1, 1.1) mg/m<sup>3</sup></p> <p><b>Copollutant correlation:</b>  PM<sub>10</sub>: r = -0.03  PM<sub>10-2.5</sub>: r = -0.30  PM<sub>2.5</sub>: r = 0.32  PM<sub>1</sub>: r = 0.39  PNC: r = 0.44  NC0.01-0.1: r = 0.43  NC0.1-1: r = 0.47  NO: r = 0.60  NO<sub>2</sub>: r = 0.44</p>	<p><b>Increment:</b> 0.2 mg/m<sup>3</sup></p> <p><b>β Coefficient (SE); lag:</b></p> <p><b>PEF Deviations (L/min)</b></p> <p><b>Morning</b>  0.27 (0.38); 0  -1.08 (0.36); 1  0.23 (0.38); 2  -1.11 (1.19); 5-day avg</p> <p><b>Afternoon</b>  -0.4 (0.43); 0  -0.13 (0.41); 1  -0.71 (0.41); 2  -3.03 (1.06); 5-day avg</p> <p><b>Evening</b>  -0.7 (0.45); 0;  -0.31 (0.44); 1  0.3 (0.44); 2  -3.62 (1.19); 5-day avg</p> <p><b>Co-pollutant models with PNC</b></p> <p><b>Morning:</b> -0.67 (0.64); 1  <b>Afternoon:</b> -0.46 (0.69); 0  <b>Evening:</b> -0.46 (0.73); 0</p>
<p><b>Author:</b> Rabinovitch et al. (2004)</p> <p><b>Period of Study:</b> 11/1999 – 3/2000; 11/2000 – 3/2001; 11/2001 – 3/2002</p> <p><b>Location:</b> Denver, CO</p>	<p><b>Health Outcome:</b> Lung function (FEV<sub>1</sub>); asthma exacerbation; bronchodilator use</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> Pulmonary function: Mixed effects model; Asthma exacerbation and medication use: GLM</p> <p><b>Population:</b> Urban poor asthmatic children:  1999-2000: 41  2000-2001: 63  2001-2002: 43</p> <p><b>Age Groups Analyzed:</b> 6-12</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.0 (0.4) ppm</p> <p><b>Range (Min, Max):</b> (0.3, 3.5)</p> <p><b>Copollutant:</b> PM<sub>2.5</sub>; PM<sub>10</sub>; NO<sub>2</sub>; SO<sub>2</sub>; O<sub>3</sub></p>	<p><b>Increment:</b> 0.4 ppm</p> <p><b>β Coefficient (SE); lag:</b> FEV<sub>1</sub> AM: -0.001 (0.008); 3-day ma  PM: 0.015 (0.01); 3-day ma</p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p><b>Asthma exacerbation:</b> 1.012 (0.913-1.123); 3-day ma</p> <p><b>Bronchodilator use:</b> 1.065 (1.001-1.133); 3-day ma</p>
<p><b>Author:</b> Ranzi et al. (2004)</p> <p><b>Period of Study:</b> 2/1999 – 5/1999</p> <p><b>Location:</b> Emilia-Romagna, Italy</p>	<p><b>Health Outcome:</b> Lung function; respiratory symptoms, medication use</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> GLM</p> <p><b>Population:</b> 120 "asthma-like" school children</p> <p><b>Age Groups Analyzed:</b> 6-11</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> Urban area: 1.54 mg  Rural area: 1.22 mg</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NO<sub>2</sub>; TSP; PM<sub>2.5</sub></p>	<p><b>The study did not present quantitative results for CO.</b></p>
<p><b>Author:</b> Rodriguez et al. (2007)</p> <p><b>Period of Study:</b> 1996 – 2003</p> <p><b>Location:</b> Perth, Australia</p>	<p><b>Health Outcome:</b> Respiratory symptoms (body temperature, cough, wheeze/rattle chest, runny/blocked nose)</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> Logistic regression, GEE</p> <p><b>Population:</b> 263 children at high risk of developing asthma</p> <p><b>Age Groups Analyzed:</b> 0-5</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 8-h avg</p> <p><b>Mean (SD) unit:</b> 1.408 ppm</p> <p><b>Range (Min, Max):</b> (0.012, 8.031)</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p><b>Body Temperature</b>  1.024 (0.911-1.151); 0  1.056 (0.943-1.184); 5  0.991 (0.962-1.021); 0-5</p> <p><b>Cough</b>  1.001 (0.996-1.005); 0  1.064 (0.941-1.02); 5  1.028 (0.996-1.061); 0-5</p> <p><b>Wheeze/Rattle Chest</b>  1.089 (0.968-1.226); 0  1.136 (1.016-1.26); 5  1.035 (1.005-1.066); 0-5</p> <p><b>Runny/Blocked Nose</b>  1.094 (0.824-1.453); 0  1.38 (1.028-1.853); 5  1.101 (1.025-1.183); 0-5</p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)
<p><b>Author:</b> Schildcrout et al. (2006)</p> <p><b>Period of Study:</b> 11/1993 – 9/1995</p> <p><b>Location:</b> 8 North American cities: Albuquerque, NM; Baltimore, MD; Boston, MA; Denver, CO; San Diego, CA; Seattle, WA; St. Louis, MO; Toronto, ON, Canada</p>	<p><b>Health Outcome:</b> Asthma symptoms; rescue inhaler use</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> Asthma symptoms: Logistic regression; Rescue Inhaler Use: Poisson regression</p> <p><b>Population:</b> 990 asthmatic children</p> <p><b>Age Groups Analyzed:</b> 5-12</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NO<sub>2</sub>; O<sub>3</sub>; PM<sub>10</sub>; SO<sub>2</sub></p>	<p><b>Increment:</b> 1.0 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p><b>Asthma Symptoms</b></p> <p>1.08 (1.01-1.14); 0</p> <p>1.07 (0.99-1.16); 1</p> <p>1.08 (1.02-1.15); 2</p> <p>1.05 (1.01-1.09); 0-2</p> <p><b>Asthma Symptoms</b></p> <p><b>+ 20 ppb increase in NO<sub>2</sub></b></p> <p>1.07 (1-1.14); 0</p> <p>1.04 (0.96-1.11); 1</p> <p>1.09 (1.02-1.16); 2</p> <p>1.04 (1-1.08); 0-2</p> <p><b>+ 25 µg/m<sup>3</sup> increase in PM<sub>10</sub></b></p> <p>1.08 (1.01-1.15); 0</p> <p>1.06 (0.99-1.14); 1</p> <p>1.08 (1.02-1.14); 2</p> <p>1.05 (1.01-1.08); 0-2</p> <p><b>+ 10 ppb increase in SO<sub>2</sub></b></p> <p>1.07 (0.99-1.16); 0</p> <p>1.06 (0.96-1.19); 1</p> <p>1.1 (1.02-1.18); 2</p> <p>1.05 (1-1.09); 0-2</p> <p><b>Rescue Inhaler Use</b></p> <p>1.07 (1.01-1.13); 0</p> <p>1.05 (0.99-1.1); 1</p> <p>1.06 (1.01-1.1); 2</p> <p>1.04 (1.01-1.07); 0-2</p> <p><b>Rescue Inhaler Use</b></p> <p><b>+ 20 ppb increase in NO<sub>2</sub></b></p> <p>1.05 (0.99-1.12); 0</p> <p>1.04 (0.98-1.11); 1</p> <p>1.07 (1.02-1.12); 2</p> <p>1.04 (1-1.07); 0-2</p> <p><b>+ 25 µg/m<sup>3</sup> increase in PM<sub>10</sub></b></p> <p>1.06 (0.99-1.13); 0</p> <p>1.05 (0.99-1.11); 1</p> <p>1.05 (1.01-1.09); 2</p> <p>1.03 (1-1.07); 0-2</p> <p><b>+ 10 ppb increase in SO<sub>2</sub></b></p> <p>1.04 (0.96-1.12); 0</p> <p>1.04 (0.97-1.1); 1</p> <p>1.08 (1.03-1.13); 2</p> <p>1.04 (1-1.08); 0-2</p>
<p><b>Author:</b> Silkoff et al. (2005)</p> <p><b>Period of Study:</b> 11/11/1999 – 3/31/2000; 11/1/2000 – 3/16/2001</p> <p><b>Location:</b> Denver, CO</p>	<p><b>Health Outcome:</b> Lung function (FEV<sub>1</sub>, PEF); recorded symptoms; rescue medication use</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> Rescue medication use and total symptom score: GEE; Lung function: Mixed effects model</p> <p><b>Population:</b> 1st winter: 16 with a history of more than 10 pack years of tobacco use, airflow limitation with FEV<sub>1</sub> of less than 70% of predicted value, and FEV<sub>1</sub>/FVC ratio of less than 60% 2nd winter: 18 with a history of more than 10 pack years of tobacco use, airflow limitation with FEV<sub>1</sub> of less than 70% of predicted value, and FEV<sub>1</sub>/FVC ratio of less than 60%</p> <p><b>Age Groups Analyzed:</b> ≥ 40</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1999-2000: 1.2 (0.555) ppm; 2000-2001: 1.1 (0.5) ppm</p> <p><b>Range (Min, Max):</b> 1999-2000: (0.340, 3.790); 2000-2001: (0.360, 2.810)</p> <p><b>Copollutant:</b> NR</p>	<p><b>The study did not present quantitative results for CO.</b></p>

Study	Design	Concentrations	CO Effect Estimates (95% CI)																																																
<b>Author:</b> Slaughter et al. (2003) <b>Period of Study:</b> 12/1994 – 8/1995 <b>Location:</b> Seattle, WA	<b>Health Outcome:</b> Asthma severity; medication use <b>Study Design:</b> Panel study <b>Statistical Analyses:</b> Asthma severity: Ordinal logistic regression; Medication use: Poisson <b>Population:</b> 133 mild-to-moderate asthmatic children <b>Age Groups Analyzed:</b> 5-13	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Median unit:</b> 1.47 ppm <b>IQR (25th, 75th):</b> (0.23, 1.87) <b>Copollutant:</b> NR	<b>Increment:</b> Increased asthma attack severity: 0.67 ppm Increased rescue inhaler use: 1.0 ppm <b>Odds Ratio (Lower CI, Upper CI); lag:</b> Increased asthma attack severity: Without transition: 1.21; 1 With transition: 1.17; 1 Increased rescue inhaler use: Without transition: 1.09 (1.03-1.16); 1 With transition: 1.06 (1.01-1.1); 1																																																
<b>Author:</b> Steerenberg et al. (2001) <b>Period of Study:</b> NR <b>Location:</b> Bilthoven and Utrecht, the Netherlands	<b>Health Outcome:</b> Lung function (PEF); exhaled nitric oxide; inflammatory nasal markers <b>Study Design:</b> Panel study <b>Statistical Analyses:</b> Restricted maximum likelihood linear model <b>Population:</b> 126 children <b>Age Groups Analyzed:</b> 8-13 <b>Notes:</b> The study was only conducted for a two month period: February and March.	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> Utrecht: 0.8 mg/m <sup>3</sup> Bilthoven: 0.5 mg/m <sup>3</sup> <b>Range (Min, Max):</b> Utrecht: (0.3, 2.3) Bilthoven: (0.3, 0.9) <b>Copollutant:</b> NR	<b>The study did not present quantitative results for CO.</b>																																																
<b>Author:</b> Timonen et al. (2002) <b>Period of Study:</b> 2/1994 – 4/1994 <b>Location:</b> Kuopio, Finland	<b>Health Outcome:</b> Exercise induced bronchial responsiveness; Lung function (FVC, FEV <sub>1</sub> , MMEF, AEFV) <b>Study Design:</b> Panel study <b>Statistical Analyses:</b> Linear regression <b>Population:</b> 33 children with chronic respiratory symptoms <b>Age Groups Analyzed:</b> 7-12	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.6 mg/m <sup>3</sup> <b>Range (Min, Max):</b> (0.1, 2.8) <b>Copollutant correlation:</b> PM <sub>10</sub> : r = 0.52 BS: r = 0.80 PNC <sub>0.01-0.03</sub> : r = 0.81 PNC <sub>0.03-0.1</sub> : r = 0.87 PNC <sub>0.1-0.3</sub> : r = 0.71 PNC <sub>0.3-1.0</sub> : r = 0.60 PNC <sub>1.0-3.2</sub> : r = 0.84 PNC <sub>3.2-10</sub> : r = 0.79 NO <sub>2</sub> : r = 0.85	<b>Increment:</b> 0.32 mg/m <sup>3</sup> <b>β Coefficient (SE); lag:</b> <b>Exercise induced responsiveness</b> <table border="0"> <tr> <td><b>ΔFEV<sub>1</sub> (%)</b></td> <td><b>FEV<sub>1</sub> (mL)</b></td> </tr> <tr> <td>-0.081 (0.647); 0</td> <td>19.2 (13.2); 0</td> </tr> <tr> <td>0.03 (0.262); 1</td> <td>-9.04 (5.45); 1</td> </tr> <tr> <td>0.087 (0.26); 2</td> <td>-9.15 (5.21); 2</td> </tr> <tr> <td>-0.091 (0.275); 3</td> <td>-11.7 (5.77); 3</td> </tr> <tr> <td>0.19 (0.599); 0-3</td> <td>-17.5 (12.5); 0-3</td> </tr> <tr> <td><b>ΔMMEF (%)</b></td> <td><b>MMEF (mL/s)</b></td> </tr> <tr> <td>0.442 (1.79); 0</td> <td>22.2 (36.9); 0</td> </tr> <tr> <td>0.52 (0.723); 1</td> <td>-23 (15.2); 1</td> </tr> <tr> <td>0.313 (0.719); 2</td> <td>-4.63 (14.7); 2</td> </tr> <tr> <td>-0.616 (0.75); 3</td> <td>-30.9 (16); 3</td> </tr> <tr> <td>0.096 (1.64); 0-3</td> <td>-24.9 (34.8); 0-3</td> </tr> <tr> <td><b>ΔAEFV (%)</b></td> <td><b>AEFV (L2/s)</b></td> </tr> <tr> <td>0.287 (1.19); 0</td> <td>-0.093 (0.088); 0</td> </tr> <tr> <td>0.281 (0.482); 1</td> <td>-0.068 (0.036); 1</td> </tr> <tr> <td>0.904 (0.474); 2</td> <td>-0.06 (0.035); 2</td> </tr> <tr> <td>0.15 (0.483); 3</td> <td>-0.05 (0.039); 3</td> </tr> <tr> <td>1.6 (1.05); 0-3</td> <td>-0.076 (0.083); 0-3</td> </tr> <tr> <td><b>FVC (mL)</b></td> <td></td> </tr> <tr> <td>0.064 (10.9); 0</td> <td></td> </tr> <tr> <td>-4.79 (4.51); 1</td> <td></td> </tr> <tr> <td>-9.78 (4.24); 2</td> <td></td> </tr> <tr> <td>-13.9 (4.7); 3</td> <td></td> </tr> <tr> <td>-29.4 (10.1); 0-3</td> <td></td> </tr> </table>	<b>ΔFEV<sub>1</sub> (%)</b>	<b>FEV<sub>1</sub> (mL)</b>	-0.081 (0.647); 0	19.2 (13.2); 0	0.03 (0.262); 1	-9.04 (5.45); 1	0.087 (0.26); 2	-9.15 (5.21); 2	-0.091 (0.275); 3	-11.7 (5.77); 3	0.19 (0.599); 0-3	-17.5 (12.5); 0-3	<b>ΔMMEF (%)</b>	<b>MMEF (mL/s)</b>	0.442 (1.79); 0	22.2 (36.9); 0	0.52 (0.723); 1	-23 (15.2); 1	0.313 (0.719); 2	-4.63 (14.7); 2	-0.616 (0.75); 3	-30.9 (16); 3	0.096 (1.64); 0-3	-24.9 (34.8); 0-3	<b>ΔAEFV (%)</b>	<b>AEFV (L2/s)</b>	0.287 (1.19); 0	-0.093 (0.088); 0	0.281 (0.482); 1	-0.068 (0.036); 1	0.904 (0.474); 2	-0.06 (0.035); 2	0.15 (0.483); 3	-0.05 (0.039); 3	1.6 (1.05); 0-3	-0.076 (0.083); 0-3	<b>FVC (mL)</b>		0.064 (10.9); 0		-4.79 (4.51); 1		-9.78 (4.24); 2		-13.9 (4.7); 3		-29.4 (10.1); 0-3	
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Study	Design	Concentrations	CO Effect Estimates (95% CI)
<b>Author:</b> von Klot et al. (2002) <b>Period of Study:</b> 9/1996 – 3/1997 <b>Location:</b> Erfurt, Germany	<b>Health Outcome:</b> Asthma symptoms; medication use <b>Study Design:</b> Panel study <b>Statistical Analyses:</b> Logistic regression <b>Population:</b> 53 adults with asthma or asthma symptoms <b>Age Groups Analyzed:</b> 37-77	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.9 mg/m <sup>3</sup> <b>Range (Min, Max):</b> (0.3, 3.0) <b>Copollutant correlation:</b> NC <sub>0.01-0.1</sub> : r = 0.66 NC <sub>0.1-0.5</sub> : r = 0.79 NC <sub>0.5-2.5</sub> : r = 0.46 MC <sub>0.1-0.5</sub> : r = 0.66 MC <sub>0.01-2.5</sub> : r = 0.65 PM <sub>2.5-10</sub> : r = 0.42 PM <sub>10</sub> : r = 0.69 NO <sub>2</sub> : r = 0.82 SO <sub>2</sub> : r = 0.32	<b>Increment:</b> 0 and 5-day avg lag: 0.6 mg/m <sup>3</sup> 14-day avg lag: 0.54 mg/m <sup>3</sup> <b>Odds Ratio (Lower CI, Upper CI); lag:</b> <b>Prevalence: Inhaled β<sub>2</sub>-agonist use</b> 0.98 (0.93-1.03); 0 1.04 (0.97-1.12); 0-4 0.93 (0.86-1.01); 0-13 <b>Prevalence: Inhaled corticosteroid use</b> 1.05 (1-1.11); 0 1.25 (1.17-1.34); 0-4 1.06 (0.97-1.15); 0-13 <b>Prevalence: Wheezing</b> 1.03 (0.97-1.08); 0 1.13 (1.05-1.22); 0-4 1.14 (1.05-1.25); 0-13 <b>Co-pollutant models</b> <b>Inhaled β<sub>2</sub>-agonist use</b> <b>CO+MC0.01-2.5:</b> 1 (0.91-1.11); 0-4 <b>CO+NC0.01-0.1:</b> 1.01 (0.91-1.11); 0-4 <b>Inhaled corticosteroid use</b> <b>CO+MC0.01-2.5:</b> 0.89 (0.81-0.98); 0-13 <b>CO+NC: 0.01-0.</b> 1: 0.81 (0.72-0.91); 0-13 <b>Wheezing</b> <b>CO+MC0.01-2.5:</b> 1.15 (1.04-1.27); 0-4 <b>CO+NC0.01-0.1:</b> 1.09 (0.98-1.22); 0-4
<b>Author:</b> Yu et al. (2000) <b>Period of Study:</b> 11/1993 – 8/1995 <b>Location:</b> Seattle, Washington	<b>Health Outcome:</b> Asthma symptoms (Wheezing, coughing, chest tightness, shortness of breath) <b>Study Design:</b> Panel study <b>Statistical Analyses:</b> Repeated measures logistic regression models (GEE) <b>Population:</b> 133 mild-to-moderate asthmatics <b>Age Groups Analyzed:</b> 5-13	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.6 ppm <b>Range (Min, Max):</b> (0.65, 4.18) <b>Copollutant correlation:</b> PM <sub>1.0</sub> : r = 0.82 PM <sub>10</sub> : r = 0.86 SO <sub>2</sub> : r = 0.31	<b>Increment:</b> 1.0 ppm <b>Odds Ratio (Lower CI, Upper CI); lag:</b> <b>Marginal GEE</b> 1.22 (1.03-1.45); 0 1.3 (1.11-1.52); 1 1.26 (1.09-1.46); 2 <b>Transition GEE</b> 1.18 (1.02-1.37); 0 1.25 (1.1-1.42); 1 1.18 (1.04-1.33); 2

**Table C-5. Studies of short-term CO exposure and respiratory hospital admissions and ED visits.**

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Anderson et al. (2001) <b>Period of Study:</b> 10/1994-12/1996 <b>Location:</b> West Midlands; U.K.	<b>Hospital Admission</b> <b>Health Outcome (ICD9):</b> Respiratory Diseases Asthma (493) COPD (490-492, 494-496) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Regression with quasi-likelihood approach and GAM <b>Age Groups Analyzed:</b> All ages 0-14 15-64 ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h avg <b>Mean (SD) unit:</b> 0.8 (0.7) ppm <b>Range (Min, Max):</b> (0.2, 10) <b>Copollutant correlation:</b> PM <sub>10</sub> : r = 0.55; PM <sub>2.5</sub> : r = 0.54; PM <sub>2.5-10</sub> : r = 0.10; BS: r = 0.77; SO <sub>4</sub> <sup>2-</sup> : r = 0.17; NO <sub>2</sub> : r = 0.73; O <sub>3</sub> : r = -0.29; SO <sub>2</sub> : r = 0.49	<b>Increment:</b> 1.0 ppm <b>% Increase (Lower CI, Upper CI); lag:</b> <b>Respiratory Diseases</b> <b>Age Group</b> <b>All ages:</b> 0.3% (-1.10 to 1.70); 0-1 <b>0-14:</b> 1.50% (-0.60 to 3.60); 0-1 <b>15-64:</b> -0.70% (-3.60 to 2.30); 0-1 <b>≥ 65:</b> 0.00% (-2.10 to 2.10); 0-1 <b>Asthma</b> <b>Age Group</b> <b>0-14:</b> 3.90% (-0.50 to 8.50); 0-1 <b>15-64:</b> -4.90% (-10.60 to 1.10); 0-1 <b>COPD</b> <b>Age Group</b> <b>≥ 65:</b> 1.00% (-2.50 to 4.60); 0-1

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Andersen et al. (2007)</p> <p><b>Period of Study:</b> 1/1999-12/2004</p> <p><b>Location:</b> Copenhagen, Denmark</p>	<p><b>Design:</b> Hospital Admission</p> <p><b>Health Outcome (ICD10):</b> Respiratory diseases: Chronic bronchitis (J41-42), Emphysema (J43), COPD (J44), Asthma (J45), Status asthmaticus (J46), Pediatric asthma (J45), Pediatric asthmaticus (J46)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> 5-18; ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.3 (0.1) ppm</p> <p><b>IQR (25th, 75th):</b> (0.22, 0.34)</p> <p><b>Copollutant; correlation:</b> PM<sub>10</sub>: r = 0.45</p>	<p><b>Increment:</b> 0.12 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>Respiratory Disease</b></p> <p><b>Age Group: ≥ 65</b> CO: 1.024 (0.997-1.053); 0-4 CO, PM<sub>10</sub>: 1.001 (0.961-1.042); 0-4</p> <p><b>Asthma</b></p> <p><b>Age Group: 5-18</b> CO: 1.104 (1.018-1.198); 0-5 CO, PM<sub>10</sub>: 1.023 (0.911-1.149); 0-5</p>
<p><b>Author:</b> Atkinson et al. (1999)</p> <p><b>Period of Study:</b> 1/1992-12/1994</p> <p><b>Location:</b> London, U.K.</p>	<p><b>Design:</b> ED Visits</p> <p><b>Health Outcome (ICD9):</b> Respiratory complaints: wheezing, inhaler request, chest infection, chronic obstructive lung disease (COLD), difficulty breathing, cough, other respiratory complaints. e.g., croup, pleurisy, noisy breathing; Asthma (493)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson</p> <p><b>Age Groups Analyzed:</b> All ages 0-14 15-64 ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.8 (0.4) ppm</p> <p><b>Range (Min, Max):</b> (0.2, 5.6)</p> <p><b>Copollutant; correlation:</b> NO<sub>2</sub> O<sub>3</sub> SO<sub>2</sub> PM<sub>10</sub> BS</p>	<p><b>Increment:</b> 0.8 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Respiratory complaints</b></p> <p><b>Age Group</b></p> <p><b>All ages:</b> 0.76% (-0.83, 2.38); 1 <b>0-14:</b> 2.92% (0.60, 5.30); 1 <b>15-64:</b> 2.15% (-0.27, 4.63); 1 <b>≥ 65:</b> 4.29% (1.15, 7.54); 0</p> <p><b>Asthma visits:</b></p> <p><b>Single-pollutant model</b></p> <p><b>Age Group:</b></p> <p><b>All ages:</b> 3.32% (0.56, 6.16); 1 <b>0-14:</b> 4.13% (-0.11, 8.54); 0 <b>15-64:</b> 4.41% (0.46, 8.52); 1</p> <p><b>Multi-pollutant model</b></p> <p><b>Age Group:</b></p> <p><b>0-14</b> CO, NO<sub>2</sub>: 2.05% (-2.25, 6.54); 0 CO, O<sub>3</sub>: 4.48% (0, 9.16); 0 CO, SO<sub>2</sub>: 2.34% (-1.94, 6.81); 0 CO, PM<sub>10</sub>: 2.93% (-1.53, 7.58); 0 CO, BS: 4.19% (-0.04, 8.60); 0</p>
<p><b>Author:</b> Bedeschi et al. (2007)</p> <p><b>Period of Study:</b> 1/2001-3/2002</p> <p><b>Location:</b> Reggio Emilia, Italy</p>	<p><b>Design:</b> ED Visits</p> <p><b>Health Outcome (ICD9):</b> Asthma (493); Asthma-like disorders, i.e., asthma, bronchiolitis, dyspnea/shortness of breath; Other respiratory disorders (i.e., upper and lower respiratory illness including sinusitis, bronchitis, and pneumonia)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, penalized splines</p> <p><b>Age Groups Analyzed:</b> &lt;15</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.4 (0.7) mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> (0.4, 4.6)</p> <p><b>Copollutant; correlation:</b> PM<sub>10</sub>: r = 0.61 TSP: r = 0.61 SO<sub>2</sub>: r = 0.71 NO<sub>2</sub>: r = 0.77</p>	<p><b>The study did not provide quantitative results for CO.</b></p>



Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Bell et al. (2008)</p> <p><b>Period of Study:</b> 1/1995-12/2002</p> <p><b>Location:</b> Taipei, Taiwan</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Pneumonia (486); Asthma (493)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SE) unit:</b> 0.9 ppm</p> <p><b>Range (Min, Max):</b> (0.3, 3.6)</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 0.5 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Asthma (avg correlation between monitor pairs = 0.75 (13 monitors))</b></p> <p>3.29% (-0.74 to 7.49); 0</p> <p>-.49% (-4.25 to 3.41); 1</p> <p>-.84% (-4.43 to 2.88); 2</p> <p>0.48% (-4.02 to 3.18); 3</p> <p>0.74% (-4.62 to 6.4); 0-3</p> <p><b>Pneumonia (avg correlation between monitor pairs = 0.75 (13 monitors))</b></p> <p>1.91% (-1.97 to 5.95); 0</p> <p>0.03% (-3.65 to 3.85); 1</p> <p>0.36% (-3.2 to 4.04); 2</p> <p>-1.29% (-4.77 to 2.32); 3</p> <p>0.21% (-5.03 to 5.73); 0-3</p> <p><b>Asthma (avg correlation between monitor pairs = 0.88 (5 monitors))</b></p> <p>1.68% (-1.68 to 5.15); 0</p> <p>-1.19% (-4.29 to 2.01); 1</p> <p>-0.83% (-3.83 to 2.26); 2</p> <p>-0.35% (-3.32 to 2.71); 3</p> <p>-0.31% (-4.9 to 4.5); 0-3</p> <p><b>Pneumonia (avg correlation between monitor pairs = 0.88 (5 monitors))</b></p> <p>1.24% (-2.02 to 4.6); 0</p> <p>-0.01% (-3.06 to 3.13); 1</p> <p>0.57% (-2.4 to 3.62); 2</p> <p>-0.85% (-3.78 to 2.16); 3</p> <p>0.31% (-4.23 to 5.06); 0-3</p> <p><b>Asthma (monitors with <math>\geq 0.75</math> between monitor correlations (11 monitors), avg correlation between monitor pairs = 0.81)</b></p> <p>2.87% (-0.91 to 6.79); 0</p> <p>-0.71% (-4.2 to 2.91); 1</p> <p>-0.73% (-4.08 to 2.73); 2</p> <p>-0.41% (-3.72 to 3.01); 3</p> <p>0.51% (-4.6 to 5.89); 0-3</p> <p><b>Pneumonia (monitors with <math>\geq 0.75</math> between monitor correlations (11 monitors) to avg correlation between monitor pairs = 0.81)</b></p> <p>0.98% (-1.68 to 5.76); 0</p> <p>-0.12% (-3.54 to 3.42); 1</p> <p>0.37% (-2.95 to 3.8); 2</p> <p>-1.08% (-4.34 to 2.3); 3</p> <p>0.3% (-4.71 to 5.57); 0-3</p>
<p><b>Author:</b> Bellini et al. (2007)</p> <p><b>Period of Study:</b> 1996-2002</p> <p><b>Location:</b> 15 Italian cities</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome:</b> Respiratory Conditions</p> <p><b>Study Design:</b> Time-series; Meta-analysis</p> <p><b>Statistical Analyses:</b></p> <ol style="list-style-type: none"> <li>GLM for city-specific estimates</li> <li>Bayesian random-effects for meta analysis</li> </ol> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> NR</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> correlation NR</p>	<p><b>Increment:</b> 1 mg/m<sup>3</sup></p> <p><b>% Increase (Lower CI, Upper CI); Lag:</b></p> <p><b>Respiratory conditions</b></p> <p><b>All ages:</b></p> <p><b>Season:</b></p> <p><b>Winter:</b> 0.58%; 0-1</p> <p><b>Summer:</b> 3.47%; 0-1</p> <p><b>All Season:</b> 1.25%; 0-3</p> <p><b>Note:</b> Estimates from Biggeri et al. (2004)</p>
<p><b>Author:</b> Braga et al. (2001)</p> <p><b>Period of Study:</b> 1/1993-11/1997</p> <p><b>Location:</b> Sao Paulo, Brazil</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory (460-519)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b></p> <p><math>\leq 2</math></p> <p>3-5</p> <p>6-13</p> <p>14-19</p> <p>0-19</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Maximum 8-h avg</p> <p><b>Mean (SD) unit:</b> 4.8 (2.3) ppm</p> <p><b>Range (Min, Max):</b> (0.6, 19.1)</p> <p><b>Copollutant:</b> correlation</p> <p><b>Copollutant: correlation</b></p> <p>PM<sub>10</sub>: r = 0.60</p> <p>O<sub>3</sub>: r = -0.07</p> <p>SO<sub>2</sub>: r = 0.47</p>	<p><b>Increment:</b> 3 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p>Respiratory</p> <p><b>Age Group:</b></p> <p><math>\leq 2</math>: 5.00% (3.30-6.80); 0-6</p> <p>3-5: 4.90% (1.40-8.50); 0-6</p> <p>6-13: 1.00% (-2.50 to 4.60); 0-6</p> <p>14-19: 11.30% (5.90-16.80); 0-6</p> <p>0-19: 4.90% (3.50-6.40); 0-6</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Burnett et al. (1999)</p> <p><b>Period of Study:</b> 1/1980-12/1994</p> <p><b>Location:</b> Toronto, ON, Canada</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493); COPD (490-492, 496); Respiratory infection (464, 466, 480-487, 494)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.18 ppm</p> <p><b>IQR (25th, 75th):</b> (0.9, 1.4)</p> <p><b>Copollutant: correlation</b>  <math>PM_{2.5}</math>: <math>r = 0.49</math>  <math>PM_{10-2.5}</math>: <math>r = 0.20</math>  <math>PM_{10}</math>: <math>r = 0.43</math>  <math>NO_2</math>: <math>r = 0.55</math>  <math>SO_2</math>: <math>r = 0.37</math>  <math>O_3</math>: <math>r = -0.23</math></p>	<p><b>Increment:</b> 1.18 ppm</p> <p><b>% Increase (t-value); lag:</b></p> <p><b>Asthma:</b> 5.35% (3.92); 0</p> <p><b>COPD:</b> 2.93% (1.48); 0</p> <p><b>Respiratory Infection:</b> 5.00% (4.25); 0</p> <p><b>Asthma:</b></p> <p><b>Multi-pollutant model</b>  <math>CO, SO_2, O_3</math>: 5.15%  <math>CO, PM_{2.5}, SO_2, O_3</math>: 4.63%  <math>CO, PM_{10-2.5}, SO_2, O_3</math>: 5.25%  <math>CO, PM_{10}, SO_2, O_3</math>: 4.80%  <math>CO, PM_{10-2.5}, O_3</math>: 4.00%</p> <p><b>COPD:</b></p> <p><b>Multi-pollutant model</b>  <math>CO, SO_2, O_3</math>: 3.02%  <math>CO, PM_{2.5}, SO_2, O_3</math>: 2.46%  <math>CO, PM_{10-2.5}, SO_2, O_3</math>: 3.00%  <math>CO, PM_{10}, SO_2, O_3</math>: 2.75%  <math>CO, PM_{10-2.5}, O_3</math>: 3.00%</p>
<p><b>Author:</b> Burnett et al. (2001)</p> <p><b>Period of Study:</b> 1/1980-12/1994</p> <p><b>Location:</b> Toronto, ON, Canada</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493); Acute bronchitis/bronchiolitis (466); Croup (464.4); Pneumonia (480-486)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> &lt;2</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 1-h avg</p> <p><b>Mean (SD) unit:</b> 1.9 ppm</p> <p><b>IQR (25th, 75th):</b> (1.3, 2.3)</p> <p><b>Copollutant: correlation</b>  <math>O_3</math>: <math>r = 0.24</math></p>	<p><b>Increment:</b> 1.9 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Respiratory problems</b>  <math>CO</math>: 19.20%; 0-1  <math>CO, O_3</math>: 14.30%; 0-1</p>
<p><b>Author:</b> Cakmak et al. (2006b)</p> <p><b>Period of Study:</b> 4/1993-3/2000</p> <p><b>Location:</b> 10 Canadian cities</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Acute bronchitis/bronchiolitis (466); Pneumonia (480-486); Chronic/ unspecified bronchitis (490, 491); Emphysema (492); Asthma (493); Bronchiectasis (494); Chronic airway obstruction (496)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b>  1. Poisson  2. Restricted Maximum Likelihood Method</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.8 ppm</p> <p><b>Range (Min, Max):</b> (0.0, 6.5)</p> <p><b>Copollutant: correlation</b>  <math>SO_2</math>  <math>NO_2</math>  <math>O_3</math></p>	<p><b>Increment:</b> 0.8 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Respiratory disease</b>  <math>CO</math>: 0.60% (0.20, 1); 2.8  <math>CO, SO_2, NO_2, O_3</math>: -0.20% (-0.70- 0.30); 2.8</p>
<p><b>Author:</b> Cheng et al. (2007)</p> <p><b>Period of Study:</b> 1996-2004</p> <p><b>Location:</b> Kaohsiung, Taiwan</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Pneumonia (480-486)</p> <p><b>Study Design:</b> Bi-directional case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.76 ppm</p> <p><b>Range (Min, Max):</b> (0.14, 1.72)</p> <p><b>Copollutant: correlation</b>  <math>PM_{10}</math>  <math>SO_2</math>  <math>NO_2</math>  <math>O_3</math></p>	<p><b>Increment:</b> 0.31 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p><b>OR for pneumonia and exposure to various pollutants for all ages in areas <math>\geq 25^\circ C</math> or <math>&lt;25^\circ C</math></b></p> <p><b>Pollutant and Temperature</b>  <math>CO, \geq 25^\circ C</math>: 1.18 (1.14-1.23); 0-2  <math>CO, &lt;25^\circ C</math>: 1.47 (1.41-1.53); 0-2  <math>CO, PM_{10}, \geq 25^\circ C</math>: 1.15 (1.11-1.2); 0-2  <math>CO, PM_{10}, &lt;25^\circ C</math>: 1.28 (1.21-1.35); 0-2  <math>CO, SO_2, \geq 25^\circ C</math>: 1.22 (1.17-1.27); 0-2  <math>CO, SO_2, &lt;25^\circ C</math>: 1.49 (1.42-1.56); 0-2  <math>CO, NO_2, \geq 25^\circ C</math>: 1.2 (1.15-1.27); 0-2  <math>CO, NO_2, &lt;25^\circ C</math>: 1.01 (0.95-1.08); 0-2  <math>CO, O_3, \geq 25^\circ C</math>: 1.16 (1.12-1.2); 0-2  <math>CO, O_3, &lt;25^\circ C</math>: 1.44 (1.38-1.5); 0-2</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Cho et al. (2000)</p> <p><b>Period of Study:</b> 1/1996-12/1996</p> <p><b>Location:</b> 3 South Korea cities:</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Bronchial asthma; COPD; Bronchitis</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> All Ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> Daejeon: 1.424 (0.611) ppm Ulsan: 0.950 (0.211) ppm Suwon: 1.270 (0.549) ppm</p> <p><b>Range (Min, Max):</b> Daejeon: (.364, 3.504) Ulsan: (.380, 1.675) Suwon: (.250, 3.616)</p> <p><b>Copollutant: correlation</b> <b>Daejeon</b> SO<sub>2</sub>: r = 0.280; NO<sub>2</sub>: r = 0.041; TSP: r = 0.193; O<sub>3</sub>: r = -0.101; O<sub>3</sub> Max: r = -0.069 <b>Ulsan</b> SO<sub>2</sub>: r = 0.108; NO<sub>2</sub>: r = 0.446; TSP: r = 0.286; O<sub>3</sub>: r = -0.195; O<sub>3</sub> Max: r = -0.107 <b>Suwon</b> SO<sub>2</sub>: r = 0.556; NO<sub>2</sub>: r = 0.291; TSP: r = 0.496; O<sub>3</sub>: r = -0.371; O<sub>3</sub> Max: r = -0.365</p>	<p><b>Increment:</b> 1,000 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Estimates obtained using dummy variables to apply environmental indicators to the model</b></p> <p><b>Daejeon</b> CO: 1.26 (1.08-1.47) TSP, SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>: 1.21 (1.02-1.44) <b>Ulsan</b> CO: 3.55 (1.65-7.63) TSP, SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>: 2.51 (1.06-5.93) <b>Suwon</b> CO: 1.24 (0.97-1.59) TSP, SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>: 1.19 (0.88-1.61) <b>Estimates obtained using actual measured integrated environmental pollution indicator values</b></p> <p><b>Daejeon</b> CO: 1.34 (1.14-1.58) <b>Ulsan</b> CO: 1.27 (0.94-1.71) <b>Suwon</b> CO: 3.55 (1.27-9.93)</p>
<p><b>Author:</b> Farhat et al. (2005)</p> <p><b>Period of Study:</b> 8/1996-8/1997</p> <p><b>Location:</b> Sao Paulo, Brazil</p>	<p><b>Hospital Visits &amp; ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Pneumonia/bronchopneumonia (480-486); Asthma (493); Bronchiolitis (466)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Maximum 8-h avg</p> <p><b>Mean (SD) unit:</b> 3.8 (1.6) ppm</p> <p><b>Range (Min, Max):</b> (1.1, 11.4)</p> <p><b>Copollutant: correlation</b> PM<sub>10</sub>: r = 0.72; SO<sub>2</sub>: r = 0.49; NO<sub>2</sub>: r = 0.59; O<sub>3</sub>: r = -0.8</p>	<p><b>Increment:</b> 1.8 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b> <b>Lower Respiratory Tract Disease ED Visits</b> CO, PM<sub>10</sub>: -0.10% (-5.60 to 5.30); 0-2 CO, NO<sub>2</sub>: -1.20% (-6.70 to 4.20); 0-2 CO, SO<sub>2</sub>: 3.70% (-1.00 to 8.40); 0-2 CO, O<sub>3</sub>: 4.80% (0.50-9.10); 0-2 CO, PM<sub>10</sub>, NO<sub>2</sub>, SO<sub>2</sub>, O<sub>3</sub>: -0.64% (-6.90 to 5.60); 0-2 <b>Pneumonia/ Bronchopneumonia Hospital Admissions</b> CO, PM<sub>10</sub>: 4.40% (-7.90 to 16.70); 0-2 CO, NO<sub>2</sub>: 4.40% (-88.70 to 17.50); 0-2 CO, SO<sub>2</sub>: 7.80% (-2.50 to 18.20); 0-2 CO, O<sub>3</sub>: 9.60% (-0.50 to 19.70); 0-2 CO, PM<sub>10</sub> to NO<sub>2</sub>, SO<sub>2</sub>, O<sub>3</sub>: 5.10% (-9.60 to 19.70); 0-2 <b>Asthma/ Bronchiolitis Hospital Admissions</b> CO, PM<sub>10</sub>: 6.10% (-14.90 to 27.10); 0-2 CO, NO<sub>2</sub>: 2.40% (-16.90 to 21.70); 0-2 CO, SO<sub>2</sub>: 10.60% (-6.60 to 27.80); 0-2 CO, O<sub>3</sub>: 12.40% (-3.60 to 28.40); 0-2 CO, PM<sub>10</sub> to NO<sub>2</sub>, SO<sub>2</sub>, O<sub>3</sub>: 8.80% (-15.60 to 33.30); 0-2</p>
<p><b>Author:</b> Fung et al. (2006)</p> <p><b>Period of Study:</b> 6/1995-3/1999</p> <p><b>Location:</b> Vancouver, Canada</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory Illness</p> <p><b>Study Design:</b> 1. Dewanji and Moolgavkar 2. Time-series 3. Bi-directional case-crossover</p> <p><b>Statistical Analyses:</b> 1. Dewanji and Moolgavkar 2. Poisson 3. Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.69 (0.25) ppm</p> <p><b>Range (Min, Max):</b> (0.28, 2.03)</p> <p><b>Copollutant: correlation</b> COH: r = 0.85; O<sub>3</sub>: r = -0.53; NO<sub>2</sub>: r = 0.74; SO<sub>2</sub>: r = 0.61; PM<sub>10</sub>: r = 0.46; PM<sub>2.5</sub>: r = 0.23; PM<sub>10-2.5</sub>: r = 0.51</p>	<p><b>Increment:</b> 0.24 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag</b></p> <p><b>Dewanji and Moolgavkar</b> 1.008 (0.997-1.02); 0 1.012 (0.999-1.025); 0-2 1.010 (0.995-1.025); 0-4 1.009 (0.991-1.026); 0-6</p> <p><b>Time-series</b> 1.012 (1.000-1.023); 0 1.017 (1.003-1.032); 0-2 1.017 (1.001-1.035); 0-4 1.016 (0.996-1.036); 0-6</p> <p><b>Bi-directional case-crossover</b> 1.010 (0.006-1.023); 0 1.012 (0.996-1.027); 0-2 1.012 (0.995-1.03); 0-4 1.010 (0.991-1.031); 0-6</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Fusco et al. (2001)</p> <p><b>Period of Study:</b> 1/1995-10/1997</p> <p><b>Location:</b> Rome, Italy</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory conditions (460-519, excluding 470-478); Acute respiratory infections plus pneumonia (460-466, 480-486); COPD (490-492, 494-496) Asthma (493)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> All ages 0-14</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 3.6 (1.2) mg/m<sup>3</sup></p> <p><b>IQR (25th, 75th):</b> (2.8, 4.3)</p> <p><b>Copollutant: correlation All Year</b></p> <p>SO<sub>2</sub>: r = 0.56 NO<sub>2</sub>: r = 0.31 O<sub>3</sub>: r = -0.57</p> <p><b>Cold Season</b></p> <p>SO<sub>2</sub>: r = 0.37 NO<sub>2</sub>: r = 0.41 O<sub>3</sub>: r = -0.44</p> <p><b>Warm Season</b></p> <p>SO<sub>2</sub>: r = 0.44 NO<sub>2</sub>: r = 0.59 O<sub>3</sub>: r = -0.38</p>	<p><b>Increment:</b> 1.5 mg/m<sup>3</sup></p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Age Group: All Ages</b></p> <p><b>Respiratory conditions</b></p> <p>2.80% (1.30-4.30); 0 1.80% (0.20-3.30); 1 0.20% (-1.30 to 1.80); 2 0.50% (-2.00 to 1.10); 3 0.70% (-0.80 to 2.20); 4 CO, NO<sub>2</sub>: 2.30% (0.60-4.00); 0</p> <p><b>Acute Respiratory Infections plus Pneumonia</b></p> <p>2.20% (0.00-4.40); 0 2.10% (-0.10 to 4.40); 0 1.70% (-0.50 to 4.00); 2 -0.90% (-3.00 to 1.30); 3 1.50% (-0.70 to 3.70); 4 CO, NO<sub>2</sub>: 0.00% (-2.30 to 2.40); 0</p> <p><b>Asthma</b></p> <p>5.50% (0.90-10.40); 0 0.80% (-3.80 to 5.70); 1 -1.30% (-5.90 to 3.50); 2 -3.00% (-7.40 to 1.60); 3 0.60% (-4.00 to 5.30); 4 CO, NO<sub>2</sub>: 4.80% (0.30-9.50); 0</p> <p><b>COPD</b></p> <p>4.30% (1.60-7.10); 0 -0.20% (-2.90 to 2.50); 1 -0.20% (-2.90 to 2.60); 2 -0.30% (-3.00 to 2.40); 3 -0.10% (-2.80 to 2.60); 4 CO, NO<sub>2</sub>: 4.80% (0.90-7.90); 0</p> <p><b>Warm Season</b></p> <p><b>Respiratory Conditions:</b></p> <p>10.80% (6.70-14.80); 0</p> <p><b>Acute respiratory infections plus pneumonia:</b></p> <p>8.60% (2.90-14.60); 0</p> <p><b>COPD:</b></p> <p>13.90% (6.80-21.50); 0</p> <p><b>Age Group: 0-14</b></p> <p><b>Respiratory conditions</b></p> <p>2.50 (-0.30 to 5.50); 0 0.80 (-2.10 to 3.80); 1 0.20 (-2.70 to 3.10); 2 -1.00 (-3.70 to 1.90); 3 3.20 (0.40- 6.20); 4 CO, NO<sub>2</sub>: 4.10 (-1.20 to 9.80); 1</p> <p><b>Acute Respiratory Infections plus Pneumonia</b></p> <p>2.50 (-0.80 to 5.80); 0 -0.10 (-3.40 to 3.20); 1 0.90 (-2.30 to 4.30); 2 -2.00 (-5.10 to 1.20); 3 3.20 (0.00-6.60); 4 CO, NO<sub>2</sub>: 6.90 (0.80-13.40); 1</p> <p><b>Asthma</b></p> <p>6.30 (-0.50 to 13.50); 0 8.20 (1.10-15.70); 1 -0.70 (-7.30 to 6.30); 2 3.50 (-3.20 to 10.60); 3 4.80 (-1.90 to 12.00); 4 CO, NO<sub>2</sub>: 3.30 (-4.20 to 11.30); 1</p>
<p><b>Author:</b> Gouveia and Fletcher (2000a)</p> <p><b>Period of Study:</b> 11/1992-9/1994</p> <p><b>Location:</b> Sao Paulo, Brazil</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> All respiratory diseases Pneumonia (480-486); Asthma (493); Bronchitis (466, 490, 491)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson</p> <p><b>Age Groups Analyzed:</b> &lt;1 &lt;5</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Maximum 8-h avg</p> <p><b>Mean (SD) unit:</b> 5.8 (2.4) ppm</p> <p><b>Range (Min, Max):</b> (1.3, 22.8)</p> <p><b>Copollutant: correlation</b></p> <p>PM<sub>10</sub>: r = 0.63 SO<sub>2</sub>: r = 0.65 NO<sub>2</sub>: r = 0.35</p>	<p><b>Increment:</b> 6.9 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>All respiratory diseases</b></p> <p><b>Age Group:</b></p> <p>&lt;5: 1.017 (0.971-1.065); 0</p> <p><b>Pneumonia</b></p> <p><b>Age Group:</b></p> <p>&lt;5: 1.015 (0.961-1.071); 0 &lt;1: 1.035 (0.975-1.099); 2</p> <p><b>Asthma</b></p> <p><b>Age Group:</b></p> <p>&lt;5: 1.081 (0.98-1.192); 0</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Hajat et al. (1999)</p> <p><b>Period of Study:</b> 1/1992-12/1994</p> <p><b>Location:</b> London, U.K.</p>	<p><b>General Practitioner Visits</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493); Lower Respiratory Diseases (464, 466, 476, 480-483, 485-487, 490-492, 494-496, 500, 501, 503-505, 510-515, 518, 519, 786)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson</p> <p><b>Age Groups Analyzed:</b> All ages 0-14 15-64 ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> All year: 0.8 (0.4) ppm Warm Season (April-September): 0.7 (0.3) ppm</p> <p><b>Cool Season (October-March):</b> 1.0 (0.5) ppm</p> <p><b>Range (10th, 90th):</b> All Year: (0.5, 1.3) Warm Season: (0.4, 1.0) Cool Season: (0.5, 1.6)</p> <p><b>Copollutant: correlation</b> <b>All Year</b> NO<sub>2</sub>: r = 0.72; SO<sub>2</sub>: r = 0.51; BS: r = 0.85; O<sub>3</sub>: r = -0.40; PM<sub>10</sub>: r = 0.56 <b>Warm Season</b> NO<sub>2</sub>: r = 0.70; SO<sub>2</sub>: r = 0.32; BS: r = 0.65; O<sub>3</sub>: r = -0.12; PM<sub>10</sub>: r = 0.58 <b>Cool Season</b> NO<sub>2</sub>: r = 0.84; SO<sub>2</sub>: r = 0.58; BS: r = 0.87</p>	<p><b>Increment:</b> 0.8 &amp; 0.7 ppm</p> <p><b>% Increase (Lower CI, Upper CI); Lag</b></p> <p><b>All Year:</b> <b>Asthma – Single Day Lags</b> <b>Increment:</b> 0.8 ppm <b>Age Group</b> <b>0-14:</b> 4.10% (-0.10 to 8.40); 2 <b>15-64:</b> 0.90% (-2.10 to 4.10); 0 <b>≥ 65:</b> 7.50% (0.50-14.90); 2 <b>All ages:</b> 1.60% (-1.20 to 4.60); 2 <b>Asthma – Cumulative exposure</b> <b>Increment:</b> 0.7 ppm <b>Age Group</b> <b>0-14:</b> 6.90% (1.30-12.90); 0-3 <b>15-64:</b> 1.00% (-3.20 to 5.40); 0-2 <b>≥ 65:</b> 8.20% (0.40-16.60); 0-2 <b>All ages:</b> 1.80% (-1.50 to 5.20); 0-2 <b>Lower Respiratory Diseases – Single Day Lags</b> <b>Increment:</b> 0.8 ppm <b>Age Group</b> <b>0-14:</b> 4.40 (1.70-7.10); 2 <b>15-64:</b> 1.10 (-0.70 to 3.00); 2 <b>≥ 65:</b> -2.60 (-4.80 to -0.30); 3 <b>All ages:</b> 2.00 (0.50-3.40); 2 <b>Lower Respiratory Diseases – Cumulative exposure</b> <b>Increment:</b> 0.7 ppm for 0-2 and 0-3; 0.8 for 0-1 <b>Age Group</b> <b>0-14:</b> 3.00% (-1.00 to 7.20); 0-3 <b>15-64:</b> -0.70% (-2.90 to 1.50); 0-1 <b>≥ 65:</b> -1.60% (-5.10 to 2.00); 0-3 <b>All ages:</b> 1.80% (0.10-3.60); 0-2 <b>Warm or Cold Seasons:</b> <b>Asthma, Increment:</b> 0.8 ppm <b>Age Group &amp; Season</b> 0-14 &amp; Warm Season: 11.40% (3.30-20.00); 2 0-14 &amp; Cold Season: 2.90% (-3.20 to 9.40); 2 15-64 &amp; Warm Season: 4.80% (-0.60 to 10.60); 0 15-64 &amp; Cold Season: -0.30% (-4.80 to 4.50); 0 ≥ 65 &amp; Warm Season: 15.60% (3.10-29.60); 2 ≥ 65 &amp; Cold Season: 4.20% (-6.00 to 15.60); 2 <b>Lower Respiratory Diseases, Increment:</b> 0.8 ppm <b>Age Group &amp; Season</b> 0-14 &amp; Warm Season: 2.70% (-2.90 to 8.60); 2 0-14 &amp; Cold Season: 6.20% (2.30-10.20); 2 15-64 &amp; Warm Season: 6.20% (2.30-10.20); 2 15-64 &amp; Cold Season: 2.40% (-1.20 to 6.10); 2 ≥ 65 &amp; Warm Season: 1.00% (-1.60 to 3.80); 2 ≥ 65 &amp; Cold Season: -2.20% (-6.50 to 2.40); 3</p>
<p><b>Author:</b> Hajat et al. (2002)</p> <p><b>Period of Study:</b> 1/1992-12/1994</p> <p><b>Location:</b> London, U.K.</p>	<p><b>General Practitioner Visits</b></p> <p><b>Health Outcome (ICD9):</b> Upper Respiratory Diseases (URD)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson, GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> 0-14 15-64 ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> All year: 0.8 (0.4) ppm Warm Season (April-September): 0.7 (0.3) ppm Cool Season (October-March): 1.0 (0.5) ppm</p> <p><b>Range (10th, 90th):</b> All Year: (0.5, 1.3) Warm Season: (0.4, 1.0) Cool Season: (0.5, 1.6)</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 0.6 ppm, 0.8 ppm, &amp; 1.1 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Warm Season, Increment:</b> 0.6 ppm <b>Age Group</b> <b>0-14:</b> 2.90% (-0.60 to 6.40); 1 <b>14-64:</b> 7.90% (4.80-11.10); 1 <b>≥ 65:</b> 4.90% (-1.80 to 12.10); 3 <b>Cold Season, Increment:</b> 1.1 ppm <b>Age Group</b> <b>0-14:</b> -2.50% (-4.90 to 0.10); 1 <b>14-64:</b> 0.60% (-1.60 to 2.90); 1 <b>≥ 65:</b> 5.60% (0.90-10.60); 3 <b>All Year, Increment:</b> 0.8 ppm <b>Age Group</b> <b>0-14:</b> -2.20% (-4.00 to -0.30); 1 <b>14-64:</b> 2.70% (0.10-5.50); 1 <b>≥ 65:</b> 5.80% (2.40 to 9.30); 3</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Hapcioglu et al. (2006) <b>Period of Study:</b> 1/1997-12/2001 <b>Location:</b> Istanbul, Turkey	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> COPD (490-492, 494-496) <b>Study Design:</b> Cross-sectional <b>Statistical Analyses:</b> Pearson Correlation Coefficient <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> Monthly <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant:</b> NR	<b>Correlation Coefficient:</b> Between CO exposure and COPD: 0.57 Between CO exposure and COPD when controlling for temperature: 0.25
<b>Author:</b> Hinwood et al. (2006) <b>Period of Study:</b> 1/1992-12/1998 <b>Location:</b> Perth, Australia	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> COPD (490.00-496.99 excluding asthma) Pneumonia/influenza (480.00-489.99); Asthma (493) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h avg <b>Mean (SD) unit:</b> All Year: 2.3 (1.3) ppm; November-April: 2.2 (1.3) ppm; May-October: 2.4 (1.2) ppm <b>Range (10th, 90th):</b> All Year: (0.9, 4.2) November-April: (0.8, 4.2) May-October: (1.1, 4.2) <b>Copollutant: correlation</b> <b>All Year:</b> NO <sub>2</sub> : r = 0.57 O <sub>3</sub> : r = 0.00 <b>November-April:</b> NO <sub>2</sub> : r = 0.55 O <sub>3</sub> : r = 0.00 <b>May-October:</b> NO <sub>2</sub> : r = 0.57 O <sub>3</sub> : r = 0.16	<b>Increment:</b> 2.3 ppm <b>Odds Ratio (Lower CI, Upper CI); Lag</b> <b>Pneumonia</b> 0.99999 (0.9737-1.0268); 0 1.00650 (0.9806-1.0331); 1 1.00351 (0.9779-1.0298); 2 1.00424 (0.9790-1.0301); 3 1.00581 (0.9752-1.0374); 0-1 1.01005 (0.9755-1.0458); 0-2 1.00805 (0.9701-1.0474); 0-3 <b>COPD</b> 0.99915 (0.9693-1.0297); 0 1.00205 (0.9727-1.0323); 1 0.98630 (0.9577-1.0158); 2 0.98970 (0.9619-1.0182); 3 0.99960 (0.9647-1.0357); 0-1 0.99260 (0.9538-1.0329); 0-2 0.99160 (0.9493-1.0357); 0-3
<b>Author:</b> Hwang and Chan (2002) <b>Period of Study:</b> 1998 <b>Location:</b> 50 communities in Taiwan	<b>Clinic Visits</b> <b>Health Outcome (ICD9):</b> Lower respiratory tract infections (466, 480-486) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> 1. General linear regression 2. Bayesian hierarchical modeling <b>Age Groups Analyzed:</b> All Ages 0-14 15-64 ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h avg <b>Mean (SD) unit:</b> 1.00 (0.30) ppm <b>Range (Min, Max):</b> (0.51, 1.71) <b>Copollutant:</b> NR	<b>Increment:</b> 0.1 ppm <b>% Increase (Lower CI, Upper CI); Lag</b> <b>Age Group: All Ages</b> 0.80% (0.60-1.00); 0 0.10% (-0.10 to 0.30); 1 0.10% (-0.10 to 0.30); 2 <b>Age Group: 0-14</b> 0.70% (0.50-1.00); 0 0.10% (-0.20 to 0.30); 1 0.20% (-0.10 to 0.40); 2 <b>Age Group: 15-64</b> 0.90% (0.60-1.10); 0 0.20% (0.00-0.50); 1 0.20% (-0.10 to 0.40); 2 <b>Age Group: ≥ 65</b> 1.10% (0.80-1.50); 0 0.60% (0.30-1.00); 1 0.40% (0.10-0.80); 2
<b>Author:</b> Ito et al. (2007) <b>Period of Study:</b> 1999-2002 <b>Location:</b> New York City, NY	<b>ED Visits</b> <b>Health Outcome (ICD9):</b> Asthma (493) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GLM <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h avg <b>Mean (SD) unit:</b> All Season: 1.31 (0.43) ppm <b>Warm Months (April-September):</b> 1.22 (0.32) ppm <b>Cold Months (October-March):</b> 1.41 (0.5) ppm <b>Range (5th, 95th):</b> All season: (0.77, 2.11) <b>Warm Months (April-September):</b> (0.75, 1.82) <b>Cold Months (October-March):</b> (0.78, 2.33) <b>Copollutant:</b> NR	<b>Increment:</b> 1.3 ppm <b>Relative Risk (Lower CI, Upper CI); Lag</b> <b>Warm months:</b> 1.15 (1.07-1.25); 0-1

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Karr et al. (2007) <b>Period of Study:</b> 1995-2000 <b>Location:</b> South Coast Air Basin, CA	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Acute bronchiolitis (466.1) <b>Study Design:</b> Matched case-control <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> Infants: 3 weeks -1 yr	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> Chronic: 1,770 ppb Subchronic: 1,720 ppb <b>Range (Min, Max):</b> Chronic: (120, 8300) Subchronic: (130, 5070) <b>Copollutant:</b> NR	<b>Increment:</b> 910 ppb, 960 ppb <b>Odds Ratio (Lower CI, Upper CI); lag:</b> <b>Increment:</b> 910 ppb <b>Subchronic bronchitis:</b> 1 (0.97-1.03) <b>Increment:</b> 960 ppb <b>Chronic bronchitis:</b> 1 (0.97-1.03)
<b>Author:</b> Karr et al. (2006) <b>Period of Study:</b> 1995-2000 <b>Location:</b> South Coast Air Basin, CA	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Acute bronchiolitis (466.1) <b>Study Design:</b> Case-Crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> Infants: 3 weeks – 1 year	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1-day lag: Index*: 1,730 ppb Referent*: 1,750 ppb 4-day lag: Index*: 1,760 ppb Referent*: 1,790 ppb <b>Range (Min, Max):</b> Lag 1: Index*: (4, 9600) Referent*: (4, 9600) Lag 4: Index* (4, 8710) Referent* (4, 9600) <b>Copollutant:</b> NR * Index days: days lagged in reference to date of hospitalization of a case. Referent days: are for each case and includes all days that are the same day of week and in the same month as the index day for that case for CO.	<b>Increment:</b> 1361, 1400 ppb <b>Odds Ratio (Lower CI, Upper CI); Lag</b> <b>Increment:</b> 1361 ppb <b>Age Group:</b> <b>Overall:</b> 0.99 (0.96-1.02); 1 <b>25-29 weeks:</b> 0.86 (0.68-1.1); 1 <b>29 1/7 – 34 weeks:</b> 1 (0.86-1.15); 1 <b>34 1/7 – 37 weeks:</b> 0.95 (0.87-1.04); 1 <b>37 1/7 – 44 weeks:</b> 1 (0.97-1.03); 1 <b>Increment:</b> 1400 ppb <b>Age Group:</b> <b>Overall:</b> 0.97 (0.94-1); 4 <b>25-29 weeks:</b> 0.93 (0.72-1.2); 4 <b>29 1/7 – 34 weeks:</b> 0.89 (0.77-1.03); 4 <b>34 1/7 – 37 weeks:</b> 0.98 (0.90-1.08); 4 <b>37 1/7 – 44 weeks:</b> 0.97 (0.94-1); 4
<b>Author:</b> Kim et al. (2007) <b>Period of Study:</b> 2002 <b>Location:</b> Seoul, Korea	<b>Hospital Admissions</b> <b>Health Outcome (ICD10):</b> Asthma (J45 and J46) <b>Study Design:</b> Bi-directional case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All Ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h avg <b>Mean (SD) unit:</b> Daily Concentration: 8.6 (4.6) ppm Relevant Concentration: 2.8 (2.8) ppm <b>Range (Min, Max):</b> Daily Concentration: (0.8, 44.0) Relevant Concentration: (0.0, 30.4) <b>Copollutant:</b> NR	<b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Individual Level SEP</b> Quintile 1: 1.06 (1.02-1.09); 1-3 ma Quintile 2: 1.05 (1.02-1.09); 1-3 ma Quintile 3: 1.05 (1.01-1.08); 1-3 ma Quintile 4: 1.07 (1.03-1.11); 1-3 ma Quintile 5: 1.05 (1.00-1.09); 1-3 ma <b>Regional Level SEP</b> Quintile 1: 0.99 (0.92-1.07); 1-3 ma Quintile 2: 1.06 (1.02-1.11); 1-3 ma Quintile 3: 1.04 (1.02-1.07); 1-3 ma Quintile 4: 1.10 (1.06-1.15); 1-3 ma Quintile 5: 1.06 (1.03-1.09); 1-3 ma <b>Overall:</b> 1.06 (1.04-1.07); 1-3 ma <b>Relative Effect Modification for SES</b> <b>Individual Level SEP</b> Quintile 1: 1 Quintile 2: 1 (0.95-1.04); 1-3 ma Quintile 3: 0.99 (0.94-1.03); 1-3 ma Quintile 4: 1.02 (0.97-1.06); 1-3 ma Quintile 5: 0.99 (0.94-1.04); 1-3 ma <b>Regional Level SEP</b> Quintile 1: 1 Quintile 2: 1.05 (0.97-1.14); 1-3 ma Quintile 3: 1.03 (0.96-1.11); 1-3 ma Quintile 4: 1.08 (1-1.16); 1-3 ma Quintile 5: 1.05 (0.97-1.13); 1-3 ma

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Kontos et al. (1999)</p> <p><b>Period of Study:</b> 1/1987-12/1992</p> <p><b>Location:</b> Piraeus, Greece</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory conditions (laryngitis, bronchiolitis, tonsillitis, acute rhinopharyngitis, otitis, bronchopneumonia, pneumonia, asthma)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Stochastic dynamical system approach</p> <p><b>Age Groups Analyzed:</b> 0-14</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean Range (SD) unit:</b> 1987: 4.2 mg/m<sup>3</sup> 1992: 3.6 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant: correlation 1987-1989</b> Smoke: r = 0.2979; SO<sub>2</sub>: r = 0.2166; NO<sub>2</sub>: r = 0.1913</p> <p><b>1990-1992</b> Smoke: r = 0.5383; SO<sub>2</sub>: r = 0.43283; NO<sub>2</sub>: 0.5223</p>	<p>This study did not present quantitative results for CO.</p>
<p><b>Author:</b> Lee et al. (2002)</p> <p><b>Period of Study:</b> 12/1997-12/1999</p> <p><b>Location:</b> Seoul, Korea</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD10):</b> Asthma (J45, J46)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> &lt;15</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 1-h maximum</p> <p><b>Mean Range (SD) unit:</b> 1.8 (0.7) ppm</p> <p><b>IQR (25th, 75th):</b> (1.2, 2.2)</p> <p><b>Copollutant: correlation</b> PM<sub>10</sub>: r = 0.598 SO<sub>2</sub>: r = 0.812 NO<sub>2</sub>: r = 0.785 O<sub>3</sub>: r = -0.388</p>	<p><b>Increment:</b> 1.0 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b> RR for asthma and exposure to various pollutants for children under 15 years old</p> <p><b>Pollutant:</b> CO: 1.16 (1.10-1.22); 2-3 avg CO, PM<sub>10</sub>: 1.13 (1.07-1.20); 2-3 avg CO, SO<sub>2</sub>: 1.17 (1.08-1.27); 2-3 avg CO, NO<sub>2</sub>: 1.04 (0.95-1.14); 2-3 avg CO, O<sub>3</sub>: 1.16 (1.11-1.22); 2-3 avg CO, O<sub>3</sub>, PM<sub>10</sub>: 1.148 (1.084-1.217); 2-3 avg CO, O<sub>3</sub>, PM<sub>10</sub>, SO<sub>2</sub>: 1.168 (1.075-1.269); 2-3 avg CO, O<sub>3</sub>, PM<sub>10</sub>, SO<sub>2</sub>, NO<sub>2</sub>: 1.098 (0.994-1.214); 2-3 avg</p>
<p><b>Author:</b> Lee et al. (2006)</p> <p><b>Period of Study:</b> 1/2002-12/2002</p> <p><b>Location:</b> Seoul, Korea</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD10):</b> Asthma (J45-46)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> GAM with stringent parameters</p> <p><b>Age Groups Analyzed:</b> &lt;15</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Maximum 2-h avg</p> <p><b>Mean (SD) unit:</b> High SES: 6.08 (2.10) ppb Moderate SES: 6.35 (2.44) ppb Low SES: 6.67 (2.59) ppb</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant: correlation</b> NO<sub>2</sub>: r = 0.55 SO<sub>2</sub>: r = 0.72 PM<sub>10</sub>: r = 0.28 O<sub>3</sub>: r = -0.36</p>	<p><b>Increment:</b> 3.01 ppb, 0.26 ppb, 4.52 ppb, 3.68 ppb</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Increment:</b> 3.01 ppb <b>Overall:</b> 1.07 (0.96-1.20); 0 <b>Increment:</b> 0.26 ppb <b>High SES:</b> 1.06 (0.96-1.17); 0 <b>Increment:</b> 4.52 ppb <b>Moderate SES:</b> 0.96 (0.84-1.10); 0 <b>Increment:</b> 3.68 ppb <b>Low SES:</b> 1.02 (0.85-1.24); 0</p>
<p><b>Author:</b> Lee et al. (2007b)</p> <p><b>Period of Study:</b> 1996-2003</p> <p><b>Location:</b> Kaohsiung, Taiwan</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> COPD (490-492, 494, 496)</p> <p><b>Study Design:</b> Bi-directional case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.77 ppm</p> <p><b>Range (Min, Max):</b> (0.23, 1.72)</p> <p><b>Copollutant:</b> PM<sub>10</sub> SO<sub>2</sub> NO<sub>2</sub> O<sub>3</sub></p>	<p><b>Increment:</b> 0.29 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b> CO &lt;25°C : 1.398 (1.306-1.496); 0-2 ≥ 25°C : 1.189 (1.123-1.259); 0-2 CO, PM<sub>10</sub> &lt;25°C : 1.257 (1.152-1.371); 0-2 ≥ 25°C : 1.149 (1.079-1.224); 0-2 CO, SO<sub>2</sub> &lt;25°C : 1.396 (1.295-1.504); 0-2 ≥ 25°C : 1.241 (1.161-1.326); 0-2 CO, NO<sub>2</sub> &lt;25°C : 0.973 (0.877-1.080); 0-2 ≥ 25°C : 1.196 (1.104-1.297); 0-2 CO, O<sub>3</sub> &lt;25°C : 1.378 (1.286-1.477); 0-2 ≥ 25°C : 1.170 (1.105-1.239); 0-2</p>



Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Lin et al. (1999) <b>Period of Study:</b> 5/1991-4/1993 <b>Location:</b> Sao Paulo, Brazil	<b>ED Visits</b> <b>Health Outcome (ICD9):</b> Respiratory illness (lower respiratory illness, upper respiratory illness, wheezing) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson <b>Age Groups Analyzed:</b> <13	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 5 ppm <b>Range (Min, Max):</b> (1, 12) <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.50 NO <sub>2</sub> : r = 0.35 SO <sub>2</sub> : r = 0.56 O <sub>3</sub> : r = 0.04	<b>Increment:</b> NR <b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Overall Respiratory Illnesses</b> CO: 1.206 (1.066-1.364); 0-5 CO, PM <sub>10</sub> , O <sub>3</sub> , SO <sub>2</sub> , NO <sub>2</sub> : 0.945 (0.808-1.105); 0-5 <b>Lower Respiratory Illness</b> CO: 1.203 (0.867-1.669); 0-5 CO, PM <sub>10</sub> , O <sub>3</sub> , SO <sub>2</sub> , NO <sub>2</sub> : 0.971 (0.641-1.472); 0-5 <b>Upper Respiratory Illness</b> CO: 1.237 (1.072-1.428); 0-5 CO, PM <sub>10</sub> , O <sub>3</sub> , SO <sub>2</sub> , NO <sub>2</sub> : 0.944 (0.785-1.135); 0-5 <b>Wheezing</b> CO: 0.813 (0.606-1.091); 0-5 CO, PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub> : 0.74 (0.505-1.085); 0-5
<b>Author:</b> Lin et al. (2003) <b>Period of Study:</b> 1/1981-12/1993 <b>Location:</b> Toronto, ON, Canada	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Asthma (493) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> 6-12	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.18 (0.50) ppm <b>Range (Min, Max):</b> (0, 6.10) <b>Copollutant: correlation</b> SO <sub>2</sub> : r = 0.37 NO <sub>2</sub> : r = 0.55 O <sub>3</sub> : r = -0.16 PM <sub>2.5</sub> : r = 0.45 PM <sub>10-2.5</sub> : r = 0.17 PM <sub>10</sub> : r = 0.38	<b>Increment:</b> 0.5 ppm <b>Odds Ratio (Lower CI, Upper CI); lag:</b> <b>Boys:</b> <b>Adjusting for Daily Weather Variables</b> 1.05 (1.1-1.11); 1 / 1.07 (1.01-1.14); 2 1.08 (1.01-1.16); 3 / 1.08 (1.1-1.17); 4 1.07 (0.99-1.16); 5 / 1.07 (0.98-1.17); 6 1.07 (0.98-1.17); 7 <b>Adjusting for PM and Daily Weather Variables</b> 1.05 (0.99-1.11); 1 / 1.08 (1.01-1.16); 2 1.09 (1.01-1.18); 3 / 1.10 (1.02-1.20); 4 1.09 (1.00-1.18); 5 / 1.09 (0.99-1.19); 6 1.09 (0.99-1.20); 7 <b>Girls:</b> <b>Adjusting for Daily Weather Variables</b> 1.00 (0.93-1.06); 1 / 1.01 (0.94-1.10); 2 1.00 (0.91-1.09); 3 / 0.98 (0.89-1.09); 4 1.01 (0.91-1.13); 5 / 1.03 (0.92-1.16); 6 1.04 (0.93-1.17); 7 <b>Adjusting for PM and Daily Weather Variables</b> 1.00 (0.93-1.07); 1 / 1.01 (0.92-1.10); 2 0.99 (0.90-1.09); 3 / 0.97 (0.87-1.08); 4 0.99 (0.89-1.11); 5 / 1.02 (0.90-1.15); 6 1.05 (0.93-1.20); 7
<b>Author:</b> Lin et al. (2004c) <b>Period of Study:</b> 1/1987-12/1998 <b>Location:</b> Vancouver, BC Canada	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Asthma (493) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> GAM, LOESS <b>Age Groups Analyzed:</b> 6-12	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.96 (0.52) ppm <b>Range (Min, Max):</b> (0.23, 4.90) <b>Copollutant: correlation</b> SO <sub>2</sub> : r = 0.67 NO <sub>2</sub> : r = 0.73 O <sub>3</sub> : r = -0.35	<b>Increment:</b> 0.5 ppm <b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Boys</b> <b>High SES:</b> 1.06 (0.98-1.14); 1 / 1.06 (0.97-1.15); 2 1.07 (0.97-1.17); 3 / 1.03 (0.93-1.14); 4 1.01 (0.91-1.12); 5 / 1.01 (0.91-1.13); 6 1.06 (0.94-1.18); 7 <b>Low SES:</b> 1.06 (0.99-1.14); 1 / 1.03 (0.95-1.12); 2 1.01 (0.93-1.11); 3 / 0.99 (0.90-1.09); 4 0.96 (0.87-1.06); 5 / 0.98 (0.88-1.08); 6 0.98 (0.88-1.09); 7 <b>Girls</b> <b>High SES:</b> 1.05 (0.94-1.16); 1 / 1.02 (0.90-1.15); 2 0.97 (0.85-1.11); 3 / 0.95 (0.83-1.10); 4 0.93 (0.80-1.08); 5 / 0.95 (0.82-1.11); 6 1.01 (0.87-1.19); 7 <b>Low SES:</b> 1.01 (0.92-1.11); 1 / 0.98 (0.89-1.10); 2 0.99 (0.88-1.11); 3 / 1.05 (0.93-1.19); 4 1.07 (0.94-1.21); 5 / 1.07 (0.94-1.23); 6 1.04 (0.91-1.20); 7

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Lin et al. (2005) <b>Period of Study:</b> 1998-2001 <b>Location:</b> Toronto, Canada	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Respiratory Infections (464, 466, and 480-487) <b>Study Design:</b> Bi-directional case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> <15	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.16 (0.38) ppm <b>Range (Min, Max):</b> (0.38, 2.45) <b>Copollutant: correlation</b> PM <sub>2.5</sub> : r = 0.10 PM <sub>10-2.5</sub> : r = 0.06 PM <sub>10</sub> : r = 0.10 SO <sub>2</sub> : r = 0.12 NO <sub>2</sub> : r = 0.20 O <sub>3</sub> : r = -0.11	<b>Increment:</b> 0.44 ppm <b>Odds Ratio (Lower CI, Upper CI); Lag</b> <b>Boys</b> <b>No adjustment:</b> 1.11 (1.01-1.22); 0-3 / 1.10 (1.00-1.22); 0-5 <b>Adjustment for weather variables:</b> 1.13 (1.03-1.24); 0-3 / 1.13 (1.02-1.25); 0-5 <b>Adjustment for weather variables and PM:</b> 1.08 (0.98-1.20); 0-3 / 1.08 (0.97-1.20); 0-5 <b>Girls</b> <b>No adjustment:</b> 0.99 (0.89-1.10); 0-3 / 1.00 (0.89-1.13); 0-5 <b>Adjustment for weather variables:</b> 1.02 (0.92-1.14); 0-3 / 1.05 (0.93-1.18); 0-5 <b>Adjustment for weather variables and PM:</b> 1.01 (0.90-1.13); 0-3 / 1.02 (0.90-1.15); 0-5 <b>Total</b> <b>No adjustment:</b> 1.06 (0.98-1.14); 0-3 / 1.06 (0.98-1.15); 0-5 <b>Adjustment for weather variables:</b> 1.09 (1.01-1.17); 0-3 / 1.10 (1.01-1.19); 0-5 <b>Adjustment for weather variables and PM:</b> 1.05 (0.97-1.14); 0-3 / 1.06 (0.97-1.15); 0-5
<b>Author:</b> Linn et al. (2000) <b>Period of Study:</b> 1992-1995 <b>Location:</b> Los Angeles, CA	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> APR-DRG Codes: Pulmonary (75-101); COPD (88) ICD9 Codes: Asthma (493) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson <b>Age Groups Analyzed:</b> 0-29, ≥ 30	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> Winter 1.7 (0.8) ppm Spring 1.0 (0.3) ppm Summer 1.2 (0.4) ppm Fall 2.1 (0.8) ppm <b>Range (Min, Max):</b> Winter: (0.5, 5.3) Spring: (0.4, 2.2) Summer: (0.3, 2.7) Fall: (0.6, 4.3) <b>Copollutant: correlation</b> <b>Winter</b> NO <sub>2</sub> : r = 0.89; PM <sub>10</sub> : r = 0.78; O <sub>3</sub> : r = -0.43 <b>Spring</b> NO <sub>2</sub> : r = 0.92; PM <sub>10</sub> : r = 0.54; O <sub>3</sub> : r = 0.29 <b>Summer</b> NO <sub>2</sub> : r = 0.94; PM <sub>10</sub> : r = 0.72; O <sub>3</sub> : r = 0.03 <b>Fall</b> NO <sub>2</sub> : r = 0.84; PM <sub>10</sub> : r = 0.58; O <sub>3</sub> : r = -0.36	<b>Increment:</b> 1.0 ppm <b>β (SE); lag:</b> <b>Pulmonary</b> <b>Age Group: ≥ 30</b> All Year: 0.007 Winter: 0.016 Spring: 0.014 Summer: 0.020 Fall: 0.020 <b>Asthma</b> <b>Age Group 0-29</b> All Year: 0.036 <b>Asthma</b> <b>Age Group: ≥ 30;</b> All Year: 0.028 Winter: 0.045 Fall: 0.039 <b>COPD</b> <b>Age Group: ≥ 30</b> All Year: 0.019 Winter: 0.035 Fall: 0.029

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Luginah et al. (2005) <b>Period of Study:</b> 4/1995-12/2000 <b>Location:</b> Windsor, ON, Canada	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Respiratory illness (460-519) <b>Study Design:</b> Time-series and Case-crossover <b>Statistical Analyses:</b> 1. Time-series: Poisson 2. Case-crossover: conditional logistic regression <b>Age Groups Analyzed:</b> All ages 0-14 15-64 ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 1-h maximum <b>Mean (SD) unit:</b> 1.3 (1.0) ppm <b>Range (Min, Max):</b> (0, 11.82) <b>Copollutant: correlation</b> NO <sub>2</sub> : r = 0.38 SO <sub>2</sub> : r = 0.16 O <sub>3</sub> : r = 0.10 CoH: r = 0.31 PM <sub>10</sub> : r = 0.21	<b>Increment:</b> 1.17 ppm <b>Relative Risk (Lower CI, Upper CI); Lag</b> <b>Females and Case-crossover study design</b> <b>Age Group: All ages:</b> 1.037 (0.968-1.111); 1 1.063 (0.976-1.158); 2 1.087 (0.982-1.203); 3 <b>Age Group: 0-14:</b> 1.147 (1.006-1.307); 1 1.186 (1.020-1.379); 2 1.221 (1.022-1.459); 3 <b>Age Group: 15-64:</b> 1.005 (0.884-1.141); 1 1.007 (0.859-1.181); 2 1.032 (0.858-1.240); 3 <b>Age Group: ≥ 65:</b> 1.014 (0.922-1.116); 1 1.024 (0.907-1.156); 2 1.035 (0.893-1.200); 3 <b>Males and Case-crossover study design</b> <b>Age Group: All Ages:</b> 0.950 (0.884-1.020); 1 0.945 (0.862-1.036); 2 0.965 (0.866-1.075); 3 <b>Age Group: 0-14:</b> 1.003 (0.904-1.113); 1 0.997 (0.871-1.141); 2 0.970 (0.824-1.141); 3 <b>Age Group: 15-64:</b> 1.036 (0.870-1.233); 1 1.033 (0.821-1.299); 2 0.991 (0.760-1.293); 3 <b>Age Group: ≥ 65:</b> 0.867 (0.775-0.970); 1 0.865 (0.752-0.994); 2 0.946 (0.807-1.109); 3 <b>Female and Time-series study design</b> <b>Age Group: All Ages:</b> 1.049 (0.993-1.108); 1 1.032 (0.993-1.188); 2 1.051 (0.993-1.112); 3 <b>Age Group: 0-14:</b> 1.077 (0.979-1.184); 1 1.068 (1.001-1.139); 2 1.100 (0.997-1.213); 3 <b>Age Group: 15-64:</b> 1.072 (0.962-1.195); 1 1.025 (0.944-1.112); 2 1.081 (0.963-1.213); 3 <b>Age Group: ≥ 65:</b> 1.029 (0.957-1.118); 1 1.030 (0.928-1.144); 2 1.013 (0.899-1.142); 3 <b>Male and Time-series study design</b> <b>Age Group: All Ages:</b> 0.989 (0.932-1.049); 1 0.986 (0.946-1.029); 2 0.987 (0.929-1.048); 3 <b>Age Group: 0-14:</b> 1.034 (0.949-1.126); 1 0.996 (0.933-1.062); 2 0.968 (0.881-1.064); 3 <b>Age Group: 15-64:</b> 0.994 (0.854-1.157); 1 0.988 (0.884-1.104); 2 0.951 (0.806-1.121); 3 <b>Age Group: ≥ 65:</b> 0.901 (0.817-0.994); 1 0.904 (0.803-1.019); 2 0.963 (0.845-1.098); 3

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Martins et al. (2002)</p> <p><b>Period of Study:</b> 5/1996-9/1998</p> <p><b>Location:</b> Sao Paulo, Brazil</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD10):</b> Chronic Lower Respiratory Disease (CLRD: J40-47) for chronic bronchitis, emphysema, other COPD, asthma, and bronchiectasia</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> &gt;64</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Maximum 8-h avg</p> <p><b>Mean (SD) unit:</b> 3.7 (1.7) ppm</p> <p><b>Range (Min, Max):</b> (1.0, 12.6)</p> <p><b>Copollutant: correlation</b>  NO<sub>2</sub>: r = 0.62;  SO<sub>2</sub>: r = 0.51;  PM<sub>10</sub>: r = 0.73;  O<sub>3</sub>: r = 0.07</p>	<p><b>Increment:</b> 1.63 ppm</p> <p><b>β (SE); lag:</b></p> <p><b>Chronic Lower Respiratory Diseases</b></p> <p><b>Age Group</b></p> <p><b>&gt;64:</b> 0.0489 (0.0274); 2</p>
<p><b>Author:</b> Masjedi et al. (2003)</p> <p><b>Period of Study:</b> 9/1997-2/1998</p> <p><b>Location:</b> Tehran, Iran</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Total acute respiratory conditions; Asthma (493); COPD (490-492, 494, 496)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Multiple stepwise regression</p> <p><b>Age Groups Analyzed:</b> Adults</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 8.85 ppm</p> <p><b>Range (Min, Max):</b> (2.15, 23.8)</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>β (p-value); lag:</b></p> <p>Asthma: -0.779 (0.12)</p> <p>COPD: 0.012 (0.71)</p> <p>Acute Respiratory conditions: -0.086 (0.400)</p> <p>Correlation coefficients:</p> <p>Mean 3-day CO levels and asthma: -0.300 (0.149)</p> <p>Mean weekly CO level and asthma: -0.14 (0.2)</p> <p>Mean 10-day CO levels and asthma: -0.05 (0.43)</p>
<p><b>Author:</b> McGowan et al. (2002)</p> <p><b>Period of Study:</b> 6/1988- 12/1998</p> <p><b>Location:</b> Christchurch, New Zealand</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Pneumonia (480-487); Acute respiratory infections (460-466); Chronic lung Diseases (491-492, 494-496); Asthma (493)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Generalized Additive Model</p> <p><b>Age Groups Analyzed:</b> &lt;15; &gt;64</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.16 (1.51) mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> (0, 15.7)</p> <p><b>Copollutant:</b> NR</p>	<p>This study did not provide quantitative results for CO.</p>
<p><b>Author:</b> Migliaretti et al. (2007)</p> <p><b>Period of Study:</b> 1/1997 - 12/1999</p> <p><b>Location:</b> Turin, Italy</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory Illness (chronic bronchitis, emphysema, and other COPD) (490-496)</p> <p><b>Study Design:</b> Case-control</p> <p><b>Statistical Analyses:</b> Multiple logistic regression</p> <p><b>Age Groups Analyzed:</b> ≥ 15 15-64 &gt;64</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 8-h median</p> <p><b>Median (SD) unit:</b> 3.36 (1.57) mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant: correlation</b> TSP</p>	<p><b>Increment:</b> 1 mg/m<sup>3</sup></p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p><b>CO</b></p> <p><b>Age Group</b></p> <p>≥ 15: 1.053 (1.030-1.070)</p> <p>15-64: 1.040 (0.987-1.085)</p> <p>&gt;64: 1.054 (1.027-1.083)</p> <p><b>CO, TSP</b></p> <p><b>Age Group</b></p> <p>≥ 15: 1.058 (1.024-1.096)</p> <p>15-64: 1.062 (0.993-1.135)</p> <p>&gt;64: 1.054 (1.011-1.099)</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Moolgavkar (2000a)</p> <p><b>Period of Study:</b> 1987 – 1995</p> <p><b>Location:</b> 3 U.S. counties: Los Angeles County, CA, Cook County, IL, Maricopa County, AZ</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> COPD plus asthma (490-496)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> All Ages 0-19 20-64 ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h median</p> <p><b>Median unit:</b> Cook: 993 ppb LA: 1347 ppb Maricopa: 1240 ppb</p> <p><b>Range (Min, Max):</b> Cook: (224, 3912) LA: (237, 5955) Maricopa: (269, 4777)</p> <p><b>Copollutant: correlation</b> <b>Cook County:</b> NO<sub>2</sub>: r = 0.63; SO<sub>2</sub>: r = 0.35; O<sub>3</sub>: r = -0.28</p> <p><b>LA County:</b> NO<sub>2</sub>: r = 0.80; SO<sub>2</sub>: r = 0.78; O<sub>3</sub>: r = -0.52</p> <p><b>Maricopa County:</b> NO<sub>2</sub>: r = 0.66; SO<sub>2</sub>: r = 0.53; O<sub>3</sub>: r = -0.61</p>	<p><b>Increment:</b> 1.0 ppm</p> <p><b>% Increase (t-statistic); lag:</b></p> <p><b>Age Group: ≥ 65</b> <b>Cook County</b> CO: 2.60 (1.9); 0; / 3.00 (2.2); 1; / 1.30 (1.0); 2; 1.40 (1.1); 3; / 1.10 (0.8); 4; / 2.30 (1.8); 5</p> <p><b>Los Angeles County</b> CO: 5.40 (11.3); 0; / 4.90 (10.1); 1; / 5.00 (10.2); 2; 4.90 (10.1); 3; / 4.00 (8.3); 4; / 4.30 (8.6); 5; CO, PM<sub>10</sub>: 4.30 (3.3); 0; / 5.30 (4.2); 1; / 5.10 (4.0); 2; 6.80 (5.6); 3; / 6.90 (5.4); 4; / 6.30 (4.7); 5; CO, PM<sub>2.5</sub>: 3.00 (1.9); 0; / 3.90 (2.5); 1; / 4.20 (2.6); 2; 6.50 (4.4); 3; / 5.80 (3.8); 4; / 5.10 (3.1); 5</p> <p><b>Maricopa County</b> CO: 1.40 (1.0); 0; / 0.80 (0.6); 1; / 1.20 (0.9); 2; 1.20 (0.9); 3; / 1.50 (1.1); 4; / 4.90 (3.8); 5</p> <p><b>Age Group: 0-19</b> <b>Los Angeles County</b> CO: 8.20 (14.4); 0; / 9.00 (15.9); 1; / 9.20 (16.4); 2; 8.50 (15.0); 3; / 7.00 (12.1); 4; / 4.80 (8.1); 5; CO, PM<sub>10</sub>: 7.50 (14.4); 0; / 7.40 (5.2); 1; / 6.40 (4.3); 2; 8.00 (5.5); 3; / 6.30 (4.0); 4; / 5.30 (3.5); 5; CO, PM<sub>10-2.5</sub>: 5.70 (3.4); 0; / 7.50 (4.9); 1; / 5.60 (3.3); 2; 5.40 (3.5); 3; / 4.40 (2.7); 4; / 1.80 (1.1); 5</p> <p><b>Age Group: 20-64</b> <b>Los Angeles County</b> CO: 3.70 (8.6); 0; / 3.90 (9.1); 1; / 4.50 (10.6); 2; 3.50 (8.3); 3; / 3.40 (7.9); 4; / 3.50 (7.9); 5; CO, PM<sub>10</sub>: 5.00 (4.6); 0; / 3.00 (2.7); 1; / 3.10 (2.8); 2; 5.20 (4.7); 3; / 5.90 (5.1); 4; / 4.90 (4.4); 5; CO, PM<sub>2.5</sub>: 3.50 (2.5); 0; / 0.60 (0.4); 1; / 1.10 (0.8); 2; 5.70 (4.1); 3; / 4.70 (3.3); 4; / 3.90 (2.8); 5; CO, PM<sub>10-2.5</sub>: 2.80 (2.2); 0; / 2.50 (2.0); 1; / 0.60 (0.5); 2; 3.90 (3.2); 3; / 3.40 (2.8); 4; / 4.00 (3.4); 5</p>
<p><b>Author:</b> Moolgavkar (2003b)</p> <p><b>Period of Study:</b> 1987 – 1995</p> <p><b>Location:</b> 2 U.S. counties: Los Angeles County, CA, and Cook County, IL</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> COPD plus asthma (490-496)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, Poisson GLM with natural splines</p> <p><b>Age Groups Analyzed:</b> All Ages; ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h median</p> <p><b>Median unit:</b> Cook: 993 ppb LA: 1347 ppb Maricopa: 1240 ppb</p> <p><b>Range (Min, Max):</b> Cook: (224, 3912) LA: (237, 5955)</p> <p><b>Copollutant: correlation</b> <b>Cook County:</b> NO<sub>2</sub>: r = 0.63; SO<sub>2</sub>: r = 0.35; O<sub>3</sub>: r = -0.28</p> <p><b>Los Angeles County:</b> NO<sub>2</sub>: r = 0.80; SO<sub>2</sub>: r = 0.78; O<sub>3</sub>: r = -0.52</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>% Increase (t-statistic); lag:</b></p> <p><b>COPD - Los Angeles County</b> CO - <b>GAM-30 (10<sup>-8</sup>):</b> 5.48 (17.67); 0; / 5.67 (18.22); 1; / 5.90 (19.01); 2; 5.28 (16.94); 3; / 4.59 (14.50); 4; / 4.10 (12.80); 5</p> <p><b>GAM-100 (10<sup>-8</sup>):</b> 2.37 (8.67); 0; / 2.41 (8.73); 1; / 2.41 (8.76); 2; 1.81 (6.58); 3; / 1.38 (4.94); 4; / 1.07 (3.82); 5</p> <p><b>NS-100:</b> 2.28 (5.65); 0; / 2.29 (5.50); 1; / 2.32 (5.33); 2; 1.74 (4.10); 3; / 1.30 (3.16); 4; / 1.00 (2.46); 5</p> <p><b>COPD - Cook County</b> CO - <b>GAM-100 (10<sup>-8</sup>):</b> 2.11 (1.62); 0; / 2.85 (2.16); 1; / 1.14 (0.86); 2; 1.05 (0.79); 3; / 0.43 (0.33); 4; / 0.34 (0.26); 5</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Neidell et al. (2004)  <b>Period of Study:</b> 1992-1998  <b>Location:</b> California</p>	<p><b>Hospital Admissions</b>  <b>Health Outcome (ICD9):</b> Asthma (493)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> Linear Regression  <b>Age Groups Analyzed:</b> 0-1 1-3 3-6 6-12 12-18</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 24-h avg  <b>Mean (SD) unit:</b> 1.777 (1.037) ppm  <b>Range (Min, Max):</b> NR  <b>Copollutant: correlation</b>  O<sub>3</sub>  PM<sub>10</sub>  NO<sub>2</sub></p>	<p><b>Increment:</b> NR  <b>β (SE); lag;</b>  <b>Single-pollutant model</b>  <b>Age Group</b>  <b>0-1:</b> -0.007 (0.009);  <b>1-3:</b> 0.027 (0.009);  <b>3-6:</b> 0.053 (0.010);  <b>6-12:</b> 0.047 (0.009);  <b>12-18:</b> 0.025 (0.008)  <b>Fixed effect controlling for O<sub>3</sub>, PM<sub>10</sub>, and NO<sub>2</sub></b>  <b>Age Group</b>  <b>0-1:</b> -0.01 (0.01);  <b>1-3:</b> 0.024 (0.011);  <b>3-6:</b> 0.049 (0.011);  <b>6-12:</b> 0.023 (0.011);  <b>12-18:</b> 0.021 (0.009)  <b>Fixed effect controlling for O<sub>3</sub>, PM<sub>10</sub>, NO<sub>2</sub> and Avoidance Behavior</b>  <b>Age Group</b>  <b>0-1:</b> -0.010 (0.010);  <b>1-3:</b> 0.027 (0.011);  <b>3-6:</b> 0.051 (0.011);  <b>6-12:</b> 0.025 (0.011);  <b>12-18:</b> 0.021 (0.009)</p>
<p><b>Author:</b> Norris et al. (1999)  <b>Period of Study:</b> 9/1995- 12/1996  <b>Location:</b> Seattle, WA</p>	<p><b>ED Visits</b>  <b>Health Outcome (ICD9):</b> Asthma (493)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> Semiparametric Poisson GAM  <b>Age Groups Analyzed:</b> &lt;18</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 24-h avg  <b>Mean (SD) unit:</b> 1.6 (0.5) ppm  <b>Range (Min, Max):</b> (0.6, 4.1)  <b>Copollutant: correlation</b>  PM<sub>10</sub>: r = 0.74  NO<sub>2</sub> (1-h max): r = 0.47  NO<sub>2</sub> (24-h avg.): r = 0.66  SO<sub>2</sub> (1-h max): r = 0.15  SO<sub>2</sub> (24-h avg.): r = 0.32</p>	<p><b>Increment:</b> 0.6 ppm  <b>Relative Risk (Lower CI, Upper CI); Lag</b>  High Utilization: 1.04 (0.93-1.16); 1  Low Utilization: 1.15 (1.05-1.28); 1  All: 1.10 (1.02-1.19); 1</p>
<p><b>Author:</b> Peel et al. (2005)  <b>Period of Study:</b> 1/1993- 8/2000  <b>Location:</b> Atlanta, GA</p>	<p><b>ED Visits</b>  <b>Health Outcome (ICD9):</b> Asthma (493, 786.09); COPD (491, 492, 496); URI (460-466, 477); Pneumonia (480-486)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> 1. Poisson GEE or asthma, URI, all respiratory 2. Poisson GLM for pneumonia and COPD  <b>Age Groups Analyzed:</b> Primary Analysis: All Ages Secondary Analysis: 2-18</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 1-h maximum  <b>Mean (SD) unit:</b> 1.8 (1.2) ppm  <b>Range (10th, 90th):</b> (0.5, 3.4)  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1.0 ppm  <b>Relative Risk (Lower CI, Upper CI); Lag</b>  <b>Health Condition</b>  All respiratory illnesses: 1.011 (1.004-1.019); 0-2  <b>URI:</b>  1.012 (1.003-1.021); 0-2 / 1.066 (1.045-1.087); 0-13  <b>Asthma:</b>  1.010 (0.999-1.022); 0-2  1.076 (1.047-1.105); 0-13  <b>Pneumonia:</b>  1.009 (0.996-1.021); 0-2  1.045 (1.011-1.080); 0-13  <b>COPD:</b>  1.026 (1.004-1.048); 0-2  1.032 (0.975-1.092); 0-13  <b>RR for asthma and exposure to CO for children age 2-18:</b>  1.019 (1.004-1.035); 0-2  <b>RR for all respiratory illnesses and CO exposure for all ages</b>  AQS (1/1/93- 8/31/00): 1.011 (1.004-1.019); 0-2  AQS (8/1/98- 8/31/00): 1.010 (1.000-1.021); 0-2  ARIES (8/1/98- 8/31/00): 1.018 (1.003-1.033); 0-2</p>
<p><b>Author:</b> Sheppard et al. (1999)  <b>Period of Study:</b> 1987-1994  <b>Location:</b> Seattle, WA</p>	<p><b>Hospital Admissions</b>  <b>Health Outcome (ICD9):</b> Asthma (493)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> Poisson  <b>Age Groups Analyzed:</b> &lt;65</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 24-h avg  <b>Mean (SD) unit:</b> 1831 ppb  <b>IQR (25th, 75th):</b> (1277, 2201)  <b>Copollutant: correlation</b>  PM<sub>10</sub>: r = 0.83; PM<sub>2.5</sub>: r = 0.78;  PM<sub>10-2.5</sub>: r = 0.56; O<sub>3</sub>: r = -0.18;  SO<sub>2</sub>: r = 0.24</p>	<p><b>Increment:</b> 924 ppb  <b>% Increase (Lower CI, Upper CI); Lag</b>  CO: 6% (3, 9); 3  CO, PM<sub>2.5</sub>: 5% (1, 8); 3</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Slaughter et al. (2005)</p> <p><b>Period of Study:</b> 1/1995 – 6/2001</p> <p><b>Location:</b> Spokane, WA</p>	<p><b>Hospital Admissions &amp; ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory causes (460-519) Asthma (493); COPD (491, 492, 494, 496) Acute respiratory tract infections not including colds and sinusitis (464-466, 490)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GLM, Natural Splines</p> <p><b>Age Groups Analyzed:</b> All ages, Adults</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (5th, 95th):</b> (1.25, 3.05)</p> <p><b>Copollutant: correlation</b>  PM<sub>11</sub>: r = 0.63  PM<sub>2.5</sub>: r = 0.62  PM<sub>10</sub>: r = 0.32  PM<sub>10-2.5</sub>: r = 0.32</p>	<p><b>Increment:</b> 1.0 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>ED Visits</b></p> <p><b>All Respiratory Illnesses</b>  <b>Age Group: All Ages:</b> 0.99 (0.96-1.02); 1 / 1.01 (0.98-1.04); 2  1.03 (1.00-1.06); 3</p> <p><b>Asthma</b>  <b>Age Group: All Ages:</b> 1.00 (0.95-1.06); 1 / 1.01 (0.96-1.07); 2  1.06 (1.00-1.11); 3</p> <p><b>COPD</b>  <b>Age Group: Adults:</b> 0.92 (0.85-1.00); 1 / 0.99 (0.91-1.08); 2  1.01 (0.93-1.10); 3</p> <p><b>Hospital Admissions:</b></p> <p><b>All Respiratory Illnesses</b>  <b>Age Group: All Ages:</b> 0.99 (0.95-1.02); 1 / 1.00 (0.96-1.04); 2  0.99 (0.96-1.03); 3</p> <p><b>Asthma</b>  <b>Age Group: All Ages:</b> 1.02 (0.92-1.13); 1 / 1.06 (0.96-1.17); 2  1.00 (0.91-1.11); 3</p> <p><b>COPD</b>  <b>Age Group: Adults:</b> 0.94 (0.86-1.03); 1 / 1.04 (0.95-1.13); 2  0.97 (0.88-1.06); 3</p>
<p><b>Author:</b> Steib et al. (2007)</p> <p><b>Period of Study:</b> 7/1992- 3/1996</p> <p><b>Location:</b> Saint John, Canada</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Asthma; COPD; Respiratory infections; All respiratory illnesses</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg 1-h maximum</p> <p><b>Mean (SD) unit:</b>  All year: 0.5 (0.3) ppm  May-September: 0.6 (0.3) ppm  All year: 1.6 (1.1) ppm,  May-September: 1.7 (0.9) ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant: correlation</b>  H<sub>2</sub>S: r = -0.10; NO<sub>2</sub>: r = 0.68;  O<sub>3</sub>: r = -0.05; SO<sub>2</sub>: r = 0.31;  TRS: r = 0.07; PM<sub>10</sub>: r = 0.28;  PM<sub>2.5</sub>: r = 0.27; H<sup>+</sup>: r = 0.23;  SO<sub>4</sub><sup>2-</sup>: r = 0.27; CoH: r = 0.55</p>	<p><b>Increment:</b> 0.5 &amp; 1.7 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p>All Respiratory Illnesses</p> <p><b>Increment:</b> 0.5 ppm</p> <p>All Year: -3.40; 7</p> <p><b>Increment:</b> 1.7 ppm</p> <p>May- September: -5.70</p>
<p><b>Author:</b> Sun et al. (2006)</p> <p><b>Period of Study:</b> 1/2004- 12/2004</p> <p><b>Location:</b> Taiwan</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493)</p> <p><b>Study Design:</b> Cross-sectional</p> <p><b>Statistical Analyses:</b> Pearson correlation analysis</p> <p><b>Age Groups Analyzed:</b> &lt;16; 16-55</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Monthly</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>Correlation Coefficient:</b></p> <p><b>Asthma</b>  <b>Age Group:</b>  &lt;16: 0.653  16-55: 0.425</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Tenias et al. (2002)</p> <p><b>Period of Study:</b> 1/1994- 12/1995</p> <p><b>Location:</b> Valencia, Spain</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> COPD (491, 492, 494, 496)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> 1. Poisson autoregressive 2. Sensitivity: GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> &gt;14</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg 1-h maximum</p> <p><b>Mean (SD) unit:</b> 24-h avg All year: 3.1 mg/m<sup>3</sup> Warm Months: 2.5 mg/m<sup>3</sup> Cold Months: 3.7 mg/m<sup>3</sup> 1-h avg All year: 6.7 mg/m<sup>3</sup> Warm Months: 5.4 mg/m<sup>3</sup> Cold Months: 8.0 mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> 24-h avg: (0.9, 7.1) 1-h maximum: (1.6, 17.2)</p> <p><b>Copollutant: correlation</b> SO<sub>2</sub>: r = 0.734; NO<sub>2</sub>: r = 0.180; O<sub>3</sub>: r = -0.517</p>	<p><b>Increment:</b> 1 mg/m<sup>3</sup></p> <p><b>Relative Risk (Lower CI, Upper CI); Lag</b></p> <p><b>24-h avg</b> All Year: 1.074 (0.998- 1156); 1 Cold Months: 1.070 (0.991-1.156); 1 Warm Months: 1.129 (0.960-1.329); 1</p> <p><b>1-h maximum</b> All Year: 1.039 (1.014-1.066); 1 Cold Months: 1.037 (1.010-1.064); 1 Warm Months: 1.058 (0.994-1.127); 1</p> <p><b>All Year: sinusoidal terms:</b> 1.039 (1.010-1.066); 1</p> <p><b>All Year: humidity and temperature variables:</b> 1.040 (1.014-1.067); 1</p> <p><b>All Year: GAM, LOESS:</b> 1.042 (1.019-1.066); 1</p>
<p><b>Author:</b> Thompson et al. (2001)</p> <p><b>Period of Study:</b> 1/1993- 12/1995</p> <p><b>Location:</b> Belfast, Northern Ireland</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson</p> <p><b>Age Groups Analyzed:</b> Children</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> Warm Season: 0.57 (0.41) ppm Cold Season: 0.74 (0.73) ppm</p> <p><b>IQR (25th, 75th):</b> Warm Season: (0.3, 0.7) Cold Season: (0.4, 0.8)</p> <p><b>Copollutant: correlation</b> SO<sub>2</sub> (log): r = 0.64; PM<sub>10</sub> (log): r = 0.57; O<sub>3</sub>: r = -0.52; NOx (log): r = 0.74; NO (log): r = 0.71; NO<sub>2</sub>: r = 0.69</p>	<p><b>Increment:</b> NR</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>Temperature included in the model:</b> 1.04 (1.00-1.09); 0 / 1.07 (1.02-1.12); 0-1 1.06 (1.00-1.12); 0-2 / 1.07 (1.00-1.14); 0-3</p> <p>Warm Season: 1.06 (0.98-1.16); NR Cold Season: 1.07 (1.01-1.14); NR</p> <p><b>Adjusted for benzene level:</b> 0.92 (0.83-.02); 0-1 avg.</p> <p><b>Note:</b> The increment the study uses to calculate effect estimates is a doubling in CO levels, but The study did not provide this value.</p>
<p><b>Author:</b> Tolbert et al. (2007)</p> <p><b>Period of Study:</b> 1/1993- 12/2004</p> <p><b>Location:</b> Atlanta, GA</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory Disease: Asthma (493, 786.07, 786.09); COPD (491, 492, 496); URI (460-465, 460.0, 477); Pneumonia (480-496); Bronchiolitis (466.1, 466.11, 466.19))</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GLM</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 1-h maximum</p> <p><b>Mean (SD) unit:</b> 1.6 ppm</p> <p><b>Range (Min, Max):</b> (0.1, 7.7)</p> <p><b>Copollutant: correlation</b> PM<sub>10</sub>: r = 0.51; O<sub>3</sub>: r = 0.27; NO<sub>2</sub>: r = 0.70; SO<sub>2</sub>: r = 0.28; Coarse PM: r = 0.38; PM<sub>2.5</sub>: r = 0.47; SO<sub>4</sub>: r = 0.14; EC: r = 0.66; OC: r = 0.59; TC: r = 0.63; OHC: r = 0.29</p>	<p><b>Increment:</b> 1.22 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>Respiratory Diseases:</b> 1.016 (1.009-1.022); 3</p> <p><b>Note:</b> The study only provides results of the multi-pollutant models in figures, not quantitatively.</p>
<p><b>Author:</b> Trapasso and Keith (1999)</p> <p><b>Period of Study:</b> 1/1994- 12/1994</p> <p><b>Location:</b> Bowling Green, KY</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Spearman Rank Correlation Coefficient</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> NR</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>Correlation Coefficient (lag)</b> CO Mean: r = 0.19;0 CO Mean: r = 0.27; 1 CO Mean: r = 0.21; 2</p> <p>CO Max: r = 0.26; 0 CO Max: r = 0.36; 1 CO Max: r = 0.24; 2</p>



Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Tsai et al. (2006a)</p> <p><b>Period of Study:</b> 1996 - 2003</p> <p><b>Location:</b> Kaohsiung, Taiwan</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493)</p> <p><b>Study Design:</b> Case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.77 ppm</p> <p><b>Range (Min, Max):</b> (0.23, 1.72)</p> <p><b>Copollutant:</b> PM<sub>10</sub> SO<sub>2</sub> NO<sub>2</sub> O<sub>3</sub></p>	<p><b>Increment:</b> 0.29 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); Lag</b></p> <p>OR for getting asthma and exposure to various pollutants for all ages at either &lt;25°C or ≥ 25°C</p> <p>CO &lt;25°C : 1.414 (1.300-1.537); 0-2 ≥ 25°C : 1.222 (1.138-1.312); 0-2</p> <p>CO, PM<sub>10</sub> &lt;25°C : 1.251 (1.125-1.393); 0-2 ≥ 25°C : 1.178 (1.088-1.274); 0-2</p> <p>CO, SO<sub>2</sub> &lt;25°C : 1.207 (1.076-1.354); 0-2 ≥ 25°C : 1.290 (1.188-1.400); 0-2</p> <p>CO, NO<sub>2</sub> &lt;25°C : 0.916 (0.807-1.039); 0-2 ≥ 25°C : 1.249 (1.127-1.384); 0-2</p> <p>CO, O<sub>3</sub> &lt;25°C : 1.396 (1.282-1.520); 0-2 ≥ 25°C : 1.195 (1.113-1.284); 0-2</p>
<p><b>Author:</b> Vigotti et al. (2007)</p> <p><b>Period of Study:</b> 1/2000- 12/2000</p> <p><b>Location:</b> Pisa, Italy</p>	<p><b>ED Visits</b></p> <p><b>Health Outcome (ICD9):</b> Respiratory Disease: Asthma (493); Dry cough (468); Acute bronchitis (466)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> &lt;10; &gt;65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.5 (0.7) ug/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> (0.3, 3.5)</p> <p><b>Copollutant: correlation</b> NO<sub>2</sub>: r = 0.62 PM<sub>10</sub>: r = 0.70</p>	<p><b>Increment:</b> 1mg/m<sup>3</sup></p> <p><b>% Increase (Lower CI, Upper CI); Lag</b></p> <p><b>Age Group</b> &lt;10: 18.60% (-6.90 to 51.10); 1 &gt;65: 26.50% (3.40-54.80); 4</p>
<p><b>Author:</b> Villeneuve et al. (2006b)</p> <p><b>Period of Study:</b> 1995-2000</p> <p><b>Location:</b> Toronto, ON, Canada</p>	<p><b>Physician Visits</b></p> <p><b>Health Outcome (ICD9):</b> Allergic rhinitis (177)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GLM</p> <p><b>Age Groups Analyzed:</b> &gt;65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.1 (0.4) ppm</p> <p><b>Range (Min, Max):</b> (0.0, 2.2)</p> <p><b>Copollutant:</b> PM<sub>2.5</sub> PM<sub>10</sub> PM<sub>10-2.5</sub> SO<sub>2</sub> NO<sub>2</sub> O<sub>3</sub></p>	<p><b>Increment:</b> 0.4 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); Lag</b></p> <p>The study did not present quantitative results for CO.</p>
<p><b>Author:</b> Xirasagar et al. (2006)</p> <p><b>Period of Study:</b> 1998- 2001</p> <p><b>Location:</b> Taiwan</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493)</p> <p><b>Study Design:</b> Cross-sectional</p> <p><b>Statistical Analyses:</b> Spearman Rank Correlations</p> <p><b>Age Groups Analyzed:</b> 0-14; &lt;2; 2-5; &gt;5</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Monthly</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>Correlation Coefficient (Lag)</b></p> <p><b>Age Group:</b> &lt;2: r = -0.208 2-5: r = -0.281 &gt;5: r = -0.134</p>
<p><b>Author:</b> Yang et al. (2007)</p> <p><b>Period of Study:</b> 1996 - 2003</p> <p><b>Location:</b> Taipei, Taiwan</p>	<p><b>Hospital Admissions</b></p> <p><b>Health Outcome (ICD9):</b> Asthma (493)</p> <p><b>Study Design:</b> Case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.33 ppm</p> <p><b>Range (Min, Max):</b> (0.32, 3.62)</p> <p><b>Copollutant:</b> PM<sub>10</sub> SO<sub>2</sub> NO<sub>2</sub> O<sub>3</sub></p>	<p><b>Increment:</b> 0.53 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); Lag</b></p> <p>CO &lt;25°C : 1.076 (1.019-1.136); 0-2 ≥ 25°C : 1.277 (1.179-1.383); 0-2</p> <p>CO, PM<sub>10</sub> &lt;25°C : 1.050 (0.983-1.122); 0-2 ≥ 25°C : 1.332 (1.216-1.459); 0-2</p> <p>CO, SO<sub>2</sub> &lt;25°C : 1.131 (1.059-1.207); 0-2 ≥ 25°C : 1.278 (1.174-1.392); 0-2</p> <p>CO, NO<sub>2</sub> &lt;25°C : 0.915 (0.839-0.997); 0-2 ≥ 25°C : 1.177 (1.049-1.320); 0-2</p> <p>CO, O<sub>3</sub> &lt;25°C : 1.169 (1.102-1.240); 0-2 ≥ 25°C : 1.275 (1.177-1.382); 0-2</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Yang et al. (2007) <b>Period of Study:</b> 1996-2003 <b>Location:</b> Taipei, Taiwan	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> COPD: (490-492, 494, 496) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.33 ppm <b>Range (Min, Max):</b> (0.32, 3.66) ppm <b>Copollutant:</b> PM <sub>10</sub> , SO <sub>2</sub> , NO <sub>2</sub> , O <sub>3</sub>	<b>Increment:</b> 0.53 ppm <b>Odds Ratio (Lower CI, Upper CI); Lag</b> CO <20°C: 0.975 (0.921-1.033); 0-2 ≥ 20°C: 1.227 (1.178-1.277); 0-2 CO, PM <sub>10</sub> <20°C: 0.925 (0.863-0.992); 0-2 ≥ 20°C: 1.177 (1.123-1.235); 0-2 CO, SO <sub>2</sub> <20°C: 0.895 (0.832-0.962); 0-2 ≥ 20°C: 1.274 (1.219-1.331); 0-2 CO, NO <sub>2</sub> <20°C: 1.000 (0.910-1.099); 0-2 ≥ 20°C: 1.061 (0.998-1.129); 0-2 CO, O <sub>3</sub> <20°C: 0.935 (0.875-0.999); 0-2 ≥ 20°C: 1.234 (1.185-1.285); 0-2
<b>Author:</b> Yang et al. (2005) <b>Period of Study:</b> 1/1994- 12/1998 <b>Location:</b> Vancouver, Canada	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> COPD (490-492, 494, 496) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson <b>Age Groups Analyzed:</b> ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> .71 (0.28) ppm <b>Range (Min, Max):</b> (0.30, 2.48) <b>Copollutant: correlation</b> O <sub>3</sub> : r = -0.56 NO <sub>2</sub> : r = 0.73 SO <sub>2</sub> : r = 0.67 PM <sub>10</sub> : r = 0.50	<b>Increment:</b> 0.3 ppm <b>Relative Risk (Lower CI, Upper CI); Lag</b> CO 1.03 (1.00-1.06); 0 / 1.04 (1.01-1.08); 0-1 1.05 (1.01-1.09); 0-2 / 1.05 (1.00-1.10); 0-3 1.06 (1.01-1.11); 0-4 / 1.07 (1.02-1.12); 0-5 1.08 (1.02-1.13); 0-6 <b>Multipollutant:</b> CO, O <sub>3</sub> : 1.11 (1.04-1.18); 0-6 CO, NO <sub>2</sub> : 1.04 (0.95-1.14); 0-6 CO, SO <sub>2</sub> : 1.11 (1.01-1.22); 0-6 CO, PM <sub>10</sub> : 1.02 (0.93-1.12); 0-6 CO, PM <sub>10</sub> , O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub> : 1.08 (0.96-1.22); 0-6 CO, O <sub>3</sub> , NO <sub>2</sub> , SO <sub>2</sub> : 1.10 (0.98-1.23); 0-6
<b>Author:</b> Yang et al. (2003) <b>Period of Study:</b> 1/1986- 12/1998 <b>Location:</b> Vancouver, BC, Canada	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Respiratory diseases (460-519) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> <3; ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.98 (0.54) ppm <b>IQR (25th, 75th):</b> (0.62, 1.16) <b>Copollutant: correlation</b> O <sub>3</sub> : r = -0.52 CoH NO <sub>2</sub> SO <sub>2</sub>	<b>Increment:</b> 0.54 ppm <b>Odds Ratio (Lower CI, Upper CI); Lag</b> <b>OR for respiratory diseases and exposure to various pollutants for people &lt;3 and ≥ 65</b> <b>Age Group: &lt;3</b> CO alone: 1.04 (1.01-1.07); 1 CO, O <sub>3</sub> : 1.04 (1.01-1.07); 1 CO, O <sub>3</sub> , CoH, NO <sub>2</sub> , SO <sub>2</sub> : 1.02 (0.96-1.08); 1 <b>Age Group: ≥ 65</b> CO alone: 1.02 (1.00-1.04); 1 CO, O <sub>3</sub> : 1.02 (1.00-1.04); 1 CO, O <sub>3</sub> , CoH, NO <sub>2</sub> , SO <sub>2</sub> : 0.96 (0.93-1.00); 1
<b>Author:</b> Yang et al. (2004b) <b>Period of Study:</b> 6/1/1995 – 3/31/1999 <b>Location:</b> Vancouver, Canada	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Respiratory diseases (460-519); Pneumonia (480-486); Asthma (493) <b>Study Design:</b> Case-control <b>Statistical Analyses:</b> Pearson's correlation coefficient <b>Age Groups Analyzed:</b> <3	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.70 (0.30) ppm <b>IQR (25th, 75th):</b> (0.50, 0.80) <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.46; PM <sub>2.5</sub> : r = 0.24; PM <sub>10-2.5</sub> : r = 0.33; O <sub>3</sub> : r = -0.53; NO <sub>2</sub> : r = 0.74; SO <sub>2</sub> : r = 0.61	<b>This study did not present quantitative results for CO.</b>
<b>Author:</b> Zanobetti and Schwartz (2006) <b>Period of Study:</b> 1995-1999 <b>Location:</b> Boston, MA	<b>Hospital Admissions</b> <b>Health Outcome (ICD9):</b> Pneumonia (480-487) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> NR <b>IQR (25th, 75th):</b> (0.39, 0.60) <b>Copollutant: correlation</b> PM <sub>2.5</sub> : r = 0.52; BC: r = 0.82; NO <sub>2</sub> : r = 0.67; O <sub>3</sub> : r = -0.30	<b>Increment:</b> 0.475 ppm <b>% Increase (Lower CI, Upper CI); lag:</b> 5.45 (1.10, 9.51); 0 5.12 (0.83, 9.16); 0-1

**Table C-6. Studies of long-term CO exposure and respiratory morbidity.**

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Goss et al. (2004)  <b>Period of Study:</b> 1999 – 2000  <b>Location:</b> U.S.</p>	<p><b>Health Outcome:</b> Lung function (FEV<sub>1</sub>, Cystic fibrosis pulmonary exacerbation)  <b>Study Design:</b> Cohort  <b>Statistical Analyses:</b> Logistic regression  <b>Population:</b> 11,484 cystic fibrosis patients  <b>Age Groups Analyzed:</b> &gt;6</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> Annual avg  <b>Mean (SD) unit:</b> 0.692 (0.295) ppm  <b>IQR (25th, 75th):</b> (0.48, 0.83)  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1.0 ppm  <b>Odds Ratio (Lower CI, Upper CI); lag:</b> Two or More Pulmonary Exacerbations During 2000  1.02 (0.85-1.22)</p>
<p><b>Author:</b> Guo et al. (1999)  <b>Period of Study:</b> 10/1995 – 5/1996  <b>Location:</b> Taiwan</p>	<p><b>Health Outcome:</b> Asthma  <b>Study Design:</b> Cohort  <b>Statistical Analyses:</b> Logistic regression  <b>Population:</b> 331,686 non-smoking children  <b>Age Groups Analyzed:</b> Middle-school children (mean age: 13.8)</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> Annual avg  <b>Mean (SD) unit:</b> 853 (277) ppb  <b>Range (Min, Max):</b> (381, 1610)  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 326 ppb  <b>% Increase (Lower CI, Upper CI); lag:</b>  <b>Boys</b>  Physician-diagnosed asthma:  1.17% (0.63-1.72)  Questionnaire-diagnosed asthma:  1.10% (0.45-1.75)  <b>Girls</b>  Physician-diagnosed asthma:  0.84% (0.45-1.22)  Questionnaire-diagnosed asthma:  1% (0.44-1.56)</p>
<p><b>Author:</b> Hirsch et al. (1999)  <b>Period of Study:</b> Population: 9/1995-6/1996  Air: 4/1994 – 4/1995  <b>Location:</b> Dresden, Germany</p>	<p><b>Health Outcome:</b> Asthma symptoms in the past 12 months (wheeze, morning cough); Doctor's diagnosis (asthma, bronchitis); Lung function (bronchial hyperresponsiveness (BHR), FEV<sub>1</sub> &lt;85% pred., FEF<sub>25-75%</sub> &lt;70% pred.)  <b>Study Design:</b> Cross-sectional  <b>Statistical Analyses:</b> Multiple logistic regression  <b>Population:</b> 5-7: 2,796; 9-11: 2,625  <b>Age Groups Analyzed:</b> 5-7 and 9-11</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> Annual avg  <b>Mean (SD) unit:</b> 0.69 mg/m<sup>3</sup>  <b>Range (Min, Max):</b> (0.32, 1.54)  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 0.2 µg/m<sup>3</sup>  <b>Prevalence Odds Ratio (Lower CI, Upper CI); lag:</b>  <b>Symptoms in the past 12 months: Wheeze</b>  Home Exposure  <b>Age Groups: 5-7; 9-11:</b> 1.05 (0.93-1.18)  Home/School Exposure  <b>Age Groups: 9-11:</b> 1.02 (0.85-1.22)  <b>Morning Cough</b>  Home Exposure  <b>Age Groups: 5-7; 9-11:</b> 1.12 (1.01-1.23)  <b>Age Group: 9-11:</b> 1.13 (0.98-1.3)  <b>Doctor's diagnosis: Asthma</b>  Home Exposure  <b>Age Groups: 5-7; 9-11:</b> 1.07 (0.94-1.21)  <b>Age Groups: 9-11:</b> 1.16 (0.97-1.38)  <b>Doctor's diagnosis: Bronchitis</b>  <b>Age Groups: 5-7; 9-11:</b> 1.19 (1.11-1.27)  <b>Age Group: 9-11:</b> 1.24 (1.12-1.38)  <b>Lung function: BHR</b>  <b>Age Groups: 5-7; 9-11:</b> 0.79 (0.63-0.99)  <b>Age Group: 9-11:</b> 0.77 (0.6-0.99)  <b>Lung function: FEV<sub>1</sub> &lt;85% pred.</b>  <b>Age Groups: 5-7; 9-11:</b> 1.09 (0.81-1.47)  <b>Age Group: 9-11:</b> 1.01 (0.73-1.41)  <b>Lung function: FEF<sub>25-75%</sub> &lt;70% pred.</b>  <b>Age Groups: 5-7; 9-11:</b> 1.15 (0.94-1.39)  <b>Age Group: 9-11:</b> 1.07 (0.86-1.34)  <b>Symptoms in the past 12 months: Wheeze</b>  <b>Age Groups: 5-7; 9-11</b>  Atopic children: 1 (0.81-1.24)  Nonatopic children: 1.05 (0.83-1.31)  <b>Morning cough</b>  <b>Age Groups: 5-7; 9-11</b>  Atopic children: 1.03 (0.82-1.29)  Nonatopic children: 1.22 (1.05-1.41)  <b>Doctor's diagnosis: Asthma</b>  Atopic children: 1.05 (0.83-1.32)  Nonatopic children: 1.29 (1.05-1.59)  <b>Doctor's diagnosis: Bronchitis</b>  <b>Age Groups: 5-7; 9-11</b>  Atopic children: 1 (0.86-1.16)  Nonatopic children: 1.21 (1.1-1.33)</p> <p><b>Notes:</b> Atopic Children were defined as those children with specific IgE to aeroallergens &gt;0.7 kU-L-1; Nonatopic</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
			Children were defined as those children with specific IgE to aeroallergens $\leq 0.7$ kU-L-1.
<b>Author:</b> Hwang et al. (2006) <b>Period of Study:</b> 2001 <b>Location:</b> Taiwan	<b>Health Outcome:</b> Allergic rhinitis <b>Study Design:</b> Cross-sectional <b>Statistical Analyses:</b> Two-stage hierarchical model (logistic and linear regression) <b>Population:</b> 32,143 Taiwanese school children <b>Age Groups Analyzed:</b> 6-15	<b>Pollutant:</b> CO <b>Averaging Time:</b> Annual avg <b>Mean (SD) unit:</b> 664 (153) ppb <b>Range (Min, Max):</b> (416, 964) <b>Copollutant: correlation</b> NOx: r = 0.88 O <sub>3</sub> : r = -0.37 PM <sub>10</sub> : r = 0.27 SO <sub>2</sub> : r = 0.40	<b>Increment:</b> 100 ppb <b>Adjusted Odds Ratio (Lower CI, Upper CI); lag:</b> <b>Physician-diagnosed allergic rhinitis</b> 1.05 (1.04-1.07) CO, SO <sub>2</sub> : 1.04 (1.02-1.06) CO, PM <sub>10</sub> : 1.05 (1.03-1.07) CO, O <sub>3</sub> : 1.07 (1.05-1.09) Male: 1.06 (1.03-1.08); Female: 1.05 (1.02-1.08) Parental atopy: Yes: 1.05 (1.02-1.08) Parental atopy: No: 1.06 (1.03-1.08) Parental Education: <6: 1 (0.91-1.09) Parental Education: 6-8: 1.07 (1.02-1.12) Parental Education: 9-11: 1.05 (1.02-1.08) Parental Education: $\geq 12$ : 1.06 (1.03-1.09) ETS: Yes: 1.06 (1.03-1.08); ETS: No: 1.05 (1.02-1.08) Visible Mold: Yes: 1.07 (1.03-1.11) Visible Mold: No: 1.05 (1.03-1.07)
<b>Author:</b> Hwang et al. (2005) <b>Period of Study:</b> 2001 <b>Location:</b> Taiwan	<b>Health Outcome:</b> Asthma <b>Study Design:</b> Cross-sectional <b>Statistical Analyses:</b> Two-stage hierarchical model (logistic and linear regression) <b>Population:</b> 32,672 Taiwanese school children <b>Age Groups Analyzed:</b> 6-15	<b>Pollutant:</b> CO <b>Averaging Time:</b> Annual avg <b>Mean (SD) unit:</b> 664 (153) ppb <b>Range (Min, Max):</b> (416, 964) <b>Copollutant: correlation</b> NOx: r = 0.88 O <sub>3</sub> : r = -0.37 PM <sub>10</sub> : r = 0.27 SO <sub>2</sub> : r = 0.40	<b>Increment:</b> 100 ppb <b>Adjusted Odds Ratio (Lower CI, Upper CI); lag:</b> <b>Physician-diagnosed asthma:</b> 1.045 (1.017-1.074) CO, SO <sub>2</sub> : 1.066 (1.034-1.099) CO, PM <sub>10</sub> : 1.079 (1.047-1.112) CO, O <sub>3</sub> : 1.063 (1.1-1.474) CO, SO <sub>2</sub> , O <sub>3</sub> : 1.111 (1.074-1.15) CO, PM <sub>10</sub> , O <sub>3</sub> : 1.119 (1.084-1.155) Male: 1.49 (1.37-1.63); Female: 1 Parental atopy: Yes: 1 Parental atopy: No: 2.72 (2.5-2.97) Parental Education: <6: 1 Parental Education: 6-8: 1.17 (0.9-1.52) Parental Education: 9-11: 1.61 (1.26-2.05) Parental Education: $\geq 12$ : 2.43 (1.9-3.09) ETS: Yes: 0.85 (0.78-0.92); ETS: No: 1 Visible Mold: Yes: 1.27 (1.16-1.4); Visible Mold: No: 1 Maternal smoking during pregnancy: Yes: 1.18 (0.89-1.56) Maternal smoking during pregnancy: No: 1 Cockroaches noted monthly: Yes: 1.15 (1.03-1.29) Cockroaches noted monthly: No: 1 Water damage: Yes: 0.96 (0.81-1.12) Water damage: No: 1
<b>Author:</b> Lee et al. (2003c) <b>Period of Study:</b> 10/1995-5/1996 <b>Location:</b> Taiwan	<b>Health Outcome:</b> Allergic rhinitis <b>Study Design:</b> Cohort <b>Statistical Analyses:</b> Multiple logistic regression <b>Population:</b> 331,686 non-smoking children <b>Age Groups Analyzed:</b> 12-14	<b>Pollutant:</b> CO <b>Averaging Time:</b> Annual avg <b>Mean (SD) unit:</b> 853 (277) ppb <b>Range (Min, Max):</b> (381, 1610) <b>Copollutant:</b> NR	The study did not present quantitative results for CO.
<b>Author:</b> Meng et al. (2007) <b>Period of Study:</b> 11/2000 – 9/2001 <b>Location:</b> Los Angeles County and San Diego County, California	<b>Health Outcome:</b> Asthma <b>Study Design:</b> Cohort <b>Statistical Analyses:</b> Logistic regression <b>Population:</b> 1,609 physician-diagnosed asthmatics <b>Age Groups Analyzed:</b> $\geq 18$	<b>Pollutant:</b> CO <b>Averaging Time:</b> Annual avg <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant: correlation</b> Traffic: r = -0.04; O <sub>3</sub> : r = -0.55; PM <sub>10</sub> : r = 0.42; PM <sub>2.5</sub> : r = 0.52; NO <sub>2</sub> : r = 0.55	The study did not present quantitative results for CO.

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Mortimer et al. (2008)</p> <p><b>Period of Study:</b> 1989 – 2000</p> <p><b>Location:</b> San Joaquin Valley, CA</p>	<p><b>Health Outcome:</b> Lung function (FVC, FEV<sub>1</sub>, PEF, FEF<sub>25-75</sub>, FEF<sub>1</sub>/FVC, FEF<sub>25-75</sub>/FVC, FEF<sub>25</sub>, FEF<sub>75</sub>)</p> <p><b>Study Design:</b> Cohort</p> <p><b>Statistical Analyses:</b> 1. DSA algorithm 2. GEE</p> <p><b>Population:</b> 232 asthmatic children</p> <p><b>Age Groups Analyzed:</b> 6-11</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 8-h maximum monthly mean</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant; correlation:</b></p> <p><b>Lifetime</b> NO<sub>2</sub> (24-h avg): r = 0.68 O<sub>3</sub> (8-h maximum): r = -0.40 PM<sub>10</sub> (24-h avg): r = 0.05</p> <p><b>Prenatal</b> CO (8-h maximum): r = 0.52 NO<sub>2</sub> (24-h avg): r = 0.37 O<sub>3</sub> (8-h maximum): r = -0.16 PM<sub>10</sub> (24-h avg): r = -0.05</p>	<p><b>Increment:</b> NR</p> <p><b>Effect Size per IQR Increase in Pollutant (SE):</b></p> <p><b>FEF<sub>25-75</sub>:</b> 24-h avg CO exposure during 1st trimester 0.90% (0.0113)</p> <p><b>FEV<sub>1</sub>/FVC</b> Daily maximum CO exposure during ages 0 to 3 -2.50% (0.0016)</p> <p><b>FEF<sub>25-75</sub>/FVC</b> 24-h avg CO exposure during ages 0 to 6 and diagnosed with asthma &lt;2 years old -4.80% (0.0446)</p> <p><b>FEF<sub>25</sub></b> 24-h avg CO exposure during ages 0 to 6 and diagnosed with asthma &lt;2 years old plus 24-h avg PM<sub>10</sub> exposure during 2<sup>nd</sup> trimester and mother smoked when pregnant -6.70% (0.015)</p> <p><b>Coefficient (SE):</b></p> <p><b>FVC</b> 24-h avg CO exposure during 2<sup>nd</sup> trimester -0.0878 (0.0415)</p> <p><b>FEF<sub>25-75</sub></b> Lifetime 24-h avg CO exposure -0.94454 (0.3975)</p> <p><b>FEF<sub>25-75</sub>/FVC</b> -0.1090 (0.0303)</p> <p><b>FEV<sub>1</sub>/FVC</b> Prenatal 8-h maximum CO exposure: 0.1711 (0.0653) Lifetime 1-h maximum CO exposure: -0.3242 (0.0919)</p> <p>24-h avg CO exposure during ages 0-3 and diagnosed with asthma &lt;2 years old: -0.1814 (0.0599)</p> <p><b>FEF<sub>25</sub></b> 24-h avg CO exposure during ages 0-6 and diagnosed with asthma &lt;2 years old: -1.0460 (0.1953)</p> <p><b>FEF<sub>75</sub></b> Lifetime 8-h maximum CO exposure: -0.4214 (0.1423)</p>
<p><b>Author:</b> Singh et al. (2003)</p> <p><b>Period of Study:</b> NR</p> <p><b>Location:</b> Jaipur, India</p>	<p><b>Health Outcome:</b> Lung function</p> <p><b>Study Design:</b> Panel study</p> <p><b>Statistical Analyses:</b> Parametric statistical methods</p> <p><b>Population:</b> Campus panel: 142 Commuter panel: 158</p> <p><b>Age Groups Analyzed:</b> ~20</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Annual avg</p> <p><b>Mean (SD) unit:</b> Roadside: 3,175 µg/m<sup>3</sup> Campus: 2,150 µg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>The study did not present quantitative results for CO.</b></p>
<p><b>Author:</b> Wang et al. (1999)</p> <p><b>Period of Study:</b> 10/1995-6/1996</p> <p><b>Location:</b> Kaohsiung and Pintong, Taiwan</p>	<p><b>Health Outcome:</b> Asthma</p> <p><b>Study Design:</b> Cross-sectional</p> <p><b>Statistical Analyses:</b> Multiple logistic regression</p> <p><b>Population:</b> 165,173 high school students</p> <p><b>Age Groups Analyzed:</b> 11-16</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Annual median</p> <p><b>Median (SD) unit:</b> 0.80 ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NR</p>	<p><b>Increment:</b> NR</p> <p><b>Adjusted Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p>CO Concentrations: &lt;0.80 ppm: 1.0 CO Concentrations ≥ 0.80 ppm: 1.23 (1.19-1.28)</p> <p>Multivariate analysis with variables for exercise, smoking, alcohol, incense use, ETS: 1.15 (1.1-1.2)</p>

**Table C-7. Studies of short-term CO exposure and mortality.**

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Anderson et al. (2001)  <b>Period of Study:</b> 10/1994 – 12/1996  <b>Location:</b> West Midlands, United Kingdom</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental) (&lt;800); Cardiovascular (390-459); Respiratory (460-519)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> Poisson GAM  <b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> Maximum 8-h moving avg  <b>Mean (SD) unit:</b> 0.8 (0.7) ppm  <b>Range (Min, Max):</b> (0.2, 10.0)  <b>Copollutant correlation:</b> PM<sub>10</sub>: r = 0.55; PM<sub>2.5</sub>: r = 0.54; PM<sub>10-2.5</sub>: r = 0.10; BS: r = 0.77; SO<sub>4</sub><sup>2-</sup>: r = 0.17; NO<sub>2</sub>: r = 0.73; O<sub>3</sub>: r = -0.29; SO<sub>2</sub>: r = 0.49</p>	<p><b>Increment:</b> 1.0 ppm  <b>% Increase (Lower CI, Upper CI); lag:</b>  All-cause  0.8% (-0.6 to 2.2); 0-1  Cardiovascular  2.5% (0.4-4.6); 0-1  Respiratory  1.2% (-2.1 to 4.6); 0-1</p>
<p><b>Author:</b> Bellini et al. (2007)  <b>Period of Study:</b> 1996 – 2002  <b>Location:</b> 15 Italian cities</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental) (&lt;800); Cardiovascular (390-459); Respiratory (460-519)  <b>Study Design:</b> Meta-analysis  <b>Statistical Analyses:</b> Poisson GLM  <b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 24-h avg  <b>Mean (SD) unit:</b> NR  <b>Range (Min, Max):</b> NR  <b>Copollutant:</b> SO<sub>2</sub>  NO<sub>2</sub>  O<sub>3</sub>  PM<sub>10</sub></p>	<p><b>Increment:</b> 1 mg/m<sup>3</sup>  <b>% Increase (Lower CI, Upper CI); lag:</b>  All-cause  1.19% (0.61-1.72); 0-1  Respiratory  0.66% (-1.46 to 2.88); 0-1  Cardiovascular  0.93% (-0.10 to 1.77); 0-1</p>
<p><b>Author:</b> Biggeri et al. (2005)  <b>Period of Study:</b> 1990-1999  <b>Location:</b> 8 Italian Cities (Turin, Milan, Verona, Bologna, Ravenna, Florence, Rome, and Palermo)</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental) (&lt;800); Cardiovascular (390-459); Respiratory (460-519); Cardio-respiratory  <b>Study Design:</b> Meta-analysis  <b>Statistical Analyses:</b> Poisson GLM, cubic splines  <b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> Maximum 8-h moving average  <b>Mean (SD) unit:</b>  Turin, 1991-1994: 5.8 mg/m<sup>3</sup>  Turin, 1995-1998: 4.0 mg/m<sup>3</sup>  Milan, 1990-1994: 5.9 mg/m<sup>3</sup>  Milan, 1995-1997: 4.0 mg/m<sup>3</sup>  Verona, 1995-1999: 2.5 mg/m<sup>3</sup>  Ravenna, 1991-1995: 1.8 mg/m<sup>3</sup>  Bologna, 1996-1998: 2.4 mg/m<sup>3</sup>  Florence, 1996-1998: 2.7 mg/m<sup>3</sup>  Rome, 1992-1994: 6.5 mg/m<sup>3</sup>  Rome, 1995-1997: 5.4 mg/m<sup>3</sup>  Palermo, 1997- 1999: 2.1 mg/m<sup>3</sup>  <b>Range (Min, Max):</b>  Turin, 1991-1994: (NR, 24.7)  Turin, 1995-1998: (NR, 19.8)  Milan, 1990-1994: (NR, 26.5)  Milan, 1995-1997: (NR, 12.3)  Verona, 1995-1999: (NR, 10.2)  Ravenna, 1991-1995: (NR, 7.0)  Bologna, 1996-1998: (NR, 11.1)  Florence, 1996-1998: (NR, 8.7)  Rome, 1992-1994: (NR, 22.3)  Rome, 1995-1997: (NR, 18.5)  Palermo, 1997- 1999: (NR, 8.0)  <b>Copollutant:</b> NR</p>	<p><b>Increment:</b> 1.0 mg/m<sup>3</sup>  <b>% Increase (Lower CI, Upper CI); lag:</b>  Non-accidental  Fixed: 0.93 (0.50-1.36); 0-1  Random: 0.93 (0.50-1.36); 0-1  Cardiovascular  Fixed: 1.29 (0.62-1.96); 0-1  Random: 1.29 (0.62-1.96); 0-1  Respiratory  Fixed: 2.44 (0.74-4.17); 0-1  Random: 2.47 (0.14-4.85); 0-1</p>
<p><b>Author:</b> Botter et al. (2002)  <b>Period of Study:</b> 1991-1993  <b>Location:</b> São Paulo, Brazil</p>	<p><b>Health Outcome (ICD9):</b> Mortality  <b>Study Design:</b> Longitudinal study  <b>Statistical Analyses:</b> State space model  <b>Age Groups Analyzed:</b> ≥ 65</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 24-h avg  <b>Mean (SD) unit:</b> NR  <b>Range (Min, Max):</b> NR  <b>Copollutant:</b> TSP; NO<sub>2</sub>; O<sub>3</sub>; SO<sub>2</sub></p>	<p><b>Increment:</b> NR  <b>β (SE):</b>  Model 1: 0.0053 (0.0036)  Model 2: 0.0046 (0.0028)  Model 3: 0.0040 (0.0028)  Model 4: 0.0032 (0.0028)</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Bremner et al. (1999)</p> <p><b>Period of Study:</b> 1/1992–12/1994</p> <p><b>Location:</b> London, U.K.</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental) (&lt;800); Cardiovascular (390-459); Respiratory (460-519)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson, cubic splines</p> <p><b>Age Groups Analyzed:</b> All ages 0-64 ≥ 65 65-74 ≥ 75</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.8 (0.4) ppm</p> <p><b>Range (Min, Max):</b> (0.2, 5.6)</p> <p><b>Copollutant:</b> NO<sub>2</sub>; O<sub>3</sub>; SO<sub>2</sub>; PM<sub>10</sub>; BS</p>	<p><b>Increment:</b> 0.8 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>All-cause Age Group:</b> All ages: 0.9% (-0.2 to 2.0); 1 0-64: 1.2% (-1.0 to 3.5); 1 ≥ 65: 0.8% (-0.4 to 1.9); 2 65-74: 0.8% (-1.2 to 2.8); 3 ≥ 75: 0.9% (-0.4 to 2.2); 2</p> <p><b>Respiratory Age Group:</b> All ages: 2.0% (-0.3 to 4.5); 3 0-64: 7.8% (0.2-15.9); 3 ≥ 65: 0.7% (-1.7 to 3.2); 3 65-74: 7.5% (2.1-13.2); 3 ≥ 75: 2.3% (-0.5 to 5.3); 0</p> <p><b>Multipollutant :</b> CO, SO<sub>2</sub>: 1.90% (0.18-3.64); 3 CO, PM<sub>10</sub>: 1.25% (0.04-2.47); 3 CO, BS: 2.41% (-0.65 to 5.57); 3</p> <p><b>Cardiovascular Age Group:</b> All ages: 1.4% (-0.1 to 3.0); 1 0-64: 2.1% (-1.7 to 6.0); 2 ≥ 65: 1.1% (-0.4 to 2.8); 2 65-74: 2.4% (-0.6 to 5.5); 2 ≥ 75: 1.9% (0.0-3.9); 2</p> <p><b>Multipollutant:</b> CO, NO<sub>2</sub>: 2.55% (0.40-4.75); 1 CO, O<sub>3</sub>: 3.98% (0.85-7.21); 1 CO, PM<sub>10</sub>: 0.62% (-0.59 to 1.85); 1 CO, BS: 1.29% (-1.53 to 4.19); 1</p>
<p><b>Author:</b> Burnett et al. (2000)</p> <p><b>Period of Study:</b> 1986-1996</p> <p><b>Location:</b> 8 Canadian cities</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental) (&lt;800)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> 1. Single-pollutant models: Poisson GAM, LOESS 2. Multi-pollutant models: Principal component regression analysis</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.9 ppm</p> <p><b>Range (Max):</b> 7.2 ppm</p> <p><b>Copollutant: correlation</b> O<sub>3</sub>: r = -0.05 SO<sub>2</sub>: r = 0.42 PM<sub>2.5</sub>: r = 0.44 PM<sub>10-2.5</sub>: r = 0.29 PM<sub>10</sub>: r = 0.45</p>	<p><b>Increment:</b> 0.9 ppm</p> <p><b>% Increase (t-value); lag:</b></p> <p>Temporally filtered daily non-accidental mortality (days in which PM<sub>10</sub> data available) CO: 0.4 (0.4); 0; 2.0 (2.3); 1 CO, PM<sub>2.5</sub>: -0.7 (-0.7); 0; 1.1 (1.1); 1 CO, PM<sub>10-2.5</sub>: 0.1 (0.2); 0; 1.8 (2.1); 1 CO, PM<sub>10</sub>: -0.5 (-0.6); 0; 1.2 (1.3); 1</p> <p>Daily filtered non-accidental mortality Single-pollutant model: 2.1 (2.1) Multi-pollutant models: <b>Model 1:</b> CO, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, O<sub>3</sub>, NO<sub>2</sub>, SO<sub>2</sub>: 0.7 (1.9) <b>Model 2:</b> CO, SO<sub>4</sub>, Ni, Fe, Zn, O<sub>3</sub>, NO<sub>2</sub>: 0.7 (1.7)</p>
<p><b>Author:</b> Burnett et al. (2004)</p> <p><b>Period of Study:</b> 1981 – 1999</p> <p><b>Location:</b> 12 Canadian cities</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental) (&lt;800)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> 1. Poisson, natural splines 2. Random effects regression model</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.02 ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> NO<sub>2</sub>; O<sub>3</sub>; SO<sub>2</sub>; PM<sub>2.5</sub>; PM<sub>10-2.5</sub></p>	<p><b>Increment:</b> 1.02 ppm</p> <p><b>% Increase (t-value); lag:</b> 0.68% (3.12); 1 CO, NO<sub>2</sub>: 0.07% (0.30); 1</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Cakmak et al. (2007)</p> <p><b>Period of Study:</b> 1/1997 – 12/2003</p> <p><b>Location:</b> Chile – 7 cities</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental) (&lt;800); Cardiovascular diseases (390-459); Respiratory diseases (460-519)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson; Random effects regression model</p> <p><b>Age Groups Analyzed:</b> All ages ≤ 64 65 – 74 75 – 84 ≥ 85</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.29 ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant correlation:</b> O<sub>3</sub>: r = -0.55 to -0.01 SO<sub>2</sub>: r = 0.31 to 0.67 PM<sub>10</sub>: r = 0.49 to 0.82</p> <p><b>Note:</b> Correlations are between pollutants for seven monitoring stations.</p>	<p><b>Increment:</b> 1.29 ppm</p> <p><b>% Increase (t-value); lag:</b></p> <p><b>Non-accidental:</b> 5.88% (6.42); 1; / 9.39% (6.89); 0-5 CO+PM<sub>10</sub>+O<sub>3</sub>+SO<sub>2</sub>: 6.13% (4.34); 1</p> <p><b>Age Group: ≤ 64</b> 4.10% (2.52); 1; / 4.76% (2.19); 0-5</p> <p><b>Age Group: 65-74</b> 6.24% (3.17); 1; / 8.12% (3.88); 0-5</p> <p><b>Age Group: 75-84</b> 8.64% (4.82); 1; / 13.12% (5.12); 0-5</p> <p><b>Age Group: ≥ 85</b> 8.58% (4.45); 1; / 13.20% (4.82); 0-5</p> <p><b>April-September</b> 7.09% (4.02); 1; / 9.65% (4.50); 0-5</p> <p><b>October-March</b> 5.45% (1.14); 1; / 7.80% (1.89); 0-5</p> <p><b>Cardiac</b> 7.79% (4.56); 1; / 11.22% (4.8); 0-5</p> <p><b>Respiratory</b> 12.93% (5.78); 1; / 21.31% (6.34); 0-5</p>



Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Chock et al. (2000)</p> <p><b>Period of Study:</b> 1989 – 1991</p> <p><b>Location:</b> Pittsburgh, PA</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Respiratory (480-486, 490-496, 507); Cardiovascular (390-448); Influenza (487)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM; Cubic B-spline basis functions</p> <p><b>Age Groups Analyzed:</b> All ages &lt;75 &gt;75</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 1-h avg</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> PM<sub>10</sub>; PM<sub>2.5</sub>; O<sub>3</sub>; SO<sub>2</sub>; NO<sub>2</sub></p>	<p><b>Increment:</b> NR</p> <p><b>β (SE); lag:</b></p> <p><b>Age Group: &lt;75</b> CO alone: 0.0080 (1.56); 0 PM<sub>10</sub>, CO: 0.0030 (0.48); 0 PM<sub>10</sub>, NO<sub>2</sub>, CO: 0.0079 (1.14); 0 PM<sub>10</sub>, O<sub>3</sub>, SO<sub>2</sub>, NO<sub>2</sub>, CO: 0.072 (1.02); 0 CO -0.00738 (-1.42); -3; / 0.00133 (0.23); -2; -0.00219 (-0.38); -1; / 0.00809 (1.48); 0; -0.00129 (-0.22); 1; / 0.00512 (0.90); 2; -0.00974 (-1.87); 3 CO, PM<sub>10</sub>, O<sub>3</sub>, SO<sub>2</sub>, NO<sub>2</sub> -0.01103 (-1.48); -3; / -0.00097 (-0.13); -2; 0.00514 (0.67); -1; / 0.00853 (1.15); 0; -0.00404 (-0.52); 1; / -0.00296 (-0.39); 2; -0.00346 (-0.46); 3</p> <p><b>Season</b> CO Winter: 0.00539 (0.78); 0 Spring: 0.01655 (1.90); 0 Summer: 0.00155 (0.14); 0 Fall: 0.00797 (1.14); 0 CO, PM<sub>10</sub> Winter: -0.00563 (-0.50); 0 Spring: 0.01233 (0.99); 0 Summer: -0.00712 (-0.48); 0 Fall: 0.00661 (0.73); 0 CO, PM<sub>10</sub>, O<sub>3</sub>, SO<sub>2</sub>, NO<sub>2</sub> Winter: -0.01326 (-0.95); 0 Spring: 0.02501 (1.54); 0 Summer: 0.01874 (0.92); 0 Fall: 0.01011 (0.88); 0</p> <p><b>Age Group:&gt;75</b> CO Alone: -0.0035 (-0.67); 0 CO, PM<sub>10</sub>: -0.0104 (-1.67); 0 CO, PM<sub>10</sub>, NO<sub>2</sub>: -0.0128 (-1.80); 0 CO, PM<sub>10</sub>, O<sub>3</sub>, SO<sub>2</sub>, NO<sub>2</sub>: -0.0144 (-1.99); 0 CO -0.00025 (-0.05); -3; / -0.00242 (-0.42); -2; -0.00238 (-0.41); -1; / -0.00302 (-0.54); 0; -0.00116 (-0.20); 1; / -0.00508 (-0.88); 2; -0.00251 (-0.48); 3 CO, PM<sub>10</sub>, O<sub>3</sub>, SO<sub>2</sub>, NO<sub>2</sub> -0.00123 (-0.17); -3; / -0.00876 (-1.13); -2; -0.00682 (-0.88); -1; / -0.01248 (-1.66); 0; -0.00672 (-0.86); 1; / -0.00181 (-0.23); 2; -0.00515 (-0.69); 3</p> <p><b>Season</b> CO Winter: -0.00304 (-0.43); 0 Spring: 0.00482 (0.54); 0 Summer: 0.01178 (1.07); 0 Fall: -0.01011 (-1.43); 0 CO, PM<sub>10</sub> Winter: -0.02303 (-2.03); 0 Spring: -0.00517 (-0.40); 0 Summer: 0.00735 (0.50); 0 Fall: -0.01042 (-1.14); 0 CO, PM<sub>10</sub>, O<sub>3</sub>, SO<sub>2</sub>, NO<sub>2</sub> Winter: -0.03370 (-2.41); 0 Spring: -0.00652 (-0.39); 0 Summer: 0.01258 (0.61); 0 Fall: -0.01250 (-1.07); 0</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Cifuentes et al. (2000)</p> <p><b>Period of Study:</b> 1988 – 1996</p> <p><b>Location:</b> Santiago, Chile</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All causes (non-accidental) (&lt;800)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, GAM with filtered variables &amp; GLM</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 1-h avg</p> <p><b>Mean (SD) unit:</b> 2.5 ppb</p> <p><b>Range (5th, 95th):</b> (0.6, 6.2)</p> <p><b>Copollutant correlation:</b>  PM<sub>2.5</sub>: r = 0.80  PM<sub>10-2.5</sub>: r = 0.47  SO<sub>2</sub>: r = 0.62  NO<sub>2</sub>: r = 0.65  O<sub>3</sub>: r = -0.01</p>	<p><b>Increment:</b>  All year: 2.5 ppm  Winter: 3.6 ppm  Summer: 1.3 ppm</p> <p><b>Relative Risk (t-ratio); Lag</b></p> <p><b>All Year</b>  CO: 1.041 (7.2); 0-1  CO, PM<sub>2.5</sub>: 1.025 (3.5); 0-1  CO, PM<sub>10-2.5</sub>: 1.035 (4.9); 0-1  CO, SO<sub>2</sub>: 1.038 (6.0); 0-1  CO, NO<sub>2</sub>: 1.026 (3.9); 0-1  CO, O<sub>3</sub>: 1.036 (4.8); 0-1</p> <p><b>Winter</b>  CO: 1.052 (5.9); 0-1  CO, PM<sub>2.5</sub>: 1.025 (2.1); 0-1  CO, PM<sub>10-2.5</sub>: 1.049 (4.3); 0-1  CO, SO<sub>2</sub>: 1.049 (5.0); 0-1  CO, NO<sub>2</sub>: 1.027 (2.6); 0-1  CO, O<sub>3</sub>: 1.051 (4.4); 0-1</p> <p><b>Summer</b>  CO: 1.053 (6.0); 0-1  CO, PM<sub>2.5</sub>: 1.053 (5.3); 0-1  CO, PM<sub>10-2.5</sub>: 1.053 (5.3); 0-1  CO, SO<sub>2</sub>: 1.050 (5.2); 0-1  CO, NO<sub>2</sub>: 1.047 (5.2); 0-1  CO, O<sub>3</sub>: 1.042 (3.6); 0-1</p> <p><b>All Year</b>  <b>GAM model</b>  CO: 1.041 (7.2); 0-1  CO, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>: 1.032 (4.6); 0-1</p> <p><b>GAM Filtered Variables</b>  CO: 1.030 (4.3); 0-1  CO, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>: 1.022 (2.4); 0-1</p> <p><b>GLM</b>  CO: 1.023 (2.4); 0-1  CO, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>: 1.013 (1.1); 0-1</p>
<p><b>Author:</b> Conceicao et al. (2001)</p> <p><b>Period of Study:</b> 1994 – 1997</p> <p><b>Location:</b> Sao Paulo, Brazil</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Respiratory diseases (460-519)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> &lt;5</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Maximum 8-h moving avg</p> <p><b>Mean (SD) unit:</b>  Total: 4.4 (2.2) ppm  1994: 5.1 (2.4) ppm  1995: 5.1 (2.4) ppm  1996: 3.9 (2.0) ppm  1997: 3.7 (1.6) ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> PM<sub>10</sub>; SO<sub>2</sub>; O<sub>3</sub></p>	<p><b>Increment:</b> NR</p> <p><b>β (SE); lag:</b>  CO: 0.0306 (0.0076); 2  CO, SO<sub>2</sub>, PM<sub>10</sub>, O<sub>3</sub>: 0.0259 (0.0116); 2</p> <p>Model 1: Pollutant concentration:  0.0827 (0.0077); 2</p> <p>Model 2: 1+loess(time):  0.0285 (0.0074); 2</p> <p>Model 3: 2+loess(temperature)+humidity:  0.0309 (0.0076); 2</p> <p>Model 4: 3+nonrespiratory counts:  0.0306 (0.0076); 2</p> <p>Model 5: 4+autoregressive parameters:  0.0292 (0.0118); 2</p>
<p><b>Author:</b> De Leon et al. (2003)</p> <p><b>Period of Study:</b> 1/1985 – 12/1994</p> <p><b>Location:</b> New York, NY</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Circulatory (390-459); Cancer (140-239)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> All ages  &lt;75  &gt;75</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 2.45 ppm</p> <p><b>IQR (25th, 75th):</b> (1.80, 2.97)</p> <p><b>Copollutant:</b> PM<sub>10</sub>; O<sub>3</sub>; SO<sub>2</sub>; NO<sub>2</sub></p>	<p><b>The study did not present quantitative results for CO.</b></p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Dominici et al. (2003b)</p> <p><b>Period of Study:</b> 1987 – 1994</p> <p><b>Location:</b> 90 U.S. cities (NMMAPS)</p>	<p><b>Health Outcome (ICD9):</b> Mortality: All-cause (non-accidental); Cardiovascular; Respiratory</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> 1. GAM with S-PLUS default convergence criteria 2. GAM with more stringent convergence criteria 3. Poisson GLM with natural cubic splines</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> O<sub>3</sub>; NO<sub>2</sub>; SO<sub>2</sub>; CO</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>% Increase (Lower CI, Upper CI); Lag</b></p> <p><b>CO</b> 0.08% (-0.18 to 0.34); 0 0.46% (0.18-0.73); 1 0.16% (-0.12 to 0.45); 2</p>
<p><b>Author:</b> Fairley et al. (1999)</p> <p><b>Period of Study:</b> 1989 – 1996</p> <p><b>Location:</b> Santa Clara, CA</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Respiratory; Cardiovascular</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg; Maximum 8-h avg</p> <p><b>Median (SD) unit:</b> 24-h avg: 1.4 (1.0) ppm Maximum 8-h avg: 2.1 (1.6) ppm</p> <p><b>Range (Min, Max):</b> 24-h avg: (0.0, 7.6) Maximum 8-h avg: (0.2, 2.5)</p> <p><b>Copollutant: correlation</b> PM<sub>10</sub>: r = 0.609; PM<sub>2.5</sub>: r = 0.435; PM<sub>10-2.5</sub>: r = 0.326; COH: r = 0.736; NO<sub>3</sub>: r = 0.270; SO<sub>4</sub>: r = 0.146; O<sub>3</sub>: r = -0.215</p>	<p><b>Increment:</b> 2.2 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>1980-1986</b> CO: 1.04; 0; CO: 1.05; 1; CO, COH: 1.00; 1; CO, NO<sub>3</sub>: 1.03; CO, NO<sub>3</sub>, O<sub>3</sub>, COH: 1.00</p> <p><b>1989-1996</b> CO: 1.02; 0; CO: 1.04; 1; CO, PM<sub>2.5</sub>: 0.98; CO, NO<sub>3</sub>: 1.01; CO, NO<sub>2</sub>, O<sub>3</sub>, NO<sub>3</sub>: 1.06</p> <p><b>Respiratory mortality:</b> CO: 1.08; 1</p> <p><b>Cardiovascular mortality:</b> CO: 1.04; 1</p>
<p><b>Author:</b> Fischer et al. (2003)</p> <p><b>Period of Study:</b> 1986 – 1994</p> <p><b>Location:</b> The Netherlands</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Non-accidental (&lt;800); Pneumonia (480-486); COPD (490-496); Cardiovascular (390-448)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> &lt;45 45-64 65-74 ≥ 75</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Median (SD) unit:</b> 406 µg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> (174, 2620)</p> <p><b>Copollutant:</b> PM<sub>10</sub>; BS; O<sub>3</sub>; NO<sub>2</sub>; SO<sub>2</sub></p>	<p><b>Increment:</b> 1,200 µg/m<sup>3</sup></p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>Cardiovascular</b></p> <p><b>Age Group:</b> &lt;45: 0.965 (0.750-1.240); 0-6 45-64: 1.029 (0.941-1.125); 0-6 65-74: 1.038 (0.972-1.108); 0-6 ≥ 75: 1.024 (0.984-1.065); 0-6</p> <p><b>COPD</b></p> <p><b>Age Group:</b> &lt;45: 1.710 (0.852-3.435); 0-6 45-64: 1.181 (0.850-1.640); 0-6 65-74: 1.377 (1.147-1.654); 0-6 ≥ 75: 1.072 (0.963-1.193); 0-6</p> <p><b>Pneumonia</b></p> <p><b>Age Group:</b> &lt;45: 0.927 (0.463-1.856); 0-6 45-64: 2.691 (1.509-4.800); 0-6 65-74: 1.118 (0.743-1.683); 0-6 ≥ 75: 1.230 (1.090-1.389); 0-6</p>
<p><b>Author:</b> Forastiere et al. (2005)</p> <p><b>Period of Study:</b> 1998 – 2000</p> <p><b>Location:</b> Rome, Italy</p>	<p><b>Health Outcome (ICD9):</b> Mortality: IHD (410-414)</p> <p><b>Study Design:</b> Time-stratified case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> &gt;35</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 2.4 (1.0) mg/m<sup>3</sup></p> <p><b>IQR (25th, 75th):</b> (1.7, 2.9)</p> <p><b>Copollutant correlation:</b> PNC: r = 0.89; PM<sub>10</sub>: r = 0.34; NO<sub>2</sub>: r = 0.54; SO<sub>2</sub>: r = 0.52; O<sub>3</sub>: r = 0.01</p>	<p><b>Increment:</b> 1.2 mg/m<sup>3</sup></p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p>6.5% (1.0-12.3); 0 4.7% (-0.9 to 10.7); 1 2.6% (-3.0 to 8.5); 2 -0.1% (-5.5 to 5.5); 3 7.0% (0.8-13.7); 0-1</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Forastiere et al. (2007) <b>Period of Study:</b> 1998 – 2001 <b>Location:</b> Rome, Italy	<b>Health Outcome (ICD9):</b> Malignant Neoplasms (140-208); Diabetes Mellitus (250); Hypertensive (401-405); Previous AMI (410, 412); IHD (410-414); Conduction disorders of the heart (426); Dysrhythmia (427); Heart Failure (428); Cerebrovascular (430-438); Peripheral Artery disease (440-448); COPD (490-496) <b>Study Design:</b> Time-stratified case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> >35	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> NR <b>IQR (25th, 75th):</b> NR <b>Copollutant:</b> PM <sub>10</sub> ; PM <sub>2.5</sub> ; NO <sub>x</sub> ; Benzene	This study did not present quantitative results for CO.
<b>Author:</b> Goldberg et al. (2001) <b>Period of Study:</b> 1984 – 1993 <b>Location:</b> Montreal, Quebec, Canada	<b>Health Outcome (ICD9):</b> Mortality: Upper respiratory diseases (472-478); Acute Upper respiratory diseases (460-465); Acute Lower Respiratory (466, 480-487, 512, 513, 518, 519) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM; LOESS <b>Age Groups Analyzed:</b> <65; ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.8 (0.5) ppm <b>Range (Min, Max):</b> (0.1, 5.1) <b>Copollutant:</b> TSP; PM <sub>10</sub> ; PM <sub>2.5</sub> ; Sulfates; COH; SO <sub>2</sub> ; NO <sub>2</sub> ; NO; O <sub>3</sub>	The study did not present quantitative results for CO.
<b>Author:</b> Goldberg et al. (2003) <b>Period of Study:</b> 1984 – 1993 <b>Location:</b> Montreal, Quebec, Canada	<b>Health Outcome (ICD9):</b> Mortality: CHF (428) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GLM, natural splines <b>Age Groups Analyzed:</b> ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.8 (0.5) ppm <b>Range (Min, Max):</b> (0.1, 5.1) <b>Copollutant:</b> PM <sub>2.5</sub> ; Sulfate; SO <sub>2</sub> ; NO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 0.50 ppm <b>% Increase (Lower CI, Upper CI); lag:</b> <b>Daily mortality from CHF</b> -0.99% (-6.31 to 4.63); 0 0.12% (-5.29 to 5.84); 1 -1.38% (-8.81 to 6.66); 0-2 <b>Daily mortality among persons classified as having CHF before death</b> 2.10% (-0.24 to 4.49); 0 2.28% (-0.09 to 4.72); 1 2.86% (-0.46 to 6.29); 0-2
<b>Author:</b> Goldberg et al. (2006) <b>Period of Study:</b> 1984 – 1993 <b>Location:</b> Montreal, Quebec, Canada	<b>Health Outcome (ICD9):</b> Mortality: Diabetes (250) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson, natural splines <b>Age Groups Analyzed:</b> ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.8 (0.5) ppm <b>Range (Min, Max):</b> (0.1, 5.1) <b>Copollutant:</b> PM <sub>2.5</sub> ; Sulfate; SO <sub>2</sub> ; NO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 0.50 ppm <b>% Increase (Lower CI, Upper CI); lag:</b> <b>Daily mortality from diabetes</b> 2.64% (-2.56 to 8.12); 0 6.54% (1.31-12.03); 1 8.08% (1.03-15.62); 0-2 <b>Daily mortality among persons classified as having diabetes before death</b> 1.15% (-1.69 to 4.07); 0 1.30% (-1.58 to 4.27); 1 2.63% (-1.42 to 6.85); 0-2
<b>Author:</b> Gouveia et al. (2000b) <b>Period of Study:</b> 1991 – 1993 <b>Location:</b> Sao Paulo, Brazil	<b>Health Outcome (ICD9):</b> Mortality: Respiratory; Cardiovascular; All other causes <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson <b>Age Groups Analyzed:</b> All ages >65 <5	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h moving avg <b>Mean (SD) unit:</b> 5.8 (2.1) ppm <b>Range (Min, Max):</b> (1.3, 16.2) <b>Copollutant:</b> PM <sub>10</sub> ; SO <sub>2</sub> ; NO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 5.1 ppm <b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Age Group: All ages:</b> All-causes 1.012 (0.994-1.031); 0 <b>Age Group: &gt;65</b> All-causes: 1.020 (0.996-1.046); 0 Respiratory: 0.981 (0.927-1.037); 2 CVD: 1.041 (1.007-1.076); 0 <b>Age Group: &lt;5</b> Respiratory: 1.086 (0.950-1.238); 0 Pneumonia: 1.141 (0.962-1.321); 2

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Gwynn et al. (2000) <b>Period of Study:</b> 5/1988 – 10/1990 <b>Location:</b> Buffalo, NY	<b>Health Outcome (ICD9):</b> Mortality: Respiratory (466, 480-486); Circulatory (401-405, 410-414, 415-417); All non-accidental causes (<800) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GLM <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant correlation:</b> H+: r = 0.15; SO <sub>4</sub> <sup>2-</sup> : r = 0.24; O <sub>3</sub> : r = -0.23; SO <sub>2</sub> : r = 0.11; NO <sub>2</sub> : r = 0.65	<b>Increment:</b> NR <b>β (SE); lag:</b> Respiratory mortality: 0.032466 (0.053802); 0 Circulatory mortality: 0.039216 (0.026544); 3 Total mortality: 0.040214 (0.015205); 3
<b>Author:</b> Hoek et al. (2001) <b>Period of Study:</b> 1986-1994 <b>Location:</b> The Netherlands	<b>Health Outcome (ICD9):</b> Mortality: Heart Failure (428); Arrhythmia (427); Cerebrovascular (430-436); Thrombotic (433, 434, 444, 452, 453); Cardiovascular (390-448) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant:</b> O <sub>3</sub> ; BS; PM <sub>10</sub> ; SO <sub>2</sub> ; NO <sub>2</sub>	<b>Increment:</b> 120 µg/m <sup>3</sup> <b>Relative Risk (Lower CI, Upper CI); Lag</b> Total CVD mortality: 1.026 (0.993-1.060); 0-6 MI and other IHD mortality: 1.050 (1.004-1.099); 0-6 Arrhythmia: 1.062 (0.937-1.203); 0-6 Heart failure mortality: 1.109 (1.012-1.216); 0-6 Cerebrovascular mortality: 1.066 (1.029-1.104); 0-6 Embolism, thrombosis: 1.065 (0.926-1.224); 0-6
<b>Author:</b> Hoek et al. (2000) <b>Period of Study:</b> 1986 – 1994 <b>Location:</b> The Netherlands	<b>Health Outcome (ICD9):</b> Mortality: Pneumonia (480-486); COPD (490-496); Cardiovascular diseases (CVD) (390-448) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, LOESS <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> Netherlands: 457 µg/m <sup>3</sup> Four Major Cities: 589 µg/m <sup>3</sup> <b>Range (Min, Max):</b> Netherlands: (174, 2620) Four Major Cities: (202, 4621) <b>Copollutant correlation:</b> PM <sub>10</sub> : r = 0.64; BS: r = 0.89; O <sub>3</sub> : r = -0.48; NO <sub>2</sub> : r = 0.89; SO <sub>2</sub> : r = 0.65; SO <sub>4</sub> <sup>2-</sup> : r = 0.55; NO <sub>3</sub> : r = 0.58	<b>Increment:</b> Single-day lag (1): 1,500 µg/m <sup>3</sup> Weekly avg (0-6): 1200 µg/m <sup>3</sup> <b>Relative Risk (Lower CI, Upper CI); Lag</b> CO Four Major Cities: 1.022 (0.995-1.050); 1 Four Major Cities: 1.044 (1.008-1.082); 0-6 Netherlands w/o Major Cities: 1.040 (1.020-1.060); 1 Netherlands w/o Major Cities: 1.051 (1.026-1.076); 0-6 avg Entire Netherlands: 1.035 (1.018-1.052); 1 Entire Netherlands: 1.046 (1.025-1.068); 0-6 <b>CVD:</b> 1.044 (1.012-1.077); 0-6 <b>COPD:</b> 1.194 (1.099-1.298); 0-6 <b>Pneumonia:</b> 1.276 (1.143-1.426); 0-6 <b>Winter:</b> 1.038 (1.013-1.063); 0-6 <b>Summer:</b> 1.199 (1.108-1.296); 0-6 <b>Multi-pollutant model</b> CO, PM <sub>10</sub> Total mortality: 0.969 (0.914-1.028); 0-6 CVD: 1.005 (0.918-1.101); 0-6 BS, CO Total mortality: 0.980 (0.933-1.030); 0-6 CVD: 0.927 (0.860-0.999); 0-6 CO, SO <sub>4</sub> <sup>2-</sup> Total mortality: 0.990 (0.951-1.030); 0-6 CVD: 0.999 (0.939-1.063); 0-6
<b>Author:</b> Honda et al. (2003) <b>Period of Study:</b> 1976-1990 <b>Location:</b> Tokyo, Japan	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson <b>Age Groups Analyzed:</b> ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Median (SD) unit:</b> 1.6 ppm <b>Range (Min, Max):</b> (0, 6.8) <b>Copollutant correlation:</b> NO: r = 0.403; NO <sub>2</sub> : r = 0.415; Oxidant: r = 0.396; SO <sub>2</sub> : r = 0.675	<b>Increment:</b> NR <b>Rate Ratio (Lower CI, Upper CI); lag:</b> <b>CO concentration</b> <1.1 ppm: 1.00 (reference category) 1.1-1.6 ppm: 1.017 (1.009, 1.026) 1.6-2.2 ppm: 1.031 (1.020, 1.041) >2.2 ppm: 1.051 (1.039, 1.063)
<b>Author:</b> Hong et al. (2002b) <b>Period of Study:</b> 1/1991-12/1997 <b>Location:</b> Seoul, Korea	<b>Health Outcome (ICD9):</b> Mortality: Hemorrhagic and ischemic stroke (431-434) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, LOESS <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.44 (0.70) ppm <b>Range (Min, Max):</b> (0.430, 5.14) <b>Copollutant:</b> TSP; SO <sub>2</sub> ; NO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 0.76 ppm <b>Relative Risk (Lower CI, Upper CI); lag:</b> 1.06 (1.02, 1.09); 1 <b>Multipollutant:</b> CO, TSP: 1.07 (1.03, 1.11); 1 CO, NO <sub>2</sub> : 1.06 (1.00, 1.11); 1 CO, SO <sub>2</sub> : 1.05 (1.01, 1.10); 1 CO, O <sub>3</sub> : 1.09 (1.05, 1.13); 1

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Hong et al. (1999a) <b>Period of Study:</b> 1/1995 - 12/1995 <b>Location:</b> Incheon, Korea	<b>Health Outcome (ICD9):</b> Mortality: Cardiovascular (400-440); Respiratory (460-519); Non-accidental causes (<800) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, LOESS <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.7 (0.8) ppm <b>Range (Min, Max):</b> (0.3, 5.1) <b>Copollutant:</b> SO <sub>2</sub> ; NO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 1 ppm <b>Relative Risk (Lower CI, Upper CI); lag:</b> Total mortality: 0.993 (0.950, 1.037); 0-4 Cardiovascular mortality: 0.965 (0.892, 1.044); 0-4
<b>Author:</b> Hong et al. (2002a) <b>Period of Study:</b> 1/1995 - 12/1998 <b>Location:</b> Seoul, Korea	<b>Health Outcome (ICD9):</b> Mortality: Stroke (160-169) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.2 (0.5) ppm <b>Range (Min, Max):</b> (0.4, 3.4) <b>Copollutant: correlation</b> PM <sub>10</sub> : r = 0.22; NO <sub>2</sub> : r = 0.64; SO <sub>2</sub> : r = 0.90; O <sub>3</sub> : r = -0.35	<b>Increment:</b> 0.3 ppm <b>% Increase (Lower CI, Upper CI); lag:</b> CO: 2.2% (0.4, 4.1); 2 CO (stratified by PM <sub>10</sub> concentration): <median concentration of PM <sub>10</sub> : 1.1; 2 ≥ median concentration of PM <sub>10</sub> : 3.6; 2
<b>Author:</b> Hong et al. (1999b) <b>Period of Study:</b> 1/1995 - 8/1996 <b>Location:</b> Incheon, South Korea	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Respiratory; Cardiovascular <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM; LOESS <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 15.2 (7.1) ppb <b>Range (Min, Max):</b> (2.9, 51.2) <b>Copollutant:</b> PM <sub>10</sub> ; NO <sub>2</sub> ; SO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 100 ppb <b>β (SE); lag:</b> <b>Total Mortality</b> CO 0.0019 (0.0015); 1 0.0024 (0.0041); 0-4 CO, PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub> -0.0009 (0.0019); 1 -0.0018 (0.0043); 0-4 <b>Cardiovascular Mortality</b> CO 0.0019 (0.0073); 1 -0.0008 (0.0028); 0-4 CO, PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub> -0.0053 (0.0078); 1 -0.0037 (0.0033); 0-4 <b>Respiratory Mortality</b> CO 0.0148 (0.0065); 1 0.0063 (0.0171); 0-4 CO, PM <sub>10</sub> , NO <sub>2</sub> , SO <sub>2</sub> , O <sub>3</sub> 0.0121 (0.0079); 1 -0.0034 (0.0183); 0-4
<b>Author:</b> Keatinge et al. (2001) <b>Period of Study:</b> 1976-1995 <b>Location:</b> London, England	<b>Health Outcome (ICD9):</b> Mortality: Non-accidental causes (<800) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Single and multiple delay regression <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant:</b> SO <sub>2</sub> ; PM <sub>10</sub>	The study did not present quantitative results for CO.
<b>Author:</b> Kettunen et al. (2007) <b>Period of Study:</b> 1998-2004 <b>Location:</b> Helsinki, Finland	<b>Health Outcome (ICD10):</b> Mortality: Stroke (I60-I61, I63-I64) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, penalized thin-plate splines <b>Age Groups Analyzed:</b> ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h moving avg <b>Median (SD) unit:</b> Cold Season: 0.5 mg/m <sup>3</sup> Warm Season: 0.4 mg/m <sup>3</sup> <b>Range (Min, Max):</b> Cold Season: (0.1, 2.4) Warm Season: (0.1, 1.1) <b>Copollutant: correlation</b> <b>Cold Season:</b> PM <sub>2.5</sub> : r = 0.32; UFP: r = 0.47 <b>Warm Season:</b> PM <sub>2.5</sub> : r = 0.24; UFP: r = 0.39	<b>Increment:</b> 0.2 mg/m <sup>3</sup> <b>% Increase (Lower CI, Upper CI); lag:</b> <b>Cold Season</b> 0.47 (-3.29 to 4.39); 0; / -0.63 (-4.39 to 3.28); 1; -2.69 (-6.46 to 1.24); 2; / -0.19 (-3.93 to 3.69); 3 <b>Warm Season</b> 3.95 (-3.78 to 12.30); 0; / 8.33 (0.63 to 16.63); 1; 6.97 (-0.59 to 15.11); 2; / 7.54 (-0.05 to 15.71); 3

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Klemm et al. (2004)</p> <p><b>Period of Study:</b> 8/1998-7/2000</p> <p><b>Location:</b> Fulton County and DeKalb County, GA (ARIES)</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Non-accidental (&lt;800); Cardiovascular (390-459); Respiratory (460-519); Cancer (140-239)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GLM, natural cubic splines</p> <p><b>Age Groups Analyzed:</b> &lt;65; ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 1-h maximum</p> <p><b>Median (SD) unit:</b> 1,310 (939.13) ppb</p> <p><b>Range (Min, Max):</b> (303.58, 7400)</p> <p><b>Copollutant:</b> PM<sub>2.5</sub>; PM<sub>10-2.5</sub>; O<sub>3</sub>; NO<sub>2</sub>; SO<sub>2</sub>; Acid; EC; OC; SO<sub>4</sub>; Oxygenated HCs; NMHCs; NO<sub>3</sub></p>	<p><b>Increment:</b> NR</p> <p><b>β (SE); lag:</b> Quarterly Knots: 0.00002 (0.00001); 0-1 Monthly Knots: 0.00002 (0.00001); 0-1 Biweekly Knots: 0.00001 (0.00002); 0-1</p>
<p><b>Author:</b> Kwon et al. (2001)</p> <p><b>Period of Study:</b> 1994-1998</p> <p><b>Location:</b> Seoul, Korea</p>	<p><b>Health Outcome (ICD9):</b> Mortality: CHF (428); Cardiovascular (390-459)</p> <p><b>Study Design:</b> 1. Time-series 2. Bi-directional case-crossover</p> <p><b>Statistical Analyses:</b> 1. Poisson GLM, LOESS 2. Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> &lt;55 55-64 65-74 75-84 ≥ 85</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 1-h avg</p> <p><b>Mean (SD) unit:</b> 12.4 ppb</p> <p><b>Range (Min, Max):</b> (4.1, 38.0)</p> <p><b>Copollutant correlation:</b> PM<sub>10</sub>: r = 0.713; NO<sub>2</sub>: r = 0.744; SO<sub>2</sub>: r = 0.843; O<sub>3</sub>: r = -0.367</p>	<p><b>Increment:</b> 0.59 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p><b>From GAM approach</b> CHF patients: 1.054 (0.991-1.121); 0; 0 General Population: 1.022 (1.017- 1.029); 0</p> <p><b>From case-crossover design</b> CHF patients: 1.033 (0.946-1.127); 0 General Population: 1.007 (0.997- .016); 0</p> <p><b>Modifiers and CHF patients (case-crossover design)</b></p> <p><b>Gender</b> Male: 1.025 (0.890-1.180); 0 Female: 1.035 (0.925-1.157); 0</p> <p><b>Age Group:</b> &lt;75: 0.948 (0.890-1.180); 0 ≥ 75: 1.116 (0.989-1.258); 0</p> <p><b>Time from admission to death</b> 4 or less weeks: 1.088 (0.907-1.306); 0 &gt;4 weeks: 1.017 (0.920-1.124); 0 Total mortality: 1.033 (0.946-1.127); 0 Cardiovascular mortality: 1.033 (0.920-1.160); 0 Cardiac death: 1.052 (0.919-1.204); 0</p> <p><b>Two-pollutant model in CHF patients (case-crossover design)</b> CO alone: 1.054 (0.991-1.121); 0 CO, PM<sub>10</sub>: 1.096 (0.981-1.224); 0 CO, NO<sub>2</sub>: 1.022 (0.932-1.122); 0 CO, SO<sub>2</sub>: 1.014 (0.909-1.131); 0 CO, O<sub>3</sub>: 1.056 (0.992-1.124); 0</p>
<p><b>Author:</b> Lee et al. (2007c)</p> <p><b>Period of Study:</b> 1/2000-12/2004</p> <p><b>Location:</b> Seoul, Korea</p>	<p><b>Health Outcome (ICD10):</b> Mortality: Non-accidental (A00-R99)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Maximum 8-h moving avg</p> <p><b>Mean (SD) unit:</b> w/ Asian dust days: 0.92 (0.42) ppm w/o Asian dust days: 0.92 (0.41) ppm Asian dust days only: 1.00 (0.47) ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> PM<sub>10</sub>; NO<sub>2</sub>; SO<sub>2</sub>; O<sub>3</sub></p>	<p><b>Increment:</b> 0.54 ppm</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p>Model with Asian Dust Days: 3.3% (2.5-4.1); 1 Model without Asian dust days: 3.3% (2.5-4.2); 1</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Lipfert et al. (2000b)</p> <p><b>Period of Study:</b> 5/1992 - 9/1995</p> <p><b>Location:</b> Philadelphia, PA, three nearby suburban Pennsylvania counties, and three nearby New Jersey counties</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Respiratory (460-519); Cardiac (390-448); Cancer; Other causes (&lt;800)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Stepwise regression</p> <p><b>Age Groups Analyzed:</b> &lt;65 ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg; 1-h maximum</p> <p><b>Mean (SD) unit:</b> Camden: 24-h avg: 0.75 (0.40) ppm Philadelphia: 24-h avg: 0.63 (0.40) ppm 1-h maximum: 1.44 (1.04)</p> <p><b>Range (Min, Max):</b> Camden: (0.10, 3.8) Philadelphia: 24-h avg: (0.10, 3.3) 1-h maximum: (0.0, 7.8)</p> <p><b>Copollutant:</b> NO; NO<sub>2</sub>; O<sub>3</sub>; SO<sub>2</sub>; SO<sub>4</sub><sup>2-</sup>; PM<sub>10</sub>; PM<sub>2.5</sub></p>	<p><b>Increment:</b> NR</p> <p><b>Attributable Risk; lag:</b></p> <p><b>Peak CO</b> All-cause Philadelphia: 0.0054; 0-1 4 Pennsylvania Counties: 0.0081; 0-1 Pennsylvania + NJ: 0.0085; 0-1</p> <p><b>CO</b> All seven counties in Pennsylvania and New Jersey <b>All ages</b> Respirator y: -0.0067; Cardiac: 0.0131; Other: 0.0078</p> <p><b>All-cause:</b> &lt;65: 0.0148; 0-1; ≥ 65: 0.0054; 0-1</p> <p><b>Joint model with CO</b> Philadelphia: 0.0059; 0-1 4 Pennsylvania Counties: 0.0089; 0-1 Pennsylvania + NJ: 0.0096; 0-1 Cardiac: 0.0135; 0-1; Other causes: 0.0084 &lt;65: 0.0154; 0-1; ≥ 65: 0.0060; 0-1</p>
<p><b>Author:</b> Lippmann et al. (2000)</p> <p><b>Period of Study:</b> 1985-1990 1992-1994</p> <p><b>Location:</b> Detroit, MI and Windsor, ON</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (&lt;800); Circulatory (390-459); Respiratory (460-519)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GLM</p> <p><b>Age Groups Analyzed:</b> ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1985-1990: 0.9 ppm 1992-1994: 0.72 ppm</p> <p><b>Range (5th, 95th):</b> 1985-1990: (.46, 1.61) 1992-1994: (0.36, 1.2)</p> <p><b>Copollutant correlation:</b> <b>1985-1990</b> PM<sub>10</sub>: r = 0.35; TSP: r = 0.28; TSP-PM<sub>10</sub>: r = 0.02; TSP-SO<sub>4</sub><sup>2-</sup>: r = 0.18; O<sub>3</sub>: r = -0.22; SO<sub>2</sub>: r = 0.36; NO<sub>2</sub>: r = 0.58</p> <p><b>1992-1994</b> PM<sub>10</sub>: r = 0.38; PM<sub>2.5</sub>: r = 0.38; PM<sub>10-2.5</sub>: r = 0.24; H+: r = 0.16; SO<sub>4</sub><sup>2-</sup>: r = 0.32; O<sub>3</sub>: r = 0.16; SO<sub>2</sub>: r = 0.42; NO<sub>2</sub>: r = 0.68</p>	<p><b>Increment:</b> 1985-1990: 11.5 ppm; 1992-1994: 8.4 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p><b>1985-1990</b> <b>Total Mortality:</b> 0.9842 (0.9667-1.002); 0 1.0103 (0.9926-1.0284); 1 1.0075 (0.9898-1.0254); 2 1.0145 (0.9967-1.0326); 3 0.9968 (0.9789-1.0151); 0-1 1.0105 (0.9925-1.0288); 1-2 1.0134 (0.9954-1.0317); 2-3 1.0003 (0.9823-1.0187); 0-2 1.0152 (0.9971-1.0336); 1-3 1.0053 (0.9873-1.0236); 0-3</p> <p><b>Circulatory Mortality:</b> 0.9818 (0.9574-1.0068); 0 0.9991 (0.9745-1.0243); 1 0.9980 (0.9735-1.0232); 2 1.0088 (0.9841-1.0341); 3 0.9888 (0.9640-1.0144); 0-1 0.9981 (0.9732-1.0237); 1-2 1.0042 (0.9792-1.0298); 2-3 0.9900 (0.9650-1.0157); 0-2 1.0029 (0.9777-1.0287); 1-3 0.9944 (0.9692-1.0202); 0-3</p> <p><b>Respiratory Mortality:</b> 0.9644 (0.9042-1.0287); 0 1.0142 (0.9518-1.0808); 1 1.0483 (0.9845-1.1164); 2 1.0468 (0.9828-1.1149); 3 0.9868 (0.9248-1.053); 0-1 1.0372 (0.9730-1.1056); 1-2 1.0554 (0.9904-1.1246); 2-3 1.0088 (0.9457-1.0762); 0-2 1.0466 (0.9817-1.1158); 1-3 1.0205 (0.9569-1.0884); 0-3</p> <p><b>Total minus respiratory and circulatory mortality:</b> 0.9939 (0.9668-1.0217); 0 1.0278 (1.0001-1.0562); 1 1.0178 (0.9902-1.0461); 2 1.0227 (0.9948-1.0514); 3 1.0127 (0.9860-1.0412); 0-1 1.0269 (0.9989-1.0556); 1-2 1.0249 (0.9968-1.0538); 2-3 1.0172 (0.9893-1.0458); 0-2 1.0322 (1.0041-1.0612); 1-3 1.0229 (0.9950-1.0516); 0-3</p> <p><b>1992-1994</b> <b>Total Mortality</b> 0.9933 (0.9636-1.024); 0</p>



Study	Design	Concentrations	Effect Estimates (95% CI)
			1.0162 (0.9860-1.0473); 1 1.0116 (0.9816-1.0426); 2 0.9947 (0.9648-1.0254); 3 1.0056 (0.9756-1.0366); 0-1 1.0165 (0.9864-1.0476); 1-2 1.0038 (0.9739-1.0476); 2-3 1.0098 (0.9796-1.0409); 0-2 1.0104 (0.9862-1.0414); 1-3 1.0064 (0.9755-1.0382); 0-3
			<b>Circulatory Mortality</b> 1.0076 (0.9640-1.0531); 0 1.0307 (0.9865-1.0768); 1 1.0142 (0.9705-1.0598); 2 0.9523 (0.9102-0.9964); 3 1.0229 (0.9788-1.0688); 0-1 1.0267 (0.9827-1.0727); 1-2 0.9802 (0.9375-1.0248); 2-3 1.0243 (0.9801-1.0726); 0-2 0.9987 (0.9553-1.0441); 1-3 1.0019 (0.9573-1.0487); 0-3
			<b>Respiratory Mortality</b> 0.9894 (0.8912-1.0984); 0 0.9474 (0.8521-1.0533); 1 0.9652 (0.8682-1.0732); 2 0.9931 (0.8934-1.1040); 3 0.9626 (0.8668-1.0691); 0-1 0.9485 (0.8535-1.0541); 1-2 0.9752 (0.8775-1.0838); 2-3 0.9555 (0.8802-1.0615); 0-2 0.9567 (0.8607-1.0635); 1-3 0.9584 (0.9604-1.0675); 0-3
			<b>Total minus respiratory and circulatory mortality:</b> 0.9769 (0.9332-1.0227); 0 1.0135 (0.9682-1.0609); 1 1.0195 (0.9747-1.0664); 2 1.0429 (0.9974-1.0905); 3 0.9940 (0.9494-1.0406); 0-1 1.0197 (0.9746-1.0670); 1-2 1.0371 (0.9918-1.0845); 2-3 1.0045 (0.9596-1.0515); 0-2 1.0353 (0.9896-1.0831); 1-3 1.0215 (0.9749-1.0702); 0-3
<b>Author:</b> Maheswaran et al. (2005a)	<b>Health Outcome (ICD9):</b> Mortality: CHD (410-414)	<b>Pollutant:</b> CO	<b>Increment:</b> NR
<b>Period of Study:</b> 1994-1998	<b>Study Design:</b> Ecological	<b>Averaging Time:</b> 24-h avg	<b>Rate Ratios (Lower CI, Upper CI):</b>
<b>Location:</b> Sheffield, United Kingdom	<b>Statistical Analyses:</b> Poisson	<b>Mean (SD) unit:</b> NR	<b>CO</b>
	<b>Age Groups Analyzed:</b> ≥ 45	<b>Range (Min, Max):</b> NR	<b>Adjusted for sex and age</b>
		<b>Copollutant:</b> NO <sub>x</sub> ; PM <sub>10</sub>	Quintile:
		<b>Notes:</b> Quintiles represent the following mean CO concentrations and category limits: 5: 482 µg/m <sup>3</sup> (≥ 455) 4: 443 µg/m <sup>3</sup> (≥ 433 to <455) 3: 426 µg/m <sup>3</sup> (≥ 419 to <433) 2: 405 µg/m <sup>3</sup> (≥ 387 to <419) 1: 360 µg/m <sup>3</sup> (<387)	5 (highest): 1.24 (1.14, 1.36) 4: 1.30 (1.19, 1.41) 3: 1.15 (1.05, 1.25) 2: 1.08 (0.99, 1.17) 1: (lowest): 1.00
			<b>CO</b>
			<b>Adjusted for sex, age, deprivation, and smoking</b>
			<b>Quintile:</b>
			5 (highest): 1.05 (0.95, 1.16); 4: 1.16 (1.06, 1.28); 3: 1.04 (0.95, 1.14); 2: 1.03 (0.94, 1.13); 1 (lowest): 1.00
			<b>CO</b>
			<b>Adjusted for sex, age, deprivation, and smoking (spatially smoothed using a 1 km radius)</b>
			<b>Quintile:</b>
			5 (highest): 1.07 (0.96, 1.18); 4: 1.13 (1.03, 1.24); 3: 1.04 (0.95, 1.14); 2: 1.01 (0.92, 1.10); 1 (lowest): 1.00

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Maheswaran et al. (2005b)</p> <p><b>Period of Study:</b> 1994-1998</p> <p><b>Location:</b> Sheffield, United Kingdom</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Stroke deaths (430-438)</p> <p><b>Study Design:</b> Ecological</p> <p><b>Statistical Analyses:</b> Poisson</p> <p><b>Age Groups Analyzed:</b> ≥ 45</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> Quintile: 5: 482 µg/m<sup>3</sup>; 4: 443 µg/m<sup>3</sup>; 3: 426 µg/m<sup>3</sup>; 2: 405 µg/m<sup>3</sup>; 1: 360 µg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant correlation :</b> PM<sub>10</sub>: r = 0.88; NO<sub>x</sub>: r = 0.87</p> <p><b>Notes:</b> Quintiles represent the following mean CO concentrations and category limits: 5: 482 µg/m<sup>3</sup> (≥ 455) 4: 443 µg/m<sup>3</sup> (≥ 433 to &lt;455) 3: 426 µg/m<sup>3</sup> (≥ 419 to &lt;433) 2: 405 µg/m<sup>3</sup> (≥ 387 to &lt;419) 1: 360 µg/m<sup>3</sup> (&lt;387)</p>	<p><b>Increment:</b> NR</p> <p><b>Rate Ratios (Lower CI, Upper CI); lag:</b> RR for mortality and CO modeled outdoor air pollution</p> <p><b>Adjusted for sex and age Quintile:</b> 5 (highest): 1.35 (1.19, 1.53); 4: 1.40 (1.24, 1.58); 3: 1.08 (0.95, 1.23); 2: 1.10 (0.97, 1.24); 1 (lowest): 1.00</p> <p><b>Adjusted for sex, age, deprivation, and smoking Quintile:</b> 5 (highest): 1.26 (1.10, 1.46); 4: 1.32 (1.15, 1.50); 3: 1.07 (0.93, 1.22); 2: 1.12 (0.99, 1.28); 1 (lowest): 1.00</p> <p><b>Not spatially smoothed CO outdoor air pollution Quintile:</b> 5 (highest): 1.26 (1.10, 1.46); 4: 1.32 (1.15, 1.50); 3: 1.07 (0.93, 1.22); 2: 1.12 (0.99, 1.28); 1 (lowest): 1.00</p> <p><b>Spatially smoothed using a 1-km radius Quintile:</b> 5 (highest): 1.16 (1.01, 1.34); 4: 1.22 (1.07, 1.39); 3: 0.95 (0.83, 1.09); 2: 0.97 (0.85, 1.11); 1 (lowest): 1.00</p>
<p><b>Author:</b> Mar et al. (2000)</p> <p><b>Period of Study:</b> 1995-1997</p> <p><b>Location:</b> Phoenix, AZ</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (&lt;800); Cardiovascular (390-449)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson</p> <p><b>Age Groups Analyzed:</b> &gt;65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.5 (0.8) ppm</p> <p><b>Range (Min, Max):</b> 1995: (0.5, 4.0) ppm 1996: (0.3, 4.0) ppm 1997: (0.3, 3.7) ppm</p> <p><b>Copollutant correlation:</b> PM<sub>2.5</sub>: r = 0.85; PM<sub>10</sub>: r = 0.53; PM<sub>10-2.5</sub>: r = 0.34; NO<sub>2</sub>: r = 0.87; O<sub>3</sub>: r = -0.40; SO<sub>2</sub>: r = 0.53</p>	<p><b>Increment:</b> 1.19 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b> Total Mortality (CO exposure): 1.06 (1.02, 1.09); 0; 1.05 (1.01, 1.09); 1</p> <p><b>Cardiovascular Mortality (CO exposure):</b> 1.05 (1.00, 1.11); 0; 1.10 (1.04, 1.15); 1; 1.07 (1.02, 1.12); 2; 1.07 (1.02, 1.12); 3; 1.08 (1.03, 1.13); 4</p>
<p><b>Author:</b> Moolgavkar et al. (2000b)</p> <p><b>Period of Study:</b> 1987-1995</p> <p><b>Location:</b> Cook County, IL Los Angeles County, CA Maricopa County, AZ</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Circulatory (390-448); Cardiovascular (390-429); Cerebrovascular (430-448); COPD (490-496); Asthma (493)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, spline smoother</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Median unit:</b> Cook county: 993 ppb Los Angeles: 1347 ppb Maricopa: 1240 ppb</p> <p><b>Range (Min, Max):</b> Cook county: (224, 3912) Los Angeles: (237, 5955) Maricopa: (269, 4777)</p> <p><b>Copollutant correlation :</b> PM<sub>10</sub>: Cook: r = 0.30; LA: r = 0.45; Maricopa: r = 0.20</p> <p>NO<sub>2</sub>: Cook: r = 0.63; LA: r = 0.80; Maricopa: r = 0.66</p> <p>SO<sub>2</sub>: Cook: r = 0.35; LA: r = 0.78; Maricopa: r = 0.53</p> <p>O<sub>3</sub>:</p>	<p><b>Increment:</b> 1 ppm</p> <p><b>% Change (Lower CI, Upper CI); lag:</b> CVD Mortality Cook County CO -1.07 (-2.67, 0.54); 0; / 1.25 (-0.36, 2.87); 1; 1.49 (-0.09, 3.07); 2; / 1.90 (0.32, 3.48); 3; 1.44 (-0.16, 3.03); 4; / 0.72 (-0.89, 2.32); 5</p> <p><b>Los Angeles County CO</b> 3.47 (2.94, 4.00); 0; / 3.93 (3.41, 4.46); 1; 4.08 (3.56, 4.60); 2; / 3.76 (3.24, 4.28); 3; 2.91 (2.37, 3.44); 4; / 2.63 (2.09, 3.17); 5</p> <p><b>CO, PM<sub>10</sub></b> 2.27 (0.88, 3.66); 0; / 4.33 (2.96, 5.69); 1; 4.72 (3.38, 6.05); 2; / 4.26 (2.90, 5.63); 3; 2.49 (1.10, 3.88); 4; / 5.93 (4.60, 7.27); 5</p> <p><b>CO and PM<sub>2.5</sub></b> 0.43 (-1.35, 2.20); 0; / 2.88 (1.16, 4.60); 1; 4.65 (2.93, 6.37); 2; / 5.93 (4.20, 7.65); 3; 3.88 (2.13, 5.63); 4; / 5.85 (4.12, 7.58); 5</p> <p><b>Maricopa County CO</b> 0.81 (-0.79, 2.39); 0; / 2.20 (0.61, 3.79); 1;</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
		Cook: r = -0.28; LA: r = -0.52; Maricopa: r = -0.61	3.05 (1.49, 4.61); 2; / 3.78 (2.27, 5.28); 3; 3.73 (2.27, 5.19); 4; / 2.25 (0.76, 3.72); 5 <b>COPD Mortality</b> <b>Cook County</b> CO -2.65 (-7.05, 1.75); 0; / 2.80 (-1.60, 7.19); 1; 0.98 (-3.34, 5.31); 2; / 2.20 (-2.12, 6.53); 3; 1.31 (-3.06, 5.68); 4; / 1.59 (-2.78, 5.97); 5 <b>Los Angeles County</b> CO 3.78 (2.31, 5.25); 0; / 5.23 (3.78, 6.69); 1; 5.71 (4.26, 7.17); 2; / 5.42 (3.95, 6.89); 3; 4.01 (2.51, 5.50); 4; / 3.82 (2.31, 5.33); 5 <b>Maricopa County</b> CO 1.29 (-2.19, 4.76); 0; / 4.63 (1.17, 8.09); 1; 0.07 (-3.36, 3.50); 2; / 3.00 (-0.30, 6.30); 3; 6.21 (3.02, 9.40); 4; / 3.27 (0.04, 6.50); 5 <b>Cerebrovascular Disease Mortality</b> <b>Cook County</b> -0.41 (-3.30, 2.47); 0; / 3.13 (0.23, 6.02); 1; 2.12 (-0.73, 4.97); 2; / 1.00 (-1.85, 3.86); 3; 2.50 (-0.36, 5.37); 4; / 1.88 (-1.00, 4.76); 5 <b>Los Angeles County</b> 3.31 (2.32, 4.31); 0; / 3.88 (2.89, 4.87); 1; 3.23 (2.25, 4.22); 2; / 2.65 (1.66, 3.65); 3; 2.11 (1.11, 3.12); 4; / 2.04 (1.02, 3.06); 5 <b>Maricopa County</b> 0.26 (-2.65, 3.16); 0; / 3.50 (0.60, 6.41); 1; 3.52 (0.66, 6.38); 2; / 4.61 (1.85, 7.37); 3; 4.78 (2.10, 7.46); 4; / 5.15 (2.45, 7.84); 5 <b>Notes:</b> Total Mortality effect estimates were not presented quantitatively.
<b>Author:</b> Moolgavkar et al. (2003b) <b>Period of Study:</b> 1987-1995 <b>Location:</b> Cook County, Illinois & Los Angeles County, California	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Circulatory (390-448) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM <b>Age Groups Analyzed:</b> All Ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Median unit:</b> Cook County: 993 ppb LA County: 1347 ppb <b>Range (Min, Max):</b> Cook County: (224, 3912) ppb LA County: (237, 5955) ppb <b>Copollutant correlation:</b> <b>Cook County:</b> NO <sub>2</sub> : r = 0.63; O <sub>3</sub> : r = -0.22; SO <sub>2</sub> : r = 0.35; PM <sub>10</sub> : r = 0.30 <b>LA County:</b> NO <sub>2</sub> : r = 0.80; O <sub>3</sub> : r = -0.52; SO <sub>2</sub> : r = 0.78; PM <sub>10</sub> : r = 0.45; PM <sub>2.5</sub> : r = 0.58	<b>Increment:</b> 1 ppm <b>% Increase (t-statistic); lag</b> <b>Total Mortality Cook County</b> <b>CO:</b> 0.6% (1.2); 0; / 2.5% (5.4); 1; / 1.2% (2.6); 2; 1.5% (3.2); 3; / 1.1% (2.5); 4; / 0.6% (1.3); 5 <b>CO, PM<sub>10</sub>:</b> -0.5% (-1.0); 0; / 2.2% (4.3); 1; / 1.1% (2.2); 2; 1.0% (1.9); 3; / 1.1% (2.1); 4; / 1.4% (2.7); 5 <b>Total Mortality Los Angeles County</b> <b>CO:</b> 1.3% (7.4); 0; / 1.9% (10.5); 1; / 1.6% (8.9); 2; 1.4% (8.1); 3; / 1.0% (5.9); 4; / 0.7% (4.1); 5 <b>CO, PM<sub>10</sub>:</b> 0% (0); 0; / 2.2% (4.8); 1; / 1.4% (3.1); 2; 0.8% (1.8); 3; / 0.7% (1.6); 4; / 1.3% (3.0); 5 <b>CO, PM<sub>2.5</sub>:</b> -0.1% (-1.5); 0; / 1.5% (2.5); 1; / 2.4% (3.8); 2; 0.3% (0.5); 3; / 1.6% (2.8); 4; / 1.5% (2.6); 5 <b>Total Mortality (Season-specific) Cook County</b> <b>Spring (CO):</b> 0.8% (0.9); 0; / 2.4% (2.9); 1; / 0% (0); 2; 1.2% (1.5); 3; / 0.8% (1.0); 4; / -0.1% (-0.2); 5 <b>Summer (CO):</b> 1.2% (1.0); 0; / 3.6% (3.0); 1; / 4.2% (3.6); 2; -0.3% (-0.2); 3; / -1.1% (-1.0); 4; / -0.7% (-0.6); 5 <b>Fall (CO):</b> 1.2% (1.5); 0; / 2.1% (2.7); 1; / 0% (0); 2; 0% (0); 3; / -0.5% (-0.6); 4; / -0.7% (-0.9); 5 <b>Winter (CO):</b> -0.7% (-1.0); 0; / 1.8% (2.3); 1; / -0.2% (-0.3); 2; 0.5% (0.6); 3; / 1.2% (1.5); 4; / 1.0% (1.3); 5 <b>Los Angeles County</b> <b>Total Mortality (Season-specific)</b> <b>Spring (CO):</b> 3.6% (6.3); 0; / 3.5% (6.2); 1; / 1.9% (3.4); 2; 0.6% (1.0); 3; / -0.5% (-0.8); 4; / -0.7% (-1.2); 5

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Ostro et al. (1999)  <b>Period of Study:</b> 1989-1992  <b>Location:</b> Coachella Valley, California</p>	<p><b>Health Outcome (ICD9):</b>  Mortality: Total (non-accidental) (&lt;800); Respiratory (460-519); Cardiovascular (393-440)  <b>Study Design:</b> Time-series  <b>Statistical Analyses:</b> Poisson GAM; LOESS  <b>Age Groups Analyzed:</b> &gt;50</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 1-h maximum  <b>Mean (SD) unit:</b> 1.35 ppm  <b>Range (Min, Max):</b> (0, 6.0)  <b>Copollutant correlation:</b>  PM<sub>10</sub>: r = -0.18; O<sub>3</sub>: r = -0.47; NO<sub>2</sub>: r = 0.65</p>	<p><b>Summer (CO):</b>  3.0% (3.0); 0; / 4.7% (4.6); 1; / 5.2% (5.1); 2;  4.1% (3.8); 3; / 1.9% (1.8); 4; / 1.4% (1.3); 5</p> <p><b>Fall (CO):</b>  1.8% (4.6); 0; / 2.0% (5.1); 1; / 1.0% (2.6); 2;  0.6% (1.5); 3; / 0.4% (1.2); 4; / 0.2% (0.6); 5</p> <p><b>Winter (CO):</b>  0% (0); 0; / 0.8% (2.5); 1; / 0.9% (3.1); 2;  1.0% (3.4); 3; / 0.5% (1.7); 4; / 0.5% (1.6); 5</p> <p><b>CVD Mortality Cook County</b>  <b>CO:</b>  -1.1% (-1.5); 0; / 1.8% (2.5); 1; / 1.5% (2.2); 2;  1.6% (2.4); 3; / 1.4% (2.1); 4; / 0.7% (1.0); 5</p> <p><b>CO, PM<sub>10</sub>:</b>  -2.1% (-2.6); 0; / 1.5% (1.8); 1; / 1.4% (1.7); 2;  0.1% (1.1); 3; / 1.4% (1.9); 4; / 1.6% (2.1); 5</p> <p><b>CVD Mortality Los Angeles County</b>  <b>CO:</b>  1.6% (6.3); 0; / 1.9% (7.6); 1; / 1.6% (6.6); 2;  1.9% (8.2); 3; / 1.6% (7.1); 4; / 1.4% (6.1); 5</p> <p><b>CO, PM<sub>10</sub>:</b>  -0.8% (-1.2); 0; / 1.9% (3.0); 1; / 2.7% (4.3); 2;  1.3% (2.2); 3; / 0.5% (0.9); 4; / 2.8% (4.7); 5</p> <p><b>CO, PM<sub>2.5</sub>:</b>  -2.2% (-2.7); 0; / 1.5% (1.8); 1; / 1.9% (2.0); 2;  1.9% (2.2); 3; / 2.1% (2.6); 4; / 3.7% (4.5); 5</p> <p><b>CVD Mortality (Season Specific) Cook County</b>  <b>Spring (CO):</b>  0.7% (0.5); 0; / 1.4% (1.1); 1; / 0.3% (0.3); 2;  1.1% (0.9); 3; / 0.4% (3.1); 4; / 0.1% (0.6); 5</p> <p><b>Summer (CO):</b>  -2.6% (-1.4); 0; / 2.5% (1.4); 1; / 6.5% (3.7); 2;  0.9% (0.5); 3; / -1.9% (-1.1); 4; / -1.0% (-0.6); 5</p> <p><b>Fall (CO):</b>  0% (0); 0; / 2.9% (2.5); / 1; 0% (0); 2;  0% (0); 3; / -0.8% (-0.7); / 4; 0% (0); 5</p> <p><b>Winter (CO):</b>  -2.5% (-2.2); 0; / 0.7% (0.6); 1; / 0% (0); 2;  1.3% (1.1); 3; / 0.8% (0.7); 4; / 0.4% (0.4); 5</p> <p><b>Los Angeles County</b>  <b>CVD Mortality (Season-specific)</b>  <b>Spring (CO):</b>  3.0% (3.7); 0; / 3.3% (4.1); 1; / 2.3% (2.9); 2;  0.7% (0.9); 3; / -1.2% (-1.6); 4; / 0% (0); 5</p> <p><b>Summer (CO):</b>  4.0% (2.8); 0; / 5.2% (3.5); 1; / 6.3% (4.3); 2;  5.0% (3.3); 3; / 3.1% (2.0); 4; / 3.6% (2.3); 5</p> <p><b>Fall (CO):</b>  2.3% (4.2); 0; / 2.1% (3.7); 1; / 1.1% (1.9); 2;  1.2% (2.2); 3; / 1.5% (2.9); 4; / 1.0% (1.8); 5</p> <p><b>Winter (CO):</b>  0.3% (0.8); / 0; 0.7% (1.7); 1; / 0.8% (2.0); 2;  1.4% (3.4); 3; / 1.0% (2.3); 4; / 1.1% (2.5); 5</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Penttinen et al. (2004) <b>Period of Study:</b> 1988-1996 <b>Location:</b> Helsinki, Finland	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Respiratory (460-519); Cardiovascular (393-440) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, LOESS <b>Age Groups Analyzed:</b> All ages 15-64 years 65-74 years ≥ 75	<b>Pollutant:</b> CO <b>Averaging Time:</b> Maximum 8-h avg <b>Median unit:</b> 1.2 mg/m <sup>3</sup> <b>Range (Min, Max):</b> (0, 12.4) <b>Copollutant correlation:</b> O <sub>3</sub> : r = -0.46; NO <sub>2</sub> : r = 0.59; SO <sub>2</sub> : r = 0.55; PM <sub>10</sub> : r = 0.45; TSP: r = 0.26; TSP Blackness: r = 0.26	<b>Increment:</b> 1 mg/m <sup>3</sup> <b>% Increase (Lower CI, Upper CI); lag:</b> <b>Total Mortality</b> -1.50% (-2.78, -0.22); 0 0.15% (-1.09, 1.39); 1 -1.00% (-2.80, 0.81); 0-3 <b>Cardiovascular Mortality</b> -2.48% (-4.30, -0.66); 0 -0.84% (-2.61, 0.93); 1 -1.87% (-4.43, 0.69); 0- <b>Respiratory Mortality</b> -0.48% (-4.84, 3.87); 0 -0.14% (-4.43, 4.15); 1 -1.49% (-7.73, 4.74); 0-3
<b>Author:</b> Peters et al. (2000a) <b>Period of Study:</b> 1982-1994 <b>Location:</b> Northern Bavaria (Rural Germany) and the Coal Basin of the Czech Republic	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Cardiovascular (390-459); Respiratory (460-519); Cancer (140-239) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> (1) Poisson Regression Models by logistic regression analyses with a cubic function; (2) Poisson GAM, natural splines <b>Age Groups Analyzed:</b> All Ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> Coal Basin: 0.58 (0.39) mg/m <sup>3</sup> Northeast Bavaria: 0.88 (0.69) mg/m <sup>3</sup> <b>Range (Min, Max):</b> Coal Basin: (-0.1, 2.88) Northeast Bavaria: (0.1, 6.2) <b>Copollutant correlation:</b> SO <sub>2</sub> : r = 0.37; TSP: r = 0.37; NO <sub>2</sub> : r = 0.32; O <sub>3</sub> : r = -0.57; PM <sub>10</sub> : r = 0.44; PM <sub>2.5</sub> : r = 0.42	<b>Increment:</b> 1 mg/m <sup>3</sup> <b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Coal Basin of the Czech Republic</b> Total Mortality: 1.016 (0.998, 1.035); 0; / 1.016 (0.998, 1.034); 1; 1.013 (0.996, 1.030); 2; / 1.012 (0.995, 1.028); 3 <b>Northeast Bavaria</b> Total Mortality: 1.014 (0.994, 1.034); 0; / 1.023 (1.005, 1.041); 1; 1.013 (0.995, 1.031); 2; / 1.003 (0.985, 1.021); 3 Cardiovascular Disease Mortality: 1.018 (0.994, 1.044); 0; / 1.012 (0.987, 1.038); 1; 1.016 (0.991, 1.041); 2; / 1.004 (0.980, 1.029); 3
<b>Author:</b> Rainham et al. (2003) <b>Period of Study:</b> 1980-1996 <b>Location:</b> Toronto, ON, Canada	<b>Health Outcome (ICD9):</b> Mortality: Cardiac (390-459); Respiratory (480-519); Total (non-accidental) (<800) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, natural cubic splines <b>Age Groups Analyzed:</b> <65 ≥ 65	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 1.0 (0.4) ppm <b>Range (Min, Max):</b> (0.0, 4.0) <b>Copollutant:</b> O <sub>3</sub> ; NO <sub>2</sub> ; SO <sub>2</sub>	<b>The study did not present quantitative results for CO.</b>
<b>Author:</b> Roemer et al. (2001) <b>Period of Study:</b> 1/1987 - 11/1998 <b>Location:</b> Amsterdam	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> Air pollution background: 836 µg/m <sup>3</sup> Air pollution traffic: 1805 µg/m <sup>3</sup> <b>Range (10th, 90th):</b> Air pollution background: (448, 1315) µg/m <sup>3</sup> Air pollution traffic: (727, 3192) µg/m <sup>3</sup> <b>Copollutant:</b> BS; PM <sub>10</sub> ; SO <sub>2</sub> ; NO <sub>2</sub> ; NO; O <sub>3</sub>	<b>Increment:</b> Lag 1 and 2: 100 µg/m <sup>3</sup> Lag 0-6: 50 µg/m <sup>3</sup> <b>Relative Risk (Lower CI, Upper CI); lag:</b> <b>Total Population using Background sites</b> 1.002 (1.000-1.004); 1; 1.001 (0.999-1.003); 2; 1.001 (1.000-1.003); 0-6 <b>Traffic Population using Background Sites</b> 1.003 (0.997-1.008); 1; 1.008 (1.003-1.013); 2; 1.003 (0.999-1.007); 0-6 <b>Total population using Traffic Sites</b> 1.000 (1.000-1.001); 1; 1.000 (0.999-1.001); 2; 1.000 (1.000-1.001); 0-6

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Samet et al. (2000b)</p> <p><b>Period of Study:</b> 1987-1994</p> <p><b>Location:</b> 20 U.S. Cities: Los Angeles, CA; New York, NY; Chicago, IL; Dallas, TX; Houston, TX; San Diego, CA; Anaheim, CA; Phoenix, AZ; Detroit, MI; Miami, FL; Philadelphia, PA; Minneapolis, MN; Seattle, WA; San Jose, CA; Cleveland, OH; San Bernardino, CA; Pittsburgh, PA; Oakland, CA; Atlanta, GA; San Antonio, TX</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Cardiovascular (390-459); Respiratory (460-519); Other (non-accidental) (&lt;800)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Two-stage log linear regression model</p> <p><b>Age Groups Analyzed:</b> &lt;65 65-74 ≥ 75</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> Los Angeles: 15.1 ppm New York: 20.4 ppm Chicago: 7.9 ppm Dallas: 7.4 ppm Houston: 8.9 ppm San Diego: 11.0 ppm Anaheim: 12.3 ppm Phoenix: 12.6 ppm Detroit: 6.6 ppm Miami: 10.6 ppm Philadelphia: 11.8 ppm Minneapolis: 11.8 ppm Seattle: 17.8 ppm San Jose: 9.4 ppm Cleveland: 8.5 ppm San Bernardino: 10.3 ppm Pittsburgh: 12.2 ppm Oakland: 9.1 ppm Atlanta: 8.0 ppm San Antonio: 10.1 ppm</p> <p><b>Range (10th, 90th):</b> Los Angeles: (5.9, 28.3) New York: (14.8, 27.6) Chicago: (4.5, 11.9) Dallas: (3.6, 12.0) Houston: (4.0, 14.2) San Diego: (4.5, 20.5) Anaheim: (3.7, 25.2) Phoenix: (5.4, 22.6) Detroit: (3.2, 11.1) Miami: (6.5, 15.9) Philadelphia: (7.0, 17.2) Minneapolis: (7.0, 17.0) Seattle: (10.5, 26.4) San Jose: (1.7, 21.3) Cleveland: (3.7, 13.8) San Bernardino: (4.0, 17.5) Pittsburgh: (6.1, 19.8) Oakland: (2.9, 17.0) Atlanta: (3.2, 14.3) San Antonio: (4.1, 17.3)</p> <p><b>Copollutant correlation:</b> PM<sub>10</sub>: r = 0.45; O<sub>3</sub>: r = -0.19; NO<sub>2</sub>: r = 0.64; SO<sub>2</sub>: r = 0.41</p>	<p>This study did not provide quantitative results for CO.</p>
<p><b>Author:</b> Samoli et al. (2007)</p> <p><b>Period of Study:</b> 1990-1997</p> <p><b>Location:</b> 19 European Cities (APHEA2)</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (&lt;800); Cardiovascular (390-459)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson and two-stage hierarchical model</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean Range (unit - mg/m3):</b> Athens: 6.1; Barcelona: 0.9; Basel: 0.6; Birmingham: 1.0; Budapest: 5.1; Geneva: 1.5; Helsinki: 1.2; Ljubljana: 1.6; London: 1.4; Lyon: 3.8; Milano: 5.4; Netherlands: 0.6; Prague: 0.9; Rome: 4.1; Stockholm: 0.8; Teplice: 0.7; Torino: 5.5; Valencia: 4.1; Zurich: 1.2</p> <p><b>Range (10th, 90th):</b> Athens: (3.5, 9.2) Barcelona: (0.4, 1.7) Basel: (0.4, 1.1) Birmingham: (0.5, 1.6) Budapest: (3.3, 7.4) Geneva: (0.8, 2.6) Helsinki: (0.7, 1.9) Ljubljana: (0.6, 3.0) London: (0.7, 2.2) Lyon: (2.0, 6.0) Milano: (2.9, 8.7) Netherlands: (0.4, 1.2) Prague: (0.5, 1.5) Rome: (2.5, 5.9) Stockholm: (0.5, 1.2)</p>	<p><b>Increment:</b> 1 mg/m<sup>3</sup></p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Non-accidental mortality</b></p> <p><b>8 Degrees of Freedom per year</b></p> <p><b>Fixed Effects:</b> CO: 0.59% (0.41-0.78); 0-1 CO, BS: 0.35% (-0.03 to 0.72); 0-1 CO, PM<sub>10</sub>: 0.48% (0.24-0.72); 0-1 CO, SO<sub>2</sub>: 0.44% (0.21-0.67); 0-1 CO, O<sub>3</sub>: 0.66% (0.46-0.86); 0-1 CO, NO<sub>2</sub>: 0.27% (0.03-0.51); 0-1</p> <p><b>Random Effects:</b> CO: 0.66% (0.27-1.05); 0-1 CO, BS: 0.45% (-0.01 to 0.92); 0-1 CO, PM<sub>10</sub>: 0.58% (0.12-1.04); 0-1 CO, SO<sub>2</sub>: 0.46% (0.07-0.85); 0-1 CO, O<sub>3</sub>: 0.76% (0.45-1.06); 0-1 CO, NO<sub>2</sub>: 0.30% (-0.11 to 0.71); 0-1</p> <p><b>PACF: (Partial Autocorrelation Function) Plot Fixed Effects:</b> CO: 1.00% (0.83-1.18); 0-1 CO, BS: 0.67% (0.30-1.04); 0-1 CO, PM<sub>10</sub>: 0.78% (0.55-1.00); 0-1 CO, SO<sub>2</sub>: 0.68% (0.47-0.90); 0-1 CO, O<sub>3</sub>: 1.12% (0.93-1.31); 0-1 CO, NO<sub>2</sub>: 0.72% (0.50-0.95); 0-1</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
		Teplice: (0.3, 1.2) Torino: (2.8, 9.1) Valencia: (2.4, 5.9) Zurich: (0.7, 2.0) <b>Copollutant correlation:</b> PM <sub>10</sub> : r = 0.16 to 0.70 BS: r = 0.67 to 0.82 SO <sub>2</sub> : r = 0.35 to 0.82 NO <sub>2</sub> : r = 0.03 to 0.68 O <sub>3</sub> : r = -0.25 to -0.65	<b>Random Effects:</b> CO: 1.20% (0.63-1.77); 0-1 CO, BS: 0.77% (0.28-1.26); 0-1 CO, PM <sub>10</sub> : 1.09% (0.36-1.83); 0-1 CO, SO <sub>2</sub> : 0.75% (0.26-1.26); 0-1 CO, O <sub>3</sub> : 1.37% (0.81-1.95); 0-1 CO, NO <sub>2</sub> : 0.88% (0.22-1.55); 0-1 <b>Cardiovascular Mortality</b> <b>8 Degrees of Freedom per Year</b> <b>Fixed Effects:</b> CO: 0.80% (0.53-1.07); 0-1 CO, BS: 0.49% (-0.04 to 1.02); 0-1 CO, PM <sub>10</sub> : 0.73% (0.39-1.07); 0-1 CO, SO <sub>2</sub> : 0.72% (0.39-1.04); 0-1 CO, O <sub>3</sub> : 0.91% (0.62-1.20); 0-1 CO, NO <sub>2</sub> : 0.44% (0.10-0.79); 0-1 <b>Random Effects:</b> CO: 0.81% (0.36-1.26); 0-1 CO, BS: 0.49% (-0.04 to 1.02); 0-1 CO, PM <sub>10</sub> : 0.73% (0.39-1.07); 0-1 CO, SO <sub>2</sub> : 0.68% (-0.03 to 1.40); 0-1 CO, O <sub>3</sub> : 1.02% (0.58-1.46); 0-1 CO, NO <sub>2</sub> : 0.43% (-0.06 to 0.93); 0-1 <b>PACF (Partial Autocorrelation Function) Fixed</b> <b>Effects:</b> CO: 1.06% (0.80-1.32); 0-1 CO, BS: 0.83% (0.31-1.35); 0-1 CO, PM <sub>10</sub> : 0.95% (0.62-1.27); 0-1 CO, SO <sub>2</sub> : 0.91% (0.59-1.22); 0-1 CO, O <sub>3</sub> : 1.28% (1.01-1.56); 0-1 CO, NO <sub>2</sub> : 0.68% (0.35-1.00); 0-1 <b>Random Effects:</b> CO: 1.25% (0.30-2.21); 0-1 CO, BS: 0.83% (0.31-1.35); 0-1 CO, PM <sub>10</sub> : 1.13% (0.60-1.67); 0-1 CO, SO <sub>2</sub> : 0.86% (0.06-1.66); 0-1 CO, O <sub>3</sub> : 1.62% (0.72-2.52); 0-1 CO, NO <sub>2</sub> : 0.84% (-0.03 to 1.71); 0-1 <b>Effect Modifiers</b> <b>Non-accidental Mortality</b> <b>8 Degrees of Freedom per Year</b> <b>Number of CO monitors:</b> 25th Percentile: 0.71% (0.48-0.94); 0-1 75th Percentile: 0.54% (0.34-0.74); 0-1 <b>Mean PM<sub>10</sub> Levels:</b> 25th Percentile: 0.37% (0.08-0.66); 0-1 75th Percentile: 0.49% (0.28-0.69); 0-1 <b>Standardized Mortality Rate:</b> 25th Percentile: 0.79% (0.55-1.03); 0-1 75th Percentile: 0.44% (0.22-0.66); 0-1 Western cities: 0.75% (0.47-1.03); 0-1 Southern cities: 0.61% (0.32-0.91); 0-1 Eastern cities: 0.03% (-0.47 to 0.53); 0-1 <b>PACF (Partial Autocorrelation Function)</b> <b>Number of CO monitors:</b> 25th Percentile: 1.18% (0.96-1.39); 0-1 75th Percentile: 0.92% (0.73-1.11); 0-1 <b>Mean PM<sub>10</sub> Levels:</b> 25th Percentile: 0.74% (0.46-1.02); 0-1 75th Percentile: 1.07% (0.87-1.27); 0-1 <b>Standardized Mortality Rate:</b> 25th Percentile: 1.29% (1.06-1.52); 0-1 75th Percentile: 0.77% (0.56-0.98); 0-1 Western cities: 1.15% (0.90-1.40); 0-1 Southern cities: 1.08% (0.79-1.38); 0-1 Eastern cities: 0.27% (-0.20 to 0.74); 0-1 <b>Cardiovascular Mortality</b> <b>8 Degrees of Freedom per Year</b> <b>Mean O<sub>3</sub>:</b> 25th Percentile: 1.04% (0.67-1.41); 0-1 75th Percentile: 0.82% (0.55-1.10); 0-1 <b>Standardized Mortality Rate:</b> 25th Percentile: 1.06% (0.71-1.42); 0-1 75th Percentile: 0.61% (0.30-0.93); 0-1 <b>Population &gt;75 years of age (%):</b> 25th Percentile: 0.58% (0.25-0.92); 0-1 75th Percentile: 0.94% (0.64-1.24); 0-1

Study	Design	Concentrations	Effect Estimates (95% CI)
			Western cities: 1.06% (0.67-1.46); 0-1 Southern cities: 0.70% (0.26-1.14); 0-1 Eastern cities: 0.21% (-0.48 to 0.90); 0-1 <b>PACF (Partial Autocorrelation Function)</b> <b>Mean O<sub>3</sub>:</b> 25th Percentile: 1.32% (0.96-1.68); 0-1 75th Percentile: 1.09% (0.83-1.14); 0-1 <b>Standardized Mortality Rate:</b> 25th Percentile: 1.40% (1.06-1.75); 0-1 75th Percentile: 0.85% (0.55-1.14); 0-1 <b>Population &gt;75 years of age (%):</b> 25th Percentile: 0.74% (0.41-1.06); 0-1 75th Percentile: 1.25% (0.96-1.54); 0-1 Western cities: 1.38% (1.00-1.76); 0-1 Southern cities: 0.90% (0.47-1.33); 0-1 Eastern cities: 0.48% (-0.14 to 1.11); 0-1
<b>Author:</b> Schwartz et al. (1999) <b>Period of Study:</b> 1989-1995 <b>Location:</b> Spokane, WA	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 1-h avg <b>Mean (SD) unit:</b> <b>Dust Storm Days:</b> 09/08/1990: 6.37 ppm 09/12/1990: 3.40 ppm 10/04/1990: 3.15 ppm 11/09/1990: 2.45 ppm 11/23/1990: 2.50 ppm 09/13/1991: 4.60 ppm 10/16/1991: 2.10 ppm 10/21/1991: 2.20 ppm 09/04/1992: 3.43 ppm 09/12/1992: 1.80 ppm 09/13/1992: 1.65 ppm 09/25/1992: 2.95 ppm 09/26/1992: 4.30 ppm 10/08/1992: 3.85 ppm 09/11/1993: 1.88 ppm 11/3/1993: 5.33 ppm 07/24/1994: 2.10 ppm 08/30/1996: 2.85 ppm <b>Range (Min, Max):</b> NR <b>Copollutant:</b> PM <sub>10</sub>	<b>The study did not present quantitative results for CO.</b>
<b>Author:</b> Sharovsky et al. (2004) <b>Period of Study:</b> 1996-1998 <b>Location:</b> Sao Paulo, Brazil	<b>Health Outcome (ICD10):</b> Mortality: Myocardial Infarction (I.21) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, LOESS <b>Age Groups Analyzed:</b> 35-109	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 3.7 (1.6) ppm <b>Range (Min, Max):</b> (1.0, 11.8) <b>Copollutant: correlation</b> SO <sub>2</sub> : r = 0.73; PM <sub>10</sub> : r = 0.51	<b>Increment:</b> NR <b>β x 100 (SE); lag:</b> CO: 1.42 (1.01) CO, SO <sub>2</sub> , PM <sub>10</sub> : 0.97 (1.27) <b>Notes:</b> The study did not present the lag used for CO.
<b>Author:</b> Slaughter et al. (2005) <b>Period of Study:</b> 1/1995 - 6/2001 <b>Location:</b> Spokane, WA	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Respiratory (460-519); Asthma (493); COPD (491, 492, 494, 496); Pneumonia (480-487); Acute Upper Respiratory Tract Infections (464-466, 490); Cardiac Outcomes (390-459) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Log-linear Poisson GLM, natural splines for calendar time <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> Areas in Spokane Hamilton St: 1.73 (0.46) ppm Backdoor Tavern: 1.29 (0.23) ppm Spokane Club: 1.41 (0.32) ppm Third and Washington: 1.82 (0.33) ppm Rockwood: 0.42 (0.15) ppm <b>Range (Min, Max):</b> NR <b>Copollutant correlation :</b> PM <sub>1</sub> : r = 0.63; PM <sub>2.5</sub> : r = 0.62; PM <sub>10</sub> : r = 0.32; PM <sub>10-2.5</sub> : r = 0.32	<b>The study did not present quantitative results for CO.</b>



Study	Design	Concentrations	Effect Estimates (95% CI)
<b>Author:</b> Stieb et al. (2003) <b>Period of Study:</b> 1985-2000 <b>Location:</b> All locations	<b>Health Outcome (ICD9):</b> Mortality: Non-accidental <b>Study Design:</b> Meta-analysis <b>Statistical Analyses:</b> NR <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> NR <b>IQR (25th, 75th):</b> NR <b>Copollutant:</b> NR	<b>Increment:</b> 1.1 ppm <b>% Excess Mortality (Lower CI, Upper CI); lag:</b> <b>Non-GAM:</b> Single-pollutant model (4 studies): 4.7% (1.1-8.4) Multi-pollutant model (1 study): 0.0% (-3.8 to 3.8) <b>GAM:</b> Single-pollutant model (18 studies): 1.6% (1.1-2.1) Multi-pollutant model (11 studies): 0.7% (-0.1 to 1.5)
<b>Author:</b> Stölzel et al. (2006) <b>Period of Study:</b> 9/1995-8/2001 <b>Location:</b> Erfurt, Germany	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Cardio-respiratory (390-459, 460-519, 785, 786) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.47 (0.39) mg/m <sup>3</sup> <b>IQR (25th, 75th):</b> (0.23, 0.57) <b>Copollutant correlation:</b> MC <sub>0.1-0.5</sub> : r = 0.58; MC <sub>0.01-2.5</sub> : r = 0.57; PM <sub>10</sub> : r = 0.50; NO: r = 0.70; NO <sub>2</sub> : r = 0.71	<b>Increment:</b> 0.34 mg/m <sup>3</sup> <b>Relative Risk (Lower CI, Upper CI); lag:</b> Total (non-accidental) 1.000 (0.977-1.023); 0; 1.002 (0.980-1.024); 1; 1.013 (0.991-1.035); 2; 1.007 (0.986-1.029); 3; 1.012 (0.990-1.034); 4; 0.995 (0.974-1.017); 5
<b>Author:</b> Sunyer et al. (2001) <b>Period of Study:</b> 1990-1995 <b>Location:</b> Barcelona, Spain	<b>Health Outcome (ICD9):</b> Mortality: COPD (491, 492, 494, 496) <b>Study Design:</b> Bi-directional case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> >35	<b>Pollutant:</b> CO <b>Averaging Time:</b> 8-h avg <b>Mean (SD) unit:</b> NR <b>Range (Min, Max):</b> NR <b>Copollutant:</b> PM <sub>10</sub> ; NO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 4.5 µg/m <sup>3</sup> <b>Odds Ratio (Lower CI, Upper CI); lag:</b> CO: 1.052 (0.990-1.117); 0-2 CO, PM <sub>10</sub> : 1.017 (0.947-1.091); 0-2
<b>Author:</b> Sunyer et al. (2002) <b>Period of Study:</b> 1985-1995 <b>Location:</b> Barcelona, Spain	<b>Health Outcome (ICD9):</b> Mortality: Respiratory mortality <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> >14 <b>Study population:</b> Asthmatic individuals: 5,610	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Median (SD) unit:</b> 7.7 µg/m <sup>3</sup> <b>Range (Min, Max):</b> (0.6, 66.0) <b>Copollutant:</b> PM <sub>10</sub> ; BS; NO <sub>2</sub> ; O <sub>3</sub> ; SO <sub>2</sub>	<b>Increment:</b> 7.2 µg/m <sup>3</sup> <b>Odds Ratio (Lower CI, Upper CI); lag:</b> Asthmatic individuals with 1 ED visit 1.127 (0.895-1.418); 0-2 Asthmatic individuals with >1 ED visit 1.125 (0.773-1.638); 0-2 Asthma/COPD individuals with >1 ED visit 0.815 (0.614-1.082); 0-2
<b>Author:</b> Tsai et al. (2003a) <b>Period of Study:</b> 1994-2000 <b>Location:</b> Kaohsiung, Taiwan	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Respiratory (460-519); Circulatory (390-459) <b>Study Design:</b> Bidirectional case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.827 ppm <b>Range (Min, Max):</b> (0.226, 1.770) <b>Copollutant:</b> PM <sub>10</sub> ; SO <sub>2</sub> ; NO <sub>2</sub> ; O <sub>3</sub>	<b>Increment:</b> 0.313 ppm <b>Odds Ratio (Lower CI, Upper CI); lag:</b> Total (non-accidental): 1.003 (0.968-1.039); 0-2 Respiratory: 1.011 (0.883-1.159); 0-2 Circulatory: 0.986 (0.914-1.063); 0-2
<b>Author:</b> Tsai et al. (2006b) <b>Period of Study:</b> 1994-2000 <b>Location:</b> Kaohsiung, Taiwan	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800) <b>Study Design:</b> Case-crossover <b>Statistical Analyses:</b> Conditional logistic regression <b>Age Groups Analyzed:</b> 27 days old to <1 yr of age	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 8.27 ppm <b>Range (Min, Max):</b> (2.26, 17.70) <b>Copollutant:</b> PM <sub>10</sub> ; SO <sub>2</sub> ; O <sub>3</sub> ; NO <sub>2</sub>	<b>Increment:</b> 0.31 ppm <b>Odds Ratio (Lower CI, Upper CI); lag:</b> Postneonatal Mortality 1.051 (0.304-3.630); 0-2
<b>Author:</b> Vedal et al. (2003) <b>Period of Study:</b> 1/1994-12/1996 <b>Location:</b> Vancouver, BC, Canada	<b>Health Outcome (ICD9):</b> Mortality: Total (non-accidental) (<800); Respiratory (460-519); Cardiovascular (390-459) <b>Study Design:</b> Time-series <b>Statistical Analyses:</b> Poisson GAM, LOESS <b>Age Groups Analyzed:</b> All ages	<b>Pollutant:</b> CO <b>Averaging Time:</b> 24-h avg <b>Mean (SD) unit:</b> 0.6 (0.2) ppm <b>Range (Min, Max):</b> (0.3, 1.9) <b>Copollutant correlation:</b> <b>Summer:</b> PM <sub>10</sub> : r = 0.71; O <sub>3</sub> : r = 0.12; NO <sub>2</sub> : r = 0.81; SO <sub>2</sub> : r = 0.67 <b>Winter:</b> PM <sub>10</sub> : r = 0.76; O <sub>3</sub> : r = -0.65; NO <sub>2</sub> : r = 0.78; SO <sub>2</sub> : r = 0.83	<b>The study did not present quantitative results for CO.</b>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Villeneuve et al. (2003)</p> <p><b>Period of Study:</b> 1986-1999</p> <p><b>Location:</b> Vancouver, BC, Canada</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Non-accidental (&lt;800); Cardiovascular (401-440); Respiratory (460-519); Cancer (140-239)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson, natural splines</p> <p><b>Age Groups Analyzed:</b> ≥ 65</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.0 ppm</p> <p><b>Range (Min, Max):</b> (0.2, 4.9)</p> <p><b>Copollutant:</b> PM<sub>2.5</sub>; PM<sub>10</sub>; PM<sub>10-2.5</sub>; TSP; SO<sub>4</sub>; CO; COH; O<sub>3</sub>; NO<sub>2</sub>; SO<sub>2</sub></p>	<p><b>Increment:</b> 1.1 ppb</p> <p><b>% Increase (Lower CI, Upper CI); lag:</b></p> <p><b>Non-accidental</b> 0.5% (-1.9 to 2.9); 0-2; / -0.3% (-2.2 to 1.7); 0; 0.6% (-1.3 to 2.6); 1; / 0.5% (-1.4 to 2.5); 2</p> <p><b>Cardiovascular</b> 2.3% (-1.6 to 6.3); 0-2; / 1.6% (-1.5 to 4.7); 0; 1.2% (-2.0 to 4.5); 1; / 1.5% (-1.5 to 4.4); 2</p> <p><b>Respiratory</b> -1.0% (-7.3 to 5.8); 0-2; / 1.3% (-4.4 to 7.3); 0; -0.1% (-5.3 to 5.4); 1; / -2.8% (-7.8 to 2.6); 2</p> <p><b>Cancer</b> -2.8% (-7.6 to 2.4); 0-2; / -3.0% (-6.9 to 1.1); 0; -1.6% (-5.6 to 2.4); 1; / -0.5% (-4.7 to 3.8); 2</p>
<p><b>Author:</b> Wichmann et al. (2000)</p> <p><b>Period of Study:</b> 9/1995 - 12/1998</p> <p><b>Location:</b> Erfurt, Germany</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Non-accidental (&lt;800); Cardiovascular (401-440); Respiratory (460-519)</p> <p><b>Study Design:</b> Time-series</p> <p><b>Statistical Analyses:</b> Poisson GAM, LOESS</p> <p><b>Age Groups Analyzed:</b> &lt;70 70-79 ≥ 80</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 0.6 (0.5) mg/m<sup>3</sup></p> <p><b>Range (Min, Max):</b> (0.10, 2.50)</p> <p><b>Copollutant correlation:</b> PM<sub>2.5</sub>: r = 0.62; PM<sub>10</sub>: r = 0.58; TSP: r = 0.57; SO<sub>2</sub>: r = 0.59; NO<sub>2</sub>: r = 0.71</p>	<p><b>Increment:</b> 0.5 ppm</p> <p><b>Relative Risk (Lower CI, Upper CI); lag:</b></p> <p>Single-Day Lag CO: 1.055 (1.003-1.110); 4</p> <p><b>Polynomial Distributed Lag</b> <b>Multi-pollutant model:</b> 1.076 (1.017-1.138); 4</p> <p><b>Total Mortality</b> CO: 1.012 (0.977-1.049); 0 Log-transformed: 1.016 (0.962-1.073); 0 1.004 (0.969-1.040); 1 Log-transformed: 1.027 (0.973-1.083); 1 1.020 (0.984-1.057); 2 Log-transformed: 1.024 (0.970-1.081); 2 1.019 (0.984-1.055); 3 Log-transformed: 1.037 (0.984-1.093); 3 1.029 (0.995-1.063); 4 Log-transformed: 1.055 (1.003-1.110); 4 0.997 (0.965-1.031); 5 Log-transformed: 1.014 (0.966-1.065); 5</p> <p><b>Total Mortality (Season-specific):</b> Log-transformed Winter: 1.002 (0.922-1.088); 4 Spring: 1.019 (0.942-1.102); 4 Summer: 1.085 (1.018-1.156); 4 Fall: 1.111 (1.039-1.188); 4</p> <p><b>Winter-specific:</b> Log-transformed 10/95-3/96: 1.046 (0.949-1.153); 4 10/96-3/97: 1.091 (0.998-1.193); 4 10/97-3/98: 1.028 (0.966-1.095); 4</p> <p><b>One-pollutant Model:</b> Log-transformed CO: 1.055 (1.003-1.110); 4</p>
<p><b>Author:</b> Yang et al. (2004a)</p> <p><b>Period of Study:</b> 1994-1998</p> <p><b>Location:</b> Taipei, Taiwan</p>	<p><b>Health Outcome (ICD9):</b> Mortality: Non-accidental (&lt;800); Circulatory (390-459); Respiratory (460-519)</p> <p><b>Study Design:</b> Bi-directional case-crossover</p> <p><b>Statistical Analyses:</b> Conditional logistic regression</p> <p><b>Age Groups Analyzed:</b> All ages</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1.16 ppm</p> <p><b>Range (Min, Max):</b> (0.24, 4.42)</p> <p><b>Copollutant:</b> PM<sub>10</sub>; SO<sub>2</sub>; NO<sub>2</sub>; O<sub>3</sub></p>	<p><b>Increment:</b> 0.52 ppm</p> <p><b>Odds Ratio (Lower CI, Upper CI); lag:</b></p> <p>Non-accidental: 1.005 (0.980-1.031); 0-2</p> <p>Respiratory: 1.014 (0.925-1.110); 0-2</p> <p>Circulatory: 0.996 (0.948-1.046); 0-2</p>

**Table C-8. Studies of long-term CO exposure and mortality.**

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Lipfert et al. (2000b)  <b>Period of Study:</b> 1975-1996  <b>Location:</b> 32 Veterans Hospitals, USA</p>	<p><b>Mortality</b>  <b>Health Outcome (ICD9):</b> Non-accidental  <b>Study Design:</b> Cohort  <b>Study Population:</b> ~90,000 hypertensive male U.S. veterans  <b>Statistical Analyses:</b> Staged regression  <b>Age Groups Analyzed:</b> NR</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> 95th Percentile Annual avg  <b>Mean (SD) unit:</b>            1960-1974: 10.82 (5.15) ppm            1975-1981: 7.64 (2.94) ppm            1982-1988: 3.42 (0.95) ppm            1989-1996: 2.36 (0.67) ppm  <b>Range (Min, Max):</b>            1960-1974: (0.94, 35.30)            1975-1981: (0.43, 22.38)            1982-1988: (0.30, 15.20)            1989-1996: (0.30, 7.10)  <b>Copollutants; correlation:</b>  <b>1960-1974:</b>            O<sub>3</sub>: r = 0.004;            NO<sub>2</sub>: r = 0.690;            SO<sub>4</sub><sup>2-</sup>: r = 0.469  <b>1975-1981:</b>            O<sub>3</sub>: r = 0.109;            NO<sub>2</sub>: r = 0.249;            SO<sub>4</sub><sup>2-</sup>: r = -0.155;            IP SO<sub>4</sub><sup>2-</sup>: r = 0.356;            PM<sub>2.5</sub>: r = 0.634;            PM<sub>10-2.5</sub>: r = 0.498;            PM<sub>15</sub>: r = 0.626  <b>1982-1988</b>            O<sub>3</sub>: r = 0.158; NO<sub>2</sub>: r = 0.413; SO<sub>4</sub><sup>2-</sup>: r = -0.518;            IP SO<sub>4</sub><sup>2-</sup>: r = 0.075;            PM<sub>2.5</sub>: r = 0.296;            PM<sub>10-2.5</sub>: r = 0.135            PM<sub>15</sub>: r = 0.284  <b>1989-1996</b>            O<sub>3</sub>: r = 0.397;            NO<sub>2</sub>: r = 0.492;            SO<sub>4</sub><sup>2-</sup>: r = -0.551</p>	<p><b>Increment:</b> NR  <b>Coefficient:</b>  <b>Baseline Model</b>  <b>Exposure Period: up to 1975</b>            Single Period: -0.000            Deaths, 1976-81: 0.0043            Deaths, 1982-88: -0.0002            Deaths after 1988: -0.0041  <b>Exposure Period: 1975-81</b>            Single Period: -0.013            Deaths, 1976-81: -0.0170            Deaths, 1982-88: -0.0217            Deaths after 1988: -0.0240  <b>Exposure Period: 1982-88</b>            Single Period: -0.028            Deaths, 1976-81: -0.0294            Deaths, 1982-88: -0.0484            Deaths after 1988: -0.0424  <b>Exposure Period: 1989-96</b>            Single Period: -0.046            Deaths, 1976-81: -0.0590            Deaths, 1982-88: -0.0581            Deaths after 1988: -0.0536  <b>Final Model w/ Ecological Variables</b>  <b>Exposure Period: up to 1975</b>            Single Period: -0.001            Deaths, 1976-81: 0.0013            Deaths, 1982-88: -0.0022            Deaths after 1988: -0.0061  <b>Exposure Period: 1975-81</b>            Single Period: -0.008            Deaths, 1976-81: -0.0128            Deaths, 1982-88: -0.0186            Deaths after 1988: -0.0203  <b>Exposure Period: 1982-88</b>            Single Period: -0.009            Deaths, 1976-81: -0.0007            Deaths, 1982-88: -0.0246            Deaths after 1988: -0.0216  <b>Exposure Period: 1989-96</b>            Single Period: -0.009            Deaths, 1976-81: -0.0106            Deaths, 1982-88: -0.0136            Deaths after 1988: -0.0078  <b>Notes:</b> Mortality risks based on mean concentrations of pollutants less estimated background weighted by the number of subjects in each county, but The study did not present this value for each pollutant.</p>
<p><b>Author:</b> Lipfert and Morris (2002)  <b>Period of Study:</b> 1960-1997  <b>Location:</b> U.S. counties</p>	<p><b>Mortality</b>  <b>Health Outcome (ICD9):</b> Non-accidental  <b>Study Design:</b> Ecological/ cross-sectional  <b>Statistical Analyses:</b> Staged regression  <b>Age Groups Analyzed:</b>            15-44            45-64            65-74            75-84            ≥ 85</p>	<p><b>Pollutant:</b> CO  <b>Averaging Time:</b> Annual avg  <b>Mean (SD) unit:</b>            1960-1969: 13.81 (8.47) ppm            1970-1974: 9.64 (5.63) ppm            1979-1981: 5.90 (3.54) ppm            1989-1991: 2.69 (1.22) ppm            1995-1997: 1.72 (0.76) ppm  <b>Range (Min, Max):</b> NR  <b>Copollutant:</b>            TSP            SO<sub>4</sub><sup>2-</sup>            SO<sub>2</sub>            NO<sub>2</sub>            O<sub>3</sub></p>	<p><b>Increment:</b> NR  <b>Attributable risk (SE):</b>  <b>Attributable Risks of mortality (1960-4)</b>  <b>Peak CO 1960-1964, All locations</b>            Ages 15-44: 0.1299 (0.0341)            Ages 45-64: 0.0340 (0.0280)            Ages 65-74: -0.0058 (0.0220)            Ages 75-84: 0.0121 (0.0188)            Ages ≥ 85: 0.0374 (0.0225)            Log Mean: 0.0365 (0.0149)  <b>Attributable Risks of mortality (1970-4)</b>  <b>Peak CO 1970-1974, All locations</b>            Ages 15-44: 0.0553 (0.0240)            Ages 45-64: 0.0181 (0.0148)            Ages 65-74: -0.0146 (0.0134)            Ages 75-84: -0.0128 (0.0098)            Ages ≥ 85: -0.0151 (0.0093)            Log Mean: 0.0038 (0.0086)</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
			<b>Attributable Risks of mortality (1979-81)</b>
			<b>Peak CO 1979-1981, All locations</b>
			Ages 15-44: 0.0054 (0.0174)
			Ages 45-64: -0.0060 (0.0141)
			Ages 65-74: -0.0251 (0.0105)
			Ages 75-84: -0.0331 (0.0086)
			Ages ≥ 85: -0.0123 (0.0079)
			Log Mean: -0.0183 (0.0077)
			<b>Peak CO 1970-1974, All locations</b>
			Ages 15-44: 0.0218 (0.0200)
			Ages 45-64: 0.0327 (0.0161)
			Ages 65-74: -0.0136 (0.0119)
			Ages 75-84: -0.0250 (0.0105)
			Ages ≥ 85: -0.0202 (0.0085)
			Log Mean: -0.0048 (0.0077)
			<b>Peak CO 1960-1969, All locations</b>
			Ages 15-44: 0.0506 (0.0478)
			Ages 45-64: 0.0704 (0.0337)
			Ages 65-74: 0.0100 (0.0211)
			Ages 75-84: -0.0124 (0.0143)
			Ages ≥ 85: 0.0187 (0.0135)
			Log Mean: 0.0084 (0.0149)
			<b>Peak CO 1979-1981, CO 1970-1974</b>
			Ages 15-44: 0.0244 (0.0209)
			Ages 45-64: 0.0016 (0.0181)
			Ages 65-74: -0.0183 (0.0128)
			Ages 75-84: -0.0382 (0.0108)
			Ages ≥ 85: -0.0201 (0.0089)
			Log Mean: -0.0165 (0.0089)
			<b>Peak CO 1979-1981, CO 1960-1969</b>
			Ages 15-44: 0.0748 (0.0679)
			Ages 45-64: 0.0844 (0.0496)
			Ages 65-74: 0.0144 (0.0259)
			Ages 75-84: -0.0158 (0.0168)
			Ages ≥ 85: -0.0073 (0.0170)
			Log Mean: 0.0109 (0.0218)
			<b>Peak CO 1979-1981, CO 1960-1969</b>
			Ages 15-44: 0.1191 (0.0709)
			Ages 45-64: 0.1163 (0.0491)
			Ages 65-74: 0.0177 (0.0310)
			Ages 75-84: -0.0120 (0.0212)
			Ages ≥ 85: -0.0040 (0.0202)
			Log Mean: 0.0211 (0.0231)
			<b>Attributable Risks of mortality (1989-91)</b>
			<b>Peak CO 1989-1991, All locations</b>
			Ages 15-44: 0.0404 (0.0322)
			Ages 45-64: -0.0262 (0.0162)
			Ages 65-74: -0.0397 (0.0115)
			Ages 75-84: -0.0464 (0.0097)
			Ages ≥ 85: -0.0209 (0.0073)
			Log Mean: -0.0178 (0.0098)
			<b>Peak CO 1979-1981, All locations</b>
			Ages 15-44: 0.0522 (0.0227)
			Ages 45-64: -0.0047 (0.0121)
			Ages 65-74: -0.0165 (0.0078)
			Ages 75-84: -0.0268 (0.0068)
			Ages ≥ 85: -0.0027 (0.0055)
			Log Mean: -0.0020 (0.0065)
			<b>Peak CO 1970-1974, All locations</b>
			Ages 15-44: 0.0685 (0.0274)
			Ages 45-64: 0.0022 (0.0148)
			Ages 65-74: -0.0051 (0.0091)
			Ages 75-84: -0.0158 (0.0079)
			Ages ≥ 85: -0.0069 (0.0060)
			Log Mean: 0.0038 (0.0077)
			<b>Peak CO 1960-1969, All locations</b>
			Ages 15-44: 0.0578 (0.0713)
			Ages 45-64: 0.0583 (0.0347)
			Ages 65-74: 0.0007 (0.0174)
			Ages 75-84: -0.0245 (0.0130)
			Ages ≥ 85: -0.0138 (0.0113)
			Log Mean: 0.0041 (0.0176)
			<b>Attributable Risks of mortality (1995-97)</b>
			<b>Peak CO 1995-1997, All locations</b>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Lipfert et al. (2006b)</p> <p><b>Period of Study:</b> 1976-2001</p> <p><b>Location:</b> 32 Veterans Hospitals, USA</p>	<p><b>Mortality</b></p> <p><b>Health Outcome (ICD9):</b> Non-accidental</p> <p><b>Study Design:</b> Cohort</p> <p><b>Study Population:</b> ~70,000 hypertensive male U.S. veterans</p> <p><b>Statistical Analyses:</b> Cox proportional-hazards model</p> <p><b>Age Groups Analyzed:</b> NR</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 95th Percentile Annual avg</p> <p><b>Mean (SD) unit:</b>  1976-1981: 7.6 (2.9) ppm  1982-1988: 3.4 (9.5) ppm  1989-1996: 2.4 (0.67) ppm  1997-2001: 1.6 (5.6) ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutants correlation:</b>  ln(VKTA): r = -0.06  Avg NO<sub>2</sub>: r = 0.43  Peak O<sub>3</sub>: r = 0.08  Peak SO<sub>2</sub>: r = -0.05  PM<sub>2.5</sub>: r = 0.08  SO<sub>4</sub><sup>2-</sup>: r = -0.16</p> <p><b>Note:</b> VKTA=annual vehicle-km traveled/km<sup>2</sup></p>	<p>Ages 15-44: 0.0344 (0.0256)  Ages 45-64: -0.0203 (0.0198)  Ages 65-74: -0.0346 (0.0146)  Ages 75-84: -0.0378 (0.0161)  Ages ≥ 85: -0.0283 (0.0119)  Log Mean: -0.0188 (0.0103)</p> <p><b>Peak CO 1989-1991, All locations</b>  Ages 15-44: 0.0289 (0.0248)  Ages 45-64: -0.0192 (0.0192)  Ages 65-74: -0.0466 (0.0140)  Ages 75-84: -0.0497 (0.0147)  Ages ≥ 85: -0.0301 (0.0108)  Log Mean: -0.0240 (0.0096)</p> <p><b>Peak CO 1979-1981, All locations</b>  Ages 15-44: 0.0336 (0.0176)  Ages 45-64: -0.0037 (0.0135)  Ages 65-74: -0.0298 (0.0096)  Ages 75-84: -0.0301 (0.0105)  Ages ≥ 85: -0.0087 (0.0078)  Log Mean: -0.0094 (0.0071)</p> <p><b>Peak CO 1970-1974, All locations</b>  Ages 15-44: 0.0464 (0.0202)  Ages 45-64: 0.0202 (0.0155)  Ages 65-74: -0.0032 (0.0112)  Ages 75-84: -0.0157 (0.0122)  Ages ≥ 85: -0.0142 (0.0084)  Log Mean: 0.0007 (0.0077)</p> <p><b>Peak CO 1960-1969, All locations</b>  Ages 15-44: 0.0679 (0.0441)  Ages 45-64: 0.0772 (0.0405)  Ages 65-74: 0.0059 (0.0173)  Ages 75-84: -0.0085 (0.0213)  Ages ≥ 85: -0.0158 (0.0162)  Log Mean: 0.0162 (0.0149)</p>
			<p><b>Increment:</b> 2 ppm</p> <p><b>Relative risk (Lower CI, Upper CI):</b>  CO: 1.032 (0.954-1.117)  CO, lnVKTA: 0.999 (0.923-1.081)  CO, lnVKTA, NO<sub>2</sub>: 1.012 (0.923-1.110)  CO, lnVKTA, NO<sub>2</sub>+O<sub>3</sub>: 1.023 (0.939-1.115)</p>

Study	Design	Concentrations	Effect Estimates (95% CI)
<p><b>Author:</b> Lipfert et al. (2006a)</p> <p><b>Period of Study:</b> 1997-2002</p> <p><b>Location:</b> 32 Veterans Hospitals, USA</p>	<p><b>Mortality</b></p> <p><b>Health Outcome (ICD9):</b> Non-accidental</p> <p><b>Study Design:</b> Cohort</p> <p><b>Study Population:</b> ~18,000 hypertensive male U.S. veterans</p> <p><b>Statistical Analyses:</b> Cox proportional-hazards model</p> <p><b>Age Groups Analyzed:</b> NR</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 95th Percentile Annual avg</p> <p><b>Mean (SD) unit:</b> 1999-2001: 1.63 (0.84) ppm 1999-2001 (STN sites only): 1.73 (0.77)</p> <p><b>Range (Min, Max):</b> 1999-2001: (0.40, 6.7) 1999-2001 (STN sites only): (0.47, 4.2)</p> <p><b>Copollutants correlation:</b> ln(traffic density): r = -0.199 PM<sub>2.5</sub>: r = 0.040; As: r = 0.148 Cr: r = 0.448; Cu: r = 0.177 Fe: r = -0.138; Pb: r = 0.420 Mn: r = 0.357; Ni: r = 0.090 Se: r = -0.110; V: r = 0.230 Zn: r = 0.472; OC: r = 0.470 EC: r = 0.234; SO<sub>4</sub><sup>2-</sup>: r = -0.123 NO<sub>3</sub><sup>-</sup>: r = -0.088 ΣPM<sub>2.5</sub> comp.: r = 0.133 NO<sub>2</sub>: r = 0.418 Peak O<sub>3</sub>: r = 0.172 Peak SO<sub>2</sub>: r = 0.405</p>	<p><b>Increment:</b> NR</p> <p><b>β coefficient (SE); t-statistic:</b> -0.00000536 (0.0000324); -0.165</p>
<p><b>Author:</b> Jerrett et al. (2003)</p> <p><b>Period of Study:</b> 1982-1989</p> <p><b>Location:</b> 107 U.S. cities</p>	<p><b>Mortality</b></p> <p><b>Health Outcome (ICD9):</b> Cardiovascular; CHD; Cerebrovascular disease</p> <p><b>Study Design:</b> Cohort</p> <p><b>Study Population:</b> 65, 893 postmenopausal women without previous cardiovascular disease</p> <p><b>Statistical Analyses:</b> Cox proportional-hazards model</p> <p><b>Age Groups Analyzed:</b> ≥ 30</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Annual avg</p> <p><b>Mean (SD) unit:</b> 1.56 ppm</p> <p><b>Range (Min, Max):</b> (0.19, 3.95)</p> <p><b>Copollutants correlation:</b> Sulfates: r = -0.07 NO<sub>2</sub> O<sub>3</sub> SO<sub>2</sub></p>	<p><b>Increment:</b> 1 ppm</p> <p><b>Relative risk (Lower CI, Upper CI):</b> CO: 0.98 (0.92-1.03) CO, Sulfates: 0.97 (0.92-1.03)</p>
<p><b>Author:</b> Miller et al. (2007)</p> <p><b>Period of Study:</b> 1994-1998</p> <p><b>Location:</b> 36 U.S. cities</p>	<p><b>Mortality</b></p> <p><b>Health Outcome (ICD9):</b> Cardiovascular; CHD; Cerebrovascular disease</p> <p><b>Study Design:</b> Cohort</p> <p><b>Study Population:</b> 65, 893 postmenopausal women without previous cardiovascular disease</p> <p><b>Statistical Analyses:</b> Cox proportional-hazards model</p> <p><b>Age Groups Analyzed:</b> 50-79</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> Annual avg</p> <p><b>Mean (SD) unit:</b> NR</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutants:</b> PM<sub>2.5</sub> PM<sub>10-2.5</sub> SO<sub>2</sub> NO<sub>2</sub> O<sub>3</sub></p>	<p><b>Increment:</b> 1 ppm</p> <p><b>Hazard ratio (Lower CI, Upper CI):</b> <b>All subjects</b> CO: 1.0 (0.81-1.22) <b>Only subjects with non-missing exposure data</b> CO: 0.92 (0.71-1.21) CO, PM<sub>2.5</sub>, PM<sub>10-2.5</sub>, SO<sub>2</sub>, NO<sub>2</sub>, O<sub>3</sub>: 0.93 (0.67, 1.30)</p>
<p><b>Author:</b> Pope et al. (2002)</p> <p><b>Period of Study:</b> 1980-1998</p> <p><b>Location:</b> All 50 States, Washington DC, and Puerto Rico (ACS-CPS-II)</p>	<p><b>Mortality</b></p> <p><b>Health Outcome (ICD9):</b> Total (non-accidental) (&lt;800); Lung Cancer (162); Cardiopulmonary (401-440, 460-519)</p> <p><b>Study Design:</b> Prospective cohort</p> <p><b>Statistical Analyses:</b> Cox proportional hazards model</p> <p><b>Age Groups Analyzed:</b> ≥ 30</p>	<p><b>Pollutant:</b> CO</p> <p><b>Averaging Time:</b> 24-h avg</p> <p><b>Mean (SD) unit:</b> 1980: 1.7 (0.7) ppm 1982-1998: 1.1 (0.4) ppm</p> <p><b>Range (Min, Max):</b> NR</p> <p><b>Copollutant:</b> PM<sub>2.5</sub>; PM<sub>10</sub>; TSP; SO<sub>2</sub>; NO<sub>2</sub>; O<sub>3</sub></p>	<p><b>The study presents results for CO graphically.</b></p>

# Annex D. Controlled Human Exposure Studies

**Table D-1. Controlled human exposure studies.**

Study	Subjects	Exposure	Findings
Adir et al. (1999)	15 healthy non-smoking males; 22-34 years old	Subjects exposed to both CO and room air for 3 min 45 sec. Actual inhaled CO concentration not provided. Subjects exposed to CO concentration required to produce a venous COHb of 4-6%. Exposures were separated by 1 month with the order of exposure randomly assigned.	Exposure to CO resulted in a decrease in post-exposure exercise duration (Bruce protocol) relative to clean air exposure in 13 out of 15 subjects (p=0.0012). Statistically significant decreases in metabolic equivalent units (METs) were also reported following CO exposure (p = 0.0001). No CO-induced changes in heart rate (HR), BP, ECG parameters, or myocardial perfusion were observed.
Bathoorn et al. (2007)	19 former smokers with COPD; 18 males/1 female; 66-70 years old	Exposures to both CO and room air conducted for 2 h on each of four consecutive days in a randomized crossover design. 9 subjects were exposed to 100 ppm and 10 subjects were exposed to 125 ppm. CO and room air exposures were separated by at least 1 week.	Following the fourth day of exposure, CO inhalation reduced sputum eosinophils relative to room air and also increased the provocative concentration of methacholine required to cause a 20% reduction in FEV <sub>1</sub> . Neither of these effects were shown to reach statistical significance. No changes in sputum neutrophils, white blood cell counts or serum C-reactive protein (CRP) were observed. Although this study appears to demonstrate some evidence of an anti-inflammatory effect of CO among subjects with COPD, it must be noted that two of these patients experienced exacerbations of COPD during or following CO exposure, with one patient requiring hospitalization two months after exposure (initial symptoms first experienced 1-week post-exposure).
Hanada et al. (2003)	20 healthy males; 26 ± 1 years old	15 subjects exposed for 20 min (10 min rest, 5 min handgrip exercise, 2 min post-exercise ischemia, 3 min recovery) under the following four conditions: (1) normoxia (inspiratory O <sub>2</sub> fraction 21.4%), (2) hypoxia (inspiratory O <sub>2</sub> fraction 10.3%), (3) CO + normoxia, (4) CO + hyperoxia (inspiratory O <sub>2</sub> fraction 95.9%). Inhaled CO administered to achieve ~20% COHb in venous blood. Each of the four conditions were separated by 20 min of rest. 5 subjects served as controls (four consecutive 20 min periods of normoxia).	Blood oxygenation, BP, HR and respiratory rate were measured during exposure. Muscle sympathetic nerve activity (MSNA) and leg hemodynamics were evaluated in two subsets of the study group (n = 8 and 7, respectively). Arterial oxygen saturation (pulse oximetry) was significantly lower, and resting HR and ventilation significantly higher during the period of hypoxia compared to the other periods; none of these measures were affected by exposure to CO. MSNA was shown to increase during hypoxia and CO exposure relative to normoxia. Neither hypoxia nor CO was found to affect leg blood flow or vasoconstriction.
Kizakevich et al. (2000)	16 healthy non-smoking males; 18-29 years old	Subjects exposed on 4 separate days to increasing CO concentrations during either upper-body exercise (hand-crank) or lower-body exercise (treadmill). Targeted COHb levels were initially attained using short term (4-6 min) exposures to CO at concentrations between 1,000 and 3,000 ppm. Chamber exposures were then conducted at CO concentrations required to maintain COHb levels of <2% (room air), 5% (27 ppm), 10% (55 ppm), 15% (83 ppm), and 20% (100 ppm).	At all levels of upper- and lower-body exercise, exposures to CO resulted in increases in HR, cardiac output, and cardiac contractility relative to clean air exposures. Increases in HR reached statistical significance at COHb concentrations ≥ 5%, and increases in both cardiac output and cardiac contractility reached statistical significance at COHb concentrations ≥ 10%. CO exposure during exercise was not observed to cause ventricular arrhythmias or affect ECG wave shape (no evidence of ST-segment depression) at COHb concentrations ≤ 20%.
Mayr et al. (2005)	13 healthy non-smoking males; 18-38 years old	Subjects exposed to both 500 ppm CO and clean air for 1 h, with exposures separated by a 6-week period. Immediately following exposure, subjects were administered an intravenous bolus dose (2 ng/kg) of lipopolysaccharide (LPS).	The average COHb concentration was 7% following the 1 h CO exposure. Infusion of LPS significantly increased plasma concentrations of TNF-α, CRP, IL-6, and IL-8, with no difference in the inflammatory response between clean air and CO exposures.
Morse et al. (2008)	12 healthy non-smoking males; 25 ± 2.9 years old	Exposures conducted on two separate occasions to both room air (6 min) and CO. Subjects were exposed to 3,000 ppm CO until COHb reached 6% (3-8 min exposures).	Leg strength and muscle fatigue were evaluated immediately following exposure. CO exposure did not affect muscle strength (maximal voluntary isometric contraction), but did cause a statistically significant increase in muscle fatigue (p <0.05)

Study	Subjects	Exposure	Findings
Ren et al. (2001)	12 healthy subjects; 11 nonsmokers; 9 males / 3 females; 20-32 years old	Each subject underwent four different 8 h experimental protocols: (1) isocapnic hypoxia (end-tidal PO <sub>2</sub> held at 55 mmHg), (2) withdrawal of 500 mL of venous blood at the start of an 8 h period, (3) CO exposure at a concentration required to maintain a COHb level of 10%, and (4) a control exposure where subjects breathed room air with no intervention.	A statistically significant increase in ventilation was observed following hypoxia, but no such increase was found following any of the other 3 protocols, including exposure to CO. One subject felt faint during the blood withdrawal protocol and did not complete the study.
Resch et al. (2005)	15 healthy non-smoking males; 27 ± 4 years old	Subjects exposed to 500 ppm CO and synthetic air for 1 h at rest in a randomized crossover study design. Exposures were separated by a period of at least 1 week.	COHb levels averaged 5.6% after 30 min and 9.4% after 60 min of exposure. Statistically significant increases in retinal blood flow, retinal vessel diameter, and choroidal blood flow were observed with CO exposure relative to synthetic air at both time points. Exposure to CO did not affect oxygen saturation of arterial blood.
Vesely et al. (2004).	10 healthy non-smoking males; 22-52 years old	Each subject was exposed to CO for 30-45 min to achieve a COHb level of 10%. Prior to and following exposure, subjects performed hypoxic and hyperoxic rebreathing tests. Four subjects were exposed to hypoxic conditions first, while six subjects were exposed to hyperoxic conditions first, both prior to and following CO exposure.	Ventilation rate was observed to significantly increase during hypoxic rebreathing relative to hyperoxic rebreathing. However, exposure to CO had no effect on ventilation under either hypoxic or hyperoxic conditions. The authors concluded that exposure to low levels of CO does not significantly affect chemoreflex sensitivity of the CO <sub>2</sub> -induced stimulation of ventilation.
Zevin et al. (2001)	12 healthy male smokers; 27-47 years old	Exposures were conducted over 21 consecutive days under three different protocols, with each protocol lasting 7 days. In one protocol, subjects smoked 20 cigarettes per day, one every 45 min. In the other two protocols, every 45 min (20 times per day) subjects breathed either air or CO (~1200 ppm) from a 1 liter bag once per min for 10 min at a time. Subjects completed all three protocols, with six subjects exposed sequentially to CO, smoking, then air, and the other six exposed sequentially to air, smoking, then CO.	COHb levels were similar during smoking and exposure to CO, with average concentrations of 6% and 5%, respectively. Blood was drawn on day 4 of each exposure and analyzed for CRP, plasma platelet factor 4, and white blood cell count. Plasma levels of CRP and platelet factor 4 were significantly elevated with smoking, but not with CO exposure, relative to air control. HR and BP were evaluated on day 3 of each protocol. Cigarette smoke, but not CO, was observed to significantly increase HR, while no difference in BP was observed between any of the three exposures.



# Annex E. Toxicological Studies

**Table E-1. Human and animal studies.**

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Acevedo and Ahmed (1998)	Human			HO-1 and HO-2 (mRNA and protein) were upregulated in pregnant myometrium when compared to non-pregnant myometrium. The HO activator hemin inhibited spontaneous and oxytocin-induced contractility of the myometrium. Progesterone induced HO-1 and HO-2 mRNA expression.
Achouh et al. (2008)	Human arteries	Until equilibrium	Approximately 30 $\mu$ M	CO induced endothelium- and NO*-independent relaxation of precontracted human ITA and RA graft by partially stimulating cGMP production. The mechanism and extent of relaxation depended upon the tissue.
Ahmed et al. (2000)	Human			Placental HO-1 was significantly higher at term. HO-1 significantly attenuated TNF $\alpha$ -dependent cellular damage in placental explants. HO-1 was significantly attenuated in pre-eclampsia pregnancies versus non-pre-eclamptic pregnancies. Placental arteries exposed to the HO activator hemin demonstrated reduced vascular tension (i.e., placental blood vessel relaxation).
Ahmed et al. (2005)	Human			The source of CO in term human placental chorionic villi was found to be the catalysis of heme by HO and not endogenous lipid peroxidation.
Alexander et al. (2007)	Rat (Sprague Dawley) Adult female			Modulation of the HO/CO system in the anterior pituitary of the female rat led to altered secretion of gonadotropins and prolactin.
Alexandreaanu et al. (2002)	Rat (Sprague Dawley)			The role of the HO/CO system in estrous cyclicity, pregnancy and lactation was evaluated using HO inhibitors and substrates. The HO inhibitor CrMP decreased time in estrous. Administering HO-inhibitors to pregnant rodents induced total litter loss. CrMP induced decreased litter weight gain during lactation, which the authors attribute to maternal milk production or ejection problems as cross-fostered pups regained weight lost during nursing on CrMP dams.
Alexandreaanu and Lawson (2003a)	Rat (Sprague Dawley) Adult female			Modulation of the HO/CO system in the anterior pituitary of the female rat led to altered secretion of gonadotropins and prolactin.
Alexandreaanu and Lawson (2003b)	Rat (Sprague Dawley) Adult female ovary			HO-1 and HO-2 were localized in the ovaries in rats and treatment of rat ovaries in vitro with CrMP, an inhibitor of HO, or with hemin, a substrate for HO induced steroidogenic changes in the ovaries.
Alonso et al. (2003)	Human muscle tissue mitochondria	5 min	50-500 ppm	CO significantly reduced muscle mitochondrial cytochrome <i>c</i> oxidase activity by 20%, 42%, and 55% after treatment with 50, 100, and 500 ppm CO respectively but did not change the activity of three other electron transport proteins.
Andresen et al. (2006)	Rat (Long Evans) Male Mouse (C57BL/6J) Male Cerebral vessels		1-100 $\mu$ M	CO did not dilate rat or mouse cerebral arteries until 100 $\mu$ M, which is not a physiological concentration. Also, the HO inhibitors constricted vessels in a nonspecific manner.
Antonelli et al. (2006)	Rat (Wistar)	GD5-GD20	75 ppm	Pups exposed to CO in utero had significant impairment of cortical neuronal glutamatergic transmission at PND1 in both neurons at rest and in neurons stimulated with depolarization.
Appleton and Marks (2002)	Human placenta			Endogenous CO production by HO in the human placenta was regulated by O <sub>2</sub> availability. Placental HO activity was directly dependent on O <sub>2</sub> availability; this does not vary between pre-eclamptic and normotensive placentas.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Ashfaq et al. (2003)	Human placenta			Placentas were collected from smokers and nonsmokers who gave birth to male infants. Premature aging and a statistically significant increase in apoptotic cells were seen in placentas from smokers vs nonsmokers.
Astrup (1972)	Rabbit (strain not identified)	Continuous CO exposure over gestation	90 or 180 ppm	Skeletal abnormalities: Three pups (from n = 123) in the 180 ppm CO group had deformities in their extremities at birth, whereas no control and no 90 ppm CO-exposed animals manifested with this malformation.
Bainbridge et al. (2002)	Human placenta			Isolated human placenta exposed to solutions containing CO demonstrated a concentration-dependent decrease in perfusion pressure further demonstrating the role of CO in maintaining basal vasculature tone.
Bainbridge et al. (2006)	Human placenta	6 h	Starting concentrations of CO: 3.9 $\mu$ M CO in cell culture media (control) and CO-exposed groups: 116 $\mu$ M, 145 $\mu$ M, 181 $\mu$ M. After 3 hours, the CO in the culture media was 3.7 $\mu$ M (control), and CO-exposed cells 10.2, 12, and 15.9 $\mu$ M.	C-section placentas were collected from healthy term pregnancies. Villous explants of placentas were cultured under hypoxia followed by reoxygenation (H/R). H/R + CO-exposed placental tissue had decreased apoptosis and decreased PARP (a protein marker of apoptosis) versus control H/R exposed cells. Secondary necrosis of the placental tissue post H/R was inhibited by CO treatment.
Bainbridge and Smith (2005)	Human placenta			The role of HO in the placenta and during pregnancy are reviewed in this article. The conflicting data on the activity, localization, and expression of HO in the placentas of pre-eclamptic women are presented.
Bamberger et al. (2001)	Human placenta			Expression and tissue localization of soluble guanylyl cyclase in human placenta using antibody localization were characterized. These tools can be used in future studies to elucidate the NO <sup>*</sup> /CO/cGMP pathway.
Barber et al. (1999)	Human			HO and NOS did not maintain human uterine quiescence during pregnancy.
Barber et al. (2001)	Human placenta			Women who had pregnancies with fetal growth restrictions (FGR) produced term placenta with significant decreases in HO-2 versus healthy pregnancies.
Baum et al. (2000)	Human			End-tidal CO measurements in women with pregnancy-induced hypertension and pre-eclampsia were significantly lower than in normotensive pregnant women.
Bergeron et al. (1998)	Rat Brain			To address the developmental changes of HO staining in the brain, immunohistochemical staining for HO-1 was performed on the developing rat brain at PND7, PND14, and PND21. HO-1 staining was most intense at PND7 and by PND21 reached its adult pattern of staining localizing to the hippocampus, thalamic and hypothalamic nuclei, with virtually no staining of endothelium, white matter and cortex. HO-2 is the dominant HO isoform in the brain.
Bing et al. (1995)	Rodent			Spatial learning in the Morris water maze was enhanced in rodents exposed to the HO inhibitor tin protoporphyrin (Sn-PP).
Burmester et al. (2000)	Human and Mouse			Nb had a high oxygen affinity similar to Mb, thus may increase the availability of O <sub>2</sub> to brain tissue.
Bye et al. (2008)	Rat (Wistar) Female	100 h/wk for 18 mo	200 ppm	CO-exposed (11-14.7% COHb) rats experienced a 24% decrease in aerobic capacity evidenced by VO <sub>2</sub> max deficits. Left ventricular cardiomyocytes were longer and wider, had increased expression of growth-related proteins, and had impaired contraction-relaxation cycles. CO increased cGMP and impaired cardiomyocyte Ca <sup>2+</sup> handling. No change in BP was observed.
Cagiano et al. (1998)	Rat	GD0-GD20	75 or 150 ppm	At 5 months of age, CO-exposed male offspring showed decrements in sexual behavior including an increase in mount to intromission latency, a decrease in mount to intromission frequency, and a decrease in ejaculation frequency. Basal extracellular dopamine concentration in the nucleus accumbens was unchanged after CO-exposure. However, when stimulated with amphetamine administration, control rats had increased release of dopamine that is absent with CO-exposed rats.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Carmines and Rajendran (2008)	Rat (Sprague Dawley)	GD6-GD19 of gestation for 2 h/day	600 ppm	Significant decreases in birth weight were reported after CO exposure. Maternal body weight was unchanged during gestation, but corrected terminal body weight (body weight minus uterine weight) was significantly elevated in CO-exposed dams at term.
Carratu et al. (1995)	Rat (Wistar)		150 ppm	Sphingoid homeostasis was disrupted in male offspring of prenatally exposed rats, without a disruption in motor function.
Carratu et al. (2000)	Rat (Wistar)	GD0-GD20	150 ppm	Maternal COHb (mean % $\pm$ SEM) was $1.9 \pm 0.04$ and $16.02 \pm 0.98$ in control and 150 ppm CO-exposed animals, respectively. Prenatal CO exposure had no effect on brain sphinganine (SA) or sphingosine (SO) levels in male offspring at 90 days of age. However, the sciatic nerve had significant increases in SO after CO exposure, no changes in SA at 90 days of age. Motor activity, which could be affected by changes in myelination, showed no differences between CO and control animals at 90 days of age.
Carratu et al. (2000)	Rat (Wistar)	GD0-GD20	75 or 100 ppm	The myelin sheath thickness of the nerve fibers was significantly decreased in CO-exposed animals (75 and 150 ppm). Axon diameter was not affected by CO exposure. Even though CO affected myelination, it did not significantly affect motor activity of CO-exposed rats at 40 and 90 days.
Carraway et al. (2002)	Rat model of hypoxic pulmonary vascular remodeling Strain of rat not stated	3 weeks	Hypobaric hypoxia $\pm$ 50 ppm	CO promoted remodeling and increased pulmonary vascular resistance in response to HH. The number of small muscular vessels was increased compared with HH alone. Changes in cell proliferation, apoptosis, actin and HO-1 gene and protein expression correlated with structural changes. COHb levels were <0.5% in controls, 1.5-2.8% in the HH treatment group and 3.5-3.9% in the HH + CO treatment group.
Cella et al. (2006)	Rat (Sprague Dawley)			HO-1 production and HO concentration were shown to be regulated by estrogen in the rat uterus.
Chung et al. (2006)	Rat (Sprague Dawley) Male		3-6%	CO inactivation of Mb does not induce any change in the respiration rate, contractile function, or high-energy phosphate levels in perfused rat hearts.
Cronje et al. (2004)	Rat (Sprague Dawley) Male 240-325 g	45 min	2,500 ppm	Results indicate that tissue and blood [CO] (66-72% COHb) dissociate during CO inhalation, but tissue [CO] does not follow blood [CO] or $1/pO_2$ as in the Warburg theory during intake or elimination. Tissue [CO] increases later during the resolution period and varies significantly among animals and tissues. The deviation from the predicted values in the brain is likely due to the release of heme and increase in NADPH stimulating endogenous CO production by HO. Immediately following exposure, tissue CO concentrations were found to be:  Blood: 27,500 (800) pmol/mg Heart: 800 (300) pmol/mg Muscle: 90 (80) pmol/mg Brain: 60 (40) pmol/mg  These values are estimates taken from a graph, with control levels in parentheses  A later report stated that these tissue CO values were too high due to a computational error (Piantadosi et al., 2006)
Cudmore et al. (2007)	Human placenta Human (HUVEC) Mouse (HO-1 deficient mouse on 129/SV x C57BL/6 background) Pig (Porcine aortic endothelial cells)			HUVEC cells, porcine aortic endothelial cells, HO-1 null mice and placental villous explants (normotensive and pre-eclamptic pregnancies) were used in this study. The HO-1/CO system inhibited sFlt-1 and sEng release, two factors upregulated in pre-eclampsia.
D'Amico et al. (2006)	Human embryonic kidney (HEK293) cells	0-30 min	20 $\mu$ M	Exogenous CO inhibited respiration in HEK293 cells under ambient $O_2$ concentration (21%). Inhibition was enhanced under hypoxic conditions. Increased endogenous CO resulting from HO-1 overexpression inhibited respiration by 12% and cytochrome <i>c</i> oxidase activity by 23%. This effect was enhanced under hypoxic conditions.
Dani et al. (2007)	Human (neonatal blood)			CO was lower at birth and 48-72 hours postpartum in infants born by elective C-section and higher in vaginally born infants.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
De Salvia et al. (1995)	Rat (Wistar)	GD0-GD20	75 or 150 ppm	Animals exposed to the higher dose of CO (150 ppm) in utero had significantly impaired acquisition (at 3 and 18 months) and reacquisition (at 18 months) of conditioned avoidance behavior.
Denschlag et al. (2004)	Human			Genetic polymorphisms in human HO-1 are linked to idiopathic recurrent miscarriages.
Dewilde et al. (2001)				Nb exists as a reversibly hexacoordinated Hb type with a His-Fe <sup>2+</sup> -His binding scheme. Dissociation of the internal ligand by O <sub>2</sub> or CO is the rate limiting step.
Dubuis et al. (2002)	Rat (Wistar) Adult female 250 g	3 weeks	530 ppm	Intrapulmonary resistance artery smooth muscle cells were isolated from control and exposed rats. Electrophysiological recordings provided evidence of increased Ca <sup>2+</sup> -activated K <sup>+</sup> current consequent to chronic CO exposure. The authors speculated that this could in part explain the vasodilatory effect of CO in the pulmonary circulation.
Dubuis et al. (2005)	Rat (Wistar) Male	21 days	50 ppm	CO attenuated PAHT by activating BK <sub>Ca</sub> channels in PA myocytes and reduced hemodynamic changes of PAHT.
Dubuis et al. (2003)	Rat (Wistar) Male	21 days	50 ppm	CO induced relaxation of pulmonary artery rings in normoxic, hypoxic, and hypoxic-CO rats and it was not endothelium dependent. Chronic hypoxia decreased acute CO sensitivity, while CO-hypoxia increased it. K <sup>+</sup> channel blocker reduced this effect while sGC blocker did not.
Favory et al. (2006b)	Rat 250-300 g Strain not stated	90 min	250 ppm	CO inhibited myocardial permeabilized fiber respiration (complex IV), increased coronary perfusion pressure and left ventricular developed pressure (LVDP) first derivative and decreased the cGMP/cAMP ratio in the heart. These changes were maintained over 24-48 h of recovery in air. Cardiac function and vasodilatory responses were evaluated at 3-h recovery in air. β-adrenergic blockade had no effect on coronary perfusion pressure or LVDP first derivative. Total inhibition of vasodilator response to acetylcholine and partial inhibition of vasodilator response to nitroprusside were observed. An increase in myofilament calcium sensitivity was also observed. Thus CO promotes abnormalities in mitochondrial respiration, coronary vascular relaxation and myocardial contractility. The authors speculated that CO may have a detrimental effect on heart O <sub>2</sub> supply-to-utilization which could potentially lead to myocardial hypoxia because of the increased O <sub>2</sub> demand resulting from increased contractility, the inhibited mitochondrial respiration and the reduced coronary blood-flow reserve resulting from the decreased vasodilatory capacity. COHb was found to be 11% immediately after exposure. COHb levels gradually returned to baseline (1.5%) over the next 96 h.
Fechter and Annau (1977)	Rat (Long Evans)	Continuous CO exposure throughout pregnancy	150 ppm CO	The authors found a 5% significantly decreased birth weights at PND1 in gestationally CO-exposed pups versus control animals with weight decrements persisting to weaning; lactational cross fostering did not ameliorate the CO-dependent reduced growth rates. Dams exposed to CO during gestation had COHb over gestation of 15% with control dams having less than 1%. Decreased birth weight and pre-weaning weight were seen in CO-exposed pups despite a lack of weight decrement in CO-exposed dams versus air-exposed control dams.
Fechter et al. (1980)	Rat (Long Evans)	Continuous CO exposure throughout pregnancy	150 ppm	CO-exposed animals had cardiomegaly at birth (wet heart weight) that dissipated by PND4.
Fechter and Annau (1980)	Rat	Continuous CO exposure throughout pregnancy	150 ppm	CO-exposed animals had decreased birth weight, impaired righting reflexes, impaired negative geotaxis, and delayed homing behavior.
Fechter et al. (1987)	Rat (Long Evans)	Continuous CO exposure throughout pregnancy or from GD0 to PND10	75, 150, or 300 ppm	The neostriatum of PND21 rat brains were collected and showed disrupted development following CO exposure (GD0-PND10 group, 300 ppm CO). Dopamine levels were also significantly elevated in CO-exposed animals (GD0-PND10, 150 and 300 ppm CO).

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Garofolo et al. (2002)	Human infants Rat	Rat: PND2-PND5		Human infants who die from SIDS show decreased brainstem muscarinic receptor binding versus infants dying from other causes. $\beta$ -adrenergic modulation of muscarinic receptors in developing heart was observed.  Rodent $\beta$ -adrenergic agonists at PND2-PND5 induced muscarinic receptor decrement in adenylyl cyclase.
Gautier et al. (2007)	Rat (Wistar) Adult male  Model of right ventricular hypertrophy secondary to chronic hypoxia	3 weeks of HH $\pm$ CO in final week Or 1 week of CO	50 ppm	CO altered the right ventricular adaptive response to pulmonary hypertension which occurs secondarily to chronic hypoxia. Right ventricular end-systolic pressure (RVESP) and right ventricular shortening fraction (RVSF) were smaller in rats treated with CO+HH compared with rats treated with HH alone. CO alone had no effect on these measures. Hypobaric hypoxia had no effect on left ventricular function while CO+HH led to an increased left ventricular shortening fraction (LVSF). CO alone led to a decrease in LVSF and the mitral E-to-A ratio, indicative of a LV filling impairment. Hypobaric hypoxia decreased the relative RV perfusion and increased the relative LV perfusion. These effects were prevented with concomitant exposure to CO although exposure to CO alone had no effects on myocardial perfusion. Morphologic and histologic analysis demonstrated RV hypertrophy in both the HH group and the CO+HH group and fibrotic lesions in the CO+HH group. The authors concluded that the 1-week exposure to 50 ppm CO had a deleterious effect on RV myocardial perfusion adaptation to chronic hypoxia and pressure overload. Although the reduced RV pressure overload was beneficial it was counterbalanced by impaired RV perfusion and redistribution of perfusion toward the LV.
Gaworski et al. (2004)	Rat (Sprague Dawley)	2 h/day, 7 days/week by nose-only inhalation  Males: 4 weeks prior to and during mating; and  Females: 2 weeks prior to mating, during mating, and through weaning to PND21	Cigarette smoke: 150, 300, or 600 mg/m <sup>3</sup> Total Particulate Matter (TPM)	Maternal exposure to high concentrations of cigarette smoke during gestation and lactation reduced pup birth weight and retarded neonatal pup growth. Developmental and neurobehavioral testing of neonates did not show any behavioral effects following parental smoke exposure.
Ghio et al. (2008)	Rat (Sprague Dawley) Adult male  Human bronchial epithelial cells (BEAS-2B)	24 h  2-24 h	50 ppm  10-100 ppm	Mild neutrophil accumulation was observed in BALF accompanied by increases in BALF MIP-2, protein and LDH. Iron status was altered since CO exposure led to an increase in BALF iron and ferritin, a decrease in lung non-heme iron and an increase in liver non-heme iron.  CO exposure for 24 h led to a dose-dependent decrease in cellular non-heme iron, with the effect at 10 ppm statistically significant and the effect at 50 ppm maximal. This effect was reversible since removing the cells after 2 h of CO and incubating them in air restored non-heme iron concentrations at 24 h. A dose-dependent decrease in cellular ferritin was observed following exposure for 24 h to 50-500 ppm CO. In addition, exposure to 50 ppm CO for 20 h blocked iron uptake by cells while exposure to 50 ppm CO for 2 h increased iron release from cells. Increased protein expression of the iron transporter DMT-1 was also noted after 24 h exposure to 50 ppm CO. Oxidative stress, mediator release and cell proliferation were also decreased by exposure to 50 ppm for 24 h. This effect was also reversible upon removal to air. Effects of CO on cell proliferation indices were mimicked by with the iron-depleting agent deferoxamine. The authors concluded that CO exposure altered lung iron homeostasis possibly by initially causing heme release from proteins.
Giustino et al. (1999)	Rat (Wistar) Male and Pregnant female	From GD0-GD20 of pregnancy	75 or 150 ppm	This study showed that CO (75 and 150 ppm) exposed male animals at 40 days of age had a significantly decreased time of exploration of novel objects. The 150 ppm CO group showed a lack of habituation after the second exposure to a previously viewed object. Blood COHb concentrations (mean % $\pm$ SEM) on GD20 were reported (0 ppm: 1.6 $\pm$ 0.1; CO 75 ppm: 7.36 $\pm$ 0.2; CO 150 ppm: 16.1 $\pm$ 0.9).

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Giustino et al. (1993)	Rat (Wistar)	GD0-20	75, or 150 ppm	CO exposure in utero led to a reversible and dose dependent loss of function of splenic macrophages with decreased killing ability, decreased phagocytosis, and decreased ROS production during the macrophage respiratory burst.
Glabe et al. (1998)	Rat (Sprague Dawley) Male, Myocardium		P <sub>CO</sub> = 0 - 107 Torr	Increased P <sub>CO</sub> and increased COMb saturation did not alter high energy phosphate signals (ATP, phosphocreatine, Pi). MV <sub>O2</sub> began to decline at 87.6% COMb and is likely not due to cytochrome <i>c</i> oxidase inhibition.
Grover et al. (2000)	Fetal lamb (mixed breed)	10 min	500 ppm	Fetal methoxyhemoglobin (COHb%) ranged from 3.8 ± 0.2 to 8.1 ± 2.0 at 0 and 500 ppm CO, respectively. Inhaled 0-500 ppm CO administered to near-term fetal lambs did not induce pulmonary vasodilation (main pulmonary artery, left pulmonary artery, aorta and left atrium) and the HO inhibitor zinc protoporphyrin IX failed to affect baseline vascular tone.
Hara et al. (2002)	Rat (Sprague Dawley) Male	40 min	1,000-3,000 ppm	CO exposure increased extracellular dopamine levels and decreased its major metabolites in a Na <sup>+</sup> -dependent pathway. CO withdrawal and reoxygenation caused levels to return to control or overshoot which may suggest an increase in oxidative metabolism of CO, mediated by MAO-A.
Harada et al. (2004)	Pig (Granulosa cells)			In this porcine model, HO was able to augment granulosa cell apoptosis allowing for proper follicular maturation.
Hendler and Baum (2004)	Human			End-tidal breath CO measurements in pregnant women with contractions (term and pre-term) were lower than those measurements in non-contracting women.
Hofmann and Brittain (1998)	Human			Partitioning of O <sub>2</sub> and CO in the human embryonic Hb is discussed.
Iheagwara et al. (2007)	Mouse (C57Bl6) Male	3 h	1,000 ppm	CO significantly reduced cytochrome <i>c</i> oxidase activity and V <sub>max</sub> but not K <sub>m</sub> in myocardial mitochondria. Cytochrome <i>c</i> oxidase protein levels and heme content were significantly decreased. The average COHb level was 61% but no tissue hypoxia was observed in the heart.
Imai et al. (2001)	HO-1 transgenic mice which specifically over-express HO-1 in smooth muscle			Transgenic mice had a significant increase in arterial pressure and impaired nitrovasodilatory aortic responses. The mice had enhanced NO <sup>*</sup> production and impaired sGC activity. The authors speculated that the effect of HO-1 overexpression was to suppress vasodilatory responses to NO <sup>*</sup> in vascular smooth muscle.
Ischiropoulos et al. (1996)	Rat (Wistar) Male 200-290 g	60 min 40-60 min	1,000-3,000 ppm 1,000 ppm	CO poisoning resulted in free NO <sup>*</sup> in brains as measured by electron paramagnetic resonance spectroscopy and in a 10-fold increase in nitrotyrosine as measured by immunohistochemical staining. These responses were blocked by pretreatment with a NOS inhibitor but not by neutrophil depletion.  Brain nitrotyrosine formation was blocked by platelet depletion following 40 min but not 60 min exposure to 1,000 ppm CO.  Following CO poisoning, myeloperoxidase activity, a measure of leukocyte sequestration, was increased in brain microvessels. This response was blocked by NOS inhibition but not by platelet depletion. Similar effects were noted for xanthine oxidase activation.  The authors concluded that perivascular reactions mediated by peroxynitrite are key to CO poisoning effects in brain.
Johnson and Johnson (2003)	Rat (Sprague Dawley) Male 250-300 g		0-100 μM	CO produced a concentration dependent, endothelium-dependent vasoconstriction in isolated gracilis muscle arterioles, evident at 1 μM CO. Pre-treatment with a NOS substrate prevented this response while pretreatment with a NOS inhibitor converted this response to a vasodilation. The authors concluded that exogenous CO was acting through NOS inhibition.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Johnson et al. (2003)	Rat (Dahl/Rapp salt-resistant and salt-sensitive model) Male			High-salt diet increased COHb, BP, and aortic HO-1 protein levels in salt-sensitive Dahl rats. Enhanced immunostaining was observed for HO-1 but not HO-2 in isolated gracilis muscle arterioles. Compared with the low-salt diet, the high-salt diet resulted in a smaller vasoconstrictor response when NOS was inhibited. Vasoconstriction was exacerbated in arterioles from both low salt- and high salt-treated rats using both NOS and HO inhibitors. Acetylcholine-induced vasodilation was diminished in the high-salt diet group compared with the low-salt diet group. This effect was not seen using the HO inhibitor. The high-salt diet did not alter endothelium-independent vasodilation. The authors concluded that HO-derived CO caused dysfunction of the NO* system in salt-sensitive rats treated with a high-salt diet.
Johnson et al. (2004)	Rat (Sprague Dawley) Male Deoxycorticosterone acetate (DOCA)-salt hypertension model WKY rats Spontaneously hypertensive rats (SHR)			Salt-sensitive DOCA rats, but not SHR, had elevated aortic HO-1 expression and blood COHb levels. Both had elevated mean arterial BP compared with controls. Acetylcholine-mediated vasodilation of isolated gracilis muscle arterioles was attenuated in DOCA rats but not SHR. Pretreatment with a HO inhibitor restored the response in DOCA rats. The authors concluded that HO-1-derived CO contributes to endothelial dysfunction in DOCA but not SHR.
Johnson et al. (2006)	Rat (Zucker) Lean and obese Male		100 µM CO	The obese rats had increased CO expiration and mean arterial pressure which was decreased by pretreatment with a HO inhibitor. No difference was observed in HO-1 protein between lean and obese rats. Acetylcholine- and flow-mediated vasodilation of isolated gracilis muscle arterioles was attenuated in obese but not lean rats. Pretreatment with a HO inhibitor restored the response in obese rats. Exogenous CO prevented the restoration of flow-induced dilation by the HO inhibitor. The authors concluded that HO-derived CO contributes to endothelial dysfunction in this model of metabolic syndrome.
Katoue et al. (2005)	Rat (Wistar)			HO activity in the aorta is significantly increased during pregnancy but aortic AVP-dependent vasoconstriction appears to be HO/CO independent.
Katoue et al. (2006)	Rat (Wistar)			Pregnancy-induced modulation of calcium mobilization and down-regulation of Rho-kinase expression contributed to attenuated vasopressin-induced contraction of the rat aorta.
Khan et al. (2006)	Nb overexpressing BDF × CD1 mice			Cerebral and myocardial infarcts were decreased in neuroglobin overexpressing mice, decreasing ischemic injury.
Kim et al. (2005)	Primary rat pulmonary artery smooth muscle cells Inbred LEW rat (Sprague Dawley) 200-250 g	24 h or pretreatment for 1-2 h followed by 24 h posttreatment	250 ppm	Exposure of cells in culture to 250 ppm CO for 24-h inhibited serum-stimulated cell proliferation, increased expression of p21 <sup>Waf1/Cip1</sup> and decreased expression of cyclin A. CO also inhibited PDGF-stimulated cell proliferation and reversed the inhibitory effect of PDGF on caveolin-1 expression. Genetic silencing of caveolin-1 using siRNA, prevented the antiproliferative effect of CO. Endogenous CO derived from HO-1 in an overexpression system was found to upregulate caveolin-1 expression. Effects of CO on caveolin-1 were found to be mediated by p38 MAPK and cGMP. Experiments in fibroblasts deficient in p38 confirmed a role for p38 in CO-mediated inhibition of cellular proliferation via effects on p21 <sup>Waf1/Cip1</sup> , cyclin A and caveolin-1. Experiments in fibroblasts deficient in caveolin-1 confirmed the role of caveolin-1 in the anti-proliferative effects of CO.  In a model of neointimal injuries induced by balloon injuries in intact animals, exposure to CO inhibited neointimal formation and increased caveolin-1 expression in the intima and media.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Kim et al. (2008)	Primary rat hepatocytes Primary mouse hepatocytes Respiration-deficient human Hep3B cells	10-60 min	250 ppm	Exposure of cells in culture to 250 CO for 1 h twice a day prevented spontaneous hepatocyte death over 6 days in culture. CO also decreased caspase-3 activity. Cell death was determined to be partly due to apoptosis. CO also increased ROS as measured by dichlorofluorescein fluorescence in rat hepatocytes, mouse hepatocytes and Hep3B cells but not in respiration-deficient Hep3B cells indicating that ROS were mitochondrial in origin. An increase in mitochondrial oxidized glutathione was noted in rat hepatocytes treated with CO for 30 min. Increased Akt phosphorylation occurred following 10-30 min CO and was diminished by treatment with antioxidants. CO was found to activate NFκB through a PI3K and oxidant-dependent pathway. CO mediated spontaneous cell death was found to be dependent on ROS and Akt phosphorylation. The authors concluded that CO prevents hepatocyte apoptosis through redox mechanisms leading to cytoprotection.
Kinobe et al. (2006)	Sheep (gravid and non-gravid sheep and their near term fetuses)			There were no significant differences in hypoxic adult and hypoxic fetal sheep when compared to their normoxic controls.
Knuckles et al. (2008)	Mouse	4 h	Diesel emissions: 350 µg/m <sup>3</sup>	Diesel exhaust enhanced vasoconstriction in veins but not arteries. It was suggested that this is through the uncoupling of eNOS.
Kreiser et al. (2004)	Human			End tidal CO concentrations were lower in pregnant women with gestational hypertension and pre-eclampsia than normotensive women.
Lash et al. (2003)	Human Term placental chorioic villi from healthy or pre-eclamptic placentas			Infarcted areas of placenta had decreased HO expression (in pre-eclamptic placenta only).
Li et al. (2008)	Mouse (ICR [CD-1]) Pregnant			The effect of maternal LPS exposure on fetal liver HO was measured. HO-1 was upregulated in fetal livers post-LPS exposure and this HO-1 upregulation was attenuated with the spin trap agent PBN, pointing to a ROS dependent HO-1 upregulation post maternal LPS treatment.
Loennechen et al. (1999)	Rat (Sprague Dawley) Female 220-240g	1 week 1 week 100 ppm and 1 week 200 ppm	100 ppm 100-200 ppm	Endothelin-1 expression increased by 53% and 54% in the left and right ventricle respectively during the 2 week exposure and by 43% and 12% in the left and right ventricle respectively during the 1 week exposure. Right ventricular to body weight ratio was increased by 18% and 16% in the 2 week and 1 week exposure groups respectively. COHb levels were 23% and 12% in the 2 week and 1 week exposure groups respectively.
Longo et al. (1999)	Rat (Sprague Dawley) Human			The addition of exogenous CO to isolated human and rat uterine tissue failed to induce relaxation of uterine tissue. Isolated rat aortic rings and tail artery rings from pregnant dams can be relaxed by submersion in exogenous CO solutions.
Lopez et al. (2008)	Rat (Sprague-Dawley)	Pregnant rats exposed to CO GD5-GD20 (Group A) or GD5-GD20 plus PND5-PND20 (Group B); Group C (control air exposure).  10 - 18 h/day	25 ppm	CO exposure induced damage to the spiral ganglia neurons and inner hair cells with oxidative stress seen in cochlear blood vessels. At PND20 groups A and B show vacuolization of afferent terminals at the base of the cochlea. At PND3, group A shows decreased synapsin-1 staining of the efferent nerve terminals. At PND20, groups A and B show decreased neurofilament-IR (staining) in type I spiral ganglia neurons and afferent nerve fibers. At PND12 and PND20, group B shows increased HO-1 and SOD-1-IR in blood vessels of the stria vasularis; group A is similar to controls. From PND3-PND20, there is increased iNOS and increased nitrotyrosine-IR in blood vessels of the cochlea.
Lopez et al. (2003)	Rat (Sprague-Dawley)	PND6 to weaning (PND19-PND20)	12 or 25 ppm	In the cochlea, atrophy or vacuolization of the nerve cells that innervate the inner (not outer) hair cells was seen. Fibers of the 8th cranial nerve (internal auditory canal of the ARCO animals, 25 ppm) had distorted myelination and vacuolization of the axoplasm. In the organ of corti and spiral ganglion neurons, cytochrome <i>c</i> oxidase and NADH-TR were significantly decreased in 25 ppm exposure group versus control. Expression of the calcium-mediated myosin ATPase in the organ of corti and spiral ganglion neurons was significantly decreased in the 25 ppm CO exposure group versus controls.



Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Lund et al. (2007)	Mouse (ApoE <sup>-/-</sup> ) Male High fat diet	6 h/day, 7 days/week, 7 weeks	8, 40, or 60 µg/m <sup>3</sup> PM whole gasoline exhaust; or filtered exhaust with gases matching the 60 µg/m <sup>3</sup> concentration. CO concentrations were 9, 50, and 80 ppm corresponding to the 8, 40, and 60 µg/m <sup>3</sup> PM whole exhaust exposures	Both whole and filtered exhaust increased aortic mRNA expression of matrix metalloproteinase-3 (MMP-3), MMP-7, and MMP-9, tissue inhibitor of metalloproteinases-2, endothelin-1 and HO-1 at 60 µg/m <sup>3</sup> . Aortas also showed increased immunostaining for MMP-9 and nitrotyrosine in 60 µg/m <sup>3</sup> PM whole exhaust and PM-filtered exhaust exposed groups. Aortic TBARS, a measure of lipid peroxidation, was also increased in all treatment groups.
Lund et al. (2009)	Mouse (ApoE <sup>-/-</sup> ) Male High fat diet	6 h/day, 1 or 7 days	Gasoline engine exhaust containing 60 µg/m <sup>3</sup> PM and 80 ppm CO	Gasoline exhaust exposure increased aortic MMP-2/9 activity at 1 and 7 days. Protein levels of aortic MMP-9, MMP-2, TMP-2 and plasma MMP-9 were also increased after 7 days. Lipid peroxidation in aorta resulting from gasoline exhaust exposure was inhibited by treatment with the antioxidant Tempol, while increases in mRNA for ET-1 and MMP-9 in aortas were inhibited by treatment with BQ-123, an antagonist of ET <sub>A</sub> receptor. Treatment with BQ-123 also reduced aortic MMP-2/9 activity in aortas following gasoline exhaust exposure. The authors concluded that ET <sub>A</sub> receptor pathway is a key mediator of gasoline engine exhaust effects in the vasculature.
Lyll and Myatt (2002)	Human			Women with pre-eclampsia, produced term placenta with significant decreases in HO-2 versus women with healthy pregnancies.
Lyll et al. (2000)	Human (placentas from 8-19 weeks pregnancy and term placentas)			The use of a HO inhibitor, ZnPP, increased placental perfusion pressure. HO-1 and HO-2 were expressed in the placenta and placental bed and vary in expression over the course of pregnancy. HO may thus be involved in trophoblast invasion, placental function and perfusion pressure.
Mactutus and Fechter (1984)	Rat (Long Evans)	Continous exposure to CO over gestation	150 ppm	Acquisition as measured in a two-way conditioned avoidance (flashing light warnings followed by mild footshock) test failed to improve with age of in utero CO-exposed (150 ppm, dam COHb 15%) rats (male and female offspring) in contrast to air-exposed controls who improved with age/maturation, indicating a failure in the associative process of learning. The authors also found impairments in reacquisition performance, an index of retention, in PND31 rats that had received continuous in utero CO exposure. Prenatal CO exposure induced learning and memory deficits in male and female offspring.
McGregor and Westcott (1998)	Guinea pig	GD23-GD25 until term (approximately 68 days) 10 h/day	200 ppm	Aberrant respiratory responses (to asphyxia and CO <sub>2</sub> ) of offspring with prenatal CO exposure. The authors hypothesized this may be related to changes in the brainstem. COHb in maternal (8.53 ± 0.6% versus 0.25 ± 0.1%) and fetal blood (13.0 ± 0.4% versus 1.6 ± 0.1%) from CO-treated versus controls.
McLaughlin et al. (2001)	Human			Various pathologies of pregnancy including IUGR and pre-eclampsia are associated with significant decreases in placental HO activity. The endogenous generation of CO in the placenta has been demonstrated in choionic villi of term placenta.
McLaughlin et al. (2000)	Human placenta			Placental regional localization of HO was explored. The chorionic plate, chorionic villi, basal plate and choorio-decidua had significantly higher HO activity than the amnion.
McLaughlin et al. (2003)	Human placenta			HO expression in various regions of term placentas was explored. Microsomal HO-2 protein content was not different between normotensive and milk pre-eclamptic pregnancies. There was increased expression of microsomal HO-1 protein in chorionic villi and fetal membranes from pre-eclamptic pregnancies versus normotensive pregnancies.
McLean et al. (2000)	Human placenta			HO activity was highest in the placenta near term.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Melin et al. (2002)	Rat (Dark Agouti) Male Model of right ventricle hypertrophy secondary to chronic hypoxia (HH 10 weeks)	10 weeks	50 ppm alone or concomitant with HH	Hb and hematocrit levels were increased above controls in HH rats, CO rats and HH+CO rats, with the increase due to the combined treatment significantly higher than the increase due to HH. COHb levels were 1.1% in controls, 1.3% in HH rats, 4.7% in CO rats and 9.1% in HH plus CO rats. HH treatment significantly increased right ventricular (RV) heart weight above controls while CO treatment had no effect on any postmortem heart weights. Combined treatment with HH+CO resulted in a significant increase in left ventricular plus septum (LV+S) weight and RV weight compared with HH treatment alone. Echocardiographic left ventricular morphology and mass also showed the greatest changes in the HH+CO group. Hemodynamic measurements of LV function demonstrated significant effects in the HH+CO group for left ventricular end diastolic pressure (LVESP), left ventricular maximal first derived pressure ( $+dP/dt_{LV}$ ), and left ventricular work (LVW) compared with controls. Hemodynamic measurements of RV function demonstrated significant effects in the HH group for right ventricular end systolic and diastolic pressure (RVESP, RVEDP), right ventricular maximal and minimal first derived pressure ( $+dP/dt_{RV}$ , $-dP/dt_{RV}$ ) and right ventricular work (RVW). CO significantly enhanced the effects of HH on RVEDP and significantly diminished the effects of HH on $dP/dt_{RV}$ and RVW. The authors concluded that CO intensified the HH-induced RV hypertrophy, increased LV weight and induced severe hematological responses that could hamper adaptation.
Melin et al. (2005)	Rat (Dark Agouti) Male and female Model of right ventricle hypertrophy secondary to chronic hypoxia (HH, 10 weeks) Half of the animals were exercise-trained to induce LV hypertrophy	10 weeks	50 ppm alone or concomitant with HH	In untrained animals, combined treatment with HH+CO led to increased LV+S and RV weights compared with HH treatment alone. HH+CO led to several changes in measured echocardiographic parameters including increased anterior and posterior wall thickness in diastole (AWTd, PWTd) and to increased fraction of shortening. These effects were not seen with HH alone. In addition RVEDP was enhanced in HH+CO compared with HH alone. HRV components were altered by HH+CO but not by CO alone.
Mereu et al (2000)	Rat (Wistar)	GD0-GD20 continuous CO exposure	150 ppm	In utero exposure to CO disrupted hippocampal LTP with concomitant HO-2 and nNOS reductions. The authors surmised that these changes may be related to the memory deficits seen in animals exposed to CO in utero.
Middendorff et al. (2000)	Human Adult males aged 65-75. Testicular tissue from orchietomy			Zn protoporphyrin (ZnPP) and Hb both significantly reduced seminiferous tubular cGMP generation, suggesting a role for CO in human testicular tissue.
Naik and Walker (2003)	Rat (Sprague-Dawley) Male		210 $\mu$ L of CO/100 mL of physiological saline solution	Endogenous CO mediated vasorelaxation involved cGMP-independent activation of vascular smooth muscle large-conductance $Ca^{2+}$ -activated $K^{+}$ channels. However exogenous CO vasodilation was cGMP dependent.
Ndisang et al. (2004)				Review of CO and hypertension. CO is a vasorelaxant due to activation of the big conductance calcium-activated potassium channels and soluble guanylate cyclase/cGMP pathway. Developmental stage and tissue type will determine which of these pathways plays more of a role in vasorelaxation.
Neggens and Singh (2006)	Mouse (CD-1)	GD7-GD18	500 ppm	Developmental toxicity of CO was attenuated by protein supplementation, i.e., protein supplemented animals (27%) showed a significantly lower incidence of fetal mortality versus 8% and 16% protein groups. Further, dietary restriction of both protein and zinc with CO-exposure to CO during gestation increased the incidence of pup mortality and malformations including gastroschisis. Zinc supplementation to protein deficient diet in CO-exposed mice decreased fetal mortality and malformation.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Newby et al. (2005)	Human (placental cells in culture)			Term human placental cells were grown in cell culture under basal and hypoxic conditions to explore changes in HO expression. HO-1 was unchanged in cytotrophoblasts under hypoxia, but HO-1 was significantly decreased in hypoxic syncytiotrophoblasts. HO-2 was unchanged in either cell type with hypoxia. These cell culture data can give insight into what cell types might be responsive to hypoxia through the HO/CO system in the human placenta.
Odrich et al. (1998)	Guinea pig			Immunohistochemical localization of HO in guinea pig placenta showed that HO-1 staining was highest near term (PND62) and lesser at term or earlier in pregnancy. HO-1 was localized in the adventitial layer of fetal blood vessels.
Ozawa et al. (2002)	Rat (Wistar) Adult male			The role of HO-1 in spermatogenesis was explored. CdCl <sub>2</sub> induced testicular HO-1 and reduced HO-2 protein in rats. Pretreatment with ZnPPiX attenuated CdCl <sub>2</sub> -dependent apoptosis. Leydig cells use HO-1 derived CO to trigger apoptosis of pre-meiotic germ cells and modulate spermatogenesis under CdCl <sub>2</sub> dependent oxidative stress.
Patel et al. (2003)	Rat (Sprague Dawley) Male 262 ± 30 g Isolated hearts	30 min	Buffer saturated with 0.01 and 0.05% CO	The ventricular glutathione content, both reduced and oxidized, decreased by 76% and 84% 90 min post-exposure to 0.01% and 0.05% CO, respectively. Treatment with antioxidants partially blocked the decreases in glutathione. Increased creatine kinase activity was observed in heart perfusate during and after treatment.
Penney et al. (1983)	Rat (strain not reported)	GD17 - GD22	157, 166 or 200 ppm	In utero CO exposure induced decreased fetal body weight, decreased placental weight, increased wet heart weight at birth, and altered cardiac enzymes at birth.
Piantadosi (2002)				Reviews the biochemical activities of CO, including various heme protein binding. The review stresses the importance of the CO/O <sub>2</sub> ratio in determining the physiological effects of CO.
Piantadosi (2008)				Reviews the physiologic responses to exogenous and endogenous CO and biochemical effects including the binding to heme proteins, the generation of reactive O <sub>2</sub> species and activation related signaling pathways.
Piantadosi et al. (2006)	Rat (Sprague Dawley) Adult male	1, 3, or 7 days	50 ppm or HH	COHb produced COHb levels of 4-5% (controls approximately 1%) and liver CO concentration of 30-40 pmol/mg wet weight (controls approximately 10 pmol/mg wet weight). Both CO and HH led to increased expression of hypoxia-sensitive proteins HO-1 and HIF-1 $\alpha$ and mitochondrial antioxidant protein SOD-2. CO caused a greater change in mitochondrial GSH/GSSG than HH. Only CO increased mitochondrial 3-nitrotyrosine and protein mixed disulfides. Mitochondria isolated from CO-exposed rats, but not from HH-exposed rats, showed an increase in the calcium sensitivity of the mitochondrial permeability transition (MPT). Exposure to CO or HH resulted in a loss of the ability of adenine nucleotides to protect mitochondria from MPT. This effect was restored in the presence of a strong reductant. The authors conclude that CO causes mitochondrial pore stress independently of its hypoxic effects
Prigge and Hochrainer (1977)	Rat (Wistar, SPF)	GD0-GD20	60, 100, 250, 500 ppm	Fetuses were collected by C-section after 21-days exposure. Significant increases in fetal heart weight were seen in fetuses exposed to CO in all dose groups. Fetal body weight was significantly decreased (NOAEL 125 ppm CO).
Raub and Benignus (2002)				Reviews the physiology of CO and the effects on the nervous system. It is estimated that COHb would have to rise to 15-20% before a 10% reduction in any behavioral or visual measurement could be observed.
Richardson et al. (2002)	Human Male		20% COHb	20% COHb did not influence O <sub>2</sub> Mb binding indicated by unaltered deoxy-myoglobin signal. Resting skeletal muscle metabolic rate was unaffected by 20% COHb. VO <sub>2</sub> max was decreased. No decrement in intracellular PO <sub>2</sub> was found. 20% COHb altered exercising bioenergetics, pH, PCr, and ATP levels.
Ryter et al. (2006)				Reviews the basic science of exogenous and endogenous CO including HO-1 regulation. It also reviews some therapeutic applications for CO.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Sartiani et al. (2004)	Rat (Wistar)	In utero inhalation exposure	150 ppm	At 4 weeks of age, the action potential duration APD of isolated cardiac myocytes from CO-exposed animals failed to shorten or mature as did the APD of control animals. Further, the two ion conduction channels $I_{to}$ (transient outward current, $K^+$ -mediated) and $I_{Ca,L}$ (L-type $Ca^{2+}$ current), which largely control the rat APD, were significantly different from control animals after CO exposure at 4 weeks of age. All of these CO-dependent changes were no longer different from controls at 8 weeks of age, showing a delayed maturation.
Schwetz et al. (1979)	Mouse (CF-1) Rabbit (New Zealand)	7 or 24-h/day GD6-GD15 (Mice) GD6-GD18 (Rabbits)	250 ppm	In mice there was a significant increase in number of skeletal abnormalities in CO-exposed mice. Decreased birth weight in mice exposed to 24 h/day CO versus control. Increased birth weight in mice exposed to 7 h/day CO versus controls. No similar effects were seen in rabbits.
Singh et al. (1992)	Mouse (CD-1)	GD8-GD18	65, 125, or 250 ppm	CO exposure concomitant with a low protein diet exacerbated the percent of skeletal malformations in offspring. The percent of dead, resorbed, or grossly malformed fetuses was directly related to CO concentration and inversely related to maternal dietary protein levels. CO and maternal dietary protein restriction have a synergistic effect on offspring survival and an additive effect on malformations.
Singh (2006)	Mouse (CD-1)	6 h/day during the first 2 weeks of pregnancy	65 or 125 ppm	Modulating dam protein intake during in utero CO exposure altered pup mortality.
Sitdikova et al. (2007)	Frog neuro-muscular junctions	20 min	96 $\mu$ M	CO induced acetylcholine release, without effects on the pre-synaptic action potential or functional properties of post-synaptic receptors in frog neuro-muscular preparations.
Song et al. (2002)	Human Primary human aortic smooth muscle cells	0-48 h	10-250 ppm	CO inhibited SMC proliferation at concentrations from 50-500 ppm. The cell cycle arrest occurred at the G0/G1 phase of the cell cycle. CO increased expression of the cell cycle inhibitor p21 <sup>Cip1</sup> at 1 h and decreased expression of cyclin D1 over 24-48 h. The antiproliferative actions of CO were found to be independent of sGC, but instead exerted through the inhibition of ERK MAPK activation since 15 min exposure to 250 ppm CO blocked serum-mediated ERK phosphorylation.
Sorhaug et al. (2006)	Rat (Wistar) Female 169 $\pm$ 4.5 g	20 h/day, x 5 days/week, x 72 weeks	200 ppm	COHb was 14.7% in CO-exposed animals and 0.3% in controls. Total Hb was also increased in following CO exposure. CO caused no changes in lung morphology or pulmonary hypertension. No atherosclerotic lesions were found in aorta or femoral artery. Weight increases of 20% and 14% were observed in the right ventricle and left ventricle plus septum, respectively, indicative of ventricular hypertrophy following chronic CO exposure.
Stockard-Sullivan et al. (2003)	Rat (Sprague-Dawley)	22 h/day, PND6-PND22	12, 25, 50, or 100 ppm	Using functional OAE testing and ABR showed that with perinatal CO exposure (50 and 100 ppm CO), there were significant decrements in OAE in CO-exposed animals. ABR showed no functional deficits with CO exposure. Using another otoacoustic test revealed significant attenuation of the AP of the 8th cranial nerve with CO exposure (12, 25, and 50 ppm CO) versus controls at PND22.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Suliman et al. (2007)	Mouse (C57BL/6 wild-type and eNOS deficient) Male Rat (embryonic cardiomyocytes H9c2 cells)	1 h	50-1,250 ppm Or HH Or 100 mM dichloromethane	<p>1-h exposure of mice to 1,250 ppm CO increased cardiac mitochondrial content of all 5 respiratory complexes 24 h later. The volume density of interfibrillar mitochondria was increased by 30% after 24 h demonstrating that CO caused cardiac mitochondrial biogenesis. The CO concentration in heart increased from 9 pmol/mg to 50-150 pmol/mg in mice exposed to 50-1,250 ppm CO for 1 h. These levels declined to baseline by 6 h. Peroxisome proliferator-activated receptor gamma coactivator 1 alpha (PGC-1<math>\alpha</math>) expression was increased 6 h following exposure to 50-1,250 ppm CO. Expression of DNA polymerase and mitochondrial transcription factor A (TFAM) was increased 6 and 24 h after exposure, while mitochondrial DNA was increased 2-3 fold 24 h after exposure. CO activated gene expression of these proteins involved in cardiac mitochondrial biogenesis beginning at 2 h post exposure for PGC-1<math>\alpha</math>, nuclear respiratory factors 1 and 2 (NRF-1 and -2) and at 6 h postexposure for TFAM. These effects were independent of NOS and not seen with HH. CO exposure resulted in phosphorylation of p38 MAPK and Akt at 2 and 6 h post-exposure to 1,250 ppm CO for 1 h. Inhibition of p38 activation failed to inhibit the CO-mediated increase in cardiac mitochondrial biogenesis.</p> <p>In cell culture experiments, CO derived from dichloromethane metabolism resulted in increased cGMP, protein levels of SOD2, TFAM, NRF-1, NRF-2, PGC-1, mitochondrial ROS, Akt phosphorylation, and mitochondrial DNA. Inhibition of GC or PI3K/Akt, but not p38, blocked the responses to CO. A role for mitochondrial H<sub>2</sub>O<sub>2</sub> in Akt regulation was demonstrated. Mitochondrial H<sub>2</sub>O<sub>2</sub> and the PI3K/Akt pathway were important mediators of TFAM expression.</p> <p>The authors concluded that CO exposure increased mitochondrial ROS which promoted mitochondrial biogenesis in the heart.</p>
Sun et al. (2001)	Mouse Neuronal cultures prepared from the cerebral hemispheres of 16-day Charles River CD1 mouse embryos			Nb expression was increased by neuronal hypoxia in vitro and focal cerebral ischemia in vivo. Inhibiting Nb reduced neuronal survival after hypoxia whereas Nb overexpression enhanced neuronal survival.
Tattoli et al. (1999)	Rat (Wistar) Male and Pregnant female	PND1-PND10	75 and 150 ppm	Cognitive function was assessed in rats after postnatal CO exposure at 3 and 18 months of age. Postnatal CO exposure did not affect the acquisition and reacquisition of an active avoidance task. This is different from previous findings by the same laboratory indicating that in utero exposure to CO (75 and 150 ppm) induced long-lasting learning and memory deficits.
Telfer et al. (2001)	Human Myometrium tissue obtained from gravid [pre-term (25-34 weeks gestation), term not in labor or term in labor] and non-gravid women			cGMP was monitored in various myometrial tissues. cGMP was significantly higher than that from nonpregnant tissue and decreased at term, especially in tissue from laboring women.
Teran et al. (2005)	Rat (Dahl/Rapp salt-sensitive rats) Male		100 $\mu$ M	A high salt diet for 1-4 weeks resulted in increased aortic HO-1 protein expression, an increase in mean arterial pressure and time-dependent inhibition of flow- and acetylcholine-mediated vasodilation in isolated gracilis muscle arterioles. A smaller degree of inhibition of acetylcholine-mediated vasodilation was observed with a low salt diet for 1-4 weeks. Pretreatment with a HO inhibitor restored these responses but this effect was reversed in the presence of exogenous CO. Mean arterial pressure was decreased in intact animals fed a high salt diet for 4 weeks and then treated with a HO inhibitor. The authors concluded that the HO-derived CO contributed to the development of hypertension and the impairment of endothelium-dependent vasodilator responses in this model.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Thom et al. (1994)	Rat (Wistar) Male Isolated blood cells	1 h OR >1 h 30 min	1,000 ppm OR 1,000-3,000 and higher ppm 0.5 mL of pure CO	CO poisoning inhibited B <sub>2</sub> integrin-dependent PMN adherence in heparinized blood obtained from rats immediately after exposure. Adherence was restored when platelet number was decreased. Adherence was also decreased when PMN from control animals were incubated with platelets from poisoned animals. Adherence of activated PMN was reduced in the presence of SOD and enhanced by NOS inhibition. Platelet production of NO* was significantly greater, while platelet NOS activity was significantly inhibited after poisoning.  When whole blood or platelet-rich plasma was incubated with CO, PMN adherence was inhibited.  The authors concluded that PMN B <sub>2</sub> integrin activity was inhibited by CO-dependent release of NO* from the platelets into the blood.
Thom and Ischiropoulos (1997)	Rat (Wistar) Male 200-290 g  Platelet-rich plasma from rats was used as the source of platelets  Bovine pulmonary artery endothelial cells	1 h 30 min or 2 h 1 h	20-1,000 ppm 10-20 ppm 10-100 ppm	Platelets isolated from rats exposed to 20-1,000 ppm CO for 1-h released NO* in a dose-dependent manner. COHb levels were 0.7% in controls, and 3.2%, 7.8% and 51.0% in 20, 100 and 1,000 ppm exposure groups respectively.  Isolated platelets released NO* when incubated for 30 min with 20-100 ppm CO. NOS activity was not enhanced by 100 ppm CO. Platelets released NO* in response to 10-100 ppm CO after 30 min pretreatment with a NOS inhibitor, suggesting that CO displaces NO* from heme-binding sites. Longer incubations (2 h) with the NOS inhibitor led to a diminished response to 100 ppm CO. There appears to be a discrepancy in the results depending on how NO* was measured (electrode versus Greiss reaction).  Endothelial cells released NO* in response to 20-100 ppm CO. NOS inhibition blocked the response to 100 ppm CO. CO was found not to affect arginine transport or NOS activity in endothelial cells. Exposure to 40-100 ppm CO resulted in the release of short-lived oxidants. This response was blocked by NOS inhibition. Lysates from cells exposed to 50 and 100 ppm CO had increased nitrotyrosine content. This response was blocked by NOS inhibition. Cellular reduced sulfhydryls were not decreased by 100 ppm CO. Dihydrorhodamine 123 oxidation, a measure of peroxynitrite formation, was increased by exposure to 100 ppm CO. This effect was blocked by NOS inhibition. Cytotoxicity of CO was evaluated by the release of <sup>51</sup> chromium. Cytotoxicity was evident 4 h following a 2-h incubation with 100 ppm CO, but not immediately after exposure. This response was not blocked by NOS inhibition, although NOS inhibition had protective effects under conditions of continuous CO exposure of 4 h. Exposure to 20 and 100 ppm CO for 2 h led to the loss of membrane integrity, measured by ethidium homodimer-1 staining, 18 h later.  Results demonstrate that 10-20 ppm CO released NO* from platelets and endothelial cells in vitro. Platelets from rats that inhaled 20 ppm CO also released NO* in vitro. The authors suggested that CO-mediated NO* release from platelets and endothelial cells resulted from disrupted intracellular scavenging for NO*. They also suggested that peroxynitrite may have been generated in response to CO.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Thom et al. (1997)	Bovine pulmonary artery endothelial cells	30 min-4 h	10-100 ppm (11-110 nM)	<p>1-h exposure to 111-110 nM CO led to a dose-dependent increase in NO<sup>*</sup> release, as measured by nitrite+nitrate. Significance was achieved at 22 nM (corresponding to an interstitial partial pressure of 20 ppm and a blood COHb level of 7%). NOS inhibition blocked the response to 110 nM CO. A dose-dependent increase in cellular nitrotyrosine was also observed following a 2-h exposure to CO, with significance achieved at 55 nM CO. NOS inhibition blocked the response to 110 nM. CO exposure failed to decrease the concentration of reduced sulfhydryls, but did result in the extracellular release of a short-lived oxidant species which was blocked by NOS inhibition. Dihydrorhodamine oxidation, a measure of peroxynitrite formation, occurred in response to 110 nM CO, an effect which blocked by NOS inhibition. Cytotoxicity of CO was evaluated by the release of <sup>51</sup>chromium. Cytotoxicity was evident 4 h following a 2-h incubation with 110 nM CO, but not immediately after exposure. This response was not blocked by NOS inhibition, although NOS inhibition had protective effects under conditions of continuous CO exposure of 4 h. Exposure to 110 nM CO for 2 h led to the loss of membrane integrity, measured by ethidium homodimer-1 staining, 18 h later. This response was blocked by NOS inhibition. Exposure to 110 nM CO had no effect on O<sub>2</sub> consumption, production of intracellular H<sub>2</sub>O<sub>2</sub> or cellular redox activity. Exposure to 110 nM did not alter arginine transport or NOS activity. NO<sup>*</sup> release from cells which had been pre-treated with a NOS inhibitor and then exposed briefly to 5% CO was measured using a NO-selective electrode suggesting that CO competed with intracellular binding sites of NO<sup>*</sup>.</p> <p>The authors concluded that endothelial cells release NO<sup>*</sup> and NO<sup>*</sup>-derived oxidants in response to CO. A delayed cell death occurred following exposures to 22 nM and higher concentrations of CO.</p>
Thom, et al. (1999b)	Rat (Wistar) Male 200-290 g Some rats fed a high cholesterol diet	1 h	50-1,000 ppm	<p>Nitrotyrosine immunoreactivity was found in aortic intima in rats exposed to CO for 1 h but not in controls. Nitrotyrosine content was quantitated and found to be increased in a dose-dependent manner following 1-h exposure to 50-1,000 ppm CO. The effect was significant at 50 ppm but the COHb content measured immediately after exposure was not different than controls. Platelet and neutrophil depletion did not alter nitrotyrosine content following CO exposure. Leukocyte adherence to the aorta occurred 18 h, but not immediately, after a 1-h exposure to 100 ppm CO. This effect was blocked by NOS inhibition. The influx of albumin from the microvasculature into skeletal muscle increased during the 3 h after exposure to 100 ppm CO but was not seen 18 h later. This effect was blocked by NOS inhibition.</p> <p>Rats fed a high cholesterol diet and exposed to 100 ppm CO for 1 h had increased aortic nitrotyrosine content which was not different than that in CO-exposed rats fed the standard diet. However, rats on the high cholesterol diet had a 6-fold increase in LDL oxidation immediately after 1-h exposure to 100 ppm CO. This effect was not blocked by NOS inhibition.</p> <p>The authors concluded that CO can alter vascular status by several mechanisms linked to NO<sup>*</sup>-derived oxidants.</p>

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Thom et al. (1999a)	Rat (Wistar) Male 200-290 g	1 h	50-1,000 ppm	Leakage of albumin into lung parenchyma occurred 18 h, but not at earlier time points, after rats were exposed to 100 ppm CO for 1 h. This response was also observed using 50 and 1,000 ppm but not 20 ppm CO. Leakage resolved by 48 h. Furthermore no leakage occurred when rats which were exposed to 100 ppm CO were pretreated with a NOS inhibitor. COHb levels were 0.9% in controls and 4.8%, 10.6% and 53.7% following 1-h exposure to 50, 100 and 1,000 ppm CO, respectively. Elevated free NO <sup>*</sup> , determined by EPR, was observed in lungs of rats exposed to 100 ppm CO for 1 h. This effect was blocked when rats were pretreated with a NOS inhibitor. Lung H <sub>2</sub> O <sub>2</sub> was elevated by exposure to 100 ppm CO for 1 h and this effect was blocked when rats were pretreated with a NOS inhibitor. Elevated nitrotyrosine content was observed in lung homogenates 2-4 h following 1-h exposure of rats to 100 ppm CO. This effect was also blocked by pretreatment with a NOS inhibitor. No leukocyte sequestration was observed in lungs 18 h following exposure to 100 ppm CO. CO-induced lung leak was not affected by neutrophil depletion. The authors concluded that CO causes lung vascular injury which is dependent on NO <sup>*</sup> .
Thom et al. (2000)	Bovine pulmonary artery endothelial cells	40 min-2 h	11-110 nM (10-100 ppm)	Increased uptake of ethidium homodimer-1, a measure of decreased membrane integrity and cell death, was observed in endothelial cells 18 h after exposure to 110 nM for 60-120 min. Exposures of 20-40 nM were ineffective in this regard. Ethidium uptake was also increased by 2-h exposure to 88 nM CO. Preincubation for 2 h with an inhibitor of eNOS, an antioxidant, and an inhibitor of peroxynitrite reactions blocked the CO-mediated cell death. Morphological changes in cells were observed 2 h following a 2-h exposure to 110 nM CO. Cell death induced by 110 nM CO was also blocked by inhibition of protein synthesis and inhibition of caspase-1 but of caspase-3. Caspase-1 activity was increased following 2-h exposure to 110 nM CO; this effect was blocked by inhibiting eNOS. Pre-exposure of cells to 11 nM CO for 40 min followed by a 3-h incubation period resulted in an increased level of MnSOD and protection against cell death 18 h following a 2-h exposure to 110 nM CO. The authors concluded that exposure to 11 nM CO led to an adaptive response which protected cells from injury and apoptosis resulting from NO <sup>*</sup> -derived oxidants.
Thom et al. (2001a)	Rat	Until lost consciousness	1,000-3,000 ppm	Neutrophils sequestration was observed in the brain vessels of rats exposed to high dose CO. CO also led to increased nitrotyrosine formation in the brain vessels. These events were blocked by pretreatment with a peroxynitrite scavenger or a PAF receptor antagonist.
Thom et al. (2006)	Human Rat (Wistar), male Mouse (C57B6J, MPO-deficient) Blood samples and brain tissue	1 h	Humans: acute CO poisoning Rats and mice: 1,000-3,000 ppm	In humans, COHb was 20-30.5%. Increased cell surface expression of CD18 and PAC1 was observed in neutrophils from people with CO poisoning. Increased surface-bound myeloperoxidase (MPO, indicative of neutrophil degranulation), increased plasma MPO and more numerous platelet-neutrophil aggregates were also observed.  Similar changes were observed in blood of CO-poisoned rats. Platelet depletion, inhibition of NOS and inhibition of platelet integrin-dependent adhesion blocked these responses. Brains from poisoned rats had significant elevations in MPO which could reflect either an increase number of neutrophils or an increase in neutrophil degranulation. Perivascular MPO and nitrotyrosine were CO-localized in brain. CO poisoning also resulted in altered brain myelin basic protein.  Similar changes were observed in blood of CO-poisoned mice. MPO deficiency blocked the CO-mediated alteration in brain myelin basic protein.  The authors concluded that exposure to CO triggers intravascular interactions between platelets and neutrophils that lead to neutrophil degranulation in experimental animals and people with CO poisoning.



Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Thorup et al. (1999)	Rat (Sprague Dawley) Males 200-250 g		0.01-10 $\mu$ M	<p>Perfusion of Isolated rat renal resistance arteries with CO-containing buffer (0.001-10 <math>\mu</math>M) led to the biphasic release of NO<sup>*</sup>, peaking at 100 nM and declining to undetectable responses at 10 <math>\mu</math>M. Sequential pulses of 100 nM resulted in a blunting of NO<sup>*</sup> release with consecutive pulses, consistent with a depletion of intracellular NO<sup>*</sup> stores. NO<sup>*</sup> release was dependent on arginine concentrations and was inhibited by pretreatment with a NOS inhibitor. Perfusion with 100 nM CO blocked carbachol-dependent NO<sup>*</sup> release from vessels.</p> <p>Rats were treated with a HO-1 inducer and renal resistance arteries were isolated 12 h later. Carbachol-induced NO<sup>*</sup> release was smaller in the HO-1 induced rats compared with controls suggesting that endogenous CO has a similar effect as 100 nM exogenous CO. This effect was reversed in the presence of excess arginine.</p> <p>Vasodilation was measured in blood-perfused afferent arterioles perfused with CO in solution. A biphasic vasodilatory response was observed as well as a blunted muscarinic vasorelaxation.</p> <p>CO (0.1-10 <math>\mu</math>M) suppressed the release of NO<sup>*</sup> from purified recombinant eNOS in solution.</p> <p>The authors concluded that low levels of CO may release NO<sup>*</sup> and elicit vasorelaxation and modulate basal vascular tone while higher levels of CO may inhibit eNOS and NO<sup>*</sup> generation.</p>
Tolcos et al. (2000b)	Guinea pig	10 h/day over the last 60% of gestation	200 ppm	<p>Fetal and maternal COHb were 13% and 8.5% respectively. Neurotransmitter systems were affected after CO exposure. The catecholaminergic system of the brainstem displayed significant decreases in immunoreactivity for tyrosine hydroxylase (TH), which was likely due to decreased cell number in specific medullar regions. The cholinergic system was also affected by prenatal CO exposure with significant increases in ChAT immunoreactivity of the medulla and no changes in muscarinic acetylcholine receptor.</p>
Tolcos et al. (2000a)	Guinea pig	10 h/day for the last 60% of gestation	200 ppm	<p>Brains were collected at 1 and 8 weeks of age. These data showed that CO exposure in utero sensitized the brain to hyperthermia at PND4 leading to generation of necrotic lesions in the brain and changes in neurotransmitter levels.</p>
Tschugguel et al. (2001)	Human HUVEC			<p>CO was generated by primary endothelial cells from human umbilical veins and uterine arteries after exogenous 17-<math>\beta</math> estradiol administration.</p>
Vallone et al. (2004)	Mouse protein			<p>The authors present the X-ray structure of CO-bound ferrous murine Nb. When CO binds, the heme group slides deeper into the protein crevice.</p>
Vreman et al. (2000)	Human umbilical cord (artery and vein) Rat Aorta, vena cavae, liver and heart			<p>HO activity was quantified in human umbilical cord and in the rat vasculature (aorta and vena cavae). Human umbilical artery and vein HO activity were equal. The rat aorta and vena cavae produced equal amounts of HO activity (wet weight/g tissue) but generated 3x greater HO than the heart and 0.2x of the liver. HO activity in rat vasculature was 3x that of the human cord tissues. Use of the HO inhibitor CrMP effectively blocked HO activity in the rat liver and heart but was less effective at blocking HO activity in the human umbilical cord or the rat vasculature (only 50% effective). The activity of HO in the umbilical vessels may provide a role for CO in control of vasculature tone during pregnancy.</p>

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Vreman et al. (2005)	Mouse (BALB/c)	30 min	500 ppm OR  Heme arginate 30 µmol/kg body weight i.v.	<p>Following CO exposure, COHb levels were 28%. Tissue concentrations of CO were as follows with control levels in parenthesis.</p> <p>Blood: 2648 ± 400 (45) pmol/mg Heart: 100 ± 18 (6) pmol/mg Muscle: 14 ± 1 (10) pmol/mg Brain: 18 ± 4 (2) pmol/mg Kidney: 120 ± 12 (7) pmol/mg Spleen: 229 ± 55 (6) pmol/mg Liver: 115 ± 31 (5) pmol/mg Lung: 250 ± 2 (3) pmol/mg Intestine: 9 ± 7 (4) pmol/mg Testes: 6 ± 3 (2) pmol/mg</p> <p>CO concentration relative to 100% blood: Lung: 9.4%, Spleen: 8.6%, Kidney: 4.5%, Liver: 4.3%, Heart: 3.8%, Brain: 0.7%, Muscle: 0.5%, Intestine: 0.3%, Testes: 0.2%</p> <p>Injection of heme arginate resulted in a 3-fold increase in CO excretion reaching a maximum at 60 min. Animals were sacrificed at 90 min. COHb levels were 0.9%. Tissue concentrations of CO were as follows with control levels in parenthesis.</p> <p>Blood: 88 ± 10 (45) pmol/mg Heart: 14 ± 3 (6) pmol/mg Muscle: 7 ± 1 (10) pmol/mg Brain: 2 ± 0 (2) pmol/mg Kidney: 7 ± 2 (7) pmol/mg Spleen: 11 ± 1 (6) pmol/mg Liver: 8 ± 3 (5) pmol/mg Lung: 8 ± 3 (3) pmol/mg Intestine: 3 ± 1 (4) pmol/mg Testes: 2 ± 0 (2) pmol/mg</p> <p>CO concentration relative to 100% blood: Heart: 16% Spleen: 13% Lung: 9% Liver: 9% Kidney: 8% Muscle: 8% Intestine: 3% Brain: 2% Testes: 2%</p>
Weaver et al. (2007)	Human		Acute CO poisoning	<p>Mean COHb in humans with acute CO poisoning was 35%. Hyperbaric O<sub>2</sub> reduces cognitive sequelae in a randomized clinical trial of CO-poisoned patients. Risk factors for cognitive sequelae without hyperbaric O<sub>2</sub> included older age and longer CO exposures. Patients with loss of consciousness or high initial COHb levels should also be treated with hyperbaric O<sub>2</sub>.</p>
Webber et al. (2003)	Rat Strain not stated	PND8-PND22	12.5, 25, or 50 ppm	<p>Immunostating of c-Fos, a marker of neuronal activation in the nervous system was followed. C-Fos immunoreactivity in the central IC was significantly decreased in the CO-exposed animals at both PND27 and PND75-PND77 over all dose groups of CO; immunostaining of other subregions of the IC were not affected by CO. These studies show exposure to CO during development can lead to permanent changes in the auditory system of rats that persist into adulthood.</p>
Webber et al. (2005)	Rat Strain not stated	PND9-PND24	25 or 100 ppm	<p>Neurofilament loss from the spiral ganglionic neurons and somas after ARCO treatment was rescued (no detectable neurofilament loss) with low iron+CO (ARIDCO); ARID (low iron) treatment induced no change in neurofilaments. CuZn superoxide dismutase (SOD1) was significantly increased with CO exposure (ARCO) and rescued in ARIDCO animals; SOD1 was unchanged in low iron only animals (ARID). Low iron treatment or CO exposure alone led to significant decreases in c-fos positive cell numbers of the central IC, but c-fos levels were unchanged after low iron diet concomitant with CO exposure (ARIDCO).</p>

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Wellenius et al. (2004)	Rat (Sprague Dawley) 250 g Diazepam-sedated Model of acute MI induced by thermocoagulation	1 h, 12-18 h after surgery	35 ppm	CO exposure decreased ventricular premature beat frequency by 60.4% during the exposure period compared to controls. 1-h exposure to CAPs (318 µg/m <sup>3</sup> ) decreased ventricular premature beat frequency in specific subgroups. Neither CAPs nor CO had an effect on heart rate. There were no significant interactions between their effects when rats were exposed to both CO and CAPs.
Wellenius et al. (2006b)	Rat (Sprague Dawley) 250 g Diazepam sedated Model of acute MI induced by thermocoagulation	1 h, 12-18 h after surgery	35 ppm	Exposure to CO failed to increase the probability of observing supraventricular ectopic beats (SVEB). Exposure to CAPs (646 µg/m <sup>3</sup> ) for 1 h decreased the frequency of SVEB. There were no significant effects observed when rats were exposed to both CO and CAPs. Among a subset of rats with one or more SVEB at baseline, a significant decrease in number of SVEB during the exposure period was observed with either CO or CAPs exposure compared with controls.
Yoshiki et al. (2001)	Human			HO localization in human endometrium and its changes in expression over the menstrual cycle were explored in this study. HO-1 was constitutively expressed throughout the menstrual cycle and HO-2 was greater in the secretory than the proliferative phase of the menstrual cycle. HO-1 was localized to the epithelial cells and macrophages. HO-2 was found in endothelial cells and smooth muscle cells of endometrial blood vessels.
Zamudio et al. (1995)	Human			Women living at high altitude had an increased risk of adverse pregnancy outcomes versus women living at lower altitudes.
Zenclussen et al. (2006)	Mouse (CBA/J x DBA/2J)			To evaluate the role of HO-1 in spontaneous abortion, a mouse model that spontaneously undergoes abortion (CBA/J x DBA/2J mice) was used with and without HO adenovirus treatment to see if pregnancy outcome could be modulated by changing HO concentration. Pregnancy outcome was significantly better (abortion rate significantly decreased) in mice overexpressing HO due to adenovirus transfer.
Zhang et al. (2005)	Rat (pulmonary artery endothelial cells)	8-28 h	15 ppm	Exposure to 15 ppm CO during anoxia resulted in decreased phosphorylation of STAT1 and increased phosphorylation of STAT3 at 8-24 h. Similar responses were observed when 24 h anoxia was followed by a period of reoxygenation (0.5-4 h). DNA binding of STAT1 was decreased while that of STAT3 was enhanced by CO treatment during anoxia/reoxygenation. Exposure to 15 ppm during 8-24 h anoxia or 24 h anoxia followed by 0.5-4 h reoxygenation resulted in increased phosphorylation of Akt and p38 MAPK. Inhibitor studies demonstrated that activation of the PI3K pathway by CO was upstream of p38 MAPK activation during anoxia/reoxygenation. Similarly, the PI3K and p38 MAPK pathways were found to be upstream of STAT modulation. The anti-apoptotic effects of 15 ppm CO during anoxia-reoxygenation involved decreased FAS expression and decreased caspase 3 activity. These effects were dependent on activation of the PI3K, p38 MAPK and STAT3 pathways.  The authors concluded that CO blocks anoxia-reoxygenation mediated apoptosis through modulation of PI3K/Akt/p38 MAPK and STAT1 and STAT3.
Zhang et al. (2007)	Mouse			A single dose of LPS administered to pregnant mice induced up-regulation of HO-1 but not HO-2 in the mouse placenta 12-48 h post-LPS treatment. Pre-treatment of mice with the spin trap agent PBN or the TNF $\alpha$ inhibitor pentoxifylline prevented the LPS-dependent HO-1 upregulation. Thus ROS may mediate the LPS-dependent upregulation of HO-1.
Zhao et al. (2008)	Mouse (FVB)			With pregnancy, there was an increased blood volume without a concurrent increase in systemic BP; this was accomplished by a decrease in total vascular resistance, to which CO contributed as determined by using HO inhibitors.
Zhuo et al. (1993)	Rodent			Hippocampal LTP of brain sections is significantly affected by CO exposure with ZnPP IX, a HO inhibitor, blocking hippocampal LTP.

Reference	Species / Model	Exposure Duration	CO Concentration	Findings
Zuckerbraun et al. (2007)	Macrophages RAW 264.7 THP-1 cells, wild-type and respiration-deficient	10 min - 24 h	50 – 500 ppm	Exposure of RAW macrophages to 250 ppm CO for 10-60 min increased ROS generation, measured as dichlorofluorescein (DCF) fluorescence. ROS generation at 1 h was dose-dependent with significant effects observed at 50, 250 and 500 ppm CO. This response was not blocked with a NOS inhibitor. A 1-h exposure to 250 ppm resulted in decreased intracellular glutathione levels. CO treatment was found to block TNF $\alpha$ production and to enhance p38 MAPK phosphorylation in LPS-stimulated cells. These effects were diminished by pretreatment with antioxidants. The source of CO-derived oxidants was determined to be mitochondrial since respiration-deficient THP-1 macrophages, unlike wild-type cells, failed to generate ROS in response to 250 ppm CO. Furthermore, treatment of RAW cells with the mitochondrial complex III inhibitor antimycin C, blocked ROS generation in response to 250 ppm CO. Exposure of RAW cells to 250 ppm CO for 1 h inhibited cytochrome <i>c</i> oxidase activity by 50%. Exposure to 250 ppm CO for 6 h had no effect on cellular ATP levels or mitochondrial membrane potential. Antimycin C treatment was found to reverse the effects of CO on LPS-mediated responses (TNF $\alpha$ and p38 MAPK), suggesting that mitochondrial-derived ROS mediated the effects of CO. The authors concluded that CO increased the generation of mitochondrial-derived ROS.

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