



# **TOXICOLOGICAL REVIEW**

## **OF**

# **LIBBY AMPHIBOLE ASBESTOS**

**In Support of Summary Information on the  
Integrated Risk Information System (IRIS)**

*August 2014*

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*(Note: This document is an assessment of the noncancer and cancer health effects associated with the inhalation route of exposure only)*

Integrated Risk Information System  
National Center for Environmental Assessment  
Office of Research and Development  
U.S. Environmental Protection Agency  
Washington, DC

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# CONTENTS—TOXICOLOGICAL REVIEW OF LIBBY AMPHIBOLE ASBESTOS

LIST OF TABLES .....	ix
LIST OF FIGURES .....	xv
LIST OF ABBREVIATIONS AND ACRONYMS .....	xvii
FOREWORD .....	xx
AUTHORS, CONTRIBUTORS, AND REVIEWERS .....	xxi
1. INTRODUCTION .....	1-1
1.1. RELATED ASSESSMENTS.....	1-2
1.1.1. Integrated Risk Information System (IRIS) Assessment for Asbestos (U.S. EPA, 1988a) .....	1-2
1.1.2. EPA Health Assessment for Vermiculite (U.S. EPA, 1991b) .....	1-4
1.2. LIBBY AMPHIBOLE ASBESTOS-SPECIFIC HUMAN HEALTH ASSESSMENT.....	1-4
2. LIBBY AMPHIBOLE ASBESTOS: GEOLOGY AND EXPOSURE POTENTIAL .....	2-1
2.1. INTRODUCTION .....	2-1
2.2. GEOLOGY AND MINERALOGY OF AMPHIBOLES .....	2-3
2.2.1. Overview .....	2-3
2.2.2. Mineralogy of Amphibole Asbestos and Related Amphibole Minerals .....	2-3
2.2.3. Morphology of Amphibole Minerals .....	2-6
2.3. METHODS FOR ANALYSIS OF ASBESTOS .....	2-9
2.3.1. Methods for Air Samples .....	2-9
2.3.2. Methods for Solid Materials .....	2-10
2.4. CHARACTERISTICS OF LIBBY AMPHIBOLE ASBESTOS.....	2-10
2.4.1. Mineralogy of Libby Amphibole Asbestos.....	2-11
2.4.2. Morphology of Libby Amphibole Asbestos .....	2-16
2.5. HUMAN EXPOSURE POTENTIAL.....	2-20
2.5.1. Exposures Pathways in the Libby Community .....	2-20
2.5.2. Exposure Pathways in Communities with Vermiculite Expansion and Processing Plants .....	2-21
2.5.3. Exposure Pathways in Other Communities .....	2-23
3. FIBER TOXICOKINETICS.....	3-1
3.1. DEPOSITION OF FIBERS IN THE RESPIRATORY TRACT .....	3-2
3.2. CLEARANCE MECHANISMS.....	3-8
3.2.1. Physical and Physicochemical Clearance of Fibers.....	3-9
3.2.1.1. Mechanical Reflex Mechanisms .....	3-9
3.2.1.2. Mucociliary Clearance .....	3-9
3.2.1.3. Phagocytosis by Alveolar Macrophages .....	3-10
3.2.1.4. Epithelial Transcytosis.....	3-11
3.2.1.5. Translocation.....	3-11
3.2.1.6. Dissolution and Fiber Breakage.....	3-13
3.3. DETERMINANTS OF TOXICITY .....	3-13
3.3.1. Dosimetry and Biopersistence .....	3-13
3.3.2. Biological Response Mechanisms .....	3-14

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## CONTENTS (continued)

3.3.2.1.	Inflammation and Reactive Oxygen Species (ROS) Production.....	3-16
3.3.2.2.	Genotoxicity.....	3-16
3.3.2.3.	Carcinogenicity .....	3-17
3.4.	FIBER DOSIMETRY MODELS .....	3-18
3.5.	SUMMARY .....	3-18
4.	HAZARD IDENTIFICATION OF LIBBY AMPHIBOLE ASBESTOS .....	4-1
4.1.	STUDIES IN HUMANS—EPIDEMIOLOGY .....	4-1
4.1.1.	Overview of Primary Studies.....	4-3
4.1.1.1.	Studies of Libby, MT Vermiculite Mining and Milling Operations Workers .....	4-3
4.1.1.2.	Studies of O.M. Scott, Marysville, OH Plant Workers.....	4-9
4.1.1.3.	Community-Based Studies Around Libby, MT Conducted by Agency for Toxic Substances and Disease Registry (ATSDR)....	4-12
4.1.2.	Respiratory Effects, Noncancer .....	4-14
4.1.2.1.	Asbestosis and Other Nonmalignant Respiratory Disease Mortality .....	4-14
4.1.2.2.	Pathological Alterations of the Parenchyma and Pleura, Pulmonary Function, and Respiratory Symptoms .....	4-17
4.1.3.	Other Effects, Noncancer .....	4-36
4.1.3.1.	Cardiovascular Disease .....	4-36
4.1.3.2.	Autoimmune Disease and Autoantibodies .....	4-37
4.1.4.	Cancer Effects .....	4-40
4.1.4.1.	Lung Cancer .....	4-40
4.1.4.2.	Mesothelioma.....	4-45
4.1.4.3.	Other Cancers.....	4-49
4.1.4.4.	Summary of Cancer Mortality Risk in Populations Exposed to Libby Amphibole Asbestos .....	4-49
4.1.5.	Comparison With Other Asbestos Studies—Environmental Exposure Settings.....	4-50
4.2.	SUBCHRONIC- AND CHONIC-DURATION STUDIES AND CANCER BIOASSAYS IN ANIMALS—ORAL, INHALATION, AND OTHER ROUTES OF EXPOSURE .....	4-52
4.2.1.	Inhalation .....	4-61
4.2.2.	Intratracheal Instillation Studies .....	4-62
4.2.3.	Injection/Implantation Studies .....	4-64
4.2.4.	Oral .....	4-65
4.2.5.	Summary of Animal Studies for Libby Amphibole and Tremolite Asbestos .....	4-66
4.3.	OTHER DURATION- OR ENDPOINT-SPECIFIC STUDIES.....	4-68
4.3.1.	Immunological .....	4-68
4.4.	MECHANISTIC DATA AND OTHER STUDIES IN SUPPORT OF THE MODE OF ACTION .....	4-70
4.4.1.	Inflammation and Immune Function .....	4-77
4.4.2.	Genotoxicity.....	4-80

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## CONTENTS (continued)

4.4.3.	Cytotoxicity and Cellular Proliferation.....	4-82
4.5.	SYNTHESIS OF MAJOR NONCANCER EFFECTS.....	4-83
4.5.1.	Pulmonary Effects.....	4-84
4.5.1.1.	Pulmonary Fibrosis (Asbestosis) .....	4-84
4.5.1.2.	Other Nonmalignant Respiratory Diseases .....	4-85
4.5.2.	Pleural Effects .....	4-85
4.5.3.	Other Noncancer Health Effects (Cardiovascular Toxicity, Autoimmune Effects).....	4-86
4.5.4.	Libby Amphibole Asbestos Summary of Noncancer Health Effects .....	4-87
4.5.5.	Mode-of-Action Information (Noncancer) .....	4-87
4.6.	EVALUATION OF CARCINOGENICITY.....	4-89
4.6.1.	Summary of Overall Weight of Evidence.....	4-89
4.6.1.1.	Synthesis of Human, Animal, and Other Supporting Evidence....	4-90
4.6.2.	Mode-of-Action Information (Cancer) .....	4-91
4.6.2.1.	Description of the Mode-of-Action Information .....	4-91
4.6.2.2.	Evidence Supporting a Mutagenic Mode of Action .....	4-92
4.6.2.3.	Evidence Supporting Mechanisms of Action of Chronic Inflammation, Cytotoxicity, and Cellular Proliferation.....	4-93
4.6.2.4.	Conclusions About the Hypothesized Modes of Action.....	4-96
4.6.2.5.	Application of the Age-Dependent Adjustment Factors.....	4-100
4.7.	SUSCEPTIBLE POPULATIONS .....	4-101
4.7.1.	Influence of Different Life Stages on Susceptibility .....	4-101
4.7.1.1.	Life-Stage Susceptibility.....	4-102
4.7.2.	Influence of Gender on Susceptibility .....	4-106
4.7.3.	Influence of Race or Ethnicity on Susceptibility .....	4-106
4.7.4.	Influence of Genetic Polymorphisms on Susceptibility.....	4-107
4.7.5.	Influence of Health Status on Susceptibility.....	4-108
4.7.6.	Influence of Lifestyle Factors on Susceptibility .....	4-109
4.7.7.	Susceptible Populations Summary.....	4-109
5.	EXPOSURE-RESPONSE ASSESSMENT.....	5-1
5.1.	ORAL REFERENCE DOSE (RfD).....	5-1
5.2.	INHALATION REFERENCE CONCENTRATION (RfC) .....	5-1
5.2.1.	Choice of Principal Study .....	5-3
5.2.1.1.	Candidate Studies.....	5-3
5.2.1.2.	Evaluation of Candidate Studies and Selection of Principal Study .....	5-7
5.2.2.	Methods of Analysis .....	5-10
5.2.2.1.	Exposure Assessment.....	5-10
5.2.2.2.	Data Sets for Modeling Analyses .....	5-11
5.2.2.3.	Selection of Critical Effect.....	5-14
5.2.2.4.	Selection of Explanatory Variables to Include in the Modeling.....	5-19
5.2.2.5.	Selection of the Benchmark Response.....	5-21
5.2.2.6.	Exposure-Response Modeling .....	5-22

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## CONTENTS (continued)

5.2.3.	Derivation of a Reference Concentration (RfC) for the Critical Effect of Localized Pleural Thickening (LPT) in the Marysville Workers Who Underwent Health Evaluations in 2002–2005 and Were Hired in 1972 or Later—Including Application of Uncertainty Factors (UFs) .....	5-42
5.2.3.1.	Derivation of a Reference Concentration (RfC) for the Alternative Endpoint of Any Pleural Thickening (APT) in the Marysville Workers Who Underwent Health Evaluations in 2002–2005 and Were Hired in 1972 or Later.....	5-44
5.2.3.2.	Derivation of a Reference Concentration (RfC) for the Alternative Endpoint of Any Radiographic Change (ARC) in the Marysville Workers Who Underwent Health Evaluations in 2002–2005 and Were Hired in 1972 or Later.....	5-45
5.2.4.	Derivation of a Reference Concentration (RfC) for Localized Pleural Thickening (LPT) in the Marysville Workers Who Underwent Health Evaluations in 2002–2005 and Were Hired in 1972 or Later Based on the Cumulative Exposure Model .....	5-45
5.2.5.	Derivation of a Reference Concentration (RfC) for the Alternative Endpoint of Any Pleural Thickening (APT) in the Marysville Cohort with Combined X-Ray Results from 1980 and 2002–2005 Regardless of Date of Hire .....	5-46
5.2.6.	Summary of Reference Concentration Values (RfCs) for the Different Health Endpoints and Different Sets of Workers in the Marysville Cohort .....	5-48
5.3.	UNCERTAINTIES IN THE INHALATION REFERENCE CONCENTRATION (RfC) .....	5-50
5.3.1.	Uncertainty in the Exposure Reconstruction .....	5-50
5.3.2.	Uncertainty in the Radiographic Assessment of Localized Pleural Thickening (LPT).....	5-55
5.3.3.	Uncertainty Due to Potential Confounding.....	5-56
5.3.4.	Uncertainty Due to Time Since First Exposure (TSFE) .....	5-60
5.3.5.	Uncertainty in the Endpoint Definition.....	5-64
5.3.6.	Summary of Sensitivity Analyses .....	5-68
5.4.	CANCER EXPOSURE-RESPONSE ASSESSMENT.....	5-69
5.4.1.	Overview of Methodological Approach .....	5-69
5.4.2.	Choice of Study/Data—with Rationale and Justification .....	5-71
5.4.2.1.	Description of the Libby Worker Cohort.....	5-72
5.4.2.2.	Description of Cancer Endpoints .....	5-74
5.4.2.3.	Description of Libby Worker Cohort Work Histories .....	5-76
5.4.2.4.	Description of Libby Amphibole Asbestos Exposures .....	5-77
5.4.2.5.	Estimated Exposures Based on Job-Exposure Matrix (JEM) and Work Histories .....	5-84
5.4.3.	Exposure-Response Modeling .....	5-89
5.4.3.1.	Modeling of Mesothelioma Exposure Response in the Libby Worker Cohort .....	5-90
5.4.3.2.	Results of the Analysis of Mesothelioma Mortality in the Full Libby Worker Cohort.....	5-92

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## CONTENTS (continued)

5.4.3.3.	Modeling and Results of Lung Cancer Exposure Response in the Full Libby Worker Cohort .....	5-95
5.4.3.4.	Rationale for Analyzing the Subcohort of Libby Workers After 1959 .....	5-100
5.4.3.5.	Results of the Analysis of Mesothelioma Mortality in the Subcohort .....	5-102
5.4.3.6.	Results of the Analysis of the Lung Cancer Mortality in the Subcohort .....	5-112
5.4.3.7.	Sensitivity Analysis of the Influence of High Exposures in Early 1960s on the Model Fit in the Subcohort .....	5-123
5.4.3.8.	Additional Analysis of the Potential for Confounding of Lung Cancer Results by Smoking in the Subcohort .....	5-125
5.4.4.	Exposure Adjustments and Extrapolation Methods .....	5-126
5.4.5.	Inhalation Unit Risk (IUR) of Cancer Mortality .....	5-127
5.4.5.1.	Unit Risk Estimates for Mesothelioma Mortality .....	5-127
5.4.5.2.	Unit Risk Estimates for Lung Cancer Mortality .....	5-130
5.4.5.3.	Inhalation Unit Risk (IUR) Derivation for Combined Mesothelioma and Lung Cancer Mortality .....	5-131
5.4.6.	Uncertainties in the Cancer Risk Values .....	5-139
5.4.6.1.	Sources of Uncertainty .....	5-139
5.4.6.2.	Summary .....	5-154
6.	MAJOR CONCLUSIONS IN THE CHARACTERIZATION OF HAZARD AND EXPOSURE RESPONSE .....	6-1
6.1.	HUMAN HAZARD POTENTIAL .....	6-1
6.1.1.	Exposure .....	6-1
6.1.2.	Fiber Toxicokinetics .....	6-2
6.1.3.	Noncancer Health Effects in Humans and Laboratory Animals .....	6-3
6.1.4.	Carcinogenicity in Humans and Laboratory Animals .....	6-5
6.1.5.	Susceptible Populations .....	6-6
6.1.6.	Mode-of-Action Information .....	6-7
6.1.7.	Weight-of-Evidence Descriptor for Cancer Hazard .....	6-7
6.2.	EXPOSURE-RESPONSE .....	6-8
6.2.1.	Noncancer/Inhalation .....	6-8
6.2.1.1.	Uncertainty and Sensitivity Analyses for Reference Concentration (RfC) Derivation .....	6-12
6.2.2.	Cancer/Inhalation .....	6-13
6.2.2.1.	Background and Methods .....	6-13
6.2.3.	Modeling of Mesothelioma Exposure Response .....	6-15
6.2.4.	Unit Risk Estimates for Mesothelioma Mortality .....	6-16
6.2.5.	Modeling of Lung Cancer Exposure Response .....	6-17
6.2.5.1.	Analysis of Potential Confounding of Lung Cancer Results by Smoking in the Subcohort .....	6-18
6.2.6.	Unit Risk Estimates for Lung Cancer Mortality .....	6-18

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## CONTENTS (continued)

6.2.7. Inhalation Unit Risk (IUR) Derivation Based on Combined Mesothelioma and Lung Cancer Mortality from Exposure to Libby Amphibole Asbestos .....	6-19
6.2.7.1. Comparison with Other Published Studies of Libby, MT Workers Cohort.....	6-21
6.2.8. Uncertainty in the Cancer Risk Values .....	6-21
7. REFERENCES .....	7-1
APPENDIX A: EPA RESPONSE TO MAJOR EXTERNAL PEER-REVIEW AND PUBLIC COMMENTS.....	A-1
APPENDIX B: PARTICLE SIZE DISTRIBUTION DATA FOR LIBBY AMPHIBOLE STRUCTURES OBSERVED IN AIR AT THE LIBBY ASBESTOS SUPERFUND SITE.....	B-1
APPENDIX C: CHARACTERIZATION OF AMPHIBOLE FIBERS FROM ORE ORIGINATING FROM LIBBY, MONTANA, LOUISA COUNTY, VIRGINIA, ENOREE, SOUTH CAROLINA, AND PALABORA, REPUBLIC OF SOUTH AFRICA .....	C-1
APPENDIX D: ANALYSIS OF SUBCHRONIC- AND CHRONIC-DURATION STUDIES AND CANCER BIOASSAYS IN ANIMALS AND MECHANISTIC STUDIES.....	D-1
APPENDIX E: EVALUATION OF EXPOSURE-RESPONSE DATA FOR RADIOGRAPHIC CHANGES IN WORKERS FROM THE MARYSVILLE, OH COHORT COMBINING DATA FROM THE 1980 AND 2002-2005 HEALTH EXAMINATIONS .....	E-1
APPENDIX F: WORKER OCCUPATIONAL EXPOSURE RECONSTRUCTION FOR THE MARYSVILLE COHORT .....	F-1
APPENDIX G: EXTRA RISK AND UNIT RISK CALCULATION .....	G-1
APPENDIX H: GLOSSARY OF ASBESTOS TERMINOLOGY .....	H-1
APPENDIX I: EVALUATION OF LOCALIZED PLEURAL THICKENING IN RELATION TO PULMONARY FUNCTION MEASURES .....	I-1

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## LIST OF TABLES

1-1.	Derivation of the current Integrated Risk Information System (IRIS) inhalation unit risk for asbestos from the lifetime risk tables in the Airborne Asbestos Health Assessment Update (AAHAU).....	1-3
2-1.	Optical and crystallographic properties of fibrous amphiboles associated with Libby Amphibole asbestos .....	2-14
2-2.	Air sampling results for asbestos from Zonolite vermiculite attic insulation (VAI) in three homes .....	2-23
3-1.	Factors influencing fiber deposition and clearance in the respiratory system .....	3-5
3-2.	Determinants of fiber toxicity .....	3-15
4-1.	Population and exposure assessment methodologies used in studies of Libby, MT vermiculite workers .....	4-5
4-2.	Source of primary samples for fiber measurements at the Libby vermiculite mining and milling operations .....	4-6
4-3.	Dimensional characteristics of fibers from air samples collected in the vermiculite mill and screening plant, Libby, MT .....	4-9
4-4.	Population and methods used in studies of O.M. Scott, Marysville, OH plant workers.....	4-10
4-5.	Summary of methods used in community-based studies of Libby, MT residents conducted by Agency for Toxic Substances and Disease Registry (ATSDR).....	4-13
4-6.	Nonmalignant respiratory mortality studies of populations exposed to Libby Amphibole asbestos .....	4-15
4-7.	Chest radiographic studies of the Libby, MT vermiculite mine workers .....	4-21
4-8.	Pulmonary function and chest radiographic studies of the O.M. Scott, Marysville, OH plant workers .....	4-24
4-9.	Prevalence of pleural pathological alterations according to quartiles of cumulative fiber exposure in 280 participants .....	4-25
4-10.	Prevalence of pleural thickening in 280 participants according to various cofactors.....	4-26
4-11.	Pathological alterations of parenchyma and pleura in community-based studies .....	4-30

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## LIST OF TABLES (continued)

4-12.	Pulmonary function and respiratory system changes in the Libby, MT community .....	4-32
4-13.	Analyses of pulmonary changes seen on radiographs in relation to pulmonary function in the Libby, MT community .....	4-34
4-14.	Pulmonary function and respiratory system changes in the Libby, MT community: clinic-based study .....	4-35
4-15.	Autoimmune-related studies in the Libby, MT community .....	4-39
4-16.	Respiratory (lung) cancer mortality and exposure-response analyses based on related studies of the vermiculite mining and milling workers in Libby, MT .....	4-41
4-17.	Mesothelioma mortality risk based on studies of the vermiculite mine workers in Libby, MT .....	4-47
4-18.	Exposure levels and health effects observed in communities exposed to tremolite, chrysotile, and crocidolite asbestos .....	4-51
4-19.	In vivo data following exposure to Libby Amphibole asbestos .....	4-54
4-20.	In vivo data following exposure to tremolite asbestos.....	4-60
4-21.	In vitro data following exposure to Libby Amphibole asbestos .....	4-73
4-22.	In vitro data following exposure to tremolite asbestos .....	4-75
4-23.	Hypothesized modes of action for carcinogenicity of Libby Amphibole asbestos in specific organs .....	4-99
5-1.	Summary of candidate principal studies on LAA for reference concentration (RfC) derivation .....	5-6
5-2.	Summary of rationale for identifying candidate principal studies on LAA for reference concentration (RfC) development.....	5-8
5-3.	Characteristics of workers at the O.M. Scott plant in Marysville, OH.....	5-12
5-4.	Characteristics of workers at the O.M. Scott plant in Marysville, OH, with health evaluations in 2002–2005 .....	5-18
5-5.	Models considered to develop a point of departure (POD) .....	5-24
5-6.	Evaluation of association between covariates and exposure, and between covariates and LPT .....	5-27

*This document is a draft for review purposes only and does not constitute Agency policy.*

## LIST OF TABLES (continued)

5-7.	Model features considered in exposure-response modeling to develop a point of departure (POD) .....	5-29
5-8.	Univariate exposure-response modeling for any LPT in the Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 ( $n = 119$ ), using a benchmark response (BMR) of 10% extra risk of any localized pleural thickening (LPT) .....	5-32
5-9.	Estimated point of departure (POD) combining information from the Marysville workers who underwent health evaluations in 2002–2005 and hired in 1972 or later (Primary), and from all workers who underwent health evaluations in 2002–2005 (regardless of hire date), using a benchmark response (BMR) of 10% extra risk of LPT in the Dichotomous Hill model with plateau fixed at 85% .....	5-39
5-10.	(Copy of Table E-11) Reference concentrations (RfCs) for the alternative endpoint of any pleural thickening (APT) in the Marysville cohort with combined x-ray results from 1980 and 2002–2005 regardless of date of hire .....	5-47
5-11.	Multiple derivations of a reference concentration from the Maryville, OH cohort. Primary RfC value in bold.....	5-49
5-12.	Exposure distribution among workers at the O.M. Scott plant in Marysville, OH.....	5-52
5-13.	Effect of truncating exposures after 1980 and of using arithmetic or geometric mean to summarize multiple fiber measurements .....	5-54
5-14.	Effect of including covariates into the final model.....	5-59
5-15.	Effect of different assumptions for the plateau parameter.....	5-61
5-16.	Exposure-response modeling for any localized pleural thickening (LPT) in the Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 ( $n = 119$ ), using a benchmark response (BMR) of 10% extra risk of any LPT, and RTW exposure .....	5-63
5-17.	Effect of using different case/noncase definitions .....	5-65
5-18.	Exposure-response modeling for any localized pleural thickening (LPT) in the Marysville workers who underwent health evaluations in 2002–2005 ( $n = 252$ ), comparing the multinomial model and logistic model with different outcome group definitions .....	5-67
5-19.	Summary of sensitivity analyses.....	5-69

*This document is a draft for review purposes only and does not constitute Agency policy.*

## LIST OF TABLES (continued)

5-20.	Demographic, mortality, and exposure characteristics of the Libby worker cohort .....	5-73
5-21.	Exposure intensity (fiber/cc) for each location operation from the beginning of operations through 1982 Amandus et al. (1987b); Table VII .....	5-78
5-22.	Demographic, mortality, and exposure characteristics of the subset of the Libby worker subcohort hired after 1959 .....	5-82
5-23.	Mesothelioma mortality rate shown by duration of exposure (yr) in the full Libby worker cohort including all hires ( $n = 1,871$ ) .....	5-92
5-24.	Mesothelioma mortality rate shown by age at first exposure in the full Libby worker cohort including all hires ( $n = 1,871$ ) .....	5-92
5-25.	Mesothelioma mortality rate shown by time since first exposure (TSFE) in the full Libby worker cohort including all hires ( $n = 1,871$ ) .....	5-93
5-26.	Comparison of model fit of various univariate exposure metrics for mesothelioma mortality in the full Libby worker cohort including all hires ( $n = 1,871$ ) .....	5-94
5-27.	Lung cancer mortality rate shown by duration of exposure (yr) in the full Libby worker cohort including all hires ( $n = 1,871$ ) .....	5-96
5-28.	Lung cancer mortality rate shown by age at first exposure in the full Libby worker cohort including all hires ( $n = 1,871$ ) .....	5-96
5-29.	Lung cancer mortality rate shown by time since first exposure (TSFE) in the full Libby worker cohort including all hires ( $n = 1,871$ ) .....	5-96
5-30.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 shown by duration of exposure (yr) .....	5-103
5-31.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 shown by age at first exposure .....	5-103
5-32.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 shown by time since first exposure (TSFE) .....	5-103
5-33.	Comparison of model fit of exposure metrics for mesothelioma mortality in the subcohort hired after 1959 .....	5-104
5-34.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the cumulative exposure (CE) with 15-year lag and 5-year half-life .....	5-106
5-35.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the cumulative exposure (CE) with 10-year lag and 5-year half-life .....	5-106

*This document is a draft for review purposes only and does not constitute Agency policy.*

## LIST OF TABLES (continued)

5-36.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the Peto model.....	5-106
5-37.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the Peto model with power $k = 3.9$ and decay $\lambda = 6.8\%/yr$ .....	5-106
5-38.	Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the Peto model with power $k = 5.4$ and decay $\lambda = 15\%/yr$ .....	5-107
5-39.	Mesothelioma mortality exposure metrics fits, slopes per day, and credible intervals in the subcohort of employees hired after 1959 .....	5-111
5-40.	Peto model and Peto model with clearance fits, slopes per year, and credible intervals in the subcohort of employees hired after 1959 .....	5-112
5-41.	Lung cancer mortality rate in the subcohort of employees hired after 1959 shown by duration of exposure (yr).....	5-113
5-42.	Lung cancer mortality rate in the subcohort of employees hired after 1959 shown by age at first exposure.....	5-113
5-43.	Lung cancer mortality rate in the subcohort of employees hired after 1959 shown by time since first exposure (TSFE).....	5-113
5-44.	Model fit comparison for different exposure metrics and lung cancer mortality associated with LAA, controlling for age, gender, race, and date of birth.....	5-115
5-45.	Lung cancer mortality exposure metrics fits, slopes, and confidence intervals (CI) for all retained metrics from Table 5-44 .....	5-119
5-46.	Sensitivity analysis of model fit comparison for different exposure metrics and mesothelioma mortality associated with LAA.....	5-124
5-47.	Sensitivity analysis of model fit comparison for different exposure metrics and lung cancer mortality associated with LAA, controlling for age, gender, race, and date of birth .....	5-125
5-48.	Unit risks for the Peto model and Peto model with clearance .....	5-127
5-49.	Mesothelioma mortality exposure metrics unit risks for the subcohort hired after 1959 .....	5-128
5-50.	Mesothelioma unit risks for the subcohort hired after 1959 adjusted for underascertainment .....	5-129

## LIST OF TABLES (continued)

5-51.	Mesothelioma unit risks for the subcohort hired after 1959 based on the Peto model and the Peto model with clearance adjusted for mesothelioma underascertainment .....	5-129
5-52.	Unit risks for subset of lung cancer models with lagged exposures that yielded statistically significant model fit ( $p < 0.05$ ) and exposure metric fit ( $p < 0.05$ ) to the epidemiologic data .....	5-130
5-53.	Estimates of the combined central estimate of the unit risk for mesothelioma and lung cancer and the combined upper-bound lifetime unit risks for mesothelioma and lung cancer risks (the Inhalation Unit Risk) for different combination of mesothelioma and lung cancer models .....	5-132
5-54.	Lung cancer regression results from different analyses of cumulative exposure in the cohort of workers in Libby, MT .....	5-135
5-55.	Mesothelioma analysis results from different analyses of cumulative exposure in the Libby workers cohort. ....	5-139
6-1.	Estimates of the combined central estimate of the unit risk for mesothelioma and lung cancer and the combined upper-bound lifetime unit risks for mesothelioma and lung cancer risks (the Inhalation Unit Risk) for different combination of mesothelioma and lung cancer models .....	6-20

## LIST OF FIGURES

2-1.	Vermiculite mining operation on Zonolite Mountain, Libby, MT .....	2-1
2-2.	Unexpanded and expanded vermiculite .....	2-2
2-3.	Structure of the silicate minerals, illustrating silicate subclasses by the linking of the basic silicon tetrahedron (A) into more complex structures (B, C, or D) .....	2-4
2-4.	Cross section of amphibole fibers showing the silicon tetrahedrons (triangles with open circles at apex) that make up each double-chain plate.....	2-5
2-5.	Comparison of crystalline forms of amphibole minerals.....	2-8
2-6.	Mineralogy of LAA structures from samples taken from the Zonolite Mountain site .....	2-12
2-7.	Solution series linking tremolite, winchite, and richterite amphibole fibers .....	2-13
2-8.	Scanning electron microscope image of amphibole mineral structures from the Libby, MT mine .....	2-17
2-9.	Fiber morphology of amphibole asbestos from the Libby, MT mine viewed under a scanning electron microscope .....	2-18
2-10.	Particle size (length, width, aspect ratio) of fibers in Libby ore and Libby air ....	2-19
2-11.	Nationwide distribution of Libby ore by county (in tons) .....	2-22
3-1.	General scheme for fiber deposition, clearance, and translocation of fibers from the lung and gastrointestinal tract .....	3-3
3-2.	Architecture of the human respiratory tract and schematic of major mechanisms of fiber deposition .....	3-4
4-1.	Investigations of populations exposed to LAA.....	4-2
4-2.	A (left). Gross appearance at autopsy of asbestos-associated pleural plaques overlying the lateral thoracic wall .....	4-19
4-3.	Lung cancer mortality risk among workers in the Libby, MT vermiculite mine and mill workers .....	4-44
4-4.	Proposed mechanistic events for carcinogenicity of asbestos fibers .....	4-71
5-1.	Candidate studies for derivation of the reference concentration (RfC) in three different study populations, with the most recent study of each population circled .....	5-5

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## LIST OF FIGURES (continued)

5-2.	Radiographic outcomes among Marysville, OH workers.....	5-13
5-3:	Plot of exposure-response models for probability of LPT as a function of mean concentration of occupational exposure in the subcohort.....	5-36
5-4.	Predicted risk of localized pleural thickening (LPT) at the benchmark concentration (BMC) and the lower limit of the BMC (BMCL), using the hybrid Dichotomous Hill model with plateau fixed at 85% .....	5-41
5-5.	Plot of the National Institute for Occupational Safety and Health (NIOSH) job-exposure matrix for different job categories over time .....	5-83
5-6.	Distribution of values of the Peto metric and Peto metric values of mesothelioma deaths (shown as inverted triangles) in the subcohort of employees hired after 1959.....	5-108
5-7.	Distribution of observed values of cumulative exposure (CE) with 15-year lag and 5-year half-life and CE with 15-yr lag and 5-yr half-life values of mesothelioma deaths (shown as inverted triangles) in the subcohort of employees hired after 1959.....	5-109
5-8.	Distribution of observed values of cumulative exposure (CE) with 10-year lag and 5-year half-life and CE with 10-yr lag and 5-yr half-life values of mesothelioma deaths (shown as inverted triangles) in the subcohort of employees hired after 1959.....	5-110
5-9.	Regression diagnostics showing model fit based on the Schoenfeld residuals with two levels of nonparametric smoothing (using cubic splines) to show any patterns of departures from the model predicted values .....	5-121



## LIST OF ABBREVIATIONS AND ACRONYMS

AAHAU	Airborne Asbestos Health Assessment Update
AIC	Akaike Information Criterion
ADAF	age-dependent adjustment factor
ANA	antinuclear antibody
APC	antigen-presenting cells
APT	any pleural thickening
ARC	any radiographic change
ATS	American Thoracic Society
ATSDR	Agency for Toxic Substances and Disease Registry
BALF	bronchoalveolar lavage fluids
BGL	$\beta$ -glucuronidase
BMI	body mass index
BMC	benchmark concentration
BMCL	lower limit of the BMC
BMR	benchmark response
C	mean exposure
CAO	costophrenic angle obliteration
CDF	cumulative distribution frequency
CE	cumulative exposure
CHEEC	cumulative human equivalent exposure concentration
CI	confidence interval
COPD	chronic obstructive pulmonary disease
COX-2	cyclooxygenase-2
CPA	costophrenic angle
CVD	cardiovascular disease
DEF	deferoxamine
$d_{eq}$	aerodynamic equivalent diameter
DIC	Deviance Information Criterion
DLCO	single breath carbon monoxide diffusing capacity
DPT	diffuse pleural thickening
dsDNA	double-stranded DNA
EcSOD	extracellular superoxide dismutase
ED	El Dorado tremolite
EDS	energy dispersive spectroscopy
EPA	U.S. Environmental Protection Agency
EPMA	electron probe microanalysis
FEV	forced expiratory volume
FVC	forced vital capacity
GOF	goodness of fit
GSH	glutathione
GST	glutathione-S-transferase
HAEC	human airway epithelial cell
HO	heme oxygenase
HTE	hamster tracheal epithelial
IARC	International Agency for Research on Cancer
ICD	International Classification of Diseases

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## LIST OF ABBREVIATIONS AND ACRONYMS (continued)

IFN	interferon
Ig	immunoglobulin
IH	industrial hygiene
IL	interleukin
ILO	International Labour Organization
IQR	interquartile range
IRIS	Integrated Risk Information System
IUR	inhalation unit risk
JEM	job-exposure matrix
KL	lung cancer slope factor
KM	mesothelioma slope factor
LAA	Libby Amphibole asbestos
LDH	lactate dehydrogenase
LEC <sub>01</sub>	95% lower confidence limit of the exposure concentration associated with 1% increased risk
LPT	localized pleural thickening
MCAA	antimesothelial cell antibodies
MCMC	Monte Carlo Markov Chain
MMP	matrix metalloproteinase
MOA	mode of action
Mppcf	million particles per cubic foot
MSHA	U.S. Mine Safety and Health Administration
NRC	National Research Council
NDI	National Death Index
Nf2	neurofibromatosis 2
NIEHS	National Institute of Environmental Health Sciences
NIOSH	National Institute for Occupational Safety and Health
ON	Ontario ferroactinolite
OR	odds ratio
PBS	phosphate buffered saline
PCM	phase contrast microscopy
PCMe	phase contrast microscopy equivalent
PG-PS	peptidoglycan-polysaccharide
PLM	polarized light microscopy
PM <sub>2.5</sub>	particulate matter 2.5 µm diameter or less
POD	point of departure
RCF-1	refractory ceramic fibers
RfC	reference concentration
RfD	reference dose
RNP	ribonucleoprotein
RNS	reactive nitrogen species
ROS	reactive oxygen species
RPM	rat pleural mesothelial
RR	relative risk
RTW	residence time-weighted
SAED	selected area electron diffraction

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## LIST OF ABBREVIATIONS AND ACRONYMS (continued)

SAID	systemic autoimmune disease
SD	standard deviation
SE	standard error
SH	spontaneously hypertensive
SHE	Syrian hamster embryo
SHHF	spontaneously hypertensive-heart failure
SIR	standardized incidence ratio
SM	Sumas Mountain chrysotile
SMR	standardized mortality ratio
SOD	superoxide dismutase
SRR	standardized rate ratio
SSA/Ro52	autoantibody marker for apoptosis
SSB	anti-La
SV40	Simian virus 40
TEM	transmission electron microscopy
TSFE	time since first exposure
TWA	time-weighted average
UCL	upper confidence limit
UF	uncertainty factor
UICC	Union for International Cancer Control
USGS	United States Geological Survey
VAI	vermiculite attic insulation
WDS	wavelength-dispersive x-ray spectroscopy
WKY	Wistar-Kyoto rat
XRCC1	x-ray repair cross complementing protein 1

## FOREWORD

The purpose of this Toxicological Review is to provide scientific support and rationale for the hazard and dose-response assessment in the Integrated Risk Information System (IRIS) pertaining to chronic inhalation exposure to Libby Amphibole asbestos, a unique mixture of asbestos fibers originating from the vermiculite mine near Libby, MT. It is not intended to be a comprehensive treatise on the agent or toxicological nature of Libby Amphibole asbestos. The purpose of this document is to establish a Libby Amphibole asbestos-specific reference concentration to address noncancer health effects and to characterize the carcinogenic potential and establish an inhalation unit risk for Libby Amphibole asbestos-related lung cancer and mesothelioma mortality.

The intent of Section 6, *Major Conclusions in the Characterization of Hazard and Exposure Response*, is to present the significant conclusions reached in the derivation of the reference dose, reference concentration, and cancer assessment where applicable, and to characterize the overall confidence in the quantitative and qualitative aspects of hazard and dose-response by addressing the quality of data and related uncertainties. The discussion is intended to convey the limitations of the assessment and to aid and guide the risk assessor in the ensuing steps of the risk assessment process.

For other general information about this assessment or other questions relating to IRIS, the reader is referred to U.S. Environmental Protection Agency's (EPA's) IRIS Hotline at (202) 566-1676 (phone), (202) 566-1749 (fax), or [hotline.iris@epa.gov](mailto:hotline.iris@epa.gov) (email address).

## **AUTHORS, CONTRIBUTORS, AND REVIEWERS**

### **CHEMICAL MANAGERS/AUTHORS**

Thomas F. Bateson, ScD, MPH  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Robert Benson, PhD  
Region 8  
Office of Partnerships and Regulatory Assistance  
U.S. Environmental Protection Agency  
Denver, CO

### **AUTHORS**

Krista Yorita Christensen, PhD  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Glinda Cooper, PhD  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Danielle DeVoney, PhD, DABT, PE  
Captain in the U.S. Public Health Service  
OSRTI Science Policy Branch  
US EPA Office of Solid Waste and Emergency Response  
U.S. Environmental Protection Agency  
Washington, DC

Maureen R. Gwinn, PhD, DABT  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Leonid Kopylev, PhD  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

*This document is a draft for review purposes only and does not constitute Agency policy.*

## **AUTHORS, CONTRIBUTORS, AND REVIEWERS (continued)**

### **CONTRIBUTING AUTHORS**

Rebecca Dzubow, MPH, MEM  
Office of Children's Health Protection  
U.S. Environmental Protection Agency  
Washington, DC

David Berry, PhD  
Region 8  
U.S. Environmental Protection Agency  
Denver, CO

Malcolm Field, PhD  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Annie M. Jarabek  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Research Triangle Park, NC

Keith Salazar, PhD  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Patricia Sullivan, ScD  
Division of Respiratory Disease Studies  
National Institute for Occupational Safety and Health  
Morgantown, WV

### **CONTRIBUTORS**

David Bussard  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Samantha J. Jones, PhD  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

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## **AUTHORS, CONTRIBUTORS, AND REVIEWERS (continued)**

### **CONTRIBUTORS (continued)**

Babasaheb Sonawane, PhD  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

Paul White  
National Center for Environmental Assessment  
U.S. Environmental Protection Agency  
Washington, DC

### **CONTRACTOR SUPPORT**

William Brattin, PhD  
Syracuse Research Corporation  
Denver, CO

Highlight Technologies, LLC, Fairfax, VA  
Dan Heing  
Debbie Kleiser  
Sandra Moore  
Ashley Price  
Kathleen Secor

CACI International, Inc, Arlington, VA  
Thomas Schaffner  
Linda Tackett

ECFlex, Inc., Fairborn, OH  
Heidi Glick  
Crystal Lewis  
Carman Parker-Lawler  
Lana Wood

IntelliTech Systems, Inc., Fairborn, OH  
Cris Broyles

## **AUTHORS, CONTRIBUTORS, AND REVIEWERS (continued)**

### **REVIEWERS**

This document was provided for review to EPA scientists, interagency reviewers from other federal agencies and the Executive Office of the President, and the public, and peer reviewed by independent scientists external to EPA. A summary and EPA's disposition of the comments received from the independent external peer reviewers and the public is included in Appendix A.

### **Science Advisory Board (SAB) Panel for Review of EPA's Draft Toxicological Review of Libby Amphibole Asbestos**

#### **CHAIR**

Dr. Agnes Kane  
Professor and Chair  
Department of Pathology and Laboratory Medicine  
Brown University  
Providence, RI

#### **MEMBERS**

Dr. John R. Balmes  
Professor  
Department of Medicine, Division of Occupational and Environmental Medicine  
University of California  
San Francisco, CA

Dr. James Bonner  
Associate Professor  
Toxicology  
North Carolina State University  
Raleigh, NC

Dr. Jeffrey Everitt  
Director  
Department of Laboratory Animal Science  
GlaxoSmithKline Pharmaceuticals  
Research Triangle Park, NC

Dr. Scott Ferson  
Senior Scientist  
Applied Biomathematics  
Setauket, NY

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## **AUTHORS, CONTRIBUTORS, AND REVIEWERS (continued)**

### **MEMBERS (continued)**

Dr. George Guthrie  
Focus Area Leader  
Geological and Environmental Sciences  
National Energy Technology Laboratory, U.S. Department of Energy  
Pittsburgh, PA

Mr. John Harris  
Principal  
LabCor Portland, Inc.  
Portland, OR

Dr. Tom Hei  
Professor and Vice-Chairman  
Radiation Oncology, College of Physicians and Surgeons  
Columbia University Medical Center  
New York, NY

Dr. David Kriebel  
Professor and Chair  
Department of Work Environment  
School of Health & Environment, University of Massachusetts  
Lowell, MA

Dr. Morton Lippmann  
Professor  
Nelson Institute of Environmental Medicine  
New York University School of Medicine  
Tuxedo, NY

Dr. John Neuberger  
Professor  
Preventive Medicine and Public Health, School of Medicine  
University of Kansas  
Kansas City, KS

Dr. Lee Newman  
Professor of Medicine  
Division of Environmental and Occupational Health Sciences  
School of Public Health, University of Colorado  
Aurora, CO

## **AUTHORS, CONTRIBUTORS, AND REVIEWERS (continued)**

### **MEMBERS (continued)**

Dr. Michael Pennell  
Assistant Professor  
Division of Biostatistics  
College of Public Health, Ohio State University  
Columbus, OH

Dr. Julian Peto  
Professor  
Department of Epidemiology and Population Health  
London School of Hygiene and Tropical Medicine  
London, UK

Dr. Carrie Redlich  
Professor of Medicine  
Internal Medicine  
School of Medicine, Yale University  
New Haven, CT

Dr. Andrew G. Salmon  
Senior Toxicologist  
Office of Environmental Health Hazard Assessment  
California Environmental Protection Agency  
Oakland, CA

Dr. Elizabeth A. (Lianne) Sheppard  
Professor  
Biostatistics and Environmental & Occupational Health Sciences  
School of Public Health, University of Washington  
Seattle, WA

Dr. Randal Southard  
Professor of Soils  
AES Dean's Office  
University of California at Davis  
Davis, CA

Dr. Katherine Walker  
Senior Staff Scientist  
Health Effects Institute  
Boston, MA

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## **AUTHORS, CONTRIBUTORS, AND REVIEWERS (continued)**

### **MEMBERS (continued)**

Dr. James Webber  
Research Scientist  
Wadsworth Center  
New York State Department of Health  
Albany, NY

Dr. Susan Woskie  
Professor  
Work Environment, Health and Environment  
University of Massachusetts Lowell  
Lowell, MA

### **SCIENCE ADVISORY BOARD STAFF**

Dr. Diana Wong  
Designated Federal Officer  
U.S. Environmental Protection Agency  
Washington, DC

# 1. INTRODUCTION

This document presents background information and justification for the Integrated Risk Information System (IRIS) summary of the hazard and exposure-response assessment of Libby Amphibole asbestos (LAA),<sup>1</sup> a mixture of amphibole fibers identified in the Rainy Creek complex and present in ore from the vermiculite mine near Libby, MT. IRIS summaries may include oral reference dose (RfD) and inhalation reference concentration (RfC) values for chronic exposure durations, and a carcinogenicity assessment. This assessment reviews the potential hazards, both cancer and noncancer health effects, from exposure to LAA and provides quantitative information for use in risk assessments: an RfC for noncancer and an inhalation unit risk (IUR) addressing cancer risk. LAA-specific data are not available to support RfD or cancer slope factor derivations for oral exposures.

A RfC is defined as “an estimate of an exposure (including sensitive subgroups) that is likely to be without an appreciable risk of adverse health effects over a lifetime.” ([U.S. EPA, 2002](#)). In the case of LAA, the RfC is expressed in terms of the lifetime exposure in units of fibers per cubic centimeter of air (fibers/cc) in units of the fibers as measured by phase contrast microscopy (PCM). The inhalation RfC for LAA considers toxic effects for both the respiratory system (portal of entry) and for effects peripheral to the respiratory system (extrarespiratory or systemic effects) that may arise after inhalation of LAA.

The carcinogenicity assessment provides information on the carcinogenic hazard potential of the substance in question, and quantitative estimates of risk from inhalation exposures are derived. The information includes a weight-of-evidence judgment of the likelihood that the agent is a human carcinogen and the conditions under which the carcinogenic effects may be expressed. Quantitative risk estimates are derived from the application of a low-dose extrapolation procedure from human data. An IUR is typically defined as a plausible upper bound on the estimate of cancer risk per  $\mu\text{g}/\text{m}^3$  air breathed for 70 years. For LAA, the RfC is expressed as a lifetime daily exposure in fibers/cc (in units of the fibers as measured by PCM), and the IUR is expressed as cancer risk per fibers/cc (in units of the fibers as measured by PCM).

Development of these hazard identification and exposure-response assessments for LAA has followed the general guidelines for risk assessment as set forth by the National Research Council ([NRC, 1983](#)). U.S. Environmental Protection Agency (EPA) Guidelines and Risk Assessment Forum technical panel reports that may have been used in the development of this assessment include the following: *Guidelines for the Health Risk Assessment of Chemical*

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<sup>1</sup>The term “Libby Amphibole asbestos” is used in this document to identify the mixture of amphibole mineral fibers of varying elemental composition (e.g., winchite, richterite, tremolite, etc.) that have been identified in the Rainy Creek complex near Libby, MT. It is further described in Section 2.2.

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Mixtures ([U.S. EPA, 1986c](#)), *Guidelines for Mutagenicity Risk Assessment* ([U.S. EPA, 1986b](#)), *Recommendations for and Documentation of Biological Values for Use in Risk Assessment* ([U.S. EPA, 1988b](#)), *Guidelines for Developmental Toxicity Risk Assessment* ([U.S. EPA, 1991a](#)), *Interim Policy for Particle Size and Limit Concentration Issues in Inhalation Toxicity* ([U.S. EPA, 1994a](#)), *Methods for Derivation of Inhalation Reference Concentrations and Application of Inhalation Dosimetry* ([U.S. EPA, 1994b](#)), *Use of the Benchmark Dose Approach in Health Risk Assessment* ([U.S. EPA, 1995](#)), *Guidelines for Reproductive Toxicity Risk Assessment* ([U.S. EPA, 1996](#)), *Guidelines for Neurotoxicity Risk Assessment* ([U.S. EPA, 1998](#)), *Science Policy Council Handbook: Risk Characterization* ([U.S. EPA, 2000b](#)), *Benchmark Dose Technical Guidance Document* ([U.S. EPA, 2012](#)), *Supplementary Guidance for Conducting Health Risk Assessment of Chemical Mixtures* ([U.S. EPA, 2000c](#)), *A Review of the Reference Dose and Reference Concentration Processes* ([U.S. EPA, 2002](#)), *Guidelines for Carcinogen Risk Assessment* ([U.S. EPA, 2005a](#)), *Supplemental Guidance for Assessing Susceptibility from Early-Life Exposure to Carcinogens* ([U.S. EPA, 2005b](#)), *Science Policy Council Handbook: Peer Review* ([U.S. EPA, 2006c](#)), and *A Framework for Assessing Health Risks of Environmental Exposures to Children* ([U.S. EPA, 2006b](#)).

The literature search strategy employed for this assessment is based on EPA's National Center for Environmental Assessment's Health and Environmental Research Online database tool (which includes PubMed, MEDLINE, Web of Science, JSTOR, and other literature sources). The key search terms included the following: Libby Amphibole, tremolite, asbestos, richterite, winchite, amphibole, and Libby, MT. The relevant literature was reviewed through July 2011. Any pertinent scientific information submitted by the public to the IRIS Submission Desk was also considered in the development of this document. It should be noted that references have been added to the Toxicological Review after the external peer review [SAB \(2013\)](#) in response to peer reviewers' comments and for the sake of completeness.

## **1.1. RELATED ASSESSMENTS**

### **1.1.1. Integrated Risk Information System (IRIS) Assessment for Asbestos ([U.S. EPA, 1988a](#))**

The IRIS assessment for asbestos was posted online in IRIS in 1988 and includes an IUR of 0.23 excess cancers per 1 fiber/cc ([U.S. EPA, 1988a](#)); this unit risk is given in units of the fibers as measured by PCM. The IRIS IUR for general asbestos (CAS Number 1332-21-4) is derived by estimating excess cancers for a continuous lifetime exposure and is based on the central tendency—not the upper bound—of the risk estimates ([U.S. EPA, 1988a](#)) and is applicable to exposures across a range of exposure environments and types of asbestos. Although other cancers have been associated with asbestos ([e.g., laryngeal, stomach, ovarian; Straif et al., 2009](#)), the IRIS IUR for asbestos accounts for only lung cancer and mesothelioma. Additionally, pleural and pulmonary effects from asbestos exposure (e.g., pleural thickening,

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asbestosis, and reduced lung function) are well documented, although currently, there is no RfC for these noncancer health effects.

The derivation of the unit risk for general asbestos is based on the *Airborne Asbestos Health Assessment Update* (AAHAU; U.S. EPA, 1986a). The AAHAU provides various cancer potency factors and mathematical models of lung cancer and mesothelioma mortality based on synthesis of data from occupational studies and presents estimates of lifetime cancer risk for continuous environmental exposures (0.0001 fiber/cc and 0.01 fiber/cc; U.S. EPA, 1986a; see Table 6-3). For both lung cancer and mesothelioma, life table analysis was used to generate risk estimates based on the number of years of exposure and the age at onset of exposure. Although various exposure scenarios were presented, the unit risk is based on a lifetime continuous exposure from birth. The final asbestos IUR is 0.23 excess cancer per 1 fiber/cc continuous exposure<sup>2</sup> and was posted on the IRIS database in 1988 (U.S. EPA, 1988a; see Table 1-1).

**Table 1-1. Derivation of the current Integrated Risk Information System (IRIS) inhalation unit risk for asbestos from the lifetime risk tables in the Airborne Asbestos Health Assessment Update (AAHAU)**

Gender	Excess deaths per 100,000 <sup>a</sup>			Risk	Unit risk (per fiber/cc)
	Mesothelioma	Lung cancer	Total		
Female	183	35	218.5	$2.18 \times 10^{-1}$	
Male	129	114	242.2	$2.42 \times 10^{-1}$	
All	156	74	230.3	$2.30 \times 10^{-1}$	0.23

<sup>a</sup>Data are for exposure at 0.01 fiber/cc for a lifetime.

Source: U.S. EPA (1988a).

The IRIS database has an IUR<sup>3</sup> for asbestos based on 14 epidemiologic studies that included occupational exposure to chrysotile, amosite, or mixed-mineral exposures (chrysotile, amosite, crocidolite; U.S. EPA, 1988a, 1986a). Some uncertainty remains in applying the resulting IUR for asbestos to exposure environments and minerals different from those analyzed in the AAHAU (U.S. EPA, 1986a). No RfC, RfD, or oral slope factor are currently derived for asbestos on the IRIS database.

<sup>2</sup>An IUR of 0.23 can be interpreted as a 23% increase in lifetime risk of dying from mesothelioma or lung cancer with each 1 fiber/cc increase in continuous lifetime exposure.

<sup>3</sup>For purposes of this document, termed "IRIS IUR."

### 1.1.2. EPA Health Assessment for Vermiculite ([U.S. EPA, 1991b](#))

An EPA health assessment for vermiculite reviewed available health data, including studies on workers who mined and processed ore with no significant amphibole fiber content. The cancer and noncancer health effects observed in the Libby, MT worker cohort were not seen in studies of workers exposed to mines with similar exposure to vermiculite but much lower exposures to asbestos fibers. Therefore, it was concluded that the health effects observed from the materials mined from Zonolite Mountain near Libby, MT, were most likely due to amphibole fibers not the vermiculite itself ([U.S. EPA, 1991b](#)). At the time, EPA recommended the application of the IRIS IUR for asbestos fibers (0.23 per fiber/cc) in addressing potential risk of the amphibole fibers entrained in vermiculite mined in Libby, MT.

## 1.2. LIBBY AMPHIBOLE ASBESTOS-SPECIFIC HUMAN HEALTH ASSESSMENT

LAA is a complex mixture of amphibole fibers—both mineralogically and morphologically (see Section 2.3). The mixture primarily includes tremolite, winchite, and richterite fibers with trace amounts of magnesioriebeckite, edenite, and magnesio-arfvedsonite. These fibers exhibit a complete range of morphologies from prismatic crystals to asbestiform fibers ([Meeker et al., 2003](#)). Epidemiologic studies of workers exposed to LAA fibers indicate increased lung cancer and mesothelioma, as well as asbestosis and other nonmalignant respiratory diseases ([Larson et al., 2010b](#); [Larson et al., 2010a](#); [Moolgavkar et al., 2010](#); [Rohs et al., 2008](#); [Sullivan, 2007](#); [McDonald et al., 2004, 2002](#); [Amandus et al., 1988](#); [Amandus et al., 1987b](#); [Amandus and Wheeler, 1987](#); [Amandus et al., 1987a](#); [McDonald et al., 1986a](#); [McDonald et al., 1986b](#); [Lockey et al., 1984](#)).

## 2. LIBBY AMPHIBOLE ASBESTOS: GEOLOGY AND EXPOSURE POTENTIAL

### 2.1. INTRODUCTION

Libby is a community in northwestern Montana that is located near a large open-pit vermiculite mine that operated from the mid 1920s to 1990 (see Figure 2-1). Vermiculite is a silicate mineral that exhibits a sheet-like structure similar to mica (see Figure 2-2, Panel A). When heated to approximately 870°C, water molecules between the sheets change to vapor and cause the vermiculite to expand like popcorn into a light, porous material (see Figure 2-2, Panel B). This process of expanding vermiculite is termed “exfoliation” or “popping.” Both unexpanded and expanded vermiculite have found a range of commercial applications, the most common of which include packing material, attic and wall insulation, various garden and agricultural products, and various cement and building products.



**Figure 2-1. Vermiculite mining operation on Zonolite Mountain, Libby, MT.**



**Panel A: Vermiculite ore sample.** Vermiculite ore sample, Zonolite Mountain, Rainy Creek complex, Libby, MT.



Source: USGS Field Collection, [Meeker \(2007\)](#)

**Panel B: Expanded vermiculite**



**Figure 2-2. Unexpanded and expanded vermiculite.**

- 1 The primary product from the mine was vermiculite concentrate, which was produced by
- 2 milling, screening, and grading the raw ore to enrich for the vermiculite mineral. In general,
- 3 mining practices sought to exclude nonvermiculite material when harvesting the ore, and
- 4 subsequent processing steps were designed to eliminate nonvermiculite materials from the

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finished product. Nevertheless, small amounts of other minerals from the ore body tended to remain in the vermiculite (Zonolite) product. This included a form of asbestos referred to as Libby Amphibole asbestos (LAA).

This chapter provides a brief description of the mineralogical characteristics of asbestos (see Section 2.2), an overview of methods used to identify and measure asbestos in air and solid materials (see Section 2.3), a review of the mineralogical characteristics of LAA in particular (see Section 2.4), and an overview of the potential for current human exposures to LAA (see Section 2.5).

## **2.2. GEOLOGY AND MINERALOGY OF AMPHIBOLES**

### **2.2.1. Overview**

Asbestos is the generic name for a group of naturally-occurring silicate minerals that crystallize in long thin fibers. The basic chemical unit of asbestos and other silicate minerals is  $[\text{SiO}_4]^{4-}$ . This basic unit consists of four oxygen atoms at the apices of a regular tetrahedron surrounding and coordinated with one silicon ion ( $\text{Si}^{4+}$ ) at the center (see Figure 2-3, Panel A). The silicate tetrahedra can bond to one another through the oxygen atoms, leading to a variety of crystal structures (see Figure 2-3, Panels B, C, and D).

There are two main classes of asbestos: serpentine and amphibole. The only member of the serpentine class is chrysotile, which is the form of asbestos that was most commonly used in the past in various man-made asbestos-containing materials (insulation, brake linings, floor tiles, etc.). Chrysotile is a phyllosilicate (see Figure 2-3, Panel D), occurring in sheets that curl into a fibrous form.

There are many different types of amphibole asbestos. This includes five types that were previously used in commerce (actinolite, tremolite, amosite, crocidolite, and anthophyllite), and these forms of asbestos are now regulated. Numerous other asbestiform amphiboles exist, even though they were never used as commercial products and are not currently named in regulations ([Gunter et al., 2007](#)). All forms of amphibole asbestos are inosilicates (see Figure 2-3, Panel C) in which the long axis of the fiber (crystallographically called the c-axis) is parallel to the direction of the chain of silicon tetrahedra.

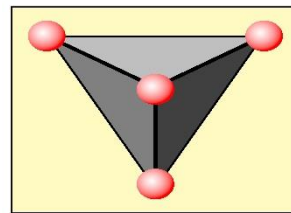
### **2.2.2. Mineralogy of Amphibole Asbestos and Related Amphibole Minerals**

Different types of amphiboles differ from each other primarily in the identity and amounts of monovalent and divalent cations that bind to sites (referred to as A, B, or C sites) along the silicate chains (see Figure 2-4). The specific cations between the two double-chain plates define the elemental composition of the mineral, while the ratio of these cations in each location is used to classify amphiboles within a solid-solution series. The general chemical formula for double-chain inosilicate amphiboles is shown below:

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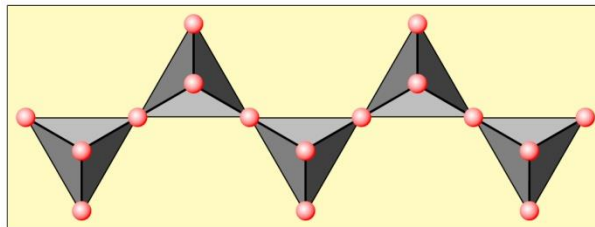
**(A) Nesosilicates or single tetrahedron.**

The single tetrahedron comprises four oxygen molecules covalently bound to the silicon, at the center of the  $[\text{SiO}_4]^{4-}$ -tetrahedron.



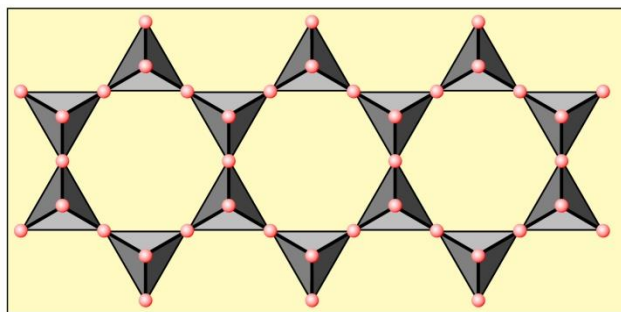
**(B) Inosilicates [*ino* (*gr.*) = thread]—Single-chain silicates.**

Chain silicates are realized by linking  $[\text{SiO}_4]^{4-}$ -tetrahedrons in a way to form continuous chains. They can be represented by a composition of  $[\text{SiO}_3]^{2-}$ . A typical example is diopside  $\text{CaMg}[\text{Si}_2\text{O}_6]$ , in which the “endless” chains are also held together by  $\text{Ca}^{2+}$  and  $\text{Mg}^{2+}$  ions.



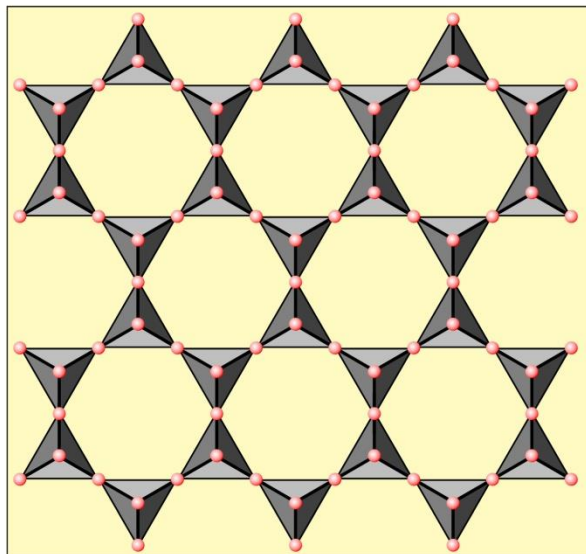
**(C) Inosilicates—Double-chain silicates.**

Two silicate chains of the inosilicates are linked at the corners, forming double-chains and yielding  $[\text{Si}_4\text{O}_{11}]^{6-}$  ions, as realized in the tremolite-ferro-actinolite series  $\text{Ca}_2(\text{Mg},\text{Fe})_5\text{Si}_8\text{O}_{22}(\text{OH},\text{F},\text{Cl})_2$ . Double-chain silicates are commonly grouped with the single-chain inosilicates.

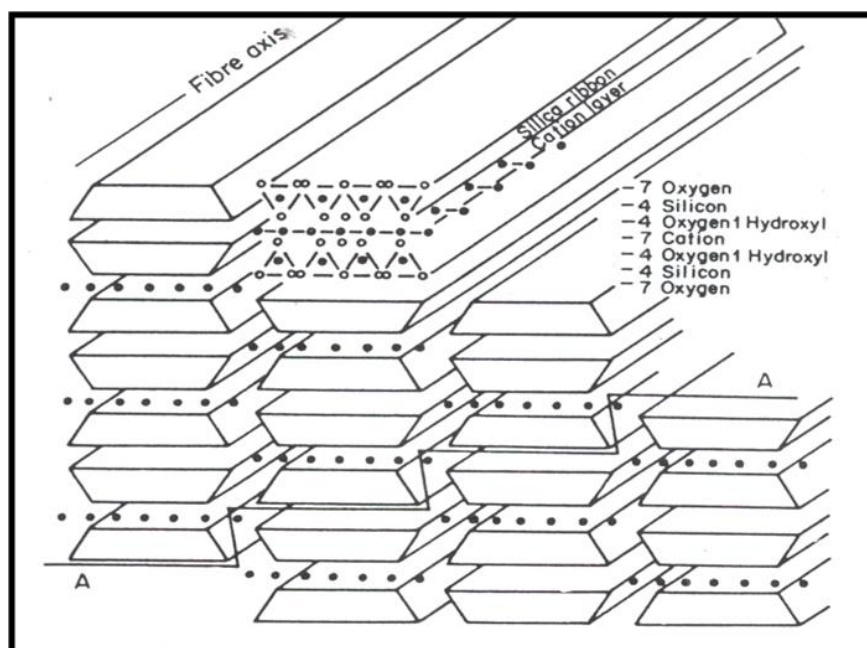


**(D) Phyllosilicates [*phyllo* (*gr.*) = sheet] or sheet silicates.**

These are formed if the double-chain inosilicate  $[\text{Si}_4\text{O}_{11}]^{6-}$  chains are linked to form continuous sheets with the chemical formula  $[\text{Si}_2\text{O}_5]^{2-}$ . Examples of sheet silicates include chrysotile  $\text{Mg}_3\text{Si}_2\text{O}_5(\text{OH})_4$  and vermiculite  $[(\text{Mg},\text{Fe},\text{Al})_3(\text{Al},\text{Si})_4\text{O}_{10}(\text{OH})_2 \bullet 4\text{H}_2\text{O}]$ .



**Figure 2-3. Structure of the silicate minerals, illustrating silicate subclasses by the linking of the basic silicon tetrahedron (A) into more complex structures (B, C, or D).**



**Figure 2-4. Cross section of amphibole fibers showing the silicon tetrahedrons (triangles with open circles at apex) that make up each double-chain plate (shown along the fiber axis). Cations (shown as the darkened dots) occur between the plates forming the basic fiber.**

Source: [Kroschwitz et al. \(2007\)](#).



where:

$A = Na, K$

$B = Na, Li, Ca, Mn, Fe^{2+}, Mg$

$C = Mg, Fe^{2+}, Mn, Al, Fe^{3+}, Ti$

$T = Si, Al.$

The mineral subgroup within amphiboles is determined by the elemental composition.

- Calcic amphiboles (tremolite)
- Sodic-calcic amphiboles (richterite, winchite)
- Sodic amphiboles (riebeckite, arfvedsonite)
- Iron-magnesium-manganese-lithium amphiboles (anthophyllite, cummingtonite-grunerite)

Because the stoichiometry of the cations is not fixed, a continuum of compositions may occur. These are referred to as “solid solution series.” The series are defined by their end

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1 members. For example, a solid solution series for the cation Site A will have one end member  
2 with 100% sodium ions and one end member with 100% potassium ions. This series would  
3 include all intervening ratios.

4 Because each cation site has multiple possibilities, the elemental composition of the  
5 amphibole silicates can be quite complex. It is the complexity of the amphiboles that has  
6 historically given rise to a proliferation of mineral names with little systematic basis ([Hawthorne,](#)  
7 [1981](#)). Currently, amphiboles are identified by a clear classification scheme based on crystal  
8 chemistry that uses well-established names based on the basic mineralogy, with prefixes and  
9 adjective modifiers indicating the presence of substantial substitutions that are not essential  
10 constituents of the end members ([Leake et al., 1997](#)). As implemented, this mineral  
11 classification system does not designate certain amphibole minerals as asbestos. However, some  
12 mineral designations have traditionally been considered asbestos (in the asbestiform habit; e.g.,  
13 tremolite, actinolite). Other commercial forms of asbestos were known by trade names (e.g.,  
14 Amosite) rather than mineralogical terminology (cummingtonite-grunerite).

### 16 **2.2.3. Morphology of Amphibole Minerals**

17 Most amphibole minerals occur in a variety of growth habits, depending on the  
18 temperature, pressure, local stress field, and solution chemistry conditions during crystallization.  
19 The nomenclature used to describe the crystal forms varies between disciplines;([field geologist,](#)  
20 [microscopist; e.g., see Lowers and Meeker, 2002](#)). Text Box 2-1 provides definitions for  
21 common terms used to describe the morphology of asbestos and other related minerals.

## Text Box 2-1: Nomenclature

**Acicular:** The shape showed by an extremely slender crystal with small cross-sectional dimensions (a special case of prismatic form). Acicular crystals may be blunt-ended or pointed. The term “needlelike” refers to an acicular crystal with pointed termination at one of both ends.

**Amphibole:** A group of silicate minerals that may occur either in massive or fibrous (asbestiform) habits.

**Asbestiform** (mineralogical): A specific type of mineral fibrosity in which the fibers and fibrils are long, thin, and possess high tensile strength and flexibility.

**Asbestiform** (regulatory): A specific type of fibrosity in which the fibers and fibrils possess high tensile strength and flexibility.

**Asbestos:** A group of highly fibrous silicate minerals that readily separate into long, thin, strong fibers that have sufficient flexibility to be woven, are heat resistant and chemically inert, are electrical insulators, and therefore are suitable for uses where incombustible, nonconducting, or chemically resistant materials are required.

**Bundle:** A group of fibers occurring side by side with parallel orientations.

**Cleavage fragment:** A fragment produced by breakage of crystal in directions that are related to the crystal structure and are always parallel to possible crystal faces.

**Cluster:** A group of overlapping fibers oriented at random.

**Fiber** (regulatory): A particle that has an aspect ratio (length of the particle divided by its width), and depending on the analytical methods used, a particle is considered a fiber if it has a greater than 3:1 (by PCM) or 5:1 (by transmission electron microscopy [TEM]) aspect ratio.

**Fiber** (mineralogical): The smallest, elongate crystalline unit that can be separated from a bundle or appears to have grown individually in that shape, and that exhibits a resemblance to organic fibers.

**Fibril:** An individual unit of structure, single, elementary fibers that have a small width. A substructure of a fiber.

**Fibrous:** The occurrence of a mineral in bundles of fibers, resembling organic fibers in texture, from which the fibers can usually be separated. Crystallized in elongated, thin, needlelike grains or fibers.

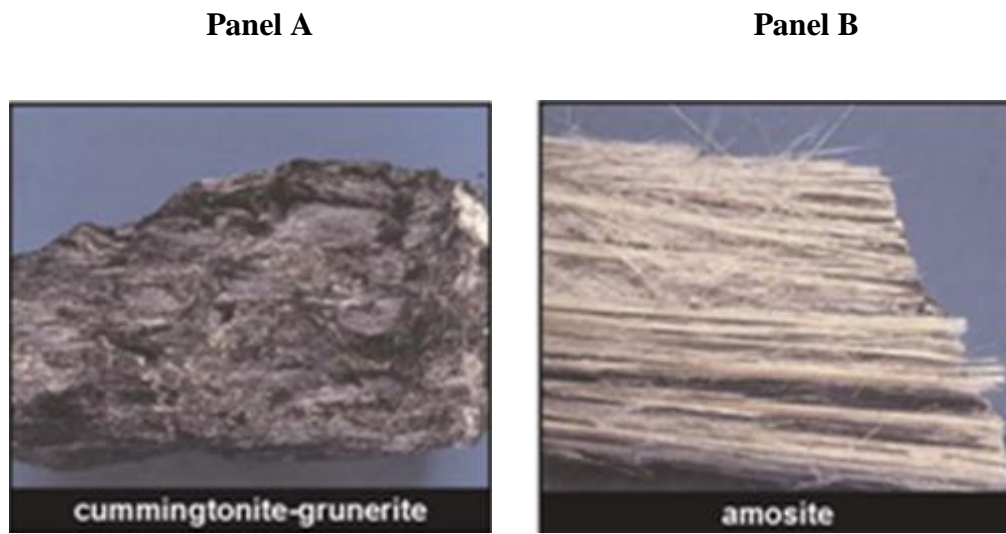
**Massive:** A mineral form that does not contain fibrous crystals.

**Matrix:** A particle of nonasbestos material that has one or more fibers associated with it.

**Prismatic:** Having blocky, pencil-like elongated crystals that are thicker than needles.

**Structure:** A term used mainly in microscopy, usually including asbestos fibers, bundles, clusters, and matrix particles that contain asbestos.

Asbestiform morphology is present where the conditions of formation allow crystals to form very long individual flexible fibers which are parallel and easily separable and may become visible to the naked eye when crushed (see Figure 2-5). Under the microscope, individual amphibole structures may be described as asbestiform, acicular, prismatic, or fibrous. Typically, a fiber is defined as a highly elongated crystal with parallel sides, where acicular crystals are “needlelike” in appearance, and prismatic crystals may have several parallel faces with a low aspect ratio (ratio of length to width, <3:1).



**Figure 2-5. Comparison of crystalline forms of amphibole minerals.** Panel A shows a specimen identified as an amphibole mineral in the cummingtonite-grunerite solid solution series. Although crystalline in form, the habit of formation did not favor formation of individual particles and fibers, hence its appearance as “massive.” Panel B shows an amphibole mineral with very similar elemental composition but formed in a habit where very long fibers were allowed to form—hence the asbestiform appearance.

Source: Adapted from [Bailey et al. \(2006\)](#).

Where conditions are not conducive to the formation of individual fibers and particles, the amphibole is described as massive—appearing as a solid contiguous sample. Mechanical forces that break amphibole crystals along the cleavage planes create smaller pieces or cleavage fragments. These fragments may be elongated and have a morphology that is generally similar to amphibole asbestos, but differ from the regulated minerals in that they did not grow in an asbestiform habit.

## 2.3. METHODS FOR ANALYSIS OF ASBESTOS

Because asbestos is a solid that does not dissolve in water or other solvents, methods for the analysis of asbestos are somewhat different than for most other chemical substances. This section provides a brief overview of the most common methods for the analysis of asbestos.

### 2.3.1. Methods for Air Samples

The exposure pathway of primary health concern for humans is inhalation of asbestos. Air is evaluated for the presence of asbestos by drawing a known volume of air through a filter that traps the solid particles in the air on the filter surface, and the number of asbestos particles are then determined. The concentration is generally expressed as fibers<sup>4</sup> per cubic centimeter of air (fiber/cc) and is computed by dividing the number of asbestos fibers on the filter by the volume of air drawn through the filter.

In all cases, the evaluation of the particles that are collected on an air filter is performed using a microscope. All methods begin with the basic shape (morphology) of a particle to classify it as a possible asbestos particle or not. In general, particles that are clearly fibrous (substantially longer than they are thick) are considered to be potential asbestos. However, other minerals besides asbestos may occur in long thin particles, and a number of nonmineral fibers may be present in a sample as well. Consequently, some techniques rely on other physical or optical properties of the particles to help distinguish asbestos from nonasbestos and to classify the type of asbestos. These differences in the ability to visualize and distinguish asbestos particles are the most important differences between the various microscopic techniques.

The most common technique in the past for analyzing asbestos in air samples was PCM, and this method remains the current industrial hygiene (IH) standard methodology, usually using National Institute for Occupational Safety and Health (NIOSH) method 7400 (<http://www.cdc.gov/niosh/docs/2003-154/pdfs/7400.pdf>). Under this method, a fiber is defined as any particle greater than 5 µm in length with an aspect ratio greater than or equal to 3:1. The limit of resolution of PCM is usually about 0.25 µm, so fibers thinner than this are usually not observable. A key attribute of PCM is that identification of countable fibers is based only on morphology, and does not consider mineralogy or crystal structure. Because of this, it is not possible to classify asbestos fibers by mineral type, or even to reliably distinguish between asbestos and nonasbestos fibers. This is not usually a significant concern when applied to air samples collected in a workplace where asbestos is present, but can become an issue in nonworkplace settings where asbestos concentrations tend to be lower and other types of fibers are more common.

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<sup>4</sup>Most techniques for analyzing air samples distinguish individual fibers from more complex structures composed of two or more fibers, including bundles, clusters and matrix particles. For simplicity, the term “fiber” is used here to include not only fibers but the more complex structures as well.



Transmission electron microscopy (TEM) has also been developed for analysis of air samples for asbestos. TEM uses a high energy electron beam rather than a beam of light to irradiate the sample, and this allows visualization of structures much smaller than can be seen under light microscopy. In addition, most TEM instruments used for asbestos analysis have equipment that allows a more detailed characterization of a particle than is possible by PCM:

- EDS (energy dispersive spectroscopy) provides data on the atomic composition of each particle being examined. This makes it easy to distinguish organic fibers from mineral fibers, and also allows for distinguishing between different types of mineral fibers.
- SAED (selected area electron diffraction) provides a diffraction pattern for crystalline particles that is helpful in distinguishing organic from mineral fibers, and in classifying the nature of the crystalline structure (serpentine, amphibole, pyroxene, etc.).
- WDS (wavelength-dispersive x-ray spectroscopy) provides x-ray spectral data from a single wavelength at a time, providing detailed atomic composition of a particle. Generally, WDS is a more precise measure of the atomic composition of a particle than EDS and is often used with an electron microprobe attached to a scanning electron microscope.

Several different standard methods have been developed for TEM analyses of air samples, the most common of which is ISO 10312 (ISO 10312:1995). Under ISO 10312 counting rules, a fiber is defined as any structure  $\geq 0.5 \mu\text{m}$  in length that has substantially parallel sides and an aspect ratio  $\geq 5:1$ . Fibers observed under TEM that meet PCM counting rules are generally referred to as PCMe (PCM-equivalent).

### **2.3.2. Methods for Solid Materials**

Measurement of asbestos in solid samples (vermiculite, building materials, soil, etc.) usually employs polarized light microscopy (PLM). There are several standard PLM methods for the analysis of asbestos in bulk materials, including NIOSH 9002, EPA/600/R-93/116, and CARB 435. PLM uses the optical properties of asbestos to identify and classify different types of asbestos fibers. In general, these methods are most reliable for materials that contain relatively high concentrations of asbestos, and results tend to become more variable as concentrations decrease below about 1% by mass. At present, the use of TEM for the analysis of bulk materials is not a well-developed procedure.

## **2.4. CHARACTERISTICS OF LIBBY AMPHIBOLE ASBESTOS**

Amphibole asbestos occurs in the Libby vermiculite ore body both in high concentration veins (>80%), as well as in lower concentrations (0.1 to 3%) within the layers of the vermiculite

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ore itself ([Lowers et al., 2012](#); [U.S. EPA, 2000a](#); [Boettcher, 1967](#); [Pardee and Larsen, 1928](#)). Analysis of historical ore samples from the Harvard and Smithsonian Museums (circa 1920s), the Butte Museum (circa 1960), and recent ore samples from the mine (circa 1999) indicate that the amphibole content of vermiculite ore from the mine has remained approximately constant over the 70-year mining history at the Rainy Creek complex ([Sanchez et al., 2008](#); [Meeker et al., 2003](#)).

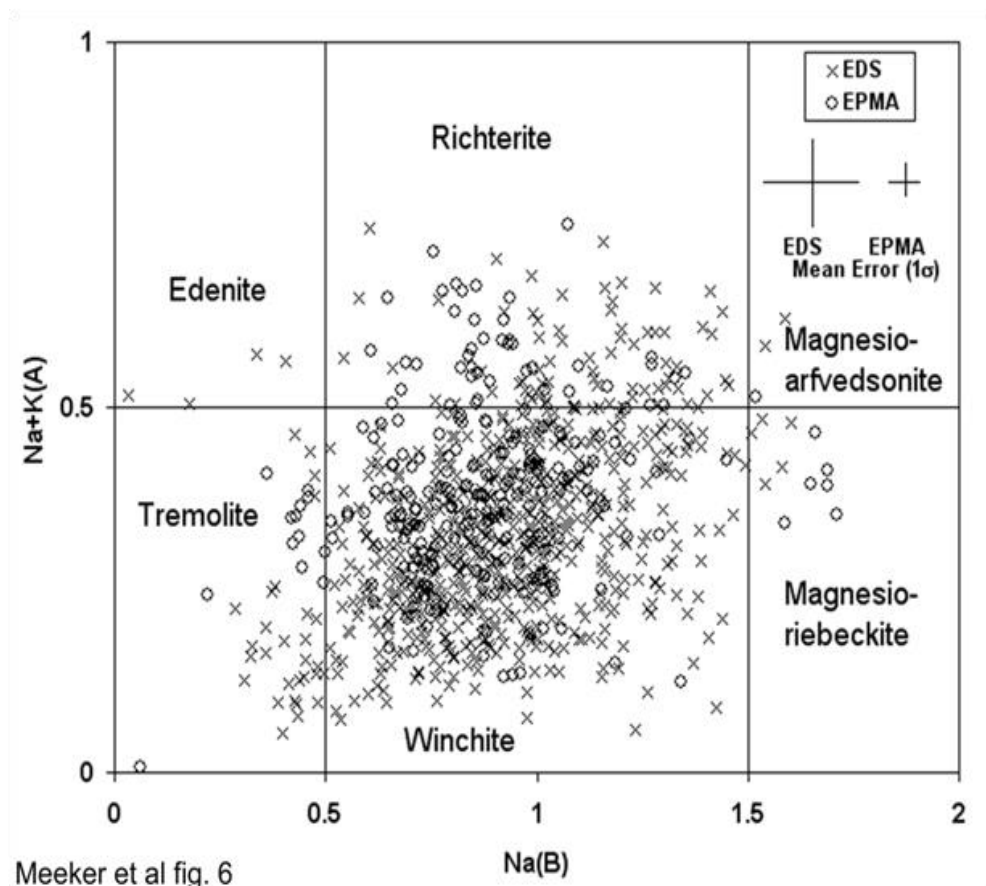
#### **2.4.1. Mineralogy of Libby Amphibole Asbestos**

Historically, the amphibole mineral fibers that occur in the Libby ore body were described as a sodium-rich tremolite ([Amandus et al., 1987b](#); [McDonald et al., 1986a](#); [Leake, 1978](#); [Boettcher, 1966](#); [Larsen, 1942](#)), although [McDonald et al. \(1986a\)](#) noted the sodium content was too high to allow classification as tremolite, and suggested that at least some fibers might be better classified as magnesio-rebeckite or richterite.

More recently, various research groups ([Gunter and Sanchez, 2009](#); [Sanchez et al., 2008](#); [Meeker et al., 2003](#); [Wylie and Verkouteren, 2000](#); [Ross et al., 1993](#); [Moatamed et al., 1986](#)) have recharacterized the mineralogical composition of amphiboles from the Libby mine using the modern classification scheme developed by [Leake et al. \(1997\)](#).

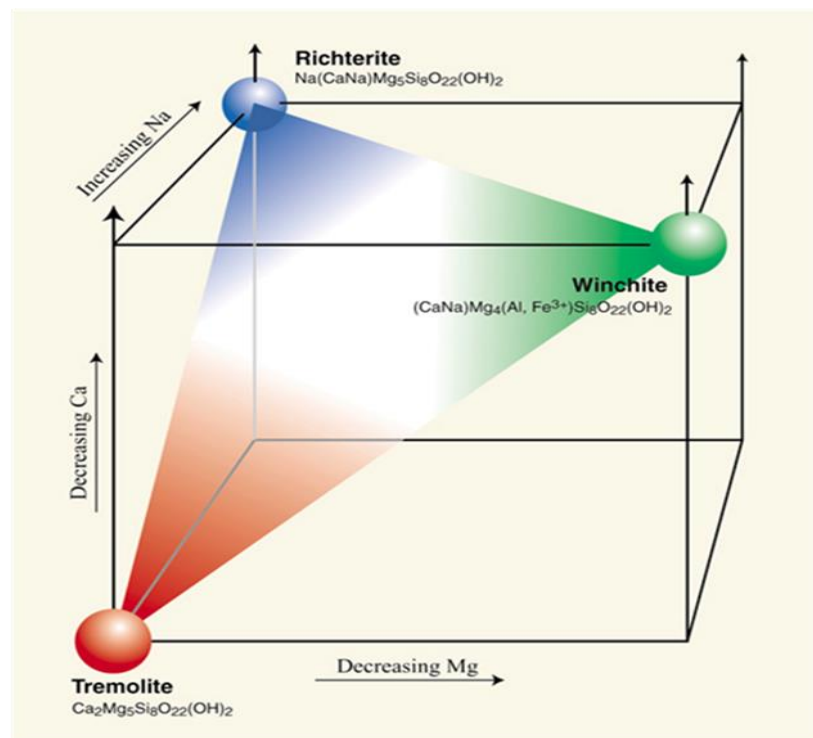
The most extensive investigation was reported by the U.S. Geological Survey ([USGS; Meeker et al., 2003](#)). In this investigation, USGS personnel collected amphibole samples from different areas of the mine to identify the range of materials present. The mineral composition of individual structures was determined using EDS and electron probe microanalysis (EPMA). The results, which are presented in Figure 2-6, show that most amphibole structures were classified as winchite (84%), with lesser amounts classified as richterite (11%) and tremolite (6%). Trace amounts of magnesio-riebeckite, magnesio-arfvedsonite, and edenite are also present. [Sanchez et al. \(2008\)](#) found a similar distribution of amphibole mineral types in a sample of ore collected from the mine in 2009. [Wylie and Verkouteren \(2000\)](#) reported the presence of asbestiform winchite and richterite in ore samples from the mine, which was consistent with the alteration of alkali igneous rocks.

The relationship between the cationic compositions of the three primary minerals is illustrated in Figure 2-7. In some fibers, the composition differed within the length of the fiber (e.g., winchite at one end and richterite at the other end). All of these minerals are within the solid solution series for tremolite-richterite-magnesio-riebeckite. Magnesio-riebeckite and magnesio-arfvedsonite fall within the range of sodic amphiboles, winchite and richterite fall within the range of sodic-calcic amphiboles, and tremolite and edenite are considered calcic amphiboles. Structural formulae and optical and crystallographic data are presented in Table 2-1.



**Figure 2-6. Mineralogy of LAA structures from samples taken from the Zonolite Mountain site.** An evaluation of the textural characteristics shows the material to include a complete range of morphologies from prismatic crystals to fibers. Each data point represents the cation composition (number of occupied sites) for a single fiber. The  $x$ -axis shows the number of sites occupied by Na, and the  $y$ -axis shows the number of sites occupied by Na or K. The data shown are a composite of the analysis of fibers taken from 30 different field samples from various locations within the mine.

Source: [Meeker et al. \(2003\)](#)



**Figure 2-7. Solution series linking tremolite, winchite, and richterite amphibole fibers.**

**Table 2-1. Optical and crystallographic properties of fibrous amphiboles associated with Libby Amphibole asbestos**

Mineral	Habit and color	Refractive indices		Birefringence	Extinction	Elongation sign
		$\alpha$	$\gamma$			
Tremolite <sup>a</sup> $\text{Ca}_2\text{Mg}_5\text{Si}_8\text{O}_{22}(\text{OH})_2$	Straight to curved fibers and bundles. Colorless to pale green.	1.600–1.628 1.604–1.612 1.599–1.612 1.6063	1.625–1.655 1.627–1.635 1.625–1.637 1.6343	0.017–0.028	Oblique up to 21°	+ (length slow)
Actinolite $\text{Ca}_2(\text{Mg},\text{Fe})_5\text{Si}_8\text{O}_{22}(\text{OH})_2$		1.600–1.628 1.612–1.668 1.613–1.628 1.6126	1.625–1.655 1.635–1.688 1.638–1.655 1.6393	0.017–0.028		+ (length slow)
Winchite $\text{CaNaMg}_4(\text{Al},\text{Fe}^{3+})\text{Si}_8\text{O}_{22}(\text{OH})_2$	Straight to curved fibers or bundles. Colorless to pale blue. Pleochroism weak to moderate: X = colorless, Y = light blue–violet, Z = light blue. <sup>d</sup>	<b>1.618–1.626<sup>b</sup></b> <b>1.618–1.621<sup>c</sup></b> 1.629 <sup>d</sup> 1.636 <sup>e</sup>	<b>1.634–1.642<sup>b</sup></b> <b>1.634–1.637<sup>c</sup></b> 1.650 <sup>d</sup> 1.658 <sup>e</sup>	<b>0.008–0.019<sup>b</sup></b> <b>0.016<sup>c</sup></b> 0.021 <sup>d</sup> 0.022 <sup>e</sup>	<b>Oblique, 22°<sup>b</sup></b> <b>15.8°<sup>c</sup></b> Oblique, 7–29° <sup>h</sup>	+ (length slow)
Richterite $\text{NaCaNa}(\text{Mg},\text{Fe})_5\text{Si}_8\text{O}_{22}(\text{OH})_2$	Straight to curved fibers or bundles. Colorless, pale yellow, brown, pale to dark green, or violet. <sup>h</sup> Pleochroism weak to strong in pale yellow, orange, and red. <sup>f</sup>	<b>1.622–1.623<sup>b</sup></b> 1.605–1.624 <sup>f</sup> 1.615 <sup>g</sup>	<b>1.638–1.639<sup>b</sup></b> 1.627–1.641 <sup>f</sup> 1.636 <sup>g</sup>	<b>0.012–0.017<sup>b</sup></b> 0.017–0.022 <sup>f</sup>	<b>Oblique, 21–22°<sup>b</sup></b> Oblique, 5–45° <sup>h</sup>	+ (length slow)
Magnesio-riebeckite $\text{Na}_2\text{Mg}_3\text{Fe}_2^{3+}\text{Si}_8\text{O}_{22}(\text{OH})_2$	Prismatic to fibrous aggregates. Blue, grey-blue, pale blue to yellow. Can be pleochroic. <sup>h</sup>	1.650–1.673 <sup>h</sup>	1.662–1.676 <sup>h</sup>	Up to 0.015 <sup>h</sup>	Oblique, 8–40° <sup>h</sup>	– (length fast) <sup>h</sup>

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**Table 2-1. Optical and crystallographic properties of fibrous amphiboles associated with Libby Amphibole asbestos (continued)**

Mineral	Habit and color	Refractive indices		Birefringence	Extinction	Elongation sign
		$\alpha$	$\gamma$			
Magnesio-arfvedsonite $\text{NaNa}_2\text{Mg}_4\text{Fe}^{3+}\text{Si}_8\text{O}_{22}(\text{OH})_2$	Prismatic to fibrous aggregates. Yellowish green, brownish green, or grey-blue. Can be pleochroic. <sup>h</sup>	1.623–1.660 <sup>h</sup>	1.635–1.680 <sup>h</sup>	0.012–0.026 <sup>h</sup>	Oblique, 18–45 <sup>oh</sup>	– (length fast)
Edenite $\text{NaCa}_2\text{Mg}_5\text{AlSi}_7\text{O}_{22}(\text{OH})_2$	Prismatic to fibrous aggregates. White, grey, pale to dark green, also brown and pale pinkish-brown. Can be pleochroic. <sup>i</sup>	1.606–1.649 <sup>i</sup>	1.631–1.672 <sup>i</sup>	0.025 <sup>i</sup>	Oblique, 12–34 <sup>oh</sup>	+ (length slow)

<sup>a</sup>Adapted from: U.S. EPA. (1993) Method for the determination of asbestos in bulk building materials. Method EPA/600/R-93/116. July 1993. (NTIS/PB93-218576).

<sup>b</sup>Bandli, BR; Gunter, ME; Twamley; et al. (2003) Optical, compositional, morphological, and x-ray data on eleven particles of amphibole from Libby, MT, U.S.A. Canadian Mineralogist 41: 1241–1253.

<sup>c</sup>Wylie, AG; Verkouteren, JR. (2000) Amphibole asbestos from Libby, MT: Aspects of nomenclature. American Mineralogist, 85: 1540–1542.

<sup>d</sup>[www.minsocam.oeg/msa/Handbook/Winchite.PDF](http://www.minsocam.oeg/msa/Handbook/Winchite.PDF).

<sup>e</sup>[www.mindat.org/min-4296.html](http://www.mindat.org/min-4296.html).

<sup>f</sup>[www.minsocam.oeg/msa/Handbook/Richterite.PDF](http://www.minsocam.oeg/msa/Handbook/Richterite.PDF).

<sup>g</sup>[www.webmineral.com/data/Richterite.shtml](http://www.webmineral.com/data/Richterite.shtml).

<sup>h</sup>Deer, WA; Howie, RA; Zussman, J. (1997)(1997)(1997) Rock Forming Minerals Volume 2B: Double Chain Silicates, 2<sup>nd</sup> Edition. The Geological Society, London.

<sup>i</sup>[www.mindat.org/min-1351.html](http://www.mindat.org/min-1351.html).

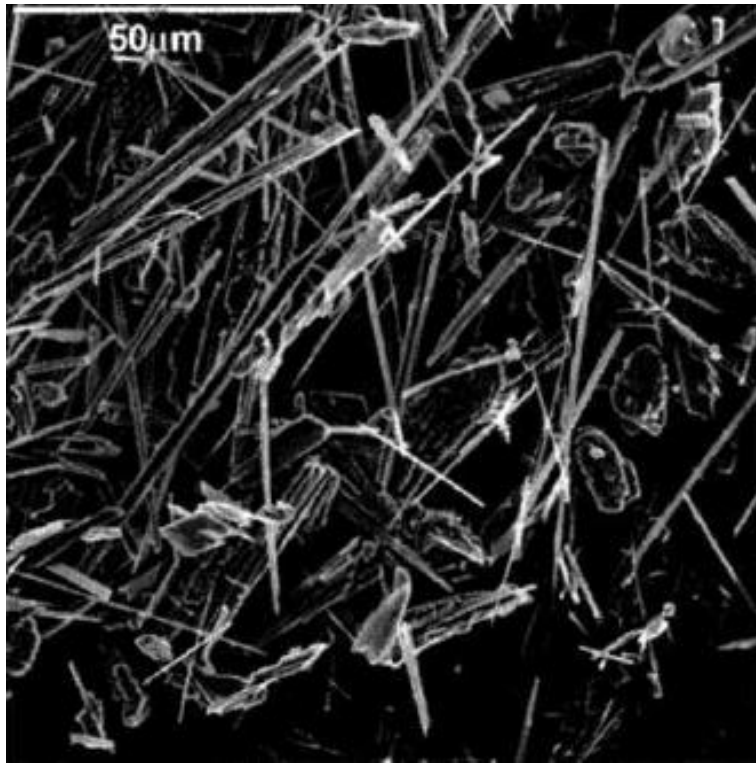
#### 2.4.2. Morphology of Libby Amphibole Asbestos

A number of investigators have reported on the morphology of LAA structures in samples from the mine site, as well as in air samples from the former mine and mill or from present-day town of Libby. [McDonald et al. \(1986a\)](#) used TEM to examine particles collected on air filters from the mine and mill in Libby. The authors reported that fibers on the filters included a range of morphologies, including straight with uniform diameter, a lath or needle shape, or curved.

[Brown and Gunter \(2003\)](#) used PLM to examine structures obtained from three different mineral samples collected at the mine in Libby. Each of the three samples was crushed and sieved through a 250 µm screen. Based on aspect ratio, 95% of the structures ranked as asbestos. Based on a more detailed evaluation of crystal structure, about one-third were judged to be asbestos, about one-third were judged to be cleavage fragments, and about one-third could not be classified with confidence.

[Meeker et al. \(2003\)](#) reported that all of the amphiboles found at the mine site, with the possible exception of magnesio-riebeckite, can occur in fibrous habit. It was observed these amphibole materials—even when originally present as massive material—can produce abundant, extremely fine fibers by gentle abrasion or crushing.

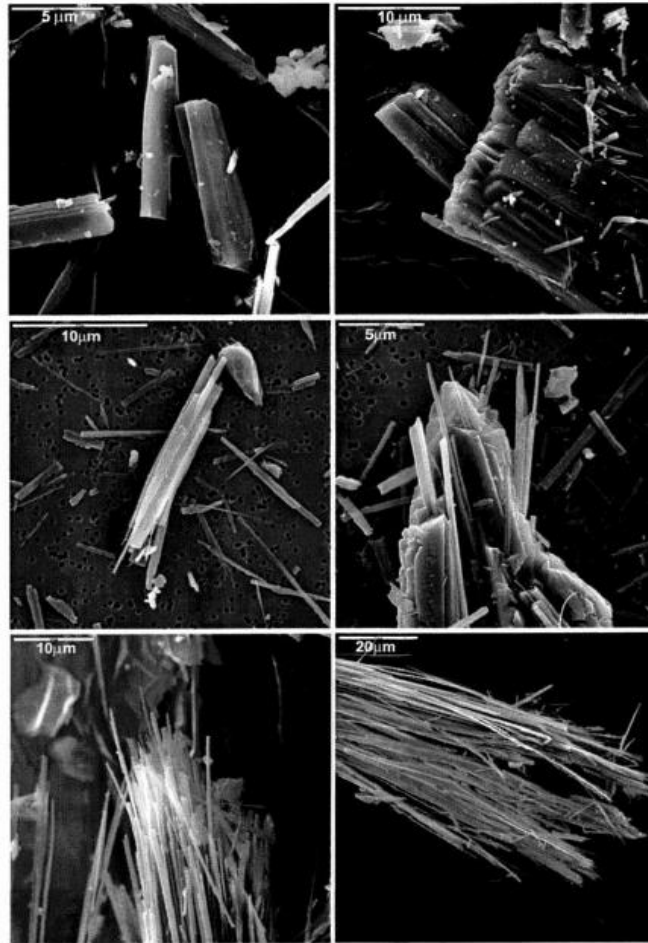
Figure 2-8 shows a scanning electron microscope image of amphibole mineral collected from the mine in Libby ([Meeker et al., 2003](#)). This image illustrates the broad range of size and morphologies that can occur in this material. As individual structures are viewed under greater magnification, the range of morphologies can be more clearly seen (see Figure 2-9). The USGS has observed structures that are fibrous, acicular, and prismatic, all within the minerals from the mine ([Meeker et al., 2003](#)).



**Figure 2-8. Scanning electron microscope image of amphibole mineral structures from the Libby, MT mine.** An evaluation of the textural characteristics shows the material to include a range of morphologies from prismatic crystals to fibers. Acicular and prismatic crystals, fibers bundles, and curved fibers are all present.

Source: [Meeker et al. \(2003\)](#).





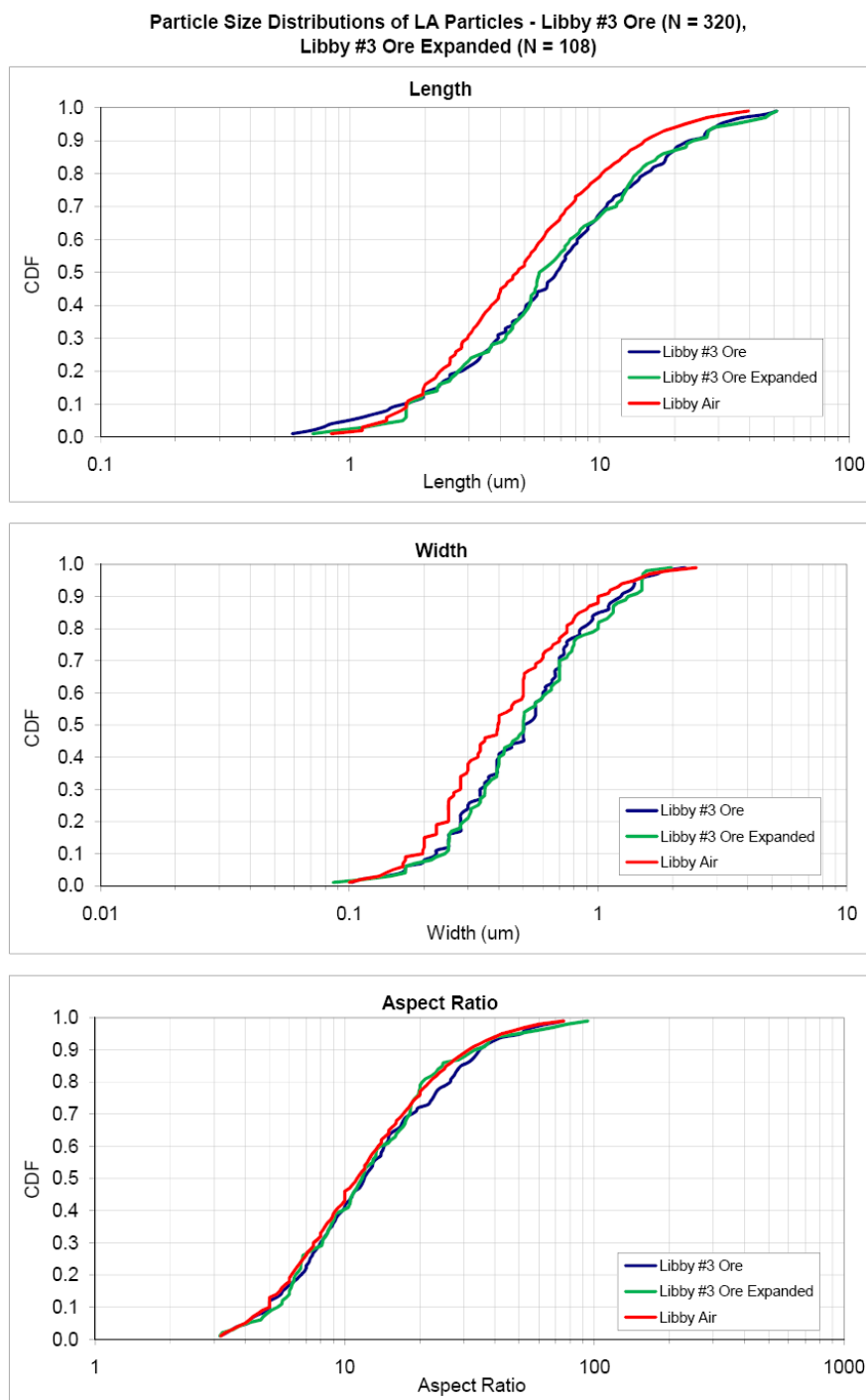
**Figure 2-9. Fiber morphology of amphibole asbestos from the Libby, MT mine viewed under a scanning electron microscope.**

Source: [Meeker et al. \(2003\)](#).

[Sanchez et al. \(2008\)](#) evaluated fiber morphology using the optical system of an electron microprobe, and reported that most structures could be classified as either prismatic or fibrous. There was no difference in the mineralogy between the two morphologies, and the authors concluded the different habits were formed at the same time.

Figure 2-10 shows cumulative particle-size-distribution frequencies (CDFs) for LAA fibers (aspect ratio  $\geq 3:1$ ) observed using TEM in Libby ore Grade 3, expanded Libby ore Grade 3, and ambient air samples collected in Libby. The data used to construct this plot are described in Appendices B and C. In general, most fibers identified as LAA have thicknesses that range from about 0.1  $\mu\text{m}$  to 1  $\mu\text{m}$ , with an average of about 0.6  $\mu\text{m}$ . Fiber lengths vary greatly, ranging from  $<1$   $\mu\text{m}$  to  $\geq 100$   $\mu\text{m}$ . Aspect ratios also range widely, from 3:1 to greater than 100:1.

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**Figure 2-10. Particle size (length, width, aspect ratio) of fibers in Libby ore and Libby air.**

CDF = cumulative distribution frequency; LA = Libby Amphibole.

Source: [U.S. EPA \(2010b\); provided as Appendix B.](#)

1 An important question is whether the mineralogy and morphology of LAA fibers  
2 observed in geological samples of amphibole material collected at the mine are similar to that  
3 observed for airborne fibers collected on filters in Libby or other locations where vermiculite  
4 was used or processed. As shown in Figure 2-10, the size distributions for fibers observed in the  
5 unexpanded and expanded Libby Grade 3 ores are very similar to each other, while the LAA  
6 fibers observed in air monitoring samples from Libby tend to be slightly thinner and shorter than  
7 in the ore samples. However, the differences are relatively minor. Mineralogical  
8 characterization by EDS and SAED of the fibers from the Libby ore Grade 3 and the expanded  
9 product provided additional confirmation of the similarity between the fibers from the Libby  
10 Grade 3 ore and Libby air samples (methodology described in Section 2.3; see also Appendix B).  
11 EDS spectra yielded an elemental fingerprint with sodium and potassium peaks that were highly  
12 consistent with values reported for the winchite-richierite solution series described for the Libby  
13 ores ([Meeker et al., 2003](#)).  
14

## 15 **2.5. HUMAN EXPOSURE POTENTIAL**

16 Several different populations have the potential for exposure to vermiculite (Zonolite)  
17 from the mine in Libby, MT, and hence the potential for exposure to the LAA associated with  
18 this material. This includes not only the former workers at the mine and mill site, but also  
19 residents in the community of Libby, MT, as well as workers at other locations who processed  
20 the vermiculite product. A brief description of these potentially exposed populations is presented  
21 below.  
22

### 23 **2.5.1. Exposures Pathways in the Libby Community**

24 When the mine in Libby, MT was active, miners, mill workers, and those working in the  
25 processing plants were exposed to vermiculite, silica dust, and amphibole structures released to  
26 air from the ore during the mining and processing operations ([Meeker et al., 2003](#); [Amandus et al., 1987b](#);  
27 [McDonald et al., 1986a](#)). In some cases, workers may have inadvertently transported,  
28 typically on their clothing, shoes, and hair, contaminated materials from the workplace to  
29 vehicles, homes, and other establishments. This transported material may have resulted in  
30 “take-home exposure” for the workers, their families, and other coresidents. The magnitude of  
31 these historic take-home exposures was not measured, so the levels to which individuals in the  
32 home might have been exposed are not known.

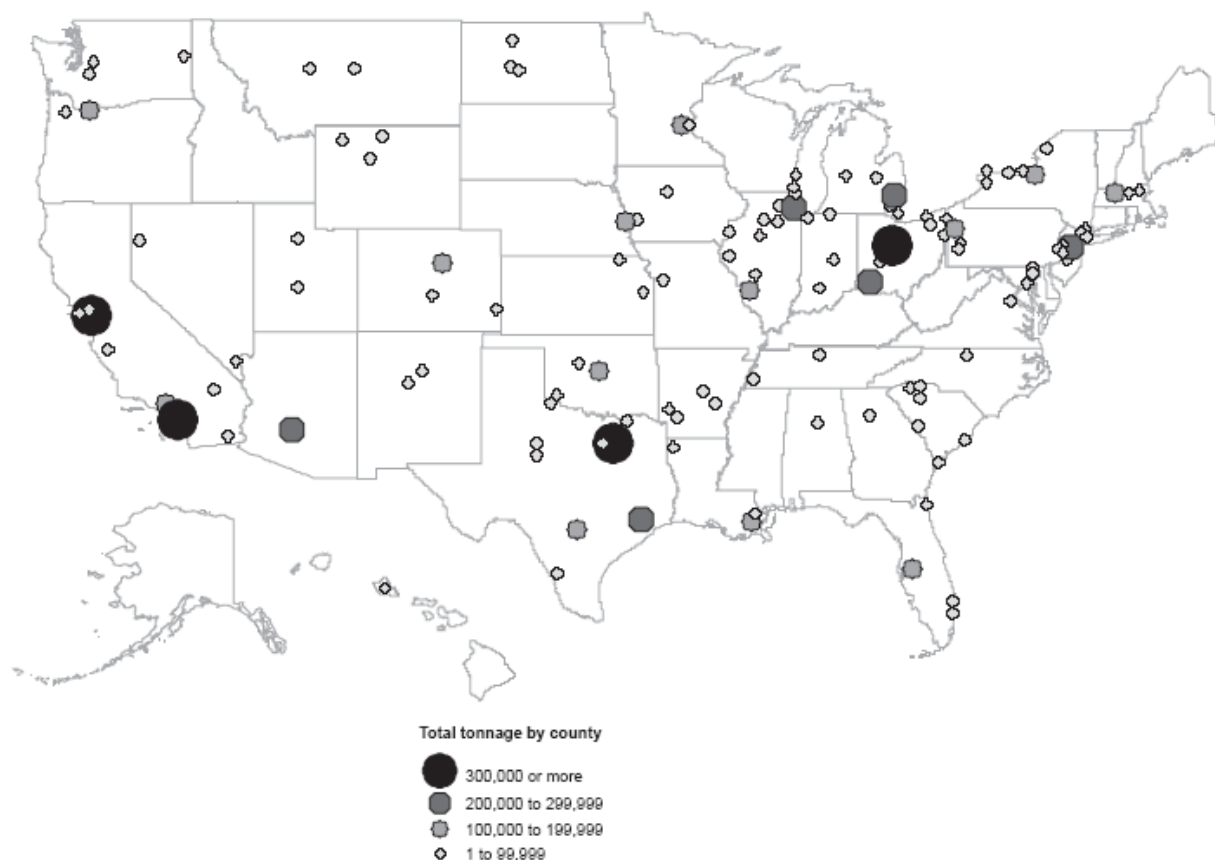
33 The Agency for Toxic Substances and Disease Registry (ATSDR) performed an exposure  
34 survey in Libby to identify activities that may have led to the exposure of residents to vermiculite  
35 and LAA. Based on the responses of survey participants, it was found that men were more likely  
36 than women to have had both occupational and nonoccupational exposures, while women were  
37 more likely to have had only household contact with exposed workers ([Peipins et al., 2003](#);

1 [ATSDR, 2001b](#)). Expanded vermiculite, as a finished product (Zonolite), was used for  
2 insulation in attics and walls in homes in Libby, and was also used as a soil amendment in homes  
3 and recreational areas. Community members may have been exposed, and are possibly still  
4 exposed, to these consumer (Zonolite) products. In a survey of Libby residents conducted by  
5 ATSDR in 2000–2001, almost 52% reported using vermiculite for gardening, 8.8% used  
6 vermiculite around the home, and 51% reported handling vermiculite attic insulation ([VAL](#),  
7 [Peipins et al., 2003](#)). Because vermiculite ore, vermiculite product, and waste stoner rock (the  
8 waste material from exfoliation) were present in the community, numerous other activities may  
9 also have resulted in exposure. Individuals also reported exposures from the following activities:  
10 participating in recreational activities along Rainy Creek Road, which is the road leading to the  
11 mine (67%); playing at the ball field near the expansion plant (66%); playing in the vermiculite  
12 piles (34%); heating the vermiculite to make it expand/pop (38%); or other activities in which  
13 contact with vermiculite occurred ([31%; Peipins et al., 2003](#)).

14 Because a number of different activities may be associated with exposure to LAA in  
15 Libby, it is important to recognize that the overall health hazard to an individual is related to the  
16 sum of the exposures across all scenarios that apply to that individual.

#### 18 **2.5.2. Exposure Pathways in Communities with Vermiculite Expansion and Processing** 19 **Plants**

20 While some vermiculite concentrate was exfoliated and used in Libby, MT, most of the  
21 concentrate was transported to expansion plants at other locations across the country where it  
22 was exfoliated and distributed. A review of company records from 1964–1990 indicates that  
23 more than 6 million tons of vermiculite concentrate was shipped to over 200 facilities outside of  
24 Libby ([ATSDR, 2008](#)). Figure 2-11 shows the locations of facilities that received and processed  
25 vermiculite from the mine in Libby.



**Figure 2-11. Nationwide distribution of Libby ore by county (in tons).** Data on the distribution of ore are based on approximately 80,000 invoices that EPA obtained from W.R. Grace that document shipments of vermiculite ore made from the Libby mine between 1964 and 1990. EPA tabulated this shipping information in a database.

Source: [U.S. GAO \(2007\)](#).

Workers in these expansion and processing facilities likely were exposed to LAA that was released during the processing operations. The 2008 ATSDR Summary Report ([ATSDR, 2008](#)) on the 28 Libby vermiculite expansion and processing facilities stated that in some cases household residents may also have been exposed by contact with vermiculite from the workers' clothes, shoes, and hair. Workers' personal vehicles likely contained vermiculite dust from facility emissions and from vermiculite that fell from their clothing and hair on the drive home after work.

Other residents living in communities near the expansion plants may also have been subjected to some of the same exposure pathways as for the Libby community. The 2008 ATSDR Summary Report observed that individuals in a community with a vermiculite expansion and processing plant could have been exposed by breathing airborne emissions from

the facility or by inhalation exposure to contaminants brought into the home on workers' clothing or from outdoor sources ([ATSDR, 2008](#)).

### 2.5.3. Exposure Pathways in Other Communities

Because expanded vermiculite from Libby was widely used in numerous consumer and construction products throughout the United States, even people not associated with Libby or other communities with expansion plants may also have the potential for exposure to LAA (see Table 2-2). Vermiculite was most notably used as attic insulation ([VAI; Versar, 2003](#)), as a soil amendment for gardening, fireproofing agent, and in the manufacturing of gypsum wallboard.

**Table 2-2. Air sampling results for asbestos from Zonolite vermiculite attic insulation (VAI) in three homes**

Activity	Personal samples		Area samples
	PCM <sup>a</sup> fibers/cc	TEM <sup>b</sup> PCMe, s/cc	TEM PCMe, s/cc
No activity	NS <sup>c</sup>	NS	<0.003
Cleaning items in the attic	1.54	<0.42	0.07
Cleaning storage area in the attic	2.87	2.58	0.47
Cutting a hole in the ceiling below the VAI	5.80	1.32	0.52
VAI removal (various methods)	2.9–2.5 <sup>d</sup>	0.98–10.3	0.53–1.47

<sup>a</sup>Air sampling results reported as fibers analyzed by PCM.

<sup>b</sup>Air sampling results reported as structures; PCMe as analyzed by TEM.

<sup>c</sup>NS—not sampled; personal samples were not taken for background levels.

<sup>d</sup>Range of results for three different removal methods (shop vacuum, homeowner method, and manufacturer-recommended method).

Source: [Ewing et al. \(2010\)](#).

### 3. FIBER TOXICOKINETICS

There are no published data on the toxicokinetics of Libby Amphibole asbestos (LAA).<sup>5</sup> However, to help inform the reader as to the expected toxicokinetics of LAA, this section contains a general summary description of the toxicokinetics of inhaled particles, with specific discussion of dosimetry differences for fibers. A more detailed discussion of fiber dosimetry is beyond the scope of this document and is reviewed elsewhere ([NIOSH, 2011](#); [ICRP, 1994](#)).

LAA includes fibers with a range of mineral compositions, including amphibole fibers primarily identified as richterite, winchite, and tremolite (see Section 2.2). Although the fiber size varies somewhat from sample to sample for LAA, a large percentage (~45%) is less than 5 µm long in bulk samples examined from the Libby mine site ([Meeker et al., 2003](#)). Limited data from air samples taken in the mill and screening plant at the Libby mine site also document a large percentage of fibers (including both respirable<sup>6</sup> fibers as well as fibers >3 µm long; see Section 4.1.1.2 and Table 4-3). Laboratory animal studies have examined the biologic response to LAA fibers from both raw samples ([Blake et al., 2007](#); [Pfau et al., 2005](#)) and rat respirable samples (<2.5 µm) following water elutriation ([Cyphert et al., 2012b](#); [Cyphert et al., 2012a](#); [Shannahan et al., 2012a](#); [Shannahan et al., 2012c](#); [Shannahan et al., 2012b](#); [Shannahan et al., 2012d](#); [Padilla-Carlin et al., 2011](#); [Shannahan et al., 2011a](#); [Shannahan et al., 2011b](#); [Shannahan et al., 2010](#); see Section 4.2, Appendix D). The mean fiber dimensions in the rat respirable fractions are in the range of length = 4.99 µm; width = 0.26 µm; aspect ratio ≥5:1 (as measured by TEM)<sup>7</sup>. The importance of the dimensions and density of fibers to their inhalation dosimetry—how they deposit and are subsequently cleared—is described below. Due to a lack of toxicokinetic data specific to LAA, these dosimetry mechanisms are discussed for inhaled fibers in general, with a specific focus on amphibole asbestos.

The main route of human exposure to mineral fibers is through inhalation. Inhaled dose of fibers to the respiratory tract tissue depends on the fiber concentration in the breathing zone, the physical (aerodynamic) characteristics of the fibers, the breathing mode (nose only or also oronasal), anatomical and physiological features of the respiratory tract (e.g., airway branching pattern and ventilation rate), and clearance mechanisms ([Oberdorster et al., 2002](#); [U.S. EPA, 1994a](#); [Oberdorster, 1991](#)). Ingestion is another pathway of human exposure and occurs mainly through the swallowing of material removed from the respiratory tract via mucociliary clearance or drinking water contaminated with asbestos, or eating, drinking, or smoking in

---

<sup>5</sup>The term “Libby Amphibole asbestos” is used in this document to identify the mixture of amphibole mineral fibers of varying elemental composition (e.g., winchite, richterite, tremolite, etc.) that have been identified in the Rainy Creek complex near Libby, MT. It is further described in Section 2.2.

<sup>6</sup>Respirable fibers are those that can penetrate into the alveolar regions and are defined by their aerodynamic diameter ( $d_a \leq 3 \mu\text{m}$ ; [NIOSH, 2011](#)).

<sup>7</sup>Detailed fiber dimension information for each study can be found in Appendix D when available.



asbestos-contaminated work environments ([Condie, 1983](#)). Handling asbestos can result in heavy dermal contact and exposure. Asbestos fibers can become lodged in the skin, producing a callus or corn—but generally with no serious health effects ([Lockey et al., 1984](#)). Because few studies have examined the deposition and clearance of fibers following ingestion or dermal exposure to fibers, the focus of this section is on the main route of exposure: inhalation.

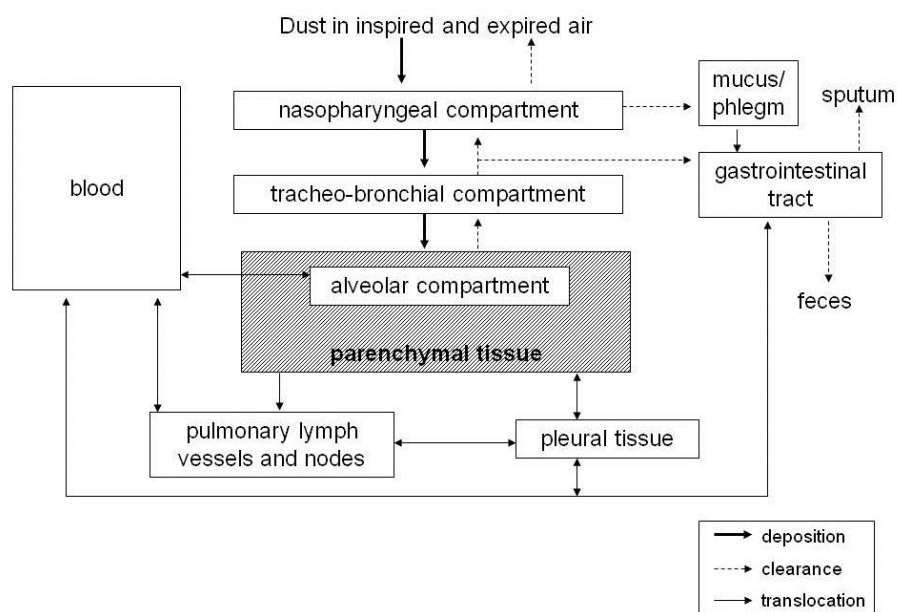
Studies useful for assessing the relationship between airborne fiber concentrations and respiratory disease must involve meaningful measurements of environmental exposure and an understanding of how to apply these measurements to the target tissue dose. Tissue dose is a more specific measure associated with disease development than is external dose. Many studies have examined the role of the physical and chemical characteristics of fibers in asbestos-induced disease in the lung and are reviewed in more depth elsewhere ([NIOSH, 2011](#); [ATSDR, 2001a](#); [Myojo and Takaya, 2001](#); [Witschi and Last, 1996](#); [Lippmann, 1990](#); [Merchant, 1990](#); [Yu et al., 1986](#); [Griffis et al., 1983](#); [Harris and Fraser, 1976](#); [Harris and Timbrell, 1975](#)). Factors influencing dose to other tissues in the body (e.g., pleura, peritoneum, stomach, and ovaries) are not as well known, but they are discussed below where data are available.

The principal components of inhaled fiber dosimetry in mammalian respiratory tract systems are (1) inhalability, (2) deposition on the epithelial surface, (3) clearance from the lung due to both physical (e.g., dissolution) and biological mechanisms (including mucociliary transport, phagocytosis, and translocation from the lung to other tissues [including the pleura]), and (4) elimination from the body (see Figure 3-1).

### **3.1. DEPOSITION OF FIBERS IN THE RESPIRATORY TRACT**

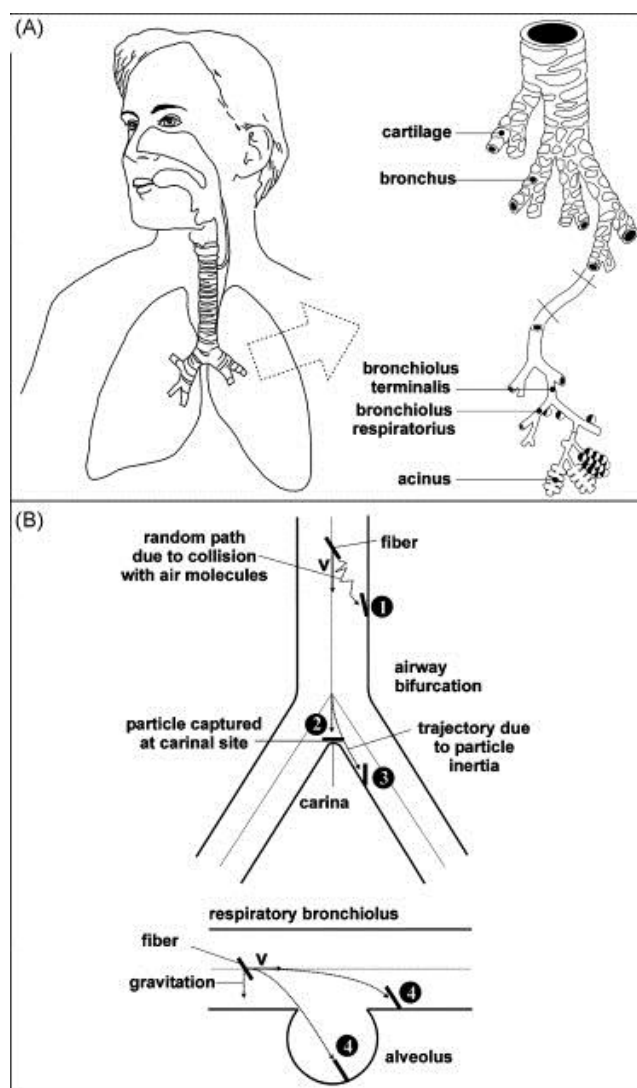
The respiratory tract encompasses the extrathoracic region (nasal passages, pharynx, and larynx), tracheobronchial region (the conducting airways [trachea bronchi, bronchioles]), and the gas-exchange or pulmonary region of the lung (respiratory bronchioles, alveolar ducts, and alveoli). Each region has unique anatomic and functional features, including dramatically different architecture, cell types, and defense mechanisms, that determine the dosimetry of inhaled agents in each region ([U.S. EPA, 1994a](#)). A full review of the anatomy and architecture of the respiratory tract is beyond the scope of this document but has been reviewed by the International Commission on Radiological Protection for its reference human respiratory tract model ([ICRP, 1994](#)). Figure 3-2 illustrates the major anatomical features of the human respiratory tract and mechanisms of fiber deposition.





**Figure 3-1. General scheme for fiber deposition, clearance, and translocation of fibers from the lung and gastrointestinal tract.** General scheme for fiber inhalation and deposition (heavy arrows), clearance (light dotted arrows), and translocation (light arrows). Diagram of [Bignon et al. \(1978\)](#) derived from International Commission on Radiological Protection lung model by the Task Group on Lung Dynamics ([Bates et al., 1966](#)), as cited in [ICRP \(1994\)](#).

Source: [ICRP \(1994\)](#).



**Figure 3-2. Architecture of the human respiratory tract and schematic of major mechanisms of fiber deposition.** Mechanisms of fiber deposition illustrated in panel (B) are as follows: (1) diffusion, (2) interception, (3) impaction, and (4) sedimentation.

Source: [Sturm and Hofmann \(2009\)](#).

Four major mechanisms determine fiber deposition: impaction, interception, sedimentation, and diffusion. Some authors also suggest electrostatic precipitation plays a role in fiber deposition, but no experimental data exist to verify its presence ([Sturm and Hofmann, 2009](#)). The relative contribution to deposition in each region of the respiratory tract depends on the fiber dimensions and density, breathing mode and ventilation rate, and the airway architecture of the species in question (e.g., rat vs. human). The deposition mechanisms and where these mechanisms are typically dominant in the human respiratory tract are described below (see Table 3-1).

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**Table 3-1. Factors influencing fiber deposition and clearance in the respiratory system**

Size of fiber (aerodynamic diameter)	Area of deposition in respiratory system	Predominant method of deposition	Mechanisms for fiber retention	Physical clearance	Dissolution	Target tissue for translocation
5–30 $\mu\text{m}$	Extrathoracic region (nasopharyngeal region, nasal passages, pharynx, larynx)	Impaction	Epithelial cell uptake	Mucous flow (mucociliary apparatus into gastrointestinal tract)  Macrophage: phagocytosis and transport	Not measured, although dissolution can occur; removal from mucous flow is fairly quick and likely predominant	Gastrointestinal tract  Nasal-associated lymphoid tissue, lymph system
1–5 $\mu\text{m}$	Thoracic Region (trachea, bronchial, and bronchiolar region)	Sedimentation, impaction, interception	Epithelial cell uptake	Mucociliary apparatus  Macrophage: phagocytosis and transport	Mucous  Macrophage	Gastrointestinal tract  Mucosa-associated lymphoid tissue, lymph system  Pleura
2 $\mu\text{m}$ or less	Gas-Exchange Region (respiratory bronchioles, alveolar ducts, alveoli)	Diffusion	Epithelial cell uptake  Translocation to other target tissues	Macrophage: phagocytosis and transport	Lung surfactant  Macrophage  Asbestos bodies	Gastrointestinal tract  Mucosa-associated lymphoid tissue, lymph system  Pleura

Source: Adapted from [Witschi and Last \(2001\)](#) in Casarett and Doull's Toxicology: The Basic Science of Poisons, 6<sup>th</sup> edition, p. 515.

- 1. Impaction:** The momentum of the fiber causes it to directly impact the airway surface as the airflow changes direction. This is the predominant method of deposition in the nasopharyngeal region where airflow is turbulent and in the larger conducting and bronchial airways at bifurcations where airflow is swift and directional changes are dramatic.
- 2. Interception:** A special case of impaction where the edge of the fiber touches the airway wall and is prevented from continuing along the airway. The longer a fiber, the higher its deposition by interception. This mechanism is important in the conducting airways (trachea and bronchi).

1       **3. Sedimentation:** Gravitational forces and air resistance cause fibers to settle out of  
2       the convective air stream onto the airway surface. For sedimentation to occur, air  
3       flow velocities must be low to allow the fiber to settle, so this is **a predominant**  
4       mechanism in the smaller conducting airways.

5       **4. Diffusion:** This method of deposition is predominant in the alveolar region where air  
6       movement is negligible. Diffusion occurs from interactions of the fibers with the  
7       movement of air molecules, and becomes important for particles  $<0.5\ \mu\text{m}$  in physical  
8       diameter.  
9

10       The aerodynamic properties of particulate aerosols, including fibers, are captured by  
11       characterizing the particle's aerodynamic diameter and its distribution, usually as the mass  
12       median aerodynamic diameter and geometric standard deviation (GSD) because aerosols tend to  
13       be log normally distributed. The aerodynamic diameter is the diameter of a unit density  
14       ( $1\ \text{g/cm}^3$ ) sphere that has the same gravitational settling velocity as the fiber of interest and the  
15       aerodynamic diameter derivation is based on fundamental laws governing fluid dynamics.  
16       However, characterizing a fiber by its aerodynamic diameter is dubious because as a fiber's  
17       aerodynamic properties depend, in addition to density, on both its length and width, as well as its  
18       orientation with respect to the convective airflow ([Asgharian and Anjilvel, 1998](#); [Cheng, 1986](#)).  
19       [Vincent \(2005\)](#) has proposed that fibers should be described by criteria that address both the  
20       aerodynamic properties that govern their regional deposition after inhalation and the biological  
21       effects and responses following deposition.

22       Computational models or algorithms to address fiber dosimetry typically derive an  
23       aerodynamic equivalent diameter ( $d_{eq}$ ) to remain consistent with the concept of aerodynamic  
24       diameter and provide some comparative context, and are based as well on fundamental fluid  
25       dynamics. Such formulae are mechanism specific (e.g., for impaction or sedimentation) and  
26       describe fibers as cylinders characterized by their density and length-to-width aspect ratio (beta),  
27       although the explicit bivariate distribution for a fiber aerosol can be described by the means and  
28       variances of the natural logarithms for length and width with correlations for their joint  
29       distribution ([Moss et al., 1994](#); [Cheng, 1986](#)). The latter is unfortunately not often done due to  
30       the lack of bivariate data when aerosols are sized in various experiments. The formulae to derive  
31        $d_{eq}$  must additionally account for the orientation of the fibers with respect to the convective  
32       airflow in the Stokes flow regime where it is necessary to describe the frictional forces  
33       encountered by an object (i.e., fiber or particle) in a fluid (i.e., air). For example, dynamic shape  
34       factors ( $\chi$ ) that relate the drag force of the cylindrical object to a sphere are derived for either  
35       perpendicular or parallel orientation with respect to the flow. Likewise, adjustments in these  
36       formulae are made for fiber orientation to the Cunningham slip correction factor which accounts  
37       for the relative velocity (or "slip") of gas molecules in air at the surface of embedded objects in  
38       that airflow. Additional assumptions regarding the orientation are typically used for each region  
39       of the respiratory tract, for example, random orientation for fibers in the upper airway subject to

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1   impaction or parallel orientation in the peripheral airways. Fibers enter the respiratory tract  
2   through the nasal and oral passages.

3       Deposition in the nasal and oral passages is mainly by impaction and diffusion. The  
4   nasal passage, from the nostril to the pharynx, serves as a filter for some fibers with diameters  
5   5–30  $\mu\text{m}$ . Clumps of fibers could deposit in these regions. Many animal species, including rats  
6   and mice, are obligate nose breathers so fibers pass only through the nasal passages and are  
7   always subject to nasopharyngeal filtering. Humans, monkeys, and dogs breathe both orally and  
8   nasally (oronasal). During oral respiration, larger fibers and clumps of fibers can bypass the  
9   filtering of the upper respiratory tract and are inhaled directly into the larynx/trachea, especially  
10   during exertion (e.g., exercise or work), thereby altering deposition as a result of the increased  
11   turbulence. This distinction is important when comparing results of inhalation studies conducted  
12   in different species.

13       Fibers in the lower respiratory tract deposit by combined mechanisms of impaction,  
14   interception, sedimentation, and diffusion. The relative contribution of each mechanism depends  
15   on the fiber characteristics and region-specific airway anatomy. Interception is heavily  
16   influenced by fiber length. Where the physical length of the fiber greatly exceeds the  
17   aerodynamic diameter, interception can be underpredicted by modeling the center of gravity of  
18   the fiber, since the length of the fiber will determine its propensity to intersect with the airway.  
19   Sedimentation is related to the mass of the fiber, as well as the aerodynamic diameter, but  
20   generally occurs at lower velocities in smaller airways. Diffusion occurs from interactions of the  
21   fibers with the movement of air molecules; this Brownian motion increases with decreasing fiber  
22   size ( $<0.5 \mu\text{m}$  diameter).

23       The conducting airways beyond the nasopharyngeal region serially bifurcate into airways  
24   of decreasing internal diameters that restrict the size of fibers deposited in these regions. Fiber  
25   length enhances bronchial deposition via interception, especially fibers exceeding lengths of  
26   10  $\mu\text{m}$  ([Sussman et al., 1991a, b](#)). The aerodynamic diameter of fibers that can deposit in the  
27   tracheobronchial region is in the range of 1–5  $\mu\text{m}$ . Fibers with an aerodynamic diameter of  
28    $<1 \mu\text{m}$  deposit in the bronchioles and the alveoli ([ICRP, 1994](#)). However, as reviewed in [Aust et al. \(2011\)](#),  
29   some studies have demonstrated that short fibers ( $<5 \mu\text{m}$ ) are present in substantially  
30   greater numbers than long fibers ( $>5 \mu\text{m}$ ) when examining the whole lung ([Churg, 1982](#)).  
31   Although information is limited on how fibers get to the pleura, fibers observed in pleural tissue  
32   from mesothelioma cases are more likely to be short ( $<5 \mu\text{m}$ ; [Suzuki et al., 2005](#)). These  
33   observations could be partly due to the increased deposition of smaller fibers or the breakage of  
34   larger fibers over time ([Bernstein et al., 1994](#); [Davis, 1994](#)).

35       Fibers with aerodynamic characteristics conducive to penetrating the peribronchiolar  
36   space and depositing in the alveoli may cause pulmonary fibrosis and other associated diseases.  
37   Regardless of shape, mineralogy, or concentration, the majority of fibers that are small enough to  
38   reach the alveoli are deposited at the alveolar duct bifurcations ([Brody and Roe, 1983](#)).

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Deposition is controlled by air flow characteristics and is greatest at the bifurcations that are closest to the terminal bronchioles ([Brody et al., 1981](#)). Furthermore, deposition in the bifurcations is consistent across laboratory animal species ([Warheit and Hartsky, 1990](#)). Alveoli deposition is limited when fiber length approaches 40  $\mu\text{m}$  ([Morgan et al., 1978](#)). However, alveolar deposition of fibers can occur with high aspect ratios and lengths ranging from <1  $\mu\text{m}$  to >200  $\mu\text{m}$  long ([Morgan et al., 1978](#)). All fibers having an aerodynamic diameter less than approximately 2  $\mu\text{m}$ , which includes LAA, meet the physical criteria necessary for deposition in the deeper regions of the respiratory tract at the level of the terminal bronchioles or alveoli.

### 3.2. CLEARANCE MECHANISMS

Once fibers deposit on the surface of the respiratory tract, they may be removed (cleared) in several ways—including physical clearance, dissolution, phagocytosis, encapsulation, or transcytosis. Some of these mechanisms, such as dissolution of the fibers or removal via the mucociliary apparatus, can result in the fibers being cleared from the body (see Figure 3-1). Other clearance mechanisms may remove fibers from the surface of the respiratory tract but result in transport of the fibers to different locations or tissues by translocation. Translocation of fibers from the terminal bronchioles and alveoli into the peribronchiolar space, lymph nodes, and pleura has been implicated in disease causation (e.g., [pleural plaques](#), [mesothelioma](#); [Dodson et al., 2001](#)). In human studies, the translocation of asbestos fibers following inhalation has been observed to varying degrees throughout the pulmonary and extrapulmonary tissues of the respiratory system ([Dodson et al., 2005](#); [Dodson et al., 2001](#); [Kohyama and Suzuki, 1991](#); [Suzuki and Kohyama, 1991](#); [Armstrong et al., 1988](#)), as well as to other organs, including the brain, kidney, liver ([Miserocchi et al., 2008](#)), and ovaries ([Langseth et al., 2007](#)). In many cases, the type of fiber is not defined, and the individual exposure information not available. Fibers that are not cleared can remain at the epithelial surface or enter the parenchymal tissue of the lung. Retention of fibers in the human thoracic region generally shows two distinct phases. The first, on the order of 24 hours, is considered to represent mucociliary clearance to the gastrointestinal tract from the conducting airways and bronchioles; the second represents clearance from the alveolar region ([ICRP, 1994](#)).

[Berry \(1999\)](#) provided a review of the animal toxicity literature specifically for fiber clearance. There are limited data on clearance patterns based on autopsy studies in humans. Two studies estimated clearance half-life for amphibole asbestos (~20 years) as compared with chrysotile asbestos (~10 years; [Finkelstein and Dufresne, 1999](#); [Churg and Vedal, 1994](#)); in evaluating the data on lung fiber burden, [Berry et al. \(2009\)](#) estimated the range of the half-life for crocidolite to be between 5 and 10 years. Generally, studies have focused on determining the size and type of asbestos retained in specific tissues ([Suzuki et al., 2005](#); [McDonald et al., 2001](#); [Suzuki and Yuen, 2001](#); [Dumortier et al., 1998](#); [Gibbs et al., 1991](#); [Dodson et al., 1990](#)) and do



not discuss changes in fiber content since exposure. [Sebastien et al. \(1980\)](#) concluded that lung fiber burden could not be used as an accurate reflection of pleural fiber burden.

### **3.2.1. Physical and Physicochemical Clearance of Fibers**

Different mechanisms of physical and physicochemical clearance of fibers depends on the fiber size, physicochemical characteristics, and site of deposition ([IOM, 2006](#)). Physical clearance includes mechanical mechanisms, including transport via the mucociliary apparatus, macrophage uptake, and translocation. Fibers can also translocate due to physical forces associated with the mechanics of respiration (e.g., [expansion, contraction of the rib cage; Davis, 1989](#)). Physicochemical clearance of fibers includes dissolution and breakage of fibers.

#### **3.2.1.1. Mechanical Reflex Mechanisms**

Fibers deposited in the nasal passages can be removed by all clearance mechanisms. When breathing occurs through the nose, many fibers are filtered by the turbulent airflow in the nasal passages, impacting against the hairs and nasal turbinates, as well as becoming entrained in mucus in the upper respiratory tract where they can be subsequently removed by mucociliary action (described below) or reflexive mechanical actions such as coughing or sneezing. Dissolution can also occur in this region, especially for soluble fibers.

#### **3.2.1.2. Mucociliary Clearance**

Physiological mechanisms include mucociliary escalator movement and how specific cells or mechanisms in various regions of the respiratory tract respond and attempt to detoxify or remove inhaled fibers.

The mucociliary escalator removes fibers through ciliary movement of the sticky mucus lining ([Wanner et al., 1996; Churg et al., 1989](#)). Fibers removed from the conducting airways through this mechanism are coughed out or swallowed and enter the digestive tract where they may adversely affect the gastrointestinal tissue, enter the blood stream, or be excreted. Clearance of fibers via mucociliary action is usually complete within hours or days ([Albert et al., 1969](#)).

The mucociliary escalator extends only down to the level of the terminal bronchioles and not to the alveoli. Thus, the particles deposited in the alveolar region of the lung cannot be cleared through this process. Particles can reach the mucociliary escalator from the alveoli either by way of surface fluids that are drawn onto the mucociliary escalator by surface tension or by travelling through lymphatic channels that empty onto the escalator at bronchial bifurcations.

Although ingestion is a potential route of exposure due to subsequent swallowing of material from the mucociliary escalator, limited research has examined clearance (e.g., translocation) of fibers following ingestion, and no clearance studies are available specific to

1 LAA. An early study to examine the tissue response to asbestos fibers is not truly representative  
2 of a natural ingestion exposure, as the researchers directly injected a suspension of amosite fibers  
3 into the duodenal wall ([Meek and Grasso, 1983](#)). This study, however, also examined oral  
4 ingestion of amosite in healthy animals and those with gastrointestinal ulcers to determine  
5 whether translocation of fibers occurs through ulcers. Following injection of amosite,  
6 granulomatous lesions were observed. Ingestion of the same material resulted in no such lesions  
7 or in any other histopathological changes in either healthy or rats compromised with ulcers.  
8 Thus, no translocation was observed from either the healthy or the compromised rat  
9 gastrointestinal tracts in this study. [Truhaut and Chouroulinkov \(1989\)](#) examined the effects of  
10 chrysotile and crocidolite ingestion in Wistar rats. No translocation was observed. No further  
11 studies have been found on clearance or translocation of fibers from the gastrointestinal tract.

12 Some fibers are not cleared from the respiratory tract, leading to an accumulation with  
13 time ([Case et al., 2000](#); [Finkelstein and Dufresne, 1999](#); [Jones et al., 1988](#)). The fibers that  
14 remain in the conducting airways and alveolar regions may undergo a number of processes  
15 including translocation, dissolution, fragmentation, splitting along the longitudinal axis, or  
16 encapsulation with protein and iron. Available data indicate prolonged clearance from the  
17 thoracic region of long ( $>5\ \mu\text{m}$ ) or short amphibole fibers ([Coin et al., 1994](#); [Tossavainen et al.,](#)  
18 [1994](#)).

19 The prolonged clearance times for long amphibole fibers have led some investigators to  
20 conclude that long fibers ( $>5\ \mu\text{m}$ ) rather than short amphibole fibers (i.e., LAA) are predominant  
21 in the cause of disease due to their persistence in the lung ([Mossman et al., 2011](#); [ATSDR,](#)  
22 [2003](#)). However, others argue that fibers of all lengths induce pathological responses and urge  
23 caution in excluding, based on length, any population of fibers from consideration as possibly  
24 contributing to the disease process ([Aust et al., 2011](#); [Dodson et al., 2003](#)). Respirable-sized  
25 fibers of LAA have been identified in air samples from Libby, MT and in airborne fibers  
26 suspended from both Libby vermiculite ore and in the exfoliated product from that ore (length  
27 range from  $1\ \mu\text{m}$  to  $20\text{--}30\ \mu\text{m}$ , with average length of  $7\ \mu\text{m}$ ; width range from  $0.1\text{--}2\ \mu\text{m}$ , with  
28 average of  $0.5\ \mu\text{m}$ ; see for details Appendix B and Appendix C). Based on fibers counted by the  
29 TEM analytical method (ISO 10312), the majority of counted fibers are respirable (see  
30 Figure 2-10).

### 31 32 **3.2.1.3. Phagocytosis by Alveolar Macrophages**

33 The principal clearance pathway for short, insoluble fibers deposited in the alveoli is  
34 through phagocytosis by macrophages. Durable fiber impaction in the deeper region of the  
35 respiratory tract stimulates activation of alveolar macrophages. In vitro and in vivo studies  
36 clearly indicate that macrophage cells play a role in the translocation of fibers ([Dodson et al.,](#)  
37 [2000a](#); [Castranova et al., 1996](#); [Brody et al., 1981](#); [Bignon et al., 1979](#)). These studies  
38 demonstrated the presence of asbestos fibers in cell cytoplasm where the fibers can be

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1 transported in association with cytoskeletal elements to the proximity of the cell nucleus.  
2 Alveolar macrophages that have phagocytized insoluble fibers migrate to the bronchoalveolar  
3 junctions where they are removed via the mucociliary escalator ([Green, 1973](#)). Alternatively,  
4 alveolar macrophages that have phagocytized insoluble fibers can also migrate through the  
5 epithelial wall into the interstitial space and enter the lymphatic system ([Green, 1973](#)).

6 A number of processes can disrupt the normal phagocytic function of alveolar  
7 macrophages, such as the overwhelming of phagocytosis and the mucociliary escalator by an  
8 excessive number of particles from a decrease in the rate of mucociliary clearance (often termed  
9 “overload”), or the attempted phagocytosis of fibers with lengths that exceed the dimensional  
10 capacity of the macrophage ([15–20 μm depending on species; often termed “frustrated  
11 phagocytosis”; NIOSH, 2011](#)). Any of these processes can induce inflammatory and fibrogenic  
12 responses. Limited inhalational laboratory animal studies exist at concentrations of fibers below  
13 overload occurred; therefore information is insufficient to determine mechanisms of  
14 inflammation at lower doses as reviewed in [Mossman et al. \(2011\)](#).

15 Fibers that are too large to be easily engulfed by the alveolar macrophage can stimulate  
16 the formation of “asbestos bodies.” Asbestos bodies are fibers that become coated with proteins,  
17 iron, and calcium oxalate as a result of prolonged residence in the lung where they can remain  
18 throughout an individual’s lifetime. Due to their iron content, histological stains for iron have  
19 long been used to identify them in tissue; thus, they are sometimes called “ferruginous bodies.”  
20 The mechanisms that result in the formation of asbestos bodies are poorly understood, although  
21 most appear to be formed around amosite fibers ([Dodson et al., 1996](#)). The iron in the coating is  
22 derived from the asbestos fiber, cells, or medium surrounding the fiber and can remain highly  
23 reactive ([Lund et al., 1994](#); [Ghio et al., 1992](#)). Asbestos bodies comprise a minor portion of the  
24 overall fiber burden of the lung. These fibers may or may not participate directly in asbestos  
25 disease once the fiber is fully coated. For instance, the presence of iron in the coating could  
26 provide a source for catalysis of reactive oxygen species (ROS) similar to that observed with  
27 fibers.

#### 28 29 **3.2.1.4. Epithelial Transcytosis**

30 In addition to phagocytosis by alveolar macrophages, fibers deposited on Type I alveolar  
31 epithelial cells may also be subjected to transcytosis with subsequent sequestration to the  
32 alveolar interstitium ([Sturm, 2011](#)). Fiber length would play a key role in this aspect of  
33 clearance, much as described above for phagocytosis by alveolar macrophages.

#### 34 35 **3.2.1.5. Translocation**

36 Translocation represents the movement of intact fibers along the epithelial surface  
37 towards the terminal bronchiole, or into and through the epithelium. Translocation typically

occurs via drainage of the alveolar macrophages to the lymphatics, but transcytosis of fibers by type I alveolar epithelial cells can also result in transport to the interstitium. The relative contribution of a specific mechanism and translocation route depends both on fiber characteristics and the tissue of deposition. Fiber translocation depends on the fibers' physicochemical characteristics, including two-dimensional size (length and width), durability, solubility, and reactivity. This translocation is aided by high durability and an inflammation-induced increase in permeability but could be hindered by fibrosis.

Translocation of fibers to extrapulmonary tissues was reported in multiple studies; however, the precise mechanism is still unknown. This process was recently reviewed by [Miserocchi et al. \(2008\)](#). Fibers were identified in all of the analyzed locations, including pleural plaques and mesothelial tissue (i.e., pleural or peritoneal) in miners, brake workers, insulation workers, and shipyard workers ([Roggli et al., 2002](#); [Dodson et al., 2000b](#); [Churg, 1994](#); [Kohyama and Suzuki, 1991](#)). However, amphibole fibers were less prevalent than chrysotile fibers in the pleura and mesothelial tissues ([Kohyama and Suzuki, 1991](#); [Sebastien et al., 1989](#); [Armstrong et al., 1988](#); [Churg, 1988](#); [Bignon et al., 1979](#)). Confocal microscopic observations of rats inhalationally exposed to amosite fibers showed that fibers were present on the parietal pleural surface 7 days postexposure and more than twofold thickening of the pleural wall was noted ([Bernstein et al., 2011](#)). [Bignon et al. \(1979\)](#) also reported increased amphibole fibers in the lymph nodes. Conflicting results from an inhalational rat study do not indicate any evidence of fiber translocation from the central to peripheral compartments, although this could be due to the short duration of the study ([29 days postexposure](#); [Coin et al., 1992](#)).

Few studies have examined the size distribution of fibers translocated to specific tissues. For example, one early study suggested that the longer amphibole fibers predominate in the lung ([Sebastien et al., 1980](#)); other studies showed that the fiber-length distribution was the same by fiber type regardless of location ([Kohyama and Suzuki, 1991](#); [Bignon et al., 1979](#)). [Dodson et al. \(1990\)](#) observed that the average-length fiber found in the lung (regardless of type) was longer than those found in the lymph nodes or plaques. Most fibers at all three sites were short (<5 µm). Similar results were observed in a later study by this group (i.e., [Dodson et al., 2000b](#)) which examined tissue from 20 individuals with mesotheliomas, most with known nonoccupational asbestos exposures.

Transplacental transfer of both asbestos (tremolite, actinolite, and anthophyllite) and nonasbestos fibers occurs in humans, as measured in the placenta and in the lungs of stillborn infants ([Haque et al., 1998](#); [Haque et al., 1996](#); [Haque et al., 1992](#); [Haque and Kanz, 1988](#)). It is hypothesized that maternal health might influence the translocation of fibers, as some of the mothers had preexisting health conditions (e.g., [hypertension, diabetes, or asthma](#); [Haque et al., 1992](#)). This group also measured transplacental translocation in a mouse study and observed early translocation of crocidolite fibers (Union for International Cancer Control [UICC]) through the placenta in animals exposed via tail-vein injection ([Haque et al., 1998](#)). These studies did not

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1 evaluate the source or levels of exposure, only the presence of fibers in the body during early life  
2 stages in mice and humans.

#### 3 4 **3.2.1.6. Dissolution and Fiber Breakage**

5 Dissolution, or the chemical breakdown of fibers, is another method of physical removal  
6 of fibers from the respiratory tract. This process varies, depending on the solubility and  
7 chemical composition of the fibers, as well as the physiological environment. Dissolution can  
8 occur in the lung's extracellular fluids or in the macrophage phagolysosome; the former can  
9 make the breakdown products available for uptake into the blood. Studies performed in vitro to  
10 determine dissolution rate of fibers attempt to mimic the extracellular lung fluids and  
11 macrophage-phagolysosome system to understand the length of time that fibers remain in the  
12 system ([Rendall and Du Toit, 1994](#)). Fibers can also be physically diminished through splitting  
13 or breakage. These smaller fragments are then more easily removed by phagocytosis or  
14 translocation.

### 15 16 **3.3. DETERMINANTS OF TOXICITY**

17 Multiple determinants of fiber toxicity, including dimension (length, diameter, aspect  
18 ratio, and surface area), chemical characteristics (solubility, charge, and surface chemistry) and  
19 durability (dissolution, breakage) have been studied relative to specific biological responses to  
20 fibers and recently reviewed ([Aust et al., 2011](#); [Broaddus et al., 2011](#); [Bunderson-Schelvan et al.,  
21 2011](#); [Case et al., 2011](#); [Huang et al., 2011](#); [Mossman et al., 2011](#)).

#### 22 23 **3.3.1. Dosimetry and Biopersistence**

24 The dosimetry factors discussed in the previous sections are major factors influencing  
25 toxicity as the initial deposition sites in the respiratory tract tissues determine the subsequent  
26 clearance mechanisms ([Brain and Mensah, 1983](#)); solubility and composition also influence the  
27 biopersistence of fibers once deposited ([Maxim et al., 2006](#); [ILSI, 2005](#)). Thus, fiber toxicity has  
28 been associated with dose, density, dimensions, and durability, and likely involves a combination  
29 of these and other factors. To the extent that a fiber and its composition are resistant to the  
30 clearance mechanisms described in Section 3.1., biopersistence becomes a determinant of toxic  
31 response. The degree of fiber durability determines the retained dose at the site of deposition  
32 and likely plays a role in chronic inflammation, fibrosis, and lung burden following chronic  
33 exposure to fibers. Biopersistence is influenced by fiber characteristics such as size (length,  
34 width) and chemistry. Hesterberg et al. ([1998b](#); [1998a](#)) observed that, in general, increased in  
35 vitro solubility decreases in vivo biopersistence. Several supporting studies reported increased  
36 levels of crocidolite, tremolite, and amosite in respiratory diseases (asbestosis, mesothelioma)  
37 compared to chrysotile and controls ([Churg and Vedal, 1994](#); [Churg et al., 1993](#)) and found that

chrysotile has a lower biopersistence than amphibole fibers ([Churg and Vedal, 1994](#); [Wagner et al., 1974](#)). The role of fiber size in biopersistence was examined by [Bernstein et al. \(2004\)](#) who found that the clearance half-time of longer fibers (>20 µm; 1.3 days) was less than that of shorter fibers (<5 µm; 23 days) for one form of chrysotile. Biopersistence is the basis of short-term in vitro testing required for new fibers introduced to the market ([Maxim et al., 2006](#); [ILSI, 2005](#)).

Fiber burden analysis of human tissues is frequently employed to determine the presence of specific fiber type and size in disease ([Aust et al., 2011](#); [Case et al., 2011](#)). The majority of these studies have focused on lung tissue, but studies have also examined fiber burden in other tissues of interest, including lymph nodes and pleural tissues ([Bunderson-Schelvan et al., 2011](#); [Dodson and Atkinson, 2006](#); [Dodson et al., 2001](#); [Dodson et al., 2000b](#); [Boutin et al., 1996](#)). While informative, this analysis has some limitations, including differences in methodologies that hinder comparisons between laboratories, as well as potential cross contamination with other tissues. A further limitation of fiber burden analysis is that it is generally performed on tissue digests, making it difficult to show fiber dimensions at specific tissue locations. The use of TEM analysis can determine length and width of fibers found in tissues from exposed individuals.

### **3.3.2. Biological Response Mechanisms**

Although numerous studies have examined specific mechanisms of toxicity for many different fiber types, the results often are contradictory or do not account for dosimetry, and thus, only limited conclusions can be made for fibers in general. Research has focused mainly on the role of length, width, and durability in fiber toxicity. The relative contribution of dimensions and chemistry that drive the toxicity of the fibers remains poorly understood due to the difficulty in experimentally evaluating each determinant independently. Further, as can be appreciated from an evaluation of Table 3-2, the determinants of toxicity induce various toxic responses that represent interrelated biological activities (e.g., chronic inflammation, oxidative stress, and genotoxicity), making a clear causal relationship or relative contribution of any individual endpoint difficult to determine. The information described below focuses on in vivo studies that have examined determinants of amphibole fiber toxicity for some major specific biological responses. A more detailed discussion of potential tissue response mechanisms can be found in the mode of action (MOA) section (see Section 4). Table 3-2 summarizes the major determinants of toxicity as fiber-host interactions along the continuum of source-exposure-dose-response used for risk assessment.

**Table 3-2. Determinants of fiber toxicity**

<b>Biological activity</b>	<b>Length</b>	<b>Width</b>	<b>Mineralogy</b>	<b>Biopersistence</b>	<b>Morphology</b>	<b>Density</b>	<b>Surface area</b>	<b>Surface chemistry (charge, metal ions)</b>
<b>ROS production</b>	+++	++	+++	+++				+++
<b>Genotoxicity (direct or indirect)</b>	+++	++	++	++	++		+	++
<b>Inflammation</b>	++	++	++	+++		++	+++	+++
<b>Carcinogenesis</b>	++	++		++			+	+

Note: This table describes the potential role for various fiber determinants in biological activity. Level of confidence is based on available literature for all fiber types. (+++, suggested role with substantial data support; ++, suggested role but data not conclusive; +, suggested role but insufficient data). Level of confidence based on recent literature reviews ([Aust et al., 2011](#); [Case et al., 2011](#); [Mossman et al., 2011](#); [ATSDR, 2003](#)).

1 Studies have also associated exposure to fibers to other biological activities, including  
2 autoimmune effects and pulmonary function impacts. Limited studies have examined the role of  
3 specific fiber determinants in autoimmune disease or pulmonary function, and therefore these  
4 endpoints are not discussed in this section.

### 6 **3.3.2.1. Inflammation and Reactive Oxygen Species (ROS) Production**

7 Inflammation is an important biologic response to fibers and is related to multiple  
8 pathways following exposure. Fiber exposure leads to ROS production, which in turn has been  
9 shown to increase the activation of inflammatory and immune signaling pathways ([Mossman et  
10 al., 2011](#)). Inflammation often occurs at the site of fiber deposition; therefore those fiber  
11 characteristics that play a role in fiber deposition (e.g., length and width) will also play a role in  
12 chronic inflammation. Further, those additional characteristics that lead to ROS production (e.g.,  
13 surface chemistry) may also contribute to induction of chronic inflammation. Acute  
14 inflammation in response to asbestos further contributes to chronic inflammation with the  
15 activation of signaling pathways (e.g., mitogen-activated protein kinase[MAPK]) that lead to the  
16 release of proinflammatory cytokines.

17 Increased ROS production is hypothesized to result from frustrated phagocytosis and  
18 activation of signaling pathways in various cell types or through iron catalysis, which may  
19 account for the differential induction of ROS due to variable intrinsic or acquired iron by  
20 different fibers ([Aust et al., 2011](#)). Either indirect or direct ROS release following exposure to  
21 fibers may in turn lead to increased damage to DNA or other biological molecules.

### 23 **3.3.2.2. Genotoxicity**

24 Genotoxicity (including mutagenicity) from exposure to fibers likely involves multiple  
25 pathways, and the role of specific fiber determinants is not completely understood. This  
26 genotoxicity is generally described as direct (e.g., fiber interference with spindle formation) or  
27 indirect (e.g., reactive oxygen species production). A recent review by [Huang et al. \(2011\)](#)  
28 examines the role of fiber determinants in genotoxicity and mutagenicity in detail. Briefly,  
29 research studies designed to examine the role of fiber dimensions or surface characteristics in  
30 genotoxicity are limited, and are mainly in vitro work. In general, fiber dimensions are expected  
31 to play a role in genotoxicity. Long thin fibers are associated with interference with the spindle  
32 apparatus during mitosis, as well as increased ROS/reactive nitrogen species (RNS) production  
33 through frustrated phagocytosis, which in turn may lead to increased genotoxicity. Similarly,  
34 increased iron associated with fibers may also lead to increased ROS/RNS production and  
35 increased genotoxicity.



### 3.3.2.3. Carcinogenicity

The work by [Stanton et al. \(1981\)](#) examined fiber type and dimension in relation to carcinogenicity in an animal model resulting in the “Stanton Hypothesis” that identifies fiber size as a major determinant of toxicity. This study focused on amphibole asbestos due to difficulties in measurements of chrysotile ([Stanton et al., 1981](#)). However, this hypothesis was formulated from the results of studies where fibers were imbedded in agar and implanted against the pleura, thereby inducing sarcomas in rats ([Stanton et al., 1981](#); [Stanton and Wrench, 1972](#)). The results of these studies led Stanton and colleagues to state that “carcinogenicity of fibers depended on dimension and durability rather than on physicochemical properties” ([Stanton et al., 1981](#)). However, the study design did not allow for the influence of pulmonary clearance mechanism, as the implant was made to the outer pleural tissue. Additionally, it is unknown how the dissolution and clearance of fibers in the agar influenced findings. All fibers tested (including mineral wool and fiber glass) induced sarcomas. While these studies showed high correlation between disease and longer ( $>8\ \mu\text{m}$ ), thinner ( $<0.25\ \mu\text{m}$ ) fibers, high correlations were noted in other size categories as well. The ability of these studies to define the dimensional aspects of fibers (length and width cutoffs) that determine toxicity is an overinterpretation of the data, since major aspects of toxicokinetics and biological activity in the lung tissue are not included in the experimental design. Additionally, these studies do not rule out the potential role of shorter ( $<4\ \mu\text{m}$ ) and wider ( $>1.5\ \mu\text{m}$ ) fibers ([Stanton et al., 1981](#)). This latter point was further confirmed by Pott et al. ([1987](#); [1974](#)) and who showed that shorter fibers ( $<10\ \mu\text{m}$  in length) could also induce tumors in rats following intraperitoneal injection. Although informative, both of these study designs bypass normal physiological clearance and deposition mechanisms that would be observed following inhalation of fibers, an important consideration when comparing these types of studies.

[Suzuki et al. \(2005\)](#) also examined the role of fiber dimensions in mesothelioma, but through direct pathological evidence from human mesothelioma tissue. Fibers were identified by high-resolution analytical electron microscopy from digested or ashed lung and mesothelial tissues samples taken from 168 cases of malignant mesothelioma. Their results were that the majority of fibers (89%) were shorter than or equal to  $5\ \mu\text{m}$  in length, and generally (92.7%) smaller than or equal to  $0.25\ \mu\text{m}$  in width, which is contrary to the “Stanton Hypothesis.” However, this study is also not without interpretation challenges, as the digestion and ashing process may lead to shorter fibers, or any longer fibers may have broken down by dissolution or fiber breakage.

Analyses of fiber dimensions in exposed humans have not led to any clear determinants of toxicity for fibers in general. [Lippmann \(1990\)](#) correlated fiber length with disease status in exposed humans and concluded that asbestosis was associated with short ( $>2\ \mu\text{m}$ ), thick fibers,  $>0.15\ \mu\text{m}$ , ); mesothelioma with longer ( $>5\ \mu\text{m}$ ), thinner fibers,  $<0.1\ \mu\text{m}$ ); and lung cancer with longer ( $>10\ \mu\text{m}$ ), thicker fibers,  $>0.15\ \mu\text{m}$ ). Throughout the years, some laboratory animal

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1 studies have demonstrated a role for longer ( $>20\ \mu\text{m}$ ) and thinner ( $<0.3\ \mu\text{m}$ ) fibers in lung cancer  
2 ([Berman et al., 1995](#)) or mesothelioma ([Miller et al., 1999](#)), while yet other laboratory animal  
3 studies have suggested a role for shorter structures ( $<0.5\text{--}5\ \mu\text{m}$ ) in disease based in part on  
4 increased numbers in dust clouds and in lung retention ([Dodson et al., 2003](#)). Some human  
5 epidemiology studies have supported the role of longer ( $>20\ \mu\text{m}$ ), thinner ( $<0.3\ \mu\text{m}$ ) fibers in  
6 lung cancer ([Loomis D, 2009](#); [Berman and Crump, 2008](#); [Dement JM, 2008](#)). However, these  
7 results have not been confirmed in other studies and are in some cases contradicted. For  
8 instance, [Churg and Vedal \(1994\)](#) did not find an association between fiber length and cancer.  
9 However, [McDonald et al. \(2001\)](#) observed that shorter fibers ( $<6\ \mu\text{m}$ ) were more abundant in  
10 diseased tissues than longer fibers ( $>10\ \mu\text{m}$ ), and [Dodson et al. \(1997\)](#) concluded that all sized  
11 fibers are associated with increased mesothelioma risk. More recently, [Berman \(2011\)](#)  
12 performed a quantitative analysis of previous studies and demonstrated that differences in  
13 biological potency among various amphibole fiber types may be due to differences in their  
14 dimensions, particularly fiber length. In all cases, the analytical methods used need to be  
15 carefully described in order to draw any conclusions across studies.

### 17 **3.4. FIBER DOSIMETRY MODELS**

18 Modeling of fiber deposition has been examined for various fiber types (e.g., refractory  
19 ceramic fibers, chrysotile asbestos; [Sturm, 2009](#); [Zhou et al., 2007](#); [Lentz et al., 2003](#); [Dai and](#)  
20 [Yu, 1998](#); [Yu et al., 1997](#); [Coin et al., 1992](#)), but not for LAA. In general, the pattern of  
21 deposition for fibers is expected to have some similarities to the well-studied deposition pattern  
22 for particles that are essentially spherical ([reviewed in ICRP, 1994](#)). For example, the multipath  
23 particle dosimetry model ([Brown et al., 2005](#); [Jarabek et al., 2005](#)) uses information on the  
24 physical properties of the particles (length, width [also called bivariate distribution] and density),  
25 the anatomy and architectural features of the airways, airflow patterns that influence the amount  
26 and the location of the deposition of the particles, and the dissolution and clearance mechanisms  
27 that are operative to estimate the retained dose in the target tissue. The site of fiber deposition  
28 within the respiratory tract has implications related to lung retention and surface dose of fibers.  
29 It should be noted that differences in airway structure and breathing patterns across life stages  
30 (i.e., children, adults) change the depositional pattern of differently sized fibers, possibly altering  
31 the site of action and causing differential clearance and health effects (see Section 4.7).

### 33 **3.5. SUMMARY**

34 Although oral and dermal exposure to fibers does occur, inhalation is considered the main  
35 route of human exposure to mineral fibers, and therefore, it has been the focus of more fiber  
36 toxicokinetic analyses in the literature. Similar to other forms of asbestos, exposure to LAA is  
37 presumed to be through all three routes of exposure; however, this assessment specifically



1 focuses on the inhalation pathway of exposure. Generally, fiber deposition in the respiratory  
2 tract is fairly well defined based on fiber dimensions and density, although the same cannot be  
3 said for fiber translocation to extrapulmonary sites (e.g., pleura). The deposition location within  
4 the pulmonary and extrapulmonary tissues plays a role in the clearance of the fibers from the  
5 organism.

6 Fiber clearance from the respiratory tract can occur through physical and biological  
7 mechanisms. Limited mechanistic information is available on fiber clearance mechanisms in  
8 general, and no information specific to clearance of LAA fibers is available. Fibers have been  
9 observed in various pulmonary and extrapulmonary tissues following exposure, suggesting  
10 translocation occurs to a variety of tissues. Studies have also demonstrated fibers may be cleared  
11 through physical mechanisms (coughing, sneezing) or through dissolution of fibers.

12 Multiple fiber characteristics (e.g., dimensions, density, and durability) play a role in the  
13 toxicokinetics and toxicity of fibers. There is extensive literature examining a variety of fiber  
14 determinants and their role in disease, with a focus on fiber length, width, and durability;  
15 however, these studies are often contradictory, making conclusions difficult for fibers in general.  
16 This is in part due to the variety of fibers analyzed, inadequate study design, and/or lack of  
17 information on fiber dimensions in earlier studies. However, due to the importance in  
18 understanding the role of these fiber determinants in the biological response, careful attention has  
19 been paid to these fiber characteristics when analyzing research studies on LAA and asbestiform  
20 tremolite, an amphibole fiber that comprises part of LAA (see Appendix D). No toxicokinetic  
21 data are currently available specific to LAA, tremolite, richterite, or winchite. When available,  
22 fiber characteristic data are presented in the discussion of each study in relation to the toxic  
23 endpoints described.

#### 4. HAZARD IDENTIFICATION OF LIBBY AMPHIBOLE ASBESTOS

This section discusses the available data derived from studies of people exposed to Libby Amphibole asbestos (LAA),<sup>8</sup> either at work or in the community, and from various laboratory studies. The effects in humans (e.g., parenchymal damage and pleural thickening, lung cancer, and mesothelioma) are supported by the available LAA experimental animal in vivo and laboratory in vitro studies. The health effects from tremolite exposure, one of the constituent minerals of LAA, reported in both human communities and laboratory animals, are consistent with the human health effects reported for LAA. Studies examining the health effects of exposure to winchite or richterite alone were not available in the published literature. The review presents noncancer and cancer health effects observed from exposures to LAA.

##### 4.1. STUDIES IN HUMANS—EPIDEMIOLOGY

The LAA epidemiologic database includes studies based in occupational settings and community-based studies of workers, family members of workers, and others in the general population. Occupational epidemiology studies have been conducted at two worksites where workers were exposed to LAA: the vermiculite mine and mill at the Zonolite Mountain operations near Libby, MT, and a manufacturing plant using the vermiculite ore in Marysville, OH. Community-based studies have also been conducted among residents around Libby, MT, ([ATSDR, 2001b, 2000](#)) and in an area around a manufacturing plant producing vermiculite insulation in Minneapolis, MN ([Alexander et al., 2012](#)).

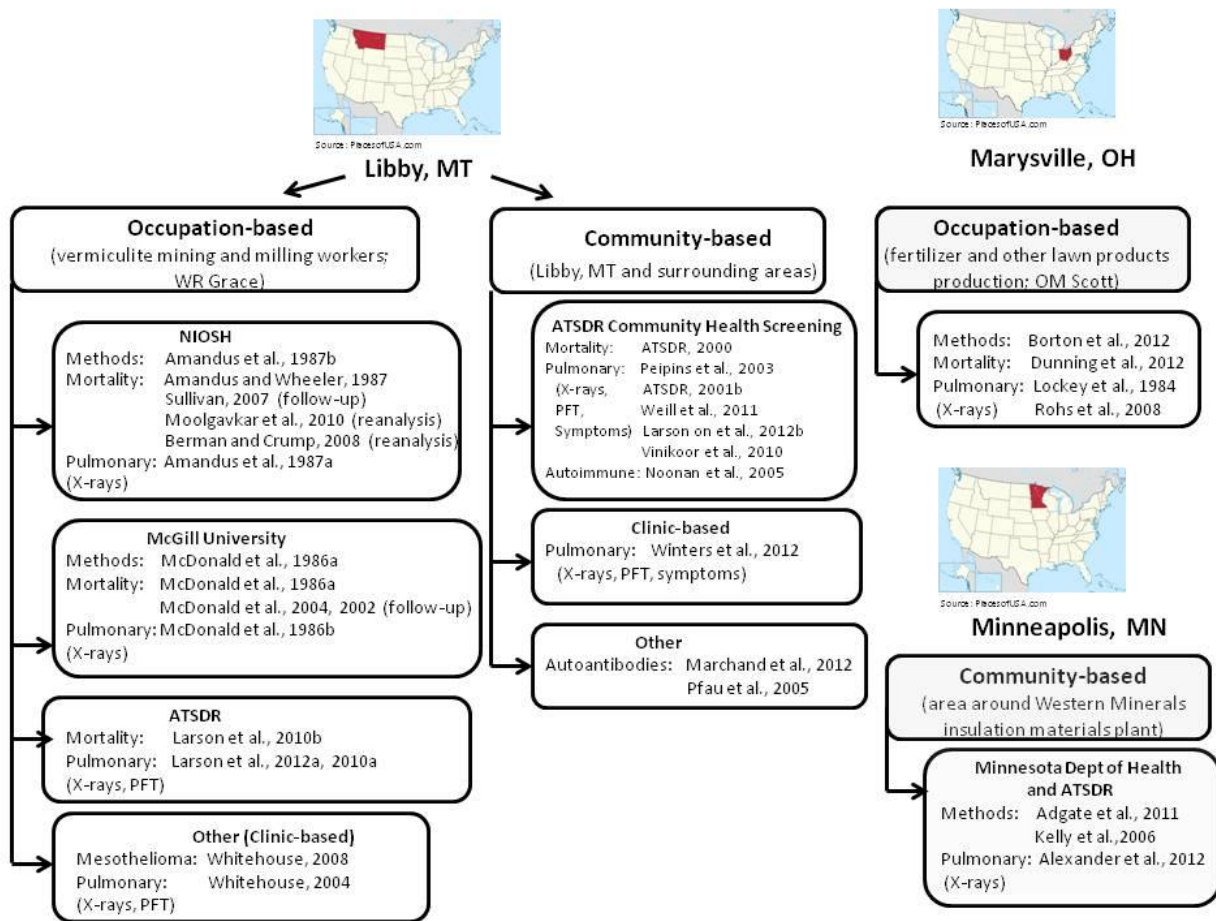
The epidemiology studies of people exposed to LAA were primarily identified through EPA's specific knowledge of the research endeavors that have taken place since recognition in the 1970s of the asbestos contamination from the vermiculite mined around Libby, MT. These studies were conducted by NIOSH, McGill University, University of Cincinnati, and the ATSDR. Analyses by other researchers using the data collected through these studies as well as other studies of people exposed to LAA were also identified through contacts with these research groups and through "forward searching" through Web of Science for references citing the key publications describing the initial studies (i.e., [Peipins et al., 2003](#); [Amandus et al., 1987b](#); [Amandus et al., 1987a](#); [McDonald et al., 1986a](#); [Lockey et al., 1984](#)).

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<sup>8</sup>The term "Libby Amphibole asbestos" is used in this document to identify the mixture of amphibole mineral fibers of varying elemental composition (e.g., winchite, richterite, tremolite, etc.) that have been identified in the Rainy Creek complex near Libby, MT. It is further described in Section 2.2.

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Figure 4-1 depicts the sets of studies conducted by different groups of researchers in Libby, MT and in two areas with plants that used Libby vermiculite in various production processes (fertilizer and other lawn products in Marysville, OH and insulation materials in Minneapolis, MN). These studies have examined cancer and noncancer mortality, pulmonary effects detected through x-ray examinations, pulmonary function tests, or respiratory symptoms, autoimmune diseases, and prevalence of autoantibodies.



**Figure 4-1. Investigations of populations exposed to LAA.** [Moolgavkar et al. \(2010\)](#) and [Berman and Crump \(2008\)](#) used the Libby worker cohort assembled by [Sullivan \(2007\)](#) to estimate cancer potency factors; these analyses are summarized in Section 5.4.5.3.1.

PFT = pulmonary function testing.

The various populations and study designs are summarized in Section 4.1.1, and the results of these studies are presented in subsequent sections: respiratory effects other than cancer (see Section 4.1.2), other noncancer effects (see Section 4.1.3), and cancer (see Section 4.1.4). A brief summary of the epidemiology studies of environmental or residential exposure to tremolite or tremolite-chrysotile mixtures and to crocidolite amphibole is presented in Section 4.1.5.

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#### 4.1.1. Overview of Primary Studies

##### 4.1.1.1. *Studies of Libby, MT Vermiculite Mining and Milling Operations Workers*

4.1.1.1.1. *Description of vermiculite mining and milling operations in Libby, MT.* The vermiculite mining and milling operations have been described in considerable detail ([ATSDR, 2000](#); [Amandus et al., 1987b](#)). Briefly, an open-pit vermiculite mine, located several miles east of Libby, began limited operations in 1923, and production increased rapidly between 1940 and 1950. The mining and milling operations continued until 1990 ([ATSDR, 2008, 2000](#)). Some of the important features of the operations that affected exposure to workers or community members are described below.

The drilling and blasting procedures used in the strip-mining operations generated considerable dust exposures, including silica dust, although the mining operations had lower intensity exposures compared to the milling operations. [Amandus et al. \(1987b\)](#) noted that in 1970, a new drill with a dust-control bagging system aimed at limiting workplace exposure was introduced to the mining operations.

Another aspect of the operations was the loading of ore for railroad shipment. From 1935–1950, railroad box cars were loaded at a station in Libby. In 1950, the loading station was moved to a loading dock on the Kootenai River, 7 miles east of town. Tank cars were used from 1950–1959 and then switched to enclosed hopper cars in 1960.

The milling operations used a screening or sifting procedure to separate vermiculite flakes from other particles and increase the concentration of vermiculite ore from approximately 20% in the bulk ore to 80–95% in the resulting product. A dry mill began operating in 1935, and a wet mill began operating in the 1950s in the same building as the dry mill. One of the primary changes in the conditions in the dry mill was the installation of a ventilation fan in 1964. Exposure to LAA inside the mill was estimated to be 4.6 times higher preceding this installation ([McDonald et al., 1986a](#)). This ventilation fan resulted in higher amphibole fiber exposures in the mill yard until 1968, when the exhaust stack for the fan was moved. Other changes to the milling operations in the 1970s included replacement of hand bagging and sewing with an automatic bagging machine (1972), pressurization of the skipper control room used for transferring the ore concentrate from the mill to a storage site (1972), and construction of a new wet mill (1974). Closing of the old dry and wet mills in 1976 had a substantial impact on exposures at the worksite. In 1974, a new screening plant used to size-sort the ore concentrate was constructed at the loading dock near the river.

Two processing plants operated within the town of Libby ([ATSDR, 2001b](#)). These expansion or exfoliation plants heated the ore concentrate, resulting in additional release of the LAA fibers in the area.

1 **4.1.1.1.2. Descriptions of cohorts of Libby, MT vermiculite mining and milling operations**  
2 **workers.** The two cohort studies conducted in the 1980s (by NIOSH and McGill University)  
3 were similar in terms of populations and study design. For example, both studies included  
4 workers who had worked for at least 1 year. [Amandus and Wheeler \(1987\)](#) included men hired  
5 before 1970 ( $n = 575$ ), with follow-up through December 31, 1981, while [McDonald et al.](#)  
6 [\(1986a\)](#) included men hired before 1963 ( $n = 406$ ) with follow-up through 1983. A subsequent  
7 analysis extended this follow-up through 1999 ([McDonald et al., 2004](#)). Another analysis of the  
8 Libby, MT workers expanded the NIOSH cohort to include all workers, regardless of duration of  
9 employment ([Sullivan, 2007](#)). The total sample ( $n = 1,672$  white men) included 808 workers  
10 who had worked for less than 1 year. These short-term workers had been excluded from the  
11 previous studies. Analyses presented in the report were based on follow-up from 1960–2001.  
12 [Larson et al. \(2010b\)](#) reconstructed a worker cohort based on company records and analyzed  
13 mortality risks through 2006. This study included 1,862 workers (including a small number of  
14 women).

15  
16 **4.1.1.1.3. Fiber exposure estimation in Libby, MT mining and milling operations.** The  
17 exposure assessment procedures used in the NIOSH and McGill University investigations of the  
18 Libby, MT mining and milling operations relied on the same exposure measurements and used  
19 similar assumptions in creating exposure estimates for specific job activities and time periods  
20 (see Table 4-1). In brief, available air sampling data were used to construct a job-exposure  
21 matrix assigning daily exposures (8-hour time-weighted average) for identified job codes based  
22 on sampling data for specific locations and activities. Various job codes and air exposures were  
23 used for different time periods as appropriate to describe plant operations. Individual exposure  
24 metrics (e.g., cumulative exposure [CE]) were calculated using the work history of each  
25 individual in conjunction with the plant job-exposure matrix.

**Table 4-1. Population and exposure assessment methodologies used in studies of Libby, MT vermiculite workers**

<b>Operation, study cohort, reference describing exposure methods</b>	<b>Asbestos fiber quantification</b>	<b>Job-exposure classification</b>	<b>Primary outcomes examined in studies using methodology</b>
Libby, MT mining and milling operations; NIOSH investigation. <i>N</i> = 575 men, hired before 1970, worked at least 1 yr <a href="#">Amandus et al. (1987a)</a>	1962–1967 (and a few earlier yr): midget impinger data ( <i>n</i> samples = 336). 1967–1982: PCM of fibers >5 µm long and aspect ratio >3:1 ( <i>n</i> samples = 4,116).	Samples assigned to 25 “occupation locations” to estimate exposures for specific jobs and time periods 1945–1982. Membrane-filter measurement to impinger conversion ratio: 4.0 fibers/cc per mppcf. Average (arithmetic mean) exposure used for >one sample per location or job task and time period. <sup>a</sup>	Mortality: <a href="#">Amandus and Wheeler (1987)</a> Pulmonary (x-rays): <a href="#">Amandus et al. (1987b)</a> <a href="#">Amandus and Wheeler (1987)</a>
Libby, MT mining and milling operations; NIOSH investigation. <i>N</i> = 1,672 men, no exclusion based on length of employment <a href="#">Sullivan (2007)</a>	Based on <a href="#">Amandus et al. (1987a)</a>	Modification to <a href="#">Amandus et al. (1987a)</a> job classification: laborers and “unknown” jobs assigned weighted-average exposure for all unskilled jobs in work area (if known) during calendar time period, rather than lower mill yard exposure. Weights based on the number of workers assigned to unskilled jobs during same calendar time period.	Mortality: <a href="#">Moolgavkar et al. (2010)</a> <a href="#">Berman and Crump (2008)</a> <a href="#">Sullivan (2007)</a>
Libby, MT mining and milling operations; McGill University investigation. <i>N</i> = 406 men, hired before 1963, worked at least 1 yr <a href="#">McDonald et al. (1986a)</a>	Similar to <a href="#">Amandus et al. (1987a)</a> , but midget impinger data was said to be available through 1969.	Similar to <a href="#">Amandus et al. (1987a)</a> . Samples assigned to 28 “occupation locations”. Conversion ratio = 4.6 used for dry mill pre- and post-1964. Mean of log-normal distributions used for >one sample per location or job task and time period. <sup>a</sup>	Mortality: McDonald et al. (2004, 2002; 1986a) Pulmonary (x-rays): <a href="#">McDonald et al. (1986b)</a>
Libby, MT mining and milling operations; ATSDR investigation. <i>N</i> = 1,862 men and women, no exclusion based on length of employment <a href="#">Larson et al. (2010b)</a>	Based on <a href="#">Amandus et al. (1987a)</a>	Extension of <a href="#">Amandus et al. (1987a)</a> exposure data (without the modification used by <a href="#">Sullivan (2007)</a> , with additional application of exposure estimates to job titles from early 1980s through 1993 (the time of the demolition of the facilities).	Mortality: <a href="#">Larson et al. (2010b)</a> Pulmonary (x-rays, pulmonary function): Larson et al. (2012a; 2010a)

<sup>a</sup>Cumulative exposure reported in units of fiber-yr (equivalent to the unit of fibers/cc-yr EPA is using for all studies).

Mppcf = million particles per cubic foot.



#### 4.1.1.1.3.1. Asbestos fiber quantification and development of job-exposure matrices.

Before 1970, exposure estimates were based on midjet impinger samples taken primarily in the dry mill by state and federal inspectors ( $n$  samples = 336). Total dust samples were measured as million particles per cubic foot (mppcf); these samples included mineral dust and vermiculite particles and scrolls as well as asbestos fibers. Membrane-filter air samples for fibers, taken at various locations within the operations, began in 1967, and data are available from company records as well as State and Federal Agencies (see Table 4-2). Stationary and short-term (i.e., 20-minute to less than 4-hour) measurements were primarily used before 1974. Air samples collected through membrane filters ( $n = 4,116$ ) were analyzed by PCM to visually count fibers greater than 5- $\mu\text{m}$  long and having an aspect ratio  $>3:1$  (Amandus et al., 1987b).<sup>9</sup> PCM methods from the 1960s allowed reliable characterization of fibers with widths greater than approximately 0.4  $\mu\text{m}$  (Amandus et al., 1987b; Rendall and Skikne, 1980). Further standardization of the PCM method and improved quality of microscopes provided better visualization of thinner fibers; a 0.25- $\mu\text{m}$  width was considered the limit of resolution for fiber width in the 1980s (IPCS, 1986), with subsequent improvements in resolution to 0.20  $\mu\text{m}$  in width.

**Table 4-2. Source of primary samples for fiber measurements at the Libby vermiculite mining and milling operations**

Source	Unit of measurement	Yr	Number of samples
State of Montana	mppcf <sup>a</sup>	1956–1969	336
NIOSH	fibers/cc <sup>b</sup>	1967–1968	48
MESA/MSHA <sup>c,d</sup>	fibers/cc	1971–1981	789
Company records	fibers/cc	1970–1982	3,279

<sup>a</sup>Million particles per cubic foot of air, sampled by a midjet impinger apparatus and examined by light microscopy.

<sup>b</sup>Fibers per cc of air drawn through a filter and examined under a phased contrast light microscope. Objects  $>5 \mu\text{m}$  and with an aspect ratio  $>3:1$  were reported as fibers (see Section 2 for details).

<sup>c</sup>MESA: U.S. Mining Enforcement and Safety Administration (former name of MSHA).

<sup>d</sup>MSHA: U.S. Mine Safety and Health Administration.

Source: Amandus et al. (1987a).

<sup>9</sup>Amandus et al. (1987b) indicate (page 12, 4<sup>th</sup> full paragraph) that fibers  $>5\text{-}\mu\text{m}$  long and with an aspect ratio  $>3:1$  were measured. The actual value of the aspect ratio used by Amandus et al. could have been  $\geq 3:1$  because the criterion for the NIOSH recommended exposure limit is based on an aspect ratio of  $\geq 3:1$ , but EPA is reporting here the information that was in the Amandus et al. (1987b) publication.

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1 The samples taken from specific work locations within the operations were used to  
2 estimate exposures in specific jobs and time periods based on industrial hygienist consideration  
3 of temporal changes in facilities, equipment, and job activities. These were defined to categorize  
4 tasks and locations across the mining, milling, and shipping operations to group-like tasks with  
5 respect to exposure potential. Both research groups established similar location operations for  
6 the Libby cohort. Because measures from sample filters were not available before 1968,  
7 different procedures had to be used to estimate exposures at the various locations for this earlier  
8 period. [Amandus et al. \(1987b\)](#) and [McDonald et al. \(1986a\)](#) provide high and low estimates for  
9 several locations to address the uncertainties in assumptions used in these estimates. Both  
10 researchers also applied a conversion factor to estimate asbestos exposures at the dry mill before  
11 1967: the conversion factor was 4.0 in [Amandus et al. \(1987a\)](#) and 4.6 in [McDonald et al.](#)  
12 [\(1986a\)](#).<sup>10</sup>

13 Jobs were mapped to operation/location based on estimated time spent in different job  
14 tasks, thus estimating an 8-hour time-weighted average exposure for each job during several  
15 calendar time periods. Job histories from date of first employment to 1982 were used with the  
16 job-exposure matrix to develop cumulative exposure estimates for each worker.

#### 17 18 *Additional considerations*

19 The resulting exposure estimates presented by both research groups, and the job-exposure  
20 matrices used in calculating cumulative exposure for the cohort, were based on fiber counts by  
21 PCM analysis of air filters. As discussed in Section 2 (see Section 2.4.4), PCM analysis does not  
22 distinguish between fiber mineralogy or morphology, and all fibers >5 µm in length with an  
23 aspect ratio of 3:1 or greater are included. Both researcher groups analyzed fibers available at  
24 the facility in order to identify the mineral fibers in the air samples.

25 TEM<sup>11</sup> analysis of airborne asbestos fibers indicated a range of fiber  
26 morphologies—including long fibers with parallel sides, needlelike fibers, and curved fibers

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<sup>10</sup>The conversion ratio used by [Amandus et al. \(1987b\)](#) was based on a comparison of 336 impinger samples taken in 1965–1969 and 81 filter samples taken in 1967–1971; both sets of air samples were taken in the dry mill. The ratio based on the average fiber counts from air samples in 1967–1971 to the average total dust measurements in 1965–1969 was 4.0 fibers/cc:1.0 mppcf. This ratio was selected because it allowed for the use of the greatest amount of data from overlapping time periods, while controlling for the reduced exposure levels after 1971 where fiber counts based on PCM—but not midget impinger data—were available. The resulting exposure concentrations in the dry mill were estimated as 168 fibers/cc in 1963 and all prior years and 35.9 fibers/cc in 1964–1967. [McDonald et al. \(1986a\)](#) used a different procedure, based on the estimated reduction in dust exposure with the installation of the ventilation system in the dry mill 1964. They observed that total dust levels dropped approximately 4.6-fold after the installation of this equipment. Exposures in the dry mill were thus calculated as 4.6 times the fiber exposures measured by PCM between 1970 and 1974 (22.1 fibers/cc), resulting in estimated dry mill exposures of 101.5 fibers/cc prior to 1965 ([McDonald et al., 1986a](#)).

<sup>11</sup>TEM utilizes a high-energy electron beam to irradiate the sample. This allows visualization of structures much smaller than can be seen under light microscopy. TEM instruments may be fitted with two supplemental instruments that allow for a more complete characterization of structure than is possible under light microscopy: EDS and SAED.



1 ([McDonald et al., 1986a](#)). Of the fibers examined by TEM, >62% were >5 µm in length and a  
2 wide range of dimensional characteristic were noted: length (1–70 µm), width (0.1–2 µm), and  
3 aspect ratios from 3:1–100:1. Energy dispersive spectroscopy used to determine the mineral  
4 analysis indicated that the fibers were in the actinolite-tremolite solid solution series, but sodium  
5 rich ([McDonald et al., 1986a](#)). This analysis is consistent with the current understanding of  
6 amphibole asbestos found in the Libby mine (see Section 2.2.3).

7       At the time of their study, when exposure concentrations were reduced to generally less  
8 than 1 fiber/cc, [Amandus et al. \(1987b\)](#) obtained eight air filters from area air samples collected  
9 in the new wet mill and screening plant (provided by the mining company). These samples were  
10 analyzed by PCM using the appropriate analytical method for the time (NIOSH Physical and  
11 Chemical Analytical Method No. 239). From early method development through current PCM  
12 analytical techniques, the Public Health Service, Occupational Safety and Health Administration  
13 and NIOSH methods have defined a fiber by PCM analysis as having an aspect ratio  $\geq 3:1$   
14 ([NIOSH, 1994](#); [Edwards and Lynch, 1968](#)). [Amandus et al. \(1987b\)](#) reported the dimensional  
15 characteristics of the fibers from these filters including aspect ratio, width, and length (see  
16 Table 4-3). Data for 599 fibers from the eight area air samples collected in the wet mill and  
17 screening plant are provided. These data are limited in one sense by the minimum diameter and  
18 length cutoffs (>4.98-µm long, >0.44-µm wide, aspect ratio >3:1);<sup>12</sup> 16% had an aspect ratio  
19  $\geq 50:1$ . Only 7% of the fibers had a width greater than 0.88 µm, with one fiber reported of the  
20 559 with a width greater than 1.76 µm. It should be noted that these data do not give the full  
21 fiber-size distribution of LAA fibers because NIOSH was examining only PCM visible fibers  
22 (see Section 2.3.1).

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<sup>12</sup>See footnote 9, page 4–6.

**Table 4-3. Dimensional characteristics of fibers from air samples collected in the vermiculite mill and screening plant, Libby, MT<sup>a</sup>**

Fiber length (µm)			Fiber width (µm)			Aspect ratio		
Range	Total counted	Percentage	Range	Total counted	Percentage	Range	Total counted	Percentage
4.98–7.04	54	9	0.44–0.62	406	68	5:1–10:1	24	4
7.04–9.96	109	18	0.62–0.88	151	25	10:1–20:1	176	29
9.96–14.08	107	18	0.88–1.24	27	5	20:1–50:1	305	51
14.08–19.91	111	19	1.24–1.76	14	2	50:1–100:1	84	14
19.91–28.16	90	15	1.76–2.49	0	0	>100:1	10	2
28.16–39.82	65	11	>2.49	1	0			
39.82–66	46	8						
66–88	10	2						
>88	7	1						

<sup>a</sup>Fibers were viewed and counted by phase contrast microscopy.

Source: [Amandus et al. \(1987b\)](#).

#### 1 **4.1.1.2. Studies of O.M. Scott, Marysville, OH Plant Workers**

2 **4.1.1.2.1. Descriptions of cohorts of O.M. Scott, Marysville, OH plant workers.** The first study  
3 of pulmonary effects in the Ohio plant workers was conducted in 1980 and involved 512 workers  
4 ([97% of the 530 workers previously identified with past vermiculite exposure; Lockey et al.,](#)  
5 [1984; see Tables 4-4 and 4-6](#)). The [Rohs et al. \(2008\)](#) study is a follow-up of respiratory effects  
6 in this cohort conducted approximately 25 years later; chest x-rays and interview data were  
7 collected from 280 of the 431 workers known to be alive at this time. [Dunning et al. \(2012\)](#)  
8 examined mortality rates in the cohort, using an updated exposure analysis described by [Borton](#)  
9 [et al. \(2012\)](#). In this analysis, vital status through June 2011 was ascertained.

**Table 4-4. Population and methods used in studies of O.M. Scott, Marysville, OH plant workers**

Reference(s)	Population	Data collection	Outcomes examined
<a href="#">Lockey (1985)</a> <a href="#">Lockey et al. (1984)</a> <sup>a</sup>	1980, $n = 512$ (from 530 identified employees with past vermiculite exposure; nonparticipants included 9 refusals and 9 unavailable due to illness or vacation). Mean age: 37.5 yr Mean duration: 10.2 yr <sup>b</sup> Ever smoked: 64.7%	Exposure estimates based on air samples taken beginning in 1972; methods described below in Section 4.1.1.2.2. Interviews: smoking history, work history at the plant, and other asbestos and fiber mineral work history data.	Respiratory effects, noncancer, based on clinical exam (listening for rales and evaluation of nail clubbing), pulmonary function (spirometry and DLCO) and chest x-rays; two B Readers, 1971 ILO classification guidelines modified with additional grading criteria (e.g., costophrenic angle blunting separated from other pleural lesions)
	Mean cumulative exposure by group (based on jobs and areas):		
	Group      Fibers/cc-yr      ( $n$ )		
	I              0.45 <sup>c</sup> (112)		
	II             1.13                (206)		
	III            6.16                (294)		
<a href="#">Rohs et al. (2008)</a>	2002–2005, $n = 280$ with interviews and readable chest x-rays (from 513 workers in the 1980 study group, 431 were alive in 2004 <sup>d</sup> ; 151 living nonparticipants included 49 refusals, 76 located but did not respond, 8 not located but presumed alive, and 18 missing either x-ray or interview). Mean age: 59.1 yr Ever smoked: 58.6% Mean (range) cumulative exposure: 2.48 (0.01–19.03) fiber/cc-yr	Exposure estimates based on <a href="#">Lockey et al. (1984)</a> . Interviews: pulmonary medical history and job history since 1980 included information on other asbestos exposure.	Respiratory effects, noncancer, based on chest x-rays; 3 B Readers, 2000 ILO classification guidelines
<a href="#">Dunning et al. (2012)</a>	Follow-up of workers identified in <a href="#">Lockey et al. (1984)</a> ; see first row of this table. Limited to $n = 465$ white men. Follow-up through June 2011; 136 deaths Mean duration: 11.0 yr Mean (range) cumulative exposure: 9.0 (<0.01–106.31)	Exposure estimates updated based on <a href="#">Borton et al. (2012)</a> .	Mortality (cancer and noncancer), based on National Death Index

<sup>a</sup>[Lockey et al. \(1984\)](#) is the published paper based on the unpublished thesis ([Lockey, 1985](#)).

<sup>b</sup>Calculated based on stratified data presented in Table 2 of [Lockey et al. \(1984\)](#).

<sup>c</sup>Characterized as similar to background levels in the community, based on an 8-h time-weighted average estimated as 0.049 fiber/cc from a single stationary sample taken outside the main facility.

<sup>d</sup>[Rohs et al. \(2008\)](#) identified one additional eligible worker from the original 512 employees identified in [Lockey et al. \(1984\)](#).

DLCO = single breath carbon monoxide diffusing capacity; ILO = International Labour Organization.

- 1 **4.1.1.2.1.1. *Exposure estimation: O.M. Scott, Marysville, OH plant.*** The plant that processed
- 2 vermiculite ore in Marysville, OH had eight main departments in the processing facility,
- 3 employing approximately 530 workers, with 232 employed in production and packaging of the
- 4 commercial products and 99 in maintenance; other divisions included research, the front office,
- 5 and the polyform plant ([Lockey, 1985](#)). Six departments were located at the main facility

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(trionizing, packaging, warehouse, plant maintenance, central maintenance, and front offices). Research and development and a polyform plant were located separately, approximately one-quarter mile from the main facility. In the trionizing section of the plant, the vermiculite ore was received by rail or truck, unloaded into a hopper, and transported to the expansion furnaces. After expansion, the vermiculite was blended with other materials (e.g., urea, potash, herbicides), packaged, and stored. Changes to the expander type and dust-control measures began in 1967, with substantial improvement in dust control occurring throughout the 1970s.

Industrial hygiene monitoring at the plant began in 1972, with measurements based on fibers  $>5\text{-}\mu\text{m}$  long, diameter  $<3\text{ }\mu\text{m}$ , aspect ratio  $\geq 3:1$ . [Lockey et al. \(1984\)](#) noted that the limited availability of data that would allow for extrapolation of exposures for earlier time periods possibly resulted in the underestimation of exposures before 1974.<sup>13</sup> Breathing-zone samples were used after 1976, with fiber analysis by PCM.

Cumulative fiber exposure indexes, expressed as fibers/cc-yr, were derived for each worker from available industrial hygiene data and individual work histories; three categories of exposure levels were defined ([97% of the 530 workers previously identified with past vermiculite exposure; Lockey et al., 1984; see Table 4-6](#)). Group I was considered to be the nonexposed group and consisted of the chemical processing, research, and front office workers, as well as other workers with an estimated cumulative exposure  $<1$  fiber/cc-yr. Group II was the low-exposure category and included central maintenance, packing, and warehouse workers. The 8-hour time-weighted average fiber exposure in this group was estimated as approximately 0.1–0.4 fiber/cc before 1974 and 0.03–0.13 fiber/cc in and after 1974. Group III was the high-exposure category and included expander, plant maintenance, and pilot plant workers. The 8-hour time-weighted average fiber exposures in this group were approximately 1.2–1.5 fibers/cc before 1974 and 0.2–0.375 fiber/cc in and after 1974. The estimated cumulative exposure for the work force, including Group I workers, ranged from 0.01 to 28.1 fibers/cc-yr using an 8-hour workday and an assumed 365 days of exposure per year.<sup>14</sup> Exposure was assumed to occur from 1957 to 1980 in this study. Exposure outside of work hours was assumed to be zero.

Additional exposure information was identified in 2009, and exposure estimates were updated and refined to reflect information (including fiber measurements) from company reports and other written materials ([Borton et al., 2012](#)). In addition, worker focus groups provided insight into plant processes—including industrial hygiene measures—and work patterns and organizations. Further details on the updated exposure assessment are included in Appendix F.

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<sup>13</sup>Subsequent exposure assessment efforts by this team of investigators are described in [Borton et al. \(2012\)](#) and in Appendix F.

<sup>14</sup>[Lockey et al. \(1984\)](#) reported the maximum value for this group as 39.9 fibers/cc-yr, but this estimate was later corrected to exclude work from 1947 to 1956, before the use of vermiculite at the plant. Information provided in [Benson \(2014\)](#).

#### 4.1.1.3. *Community-Based Studies Around Libby, MT Conducted by Agency for Toxic Substances and Disease Registry (ATSDR)*

Analyses using data from community-based studies in Libby, MT conducted by ATSDR are summarized in Table 4-5. [ATSDR \(2000\)](#) includes a mortality analysis based on death certificate data from 1979–1998, with residence at time of death geocoded to areas corresponding to Libby, MT and its surroundings. The estimated population size in 1991 for the areas used in the standardized mortality ratio (SMR) calculations ranged from 2,531 in the Libby city limits to 9,512 for the central Lincoln County area (based on a 10-mile radius around downtown Libby). Cause-specific standardized mortality ratios were computed based on Montana and United States comparison rates; asbestosis SMRs were somewhat higher using the U.S. referent group, but the choice of referent group had little difference on SMRs for most diseases.

ATSDR also conducted a community health screening from July–November 2000 and July–September 2001, with 7,307 total participants ([ATSDR, 2001b](#)). Eligibility was based on residence, work, or other presence in Libby for at least 6 months before 1991. The total population eligible for screening is not known, although the population of Libby, MT in 2000 was approximately 10,000. Other studies ([Larson et al., 2012b](#); [Weill et al., 2011](#); [Vinikoor et al., 2010](#); [Noonan et al., 2006](#)) used data collected during this community health screening.

Two additional community-based studies, using data sources other than the ATSDR community health screening ([Marchand et al., 2012](#); [Pfau et al., 2005](#)) are discussed in Section 4.1.3.2. (Autoimmune disease and autoantibodies), and two clinic-based studies from Libby, MT ([Winters et al., 2012](#)) are discussed in Section 4.1.2.2.4 (Clinic-based reports and case reports of respiratory disease [noncancer]).

**Table 4-5. Summary of methods used in community-based studies of Libby, MT residents conducted by Agency for Toxic Substances and Disease Registry (ATSDR)**

Reference(s)	Methods	Data collection	Outcomes examined
<a href="#">ATSDR (2000)</a>	1979–1998 mortality analysis, underlying cause of death from death certificates: 419 decedents identified, 418 death certificates obtained, 413 geocoded; 16 of 91 residents of elderly care facilities reclassified to nonresident.	Geocoding of street locations (residence at time of death) within six geographic boundaries (ranging from 2,532 residents in Libby city limits to 9,521 in central Lincoln County in 1990)	Mortality (cancer and noncancer); underlying cause of death
<a href="#">Vinikoor et al. (2010)</a> <a href="#">Peipins et al. (2003)</a> <a href="#">ATSDR (2001b)</a>	ATSDR community health screening, July–November 2000 ( <a href="#">Peipins et al. (2003)</a> ; <a href="#">ATSDR, 2001b</a> ) also included July–September 2001 participants) Eligibility: resided, worked, attended school, or participated in other activities in Libby for at least 6 mo before 1991 (including vermiculite mine and mill workers). <i>N</i> = 7,307 interviews and <i>n</i> = 6,668 chest x-rays.	Standardized interview: medical history, symptoms, work history, and other potential exposures. Exposure based on information on “exposure pathways” (e.g., worked at vermiculite mining or milling operations, other asbestos-related work history, lived with worker at the vermiculite mining or milling operations, use of vermiculite products, played in vermiculite piles)	Chest x-rays (posterior-anterior, oblique), 1980 ILO classification guidelines; pulmonary function, respiratory symptoms
<a href="#">Weill et al. (2011)</a>	ATSDR community health screening ( <a href="#">see ATSDR, 2001b</a> ). <i>n</i> = 4,397, ages 25 to 90 yr, excluding individuals with history of other asbestos-related work exposures.	Analysis based on five exposure categories: <ul style="list-style-type: none"> <li>- Worked at vermiculite mining or milling operations</li> <li>- Other vermiculite occupation</li> <li>- Other dusty (asbestos-related) occupations</li> <li>- Lived with vermiculite/dusty/asbestos worker</li> <li>- Environmental (did not work or live with dusty/asbestos worker)</li> </ul>	Chest x-rays (posterior-anterior), 1980 ILO classification guidelines; pulmonary function in relation to chest x-ray findings
<a href="#">Larson et al. (2012b)</a>	ATSDR community health screening ( <a href="#">see ATSDR, 2001b</a> ). <i>n</i> = 6,476, ages ≥18 yr, excluding individuals without interpretable spirometry and chest x-ray data.	Exposure pathways as described in <a href="#">Peipins et al. (2003)</a>	Chest x-rays (posterior–anterior), 1980 ILO classification guidelines modified such that plaques definition was equivalent to ILO 2000 LPT guidelines; pulmonary function in relation to chest x-ray findings
<a href="#">Noonan et al. (2006)</a>	ATSDR community health screening ( <a href="#">see ATSDR, 2001b</a> ). Nested case-control study of rheumatoid arthritis, scleroderma, and systemic lupus erythematosus cases ( <i>n</i> = 161 cases, 1,482 controls); initial self-report confirmed in second interview.	Exposure pathways as described in <a href="#">Peipins et al. (2003)</a>	Systemic autoimmune diseases

LPT = localized pleural thickening.

## 4.1.2. Respiratory Effects, Noncancer

### 4.1.2.1. Asbestosis and Other Nonmalignant Respiratory Disease Mortality

Several studies described previously reported noncancer respiratory disease mortality data. Nonmalignant respiratory disease is a broad category (International Classification of Diseases [ICD]-9 codes 460–519) that includes chronic obstructive pulmonary disease, asthma, pneumonia and respiratory infections, asbestosis (ICD-9 code 501), and various forms of pneumoconiosis. A greater specificity of effects due to asbestos would be expected using the narrower category of asbestosis compared with nonmalignant respiratory disease.

The initial studies of the Libby, MT vermiculite mining and milling worker cohorts were based on a relatively small number of nonmalignant respiratory-related deaths (<25); more than 50 deaths in this category were seen in later studies (see Table 4-6). The analytic strategy (e.g., use of a latency period to exclude cases that occurred before the effect of exposure would be expected to be manifested, or use of a lag period to exclude exposures that occurred after the onset of disease) and the cutpoints for exposure categories varied among the studies, but a pattern of increasing risk with increasing cumulative exposure is seen, with more than a 10-fold increased risk of death due to asbestosis and a 1.5- to 3-fold increased risk of nonmalignant respiratory disease in the analyses using an internal referent group ([Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#)). [Larson et al. \(2010b\)](#) used a Monte Carlo simulation to estimate the potential bias in nonmalignant respiratory disease risk that could have been introduced by differences in smoking patterns between exposed and unexposed workers in the cohort. The bias-adjustment factor (relative risk  $[RR]_{unadjusted}/RR_{adjusted} = 1.2$ ) reduced the overall RR estimate for nonmalignant respiratory mortality from 2.1 to 1.8. Asbestosis risk was also increased in the ATSDR geographic-based analysis, with SMRs of approximately 40 based on Montana rates and 65 based on U.S. comparison rates ([ATSDR, 2000](#)). Only one asbestosis-related death was observed in the Marysville, OH worker cohort, resulting in a very imprecise risk estimate ([Dunning et al., 2012](#)).



**Table 4-6. Nonmalignant respiratory mortality studies of populations exposed to Libby Amphibole asbestos<sup>a</sup>**

Reference(s)	Respiratory disease (SMR, 95% CI)	Dose-response analyses: nonmalignant respiratory diseases and asbestosis			
Occupational studies of Libby, MT mining and milling operations workers					
<a href="#">Amandus and Wheeler (1987)</a> (NIOSH)	No exclusions: nonmalignant respiratory diseases ( <i>n</i> = 20) SMR: 2.4 (1.5, 3.8)  20-yr latency: nonmalignant respiratory diseases ( <i>n</i> = 12) SMR: 2.5 ( <i>p</i> < 0.05)	No exclusions: nonmalignant respiratory diseases			
		Cumulative exposure	<i>n</i>	SMR (95% CI) <sup>b</sup>	
		0.0–49 fibers/cc-yr	8	2.2 (not reported)	
		50–99 fibers/cc-yr	2	1.7 (not reported)	
		100–399 fibers/cc-yr	3	1.8 (not reported)	
		≥400 fibers/cc-yr	10	4.0 (not reported, but <i>p</i> < 0.01)	
		20 or more yr since first hire (latency): nonmalignant respiratory diseases			
		Cumulative exposure	<i>n</i>	SMR (95% CI) <sup>b</sup>	
		0.0–49 fibers/cc-yr	7	3.3 (not reported, but <i>p</i> < 0.05)	
		50–99 fibers/cc-yr	2	2.8 (not reported)	
		100–399 fibers/cc-yr	0	0 (not reported)	
		≥400 fibers/cc-yr	3	2.8 (not reported)	
McDonald et al. (2004; 1986a) (McGill)	Nonmalignant respiratory diseases ( <i>n</i> = 51) SMR: 3.1 (2.3, 4.1)	Excluding first 10 yr of follow-up: nonmalignant respiratory diseases			
		Cumulative exposure	<i>n</i>	RR (95% CI) <sup>d</sup>	
		0.0–11.6 fibers/cc-yr	5	1.0 (referent)	
		11.7–25.1 fibers/cc-yr	13	2.5 (0.88, 7.2)	
		25.2–113.7 fibers/cc-yr	14	2.6 (0.93, 7.3)	
		≥113.8 fibers/cc-yr	19	3.1 (1.2, 8.4)	
		per 100 fibers/cc-yr	-	0.38 (0.12, 0.96) ( <i>p</i> = 0.0001)	
<a href="#">Sullivan (2007)</a> (NIOSH)	15-yr exposure lag: Asbestosis ( <i>n</i> = 22) SMR: 166 (104, 251) Nonmalignant respiratory diseases ( <i>n</i> = 111) SMR: 2.4 (2.0, 2.9) Chronic obstructive pulmonary disease ( <i>n</i> = 53) SMR: 2.2 (1.7, 2.9) Other nonmalignant respiratory diseases ( <i>n</i> = 19) SMR: 2.7 (1.6, 4.2)	15-yr exposure lag: asbestosis			
		Cumulative exposure	<i>n</i>	SMR (95% CI) <sup>b</sup>	SRR (95% CI) <sup>c</sup>
		0.0–49.9 fibers/cc-yr	3	37 (7.5, 122)	1.0 (referent)
		50.0–249.9 fibers/cc-yr	8	213 (91.6, 433)	7.3 (1.9, 28.5)
		≥250 fibers/cc-yr	11	749 (373, 1,368)	25.3 (6.6, 96.3)
		linear trend test			( <i>p</i> < 0.01)
		15-yr exposure lag: nonmalignant respiratory diseases			
		Cumulative Exposure	<i>n</i>	SMR (95% CI) <sup>b</sup>	SRR (95% CI) <sup>c</sup>
		0.0–4.49 fibers/cc-yr	18	1.8 (1.1, 2.8)	1.0 (referent)
		4.5–19.9 fibers/cc-yr	24	2.0 (1.3, 3.0)	1.2 (0.6, 2.3)
		20.0–84.9 fibers/cc-yr	26	2.2 (1.5, 3.3)	1.5 (0.8, 2.9)
		85.0–299.9 fibers/cc-yr	20	2.6 (1.6, 4.0)	1.4 (0.7, 2.7)
		≥300 fibers/cc-yr	23	4.8 (3.1, 7.3)	2.8 (1.3, 5.7)

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**Table 4-6. Nonmalignant respiratory mortality studies of populations exposed to Libby Amphibole asbestos<sup>a</sup> (continued)**

Reference(s)	Respiratory disease (SMR, 95% CI)	Dose-response analyses: nonmalignant respiratory diseases and asbestosis				
<a href="#">Larson et al. (2010b)</a>	Asbestosis ( <i>n</i> = 69) SMR: 143 (111, 181) Nonmalignant respiratory diseases ( <i>n</i> = 425) SMR: 2.4 (2.2, 2.6) Chronic obstructive pulmonary disease ( <i>n</i> = 152) SMR: 2.2 (1.9, 2.6) Other nonmalignant respiratory ( <i>n</i> = 120) SMR: 2.8 (2.3 3.4)	20-yr exposure lag: asbestosis				
		Cumulative exposure	<i>n</i>	SMR (95% CI) <sup>b</sup>	RR (95% CI) <sup>e</sup>	
		<1.4 fibers/cc-yr	4	(not reported)	1.0 (referent)	
		1.4–<8.6 fibers/cc-yr	8	(not reported)	2.8 (1.0, 7.6)	
		86–<44.0 fibers/cc-yr	25	(not reported)	8.0 (3.2, 19.5)	
		≥44.0 fibers/cc-yr	32	(not reported)	11.8 (4.9, 28.7)	
		Per 100 fibers/cc-yr increase			1.18 (1.12, 1.23)	
					( <i>p</i> < 0.001)	
		20-yr exposure lag: nonmalignant respiratory diseases				
		Cumulative exposure	<i>n</i>	SMR (95% CI) <sup>b</sup>	RR (95% CI) <sup>e</sup>	
		<1.4 fibers/cc-yr	43	(not reported)	1.0 (referent)	
		1.4–<8.6 fibers/cc-yr	46	(not reported)	1.4 (0.9, 2.1)	
		86–<44.0 fibers/cc-yr	56	(not reported)	1.8 (1.3, 2.7)	
		≥44.0 fibers/cc-yr	58	(not reported)	2.5 (1.7, 3.6)	
		Per 100 fibers/cc-yr increase			1.08 (1.03, 1.13)	
					( <i>p</i> = 0.0028)	
<i>Community-based studies in Libby, MT</i>						
<a href="#">ATSDR (2000)</a>	Asbestosis ( <i>n</i> = 11)	SMR	(95% CI)	SMR		(95% CI)
	Comparison area (Montana reference rates):			Comparison area (U.S. reference rates):		
	Libby city limits	40.8	(13.2, 95.3)	Libby city limits	63.5	(20.5, 148)
	Extended Libby boundary	47.3	(18.9, 97.5)	Extended Libby boundary	74.9	(30.0, 154)
	Air modeling	44.3	(19.1, 87.2)	Air modeling	71.0	(30.6, 140)
	Medical screening	40.6	(18.5, 77.1)	Medical screening	66.1	(30.2, 125)
	Libby valley	38.7	(19.3, 69.2)	Libby valley	63.7	(31.7, 114)
	Central Lincoln County	36.3	(18.1, 64.9)	Central Lincoln County	59.8	(29.8, 107)
<i>Occupational studies of O.M. Scott, Marysville, OH plant workers</i>						
<a href="#">Dunning et al. (2012)</a>	Asbestosis ( <i>n</i> = 1) SMR: 15.4 (0.4, 86)					

**Table 4-6. Nonmalignant respiratory mortality studies of populations exposed to Libby Amphibole asbestos<sup>a</sup> (continued)**

CI = confidence interval, SRR = standardized rate ratio.

<sup>a</sup>Libby, MT mining and milling operations includes miners, millers, and processors; workers in the screening plant, loading docks, and expansion plants; and office workers.

<sup>b</sup>SMR based on external referent group.

<sup>c</sup>In [Sullivan \(2007\)](#), the SRR is a ratio of sums of weighted rates in which the weight for each stratum-specific rate is the combined person-yr for the observed cohort across all duration (or cumulative level of exposure) categories. The Life Table Analysis System provides the SRR for each duration (or cumulative level of exposure) group compared to the referent group. The cutoff points for the categories are specified by the user. Taylor-series-based confidence intervals ([Rothman, 1986](#)) are given for each specific SRR.

<sup>d</sup>In [McDonald et al. \(2004\)](#), the RR is based on Poisson analysis using an internal referent group.

<sup>e</sup>In [Larson et al. \(2010b\)](#), the RR is based on Cox proportional hazards modeling using an internal referent group.

- 1 **4.1.2.2. *Pathological Alterations of the Parenchyma and Pleura, Pulmonary Function, and***
- 2 ***Respiratory Symptoms***
- 3 **4.1.2.2.1. *Definition of outcomes***
- 4 **4.1.2.2.1.1. *Pathological alterations of the parenchyma and pleura***

#### **Text Box 4-1. Pathological Alterations of the Parenchyma and Pleura**

**Parenchymal changes in the lung (small opacities):** The small opacities viewed within the lung (interstitial changes) are indicative of pneumoconiosis and are associated with exposure to not only mineral fibers, but also mineral dust and silica. The radiographic signs of pneumoconiosis begin as small localized areas of scarring in the lung tissue and can progress to significant scarring and lung function deficits. The ILO classification guidelines provide a scheme for grading the severity of the small opacities; the size, shape, and profusion of the small opacities are recorded, as well as the affected zone of the lung.

**Obliteration of the costophrenic angle:** The costophrenic angle is measured as the angle between the ribcage and the diaphragm on a posterior anterior-viewed radiograph (the costophrenic recess). When blunting or obliteration is noted on a radiograph, it is recorded as present or absent. Obliteration of the costophrenic angle may occur in the absence of other radiographic signs.

**Pleural thickening:** The pleural lining around the lungs (visceral pleura) and along the chest wall and diaphragm (parietal pleura) may thicken due to fibrosis and collagen deposits. Pleural thickening (all sites) is reported as either LPT or diffuse pleural thickening (DPT). DPT of the chest wall may be reported as in-profile or face-on, and is recorded on the lateral chest wall “only in the presence of and in continuity with, an obliterated costophrenic angle.” Localized pleural thickening may also be viewed in-profile or face-on and is generally a pleural plaque (parietal). Calcification is noted where present.

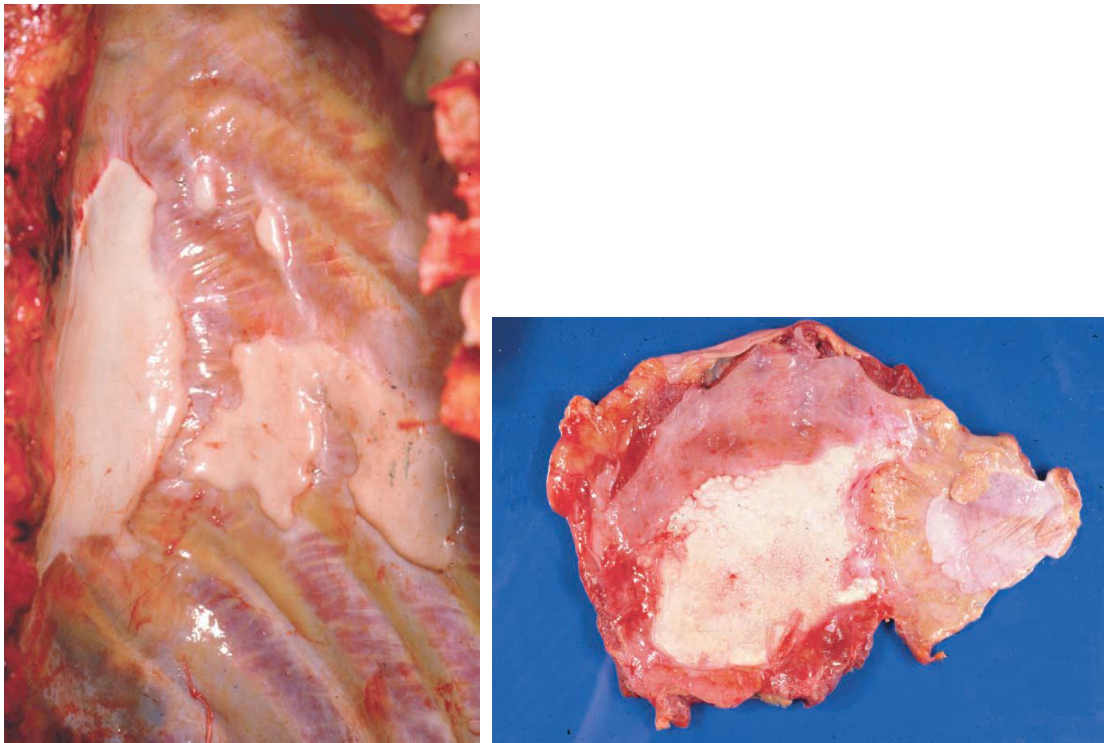
Source: [ILO \(2002\)](#) (describing the ILO 2000 revision).

1 Respiratory disease risk can be evidenced by pleural and parenchymal abnormalities  
2 (pathological, structural alterations) detected through radiographic or other types of imaging (see  
3 Text Box 4-1). These types of effects are usually classified using criteria developed by the  
4 International Labour Organization (ILO) of the United Nations to standardize descriptions of  
5 effects and improve inter-rater agreement and accuracy for reading chest radiographs in  
6 pneumoconiosis. The guidelines were initially developed in 1950 with several subsequent  
7 revisions. A key component of the guidelines is the use of a set of standard films illustrating  
8 different types of findings; these films are used by “B Readers” as a reference for comparison to  
9 films collected in a research or clinical setting. The B Reader program was initiated in 1974 to  
10 reduce variability in readings; B Readers are physicians who pass an examination, recertifying  
11 every 4 years, in the adherence to detailed criteria when reading radiographs in individuals with  
12 pneumoconiosis. The criteria provide for the exclusion of anomalies potentially due to  
13 nonasbestos-related causes (e.g., trauma, tuberculosis).

14 Parenchymal (the inner structure of the lungs) abnormalities include opacities; these  
15 abnormalities are defined as small ( $\leq 10$  mm diameter) or large ( $> 10$  mm diameter). Small  
16 opacities are assigned a score based on the concentration of opacities in a given area (profusion),  
17 zone(s) of the lung(s) affected, shape, and size. Small opacity profusion is graded on a 4-point  
18 scale (0 = absence of opacity, 3 = highest level of opacity). Two ratings are given (e.g., 0/1, or  
19 2/2), with the second number indicating a grade that was seriously considered as an alternative to  
20 the first grade. Large opacities are scored based on dimensions within the lung zone(s) they  
21 occupy. The scarring of the parenchymal tissue of the lung contributes to measurable  
22 decrements in pulmonary function, including obstructive pulmonary deficits from narrowing  
23 airways, restrictive pulmonary deficits from the decreased elasticity of the lung, and decrements  
24 in gas exchange ([ATS, 2004](#)).

25 According to the 2000 ILO guidelines ([ILO, 2002](#)), pleural abnormalities are classified as  
26 (a) localized pleural thickening (LPT) or (b) diffuse pleural thickening (DPT), defined as pleural  
27 thickening that is present “only in the presence of and in continuity with an obliterated  
28 costophrenic angle (CPA).” Previous ILO guidelines ([ILO, 1980, 1971](#)) defined DPT without  
29 the requirement for CPA obliteration; thus, under the 2000 ILO guidelines, the LPT category  
30 includes what was previously defined as DPT without an obliterated costophrenic angle. The  
31 2000 ILO category of LPT also includes what previous ILO guidelines defined as “plaques.”  
32 Different researchers implementing the earlier ILO guidelines variously used terms such as  
33 “discrete” or “circumscribed” or “pleural” to describe these plaques. Both LPT and DPT are  
34 scored based on their location, extent, and whether calcification is seen. LPT is a change in  
35 tissue structure, and is not known to be an adaptive response to toxicity generally or to asbestos  
36 specifically. Examples of pleural plaques (a subset of LPT) as visualized on autopsy are shown  
37 in Figures 4-2A and 4-2B from Official Journal of the [ATS \(2004\)](#). Additional discussion of the  
38 adversity of LPT is included in Section 5.2.2.3 (*Selection of Critical Effect*) and Appendix I.

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**Figure 4-2. A (left). Gross appearance at autopsy of asbestos-associated pleural plaques overlying the lateral thoracic wall. ([ATS, 2004, Figure 12](#))**

**Figure 4-2B (right). Gross appearance of large asbestos-related pleural plaque over the dome of the diaphragm. ([ATS, 2004, Figure 13](#))**

Source: [ATS \(2004\)](#). Reprinted with permission of the American Thoracic Society. Copyright © 2014 American Thoracic Society.

1        The latency period for the initial detection of pleural or parenchymal abnormalities varies  
 2 by type of lesion. [Larson et al. \(2010a\)](#) examined x-rays of 84 workers from the Libby, MT  
 3 mining and milling operations for whom pleural and/or parenchymal abnormalities were seen  
 4 and who had one or more previous x-rays covering a span of at least 4 years available for  
 5 comparison. Circumscribed pleural plaques was seen in 83 of these 84 workers at a median  
 6 latency of 8.6 years. Any pleural calcification was seen in 37 workers, with a median latency of  
 7 17.5 years, and DPT was seen in 12 workers (median latency: 27.0 years). The latency period  
 8 for small opacities indicating parenchymal changes (e.g., asbestosis) increased with increasing  
 9 profusion categories, from a median of 18.9 years for  $\geq 1/0$ , 33.3 years for progression to  $\geq 2/1$ ,  
 10 and 36.9 years for progression to  $\geq 3/2$ .

## *Pulmonary function*

Pulmonary function, commonly measured by spirometry, is used as an indicator of respiratory health and lung disease. Spirometric measurements involve assessment of lung volume and of air flow ([Pellegrino et al., 2005](#)). Forced vital capacity (FVC) is a measure of the maximum amount of air that can be exhaled. Forced expiratory volume (FEV) is the maximum amount of air exhaled in a given time period; for example, FEV<sub>1</sub> refers to the amount of air exhaled in the first second of the test procedure. Standardization of test procedures is very important in these tests, and multiple measurements ( $\geq 3$ ) are typically needed. Values are compared to “reference values” based on age, gender, height (and sometimes race).

Combinations of various functional measurements may be indicative of specific types of abnormalities affecting lung function. For example, restrictive lung function (or restrictive ventilatory defect) refers to reduced lung volume. Both FEV<sub>1</sub> and FVC would be reduced, but the reduction in FVC would be greater than that for FEV<sub>1</sub> (e.g., FEV<sub>1</sub>/FVC ratio  $>0.8$ ). Restrictive lung function can result from inflammation or scarring of the parenchyma, interstitial lung disease, fibrosis, or other conditions that restrict the ability of the lungs to expand. Obstructive lung function (or obstructive ventilatory defect) refers to reduced airflow, and is characterized by inflammation or obstruction of the airways. It is indicated by a reduction in FEV<sub>1</sub> without an accompanying change in FVC (e.g., the ratio of FEV<sub>1</sub>/FVC  $<0.7$ , or FEV<sub>1</sub>/FVC  $<5^{\text{th}}$  percentile). Both restrictive and obstructive conditions can result in dyspnea (shortness of breath), cough, and chest pain.

### **4.1.2.2.2. Results: *pathological alterations of parenchyma and pleura: occupational studies*** *Libby, MT vermiculite mine and mill workers*

Studies examining pleural and parenchymal abnormalities in the Libby, MT worker cohorts are shown in Table 4-7. In the [McDonald et al. \(1986b\)](#) and [Amandus et al. \(1987a\)](#) studies, x-ray films for each worker, which NIOSH obtained from the Libby hospital that performed the screening, were independently read by three qualified readers using the 1980 ILO classification system. For the analysis, the classification indicating pleural abnormalities by at least two of the three readers was used to determine the presence of pleural abnormalities, while the median reading was used to determine the profusion category of small opacities. Although both research groups used the ILO 1980 guidelines, [McDonald et al. \(1986b\)](#) reported pleural thickening on the chest wall (both pleural plaques and diffuse) but not in other sites. [Amandus et al. \(1987a\)](#) examined “any unilateral or bilateral pleural change”, which included “...pleural plaque, diffuse pleural thickening of the chest wall, diaphragm or other site, but excluded costophrenic angle obliteration.” This classification would be equivalent to the LPT classification used in the revised ILO guidelines ([ILO, 2002](#)); however, the results reported in the paper are for thickening on the chest wall only (rather than including other sites) are not equivalent to the 2000 ILO LPT classification.

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**Table 4-7. Chest radiographic studies of the Libby, MT vermiculite mine workers**

Reference(s)	Inclusion criteria and design details	Results				
<a href="#">McDonald et al. (1986b)</a>	Men employed July 1, 1983 ( <i>n</i> = 164). Former employees living within 200 miles; hired before 1963 ( <i>n</i> = 80), worked at least 1 yr (80 participants from 110 eligible). Comparison group—men without known occupational dust exposure ( <i>n</i> = 47); x-rays taken for other reasons (mostly employment related) at same place during study period.	Prevalence (%)	Current workers	Former workers	Comparison group	
		Pleural thickening	15.9	52.5	8.5	
		Small opacities (≥1/0)	9.1	37.5	2.1	
		Both abnormalities increased with age, and with increasing cumulative exposure in age-adjusted and stratified (>60 yr old) analyses.				
<a href="#">Amandus et al. (1987a)</a>	Men, employed 1975–1982 for ≥5 yr ( <i>n</i> = 191); 184 with previous chest x-rays; 121 with smoking questionnaires. Annual radiographs since 1964; most recent radiograph evaluated. Duration: mean 14 yr Cumulative exposure: mean 123 (all workers), 119 (with radiographs) fiber-yr	Pleural thickening of the chest wall observed in 13%.				
		Small opacities (≥1/0) observed in 10%.				
		Beta ( <i>p</i> -value), cumulative exposure in relation to:				
		Small opacities		0.0026	( <i>p</i> < 0.05)	
		Any pleural change		0.0008	( <i>p</i> ≥ 0.05)	
		Pleural calcification		−0.0010	( <i>p</i> ≥ 0.05)	
		Pleural change on wall		0.0008	( <i>p</i> ≥ 0.05)	
		Effect of age was significant in all models, controlling for exposure.				
<a href="#">Larson et al. (2012a)</a>	<i>N</i> = 336 participants in community screening (see Table 4-5 for more details) who reported working at facility, confirmed by company records. Mean age 55.6 yr, 93.6% male. Duration: median 1.5 yr Cumulative exposure: median 3.6 fibers/cc-yr Restrictive spirometry defined as FVC < lower limit of normal and FEV <sub>1</sub> /FVC > lower limit of normal.				Association with cumulative exposure (cc/f-yr) <sup>a</sup>	
			<i>n</i>	(%)	Starting at	Statistically significant at
		DPT, CAO	18	(5)	5	>200
		Profusion ≥1/0	18	(5)	1	108
		Localized pleural thickening	117	(35)	0.5	1.0
		Restrictive spirometry	45	(16)	26	166
		<sup>a</sup> Logistic regression with continuous cumulative exposure; restricted cubic spline functions used to assess shape of exposure-response. “Starting at” refers to the cumulative exposure level reflecting the beginning of the increasing risk pattern; “statistically significant at” refers to the cumulative exposure level at which the relative risk estimate was statistically significant				

1 [Amandus et al. \(1987a\)](#) reported pleural thickening of the chest wall in 13% and small  
2 opacities ( $\geq 1/0$ ) in 9.1% of current employees. Similar data were reported by [McDonald et al.](#)  
3 [\(1986b\)](#), with 15.9 and 10% with pleural thickening of the chest wall and small opacities,  
4 respectively. In both studies, prevalence of these abnormalities increased with increasing  
5 cumulative exposure. [McDonald et al. \(1986b\)](#) also included 80 former employees in their  
6 study. The prevalence of pleural thickening of the chest wall (52.5%) and small opacities  
7 (37.5%) was higher in former employees compared with current workers. These groups differed  
8 by age, with only one of the 80 former workers being less than 40 years of age while 80 of  
9 164 current workers were under 40 years of age. Both overall and within age categories,  
10 however, the prevalence is higher among former employees, and this is attributed to higher  
11 cumulative exposure in this group.

12 Both [Amandus et al. \(1987a\)](#) and [McDonald et al. \(1986b\)](#) provided categorical  
13 exposure-response data as well as logistic models for various endpoints (e.g., small opacities,  
14 pleural calcification, pleural thickening of the chest wall, and “any pleural change”). In  
15 [McDonald et al. \(1986b\)](#), exposure and age were both predictive of pleural thickening along the  
16 chest wall; the regression coefficient for cumulative exposure (fibers/cc-yr) was 0.0024 per unit  
17 increase in cumulative exposure for the log odds of the presence of pleural thickening, adjusting  
18 for age and smoking. Cumulative exposure, age, and smoking status were all predictive of small  
19 opacities; the parameter for cumulative exposure had a regression coefficient of 0.0035 per unit  
20 increase in cumulative exposure. In contrast, although the categorical analysis reported by  
21 [Amandus et al. \(1987a\)](#) indicated a positive exposure response relationship for both “any pleural  
22 change” and pleural thickening along the chest wall, cumulative exposure was not a significant  
23 predictor in regression analysis adjusting for age (regardless of smoking status). The lack of  
24 statistical significance in these models may reflect a nonlinearity resulting from the lower  
25 prevalence in exposure Category 2 compared to exposure Category 1. The estimated relationship  
26 between exposure and prevalence of small opacities in [Amandus et al. \(1987a\)](#) was similar to  
27 that reported by [McDonald et al. \(1986b\)](#).

28 [Larson et al. \(2012a\)](#) used data collected as part of the community screening program  
29 conducted in 2001 ([ATSDR, 2001b; see Section 4.1.1.3](#)) to examine the pleural and pulmonary  
30 outcomes based on chest radiographs, spirometry results, and self-reported symptoms in relation  
31 to cumulative exposure among 336 workers. Diffuse pleural thickening (in the presence of  
32 costophrenic angle obliteration) and parenchymal lesions (profusion  $\geq 1/0$ ) were each detected in  
33 5% of the workers. Risk increased monotonically with increasing cumulative exposure for each  
34 of these outcomes; however, the slope was shallower for diffuse pleural thickening and was not  
35 statistically significant. Localized pleural thickening (only) was found in 35% of the workers  
36 with an elevated risk associated with cumulative exposures as low as 1 fiber/cc-yr. For a  
37 diagnosis of restrictive spirometry (prevalence = 16%), risk began to increase at 26 fibers/cc-yr  
38 and reached statistical significance at 166 fibers/cc-yr. Chronic bronchitis defined as coughing

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up phlegm “for at least 3 months of the year for the past 2 years” was reported in 8% of the workers, and a statistically significant increased risk was calculated at 24 fibers/cc-yr.

#### *O.M. Scott, Marysville, OH plant workers*

The first study of the O.M. Scott, Marysville, OH plant workers was conducted by [Lockey et al. \(1984\)](#); see [Table 4-12](#). Physical examination (for detection of pulmonary rales and nail clubbing), pulmonary function (spirometry and DLCO), and chest x-rays were performed, and information pertaining to smoking history, work history at the plant, and other relevant work exposures was collected using a trained interviewer. Approximately 44% of the 512 workers in the study were current smokers, 20% former smokers, and 35% lifetime nonsmokers, but smoking history (i.e., smoking status, pack-years) did not differ by exposure group. An increased risk of costophrenic angle blunting ( $n = 11$ ), other pleural and parenchymal abnormalities ( $n = 11$ ), or any of these outcomes ( $n = 22$ ) was observed in relation to exposure assessed by job title and area (see description of exposure groups in Section 4.1.1.2.2) and categorized into groups based on the cumulative fiber estimates. The prevalence of any radiographic change was 2.8% in Group I, 3.9% in Group II, and 5.8% in Group III. Using the cumulative fiber metric, the prevalence of any radiographic change was 2.4% in the <1 fiber/cc-yr group, 5.0% in 1–10 fibers/cc-yr group, and 12.5% in the >10 fibers/cc-yr group. [Lockey et al. \(1984\)](#) used a modification of the ILO 1971 guidelines; one modification was that costophrenic angle blunting was considered a category separate from other pleural lesions. The results in [Lockey et al. \(1984\)](#) are presented in sufficient detail to allow interpretation according to the ILO 2000 guidelines ([ILO, 2002](#)).

A follow-up study of this cohort was conducted in 2002–2005 ([Rohs et al., 2008](#); see [Table 4-8](#)). This study included 298 workers, of which 280 completed the study interview (with work history and smoking history) and chest x-ray. The evaluation of each worker included an interview to determine work and health history, pulmonary examination, and chest x-ray. Exposure was estimated using the procedure previously described ([Lockey et al., 1984](#)). Exposure was assumed to occur from 1963 to 1980 in this study, assuming an 8-hour workday and 365 days of exposure per year ([Benson, 2014](#)). Each worker supplied a detailed work history (start and end date for each area within the facility). The exposure reconstruction resulted in a cumulative exposure estimate for each individual. The estimated cumulative exposure for this follow-up study ranged from 0.01 to 19.03 fibers/cc-yr (mean = 2.48). The time from first exposure ranged from 23 to 47 years. Exposure outside of work was assumed to be zero.



**Table 4-8. Pulmonary function and chest radiographic studies of the O.M. Scott, Marysville, OH plant workers**

Reference(s)	Inclusion criteria and design details			Results
Lockey (1985) Locket et al. (1984) <sup>a</sup>	1980, <i>n</i> = 512 Three exposure groups, based on jobs and area: Mean cumulative exposure <sup>b</sup>			Cumulative fiber exposure related to history of pleuritic chest pain and shortness of breath. No relation between cumulative exposure and forced vital capacity, forced expiratory volume, or diffusing capacity. Costophrenic angle blunting ( <i>n</i> = 11); other pleural thickening or plaques in ( <i>n</i> = 10); bilateral, small opacities ( <i>n</i> = 1). Abnormality (combined outcomes) increased with increasing cumulative exposure.
	Group I	0.45 fiber/cc-yr	( <i>n</i> = 112)	
	Group II	1.13 fibers/cc-yr	( <i>n</i> = 206)	
	Group III	6.16 fibers/cc-yr	( <i>n</i> = 194)	
	Radiographs read independently by two board-certified radiologists (B Readers), with a reading by a third reader when the initial two readings did not agree. Modification of ILO 1971 classification guidelines (e.g., separated costophrenic angle blunting from other pleural thickening) (see Table 4-4 for additional details)			
<a href="#">Rohs et al. (2008)</a>	2002–2005, interviews and chest x-rays conducted, <i>n</i> = 298; 280 with interviews and readable chest x-rays; Three B Readers based on 2000 ILO classification guidelines (see Table 4-4 for additional details)			Pleural abnormalities in 80 workers (28.7%). Small opacities ( $\geq 1/0$ ) in eight workers (2.9%). Increasing risk of pleural abnormalities with increasing cumulative fiber exposure: odds ratios (adjusting for age, date of hire, body mass index) by exposure quartile were 1.0 (referent), 2.7, 3.5, and 6.9.

<sup>a</sup>[Lockey et al. \(1984\)](#) is the published paper based on the unpublished thesis ([Lockey, 1985](#)).

<sup>b</sup>Calculated based on stratified data presented in Table 2 of [Lockey et al. \(1984\)](#).

Three board-certified radiologists, blinded to all identifiers, independently classified the radiographs using the 2000 ILO classification system ([ILO, 2002](#)). Pleural thickening (all sites) was reported as either localized pleural thickening or diffuse pleural thickening. Diffuse pleural thickening of the chest wall was recorded on the lateral chest wall “only in the presence of and in continuity with, an obliterated costophrenic angle” ([ILO, 2002](#)). Localized pleural thickening was described by [Rohs et al. \(2008\)](#) as “...(pleural) thickening with or without calcification, excluding solitary costophrenic angle blunting,” consistent with current ILO classification. Interstitial abnormalities indicative of asbestosis were considered present if the reader identified irregular opacities of profusion 1/0 or greater. A chest x-ray was defined as positive for pleural abnormality and/or interstitial abnormality when the median classification from the three readings was consistent with such effects. Radiographs classified as unreadable were not used (*n* not reported).

Pleural thickening was observed in 80 workers (28.7%), and small opacities ( $\geq 1/0$ ) were observed in 8 (2.9%). The 80 workers with pleural thickening include 68 with LPT (85%) and 12 with DPT (15%). Six of the eight participants with small opacities also had pleural

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thickening (four as LPT, two as DPT). The prevalence of pleural thickening increased across exposure quartiles from 7.1% in the first quartile to 24.6, 29.4, and 54.3% in the second, third, and fourth quartiles, respectively ([see Table 4-9; Rohs et al., 2008](#)).

**Table 4-9. Prevalence of pleural pathological alterations according to quartiles of cumulative fiber exposure in 280 participants**

Exposure quartile	Exposure range, fiber/cc-yr, (mean)	Number of workers	Number of workers with pleural thickening (%) <sup>b</sup>	Crude OR (95% CI)	Age-adjusted OR (95% CI)	BMI-adjusted OR (95% CI)	Number of workers with small opacities (%)
First	0.01–0.28 (0.12)	70	5 (7.1)	1.0 (referent)	1.0 (referent)	1.0 (referent)	0 (0)
Second	0.29–0.85 (0.56)	72 <sup>a</sup>	17 (24.6)	4.0 (1.4–11.6)	3.2 (1.0–9.7)	4.9 (1.3–18.2)	0 (0)
Third	0.86–2.20 (1.33)	68 <sup>a</sup>	20 <sup>c</sup> (29.4)	5.4 (1.9–15.5)	4.0 (1.3–12.8)	7.6 (2.1–27.5)	1 (1.5)
Fourth	2.21–19.03 (7.93)	70	38 (54.3)	15.4 (5.6–43)	10.0 (3.1–32)	17.0 (4.8–60.4)	7 (10)
Total	(2.48)	280	80 (28.6)				8 (2.9)

<sup>a</sup>Two observations in the second quartile and two in the third quartile had exact exposure values at the 50<sup>th</sup> percentile cutoff point. Rounding put these four observations in the second quartile.

<sup>b</sup>Statistically significant trend across exposure groups,  $p < 0.001$ .

<sup>c</sup>Typographical error in publication corrected.

OR = odds ratio; BMI = body mass index.

Source: [Rohs et al. \(2008\)](#), Table 3 and Figure 2; mean exposure levels and number of workers with parenchymal abnormalities by quartile obtained from J. Lockey, University of Cincinnati ([Benson, 2014](#)).

Pleural thickening was strongly associated with hire on or before 1973 and age at time of interview, but not with body mass index (BMI) or smoking history (ever smoked; see Table 4-10); BMI is a potentially important confounder because fat pads can sometime be misclassified as localized pleural thickening. A hire date of on or before 1973 and age at time of interview are each highly correlated with cumulative exposure to fibers. The small number of females ( $n = 16$ ) in the cohort limits the analysis of the association with gender.

**Table 4-10. Prevalence of pleural thickening in 280 participants according to various cofactors**

Variable	Number of workers	Number with pleural thickening (%)	Crude OR	95% CI	p-value
Hired after 1973	94	10 (10.6)	Reference		
Hired on or before 1973	186	70 (37.6)	5.07	2.47–10.41	<0.001
Body Mass Index, <sup>a</sup> kg/m <sup>2</sup>					
≤24.9	28	8 (28.6)	Reference		
25–29.9	101	31 (30.7)	1.11	0.44–2.79	0.52
≥30	110	27 (24.5)	0.81	0.32–2.06	0.43
Ever smoked <sup>b</sup>					
No	96	25 (26.04)	Reference		
Yes	184	55 (29.9)	1.21	0.70–2.11	0.50
Age at time of interview					
40–49	55	5 (9.1)	Reference		
50–59	116	28 (24.1)	3.18	1.16–8.76	0.03
≥60	109	47 (43.1)	7.58	2.80–20.49	<0.001
Female	16	1 (6.3)	Reference		
Male	264	79 (29.9)	6.40	0.83–49.32	0.07

<sup>a</sup>n = 239 for BMI due to 38 persons undergoing phone interview and 3 persons with onsite interviews who were not measured for height and weight.

<sup>b</sup>Smoking history as recorded in 2004 questionnaire. Of these 280 participants, 20 persons reported never smoking in the 1980 questionnaire but subsequently reported a history of smoking in the 2004 questionnaire (either current or ex-smoker).

Source: [Rohs et al. \(2008\)](#).

1 Odds ratios (ORs) for quartiles of cumulative fiber exposure were also estimated  
2 including various cofactors (age, hired before 1973, or BMI). Each model demonstrated the  
3 same trend: increased prevalence of pleural thickening with increasing cumulative exposure to  
4 fibers. Adjusting for age, date of hire, and BMI resulted in odds ratios of 2.7, 3.5, and 6.9 for the  
5 second, third, and fourth quartiles, respectively. There was no evidence of significant  
6 interactions using this modeling.

1       There was potential coexposure to a number of herbicides, pesticides, and other  
2 chemicals in the facility ([Smith, 2014](#)).<sup>15</sup> No quantitative information on exposure to these  
3 chemicals is available. However, the addition of the other chemicals to the vermiculite carrier  
4 occurred in a different part of the facility after expansion of the vermiculite ore. Industrial  
5 hygiene monitoring in these areas showed very low levels of fibers in the air. In addition, none  
6 of these other chemicals is volatile. Thus, it is unlikely that workers would be coexposed by  
7 inhalation to these other chemicals. In addition, EPA has no information indicating that  
8 exposure to any of these individual chemicals causes pleural thickening or small opacities typical  
9 of those found in workers employed in the Marysville facility, and thus EPA does not consider  
10 the presence of these coexposures likely to produce any confounding in the observed  
11 associations between LAA exposure and the pulmonary effects seen in this cohort.

12       The [Rohs et al. \(2008\)](#) study demonstrates that exposure to LAA can cause radiographic  
13 evidence of pleural thickening and parenchymal abnormalities (small opacities) in exposed  
14 workers. The prevalence of pleural pathological alterations was 28.7% in 2004 (80/280),  
15 compared to a 2% prevalence observed in 1984 (10/501). This increase in prevalence is most  
16 likely due to the additional time between the two studies giving additional time for the  
17 abnormalities to become apparent in conventional x-rays. The follow-up study also shows an  
18 increasing prevalence of pleural thickening with increasing cumulative exposure to LAA.

19       The influence of some potential sources of selection bias in [Rohs et al. \(2008\)](#) is difficult  
20 to qualitatively or quantitatively assess. One type of conceivable selection bias is the loss of  
21 participants due to the death of 84 of the 513 (16%) workers in the first study; this group may  
22 represent a less healthy or more susceptible population. Exclusion of the very sick or susceptible  
23 may imply that the population of eligible participants was somewhat healthier than the whole  
24 population of workers; this exclusion may result in an underestimation of risk. Another type of  
25 selection is the loss due to nonparticipation among the 431 individuals identified as alive in 2004  
26 ( $n = 135$  refusals and nonresponders; 31%). Participation rates in epidemiologic studies can be  
27 associated with better health status, and participation is often higher among nonsmokers  
28 compared with smokers. This type of selection of a relatively healthier group (among the living)  
29 could also result in an underestimation of the risk of observed abnormalities within the whole  
30 exposed population. However, if participation was related differentially based on exposure and  
31 outcome (i.e., if workers experiencing pulmonary effects and who were more highly exposed  
32 were more likely to participate than the highly exposed workers who were not experiencing

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<sup>15</sup>The herbicides and pesticides used during the time when Libby ore was used included atrazine, benomyl, bensulide, chloroneb, chlorothalonil, chlorpyrifos, 2,4-D, dacthal, diazinon, dicamba, dephenamid, disodium methanearsonate, dyrene, ethoprop, linuron, MCPP, monuron, neburon, oxadiazon, terrachlor, pentachlorophenol, phenylmercuric acetate, siduron, terrazole, thiophannate-methyl, and thiram. Other chemicals used included ammonium hydroxide, brilliant green crystals, caustic soda, corncobs, ferrous ammonium sulfate, ferrous sulfate, florex RVM, frit-504, frit-505, hi sil, lime, magnesium sulfate, mon-a-mon, potash, potassium sulfate, sudan orange, sudan red, sulfur, sulfuric acid, UFC, urea, and Victoria green liquid dye.

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pulmonary effects), the result would be to overestimate the exposure response relationship. This latter scenario is less likely to occur for asymptomatic effects (i.e., abnormalities detected by chest x-ray), such as those that are the focus of this study, than for symptoms such as shortness of breath or chest pain.

Some information is available on differences by participation status in the [Rohs et al. \(2008\)](#) study. Although current age was similar (mean: 59.1 and 59.4 years, respectively, in participants and living nonparticipant groups,  $p = 0.53$ ), participants were more likely to have been hired before or during 1973 (66.4 and 49.7%, respectively,  $p = 0.001$ ) and were also somewhat less likely to ever be smokers (58.6%) compared with the living nonparticipants (66.2%). Participants had higher mean exposure levels (mean cumulative exposure: 2.48 and 1.76 fibers/cc-yr, respectively, in participants and nonparticipants,  $p = 0.06$ ), but when combining living and deceased nonparticipants, there is no evidence of major differences in exposure distribution in participants compared with the original full population.

#### **4.1.2.2.3. Results: pathological alterations of parenchyma and pleura, pulmonary function, and respiratory symptoms—community-based studies**

##### *Pathological alterations of parenchyma and pleura*

In the ATSDR community health screening ([ATSDR, 2001b](#); see Table 4-15), two board-certified radiologists (B Readers) examined each radiograph, and a third reader was used in cases of disagreement (see Tables 4-5 and 4-11). Readers were aware that the radiographs were from participants in the Libby, MT health screening but were not made aware of exposure histories and other participant characteristics ([Peipins et al., 2004a](#); [Price, 2004](#); [Peipins et al., 2003](#)). The radiographs revealed pleural abnormalities in 17.9% of participants, with prevalence increasing with increasing number of “exposure pathways” (defined on the basis of potential work-related and residential exposure to asbestos within Libby and from other sources). The authors noted that the relationship between number of exposure pathways and increasing prevalence of pleural abnormalities was somewhat attenuated after excluding former workers from the vermiculite mining and milling operations. The prevalence of pleural anomalies decreased from approximately 35 to 30% in individuals with 12 or more exposure pathways when these workers were excluded from the analysis. Among individuals with no defined exposure pathways, the prevalence of pleural anomalies was 6.7%, which the authors report is higher than reported in other population studies ([Peipins et al., 2004a](#); [Price, 2004](#)). The direct comparability between study estimates is difficult to make; the possibility of over- or under ascertainment of findings from the x-rays based on knowledge of conditions in Libby was not assessed in this study. No information is provided regarding analyses excluding all potential work-related asbestos exposures.

[Weill et al. \(2011\)](#) used the ATSDR community health screening data to analyze the prevalence of x-ray abnormalities in relation to age, smoking history, and types of exposures (see

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1 Tables 4-5 and 4-11). Analysis was based on five exposure categories in  $n = 4,397$  participants  
2 ages 25 to 90 years. The prevalence of x-ray abnormalities (plaques, or diffuse pleural  
3 thickening, and/or costophrenic angle obliteration) also generally increased with age (divided  
4 into 25–40, 41–50, 51–60, and 61–90 years) within each of the exposure categories, with the  
5 highest prevalence seen among former workers in the vermiculite mining and milling operations.  
6 Among those with environmental exposure only (i.e., no household or occupational exposures),  
7 the prevalence increased from approximately 2% at ages 41–50 years to 12% at ages  
8 61–90 years.

9 The community-based study by [Alexander et al. \(2012\)](#) was conducted in an area other  
10 than Libby, MT. The Western Minerals plant in Minneapolis, MN processed Libby vermiculite  
11 ore to produce insulation material from 1939 to 1989. The plant was surrounded by residential  
12 neighborhoods, and the waste material from the plant was offered to community residents for use  
13 as filler in their yards and driveways. The Minnesota Department of Health and ATSDR  
14 initiated a study of community exposures in 2000, including a baseline survey of  
15 >6,400 residents. Residential history information was combined with period-specific air  
16 dispersion models and data on facility emissions to classify the level of background exposure  
17 ([Kelly et al., 2006](#)); details pertaining to the input parameters and modeling assumptions are  
18 limited and result in considerable uncertainty in exposure estimates.<sup>16</sup> Intermittent high  
19 exposures were estimated for specific activities (e.g., playing on waste piles, moving waste from  
20 the plant), based on experiments reconstructing exposure occurring during these activities  
21 ([Adgate et al., 2011](#)). In a follow-up study of people who had not worked in the plant (or lived  
22 with a worker), measures of background exposure and activity-based (intermittent) exposure  
23 were associated with increased prevalence of pleural abnormalities ([Alexander et al., 2012; see](#)  
24 [Table 4-11](#)).

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<sup>16</sup>Based on review of supporting documentation provided by Gregory Pratt (Minnesota Department of Pollution Control).

**Table 4-11. Pathological alterations of parenchyma and pleura in community-based studies**

Reference(s)	Inclusion criteria and design details	Results				
Libby, MT community						
<a href="#">Peipins et al. (2003)</a> <a href="#">ATSDR (2001b)</a>	Participants in ATSDR community health screening, <i>n</i> = 6,668 with chest x-rays (see Table 4-5). 19 “exposure pathways” including Libby mining company work, contractor work, dust exposure at other jobs, vermiculite exposure at other jobs, potential asbestos exposure at other jobs or in the military, cohabitation with Libby mining company worker, and residential and recreational use of vermiculite. Pleural abnormality: (a) any unilateral or bilateral pleural calcification on the diaphragm, chest wall, or other site or (b) any unilateral or bilateral pleural thickening or plaque on the chest wall, diaphragm, or costophrenic angle site, consistent with asbestos-related pleural disease.	<a href="#">Peipins et al. (2003)</a> and <a href="#">ATSDR (2001b)</a> : Pleural abnormalities seen in 17.9% of participants; increasing prevalence with increasing number of exposure pathways (6.7% among those with no specific pathways, 34.6% among those with 12 or more pathways).  <a href="#">ATSDR (2001b)</a> : Moderate-to-severe FVC <sub>1</sub> restriction (FVC <70% predicted): 2.2% of men >17 yr old; 1.6% of women >17 yr old				
<a href="#">Weill et al. (2011)</a>	Participants in ATSDR community health screening, <i>n</i> = 4,397 ages 25 to 90 yr (see Table 4-5). Analysis based on five exposure categories: (1) Vermiculite mining or milling workers employed directly by the company (W.R. Grace; <i>n</i> = 255), (2) other vermiculite worker (contractor work; <i>n</i> = 664), (3) dusty occupation ( <i>n</i> = 831), (4) household (combination of three household categories; <i>n</i> = 880), and (5) environment (“no” to work and household exposures in Categories 1–4; <i>n</i> = 1,894). Profusion ≥1/0: defined as “any two readers reporting any profusion ≥1/0” Plaque: defined as “any two readers reporting any diaphragm or wall, or other site plaques, even if the readers did not agree on specifics.” DPT or CAO: defined as “any two readers reporting any DPT or CAO, even if the readers did not agree on specifics.”	Exposure Source	Prevalence (%)			
			Profusion			
			≥1/0	Plaque	DPT/CAO	
		Age 25–40 ( <i>n</i> = 1,075)				
		Vermiculite worker by company	0.0	20.0	5.0	
		Other vermiculite worker	0.8	0.8	0.0	
		Dusty work	0.0	3.8	0.4	
		Household	0.0	2.2	0.0	
		Environment	0.0	0.4	0.0	
		Age 41–50 ( <i>n</i> = 1,187)				
		Vermiculite worker by company	0.0	26.2	5.0	
		Other vermiculite worker	0.5	7.8	1.0	
		Dusty work	0.0	3.8	0.9	
Household	0.0	11.1	0.4			
Environment	0.0	1.9	0.2			



**Table 4-11. Pathological alterations of parenchyma and pleura in community-based studies (continued)**

Reference(s)	Inclusion criteria and design details	Results				
<a href="#">Weill et al. (2011)</a> (continued)		Exposure Source	Prevalence (%)			
			Profusion			
			≥1/0	Plaque	DPT/CAO	
		Age 51–60 (n = 1,034 )				
		Vermiculite worker by company	3.2	34.9	3.2	
		Other vermiculite worker	0.6	13.7	0.6	
		Dusty work	1.0	12.6	0.0	
		Household	1.0	20.1	1.5	
		Environment	0.0	7.7	0.9	
		Age 61–90 (n = 1,101 )				
		Vermiculite worker by company	11.0	45.7	8.6	
		Other vermiculite worker	0.6	24.8	8.5	
		Dusty work	1.1	21.9	3.3	
		Household	2.4	38.3	5.7	
		Environment	1.3	12.7	2.2	
Minneapolis, MN community						
<a href="#">Alexander et al. (2012)</a>	Participants with personal or family work history at the plant; 1,765 of 2,222 individuals randomly chosen within three strata based on exposure scenarios, (intense intermittent, long-term high ambient background, and low ambient background); n = 461 completed the study. Clinical examination, chest x-rays read by 2000 ILO classification guidelines (posterior-anterior). Participants more likely than nonparticipants to report exposure-related activities (48 and 32%, respectively), but similar history of occupational asbestos exposure (28 and 27%, respectively). Exposure based on modeling by <a href="#">Kelly et al. (2006)</a> and <a href="#">Adgate et al. (2011)</a> .	-		<u>Prevalence</u>		
Pleural abnormality (any)		49 (10.6%)				
DPT		5 (1.1%)				
LPT (referred to as “pleural plaques” by study authors		45 (9.9%)				
Regression analysis:						
Exposure type		Beta(±SE)	OR (95% CI)			
Background		0.322 (±0.125)	1.38 (1.08, 1.77)			
Intermittent		0.063 (±0.039)	1.07 (0.99, 1.15)			
Per unit increase in exposure measure; adjusted for yr of birth, history of asbestos-related job, gender, and the other exposure measure.						

SE = standard error

## Respiratory symptoms

[Vinikoor et al. \(2010\)](#) used the 2000–2001 health screening data to examine respiratory symptoms and pulmonary function results among 1,003 adolescents and young adults (≤18 years in 1990 when the mining/milling operations closed), excluding individuals with a work history

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that could result in vermiculite or dust exposure (see Tables 4-5 and 4-12). The potential for vermiculite exposure outside of the workplace was classified based on responses to questions about six activities (e.g., handling vermiculite insulation, playing in vermiculite piles, “popping” vermiculite by heating it to make it expand). The medical history questionnaire included information on three respiratory symptoms: (1) usually have a cough ( $n = 108$ , 10.8%); (2) troubled by shortness of breath when walking up a slight hill or when hurrying on level ground ( $n = 145$ , 14.5%); or (3) coughed up phlegm that was bloody in the past year ( $n = 59$ , 5.9%). A question on history of physician-diagnosed lung disease ( $n = 51$ , 5.1%) was also included. The pulmonary function results were classified as normal in 896 (90.5%), obstructive in 62 (6.3%), restrictive in 30 (3.0%), and mixed in 2 (0.2%). There was little variation in prevalence of shortness of breath, physician-diagnosed lung disease, or abnormal spirometry across the exposure categories; for two other symptoms, the highest relative risk was seen in the highest exposure group, but neither of these estimates was statistically significant (OR 2.93, 95% confidence interval (CI) 0.93, 9.25 for usually having a cough and OR 1.49, 95% CI: 0.41, 5.43 for coughing up bloody phlegm).

**Table 4-12. Pulmonary function and respiratory system changes in the Libby, MT community**

Reference(s)	Inclusion criteria and design details	Results			
<a href="#">Vinikoor et al. (2010)</a>	Participants in the ATSDR community health screening (see Table 4-5); limited to $n = 1,003$ , ages 10–29 yr when screened (age $\leq 18$ yr in 1990 when the mining/milling operations closed). Excluded if employed in vermiculite mining or milling operations, exposed to dust at other jobs, or exposed to vermiculite at other jobs. Analysis of respiratory symptoms and spirometry in relation to six vermiculite exposure activities (handling vermiculite insulation, recreational activities on a vermiculite-contaminated gravel road leading to the mine, playing at ball fields near the expansion plant, playing in or around the vermiculite piles, heating the vermiculite to “pop” it, other activities involving vermiculite).		OR (95% CI) <sup>a</sup>		
			Sometimes	Frequently 1–2 activities	Frequently $\geq 3$ activities
		Usual cough	1.88 (0.71, 5.00)	2.00 (0.76, 5.28)	2.93 (0.93, 9.25)
		Shortness of breath	1.16 (0.55, 2.44)	1.27 (0.61, 2.63)	1.32 (0.51, 3.42)
		Bloody phlegm	0.85 (0.31, 2.38)	1.09 (0.41, 2.98)	1.49 (0.41, 5.43)
		Physician-diagnosed lung disease	1.95 (0.57, 6.71)	1.51 (0.43, 5.24)	1.72 (0.36, 8.32)
		Abnormal spirometry <sup>b</sup>	1.34 (0.60, 2.96)	1.20 (0.53, 2.70)	1.33 (0.42, 4.19)
		<sup>a</sup> Adjusted for age, gender, personal smoking history, and living with a smoker; referent group = “never” response to each of the six vermiculite exposure activities.			
<sup>b</sup> Obstructive ( $FEV_1/FVC < LLN$ and $FVC \geq LLN$ ), restrictive ( $FEV_1/FVC \geq LLN$ and $FVC < LLN$ ) or mixed ( $FEV_1/FVC < LLN$ and $FVC < LLN$ ) compared with normal ( $FEV_1/FVC \geq LLN$ and $FVC \geq LLN$ ).					

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1 *Pathological alterations of parenchyma and pleura in relation to pulmonary function*

2 Two studies that examined LAA specifically provide insight into the question of the  
3 relation between specific types pleural or parenchymal lesions and pulmonary function (see  
4 Table 4-13). These studies, based on data from the community health screening conducted by  
5 ATSDR, reported an association between the presence of pleural plaques and a mean decrement  
6 (approximately 5%) in FVC measures ([Weill et al., 2011](#)), and of an increased risk of restrictive  
7 pulmonary function ([Larson et al., 2012b](#)). The authors of the first study ([Weill et al., 2011](#))  
8 concluded that the change in FVC was “probably clinically insignificant.” The second study  
9 ([Larson et al., 2012b](#)), focused on the likelihood of an “abnormal” pulmonary function test  
10 (restrictive lung function), rather than on a difference in the mean of the distribution. Although  
11 the association with a restrictive pulmonary function was weaker for circumscribed pleural  
12 plaques (OR = 1.4) than for DPT (OR = 4.1), increasing risk was seen with increasing size and  
13 across categories of pulmonary impairment (OR 1.7, 2.1, and 2.3, respectively, for mild,  
14 moderate and severe). These two analyses, using essentially the same data set, illustrate that the  
15 clinical perspective of an “insignificant” decrement in lung function (i.e., a small mean  
16 difference) is compatible with a population perspective of an increased risk of an adverse  
17 outcome.

18  
19 **4.1.2.2.4. Clinic-based reports and case reports of respiratory disease (noncancer).**

20 [Whitehouse \(2004\)](#) examined changes in pulmonary function measures in 123 patients  
21 (86 former employees of the vermiculite operations, 27 family members of employees, 10 Libby  
22 residents with only environmental exposures) seen in a pulmonary disease practice serving the  
23 Libby, MT area. The mean age of study participants was 66 years, and the mean follow-up time  
24 was 35 months. Chest x-rays or high resolution computed tomography scans revealed no  
25 evidence of interstitial changes in 67 (55%) of the 123 patients, and 56 patients (45%) were  
26 found to have interstitial changes at profusion Category 0/1 or 1/0. The average yearly loss of  
27 pulmonary function was 2.2% for FVC, 2.3% for total lung capacity, and 3.0% for DLCO.

**Table 4-13. Analyses of pulmonary changes seen on radiographs in relation to pulmonary function in the Libby, MT community**

Reference(s)	Inclusion criteria and design details	Results			
		Radiographic results	% Predicted FVC		
<a href="#">Weill et al. (2011)</a>	Participants in ATSDR community health screening, $n = 4,397$ , ages 25 to 90 yr. ILO 1980 classification guidelines. Profusion $\geq 1/0$ : any two readers reporting any profusion $\geq 1/0$ . Plaque: any two readers reporting any diaphragm or wall, or other site plaques, even if the readers did not agree on specifics. DPT or CAO: defined as any two readers reporting any DPT or CAO, even if the readers did not agree on specifics.		$n$	Mean	( $\pm$ SE)
		DPT, CAO	33	78.76	( $\pm 3.64$ )
		Profusion $\geq 1/0$	40	82.16	( $\pm 3.34$ )
		Other pleural abnormality	482	95.63	( $\pm 0.76$ )
		None of above	4,065	103.15	( $\pm 0.25$ )
<a href="#">Larson et al. (2012b)</a>	Participants in the ATSDR community health screening, $n = 6,476$ , ages $\geq 18$ yr. Pulmonary function classified as normal, restrictive only ( $FVC < \text{lower limit of normal}$ and $FEV_1/FVC > \text{lower limit of normal}$ ), obstructive only, ( $FVC \geq \text{lower limit of normal}$ and $FEV_1/FVC < \text{lower limit of normal}$ ) or mixed based on reference values for FVC and $FEV_1/FVC$ for the U.S. population ( <a href="#">Hankinson et al., 1999</a> ). Analysis adjusted for parenchymal abnormalities, age, gender, smoking history, BMI, exposure group, number of exposure pathways, duration of residence in Libby, and shortness of breath; referent group = “normal” pulmonary function. ILO 1980 classification guidelines modified such that plaques definition was equivalent to ILO 2000 LPT guidelines.	Radiographic results	Risk of restriction		
			$n$	OR	(95% CI)
		DPT, CAO	58	4.1	(2.1, 7.8)
		Profusion $\geq 1/0$	50	2.9	(1.4, 6.0)
		Calcification	254	2.7	(1.2, 2.4)
		Circumscribed plaques	708	1.4	(1.1, 1.8)
		<i>By index of degree of abnormality (median = 3.0 for DPT, 2.5 for LPT)</i>			
		DPT Only			
		$\leq \text{median}$	78	2.1	(1.1, 3.7)
		$> \text{median}$	57	5.6	(2.7, 11.6)
		Index of plaque size			
		$\leq \text{median}$	562	1.3	(1.0, 1.7)
		$> \text{median}$	499	1.9	(1.5, 2.5)
		<i>By severity of impairment, for index of plaque size <math>&gt; \text{median}</math></i>			
		Mild ( $FEV_1 > 70\%$ )	63	1.7	(1.3, 2.5)
		Moderate ( $50\% \leq FEV_1 < 69\%$ )	50	2.1	(1.4, 3.2)
		Severe ( $FEV_1 < 50\%$ )	6	2.3	(0.8, 6.7)

1 A study by [Winters et al. \(2012\)](#) was conducted among patients seen for annual  
2 examinations at a clinic in Libby, MT specializing in the diagnosis and treatment of  
3 asbestos-related disease (see Table 4-14). The x-rays (posterior-anterior and lateral views) were  
4 read by one radiologist. In this clinic sample, 60 individuals were considered to have no  
5 abnormality, 182 had pleural abnormality only, 18 had interstitial abnormality, and 69 had both  
6 pleural and interstitial abnormalities. FVC was lower among those with pleural abnormalities

compared with the no abnormalities group, and the decrement in FEV<sub>1</sub> was similar for the pleural and the interstitial abnormalities groups. Higher scores on the respiratory quality of life scale (indicating increased impairment) were also seen in relation the presence of pleural abnormalities. One limitation of this study is that the ILO classification criteria were not used and a description or definition of the classification categories was not provided; in addition, factors influencing the decision of residents (and past residents) to receive asbestos-related health care through this clinic introduces additional challenges to the interpretation of these data.

**Table 4-14. Pulmonary function and respiratory system changes in the Libby, MT community: clinic-based study**

Reference(s)	Inclusion criteria and design details	Results					
<a href="#">Winters et al. (2012)</a>	Patients seen at Center for Asbestos-Related Disease clinic in Libby, MT. N = 329 (2/3 local, 1/3 distant), seen for annual examination; 70% between ages 50–69 yr. (156 other patients excluded because of missing data). Analysis of chest x-ray (diagnostic criteria not provided), pulmonary function, and respiratory health quality of life (questionnaire).		Mean (SD), by group				
			Abnormalities				
		Normal (n = 60)	Pleural (n = 182)	Interstitial (n = 18)	Both (n = 69)		
		Pulmonary function					
		FVC	103.8 (15.0)	94.9 (20.2) <sup>a</sup>	95.6 (12.0)	88.8 (16.3)	
		FEV <sub>1</sub>	92.8 (21.8)	87.7 (20.0)	86.4 (16.5)	80.4 (18.6)	
		FEV <sub>1</sub> /FVC <sub>1</sub>	94.8 (13.6)	95.9 (10.8)	94.7 (12.6)	95.6 (13.5)	
		DLCO	90.2 (19.5)	85.7 (20.1)	68.1 (24.5)	73.7 (22.4)	
		Respiratory symptoms and quality of life <sup>b</sup>					
		Total score	29.8 (20.8)	39.1 (22.5) <sup>a</sup>	41.3 (21.5)	43.8 (20.4)	
		Symptoms	47.1 (24.2)	51.8 (25.1)	54.5 (27.0)	56.3 (23.8)	
		Activity	36.5 (26.8)	49.9 (27.4) <sup>a</sup>	54.6 (22.3)	57.7 (23.0)	
		Impact	20.1 (19.2)	28.4 (21.2) <sup>a</sup>	28.5 (23.4)	31.4 (21.8)	
		<sup>a</sup> p < 0.05 for comparison with no abnormality (“normal”) group. <sup>b</sup> Measured with St. George Respiratory Questionnaire; total score from 0 (better health) to 100 (worse health); symptoms = frequency and severity of respiratory symptoms; activity = activities that cause or are limited by breathlessness; impact = social function and psychological disturbances related to respiratory problems.					

Additional studies in the form of case reports provide ancillary evidence that exposure to LAA may lead to respiratory disease. Progressive disease from exposure to LAA was noted in a case report of fatal asbestosis in an individual who died 50 years after working at a vermiculite processing plant for a few months at about age 17 ([Wright et al., 2002](#)). In another case report, exposures that stemmed from playing for a few years as a child in contaminated vermiculite waste materials around a former Libby vermiculite processing facility was reportedly associated

with the development of asbestosis and fatal lung cancer ([Srebro and Roggli, 1994](#)). Although these case reports do not provide quantitative exposure measures, they do illustrate the potential for demonstrable health effects from exposures of short duration and from exposures in a community setting.

**4.1.2.2.5. Summary of respiratory effects, other than cancer.** Epidemiology studies demonstrate consistent results pertaining to the association between LAA exposure and various forms of respiratory effects, with effects seen in worker populations and in populations with residential (nonoccupational) routes of exposure. The risk of mortality related to asbestosis and other forms of nonmalignant respiratory disease is elevated in the Libby vermiculite mining and processing operations, with a pattern of increasing risk with increasing cumulative exposure (more than a 10-fold increased risk of asbestosis and a 1.5- to 3-fold increased risk of nonmalignant respiratory disease) in the analyses using internal, referent groups in [McDonald et al. \(2004\)](#), [Sullivan \(2007\)](#), and [Larson et al. \(2010b\)](#). Radiographic evidence of small opacities (evidence of parenchymal damage) and pleural thickening (pleural plaques, LPT, and DPT) has also been shown in studies of Libby workers ([Larson et al., 2012a](#); [Larson et al., 2010a](#); [Whitehouse, 2004](#); [Amandus et al., 1987a](#); [McDonald et al., 1986b](#)), and in the studies of workers in the Marysville, OH plant ([Rohs et al., 2008](#); [Lockey et al., 1984](#)). In the Marysville cohort, the prevalence of small opacities (interstitial changes in the lung) increased from 0.2% in the original study to 2.9% in the follow-up study, and the prevalence of pleural thickening increased from 2 to 28.6%. No effects on lung function were found in the original study ([Lockey et al., 1984](#)), and lung function was not reported for the [Rohs et al. \(2008\)](#) analysis of the cohort follow-up. Data from the ATSDR community health screening study in Libby, MT indicate that the prevalence of pleural abnormalities, identified by radiographic examination, increases substantially with increasing number of exposure pathways ([Peipins et al., 2003](#)). The presence of pleural plaques is associated with a small decrement in lung function (approximately 5%) when evaluated based on mean values ([Weill et al., 2011](#)), and presence of LPT is associated with an increased risk of restrictive lung function ([Larson et al., 2012b](#)). Additional evidence of respiratory effects of LAA exposure comes from the study of residents in an area surrounding a processing plant in Minneapolis, MN ([Alexander et al., 2012](#)).

### **4.1.3. Other Effects, Noncancer**

#### **4.1.3.1. Cardiovascular Disease**

[Larson et al. \(2010b\)](#) presents data on mortality due to cardiovascular diseases (CVDs) among the Libby cohort of vermiculite workers, with SMRs of 0.9 (95% CI: 0.9, 1.0) for heart disease ( $n = 552$ ) and 1.4 (95% CI: 1.2, 1.6) for circulatory system diseases ( $n = 258$ ). Deaths due to heart diseases were further categorized into ischemic heart disease ( $n = 247$ ) and other

heart disease ( $n = 120$ , for pericarditis, endocarditis, heart failure, and ill-defined descriptions and complications of heart disease), with SMRs of 0.7 (95% CI: 0.6, 0.8) and 1.5 (95% 1.2, 1.8), respectively. Circulatory diseases included hypertension without heart disease ( $n = 42$ ), with an SMR of 1.7 (95% CI: 1.2, 2.4) and diseases of arteries, veins, or lymphatic vessels ( $n = 136$ ), SMR = 1.6 (95% CI: 1.4, 2.0). The combined category of cardiovascular-related mortality resulted in modestly increased risks across quartiles of exposure, with RRs of 1.0 (referent), 1.3 (95% CI: 1.0, 1.6), 1.3 (95% CI: 1.0, 1.6), and 1.5 (95% CI: 1.1, 2.0) with exposure groups of  $<1.4$ , 1.4 to  $<8.6$ , 8.6 to  $<44.0$ , and  $\geq 44.0$  fibers/cc-yr, respectively. In the Monte Carlo simulation used to estimate the potential bias in cardiovascular disease risk that could have been introduced by differences in smoking patterns between exposed and unexposed workers in the cohort, the bias adjustment factor was relatively small ( $RR_{unadjusted}/RR_{adjusted} = 1.1$ ), reducing the overall RR estimate from 1.6 to 1.5. The observed association between asbestos exposure and cardiovascular disease-related mortality may reflect, at least in part, a consequence of an underlying respiratory disease.

#### **4.1.3.2. Autoimmune Disease and Autoantibodies**

Three epidemiology studies have examined the potential role of LAA and autoimmunity. [Noonan et al. \(2006\)](#) used the data from the community health screening to examine self-reported histories of autoimmune diseases (rheumatoid arthritis, scleroderma, or lupus) in relation to the asbestos exposure pathways described above (see Tables 4-5 and 4-15). To provide more specificity in the self-reported history of these diseases, a follow-up questionnaire was mailed to participants to confirm the initial report and obtain clarifying information regarding the type of disease, whether the condition had been diagnosed by a physician, and whether the participant was currently taking medication for the disease. Responses were obtained from 208 (42%) of the 494 individuals who had reported these conditions. Of these 208 responses, 129 repeated the initial report of the diagnosis of rheumatoid arthritis, and 161 repeated the initial report of the diagnosis of one of the three diseases (rheumatoid arthritis, scleroderma, or lupus); approximately 70% of those confirming the diagnosis also reported taking medication for the condition. Among people aged 65 and over ( $n = 34$  rheumatoid arthritis cases, determined using responses from the follow-up questionnaire), a twofold to threefold increase in risk was observed in association with several measures reflecting potential exposure to asbestos (e.g., asbestos exposure in the military) or specifically to LAA (e.g., past work in mining and milling operations, use of vermiculite in gardening, and frequent playing on vermiculite piles when young). Restricted forced vital capacity (defined as FVC  $<80\%$  predicted and a ratio of FEV<sub>1</sub> to FVC  $\geq 70\%$  predicted), presence of parenchymal abnormalities, playing on vermiculite piles, and other dust or vermiculite exposures were also associated with rheumatoid arthritis in the group younger than 65 years ( $n = 95$  cases). For all participants, an increased risk of rheumatoid arthritis was observed with increasing number of exposure

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1 pathways. Although the information gathered in the follow-up questionnaire and repeated  
2 reports of certain diagnoses decreased the false-positive reports of disease, the reliance of  
3 self-reported data is a limitation of this study. Considerable misclassification (over-reporting  
4 and under-reporting) is likely, given the relatively low confirmation rate of self-reports of  
5 physician-diagnosed rheumatoid arthritis (and other autoimmune diseases) seen in other studies  
6 ([Karlson et al., 2003](#); [Rasch et al., 2003](#); [Ling et al., 2000](#)).

7 Another study examined serological measures of autoantibodies in 50 residents of Libby,  
8 MT, and a comparison group of residents of Missoula, MT ([Pfau et al., 2005](#); see Table 4-16).  
9 The Libby residents were recruited for a study of genetic susceptibility to asbestos-related lung  
10 disease, and the Missoula residents were participants in a study of immune function. None of the  
11 50 Missoula residents and three of the Libby participants reported a history of a rheumatoid  
12 arthritis, systemic lupus erythematosus, or other systemic autoimmune disease (SAID). Libby  
13 residents exhibited an increased prevalence (22%) of high-titer ( $\geq 1:320$ ) antinuclear antibodies  
14 when compared to Missoula residents (6%), and similar increases were seen in the Libby  
15 samples for rheumatoid factor, antiribonucleoprotein (RNP), anti-Scl-60, anti-Sm, anti-Ro  
16 (SSA), and anti-La (SSB) antibodies. Although neither sampling approach was based on a  
17 random selection from the community residents, an individual's interest in participating in a gene  
18 and lung-disease study would not likely be influenced by the presence of autoimmune disease or  
19 autoantibodies in that individual. Thus selection bias would not be considered likely in this  
20 study.

21 In a follow-up study, [Marchand et al. \(2012\)](#) examined the association of autoantibodies  
22 with asbestos-related lung disease in 124 Libby residents (65 female, 59 male). Serum samples  
23 were tested for the presence of antimesothelial cell antibodies (MCAA) to determine if the  
24 mesothelial cells of the pleural lining are targets for autoimmune responses. Mean  
25 concentrations of MCAA were increased in Libby residents, particularly in those with lung or  
26 pleural lesions compared to a reference population of Missoula residents. In addition, Libby  
27 residents positive for antinuclear antibodies or MCAA had an increased odds ratio of also  
28 presenting with pleural abnormalities (odds ratio 3.60 and 4.88, respectively). However, the  
29 cross-sectional nature of this study design makes it difficult to determine if these autoantibodies  
30 were a principal mechanism for inducing pleural disease, or if their presence is an indication of  
31 tissue damage.

**Table 4-15. Autoimmune-related studies in the Libby, MT community**

Reference(s)	Inclusion criteria and design details	Results		
<a href="#">Noonan et al. (2006)</a>	Nested case-control study among 7,307 participants in 2000–2001 community health screening. Conducted interviews, gathered self-reported history of rheumatoid arthritis, scleroderma, or lupus. Follow-up questionnaire mailed to participants concerning self-report of “physician-diagnosis” of these diseases and medication use.	Association with work in Libby mining/milling operations (ages 65 and older): Rheumatoid arthritis OR: 3.2 (95% CI: 1.3, 8.0) Rheumatoid arthritis, lupus, scleroderma OR: 2.1 (95% CI: 0.90, 4.1)		
		Risk increased with increasing number of asbestos exposure pathways:		
		Zero pathways:	1.0	(Referent)
		One pathway:	1.02	
		Two to three pathways:	1.79	
		Four to five pathways:	2.51	
		≥Six pathways:	3.98	
		(trend $p < 0.001$ , adjusting for restrictive spirometry, parenchymal abnormalities, and smoking history)		
<a href="#">Pfau et al. (2005)</a>	Libby residents ( $n = 50$ ) recruited for study of genetic susceptibility to asbestos-related lung disease. Missoula, MT comparison group ( $n = 50$ ), recruited for study of immune function; age- and gender-matched to Libby participants. Serum samples obtained: IgA levels; prevalence of antinuclear, anti-dsDNA, anti-rheumatoid factor, anti-Sm, anti-RNP, anti-Ro, anti-La, and anti-Scl-70 antibodies.	Increased prevalence of high-titer ( $\geq 1:320$ ) antinuclear antibodies in Libby sample (22%) compared to Missoula sample (6%). Similar increases for rheumatoid factor, anti-RNP, anti-Scl-60, anti-Sm, anti-R <sub>o</sub> (SSA), and anti-La (SSB) antibodies observed in Libby sample.		
<a href="#">Marchand et al. (2012)</a>	Follow-up to <a href="#">Pfau et al. (2005)</a> study (see row above). Randomly selected 124 out of 318 banked samples from Libby residents, mean age 50 yr (ranging from 14 to 84 yr). Compared with 25 samples from Missoula, MT, mean age 45 yr (ranging from 19 to 78 yr); positive autoantibody test for mesothelial cells defined based on mean + 3SD of Missoula samples. Results of chest radiographs for Libby residents obtained from community screening program described in <a href="#">ATSDR (2001b)</a> ; see <a href="#">Section 4.1.2, Table 4 11</a> .	Prevalence in Libby residents:		
		Pleural abnormalities		25%
		Interstitial abnormalities only		52%
		No abnormalities		23%
		Association with pulmonary abnormality:		
	$n$	(%)	OR	
ANA	76	(61.3)	3.6	
MCAA	23	(18.5)	3.8	

ANA = antinuclear antibody; dsDNA = double-stranded DNA; MCAA = antimesothelial cell autoantibodies; RNP = ribonucleoprotein.



#### 4.1.4. Cancer Effects

##### 4.1.4.1. Lung Cancer

Several analyses of the mortality experience of Libby vermiculite workers have been conducted (see Section 4.1.1 for a summary of exposure measures and cohort descriptions for these studies). The studies of the Libby worker cohort by [Amandus and Wheeler \(1987\)](#), [Sullivan \(2007\)](#), and [Larson et al. \(2010b\)](#) defined lung cancer mortality based on a more specific cause of death codes (e.g., cancers of the trachea, bronchus, and lung) compared to the broader classification of “all respiratory cancer” used by McDonald et al. (2004; 1986a), which would include laryngeal and “other” respiratory cancers. In the national Surveillance, Epidemiology, and End Results cancer data from 2003–2007, the age-adjusted mortality rate for cancer of the larynx was 1.2 per 100,000 person-year, compared to 52.5 per 100,000 person-year for lung and bronchial cancer ([NCI, 2011](#)). Thus, these additional categories (larynx and “other” respiratory cancers) represent a relatively small proportion of respiratory cancers. Although they could also be a source of some misclassification of the outcome if these other cancers are not related to asbestos exposure, the magnitude of this bias would be small.

In the more recent study by [McDonald et al. \(2004\)](#), 44 respiratory cancers were observed among 406 men who had worked at least 1 year in the vermiculite mining and milling facilities. [Sullivan \(2007\)](#) and [Larson et al. \(2010b\)](#) included workers with less than 1 year of work, resulting in a larger sample size (approximately 1,700) and more than 80 lung cancer deaths. Each of these studies observed an increased overall risk, with SMRs of 1.4, 16, and 2.4, respectively in [Sullivan \(2007\)](#), [Larson et al. \(2010b\)](#), and [McDonald et al. \(2004\)](#). Exposure-response analyses from these studies demonstrated increasing mortality with increasing exposure, using categorical and continuous measures of exposure, different lag periods, and different exposure metrics, with approximately a twofold to threefold increased risk in the highest exposure group (see Table 4-16 and Figure 4-3).

**Table 4-16. Respiratory (lung) cancer mortality and exposure-response analyses based on related studies of the vermiculite mining and milling workers in Libby, MT<sup>a</sup>**

Reference(s)	Inclusion criteria and design details	Standardized mortality ratio (95% CI)	Exposure-response analyses—lung cancer		
<a href="#">Amandus and Wheeler (1987)</a>	Men, hired before 1970, worked at least 1 yr, follow-up through 1982 ( $n = 575$ ); 161 deaths (159 with death certificates) Mean duration: 8.3 yr Mean fiber-yr: 200.3 12 female workers not included in this analysis	<i>No exclusions:</i> All cancer ( $n = 38$ ) SMR: 1.3 (0.9, 1.8) Lung ( $n = 20$ ) SMR: 2.2 (1.4, 3.4)  <i>20 or more yr since first hire (latency):</i> Lung ( $n = 12$ ) SMR: 2.3 ( $p < 0.05$ )	<i>No exclusions:</i>		
			Cumulative exposure	$n$	SMR (95% CI) <sup>b</sup>
			0.0–49 fibers/cc-yr	6	1.5 (not reported)
			50–99 fibers/cc-yr	2	1.6 (not reported)
			100–399 fibers/cc-yr	2	1.1 (not reported)
			≥400 fibers/cc-yr	10	5.8 (not reported, but $p < 0.01$ )
			<i>20 or more yr since first hire (20-yr latency)</i>		
			Cumulative exposure	$n$	SMR (95% CI) <sup>b</sup>
			0.0–49 fibers/cc-yr	2	0.85 (not reported)
			50–99 fibers/cc-yr	2	2.3 (not reported)
			100–399 fibers/cc-yr	1	1.1 (not reported)
			≥400 fibers/cc-yr	7	6.7 (not reported, but $p < 0.01$ )
			In linear regression analysis of data with at least 20-yr latency, results per fiber-yr: beta (standard error) = 0.60 (0.13) and 0.58 (0.08), respectively, for threshold and nonthreshold models. Using a survival (Cox) model, the corresponding estimate is 0.11 (0.04). All estimates are statistically significant ( $p < 0.05$ ).		

**Table 4-16. Respiratory (lung) cancer mortality and exposure-response analyses based on related studies of the vermiculite mining and milling workers in Libby, MT<sup>a</sup> (continued)**

Reference(s)	Inclusion criteria and design details	Standardized mortality ratio (95% CI)	Exposure-response analyses—lung cancer			
McDonald et al. (2004; 1986a)	Men, hired before 1963, worked at least 1 yr ( <i>n</i> = 406); follow-up through 1999 (McDonald et al., 2004); 165 deaths before July 1983 (163 with death certificates); 120 deaths July 1983–1998 coded by nosologists using ICD-8 classifications; cause of death for deaths from 1983–1998 obtained from National Death Index. Mean duration: 8.7 yr Mean fiber-yr: 144.6	Respiratory ( <i>n</i> = 44) SMR: 2.4 (1.7, 3.2)	<i>Excluding first 10 yr of follow-up:</i>			
			Cumulative exposure	<i>n</i>	RR (95% CI) <sup>d</sup>	
			0.0–11.6 fibers/cc-yr	5	1.0 (referent)	
			11.7–25.1 fibers/cc-yr	9	1.7 (0.58, 5.2)	
			25.2–113.7 fibers/cc-yr	10	1.9 (0.63, 5.5)	
			≥113.8 fibers/cc-yr	16	3.2 (1.2, 8.8)	
			per 100 fibers/cc-yr increase (linear model, RR = 1 + b*exposure) 0.36 (0.03, 1.2) ( <i>p</i> = 0.02) Similar patterns were reported for analyses of intensity and residence-weighted exposure, but results not presented in paper.			
Sullivan (2007)	White men, enumerated in 1982, alive in 1960 or hired after 1960, worked at least 1 d, follow-up 1960–2001 ( <i>n</i> = 1,672); 767 deaths (95% with known cause of death) Mean duration: 4.0 yr (808, ~50% worked less than 1 yr) Median fibers/cc-yr: 8.7 Underlying cause of death data from death certificates or National Death Index-Plus.	15-yr exposure lag: All cancer ( <i>n</i> = 202) SMR: 1.4 (1.2, 1.6) Lung ( <i>n</i> = 89) SMR: 1.7 (1.4, 2.1)	15-yr exposure lag:			
			Cumulative exposure	<i>n</i>	SMR (95% CI) <sup>b</sup>	SRR (95% CI) <sup>c</sup>
			0.0–4.49 fibers/cc-yr	19	1.5 (0.9, 2.3)	1.0 (referent)
			4.5–22.9 fibers/cc-yr	24	1.6 (1.1, 2.5)	1.1 (0.6, 2.0)
			23.0–99.0 fibers/cc-yr	23	1.8 (1.1, 2.7)	1.4 (0.7, 2.7)
			≥100 fibers/cc-yr	23	1.9 (1.2, 2.9)	1.5 (0.8, 2.8)
			Linear trend test			( <i>p</i> < 0.001)
			Duration	<i>n</i>	SMR (95% CI) <sup>b</sup>	SRR (95% CI) <sup>c</sup>
			<1 yr	41	1.6 (1.1, 2.1)	1.0 (referent)
			1–9.9 yr	34	1.7 (1.1, 2.3)	1.1 (0.7, 1.8)
			≥10 yr	14	2.5 (1.4, 4.3)	1.8 (0.9, 3.4)

*This document is a draft for review purposes only and does not constitute Agency policy.*

**Table 4-16. Respiratory (lung) cancer mortality and exposure-response analyses based on related studies of the vermiculite mining and milling workers in Libby, MT<sup>a</sup> (continued)**

Reference(s)	Inclusion criteria and design details	Standardized mortality ratio (95% CI)	Exposure-response analyses—lung cancer			
<a href="#">Larson et al. (2010b)</a>	Inclusion criteria not described ( $n = 1,862$ ); follow-up through 2006; 952 deaths (80% with known cause of death). Median duration: 0.8 yr Median fibers/cc-yr = 4.3 Immediate and underlying cause of death data (i.e., multiple cause of death) from death certificates or National Death Index-Plus.	Lung ( $n = 104$ ) SMR: 1.6 (1.3, 2.0)	<i>20-yr exposure lag:</i>			
			Cumulative exposure	$n$	SMR (95% CI) <sup>b</sup>	RR (95% CI) <sup>c</sup>
			0.0–<1.4 fibers/cc-yr	19	(not reported)	1.0 (referent)
			1.4 to <8.6 fibers/cc-yr	20	(not reported)	1.1 (0.6, 2.1)
			8.6 to <44.0 fibers/cc-yr	21	(not reported)	1.7 (1.0, 3.0)
			≥44.0 fibers/cc-yr	38	(not reported)	3.2 (1.8, 5.3)
			Per 100 fibers/cc-yr increase			1.11 (1.05, 1.18)
						( $p = 0.006$ )

<sup>a</sup>Includes miners, millers, and processors; workers in the screening plant, loading docks, and expansion plants; and office workers.

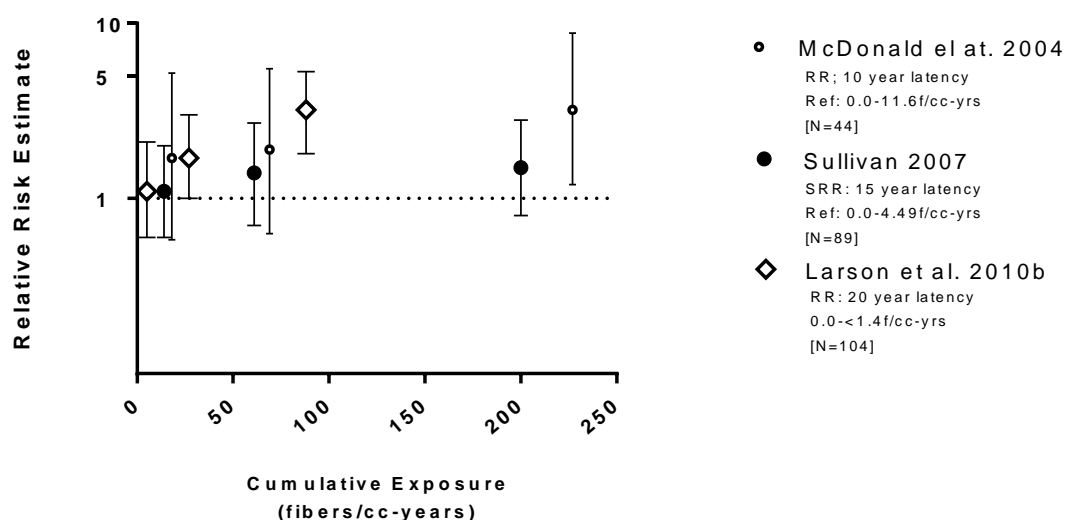
<sup>b</sup>SMR based on external referent group.

<sup>c</sup>In [Sullivan \(2007\)](#), the SRR is a ratio of sums of weighted rates in which the weight for each stratum-specific rate is the combined person-yr for the observed cohort across all duration (or cumulative level of exposure) categories. The Life Table Analysis System provides the SRR for each duration (or cumulative level of exposure) group compared to the referent group. The cutoff points for the categories are specified by the user. Taylor-series-based confidence intervals are given for each specific SRR.

<sup>d</sup>In [McDonald et al. \(2004\)](#), the RR is based on Poisson analysis using an internal referent group.

<sup>e</sup>In [Larson et al. \(2010b\)](#), the RR is based on Cox proportional hazards modeling using an internal referent group.

SMR = standardized mortality ratio; CI = confidence interval; SRR = standardized rate ratio; RR = relative risk.



**Figure 4-3. Lung cancer mortality risk among workers in the Libby, MT vermiculite mine and mill workers.** Data from the three studies with updated follow-up (through 1998, 2001, and 2006, respectively, in [McDonald et al. \(2004\)](#), [Sullivan \(2007\)](#), and [Larson et al. \(2010b\)](#)). Size of symbols is proportional to number of observed cases. Midpoint of the highest exposure category in each group is estimated as twice the value of the lower cut-point.

Two of these studies included data addressing the question of the extent to which the results could be confounded by smoking ([Larson et al., 2010b](#); [Amandus and Wheeler, 1987](#)). [Amandus and Wheeler \(1987\)](#) provide some information on the smoking history of a sample of 161 male workers employed during 1975–1982 with at least 5 years of employment in the Libby cohort study and comparison data based on surveys conducted in the United States from 1955–1978. Among the workers, 35% were current smokers and 49% were former smokers. This smoking information was obtained from questionnaires the company administered to workers after 1975. Assuming the definitions are similar to those of the national surveys, the prevalence of current smokers is similar in the worker cohort compared to the U.S. white male population data (ranging from 37.5–41.9% current smokers between 1975 and 1978). The only year in this range with data on former smokers in the national survey is 1975, and at that time, the prevalence of former smokers in the population data was 29.2%, about 20% lower than among the workers. Using an estimated RR of lung cancer of 14 among smokers, [Amandus and Wheeler \(1987\)](#) estimated that the difference in smoking rates between workers and the comparison population could have resulted in a 23% increase in the observed risk ratio and commented that the increased risk observed in the lower dose range (<50 fiber-year) could be the result of confounding by smoking status.

Smoking patterns in the U.S. population changed considerably over the period corresponding to the data reported by [Amandus and Wheeler \(1987\)](#). In the National Health

Interview Surveys conducted between 1974 and 1983, the prevalence of smoking in males age 20 and older decreased from 42.1 to 35.5% ([HHS, 1990](#)). In addition, the prevalence of former smokers can depend on the definition used. Based on 1986 survey data, the percentage of adults age 17 and older classified as former smokers varied between 14.7 and 25.8% using different definitions for time since last smoked (e.g., [from quitting 5 or more years ago to quitting within the past 3 months](#); [HHS, 1990](#)). Thus, given the lack of information pertaining to the period in which smoking information was collected and the specifics of the sources that were used, EPA concludes there is considerable uncertainty regarding the evidence for differences in smoking rates between the workers and the external comparison population.

[Larson et al. \(2010b\)](#) used data from the ATSDR community health screening in Libby (described in Section 4.1.1.3) pertaining to smoking history to estimate that the proportion of smokers ranged from 50 to 66% in the unexposed group (defined as exposure <8.6 fibers/cc-yr) and between 66 and 85% among the exposed (defined as  $\geq 8.6$  fibers/cc-yr). [Larson et al. \(2010b\)](#) used these estimates in a Monte Carlo simulation to estimate the potential bias in lung cancer risks that could have been introduced by differences in smoking patterns. The bias adjustment factor ( $RR_{unadjusted}/RR_{adjusted} = 1.3$ ) reduced the overall RR estimate for lung cancer from 2.4 to 2.0.

#### *O.M. Scott, Marysville, OH plant workers*

There was no evidence of an increased risk of lung cancer in the analysis of mortality among the 465 Marysville, OH plant workers ([Dunning et al.](#)). The SMR was 0.9 (95% CI: 0.5–1.5), based on 16 observed lung cancer deaths, and there was no indication of an increased risk in analyses stratifying by tertiles of cumulative exposure (SMRs varying between 0.8 and 1.0, and standardized rate ratios (SRRs) varying between 0.9 and 1.0).

#### *Geographic mortality analysis*

In the geographic mortality analysis (1979–1998) conducted by [ATSDR \(2000\)](#), the SMR for lung cancer ranged from 0.9–1.1 and 0.8–1.0 for each of the six geographic boundaries using Montana and U.S. reference rates, respectively. These analyses did not distinguish between deaths among workers and deaths among other community members.

#### **4.1.4.2. Mesothelioma**

Prior to the 10<sup>th</sup> revision of the ICD, which was implemented in the United States in 1999, there was no unique ICD code for mesothelioma. The updated NIOSH study by [Sullivan \(2007\)](#) identified 15 deaths for which mesothelioma was mentioned on the death certificate. Only two of these deaths occurred between 1999 and 2001, which were coded using the ICD-10 mesothelioma coding (C45). [Larson et al. \(2010b\)](#) classified all death certificates listing mesothelioma as ICD-10 code C45. The updated McGill study ([McDonald et al., 2004](#); with

1 [analysis through 1998](#)) noted that the classification of mesothelioma was based on a nosologist's  
2 review of death certificates; only 5 of the 12 cases classified as mesothelioma had a cause of  
3 death listed as pleural cancer (ICD-9 code 163).

4 Data pertaining to mesothelioma risk from the available occupational studies are  
5 summarized in Table 4-17. [McDonald et al. \(2004\)](#) presented dose-response modeling using  
6 Poisson regression of mesothelioma risk based on 12 cases. Note that the referent group was  
7 also at excess risk of dying from mesothelioma; that is, one to three cases of mesothelioma were  
8 observed in the referent group, depending on the exposure index. Three exposure indices were  
9 used in the analysis: average intensity over the first 5 years of employment, cumulative  
10 exposure, and residence-weighted cumulative exposure. Because of the requirement for 5 years  
11 of employment data, 199 individuals (including three mesothelioma cases) were excluded from  
12 the analysis of average intensity. The residence-weighted cumulative exposure was based on the  
13 summation of exposure by year, weighted by years since the exposure. This metric gives greater  
14 weight to exposures that occurred a longer time ago. Although evidence of an excess risk of  
15 dying from mesothelioma was seen in all groups, only the residence-weighted cumulative  
16 exposure metric exhibited a monotonically increasing pattern, with an RR of 1.57 among those  
17 with 500.1–1,826.8 fibers/cc-yr exposure, and an RR of 1.95 among workers with higher  
18 residence-weighted cumulative exposure. In the study by [Sullivan \(2007\)](#), which identified  
19 15 deaths from mesothelioma through a manual review of death certificates, the SMR for  
20 mesothelioma was 14.1 (95% CI: 1.8, 54.4), based on the two mesothelioma deaths occurring  
21 between 1999 and 2001, the period for which comparison data using the ICD-10 classification  
22 criteria were available.

**Table 4-17. Mesothelioma mortality risk based on studies of the vermiculite mine workers in Libby, MT<sup>a</sup>**

Reference(s)	Inclusion criteria and design details	Results				
<a href="#">Amandus and Wheeler (1987)</a>	Men, hired before 1970, worked at least 1 yr, follow-up through 1982 ( <i>n</i> = 575); 161 deaths (159 with death certificates). Mean duration: 8.3 yr (0 worked less than 1 yr) Mean fiber-yr: 200.3. Twelve female workers not included in this analysis.	Two mesothelioma deaths observed (hired in 1946, 33-yr latency, exposure >300 fibers/cc-yr); 1.2% of all deaths.				
McDonald et al. (2004; 1986a)	Men, hired before 1963, worked at least 1 yr ( <i>n</i> = 406), follow-up through 1999 ( <a href="#">McDonald et al., 2004</a> ); 165 deaths before July 1983 (163 with death certificates); 120 deaths from July 1983–1998 coded by nosologists using ICD-8 classifications; cause of death for deaths from 1983–1998 obtained from National Death Index. Mean duration: 8.7 yr (0 worked less than 1 yr). Mean fiber-yr: 144.6.	12 mesothelioma deaths observed; 4.2% of all deaths				
		<i>Excluding first 10 yr of follow-up:</i>				
		Cumulative exposure	<i>n</i>	RR (95% CI) <sup>b</sup>		
		0.0–11.6 fibers/cc-yr	1	1.0 (referent)		
		11.7–25.1 fibers/cc-yr	4	3.7 (0.41, 33.5)		
		25.2–113.7 fibers/cc-yr	3	3.4 (0.35, 33.2)		
		≥113.8 fibers/cc-yr	4	3.7 (0.41, 33.2)		
		per 100 fibers/cc-yr increase_		0.10 (<0, 1.81)		
				<i>(p</i> > 0.20)		
		Intensity category				
		0.0–11.6 fibers/cc-yr	<i>n</i>	RR (95% CI) <sup>b</sup>		
		11.7–25.1 fibers/cc-yr	1	1.0 (referent)		
		25.2–113.7 fibers/cc-yr	4	3.4 (0.37, 30.9)		
		≥113.8 fibers/cc-yr	2	2.3 (0.21, 26.1)		
		per 100 fibers/cc-yr increase		2	2.1 (0.19, 23.9)	
					0.02 (<0, 1.08)	
					<i>(p</i> > 0.20)	
		Residence-weighted				
		0.0–25.1 fibers/cc-yr	<i>n</i>	RR (95% CI) <sup>b</sup>		
		25.2–113.7 fibers/cc-yr	3	1.0 (referent)		
		≥113.8 fibers/cc-yr	4	1.57 (0.35, 7.07)		
		per 100 fibers/cc-yr increase		5	1.95 (0.41, 8.51)	
					0.03 (<0, 6.4)	



**Table 4-17. Mesothelioma mortality risk based on studies of the vermiculite mine workers in Libby, MT<sup>a</sup> (continued)**

Reference(s)	Inclusion criteria and design details	Results		
<a href="#">Sullivan (2007)</a>	White men, enumerated in 1982, alive in 1960 or hired after 1960, worked at least 1 d, follow-up 1960–2001 ( <i>n</i> = 1,672); 767 deaths (95% with known cause of death) Mean duration: 4.0 yr (808, ~50% worked less than 1 yr) Median fibers/cc-yr: 8.7 Underlying cause of death data from death certificates or National Death Index-Plus. SMR analysis limited to 1999–2001 because this is the period for which comparison data from ICD-10 are available.	15 mesothelioma deaths observed; 2% of all deaths <i>N</i> = 2 for 1999–2001: SMR: 15.1 (95% CI: 1.8, 54.4) Pleural ( <i>n</i> = 4) SMR: 23.3 (95% CI: 6.3, 59.5)		
<a href="#">Larson et al. (2010b)</a>	Inclusion criteria not described ( <i>n</i> = 1,862); follow-up through 2006; 952 deaths (80% with known cause of death). Median duration: 0.8 yr Median fibers/cc-yr = 4.3 Immediate and underlying cause of death data (i.e., multiple causes of death) from death certificates or National Death Index-Plus.	19 mesothelioma deaths observed		
		<i>20-yr exposure lag:</i>		
		Cumulative exposure	<i>n</i>	RR (95% CI) <sup>c</sup>
		<1.4 fibers/cc-yr	1	1.0 (referent)
		1.4 to <8.6 fibers/cc-yr	2	1.9 (0.31, 13.6)
		8.6 to <440 fibers/cc-yr	5	4.5 (0.8, 24.6)
		≥44.0 fibers/cc-yr	11	17.1 (3.7, 78.1)
		per 100 fibers/cc-yr increase		1.15 (1.03, 1.28) ( <i>p</i> = 0.0134)

<sup>a</sup>Includes miners, millers, and processors; workers in the screening plant, loading docks, and expansion plants; and office workers.

<sup>b</sup>In [McDonald et al. \(2004\)](#), the RR is based on Poisson analysis using an internal referent group.

<sup>c</sup>In [Larson et al. \(2010b\)](#), the RR is based on Cox proportional hazards modeling using an internal referent group.

A more descriptive presentation of a collection of mesothelioma cases was reported by [Whitehouse et al. \(2008\)](#). This report reviewed 11 cases of mesothelioma diagnosed between 1993 and 2006 in residents in or around Libby, MT (*n* = 9) and in family members of workers in the mining operations (*n* = 2). Three cases were men who might have had occupational asbestos exposure through construction work (Case 1), working in the U.S. Coast Guard and as a carpenter (Case 5), or through railroad work involving sealing railcars in Libby (Case 7). One case was a woman whose father had worked at the mine for 2 years; although the family lived 100 miles east of Libby, her exposure may have come through her work doing the family laundry, which included laundering her father's work clothes. The other seven cases (four women, three men) had lived or worked in Libby for 6–54 years and had no known occupational or family-related exposure to asbestos. Medical records were obtained for all 11 patients;

pathology reports were obtained for 10 of the 11 patients. The Centers for Disease Control and Prevention estimated the death rate from mesothelioma, using 1999 to 2005 data, as approximately 14 per million per year ([CDC, 2009](#)), approximately three to four times lower than the Libby-area rate estimated by EPA using the observation of seven cases without known personal or familial occupational exposure, over a 15-year observation period and an estimated Libby area (Lincoln County) population of 9,500 (142,000 person-year). [Whitehouse et al. \(2008\)](#) stated that a W.R. Grace unpublished report of measures taken in 1975 indicated that exposure levels of 1.1 fibers/cc were found in Libby, and 1.5 fibers/cc were found near the mill and railroad facilities. Because the mining and milling operations continued to 1990, and because of the expected latency period for mesothelioma, [Whitehouse et al. \(2008\)](#) suggested that additional cases can be expected to occur within this population, as well as in transitory workers and in workers who had left the area.

**4.1.4.2.1. O.M. Scott, Marysville, OH plant workers.** In the analysis of mortality among the Marysville, OH plant workers, 2 of the 465 workers died of mesothelioma compared to an expected 0.2 cases ([SMR 10.5, 95% CI 1.3, 38; Dunning et al., 2012](#)). The cumulative exposure for each of these two cases was approximately 45 fibers/cc-yr. One other incident mesothelioma case was identified in the cohort. This case was alive at the time of the study, and so is not included in the mortality analysis, but the cumulative exposure of the individual was 5.73 fibers/cc-yr.

#### **4.1.4.3. Other Cancers**

[Larson et al. \(2010b\)](#) presented data on cancers other than respiratory tract and mesothelioma. The category of malignant neoplasms of digestive organs and peritoneum included 39 observed deaths, for an SMR of 0.8 (95% CI: 0.6, 1.1). No risk in relation to asbestos exposure was seen with a 20-year lag.

#### **4.1.4.4. Summary of Cancer Mortality Risk in Populations Exposed to Libby Amphibole Asbestos**

The studies conducted in the 1980s ([Amandus and Wheeler, 1987; McDonald et al., 1986a](#)) as well as the extended follow-up studies published in more recent years ([Larson et al., 2010b; Sullivan, 2007; McDonald et al., 2004](#)) provide consistent evidence of an increased risk of lung cancer mortality and of mesothelioma mortality among the workers in the Libby vermiculite mining and processing operations. The lung cancer analyses using an internal referent group in the larger follow-up studies ([Larson et al., 2010b; Sullivan, 2007; McDonald et al., 2004](#)) observed increasing risks with increasing cumulative exposure when analyzed using quartiles or as a continuous measure. Increased risks are also seen in the studies reporting analyses using an external referent group (i.e., standardized mortality ratios; [Sullivan, 2007](#);

1 [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). Although an increased lung cancer risk  
2 was not observed in the Marysville, OH plant workers, three cases of mesothelioma (two of  
3 which resulted in death) have been identified as of June, 2011. These observations further  
4 support the identification of cancer (specifically, lung cancer and mesothelioma) as a hazard of  
5 LAA.

#### 6 7 **4.1.5. Comparison With Other Asbestos Studies—Environmental Exposure Settings**

8 The literature pertaining to risks of asbestos is extensive; of particular interest is the set of  
9 studies examining environmental exposures to constituents of LAA (e.g., tremolite) or other  
10 amphiboles. This literature provides findings consistent with those identified for LAA.

11 Several communities have been exposed in environmental or residential settings to  
12 tremolite or tremolite-chrysotile mixtures from natural soils and outcroppings as well as  
13 construction materials found in the home (see Table 4-18). Studies on these affected populations  
14 (published as early as 1979) reported an increased risk of pleural and peritoneal malignant  
15 mesothelioma ([Sichletidis et al., 1992b](#); [Baris et al., 1987](#); [Langer et al., 1987](#); [Baris et al., 1979](#)).  
16 Clinical observations include a bilateral increase in pleural calcification accompanied by  
17 restrictive lung function decrements as the disease progresses, a condition known as “Metsovo  
18 lung,” named after a town in Greece ([Constantopoulos et al., 1985](#)). These health effects are  
19 consistent with the health effects documented for workers exposed to commercial forms of  
20 asbestos.

**Table 4-18. Exposure levels and health effects observed in communities exposed to tremolite, chrysotile, and crocidolite asbestos**

Area, population	Fiber type, exposure levels, and fiber size	Effects observed			References
Tremolite and tremolite-chrysotile mixtures: whitewash material used in homes					
Turkey—Anatolia (Eshisehir district) ~2,000	Fiber: tremolite, tremolite/chrysotile mixtures Exposure: indoor 0.089 f/mL; outdoor 0.013 f/mL Size: not available	Mesothelioma			Metintas et al. (2005; 2002) Yazicioglu et al (1980; 1976) Baris et al. (1979)
		Men	SIR	53	
		Women	SIR	144	
		Pleural plaques prevalence ~14% Diffuse pleural thickening prevalence ~10%			
Greece—Metsovo ~5,000	Fiber: tremolite Exposure: Variable (1 to >200 f/mL) Size: length ≤10 μm, diameter 0.2 μm	Mesothelioma SIR ~280  Pleural plaques prevalence ~45%			Constantopoulos et al. (1987; 1985) Bazas et al. (1985)
Greece—Almopa ~4,000	Fiber: tremolite, chrysotile Exposure: indoors 0.01–17.9 f/cc Size: not available	Mesothelioma four incident cases among 198 people with pleural plaques over 15-yr follow-up period. Pleural plaques ~24% among people over age 40 yr; increases with age; extent of plaques (surface area) increased between 1988–2003.			Sichletidis et al. (2006; 1992b; 1992a)
New Caledonia ~40,000	Fiber: tremolite Exposure: not available Size: not available	Mesothelioma SMR 41 Lung cancer			Luce et al. (2000)
		Men	SMR	~1.0	
		Women	SMR	2.4	

**Table 4-18. Exposure levels and health effects observed in communities exposed to tremolite, chrysotile, and crocidolite asbestos (continued)**

Area, population	Fiber type, exposure levels, and fiber size	Effects observed			References
Crocidolite: communities surrounding amphibole asbestos mines or mills					
Australia (Wittenoom) ~4,700 nonworker residents	Fiber: Crocidolite Cumulative exposure: 76% ≤7 f/mL-yr, 5.5% >20 f/mL-yr Size: length >5 μm	2,500 women follow-up through 2004:			Reid et al. ( <a href="#">2013</a> ; <a href="#">2007</a> ) Hansen et al. ( <a href="#">1998</a> ; <a href="#">1997</a> ; <a href="#">1993</a> )
			(n)	SMR	
		Mesothelioma	(30)	Not reported	
		Lung cancer	(30)	~1.9	
		Ovarian cancer	(9)	~1.4	
		Pneumoconiosis	(2)	~11	
		2,460 people initially exposed age <15 yr:			
		Cancer type	(n) SIR		
			Women	Men	
		Mesothelioma	(13) ~ 90	(29) ~ 60	
		Lung	(5) ~ 2.0	(3) ~ 1.0	
		Brain	(4) ~ 3.6	(5) ~ 3.4	
		Leukemia	(4) ~ 3.0	(7) ~ 4.2	
		Ovary	(6) ~ 3.3	--	
		(SMRs and SIRs based on means of two different censoring methods)			

SIR = standardized incidence ratio; SMR = standardized mortality ratio.

Although it is not a constituent of LAA, crocidolite is another type of amphibole asbestos that has been studied with respect to health effects arising from environmental exposures. Several studies have examined cancer risk and pneumoconiosis risk among nonworker residents of Wittenoom, Australia, an area surrounding a crocidolite asbestos mine and mill (Reid et al., 2007; Hansen et al., 1998; see Table 4-18). Increased risk of mesothelioma and pneumoconiosis and more modestly increased risk of lung cancer were reported in these studies.

#### 4.2. SUBCHRONIC- AND CHONIC-DURATION STUDIES AND CANCER BIOASSAYS IN ANIMALS—ORAL, INHALATION, AND OTHER ROUTES OF EXPOSURE

Laboratory animal studies of exposure to Libby Amphibole or tremolite asbestos show effects similar to those observed in occupationally exposed human populations, including pleural pathology, mesothelioma, and lung cancer. Tremolite is an amphibole asbestos fiber that is a component of LAA (~6%). Also, in early studies, LAA was defined as tremolite. Therefore, laboratory animal studies examining the effect of tremolite exposure have been reviewed and are

1 summarized below to potentially increase understanding of the effects and mechanisms of LAA.  
2 Detailed study summaries can be found in Appendix D and summarized in Tables 4-19 and 4-20.  
3 As noted in Section 3, the primary route of human exposure is inhalation. Thus, studies that  
4 expose animals through a pulmonary route are the most relevant for hazard identification. No  
5 inhalation studies have been performed for LAA, but chronic intrapleural injection studies in  
6 hamsters demonstrate carcinogenicity following exposure. The chronic inhalation and  
7 intrapleural injection laboratory animal studies with tremolite asbestos demonstrated pleural  
8 pathology and carcinogenicity in rats. These studies support the epidemiology studies of LAA  
9 exposure (see Section 4.1) and aid in informing the mechanisms of LAA-induced disease.

**Table 4-19. In vivo data following exposure to Libby Amphibole asbestos**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
LVG:LAK Hamsters (M) (n ~ 60/group)	Intraperitoneal injection (once) 25 mg/0.5 mL 0.9% NaCl solution	Tremolite (Sample 60) and tremolite + vermiculite (Sample 63)	n/a	n/a	Pleural adhesions (fibrosis): examined 10 animals/group at ~3 mo post exposure: Sample 60: 10/10; Sample 63: 10/10; Control: 0/10  Mesothelioma: Sample 60: 5/66; Sample 63: 5/64; Control: 0/60	<a href="#">Smith (1978)</a> (W.R. Grace study)
C57Bl/6 mice (M, F) (n = 7/group)	Intratracheal instillation (once) 1 wk, 1 mo, 3 mo 100 µg of sample in 30 µL saline	LAA (Six Mix) and crocidolite	LAA: 7.21 ± 7.01 µm Crocidolite: 4.59 ± 4.22 µm	LAA: 0.61 ± 1.22 µm Crocidolite: 0.16 ± 0.09 µm	Altered gene expression in mice exposed to both samples; increase in collagen in exposed animals	<a href="#">Putnam et al. (2008)</a>
C57Bl/6 mice (M, F) (n = 7/group)	Intratracheal instillation (once) 1 wk, 1 mo, 3 mo 100 µg of sample in 30 µL saline	LAA (Six Mix) and crocidolite	LAA: 7.21 ± 7.01 µm Crocidolite: 4.59 ± 4.22 µm	LAA: 0.61 ± 1.22 µm Crocidolite: 0.16 ± 0.09 µm	Collagen gene expression and protein levels increased following exposure to both forms of asbestos (~1 mo post exposure).	<a href="#">Smartt et al. (2010)</a>
Wistar-Kyoto rats (M) (n = 12/group) SH (n = 6/group) SHHF (n = 6/group)	Intratracheal instillation (once) 1 d, 1 wk, 1 mo 0.25 or 1.0 mg/rat	LAA (Six Mix)	5.0 ± 4.5 µm	0.29 ± 0.19 µm	Strain-related differences observed in biomarkers of inflammation following exposure to LAA.  No differences were observed in histopathology.	<a href="#">Shannahan et al. (2011a)</a>

**Table 4-19. In vivo data following exposure to Libby Amphibole asbestos (continued)**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
SH (M) (n = 8/group)	Intratracheal instillation (once) 4 h, 1 d 1.0 mg DEF; 21 µg FeCl <sub>3</sub> ; 0.5 mg LAA, 0.5 mg FeLAA; 0.5 mg LAA + 1 mg DEF in 300 µL saline	LAA (Six Mix)	5.0 ± 4.5 µm	0.29 ± 0.19 µm	Statistically significant increases in neutrophils was observed in BALF in animals exposed to LAA, FeLAA, and LAA + DEF with the greatest increase observed in the LAA + DEF animals.	<a href="#">Shannahan et al. (2011b)</a>
Fischer 344 rats (M) (n = 8/group)	Intratracheal instillation (once) 1 d, 3 d, 7 d, 2 wk, 3 mo 0.65 or 6.5 mg/rat LAA; 0.65 mg amosite in 250 µL saline	LAA (Six Mix)  Amosite	5.0 ± 4.5 µm	0.29 ± 0.19 µm	Statistically significant increases in inflammatory markers were observed following exposure to LAA and amosite, including increased neutrophils and inflammatory gene expression, with the greatest increase in amosite-exposed rats.	<a href="#">Padilla-Carlin et al. (2011)</a>



**Table 4-19. In vivo data following exposure to Libby Amphibole asbestos (continued)**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
Four separate study designs: (A) WKY rats (M) (n = 12/group) SH (M) (n = 6/group) SHHF (M) (n = 6/group) (B) F344 rats (M) (n = 8–12/group) (C) F344 rats (M) (n = 8/group) (D) WKY rats (M) (n = 5/group)	(A) Intratracheal instillation (once) 1 d, 1 wk, 1 mo, 3 mo 0.25 or 1.0 mg/rat (B) Intratracheal instillation (once) 3 mo, 1 yr 1.0 or 5.0 mg/rat (C) Intratracheal instillation (once every other wk for 13 wk) 1 d, 2 wk Cumulative dose of 1.0 or 5.0 mg/rat (D) Intratracheal instillation (once every wk for 4 wk) 1 d, 1 mo 0.25 or 0.5 mg/rat LA or 0.5 mg/rat of diesel exhaust particles	LAA (Six Mix)	5.0 ± 4.5 µm	0.29 ± 0.19 µm	Analysis of biomarker expression following exposure to LAA in healthy (WKY, F344) or susceptible (SH, SHHF) rats demonstrated increases in acute phase proteins associated with inflammatory response; biomarkers associated with cancer (e.g., mesothelin) were increased only at 1 d postexposure. Biomarker expression in all four studies occurred rapidly and returned to homeostatic levels after 1 d postinstillation.	<a href="#">Shannahan et al. (2012a)</a>

**Table 4-19. In vivo data following exposure to Libby Amphibole asbestos (continued)**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
Three separate study designs: (A) WKY rats (M) (n = 12/group) SH (M) (n = 6/group) SHHF (M) (n = 6/group)  (B) F344 rats (M) (n = 8–24/group)  (C) F344 rats (M) (n = 24/group)	(A) Intratracheal instillation (once) 1 d, 1 wk, 1 mo, 3 mo 0.25 or 1.0 mg/rat  (B) Intratracheal instillation (once) 1 d, 1 wk, 1 mo, 3 mo, 1 yr, 2 yr 0.15, 0.5, 1.5 or 5.0 mg/rat  (C) Intratracheal instillation (once every other wk for 13 wk) 1 d, 2 wk, 2 yr Cumulative dose of 0.15, 0.5, 1.5 or 5.0 mg/rat	LAA (Six Mix)	5.0 ± 4.5 µm	0.29 ± 0.19 µm	LAA exposure in healthy rats (WKY, F344) increased expression of biomarkers of oxidative stress, thrombosis and vasoconstriction in the aorta. These levels were similar to CVD-sensitive rats at baseline.	<a href="#">Shannahan et al. (2012d)</a>
SH (M) (n = 8/group)	Intratracheal instillation (once) 4 h, 1 d 1.0 mg DEF; 21 µg FeCl <sub>3</sub> ; 0.5 mg LA, 0.5 mg FeLA; 0.5 mg LA + 1 mg DEF in 300 µL saline	LAA (Six Mix) LAA + Fe (iron-loaded LA)	5.0 ± 4.5 µm	0.29 ± 0.19 µm	LAA exposure increased expression of inflammasome-related molecules, inflammatory cytokines and upstream regulators of the inflammasome. These changes were not impacted by iron levels.	<a href="#">Shannahan et al. (2012b)</a>
WKY rats (M) (n = 12/group) SH (M) (n = 6/group) SHHF (M) (n = 6–36/group)	Intratracheal instillation (once) 1 wk, 1 mo, 3 mo 0.25 or 1.0 mg/rat	LAA (Six Mix)	5.0 ± 4.5 µm	0.29 ± 0.19 µm	Gene expression analysis demonstrated that LAA exposure upregulated inflammatory-related genes in healthy rats (WKY) but downregulated inflammatory-related genes in CVD-susceptible rats (SH, SHHF).	<a href="#">Shannahan et al. (2012c)</a>

**Table 4-19. In vivo data following exposure to Libby Amphibole asbestos (continued)**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
Fischer 344 rats (M) ( <i>n</i> = 8/group)	Intratracheal instillation (once) 1 d, 3 d, 7 d, 2 wk, 3 mo 0.65 or 6.5 mg/rat LA; 0.65 mg amosite in 250 µL saline	LAA (Six Mix)	1.9 ± 3.0 µm	0.29 ± 0.23 µm	LAA exposure induced significant fibrogenic (but not carcinogenic) effects up to 2 yr postexposure. This response differed from that of amosite exposure in the same study, with LAA being less potent than amosite on a mass basis.	<a href="#">Cyphert et al. (2012a)</a>
Fischer 344 rats (M) ( <i>n</i> = 8/group)	Intratracheal instillation (once) 1 d, 3 mo 0.5 or 1.5 mg/rat LA, SM, ED, ON; 250 µL saline	LAA (Six Mix)  Sumas Mountain chrysotile (SM)  El Dorado tremolite (ED)  Ontario ferroactinolite (ON)	1.9 ± 3.0 µm  2.0 ± 2.4 µm  0.9 ± 0.9 µm  1.1 ± 0.9 µm	0.39 ± 0.3 µm  0.31 ± 0.4 µm  0.42 ± 0.4 µm  0.40 ± 0.3 µm	Inflammatory markers were increased in BALF at 1 d postexposure, but returned to control levels by 3 mo; development of fibrosis persisted at 3 mo and was greatest in SM-exposure rats (SM > LA > ON > ED). This correlated with fiber length and AR of the different fiber types.	<a href="#">Cyphert et al. (2012b)</a>
Lewis rats (F) ( <i>n</i> = 8/group)	Intratracheal instillation (biweekly for 13 wk) 19 wk 0.15, 0.5, 1.5, or 5 mg/rat LA; 0.5 or 1.5 mg amosite in 250 µL saline	LAA (Six Mix)  Amosite	5.0 ± 4.5 µm	0.29 ± 0.19 µm	Results failed to show a positive correlation between LA exposure and rheumatoid arthritis in two animal models. Upregulated ANA following exposure suggest an altered immunological profile similar to other systemic autoimmune diseases.	<a href="#">Salazar et al. (2012)</a>

**Table 4-19. In vivo data following exposure to Libby Amphibole asbestos (continued)**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
Lewis rats (F) (n = 8/group)	Intratracheal instillation (biweekly for 13 wk) 0.15, 0.5, 1.5, or 5 mg/rat LA; 0.5 or 1.5 mg amosite in 250 µL saline	LAA (Six Mix)	5.0 ± 4.5 µm	0.29 ± 0.19 µm	ANA in serum increased at all doses of LA except 1.5 mg by Week 28 postexposure. No dose-response related histopathological effects were observed in the kidney.	<a href="#">Salazar et al. (2013)</a>

BALF = bronchoalveolar lavage fluids; DEF = deferoxamine; SH = spontaneously hypertensive; SHHF = spontaneously hypertensive-heart failure; WKY = Wistar-Kyoto rat; FeLAA = LA loaded with Fe; AR = aspect ratio.

<sup>a</sup>When available, results are shown as number of animals with tumors/total number of animals examined.

**Table 4-20. In vivo data following exposure to tremolite asbestos**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
F344 rats (M, F) ( <i>n</i> = 100 to 250/group)	Oral 1% bw in feed pellets; lifetime exposure starting in dam	Tremolite-nonfibrous (Gouverneur Talc Co., Gouverneur, NY)	n/a	n/a	Offspring from exposed mothers were smaller at weaning and throughout life; No toxicity or increase in neoplasia in tremolite rats as compared to controls.	<a href="#">McConnell et al. (1983a)</a>
Wistar rats (M) ( <i>n</i> = 48)	Inhalation 10 mg/m <sup>3</sup> (7 h each d, 5 d per wk, total of 224 d)	South Korean tremolite and brucite	>5 µm	<3 µm	Increased fibrosis (19/39) and carcinogenesis (18/39).	<a href="#">Davis et al. (1985)</a>
AF/Han rats ( <i>n</i> = 33–36/group)	Intraperitoneal injection 10 mg/2 mL PBS; single exposure	Tremolite (six samples)	n/a	n/a	All six fibers could induce mesothelioma: California: 36/36 <sup>b</sup> Swansea: 35/36 <sup>b</sup> Korea: 32/36 <sup>b</sup> Italy: 24/36 Carr Brae: 4/33 Shininess: 2/36	<a href="#">Davis et al. (1991)</a>
Hamsters ( <i>n</i> ≤ 35/group)	Intrapleural injection 10 or 25 mg	Four types of tremolite (Sample FD-14; 275; 31; 72)	FD-14: 5.7 µm 275: N/A 31: >20 µm 72: >20 µm	FD-14: 1.6 µm 275: N/A 31: <0.4 µm 72: <0.4 µm	Tumors/survivors at 350 d Sample FD-14: 0/35 Sample 275: 0/34 (10 mg); 0/31 (25 mg) Samples 31: 3/41 (10 mg); 12/28 (25 mg) Sample 72: 4/13 (10 mg); 13/20 (25 mg)	<a href="#">Smith et al. (1979)</a>
Rats ( <i>n</i> = 32 Wistar rats—Sample A and <i>n</i> = 48 Sprague-Dawley rats—Samples B and C)	Intrapleural injection 20 mg/rat	Tremolite (three samples)	California: <6 µm Greenland: <3 µm Korea: >8 µm	California: <0.8 µm Greenland: <1.2 µm Korea: <1.5 µm	No tumors following exposure to Samples A and B; Sample C: 14/47	<a href="#">Wagner et al. (1982)</a>
Osborne-Mendel rats ( <i>n</i> = 28/group)	Hardened gelatin technique 40 mg	Tremolite (two samples)	N/A	N/A	Sample 1: 21/28 pleural sarcomas Sample 2: 22/28 pleural sarcomas	<a href="#">Stanton et al. (1981)</a>

**Table 4-20. In vivo data following exposure to tremolite asbestos (continued)**

Species (gender)	Exposure route	Fiber type	Mean fiber length	Mean fiber diameter	Effects <sup>a</sup>	Reference
Wistar rats (F) ( <i>n</i> = 40/group)	Intraperitoneal injection 1 × 3.3 and 1 × 15 mg, lifetime observation	Tremolite	N/A  22% of fibers >5 µm	N/A	Limited details in text. Increase in mesothelioma following exposure to tremolite: 3.3 mg sample: 9/29; 15 mg sample: 30/37	Roller et al. (1997, 1996)
Wistar rats (M) ( <i>n</i> = 56)	Inhalation (flow-past nose only) 100 fibers/cm <sup>3</sup> longer than 20 µm, 5 d, follow-up 1 yr later	Tremolite	5.49 ± 13.97 µm	0.32 ± 3.52 µm	Tremolite had a pronounced inflammatory response with rapid granuloma development (1 d postexposure);  Slight interstitial fibrosis observed at 90 and 180 d postexposure.	Bernstein et al (2005; 2003)
C57Bl/6 mice (F) ( <i>n</i> = 10/group)	Intratracheal instillation Two doses of 60 µg each given wk apart in the first and second wk of a 7-mo experiment	Tremolite and wollastonite	Wollastonite: 4.46 ± 7.1 µm  Tremolite: N/A	Wollastonite: 0.75 ± 1.02 µm  Tremolite: N/A	Tremolite-exposed mice demonstrated increased IgG immune complex deposition in the kidneys, increased size of local lymph nodes, and increased total cell count.	Pfau et al. (2008)

BW = body weight; PBS = phosphate buffer saline.

<sup>a</sup>When available, results are shown as number of animals with tumors/total number of animals examined.

<sup>b</sup>Asbestiform types led to mesothelioma in most if not all exposed animals in this study.

#### 4.2.1. Inhalation

There are no laboratory animal studies following inhalation exposure to LAA; however, three studies have examined the effect of inhalation exposure to tremolite in Wistar rats (Bernstein et al., 2005; Bernstein et al., 2003; Davis et al., 1985). Davis et al. (1985) performed a chronic inhalation study examining response in male Wistar rats exposed in a chamber to 10 mg/m<sup>3</sup> (~1,600 fibers/mL, >5 µm) of commercially mined tremolite over a 12-month period. Bernstein et al. (2005) and Bernstein et al. (2003) exposed Wistar rats to tremolite (100 fibers/cm<sup>3</sup>) and chrysotile for 13 consecutive weeks (6 hours per day, 5 days per week) with 1-year follow-up. The results of these inhalation studies produced pronounced inflammation and very high levels of pulmonary fibrosis. Davis et al. (1985) also demonstrated an increase in carcinomas and mesotheliomas following exposure to tremolite, with no pulmonary tumors

observed in the controls. These results show that Wistar rats exposed to tremolite exhibited increased numbers of pulmonary lesions and possibly tumors.

#### 4.2.2. Intratracheal Instillation Studies

Intratracheal instillation has been used to examine the effect of exposure to Libby Amphibole ([Cyphert et al., 2012b](#); [Cyphert et al., 2012a](#); [Shannahan et al., 2012a](#); [Shannahan et al., 2012c](#); [Shannahan et al., 2012b](#); [Shannahan et al., 2012d](#); [Padilla-Carlin et al., 2011](#); [Shannahan et al., 2011a](#); [Shannahan et al., 2011b](#); [Smartt et al., 2010](#); [Putnam et al., 2008](#)) and tremolite asbestos ([Blake et al., 2008](#); [Pfau et al., 2008](#); [Sahu et al., 1975](#)). These studies exposed C57Bl/6 mice (100 µg/mouse), Wistar-Kyoto (WKY) rats (0.25 or 1 mg/rat) or Fischer 344 rats (0.25 or 6.5 mg/rat) once to LAA and analyzed the results up to 2 years postexposure. [Putnam et al. \(2008\)](#) observed statistically nonsignificant increases in collagen following exposure to LAA, as well as gene expression alterations related to membrane transport, signal transduction, epidermal growth factor signaling, and calcium regulation. [Smartt et al. \(2010\)](#) followed up this study by analyzing specific genes by quantitative reverse transcription polymerase chain reaction for genes involved in collagen accumulation and scar formation (Col1A1, Col1A2, Col3A1). LAA exposure led to increased gene expression of Col1A2 at 1 week postinstillation and Col3A1 at 1 month postexposure. Both studies observed increased inflammation; however, LAA exposure demonstrated minimal inflammation that did not progress in the time points examined. These studies demonstrate that exposure to LAA may lead to inflammation and fibrosis.

[Shannahan et al. \(2011a\)](#) exposed two rat models of human cardiovascular disease to LAA to determine if the preexisting CVD in these models would impact the lung injury and inflammation following exposure. Healthy WKY rats were compared to spontaneously hypertensive (SH) and spontaneously hypertensive-heart failure (SHHF) rats following exposure. All rats (male only) were exposed to 0, 0.25, or 1.0 mg/rat via intratracheal instillation and were examined at 1 day, 1 week, and 1 month postexposure. No changes were observed histopathologically; however, changes were observed in markers of homeostasis, inflammation, and oxidative stress. While inflammation and cell injury were observed in all strains, no strain-related differences were observed following exposure to LAA ([Shannahan et al., 2011a](#)).

A series of studies further examined SH, SHHF, and WKY rats over several durations of exposure to identify potential biomarkers of LAA exposure and determine if asbestos exposure shifts biomarker expression in healthy rats to resemble CVD ([Shannahan et al., 2012a](#); [Shannahan et al., 2012d](#)). Acute-phase response molecules involved in inflammatory responses such as  $\alpha$ 2-macroglobulin and  $\alpha$ 1-acid glycoprotein, as well as the metabolic molecule lipocalin-2 were generally increased 1 day after exposure regardless of duration ([Shannahan et al., 2012a](#)). In addition, LAA generally did not change biomarker expression similarly to the CVD rat strains ([Shannahan et al., 2012b](#)). However, the expression of two vasoconstriction

genes, eNOS and ETR-A, were altered in Libby Amphibole-exposed WKY rats to resemble untreated SH and SHHF rats ([Shannahan et al., 2012b](#)). Biomarkers for cancer were largely unaffected in all three strains following LAA exposure ([Shannahan et al., 2012a](#)).

In a follow-up study to further examine the role of iron in the inflammatory response to LAA exposure, [Shannahan et al. \(2011b\)](#) exposed SH rats to LAA alone and with bound Fe as well as with an iron chelator (deferoxamine [DEF]). Exposure to LAA led to significant increases in inflammatory markers (e.g., neutrophils, interleukin [IL]-8) with the greatest increase occurring in the presence of DEF. Iron bound to LAA was not released following instillation except in the presence of DEF as supported by the lack of increase of iron in bronchoalveolar lavage fluid (BALF). These results suggest that chelation of iron bound to LAA as well as endogenous proteins increases the toxicity of LAA in vivo.

A pair of studies further examined the effect of iron in the context of Libby Amphibole-induced lung injury and inflammasome activation. DEF treatment in addition to LAA significantly affected cyclooxygenase-2 (COX-2), IL-6, and CCL-11 in lung tissue compared to LAA treatment alone ([Shannahan et al., 2012c](#)). Addition of iron to LAA significantly altered NF- $\kappa$ B and IL-1 $\beta$  compared to LAA alone ([Shannahan et al., 2012c](#)). However, iron overload and DEF treatment generally were not significantly changed from each other, suggesting that iron has little impact on the inflammasome cascade. Histological examination and gene array analysis of inflammatory genes in WKY, SH, and SHHF rats did not identify significant differences in the progression of pulmonary fibrosis between the three strains ([Shannahan et al., 2012d](#)). These data do not indicate that the iron overload conditions that are characteristic of the cardiovascular disease-rat strains amplify the pulmonary effects of LAA. [Padilla-Carlin et al. \(2011\)](#) exposed Fischer 344 rats (male only) to LAA (0.65 or 6.5 mg/rat) or amosite (0.65 mg/rat; positive control) by intratracheal instillation to examine inflammatory response for 3 months postexposure. LAA exposure led to statistically significant increases of neutrophils in BALF as early as 1 day postexposure, with other inflammatory markers (e.g., protein, lactate dehydrogenase [LDH], gamma-glutamyl transpeptidase [GGT]) increased statistically significantly at different time points during the 3-month-period postexposure. However, on a mass basis, amosite produced a greater inflammatory response as measured by inflammatory markers (e.g., neutrophil influx, gene expression changes) and histopathological analysis demonstrating interstitial fibrosis. Examination of male Fischer 344 rats from this study at 2 years postexposure demonstrated that LAA induced a significant fibrogenic (but not carcinogenic) effect ([Cyphert et al., 2012b](#)). Response to LAA exposure in this study was less potent than amosite on a mass basis. Further comparison of LAA to other fiber types (Sumas Mountain chrysotile [SM], El Dorado tremolite [ED], and Ontario ferroactinolite [ON]) demonstrated that LAA exposure increased inflammatory markers at 1 day postexposure which returned to control levels by 3 months ([Cyphert et al., 2012a](#)). LAA exposure also led to an increased fibrogenic response at 3 months postexposure. As compared to other fibers tested,



fibrogenic response was correlated with the fiber length and width of each fiber, with SM-exposed rats demonstrating the greatest fibrogenic response (SM > LAA > ON > ED). These studies demonstrate a statistically significant increase in inflammatory response to LAA in mice and rats as measured in BALF by cytology, histopathology, and gene expression analysis. Follow-up studies are needed to inform the chronic effects of exposure to LAA.

Laboratory animal studies of tremolite intratracheal instillation exposure have been performed in mice in doses ranging from 60 µg to 5 mg. Male Swiss albino mice exposed to tremolite (5 mg) via intratracheal instillation demonstrated histological changes ([Sahu et al., 1975](#)). Microscopic results following exposure to tremolite showed acute inflammation of the lungs at 7 days postexposure, including macrophage proliferation and phagocytosis similar to that observed with amosite and anthophyllite. Limited progression of fibrotic response was observed at 60 and 90 days postexposure, with no further progression of fibrotic response.

#### 4.2.3. Injection/Implantation Studies

There are no laboratory animal studies examining intraperitoneal injection or implantation of LAA. Biological effects following exposure to tremolite have been examined in five intraperitoneal injection studies ([Roller et al., 1997, 1996](#); [Davis et al., 1991](#); [Wagner et al., 1982](#); [Smith et al., 1979](#); [Smith, 1978](#)) and one implantation study ([Stanton et al., 1981](#)).

Studies by Smith and colleagues ([Smith et al., 1979](#); [Smith, 1978](#)), [Wagner et al. \(1982\)](#), [Davis et al. \(1991\)](#), and [Roller et al. \(1997, 1996\)](#) demonstrated that intrapleural injections of tremolite asbestos<sup>17</sup> is associated with an increase in pleural fibrosis and mesothelioma in hamsters and rats compared to controls or animals injected with less fibrous materials. Doses ranged from 10–25 mg/animal for each study, and although carcinogenesis was observed in these studies, there was a variable level of response to the different tremolite forms examined. Although these studies clearly show the carcinogenic potential of Libby Amphibole or tremolite asbestos fibers, intrapleural injections bypass the clearance and dissolution of fibers from the lung after inhalation exposures. Further, limited information was provided confirming the presence or absence of particles or fibers less than 5 µm in length in these studies, limiting the interpretation of results.

One laboratory animal study examined the effect of tremolite exposure following implantation of fibers in the pleural cavity. [Stanton et al. \(1981\)](#) examined tremolite and described a series of studies on various forms of asbestos. Fibers embedded in hardened gelatin were placed against the lung pleura. As an intrapleural exposure, results might not be comparable to inhalation exposures, as the dynamics of fiber deposition and pulmonary clearance mechanisms are not accounted for in the study design. Studies using two tremolite

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<sup>17</sup>[Smith \(1978\)](#) used tremolite from Libby, MT; [Smith et al. \(1979\)](#) may also have used tremolite from Libby, MT (i.e., Libby Amphibole asbestos).

asbestos samples from the same lot were described as being in the optimal size range for carcinogenesis; the fibers were distinctly smaller in diameter than the tremolite fibers [Smith et al. \(1979\)](#) used. These samples both had a high number of fibers in the size range (>8-μm long and <0.25 μm in diameter; i.e., “Stanton fibers”). Exposure to both tremolite samples led to mesotheliomas in 21 and 22 of 28 rats exposed. The [Stanton et al. \(1981\)](#) study also used talc, which did not lead to mesothelioma production.

There are no studies currently available in laboratory animals exposed to LAA by inhalation. However, the chronic intraperitoneal injection study in hamsters ([Smith et al., 1979](#); [Smith, 1978](#)) demonstrated tumor formation following exposure to tremolite obtained from the Libby, MT mine. No other chronic inhalation studies of LAA are available. A recent study in rats examining the impact of preexisting cardiovascular disease on pulmonary inflammation demonstrated an increase in inflammatory markers following exposure to LAA via intratracheal instillation in SH rats as compared to normal healthy controls exposed to the same dose ([Shannahan et al., 2011b](#)). More recent studies examined gene expression changes ([Hillegass et al., 2010](#); [Putnam et al., 2008](#)) and early protein markers of fibrosis ([Smartt et al., 2010](#)) in mice exposed to LAA via intraperitoneal injection. These studies demonstrated an increase in gene and protein expression related to fibrosis following exposure to LAA. Tremolite fibers, although obtained from different locations throughout the world, consistently led to pulmonary lesions and/or tumor formation with various routes of exposure (inhalation, injection, instillation) and in multiple species (rats, hamsters, and mice; [Bernstein et al., 2005](#); [Bernstein et al., 2003](#); [Roller et al., 1997, 1996](#); [Davis et al., 1985](#); [Wagner et al., 1982](#); [Stanton et al., 1981](#)). Although comparing potency of the various forms of tremolite is difficult given the limited information on fiber characteristics and study limitations (e.g., length of follow-up postexposure), these results show potential increased risk for cancer (lung and mesothelioma) following exposure to tremolite asbestos.

The results of the studies described above show the fibrogenic and carcinogenic potential of Libby Amphibole and tremolite asbestos. Further, the more recent studies by Salazar ([2013](#); [2012](#)), Blake et al. ([2008](#)), and Pfau et al. ([2008](#)) support human studies demonstrating potential autoimmune effects of asbestos exposure (see Section 4.3.1).

#### 4.2.4. Oral

No studies in laboratory animals with oral exposure to Libby Amphibole were found in the literature. However, one chronic cancer bioassay was performed following oral exposure to tremolite. [McConnell et al. \(1983a\)](#) describe part of a National Toxicology Program study ([NTP, 1990b](#)) performed to evaluate the toxicity and carcinogenicity of ingestion of several minerals, including tremolite. The tremolite (Gouverneur Talc Co, Gouverneur, NY) used was not fibrous. No significant tumor induction was observed in the animals with oral exposure to tremolite animals. Although nonneoplastic lesions were observed in many of the aging rats, these were

1 mostly in the stomach and occurred in both controls and exposed animals. The observed lesions  
2 included chronic inflammation, ulceration, and necrosis of the stomach ([McConnell et al.,  
3 1983a](#)). [McConnell et al. \(1983a\)](#) suggested that nonfibrous nature of this tremolite sample  
4 could account for the lack of toxicity following exposure in this group of animals. Also, oral  
5 studies of asbestos, in general, show decreased toxicity and carcinogenicity as compared to  
6 inhalation and implantation/injection studies ([Condie, 1983](#)).  
7

#### 8 **4.2.5. Summary of Animal Studies for Libby Amphibole and Tremolite Asbestos**

9 Tables 4-19 and 4-20 summarize the studies described in this section, with full study  
10 details available in Appendix D. Limited in vivo studies have been performed exposing  
11 laboratory animals to LAA. One intrapleural injection study using tremolite from the Libby, MT  
12 area is included in this section under LAA since earlier terminology for LAA was often tremolite  
13 ([Smith, 1978](#)). Hamsters in this study exposed to LAA developed fibrosis and mesothelioma  
14 following exposure. Intratracheal instillation studies of LAA in rats showed increased collagen  
15 gene expression at 2-years postexposure ([Cyphert et al., 2012a](#)). Subchronic-duration studies in  
16 mice ([Smartt et al., 2010](#); [Putnam et al., 2008](#)) demonstrated gene and protein expression  
17 changes related to fibrosis production following exposure to LAA. Finally, short-term-duration  
18 studies in rats demonstrated an increase in inflammatory and cardiovascular disease markers  
19 following exposure to LAA ([Padilla-Carlin et al., 2011](#); [Shannahan et al., 2011a](#); [Shannahan et  
20 al., 2011b](#)).

21 Because tremolite is part of LAA, results from tremolite studies were also described. In  
22 general, fibrous tremolite has been shown to cause pulmonary inflammation, fibrosis, and/or  
23 mesothelioma or lung cancer in rats ([Bernstein et al., 2005](#); [Bernstein et al., 2003](#); [Davis et al.,  
24 1991](#); [Davis et al., 1985](#); [Wagner et al., 1982](#)) and hamsters ([Smith et al., 1979](#)). The single  
25 short-term-duration study on mice showed limited response to tremolite ([Sahu et al., 1975](#)). The  
26 one chronic-duration oral study ([McConnell et al., 1983a](#)) did not show increased toxicity or  
27 carcinogenicity; this study, however, used only nonfibrous tremolite, which later studies showed  
28 to be less toxic and carcinogenic than fibrous tremolite ([Davis et al., 1991](#)).

29 Chronic inflammation is hypothesized to lead to a carcinogenic response through the  
30 production of reactive oxygen species and increased cellular proliferation ([Hanahan and  
31 Weinberg, 2011](#)). Although limited, the data described in Section 4.2 suggest an increase in  
32 inflammatory response following exposure to LAA and tremolite asbestos similar to that  
33 observed for other durable mineral fibers ([reviewed in Mossman et al., 2007](#)). Whether this  
34 inflammatory response then leads to cancer is unknown. Studies examining other types of  
35 asbestos (e.g., crocidolite, chrysotile, and amosite) have demonstrated an increase in chronic  
36 inflammation as well as respiratory cancer related to exposure ([reviewed in Kamp and  
37 Weitzman, 1999](#)). Chronic inflammation has also been linked to genotoxicity and mutagenicity  
38 following exposure to some particles and fibers ([Driscoll et al., 1997](#); [Driscoll et al., 1996](#);

1 [Driscoll et al., 1995](#)). The evidence described above suggests chronic inflammation is observed  
2 following LAA and tremolite asbestos exposure; however, the role of inflammation and whether  
3 it leads to lung cancer or mesothelioma following exposure to LAA is unknown.

4 ROS production has been measured in response to both LAA and tremolite asbestos  
5 exposure. [Blake et al. \(2007\)](#) demonstrated an increase in the production of superoxide anions  
6 following exposure to LAA. [Blake et al. \(2007\)](#) also demonstrated that total superoxide  
7 dismutase (SOD) was inhibited, along with a decrease in intracellular glutathione (GSH), both of  
8 which are associated with increased levels of ROS. These results are supported by a recent study  
9 in human mesothelial cells ([Hillegass et al., 2010; described in Section 4.4 and Appendix D](#)).  
10 Increased ROS production was also observed in human airway epithelial cells (HAECs)  
11 following exposure to LAA ([Duncan et al., 2010; described in Section 4.4 and Appendix D](#)).  
12 This increase in ROS and decrease in glutathione are common effects following exposure to  
13 asbestos fibers and particulate matter. [Pfau et al. \(2012\)](#) examined the role of the amino acid  
14 transport system  $x_c^-$ , which is one of the pathways murine macrophages use to detect and  
15 respond to stressful conditions. This study demonstrated that ROS production increase system  
16  $x_c^-$  activity. Although ROS production is relevant to humans, based on similar human responses  
17 as compared to animals, information on the specifics of ROS production following exposure to  
18 LAA is limited to the available data described here. Therefore, the role of ROS production in  
19 lung cancer and mesothelioma following exposure to LAA is unknown.

20 Recent studies have also examined the role of the inflammasome and iron in the  
21 development of fibrosis in male SH-rats. The role of inflammasome activation and iron in the  
22 development of LAA-induced fibrosis was studied in [Shannahan et al. \(2012c\)](#). Lung tissue  
23 expression of inflammatory cytokines CCL-7, Cox-2, CCL-2, and CXCL-3 was increased  
24 4 hours following LAA exposure. Conversely, LAA exposure reduced IL-4 and CXCL-1 in the  
25 BALF. Finally, the ratio of phosphorylated ERK (pERK)/extracellular signal-related kinases  
26 (ERK), which is an upstream activator of the inflammasome cascade, was increased in the lung  
27 of LAA exposed rats 1 day post exposure. Rats treated with LAA + DEF or LAA + Fe had  
28 significantly different levels of Cox-2 in the BALF and IL-6 in lung tissue but all other endpoints  
29 were not significantly different. These data suggest that the concentration of iron does not  
30 impact the activation of the inflammasome cascade and cytokines downstream of the pathway in  
31 LAA-exposed animals.

32 In another study examining the role of iron in lung disease, [Shannahan et al. \(2012d\)](#)  
33 evaluated the effect of Fe overload on LAA-induced lung injury in rats with cardiovascular  
34 disease. Gene array analysis demonstrated that LAA exposure upregulated inflammatory-related  
35 genes such as NF- $\kappa$ B and cell cycle regulating genes such as matrix metalloproteinase-9 in WKY  
36 rats but inhibited these same cluster of genes in SH and SHHF animals 3 months after  
37 instillation. Histological examination of lung sections observed greater Fe staining of  
38 macrophages in SHHF rats compared to WKY and SH rats 1 and 3 months post exposure;

1 however, no differences in the progression of pulmonary fibrosis were noted between the three  
2 strains. Altogether, these data do not suggest that the iron overload conditions that are  
3 characteristic of the cardiovascular disease strains amplify the pulmonary effects of LAA.

### 4.3. OTHER DURATION- OR ENDPOINT-SPECIFIC STUDIES

#### 4.3.1. Immunological

4 Salazar et al. (2013; 2012), Rasmussen and Pfau (2012), Blake et al. (2008), Pfau et al.  
5 (2008), Serve et al. (2013) and Hamilton et al. (2004) examined the role of asbestos in  
6 autoimmunity in laboratory animal or in vitro studies. Blake et al. (2008) performed in vitro  
7 assays with LAA (see Section 4.4), and both studies performed the in vivo assays with tremolite.  
8 C57BL/6 mice were instilled intratracheally for a total of two doses each of 60-μg saline and  
9 wollastonite or Korean tremolite sonicated in sterile phosphate buffered saline (PBS), given  
10 1 week apart in the first 2 weeks of a 7-month experiment. Sera from mice exposed to tremolite  
11 showed antibody binding colocalized with SSA/Ro52 on the surface of apoptotic blebs (Blake et  
12 al., 2008). In Pfau et al. (2008), by 26 weeks, the tremolite-exposed animals had a significantly  
13 higher frequency of positive antinuclear antibody (ANA) tests compared to wollastinate and  
14 saline. Most of the tests were positive for double-stranded DNA (dsDNA) and SSA/Ro52.  
15 Serum isotyping showed no major changes in immunoglobulin (Ig) subclasses (IgG, IgA, IgM),  
16 but serum IgG in tremolite-exposed mice decreased overall. Further, IgG immune complex  
17 deposition in the kidneys increased, with abnormalities suggestive of glomerulonephritis. No  
18 increased proteinuria was observed during the course of the study. Local immunologic response  
19 was further studied on the cervical lymph nodes. Although total cell numbers and lymph-node  
20 sizes were significantly increased following exposure to tremolite, percentages of T- and B-cells  
21 did not significantly change.

22 Hamilton et al. (2004) investigated the ability of Libby Amphibole, crocidolite, and  
23 particulate matter 2.5 μm in diameter or less (PM<sub>2.5</sub>, collected over a 6-month period in Houston,  
24 TX, from EPA site 48-201-1035) to alter the antigen-presenting cell (APC) function in cultured  
25 human alveolar macrophages. Asbestos exposure (regardless of type) and PM<sub>2.5</sub> up-regulated a  
26 Th1 lymphocyte-derived cytokine, interferon gamma (IFNγ), and the Th2 lymphocyte-derived  
27 cytokines IL-4 and IL-13. However, extreme variation among subjects was noted in the amount  
28 of response. In addition, no correlation was present between an individual's cells' response to  
29 asbestos versus particulate matter, suggesting that more than one possible mechanism exists for a  
30 particle-induced APC effect and individual differential sensitivities to inhaled bioactive particles.  
31 Rasmussen and Pfau (2012) examined the role of macrophages in the development of  
32 autoantibody production following exposure to LAA. LAA exposure alone did not affect cell  
33 proliferation or antibody production; however, culturing lymphocytes with macrophage medium  
34 following exposure to LAA did lead to increased cellular proliferation and antibody production.



[Serve et al. \(2013\)](#) examined a possible role of autoimmunity in fibrosis by an in vitro examination of potential mechanisms of MCAA leading to collagen deposition, a precursor to fibrosis. This study demonstrated MCAA binding leads to increased collagen deposition through altering matrix metalloproteinase (MMP) expression.

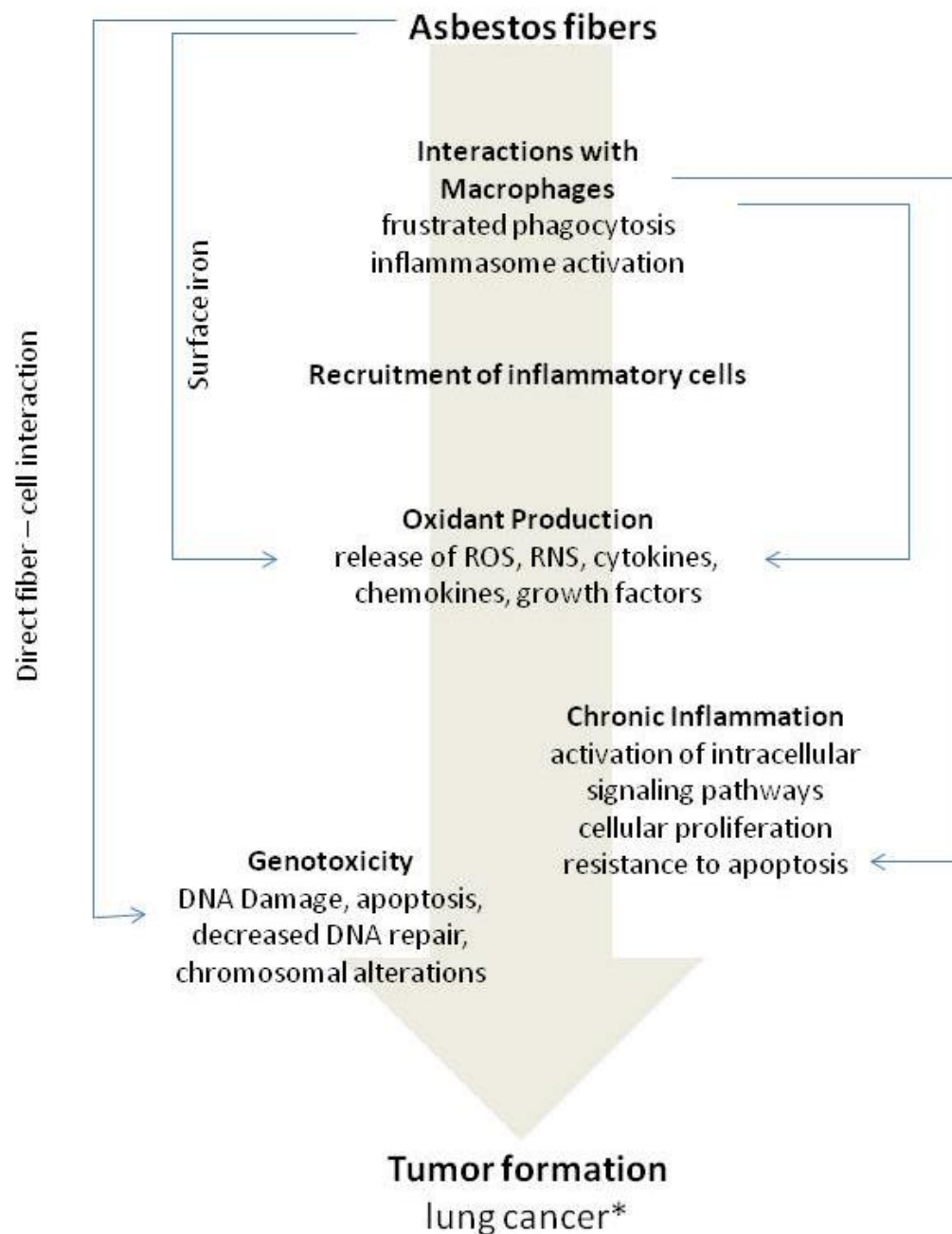
In two studies examining potential autoimmune effects of LAA exposure, Salazar et al. ([2013](#); [2012](#)) examined the potential impact of LAA exposure on rheumatoid arthritis and on ANA increases associated with systemic autoimmune disease. Salazar et al. ([2013](#); [2012](#)) conducted a series of studies to establish the effects of LAA exposure on autoimmune disease. The first set of studies utilized the collagen-induced arthritis and peptidoglycan-polysaccharide (PG-PS) models of rheumatoid arthritis to determine whether LAA exposure increased the onset, prolonged, or intensified the joint inflammation characteristic of the disease ([Salazar et al., 2012](#)). Female Lewis rats were instilled biweekly for 13 weeks with a total dose of 0, 0.15, 0.5, 1.5, and 5.0 mg LAA followed by induction with either model of arthritis. LAA 5.0 mg reduced the magnitude of the swelling response in the cell-mediated PG-PS model; however, neither the onset nor the duration of swelling was affected by LAA exposure. LAA 1.5 and 5.0 mg and amosite 0.5 and 1.5 mg reduced total serum IgM. LAA 5.0 mg and amosite 1.5 mg reduced anti-PG-PS IgG in the serum 17 weeks after the final instillation. Finally, the number of rats positive for ANA was increased only at the low exposure concentrations of LAA in PG-PS-treated and nonarthritic rats. These results suggest that LAA may have a modest inhibitory effect on the PG-PS rat model but may enhance responses to other systemic autoimmune diseases.

In a follow-up study, [Salazar et al. \(2013\)](#) explored in greater detail the effect of LAA exposure on ANA over time and the antigen specificity of the ANA. Female Lewis rats were intratracheally instilled under the conditions in the previous study ([Salazar et al., 2012](#)). Serum samples were analyzed every 4 weeks from the beginning of the instillations up to termination at Week 28. Since elevated ANA are commonly associated with kidney disease, proteinuria was assessed every 3 weeks beginning at Week 6 until termination of the experiment. Histopathological analysis was also performed on the kidneys. ANA was increased 8 weeks postexposure to LAA 5.0 mg. By Week 28, all doses of LAA except 1.5 mg increased ANA in the serum. Analysis of the antigen specificity found that only the LAA at 1.5 mg significantly increased antibodies specific for extractable nuclear antigens and the Jo-1 antigen. Urinalysis found that all doses of LAA exposure induce moderate levels of proteinuria, but this effect was not dose responsive. No dose-related histopathological effects were observed. Altogether, these data suggest that LAA exposure increases autoimmune antibodies in the serum, but no evidence of autoimmune disease was identified. However, the lack of SAID in the Lewis rat may be due to strain-specific factors and suggests that other animal models may be more appropriate for studying autoimmune effects of LAA.

1 Although the number of studies is limited, the results suggest a possible effect on  
2 autoimmunity following exposure to LAA. Further studies are needed to increase understanding  
3 of this potential effect.  
4

#### 5 **4.4. MECHANISTIC DATA AND OTHER STUDIES IN SUPPORT OF THE MODE OF** 6 **ACTION**

7 For asbestos in general, International Agency for Research on Cancer (IARC) has  
8 proposed a mechanism for the carcinogenicity of asbestos fibers ([see Figure 4-4; IARC, 2012](#)).  
9 Asbestos fibers lead to oxidant production through interactions with macrophages and through  
10 hydroxyl radical generation from surface iron. Inhaled fibers that are phagocytosed by  
11 macrophages may be cleared or lead to frustrated phagocytosis, which results in macrophage  
12 activation, release of oxidants, and increased inflammatory response, in part due to  
13 inflammasome activation. Free radicals may also be released by interaction with the iron on the  
14 surface of fibers. Increased oxidant production may result in epithelial cell injury, including  
15 DNA damage. Frustrated phagocytosis may also lead to impaired clearance of fibers, with fibers  
16 being available for translocation to other sites (e.g., pleura). Mineral fibers may also lead to  
17 direct genotoxicity by interfering with the mitotic spindle and leading to chromosomal  
18 aberrations. Asbestos exposure also leads to the activation of intracellular signaling pathways,  
19 which in turn may result in increased cellular proliferation, decreased DNA damage repair, and  
20 activation of oncogenes. Research on various types of mineral fibers supports a complex  
21 mechanism involving multiple biologic responses following exposure to asbestos (i.e.,  
22 genotoxicity, chronic inflammation/cytotoxicity leading to oxidant release, and cellular  
23 proliferation) in the carcinogenic response to mineral fibers ([see Figure 4-4, reviewed in IARC,](#)  
24 [2012](#)). These complexities of fiber toxicity need to be considered when analyzing MOA for  
25 asbestos (as reviewed in [Aust et al., 2011](#); [Broadbush et al., 2011](#); [Bunderson-Schelvan et al.,](#)  
26 [2011](#); [Huang et al., 2011](#); [Mossman et al., 2011](#)).



\*similar mechanisms expected for mesothelioma following translocation of fibers from the lung to the pleura

**Figure 4-4. Proposed mechanistic events for carcinogenicity of asbestos fibers.**

Adapted from [IARC \(2012\)](#).



1 Important considerations in evaluating the available mechanism and MOA data are fiber  
2 characteristics, route of exposure, dose metric, as well as study design and interpretation.  
3 Specific fiber characteristics impact the fiber toxicokinetics (reviewed in Section 3), and in turn  
4 the biologic response to fibers. For example, fiber length is an important determinant of fiber  
5 clearance, with shorter fibers generally being cleared more efficiently as longer fibers result in  
6 frustrated phagocytosis. Mechanisms of carcinogenesis may accordingly vary based on fiber  
7 characteristics. The biologic response to respirable fibers is also influenced by the route of  
8 exposure. Inhalation exposure studies are the most informative. Intratracheal instillation and  
9 aspiration exposures bypass normal clearance mechanisms, and therefore affect fiber dosimetry.  
10 Concerning dose metric, some studies suggest that the dose should be determined based on fiber  
11 length, width, number, or surface area ([Case et al., 2011](#); [Mossman et al., 2011](#)). However, the  
12 majority of studies of fibers have been performed using mass as a dose. Finally, an important  
13 consideration in analysis of in vitro studies is the cell types used, particularly related to the  
14 ability to internalize fibers and produce an oxidative stress response. The discussions below  
15 highlight these considerations in presenting the available mechanistic evidence for LAA.

16 Limited in vitro studies have been conducted with LAA from the Zonolite Mountain  
17 mine. These studies demonstrated an effect of LAA on inflammation and immune function  
18 ([Duncan et al., 2014](#); [Duncan et al., 2010](#); [Blake et al., 2008](#); [Blake et al., 2007](#); [Hamilton et al.,](#)  
19 [2004](#)), oxidative stress ([Hillegass et al., 2010](#)), and genotoxicity ([Pietruska et al., 2010](#)). Similar  
20 endpoints have been examined in vitro following exposure to tremolite asbestos ([Okayasu et al.,](#)  
21 [1999](#); [Wylie et al., 1997](#); [Suzuki and Hei, 1996](#); [Athanasίου et al., 1992](#); [Wagner et al., 1982](#)).  
22 Results from in vitro studies have demonstrated potential biological mechanisms of oxidative  
23 stress and inflammation in response to exposure to Libby Amphibole and tremolite asbestos. As  
24 discussed in Section 4.2, laboratory animal studies examining the effects of tremolite exposure  
25 have been reviewed and are summarized to potentially increase understanding of the effects and  
26 mechanisms of LAA, because tremolite is a component of LAA (~6%). These studies are  
27 summarized below and in Tables 4-21 and 4-22, with detailed study descriptions available in  
28 Appendix D.

**Table 4-21. In vitro data following exposure to Libby Amphibole asbestos**

Test system	Fiber type	Dose/exposure duration	Effects	Reference
Primary human alveolar macrophages and lymphocytes	LAA or crocidolite	0, 25, 50 µg/mL 24 h	Upregulated Th1 and Th2 cytokines (IFNγ, IL-4, IL-13)	<a href="#">Hamilton et al. (2004)</a>
Murine macrophages (primary and RAW264.7) <sup>a</sup>	LAA and crocidolite	Internalization: 0, 5, 62.5 µg/cm <sup>2</sup> 3–24 h	Internalized LAA fibers were mostly less than 2 µm in length	<a href="#">Blake et al. (2007)</a>
		Oxidative stress: 0, 6.25, 32.5, 62.5 µg/cm <sup>2</sup> 3, 7, 12, and 24 h	Increased ROS over control (wollastonite) and crocidolite Decreased GSH	
		Cell viability: 0, 6.25, 32.5, 62.5 µg/cm <sup>2</sup> 3, 7, 12, and 2 h	No effect was observed on cell viability	
		DNA damage: 0, 6.25, 32.5, 62.5 µg/cm <sup>2</sup> 3, 7, 12, and 24 h	No increase in DNA damage and adduct formation	
Murine macrophages (primary and RAW264.7)	LAA or crocidolite	0, 62.5 µg/cm <sup>2</sup> 0–72 h	Time-course dose-response for apoptosis; redistribution of autoantigen on cell surface	<a href="#">Blake et al. (2008)</a>
Human lung epithelial cells (wild-type and XRCC1-deficient)	LAA or crocidolite	5 µg/cm <sup>2</sup> 24 h	Dose-dependent increase in micronuclei in both cell types, but increased in the XRCC1-deficient cells as compared to wild-type	<a href="#">Pietruska et al. (2010)</a>
Human mesothelial cell lines (LP9/TERT-1 and HKNM-2)	LAA or crocidolite	0, 15 × 10 <sup>6</sup> µm <sup>2</sup> /cm <sup>2</sup> (nontoxic) and 75 × 10 <sup>6</sup> µm <sup>2</sup> /cm <sup>2</sup> (toxic) for 8 or 24 h	Alterations in genes related to oxidative stress at cytotoxic doses, particularly <i>SOD2</i>	<a href="#">Hillegass et al. (2010)</a>
Primary HAECs	LAA (fractionated and unfractionated), amosite (fractionated and unfractionated), crocidolite	0, 2.64, 13.2 or 26.4 µg/cm <sup>2</sup> 2, 4 or 24 h	Increases in proinflammatory gene expression and ROS production	<a href="#">Duncan et al. (2010)</a>

**Table 4-21. In vitro data following exposure to Libby Amphibole asbestos (continued)**

Test system	Fiber type	Dose/exposure duration	Effects	Reference
CH12.LX B-lymphocytes	LAA	35 µg/cm <sup>2</sup> 48 h	Data from macrophage-conditioned media demonstrate that asbestos leads to immunologic changes consistent with activation of B1a B-lymphocytes.	<a href="#">Rasmussen and Pfau (2012)</a>
MeT-5A human mesothelial cells	LAA	Cells exposed to sera from exposed individuals that were MCAA+ or MCAA- (1:100).	Data demonstrated that MCAA binding leads to increased collagen deposition through altering MMP expression.	<a href="#">Serve et al. (2013)</a>
Primary human airway epithelial cells (HAEC)	LAA (2000, 2007); amosite (RTI, UICC)	0, 2.64, 13.2 or 26.4 µg/cm <sup>2</sup> 24 h	Exposure to all fibers at the highest doses led to increased LDH levels (cytotoxicity) and increased mRNA expression of IL-8, IL-6, COX-2, and TNFα (inflammatory markers). On an equal mass basis LA is as potent as UICC amosite at inducing a proinflammatory response in HAEC but less potent than RTI amosite.	<a href="#">Duncan et al. (2014)</a>
THP-1 cells (macrophage cell line)	Libby six-mix Chrysotile	0, 20, 40 µg/ml 24 h	LAA activated the NLRP3 inflammasome but to a lesser degree than chrysotile, but LAA exposure generated more ROS production compared to chrysotile.	<a href="#">Li et al. (2012)</a>

XRCC1 = x-ray repair cross complementing protein 1.

<sup>a</sup>All results for RAW264.7. Data not shown for primary cells though authors state similar response to RAW264.7.

**Table 4-22. In vitro data following exposure to tremolite asbestos**

Test system/species	Fiber type	Dose/exposure duration	Effects	Reference
Primary murine macrophages	Sample A (flake-like from California talc deposits); Sample B (medium-sized fibrous from Greenland); Sample C (fine-fiber material from S. Korea); Positive Control (crocidolite)	0, 50, 100, and 150 µg/mL 18 h	LDH and BGL levels increased following exposure to Sample C (longer, thinner fibers) and crocidolite (positive control). Sample C led to the greatest increases in giant cell formation (i.e., cells > 2 µm in diameter) and cytotoxicity of samples tested. Sample B also led to some increased cytotoxicity.	<a href="#">Wagner et al. (1982)</a>
TA98, TA100, TA102 <i>S. typhimurium</i>	Metsovo tremolite	TA98, TA100, and TA102: 0–500 µg/per plate 2 d	No significant revertants were observed in any of the three <i>Salmonella</i> strains tested.	<a href="#">Athanasίου et al. (1992)</a>
V79 and BPNi cells		V79 and BPNi: 0–4 µg/cm <sup>2</sup> 6, 24, and 48 h	No affect was observed on gap-junctional intercellular communication.	
BPNi cells		BPNi: 0–2 µg/cm <sup>2</sup> 24 h	Tremolite led to a dose-dependent increase in micronuclei induction.	
SHE cells		SHE: 0–3 µg/cm <sup>2</sup> 24 h	Tremolite exposure led to increased chromosomal aberrations but not in a dose-dependent fashion.	
A <sub>L</sub> cells (hamster hybrid cells containing human chromosome 11)	UICC chrysotile, crocidolite, Metsovo tremolite, erionite	0, 2.5–40 µg/mL 24 h	Relative increase in heme oxygenase as compared to control.	<a href="#">Suzuki and Hei (1996)</a>
HTE and RPM cell lines	NIEHS chrysotile, NIEHS crocidolite, FD14, S157, CPS 183 (talc fibers containing tremolite)	Varied (based on weight, fiber length, and surface area).	Fibrous talc exposure led to limited proliferation of cells.	<a href="#">Wylie et al. (1997)</a>

**Table 4-22. In vitro data following exposure to tremolite asbestos (continued)**

Test system/species	Fiber type	Dose/exposure duration	Effects	Reference
A <sub>L</sub> cells (hamster hybrid cells containing human chromosome 11)	Tremolite, erionite, RCF-1	0–400 µg/mL 24 h	No significant increase in hypoxanthine-guanine phosphoribosyltransferase mutations for these three fibers. Dose-dependent induction of S1 <sup>+</sup> mutations in Chromosome 11 occurs for erionite and tremolite.	<a href="#">Okayasu et al. (1999)</a>

A[L] cells = hamster hybrid cells containing human chromosome 11; BGL =  $\beta$ -glucuronidase; HTE = hamster tracheal epithelial; NIEHS = National Institute of Environmental Health Sciences; RCF-1 = refractory ceramic fibers ; RPM = rat pleural mesothelial; SHE = Syrian hamster embryo.

#### 4.4.1. Inflammation and Immune Function

Chronic inflammation following inhalation exposure to asbestos has been studied for decades in both humans and animals ([reviewed in Mossman et al., 2011](#)). This inflammatory response has been attributed to the role of alveolar macrophages, which attempt to engulf asbestos fibers to clear them from the respiratory tract. Mechanistic studies have focused on the chemokine/cytokine response of these macrophages, and subsequent signaling pathways that are activated. More recently, studies have also examined the role of inflammasome activation following exposure to asbestos in immune activation and fiber clearance ([Hillegass et al., 2013](#); [Biswas et al., 2011](#); [Dostert et al., 2008](#)).

Increased cytokine and chemokine production has been observed following exposure to LAA as well as other amphibole asbestos. [Hamilton et al. \(2004\)](#) showed an increase in Th1 and Th2 cytokines following exposure to LAA, crocidolite, and particulate matter, suggesting a similar effect of exposure to these materials on immune function. Analysis of these results is limited, as the use of primary cells in culture led to an extremely variable response. Two studies by [Blake et al. \(2008\)](#) and [Blake et al. \(2007\)](#) further examined the effect of LAA on immune response in vitro in mouse macrophages. These studies demonstrated that the size of the LAA material is such that it was able to be internalized by macrophages ( $<10\text{ }\mu\text{m}$ ), and this internalization resulted in an increase in ROS production. These studies also showed a variable cytotoxic response, because LAA exposure did not result in a statistically significant increase in cytotoxicity, while crocidolite did. DNA damage also was increased in crocidolite-exposed cells but not in LAA-exposed cells. An increase (relative to controls) in autoantibody formation following exposure to LAA was also observed. Studies that examined cellular response to tremolite also found that fiber characteristics (length and width) play a role in determining ROS production, toxicity, and mutagenicity ([Okayasu et al., 1999](#); [Wagner et al., 1982](#)).

Gene expression alterations of *IL-8*, *COX-2*, *heme oxygenase (HO)-1*, as well as other stress-responsive genes as compared to amosite was observed in primary HAEC following exposure to LAA. Comparisons were made with both fractionated (aerodynamic diameter  $\leq 2.5\text{ }\mu\text{m}$ ) and unfractionated fiber samples ([Duncan et al., 2010](#)). Crocidolite fibers (UICC) were also included in some portions of this study for comparison. Primary HAECs were exposed to 0, 2.64, 13.2, and  $26.4\text{ }\mu\text{g}/\text{cm}^2$  of crocidolite, amosite, amosite 2.5 (fractionated), LAA, or LAA 2.5 (fractionated) for 2 or 24 hours in cell culture. Minimal increases in gene expression of *IL-8*, *COX-2*, or *HO-1* were observed at 2 hours postexposure to all five fiber types; at 24 hour postexposure, however, a dose-response was observed following exposure to all fiber types with the results showing a proinflammatory gene expression response ([Duncan et al., 2010](#)). Cytotoxicity was determined by measuring LDH from the maximum dose ( $26.4\text{ }\mu\text{g}/\text{cm}^2$ ) of both amosite and LAA samples, with less than 10% LDH present following exposure to all four samples. A follow-up study with the same design by ([Duncan et al. \(2014\)](#)) was performed to examine the in vitro determinants of asbestos fiber toxicity, comparing two samples each of

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LAA (LA2000, LA2007) and amosite (UICC, RTI) asbestos. Primary human airway epithelial cells (HAEC) were exposed for 24 hours to 2.64, 13.2 or 26.4 ug/cm<sup>2</sup> LAA and amosite asbestos that had been analyzed for fiber size distribution, surface area, and surface-conjugated iron (see Table D-11). Most characteristics were similar, except RTI amosite consisted of longer fibers. Fiber toxicity was measured by cytotoxicity (LDH assay), levels of ROS production, as well as IL-8 mRNA levels as a measure of relative proinflammatory responses. Cytotoxicity levels were similar for all four samples at the highest dose, but statistically significant compared to the no-treatment control. Results on an equal mass basis demonstrated a statistically significant increase in IL-8, IL-6, COX-2, and TNF mRNA levels for all four amphiboles at the highest two doses. The greatest increase in IL-8 mRNA levels was following exposure to the RTI amosite sample, while both LAA samples and the UICC amosite resulted in a similar level of response to each other. Therefore, IL-8 was used to further analyze the dose metrics for this response. Surface iron concentrations and surface reactivity was quantified with respect to hydroxyl radical production to assess the effect of these properties on IL-8 mRNA expression. Surface iron concentrations were similar for the two LAA samples and for the two amosite samples, but the amosite samples had much greater surface iron as compared to the LAA samples. UICC amosite had slightly greater iron as compared to RTI. A strong correlation was observed between fiber dose metrics of length and external surface area. When these metrics were used in place of equal mass dose, the differential IL-8 mRNA expression following exposure to these four samples was eliminated. These results support a limited cytotoxicity and increased inflammatory cytokine response of both amosite and LAA under these concentrations and time frames.

The role of macrophages in the development of autoantibody production following exposure to asbestos was examined in a study performed by ([Rasmussen and Pfau \(2012\)](#)) culturing CH12.LX B-lymphocytes, a murine B1 lymphocyte cell line, with LAA alone did not affect proliferation or antibody production. However, culturing RAW264.7 macrophages with LAA, collecting the macrophage medium, and culturing CH12.LX lymphocytes in the conditioned medium reduced CH12.LX proliferation and increased IgG1, IgG3, and IgA production. Further analysis found that both IL-6 and tumor necrosis factor (TNF)- $\alpha$  were elevated in the medium of Libby Amphibole-treated macrophages, but only IL-6 increased IgG and IgA production. However, these data also indicate that activated macrophages may regulate CH12.LX antibody production. Altogether, these data suggest a potential mechanism for macrophages to regulate asbestos-induced autoantibody production in Libby-exposed residents.

Chronic inflammation is also associated with oxidative stress, mechanisms of which following exposure to LAA were also studied in human mesothelial cells ([Hillegass et al., 2010](#)). Gene expression changes related to oxidative stress following exposure to  $15 \times 10^6 \mu\text{m}^2/\text{cm}^2$

LAA<sup>18</sup> as compared to the nonpathogenic control ( $75 \times 10^6 \mu\text{m}^2/\text{cm}^2$  glass beads) in the human mesothelial cell line LP9/TERT-1 for 8 and 24 hours. Gene ontology of these results demonstrated alterations in genes related to signal transduction, immune response, apoptosis, cellular proliferation, extracellular matrix, cell adhesion, and motility, and only in one gene related to reactive oxygen species processing. Oxidative stress was observed to be both dose- and time-dependent in cells exposed to LAA. GSH levels were transiently depleted following 2–8 hours exposure to the higher dose of LAA, with a gradual recovery up to 48 hours in LP9/TERT-1 cells (HKNM-2 not analyzed). These studies demonstrate that LAA exposure leads to increases in oxidative stress as measured by ROS production, gene expression, protein and functional changes in oxidative stress proteins (SOD), and GSH level alterations in human mesothelial cells.

The role of inflammasome activation and iron in the development of LAA-induced fibrosis was studied in [Shannahan et al. \(2012d\)](#). Male SH rats were instilled with a single exposure to 0 or 0.5 mg LAA, DEF, 21  $\mu\text{g}$   $\text{FeCl}_3$ , 0.5 mg LAA + 21  $\mu\text{g}$   $\text{FeCl}_3$ , or 0.5 mg LAA + 1 mg DEF. Tissues were collected 4 hours and 1 day postexposure. LAA instillation increased lung expression of the inflammasome-related molecules cathepsin B, Nalp3, NF- $\kappa\text{B}$ , apoptosis-associated speck-like protein containing a CARD (ASC), IL-1 $\beta$ , and IL-6 expression 4 hours postexposure. Lung tissue expression of inflammatory cytokines CCL-7, Cox-2, CCL-2, and CXCL-3 was increased 4 hours following LAA exposure. These data suggest that LAA exposures leads to the activation of the inflammasome cascade and cytokines downstream of the pathway in LAA-exposed animals. [Li et al. \(2012\)](#) also studied LAA inflammasome activation and demonstrated in vitro in THP-1 cells that LAA activated the NLRP3 inflammasome but to a lesser degree than chrysotile. However, this study showed that LAA exposure generated more ROS production as compared to chrysotile. Although not studied, the authors suggest that differences in fiber length and surface area may play a role in this differential inflammatory response.

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<sup>18</sup>Libby Amphibole asbestos samples were characterized for this study with analysis of chemical composition and mean surface area ([Meeker et al., 2003](#)). Doses were measured in surface area and described based on viability assays as either the nontoxic ( $15 \times 10^6 \mu\text{m}^2/\text{cm}^2$ ) or the toxic dose ( $75 \times 10^6 \mu\text{m}^2/\text{cm}^2$ ).



#### 4.4.2. Genotoxicity

Genotoxicity, and more specifically, mutagenicity, are associated with tumor formation through alterations in genetic material.<sup>19</sup> Mutagenicity refers to a permanent effect on the structure and/or amount of genetic material that can lead to heritable changes in function, while genotoxicity is a broader term including all adverse effects on the genetic information ([Eastmond et al., 2009](#)). Results of standard mutation assays like the Ames test, which analyze for point mutations, have found asbestos and other mineral fibers to be negative or only marginally positive ([Walker et al., 1992](#)). Several other studies, however, have shown that asbestos exposure can result in a variety of chromosomal alterations, which are briefly discussed below. Genotoxicity following exposure to asbestos fibers has been described as the result of two distinct mechanisms, either ROS production leading to DNA damage, or physical interference of mitosis by the fibers. For both DNA damage and mitotic interference, the fibers must first enter the cell. Some studies have shown that a direct interaction between fibers and cellular receptors might also lead to increased ROS production. ROS production is also related to surface iron on fibers, with increased surface iron leading to increased ROS production ([IARC, 2012](#)). ROS production is possibly a key event in fiber-induced direct DNA damage, as observed following exposure to other forms of asbestos, while the indirect DNA damage requires fiber interaction with cellular components (e.g., mitotic spindle, chromosomes).

ROS production and genotoxicity (micronuclei induction) following exposure to LAA has been demonstrated in x-ray repair cross complementing protein 1 (XRCC1)-deficient human lung epithelial H460 cells ([Pietruska et al., 2010](#)). XRCC1 is involved in the repair mechanisms for oxidative DNA damage, particularly single strand breaks. Micronuclei induction was measured following treatment of cells by controls (positive, hydrogen peroxide; negative, paclitaxel) and by 5 µg/cm<sup>2</sup> fibers or TiO<sub>2</sub> particles for 24 hours. Spontaneous micronuclei induction was increased in XRCC1-deficient cells in a dose-dependent manner following exposure to crocidolite and LAA as compared to unexposed cells. These results support a potential genotoxic effect of exposure to both crocidolite and LAA.

[Athanasίου et al. \(1992\)](#) performed a series of experiments to measure genotoxicity following exposure to tremolite, including the Ames mutagenicity assay, micronuclei induction, chromosomal aberrations, and gap-junction intercellular communication. Although a useful test system for mutagenicity screening for many agents, the Ames assay is not the most effective test

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<sup>19</sup>*Genotoxicity*: a broad term that refers to potentially harmful effects on genetic material, which may be mediated directly or indirectly, and which are not necessarily associated with mutagenicity. Thus, tests for genotoxicity include tests that provide an indication of induced damage to DNA (but not direct evidence of mutation) via effects such as unscheduled DNA synthesis, sister chromatid exchange, or mitotic recombination, as well as tests for mutagenicity; *Mutagenicity*: refers to the induction of permanent transmissible changes in the amount or structure of the genetic material of cells or organisms. These changes, “mutations,” may involve a single gene or gene segment, a block of genes, or whole chromosomes. Effects on whole chromosomes may be structural and/or numerical (as defined in the European Union Technical Guidance on Risk Assessment ([CEC, 1996](#))).

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1 to detect mutations induced by mineral fibers. Mineral fibers can cause mutation through  
2 generation of ROS or direct disruption of the spindle apparatus during chromatid segregation.  
3 Fibers do not induce ROS in the Ames system, however, and the *Salmonella typhimurium* strains  
4 do not endocytose the fibers. Only one study was found in the published literature that used the  
5 Ames assay to measure mutagenicity of tremolite. Metsovo tremolite asbestos has been shown  
6 to be the causative agent of endemic pleural calcification and an increased level of malignant  
7 pleural mesothelioma (see Section 4.1). To measure the mutagenicity of Metsovo tremolite,  
8 *S. typhimurium* strains (TA98, TA100, and TA102) were exposed to 0–500 µg/plate of asbestos  
9 ([Athanasίου et al., 1992](#)). Metsovo tremolite did not yield a statistically significant increase in  
10 revertants in the Ames assay, including in the TA102 *Salmonella* strain, which is generally  
11 sensitive to oxidative damage. This study demonstrated clastogenic effects of tremolite,  
12 including chromosomal aberrations and micronuclei induction. Tremolite exposure in Syrian  
13 hamster embryo (SHE) cells did lead to a dose-dependent increase in chromosome aberrations  
14 that was statistically significant at the highest doses tested ( $1.0\text{--}3.0\text{ }\mu\text{g}/\text{cm}^2$ ;  $p < 0.01$ ;  
15 [Athanasίου et al., 1992](#)). A statistically significant, dose-dependent increase in levels of  
16 micronuclei was demonstrated following tremolite exposure at concentrations as low as  
17  $0.5\text{ }\mu\text{g}/\text{cm}^2$  ( $p < 0.01$ ) in BPNi cells after 24-hour exposure. Literature searches did not find  
18 tremolite tested for clastogenicity in other cell types, but the results of this study suggest  
19 interference with the spindle apparatus by these fibers. No analysis was performed to determine  
20 whether fiber interference of the spindle apparatus could be observed, which would have  
21 supported these results. No effect on the gap-junctional intercellular communication following  
22 tremolite exposure was observed in both Chinese hamster lung fibroblasts (V79) and Syrian  
23 hamster embryo BPNi cells, which are sensitive to transformation ([Athanasίου et al., 1992](#)).  
24 [Okayasu et al. \(1999\)](#) analyzed the mutagenicity of Metsovo tremolite, erionite, and the  
25 man-made refractory ceramic fiber (RCF-1). Human-hamster hybrid A<sub>L</sub> cells contain a full set  
26 of hamster chromosomes and a single copy of human chromosome 11. Mutagenesis of the CD59  
27 locus on this chromosome is quantifiable by an antibody complement-mediated cytotoxicity  
28 assay. The study authors state that this is a highly sensitive mutagenicity assay, and previous  
29 studies have demonstrated mutagenicity of both crocidolite and chrysotile ([Hei et al., 1992](#)). The  
30 cytotoxicity analysis for mutagenicity was performed by exposing  $1 \times 10^5$  A<sub>L</sub> cells to a range of  
31 concentrations of fibers as measured by weight (0–400 µg/mL or 0–80 µg/cm<sup>2</sup>) for 24 hours at  
32 37°C. A dose-dependent increase in CD59 mutant induction was observed following exposure  
33 to erionite and tremolite, but not RCF-1.

34 In summary, one in vitro study examined genotoxicity of LAA by measuring DNA  
35 adduct formation following exposure via murine macrophages ([primary and immortalized; Blake  
36 et al., 2007](#)). The data showed no increase in adduct formation as compared to unexposed  
37 controls. A second study observed increases in micronuclei induction in both normal human  
38 lung epithelial cells and XRCC1-deficient cells for both LAA and crocidolite asbestos ([Pietruska](#)

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et al., 2010). Two studies of tremolite examined genotoxicity. The first found no significant increase in revertants in the Ames assay (Athanasίου et al., 1992), which is similar to results obtained for other forms of asbestos. This study did find, however, that tremolite exposure led to a dose-dependent increase in chromosome number and micronuclei formation, which has also been described for other asbestos fibers (as reviewed in Hei et al., 2006; and Jaurand, 1997). Hei and colleagues (Okayasu et al., 1999) performed mutation analysis with tremolite and found a dose-dependent increase in mutations in CD59 in hamster hybrid cells. Genotoxicity analysis in humans, following exposure to LAA or tremolite, has not been measured, although other types of asbestos fibers have led to increases in genotoxicity in primary cultures and lymphocytes (Dopp et al., 2005; Poser et al., 2004). In general, these studies have examined genotoxicity with a focus on ROS production as a key event. Although LAA- and tremolite-specific data are limited to in vitro studies, given the similarities in response to other forms of asbestos, there is some evidence to suggest genotoxicity following exposure to Libby Amphibole and tremolite asbestos. However, the potential role of this genotoxicity in lung cancer or mesothelioma following exposure to LAA is unknown.

#### 4.4.3. Cytotoxicity and Cellular Proliferation

The initial stages of tumorigenicity may be an increased cellular proliferation at the site of fiber deposition, which can increase the chance of cancer by increasing the population of spontaneous mutations, thereby affording genotoxic effects an opportunity to multiply. Increased cell proliferative regeneration may be associated with tumor clonal expansion and can occur in response to increased apoptosis. In macrophages, increased cytotoxicity leads to an increased oxidant release, which in turn may lead to increased cell damage, signaling activation and inflammatory cell recruitment.

Wagner et al. (1982) examined the in vitro cytotoxicity of three forms of tremolite used in their in vivo studies. LDH and  $\beta$ -glucuronidase were measured in the medium following incubation of unactivated primary murine macrophages to 50, 100, and 150  $\mu\text{g/mL}$  of each sample for 18 hours. The Korean tremolite (Sample C) produced results similar to the positive control: increased toxicity of primary murine macrophages, increased cytotoxicity of Chinese hamster ovary cells, and increased formation of cells  $>25\text{ }\mu\text{m}$  diameter from the A549 cell line. The tremolite sample from Greenland (Sample B) did result in increased toxicity over controls; although to a lesser degree (statistics are not given). Although differential toxicity of these samples was noted on a mass basis, data were not normalized for fiber content or size. The inference is that differential results may be due, at least in part, to differential fiber counts.

Wylie et al. (1997) examined the mineralogical features associated with cytotoxic and proliferative effects of asbestos in hamster tracheal epithelial (HTE) and rat pleural mesothelial (RPM) cells with a colony-forming efficiency assay. HTE cells are used because they give rise to tracheobronchial carcinoma, while RPM cells give rise to mesotheliomas. The results of the

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analysis with fiber exposure by mass ( $\mu\text{g}/\text{cm}^2$ ) show elevated colonies in HTE cells following exposures to both asbestos fibers ( $p < 0.05$ ) at the lowest concentrations, while significant decreases were observed for both asbestos fibers at the higher concentrations ([0.5  \$\mu\text{g}/\text{cm}^2\$ ,  \$p < 0.05\$ ; Wylie et al., 1997](#)). No proliferation was observed for either chrysotile or crocidolite asbestos fibers in RPM cells, but cytotoxicity was observed at concentrations greater than  $0.05 \mu\text{g}/\text{cm}^2$  ( $p < 0.05$ ). All talc samples were less cytotoxic in both cell types. Analyzing the data for cytotoxicity and proliferation based on the exposure measurement demonstrated differences in response depending solely on how the fibers were measured: by mass, number, or surface area. These results show variability in interpreting the results of the same assay based on the defined unit of exposure. Most early studies used mass as the measurement for exposure, which can impact how the results are interpreted. When possible, further analysis of fiber number and surface area would help elucidate the role of these metrics, particularly for in vivo studies.

Tremolite and LAA exposure led to increases in both fibrosis and tumorigenicity in all but one animal study, supporting a possible role for proliferation in response to these fibers (see Tables 4-19 and 4-20). However, there are limited data to demonstrate that increased cytotoxicity and cellular proliferation following exposure to LAA leads to lung cancer or mesothelioma.

#### *Summary*

The review of these studies clearly highlights the need for more controlled studies examining LAA in comparison with other forms of asbestos and for examining multiple endpoints—including ROS production, DNA damage, inflammasome activation, and proinflammatory gene expression alterations—to improve understanding of mechanisms involved in cancer and other health effects. Data gaps still remain to determine specific mechanisms involved in LAA-induced disease. Studies that examined cellular response to tremolite also found that tremolite exposure may lead to increased ROS production, toxicity, and genotoxicity ([Okayasu et al., 1999](#); [Wagner et al., 1982](#)). As with the in vivo studies, the definition of fibers and how the exposures were measured varies among studies.

### **4.5. SYNTHESIS OF MAJOR NONCANCER EFFECTS**

The predominant noncancer health effects observed following inhalation exposure to LAA are effects on the lungs and pleural lining surrounding the lungs. These effects have been observed primarily in studies of exposed workers and community members, and are supported by laboratory animal studies. Recent studies have also examined other noncancer health effects following exposure to Libby Amphibole, including autoimmune effects and cardiovascular disease; this research base is currently not as well developed as that of respiratory noncancer effects. Adequate data are not available to differentiate the health effects of the predominant

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mineralogical forms composing LAA. Although the adverse effects of tremolite are reported in the literature, the contribution of winchite and richterite to the aggregate effects of LAA has not been determined.

#### **4.5.1. Pulmonary Effects**

##### **4.5.1.1. Pulmonary Fibrosis (Asbestosis)**

Asbestosis is the interstitial pneumonitis (inflammation of lung tissue) and fibrosis caused by inhalation of asbestos fibers and is characterized by a diffuse increase of collagen in the alveolar walls (fibrosis) and the presence of asbestos fibers, either free or coated with a proteinaceous material and iron (asbestos bodies). Fibrosis results from a sequence of events following lung injury, which includes inflammatory cell migration, edema, cellular proliferation, and accumulation of collagen. Asbestosis is associated with dyspnea (shortness of breath), bibasilar rales, and changes in pulmonary function: a restrictive pattern, mixed restrictive-obstructive pattern, and/or decreased diffusing capacity ([ATS, 2004](#)). Radiographic evidence of small opacities in the lung is direct evidence of scarring of the lung tissue and is the fibrotic scarring of lung tissue consistent with mineral dust and mineral fiber toxicity. The scarring of the parenchymal tissue of the lung contributes to measured changes in pulmonary function, including obstructive pulmonary deficits from narrowing airways, restrictive pulmonary deficits from impacting the elasticity of the lung, as well as decrements in gas exchange.

Workers exposed to LAA from vermiculite mining and processing facilities in Libby, MT, as well as plant workers in Marysville, OH, where vermiculite ore was exfoliated and processed, have an increased prevalence of small opacities on chest x-rays, which is indicative of fibrotic damage to the parenchymal tissue of the lung ([Rohs et al., 2008](#); [Amandus et al., 1987a](#); [McDonald et al., 1986b](#); [Lockey et al., 1984](#)). Significant increases in asbestosis as a cause of death have been documented in studies of the Libby worker cohort report (see Table 4-6 for details; [Larson et al., 2010b](#); [Sullivan, 2007](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). For both asbestosis mortality and radiographic signs of asbestos (small opacities), positive exposure-response relationships are described where these effects are greater with greater cumulative exposure to LAA.

Deficits in pulmonary function consistent with pulmonary fibrosis have been reported in individuals exposed to LAA in community-based studies.<sup>20</sup> Data from the ATSDR community screening, which included workers, provide support for functional effects from parenchymal changes. The original report of the health screening data indicated moderate to severe

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<sup>20</sup>The initial study of the Marysville, OH cohort measured but reported no change in pulmonary function ([Lockey et al., 1984](#)). Pulmonary function was not reported for the cohort follow-up, although prevalence of pleural and parenchymal abnormalities was increased ([Rohs et al., 2008](#)). The initial studies of the occupational Libby worker cohort do not include assessment of pulmonary function ([Amandus et al., 1987a](#); [McDonald et al., 1986b](#)).



pulmonary restriction in 2.2% of men ([Peipins et al., 2003](#); [ATSDR, 2001b](#)). A recent reanalysis of these data show that for study participants with small opacities viewed on the radiographs (grade 1/0 or greater) and DPT, the mean FVC is reduced to 78.76 ( $\pm 3.64$ ), 82.16 ( $\pm 3.34$ ), respectively of the expected value ([Weill et al., 2011](#)). A mean FVC of 95.63 ( $\pm 0.76$ ) was reported for those with other pleural abnormalities versus 103.15 ( $\pm 0.25$ ) in participants with no radiographic abnormalities. The strongest effects of diffuse pleural thickening and/or costophrenic angle obliteration on FVC were seen among men who had never smoked ( $-23.77$ ,  $p < 0.05$ ), with smaller effects seen among men who had smoked ( $-9.77$ ,  $p < 0.05$ ) and women who had smoked ( $-6.73$ ,  $p < 0.05$ ). Laboratory animal and mechanistic studies of LAA are consistent with the noncancer health effects observed in both Libby workers and community members. Pleural fibrosis was increased in hamsters after intrapleural injections of LAA ([Smith, 1978](#)). More recent studies have demonstrated increased collagen deposition consistent with fibrosis following intratracheal instillation of LAA fibers in mice and rats ([Padilla-Carlin et al., 2011](#); [Shannahan et al., 2011a](#); [Shannahan et al., 2011b](#); [Smartt et al., 2010](#); [Putnam et al., 2008](#)). Pulmonary fibrosis, inflammation, and granulomas were observed after tremolite inhalation exposure in Wistar rats ([Bernstein et al., 2005](#); [Bernstein et al., 2003](#)) and intratracheal instillation in albino Swiss mice ([Sahu et al., 1975](#)). [Davis et al. \(1985\)](#) also reported pulmonary effects after inhalation exposure in Wistar rats, including increases in peribronchiolar fibrosis, alveolar wall thickening, and interstitial fibrosis.

#### **4.5.1.2. Other Nonmalignant Respiratory Diseases**

Mortality studies of the Libby workers indicate increased mortality not only from asbestosis, but also from other respiratory diseases. Deaths attributed to chronic obstructive respiratory disease and deaths attributed to “other” nonmalignant respiratory disease were elevated more than twofold (see Table 4-6; [Larson et al., 2010b](#); [Sullivan, 2007](#)). These diseases are consistent with asbestos toxicity, and the evidence of a positive exposure-response relationship for mortality from all nonmalignant respiratory diseases, supports this association.

#### **4.5.2. Pleural Effects**

Pleural thickening caused by mineral fiber exposure includes two distinct biological lesions: discrete pleural plaques in the parietal (outer) pleura and diffuse pleural thickening of the visceral (inner) pleura. Both of these forms of pleural thickening can be viewed on standard radiographs, however smaller lesions may not be detected (high resolution computed tomography is a method that can detect smaller lesions than are visible on x-rays). Current ILO guidelines (2002) state that pleural thickening on x-ray can be Localized (LPT) or Diffuse (DPT). LPT was not defined by the ILO until the 2000 guidelines were published ([ILO, 2002](#)). Previously, the 1980 ILO guidelines defined only circumscribed pleural thickening (plaques) and

diffuse pleural thickening (DPT) either with or without costophrenic angle obliteration. The 2000 ILO revision defines LPT as the union of what was previously defined as plaques found on the chest wall or in other locations (e.g., diaphragm) in the 1980 guidelines, and what was previously defined as DPT without costophrenic angle obliteration. Neither classification for pleural thickening (LPT or DPT) in the 2000 ILO guidelines corresponds with the previous ILO classification systems for pleural thickening; LPT is defined more broadly than the previous category of pleural plaques, while DPT is defined more narrowly due to the requirement for costophrenic angle obliteration. Different researchers have used different terminology for circumscribed pleural thickening or plaques when implementing the 1980 ILO guidelines, most often using the terms “pleural plaques.”

Data from the ATSDR community health screening study indicate that the prevalence of pleural abnormalities, identified by radiographic examination, increases substantially with increasing number of exposure pathways ([Peipins et al., 2003](#)). A reanalysis of these data also considered age, smoking history, and types of exposures ([Weill et al., 2011](#)). Increased pleural thickening is reported for Libby workers, those with other vermiculite work, and those who had worked in other jobs with dust exposures (in locations other than Libby, MT). The prevalence of pleural plaques increased with age; in the 61–90 age group the prevalence was 38.3 and 12.7%, respectively among those exposed only through household contacts and those exposed through environmental exposure pathways. The community-based study in Minneapolis, MN also provides evidence of increased risk of pleural abnormalities among residents surrounding an exfoliation plant, with positive associations seen with measures of background and of intermittent (activity-based) exposures ([Alexander et al., 2012](#)).

Increased pleural thickening (including LPT) is reported for both of the studied worker cohorts, with evidence of positive exposure-response relationships ([Larson et al., 2012a](#); [Larson et al., 2010a](#); [Rohs et al., 2008](#); [Amandus et al., 1987a](#); [McDonald et al., 1986b](#); [Lockey et al., 1984](#)). Both [McDonald et al. \(1986b\)](#) and [Amandus et al. \(1987a\)](#) indicate that age is also a predictor of pleural thickening in exposed individuals, which may reflect the effects of time from first exposure. Smoking data were limited on the Libby workers and analyses do not indicate clear relationships between smoking and pleural thickening ([Amandus et al., 1987a](#); [McDonald et al., 1986b](#)). Pleural thickening in workers at the Scott Plant (Marysville, OH) was associated with hire on or before 1973 and age at time of interview but was not associated with BMI or smoking history ([ever smoked](#); [Rohs et al., 2008](#)).

#### **4.5.3. Other Noncancer Health Effects (Cardiovascular Toxicity, Autoimmune Effects)**

There is limited research available on noncancer health effects occurring outside the respiratory system and pleura. [Larson et al. \(2010b\)](#) examined cardiovascular disease-related mortality in the cohort of exposed workers from Libby (see Section 4.1). Mechanistic studies have examined the potential role of iron and the associated inflammation for both respiratory and

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cardiovascular disease ([Shannahan et al., 2012a](#); [Shannahan et al., 2012c](#); [Shannahan et al., 2012b](#); [Shannahan et al., 2012d](#); [Shannahan et al., 2011b](#)). Other studies examined the association between asbestos exposure and autoimmune disease ([Noonan et al., 2006](#)) or autoantibodies and other immune markers ([Pfau et al., 2005](#); see Table 4-16). However, limitations in the number, scope, and design of these studies make it difficult to reach conclusions as to the role of asbestos exposure in either cardiovascular disease or autoimmune disease.

#### **4.5.4. Libby Amphibole Asbestos Summary of Noncancer Health Effects**

The studies in humans summarized in Section 4.1 have documented an increase in mortality from nonmalignant respiratory disease, including asbestosis, in workers exposed to LAA ([Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus and Wheeler, 1987](#)). Additional studies have documented an increase in radiographic changes in the pleura (pleural thickening) and parenchyma among employees of a manufacturing facility in Marysville, OH that used vermiculite ore contaminated with LAA ([Rohs et al., 2008](#); [Lockey et al., 1984](#)). Radiographic evidence of pleural thickening and interstitial damage (small opacities) are also well documented among employees of the Libby vermiculite mining operations ([Larson et al., 2012a](#); [Larson et al., 2010a](#); [Amandus et al., 1987a](#); [McDonald et al., 1986b](#)). Positive exposure-response relationships for these health effects for both occupational cohorts studied, as well as the observed latency, support an association between exposure to LAA and these pleural and/or pulmonary effects. Studies of community members exposed to LAA have documented similar pleural abnormalities and pulmonary deficits consistent with parenchymal damage ([Weill et al., 2011](#); [Whitehouse, 2004](#); [Peipins et al., 2003](#)). Although limited, animal studies support the toxicity of LAA to pleural and pulmonary tissues. Developing research supports a role of inflammatory processes in the toxic action of LAA, consistent with the observed health effects ([Cyphert et al., 2012b](#); [Shannahan et al., 2012c](#); [Shannahan et al., 2011b](#); [Duncan et al., 2010](#); [Hamilton et al., 2004](#)). Taken together, the strong evidence in human studies, defined exposure-response relationships, and supportive animal studies provide compelling evidence that exposure to LAA causes nonmalignant respiratory disease, including asbestosis, pleural thickening, and deficits in pulmonary function associated with mineral fiber exposures. Existing data regarding cardiovascular effects and the potential for autoimmune disease are limited.

#### **4.5.5. Mode-of-Action Information (Noncancer)**

The precise mechanisms causing toxic injury from inhalation exposure to LAA have not been established. However, nearly all durable mineral fibers with dimensional characteristics that allow penetration to the terminal bronchioles and alveoli of the lung have the capacity to induce pathologic response in the lung and pleural cavity ([ATSDR, 2001a](#); [Witschi and Last,](#)



1996). The physicochemical attributes of mineral fibers are important in determining the type of toxicity observed. Fiber dimension (width and length), density, and other characteristics, such as chemical composition, surface area, solubility in physiological fluids, and durability, all play important roles in both the type of toxicity observed and the biologically significant dose. As described in Section 3, these characteristics also play a role in fiber dosimetry. Fibrosis results from a sequence of events following lung injury, which includes inflammatory cell migration, edema, cellular proliferation, and accumulation of collagen. Fibers do migrate to the pleural space, and it has been hypothesized that a similar cascade of inflammatory events may contribute to fibrotic lesions in the visceral pleura. Thickening of the visceral pleura is more often localized to lobes of the lung with pronounced parenchymal changes, and it has also been hypothesized that the inflammatory and fibrogenic processes within the lung parenchyma in response to asbestos fibers may influence the fibrogenic process in the visceral pleura. The etiology of parietal plaques is largely unknown with respect to mineral fiber exposure.

There is currently insufficient evidence to establish the noncancer MOA for LAA. Limited in vitro studies have demonstrated oxidative stress following LAA exposures in various cell types (Duncan et al., 2010; Hillegass et al., 2010; Pietruska et al., 2010; Blake et al., 2007). LAA fibers increased intracellular ROS in both murine macrophages and human epithelial cells (Duncan et al., 2010; Blake et al., 2007). Surface iron and inflammatory marker gene expression was increased following exposure to LAA in human epithelial cells (Shannahan et al., 2012a; Shannahan et al., 2012c; Shannahan et al., 2012b; Shannahan et al., 2012d; Shannahan et al., 2011b; Duncan et al., 2010; Pietruska et al., 2010; see Table 4-18). Tremolite studies demonstrate cytotoxicity in various cell culture systems (see Table 4-22).

The initial stages of any fibrotic response involve cellular proliferation, which may be compensatory for cell death due to cytotoxicity. Analysis of cellular proliferation has demonstrated both increases and decreases following exposure to asbestos fibers in vitro and in vivo depending on the specific fiber or cell type (Mossman et al., 1985; Topping and Nettesheim, 1980). Other studies have focused on the activation of cell-signaling pathways that lead to cellular proliferation following exposure to asbestos (Scapoli et al., 2004; Shukla et al., 2003; Ding et al., 1999; Zanella et al., 1996).

Although slightly increased compared to controls, cytotoxicity in murine macrophage cells exposed to LAA was decreased compared to other fiber types (Blake et al., 2008). Cytotoxicity was slightly, but statistically significantly, increased compared to an unexposed control at 24 hours postexposure to LAA, while crocidolite exposure resulted in even higher levels of cytotoxicity. No other in vitro study examined cytotoxicity following exposure to LAA, although an increase in apoptosis was demonstrated in this same cell system (Blake et al., 2008). Recent studies in mice exposed to LAA demonstrated increased collagen deposition and collagen gene expression, markers of fibrosis (Smartt et al., 2010; Putnam et al., 2008). Short-term-duration studies in rats also demonstrated an increased inflammatory response

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([Padilla-Carlin et al., 2011](#); [Shannahan et al., 2011a](#); [Shannahan et al., 2011b](#)). Tremolite or LAA exposure both led to increases in fibrosis in all but one animal study, supporting a role for proliferation in response to these fibers. Taken together with studies on other asbestos fibers, these data suggest that cytotoxicity and cell proliferation may play a role in the noncancer health effects following exposure to LAA.

Although continued research demonstrates that the LAA has biologic activity consistent with the inflammatory action and cytotoxic effects seen with other forms of asbestos, the data are not sufficient to establish a MOA for the pleural and/or pulmonary effects of exposure to LAA.

## 4.6. EVALUATION OF CARCINOGENICITY

### 4.6.1. Summary of Overall Weight of Evidence

Under the EPA *Guidelines for Carcinogen Risk Assessment* ([U.S. EPA, 2005a](#)), LAA is *carcinogenic to humans* following inhalation exposure based on epidemiologic evidence that shows a convincing association between exposure to LAA fibers and increased lung cancer and mesothelioma mortality ([Larson et al., 2010b](#); [Moolgavkar et al., 2010](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). These results are further supported by animal studies that demonstrate the carcinogenic potential of LAA fibers and tremolite fibers in rodent bioassays (see Section 4.1, 4.2, Appendix D). As a durable mineral fiber of respirable size, this conclusion is consistent with the extensive published literature that documents the carcinogenicity of amphibole fibers (as reviewed in [Aust et al., 2011](#); [Broaddus et al., 2011](#); [Bunderson-Schelvan et al., 2011](#); [Huang et al., 2011](#); [Mossman et al., 2011](#)).

EPA's *Guidelines for Carcinogenic Risk Assessment* ([U.S. EPA, 2005a](#)) indicate that for tumors occurring at a site other than the initial point of contact, the weight of evidence for carcinogenic potential may apply to all routes of exposure that have not been adequately tested at sufficient doses. An exception occurs when there is convincing information (e.g., toxicokinetic data) that absorption does not occur by other routes. Information on the carcinogenic effects of LAA via the oral and dermal routes in humans or animals is absent. The increased risk of lung cancer and mesothelioma following inhalation exposure to LAA has been established by studies in humans, but these studies do not provide a basis for determining the risk from other routes of exposure. Mesothelioma occurs in the pleural and peritoneal cavities and, therefore, is not considered a portal-of-entry effect. However, the role of indirect or direct interaction of asbestos fibers in disease at these extrapulmonary sites is still unknown. There is no information on the translocation of LAA to extrapulmonary tissues following either oral or dermal exposure, and limited studies have examined the role of these routes of exposure in cancer. Therefore, LAA is considered *carcinogenic to humans* by the inhalation route of exposure.

#### 4.6.1.1. *Synthesis of Human, Animal, and Other Supporting Evidence*

Libby, MT workers have been the subject of multiple mortality studies demonstrating increased cancer mortality in relation to estimated fiber exposure. Occupational studies conducted in the 1980s ([Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)) as well as the extended follow-up studies published in more recent years ([Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#)) and additional analyses of the extended follow-up ([Moolgavkar et al., 2010](#)) provide evidence of an increased risk of lung cancer mortality and of mesothelioma mortality among the workers exposed to LAA in the Libby vermiculite mining and processing operations. This pattern is seen in the lung cancer analyses using an internal referent group in the larger follow-up studies ([Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#)), with cumulative exposure analyzed using quartiles or as a continuous measure, and in the studies reporting analyses using an external referent group (i.e., standardized mortality ratios; [Sullivan, 2007](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). [McDonald et al. \(2004\)](#) also reported increasing risk of mesothelioma across categories of exposure; the more limited number of cases available in earlier studies precluded this type of exposure-response analysis. This association is also supported by the case series of 11 mesothelioma patients among residents in or around Libby, MT, and among family members of workers in the mining operations ([Whitehouse et al., 2008](#)), and by the observation of three cases of mesothelioma (two of which resulted in death) in the Marysville, OH worker cohort identified as of June 2011 ([Dunning et al., 2012](#)).

Although experimental data in animals and data on toxicity mechanisms are limited for LAA, tumors were observed in tissues similar to those in humans (e.g., mesotheliomas, lung cancer) indicating the existing data are consistent with the cancer effects observed in humans exposed to LAA. [Smith \(1978\)](#) reported increased incidence of mesotheliomas in hamsters after intrapleural injections of LAA. Additionally, studies in laboratory animals (rats and hamsters) exposed to tremolite via inhalation ([Bernstein et al., 2005](#); [Bernstein et al., 2003](#); [Davis et al., 1985](#)), intrapleural injection ([Roller et al., 1997, 1996](#); [Davis et al., 1991](#); [Wagner et al., 1982](#); [Smith et al., 1979](#)), or implantation ([Stanton et al., 1981](#)) have shown increases in mesotheliomas and lung cancers. Tremolite from various sources was used and varied in fiber content and potency (see Section 4.2, Appendix D). The most sensitive model for mesothelioma induction is the Syrian golden hamster following asbestos inhalation, with different susceptibility between species attributed to more rapid translocation to the pleural space ([Donaldson et al., 2010](#)). Although [McConnell et al. \(1983a\)](#) observed no increase in carcinogenicity following oral exposure to nonfibrous tremolite, the ability of this study to inform the carcinogenic potential of fibrous tremolite through inhalation is unclear, and these study results contribute little weight to the evaluation of the carcinogenicity of fibrous LAA.

The available mechanistic information suggests LAA induces effects that may play a role in carcinogenicity (see Section 4.2 Appendix D). Several in vitro studies have demonstrated

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oxidative stress and genotoxicity following LAA exposures in various cell types ([Duncan et al., 2010](#); [Hillegass et al., 2010](#); [Pietruska et al., 2010](#); [Blake et al., 2007](#)). LAA increased intracellular ROS in both murine macrophages and human epithelial cells ([Duncan et al., 2010](#); [Blake et al., 2007](#)). Additionally, surface iron, inflammatory marker gene expression, and aneugenic micronuclei were increased following exposure to LAA in human epithelial cells ([Duncan et al., 2010](#); [Pietruska et al., 2010](#)). Tremolite studies demonstrate cytotoxic and clastogenic effects (e.g., micronucleus induction and chromosomal aberrations) of the fibers in various cell culture systems.

In summary, the epidemiologic data demonstrate an association between exposure to LAA and increased cancer risk. Supporting evidence of carcinogenic potential was observed in the limited number of laboratory animal studies exposed to LAA or tremolite (see Tables 4-15 and 4-16 summarizing in vivo studies). Overall, the available evidence supports the conclusion that LAA is *carcinogenic to humans*.

#### **4.6.2. Mode-of-Action Information (Cancer)**

##### **4.6.2.1. Description of the Mode-of-Action Information**

EPA guidance provides a framework for analyzing the potential mode(s) of action by which physical, chemical, and biological information is evaluated to identify key events in an agent's carcinogenicity ([U.S. EPA, 2005a](#)). Agents can work through more than one MOA, and MOA can differ for various endpoints (e.g., lung cancer vs. mesothelioma). Reasonably, the analysis of a MOA would start with some knowledge of an agent's biological activity that leads to cellular transformation resulting in carcinogenicity. Although early steps in the process often can be identified, carcinogenicity is a complex process resulting from multiple changes in cell function. Due to the limited data available specific to LAA, the MOA of LAA for lung cancer and mesothelioma following inhalation exposure cannot be established.

Occupational studies demonstrate human health effects (e.g., lung cancer, mesothelioma) following exposure to LAA. Although the limited mechanistic data demonstrate biological effects similar to those of other mineral fibers following exposure to LAA, the existing literature is insufficient to establish a MOA for LAA for lung cancer or mesothelioma. These biological effects following exposure to LAA and/or tremolite are demonstrated in a limited number of laboratory animal and in vitro studies. Multiple key events for one particular MOA have not been identified; therefore, the MOA for LAA carcinogenicity cannot be established. However, multiple mechanisms of action (e.g., mutagenicity, chronic inflammation, cytotoxicity, and regenerative proliferation) can be hypothesized based on the available asbestos literature. These are described in Section 4.4, and discussed below.

#### 4.6.2.2. Evidence Supporting a Mutagenic Mode of Action

##### *Strength, consistency and specificity of the association*

Only limited genotoxicity analysis following exposure to LAA or tremolite has been reported, although studies of other types of asbestos fibers have shown increases in genotoxicity both in vitro and in vivo ([reviewed in Huang et al., 2011](#)). One in vitro study examined genotoxicity of LAA by measuring DNA adduct formation following exposure via murine macrophages ([primary and immortalized; Blake et al., 2007](#)). The data showed no increase in adduct formation as compared to unexposed controls. A second study observed increases in micronuclei induction in both normal human lung epithelial cells and XRCC1-deficient cells for both Libby Amphibole and crocidolite asbestos ([Pietruska et al., 2010](#)). Two studies of tremolite examined genotoxicity. The first found no significant increase in revertants in the Ames assay ([Athanasίου et al., 1992](#)), which is similar to results obtained for other forms of asbestos. This study did find, however, that tremolite exposure led to a dose-dependent increase in chromosome number and micronuclei formation, which has also been described for other asbestos fibers (as reviewed in [Hei et al., 2006](#); [Jaurand, 1997](#)). Hei and colleagues ([Okayasu et al., 1999](#)) performed mutation analysis with tremolite and found a dose-dependent increase in mutations in CD59 in hamster hybrid cells. Although LAA- and tremolite-specific data are limited to in vitro studies, given the similarities in response to other forms of asbestos, there is some evidence to suggest genotoxicity following exposure to Libby Amphibole and tremolite asbestos.

##### *Dose-response concordance and temporal relationship*

A dose-response concordance has not been established between the development of genotoxicity and exposure to LAA or other amphibole asbestos. Genotoxicity studies of other amphibole asbestos have examined gene mutations and chromosomal mutations, as well as DNA damage resulting from ROS production following exposure. As recently reviewed by [Huang et al. \(2011\)](#), there are a large number of in vitro studies that support possible genotoxic mechanisms following exposure to fibers. There are fewer in vivo studies of the genotoxicity of amphibole asbestos, and a very limited number of these were following inhalation exposure. Some of these studies were performed in nonrelevant cell types for inhalation endpoints, and some also were performed at doses higher than observed in environmental or occupational asbestos exposures. Temporal relationship would be impacted by direct or indirect genotoxic mechanism playing a role in asbestos-induced tumorigenesis. There is insufficient data to conclude whether the observed genotoxic effects following exposure to amphibole asbestos result from direct (e.g., spindle interference) or indirect (e.g., reactive oxygen species production) mechanisms. The available evidence suggests a role for both direct and indirect genotoxicity, but requires further research ([Huang et al., 2011](#)). Therefore, although these results suggest a possible role for genotoxicity in the MOA of LAA, dose-response concordance and a temporal relationship are difficult to determine.

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#### 1 *Biological plausibility and coherence*

2 Although only limited genotoxicity studies of LAA and tremolite have been published,  
3 these studies are supported by similar results for other amphibole asbestos studies that  
4 demonstrate genotoxicity in both in vitro and in vivo studies ([reviewed in Huang et al., 2011](#)).  
5 Although studies of asbestos genotoxicity need to be carefully reviewed to determine relevance  
6 of routes of exposure, target cell, and dose, taking these parameters into account, this review of  
7 these studies supports the biological plausibility of the genotoxicity of LAA.

#### 9 **4.6.2.3. Evidence Supporting Mechanisms of Action of Chronic Inflammation, Cytotoxicity,** 10 **and Cellular Proliferation**

11 Chronic inflammation has been observed following fiber exposure, which is often  
12 followed by fibrosis at the site of inflammation if the fibers persist ([reviewed in Mossman et al.,](#)  
13 [2011](#)). Macrophages phagocytose fibers and particulate matter and are activated to trigger the  
14 release of inflammatory cytokines, ROS, and growth factors. These responses lead to a sustained  
15 inflammatory response that can result in fibrosis at the site of fiber deposition. Chronic  
16 inflammation is hypothesized to contribute to a carcinogenic response through the production of  
17 ROS and increased cellular proliferation ([Hanahan and Weinberg, 2011](#)).

18 The initial stages of any fibrotic response involve cellular proliferation, which may be  
19 compensatory for cell death due to cytotoxicity. The same may be true for tumorigenicity, as  
20 increased cell proliferation can increase the chance of cancer by increasing the population of  
21 spontaneous mutations affording genotoxic effects an opportunity to multiply. Analysis of  
22 cellular proliferation of epithelial cells has demonstrated both increases and decreases following  
23 exposure to asbestos fibers in vitro and in vivo ([Mossman et al., 1985](#); [Topping and Nettesheim,](#)  
24 [1980](#)). Other studies have focused on the activation of cell-signaling pathways that lead to  
25 cellular proliferation following exposure to asbestos ([Scapoli et al., 2004](#); [Shukla et al., 2003](#);  
26 [Ding et al., 1999](#); [Zanella et al., 1996](#)).

27 The inflammatory response to fibers in vivo has been studied following inhalation  
28 exposure to many types of fibers but not for LAA (reviewed in [Broadus et al., 2011](#); [Mossman](#)  
29 [et al., 2011](#); [Mossman et al., 2007](#)). Results following inhalation exposure to tremolite have  
30 demonstrated increased inflammatory response as early as 1 day postexposure ([Bernstein et al.,](#)  
31 [2005](#); [Bernstein et al., 2003](#)). Earlier data from [Davis et al. \(1985\)](#) following inhalation exposure  
32 to other forms of tremolite showed increased fibrosis and carcinogenesis; however, inflammatory  
33 response was not described. In vivo studies of LAA and tremolite through other routes of  
34 exposure have demonstrated increased inflammation following exposure ([Padilla-Carlin et al.,](#)  
35 [2011](#); [Shannahan et al., 2011a](#); [Shannahan et al., 2011b](#)). Inhalation studies examining other  
36 types of asbestos (crocidolite, chrysotile, and amosite) have clearly demonstrated an increase in  
37 chronic inflammation and respiratory cancer related to exposure ([reviewed Mossman et al.,](#)  
38 [2011](#)). This effect is observed in animal studies for LAA and tremolite and is relevant to humans

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1 based on similar responses in cohorts analyzed ([Musk et al., 2008](#); [Hein et al., 2007](#); [Levin et al.,](#)  
2 [1998](#)).

3 Although limited, the data described here for LAA, and to a greater extent for tremolite,  
4 suggest a similar response as to other amphibole asbestos. In vivo exposure to tremolite led to an  
5 increase in inflammation for all studies where it was measured. This increase appeared in some  
6 cases to depend on fiber size and morphology ([Davis et al., 1991](#); [Smith et al., 1979](#)). In vitro  
7 analysis of LAA showed increases in inflammatory cytokines ([Hamilton et al., 2004](#)) and in  
8 proinflammatory gene expression ([Duncan et al., 2010](#)). Bernstein et al. (2005; 2003) observed  
9 that exposure to tremolite led to pronounced inflammation as soon as 1 day after inhalation  
10 exposure in male Wistar rats. Inflammation also occurred in male albino Swiss mice in an  
11 acute-duration study that did not lead to fibrosis or carcinogenesis, possibly due to the short  
12 study duration (150 days; [Sahu et al., 1975](#)).

13 Chronic inflammation has also been associated with increased ROS production (reviewed  
14 in [Aust et al., 2011](#); [Kamp et al., 1992](#)). Fibers can directly lead to the production of ROS by  
15 iron-catalyzed generation through the Fenton reaction. ROS are also produced following  
16 phagocytosis of fibers. ROS production following exposure to asbestos has been shown to be  
17 associated with DNA damage (described below), chronic inflammation, and lipid peroxidation.  
18 As described in the previous section, chronic inflammation may lead to increased cell  
19 proliferation and DNA damage, which in turn may lead to tumor formation. The hydroxyl  
20 radical produced has been shown to directly interact with DNA ([Leanderson et al., 1988](#)).

21 ROS production has been measured in response to both LAA and tremolite exposure.  
22 The study of LAA ([Blake et al., 2007](#)) demonstrated an increase in superoxide anions, not  
23 hydrogen peroxide, as has been demonstrated with crocidolite. [Blake et al. \(2007\)](#) also  
24 demonstrated that total SOD was inhibited following exposure to LAA, along with a decrease in  
25 intracellular glutathione. These results are supported by a recent study in human mesothelial  
26 cells ([Hillegass et al., 2010](#)). Further, increased ROS production was also observed in human  
27 airway epithelial cells following exposure to LAA ([Duncan et al., 2010](#)). This increase in ROS  
28 and decrease in glutathione are common effects following exposure to asbestos fibers and  
29 particulate matter. Limited studies, however, have examined the specific type of ROS produced  
30 following exposure to each type of asbestos.

31 A dose-response concordance has not been established between the development of  
32 chronic inflammation and exposure to LAA. Dose-response information is limited to inhalation  
33 studies of other amphibole asbestos, which were recently reviewed ([Case et al., 2011](#); [Mossman](#)  
34 [et al., 2011](#)). Many of the early studies of amphiboles described above were performed using  
35 only one dose. Therefore, while these studies demonstrate an exposure-response relationship  
36 between amphibole asbestos and chronic inflammation, dose-response concordance cannot be  
37 determined.

1 A temporal relationship has not been established between the development of chronic  
2 inflammation and inhalation exposure to LAA. Chronic inhalation studies of tremolite  
3 demonstrate an increase in chronic inflammation over time ([Bernstein et al., 2005](#); [Bernstein et](#)  
4 [al., 2003](#)) that may lead to fibrosis, or possibly tumor formation. A similar pattern has also been  
5 observed in inhalation studies with other amphibole asbestos ([as reviewed in Mossman et al.,](#)  
6 [2011](#)).

7 Chronic inflammation following exposure to fibers has been associated with the  
8 development of both malignant and nonmalignant lung and pleural diseases ([Bringardner et al.,](#)  
9 [2008](#); [Mossman and Churg, 1998](#)). In vivo and in vitro studies have shown increases in  
10 inflammation and inflammatory markers following exposure to LAA and tremolite up to 1 month  
11 following single intratracheal instillation exposures in animal models (see Section 4.2,  
12 Appendix D). Although inhalation studies are limited, the results of those studies demonstrate an  
13 increase in chronic inflammation over time, similar to studies of other amphibole asbestos fibers  
14 ([Mossman et al., 2011](#)). Overall, the evidence described above suggests chronic inflammation is  
15 observed following Libby Amphibole and tremolite asbestos exposure.

16 Although slightly increased compared to controls, cytotoxicity in murine macrophage  
17 cells exposed to LAA was decreased compared to other fiber types ([Blake et al., 2008](#)). No other  
18 in vitro study examined cytotoxicity following exposure to LAA, although an increase in  
19 apoptosis was demonstrated in this same cell system ([Blake et al., 2008](#)).

20 Compensatory proliferation in epithelial cells following cytotoxicity can lead to an  
21 increase in mutations (both spontaneous and induced). This increase is generally offset by  
22 increased levels of apoptosis, as in [Blake et al. \(2008\)](#). Recent studies in mice exposed to LAA  
23 demonstrated increased collagen deposition and collagen gene expression, markers of fibrosis  
24 ([Smartt et al., 2010](#); [Putnam et al., 2008](#)). Tremolite and LAA exposure led to increases in both  
25 fibrosis and tumorigenicity in all but one animal study, supporting a role for proliferation in  
26 response to these fibers. Taken together with studies on other asbestos fibers, these data suggest  
27 that cytotoxicity and cell proliferation may play a role in tumor formation.

28 Neither a dose-response concordance nor temporal relationship has been established  
29 between the development of cytotoxicity and regenerative cellular proliferation and exposure to  
30 LAA. However, cytotoxicity and regenerative cellular proliferation has been observed following  
31 exposure to LAA as well as other amphibole asbestos through other routes of exposure in in vivo  
32 assays (e.g., intratracheal instillation). Also, increases in markers of proliferative response have  
33 been observed in in vitro studies of LAA and other amphibole asbestos in epithelial cells. These  
34 results suggest exposure to LAA may lead to increases in cytotoxicity and regenerative  
35 proliferation; however, the data are not sufficient to determine a dose-response relationship.

36 It is generally accepted that sustained cell proliferation in response to cytotoxicity can be  
37 a significant risk factor for cancer ([Correa, 1996](#)). Sustained cytotoxicity and regenerative cell  
38 proliferation may result in the perpetuation of mutations (spontaneous or directly or indirectly

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induced by the chemical), resulting in uncontrolled growth. It is also possible that continuous proliferation may increase the probability that damaged DNA will not be repaired. Reparative proliferation alone is not assumed to cause cancer. Tissues with naturally high rates of turnover do not necessarily have high rates of cancer, and tissue toxicity in animal studies does not invariably lead to cancer. Nevertheless, regenerative proliferation associated with persistent cytotoxicity appears to be a risk factor of consequence.

#### **4.6.2.4. Conclusions About the Hypothesized Modes of Action**

*Is the hypothesized mode of action sufficiently supported in the test animals?*

There are a limited number of studies on the genotoxicity of LAA and/or tremolite. However, the studies described in Section 4 suggest a possible role for mutagenicity in asbestos-induced carcinogenicity. These studies showed chromosomal aberrations, increases in micronuclei induction, and increased reactive oxygen species production which has been shown to lead to mutagenicity (see Section 4, Table 4-21). One study of DNA adduct formation did not show any DNA damage or adducts following exposure to LAA. Laboratory animal studies of other amphibole asbestos have demonstrated similar results ([Huang et al., 2011](#)). Further research in this area is needed in order to inform the possibility of a mutagenic MOA for LAA.

Chronic inflammation is observed following exposure to most fibers studied ([Mossman et al., 2011](#)). Laboratory animal studies of LAA and tremolite demonstrated increases in inflammation, inflammatory markers, and increases in inflammatory cells. Further, in vitro studies have shown that exposure to LAA and tremolite lead to increases in expression of inflammatory cytokines. Available data are limited but consistent with the hypothesis that a MOA involving chronic inflammation contributes to asbestos-induced pulmonary and pleural tumors, either independently or in combination with a mutagenic MOA. However, it has not been determined whether chronic inflammation is a necessary precursor of carcinogenesis, and experimental support for causal links, such as compensatory cellular proliferation or clonal expansion of initiated cells, is lacking between toxicity and pulmonary or pleural tumor formation. However, further research is needed to determine if this MOA could be established for LAA and/or tremolite.

As reviewed in Section 4.2, in vivo and in vitro studies have shown a consistent cytotoxic and proliferative response to LAA and/or tremolite. Therefore, it has been proposed that cytotoxicity following pulmonary exposure to LAA and/or tremolite is a precursor to carcinogenicity. A more biologically plausible MOA may involve a combination of chronic inflammation, genotoxicity, and cytotoxicity, with genotoxicity increasing the rate of mutation and regenerative proliferation enhancing the survival or clonal expansion of mutated cells. However, this hypothesis has yet to be tested experimentally.

1 *Is the hypothesized mode of action relevant to humans*

2 Although limited for LAA, the evidence discussed above demonstrates that LAA  
3 exposure results in genotoxicity in in vitro and in vivo studies in test animal species. Therefore,  
4 the presumption is LAA would be genotoxic in humans. The few available data from human and  
5 in vivo laboratory animal studies concerning the genotoxicity of amphibole asbestos suggest  
6 consistency with this mechanism, but the studies are not sufficiently conclusive to provide direct  
7 supporting evidence for a mutagenic MOA. This MOA is considered relevant to humans.

8 The evidence discussed above demonstrates that exposure to LAA and/or tremolite  
9 asbestos lead to chronic inflammation. The available human data following exposure to LAA  
10 and other amphibole asbestos suggest consistency with this mechanism being relevant to  
11 humans. Data are inadequate to determine that a cytotoxic mechanism is operative following  
12 exposure to LAA in exposed populations; however, none of the available data suggest that this  
13 mechanism is biologically precluded in humans. Furthermore, both animal and in vitro studies  
14 suggest that LAA causes cytotoxicity at exposures that may induce pulmonary cancers,  
15 constituting positive evidence of the human relevance of this hypothesized MOA.

16  
17 *Which populations or life stages can be particularly susceptible to the hypothesized mode of*  
18 *action*

19 A mutagenic MOA is considered relevant to all populations and life stages. According to  
20 EPA's Cancer Guidelines ([U.S. EPA, 2005a](#)) and Supplemental Guidance ([U.S. EPA, 2005b](#)),  
21 there may be increased susceptibility to early-life exposures for carcinogens with a mutagenic  
22 MOA. The weight of evidence is insufficient to support a mutagenic MOA for LAA  
23 carcinogenicity and in the absence of chemical-specific data to evaluate differences in  
24 susceptibility, according to EPA's *Supplemental Guidance for Assessing Susceptibility from*  
25 *Early-Life Exposure to Carcinogens* ([U.S. EPA, 2005b](#)), the application of the age-dependent  
26 adjustment factors is not recommended.

27 Populations that may be more susceptible include those that may have varied fiber  
28 toxicokinetics related to potential anatomical, physiological, and biochemical differences which  
29 may impact fiber dosimetry (see Section 4.7). No data are available as to whether other factors  
30 may lead to different populations or life stages being more susceptible to a chronic inflammation  
31 MOA for LAA-induced tumors. For instance, it is not known how the hypothesized key events  
32 in chronic inflammatory response (e.g., increased oxidative stress) to fibers interact with known  
33 risk factors for human pulmonary or pleural carcinomas.

34 As with chronic inflammation, populations that may be more susceptible to increased  
35 cytotoxicity following exposure to LAA include those that may have varied fiber toxicokinetics  
36 related to potential anatomical, physiological, and biochemical differences which may impact  
37 fiber dosimetry (see Section 4.7). No data are available as to whether other factors may lead to  
38 different populations or life stages being more susceptible to a cytotoxic MOA for LAA-induced

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1 tumors. For instance, it is not known how the hypothesized key events (e.g., interference with  
2 the spindle apparatus) in this MOA interact with known risk factors for human pulmonary or  
3 pleural carcinomas.

#### 4 5 *Summary*

6 Research on multiple types of elongate mineral fibers supports the role of multiple modes  
7 of action following exposure to LAA. Of the MOAs described above, the evidence that chronic  
8 inflammation, genotoxicity and cytotoxicity, and cellular proliferation may all play a role in the  
9 carcinogenic response to LAA is only suggestive (see Table 4-23). In vitro studies provide  
10 evidence that amphibole asbestos is capable of eliciting genotoxic and mutagenic effects in  
11 mammalian respiratory cells; however, direct evidence linking mutagenicity to respiratory cells  
12 following inhalation exposure is lacking. Results of the in vivo studies described here are  
13 consistent with the hypothesis that some forms of amphibole asbestos act through a MOA  
14 dependent on cellular toxicity. This is largely based on the observations that cytotoxicity and  
15 reparative proliferation occur following subchronic exposure and bronchiolar tumors are  
16 produced at exposure levels that produce cytotoxicity and reparative proliferation. However,  
17 dose-response data in laboratory animal studies for damage/repair and tumor development are  
18 limited because a limited number of inhalation studies exist that used multiple doses of fibers.  
19 Although evidence is generally supportive of a MOA involving chronic inflammation or cellular  
20 toxicity and repair, there is insufficient evidence to support these hypotheses; thus, a linear  
21 approach is used to calculate the inhalation cancer unit risk in accordance with the default  
22 recommendation of the 2005 *Guidelines for Carcinogen Risk Assessment* ([U.S. EPA, 2005a](#)). It  
23 is possible that multiple MOAs discussed above, or an alternative MOA, may be responsible for  
24 tumor induction.

**Table 4-23. Hypothesized modes of action for carcinogenicity of Libby Amphibole asbestos in specific organs**

Potential MOA	Evidence for MOA	Limitations/evidence against MOA	Weight of evidence
<b>Lung cancer</b>			
Chronic inflammation	Inflammatory response demonstrated at the site of fiber deposition and has been linked to genotoxicity and mutagenicity.	Limited analysis of inflammation/tumor site concordance. Genotoxicity is commonly assumed to contribute to carcinogenesis. Inflammation can occur without progressing to cancer.	Some inconclusive evidence for this MOA.
ROS	ROS known to be produced following exposure to multiple types of fibers. ROS are associated with DNA damage, lipid peroxidation, and chronic inflammation.	ROS lead to DNA adduct formation, which in turn can lead to mutation. Limited studies have examined the production of ROS following exposure to LAA.	Suggestive evidence for this MOA for LAA (strong for other fiber types).
<b>Lung cancer—genotoxicity</b>			
Direct	Fibers directly interact with spindle apparatus and can interfere during mitosis leading to clastogenicity.	Ames assay inconclusive for fiber analysis (cell type unable to show ROS production and then possible mutations).	Suggestive evidence for this MOA for LAA (strong for other fiber types).
Indirect	Fibers lead to ROS production, which leads to DNA damage.	ROS lead to DNA adduct formation, which in turn can lead to mutation. Limited studies have examined the production of ROS following exposure to LAA (cell type unable to show ROS production).	Suggestive evidence for this MOA for LAA (strong for other fiber types).
Cytotoxicity and cellular proliferation	Increased cellular proliferation can increase the chance of cancer by increasing the population of mutations. Many fibers activate signaling pathways that lead to cellular proliferation.	Limited analysis of cell types/target tissues where cell proliferation occurs without chronic inflammation.	Suggestive evidence for this MOA for asbestos fibers.

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**Table 4-23. Hypothesized modes of action for carcinogenicity of Libby Amphibole asbestos in specific organs (continued)**

Potential MOA	Evidence for MOA	Limitations/Evidence against MOA	Weight of evidence
<b>Mesothelioma</b>			
Chronic inflammation	Inflammatory response demonstrated at the site of fiber deposition and has been linked to genotoxicity and mutagenicity.	Limited analysis of inflammation/tumor site concordance. Genotoxicity is commonly assumed to contribute to carcinogenesis. Inflammation can occur without progressing to cancer.	Insufficient evidence for this MOA.
ROS	ROS known to be produced following exposure to multiple types of fibers. ROS are associated with DNA damage, lipid peroxidation, and chronic inflammation.	Limited analysis in this target tissue. ROS lead to DNA adduct formation which in turn can lead to mutation. Limited studies have examined the production of ROS following exposure to LAA.	Insufficient evidence for this MOA.
<b>Mesothelioma—genotoxicity</b>			
Direct	Fibers directly interact with spindle apparatus and can interfere during mitosis, leading to clastogenicity.	Limited analysis in this target tissue. Ames assay inconclusive for fiber analysis (cell type unable to show ROS production followed by possible mutations).	Insufficient evidence for this MOA.
Indirect	Fibers lead to ROS production, which leads to DNA damage.	Limited analysis in this target tissue. ROS lead to DNA adduct formation which in turn can lead to mutation. Limited studies have examined the production of ROS following exposure to LAA.	Insufficient evidence for this MOA.
Cytotoxicity and cellular proliferation	Increased cellular proliferation can increase chance of cancer by increasing the population of mutations. Many fibers activate signaling pathways that lead to cellular proliferation.	Limited analysis in this target tissue. Limited analysis of cell types/target tissues where cell proliferation occurs without chronic inflammation.	Insufficient evidence for this MOA.
<b>Lymphatic system and other organs</b>			
Data not available	Data not available	Limited analysis in these target tissues.	Insufficient evidence for any MOA.

#### 4.6.2.5. Application of the Age-Dependent Adjustment Factors

As described above, the MOA for LAA is unknown. The weight of evidence does not support a mutagenic MOA for LAA carcinogenicity. Therefore, according to EPA's *Supplemental Guidance for Assessing Susceptibility from Early-Life Exposure to Carcinogens* ([U.S. EPA, 2005b](#)), the application of the age-dependent adjustment factors is not recommended.

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## 4.7. SUSCEPTIBLE POPULATIONS

Certain populations may be more susceptible to adverse health effects from exposure to LAA. Because the adverse health effects resulting from exposure to LAA have been primarily studied in occupational cohorts of adult white men (see Sections 4.1.1 and 4.1.3), there is limited information on the effects to a broader population. A few studies, however, have examined health effects resulting from nonoccupational exposure in other age groups, genders (i.e., females), and races or ethnicity groups. The data from these studies could inform whether any differential risk exists for these groups (see Sections 4.1.2 and 4.1.4). However, it should be noted that distinguishing true differences from chance variation in effect estimates is related to the sample size and statistical power, which is usually limited in these studies. In addition, genetic polymorphisms, preexisting health conditions, and differences in nutritional status may alter an individual's response to LAA. Finally, coexposures to other substances (e.g., tobacco smoke or particulate matter) may increase an individual's risk of adverse health effects from exposure to LAA. When data are available, each of these factors is discussed below with respect to increased susceptibility to cancer and noncancer effects from exposure to LAA. When information specific to LAA is not available, the general literature on the toxicity of mineral fibers is briefly referenced.

There are also factors that may influence one's exposure potential to asbestos based on life stage or other characteristics. For example, children spend more hours outside and may engage in activities which impact exposure potential compared to adults ([U.S. EPA, 2006b](#); [NRC, 1993](#)). Because life stage and activity patterns can increase the potential for health effects from exposure, these factors define who may be more susceptible to health effects due to greater exposure. Section 2.3 discusses this exposure potential, including how children, workers, household contacts, and residents may be exposed to LAA.

### 4.7.1. Influence of Different Life Stages on Susceptibility

Individuals at different life stages differ from one another physiologically, anatomically, and biochemically. Individuals in early and later life stages differ markedly from adulthood in terms of body composition, organ function, and many other physiological parameters, which can influence the toxicokinetics and toxicodynamics of chemicals and their metabolites in the body ([Guzelian et al., 1992](#)). This also holds true for mineral fibers, including asbestos fibers (see Section 3). This section presents and evaluates the literature on how individuals in early or later life stages might respond differently and thus potentially be more susceptible to adverse health effects of LAA exposure.

#### 4.7.1.1. Life-Stage Susceptibility

Humans in early life stages (i.e., conception through adolescence) can have unique susceptibilities compared to those in later life stages because they undergo rapid physiological changes during critical periods of development ([Selevan et al., 2000](#)). Furthermore, young people are often exposed to xenobiotics via unique exposure pathways (i.e., transplacental transfer and breast milk ingestion; [U.S. EPA, 2006b](#); [NRC, 1993](#)). The nature of these alternate exposure pathways, and the lack of studies that accurately document exposure levels and outcomes in the very young, contribute to the difficulty in assessing the relative susceptibility of early life stage exposure to amphibole asbestos.

No in utero exposure data exist for LAA but limited observations in stillborn infants indicate transplacental transfer of tremolite ([Haque et al., 1998](#); [Haque et al., 1996](#)) and other asbestos and nonasbestos fibers does occur ([Haque et al., 1998](#); [Haque et al., 1996](#); [Haque et al., 1992](#); [Haque et al., 1991](#)). Transplacental transfer of asbestos was also demonstrated in animals following maternal exposure by gavage ([Haque et al., 2001](#)) or injection ([Haque and Vrazel, 1998](#); [Cunningham and Pontefract, 1974](#); see Section 3). These studies did not evaluate the sources or levels of exposure, and injection studies are a less relevant route of exposure compared to inhalation. Based on these studies, LAA fibers may be transferred through the placenta, resulting in prenatal exposure at any stage of fetal development.

A number of studies have attempted to determine the impact of in utero and early life exposure on the developing child. Those analyses performed in the very young include reports of stillbirth ([Haque et al., 1998](#); [Haque et al., 1996](#)) and death among infants and young children (age 1–27 months) due to sudden infant death syndrome and bronchopulmonary dysplasia ([Haque and Kanz, 1988](#)). These studies found higher levels of asbestos in the lungs of those who died compared to unexposed individuals. In an infant study, the authors speculate that there was either a preexisting abnormal lung physiology in these children that contributed to a reduced ability to clear fibers from the lung, or the children had an increased exposure to asbestos ([Haque and Kanz, 1988](#)). Those studies conducted in older children include reports of pleural and diaphragmatic calcifications ([Epler et al., 1980](#)) and altered immune and respiratory conditions ([Shtol' et al., 2000](#)). Although the data are suggestive of increased sensitivity in infants, no definitive conclusion can be reached.

In experimental animal studies, the effects of in utero and early life exposure to asbestos are equivocal. Rats' offspring that were exposed to tremolite had decreased body-weight gain at weaning and 8-weeks old compared to controls ([NTP, 1990b](#); [McConnell et al., 1983a](#)). This finding was observed in similar studies with other forms of asbestos ([NTP, 1990a, 1988, 1985](#); [McConnell et al., 1983a](#)) but not replicated in others ([McConnell et al., 1983b](#); [NTP, 1983](#)). Embryonic toxicity was noted in a few experimental animal studies. Crocidolite injected into pregnant mice resulted in altered limb differentiation in cultured embryos ([Krowke et al., 1983, abstract](#)), and chrysotile suspended in drinking water and given to pregnant mice resulted in

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1 decreased postimplantation survival in cultured embryos ([Schneider and Maurer, 1977](#)).  
2 However, chrysotile ingested via drinking water did not affect embryonic survival in vivo in  
3 pregnant mice ([Schneider and Maurer, 1977](#)). Altogether, the data provide no clear evidence for  
4 increased susceptibility following early life or in utero asbestos exposure.

5 Several studies have examined the susceptibility of asbestos exposure on young children,  
6 including how fiber deposition is affected by the physiological differences in children's lungs.  
7 Evidence suggests that fiber deposition is increased in the lungs of children compared with adults  
8 ([Bennett et al., 2008](#); [Isaacs and Martonen, 2005](#); [Asgharian et al., 2004](#); [Phalen and Oldham,](#)  
9 [2001](#); [Oldham et al., 1997](#); [Schiller-Scotland et al., 1994](#); [Phalen et al., 1985](#)). Nasal deposition  
10 of particles was lower in children compared to adults—particularly during exercise ([Becquemin](#)  
11 [et al., 1991](#)). The lung and nasal depositional differences are partially due to structural  
12 differences across life stages that change the depositional pattern of different fiber sizes, possibly  
13 altering the site of action, and resulting in differential clearance and subsequent health effects.  
14 However, it is unclear whether the lung surface, body weight, inhalation volume, or exposure  
15 patterns are most determinative of dose.

16 There are a few studies analyzing noncancer outcomes in children exposed to Libby  
17 Amphibole. A Libby medical screening program collected data on 7,307 participants, including  
18 600 children aged 10–17 years, which represents 8.2% of the cohort ([Peipins et al., 2003](#)).  
19 Pulmonary function tests showed that none of these children had moderate or severely restricted  
20 lung function ([ATSDR, 2002, 2001b](#)). This program also studied chest radiographs for those  
21 18 years or older ([Noonan et al., 2006](#); [Peipins et al., 2003](#); [ATSDR, 2001b](#)), but x-rays were not  
22 conducted on children. Among 1,003 adolescents and young adults (ages 10 to 29) who were  
23 ≤age 18 in 1990 when the mining/milling operations closed ([Vinikoor et al., 2010](#)), there was  
24 little variation in prevalence of shortness of breath, physician-diagnosed lung disease, or  
25 abnormal pulmonary function tests (restrictive, obstructive, or mixed, based on FEV<sub>1</sub> and FVC  
26 values) across the exposure categories. This analysis does not directly address the issue of  
27 susceptibility by age, however, because a comparison with people exposed only at older ages is  
28 not included.

29 Based on limited studies described below, it is possible that early life stage exposure may  
30 increase the risk of noncancer outcomes in adulthood. Altered immunity ([Zerva et al., 1989](#)) and  
31 asbestosis ([Voisin et al., 1994](#)) were observed in adults following tremolite exposure during  
32 childhood. No other studies of noncancer outcomes in early life stages of humans or  
33 experimental animals exposed to LAA have been reported. Thus, additional research is needed  
34 to establish the clinical significance of these findings and to expand understanding of the  
35 progression of the adverse health effects in the community.

36 To address the potential for increased susceptibility to cancer from early lifetime  
37 exposures, one needs to consider if there is evidence of differential health effects such as  
38 increased potency from early lifetime exposure, changes in latency based on the age of exposure,

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1 or cancers observed with early lifetime exposures not seen with adult exposures. There are no  
2 published reports that can directly answer these questions for exposure to LAA. Few cancers  
3 occurring in childhood have been documented in children exposed to any form of asbestos.  
4 Examples of cases include a 17-year-old exposed to chrysotile and tremolite ([Andrion et al.,](#)  
5 [1994](#)) and a 3-year-old exposed to chrysotile ([Lieben and Pistawka, 1967](#)), both of whom  
6 developed mesothelioma. Notably, childhood mesothelioma may have an etiology that is  
7 different from that of the disease seen in adults, further confounding interpretation of these data  
8 ([Cooper et al., 1989](#)).

9       Studies involving populations exposed to other types of asbestos have yielded equivocal  
10 results on the carcinogenic effects following exposures occurring earlier in life, and few evaluate  
11 very early life exposures. One study in the United Kingdom described occupational exposure to  
12 chrysotile, crocidolite, and amosite for a group of 900 women. First exposure from ages  
13 15–24 years led to a higher relative mortality risk for lung and pleural cancer compared with  
14 women who were first exposed at older ages ([SMR 30 based on 12 observed and 0.4 expected,](#)  
15 [SMR 8 based on 4 observed and 0.5 expected, and SMR 6.7 based on 6 observed and 0.9](#)  
16 [expected in the first exposure at ages 15–24, 25–34, and ≥35 years, respectively; Newhouse et](#)  
17 [al., 1972](#)). In a study in Wittenoom, Western Australia, 27 individuals were diagnosed with  
18 mesothelioma who had been environmentally exposed to crocidolite (i.e., residents of the town  
19 but not directly employed in the area’s crocidolite mining and milling industry); 11 of these  
20 subjects were <15 years old at the time of exposure ([Hansen et al., 1998](#)). One-third of all the  
21 subjects were younger than 40 years old when diagnosed, but the authors found no increase in  
22 mesothelioma mortality rates when analyzed by age at first exposure. However, risk was  
23 significantly increased based on time from the first exposure, duration of exposure, and  
24 cumulative exposure ([Hansen et al., 1998](#)). Additional studies of this cohort found that the  
25 mesothelioma mortality rate was lower for those first exposed (based on age residence in the area  
26 began) to crocidolite at ages <15 years ( $n = 24$ ; mesothelioma mortality rate 47 per  
27 100,000 person-year) compared with those first exposed at ages ≥15 years ( $n = 43$ ; [mesothelioma](#)  
28 [mortality rate 112 per 100,000 person-year; Reid et al., 2007](#)). The hazard ratio for age at first  
29 residential exposure of ≥15 years compared with <15 years was 3.83 (95% CI: 2.19, 6.71),  
30 adjusting for cumulative exposure, gender, and an interaction term for gender and cumulative  
31 exposure. Altogether, these studies do not clarify whether exposure during childhood yields  
32 different adverse health effects compared with exposure during adulthood.

33       Relatively few studies have examined the effects of asbestos exposure in juvenile  
34 animals. Oral exposure to nonfibrous tremolite did not increase tumors in the offspring of rats  
35 compared to controls ([NTP, 1990b; McConnell et al., 1983a](#)). Similar studies of other forms of  
36 asbestos reported an increase of various neoplasms in the offspring ([NTP, 1990a, 1988, 1985;](#)  
37 [McConnell et al., 1983b; McConnell et al., 1983a](#)), but another study reported none ([NTP, 1983](#)).  
38 No cancer bioassays have been performed in juvenile animals exposed to LAA. Based on these

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1 very limited and inconclusive studies on other forms of asbestos, no conclusions can be drawn  
2 about differential risk of adverse health effects after early life stage exposure to LAA compared  
3 to exposure during adulthood. It is unknown whether early life stage exposure compared to adult  
4 exposure increases susceptibility for adult cancers, as measured by increased incidence, severity,  
5 or disease progression, or by decreased latency.

6 Later life stage is generally defined as  $\geq 65$  years old. Because pulmonary function  
7 (volume and rate of breathing) decreases with age ([Weiss, 2010](#)), increased deposition of fibers  
8 in the lung from exposures in later life stages is unlikely. Older adults could be more susceptible  
9 to the effects of LAA due to the gradual age-related decline in physiological processes. For  
10 instance, clearance of fibers from the lung might be reduced since cough reflex and strength of  
11 older adults is less effective and the cilia are less able to move mucus out of the airway ([U.S.  
12 EPA, 2006a](#)). Additionally, decreased immune function, increased genetic damage, and  
13 decreased DNA repair capacity can result in increased susceptibility with age ([U.S. EPA, 2006a](#)).  
14 These age-associated alterations could decrease fiber-induced DNA damage repair but might  
15 also reduce the incidence of fiber-induced DNA damage due to decreased phagocytosis or  
16 inflammation. Specific data pertaining to age-varying effects of LAA on these processes are not  
17 available.

18 Because the risk of many types of noncancer effects increases with age, an increasing rate  
19 of specific diseases with increasing age can be expected among individuals exposed at some  
20 point in their lives to LAA. Radiographic tests among those exposed to Libby Amphibole show  
21 that older age, which in some occupational settings may be highly correlated with time since first  
22 exposure (TSFE), is one of the factors most associated with pleural or interstitial abnormalities  
23 ([Rohs et al., 2008](#); [Horton et al., 2006](#); [Muravov et al., 2005](#); [Peipins et al., 2003](#); [ATSDR,  
24 2001b](#); [Amandus et al., 1987a](#); [McDonald et al., 1986b](#); [Lockey et al., 1984](#)). Abnormal  
25 radiographs also increase with age in general population studies ([Pinsky et al., 2006](#)). In a  
26 community health screening study, an increased risk of rheumatoid arthritis among individuals  
27 ages  $\geq 65$  years was observed in relation to several measures reflecting exposure to LAA ([e.g.,  
28 worked for W.R. Grace, used vermiculite for gardening; Noonan, 2006](#)). However, the available  
29 studies do not provide a basis for evaluating the timing of the exposure in relation to these  
30 outcomes. No conclusions can be drawn about differential risk of noncancer after later life stage  
31 exposure to Libby Amphibole compared to exposure earlier in life.

32 No studies assessing the carcinogenic effect of exposures occurring in older age groups  
33 are available for LAA or other amphiboles. It should be noted that health effects observed  
34 among individuals exposed to LAA are likely to increase with age due to the long latency period  
35 for the exposure response for asbestos and lung cancer and other chronic diseases. However, this  
36 type of observation would not directly address the question of whether exposures at older ages  
37 have a stronger or weaker effect compared with exposures at younger ages.

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#### 4.7.2. Influence of Gender on Susceptibility

A discussion of gender-related differences in risk from asbestos exposure raises several important issues, such as gender-related differences in exposure patterns, physiology, and dose-response ([Smith, 2002](#)). For example, nasal breathing filters out particles, and men tend to breathe less through their nose during exercise than women do ([Bennett et al., 2003](#)). [Bennett et al. \(1996\)](#) showed a gender difference in fractional deposition (defined as the ratio of particles not exhaled to total particles inhaled) of particles 2 µm in mass median aerodynamic diameter. This particle diameter is within the range of LAA particles reported in Table 2-2. This study found that, in general, women had a greater retention of particles compared to men because men had higher ventilation rates compared to women; however, the overall deposition rate was higher in the men ([Bennett et al., 1996](#)).

Most occupational studies for LAA have examined the effects of exposure only in men ([Moolgavkar et al., 2010](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus et al., 1988](#); [Amandus et al., 1987b](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#); [McDonald et al., 1986b](#)). There is limited information specifically on women exposed to LAA. In the Libby, MT community studies, no gender-related trends in mortality due to lung or digestive cancer were observed ([ATSDR, 2000](#)). These limited data do not provide a basis for drawing conclusions regarding gender-related differences in adverse health effects from LAA.

#### 4.7.3. Influence of Race or Ethnicity on Susceptibility

Race and ethnicity often are used in medical and epidemiological studies to define various groups of the population. These categories could be surrogates for differences in exposure (e.g., occupation, socioeconomic, behavior) or biology (e.g., physiology, genetics), in which case these factors may play a role in susceptibility as well. Nasal structure and lung architecture can influence the depositional patterns for both particles and fibers. One study of 18 Caucasians (ages 8 to 30 years) and 14 African Americans (ages 8 to 25 years) reported increased ventilation rates during exercise in the African Americans ([matched on gender, age, height, and weight; Cerny, 1987](#)). Another study (11 Caucasians and 11 African Americans, ages 18 to 31 years) reported decreased nasal deposition efficiency (for particle sizes of 1–2 µm, which is in the range of those for LAA reported in Table 2-2) in African Americans compared to Caucasians ([Bennett and Zeman, 2005](#)). Furthermore, nasal breathing during exercise occurred less in Caucasians compared to African Americans in this study ([Bennett et al., 2003](#)).

Of the occupational and residential studies for LAA, the vast majority of subjects with known race were white, precluding the ability to conduct an analysis of racial and ethnicity-related differences in the mortality risks within the Libby worker cohort. In a study of occupational exposure to chrysotile asbestos in a textile factor, lung cancer mortality risk in relation to exposure was lower in nonwhite males (0.84, 95% CI: 0.52–1.27) compared to white

1 males (2.34, 95% CI: 1.94–2.79), although a statistically significant increase in SMR was  
2 observed for nonwhite males at high exposure levels ( $\geq 120$  fiber-yr/mL; [Hein et al., 2007](#)). This  
3 observed difference could be due to a lower prevalence of smoking among nonwhite compared  
4 with white males ([Hein et al., 2007](#)).

#### 6 **4.7.4. Influence of Genetic Polymorphisms on Susceptibility**

7 *XRCC1* is a DNA damage repair gene. A recent study demonstrated that  
8 *XRCC1*-deficient cells exposed to Libby Amphibole or crocidolite asbestos demonstrated  
9 increased levels of micronuclei induction ([Pietruska et al., 2010](#)). Two other studies examined  
10 *XRCC1* polymorphisms in relation to disease risk with other types of asbestos exposure. [Zhao et](#)  
11 [al. \(2006\)](#) found no association between *XRCC1* polymorphisms and asbestosis in  
12 asbestos-exposed workers. A study by [Dianzani et al. \(2006\)](#), however, did find an association  
13 between *XRCC1* and asbestos-induced lung disease in a population exposed to asbestos  
14 pollution. Further work is necessary, with clear definitions of patient populations and their  
15 exposure levels, so that these studies and others can be compared to determine if *XRCC1*  
16 polymorphisms increase susceptibility to adverse health effects following exposure to LAA.

17 Superoxide dismutases are free radical scavengers that dismutate superoxide anions to  
18 oxygen and hydrogen peroxide. SODs are expressed in most cell types exposed to oxygen.  
19 Several common forms of SODs occur and are named by the protein cofactor: copper/zinc,  
20 manganese, iron, or nickel. A recent study observed no significant alterations in levels of  
21 intracellular SOD following a 3-hour exposure to LAA in mice ([Blake et al., 2007](#)). Other  
22 studies in humans and mice have examined SOD expression in relation to other types of asbestos  
23 exposure. Manganese SOD activity was elevated in biopsies of human asbestos-associated  
24 malignant mesothelioma, although no genotypic differences were found to be related to this  
25 change in activity ([Hirvonen et al., 2002](#)). Other studies have focused on the role of extracellular  
26 superoxide dismutase (EcSOD) and asbestos-induced pulmonary disease ([Kliment et al., 2009](#);  
27 [Gao et al., 2008](#); [Fattman et al., 2006](#); [Tan et al., 2004](#)). These studies have suggested a  
28 protective effect of EcSOD, because mice that lack this form of SOD have increased sensitivity  
29 to asbestos-induced lung injury ([Fattman et al., 2006](#)). Familial studies showing an unusually  
30 high incidence of mesothelioma suggest that genetic factors might play a role in the etiology of  
31 mesothelioma ([Ugolini et al., 2008](#); [Huncharek, 2002](#); [Roushdy-Hammady et al., 2001](#)), although  
32 whether a genetic factor or a common environmental element leads to the similar responses in  
33 these families is difficult to determine. Increased interest in the role of genetic factors in  
34 asbestos-related health outcomes has led to several analytical studies on specific genetic  
35 polymorphisms. A review of 24 published reports (19 studies) discusses the current state of  
36 knowledge regarding genetic susceptibility associated with asbestos-related diseases (in  
37 particular, malignant pleural mesothelioma). Results from several studies demonstrated an  
38 association between asbestosis-related diseases and *GSTM1*-null polymorphism, whereas results

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for other polymorphisms were conflicting ([Neri et al., 2008](#)). Some polymorphisms discussed in [Neri et al. \(2008\)](#) are in genes for *N-acetyl-transferase 2*; *glutathione-S-transferases (GSTs)*; *SOD*; *CYP1A1*, *CYP2D6*; *neurofibromatosis 2 (Nf2)*; *p53*; and *XRCC1*. Although occupational asbestos exposure was assessed, the type of asbestos is generally unknown in these studies.

Limited animal studies have examined the role of genetic variations related to asbestos exposure, including specific signaling pathways ([Shukla et al., 2007](#)), DNA damage repair ([Lin et al., 2000](#); [Ni et al., 2000](#)), and tumor suppressor genes ([Vaslet et al., 2002](#); [Kleymenova et al., 1997](#); [Marsella et al., 1997](#)). Genetic alterations of particular interest for mesothelioma include those involved in tumor suppression (*p53*, *Nf2*) and oxidative stress (*SOD*, *GSTs*). *Nf2* and *p53* are frequently altered in mesotheliomas, but no consistent mutations have been found ([Cheng et al., 1999](#); [Mayall et al., 1999](#); [Bianchi et al., 1995](#)). Alterations in expression of antioxidant enzymes like SOD and GST in mesothelioma can yield cells more resistant to oxidative stress as compared to normal cells due to increased antioxidant activity ([Ramos-Nino et al., 2002](#); [Rahman and MacNee, 1999](#)). No studies that examine the role of cell-cycle control genes were found following exposure to LAA. Additionally, no information on other genetic polymorphisms in relation to disease risk among those exposed to LAA was identified in the available literature.

#### 4.7.5. Influence of Health Status on Susceptibility

Preexisting health conditions could potentially alter the biological response to asbestos exposure. Mesothelioma risk has been hypothesized to be related to immune impairment ([Bianchi and Bianchi, 2008](#)) and Simian virus 40 (SV40) exposure in humans ([Carbone et al., 2007](#); [Kroczyńska et al., 2006](#); [Cristaudo et al., 2005](#); [Foddìs et al., 2002](#); [Bocchetta et al., 2000](#); [Mayall et al., 1999](#)). Coexposure to asbestos and SV40 has been associated with p53-related effects in vitro ([Foddìs et al., 2002](#); [Bocchetta et al., 2000](#); [Mayall et al., 1999](#)), and cell signaling aberrations in vivo ([Kroczyńska et al., 2006](#); [Cristaudo et al., 2005](#)). However, the influence on cancer risk is unknown, as these lines of research are not fully developed and have not been applied specifically to LAA.

Obesity can compromise inhalation exposure, as increased particle deposition in the lungs of overweight children ([Bennett and Zeman, 2004](#)) and adults ([Graham et al., 1990](#)) has been observed. Individuals with respiratory diseases could have compromised lung function that alters inhalation exposure to LAA. For example, individuals with chronic obstructive pulmonary disease (COPD) have increased inhalation volume ([Phalen et al., 2006](#)) and increased fine particle deposition ([Phalen et al., 2006](#); [Bennett et al., 1997](#); [Kim and Kang, 1997](#)) and retention ([Regnis et al., 2000](#)). Similarly, studies have reported an increase in coarse particle (aerodynamic diameter >5 µm) deposition in individuals with cystic fibrosis ([Brown and Bennett, 2004](#); [Brown et al., 2001](#)). For people exposed to LAA, an increased risk for interstitial lung abnormalities was observed for those with a history of pneumonia ([Peipins et al., 2003](#)). In

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another study, bronchial asthma was examined as a potential confounding variable for asbestos-related effects on pulmonary function, although no confounding was observed ([Whitehouse, 2004](#)).

#### 4.7.6. Influence of Lifestyle Factors on Susceptibility

Smoking can impair clearance of particles from the lung ([Camner, 1980](#); [Cohen et al., 1979](#)) and increase deposition of asbestos fibers ([Sekhon et al., 1995](#); [McFadden et al., 1986](#)). These effects could lead to the retention of more inhaled asbestos fibers and for a longer period of time in smokers compared to nonsmokers, even when controlling for initial exposure. Evidence of smoking-related susceptibility to pulmonary effects of asbestos was reported by [Christensen and Kopylev \(2012\)](#) using data from the O.M. Scott, Marysville, OH plant cohort described by [Rohs et al. \(2008\)](#). The amount of LAA exposure required to elicit the same increase in risk of localized pleural thickening was considerably lower (sixfold) for smokers compared with nonsmokers.

No studies were identified that examined lifestyle factors specifically with respect to LAA and cancer susceptibility. Lifestyle factors such as exercise, nutritional status, and smoking habits could affect the biological effects of asbestos exposure through various mechanisms. For example, those with more physically demanding jobs or those who regularly engage in vigorous exercise might experience increased lung deposition from fine particles or fibers compared to those with a more sedentary lifestyle ([Phalen et al., 2006](#); [Becquemin et al., 1991](#)). Randomized controlled trials of vitamin supplementation (beta-carotene and retinol) have been conducted for asbestos-related lung cancer, but results do not support a protective effect ([Cullen et al., 2005](#)).

For lung cancer, a synergistic relationship between cigarette smoking and asbestos exposure has been demonstrated ([Wraith and Mengersen, 2007](#); [Hammond et al., 1979](#); [Selikoff and Hammond, 1979](#)). Research has suggested that asbestos fibers might also enhance the delivery of multiple carcinogens in cigarette smoke, and that cigarette smoking decreases the clearance mechanisms in the lungs and could, therefore, lead to an increase in fiber presence in the lungs ([Nelson and Kelsey, 2002](#)). Smoking likely causes genetic alterations associated with lung cancer ([Landi et al., 2008](#)) that might increase the carcinogenic risk from exposure to asbestos. Benzo[a]pyrene, a component of tobacco, also has been observed to enhance the carcinogenic effects of asbestos ([Loli et al., 2004](#); [Kimizuka et al., 1987](#); [Mossman et al., 1984](#); [DiPaolo et al., 1983](#); [Mossman et al., 1983](#); [Reiss et al., 1983](#)).

#### 4.7.7. Susceptible Populations Summary

A very limited amount of information is available on exposure to LAA early in life that could lead to increased risk of asbestos-induced disease later in life. Due to the long latency period of some diseases in relation to asbestos exposure in general, adverse effects may be more



1 likely to be observed with an increase in age. This assumption requires further investigation.  
2 The number of women who have been occupationally exposed to LAA is very small, and health  
3 risks have not been evaluated specifically for this group. Differences between men and women  
4 in residential sources and types of exposure (e.g., types of activities done in the household) also  
5 preclude the possibility of drawing conclusions regarding the relative susceptibility of women  
6 compared with men to health effects of exposure to LAA. Similarly, sufficient data are not  
7 available to draw conclusions regarding racial or ethnic variation in susceptibility to diseases  
8 caused by exposure to LAA. In addition, the potential modifying effects of genetic  
9 polymorphisms, preexisting health conditions, nutritional status, and other lifestyle factors have  
10 not been studied, specifically as related to exposure of LAA and health outcomes.

## 5. EXPOSURE-RESPONSE ASSESSMENT

### 5.1. ORAL REFERENCE DOSE (RfD)

An oral reference dose is not derived due to a lack of data on the toxic effects of Libby Amphibole asbestos<sup>21</sup> (LAA) following oral exposure.

### 5.2. INHALATION REFERENCE CONCENTRATION (RfC)

An RfC is defined as “an estimate of an exposure (including sensitive subgroups) that is likely to be without an appreciable risk of adverse health effects over a lifetime” ([U.S. EPA, 2002](#)). Consequently, studies that relate these adverse health effects to exposure levels are necessary for RfC derivation. Preferred study characteristics for RfC derivation include adequate exposure-response information, ideally with quantitative exposure estimates to distinguish exposure levels in the study subjects, and adequate duration of follow-up to identify health effects of interest.

#### *Overview of the Methodological Approach*

The noncancer effects which were evaluated in populations with exposure to LAA (see Sections 4.1.2 and 4.1.3) are pulmonary effects (including asbestosis, pleural thickening [localized or diffuse], and other nonmalignant respiratory disease), cardiovascular disease-related mortality, and autoimmune effects. Localized pleural thickening (LPT) was deemed the most sensitive and was thus selected as the critical effect to derive the RfC (see Section 5.2.2.3). A benchmark response (BMR) of 10% extra risk was selected for exposure-response modeling (see Section 5.2.2.5).

RfCs are based on human data when appropriate epidemiologic studies are available. The general approach to developing an RfC from human epidemiologic data is to quantitatively evaluate the exposure-response relationship for that agent to derive a specific estimate of its effect on the risk of the selected outcome in the studied population. For the current assessment, the first step was to identify the most appropriate data set available to quantitatively estimate the effects of LAA exposure on pleural effects. Studies of three different cohorts provide such quantitative exposure-response information. Two are of occupationally exposed cohorts. The first one is Libby Workers ([Larson et al., 2012a](#)) and the second one is Marysville workers ([Rohs et al., 2008](#)). The third consisted of community members with nonoccupational exposure, who resided around the Western Minerals plant in Minneapolis ([Alexander et al., 2012](#)). Upon evaluating these three cohorts, the Marysville workers were selected as the most appropriate for

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<sup>21</sup>The term “Libby Amphibole asbestos” is used in this document to identify the mixture of amphibole mineral fibers of varying elemental composition (e.g., winchite, richterite, tremolite, etc.) that have been identified in the Rainy Creek complex near Libby, MT. It is further described in Section 2.2.

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1 derivation of the RfC. The Libby workers had generally higher levels of occupational exposure  
2 and additional, unquantified exposures outside of the workplace (i.e., residential exposures).  
3 While data on a critical predictor of the risk of pleural thickening (the time since first exposure  
4 [TSFE]) was available for the occupational exposures, information on TSFE for the residential  
5 exposures from living in Libby, MT prior to working in the mining or related operations is not  
6 known. Substantial uncertainty also exists in the exposure estimates for the Minneapolis study,  
7 and although some information on residential history was used by the investigators, it is unclear  
8 whether this information applied just to the residence or whether there was information on TSFE  
9 for the different exposure routes ([see Kelly et al., 2006](#)).

10 Among the Marysville workers, there were differences in the availability of exposure and  
11 health outcome data over time. No industrial hygiene measurements were available before 1972.  
12 Health examinations were performed at two time points, 1980 and 2002–2005, using different  
13 x-ray reading protocols and different film readers. Thus, the subgroup of workers with the  
14 highest quality exposure and outcome evaluation information, was determined to be those  
15 workers who were hired in 1972 or later, and who had health examinations performed in  
16 2002–2005; this group was selected as the primary analytic data set for derivation of the RfC  
17 (see Section 5.2.2.2). Once the relevant data describing a well-defined group of individuals  
18 along with their exposures and health outcomes were selected, a suite of appropriate statistical  
19 model forms was evaluated. Before performing any modeling, biological and epidemiological  
20 features were considered to determine a priori which variables and which models would be most  
21 suitable for the given exposure and health outcome (see Section 5.2.2.6.1). Based on these  
22 considerations, the Dichotomous Hill model was considered to be the most flexible and  
23 potentially most suitable model form; however, all model forms suitable for dichotomous  
24 epidemiological data were examined. Each model was evaluated for adequate fit to the data,  
25 with each person's individual-level exposures and outcomes modeled using a variety of exposure  
26 metrics. Appropriate covariates, which may be important predictors of LPT risk, were evaluated  
27 for potential confounding in the statistical model.

28 In the primary analytic data set (the subcohort of workers hired in 1972 or later), all  
29 univariate models examined had adequate fit and for each model form, mean exposure was  
30 shown to have the best relative fit (compared with either cumulative or residence time-weighted  
31 exposure metrics). Among the model forms, the relative fits were comparable, and thus the  
32 Dichotomous Hill model (with plateau fixed at 85%) using the mean exposure metric was  
33 selected as the primary model for RfC derivation. When evaluating nonexposure-related  
34 covariates, none were found to fit the criteria for a confounder (i.e., they were not associated  
35 with both the outcome and the exposure) and were not significant predictors of LPT risk when  
36 included in the final model. Time since first exposure, one of the key covariates evaluated, was  
37 associated with the exposure in the primary analytic data set but not the outcome. This is likely  
38 because there was a relatively narrow range of TSFE values (i.e., low variability) in the primary

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analytic data set. However, based on the epidemiological literature, TSFE is expected to be a major predictor of LPT risk and thus an important variable in evaluating the exposure-response relationship with LAA. Because inclusion of TSFE in the model did not improve model fit (and it was not a significant predictor), alternative strategies were explored to incorporate the effect of TSFE into the exposure-response model. EPA decided to use a hybrid modeling strategy. First, the effect of TSFE was estimated in a larger subset of the Marysville workers (all those with health evaluations in 2002–2005, regardless of hire date) with a broader range of TSFE values, using the same model as for the primary analytic data set. Next, this estimated effect was carried over to the model for the primary analytic data set (workers hired in 1972 or later with health examinations in 2002–2005) as a fixed regression coefficient, and the benchmark concentration and lower limit of the benchmark concentration (BMCL) estimated.

The BMCL from the “hybrid” modeling approach was used as the point of departure. Uncertainty factors were then applied to derive an RfC (see Section 5.2.3). Alternative analyses are presented in Section 5.2.4 and 5.2.5 with a summary in Section 5.2.6. Uncertainties in this noncancer assessment are described in detail in Section 5.3.

## **5.2.1. Choice of Principal Study**

### **5.2.1.1. Candidate Studies**

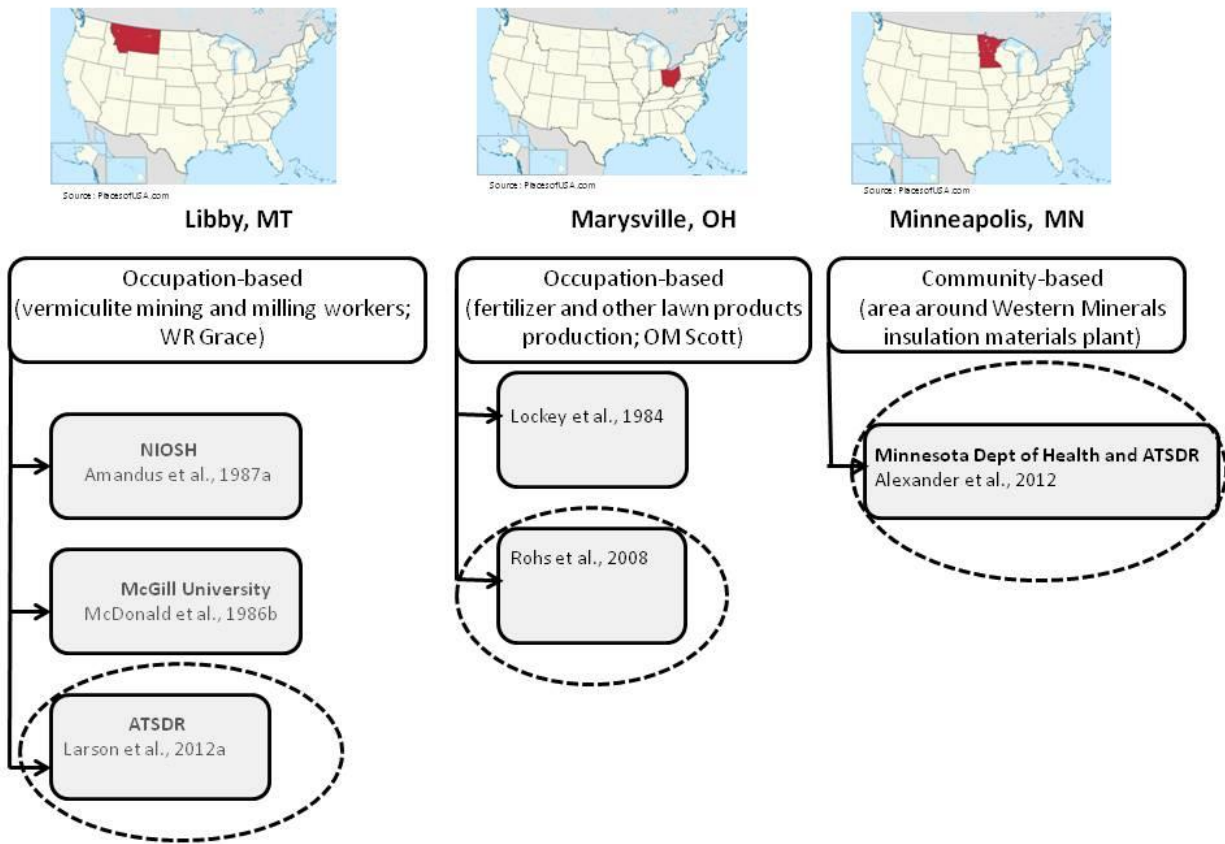
While there are studies of health effects in humans, no studies in laboratory animals on the inhalation route of exposure are suitable for derivation of an RfC because the available animal studies lack adequate LAA exposure-response information and are of a short-term duration.

Multiple studies have identified several noncancer health effects in humans that could be considered as potential critical effects for the derivation of an RfC. The noncancer health effects range in severity from mortality to pleural abnormalities. Five mortality studies of cohorts of workers who mined, milled, and processed Libby vermiculite identified increased risk of mortality from noncancer causes including nonmalignant respiratory disease—especially asbestosis and chronic obstructive pulmonary disease (COPD) ([Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#))—as well as cardiovascular disease ([Larson et al., 2010b](#)). Because an RfC is intended to be a level that is likely to be without appreciable risk of deleterious effects, these mortality studies were not considered as candidates for RfC derivation because other human studies exist that provide evidence of an association between LAA and less severe outcomes generally occurring at lower levels of exposure, such as parenchymal and pleural abnormalities. More detailed discussion of the choice of the critical effect for the RfC is presented in Section 5.2.2.3 and Appendix I.

Studies conducted among two cohorts of occupationally exposed workers have shown radiographic evidence of health effects on the lung and pleura (a thin tissue surrounding the lung

1 and lining the chest cavity). These effects include pleural thickening and fibrosis of the lung  
2 ([Larson et al., 2012a](#); [Larson et al., 2012b](#); [Larson et al., 2010b](#); [Rohs et al., 2008](#); [Amandus et](#)  
3 [al., 1987a](#); [McDonald et al., 1986b](#); [Lockey et al., 1984](#)). Studies of exposed community  
4 members in Libby, MT and Minneapolis, MN have also reported evidence of health effects on  
5 the lung and pleura ([Alexander et al., 2012](#); [Weill et al., 2011](#); [Muravov et al., 2005](#); [Peipins et](#)  
6 [al., 2004b](#); [Whitehouse, 2004](#); [Peipins et al., 2003](#), see Section 4.1.2).

7 Although data exist that define exposures from some activities in the Libby, MT  
8 community studies (see Section 2.3), the available exposure data were insufficient to estimate  
9 exposure at the individual level. Only studies that include exposure measurement data allowing  
10 estimation of individual exposures and identify appropriate health effects are considered for RfC  
11 derivation ([Alexander et al., 2012](#); [Larson et al., 2012a](#); [Rohs et al., 2008](#); [Amandus et al., 1987a](#);  
12 [McDonald et al., 1986b](#); [Lockey et al., 1984](#)). Among these six candidate principal studies (see  
13 Figure 5-1), one study was of the community surrounding a vermiculite processing facility in  
14 Minneapolis, MN ([Alexander et al., 2012](#)), three were occupational studies of exposed workers  
15 in Libby, MT ([Larson et al., 2012a](#); [Amandus et al., 1987a](#); [McDonald et al., 1986b](#)), and two  
16 were studies in workers from the Marysville, OH facility ([Rohs et al., 2008](#); [Lockey et al., 1984](#)).  
17 The studies by [Larson et al. \(2012a\)](#) and [Rohs et al. \(2008\)](#) represent the most recent evaluations  
18 of the occupational studies of exposed workers in Libby, MT and Marysville, OH workers,  
19 respectively, and were considered as candidate principal studies for the derivation of the RfC,  
20 along with the study of the Minneapolis community by [Alexander et al. \(2012\)](#).



**Figure 5-1. Candidate studies for derivation of the reference concentration (RfC) in three different study populations, with the most recent study of each population circled.**

Each study has adequate reporting of the studied populations, methods of assessment of health outcome(s) of interest, and statistical analyses. Each study also demonstrated associations between exposure to LAA and radiographic signs of nonmalignant respiratory effects, specifically pleural thickening (circumscribed and/or localized and/or diffuse) and small interstitial opacities (indicative of parenchymal damage) ([ILO, 2002](#), [1980](#), [1971](#)). Table 5-1 summarizes the candidate principal studies. See Section 4.1.1 for detailed study information and results.

**Table 5-1. Summary of candidate principal studies on LAA for reference concentration (RfC) derivation**

	<b>Libby, MT</b> <a href="#"><u>Larson et al. (2012a)</u></a>	<b>Marysville, OH</b> <a href="#"><u>Rohs et al. (2008)</u></a>	<b>Minneapolis, MN</b> <a href="#"><u>Alexander et al. (2012)</u></a>
<b>Study population</b>	Occupationally exposed ( <i>n</i> = 336) 93.2% male, median age 55.6 (interquartile range 47.4–65.8) yr	Occupationally exposed ( <i>n</i> = 280) 94.3% male, mean age 59.1 (age range 44–87) yr	Community residents not occupationally exposed ( <i>n</i> = 461) 52.3% male, median birth yr 1951–1960 (19.3% born ≤1940, 18.4% born ≥1960)
<b>Time of health assessment</b>	2000–2001	2002–2005	2001–2003
<b>Health outcome assessment</b>	Films independently read by two readers using 1980 ILO standards, with a third reader if the two primary readers disagreed  Film quality not reported Spirometry Self-reported respiratory symptoms	Films independently read by three board-certified radiologists (B Readers) using 2000 ILO standards  Seven employees had unreadable films and are not included in the cohort of 280 participants	Films independently read by two readers using 2000 ILO standards, with a third reader if the two primary readers disagreed  Seven participants had unreadable films
<b>Health outcomes evaluated</b>	(1) Parenchymal changes (small interstitial opacities ≥1/0)  (2) Pleural changes: LPT <sup>a</sup> ("presence of circumscribed plaque on the chest wall [as indicated on the International Labor Office form] or diaphragm without the presence of DPT or parenchymal abnormalities"); DPT (as indicated by ILO form and accompanied by costophrenic angle obliteration)  (3) Self-reported symptoms (shortness of breath, excess cough, chronic bronchitis)  (4) Spirometry: FVC, FEV <sub>1</sub> , FEV <sub>1</sub> /FVC ratio, and obstructive spirometry (defined as FVC ≥ lower limit of normal and FEV <sub>1</sub> /FVC < lower limit of normal) and restrictive (defined as FVC < lower limit of normal and FEV <sub>1</sub> /FVC > lower limit of normal)	(1) Parenchymal changes (irregular interstitial opacities, profusion score >1/0)  (2) Pleural changes: LPT (any pleural thickening with or without calcification, excluding solitary costophrenic angle blunting); DPT (any pleural thickening, including costophrenic angle blunting, with or without calcification)	(1) Parenchymal changes  (2) Pleural changes <sup>b</sup> : pleural plaques, DPT

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**Table 5-1. Summary of candidate principal studies on LAA for reference concentration (RfC) derivation (continued)**

	<b>Libby, MT</b> <a href="#"><u>Larson et al. (2012a)</u></a>	<b>Marysville, OH</b> <a href="#"><u>Rohs et al. (2008)</u></a>	<b>Minneapolis, MN</b> <a href="#"><u>Alexander et al. (2012)</u></a>
<b>Exposure Assessment</b>	1945–1993 Industrial hygiene measurements and work history (JEM); measurements made using midge impinger (pre-1970) and PCM (post-1970)	1963–1980 <sup>c</sup> Industrial hygiene measurements and work history (JEM); measurements made using PCM (1971 onwards)	1980–1989 Emissions-based modeling and self-reported activities; based on air dispersion modeling based on stack emissions and activity-based sampling
<b>Exposure levels</b>	Median: 3.6 fibers/cc-yr (IQR: 0.4–15.8)	Mean (standard deviation): 2.48 fibers/cc-yr (4.19)	Median: 2.42 fibers/cc-yr (cases) and 0.59 fiber/cc-yr (noncases)

<sup>a</sup>Although ILO 1980 guidelines were used, modifications were made such that the radiographic abnormalities were equivalent to ILO 2000 guidelines.

<sup>b</sup>Radiographic abnormalities were evaluated together as a group, and LPT was not modeled separately. However, in the lower exposure group, all 17 cases had pleural plaques (either alone or with another abnormality; personal communication from Bruce Alexander, 7 June 2013).

<sup>c</sup>Dates used in analysis by [Rohs et al. \(2008\)](#) are reported to be based on [ATSDR \(2005\)](#).

JEM = job-exposure matrix; IQR = interquartile range; PCM=phase contrast microscopy.

#### 1 **5.2.1.2. Evaluation of Candidate Studies and Selection of Principal Study**

2           The candidate studies were further evaluated in terms of quality attributes that would  
3 support their use as a principal study in the derivation of an RfC. When selecting among  
4 candidate principal studies, several factors, summarized in Table 5-2, are generally considered.

**Table 5-2. Summary of rationale for identifying candidate principal studies on LAA for reference concentration (RfC) development**

Attribute	Preferred characteristics for candidate principal studies for the Libby Amphibole asbestos RfC
Relevance of exposure paradigm	<p>Studies of subchronic or chronic duration are preferred over studies of acute exposure duration because they are most relevant to environmental exposure scenarios (potentially including both continuous exposure from ambient conditions and episodic activity-related exposures).</p> <p>When available studies observe occurrence of effect at both lower and higher doses, relatively low exposure intensities that may represent conditions more similar to environmental exposures are preferred as there may be less uncertainty in extrapolation of the results to lower exposure levels.</p>
Study design characteristics	<p>Sufficient follow-up time for outcomes to develop (this can depend on the health outcome being addressed).</p> <p>Study size and participation rates that are adequate to detect and quantify health outcomes being studied (without influential biases in study population selection) are preferred.</p> <p>Use of a study design or analytic approach that adequately addresses the relevant sources of potential confounding, including age, gender, smoking, and exposure to other risk factors (such as non-Libby asbestos).</p>
Measurement of exposure	<p>Emphasis is placed on the specificity of exposure assessment in time and place with a preference for greater detail where possible. Exposure measurements that are site and task specific provide generally preferred exposure information. Where available, individual-level measurements are generally preferred. Measurement techniques that are more specific to the agent of concern are preferred over less specific analytical methods. Better characterization of fibers is preferred. For asbestos fibers, transmission electron microscopy (TEM) analysis, which can identify the mineral fibers present, provides the most specific information; PCM identifies fibers as defined by that method (NIOSH 7400), and thus, is useful but does not confirm the mineral nature of the counted fibers. Total dust measurements are the least informative of those available.</p> <p>Stronger studies will often be based upon knowledge of individual work histories (job titles/tasks with consideration of changes over time); however, appropriate group-based exposure estimates may also be relevant.</p> <p>Exposure reconstruction and estimating exposures based on air sampling from other time periods and/or operations are less preferred methods of exposure estimation.</p> <p>Fibrosis in the pleural tissues needs time to develop and become visible on an x-ray (<a href="#">Larson et al., 2010a</a>). It has been shown that the prevalence of fibrotic lesions progresses as a function of time (<a href="#">Rohs et al., 2008</a>) and can appear long after the initial exposure (<a href="#">Lilis et al., 1991</a>). Many investigations of the exposure-response relationship for pleural plaques has found that time since first exposure (TSFE) is a significant explanatory variable (<a href="#">Paris et al., 2009</a>; <a href="#">Paris et al., 2008</a>; <a href="#">Järvholm, 1992</a>).</p> <p>Stronger studies will have data on TSFE for the relevant exposures.</p>

**Table 5-2. Summary of rationale for identifying candidate principal studies on LAA for reference concentration (RfC) development (continued)**

Attribute	Preferred characteristics for candidate principal studies for the Libby Amphibole Asbestos RfC
Measurement of effect(s)	<p>Emphasis is placed on the more sensitive health outcome endpoints that are available. For the parenchymal and pleural effects considered here, the radiographic abnormalities are more sensitive than the corresponding mortality causes. An RfC is intended to be a level at which no category of adverse health outcome would occur.</p> <p>Pleural and parenchymal abnormalities assessed using good-quality radiographs or high-resolution computed tomography and independently evaluated by multiple qualified readers according to ILO standards.</p> <p>Evaluation of radiographs should not be influenced by knowledge of exposure status.</p>

Two of the studies were conducted in occupationally exposed populations ([Larson et al., 2012a](#); [Rohs et al., 2008](#)), while the third was conducted in community residents without occupational exposure ([Alexander et al., 2012](#)). Each of the studies provided estimates of cumulative LAA exposure (in fiber/cc-yr). However, there were differences in exposure sources and intensity. Of the two occupational studies, one ([Larson et al., 2012a](#)) occurred in a setting where both occupational and nonoccupational exposures were relevant due to the close proximity of the local vermiculite mining and milling operations to the Libby, MT community. Nonoccupational exposures in the Libby, MT community were not quantified and thus were not accounted for in the overall estimates of individual exposure. In the other study ([Rohs et al., 2008](#)), exposures were generally lower and considered to be limited to the occupational setting because most of the employees showered and changed into civilian clothes at the end of the work shift. Therefore, nonoccupational exposure in the Marysville workers was assumed to be minimal. However, in both cases, the exposure estimates for earlier years are subject to uncertainty. For example, data on job and department were missing for the majority of the workers in the Libby facility hired before 1960 ([Larson et al., 2012a](#)). In the Marysville facility, no fiber measurements exist before 1972 ([Rohs et al., 2008](#)), although exposure estimates for this period were constructed based on measurements taken in subsequent years (see Appendix F). The third study, by [Alexander et al. \(2012\)](#), was conducted among Minneapolis community residents (including a higher proportion of women compared to the other studies). The researchers attempted to estimate individual community members' exposure based on facility emissions and the individual's specific activities that were considered to be related to exposure (e.g., installing or removing vermiculite insulation or playing in or around waste piles). However, exposure estimates were constructed from modeled emissions based on very sparse data from the facility's discharge stacks and activity-based exposure reconstruction, and as a result are considered to have greater uncertainties.

As noted in Table 5-2, relatively lower exposure levels are advantageous for developing an RfC (given sufficient numbers of individuals with the health effect of interest) due to uncertainties inherent in extrapolating from high-intensity (e.g., occupational) exposure levels to low-intensity (e.g., environmental) exposure levels. A limitation of the studies conducted among workers at the Libby facility is that the exposure levels experienced for some job codes are high compared with those in the other two studies (see Table 5-1; e.g., based on the interquartile range (IQR) of exposure from the 25<sup>th</sup> percentile value of 0.4 fibers/cc-yr to the 75<sup>th</sup> percentile value of 15.8 fibers/cc-yr, 25% of participants had cumulative exposures above 15.8 fibers/cc-yr). Another limitation of these studies for conducting exposure-response analysis for LPT is that many of the Libby workers were likely to have also been residents in Libby both before and during their employment at the mining and related operations, so their actual TSFE to any LAA exposure may be longer than their TSFE to occupational exposure to LAA. Therefore, data on this important variable is uncertain. Thus, the Libby workers study ([Larson et al., 2012a](#)) is less preferable for RfC derivation. The other two studies ([Alexander et al., 2012](#); [Rohs et al., 2008](#)) had generally lower exposure levels in comparison; however, greater uncertainty exists in the exposure estimates for the Minneapolis cohort because few measurements of facility emissions into the ambient air ([Adgate et al., 2011](#)). Indeed, the authors estimate that the numerical uncertainty in exposure estimates is likely to be at least an order of magnitude, perhaps much greater. Further, it is unclear whether TSFE is well characterized for the nonoccupational exposures in the Minneapolis cohort. In contrast, the study of workers at the O.M. Scott plant in Marysville, OH ([Rohs et al., 2008](#)) used exposure estimates based on extensive industrial hygiene sampling data, individual worker histories, and employee focus interviews. Thus, [Rohs et al. \(2008\)](#) is the preferred study for derivation of the RfC.

## **5.2.2. Methods of Analysis**

### **5.2.2.1. Exposure Assessment**

EPA collaborated with a research team at the University of Cincinnati to update the exposure reconstruction for use in the job-exposure matrix (JEM) for all workers in the Marysville, OH cohort, taking into account additional industrial hygiene data that were unavailable for previous studies conducted in this cohort ([Rohs et al., 2008](#); [Lockey et al., 1984](#)). Exposure estimates for each worker in the O.M. Scott Marysville, OH plant were developed based on the arithmetic mean of the available industrial hygiene data from the plant. The exposure assessment procedure is described in Appendix F. In brief, occupational exposure was estimated for each worker and adjusted to a cumulative human equivalent exposure for continuous exposure, incorporating adjustments for different inhalation rates in working versus nonworking time. These adjustments take into account the extensive seasonal changes in work hours at the Marysville facility (see Appendix F).

#### 5.2.2.2. Data Sets for Modeling Analyses

Section 5.2.1.2 describes the selection of the [Rohs et al. \(2008\)](#) cohort as the principal study. As explained below, EPA further considered the differential quality of exposure data for different years within this data set and concluded that estimation of an RfC would be improved if the primary data set for exposure-response modeling was restricted to the subset of workers hired in 1972 and later, when higher quality exposure information was available.

As described in Section 5.2.1.2, the Marysville workers evaluated by [Rohs et al. \(2008\)](#) formed the principal analytic group for derivation of the RfC, with exposure information updated and augmented by the University of Cincinnati in collaboration with EPA. As noted in Section 4.1.1.2.2 and Appendix F, the more reliable exposure estimates are considered to be those from 1972 and later, as these data were based on analytical measurements. Therefore, the primary modeling to derive a point of departure (POD) was conducted among the subgroup of workers evaluated by [Rohs et al. \(2008\)](#) that began work in 1972 or later and had no previous occupational exposure to asbestos (119 workers: 13 cases of localized pleural thickening and 106 unaffected individuals). However, information from workers who were hired before 1972, as well as from workers who were evaluated only in the earlier study by [Lockey et al. \(1984\)](#), were also considered in separate analyses (details and results of the analysis are in Appendix E). In each case, to avoid any potential bias from previous unmeasured occupational exposure to asbestos, only the data from those who did not report any previous occupational exposure to asbestos were used.

Table 5-3 and Figure 5-2 present summary characteristics for the three analytic groups. The first is the combined information for the 1980 ([Lockey et al., 1984](#)) and 2002–2005 ([Rohs et al., 2008](#)) evaluations, comprising all workers without previous exposure to asbestos; a detailed description of how these data were combined is in Appendix E. The second group is all workers evaluated in 2002–2005 without previous exposure to asbestos (as described by ([Rohs et al., 2008](#))). The third group is a subset of the workers evaluated in 2002–2005, hired in 1972 or later, without previous exposure to asbestos (primary analytic group). For the groups comprising only individuals evaluated in 2002–2005, exposure estimates covered the period from start of work through the date of job stop or at the time vermiculite ceased to be used in 2000, whichever occurred earlier.

**Table 5-3. Characteristics of workers at the O.M. Scott plant in Marysville, OH**

	All individuals evaluated in 1980 and/or in 2002–2005 <sup>a</sup>		Individuals evaluated in 2002–2005		Individuals evaluated in 2002–2005, hired in 1972 or later	
Demographic characteristics	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%
Total ( <i>n</i> )	434	100	252	100	119	100
Gender						
Male	403	92.86	236	93.65	106	89.08
Female	31	7.14	16	6.35	13	10.92
Smoking status <sup>b</sup>						
Never smoker	157	36.60	95	37.70	48	40.34
Ever smoker	272	63.40	157	62.30	71	59.66
Current	114	26.57	39	15.48	29	24.37
Former	158	36.83	118	46.83	42	35.29
	Mean (SD)	Median (25 <sup>th</sup> –75 <sup>th</sup> percentiles)	Mean (SD)	Median (25 <sup>th</sup> –75 <sup>th</sup> percentiles)	Mean (SD)	Median (25 <sup>th</sup> –75 <sup>th</sup> percentiles)
Age at x-ray (yr)	50.73 (14.88) Range: 19–86	52 (43–60)	58.66 (10.53) Range: 42–86	56 (50–66)	52 (7.1) Range: 42–82	50 (47–55)
Time from first exposure (yr)	24.42 (13.59) Range: 0.42–47.34	25.96 (11.75–34.77)	34.40 (7.12) Range: 23.14–47.34	33.51 (28.70–38.47)	28.24 (2.54) Range: 23.14–32.63	28.39 (25.81–30.29)
Exposure duration (yr)—duration of exposed time (i.e., accounting for gaps)	18.93 (11.44) Range: 0.41–44.00	20.75 (8.75–27.41)	24.96 (10.17) Range: 0.67–44.00	26.46 (19.75–32.17)	18.23 (8.61) Range: 0.67–29.00	21.75 (9.50–25.59)
Body mass index <sup>b</sup>	30.80 (6.25) Range: 17.30–61.97	29.44 (26.93–33.33)	30.80 (6.25) Range: 17.30–61.97	29.44 (26.93–33.33)	31.30 (6.90) Range: 20.08–61.97	30.11 (27.23–33.85)
Cumulative exposure (fiber/cc-yr)	7.9232 (17.9598) Range: 0.003–96.91	1.1252 (0.3414–3.7684)	8.75 (19.12) Range: 0.005–96.91	1.26 (0.51–5.20)	1.439 (2.5479) Range: 0.005–17.33	0.5048 (0.2188–1.5519)
Mean exposure (fiber/cc)	0.3733 (0.7942) Range: 0.007–4.34	0.0566 (0.0267–0.2364)	0.31 (0.65) Range: 0.007–4.10	0.05 (0.02–0.20)	0.0716 (0.1239) Range: 0.007–0.77	0.0234 (0.0133–0.074)
Residence time-weighted (RTW) exposure (fiber/cc-yr <sup>2</sup> ) <sup>c</sup>	193.3093 (519.3874) Range: 0.0007–3500.66	19.4767 (4.2550–78.0944)	294.38 (687.95) Range: 0.12–3500.66	34.31 (11.07–154.36)	33.7415 (69.2231) Range: 0.12–474.01	10.2075 (3.9055–29.1246)

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**Table 5-3. Characteristics of workers at the O.M. Scott plant in Marysville, OH (continued)**

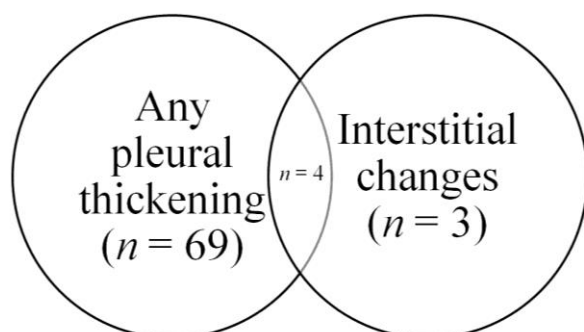
<sup>a</sup>See Appendix E for details of how the individual health outcome data for all workers who participated in the [Lockey et al. \(1984\)](#) study and the follow-up study by [Rohs et al. \(2008\)](#) were combined.

<sup>b</sup>Data on smoking status were missing for five individuals in the full cohort. Data on body mass index (BMI) was unavailable for 216 individuals in the full cohort, 34 individuals examined in 2002–2005, and 21 individuals examined in 2002–2005 who were hired in 1972 or later.

<sup>c</sup>RTW exposures are calculated using midpoint of each work season.

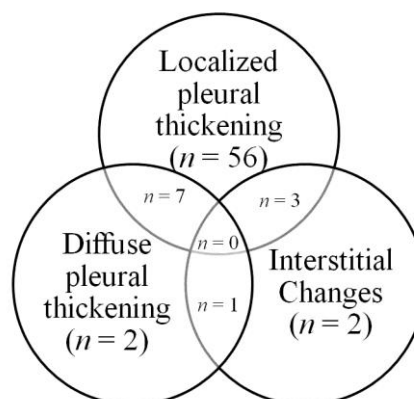
**All individuals evaluated in 1980 and/or in 2002-2005**

Total  $n = 434$



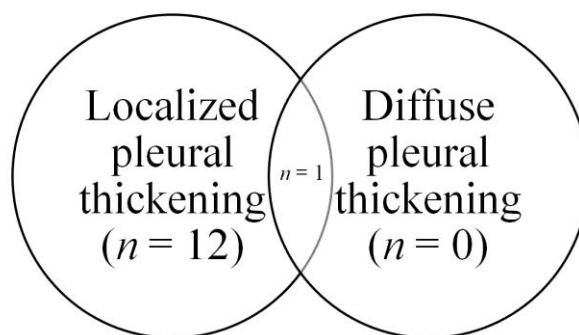
**Individuals evaluated in 2002-2005**

Total  $n = 252$



**Individuals evaluated in 2002-2005, hired in 1972 or later**

Total  $n = 119$



**Figure 5-2. Radiographic outcomes among Marysville, OH workers.**

Numbers of individuals in each category are exclusive (e.g., there are 69 individuals among the total  $n = 434$  with pleural thickening only, and an additional four individuals have pleural thickening in addition to interstitial changes).



As stated previously, fiber measurements started in the Marysville plant in 1972, and exposures before this time were estimated by University of Cincinnati scientists, based on focus group interviews with 15 long-term former workers and the times when engineering changes were made to control dust in the facility (see Appendix F). Exposure estimates for the period before 1972 can be considered less certain compared with those estimates more directly based on industrial hygiene data. The University of Cincinnati analysis assumed that early exposure levels in the plant are twice those measured in 1972 (see Appendix F). The greater uncertainty of the pre-1972 exposure estimates led to EPA's decision to focus the analysis on the group of workers hired in 1972 or later. Although it is generally true that the use of more data is an advantage for statistical analyses because it allows for the computation of more statistically precise effect estimates, this increased precision can be offset by a negative impact on the accuracy of the effect estimate if an increase in sample size is accompanied by greater exposure misclassification or other biases.

In summary, the primary analytic group was the Marysville workers evaluated by [Rohs et al. \(2008\)](#) who were hired in 1972 or later; however, additional information from workers hired before that date was also used in modeling and sensitivity analyses.

#### **5.2.2.3. Selection of Critical Effect**

A critical effect is defined as "The first adverse effect, or its known precursor, that occurs to the most sensitive species as the dose rate of an agent increases" ([U.S. EPA, 2011](#)). Three endpoints are suitable for consideration as critical effects for the derivation of an RfC for LAA where health effects data and exposure information are available in the principal study ([Rohs et al., 2008](#)): (1) parenchymal changes viewed as small interstitial opacities in the lung, (2) localized pleural thickening (LPT), or (3) diffuse pleural thickening (DPT) as defined in ILO (2000). Each of these represents persistent changes to normal tissue structure.

Small interstitial opacities (asbestosis) are widely accepted as adverse; the American Thoracic Society (ATS) states that "asbestosis is usually associated with dyspnea, bibasilar rales, and changes in pulmonary function: a restrictive pattern, mixed restrictive-obstructive pattern, and/or decreased diffusing capacity" ([ATS, 2004](#)). Similarly, DPT is also widely accepted as adverse, with the ATS stating that "decrements associated with diffuse pleural thickening reflect pulmonary restriction as a result of adhesions of the parietal with the visceral pleura. Restrictive impairment is characteristic, with relative preservation of diffusing capacity (pattern of entrapped lung)" ([ATS, 2004](#)).

Statements from the consensus groups vary as to whether pleural plaques (a subset of LPT) impact lung function. Regarding pleural plaques, the ATS notes that this endpoint is also associated with decrements in lung function:

1 Although pleural plaques have long been considered inconsequential markers of  
2 asbestos exposure, studies of large cohorts have shown a significant reduction in  
3 lung function attributable to the plaques, averaging about 5% of FVC, even when  
4 interstitial fibrosis (asbestosis) is absent radiographically... The presence of  
5 circumscribed plaques can be associated with restrictive impairment and  
6 diminished diffusing capacity on pulmonary function testing, even in the absence  
7 of radiographic evidence of interstitial fibrosis. ([ATS, 2004](#)).  
8

9 However, the statement goes on to note that findings of significant pulmonary deficits are  
10 not consistent, and that “most people with pleural plaques alone have well-preserved lung  
11 function.” In addition to the ATS document, the American College of Chest Physicians ([ACCP;](#)  
12 [Banks et al., 2009](#)) published a Delphi study conducted to gauge consensus among published  
13 asbestos researchers, and found that these researchers statistically rejected the statement that  
14 “Pleural plaques alter pulmonary function to a clinically significant degree” (although noting that  
15 some researchers strongly agreed with the statement, and the response rate was relatively low at  
16 <40%). Therefore, EPA undertook a systematic review to evaluate the magnitude and extent of  
17 the pulmonary function deficits associated with LPT, described in Appendix I. The review  
18 demonstrates that these deficits can be considered adverse. Based on the association of LPT with  
19 pulmonary function decrease, LPT is an appropriate health effect for derivation of an RfC.  
20 Because interstitial opacities, DPT, and LPT are all appropriate candidate endpoints, the critical  
21 effect was chosen as that which is the first to appear, or which occurs at the lowest levels of  
22 exposure. A summary of the systematic review of the pleural plaque data is discussed below.

23 [Larson et al. \(2012a\)](#) evaluated the timing of appearance and exposure levels at which  
24 pleural and parenchymal abnormalities occur on chest radiographs of vermiculite workers at the  
25 Libby facility relative to hire date (i.e., time since first occupational exposure). In this  
26 retrospective analysis, the study authors reported that the health endpoint with the shortest  
27 median time to appearance was circumscribed pleural plaques (a subset of LPT) with a median  
28 latency of 8.6 years, compared to median latency times of 27.0 years for DPT and 18.9 years for  
29 parenchymal changes (small interstitial opacity profusion scores of 1/0 or greater). Although all  
30 workers experienced generally high exposure, cumulative fiber levels were lowest for those with  
31 circumscribed pleural plaques (median of 44.1 fibers/cc-yr), compared to those with DPT  
32 (median of 317.8 fibers/cc-yr) and highest for those with parenchymal changes (median of  
33 235.7 fibers/cc-yr for those with major profusion Category 1 abnormalities, 678.4 for  
34 Category  $\geq 2/1$ , and 1,303.4 for Category  $\geq 3/2$ ). Similarly, [Rohs et al. \(2008\)](#) found that for all  
35 workers in that study, on average the cumulative fiber exposure for those workers with LPT only  
36 (3.45 fibers/cc-yr) was lower compared to those with DPT only (8.99 fibers/cc-yr) or with any  
37 interstitial changes (alone or with either LPT or DPT; 11.86 fibers/cc-yr). These results indicate  
38 that LPT may be the most sensitive of the effects examined, as the radiographic outcome most  
39 likely to occur soonest after first occupational exposure, and the outcome most likely to appear at

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1 relatively lower cumulative exposure levels. The clinical perspective suggests that pleural  
2 plaques do not clinically impede lung function for most people. This perspective was stated by  
3 the American College of Chest Physicians ([Banks et al., 2009](#)) regarding the 2004 ATS  
4 statement: “Data were cited showing that large studies of workers with pleural plaques had  
5 approximately a 5% mean decline in FVC compared to asbestos workers without pleural plaques.  
6 In this report, the experts concluded that the presence of pleural plaques did not decrease lung  
7 function to a significant extent”—that is, they concluded that the observed decrements were not  
8 clinically significant to an individual patient. EPA’s systematic review of the literature and  
9 formal meta-analysis found decrements in the same range—statistically significant decreases in  
10 FVC of 4.09% (4.08% when studies without limitations are used) and in FEV<sub>1</sub> of 1.99% (3.87%  
11 when studies without limitations are used). In addition, although few of the studies evaluated  
12 reported results for diffusing capacity (as evaluated by DL<sub>CO</sub>, the diffusing capacity of the lung  
13 for carbon monoxide), these studies did observe statistically significant or nearly significant  
14 decreases between those with no radiographic abnormalities and those with LPT or pleural  
15 plaques.

16 As stated by [ATS \(2004\)](#), the majority of individuals with pleural plaques (subset of  
17 LPT) may have well-preserved lung function. However, this may not be the case for individuals  
18 who are already at the lower end of the “normal” range of function, already have compromised  
19 function, or have increased vulnerability or susceptibility due to other factors (such as chronic  
20 disease, other environmental exposures, smoking, etc.). For any of these individuals, even a  
21 small decrease in lung function may be important, but once averaged into the whole study  
22 population (i.e., looking at only average changes in the whole group) the sensitive individuals’  
23 contribution to the population-wide change in mean pulmonary function measures is muted.

24 Accordingly, there is a difference in considering what is significant from a clinical  
25 perspective compared to an epidemiological perspective. The clinician’s focus is the individual  
26 patient, and decisions made in that context (i.e., benefits/risks of medical treatments or tests). In  
27 contrast, the population-level (risk assessment) perspective considers any changes in the  
28 population distribution of pulmonary function and the potentially increased risks of adversity to  
29 subpopulations of the general population. When considering an entire population with a  
30 distribution of lung function parameters, even small changes in the average of that distribution  
31 means that a much larger proportion of the exposed population is shifted down into the lower  
32 “tail” of the lung function distribution. This line of thinking is well understood in the recent  
33 examples of lead and IQ ([U.S. EPA, 2013a](#)) and respiratory function and ozone ([U.S. EPA,](#)  
34 [2013b](#)). Early childhood exposure to lead can lead to decrements in intelligence as measured by  
35 IQ. Depending on the exposure level to lead, a mean deficit of 2 IQ points would not be  
36 measurable nor lead to a clinical finding of harm in individuals, but from an epidemiologic or  
37 population-level perspective, a downward shift in a portion of the entire IQ distribution by 2 IQ  
38 points would be expected to push many individuals already in deficit further into a more clearly

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1 “adverse” state. Similarly, even small decrements in lung function on the population level could  
2 push “borderline” individuals into a state of clinically significant decreased lung function.

3 In addition, [ATS \(2004\)](#) stated “The presence of plaques is associated with a greater risk  
4 of mesothelioma and of lung cancer compared with subjects with comparable histories of  
5 asbestos exposure who do not have plaques.” While references provided in the [ATS \(2004\)](#)  
6 statement ([Hillerdal and Henderson, 1997](#); [Hillerdal, 1994b](#)) do not directly address the [ATS](#)  
7 [\(2004\)](#) statement, a recent large (5,287 retired workers, 17 mesothelioma cases) study ([Pairol et](#)  
8 [al., 2013](#)) found a statistically elevated risk of mesothelioma in a group with plaques (parietal  
9 and diaphragm) compared to a no-plaques group, using computer tomography (CT). The study  
10 authors found that, after adjusting for cumulative exposure index and TSFE, the risk of  
11 mesothelioma in the plaques (parietal or diaphragm) group was statistically elevated (hazard rate  
12 (HR) = 6.8, 95% CI 2.2–21.4) compared to the risk of mesothelioma in the exposed workers  
13 without pleural plaques.

14 In the Marysville workers evaluated in 2002–2005, differences in exposure patterns are  
15 also apparent among outcome groups (see Table 5–4). Exposure to LAA was lower among those  
16 with no radiographic abnormalities compared to those with LPT and those with DPT and/or  
17 interstitial opacities. Of the candidate critical effects, LPT has the shortest TSFE, and is more  
18 likely to appear at lower levels of LAA exposure. LPT is associated with adverse decrements on  
19 pulmonary function. Thus, LPT is selected as the critical effect for RfC derivation.

**Table 5-4. Characteristics of workers at the O.M. Scott plant in Marysville, OH, with health evaluations in 2002–2005**

	No radiographic abnormalities	All LPT cases (with or without DPT/interstitial changes)	LPT cases without DPT/interstitial changes	LPT cases with DPT/interstitial changes	DPT/interstitial changes cases, without LPT
N	181	66	56	10	5
Time (years) since first exposure (TSFE), range	23.14–47.34	24.46–47.30	24.46–47.30	31.52–42.22	36.15–45.56
TSFE, median (interquartile range, IQR)	31.14 (27.56–36.30)	38.21 (34.38–45.81)	37.41 (34.36–45.53)	42.03 (37.58–46.22)	37.22 (37.16–45.04)
Mean exposure (f/cc), range	0.0067–3.7396	0.0068–4.1000	0.0068–2.6230	0.0571–4.1000	0.0692–3.0463
Median (IQR)	0.0372 (0.0167–0.0943)	0.1634 (0.0431–0.7532)	0.0953 (0.0421–0.3958)	1.9197 (0.6623–2.2848)	0.2378 (0.0954–1.7186)
Cumulative exposure (f/cc-yr), range	0.0050–95.0386	0.0233–96.9072	0.0233–96.5450	1.9080–96.9072	1.7986–81.4815
Median (IQR)	1.0188 (0.3162–2.0743)	4.5210 (1.3355–22.9072)	3.1710 (1.2472–10.0445)	47.3994 (22.9072–61.4999)	3.6659 (2.2890–28.2097)
Residence time-weighted (RTW) exposure (f/cc-yr <sup>2</sup> ), range	0.1196–3468.28	0.5313–3500.66	0.5314–3477.33	65.3771–3500.66	62.7209–3011.68
Median (IQR)	23.7297 (7.6204–50.7792)	127.4075 (36.5668–770.4642)	88.2670 (34.8638–302.5802)	1693.86 (770.4642–2365.89)	120.5295 (74.3389–944.9676)

Table 5-4 and Figure 5-2 both highlight a complexity in that these radiographic changes are not mutually exclusive—individuals may have one or more changes simultaneously, in any combination. Among the 66 individuals with LPT, 10 also had DPT or interstitial opacities, and these 10 individuals are noticeably different with regards to TSFE and exposure compared to LPT cases without other radiographic changes, consistent with the results of [Larson et al. \(2012a\)](#). When restricting to the subgroup of individuals hired in 1972 or later, there are 106 individuals with no radiographic abnormalities, 12 individuals with LPT only, and one individual with both LPT and DPT. The individual with both LPT and DPT had a TSFE of 31.52 years, similar to the median TSFE of 29.71 years among the 12 individuals with LPT only. However, this individual had higher estimated mean exposure (0.46 fiber/cc, compared to a median of 0.08 fiber/cc) and cumulative exposure (9.13 fibers/cc-yr, compared to a median of 1.82 fibers/cc-yr), compared to the other 12 LPT cases. The primary analysis considers as the critical effect all LPT cases together, contrasted to those without radiographic abnormalities, but the effect of separating out those with multiple radiographic outcomes is examined in the sensitivity analyses (see Section 5.3.5).

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In addition, an alternative critical effect of “any pleural thickening” (APT) is considered. Note that in the case of the subcohort of workers evaluated in 2002–2005 and hired in 1972 or later, this definition is equivalent to a critical effect of LPT because no individuals had DPT alone.

#### **5.2.2.4. Selection of Explanatory Variables to Include in the Modeling**

As with the mode of action (MOA) for carcinogenicity, the MOA for LPT and the results of other asbestos epidemiology studies could potentially inform noncancer modeling decisions and suggest exposure metrics to use in modeling. The following text discusses the plausibility of exposure metrics proposed in the MOA/epidemiology literature.

As noted in Section 4.4, important considerations in evaluating the available mechanism and MOA data are fiber characteristics, route of exposure, dose metric, as well as study design and interpretation. Specific fiber characteristics impact the fiber toxicokinetics (reviewed in Section 3), and in turn, the biologic response to fibers. Fiber dimensions play a role in translocation, a clearance mechanism that may lead to inhaled fibers moving from the lung to the pleura. Data gaps still remain to determine specific mechanisms involved in LAA-induced pleural disease. The review of studies in Section 4.4 clearly highlights the need for more controlled studies examining LAA in comparison with other forms of asbestos and for examining multiple endpoints—including reactive oxygen species (ROS) production and proinflammatory gene expression alterations—to improve understanding of mechanisms involved in noncancer health effects. Although research demonstrates that the LAA has biologic activity consistent with the inflammatory action and cytotoxic effects seen with other forms of asbestos, the conclusion of Section 3 of this assessment is that the data are not sufficient to establish an MOA for the pleural and/or pulmonary effects of exposure to LAA.

A general understanding of the biology and the epidemiology of LPT can still inform the modeling as to which explanatory variables should be considered in the models, how the variables should be considered or statistically parameterized, and whether they should be retained in the model. From a general understanding of the respiratory effects of asbestos, the intensity of exposure (i.e., concentration), the duration of exposure, and the timing of exposure in relation to subsequent diagnosis of LPT (i.e., TSFE) have been shown to be univariate predictors of pleural plaques (a subset of LTP) in multiple epidemiologic studies as discussed below and, therefore, merit specific consideration in this modeling effort.

#### *Timing of exposure*

Fibrosis in the pleural tissues needs time to develop and become visible by x-ray ([Larson et al., 2010a](#)). It has been shown that the prevalence of fibrotic lesions progresses as a function of time ([Rohs et al., 2008](#)) and can appear long after the initial exposure ([Lilis et al., 1991](#)). Many investigations of the exposure-response relationship for pleural plaques has found that

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TSFE is a significant explanatory variable ([Paris et al., 2009](#); [Paris et al., 2008](#); [Järvholm, 1992](#)). This suggest that TSFE should be considered as a potential explanatory variable in the modeling. It is important to understand that even when an individual explanatory variable may be an important univariate predictor of the risk of LPT, in more complex modeling with two or more explanatory variables, the relationship observed in univariate modeling may no longer hold. One reason for this is that two variables can be highly correlated in many occupational cohorts, thus, the regression modeling may indicate that, for the data at hand, there is more unique information to explain the risk of LPT in one variable than in the other.

#### *Intensity of exposure*

A general understanding of toxicology suggests that, for a given duration of exposure, exposure at higher intensities (concentrations) will likely results in higher burdens of fibers in the alveolar region of the lung, and potentially in the pleural tissue as well. Therefore, for a given duration of exposure, there is a reasonable expectation that people exposed at higher intensities of LAA would experience greater risk of being diagnosed with LPT than people exposed at lower exposure intensities. Similarly, from general principles, at a given intensity of exposure, greater duration of exposure results in higher tissue concentrations of fibers. Epidemiologic evidence from a large cohort of asbestos-exposed workers has reported that exposure intensity (concentration) can be an important predictor of being diagnosed with pleural plaques (a subset of LPT)—even in a multivariate model along with TSFE ([Paris et al., 2009](#); [Paris et al., 2008](#)). [Järvholm \(1992\)](#) fit a mathematical model for the incidence of pleural plaques based on concentration and TSFE, which the author considered to have a biological interpretation. This suggests that exposure intensity should be considered as a potential explanatory variable in the modeling.

#### *Duration of exposure*

Important characteristics of amphibole fibers are their biodurability and biopersistence. Due to the slow clearance of amphibole fibers from the lung, the fiber burden in the alveolar region of the lung is expected to increase for a given exposure intensity as the duration of exposure increases. This may be true of the pleural tissues as well—but little scientific information is available on the time course of potential fiber accumulation in pleura. Amphibole fibers may remain biologically active for many years while the fibers are in residence in the tissues, although this biological activity may vary with time. For example, depending on the composition and structure of the fiber, certain fibers may cease to have surface activity in biological media, or may have different biological activity in this media ([Pezerat, 2009](#)). Further, some asbestos fibers can become covered with an iron-protein coat in the lungs of exposed individuals (*i.e., forming ferruginous bodies*; [Dodson et al., 1993](#); [Churg and Warnock, 1981](#)). The biological effect of this coating is unclear, but may alter the activity of the fibers.

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Epidemiologic evidence from studies of asbestos-exposed workers ([e.g., Clin et al., 2011](#)) indicates that cumulative exposure, a metric of exposure that encompasses both exposure intensity as well as duration, can be an important predictor of the probability of being diagnosed with pleural plaques. This suggests that duration of exposure should be considered in modeling. Cumulative exposure (as an expression of concentration and duration) should be considered as a potential explanatory variable in the modeling. Therefore, modeling results using both exposure intensity, C, and cumulative exposure, CE, as the exposure metric are considered. Another exposure metric related to both exposure intensity and duration is called residence time-weighted exposure, a metric of exposure that can be used to more heavily weigh earlier exposures. Residence time-weighted exposure is also considered for modeling.

#### *Other explanatory variables*

Other explanatory variables of interest include those that may be confounders of the explanatory variables' statistical relationships with the risk of LPT. These include body mass index (BMI), age, and smoking (complete list from the table of potential confounders). Each of these was assessed as a potential confounder prior to modeling the main explanatory variables of interest.

#### **5.2.2.5. Selection of the Benchmark Response**

Selecting a benchmark response ([BMR](#)) involves making judgments about the statistical and biological characteristics of the data set and about the applications for which the resulting benchmark concentration (BMCs)/lower limit of the BMC (BMCLs) will be used. An extra risk of 10% is recommended as a standard reporting level for quantal data. Biological considerations may warrant the use of a BMR of 5% or lower for some types of effects (e.g., frank effects), or a BMR greater than 10% (e.g., for early precursor effects) as the basis of the POD for a reference value ([U.S. EPA, 2012](#)).

LPT is a persistent change to normal tissue structure and is associated with a decrement in lung function on a population level (~5 and ~2.5% decrements in percentage predicted FVC and FEV<sub>1</sub>, respectively). [Larson et al. \(2012a\)](#) showed a statistically significant increased risk of people with LPT having “restrictive spirometry” and concluded that this abnormality may result in lung function impairment. However, the available data do not lead EPA to conclude LPT should be considered a frank effect and thus EPA selects a BMR of 10% extra risk for this endpoint.

As noted in Section 5.2.3.1, an alternative critical effect of APT was also considered as an alternative analysis. For this outcome, a BMR of 10% was also used, given that (as shown in Figure 5-2) a significant majority of cases were LPT.

#### 5.2.2.6. *Exposure-Response Modeling*

LPT was selected as the critical effect based on the adverse health effects associated with pleural thickening specific to this diagnosis ([ILO, 2002](#)). Note that for the primary analytic data set (workers evaluated in 2002–2005 and hired in 1972 or later), the number of individuals with LPT is the same as the number with either “any pleural thickening” or “any radiographic change” (i.e., there is no difference in the number of cases or the estimates of risk) because the single case with DPT also had LPT (and thus is included as an LPT case) and no cases of interstitial abnormalities occurred. However, in the larger cohort of workers (all workers, and those evaluated in 2002–2005 regardless of hire date), there were individuals with these more severe outcomes as well as LPT (see Figure 5-2).

The exposure-response relationship was modeled as described below, and PODs were estimated using BMC methodology. For inhalation data, the BMC is defined as the exposure level that results in a specified BMR. The RfC is derived from the lower 95% confidence limit of the BMC, referred to as the BMCL, which accounts for statistical uncertainty in the model fit to the data. All analyses were performed using SAS<sup>®</sup> statistical software v. 9.3. BMCLs were obtained by the profile likelihood method as recommended by [Crump and Howe \(1985\)](#) using the nonlinear mixed modeling procedure (PROC NL MIXED) in SAS ([Wheeler, 2005](#)).

**5.2.2.6.1. *Considerations of appropriate model forms and explanatory variables.*** The process and considerations for exposure-response modeling of the Marysville data were guided by EPA’s 2012 Benchmark Dose Technical Guidance ([U.S. EPA, 2012](#)). As outlined in that document, there are several stages of exposure-response modeling. Once the appropriate data set(s), endpoint(s), explanatory variables(s), and BMR are determined, the next step is to choose an appropriate statistical model form or set of model forms to evaluate (e.g., logistic, probit, Dichotomous Hill, etc.). Among this set of models, the overall model fit and the fit in the region of the BMR are evaluated to determine which models adequately represent the data. Finally, one or more models are selected from the group of adequately fitting models to derive a POD for the reference value. Regarding the selection of models to evaluate, the Benchmark Dose Technical Guidance (see p. 26) states: “The initial selection of a group of models to fit to the data is governed by the nature of the measurement that represents the endpoint of interest and the experimental design used to generate the data. In addition, certain constraints on the models or their parameter values sometimes need to be observed and may influence model selection.” In the Marysville data, a number of factors must be considered to determine an appropriate modeling strategy: the nature of the data set, ability to estimate the effects of exposure and of important covariate(s), the existence of a plateau or theoretical maximum response rate in a population, and the ability to estimate a background rate of the outcome in a population. Each factor is described below, and consideration of these factors in total resulted in a preference for

1 the Dichotomous Hill model, with a set of additional model forms suitable for evaluation of  
2 sensitivity to model selection.

- Nature of the data set: For the Marysville workers data set, the outcome data are dichotomous (presence or absence of an effect), and thus, appropriate models are those suitable for dichotomous endpoints. The Marysville workers underwent radiographic evaluation in 2002–2005 to ascertain the presence or absence of radiographic abnormalities (i.e., prevalence data). Radiographic outcomes are coded as present or absent, leading to a dichotomous response structure. Appropriate models for this type of data include models such as logistic, probit, log-logistic, log-probit, Dichotomous Hill, and Michaelis-Menten (see Table 5-5). Goodness of model fit for these models may be evaluated using the Hosmer-Lemeshow goodness-of-fit statistic ([Hosmer and Lemeshow, 2000](#)); a low  $p$ -value ( $<0.05$ ) indicates poor fit, while a higher  $p$ -value indicates adequate fit. Note that the computation of this statistic involves dividing the data set into bins, based on the predicted probability of the (dichotomous) outcome. The standard procedure is to use 10 bins (i.e., deciles), and this approach was used for all analyses shown here.

**Table 5-5. Models<sup>a</sup> considered to develop a point of departure (POD)**

Name	Equation	Fitting parameters	
		N	Description
Logistic	$p(x) = \frac{1}{1 + \exp[-a - b \times x]}$	2	$a$ = Intercept $b$ = Slope
Probit	$p(x) = \Phi(a + b \times x)$	2	$a$ = Intercept $b$ = Slope
Log-logistic	$p(x) = bkg + \frac{1 - bkg}{1 + \exp[-a - b \times \ln(x)]}$	3	$a$ = Slope $b$ = Shape $bkg$ = Background
Log-probit	$p(x) = bkg + (1 - bkg)\Phi(a + b \times \ln(x))$	3	$a$ = Slope $b$ = Shape $bkg$ = Background
Dichotomous Hill	$p(x) = bkg + \frac{Plateau - bkg}{1 + \exp[-a - b \times \ln(x)]}$	4	$a$ = Slope $b$ = Shape $bkg$ = Background Plateau
Michaelis-Menten (or Dichotomous Hill model, with $b$ fixed at 1)	$p(x) = bkg + \frac{Plateau - bkg}{1 + \exp[-a - \ln(x)]}$	3	$a$ = Slope $bkg$ = Background Plateau
Bivariate Dichotomous Hill with time ( $T$ )	$p(x, T) = bkg + \frac{Plateau - bkg}{1 + \exp[-a - b \times \ln(x) - c \times T]}$	4	$a$ = Slope of exposure $b$ = Shape $c$ = Slope of time $bkg$ = Background

<sup>a</sup>Equations used to derive the BMC for each model from *Benchmark Dose Technical Guidance* ([U.S. EPA, 2012](#)) are shown below:

Logistic:  $BMC = -\ln [(1 - BMR)/(1 + BMR \times \exp(-a))]/b$

Probit:  $BMC = [\Phi^{-1}(BMR \times (1 - \Phi(a)) + \Phi(a)) - a]/b$

Log-logistic:  $BMC = \exp[((-\ln((1/BMR) - 1)) - a)/b]$

Log-probit:  $BMC = \exp[(\Phi^{-1}(BMR) - a)/b]$

Michaelis-Menten:  $BMC = \exp[(-\ln((Plateau - bkg)/((1 - bkg) \times BMR) - 1) - a)]$

Dichotomous Hill:  $BMC = \exp[(-\ln((Plateau - bkg)/((1 - bkg) \times BMR) - 1) - a)/b]$

Dichotomous Hill with time( $T$ ) covariate:  $BMC = \exp [(-\ln((Plateau - bkg)/((1 - bkg) \times BMR) - 1) - a - c \times T)/b]$

- Effect of exposure: Because the data set include estimates of individual exposure and the goal is to derive an RfC, appropriate models need to include an independent exposure variable. All the models listed above can estimate the effect of changes in exposure on risk of the outcome, although the parameter that reflects the magnitude of that effect changes across models. In models where exposure is included without logarithmic transformation, the  $b$  parameter corresponding to exposure is interpreted as a “slope” and represents the change in outcome per unit change in exposure. In models where exposure is natural log transformed, the interpretation is somewhat different; both the  $a$  and  $b$  parameters determine the shape of the exposure-response relationship (e.g.,  $a + b \times \ln(x) = \ln(\exp(a) \times x^b)$ ). Thus,  $b$  is a power parameter and behaves more like a shape parameter in this context, while  $\exp(a)$  behaves like a

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traditional slope. This is an important distinction in interpreting models including natural log transformation of the exposure (i.e., log-logistic, log-probit, Michaelis-Menten, Dichotomous Hill).

- Plateau in models: Some model forms have an explicit parameter representing a plateau (e.g., Dichotomous Hill and Michaelis Menten), which is an asymptotic quantity interpretable as the maximum prevalence of the outcome that would ever be observed at very high levels of the predictor variables in the model (e.g., high levels of LAA exposure). In contrast, certain model forms (e.g., log-logistic and log-probit) do not estimate a separate plateau parameter, but instead have maximum asymptotic values of 100%. EPA wanted to understand whether the considered models implied an asymptotic maximum incidence for the endpoint (i.e., a “plateau”) and to evaluate the sensitivity to alternative specified values for that plateau. Thus, EPA selected a model with an explicit plateau term for the option of fitting that plateau term to the data and for the ability to do sensitivity analysis with alternative fixed plateau values to evaluate the sensitivity of results.
- Plateau—further considerations: For models with an explicit plateau parameter, EPA considered whether to let the plateau term be fit to the data or to select a fixed value prior to fitting the model. As described below, EPA chose to fix the value of the plateau prior to fitting these models to better consider a broader set of data on pleural thickening. However, EPA also conducted sensitivity analysis on the impact of this assumption for model results. Importantly, this plateau parameter of an asymptotic maximum prevalence cannot be directly observed in, and is not well estimated from the Marysville data because none of the workers experienced high enough exposure and follow-up. In the group of workers defined for primary exposure-response analysis (i.e., those hired in or after 1972 with radiographs performed in 2002–2005), the TSFE averaged 28.4 years and ranged from 23.14 to 32.63 years. Exposure-response models that include TSFE or otherwise incorporate the timing between exposure periods and observation, such as models using the residence time-weighted (RTW) exposure metric, could allow for the estimation of a plateau, but the limited data on the effect of elapsed time in the workers hired in or after 1972 does not support a reliable estimate of the asymptotic maximum prevalence. In addition, standard radiographs may not have perfect sensitivity or specificity to identify the outcomes of interest (thus “observed” prevalence may differ from “actual” prevalence). Models that do not include time from exposure to the x-ray observation would be estimating a plateau that might similarly be extrapolating on dose and might not appropriately estimate the impact of a longer follow-up period. For the RfC, the question is what happens when individuals are exposed over a lifetime (assumed to be 70 years). This may be difficult to answer if a given model results in a plateau significantly lower than what might result from sufficient duration or follow-up time. One option is to fix the plateau parameter at a value informed by the existing literature on observed prevalence in populations that had higher exposures and longer TSFE values. In a cross-sectional study of Libby workers and residents seen at a clinic in Libby, MT, [Winters et al. \(2012\)](#) observed a prevalence of 76% for pleural thickening, although the maximum TSFE was not known. Previous studies in populations exposed to asbestos (potentially amphibole and/or nonamphibole) have reported prevalences of pleural thickening of 82.4% among U.S. insulators with  $\geq 40$  years since first exposure ([Lilis et al., 1991](#)) and prevalence of

1 pleural plaques of 85.7% among Swedish shipyard workers with 50–54 years since  
2 first exposure ([Järvholm, 1992](#)). Thus, a reasonable option would be to use the  
3 Dichotomous Hill or Michaelis-Menten model and fix the plateau term at a value (i.e.,  
4 85%) consistent with maximum observed prevalence rates in the asbestos literature.  
5 EPA performed a sensitivity analysis (see Section 5.3.4) to evaluate the effect of  
6 assumptions regarding the plateau on the POD.

- 7 • Effect of covariates: As with the discussion for effect of exposure, a desirable model  
8 attribute is the ability to estimate the effect of additional covariates. EPA evaluated a  
9 variety of possible covariates and determined that in the primary analytic data set of  
10 individuals evaluated in 2002–2005 and hired in 1972 or later, none showed evidence  
11 of potential confounding of the LAA exposure-LPT relationship. Each of the models  
12 listed above (logistic, probit, log-logistic, log-probit, Dichotomous Hill, and  
13 Michaelis-Menten) allow for the inclusion of covariates. Specifically for modeling  
14 LAA exposure and risk of LPT, one of the most important covariates to consider is  
15 TSFE. As described above, the prevalence of pleural plaques (a subset of LPT) has  
16 been shown to increase as TSFE increases, even in the absence of continued asbestos  
17 exposure. Although the literature indicates TSFE is the most important time-related  
18 factor, other factors may be important to consider, including age at examination, hire  
19 year, job tenure (time elapsed from job start to job stop), and exposure duration  
20 (taking into account gaps in exposure). There are also nontime-related factors which  
21 may influence the association between LAA exposure and risk of LPT. These include  
22 gender, smoking status, and BMI. Smoking is a particularly important variable to  
23 consider when evaluating respiratory health outcomes. Each of these factors was  
24 investigated in the primary data set. To be a potential confounder, the factor must be  
25 associated with both LAA exposure and LPT, and must not be an intermediate in the  
26 causal pathway between exposure and outcome. The association with natural  
27 log-transformed LAA exposure in the subcohort was assessed using a linear  
28 regression model, and the association with LPT was assessed using a logistic  
29 regression model (see Table 5-6). While many of the time-related factors (with the  
30 exception of age at x-ray examination) as well as male gender and former smoker  
31 status were associated with each of the three exposure metrics, none were associated  
32 with risk of LPT. Thus, none of the factors met the criteria of being associated with  
33 both LAA exposure and LPT, and none were considered as potential confounders.  
34 Further consideration of potential confounding and effect modification is addressed in  
35 the uncertainty analyses described in Section 5.3.3.

**Table 5-6. Evaluation of association between covariates and exposure, and between covariates and LPT.<sup>a</sup> Cells display beta coefficient (standard error), *p*-value for predictor**

	Association with cumulative exposure	Association with mean exposure	Association with RTW exposure	Association with LPT
<i>Time-related</i>				
Hire yr	−0.3162 (0.0473), <0.0001	−0.1772 (0.0374), <0.0001	−0.3653 (0.0441), <0.0001	−0.1645 (0.1247), 0.1870
TSFE	0.2703 (0.0499), <0.0001	0.1564 (0.0383), <0.0001	0.3273 (0.0469), <0.0001	0.1702 (0.1237), 0.1690
Job tenure	0.1189 (0.0123), <0.0001	0.0397 (0.0115), 0.0008	0.1091 (0.0130), <0.0001	0.0038 (0.0346), 0.9124
Exposure duration	0.1186 (0.0122), <0.0001	0.0386 (0.0115), 0.0011	0.1102 (0.0129), <0.0001	0.0111 (0.0350), 0.7520
Age at x-ray	0.0185 (0.0199), 0.3551	0.0155 (0.0146), 0.2915	0.0265 (0.0199), 0.1840	0.0084 (0.0402), 0.8349
<i>Other covariates</i>				
Male gender	1.3638 (0.4337), 0.0021	0.9517 (0.3199), 0.0036	1.3264 (0.4348), 0.0028	0.4265 (1.0850), 0.6943
Ever smoker	0.4435 (0.2843), 0.1214	0.2804 (0.2094), 0.1830	0.4044 (0.2848), 0.1582	0.8997 (0.6870), 0.1903
Current	−0.0007 (0.3529), 0.9984	0.0566 (0.2624), 0.8297	−0.0337 (0.3537), 0.9243	0.5485 (0.8528), 0.5201
Former	0.7502 (0.317), 0.0196	0.4350 (0.2358), 0.0676	0.7069 (0.3178), 0.0280	1.0986 (0.7259), 0.1302
BMI <sup>b</sup>	−0.0049 (0.0227), 0.8309	0.0050 (0.0165), 0.7621	−0.0016 (0.0228), 0.9456	0.0309 (0.0426), 0.4690

<sup>a</sup>Association with exposure assessed using a linear regression model, where the outcome is natural log-transformed exposure and the predictor is the covariate of interest. Association with outcome assessed using a logistic model, where the outcome is LPT status and the predictor is the covariate of interest.

<sup>b</sup>Data on BMI were missing for 21 individuals. Thus, the AIC for this model cannot be compared with the AIC for other models in the table.



- Background rate: There may be a nonzero background rate of LPT in the population, and it may be desirable to estimate this rate explicitly rather than using a model that implicitly assumes a background rate of zero. Certain model forms (e.g., log-logistic, log-probit, Dichotomous Hill, Michaelis-Menten) include an explicit parameter representing the background rate of the response while others (e.g., logistic, probit) do not include this parameter. Establishing a background rate for LPT prevalence in the population is challenging, as estimates from previous studies in a variety of populations vary widely ([Weill et al., 2011](#); [Rogan et al., 2000](#); [Zitting, 1995](#); [Cordier et al., 1987](#); [Rogan et al., 1987](#); [Castellan et al., 1985](#); [Anderson et al., 1979](#)); however, these previous studies do indicate that the background rate is unlikely to be zero. Because there is not a clear indication of what the background rate is in an unexposed population, models that allow estimation of a background rate rather than assuming it to be zero, were considered to have greater weight.

Based on the considerations outlined above, EPA developed the following list of desirable model features (see Table 5-7):

- a) Models suitable for a dichotomous outcome
- b) Ability to estimate the effect of exposure via inclusion of slope (models using untransformed exposure), or slope and shape parameters (models using natural log-transformed exposure)
- c) Ability to estimate effect of covariates
- d) Ability to estimate or specify the plateau
- e) Ability to estimate the background rate of LPT in the study population

**Table 5-7. Model features considered in exposure-response modeling to develop a point of departure (POD)**

	Model properties				
	Models suitable for a dichotomous outcome	Allows estimation of slope and shape parameters (where present)	Allows estimation of effect of covariates	Allows estimation or specification of plateau	Allows estimation of background rate of LPT
Probit/logistic	Yes	Yes for slope No for shape	Yes	No	No
Log-probit/log-logistic	Yes	Yes for slope No for shape	Yes	No	Yes
Michaelis-Menten	Yes	Yes for slope No for shape	Yes	Yes	Yes
Dichotomous Hill	Yes	Yes for slope Yes for shape	Yes	Yes	Yes

The epidemiological models described above (logistic, probit, log-logistic, log-probit, Dichotomous Hill, and Michaelis-Menten) are suited for dichotomous outcomes, and all allow for the inclusion of covariates. However, the Michaelis-Menten model does not allow for estimation of a separate shape parameter ( $b$  parameter is implicitly fixed at 1). The ability to estimate both  $a$  and  $b$  (rather than imposing a preassumed shape) provides greater flexibility in exposure-response modeling, and thus the Dichotomous Hill model is preferred over the Michaelis-Menten model. The logistic and probit models do not include separate parameters for either the background rate or a plateau. The three remaining models—log-logistic, log-probit, Michaelis-Menten, and Dichotomous Hill—do allow for estimation of the effect of exposure and the background rate, but only the Dichotomous Hill model includes a separate plateau parameter. Therefore, the Dichotomous Hill model is considered to be the most flexible and potentially the most suitable based on biological and epidemiologic properties in the absence of information on actual model fit, which is also an important consideration in model selection. The Michaelis-Menten, log-probit, and log-logistic models are also reasonable alternatives (note that latter two models implicitly fix the plateau at 100%, above the maximum observed prevalence in reported studies). These models are also evaluated for sensitivity to modeling properties and assumptions. As described above, the plateau is an asymptotic parameter, and it may not be possible to reliably estimate given the limitations of the data. The Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 had relatively low levels of exposure and a narrow range of TSFE, and it is likely that estimation of the maximum prevalence of LPT by radiograph is not well supported in these data. One option to address this difficulty is to fix the plateau parameter at a value consistent with the asbestos literature reviewed above. This assessment will use a plateau value of 85%, based on

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the literature regarding maximum observed prevalence in populations reported to have had long follow-up periods and significant exposure to asbestos, although the sensitivity of the model to this assumption is examined in Section 5.3.4.

#### **5.2.2.6.2. Considerations of exposure metric, statistical model fit and selection of**

**exposure-response model.** While the above description in Section 5.2.2.6.1 explains EPA's preference to use the Dichotomous Hill model for exposure-response modeling in this data set (along with a reasonable set of models to evaluate sensitivity to model form), the text below explains EPA's evaluation of options for the specific exposure metrics to use, and the results of the analysis of statistical fit among the considered models. As noted in Section 5.2.2.4, in the absence of an MOA for pleural health effects, an understanding of the general biology and the epidemiologic literature may help to inform the consideration of exposure metrics for exposure-response modeling. For pleural effects, the timing of exposure, the intensity of exposure, and the duration of exposure may all be important variables to predict the risk of LPT and were considered in the regression modeling.

In the Marysville data, exposure information was collected for each individual based on specific work location, season, and year. There are several ways in which these estimates of exposure by person by year can be aggregated into a single measure. Exposure estimates can be summed over each individual's work history to yield a cumulative exposure estimate for each individual. The cumulative exposure may also be divided by duration of (occupational) exposure to yield an average intensity of exposure (mean exposure). A third option for an exposure metric is to weigh more heavily exposures occurring in the more distant past, using a RTW<sup>22</sup> exposure metric for which each year is weighted by the number of years it occurs prior to the year in which prevalence is evaluated. These three expressions of exposure can be used to derive a POD with appropriate adjustment of the units to arrive at the RfC.

Table 5-8 shows the univariate model results for each of the model forms evaluated, using each of the three different parameterizations of exposure. All models had adequate goodness of fit (GOF), as indicated by the Hosmer-Lemeshow *p*-values (all had *p*-values  $\geq 0.7$ , substantially higher than the standard cutoff value of 0.10, below which a model is considered to fit the data poorly), and were carried forward for further consideration. Because each of the models shown in Table 5-8 was evaluated on the same data set (same number of observations, and same response variable), it is appropriate to compare relative fit among them using the Akaike's Information Criterion (AIC). The AIC for the models ranged from a low value of 73.8

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<sup>22</sup>The RTW exposure value associated with a constant concentration (*c*), with 70-years duration and the evaluation of response at age 70 is the sum of  $1 \times C + 2 \times C + \dots 70 \times C$ , which is equal to about  $C \times 71 \times (70/2)$  or  $C \times 2485$ . This yields a concentration that is about eightfold lower than the concentration that would yield the same RTW exposure, and hence the same modeled risk, for a median experience in the cohort of 20-years duration starting about 30 years prior to exposure, which would yield a RTW exposure for a constant concentration of the sum of  $11 \times C + 12 \times C + \dots 30 \times C$ , or  $C \times 31 \times (20/2)$ , or 310.

1 (probit and Michaelis-Menten models using mean exposure) to a high of 79.3 (logistic model  
2 using RTW exposure). Note that the Dichotomous Hill model with estimated plateau yielded  
3 unrealistic parameter estimates when using either cumulative or RTW exposure (e.g., slope  
4 estimates >100) and were considered less reliable. Within each model form, mean exposure  
5 consistently provided a superior fit (as evidenced by lower AIC) compared to either cumulative  
6 or RTW exposure.

**Table 5-8. Univariate exposure-response modeling for any LPT in the Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 ( $n = 119$ ), using a benchmark response (BMR) of 10% extra risk of any localized pleural thickening (LPT)**

Model form	Exposure metric <sup>a</sup>	Hosmer-Lemeshow GOF $p$ -value	AIC	Intercept (SE)	Background rate (SE)	$b$ (SE), $p$ -value	Benchmark value	Lower limit of benchmark value
Logistic	Mean	0.8035	74.0	−2.7529 (0.3976)	--	6.1969 (1.9469), 0.0015	0.16925	0.11299
	CE	0.8053	79.2	−2.5622 (0.3760)	--	0.2291 (0.0890), 0.0100	4.08863	2.62151
	RTW	0.7012	79.3	−2.5039 (0.3624)	--	0.0082 (0.0032), 0.0103	109.910	69.0617
Probit	Mean	0.8070	73.8	−1.5902 (0.1980)	--	3.6117 (1.0972), 0.0010	0.15369	0.10470
	CE	0.8147	78.7	−1.4978 (0.1908)	--	0.1320 (0.0510), 0.0096	3.82454	2.50223
	RTW	0.6996	78.8	−1.4632 (0.1838)	--	0.0047 (0.0019), 0.0105	102.910	67.3292
Log-logistic	Mean	0.7895	75.3	1.032 (1.0973)	0.0375 (0.0394)	1.3272 (0.6979), 0.0596	0.087768	0.024088
	CE	0.8727	77.0	−2.8335 (0.9114)	0.0376 (0.03)	1.1839 (0.5311), 0.0277	1.71154	0.46974
	RTW	0.7576	77.0	−5.7331 (2.0944)	0.0342 (0.0331)	1.0073 (0.4394), 0.0236	33.4520	7.50156
Log-probit	Mean	0.7745	75.4	0.5262 (0.6337)	0.0407 (0.0359)	0.7311 (0.3578), 0.0432	0.084358	0.024599
	CE	0.8752	76.9	−1.6462 (0.5098)	0.0417 (0.0298)	0.6685 (0.3101), 0.0331	1.72526	0.56792
	RTW	0.7548	76.8	−3.2071 (1.0398)	0.037 (0.0321)	0.5546 (0.2264), 0.0158	32.2062	9.39139

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**Table 5-8. Univariate exposure-response modeling for any LPT in the Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 ( $n = 119$ ), using a benchmark response (BMR) of 10% extra risk of any localized pleural thickening (LPT) (continued)**

Model form	Exposure metric <sup>a</sup>	Hosmer-Lemeshow GOF $p$ -value	AIC	Intercept (SE)	Background rate (SE)	B (SE), $p$ -value	Benchmark value	Lower limit of benchmark value
Dichotomous Hill, plateau = 85%	Mean	0.7854	75.4	1.4136 (1.2953)	0.0384 (0.0391)	1.4043 (0.7769), 0.0732	0.087535	0.024024
	CE	0.8803	76.8	-2.7993 (1.0922)	0.0404 (0.03)	1.3749 (0.7212), 0.059	1.78013	0.51909
	RTW	0.7527	76.8	-5.9883 (2.5304)	0.0366 (0.0338)	1.1266 (0.5493), 0.0425	34.2516	8.32733
Dichotomous Hill <sup>b</sup>	Mean	0.7895	77.3	1.032 (1.0973)	0.0375 (0.0394)	1.3272 (0.6979), 0.0596 Plateau = 1 (--)	0.087768	0.024088
	CE	1.0000	76.3	-129.53 (0.2141)	0.0654 (0.0239)	109.21 (--) Plateau = 0.4999 (0.1443)	3.23559	0.50437
	RTW	0.9999	76.4	-477.74 (0.9458)	0.0655 (0.0239)	108.91 (--) Plateau = 0.5047 (0.1489)	79.4159	0.000002882
Michaelis-Menten, plateau = 85%	Mean	0.7702	73.8	0.7728 (0.5074)	0.0201 (0.0292)	--	0.061813	0.029272
	CE	0.8315	75.1	-2.3490 (0.4933)	0.0310 (0.0258)	--	1.40562	0.67922
	RTW	0.7528	74.8	-5.4305 (0.5333)	0.0320 (0.0271)	--	30.6380	14.2184
Michaelis-Menten	Mean	0.7800	75.6	0.5494 (0.4847)	0.0214 (0.0287)	-- Plateau = 1.00 (--)	0.064145	0.028018
	CE	0.8298	77.1	-2.3217 (1.3330)	0.0309 (0.0262)	-- Plateau = 0.8342 (0.7120)	1.39843	0.57489
	RTW	0.7458	76.8	-5.2131 (1.2187)	0.0306 (0.0277)	-- Plateau = 0.7418 (0.5143)	28.9826	11.2141

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**Table 5-8. Exposure-response modeling for any LPT in the Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 ( $n = 119$ ), using a benchmark response (BMR) of 10% extra risk of any localized pleural thickening (LPT ) (continued)**

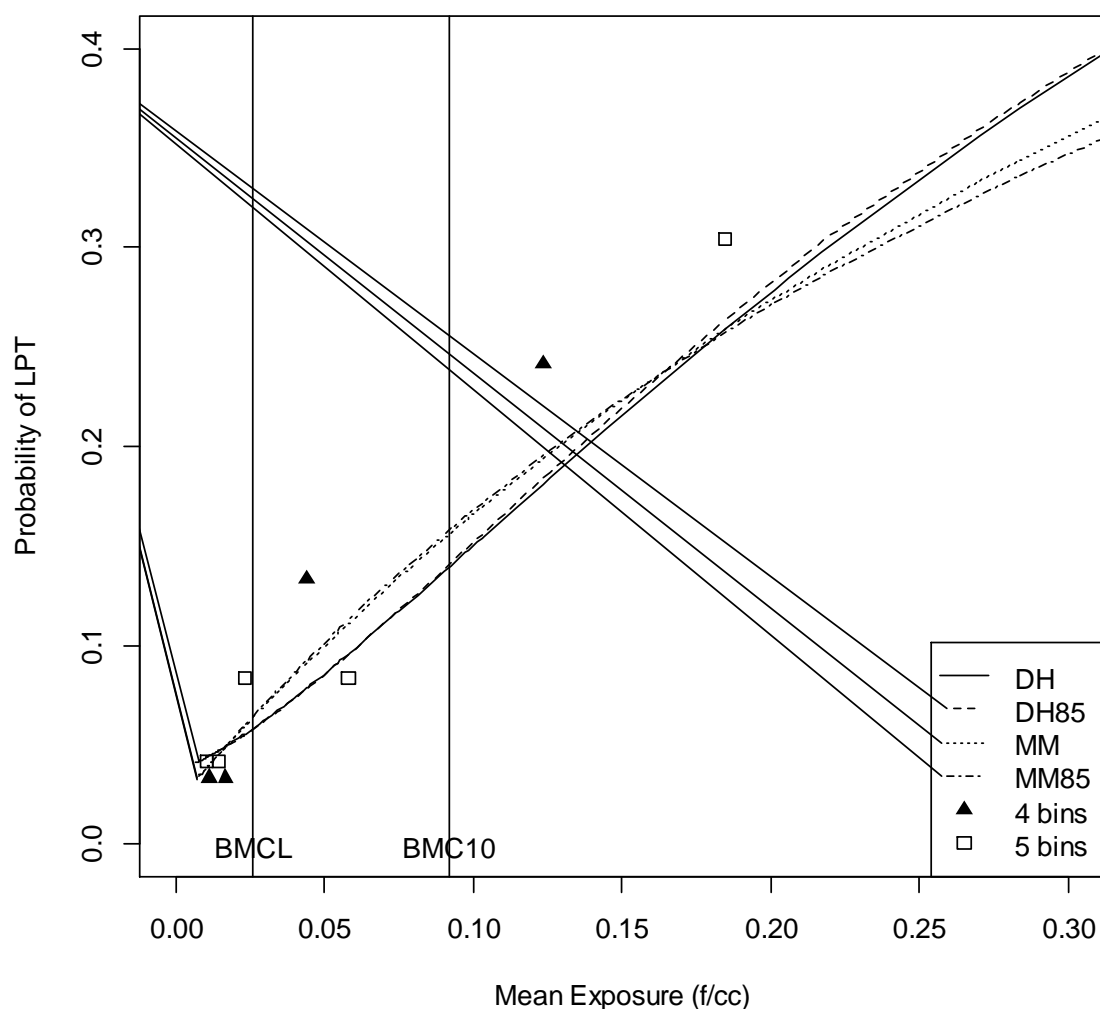
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<sup>a</sup>CE indicates cumulative exposure (fiber/cc-yr), Mean indicates mean exposure (fiber/cc), RTW indicates residence time weighted exposure (fiber/cc-yr<sup>2</sup>), calculated using the midpoint of each work season.

<sup>b</sup>Shaded cell indicates the model did not yield a reasonable estimate for one or more parameters.



1           Each of the different candidate models shown in Table 5-8 is similar in general form, and  
2 comparison of model fit informed by the biological and epidemiologic features of the models  
3 does not strongly imply a preference for one model form over the others. For the mean exposure  
4 model, the AIC values for the logistic, probit, log-logistic, log-probit, Dichotomous Hill model  
5 with plateau fixed at 85%, and Michaelis-Menten model with plateau fixed at 85% ranged from  
6 73.8 to 75.4, a difference of only 1.6 AIC units for the mean exposure model which indicates  
7 essentially equivalent fits. Figure 5-3 shows a plot of the fit for the Dichotomous Hill model  
8 with plateau fixed at 85% and the Michaelis Menten model with plateau fixed at 85%.



**Figure 5-3: Plot of exposure-response models for probability of LPT as a function of mean concentration of occupational exposure in the subcohort.**

Based on the results in Table 5-8, the four lines show the predicted exposure-response shapes of the following models: Dichotomous-Hill model with estimated plateau (DH), Dichotomous-Hill model with plateau fixed at 85% (DH85), Michaelis-Menten model with estimated plateau (MM), and Michaelis-Menten model with plateau fixed at 85% (MM85). The Dichotomous-Hill with plateau fixed at 85% is the selected model and Michaelis-Menten with plateau fixed at 85% is the best fitted model according to AIC. The full range of observed exposure data extended to 0.77 fiber/cc; however, as interest in the fit is in the range of the BMC<sub>10</sub>, the range of exposure values shown here is restricted to 0.3 fiber/cc. The two sets of unconnected symbols show the categorical probability estimates based on quartiles and quintiles of exposure plotted in the median concentration for each category. Data are aggregated in four bins based on quartiles (1,1,4,7 cases in each bin) and five bins based on quintiles (1,1,2,2,7 cases in each bin). Vertical lines at the BMC<sub>10</sub> and BMCL are drawn at the corresponding estimates from Dichotomous-Hill model with plateau fixed at 85%. BMCs and BMCLs for other models are in Table 5-8.

As described above, the Dichotomous Hill model with plateau fixed at 85% possesses desirable biological/epidemiological properties and is the most flexible of the evaluated models. This model had an AIC at the lower range of the models evaluated (75.4 when using mean exposure or 76.8 when using cumulative or RTW exposure); the AIC for the model using mean exposure was within two AIC units of the best-fitting models, a difference that is generally considered indistinguishable with respect to relative model fit ([Burnham and Anderson, 2002](#)). Because it was considered the most flexible model, the Dichotomous Hill model with plateau fixed at 85% was selected as the primary model for RfC derivation. This model was carried forward through the extensive sensitivity analyses (see Section 5.3) because it was considered to be the model most likely to be able to detect sensitivity to covariates and alternative model parameterization.

The RfC is “an estimate of an exposure (including sensitive subgroups) that is likely to be without an appreciable risk of adverse health effects over a lifetime.” ([U.S. EPA, 1994b](#)), where a lifetime is commonly assumed to be 70 years. Thus, consideration of the effect of time (specifically time elapsed from exposure to outcome evaluation) is an important aspect of deriving an RfC. The literature on pleural abnormalities and asbestos generally supports a conclusion that the amount of time elapsed between exposure and evaluation has a major impact on observed response. The model form and/or the selection of an exposure metric should incorporate considerations of time factors. As described above, in the primary data set (which had a very limited range of TSFE values) neither TSFE nor any of the other covariates evaluated were significantly associated with LPT. However, the epidemiologic literature is clear that the timing of exposure is an important factor in evaluating risk over a lifetime, and it was considered critical to address the time-course of exposure and LPT in deriving the RfC. Thus, one option would be to incorporate TSFE through the choice of exposure metric. Neither mean nor cumulative exposure takes into account the TSFE or the timing of subsequent exposures. The RTW exposure metric does incorporate information for each year’s exposure of the time between that exposure and evaluation of the prevalence—exposure in a given time interval is weighted according to time elapsed, with exposure occurring earlier given greater weight. This approach might be preferable to a model including TSFE as a covariate along with mean or cumulative exposure because the RTW metric weights each year of exposure according to the time elapsed until health evaluation. However, while time prior to evaluation is “considered” when using the RTW exposure metric, the subcohort with better exposure data is still limited in that the range of TSFE in this group of workers is relatively narrow (from 23.14 to 32.63 years), which may explain the lack of predictive value of TSFE in this group. Thus, utilizing this approach to estimate a concentration yielding the same risk for a 70-year exposure is informed by a relatively small range of TSFE values.

Therefore, EPA considered explicitly including TSFE as a covariate in exposure-response modeling, along with either mean or cumulative exposure. As noted above, less variability is

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present in TSFE (and other time-related factors) for the subgroup of workers hired in 1972 or later, compared to the larger study population of Marysville workers who underwent health evaluations in 2002–2005 (regardless of hire date,  $n = 252$ ). In this group, the median TSFE was 33.5 years (SD = 7.12 years) and the range was 23.1 to 47.3 years, considerably longer than the subset of these workers whose job start date was on or after 1/1/1972. EPA selected the smaller subgroup for primary exposure analysis because the exposure data after 1972 is of higher quality even though the range in TSFE is more limited. However, it is possible to use a larger subset of the Marysville workers with a wider range of time-related factors (but more uncertain exposure information) to model the effect of TSFE, then include this effect as a fixed parameter when modeling the exposure-response relationship among the primary analytic group of workers who underwent health evaluations in 2002–2005 and were hired  $\geq 1972$ . This hybrid model may be used to calculate a BMCL for a scenario of lifetime exposure (70-years duration and 70-years TSFE). This procedure could use either mean exposure or cumulative exposure. The use of RTW may not be appropriate because RTW includes consideration of the timing of exposures and thus a model using RTW in the exposure metric and TSFE as a covariate might have two variables both reflecting the timing of exposure.

#### *Modeling procedure*

- 1) Fit the model (Dichotomous Hill with plateau fixed at 85%) in all the workers evaluated in 2002–2005 (regardless of hire date) with TSFE and LAA exposure (either represented as cumulative exposure [CE] or as mean exposure [C]) as predictor variables.

$$p(x, T) = bkg + \frac{\text{Plateau} - bkg}{1 + \exp[-a - b \times \ln(x) - c \times T]} \quad (5-1)$$

- 2) Use the regression coefficient for TSFE calculated in (1), represented by “ $c$ ”, as a fixed parameter in the model for workers who underwent health evaluations in 2002–2005, using data only on those hired in 1972 or later; fit the model to the data on this subcohort using the individual data on both TSFE and LAA exposure as independent variables—note that as in the larger cohort of all workers evaluated in 2002–2005, LAA exposure may be modeled as either mean exposure or cumulative exposure.
- 3) Use some fixed value of TSFE to estimate the benchmark value and the lower limit conditional upon that TSFE.

$$\text{Benchmark value} = \exp\left(\frac{-\ln\left(\frac{\text{Plateau} - bkg}{(1 - bkg) \times BMR} - 1\right) - a - c \times T}{b}\right) \quad (5-2)$$

The bivariate modeling results for this hybrid model option are shown in Table 5-9. Both mean and cumulative exposure metrics were evaluated on the basis of two measures of model fit.

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1 The AIC allows the comparison of the fit among two or more models of the same dependent  
2 variable, and AIC results within approximately two units can be considered to be of equivalent  
3 fit. The results in Table 5-9 show similar AIC values for the two hybrid models with exposure  
4 characterized as mean concentration or as cumulative exposure. However, two studies in the  
5 epidemiologic literature also compared mean concentration and cumulative exposure in relation  
6 to the risk of pleural plaques and found that when also including TSFE as an explanatory  
7 variable (as in the results shown in Table 5-9), mean exposure provided a significantly better  
8 model fit ([Paris et al., 2008](#)). In addition, another study ([Järvholm, 1992](#)) proposed model  
9 including TSFE and intensity of exposure and stated that this model is biologically interpretable;  
10 statistical modeling of pleural thickening (18% diffuse) showed that duration of exposure does  
11 not matter when TSFE is included in the model ([Lilis et al., 1991](#)).

**Table 5-9. Estimated point of departure (POD) combining information from the Marysville workers who underwent health evaluations in 2002–2005 and hired in 1972 or later (Primary), and from all workers who underwent health evaluations in 2002–2005 (regardless of hire date), using a benchmark response (BMR) of 10% extra risk of LPT in the Dichotomous Hill model with plateau fixed at 85%. Models include LAA exposure as well as TSFE.**

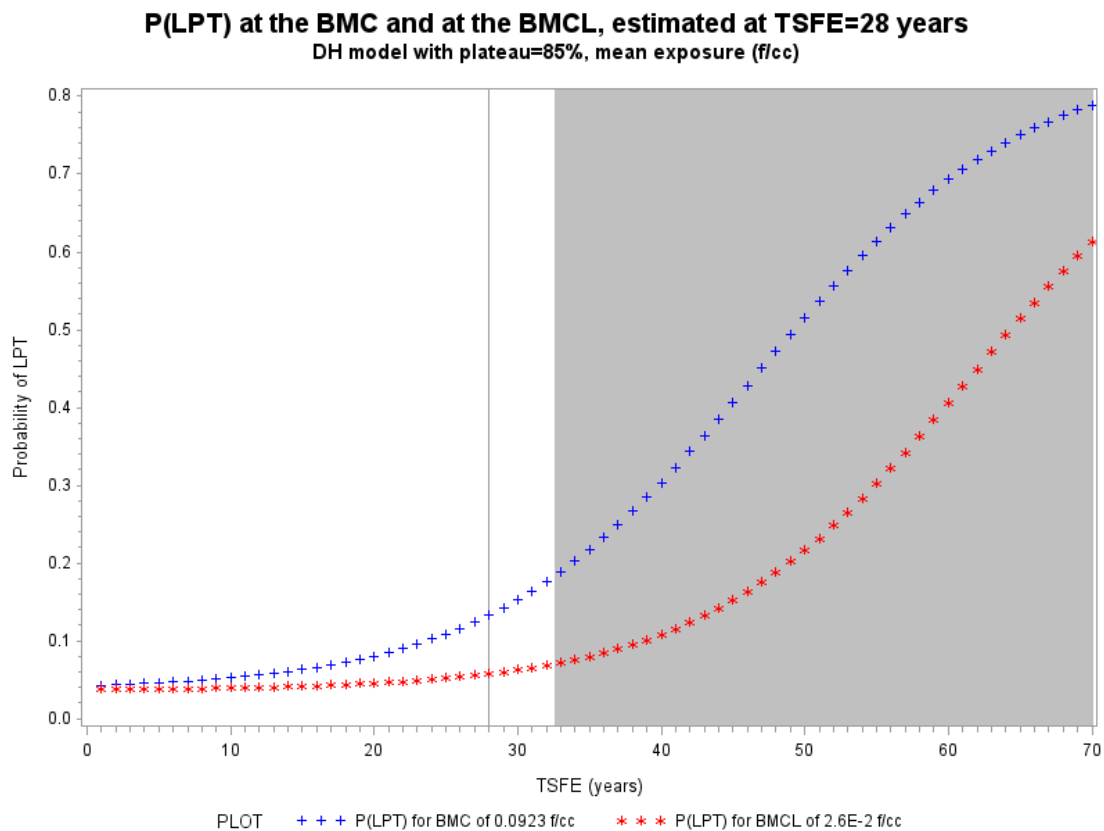
	Mean exposure		Cumulative exposure	
	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)
Hosmer-Lemeshow GOF <i>p</i> -value	0.73626	0.34206	0.76267	0.037763
Model AIC <sup>a</sup>	75.4	242.7	76.8	242.6
Alpha (intercept)	–1.9798 (SE = 1.2270)	–3.4130 (SE = 1.1368)	–5.4574 (SE = 1.0644)	–4.6279 (SE = 1.2668)
Bkg (background)	0.03682 (SE = 0.04037)	0 (–)	0.0388 (SE = 0.0321)	0.0133 (SE = 0.0314)
Beta for TSFE	0.1075 (fixed)	0.1075 (SE = 0.0281), <i>p</i> = 0.0002	0.0957 (fixed)	0.0957 (SE = 0.0326), <i>p</i> = 0.0036
Beta for ln(exposure)	1.2750 (SE = 0.7159), <i>p</i> = 0.0775	0.4819 (SE = 0.1390), <i>p</i> = 0.0006	1.2400 (SE = 0.6809), <i>p</i> = 0.0711	0.4917 (SE = 0.1588), <i>p</i> = 0.0022
BMC/BMCL at 28 yr	0.0923/0.026 f/cc (Ratio = 3.5)		1.8622/0.5770 f/cc-yr (Ratio = 3.2)	

<sup>a</sup>Note that the results in this table are from two different datat sets (Primary, *n* = 119 and Individuals with radiographs in 2002–2005, *n* = 252). The AICs for the models cannot be compared between the two data sets.

1 For the Marysville data, the model with cumulative exposure did not fit well on the other  
2 measure of model fit. In the larger group of workers evaluated in 2002–2005 (regardless of hire  
3 date), fit of cumulative exposure had a low  $p$ -value (i.e.,  $<0.10$ ) for the Hosmer-Lemeshow  
4 goodness-of-fit statistic. Mean exposure did provide an adequate fit and was, therefore, carried  
5 forward in the analysis as the primary basis of the derivation of the lifetime RfC in Section 5.2.3.  
6 Although the model using cumulative exposure did not show an adequate fit based on the  
7 Hosmer-Lemeshow goodness-of-fit test, the model fit as measured by the AIC was nearly  
8 equivalent to the AIC for the mean exposure model, and the beta estimated for the effect of  
9 TSFE was similar to that estimated in the model using mean exposure. Therefore, the derivation  
10 of a chronic RfC based on the CE model is also shown in Section 5.2.3 for comparison.

11 In the model using mean exposure intensity in the larger data set of individuals evaluated  
12 in 2002–2005, the beta coefficient for TSFE was 0.1075. This coefficient value was transported  
13 to the equivalent model in the subset of workers hired in 1972 or later; at a TSFE of 28 years (the  
14 median in this group of workers), the BMC and BMCL were 0.092 fiber/cc and 0.026 fiber/cc,  
15 respectively. Ideally, the objective of the RfC derivation is to compute the BMCL for 70 years  
16 (i.e., a lifetime) of exposure. However, the maximum observed TSFE in the primary cohort was  
17 32.6 years, which while clearly a chronic exposure, cannot be expected to approximate a lifetime  
18 exposure in this particular circumstance when TSFE is the most important predictive factor for  
19 the prevalence of LPT. The BMCL values calculated at longer TSFEs were very low (e.g., a  
20 BMCL of  $2.7 \times 10^{-6}$  at TSFE of 70 years) and the ratio of BMC to BMCL grows exponentially  
21 such that these values at 70 years are considered to be unreliable (e.g., the BMC:BMCL ratio is  
22 1,000 at TSFE of 70 years). This is likely because longer TSFE values are extrapolated well  
23 outside of the range of the data, and attempting to extrapolate beyond ~30 years leads to greater  
24 statistical uncertainty. Consequently, the BMC and BMCL values corresponding to 28 years  
25 were selected as the primary modeling result, with the BMCL of 0.026 fiber/cc serving as the  
26 POD for RfC derivation with some additional accommodation of the uncertainty in using less  
27 than lifetime exposure data to estimate lifetime risks.

28 At the BMCL ( $2.6 \times 10^{-2}$  fibers/cc), the final model leads to an estimated probability of  
29 LPT of 0.06 at 28-years TSFE, and 0.61 at 70-years TSFE (see Figure 5-4). This 10-fold  
30 increase in the probability of the critical effect going from the median TSFE in the cohort, to the  
31 full lifetime, is used to derive a data-based subchronic to chronic uncertainty factor (UF) for RfC  
32 derivation in the following section.



**Figure 5-4. Predicted risk of localized pleural thickening (LPT) at the benchmark concentration (BMC) and the lower limit of the BMC (BMCL), using the hybrid Dichotomous Hill model with plateau fixed at 85%.** The parameters for this were taken from Table 5-8. Note that the vertical reference line indicates TSFE = 28 years, used to calculate the BMC and BMCL. The shaded region indicates TSFE beyond that observed in the cohort of workers evaluated in 2002–2005.



### 5.2.3. Derivation of a Reference Concentration (RfC) for the Critical Effect of Localized Pleural Thickening (LPT) in the Marysville Workers Who Underwent Health Evaluations in 2002–2005 and Were Hired in 1972 or Later—Including Application of Uncertainty Factors (UFs)

Among the available studies that could provide exposure-response data for the relationship between LAA exposure and risk of LPT, consideration of study attributes led to the selection of a study of the Marysville, OH workers evaluated in 2002–2005 as the primary data set for RfC derivation ([Rohs et al., 2008, see Section 5.2.1](#)). An updated job-exposure matrix was available for this follow-up of the original study group, with a refined understanding of exposure to LAA throughout plant operation (see Section 5.2.3.1 and Appendix F). However, due to remaining uncertainties in exposures prior to 1972, EPA elected to focus on exposure-response modeling for the subgroup of plant employees hired in 1972 or later (see Section 5.2.3.2). The critical effect selected for derivation of the RfC is LPT, a persistent change to normal tissue structure associated with decreased pulmonary function.

Using a 10% BMR for LPT, a BMC of 0.092, and a BMCL<sub>10</sub> of 0.026 fiber/cc were calculated for the mean exposure model (see Table 5-9). Following EPA practices and guidance ([U.S. EPA, 2002, 1994b](#)), application of the following UFs was evaluated resulting in a composite UF of 300.

- An interspecies uncertainty factor, UF<sub>A</sub>, of 1 is applied for extrapolation from animals to humans because the POD used as the basis for the RfC was based on human data.
- An intraspecies uncertainty factor, UF<sub>H</sub>, of 10 was applied to account for human variability and potentially susceptible individuals. Only adults sufficiently healthy for full-time employment were included in the principal study and the study population was primarily male. Other population groups, such as the elderly, children, and those with preexisting health conditions, were not evaluated in the principal study but may have a different response to LAA exposure.
- An uncertainty factor for extrapolating from a lowest observed adverse effect level (LOAEL) to no observed adverse effect level (NOAEL), UF<sub>L</sub>, of 1 was applied because the current approach is to address this factor as one of the considerations in selecting a BMR for BMC modeling.
- A subchronic-to-chronic uncertainty factor, UF<sub>S</sub>, of 10 was applied because while the selected POD is from a study population including workers with chronic exposure defined as more than 10% of a lifetime (i.e., more than 7 years), for this particular health endpoint, even ~30 years of observation ([Rohs et al., 2008](#)) is insufficient to describe lifetime risks.

Although data do exist to define an exposure-response relationship for radiographic abnormalities in the Marysville, OH worker cohort, these data are limited by the dates of the available radiographs. The data for the subcohort of workers exposed post-1972 allowed for assessing prevalence of LPT up to approximately 30 years after

first exposure. EPA used information from the larger group of all workers evaluated in 2002–2005 to estimate the effect of TSFE because this group had greater variability in TSFE (with a maximum TSFE of 47 years). However, the Marysville data did not have information on effects after a full lifetime of exposure (i.e., 70-years TSFE), and evidence indicates that the prevalence of pleural plaques is likely to continue to increase over the life span ([Paris et al., 2009](#); [Paris et al., 2008](#); [Jakobsson et al., 1995](#); [Hillerdal, 1994a](#); [Järvholm, 1992](#); [Lilis et al., 1991](#)). As the RfC is intended for a lifetime of exposure, and pleural thickening is known to progress across the lifetime (even with less-than-lifetime exposures), the lack of health data assessed at end of lifetime is a data gap. Using the model selected for derivation of the RfC, the probability of LPT increases 10-fold between 28-years TSFE (the median in the population of workers used for analysis) and 70-years TSFE. Thus, use of a 10-fold UF for subchronic-to-chronic uncertainty should capture the uncertainty due to increasing risk of LPT over the life course.

- A database uncertainty factor,  $UF_D$ , of 3 was applied to account for database deficiencies in the available literature for the health effects of LAA.

Although a large database exists for asbestos in general, only four study populations exist for LAA specifically: the Minneapolis community study, the Marysville, OH worker cohort, the Libby worker cohort, and the ATSDR community screening (which includes some Libby worker cohort participants). Studies conducted in three of these populations, the Libby worker cohort ([Larson et al., 2012a](#)), Minneapolis community study ([Alexander et al., 2012](#)), and Marysville workers ([Rohs et al., 2008](#)), have all demonstrated substantial numbers of LPT cases occurring at the lowest exposure levels examined in each study ([Christensen et al., 2013](#)), lending confidence to the use of LPT as a critical effect and ([Rohs et al., 2008](#)) as the principal study for RfC derivation.

However, studies in the Libby population have also demonstrated an association between exposure to LAA and autoimmune effects (i.e., self-reported autoimmune disease and autoimmune markers in Libby residents ([Marchand et al., 2012](#); [Noonan et al., 2006](#); [Pfau et al., 2005](#)). Because these studies did not provide exposure-response information, it is unknown whether a lower POD or RfC would be derived for these effects. For other (non-Libby) forms of amphibole asbestos, there is evidence regarding autoimmune effects from a study of individuals in a community exposed to tremolite. In this population, there were changes in immune parameters in exposed individuals without pleural plaques, and additional immune markers (including autoantibodies) were increased in individuals with pleural plaques ([Zerva et al., 1989](#)). Also it has been hypothesized that shorter asbestos fibers reach the pleura via passage through lymphatic channels ([Peacock et al., 2000](#)), although experimental evidence is lacking for this or alternative potential mechanisms of fiber migration. This uncertainty in the sequence of health effects (pleural or autoimmune) is the basis for selecting a  $UF_D$  of 3.

The derivation of the RfC from the morbidity studies of the Marysville, OH worker cohort ([i.e., Rohs et al., 2008](#)) was calculated from a POD,  $BMCL_{10}$  for LPT of 0.026 fiber/cc,

and dividing by a composite UF of 300. As derived below, the chronic RfC is  $8.67 \times 10^{-5}$  fibers/cc for LAA:

$$\begin{aligned}\text{Chronic RfC for LPT} &= \text{BMCL}_{10} \div \text{UF} \\ &= 0.026 \text{ fiber/cc} \div 300 \\ &= 8.67 \times 10^{-5} \text{ fibers/cc, rounded to } 9 \times 10^{-5} \text{ fibers/cc}\end{aligned}\tag{5-3}$$

It should be noted that for the primary RfC and all the alternative RfCs, the fiber concentrations are presented here as continuous lifetime exposure in fiber/cc where exposure measurements are based on analysis of air filters by PCM. Current analytical instruments used for PCM analysis have resulted in a standardization of minimum fiber width considered visible by PCM between 0.2 and 0.25  $\mu\text{m}$ . Historical PCM analysis (1960s and early 1970s) generally had less resolution, and fibers with minimum widths of 0.4 or 0.44  $\mu\text{m}$  were considered visible by PCM ([Amandus et al., 1987b](#); [Rendall and Skikne, 1980](#)). Methods are available to translate exposure concentrations measured in other units into PCM units for comparison.

#### **5.2.3.1. Derivation of a Reference Concentration (RfC) for the Alternative Endpoint of Any Pleural Thickening (APT) in the Marysville Workers Who Underwent Health Evaluations in 2002–2005 and Were Hired in 1972 or Later**

As shown in the uncertainty analyses in Section 5.3.5 (see Table 5-17), use of an alternative critical effect of APT results in an almost identical  $\text{BMCL}_{10}$  as derived for the primary analysis using LPT as the critical effect. In the subcohort, the number of cases of APT is identical to the number of cases of LPT, and in the larger group of workers evaluated in 2002–2005 used to estimate the effect of TSFE, the number of cases of APT is very similar to the number of LPT cases ( $n = 69$  cases of APT and  $n = 66$  cases of LPT). In this larger group, the regression coefficient for the effect of TSFE was very close when using either of these two endpoints—values of 0.1108 for APT and 0.1075 for LPT. Thus, using the same uncertainty factors as described above, the chronic RfC is  $9.0 \times 10^{-5}$  fibers/cc for LAA, using an alternative endpoint of APT:

$$\begin{aligned}\text{Chronic RfC for APT} &= \text{BMCL}_{10} \div \text{UF} \\ &= 0.027 \text{ fiber/cc} \div 300 \\ &= 9.0 \times 10^{-5} \text{ fibers/cc}\end{aligned}\tag{5-4}$$

This value is identical to the primary RfC derived using a critical effect of LPT, when rounded to one significant digit. These results provide additional support to further substantiate the primary RfC.

1 **5.2.3.2. Derivation of a Reference Concentration (RfC) for the Alternative Endpoint of Any**  
2 **Radiographic Change (ARC) in the Marysville Workers Who Underwent Health**  
3 **Evaluations in 2002–2005 and Were Hired in 1972 or Later**

4 As shown in the uncertainty analyses in Section 5.3.5 (see Table 5-17), use of an  
5 alternative critical effect of any radiographic change (ARC) results in an almost identical  
6 BMCL<sub>10</sub> as derived for the primary analysis using LPT as the critical effect. In the subcohort,  
7 the number of cases of ARC is identical to the number of cases of LPT, and in the larger group  
8 of workers evaluated in 2002–2005 used to estimate the effect of TSFE, the number of cases of  
9 APT is very similar to the number of LPT cases ( $n = 71$  cases of ARC, and  $n = 66$  cases of LPT).  
10 In this larger group, the regression coefficient for the effect of TSFE was very close when using  
11 either of these two endpoints—values of 0.1115 for ARC and 0.1075 for LPT. Thus, using the  
12 same uncertainty factors as described above, the chronic RfC is  $9.0 \times 10^{-5}$  fibers/cc for LAA,  
13 using an alternative endpoint of ARC:  
14

$$\begin{aligned}\text{Chronic RfC for ARC} &= \text{BMCL}_{10} \div \text{UF} & (5-5) \\ &= 0.027 \text{ fiber/cc} \div 300 \\ &= 9.0 \times 10^{-5} \text{ fibers/cc}\end{aligned}$$

15  
16  
17  
18  
19 This value is identical to the primary RfC derived using a critical effect of LPT, when  
20 rounded to one significant digit. These results provide additional support to further substantiate  
21 the primary RfC.  
22

23 **5.2.4. Derivation of a Reference Concentration (RfC) for Localized Pleural Thickening**  
24 **(LPT) in the Marysville Workers Who Underwent Health Evaluations in 2002–2005**  
25 **and Were Hired in 1972 or Later Based on the Cumulative Exposure Model**

26 As shown in the analyses in Section 5.2.2 (see Table 5-9), the model with TSFE and CE  
27 yields a similar overall fit to the model based on TSFE and mean exposure as judged by AIC  
28 values within 2 units (242.6 vs. 242.7). However, the CE model yielded an unacceptably low  
29 value for the Hosmer-Lemeshow goodness-of-fit test ( $p < 0.04 < 0.1$ ) when also including TSFE  
30 in the model, and therefore, the mean exposure model was selected for the derivation of the RfC.  
31 This same pattern was corroborated in the analysis of the combined cohort in Appendix E where  
32 the same bivariate Dichotomous Hill model with plateau fixed at 85% passed the  
33 Hosmer-Lemeshow goodness-of-fit test for the model with TSFE and mean exposure but yielded  
34 an unacceptably low value for the model with TSFE and CE ( $p = 0.006 < 0.1$ ; see Table E-4  
35 Row: BV DH [CE, TSFE]).

36 However, as CE has been a traditional exposure metric of asbestos exposure, there may  
37 be interest in an estimate of the RfC for such a model based on TSFE and CE. The use of an  
38 alternative model based on the TSFE and CE data yields a BMCL<sub>10</sub> of 0.577 fiber/cc-yr using

LPT as the critical effect among individuals evaluated in 2002–2005 and hired in 1972 or later. In order to adjust the POD to units of concentration (fiber/cc), this POD was divided by the mean duration of occupational exposure (18.23 years) to yield an adjusted BMCL<sub>10</sub> of 0.0317 fiber/cc. Using the same uncertainty factors as described above, the value of an RfC for LAA based on TSFE and CE would be  $1.0 \times 10^{-4}$  fibers/cc:

$$\begin{aligned}\text{Chronic RfC for LPT} &= \text{BMCL}_{10} \div \text{UF} \\ &= 0.0317 \text{ fiber/cc} \div 300 \\ &= 1.0 \times 10^{-4} \text{ fibers/cc}\end{aligned}\tag{5-6}$$

This value is approximately 11% higher than the primary RfC derived using a critical effect of LPT. As noted previously, this result is included for comparative purposes with the primary RfC, but was not selected as the primary RfC as the model based on CE did not indicate an adequate fit to the Marysville subcohort of workers evaluated in 2002–2005 and hired in 1972 or later when also including TSFE in the model.

#### **5.2.5. Derivation of a Reference Concentration (RfC) for the Alternative Endpoint of Any Pleural Thickening (APT) in the Marysville Cohort with Combined X-Ray Results from 1980 and 2002–2005 Regardless of Date of Hire**

EPA also conducted modeling of the full Marysville cohort to substantiate the derivation of the RfC in the primary analytic subset of workers evaluated in 2002–2005 and hired in 1972 or later and considered in the previous sections. The “combined cohort” was assembled using all individuals who participated in the health examination in 1980 ([Lockey et al., 1984](#)) and 2002–2005 ([Rohs et al., 2008](#)), and who were not exposed to asbestos from a source outside of the Marysville facility (see Table 5-3 and Appendix E for details). Due to differences in the 1980 x-ray evaluations compared with the 2002–2005 x-ray evaluations, an alternative critical effect of APT was used in the combined cohort modeling.

The modeling of the combined cohort is described in detail in Appendix E. A suite of model forms (univariate and bivariate) and exposure metrics (mean exposure, cumulative exposure, RTW exposure) were evaluated using the combined cohort of 434 individuals. Two models (Cumulative Normal Dichotomous Hill and Cumulative Normal Michaelis Menten models), which incorporated TSFE into the plateau term rather than as an independent predictor alongside the exposure metric, were also evaluated as a supplement to the standard suite of models.

The results for the BMC and BMCL for the three endpoints (LPT—defined as LPT diagnosed in 2002–2005 and PT in 1980, APT, and ARC) are presented in Tables E-3, E-4, and E-5. Any pleural thickening is selected as the preferred endpoint for the alternative derivation of the reference concentration (RfC) for the combined cohort because the endpoint of APT is more

inclusive (it includes those with LPT and those with DPT in the absence of LPT) and eliminates the uncertainty regarding the type of pleural thickening observed in the 1980 study (Lockey et al., 1984) using the 1971 ILO guidance. The two models selected for derivation of an RfC were (1) the same bivariate Dichotomous Hill model with fixed plateau used in the primary analysis based on TSFE and mean concentration and (2) the cumulative normal Dichotomous Hill model with fixed background rate of APT based on TSFE and CE. Additionally, in keeping with the primary analysis of the subcohort hired in 1972 or later, an analogous hybrid analysis of the subcohort based on the combined cohort was also conducted based on the two selected models. All of these results are presented at a value of TSFE = 70 years for those models that were able to reasonably extrapolate the TSFE value outside the observed range (47 years maximum) and for the median value of TSFE in the combined cohort (25 years) for models where the extrapolation to 70 years was not reasonable.

Following EPA practices and guidance (U.S. EPA, 2002, 1994b) as discussed in Section 5.2.3, a composite uncertainty factor (UF) of 300 is used when deriving the RfC from the POD calculated at the median TSFE (25 years). This includes an uncertainty factor of 10 to account for intraspecies variability ( $UF_H = 10$ ), a factor of three to account for database uncertainty ( $UF_D = 3$ ) and an extra factor ( $UF_S$ ) of 10 to account for the lack of information on people at risk for a complete lifetime ( $UF_S = 10$ ). When using the POD based on the BMCL calculated at TSFE = 70 years, the additional adjustment factor of 10 is not necessary and a composite UF of 30 is used ( $UF_H = 10$  and  $UF_D = 3$ ). The calculations of the RfC for the combined cohort and the Rohs subcohort using both options are shown in Table 5-10. The RfCs are rounded to one significant digit.

**Table 5-10. (Copy of Table E-11) Reference concentrations (RfCs) for the alternative endpoint of any pleural thickening (APT) in the Marysville cohort with combined x-ray results from 1980 and 2002–2005 regardless of date of hire**

Cohort	Starting from	Model (parameters)	Calculation
Combined cohort	TSFE = 25 yr	CN DH (CE, TSFE)	$RfC = (3.4 \times 10^{-2})/300 = 1 \times 10^{-4}$ fibers/cc
Combined cohort	TSFE = 25 yr	BV DH FP (C, TSFE)	$RfC = (6.3 \times 10^{-2})/300 = 2 \times 10^{-4}$ fibers/cc
Rohs subcohort	TSFE = 25 yr	CN DH (CE, TSFE) <sup>a</sup>	$RfC = (3.5 \times 10^{-2})/300 = 1 \times 10^{-4}$ fibers/cc
Combined cohort	TSFE = 70 yr	CN DH (CE, TSFE)	$RfC = (7.5 \times 10^{-4})/30 = 3 \times 10^{-5}$ fibers/cc
Rohs subcohort	TSFE = 70 yr	CN DH (CE, TSFE) <sup>a</sup>	$RfC = (8.4 \times 10^{-4})/30 = 3 \times 10^{-5}$ fibers/cc

Abbreviations: TSFE (time since first exposure), C (mean exposure), CE (cumulative exposure), CN DH (cumulative normal Dichotomous Hill), BV DH FP (bivariate Dichotomous Hill with fixed plateau).

<sup>a</sup>Hybrid model informed by the combined cohort; background rate of APT was fixed at 3% to facilitate model convergence.

1 For comparison, the above values all fall within approximately threefold when compared  
2 to the primary RfC for LPT of  $9 \times 10^{-5}$  fibers/cc derived in Section 5.2.3 from the Marysville  
3 workers who underwent health evaluations in 2002–2005 and were hired in 1972 or later. These  
4 results provide additional support to further substantiate the primary RfC.

5  
6 **5.2.6. Summary of Reference Concentration Values (RfCs) for the Different Health**  
7 **Endpoints and Different Sets of Workers in the Marysville Cohort**

8 The primary derivation of the reference concentration is based on the critical effect of LPT  
9 in the Marysville workers who underwent health evaluations in 2002–2005 and were hired in  
10 1972 or later. Multiple alternative values were derived using other health endpoints, other  
11 regression models, and other sets of workers from the Marysville cohort. The results of the  
12 different derivations are shown in Table 5-11. The range of values is from 3E-5 to 2E-4.



**Table 5-11. Multiple derivations of a reference concentration from the Marysville, OH cohort. Primary RfC value in bold.**

Location	Study population	Health endpoint	Occupational LAA exposure	Residential LAA exposure	Occupational TSFE	Residential TSFE	Model	Exposure metrics	RfC	Section
Cohorts										
Libby, MT	Workers	LPT	Measured	Unknown	Measured	Unknown	---	---	---	5.2.1
Minneapolis, MN	Residents	LPT	---	Modeled	---	Unclear	---	---	---	5.2.1
Marysville, OH	Workers	LPT, APT, ARC	Measured $\geq 1972$	---	Measured	---	---	---	---	5.2.1
RfC estimates										
<b>Marysville, OH</b>	<b>Hired <math>\geq 1972</math></b>	<b>LPT</b>	<b>Measured</b>	<b>None</b>	<b>Measured</b>	<b>---</b>	<b>Hybrid Rohs DH<sub>85</sub>; TSFE = 28 yr</b>	<b>TSFE and C</b>	<b>9E-5</b>	<b>5.2.3</b>
Marysville, OH	Hired $\geq 1972$	APT	Measured	None	Measured	---	Hybrid Rohs DH <sub>85</sub> ; TSFE = 28 yr	TSFE and C	9E-5	5.2.3.1
Marysville, OH	Hired $\geq 1972$	ARC	Measured	None	Measured	---	Hybrid Rohs DH <sub>85</sub> ; TSFE = 28 yr	TSFE and C	9E-5	5.2.3.2
Marysville, OH	Hired $\geq 1972$	LPT	Measured	None	Measured	---	Hybrid Rohs DH <sub>85</sub> ; TSFE = 28 yr	TSFE and CE	1E-4	5.2.4
Marysville, OH	All hires	APT	Measured $\geq 1972$	None	Measured	---	Combined cohort CN DH; TSFE = 25 yr	TSFE and CE	1E-4	5.2.5 App. E
Marysville, OH	All hires	APT	Measured $\geq 1972$	None	Measured	---	Combined cohort CN DH; TSFE = 25 yr	TSFE and C	2E-4	5.2.5 App. E
Marysville, OH	All hires	APT	Measured $\geq 1972$	None	Measured	---	Hybrid combined cohort DH <sub>85</sub> ; TSFE = 25 yr	TSFE and CE	1E-4	5.2.5 App. E
Marysville, OH	All hires	APT	Measured $\geq 1972$	None	Measured	---	Combined cohort CN DH; TSFE = 70 yr	TSFE and CE	3E-5	5.2.5 App. E
Marysville, OH	All hires	APT	Measured $\geq 1972$	None	Measured	---	Hybrid combined cohort DH <sub>85</sub> ; TSFE = 70 yr	TSFE and CE	3E-5	5.2.5 App. E

Abbreviations: Health endpoint (LPT = localized pleural thickening, APT = any pleural thickening, ARC = any radiographic change).

Models (hybrid Rohs = 2-step model fit to full Rohs cohort then Rohs subcohort, hybrid combined cohort = 2-step model fit to combined cohort then to the Rohs subcohort, DH<sub>85</sub> = Dichotomous Hill with plateau fixed at 85%, CN DH = cumulative normal Dichotomous Hill).

Exposure metrics (TSFE = time since first exposure, C = mean concentration, CE = cumulative exposure).

### 5.3. UNCERTAINTIES IN THE INHALATION REFERENCE CONCENTRATION (RFC)

Some sources of uncertainty remain in the derivation of the RfC. This section identifies the major sources of uncertainty and, where possible, uses a sensitivity analysis to evaluate the potential impact of these uncertainties on the point of departure.

#### 5.3.1. Uncertainty in the Exposure Reconstruction

As in all epidemiologic studies, uncertainties are present in the exposure assessment. In this case, some uncertainty lies in the employment history, and some individuals had extensive overtime work. Employment history was self-reported during interviews with each individual for the original study ([Lockey et al., 1984](#)), and any errors in this process could affect assigned LAA exposure estimates. While the uncertainties related to a lack of quantitative measurements are not relevant to the analysis of workers hired in 1972 or later, it is important to recognize that exposure assessment post-1972 also has some limitations. The main source of uncertainty is incomplete exposure measurements for some of the occupations/tasks before industrial hygiene improvements that started about 1973 or 1974 and continued throughout the 1970s (see Appendix F, Figure F-1).

Some uncertainty exists when the Libby ore was first used in the facility. Company records indicated that the date was between 1957 and 1960, and the University of Cincinnati used the best available information from focus group interviews to assign 1959 as the year of the first usage of Libby vermiculite ore (see Appendix F). In 1957 and 1958, only vermiculite ore from South Carolina was thought to be used. From 1959 to 1971, vermiculite ores from both Libby, MT and South Carolina were used. From 1972 to 1980, vermiculite ores from Libby, MT, South Carolina, South Africa, and Virginia were used, with Libby vermiculite ore being the major source. Libby vermiculite ore was not used in the facility after 1980. However, industrial hygiene measurements (based on PCM) collected after 1980 showed low levels of fibers in the facility. PCM analysis does not determine the mineral/chemical make-up of the fiber and, thus, cannot distinguish among different kinds of asbestos.

Uncertainty also exists in the data regarding the asbestos content in other vermiculite ore sources before and after the Libby ore was used. As reported in Appendix C, EPA analysis of bulk vermiculite ores from Virginia and South Africa showed the presence of only a few or no amphibole asbestos fibers. The South Carolina vermiculite ore contained relatively more fibers than the Virginia and South African vermiculite ores but still far fewer fibers than the Libby vermiculite ore. Using the industrial hygiene data, the University of Cincinnati estimated that the fiber content of the South Carolina ore was about 8.7% of that of the Libby ore (see Appendix F). This result is consistent with data comparing South Carolina and Libby vermiculite ores from samples tested in 1982 ([U.S. EPA, 2000c](#)). Based on the industrial hygiene data, the concentration of fibers detected in the workplace was near background after

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1 1980. The exposure distribution in Marysville workers is summarized in Table 5-12, which  
2 shows the mean, cumulative, and RTW exposure metrics for the full cohort of workers, the  
3 subset of those evaluated in 2002–2005, and the primary analytic group of workers evaluated in  
4 2002–2005 and hired in 1972 or later. To evaluate the potential impact of assumptions regarding  
5 exposure after 1980 when the use of Libby vermiculite ores was considered to have ceased, EPA  
6 conducted a sensitivity analysis using the hybrid Dichotomous Hill model with plateau fixed at  
7 85%, but truncating exposures at 12/31/1980 (see Table 5-13). Note that except where stated  
8 otherwise, this model (hybrid Dichotomous Hill model with plateau fixed at 85%) was used for  
9 all sensitivity analyses. Using the truncated exposures, the POD increased from  
10  $2.6 \times 10^{-2}$  fibers/cc, to  $3.9 \times 10^{-2}$  fibers/cc. However, note that for the model with truncated  
11 exposures, the modeling in the larger group of workers evaluated in 2002–2005 showed a poor  
12 fit (Hosmer-Lemeshow GOF  $p$ -value  $< 0.10$ ). These results are included for comparative  
13 purposes but should be interpreted with caution.

**Table 5-12. Exposure distribution among workers at the O.M. Scott plant in Marysville, OH**

	All individuals evaluated in 1980 and/or in 2002–2005 <sup>a</sup>		Individuals evaluated in 2002–2005		Individuals evaluated in 2002–2005, hired in 1972 or later	
Exposure metrics, based on arithmetic mean	Mean (SD)	Median (25 <sup>th</sup> –75 <sup>th</sup> percentiles)	Mean (SD)	Median (25 <sup>th</sup> –75 <sup>th</sup> percentiles)	Mean (SD)	Median (25 <sup>th</sup> –75 <sup>th</sup> percentiles)
Cumulative exposure, all yr (fiber/cc-yr)						
Arithmetic mean	7.9232 (17.9598) Range: 0.003–96.91	1.1252 (0.3414–3.7684)	8.75 (19.12) Range: 0.005–96.91	1.26 (0.51–5.20)	1.439 (2.5479) Range: 0.005–17.33	0.5048 (0.2188–1.5519)
Geometric mean	2.9258 (7.0248) Range: 0.001–37.73	0.2132 (0.1004–1.2635)	3.24 (7.48) Range: 0.002–37.73	0.29 (0.14–1.78)	0.4756 (0.8734) Range: 0.002–6.05	0.1785 (0.087–0.4632)
Mean exposure, all yr (fiber/cc)						
Arithmetic mean	0.3733 (0.7942) Range: 0.007–4.34	0.0566 (0.0267–0.2364)	0.31 (0.65) Range: 0.007–4.10	0.05 (0.02–0.20)	0.0716 (0.1239) Range: 0.007–0.77	0.0234 (0.0133–0.074)
Geometric mean	0.1366 (0.3101) Range: 0.003–1.70	0.0111 (0.0068–0.0719)	0.11 (0.25) Range: 0.003–1.59	0.01 (0.006–0.07)	0.0236 (0.0422) Range: 0.003–0.26	0.0085 (0.0062–0.0222)
Residence time-weighted exposure, all yr (fiber/cc-yr), calculated using midpoint of season dates						
Arithmetic mean	193.3093 (519.3874) Range: 0.0007–3500.66	19.4767 (4.2550–78.0944)	294.38 (687.95) Range: 0.12–3500.66	34.31 (11.07–154.36)	33.7415 (69.2231) Range: 0.12–474.01	10.2075 (3.9055–29.1246)
Geometric mean	72.2260 (204.1052) Range: 0.0003–1373.22	4.1835 (0.9229–25.2179)	110.14 (270.86) Range: 0.05–1373.22	6.24 (3.06–49.81)	11.1415 (24.2783) Range: 0.05–168.17	3.3526 (1.6814–8.5014)
Cumulative exposure, through 1980 (fiber/cc-yr)						
Arithmetic mean	7.6544 (17.8658) Range: 0.003–95.09	0.9708 (0.2120–3.3925)	8.27 (18.95) Range: 0.005–95.09	0.97 (0.22–4.65)	0.9638 (2.2774) Range: 0.005–15.37	0.212 (0.0564–0.5849)
Geometric mean	2.8324 (7.0007) Range: 0.001–37.20	0.1277 (0.0502–1.0805)	3.08 (7.43) Range: 0.002–37.20	0.13 (0.05–1.55)	0.3119 (0.8027) Range: 0.002–5.48	0.0478 (0.0244–0.1279)
Mean exposure, through 1980 (fiber/cc)						
Arithmetic mean	0.4863 (0.9568) Range: 0.008–4.34	0.0766 (0.0375–0.2950)	0.51 (0.98) Range: 0.008–4.33	0.07 (0.04–0.40)	0.1422 (0.2911) Range: 0.008–1.84	0.0352 (0.0193–0.1027)

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**Table 5-12. Exposure distribution among workers at the O.M. Scott plant in Marysville, OH (continued)**

	All individuals evaluated in 1980 and/or in 2002–2005 <sup>a</sup>		Individuals evaluated in 2002–2005		Individuals evaluated in 2002–2005, hired in 1972 or later	
Geometric mean	0.1767 (0.3732) Range: 0.003–1.70	0.0110 (0.0074–0.0968)	0.19 (0.38) Range: 0.003–1.68	0.01 (0.007–0.11)	0.0452 (0.1013) Range: 0.003–0.66	0.0102 (0.0063–0.0261)
Residence time-weighted exposure, through 1980 (fiber/cc-yr), calculated using midpoint of season dates						
Arithmetic mean	189.5463 (517.2418) Range: 0.0007–3475.17	16.0931 (2.7315–69.9986)	287.78 (685.61) Range: 0.12–3475.17	30.08 (6.07–137.64)	27.2008 (65.8222) Range: 0.12–448.51	5.9930 (1.4107–16.0968)
Geometric mean	70.9061 (203.4978) Range: 0.0003–1365.69	2.4223 (0.6805–23.5698)	107.84 (270.24) Range: 0.05–1365.69	4.05 (1.34–43.96)	8.8655 (23.3699) Range: 0.05–160.64	1.3228 (0.6124–3.6592)

<sup>a</sup>See Appendix E for details of how the individual health outcome data for all workers who participated in the [Lockey et al. \(1984\)](#) study and the follow-up study by [Rohs et al. \(2008\)](#) were combined.

**Table 5-13. Effect of truncating exposures after 1980 and of using arithmetic or geometric mean to summarize multiple fiber measurements**

	Exposures based on arithmetic mean—Primary Analysis		Exposures based on arithmetic mean, truncated at 1980	
	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)
Hosmer-Lemeshow GOF <i>p</i> -value	0.73626	0.34206	0.63257	0.00790
Model AIC	75.5	242.7	78.0	244.6
Alpha (intercept)	–1.9798 (SE = 1.2270)	–3.4130 (SE = 1.1368)	–3.4612 (SE = 0.9790)	–4.14881 (SE = 1.4937)
Bkg (background)	0.03682 (SE = 0.04037)	0 (–)	0.0594 (SE = 0.0360)	0.0079 (SE = 0.0491)
Beta for TSFE	0.1075 (fixed)	0.1075 (SE = 0.0281), <i>p</i> = 0.0002	0.1167 (fixed)	0.1167 (SE = 0.0379), <i>p</i> = 0.0023
Beta for ln (mean exposure)	1.2750 (SE = 0.7159), <i>p</i> = 0.0775	0.4819 (SE = 0.1390), <i>p</i> = 0.0006	1.4054 (SE = 1.3328), <i>p</i> = 0.2938	0.4223 (SE = 0.1517), <i>p</i> = 0.0058
BMC and BMCL at 28 yr (f/cc)	0.0923 and $2.6 \times 10^{-2}$		0.2761 and $3.9 \times 10^{-2}$	
	Exposures based on geometric mean		Exposures based on geometric mean, truncated at 1980	
	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)
Hosmer-Lemeshow GOF <i>p</i> -value	0.23188	0.07022	0.68399	0.02812
Model AIC	75.7	244.1	78.1	245.6
Alpha (intercept)	–1.0630 (SE = 1.9499)	–3.6638 (SE = 1.1191)	–2.4799 (SE = 1.9039)	–4.1735 (SE = 1.0727)
Bkg (background)	0.0367 (SE = 0.0461)	0 (–)	0.0578 (SE = 0.0401)	0 (–)
Beta for TSFE	0.1234 (fixed)	0.1234 (SE = 0.0276), <i>p</i> < 0.0001	0.1266 (fixed)	0.1266 (SE = 0.0276), <i>p</i> < 0.0001
Beta for ln (mean exposure)	1.2527 (SE = 0.7173), <i>p</i> = 0.0833	0.4012 (SE = 0.1213), <i>p</i> = 0.0011	1.2118 (SE = 1.0841), <i>p</i> = 0.2659	0.3256 (SE = 0.1015), <i>p</i> = 0.0015
BMC and BMCL at 28 yr (f/cc)	0.0298 and $9.1 \times 10^{-3}$		0.0796 and $9.9 \times 10^{-3}$	

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Another potential source of uncertainty in the exposure reconstruction is the method used to average multiple fiber measurements for a given location (within the Marysville facility) and time period. The arithmetic mean of multiple measurements was used here, but an alternative approach would be to use the geometric mean, which has the effect of “dampening” outliers or extreme values. EPA conducted a sensitivity analysis for exposures estimated using the geometric mean of multiple fiber measurements and found that the PODs were lower than those estimated for the primary analysis ( $9.1 \times 10^{-3}$  fibers/cc considering all years of exposure, and  $9.9 \times 10^{-3}$  fibers/cc if exposures were truncated after 1980) (see Table 5-13). The lower PODs are a result of the approximately threefold lower exposures estimated for each individual worker when using the geometric mean; in the Marysville workers evaluated in 2002–2005 and hired in 1972 or later, the median of each individual’s mean exposure intensity estimated using the arithmetic versus the geometric mean of multiple fiber measurements in the plant were 0.0234 fiber/cc and 0.0085 fiber/cc, respectively. Note that for both models utilizing geometric mean exposures, the modeling in the larger group of workers evaluated in 2002–2005 showed poor fit (Hosmer-Lemeshow GOF  $p$ -value < 0.10). These results are included for comparative purposes but should be interpreted with caution.

Potential coexposure to other chemicals was present in the Marysville facility (see Section 4.1.2.2.2). These other chemicals were used after expansion of vermiculite ore in another area of the facility. Industrial hygiene data showed very low levels of fibers in the areas where the additional chemicals were added to the expanded vermiculite. In addition, none of these chemicals are volatile. The most likely route of exposure to these chemicals is through dermal contact. It is unlikely that any coexposure to these particular chemicals would alter the exposure-response relationship of LAA in the respiratory system (see Section 4.1.2.2.2).

The University of Cincinnati research team assumed no exposure to LAA occurred outside of the workplace for the Marysville workers. The interviews with the Marysville workers revealed that about 10% of the workers reported bringing raw vermiculite home. These interviews also revealed that changing clothes before leaving the workplace was standard practice at the end of the shift, and approximately 64% of the workers showered before leaving the workplace. For these workers, it is likely that additional exposure outside the workplace was minimal. However, for the remainder of the workers, it is reasonable to assume that additional exposure could have occurred at home. Additional data collected by the University of Cincinnati research team document that no increased prevalence of pleural or parenchymal change consistent with asbestos exposure was observed in household contacts of the workers from the Marysville facility ([Hilbert et al., 2013](#)).

### **5.3.2. Uncertainty in the Radiographic Assessment of Localized Pleural Thickening (LPT)**

The use of conventional radiographs to diagnose pleural thickening has several limitations. The localized thickening must be of sufficient size and thickness to be viewed on the

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x-ray, and small lesions may exist but not reported. More severe and larger lesions are more reliably detected on radiographs. There are also potential interferences. Fat pads may be mistaken for pleural plaques because they generally occur against the ribcage in a similar location ([Gilmartin, 1979](#)); this is one source of potential disagreement among x-ray readers. Often signs of trauma (e.g., fractured ribs) and radiographic signs of past tuberculosis infection can be seen and are noted by the reader. In these cases, LPT would not be diagnosed. There is a certain amount of subjectivity when viewing the x-rays in determining which features are representative of pleural thickening and whether signs of alternative etiology can be noted; thus, multiple certified readers are generally consulted, and a consensus of opinions determines the diagnosis. Regardless, the potential for outcome misclassification still exists. However, uncertainty in the presence or absence of localized pleural thickening in each individual is decreased by the use of three highly qualified chest radiologists evaluating the radiographic films and the use of the majority vote of the readers for the diagnosis.

BMI was investigated as a potential explanatory variable because fat pads can sometimes be misdiagnosed as pleural thickening. The effect of such outcome misclassification would be to attenuate the observed association between exposure and outcome. In the Marysville data, BMI was not measured in the 1980 examination but was available for most participants of the 2000s examination (available for most of those in the data sets used to derive the RfC). To address whether fat deposits may affect outcome classification, EPA considered the effect of adding BMI as a covariate in the model. However, BMI did not display an association with odds of localized pleural thickening in this population ( $p = 0.6933$ ).

### 5.3.3. Uncertainty Due to Potential Confounding

Along with the effect of BMI, other covariates were also evaluated for potential confounding of the association between LAA exposure and LPT in the Marysville workers (see Section 5.2.2.6.1). Covariates included both demographic characteristics (gender, smoking status, BMI) as well as potentially exposure-related factors (hire year, job tenure, exposure duration, and age at x-ray).

Smoking is a particularly important variable to consider when evaluating respiratory health outcomes. Although data are mixed, a few studies suggest smoking may affect the risk of developing asbestos-related pleural thickening or timing of such pleural thickening development. However, no studies were identified that assessed the relationship between LPT specifically and any measure of smoking status. Plaques as defined in earlier ILO classification systems have not been associated with smoking in asbestos-exposed workers ([Mastrangelo et al., 2009](#); [Paris et al., 2009](#); [Koskinen et al., 1998](#)).

Some evidence indicates that small interstitial opacities (asbestosis) and asbestos-related DPT may be associated with smoking. Studies among populations exposed to other general types of asbestos have reported mixed effects on the impact of smoking on risk of radiographic

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1 abnormalities; two studies reported a significant association between risk of all pleural  
2 thickening, including both pleural plaques and diffuse pleural thickening ([McMillan et al., 1980](#)),  
3 or any pleural abnormality ([Welch et al., 2007](#)) and smoking after controlling for some measure  
4 of asbestos exposure. A larger number of studies reported borderline associations when  
5 examining risk of pleural changes ([Adgate et al., 2011](#); [Paris et al., 2008](#); [Dement et al., 2003](#);  
6 [Zitting et al., 1996](#); [Yano et al., 1993](#); [Lilis et al., 1991](#); [Baker et al., 1985](#)) or no association with  
7 smoking ([Soulat et al., 1999](#); [Neri et al., 1996](#); [Ehrlich et al., 1992](#); [Delclos et al., 1990](#);  
8 [Rosenstock et al., 1988](#)). Possible reasons for the different findings include varying quality of  
9 smoking information (some used categories of ever/never or former/current/never, while others  
10 used pack-years) and differences in the specific outcome studied.

11 As the current classification of LPT includes cases that would have been classified as  
12 diffuse pleural thickening without costophrenic angle involvement in previous ILO guidelines,  
13 investigation of the potential for smoking to modify the risk of LPT is warranted. In the Libby  
14 workers cohort, [McDonald et al. \(1986b\)](#) assessed pleural thickening of the chest wall (both  
15 discrete and diffuse regardless of CPA involvement) and found smoking status (current, former,  
16 or never smoker) was of borderline statistical significance ( $p = 0.10$ ) in a regression model,  
17 controlling for LAA exposure and age. This result is consistent with the broader asbestos  
18 literature, addressing all pleural thickening or all pleural abnormalities. [Amandus et al. \(1987a\)](#)  
19 evaluated radiographic abnormalities consistent with the current LPT designation; the authors  
20 took a different analytic approach to assess smoking effects, constructing separate models for the  
21 full cohort and restricting to current and former smokers. The parameter estimates were not  
22 statistically significant for the two models, although the coefficients corresponding to LAA  
23 exposure were slightly higher for the full cohort model. In the Marysville workers cohort,  
24 smoking was characterized using pack-years in the original study ([Lockey et al., 1984](#)) and as  
25 ever or never smoking in the follow-up study ([Rohs et al., 2008](#)). [Lockey et al. \(1984\)](#) reported  
26 that the pack-years variable was statistically significantly associated with risk of all radiographic  
27 changes using discriminate analysis (any pleural thickening, small interstitial opacities, and  
28 blunting of the CPA) but did not present results for effect of smoking controlling for LAA  
29 exposure. [Rohs et al. \(2008\)](#) did not find a difference in smoking prevalence among those with  
30 and without any radiographic changes but also did not report results controlling for LAA  
31 exposure, or for LPT specifically.

32 None of the potential confounding factors examined were significantly associated with  
33 LPT after controlling for LAA exposure in the primary data set of workers evaluated in  
34 2002–2005 and hired in 1972 or later. However, the effect of each covariate was reexamined in  
35 the primary (hybrid) model which utilized information from the larger set of workers evaluated  
36 in 2002–2005, regardless of hire date (see Table 5-14). Each covariate was included (one at a  
37 time) in the primary model, and the statistical significance and effect on the model parameters  
38 evaluated. None were statistically significantly related to risk of LPT, with  $p$ -values for the

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1 corresponding beta coefficients ranging from 0.1533 (gender) to 0.9858 (hire year). Note that  
2 because these main effects were not statistically significantly associated with LPT, they would  
3 not be expected to modify the association between LAA exposure (controlling for TSFE) and  
4 LPT. It is unlikely that there would be opposing effects for exposure and the covariates  
5 examined that would “cancel each other out” and mask true effect measure modification. Thus,  
6 effect modification was not considered to be an issue in these analyses. It is not surprising that  
7 the time-related factors were not statistically significant (and had little effect on the estimated  
8 beta coefficient for exposure) because these factors were highly correlated with TSFE (which  
9 was already included in the primary model). The effect of exposure was higher when including  
10 gender and smoking status (current, ex- or never smoker) and lower when including ever-smoker  
11 status and BMI. Although the effect of gender was not significant, there were relatively few  
12 women in the Marysville workers population ( $n = 16$  in the workers evaluated in 2002–2005,  
13 and  $n = 13$  in the subgroup hired in 1972 or later). In the primary analytic group, the prevalence  
14 of LPT was not very different by gender, but the comparison is limited by sample size—there  
15 was 1 woman among the total of 13 with LPT (prevalence of 7.7%), while the other 12 cases  
16 (including the individual with LPT and DPT) were among the 106 men (prevalence of 11.3%).  
17 However, women also tended to have lower LAA exposure in this study population—for  
18 example, median cumulative exposure was 0.17 (interquartile range: 0.05, 0.26) fiber/cc-yr  
19 among women, compared with 0.56 (interquartile range: 0.23, 1.78) fibers/cc-yr among  
20 men—which could explain the lower prevalence. [Larson et al. \(2012a\)](#) found that among Libby  
21 workers (93.2% of whom were male), the prevalence of LPT was 37% among men and 9%  
22 among the 23 women included in the study, but exposure levels by gender were not provided in  
23 the published report. In the Minneapolis community study, the prevalence of all pleural  
24 abnormalities was 16.5% among men, compared to 4.6% among women ([adjusted odds ratio and](#)  
25 [95% confidence interval of 3.8 \[1.6, 8.9\]; Alexander et al., 2012](#)), but again, exposure levels by  
26 gender were not reported. Thus, the potential for different effects by gender merits further  
27 investigation.

28 In the Marysville workers, the variables representing smoking history (either current  
29 versus ex- versus never smoker, or ever smoker versus never smoker) were not statistically  
30 significant. However, the limited sample size (only three cases were never smokers) and limited  
31 nature of the smoking information precluded further analysis of smoking; thus, further research  
32 is needed on the effect of smoking in relation to LPT risk among asbestos-exposed populations.

**Table 5-14. Effect of including covariates into the final model**

	<b>Primary analysis</b>	<b>Including gender</b>	<b>Including ever smoker status</b>	<b>Including smoking status (compared to never smoker)</b>	<b>Including BMI</b>
Hosmer-Lemeshow GOF <i>p</i> -value	0.73625	0.72137	0.43468	0.88918	0.95082
Model AIC	75.5	75.0	76.1	78.5	60.7†
Alpha (intercept)	−1.9798 (SE = 1.2270)	4.8281 (SE = 5.1146)	−2.8954 (SE = 1.2344)	1.2111 (SE = 4.0064)	−3.0224 (SE = 2.1693)
Bkg (background)	0.03682 (SE = 0.04037)	0.0603 (SE = 0.0258)	0.0212 (SE = 0.0246)	0.0677 (SE = 0.0262)	0 (–)
Beta for TSFE	0.1075 (fixed)	0.1075 (fixed)	0.1075 (fixed)	0.1075 (fixed)	0.1075 (fixed)
Beta for ln (mean exposure)	1.2750 (SE = 0.7159), <i>p</i> = 0.0775	2.8080 (SE = 1.5440), <i>p</i> = 0.0715	1.1182 (SE = 0.4558), <i>p</i> = 0.0156	3.6756 (SE = 2.6246), <i>p</i> = 0.1640	0.9723 (SE = 0.3583), <i>p</i> = 0.0079
Beta for covariate	--	−5.0804 (SE = 3.5349), <i>p</i> = 0.1533	1.1701 (SE = 0.9993), <i>p</i> = 0.2440	Ex-smoker: −2.2659 (SE = 2.9748), <i>p</i> = 0.4477 Current smoker: 2.2180 (SE = 2.7013), <i>p</i> = 0.4132	0.0238 (SE = 0.0600), <i>p</i> = 0.6933
		<b>Including hire yr</b>	<b>Including job tenure (yr)</b>	<b>Including exposure duration (yr)</b>	<b>Including age at x-ray (yr)</b>
Hosmer-Lemeshow GOF <i>p</i> -value		0.73619	0.46477	0.70317	0.41903
Model AIC		77.5	77.2	77.4	77.3
Alpha (intercept)		−2.0004 (SE = 0.0006)	−1.2689 (SE = 1.9188)	−1.5075 (SE = 1.9288)	0.4511 (SE = 8.4602)
Bkg (background)		0.0368 (SE = 0.0404)	0.0350 (SE = 0.0426)	0.0368 (SE = 0.0421)	0.0437 (SE = 0.0558)
Beta for TSFE		0.1075 (fixed)	0.1075 (fixed)	0.1075 (fixed)	0.1075 (fixed)
Beta for ln (mean exposure)		1.2757 (SE = 0.7175), <i>p</i> = 0.0779	1.2498 (SE = 0.6919), <i>p</i> = 0.0734	1.2747 (SE = 0.7179), <i>p</i> = 0.0784	1.4362 (SE = 1.2550), <i>p</i> = 0.2548
Beta for covariate		0.00001 (SE = 0.0006), <i>p</i> = 0.9858	−0.0358 (SE = 0.0665), <i>p</i> = 0.5914	−0.0234 (SE = 0.0681), <i>p</i> = 0.7321	−0.0422 (SE = 0.1391), <i>p</i> = 0.7621

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#### 5.3.4. Uncertainty Due to Time Since First Exposure (TSFE)

Some uncertainty is associated with the length of follow-up of the Marysville cohort. There was relatively little variation in TSFE among the workers evaluated in 2002–2005 and hired in 1972 or later. It is anticipated that the prevalence of localized pleural thickening in the study population will likely continue to increase with passage of time. However, EPA took that into account both in modeling and in its application of uncertainty factors. EPA utilized information from the broader group of workers evaluated in 2002–2005 (i.e., regardless of hire date) and with a wider range of TSFE, to estimate the effect of time. However, because even this larger group lacked information on full lifetime exposure (maximum TSFE of 47 years), the modeling approach may not accurately reflect the exposure-response relationship that would be seen with a longer follow-up time. As one approach to gauge the sensitivity of the model, the plateau parameter—representing theoretical maximum prevalence of LPT when both exposure and TSFE are very large—was investigated further (see Table 5-15). As described above, the plateau parameter for the primary modeling was fixed at a literature-derived value of 85% ([Järholm, 1992](#); [Lilis et al., 1991](#)). The sensitivity of the POD to this assumption was investigated by looking at two alternative fixed values, 70 and 100% (i.e., the log-logistic model), as well as estimating the plateau parameter from the data. The value of 70% was selected for sensitivity analysis because in a cross-sectional study of Libby workers and residents seen at a clinic in Libby, [Winters et al. \(2012\)](#) observed a prevalence of 72% for pleural thickening, although the maximum TSFE was not known. Note that for the model with estimated rather than fixed plateau, the modeling in the larger group of workers evaluated in 2002–2005 showed poor fit (Hosmer-Lemeshow GOF  $p$ -value < 0.10); these results are included for comparative purposes but should be interpreted with caution.

**Table 5-15. Effect of different assumptions for the plateau parameter**

	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)	Primary	Individuals with radiographs in 2002–2005 (to get beta for TSFE)
	Plateau fixed at 85%—Primary analysis		Plateau fixed at 100% (log-logistic model)	
Hosmer-Lemeshow GOF <i>p</i> -value	0.73626	0.34206	0.74177	0.13912
Model AIC	75.5	242.7	75.3	243.3
Plateau	0.85 (fixed)	0.85 (fixed)	1.0 (fixed)	1.0 (fixed)
Alpha (intercept)	–1.9798 (SE = 1.2270)	–3.4130 (SE = 1.1368)	–2.0260 (SE = 1.0437)	–3.5167 (SE = 1.0092)
Bkg (background)	0.03682 (SE = 0.04037)	0 (–)	0.0359 (SE = 0.0403)	0 (–)
Beta for TSFE	0.1075 (fixed)	0.1075 (SE = 0.0281), <i>p</i> = 0.0002	0.0969 (fixed)	0.0969 (SE = 0.0245), <i>p</i> = 0.0001
Beta for ln (mean exposure)	1.2750 (SE = 0.7159), <i>p</i> = 0.0775	0.4819 (SE = 0.1390), <i>p</i> = 0.0006	1.2109 (SE = 0.6454), <i>p</i> = 0.0631	0.4007 (SE = 0.1093), <i>p</i> = 0.0003
BMC/BMCL at 28 yr (f/cc)	0.0923/2.6 × 10 <sup>–2</sup>		0.0924/2.6 × 10 <sup>–2</sup>	
	Plateau fixed at 70%		Plateau estimated from the Marysville data rather than fixed	
Hosmer-Lemeshow GOF <i>p</i> -value	0.72544	0.26136	0.73285	0.07394
Model AIC	75.7	242.4	77.5	244.3
Plateau	0.70 (fixed)	0.70 (fixed)	1 (–)	0.6263 (SE = 0.2611)
Alpha (intercept)	–2.0159 (SE = 1.5078)	–3.2410 (SE = 1.5622)	–3.5258 (SE = 1.0451)	–3.2538 (SE = 1.8317)
Bkg (background)	0.0378 (SE = 0.0407)	0.0027 (SE = 0.0386)	0.0385 (SE = 0.0401)	0.0130 (SE = 0.0483)
Beta for TSFE	0.1247 (fixed)	0.1247 (SE = 0.0417), <i>p</i> = 0.0060	0.1458 (fixed)	0.1458 (SE = 0.1087), <i>p</i> = 0.1812
Beta for ln (mean exposure)	1.3610 (SE = 0.8225), <i>p</i> = 0.1006	0.6385 (SE = 0.2474), <i>p</i> = 0.0104	1.2073 (SE = 0.6537), <i>p</i> = 0.0673	0.8311 (SE = 0.9541), <i>p</i> = 0.3845
BMC/BMCL at 28 yr (f/cc)	0.0920/2.7 × 10 <sup>–2</sup>		0.1022/3.0 × 10 <sup>–2</sup>	

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1 The effects on the POD resulting from different assumptions regarding the plateau were  
2 small. When assuming different fixed values (70, 85, and 100%) the POD only ranged from  
3  $2.6 \times 10^{-2}$  to  $2.7 \times 10^{-2}$  fibers/cc. Estimating the plateau from the data led to a slightly higher  
4 POD of  $3.0 \times 10^{-2}$  fibers/cc. These results lend confidence that assumptions regarding maximum  
5 prevalence of LPT in the population do not have a substantial impact on the estimated POD.

6 As described in Section 5.2.2.6.1, one option to incorporate TSFE would be to utilize the  
7 RTW exposure metric, which incorporates aspects of both exposure duration and TSFE. This  
8 option was not selected for RfC derivation due to the narrow range of TSFE among the primary  
9 analytic group of Marysville workers who underwent health evaluations in 2002–2005 and  
10 whose job start date was on or after 1/1/1972. However, this approach was used as a sensitivity  
11 analysis, estimating the concentration that, if experienced over 70 years, would yield the BMR.  
12 The model with RTW as the metric derives as its POD a “benchmark residence time-weighted”  
13 quantity in units of fibers/cc-yr<sup>2</sup> and its associated confidence interval. In order to convert the  
14 benchmark quantity in units of fibers/cc-yr<sup>2</sup> and its associated lower limit into a 70-year  
15 exposure concentration (in units of fiber/cc), the constant 70-year concentration yielding that  
16 RTW should be determined where 70 years is both the duration and the time elapsed between the  
17 first year of exposure and the health evaluation. That concentration is equal to the benchmark  
18 RTW (or its lower limit) divided by the residence time-weighted value for exposures across  
19 70 years:  $1 + \dots + 70 = [(70 \times 71 \text{ years})/2]$ , as sum of first  $N$  natural numbers is equal to  
20  $(N \times (N + 1))/2$ . The results of using RTW exposure with the preferred model (Dichotomous  
21 Hill with plateau fixed at 85%) are shown in Table 5-16. For comparison, results also using the  
22 RTW exposure metric but using alternate model forms are also shown; the model fits and results  
23 are very similar for the Dichotomous Hill, log-logistic, and log-probit models, with the PODs all  
24  $\sim 0.003$  fiber/cc for a scenario of 70-years exposure duration and 70-years TSFE. The  
25 Michaelis-Menten model provided a lower AIC (2 units lower than the Dichotomous Hill  
26 model), and the POD was slightly higher (0.0057 versus 0.0034 fiber/cc).



**Table 5-16. Exposure-response modeling for any localized pleural thickening (LPT) in the Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 ( $n = 119$ ), using a benchmark response (BMR) of 10% extra risk of any LPT, and RTW exposure**

	Dichotomous Hill, plateau = 85%	Michaelis-Menten	Log-logistic	Log-probit
Hosmer-Lemeshow GOF $p$ -value	0.7527	0.7528	0.7576	0.7548
AIC	76.8	74.8	77.0	76.8
Intercept (SE)	−5.9883 (2.5304)	−5.4305 (0.5333)	−5.7331 (2.0944)	−3.2071 (1.0398)
Background rate (SE)	0.0366 (0.0338)	0.0320 (0.0271)	0.0342 (0.0331)	0.0370 (0.0321)
B (SE), $p$ -value	1.1266 (0.5493), 0.0425	--	1.0073 (0.4394), 0.0236	0.5546 (0.2264), 0.0158
Benchmark RTW (f/cc-yr <sup>2</sup> )	34.2516	30.6380	33.4520	32.2062
Benchmark RTW lower limit (f/cc-yr <sup>2</sup> )	8.32733	14.2184	7.50156	9.39139
BMC for TSFE = 28 yr (f/cc) <sup>a</sup>	0.08436	0.075463	0.082394	0.079326
BMCL for TSFE = 28 yr (f/cc) <sup>a</sup>	0.020511	0.035021	0.018477	0.023132

<sup>a</sup>BMCs and BMCLs are expressed in fiber/cc, and are estimated as benchmark value or its lower limit divided by  $[(70 \times 71)/2]$  yr<sup>2</sup> or divided by  $[(28 \times 29)/2]$  yr<sup>2</sup>.

Advantages of this approach using the RTW exposure metric in the subcohort are that it relies solely on the individuals with higher quality exposure information and consistent radiograph evaluation, and uses an exposure metric that weights more heavily exposure occurring in the more distant past. However, the modeling still relies solely on the subgroup of workers with little variation in TSFE, and whose TSFE values are for less than a full lifetime. This lack of variability in TSFE limits the ability to explore how risk of LPT varies across the life span. However, it does provide an important comparison for the primary RfC; the BMCLs estimated using this approach range from 0.018 to 0.035 fiber/cc, similar to the BMCL estimated in the primary analysis (0.026 fiber/cc).

Another source of information regarding TSFE comes from the study by [Larson et al. \(2010a\)](#) which examined serial radiographs conducted on a group of Libby vermiculite workers

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with pleural or parenchymal changes. The mean follow-up time was 21.6 years, with a maximum of 44.9 years. They found that among those workers with localized pleural thickening, all cases were identified within 30 years, and that the median time from hire to the first detection of localized pleural thickening was 8.6 years. Although the retrospective evaluation of radiographs is a different and more sensitive procedure, these findings indicate that the range of follow-up time in the Marysville subcohort is likely sufficient to support the exposure-response modeling developed in this current assessment. Note that the likelihood that prevalence of localized pleural thickening is expected to increase over the life span is a principal rationale cited for the selection of a subchronic-to-chronic UF of 10 in this current assessment.

### **5.3.5. Uncertainty in the Endpoint Definition**

The critical effect selected for RfC derivation is localized pleural thickening. As a sensitivity analysis, an alternative critical effect of any radiographic change was also investigated and found to yield an essentially identical POD (i.e., a BMCL of  $2.7 \times 10^{-2}$ , compared with  $2.6 \times 10^{-2}$  in the primary analysis), as that using the same modeling approach in the primary analysis. Almost no information existed on radiographic changes other than LPT in the primary analytic group of workers (evaluated in 2002–2005, hired in 1972 or later) because only one case of DPT was reported and that individual also had LPT. No individuals had interstitial changes. However, some individuals had DPT and/or interstitial changes (with and without LPT) in the larger group of workers evaluated in 2002–2005, which allowed investigation of the effect of TSFE considering alternative endpoint definitions.

The primary analysis contrasted individuals with LPT (with or without other radiographic endpoints) to those without any radiographic changes. In the group of workers evaluated in 2002–2005, this had the effect of excluding five individuals with DPT and/or interstitial changes, but without LPT. In the subgroup of workers hired in 1972 or later, this distinction had no effect because the single case of DPT also had LPT (and thus was included as a case), and no individuals showed interstitial changes. As a sensitivity analysis, the modeling procedure was repeated, using three alternative endpoint definitions (see Table 5-17). The first contrasts all those with LPT (with or without other endpoints) to those without LPT. Thus, the comparison group could include those with DPT and/or interstitial changes but without LPT. The second model used an endpoint of “any radiographic change.” The third sensitivity model contrasted those with LPT only to those with no radiographic abnormalities. Note that for two of the alternative models (those contrasting LPT with no LPT, and any radiographic change with no radiographic change) the modeling in the larger group of workers evaluated in 2002–2005 showed poor fits (Hosmer-Lemeshow GOF  $p$ -values  $< 0.10$ ); these results are included for comparative purposes, but should be interpreted with caution.

**Table 5-17. Effect of using different case/noncase definitions**

	<b>LPT vs. no radiographic abnormalities— primary analysis</b>		<b>LPT vs. no LPT</b>		<b>Any radiographic abnormality vs. no radiographic abnormalities</b>		<b>LPT alone vs. no radiographic abnormalities</b>		<b>Any pleural thickening vs. no radiographic abnormalities</b>	
	All individuals with radiographs in 2002–2005 (n = 247)	PRIMARY (hired ≥1972, n = 119)	All individuals with radiographs in 2002–2005 (n = 252)	PRIMARY (hired ≥1972, n = 119)	All individuals with radiographs in 2002–2005 (n = 252)	PRIMARY (hired ≥1972, n = 119)	All individuals with radiographs in 2002–2005 (n = 237)	PRIMARY (hired ≥1972, n = 118)	All individuals with radiographs in 2002–2005 (n = 237)	PRIMARY (hired ≥1972, n = 118)
Hosmer-Lemeshow GOF <i>p</i> -value	0.34206	0.73626	0.05378	0.73640	0.09052	0.73552	0.18854	0.73497	0.10716	0.73565
AIC	242.7	75.5	249.9	75.5	250.0	75.5	230.9	74.1	247.0	75.5
Alpha (intercept)	−3.4130 (SE = 1.1368)	−1.9798 (SE = 1.2270)	−3.4421 (SE = 1.1192)	−1.8537 (SE = 1.2275)	−3.4126 (SE = 1.1376)	−2.1027 (SE = 1.2268)	−3.7837 (SE = 1.1633)	−2.2944 (SE = 1.1804)	−3.4334 (SE = 1.1347)	−2.0812 (SE = 1.2268)
Bkg (background)	0 (--)	0.03682 (SE = 0.04037)	0 (--)	0.03669 (SE = 0.04038)	0 (--)	0.03697 (SE = 0.04037)	0 (--)	0.0341 (SE = 0.0409)	0 (--)	0.03694 (SE = 0.04037)
Beta for TSFE	0.1075 (SE = 0.0281), <i>p</i> = 0.0002	0.1075 (fixed)	0.1034 (SE = 0.0275), <i>p</i> = 0.0002	0.1034 (fixed)	0.1115 (SE = 0.0282), <i>p</i> = 0.0001	0.1115 (fixed)	0.1056 (SE = 0.0282), <i>p</i> = 0.0002	0.1056 (fixed)	0.1108 (SE = 0.0282), <i>p</i> = 0.0001	0.1108 (fixed)
Beta for ln(mean exposure)	0.4819 (SE = 0.1390), <i>p</i> = 0.0006	1.2750 (SE = 0.7159), <i>p</i> = 0.0775	0.4342 (SE = 0.1324), <i>p</i> = 0.0012	1.2764 (SE = 0.7161), <i>p</i> = 0.0772	0.5125 (SE = 0.1415), <i>p</i> = 0.0004	1.2739 (SE = 0.7160), <i>p</i> = 0.0778	0.3507 (SE = 0.1404), <i>p</i> = 0.0132	1.1247 (SE = 0.6467), <i>p</i> = 0.0846	0.5042 (SE = 0.1409), <i>p</i> = 0.0004	1.2741 (SE = 0.7160), <i>p</i> = 0.0777
BMC/BMCL at 28 yr (f/cc)	--	0.0923/ $2.6 \times 10^{-2}$ (Ratio = 3.5)	--	0.0918/ $2.6 \times 10^{-2}$ (Ratio = 3.5)	--	0.0929/ $2.7 \times 10^{-2}$ (Ratio = 3.5)	--	0.0931/ $2.5 \times 10^{-2}$ (Ratio = 3.8)	--	0.0928/ $2.7 \times 10^{-2}$ (Ratio = 3.4)

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1 The effect of TSFE in these sensitivity analyses using different endpoint definitions was very  
2 similar to that in the primary model and led to nearly identical PODs ( $2.5$  to  $2.7 \times 10^{-2}$  fibers/cc).  
3 These results lend confidence to the primary analysis.

4 Additionally, EPA conducted a sensitivity analysis using a multinomial model to incorporate  
5 information from all outcome groups. The multinomial logistic model is similar to the logistic  
6 model and compares each outcome group to the referent group (i.e., no radiographic change) and  
7 estimates separate model parameters (intercepts and beta coefficients) for each comparison.  
8 With only two outcome groups, the logistic and multinomial models are equivalent. As  
9 described earlier, there were noticeable differences when contrasting individuals who had LPT  
10 alone, with those who had LPT along with DPT and/or interstitial changes. Thus, two different  
11 multinomial models were considered. In the first, the outcome groups were (1) no radiographic  
12 change (referent), (2) any LPT (with or without other radiographic changes), and (3) DPT and/or  
13 interstitial changes (without LPT) (see Table 5-18). In the second model, the outcome groups  
14 were defined as (1) no radiographic change (referent), (2) LPT alone, (3) LPT along with other  
15 radiographic changes, and (4) DPT and/or interstitial changes (without LPT). Each of these was  
16 contrasted to a logistic model, which is equivalent except that those with DPT and/or interstitial  
17 changes without LPT are excluded (i.e., as was done for the primary analysis).

**Table 5-18. Exposure-response modeling for any localized pleural thickening (LPT) in the Marysville workers who underwent health evaluations in 2002–2005 ( $n = 252$ ), comparing the multinomial model and logistic model with different outcome group definitions<sup>a</sup>**

	<b>Logistic</b>	<b>Multinomial Model 1</b>	<b>Multinomial Model 2</b>
Pearson GOF $p$ -value	0.5170	0.9653	0.9999
AIC <sup>b</sup>	249.6	299.3	348.5
Alpha <sub>1</sub> (intercept)	−5.1351 (SE = 0.8638)	−5.1980 (SE = 0.8695)	−5.3174 (SE = 0.8946)
Alpha <sub>2</sub> (intercept)	--	−8.4598 (SE = 2.9194)	−7.4075 (SE = 2.4490)
Alpha <sub>3</sub> (intercept)	--	--	−8.3324 (SE = 2.9481)
Beta <sub>1</sub> for mean exposure	0.5878 (SE = 0.2596), $p = 0.0236$	0.5914 (SE = 0.2595), $p = 0.0225$	0.3208 (SE = 0.2957), $p = 0.2779$
Beta <sub>2</sub> for mean exposure	--	0.9443 (SE = 0.4625), $p = 0.0412$	1.5242 (SE = 0.4097), $p = 0.0002$
Beta <sub>3</sub> for mean exposure	--	--	1.0483 (SE = 0.4988), $p = 0.0356$
Beta <sub>1</sub> for TSFE	0.1103 (SE = 0.0237), $p < 0.0001$	0.1120 (SE = 0.0239), $p < 0.0001$	0.1144 (SE = 0.0245), $p < 0.0001$
Beta <sub>2</sub> for TSFE—DPT/interstitial changes vs. no change	--	0.1221 (SE = 0.0753), $p = 0.1050$	0.0958 (SE = 0.0651), $p = 0.1413$
Beta <sub>3</sub> for TSFE	--	--	0.1173 (SE = 0.0768), $p = 0.1266$

<sup>a</sup>The multinomial model is a generalized form of the logistic regression for >2 outcome categories (not ordered). The model is of the form

$$p_i(x, t) = 11 + \exp[-a_i - b_i \times x - c_i \times t]$$

Where  $p_i$  is the probability of being in the  $i^{\text{th}}$  outcome group, and separate intercepts ( $a$ ) and beta coefficients ( $b$ ,  $c$ ) are estimated for effect of predictors on probability of being in each group. Multinomial Model 1 contrasts no radiographic change (referent, Group 0) to those with any LPT (Group 1) and to those with DPT and/or interstitial changes but without LPT (Group 2). Multinomial Model 2 contrasts no radiographic change (referent, Group 0) to those with LPT alone (Group 1), to those with LPT along with DPT and/or interstitial changes (Group 2), and to those with DPT and/or interstitial changes but without LPT (Group 3).

<sup>b</sup>AIC not comparable between multinomial model and logistic model because the number of individuals is different (multinomial,  $n = 252$  compared to logistic,  $n = 247$ ).

1           The effect of TSFE was very similar across all the models and outcome groups, with the  
2 corresponding beta coefficient ranging from 0.0958 to 0.1221 (compared to 0.1075 in the primary  
3 analysis). The effect of mean exposure was much more variable, and the corresponding beta  
4 coefficients were notably higher for those with DPT and/or interstitial changes (either alone or  
5 along with LPT) compared to those for LPT alone. These results are in accordance with the  
6 descriptive statistics shown in Table 5-4, which highlighted that while exposure patterns were

different among outcome groups, there was relatively less variation in TSFE. These results lend confidence to the effect of TSFE used in the primary analysis.

#### **5.3.6. Summary of Sensitivity Analyses**

EPA conducted numerous sensitivity analyses for comparison with the primary analysis used to derive the RfC.<sup>23</sup> These included analyses to explore the effect of exposure assessment decisions (e.g., use of the cumulative exposure metric, truncation of exposures post-1980, and use of arithmetic versus the geometric mean for exposure reconstruction); potential confounding factors (time-related and nontime-related); the effect of TSFE (e.g., assumptions regarding the plateau and use of the RTW exposure metric); and the definition of cases and noncases (e.g., varying case/noncase groups, use of the multinomial model). The results of other sensitivity analyses are summarized in Table 5-19. In each case, the estimated BMCL was within an order of magnitude of the POD. The biggest impacts came from using cumulative exposure (rather than mean exposure), truncating exposures after 1980, and using the geometric mean versus the arithmetic mean for exposure reconstruction (differences of -68 to +50% from the POD). Assumptions regarding the plateau parameter (or estimating the plateau rather than fixing its value) had a very small effect on the BMCL (differences of 0 to +5% from the POD). Similarly, small differences in case and noncase definition led to small changes in the BMCL (differences of -3.9 to +3.9%). Finally, the use of RTW exposure alone (rather than mean exposure and TSFE) as the predictor in the subset of workers evaluated in 2002–2005 and hired in 1972 or later, led to a difference of -19.2% in the BMCL.

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<sup>23</sup>The primary analysis used a hybrid Dichotomous Hill model with plateau fixed at 85%, with mean exposure as the exposure metric. The effect of TSFE was estimated in the set of Marysville workers evaluated in 2002–2005, and carried over to the modeling performed in the subset of these workers who were hired in 1972 or later. The BMCL was estimated for a TSFE of 28 years, which served as the POD for RfC derivation. A composite UF of 300 was applied to account for various sources of uncertainty, leading to an RfC of  $9 \times 10^{-5}$  fibers/cc.

**Table 5-19. Summary of sensitivity analyses.** Exposure-response modeling performed using mean exposure in the hybrid Dichotomous Hill model with plateau fixed at 85%, Marysville workers who underwent health evaluations in 2002–2005 and whose job start date was on or after 1/1/1972 ( $n = 119$ ). Effect of TSFE estimated in workers evaluated in 2002–2005 regardless of hire date.

Sensitivity analysis	BMC/BMCL at 28 yr (fiber/cc)	Percentage difference in BMCL from primary analysis <sup>a</sup> [(sensitivity analysis-primary)/primary] × 100
Primary modeling	0.0923/2.6 × 10 <sup>-2</sup>	--
Use of cumulative exposure rather than mean exposure	0.0266/8.2 × 10 <sup>-3*</sup>	-68.46
Exposures based on arithmetic mean, truncated at 1980	0.2761/3.9 × 10 <sup>-2</sup>	+50.00
Exposures based on geometric mean	0.0298/9.1 × 10 <sup>-3</sup>	-65.00
Exposures based on geometric mean, truncated at 1980	0.0796/9.9 × 10 <sup>-3</sup>	-61.92
Plateau fixed at 70%	0.0920/2.7 × 10 <sup>-2</sup>	+3.85
Plateau fixed at 100%	0.0924/2.6 × 10 <sup>-2</sup>	0.00
Plateau estimated	0.1022/3.0 × 10 <sup>-2</sup>	+15.38
Contrast any LPT vs. no LPT	0.0918/2.6 × 10 <sup>-2</sup>	0.00
Contrast any radiographic change vs. no radiographic change	0.0929/2.7 × 10 <sup>-2</sup>	+3.85
Contrast LPT only vs. no radiographic change	0.0931/2.5 × 10 <sup>-2</sup>	-3.85
Alternative modeling using RTW exposure in the subgroup of workers evaluated in 2002–2005, hired in 1972 or later	0.0844/2.1 × 10 <sup>-2</sup>	-19.23

<sup>a</sup>The BMC and BMCL are 1.8622 and 0.5770 fibers/cc-yr, respectively. These values were divided by 70 yr to obtain the BMC and BMCL in terms of fiber/cc.

Multiple statistical model forms applied to different sets and subsets of the principal study population all yield results within less than an order of magnitude around the BMCL. Each of these sensitivity analyses further substantiates the BMCL used to derive the RfC.

## 5.4. CANCER EXPOSURE-RESPONSE ASSESSMENT

### 5.4.1. Overview of Methodological Approach

The inhalation unit risk (IUR) is defined as an upper-bound excess lifetime cancer risk estimated to result from continuous exposure to an agent at a concentration of 1 µg/L in water, or 1 µg/m<sup>3</sup> in air. However, current health standards for asbestos are based on health effects observed in occupational cohorts and are given in fiber/cc of air as counted by PCM ([OSHA](#),

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1994; U.S. EPA, 1988a). Thus, when examining the available health effects data on cancer for LAA, the best available studies at this time report exposure concentration in terms of fiber/cc counted by PCM (see Section 4.1.4). The cancer effects identified in populations with exposure to LAA (see Section 4.1.4) are cancer mortality from mesothelioma and lung cancer (see Section 5.4.2.2 for other cancers identified in populations exposed to asbestos in general). Therefore, the IUR represents the upper-bound excess lifetime risk of mortality from either mesothelioma or lung cancer in the general U.S. population from chronic inhalation exposure to LAA at a concentration of 1 fiber/cc of air.

IURs are based on human data when appropriate epidemiologic studies are available. The general approach to developing an IUR from human epidemiologic data is to first quantitatively evaluate the exposure-response relationship (slope) for that agent in the studied population. For the current assessment, the first step was to identify the most appropriate data set available to quantitatively estimate the effects of LAA exposure on cancer mortality. Once the relevant data describing a well-defined group of individuals along with their exposures and health outcomes were selected (see Section 5.4.2), an appropriate statistical model form (i.e., Poisson or Cox) was selected that adequately fit the specific nature of the data, and then each person's individual-level exposures were modeled using a variety of possible exposure metrics informed by the epidemiologic literature. Exposure-response modeling was conducted for each cancer mortality endpoint individually (see Section 5.4.3). In some cases, the statistical model forms and the specific metrics of exposure used for each cancer endpoint may have been different. For example, the 1988 EPA general asbestos assessment found different model forms/metrics for mesothelioma and lung cancer. Appropriate covariates, which may be important predictors of cancer mortality, were included in the statistical models. These models were then evaluated to assess how the different exposure metric representing estimated occupational exposures fit the observed epidemiologic data. The empirical model fits were compared against those models suggested by the epidemiologic literature before selecting one model for mesothelioma mortality and one for lung cancer mortality.

The selected cancer exposure-response relationships (slopes) for mesothelioma (KM) and lung cancer (KL), which were estimated from the epidemiologic data on the Libby workers cohort, were then applied to the general U.S. population in a life-table analysis using age-specific mortality statistics to determine the exposure level that would be expected to result in a specified level of response over a lifetime of continuous exposure. EPA typically selects a response level of 1% extra risk because this response level is generally near the low end of the observable range for such data. Extra risk is defined as equaling  $(R_x - R_o) \div (1 - R_o)$ , where  $R_x$  here is the lifetime cancer mortality risk in the exposed population and  $R_o$  is the lifetime cancer mortality risk in an unexposed population (i.e., the background risk). In the case of lung cancer, the expected lifetime risk of lung cancer mortality in the unexposed general U.S. population is approximately 5%; thus, this human health assessment seeks to estimate the level of exposure to LAA that

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would be expected to result in a 1% extra lifetime risk of lung cancer mortality equivalent to a lifetime risk of lung cancer mortality of 5.95%:  $[(0.0595 - 0.05) \div (1 - 0.05) = 0.01]$ . This corresponds to a relative risk ( $R_x/R_o$ ) of about 1.2, which is near the low end of the observable range for most epidemiologic studies of cancer. For mesothelioma mortality, an absolute risk was considered, rather than extra risk, for two reasons: (1) mesothelioma is very rare in the general population and (2) mesothelioma is almost exclusively caused by exposure to asbestos and other mineral fibers, including LAA. Because the background rate of mesothelioma is negligible, absolute risk models of exposure-response were considered more appropriate than relative risk models, thereby justifying the definition of the target response rate in absolute terms rather than in relative terms.

A life-table analysis (see Appendix G for details) was used to compute the 95% lower bound on the level of LAA at which a lifetime exposure corresponds to a 1% extra risk of lung cancer mortality (1% absolute risk for mesothelioma) in the general U.S. population using age-specific mortality statistics and the exposure-response relationships for each cancer endpoint as estimated in the Libby worker cohort. This lower bound on the level of exposure serves as the POD for extrapolation to lower exposures and for deriving the unit risk. Details of this analysis are presented in Section 5.4.5. Cancer-specific unit risk estimates were obtained by dividing the extra risk (1%) by the POD. The cancer-specific unit risk estimates for mortality from either mesothelioma or lung cancer were then statistically combined to derive the final IUR (see Section 5.4.5.3). Uncertainties in this cancer assessment are described in detail in Section 5.4.6.

#### **5.4.2. Choice of Study/Data—with Rationale and Justification**

This human health assessment is specific to LAA. The current assessment does not seek to evaluate quantitative exposure-response data on cancer risks from studies of asbestos that did not originate in Libby, MT. However, this assessment does draw upon the exposure-response models developed for other kinds of amphibole asbestos, as described in the epidemiologic literature, to address uncertainty in model selection.

The available sources of cancer data include the cohort of workers employed at the vermiculite mining and milling operation in and around Libby, MT. This cohort has been the subject of several epidemiologic analyses of cancer risks, described in detail in Section 4.1.4 ([Larson et al., 2010b](#); [Moolgavkar et al., 2010](#); [Sullivan, 2007](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). There have also been published reports on cases of mesothelioma in the Libby, MT area ([Whitehouse et al., 2008](#)) and mortality data published by the [ATSDR \(2000\)](#). However, published mortality data on Libby, MT residents ([Whitehouse et al., 2008](#); [ATSDR, 2000](#)) could not be used in exposure-response modeling due to lack of quantitative exposure data.

The most appropriate available data set with quantitative exposure data for deriving quantitative cancer mortality risk estimates based on LAA exposure in humans is the cohort of

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workers employed at the vermiculite mining and milling operation in and around Libby, MT (hereafter referred to as the Libby worker cohort). These data are considered the most appropriate to inform this human health assessment for several reasons: (1) these workers were directly exposed to LAA, (2) detailed work histories and job-specific exposure estimates are available to reconstruct estimates of each individual's occupational exposure experience, (3) the cohort is sufficiently large and has been followed for a sufficiently long period of time for cancer to develop (i.e., cancer incidence) and cause mortality, and (4) the broad range of exposure experiences in this cohort provided an information-rich data set, which allowed evaluation of several different metrics, or mathematical expressions, of exposure. Uncertainties in these data are discussed in Section 5.4.6.

The only other available cohort exposed to LAA was the cohort of workers from a Marysville, Ohio vermiculite processing plant (see Section 4.1.1.2; [Rohs et al., 2008](#); [Lockey et al., 1984](#)). The study of pleural changes in this population was the basis of the noncancer exposure-response analyses (see Section 5.3). Regarding mortality among the Marysville workers, [Dunning et al. \(2012\)](#) reported 2 mesothelioma deaths and 16 lung cancer deaths. The Libby worker cohort was a more suitable candidate for cancer exposure-response modeling than the Marysville worker cohort due to the larger number of cases (see Table 5-3 compared to Tables 5-20 and 5-22).

#### **5.4.2.1. Description of the Libby Worker Cohort**

Cancer mortality in the Libby worker cohort has been extensively studied (see Section 4.1.4). McDonald et al. ([2004](#), [2002](#); [1986a](#)) published three studies on a subset of the cohort. Scientists from NIOSH conducted two epidemiologic investigations, resulting in several published reports on different subsets of the cohort ([Sullivan, 2007](#); [Amandus et al., 1988](#); [Amandus and Wheeler, 1987](#)). [Berman and Crump \(2008\)](#) and [Moolgavkar et al. \(2010\)](#) reanalyzed the [Sullivan \(2007\)](#) data with mortality follow-up through 2001. [Larson et al. \(2010b\)](#) analyzed an ATSDR reconstruction of the Libby worker cohort from company records with exposure estimates obtained from NIOSH with mortality follow-up through 2006.

According to [Sullivan \(2007\)](#), nearly all of these study subjects were workers on-site at the Libby, MT vermiculite mine, mill, or processing plant. Although the mine and other facilities were several miles from downtown Libby, MT, some of the study subjects worked at vermiculite ore expansion plants, at the Export Plant, or at offices in the town (see Section 4.1.1.2). Workers may have also been assigned jobs as truck drivers or jobs working in the screening plant, railroad loading dock, expansion plants, or an office. Individuals' demographic and work history data were abstracted from company personnel and pay records. A database created by NIOSH in the 1980s contained demographic data and work history starting from September 1935 and vital status at the end of 1981 for 1,881 workers. NIOSH compared these data with company records on microfilm, and work history data were reabstracted to ensure

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data quality. One person was removed from the cohort because company records stated that he was hired but never worked ([Sullivan, 2007](#)). Nine workers with Social Security numbers listed in company records were excluded because demographic and work history data were not available, leaving 1,871 workers in the cohort available for epidemiologic analysis. Table 5-20 shows the demographic, mortality, and exposure characteristics of this cohort.

**Table 5-20. Demographic, mortality, and exposure characteristics of the Libby worker cohort**

Characteristic	All workers
Number of workers	1,871
Number of deaths from all causes	1,009
Number of deaths from mesothelioma	18
Number of deaths from lung cancer	111
Number of deaths from laryngeal cancer	2
Number of deaths from ovarian cancer	0
Number of deaths from intestinal or colorectal cancer	15
Number of deaths from chronic obstructive pulmonary disease	71
Mean yr of birth	1929
Mean yr of hire	1959
Mean age at hire (yr)	30.2
Mean person-yr of follow-up (no lag)	35.9
Total person-yr of follow-up (no lag)	67,101
Mean employment duration (yr)	2.6
Mean cumulative exposure (fiber/cc-yr)	96.0
Median cumulative exposure (fiber/cc-yr)	9.8
Range of cumulative exposures (no lag) (fiber/cc-yr) <sup>a</sup>	0–1,722

<sup>a</sup>According to the work histories and JEM, there were 26 workers who had zero exposure. These individuals (7 men and 19 women) all worked at the office downtown.

NIOSH updated the cohort vital status through 2006 using the National Death Index [NDI-Plus; Bilgrad \(1999\)](#), and these data were used for this analysis. Workers known to be alive on or after January 1, 1979 (the date NDI began tracking deaths nationwide), but not found in the NDI search, were assumed to have been alive on December 31, 2006 ([Sullivan, 2007](#)). Nearly 54% of workers in the cohort ( $n = 1,009$ ) had died by December 31, 2006. NIOSH researchers obtained death certificates from across the United States (while exposure occurred in and around Libby, deaths could have occurred elsewhere) for deaths prior to 1979, and the causes of death were coded to the ICD revision that was in effect at the time of death by a single National Center for Health Statistics-trained nosologist. After 1979, ICD codes were obtained from the NDI-Plus. For workers known to be deceased, the underlying cause of death was determined from death certificates and coded to the ICD codes using the rubrics of the ICD revision in effect at the time of death (ICD-5 ([WHO, 1938](#)), ICD-6 ([WHO, 1948](#)), ICD-7 ([WHO, 1957](#)), ICD-8 ([WHO, 1967](#)), ICD-9 ([WHO, 1977](#)), or ICD-10 ([WHO, 1992](#))).

Basic demographic information on the occupational cohort members was largely complete. However, when data were missing, they were statistically imputed (i.e., assigned) by NIOSH based on several reasonable assumptions regarding gender, race, and date of birth. For example, seven workers with unknown gender were assumed to be male because 96% of the workforce was male, and NIOSH's review of names did not challenge that assumption ([Sullivan, 2007](#)). Workers of unknown race ( $n = 935$ ) were assumed to be white because workers at this facility were known to be primarily white, and U.S. Census Bureau data from 2004 indicate that 90–95% of the local population identify themselves as white ([Sullivan, 2007](#)). Date of birth was estimated for four workers with unknown birth dates by subtracting the cohort's mean age at hire from the worker's hire date. The impact of this imputation procedure on the analytic results can reasonably be expected to be minimal.

#### **5.4.2.2. Description of Cancer Endpoints**

The cancer exposure-response assessment focuses on two cancer endpoints, mesothelioma and lung cancer, although there is evidence that other cancer endpoints may also be associated with exposure to asbestos in general. The IARC has concluded that sufficient evidence in humans is present that other types of asbestos (chrysotile, crocidolite, amosite, tremolite, actinolite, and anthophyllite) are causally associated with mesothelioma and lung cancer, as well as cancer of the larynx and the ovary ([Straif et al., 2009](#)). Among the entire Libby worker cohort, only two deaths were found to be due to laryngeal cancer, and no deaths from ovarian cancer occurred among the 84 female workers. Therefore, EPA did not evaluate these other outcomes as part of this current assessment. The limited number of female workers in this cohort is discussed later as a source of uncertainty in the derived estimates (see Section 5.4.6).

1 The endpoint for both mesothelioma and lung cancer was mortality, not incidence.  
2 Incidence data are not available for the Libby worker cohort. Nevertheless, mortality rates  
3 approximate incidence rates for cancers such as lung cancer and mesothelioma because the  
4 survival time between cancer incidence and cancer mortality is short. According to data from the  
5 National Cancer Institute's Surveillance Epidemiology and End Results (SEER ) data on cancer  
6 incidence, mortality, and survival ([Howlader et al., 2013](#)), the median length of survival with  
7 mesothelioma is less than 1 year, with 2-year survival for males about 20%, and 5-year survival  
8 for males about 6%. For lung cancer, the median length of survival is less than 1 year, with  
9 2-year survival for males about 25% and 5-year survival for males about 17%. Therefore, while  
10 the absolute rates of cancer mortality at follow-up may underestimate the rates of cancer  
11 incidence, it is considered to be unlikely that such discrepancies would be of significant  
12 magnitude. The use of mortality statistics instead of incidence statistics as a source of  
13 uncertainty in the derived estimates is further discussed in Section 5.4.6.

14 It is well established in the literature that mortality rates calculated from death certificates  
15 are lower than true mortality rates due to both lung cancer and, to a larger degree, mesothelioma.  
16 These discrepancies are due mainly to misdiagnoses and imperfect sensitivity of the coding  
17 system. Lung cancer sensitivity<sup>24</sup> ranges from 86% in an asbestos-exposed cohort ([Selikoff and](#)  
18 [Seidman, 1992](#)), to 95% in the general population ([Percy et al., 1981](#)); mesothelioma sensitivity  
19 ranges from 40% for ICD-9 ([Selikoff and Seidman, 1992](#)) to about 80% for ICD-10 ([Camidge et](#)  
20 [al., 2006](#); [Pinheiro et al., 2004](#)). This underestimation of the true mortality rate results in a lower  
21 estimated risk compared with that which would be estimated based on the true rate. EPA  
22 modeled the risk of mesothelioma mortality using an absolute risk model, while the risk of lung  
23 cancer mortality is modeled using a relative risk model. The underestimation of risk is much  
24 more pronounced for the absolute risk model (mesothelioma) than for the relative risk model  
25 (lung cancer). For lung cancer risks, misdiagnosis rates would need to be different with respect  
26 to exposure levels, and this is unlikely among the Libby workers that were included in the lung  
27 cancer analysis because nosologists are blinded to exposure levels when coding lung cancer as a  
28 cause of death. Therefore, EPA considered use of a procedure to adjust risks for mesothelioma  
29 underascertainment (see Section 5.4.5.1.1)—but not for lung cancer.

30 Mesothelioma did not have a distinct ICD code prior to introduction of the 10<sup>th</sup> revision  
31 (ICD-10), which although released in 1992, was not implemented in United States until 1999.  
32 Death certificates from 1940 to 1978 were reviewed by the NIOSH principal investigator  
33 ([Sullivan, 2007](#)) to identify any mention of mesothelioma on the death certificate, as is the  
34 standard procedure for assessing mesothelioma mortality and has been used in other analyses of  
35 Libby worker cohort mesothelioma mortality ([Larson et al., 2010b](#); [McDonald et al., 2004](#)). For  
36 deaths in the Libby worker cohort occurring from 1979 to 1998, death certificates were obtained

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<sup>24</sup>Sensitivity is measured by the percentage of actual lung cancer deaths that are detected.

1 if the NDI identified the cause of death as one of the possible mesothelioma codes identified by  
2 [Marsh et al. \(2001\)](#), as respiratory cancer, nonmalignant respiratory disease, digestive cancer, or  
3 unspecified cancer. For deaths in the Libby worker cohort that occurred after 1998, the ICD-10  
4 code for mesothelioma was used. In total, 18 mesothelioma deaths were identified by NIOSH  
5 using the methods of [Sullivan \(2007\)](#), which serve as the basis for this current assessment;  
6 19 mesothelioma deaths were identified by [Larson et al. \(2010b\)](#) for the same cohort from all  
7 death certificates rather than from death certificates with one of the specifically targeted set of  
8 causes of death identified above in [Sullivan \(2007\)](#).

9 [Whitehouse et al. \(2008\)](#) identified four mesothelioma cases among workers that, as the  
10 authors suggested, were not included in the [Sullivan \(2007\)](#) study with mortality follow-up  
11 through 2001; no other information was provided. Three mesothelioma cases from these four  
12 were most likely accounted for during the update of the NIOSH cohort from 2001 to 2006, which  
13 serves as the basis for this current assessment. [Whitehouse et al. \(2008\)](#) also provided detailed  
14 information on 11 residential cases, but this information could not be used in exposure-response  
15 analyses for this current assessment because there is no quantitative exposure information for  
16 these cases and no information defining or enumerating the population from which these cases  
17 arose.

18 Lung cancer mortality was based on the underlying cause of death identified by the ICD  
19 code on death certificates according to the ICD version in use at the time of death. Based on  
20 these different ICD codes, lung cancer mortality included malignant neoplasms of the trachea,  
21 bronchus, and lung, and was identified by the following codes: ICD-5 code “047” (excluding  
22 “47c, Cancer of unspecified respiratory organs”), ICD-6 codes “162” or “163,” ICD-7 codes  
23 “162” or “163” (excluding “162.2, Cancer of the pleura”), ICD-8 and ICD-9 code “162,” and  
24 ICD-10 codes “C33” or “C34.” In all, there were 111 deaths with an underlying cause attributed  
25 to lung cancer. All deaths after 1960 were coded as bronchus or lung because the ICD versions  
26 in use at that time distinguished malignant neoplasms of the trachea as distinct from neoplasms  
27 of the bronchus and lung. Other investigators of this cohort have used slightly different  
28 definitions of lung cancer or used different follow-up periods, as described in Section 4.1.1.1  
29 (Studies of Libby, MT Vermiculite Mining and Milling Operations Workers).

### 31 **5.4.2.3. Description of Libby Worker Cohort Work Histories**

32 NIOSH staff abstracted demographic data and work history data from company personnel  
33 and payroll records. An individual’s work history was determined from job change slips, which  
34 recorded any new job assignment, date of change, and change in hourly pay rate (which differed  
35 by the job assignment). Work history records span the time period from September 1935 to May  
36 1982. Dates of termination were unknown for 58 of 640 workers (9%) who left employment  
37 before September 1953. EPA adopted the assumption used by NIOSH ([Sullivan, 2007](#)) that  
38 these people worked for 384 days, based on the mean duration of employment among all workers

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1 with known termination dates before September 1953. The majority of workers in this cohort as  
2 a whole and those hired on or after January 1, 1960 worked at multiple jobs; many of the  
3 workers switched jobs repeatedly, and the changes in exposures associated with changes in job is  
4 accounted for through the use of the job- and time-specific JEM described in the following  
5 section.

#### 6 7 **5.4.2.4. Description of Libby Amphibole Asbestos Exposures**

8 The operations at the mine and in and around Libby, the conditions of exposure, and the  
9 job-specific estimates of exposure intensity have been thoroughly described in Section 4.1  
10 ([Sullivan, 2007](#); [Amandus et al., 1987b](#); [McDonald et al., 1986a](#)). Briefly, miners extracted  
11 vermiculite ore from an open-pit mine that operated on Zonolite Mountain outside the town of  
12 Libby, MT. The ore was processed locally in a dry mill (1935–1974) and/or two wet mills  
13 (1950–1974 and 1974–1990). The resulting concentrate was transported by railroad to  
14 processing plants around the United States where the vermiculite was expanded for use in  
15 loose-fill attic insulation, gardening, and other products (see Section 2.1). EPA adopted the JEM  
16 developed and used by [Sullivan \(2007\)](#), which was in turn based on that used in the earlier  
17 NIOSH study for jobs through 1982 ([Amandus et al., 1987b](#); [Amandus and Wheeler, 1987](#)). As  
18 discussed in more detail in Section 4.1, [Amandus et al. \(1987b\)](#) defined 25 location operations in  
19 the Libby facilities to which they assigned exposure intensity based on available information (see  
20 Table 5-21). A job category may have involved more than one location operation, and the  
21 8-hour time-weighted average (TWA) exposure (8-hour TWA) for each job category in the JEM  
22 was calculated from the exposure intensity and time spent at each location operation ([Amandus](#)  
23 [et al., 1987b](#)).

**Table 5-21. Exposure intensity (fiber/cc) for each location operation from the beginning of operations through 1982** [Amandus et al. \(1987b\); Table VII](#)

Location operation	Yr									
	<1950	1950–59	1960–63	1964–67	1968–70	1971	1972–74	1975–76	1977–79	1980–82
Downtown office building	0	0	0	0	0	0	0	0	0	0
Bus ride	1.2	1.2	1.2	1.2	1.2	1.2	1.2	0	0	0
Mine office	1.0	1.0	1.0	1.0	1.0	1.0	1.0	0.5	0.5	0.5
Mine misc.	1.6	1.6	1.6	1.6	1.6	1.6	1.6	0.8	0.8	0.8
Mine—nondrilling	2.6	2.6	2.6	2.6	2.6	2.6	2.6	0.6	0.6	0.6
Transfer point	2.2	2.2	2.2	2.2	2.2	2.2	2.2	0.6	0.6	0.6
Quality control lab	13.1	13.1	13.1	2.6	2.6	2.6	2.6	0.6	0.6	0.6
Service area by mill	1.9	1.9	1.9	3.8	1.9	1.9	1.9	0.2	0.2	0.2
Dry mill	168.4	168.4	168.4	33.2	33.2	33.2	16.6	--	--	--
Dry mill sweeping	182.1	182.1	182.1	35.9	35.9	35.9	19	--	--	--
Old and new wet mill—millwright	--	7.0	7.0	7.0	7.0	7.0	7.0	0.6	0.6	0.6
Old wet mill—nonmillwright	--	3.7	3.7	3.7	3.7	3.7	3.7	--	--	--
New wet mill—nonmillwright	--	--	--	--	--	--	3.2	2.0	0.8	0.8
Skip area	88.3	88.3	88.3	17.4	17.4	17.4	4.8	0.6	0.6	0.6
Concentrate hauling	5.5	5.5	5.5	5.5	5.5	5.5	5.5	0.4	0.4	0.4
River station binside	21.2	21.2	21.2	21.2	21.2	21.2	21.2	0.7	0.7	0.7
River conveyor tunnel	112.5	112.5	112.5	112.5	112.5	112.5	112.5	0.3	0.3	0.3
River office binside	10.6	10.6	10.6	10.6	10.6	10.6	10.6	0.2	0.2	0.2
Verxite plant	22.6	22.6	2.8	2.8	2.8	--	--	--	--	--

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**Table 5-21. Exposure intensity (fiber/cc) for each location operation from the beginning of operations through 1982 [Amandus et al. \(1987b\)](#); Table VII (continued)**

Location operation		Yr									
		<1950	1950–59	1960–63	1964–67	1968–70	1971	1972–74	1975–76	1977–79	1980–82
Tails belt		7.3	7.3	7.3	7.3	7.3	7.3	7.3	0.7	0.7	0.7
Screen plant		--	--	--	--	--	--	--	0.5	0.5	0.5
Drilling	High	23	23	23	23	9.2	9.2	9.2	0.6	0.6	0.6
	Low	6.7	6.7	6.7	6.7	6.7	9.2	9.2	0.6	0.6	0.6
Ore loading	High	82.5	27.7	10.7	10.7	3.2	3.2	3.2	0.2	0.2	0.2
	Low	24	15	9	9	3.2	3.2	3.2	0.2	0.2	0.2
River dock	High	116.9	42.5	17	17	17	5.1	5.1	0.5	0.5	0.5
	Low	38	19	6.4	6.4	5.1	5.1	5.1	0.5	0.5	0.5
Bagging plant	High	12.9	12.9	12.9	12.9	12.9	12.9	4.3	1.2	1.2	1.2
	Low	4.6	4.6	4.6	4.6	4.6	4.6	4.3	1.2	1.2	1.2

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1 For the later data in Table 5-21 from 1967 through 1982, over 4,000 air samples analyzed  
2 for fibers by PCM analysis were available to inform the exposure intensity for the 25 location  
3 operations. Therefore, the JEM for 1967–1982 is based on direct analytic measurements in air  
4 for each location operation ([Amandus et al., 1987b](#)). With the exception of two location  
5 operations in the dry mill, no air samples were available for other location operations at the mine  
6 and processing facilities before 1967. In order to estimate exposures that occurred before that  
7 time, the NIOSH researchers interviewed plant employees and based estimates of exposure  
8 intensities on known changes in operations over the years and professional judgments regarding  
9 the relative intensity of exposure. Exposure intensity for 23 of the 25 pre-1967 location  
10 operations was extrapolated from post-1967 measurements based on reasoned assumptions for  
11 each location operation ([Amandus et al., 1987b](#)).

12 In contrast to the exposure information available for 1967 through 1982, the amount and  
13 quality of measurement data in the facility in earlier years were much more limited ([Amandus et](#)  
14 [al., 1987b](#)). A total of 40 dust samples were taken, exclusively in the dry mill, over the years  
15 1950–1964. Using these measurements, higher exposures were inferred to occur before 1964  
16 than in later years.

17 Air samples collected by the State of Montana were available for the dry mill  
18 from 1956–1969, but these were analyzed for total dust, not asbestos fibers. Total dust samples  
19 (collected by a midget impinger) were examined by light microscopy, but no distinction was  
20 made among mineral dusts, debris, and asbestos fibers. All objects were counted and reported in  
21 the units of million particles per cubic foot (mppcf). [Amandus et al. \(1987b\)](#) developed a range  
22 of conversion ratios between total dust and asbestos fiber counts based on the comparison of  
23 contemporaneous air sampling in the dry mill (see Section 4.1.1.2) and selected a conversion  
24 ratio of 4.0 fibers/cc per mppcf to estimate exposure intensity for two location operations in the  
25 dry mill for the years prior to 1967. Uncertainties in the selection of this conversion ratio are  
26 described in detail in Section 5.4.6.1.2.1.

27 The exposure intensity (fiber/cc) for each of the location operations (see Table 5-21) was  
28 used to calculate an estimate of daily occupational exposure for each job category in the JEM.  
29 For each job, the time spent at each location operation and the exposure intensity for each  
30 location operation was averaged to derive an estimate of the 8-hour TWA. The resulting JEM  
31 available for this current assessment and previous epidemiologic studies of the Libby worker  
32 cohort is based on the air concentration of fibers as enumerated by PCM, which measures fibers  
33 longer than 5 µm with an aspect ratio >3:1 (i.e., the fiber size regulated under the Occupational  
34 Safety & Health Administration [OSHA] standard ([OSHA, 2006](#))). Additionally, only fibers that  
35 are wide enough to be viewed on PCM can be detected with this method. [Amandus et al.](#)  
36 [\(1987b\)](#) considered fibers >0.44 µm in diameter to be visible by PCM in the historical filter  
37 analysis. More recent techniques have refined the PCM method, and fibers greater than 0.25 µm

in diameter are now considered PCM fibers ([IPCS, 1986](#)). Uncertainties related to difference in defining PCM fibers are discussed in Section 5.4.6.1.2.1.

[Amandus et al. \(1987b\)](#) recognized the uncertainty in the pre-1968 exposures assigned to the cohort. Although there is some uncertainty in the dust-to-fiber conversion, this conversion (4.0 fibers/cc per mppcf) was based on dust and fiber data contemporaneously collected in the dry mill and only applied to the dry mill environment. [Amandus et al. \(1987b\)](#) considered a range of possible conversion factors (1.2–11.5 fibers/cc per mppcf). Greater uncertainty may lie with the reasoned assumptions used to extrapolate exposures to the early decades for all location operations considered. For example, there were four location operations for which [Amandus et al. \(1987b\)](#) estimated a range of possible exposure intensities—drilling, ore loading, the river dock, and the bagging plant, where intensity of exposure may vary as much as threefold between the low and high estimates (see Table 5-21). Finally, some workers were employed after 1982 and up until 1993, when demolition of the facilities was completed ([Larson et al., 2010b](#)). These exposures were not evaluated by [Sullivan \(2007\)](#) and were not included in the NIOSH JEM. Because exposure concentrations in 1982 (see Table 5-21) were generally at or below 1.2 fibers/cc, it is unlikely that the overall cumulative exposures of this limited set of workers were significantly underestimated by not including exposures during this time. Uncertainties in all aspects of the JEM and the associated exposure assessment are described in Section 5.4.6.1.2.

There was one important limitation of the NIOSH work history data in assigning exposure levels for each job. In the earlier study ([Amandus and Wheeler, 1987](#)), workers with “common laborer” job assignments and some workers with unknown job assignments hired between 1935 and 1959 were all assigned the same, relatively low exposure levels estimated for the mill yard ([Sullivan, 2007](#)). In addition, reabstracting work histories for the more recent study ([Sullivan, 2007](#)) identified several job assignments not mentioned in the earlier publications. [Sullivan \(2007\)](#) estimated exposure for the additional job and calendar time period-specific combinations based on professional experience and review of exposure records from earlier studies of the Libby worker cohort ([Amandus et al., 1987b](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). EPA found that of the 991 workers hired before 1960, 811 workers (82%) had at least one job with an unknown job assignment, with 706 (71%) listing neither job department nor job assignment. In the more recent study by [Sullivan \(2007\)](#), these same workers were all assigned the same, relatively high TWA exposure intensity estimated for all jobs during that time period (66.5 fibers/cc). The lack of information on specific job assignments for such a large portion of these early workers, during the time period when exposures were higher, resulted in significant exposure misclassification and effectively yielded exposure estimates that were differentiated only by the duration of each worker’s employment. Because of the lack of more specific measured fiber exposure data during this early period, EPA experienced difficulties in identifying an adequate exposure-response model fit for the complete cohort including all hires. These difficulties are described in detail in Section 5.4.3.4. As a result, the IUR analyses were

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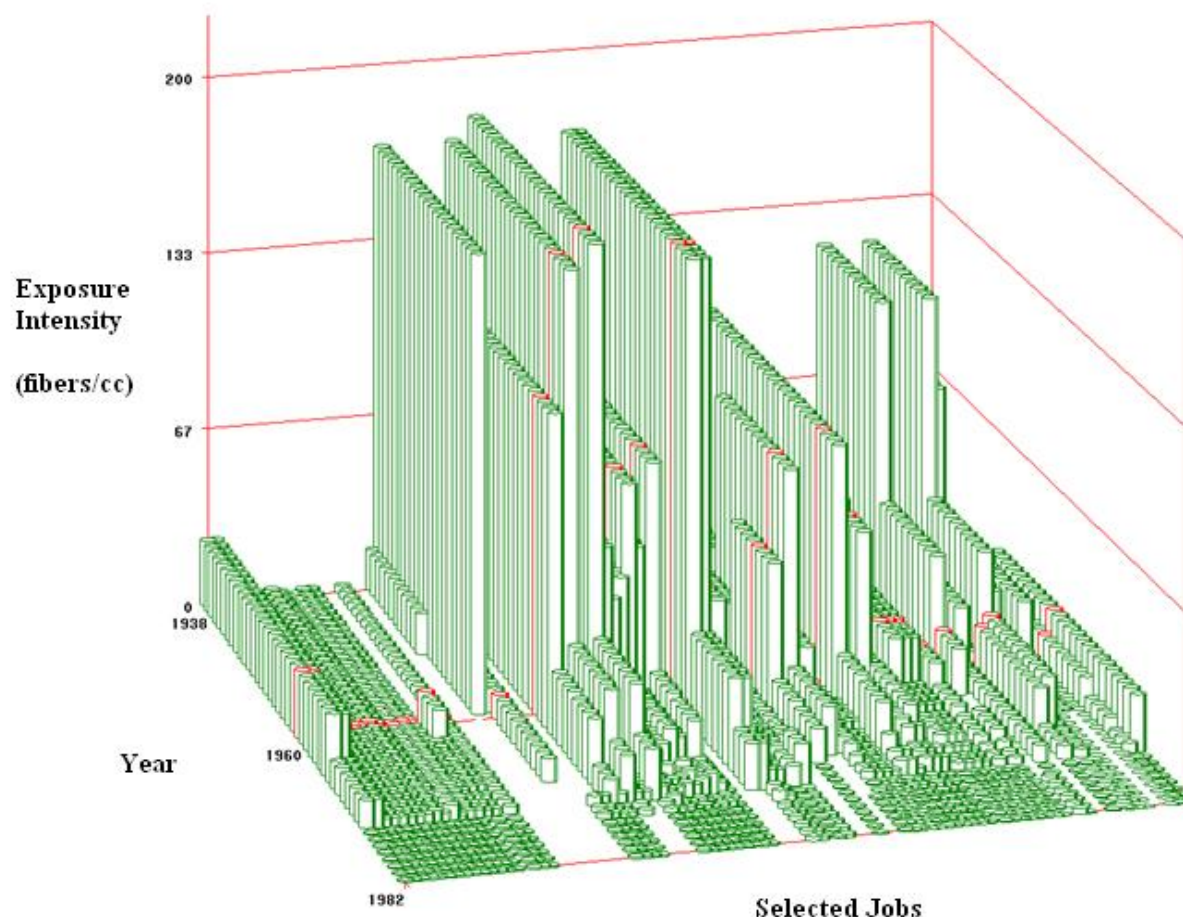
based on the subset of workers hired after 1959 (i.e., on or after January 1, 1960), totaling 880 workers (i.e., full cohort [ $n = 1,871$ ] minus those hired before 1960 [ $n = 991$ ]). Of these 880 workers hired after 1959, 28 workers had at least one job with an unknown job assignment with nine having all job and department assignments between 1960–1963 listed as unknown. As described in [Sullivan \(2007\)](#), NIOSH assigned these workers a TWA estimated exposure intensity of 66.5 fibers/cc. Uncertainties in the exposure assessment for this subcohort are described in Section 5.4.6.1.2.4. While the [Sullivan \(2007\)](#) study was limited to the white male workers, EPA’s analysis includes all workers regardless of race or gender. Table 5-22 shows the demographic, mortality, and exposure characteristics of the subcohort hired after 1959.

**Table 5-22. Demographic, mortality, and exposure characteristics of the subset of the Libby worker subcohort hired after 1959**

Characteristic	Subcohort hired after 1959
Number of workers	880
Number of deaths from all causes	230
Number of deaths from mesothelioma	7
Number of deaths from lung cancer	32
Number of deaths from laryngeal cancer	2
Number of deaths from ovarian cancer	0
Number of deaths from colorectal cancer	5
Number of deaths from chronic obstructive pulmonary disease	18
Mean yr of birth	1942
Mean yr of hire	1971
Mean age at hire (yr)	28.6
Mean person-yr of follow-up (no lag)	32.2
Total person-yr of follow-up (no lag)	28,354
Mean employment duration (yr)	2.4
Mean cumulative exposure (fiber/cc-yr)	19.2
Median cumulative exposure (fiber/cc-yr)	3.4
Range of cumulative exposures (no lag) (fiber/cc-yr) <sup>a</sup>	0–462

<sup>a</sup>According to the work histories and JEM, there were 21 subcohort workers who had zero cumulative exposure. These 21 individuals all worked at the office downtown.

Figure 5-5 shows a three-dimensional representation of the JEM used by [Sullivan \(2007\)](#) and in this cancer exposure-response assessment (note that the figure does not include all jobs and is meant to be illustrative rather than comprehensive). The three axes show the intensity of fiber exposure as an 8-hour TWA (fiber/cc, vertical axis) for selected job categories over time (horizontal axes). For several jobs, the estimated 8-hour TWA was greater than 100 fibers/cc for the decades prior to 1963.



**Figure 5-5. Plot of the National Institute for Occupational Safety and Health (NIOSH) job-exposure matrix for different job categories over time.** The height of each bar represents the intensity of exposure as an 8-hour TWA (fiber/cc) for a job in a particular year. Each row for “Selected Jobs” represents a specific job category. The line at 1960 marks the beginning of jobs included in the subcohort of Libby workers used to derive the inhalation unit risk.

#### 5.4.2.5. *Estimated Exposures Based on Job-Exposure Matrix (JEM) and Work Histories*

Exposure-response modeling of epidemiologic data is based on several considerations as summarized by [Finkelstein \(1985\)](#):

After identification of an occupational hazard one of the goals of occupational epidemiology is to quantify the risks by determining the dose-response relations for the toxic agent. In many circumstances little is known about the dose received by target tissues; the data available usually pertain only to exposure to various concentrations of the toxic material in the workplace. The calculation of dose requires additional physiological and chemical information relating to absorption, distribution, biochemical reactions, retention, and clearance.

In asbestos epidemiology the usual measure of exposure is the product of the concentration of asbestos dust in the air (fibers or particles per mL) and the duration of exposure to each concentration summed over the entire duration of exposure (years).

Cumulative exposure (CE) has been the traditional method of measuring exposure in epidemiologic analyses of many different occupational and environmental exposures and was the exposure metric applied to the risk of lung cancer mortality in the IRIS assessment for general asbestos ([U.S. EPA, 1988a](#)). That said, different health outcomes may be best described using different exposure metrics. The risk of mesothelioma mortality in the IRIS assessment for general asbestos ([U.S. EPA, 1988a](#)) used a different exposure metric based on a linear function of concentration added to a function of TSFE and duration of exposure. Additional exposure metrics were also assessed for both mesothelioma and lung cancer mortality risks.

Two alternative approaches to developing exposure metrics to describe the effects of concentrations of asbestos dust in the air on the risks of mortality have also been proposed. The first alternative was proposed by [Jahr \(1974\)](#), who studied silica-induced pneumoconiosis and suggested that exposures to occupational dusts could be weighted by the time since exposure. This yields an exposure metric that gives greater weight to earlier exposures. The second alternative was proposed by [Berry et al. \(1979\)](#) who subsequently suggested the application of exposure metrics that allowed for the clearance of dust or fibers by using a decay term on exposures.

For the evaluation of mortality risk from mesothelioma for general asbestos, [U.S. EPA \(1988a\)](#) used a different exposure metric than was used for lung cancer mortality, which factored in the TSFE. As observed in [U.S. EPA \(1988a\)](#), it is important to note that different characterizations of estimated occupational exposures may be reasonably expected to be associated with different endpoints.

Many studies have been limited in the availability of detailed exposure data—especially at the individual level. In the Libby worker cohort, detailed work histories were matched with job-specific exposure estimates, allowing for the reconstruction of each individual's estimated

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occupational exposure over time. From this information-rich, individual-level data set from NIOSH, EPA constructed a suite of the different metrics of occupational exposure which had been proposed in the asbestos literature or used in the IRIS asbestos assessment ([U.S. EPA, 1988a](#)). This suite of metrics was defined a priori to encompass a reasonable set of proposed exposure metrics to allow sufficient flexibility in model fit to these data. The types of exposure metrics evaluated were intended to allow for more or less weight to be placed on earlier or later exposures. These simulated exposure metrics were derived mathematically to approximate underlying processes that are not well understood (see Section 5.4.6). Thus, the empirical fit of various exposure metrics to the observed epidemiologic data is evaluated statistically, and the exposure metrics have epidemiological interpretation but do not necessarily have direct biological interpretations.

The first exposure metric—CE—is a simple addition of each day’s exposure across time (see eq 5-7). CE has been widely used in modeling risk of cancer in occupational epidemiology and has been used for modeling lung cancer ([Larson et al., 2010b](#); [Moolgavkar et al., 2010](#); [Berman and Crump, 2008](#); [Sullivan, 2007](#); [McDonald et al., 2004](#)) and mesothelioma ([McDonald et al., 2004](#)) in the Libby worker cohort. When using this exposure metric in the risk model, all exposures (other than for years removed from consideration based on a lag assumption) have equal weight regardless of when they occurred and lead to the same estimated cancer risk whether exposure happened early or later in life.

EPA calculated each individual’s occupational CE to LAA over time from their date of hire until the date they ceased to be employed in the Libby operations or until the date NIOSH collected the work history data for those still employed in May 1982. Workers were assumed to remain at their CE on the last day of work until death or the end of the follow-up period on December 31, 2006. Each worker’s CE at any time point (daily increment) since their date of hire was computed as the sum of their exposure intensity (fiber/cc) on each specific occupational day ( $x_t$ ) from Day 1 through Day  $k$ . Mathematically, this was defined as

$$\text{CE at time } t_k = \sum_{j=1}^k x_{t_j} \quad (5-7)$$

Where

$x_{t_j}$  = the estimated job-specific exposure intensity for the day  $t_j$ , and

$t_k$  = the day on which the exposure is estimated.

A second exposure metric—RTW exposure—gives additional weight to early exposures. By doing so, the RTW exposure metric allows the possibility that early exposures are more influential on cancer mortality predictions in the model. Unlike many chemicals that are rapidly

1 metabolized in the body and excreted, asbestos fibers are durable, and some remain in the body  
2 for years. Fibers that remain in the lung may continue to damage lung cells and tissue unless  
3 they are removed or cleared (see Section 3.2). Similarly, fibers that translocate to the pleura may  
4 damage cells as long as they remain in this tissue. Therefore, a fiber exposure may not only  
5 damage tissue during the initial exposure, but fibers may remain in these tissues, with tissue fiber  
6 concentration as well as cellular and tissue damage accumulating over time. While this  
7 represents a biological point of view, in an epidemiologic context in which the exposure is  
8 ambient fiber concentration and the event of interest is simply a cause of death (rather than  
9 survival time), it is uncertain what metrics of exposure might fit the observed data and could be  
10 considered most appropriate.

11 The RTW exposure metric in this current assessment is sometimes called the cumulative  
12 burden, or the area under the curve. A type of RTW metric was proposed for modeling of  
13 mesothelioma mortality by [Newhouse and Berry \(1976\)](#) based on a general understanding of the  
14 relationship between tumor incidence rate and time to cancer ([Cook et al., 1969](#)) as well as  
15 animal models of mesothelioma ([Berry and Wagner, 1969](#)).

16 A similar type of RTW metric was proposed in [Peto \(1978\)](#) and was subsequently applied  
17 by [Peto et al. \(1982\)](#), discussed by [Finkelstein \(1985\)](#), and applied in the derivation of the IUR in  
18 the 1988 IRIS assessment for asbestos ([U.S. EPA, 1988a](#)). [McDonald et al. \(2004\)](#) and  
19 [Moolgavkar et al. \(2010\)](#) used RTW-type metrics for modeling mesothelioma in the Libby  
20 worker cohort, and [McDonald et al. \(2004\)](#) applied an RTW metric for modeling lung cancer  
21 mortality in the Libby worker cohort.

22 In calculating RTW, each day's exposure is multiplied by the time since exposure  
23 occurred up to the time  $t_k$  when RTW is estimated (see eq 5-8). The intent of RTW CE is to  
24 allow for earlier exposures to contribute greater weight.

$$\text{RTW CE at time } t_k = \sum_{j=1}^k x_{t_j} \times (t_k - t_j) \quad (5-8)$$

28 Where

30  $x_{t_j}$  = the estimated job-specific exposure intensity for the day  $t_j$ , and

31  $t_k$  = the day  $k$  on which the exposure is estimated.

33 The CE and RTW exposure metrics result in increasing or sustained metrics of exposure  
34 across time. However, some cellular and genetic damage can be repaired over time after  
35 exposure, decreasing the cancer risk from exposure over time. Additionally, asbestos fibers are  
36 cleared (removed) from the lung through natural processes and translocated to other tissues (see



Section 3.2.1.1). Therefore, when considering lung cancer, it is possible that removal of asbestos fibers from the lung could reduce lung cancer risk over time. Although less is known about removal of asbestos from the pleura, clearance mechanisms may be operative in that tissue as well (see Section 3.2.1.2). As noted earlier, [Berry et al. \(1979\)](#) proposed the use of exposure metrics which addressed the issue of clearance through a mathematical decay term that modified estimated occupational exposures. For mesothelioma, modeling a decay term on exposure has been proposed by [Berry \(1999\)](#). Based on this proposal, several studies applied a decay term to modeling mesothelioma mortality ([Berry et al., 2009](#); [Reid et al., 2009](#); [Barone-Adesi et al., 2008](#); [Gasparrini et al., 2008](#); [Clements et al., 2007](#); [Hodgson et al., 2005](#); [Berry et al., 2004](#)). Similarly, publications indicate that the relative risk of lung cancer due to asbestos exposure declines 15–20 years after the cessation of exposure to asbestos ([Magnani et al., 2008](#); [Hauptmann et al., 2002](#)).

Mathematically allowing for the magnitude of earlier exposures to diminish with advancing time was considered to be a method of giving less weight in the analyses to earlier exposures compared to the previous two exposure metrics. Therefore, two additional exposure metrics were considered, in which a decay rate was applied to the CE and RTW exposure metrics (see eq 5-9 and 5-10).

For each exposure metric, the application of a half-life was calculated by depreciating each time interval's ( $t_{j-1}$ ;  $t_j$ ) exposure according to a model of exponential decay with various half-lives ( $T_{1/2}$ ) of 5, 10, 15, and 20 years. Note that the particular kinetics of LAA fibers are not fully understood, and the relevance of these particular half-lives was determined from the statistical fit of these exposure metrics to the risk of cancer mortality, rather than the biological half-life of the fibers. For a very large half-life, decay is very slow, and these metrics would be very similar to the CE and RTW exposure metrics.

$$\text{CE with half-life at time } t_k = \sum_{j=1}^k \left\{ x_{t_j} \times \exp \left[ \frac{\ln(0.5) \times (t_k - t_j)}{T_{1/2}} \right] \right\} \quad (5-9)$$

Where

$x_{t_j}$  = the estimated job-specific exposure intensity for the day  $t_j$ ,

$t_k$  = the day  $k$  on which the exposure is estimated, and

$T_{1/2}$  = half-life of 5, 10, 15 or 20 years.

$$\text{RTW with half-life at time } t_k = \sum_{j=1}^k \left\{ x_{t_j} \times (t_k - t_j) \times \exp \left[ \frac{\ln(0.5) \times (t_k - t_j)}{T_{1/2}} \right] \right\} \quad (5-10)$$

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1 In addition to the considerations described above for selecting metrics to represent  
2 estimated ambient exposure to LAA for use in predicting the risk of mortality, there is the  
3 important issue of potentially modifying the exposure metrics to account for cancer latency.  
4 Without knowledge of the specific timing of etiologically relevant exposure that may initiate and  
5 promote cancers that ultimately result in mortality, any exposure metric may include exposures  
6 during some time period that do not have bearing on the risk of mortality. In the absence of such  
7 information on the specific cancer latency associated with a specific exposure, [Rothman \(1981\)](#)  
8 suggested that the most relevant exposure period could be identified by comparing the fit of  
9 exposure metrics across multiple lag periods to allow for the identification of the optimal latency  
10 period as an expression of a lag time between exposure and mortality. This has since become a  
11 standard practice in occupational and environmental epidemiology. Accordingly, exposure  
12 estimates for all exposure metrics were adjusted to account for the time period between the onset  
13 of cancer and mortality. The lag period defines an interval before death, or end of follow-up,  
14 during which any exposure is excluded from the calculation of the exposure metric. Cohort  
15 members who died or were lost within the initial years of follow-up were assigned lagged  
16 exposure values of zero if they had not been followed for longer than the lag time. The various  
17 exposure metrics were lagged at 10, 15, and 20 years to account for different potential cancer  
18 latencies within the limitations of the available data. Metrics without a lag were fit for  
19 comparison purposes but were not considered to be biologically reasonable, given that the  
20 outcome under analysis is cancer mortality (specifically, mesothelioma and lung cancer), for  
21 which latency periods of 10 years or more have been suggested for asbestos ([U.S. EPA, 1988a](#)).  
22 Consequently, metrics that were not adjusted by lagging exposure in the final years before  
23 mortality (or the end of follow-up) were not considered further in the development of an IUR for  
24 LAA.

25 In addition to the exposure metrics used in the lung cancer mortality analysis, modeling  
26 of mesothelioma mortality (see Section 5.4.3.1) included additional exposure models. The Peto  
27 model ([Peto et al., 1982](#); [Peto, 1979](#)) uses a cubic power function of TSFE and a linear function  
28 of exposure concentration. The model developed by Peto was then adapted in the IRIS ([U.S.](#)  
29 [EPA, 1988a](#)) asbestos assessment. The linear function of concentration was developed based on  
30 estimated average workplace concentrations over several asbestos cohorts exposed to chrysotile,  
31 amphiboles, and mixed fibers. The two amphibole-exposed cohorts were U.S workers exposed  
32 to amosite and Australian workers exposed to crocidolite; for both cohorts there was little  
33 exposure information at the time these reports were published. As Health Effects  
34 Institute-Asbestos Research ([HEI, 1991](#)) noted “No extensive measurements of historical  
35 exposure levels are available for the cohorts exposed predominantly to crocidolite or amosite.”  
36 The only other mesothelioma models proposed in the literature for amphibole asbestos were  
37 developed by [Berry \(1991\)](#) for crocidolite. [Berry et al. \(2012\)](#) found that models that allow for  
38 clearance (mathematically, multiplicative  $\exp[-\lambda \times T]$ ) better match the actual mortality

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experience of the Wittenoom, Australia crocidolite cohort with more than 50 years follow-up, compared with Peto-type models. In particular, they found that models with a higher power of TSFE of 5.4 (compared to power = 3 for Peto) and a decay rate of 15%/year (half-life of approximately 5 years) fits the observed data best, followed by models with a power of TSFE of 3.9 and a decay rate of 6.8%/year (half-life of approximately 10 years). However, the exposure data calculated for the Berry analysis was based on just one study of airborne levels. Nevertheless, cumulative exposures calculated from these have been shown to be internally valid based on association with fiber lung burden ([Berry et al., 2012](#)).

Peto's model (also used in the 1988 IRIS assessment for asbestos ([U.S. EPA, 1988a](#)) is

$$I_m = C \times Q_k \times KM \quad (5-11)$$

where

$I_m$  = the observed deaths from mesothelioma/person-years (i.e., the mesothelioma mortality rate),

$C$  = the average concentration of asbestos in the air,

$KM$  = an estimated slope describing the relationship between LAA exposure and mesothelioma mortality, and

$Q_k$  = the function of the TSFE ( $t$ ) and the duration of exposure ( $d$ ):

$$\text{For } t \leq 10, Q_k = 0$$

$$\text{For } 10 < t \leq d + 10, Q_k = (t - 10)^k$$

$$\text{For } t > d + 10, Q_k = (t - 10)^k - (t - 10 - d)^k.$$

Alternatively,  $I_m = C \times Q_k \times KM \times \exp(-\lambda \times t)$  defines the Peto model with clearance. Possible values of  $\lambda$  and  $k$  suggested in the literature ([Berry et al., 2012](#)) are  $\lambda = 0.068$  or  $0.15$  and  $k = 3.9$  or  $5.4$ . As the Peto model and the Peto model with clearance were both proposed in the amphibole asbestos literature, these models were carried forward in the analysis below.

### 5.4.3. Exposure-Response Modeling

As discussed above, consideration of biology and previous epidemiologic studies informed the range of models considered. There is not sufficient information to select models for the epidemiology data on the basis of the biological mechanism of action for lung cancer or mesothelioma (see Section 3). In this situation, EPA's practice is to investigate several modeling options to determine how to best empirically model the exposure-response relationship in the range of the observed data as well as consider exposure-response models suggested in the

1 epidemiologic literature. For LAA, possible exposure metrics were explored for model fit to the  
2 chosen modelforms. The exposure metric options were selected to provide a range of shapes that  
3 was sufficiently flexible to allow for a variety of ways that time and duration might relate to  
4 cancer risk in the data being modeled.

5 The following sections provide information about modeling of the full cohort first, the  
6 difficulties in identifying adequately fitting models to these data, and the decision to base the  
7 analysis on a subcohort of workers that allowed for identifying adequately fitting models.

#### 9 **5.4.3.1. Modeling of Mesothelioma Exposure Response in the Libby Worker Cohort**

10 The background incidence of mesothelioma is extremely low ([Hillerdal, 1983](#)). The  
11 evaluated exposure-response models examine the relationship of the absolute risk of  
12 mesothelioma mortality attributable to LAA exposure, because it is not clear that a background  
13 risk of mesothelioma mortality exists among people who were truly unexposed to asbestos (as  
14 opposed to the relative risk model, which is used for lung cancer mortality; see Section 5.4.3.3).  
15 EPA does not have a specific technical guidance for model selection based on human cancer  
16 data, but as a general consideration, EPA's BMD Technical Guidance ([U.S. EPA, 2012](#)) states  
17 that "The initial selection of a group of models to fit to the data is governed by the nature of the  
18 measurement that represents the endpoint of interest and the experimental design used to  
19 generate the data." Here, the most prominent feature of the data is the rarity of mesothelioma  
20 deaths. Correspondingly, Poisson models are employed to estimate the absolute risk of  
21 mesothelioma, as the Poisson distribution is an appropriate model for use with data that are  
22 counts of a relatively rare outcome, such as observed mesothelioma deaths in the Libby worker  
23 cohort. Other parametric survival models, such as the Weibull model have been used for  
24 absolute risk calculation, but they are not generally used for data with rare outcomes.  
25 Consequently, there are no examples in the literature of the Weibull or other parametric survival  
26 model ever being used for modeling mesothelioma mortality. Previous analyses of  
27 mesothelioma mortality in the Libby worker cohort also used the Poisson model ([Moolgavkar et](#)  
28 [al., 2010](#); [McDonald et al., 2004](#)). Mathematically, the Poisson distribution specifies the  
29 probability of  $k$  events occurring as

$$P(k) = \frac{\lambda^k e^{-\lambda}}{k!} \quad (5-12)$$

33 where  $\lambda$  is parameterized with the exposure metric (defined in Section 5.4.2.5). Then, life-table  
34 analysis is used to estimate risks in the general U.S. population for the derivation of the unit risk  
35 of mesothelioma mortality (see Section 5.4.5.1). In the standard Poisson distribution, the  
36 assumption is that the mean is equal to the variance. However, actual count data often exhibit

1 overdispersion, a statistical consideration when the variance is larger than the mean; thus, EPA  
2 evaluated potential for overdispersion.

3 Estimation of the exposure-response relationship for mesothelioma mortality was  
4 performed using a Monte Carlo Markov Chain (MCMC) Bayesian approach with an  
5 uninformative or diffuse (almost flat) prior ([WinBUGS Version 1.4 Spiegelhalter et al., 2003](#)).  
6 Use of diffuse priors is a standard procedure in Bayesian analysis, in situations like this one,  
7 when there is no prior knowledge about the toxicity of LAA under a particular model. Because  
8 this analysis focuses only on the Libby worker cohort and does not try to factor in data from  
9 other sources in estimating potency, use of a diffuse prior is considered appropriate for this  
10 analysis.

11 The benefit of using the WinBUGS software is its computational ease and that it provides  
12 a posterior distribution of the mesothelioma coefficient (KM) rather than just a point estimate. A  
13 diffuse (high variance) Gaussian distribution, truncated to exclude negative parameter values, is  
14 used as a diffuse prior. With such a prior, results of MCMC analysis are expected to be similar  
15 to maximum likelihood estimation in a non-Bayesian analysis. Standard practices of MCMC  
16 ([Spiegelhalter et al., 2003](#)) analysis were followed for verifying convergence and sensitivity to  
17 the choice of initial values. The posterior distribution is based on three chains with a burn-in of  
18 10,000 (i.e., the first 10,000 simulations are dropped so that remaining samples are drawn from a  
19 distribution close enough to the true stationary distribution to be usable for estimation and  
20 inference) and thinning rate of 10 (i.e., only each 10<sup>th</sup> simulation is used—thus reducing  
21 autocorrelation), such that 3,000 total simulations constitute the posterior distribution of KM.  
22 The mean of the posterior distribution served as a central estimate, and the 90% credible  
23 interval<sup>25</sup> defined the 5<sup>th</sup> percentile and the 95<sup>th</sup> percentile of the distribution, which served as  
24 bounds for the 95<sup>th</sup> lower and upper one-sided confidence intervals, respectively.

25 The fit of multiple metrics of exposure, the Peto model and the Peto model with clearance  
26 (see Section 5.4.2.5), as well as exposure intensity, duration of exposure, age at death or loss to  
27 follow-up, and TSFE were compared using the Deviance Information Criterion (DIC). The DIC  
28 ([Spiegelhalter et al., 2002](#)) is used in Bayesian analysis and is an analogue of the AIC, with  
29 smaller values indicating a better statistical fit to the data. Use of the DIC and AIC is standard  
30 practice in comparing the fit of nonnested models to the same data set with the same dependent  
31 outcome variable but different independent covariates. According to [Burnham and Anderson](#)  
32 ([2002](#)), “These methods allow the data-based selection of a ‘best’ fitting model and a ranking  
33 and weighting of the remaining models in a predefined set.” Because of the small number of  
34 deaths from mesothelioma in absolute terms, only uni- and bivariate models (with age or TSFE  
35 as the second covariate) were considered. Gender and race were not used as covariates because

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<sup>25</sup>A credible interval is the Bayesian analogue of a confidence interval.

all mesothelioma deaths were observed in men assumed to be white (Sullivan, 2007). Each exposure metric was lagged by 0, 10, 15, or 20 years, where appropriate.

#### **5.4.3.2. Results of the Analysis of Mesothelioma Mortality in the Full Libby Worker Cohort**

While the final analytic exposure-response modeling for mesothelioma is based on the subcohort of workers hired after 1959 when exposure data were considered to be superior (see Section 5.4.3.5), it is important to understand how different metrics of exposure in the epidemiologic literature are related to risk in the full cohort as these age- and time-related variables are well characterized in the full cohort. A parallel set of tables is provided for the subcohort of workers hired after 1959 in Section 5.4.3.5.

Tables 5-23 to 5-25 show rates of mesothelioma mortality in the full cohort by duration of exposure, age of first exposure, and TSFE. Mesothelioma rates look to be independent of the age of first exposure, but duration of exposure and TSFE both show relationships with mesothelioma mortality rate. EPA also evaluated the potential for overdispersion of the counts of mesothelioma deaths. In the Libby worker cohort, mean and variance of exposure are nearly identical at  $9.62 \times 10^{-3}$  and  $9.53 \times 10^{-3}$ , respectively, making overdispersion very unlikely.

**Table 5-23. Mesothelioma mortality rate shown by duration of exposure (yr) in the full Libby worker cohort including all hires ( $n = 1,871$ )**

	Duration				
	0–1 yr	1–2 yr	2–3 yr	3–5 yr	5+ yr
Deaths/PY	3/40,417	1/7,493	3/4,429	2/4,984	9/9,778
Rate $\times 10^{-4}$	0.7	1.3	6.8	4.0	9.2

PY = Person-yr

**Table 5-24. Mesothelioma mortality rate shown by age at first exposure in the full Libby worker cohort including all hires ( $n = 1,871$ )**

	Age		
	15–25 yr old	25–35 yr old	35+ yr old
Deaths/PY	5/30,872	11/22,447	2/13,782
Rate $\times 10^{-4}$	1.6	4.9	1.5

**Table 5-25. Mesothelioma mortality rate shown by time since first exposure (TSFE) in the full Libby worker cohort including all hires ( $n = 1,871$ )**

	Time since first exposure					
	<15 yr	15–25 yr	25–35 yr	35–45 yr	45–55 yr	55–68.1 yr
Deaths/PY	0/27,186	2/16,553	5/12,775	8/6,818	2/3,025	1/744
Rate $\times 10^{-4}$	0	1.2	3.9	11.7	6.6	13.4

For the full Libby worker cohort ( $n = 1,871$ ), in the continuous analysis examining one explanatory variable at a time (see Table 5-26), the duration of exposure provided a considerably better model fit than the other possible exposure metrics, indicating that this exposure metric was the best single predictor of mesothelioma mortality in the full Libby worker cohort. A model, which included duration of exposure and age at death or censoring, provided the overall best fit (DIC = 196). Counterintuitively, the inclusion of information on the concentration of exposure in addition to the duration of exposure (as expressed by CE, which is the product of duration and concentration) resulted in a degradation in model fit compared to the model with just the duration of exposure (see Table 5-26). From the models proposed in the amphibole asbestos literature, the Peto model (see eq 5-11) had a much higher DIC of 233.7 in the analysis of the full cohort. For the Peto model, KM was estimated to be  $1.85 \times 10^{-9}$  and its 95<sup>th</sup> upper bound was  $2.59 \times 10^{-9}$ . The Peto model with power terms on time-since-first-exposure  $k = 3.9$  and  $5.4$  and clearance terms of  $6.8$  and  $15\%$  per year, respectively, did not improve fit over the standard Peto model.



**Table 5-26. Comparison of model fit of various univariate exposure metrics for mesothelioma mortality in the full Libby worker cohort including all hires ( $n = 1,871$ ).<sup>a</sup> Only models with DIC within 10 units of the DIC of the model with the lowest DIC are shown.<sup>b</sup>**

Variable	DIC
Duration of exposure	202.9
Age at death or censoring	209.2
CE lagged 15 yr	209.5
CE lagged 10 yr	209.9
RTW lagged 10 yr with 5-yr half-life	210.4
CE lagged 10 yr with 20-yr half-life	210.6
RTW with 5-yr half-life	210.7
RTW with 10-yr half-life	211.0
CE	211.4
Time since first exposure	211.4

<sup>a</sup>Because one of the mesothelioma deaths occurred less than 20 yr from start of the exposure, lag 20 metrics assigned no exposure to this case, which resulted in the very poor fit of exposure metrics lagged 20 yr.

<sup>b</sup>Lower DIC values represent better fits.

It is likely that the poorer fit seen when using information on exposure concentration is the result of the fact that duration of exposure is measured with comparatively little error, while derivation of specific exposure concentrations may be subject to a sizable measurement error. Moreover, as described in Section 5.4.2.3, for 706 of 991 (71%) workers hired from 1935 to 1959, only the duration of exposure was known, but not the job category or department code. Thus, the same time-weighted average estimated exposure intensity for that time period had been assigned to 653 of these workers<sup>26</sup> (Sullivan, 2007). Particularly large exposure measurement error, among more than two-thirds of the workers hired prior to 1960 who were assigned the same exposure intensity, resulted in the duration of exposure being the best predictor of mesothelioma mortality. Additionally, estimates of exposure intensity prior to 1968 have greater uncertainty associated with them than more recent exposure measurements, which are based on fiber counts in air samples analyzed by PCM. For the majority of job locations (23 of 25), no exposure measurements were available before 1968, and exposures were estimated based on employee interviews (in 1982) and what was known about major changes in operations between 1935 and 1967. For two exposure locations, the dust-to-fiber conversion ratio is based on measurements taken in the late 1960s, so extrapolations from the mid 1960s to the early 1960s is

<sup>26</sup>Note that Sullivan (2007) analyzed the population of 1,672 white male workers rather than all 1,871 workers so the numbers of workers with missing job category and department information were different.



likely to be more certain than extrapolation further back in time. The metric using only duration of exposure fit best and the additional incorporation of exposure intensity information, as expressed as the CE, only worsened the fit. Therefore, it is unlikely that IUR estimates can be developed using the full cohort data because the early exposure values (which were predominantly inferred from later data and based on missing job information) were not predictive of mesothelioma mortality.

#### **5.4.3.3. Modeling and Results of Lung Cancer Exposure Response in the Full Libby Worker Cohort**

As noted in the previous section, while the final analytic exposure-response modeling is based in the subcohort of workers hires after 1959, it is important to understand how different metrics of exposure that appear in the epidemiologic literature are related to risk in the full cohort. A parallel set of tables is provided for the subcohort of workers hired after 1959 in Section 5.4.3.6.

Tables 5-27 to 5-29 show the mortality rates of lung cancer mortality by duration of exposure, age of first exposure, and TSFE for the full Libby worker cohort ( $n = 1,871$ ). Lung cancer rates in the Libby worker cohort are substantially higher than mesothelioma rates (see Section 5.4.3.2). Basic stratified models of lung cancer rates and standardized mortality ratios (SMRs) in this population show increased rate with increases in duration of exposure greater than 5 years, age at first exposure and TSFE.

**Table 5-27. Lung cancer mortality rate shown by duration of exposure (yr) in the full Libby worker cohort including all hires ( $n = 1,871$ )**

	Duration				
	0–1 yr	1–2 yr	2–3 yr	3–5 yr	5+ yr
Deaths/PY	60/40,417	9/7,493	6/4,429	5/4,984	31/9,778
Rate $\times 10^{-4}$	14.8	12.0	13.5	10.0	31.7
White male deaths/white male PY	57/37,761	9/7,030	6/4,168	5/4,767	31/9,610
White male rate $\times 10^{-4}$	15.1	12.8	14.4	10.5	32.3
White male SMR <sub>Montana</sub>	2.3	2.0	2.2	1.6	5.0
White male SMR <sub>U.S.</sub>	2.0	1.7	1.9	1.4	4.2

SMR standardized to white male lung cancer mortality rates obtained from [NCI \(2012\)](#).

**Table 5-28. Lung cancer mortality rate shown by age at first exposure in the full Libby worker cohort including all hires ( $n = 1,871$ )**

	Age		
	15–25 yr old	25–35 yr old	35+ yr old
Deaths/PY	28/30,872	42/22,447	41/13,872
Rate $\times 10^{-4}$	9.1	18.7	29.7

SMR not computed due to lack of comparable rates.

**Table 5-29. Lung cancer mortality rate shown by time since first exposure (TSFE) in the full Libby worker cohort including all hires ( $n = 1,871$ )**

	Time since first exposure (yr)					
	<15	15–25	25–35	35–45	45–55	55–68.1
Deaths/PY	12/27,186	19/16,553	35/12,775	21/6,818	21/3,025	3/744
Rate $\times 10^{-4}$	4.4	11.5	27.4	30.8	69.4	40.3
White male deaths/white male PY	11/25,651	19/15,569	33/12,112	21/6,482	21/2,843	3/680
White male rate $\times 10^{-4}$	4.3	12.2	27.2	32.4	73.9	44.1
White male SMR <sub>Montana</sub>	0.7	1.9	4.2	5.0	11.5	6.9
White male SMR <sub>U.S.</sub>	0.6	1.6	3.5	4.2	9.6	5.7

SMR standardized to white male lung cancer mortality rates obtained from [NCI \(2012\)](#).

1 EPA does not currently have specific technical guidance for model selection based on  
2 human cancer data. However, below we explain the process and criteria used in the assessment.  
3 Standard models from similar exposure-response analyses available in the epidemiologic  
4 literature may be candidate models for exposure-response analyses when they are appropriate to  
5 the epidemiologic data at hand.

6 As noted above for mesothelioma, models are selected and evaluated based on the nature  
7 of the data set, which for lung cancer, warrants the ability to use the time-dependent data. The  
8 mesothelioma mortality data were modeled using the Poisson model within a Bayesian  
9 framework to estimate the absolute risk, because mesothelioma is very rare in the general  
10 population ([Hillerdal, 1983](#)). While the Poisson model is appropriate for modeling very rare  
11 events, the standard form does not allow for inclusion of the time-varying nature of exposure.  
12 Lung cancer is more common than mesothelioma and does have a known background risk.  
13 Thus, modeling of lung cancer mortality is based on the relative risk rather than the absolute risk  
14 and was conducted in a frequentist framework, which is the standard methodology for  
15 epidemiologic analyses. A frequentist framework is an alternative method of inference, drawing  
16 conclusions from sample data with the emphasis on the observed frequencies of the data.

17 Standard epidemiologic models for relative risk include Poisson, logistic, conditional  
18 logistic, and Cox models. Multistage clonal expansions models are also available. However,  
19 only the Cox models and clonal expansion models can accommodate the analysis of  
20 time-varying covariates as in the case of the Libby worker cohort. While different researchers  
21 have used two-stage clonal expansion models to model asbestos-related health endpoints from an  
22 occupational cohort of asbestos textile workers in South Carolina ([Zeka et al., 2011](#); [Richardson,  
23 2009](#)), divergent model results raise questions about the resilience of this method when applied  
24 to epidemiologic cohorts. Specifically, the two-stage clonal expansion analysis by [Richardson  
25 \(2009\)](#) fit the data well and was complementary and consistent with his accompanying Cox  
26 regression analysis, while the two-stage clonal expansion analysis by [Zeka et al. \(2011\)](#) on the  
27 same cohort population, but with a different length of mortality follow-up, did not completely  
28 converge, indicating poor model fit. One issue is that epidemiologic cohorts may be less regular  
29 in nature than toxicological studies in the sense that epidemiologic cohorts can be dynamic, with  
30 people joining at different times, possibly leaving and then rejoining. By comparison, in animal  
31 studies, it is more typical for all the subjects to undergo the identical exposure protocol.  
32 Additionally, the degree to which the results of two-stage clonal expansion models depends upon  
33 multiple additional assumptions is not yet well understood and EPA does not have reliable  
34 information available on which to make the required assumptions for the Libby worker cohort  
35 (e.g., the number of cells at risk, constraints on the spontaneous rates of first and second  
36 mutations, allowing for a fixed lag between malignant transformation of a cell and death from  
37 cancer, etc.). Therefore, in addition to the basic stratified models of lung cancer risk by duration  
38 of exposure, by age at first exposure, and by TSFE, EPA selected the Cox model as the most

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appropriate model for exposure-response modeling based on the suitability of this model to the nature of the data set (i.e., time-dependent exposure information), the long history of usage in analyses of occupational cohorts, and the commonality of usage in other epidemiologic analyses of the Libby workers cohort.

No other standard epidemiological model formulations allow for the analysis of time-varying exposures in the manner achieved by the Cox proportional hazards model. The exposure-response relationship (proportional hazards ratio) determined in this model intrinsically takes into account the effects of other causes of mortality that are unrelated to exposure (i.e., independent censoring). Further, all comparisons are made within the cohort by comparing the mortality experience of people with different exposures within the same cohort population. Nonetheless, the issue of competing risks that are dependent on exposure (e.g., asbestosis or nonmalignant respiratory disease) is an acknowledged uncertainty for this and other types of analyses (see Section 5.4.6).

The Cox proportional hazards model ([Cox, 1972](#)) is one of the most commonly used statistical models for the epidemiologic analysis of survival and mortality in cohort studies with extensive follow-up, including studies of the Libby worker cohort ([Larson et al., 2010b](#); [Moolgavkar et al., 2010](#)). In the Cox proportional hazards model, the conditional hazard function, given the covariate array  $Z$ , is assumed to have the form

$$\lambda(t | Z) = \lambda_0(t) \exp(\beta^T Z) \quad (5-13)$$

where  $\beta$  is the vector of regression coefficients,  $\lambda_0(t)$  denotes the baseline hazard function, and  $T$  denotes transposition of the vector. One of the strengths of this model is that knowledge of the baseline risk function is not necessary, and no particular shape is assumed for the baseline hazard; rather, it is estimated nonparametrically. The contributions of covariates to the hazard are multiplicative. When  $Z$  represents exposure and  $\beta^T Z$  is small, the Cox proportional hazards model is consistent with linearity of the dose-response relationship for low doses.

The Cox proportional hazards model assumes that a function of covariates (i.e., exposures) result in risks that are a constant multiple of the baseline hazard in unexposed individuals over some timescale, typically calendar time or age. This proportionality is assumed to be constant across the range of observed exposures, given the set of modeled covariates, and can be evaluated across time. When the proportional hazards assumption holds, it is possible to estimate the hazard ratio of exposure (relative risk) without estimating the hazard function in the unexposed (or in the lowest exposures seen within the study group), because this baseline hazard function drops out of the calculations.

Other methods common to occupational epidemiology, such as the use of standardized mortality ratios (results shown above in Tables 5-27 through 5-29) typically rely upon

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1 comparisons of the mortality experience in an exposed population group compared to that in the  
2 general population. However, the comparison population may not always be appropriate due to  
3 differences in general health status (e.g., the healthy worker effect) and differences in exposure  
4 to other risk factors for a specific disease (e.g., smoking history). The lack of comparability  
5 between the study population and the comparison population can lead to confounding by other  
6 measured or unmeasured characteristics that may be statistically associated with both the  
7 exposure of interest and the endpoint. The Cox proportional hazards model controls for such  
8 potentially confounding characteristics by using a comparison group from within the study  
9 population (i.e., internal controls). Internal controls are a statistically appropriate comparison  
10 group because they are expected to be more similar in potentially confounding characteristics to  
11 the remainder of the cohort, thereby controlling for both measured and unmeasured confounding  
12 and helping ensure that comparisons are more statistically valid.

13  
14 **5.4.3.3.1. Lung cancer mortality analysis in the Libby worker cohort.** As described in the  
15 previous section, quantitative exposure-response relationships for lung cancer mortality were  
16 evaluated using the Cox proportional hazards model. Cox proportional hazards models of this  
17 type require the specification of a timescale. Age is typically the time-related variable with the  
18 strongest relationship to cancer mortality and was used as the timescale in these analyses. Use of  
19 age as the timescale in a time-varying Cox proportional hazards model controls for age as a risk  
20 factor by design rather than by parametric modeling and effectively rules out age as a potential  
21 confounder. Individual covariates available to EPA in the complete analytic data set compiled  
22 from the NIOSH data were evaluated for their ability to explain lung cancer mortality. These  
23 included gender, race, birth year, age at hire, and various exposure-related variables including  
24 TWA workplace intensity of exposure in fiber/cc, job type, and the start and stop date of each  
25 different job. These data allowed for the computation of cumulative exposure, cumulative  
26 exposure with application of a half-life, and RTW cumulative exposure, with and without  
27 application of a half-life (see Section 5.4.2.5). Each exposure metric was also lagged by 0, 10,  
28 15, or 20 years. The use of a lag period aims to account for the latency period between the onset  
29 of lung cancer (which occurs some time before clinical diagnosis) and lung cancer mortality.

30 All lung cancer mortality analyses were conducted using SAS software version 9.1 (SAS,  
31 Cary, NC). EPA fit the extended Cox proportional hazards model ([Tableman and Kim, 2004](#);  
32 [Kleinbaum and Klein, 1996](#)), which included both time-independent factors such as gender, race,  
33 and date of birth, as well as time-dependent measures of LAA exposure over the entire time  
34 course of each individual's lifetime from his or her date of hire until death or loss to follow-up.  
35 The inclusion of date of birth in these analyses controls for potential birth cohort effect.

36 EPA's analyses of time-dependent exposure data included goodness-of-fit testing of the  
37 proportionality assumption for the Libby worker cohort. Because Cox proportional hazard  
38 models rely on the assumption that the hazard rate among the exposed is proportional to the

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hazard rate among the unexposed, it is important to evaluate the model against this assumption. Therefore, analyses of extended Cox proportional hazards models tested this assumption using a Wald test on the model interaction term between the LAA exposure metric and the timescale (i.e., age). As a general rule, a nonzero slope that is either increasing or decreasing indicates a violation of the proportional hazards assumption. Wald tests for the complete cohort consistently showed that the interaction term was a statistically significant predictor of lung cancer mortality ( $p < 0.05$ ) and was interpreted as evidence that the hazards did not remain proportional over time. The cause of the lack of proportionality is unknown, but several likely explanations are discussed in Section 5.4.3.4 below and in the discussion of uncertainties in Section 5.4.6.1.

#### **5.4.3.4. Rationale for Analyzing the Subcohort of Libby Workers After 1959**

Several possible explanations exist for the finding that duration of exposure was the best fitting exposure metric for mesothelioma mortality, as well as the finding of the lack of proportionality of hazards in the lung cancer mortality modeling.

- Duration of exposure, but not department code or job category, was known for 706 of 991 (71%) workers hired from 1935 to 1959. Without knowledge of the job category, the same exposure concentration had been assigned to almost all of these workers, likely resulting in a particularly large measurement error for exposure in approximately one-third of the total cohort of 1,871 workers. Assigning the same exposure concentration to so many of the workers hired before 1960, regardless of job, likely resulted in significant exposure misclassification and may explain the superior fit for duration of exposure in modeling of mesothelioma mortality relative to the other exposure metrics based on measured exposures.
- Even where the job category was identified, few exposure data exist prior to 1968. For the majority of job locations (23 of 25), no exposure measurements were available prior to 1967, and so exposures were estimated based on employee interviews (conducted in 1982) to determine what was known about major changes in operations between 1935 and 1967. For two job locations, dust-to-PCM extrapolations are based on measurements taken in the late 1960s; thus, extrapolating from the mid 1960s to the early 1960s is likely to be more certain than extrapolating further back in time. Random error in these exposure measurements would also generally attenuate the strength of association between exposure and observed effect during the earlier years of mine operation, and thus, a greater degree of measurement error in the earlier years could have resulted in the lack in proportionality of the hazard ratios for lung cancer over time. A greater degree of measurement error in the earlier years could also provide an explanation for the worse fit of the mesothelioma models that incorporated these exposure measures.
- Another explanation for the lack of proportional hazards in modeling lung cancer mortality may be that this cohort has an anomalous age structure due to the hiring of much older individuals during the time of the Second World War. Among those workers in the cohort hired prior to 1960, 9% were older than 50 years at the time of hire, and 22% were older than 40 years. Among those workers hired in 1960 or afterwards, only 4% were older than 50 years, and 14% were older than 40 years. Older workers differ

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from younger workers in several potentially important ways that could alter their response to exposures. Older workers were born in a different era, with different nutritional and public health standards, which may influence mortality patterns.

- The lack of proportional hazards in modeling lung cancer mortality may also be a reflection of confounding or effect modification, which can change in magnitude over time. The most likely candidate for confounding or effect modification is smoking. NIOSH records show that of the 1,871 workers in the full Libby workers cohort, 1,121 workers (60%) were missing smoking status data, while 750 (40%) had data with values “S” (Smoker), “Q” (Former Smoker), or “N” (Nonsmoker). Given this high percentage of missing values, EPA did not consider these smoking data to be adequate for use in the evaluation of confounding or effect modification. Effect modification by age is another possibility and may also explain the lack of proportionality in the modeling of lung cancer mortality as has been noted by [Richardson \(2009\)](#) in a two-stage clonal expansion model of lung cancer risk in a cohort of asbestos exposed workers; however, similar two-stage clonal expansion modeling of lung cancer risk in the same cohort of workers was unable to replicate that finding, which may be due to the reliance of this methodology on additional assumptions which EPA does not have a basis for making (see discussion of two-stage modeling in Section 5.4.3.3).
- Smoking rates and patterns among the subcohort of workers hired after 1959 are likely to have been more similar because smoking rates change more slowly over shorter periods of time than over longer ones. This restriction in time period of hiring would also result in less variation by birth year cohort, which is strongly related to smoking patterns as people of different generations develop different smoking rates. Thus, this restriction in the time period of hiring may make the cohort members more similar to each other, thereby possibly reducing the potential impact of any smoking-related confounding. Further discussion of the relevance of smoking can be found in the section on uncertainties (see Section 5.4.6).

When the assumption of proportionality is not met, the potential influence of confounding factors in the full-cohort analysis of lung cancer mortality is of concern. Additionally, the lack of job category information for 71% of the workers hired prior to 1960 and greater measurement error in early exposures may result in significant random exposure measurement error, which may bias the observed exposure-response relationships towards the null.

Although duration of exposure was the best exposure metric for modeling mesothelioma mortality in the full cohort, it does not allow quantitatively estimating an exposure-response relationship to support IUR. In addition, violation of the underlying statistical assumptions adversely affected modeling of lung cancer mortality in the full cohort. Therefore, EPA chose to undertake a subcohort analysis of workers hires after 1959.

While it is generally true that the use of more data is an advantage in statistical analyses because it allows for the computation of more statistically precise effect estimates, this advantage could not be realized, because of the difficulty in deriving risks from the full cohort analysis (see



next section on uncertainties remaining in the subcohort). The reasons stated in Section 5.4.2 for choice of Libby worker cohort data over other study populations are still valid for the subcohort. In particular, (1) these workers were directly exposed to LAA, (2) detailed work histories and job-specific exposure estimates are available to reconstruct estimates of each individual's occupational exposure experience with only nine workers completely missing job and department codes during the period when relatively high average time-weighted estimated exposure intensity was assigned, (3) the subcohort is still sufficiently large and has been followed for a sufficiently long period of time for cancer to develop (i.e., cancer incidence) resulting in mortality, and (4) the broad range of exposure experiences in the subcohort provided an information-rich data set.

EPA initially examined the fit of these models using several exposure metrics to predict mortality from mesothelioma and found that in this subcohort, the exposure metrics that included information on exposure concentration provided superior statistical fits to the exposure metrics based only on employment duration. In this same subcohort, the assumptions of the Cox proportional hazards model for analysis of lung cancer were also satisfied for the modeling of time-varying exposure.

On the other hand, there are quantitative uncertainties related to the choice of the subcohort. First of all, the numbers of cases of both lung cancer and mesothelioma are lower than in the whole cohort. Second, the follow-up of subcohort, while in excess of 40 years, may not be sufficiently long to encompass all potential lung cancer, especially, mesothelioma mortality related to LAA exposures. Third, the subcohort is younger and overall mortality is lower than in the full cohort. However, the choice of the subcohort is appropriate because of the superior exposure information based on a higher percentage of assigned exposure from actual measurements as opposed to inferred exposure values. The higher percentage of actual measurements allows a more accurate dose-response evaluation (see the discussion in [Lenters et al., 2012](#); [Lenters et al., 2011](#)) on the impact the quality of the exposure information has on estimates of dose-response relationships ([see Bateson and Kopylev, 2014](#)).

#### **5.4.3.5. Results of the Analysis of Mesothelioma Mortality in the Subcohort**

Of the 880 workers hired after 1959, 230 (26%) had died by December 31, 2006. The number of mesothelioma deaths in the subcohort is seven (two deaths coded in ICD-10 and five deaths coded in ICD-9). The mesothelioma death rate of 2.47 per 10,000 person-years for the subcohort is similar to the mesothelioma death rate of 2.68 per 10,000 person-years for the full cohort (18 mesothelioma deaths), with a difference of less than 10%.

Tables 5-30 to 5-32 show the mesothelioma mortality rate by duration of exposure, age of first exposure, and TSFE. As in the full cohort, both duration of exposure and TSFE show a relationship with mesothelioma mortality rate. However, unlike the full cohort, where there was no relationship with age at first exposure, in the subcohort, ages greater than 25 may be

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associated with higher risk than those below 25. EPA again evaluated the potential for overdispersion of the counts of mesothelioma deaths. In the subcohort, mean and variance of exposure are  $7.95 \times 10^{-3}$  and  $7.90 \times 10^{-3}$ , respectively. Therefore, as in the full cohort, overdispersion is very unlikely.

**Table 5-30. Mesothelioma mortality rate in the subcohort of employees hired after 1959 shown by duration of exposure (yr)**

	Duration			
	0–1 yr	1–2 yr	2–5 yr	5+ yr
Deaths/PY	1/14,942	0/4,129	1/4,614	5/4,669
Rate $\times 10^{-4}$	0.7	0	2.2	10.7

**Table 5-31. Mesothelioma mortality rate in the subcohort of employees hired after 1959 shown by age at first exposure**

	Age		
	15–25 yr old	25–35 yr old	35+ yr old
Deaths/PY	1/14,104	4/9,029	2/5,222
Rate $\times 10^{-4}$	0.7	4.4	3.8

**Table 5-32. Mesothelioma mortality rate in the subcohort of employees hired after 1959 shown by time since first exposure (TSFE)**

	Time since first exposure			
	<15 yr	15–25 yr	25–35 yr	35+ yr
Deaths/PY	0/12,954	2/8,155	3/5,731	2/1,514
Rate $\times 10^{-4}$	0	2.5	5.2	13.2

It is important to note that these marginal analyses, as well as the marginal analyses in the full cohort (see Section 5.4.3.2), do not specifically include the quantitative effects of the exposure—only the timing of exposure. Therefore, these marginal analyses provide an incomplete understanding of the quantitative exposure-response relationship. To more fully understand the effect of the timing of exposure, the quantitative effect of exposure must be modeled. Unlike the full cohort where personal exposure information is mostly missing, subcohort personal exposure information is available. Therefore, EPA next investigated the overall fit of different exposure models and then tabulated and represented graphically the mesothelioma mortality rate as predicted by several models that include personal exposure information.

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Table 5-33 shows the relative fit of various exposure metrics for mesothelioma mortality in the subcohort hired after 1959, including only those exposure metrics with information weights greater than 0.01. Information weights are computed from the DICs ([Burnham and Anderson, 2002](#)), and commonly used in Bayesian analyses. Information weights are computed by first assessing the differences between the best DIC and each of the others ( $\Delta DIC_i$ ) (see eq 5-14).

$$DIC\ w_i = \exp\left(-\frac{1}{2}\Delta DIC_i\right) / \sum_{r=1}^R \exp\left(-\frac{1}{2}\Delta DIC_i\right) \quad (5-14)$$

where

$R$  is the number of models and

$DIC\ w_i$  is information weight of the  $i^{th}$  model.

**Table 5-33. Comparison of model fit of exposure metrics for mesothelioma mortality in the subcohort hired after 1959.<sup>a</sup> Only the model fits with information weights greater than 0.010 are shown.<sup>b</sup>**

Exposure metric	Lag(yr)	DIC	Information weight
CE with 5-yr half-life	15	70.6	0.428
CE with 5-yr half-life	10	72.8	0.143
CE with 10-yr half-life	10	73.9	0.082
CE with 10-yr half-life	15	74.0	0.078
CE with 10-yr half-life	0	74.5	0.061
CE with 5-yr half-life	0	75.0	0.047
CE with 15-yr half-life	10	75.7	0.033
CE with 15-yr half-life	0	76.0	0.029
CE with 15-yr half-life	15	76.1	0.028
CE with 20-yr half-life	10	76.7	0.020
CE with 20-yr half-life	0	77.0	0.017
CE with 20-yr half-life	15	77.2	0.016

<sup>a</sup>Because one of mesothelioma deaths occurred in less than 20 yr from start of the exposure, lag 20 metrics assigned no exposure to this case, and the very poor fit of lag 20 metrics is a result.

<sup>b</sup>As discussed in Section 5.4.2.4, models with lag 0 were not considered further in derivation of unit risks.

Metrics with higher DICs and lower information weights indicate a poorer model fit and are not included in Table 5-33. The other exposure metrics that were evaluated included those metrics used in the full cohort analysis (duration of exposure, TSFE, age at death or censoring, RTW metrics, and CE with lag metrics), but none of these metrics fit as well as the metrics in Table 5-33.

The two metrics with cumulative exposure lagged 15 and 10 years, both with 5-year half-life, provided the two best fits as indicated by their lower DIC values and higher information weights (see Table 5-33). Cumulative exposures lagged 10 or 15 years, both with 10-year half-life, provided the next two best fits according to DIC values, but models including each of these metrics exhibited noticeably lower information weights than the best metric. All metrics in Table 5-33 contain a decay term and have the same number of parameters in their corresponding model, allowing for a direct comparison of the DIC values and information weights.

For models from the amphibole asbestos literature, in the subcohort hired after 1959, the DIC value for mesothelioma using the Peto metric (see eq 5-11) is substantially higher (DIC = 98.4) than for any of the metrics in Table 5-33. This indicates that the metric of exposure used in the previous IRIS IUR ([U.S. EPA, 1988a](#)) does not provide as good a fit for the LAA worker cohort as the other metrics of exposure in Table 5-33. Setting the power term on time since first exposure ( $k$  in eq 5-11) in the IRIS IUR ([U.S. EPA, 1988a](#)) metric to the values of 2 and 4, as suggested by [U.S. EPA \(1986a\)](#), continues to yield substantially higher DIC values compared to the fit values of the exposure metrics in Table 5-33 (DIC = 89.2 and 107.9, respectively). For the Peto model with clearance, increasing the power term in the Peto model with clearance to  $k = 3.9$  with decay  $\lambda = 0.068$  and  $k = 5.4$  with decay  $\lambda = 0.015$  decreased DIC slightly from the standard Peto model itself (DIC = 95.4 and 95.3, respectively). Using CE instead of C in decay models, as discussed in [Berry et al. \(2012\)](#), made the fit much worse, as measured by DIC. The fit also degraded when using the [Berry et al. \(2012\)](#) models of the form  $(C \text{ or } CE) \times (T - 5)^k$ .

Next, EPA considered which covariates should be added to the model with the exposure metric that provided the best fit. The addition of covariates “age at death or censoring” and “TSFE” did not improve the fit, as measured by DIC (results not shown).

As discussed above, EPA tabulated the mesothelioma rates for the two best fitting metrics in Table 5-33 and for alternative models proposed in the amphibole asbestos literature (i.e., Peto model and Peto model with clearance) in Tables 5-34 to 5-38. These tables show information by quintiles for each metric of exposure.

The first two tables (see Tables 5-34, 5-35) show a dose-response relationship between mesothelioma deaths and values of each exposure metric—while the mesothelioma data are sparse, higher values of metric correspond to higher rate of mesothelioma. Tables 5-36 to 5-38 show a somewhat less clear dose-response relationship for the Peto model and the Peto model with clearance, as the relationship between metric and rate appears to be somewhat parabolic.

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**Table 5-34. Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the cumulative exposure (CE) with 15-year lag and 5-year half-life**

	CE with 15-yr lag and 5-yr half-life				
	0–0.024	0.024–0.094	0.094–0.27	0.27–0.97	0.97+
Deaths/PY	1/4,858	0/5,975	0/5,827	0/5,494	6/5,751
Rate $\times 10^{-4}$	2.1	0	0	0	10.4

**Table 5-35. Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the cumulative exposure (CE) with 10-year lag and 5-year half-life**

	CE with 10-yr lag and 5-yr half-life				
	0–0.015	0.015–0.05	0.05–0.15	0.15–0.55	0.55+
Deaths/PY	1/5,315	0/5,626	0/5,953	1/5,995	5/5,465
Rate $\times 10^{-4}$	1.9	0	0	1.7	9.1

**Table 5-36. Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the Peto model**

	Peto metric				
	0–130	130–760	760–3,530	3,530–18,070	18,070+
Deaths/PY	1/4,585	0/5,460	0/5,639	4/5,943	2/6,727
Rate $\times 10^{-4}$	2.2	0	0	6.7	3.0

**Table 5-37. Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the Peto model with power  $k = 3.9$  and decay  $\lambda = 6.8\%/yr$**

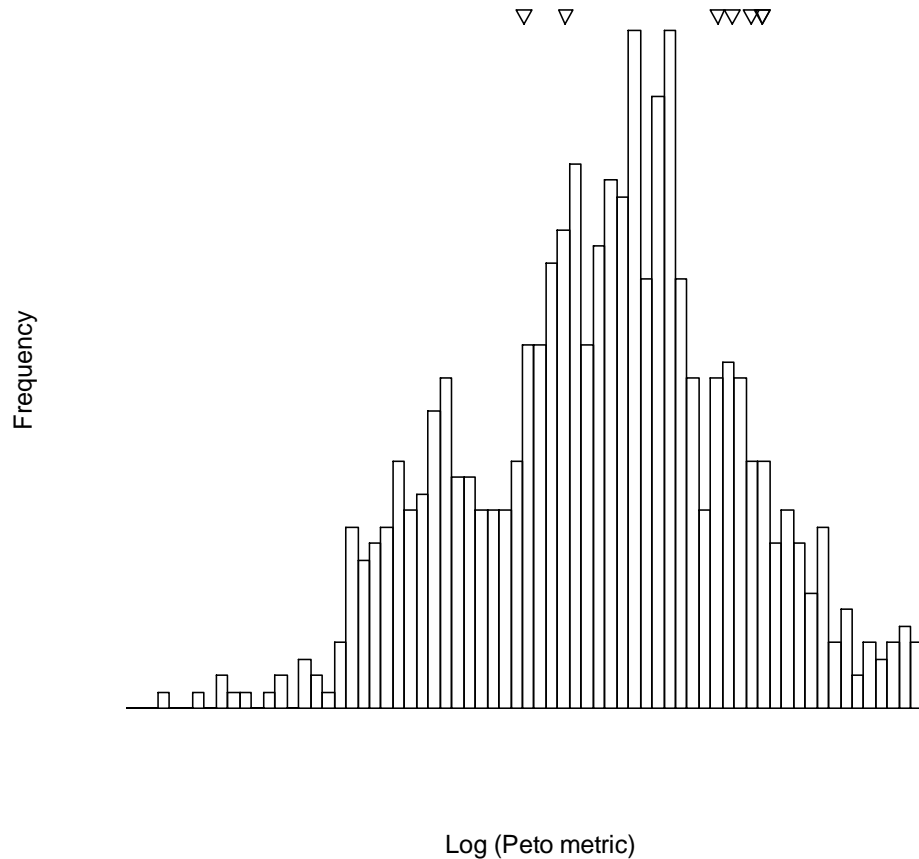
	Peto metric with power $k = 3.9$ and decay $\lambda = 6.8\%/yr$				
	0–311	311–1,837	1,837–7,400	7,400–35,330	35,330+
Deaths/PY	1/4,515	0/5,531	0/5,718	4/5,988	2/6,603
Rate $\times 10^{-4}$	2.4	0	0	6.7	3.0

**Table 5-38. Mesothelioma mortality rate in the subcohort of employees hired after 1959 for the Peto model with power  $k = 5.4$  and decay  $\lambda = 15\%/yr$**

	Peto metric with power $k = 5.4$ and decay $\lambda = 15\%/yr$				
	0–2,883	2,883–17,029	17,029–67,762	67,762–287,614	287,614+
Deaths/PY	1/4,492	0/5,588	0/5,710	4/5,941	2/6,624
Rate $\times 10^{-4}$	2.2	0	0	6.7	3.0

To further illustrate the fit of the Peto model to the Libby mesothelioma data, the frequency of the values of exposure computed using the Peto method and two best-fitting exposure metrics, which were CE with 5-year half-life and 10- or 15-year lag, were plotted and the mesothelioma cases were noted (see Figures 5-6 to 5-8). These figures provide more information than Tables 5-34 through 5-36. The histograms show in a relative scale the frequency of occurrence of each value of the specific exposure metric, thereby revealing the complete distribution and showing where exposure values of the cases were. In these figures, the exposure metric has been transformed to the natural log scale which yields a more normalized distribution. The figures show how the exposure values of the mesothelioma deaths relate to the exposure values of the subcohort of workers hired after 1959. Better fitting models are expected to show higher exposure values for the mesothelioma cases relative to those who did not die from mesothelioma, while poorer fitting models are expected to show mesothelioma cases scattered with equivalent density to the distribution as a whole and, thus, closer to the center of the distribution of exposure metric.

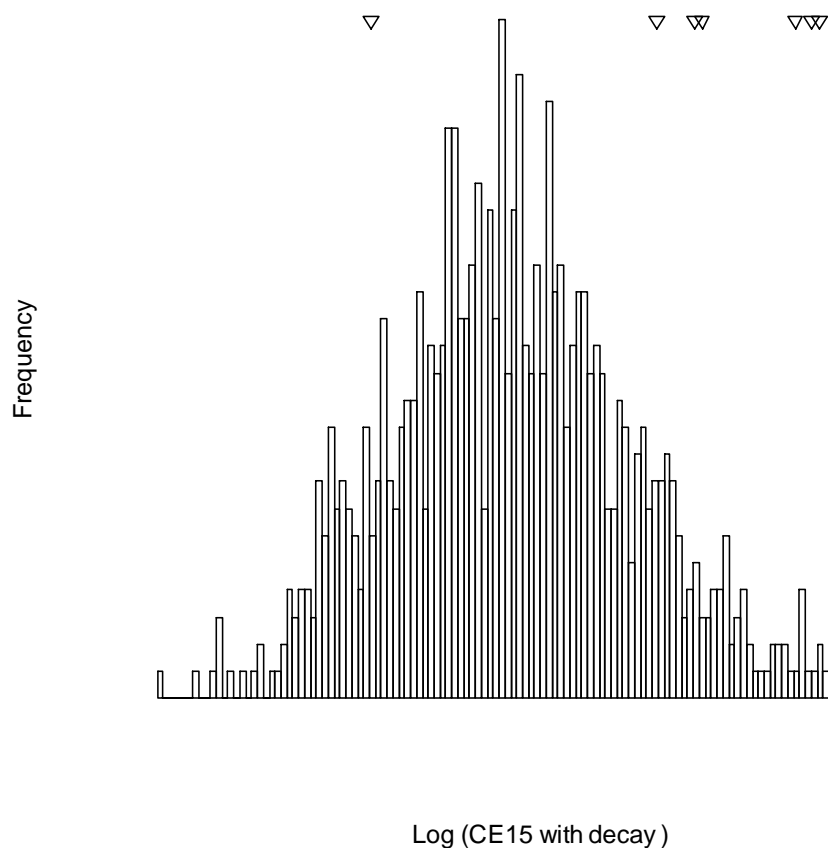
### Peto metric and mesothelioma deaths



**Figure 5-6. Distribution of values of the Peto metric and Peto metric values of mesothelioma deaths (shown as inverted triangles) in the subcohort of employees hired after 1959.**

From comparing Figures 5-6 (above) and Figures 5-7, and 5-8 (below), exposure values of the mesothelioma deaths based on the Peto exposure metric are clearly closer to the center of the distribution and, therefore, more like the values of those that did not die of mesothelioma. Thus Peto exposure metric does not reveal a higher likelihood of death from mesothelioma compared to the other exposure metrics. This is consistent with what was observed in Tables 5-34 through 5-38—the Peto model does not fit the subcohort data as well as CE metrics with decay fit. For the CE-based exposure metrics (see Figures 5-7 and 5-8), unlike the Peto exposure metric (see Figure 5-6), mesothelioma deaths are concentrated at high values of the exposure metrics and not as close to the center of the distribution.

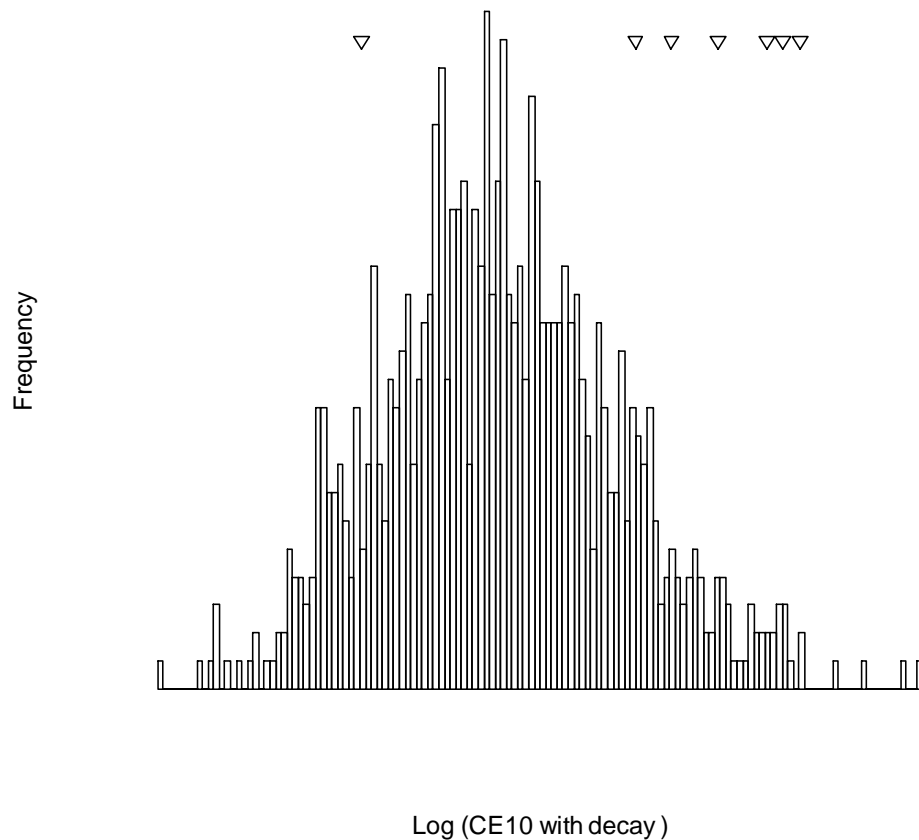
### CE with lag 15 and half-life 5 years and mesothelioma deaths



**Figure 5-7. Distribution of observed values of cumulative exposure (CE) with 15-year lag and 5-year half-life and CE with 15-yr lag and 5-yr half-life values of mesothelioma deaths (shown as inverted triangles) in the subcohort of employees hired after 1959.**



### CE with lag 10 and half-life 5 years and mesothelioma deaths



**Figure 5-8. Distribution of observed values of cumulative exposure (CE) with 10-year lag and 5-year half-life and CE with 10-yr lag and 5-yr half-life values of mesothelioma deaths (shown as inverted triangles) in the subcohort of employees hired after 1959.**

As discussed above and seen in Tables 5-34 through 5-38 and Figures 5-6 through 5-8:

- 1) While TSFE by itself, without personal exposure information, shows a clear relationship with increases in mesothelioma rates with increasing TSFE (see Table 5-32), adding information on concentration of fibers to the model (as the Peto model does) appears to degrade the fit (compare Table 5-36 to Table 5-32). The amphibole cohorts (amosite and crocidolite), which were used to derive the Peto model and the Peto model with clearance had little, if any, personal exposure data. Therefore, those models were predominantly based on the power of TSFE and demonstrated a good fit to the exposure timing data from those cohorts. Lack of good exposure data, and in particular personal exposure information, was a limitation of these analyses. In the case of the Libby workers subcohort, personal exposure data is available and when personal exposure is taken into account, the models do not appear to fit as well.

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2) The tabular fit of the best models from Table 5-33 (see Tables 5-34 to 5-35) compared to the fit in Tables 5-36 to 5-38, demonstrates somewhat better alignment for the subcohort. The tabular results are consistent with overall model fit statistics.

Comparing the relative fits of the empirical models based on individual-level exposure estimates (but just seven cases) with those literature-based models (Peto and Peto with clearance based on hundreds of cases but little actual individual-level exposure data) reveals uncertainties related to model selection. While there is understandable uncertainty in using empirical models based on a small number of cases, there is also uncertainty in applying literature-based models based on different type of amphibole asbestos without individual-levels exposure estimates. This uncertainty is discussed below in the section describing the derivation of the IUR (see Section 5.4.5.3). As described in Section 5.4.2.5, only metrics with nonzero lag were retained for derivation of unit risks. Table 5-39 shows KM (slope) and credible intervals for all metrics retained from Table 5-33.

**Table 5-39. Mesothelioma mortality exposure metrics fits, slopes per day, and credible intervals in the subcohort of employees hired after 1959**

Exposure metric	Lag yr	DIC	Slope $\times 10^{-5}$	90% Credible interval for slope $\times 10^{-5}$
CE—5-yr half-life	15	70.6	20.6	(10.2, 34.3)
CE—5-yr half-life	10	72.8	31.1	(15.2, 50.8)
CE—10-yr half-life	10	73.9	9.93	(5.00, 16.3)
CE—10-yr half-life	15	74.0	7.78	(3.72, 12.9)
CE—15-yr half-life	10	75.7	6.17	(3.04, 10.1)
CE—15-yr half-life	15	76.1	5.30	(2.63, 8.69)
CE—20-yr half-life	10	76.7	4.71	(2.34, 7.71)
CE—20-yr half-life	15	77.2	4.27	(2.12, 6.98)

Table 5-40 shows fits (DIC), KM (slope), and credible intervals for the Peto model and the Peto model with clearance (note that these slopes are not directly comparable with those in Table 5-39 because of difference in the units of time due to the use of different powers).

**Table 5-40. Peto model and Peto model with clearance fits, slopes per year, and credible intervals in the subcohort of employees hired after 1959**

Power (k)	Decay	DIC	Slope $\times 10^{-8}$	90% Credible interval for slope $\times 10^{-8}$
5.4	0.15	95.3	0.09	(0.04, 0.15)
3.9	0.068	95.4	0.66	(0.34, 1.09)
3	No	98.4	1.06	(0.52, 1.72)

Issues related to uncertainty in the choice of exposure metric are described further in the section on the derivation of the combined IUR of mesothelioma and lung cancer (see Section 5.4.5.3).

#### **5.4.3.6. Results of the Analysis of the Lung Cancer Mortality in the Subcohort**

EPA based its final analyses for lung cancer mortality on the subset of workers hired after 1959. Thus, this analysis is based on 32 deaths from lung cancer<sup>27</sup> (ICD-8: 2 deaths with the code 162.1; ICD-9: 1 death with the code 162.2, 20 deaths with the code 162.9; ICD-10: 9 deaths with the code C349) out of 230 total deaths that occurred in the subcohort of 880 workers.

Tables 5-41 to 5-43 show lung cancer mortality rates by duration of exposure, age of first exposure, and TSFE. As in the full cohort, duration of exposure, age at first exposure, and TSFE all show relationships with lung cancer mortality rate.

<sup>27</sup>Note that in the full cohort, it was unclear whether cases of tracheal cancer were included in the definition of lung cancer as many of the recorded ICD codes on death certificates did not provide sufficient detail to distinguish tracheal cancer cases from lung cancer cases. However, among the subcohort of workers hired after 1959, all the deaths from the broader category of cancers of the lung, bronchus, and trachea did provide sufficient detail to show that no deaths occurred from tracheal cancer.

**Table 5-41. Lung cancer mortality rate in the subcohort of employees hired after 1959 shown by duration of exposure (yr)**

	Duration			
	0–1 yr	1–2 yr	2–5 yr	5+ yr
Deaths/PY	13/14,942	5/4,129	2/4,614	12/4,669
Rate $\times 10^{-4}$	8.7	12.1	4.3	25.7
White male deaths/white male PY	12/13,779	5/3,848	2/4,251	12/4,601
White male rate $\times 10^{-4}$	8.7	13.0	4.7	26.1
White male SMR <sub>Montana</sub>	1.4	2.0	0.7	4.1
White male SMR <sub>United States</sub>	1.1	1.7	0.6	3.4

SMR standardized to white male lung cancer mortality rates obtained from [NCI \(2012\)](#).

**Table 5-42. Lung cancer mortality rate in the subcohort of employees hired after 1959 shown by age at first exposure**

	Age		
	15–25 yr old	25–35 yr old	35+ yr old
Deaths/PY	1/14,104	12/9,029	19/5,222
Rate $\times 10^{-4}$	0.7	13.3	36.4

SMR not computed due to lack of comparable rates.

**Table 5-43. Lung cancer mortality rate in the subcohort of employees hired after 1959 shown by time since first exposure (TSFE)**

	Time since first exposure			
	<15 yr	15–25 yr	25–35 yr	35+yr
Deaths/PY	4/12,954	11/8,155	13/5,731	4/1,514
Rate $\times 10^{-4}$	3.1	13.5	22.7	26.4
White male deaths/white male PY	4/12,054	11/7,560	12/5,404	4/1,461
White male rate $\times 10^{-4}$	3.3	14.6	22.2	27.4
White male SMR <sub>Montana</sub>	0.5	2.3	3.5	4.3
White male SMR <sub>United States</sub>	0.4	1.9	2.9	3.5

SMR standardized to white male lung cancer mortality rates obtained from [NCI \(2012\)](#).

1           As noted in Section 5.4.3.5, these marginal analyses do not specifically include the  
2 effects of the exposure as well as both the duration and the TSFE. Therefore, EPA investigated  
3 the overall fit of different exposure models and tabulated the results of several models that  
4 include personal exposure information.

5           All multivariate Cox proportional hazards models with time-varying exposures were  
6 initially fit, using one exposure metric at a time, to the subcohort hired after 1959 with covariates  
7 for gender, race, and date of birth. Lung cancer mortality was modeled using CE and RTW  
8 exposure, where each metric was potentially modified by four different half-lives (5, 10, 15, or  
9 20 years). Each of these exposure metrics was also evaluated with four different lag periods to  
10 allow for cancer latencies of 0, 10, 15, or 20 years. In all, 40 multivariate exposure-response  
11 models were evaluated for the adequacy of the exposure metric to fit the epidemiologic data.  
12 Each model and the comparative model fit statistics are presented in Table 5-44.

**Table 5-44. Model fit comparison for different exposure metrics and lung cancer mortality associated with LAA, controlling for age, gender, race, and date of birth. Results ordered at left by exposure metric and at right by model fit.**

Ordered by exposure metric			Ordered by model fit				
Exposure metric	Lag (yr)	AIC	Exposure metric	Lag (yr)	AIC	Multivariate model <i>p</i> -value†	Exposure <i>p</i> -value
CE	0	361.610	CE—10-yr half-life	10	358.400	0.0071	0.0009
CE	10	361.073	CE—5-yr half-life	10	358.502	0.0075	0.0010
CE	15	363.124	CE—15-yr half-life	10	358.777	0.0084	0.0015
CE	20	364.964	CE—20-yr half-life	10	359.122	0.0098	0.0022
CE—20-yr half-life	0	361.123	CE—5-yr half-life	15	359.910	0.0138	0.0032
CE 20-yr half-life	10	359.122	CE—10-yr half-life	15	360.543	0.0181	0.0079
CE—20-yr half-life	15	361.533	CE	10	361.073	0.0227	0.0188
CE—20-yr half-life	20	364.703	CE—20-yr half-life	0	361.123	0.0232	0.0155
CE—15-yr half-life	0	361.382	CE—15-yr half-life	15	361.129	0.0232	0.0162
CE—15-yr half-life	10	358.777	CE—15-yr half-life	0	361.382	0.0258	0.0184
CE—15-yr half-life	15	361.129	CE—20-yr half-life	15	361.533	0.0276	0.0254
CE—15-yr half-life	20	364.588	RTW 5-yr half-life	0	361.593	0.0283	0.0309
CE—10-yr half-life	0	362.169	CE	0	361.610	0.0285	0.0307
CE—10-yr half-life	10	358.400	CE—10-yr half-life	0	362.169	0.0360	0.0358
CE—10-yr half-life	15	360.543	RTW 10-yr half-life	0	362.283	0.0378	0.0588
CE—10-yr half-life	20	364.342	RTW 15-yr half-life	0	362.714	0.0452	0.0863
CE—5-yr half-life	0	364.225	RTW 20-yr half-life	0	362.973	0.0503	0.1084
CE—5-yr half-life	10	358.502	CE	15	363.124	0.0535	0.1215
CE—5-yr half-life	15	359.910	RTW 5-yr half-life	10	363.224	0.0558	0.1343
CE—5-yr half-life	20	363.644	CE—5-yr half-life	20	363.644	0.0662	0.1751
RTW	0	363.869	RTW	0	363.869	0.0726	0.2397
RTW	10	364.835	RTW 10-yr half-life	10	364.041	0.0778	0.2810
RTW	15	364.990	CE—5-yr half-life	0	364.225	0.0838	0.2908
RTW	20	364.502	RTW 15-yr half-life	10	364.336	0.0876	0.3733
RTW 20-yr half-life	0	362.973	CE—10-yr half-life	20	364.342	0.0878	0.3661
RTW 20-yr half-life	10	364.477	RTW 20-yr half-life	10	364.477	0.0927	0.4314
RTW 20-yr half-life	15	365.011	RTW	20	364.502	0.0936	0.5307
RTW 20-yr half-life	20	364.628	CE—15-yr half-life	20	364.588	0.0969	0.4815
RTW 15-yr half-life	0	362.714	RTW 20-yr half-life	20	364.628	0.0985	0.5763
RTW 15-yr half-life	10	364.336	RTW 15-yr half-life	20	364.662	0.0998	0.5909

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**Table 5-44. Model fit comparison for different exposure metrics and lung cancer mortality associated with LAA, controlling for age, gender, race, and date of birth. Results ordered at left by exposure metric and at right by model fit. (continued)**

Ordered by exposure metric			Ordered by model fit				
Exposure metric	Lag (yr)	AIC	Exposure metric	Lag (yr)	AIC	Multivariate model <i>p</i> -value†	Exposure <i>p</i> -value
RTW 15-yr half-life	15	365.001	CE—20-yr half-life	20	364.703	0.1014	0.5530
RTW 15-yr half-life	20	364.662	RTW 10-yr half-life	20	364.719	0.1021	0.6188
RTW 10-yr half-life	0	362.283	RTW 5-yr half-life	15	364.768	0.1041	0.6021
RTW 10-yr half-life	10	364.041	RTW 5-yr half-life	20	364.831	0.1067	0.6884
RTW 10-yr half-life	15	364.962	RTW	10	364.835	0.1069	0.6586
RTW 10-yr half-life	20	364.719	RTW 10-yr half-life	15	364.962	0.1124	0.8173
RTW 5-yr half-life	0	361.593	CE	20	364.964	0.1125	0.8204
RTW 5-yr half-life	10	363.224	RTW	15	364.990	0.1136	0.8809
RTW 5-yr half-life	15	364.768	RTW 15-yr half-life	15	365.001	0.1141	0.9100
RTW 5-yr half-life	20	364.831	RTW 20-yr half-life	15	365.011	0.1146	0.9599

CE = Cumulative exposure with or without exponential decay modeled with different half-lives.

RTW = Residence time-weighted exposure with or without exponential decay with different half-lives.

†Likelihood ratio test (overall model fit where  $p < 0.05$  indicated an adequate fit).

The assumptions of the Cox proportional hazards model were reevaluated for the subcohort. Restricting the cohort addressed each of the previously listed potential explanations for the lack of hazard proportionality (see Section 5.4.3.3). First, measurement error for exposures is likely to have been smaller after 1959 for several reasons. One reason is that 706 workers were removed from the analysis because job category and department code information were missing during all of their employment prior to 1960. Also, beginning in 1968, fiber concentrations by PCM analysis of site-specific air samples were available for all location operations to inform the JEM.

Second, prior to 1968, the exposure intensity for 23 of 25 location operations was estimated based on assumptions informed by employee interviews in the early 1980s. It is likely the uncertainty of these assumptions increased the farther back in time that exposures were estimated, making the earliest exposure estimates (1940s and 1950s) less certain than those only

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1 a few years before fiber count data were available. Further, between 1956 and 1967,  
2 dust-to-PCM extrapolation data were used to estimate exposures in the dry mill based on  
3 measurements taken in the late 1960s. Although there is some uncertainty in the conversion ratio  
4 selected by [Amandus et al. \(1987b\)](#), dust-to-fiber conversions are likely to be less uncertain than  
5 extrapolations further back in time to the 1950s and 1940s, where only one air sample for dust  
6 was available in 1944. Thus, the potential attenuation effect of nondifferential measurement  
7 error is likely to be reduced by examining the post-1959 cohort alone compared to the entire  
8 cohort.

9 Third, smoking rates among this more narrowly defined subcohort are likely to have been  
10 more homogeneous; thus, restricting analysis to this subcohort would help to limit any potential  
11 confounding due to smoking.

12 Finally, EPA conducted goodness-of-fit testing of the extended Cox proportional hazards  
13 model as applied to the subcohort hired post-1959. There was no evidence to reject the  
14 hypothesis of proportionality, and the exposure models demonstrated adequate fits to the data,  
15 with statistically significant effect estimates. In each of the Cox proportional hazards model  
16 analyses with time-varying exposures—across all the exposure metrics and across all the lag  
17 lengths—no violations of the assumption of proportionality of hazards were found.

18 As the exposure-response models cannot strictly be considered to be nested, a standard  
19 measure of fit called the AIC ([Burnham and Anderson, 2002](#)) was used for comparison of  
20 goodness of fit across models based on the same data set. In their text on model selection,  
21 [Claeskens and Hjort \(2008\)](#) state that “...for selecting a model among a list of candidates, AIC is  
22 among the most popular and versatile strategies.” [Claeskens and Hjort \(2008\)](#) also state that the  
23 model yielding the smallest AIC is judged the best fitting and it is a common practice in  
24 environmental epidemiology to simply select the single model with the best statistical fit (i.e., the  
25 lowest AIC) among the models evaluated. While large differences in AIC values can reveal  
26 important differences in model fit, small differences are less conclusive. For example, in a set of  
27 models differing in AIC by two or fewer units, each can be considered to have a substantially  
28 similar level of empirical support ([Burnham and Anderson, 2002 ; p. 70](#)).

29 Table 5-44 shows the models and exposure metrics ordered by fit. Of interest is whether  
30 there are models with distinct exposure metrics that adequately fit these data (as measured by  
31 statistical significance of the model  $p$ -value) and then whether a measure of relative fit exists  
32 among these adequately fitting models. Of the 40 exposure-response metrics, 14 demonstrated  
33 an adequate fit to the data as measured by the overall model fit, with the standard likelihood ratio  
34 test being statistically significant ( $p < 0.05$ ), as well as having statistically significant exposure  
35 metrics ( $p < 0.05$ ). However, note that only the nine models that demonstrated adequate model  
36 and exposure metric fit and incorporated a lag period to account for lung cancer mortality latency  
37 were advanced for potential use in developing a unit risk. While metrics that did not include an  
38 adjustment for lag on the exposure metric to account for cancer mortality latency were fit to

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1 these data for the sake of completeness, they were dropped from further consideration because  
2 they implicitly assume no passage of time between the initiation of cancer, subsequent promotion  
3 of that cancer, and mortality.

4 Several general patterns were discernible with respect to which exposure metric(s) best  
5 predicted lung cancer mortality when comparing AICs for relative model fit. The data show that  
6 lagging exposure by 10 years best predicts lung cancer mortality compared to other lags. This  
7 trend is seen across both the cumulative exposure without decay and the various half-life  
8 cumulative exposure metrics where a 10-year lag of exposure best predicts lung cancer mortality  
9 for all cumulative exposure metrics compared to other lags. Metrics with 15-year lags were  
10 generally the next best in terms of fit. Another conclusion is that the models that included RTW  
11 exposure metrics, regardless of half-life or lag, did not fit as well as the models that employed  
12 cumulative exposure with different half-lives and lags.

13 Among the 40 exposure metric models that were evaluated, the exposure model with the  
14 lowest AIC value was for cumulative exposure with a 10-year half-life for decay and a 10-year  
15 lag for cancer mortality latency and had a model  $p$ -value based on the likelihood ratio test of  
16 0.0071 (see Table 5-44). This multivariate model controlled for age, gender, race, and date of  
17 birth. This model estimated a KL (slope) of  $1.26 \times 10^{-2}$  per fibers/cc-yr based on a 365-day  
18 calendar year,<sup>28</sup> and the 95<sup>th</sup> percentile upper bound on this parameter was  $1.88 \times 10^{-2}$  per  
19 fibers/cc-yr. The  $p$ -value for the LAA regression coefficient (slope) was  $<0.001$ , indicating that  
20 this parameter was statistically significantly greater than zero. Table 5-45 shows the slopes and  
21 confidence intervals for all retained metrics from Table 5-44. Figure 5-9 shows the model  
22 residuals (in this case, the Schoenfeld residuals for Cox models) for the retained models in Table  
23 5-45. Patterns in such residuals such as an increasing or decreasing slope overall can indicate  
24 lack of fit. Attention is directed at the pattern of residuals with respect to age at death (the  
25  $x$ -axis). None of the plots appears to show a meaningful deviation from a linear function of age  
26 which indicates a lack of interaction between the model-predicted effect of exposure on lung  
27 cancer mortality risk and age. That is, the risk of lung cancer mortality does not appear to vary  
28 by age within the subcohort of workers at the Libby facility. There is no indication of a  
29 systematic lack of fit across these models excepting the nonlinear departure in the center of each  
30 residual plot which appears to be random and is minimized with the pattern in the residuals if  
31 smoothed out to a greater extent. The model fit residuals are consistent with the similarity of the  
32 AIC values in demonstrating similar model fit.

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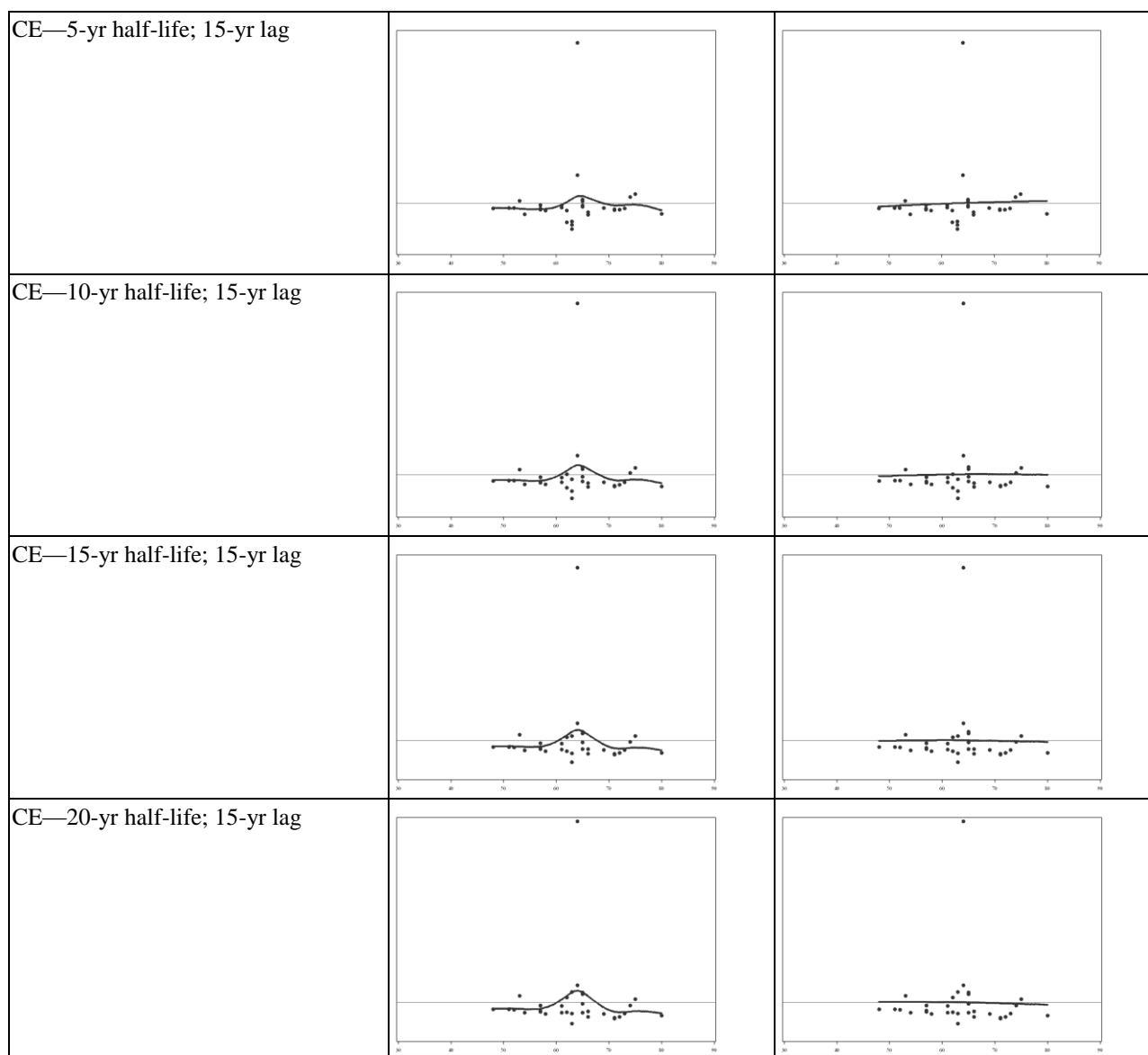
<sup>28</sup>The two-sided 90% confidence interval is  $(6.00 \times 10^{-3}, 1.88 \times 10^{-2})$ .

**Table 5-45. Lung cancer mortality exposure metrics fits, slopes, and confidence intervals (CI) for all retained metrics from Table 5-44.** Subset of lung cancer models with lagged exposures that yielded statistically significant model fit ( $p < 0.05$ ) and exposure metric fit ( $p < 0.05$ ) to the epidemiologic data

Exposure metric	Lag yr	AIC	Slope	SE	Exposure $p$ -value	90% CI for the slope
CE—10-yr half-life	10	358.400	0.0126	0.0038	0.0009	(0.0063, 0.0188)
CE—5-yr half-life	10	358.502	0.0179	0.0055	0.0010	(0.0089, 0.0269)
CE—15-yr half-life	10	358.777	0.0106	0.0033	0.0015	(0.0052, 0.0160)
CE—20-yr half-life	10	359.122	0.0095	0.0031	0.0022	(0.0044, 0.0146)
CE—5-yr half-life	15	359.910	0.0155	0.0052	0.0032	(0.0069, 0.0241)
CE—10-yr half-life	15	360.543	0.0115	0.0043	0.0079	(0.0044, 0.0186)
CE	10	361.073	0.0058	0.0025	0.0188	(0.0017, 0.0099)
CE—15-yr half-life	15	361.129	0.0097	0.0040	0.0162	(0.0031, 0.0163)
CE—20-yr half-life	15	361.533	0.0087	0.0039	0.0254	(0.0023, 0.0151)

Exposure Metric	Less smoothing	More smoothing
CE—no half-life; 10-yr lag		
CE—5-yr half-life; 10-yr lag		
CE—10-yr half-life; 10-yr lag		
CE—15-yr half-life; 10-yr lag		
CE—20-yr half-life; 10-yr lag		

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**Figure 5-9. Regression diagnostics showing model fit based on the Schoenfeld residuals with two levels of nonparametric smoothing (using cubic splines) to show any patterns of departures from the model predicted values.** In each plot, age at lung cancer mortality is shown against the model residuals. The  $x$ -axis shows the age of death while the  $y$ -axis shows the scaled residuals (predicted minus observed) according to the scale of the specific exposure metric.

1 According to the model results presented in Table 5-44, there were multiple exposure  
2 metrics that predicted lung cancer mortality and exhibited statistically significant effect  
3 estimates. Several other metrics were considered to fit nearly as well as the model with the  
4 smallest AIC because their AIC values were within two units of the exposure model with the  
5 lowest AIC, a proximity that can be considered to be a range that cannot clearly differentiate  
6 among models ([Burnham and Anderson, 2002](#)). As each of the other exposure metrics was  
7 based on a different configuration of the same exposure data, the different slopes (KLs) are not  
8 directly comparable, but all adequately fitting lagged models also produce statistically significant  
9 slopes for the exposure-response relationship ( $p < 0.05$ ). Of particular note are the results of the  
10 cumulative exposure model with a 10-year lag for latency but without a decay function because  
11 this model showed the lowest AIC among nondecay models.

12 The AIC values for models that included lag and/or half-life adjustments to the exposure  
13 metrics were not penalized in the regression analyses for using these extra parameters because  
14 these factors were not represented as covariates but rather were embedded in the computation of  
15 the exposure metric. While these results were obtained using each instance of lag and/or half-life  
16 terms in separate model fit, it may be appropriate to mathematically penalize the AICs for  
17 inclusion of these additional parameters. AIC values, as typically computed by regression  
18 software, include the addition of a penalty for model complexity as measured by the number of  
19 parameters that are fit in the regression model (thereby increasing the AIC). In the AIC  
20 calculations presented in Table 5-44, the models are treated as having the same number of  
21 parameters because each model represents the same individual's time-varying exposures in a  
22 different way but with a single exposure parameter in the regression models. For that reason, the  
23 models are equally penalized in the software's AIC calculation. However, because an argument  
24 can be made that exposure metrics that do not include a decay function, with an explicit half-life  
25 term, are implicitly more parsimonious (simpler), a comparison of the AICs is not  
26 straightforward. If the decay model fits were penalized for the inclusion of the decay function in  
27 the computation of the exposure metric, then with such an adjustment, the relative fit of the CE  
28 models would be somewhat improved in terms of their comparison with the values in Table 5-44  
29 (AICs are generally penalized two units for each additional parameter).

30 Table 5-45 displays the lagged exposure-response models and metrics with adequate  
31 model fit ( $p < 0.05$ ) to the epidemiologic data that were further considered. The units of the  
32 slopes are fiber/cc-yr. These slopes and confidence intervals represent calendar year continuous  
33 environmental exposure as described above and define the "Exposed Hazard Rate" in the  
34 life-table procedure when multiplied by the exposure level (see Appendix G for details). The  
35 plots in Figure 5-9 do not suggest a meaningful difference in model fits among the nine different  
36 parameterizations of exposure with adequate fit.

37 As presented in Table 5-45, the CE model with 10-year half-life and lag provided an  
38 adequate fit to the data based on the likelihood ratio test ( $p < 0.05$ ) and had the lowest AIC value.

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1 The cumulative exposure model with a 10-year lag also yielded a statistically adequate fit to  
2 these data ( $p < 0.05$ ), as did several decay models with a 15-year lag. These results demonstrate  
3 reasonable uncertainty in the metric of exposure such that no single exposure model can be  
4 definitively selected based on goodness of fit alone. Issues related to uncertainty in the choice of  
5 exposure metric are described further in the section explaining the derivation of the combined  
6 IUR of mesothelioma and lung cancer (see Section 5.4.5.3).

#### 7 8 **5.4.3.7. Sensitivity Analysis of the Influence of High Exposures in Early 1960s on the Model** 9 ***Fit in the Subcohort***

10 As discussed in Section 5.4.2.5, the comparison of model fit among various exposure  
11 metrics is an empirical process and does not necessarily reflect a specific biological or other  
12 factor as an underlying cause for model fit. Although data do not exist to evaluate biological  
13 bases for model fit, other potential factors can be explored where data allow. For example,  
14 because of concerns that very high ( $>100$  fibers/cc) 8-hour TWA exposures during 1960–1963  
15 (see Table 5-21) could have influenced the relative fit of the various exposure metrics, EPA  
16 conducted a sensitivity analysis of the impact on the relative model fit of reducing all estimated  
17 exposure intensities for 1960–1963 by 50%.

18 For modeling mesothelioma mortality on this revised data set, one change occurred in the  
19 relative fit of 3<sup>rd</sup> and 4<sup>th</sup> best fit decay models, but the observation that exposure metrics  
20 including decay fit better than exposure metrics without decay was unchanged (see Table 5-46).  
21 However, the fit of all the metrics decreased slightly, with each DIC increased between 0.3 and  
22 1.1. The metrics without decay and RTW metrics had DIC values higher than those in  
23 Table 5-46. The revised data set DIC for the model used in IRIS IUR ([U.S. EPA, 1988a](#))  
24 was 97.9.

**Table 5-46. Sensitivity analysis of model fit comparison for different exposure metrics and mesothelioma mortality associated with LAA.**

Estimated exposure intensities for all jobs during 1960–1963 were reduced by 50%.

Exposure metric	Lag (yr)	All workers hired after 1959 ( <i>n</i> = 880). Based on seven mesothelioma deaths (as shown in Table 5-33).	All workers hired after 1959 ( <i>n</i> = 880). Based on seven mesothelioma deaths. Exposures during 1960–1963 at 50%.
		DIC	DIC
CE—5-yr half-life	15	70.6	71.2
CE—5-yr half-life	10	72.8	73.9
CE—10-yr half-life	10	73.9	74.9
CE—10-yr half-life	15	74.0	74.6
CE—15-yr half-life	10	75.7	76.4
CE—15-yr half-life	15	76.1	76.7
CE—20-yr half-life	10	76.7	77.3
CE—20-yr half-life	15	77.2	77.7

CE = Cumulative exposure with exponential decay modeled with different half-lives.

- 1 For modeling lung cancer mortality on this revised data set, no difference was present in
- 2 the order of the relative fit between the same exposure metrics that fit the subcohort of workers
- 3 hired after 1959 and those exposures estimated by [Amandus et al. \(1987b\)](#) for 1960–1963 (see
- 4 Table 5-47). The metrics based on the revised data set fit marginally better based on AIC.

**Table 5-47. Sensitivity analysis of model fit comparison for different exposure metrics and lung cancer mortality associated with LAA, controlling for age, gender, race, and date of birth.** Estimated exposure intensities for all jobs during 1960–1963 were reduced by 50%. Lung cancer models presented include those with statistically significant multivariate model *p*-value and nonzero lag in exposure.

Exposure metric	Lag (yr)	All workers hired after 1959 ( <i>n</i> = 880) based on 32 deaths from lung cancer (as shown in Table 5-44)			All workers hired after 1959 ( <i>n</i> = 880) based on 32 deaths from lung cancer. Exposures during 1960–1963 at 50%.		
		AIC	Multivariate model <i>p</i> -value	Exposure <i>p</i> -value	AIC	Multivariate model <i>p</i> -value	Exposure <i>p</i> -value
CE—10-yr half-life	10	358.400	0.0071	0.0009	357.644	0.0051	0.0004
CE—5-yr half-life	10	358.502	0.0075	0.0010	357.781	0.0054	0.0005
CE—15-yr half-life	10	358.777	0.0084	0.0015	357.966	0.0059	0.0006
CE—20-yr half-life	10	359.122	0.0098	0.0022	358.283	0.0068	0.0009
CE—5-yr half-life	15	359.910	0.0138	0.0032	359.456	0.0113	0.0025
CE—10-yr half-life	15	360.543	0.0181	0.0079	360.167	0.0154	0.0067
CE	10	361.073	0.0227	0.0188	360.238	0.0159	0.0086
CE—15-yr half-life	15	361.129	0.0232	0.0162	360.810	0.0203	0.0138
CE—20-yr half-life	15	361.533	0.0276	0.0254	361.245	0.0244	0.0217

CE = Cumulative exposure with or without exponential decay modeled with different half-lives.

This sensitivity analysis reduces some of the potential uncertainty in the results that may have been attributed to exposure measurement error specific to the 1960–1963 time period when some of the estimated exposures were particularly high.

#### **5.4.3.8. Additional Analysis of the Potential for Confounding of Lung Cancer Results by Smoking in the Subcohort**

In the full cohort analysis, the proportional hazard assumption was not found to hold, and one of the reasons for this failure was the possible presence of confounding by smoking, which altered the proportionality of the hazard rate in the exposed workers compared to the baseline hazard rate over time. Confounding, which can bias observed results when there is an uncontrolled variable that is correlated with both the explanatory variable and the outcome variable, is a distinct concept from effect-measure modification (e.g., synergy), which might reflect different observed effects of exposure to LAA among smokers as compared to nonsmokers. The extent of effect-measure modification cannot be assessed without adequate data on smoking; however, the potential for effect-measure modification is discussed in Section 5.4.6.

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As an additional check on the potential for confounding, a novel method was evaluated to test for confounding by smoking in occupational cohorts that do not have data on smoking. A method has been described by [Richardson \(2010\)](#) to determine if an identified exposure relationship with lung cancer is confounded by unmeasured smoking in an occupational cohort study. [Richardson \(2010\)](#) demonstrated that an exposure of interest (i.e., LAA) can be used to predict an outcome other than lung cancer such as COPD, which is known to be caused by smoking, but not thought to be related to the exposure of concern.<sup>29</sup> If a positive relationship is identified where no causal association is suspected, this would suggest that smoking and the exposure metric (LAA) were positively correlated and that the identified exposure-response relationship was, in fact, confounded by smoking. EPA implemented this methodology to model the potential effects of LAA on the risk of COPD mortality ( $n = 18$ ) on the subcohort of workers hired after 1959. Using the exposure metric defined as cumulative exposure with a 10-year lag, the extended Cox proportional hazards model with time-varying exposures estimated a slope (beta) for COPD of  $-0.056$  per fiber/cc-yr based on a 365-day calendar year. The  $p$ -value for the coefficient (slope) was 0.102, indicating that this parameter was not statistically significantly different from zero. Using the exposure metric defined as cumulative exposure with a 10-year half-life for decay and a 10-year lag for cancer latency, the extended Cox proportional hazards model with time-varying exposures estimated a slope (beta) of  $-0.135$  per fiber/cc-yr based on a 365-day calendar year. The  $p$ -value for the coefficient (slope) was 0.116, indicating that this parameter was not statistically significantly different from zero.

Summarizing these findings, EPA used the method described by [Richardson \(2010\)](#) to evaluate whether exposures to LAA predicted mortality from COPD as an indication of potential confounding by smoking and found a nonsignificant negative relationship, which was inconsistent with confounding by smoking in the subcohort of workers hired after 1959.

#### **5.4.4. Exposure Adjustments and Extrapolation Methods**

The estimated exposures based on the JEM and work histories are discussed in Section 5.4.2.5. Note that all potency estimates (i.e., KM or KL) presented with units of fiber/cc-yr are for calendar year and not for occupational year, so no additional adjustment is needed to address this difference as may have been found in other evaluations based on occupational epidemiology cohort analyses. Adjustments for differences in breathing rates and the number of hours of exposure in an occupational (8-hour) day as compared to a whole (24-hour) day are not incorporated directly into the slope but rather applied in the derivation of the central risk and unit risk estimates.

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<sup>29</sup>[Richardson \(2010\)](#) cited literature with possible associations between asbestos and COPD which, if true, would have explained a positive association among the Libby workers cohort but should not detract from the use of the Richardson method as applied to these Libby workers, where a negative association is found.

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#### 5.4.5. Inhalation Unit Risk (IUR) of Cancer Mortality

The derivation of the unit risk estimates, defined as the lifetime risk of mortality from either mesothelioma or lung cancer from chronic inhalation of LAA at a concentration of 1 fiber/cc of air, is presented in the following subsections.

##### 5.4.5.1. Unit Risk Estimates for Mesothelioma Mortality

Computational details of the methodology and tables for deriving the lifetime unit risk for mesothelioma mortality are presented in Appendix G. For mesothelioma, the life-table procedure involves applying the absolute rates of mesothelioma mortality estimated in the Libby workers to the age-specific survival distribution of the general population to compute the age-specific risks of mesothelioma mortality expected at specific LAA exposure concentrations. The modeling analysis presented above showed that metrics including lag and half-life parameters provided the best empirical fit to the Libby worker subcohort data. Although there is uncertainty in applying these models for occupational mortality to estimate risks for different exposure levels and time patterns (see Section 5.4.6), following the recommendations of the *Guidelines for Carcinogen Risk Assessment* (U.S. EPA, 2005a), a linear low-dose extrapolation below the POD was used because the mode of action for LAA for mesothelioma is largely unknown. Lifetime unit risk estimates from the Peto model and the Peto model with clearance are presented in Table 5-48.

**Table 5-48. Unit risks for the Peto model and Peto model with clearance**

Model	Power	Decay	DIC	Central risk estimate	Unit risk
Peto with clearance	5.4	0.15	95.3	0.015	0.025
	3.9	0.068	95.4	0.035	0.058
Peto	3	No	98.4	0.117	0.191

The mesothelioma unit risks for model results presented in Table 5-39 and discussed in Section 5.4.3.5 are presented in Table 5-49. All of the metrics in Table 5-49 are CE metrics lagged 10–15 years (the fit of 20-year lag models was much worse because one of seven mesothelioma deaths occurred before 20 years; lags longer than 15 years are possible, and this is an uncertainty described in Section 5.4.6). Issues related to uncertainty in the choice of exposure metric are described further in the section on the derivation of the combined IUR of mesothelioma and lung cancer (see Section 5.4.5.3).

**Table 5-49. Mesothelioma mortality exposure metrics unit risks for the subcohort hired after 1959**

Exposure metric	Lag yr	DIC	Central risk estimate	Unit risk
CE—5-yr half-life	15	70.6	0.032	0.053
CE—5-yr half-life	10	72.8	0.054	0.088
CE—10-yr half-life	10	73.9	0.028	0.047
CE—10-yr half-life	15	74.0	0.020	0.032
CE—15-yr half-life	10	75.7	0.022	0.036
CE—15-yr half-life	15	76.1	0.017	0.027
CE—20-yr half-life	10	76.7	0.020	0.032
CE—20-yr half-life	15	77.2	0.015	0.025

**5.4.5.1.1. Adjustment for mesothelioma underascertainment.** For mesothelioma, the undercounting of cases (underascertainment) is a particular concern given the limitations of the ICD classification systems used prior to 1999. In practical terms, this means that some true occurrences of mortality due to mesothelioma are missed on death certificates and in almost all administrative databases such as the National Death Index. Even after the introduction of a special ICD code for mesothelioma with the introduction of ICD-10 in 1999, detection rates are still imperfect ([Camidge et al., 2006](#); [Pinheiro et al., 2004](#)), and the reported numbers of cases typically reflect an undercount of the true number.

[Kopylev et al. \(2011\)](#) reviewed the literature on this underascertainment and developed a general methodology to account for the likely numbers of undocumented mesothelioma deaths using the Libby worker cohort as an example. Because the analysis of mesothelioma mortality was based on absolute risk, it was possible to compensate for mesothelioma underascertainment in the Libby worker subcohort. [Kopylev et al. \(2011\)](#) considered analyses when mesothelioma type (i.e., pleural and peritoneal) is unknown and when it is known. As the number of peritoneal mesotheliomas is partially known in the Libby worker subcohort, the method for known proportion of pleural and peritoneal mesothelioma deaths is briefly described here.

[Selikoff and Seidman \(1992\)](#) provided information on the likelihood that individuals who have been diagnosed as having mesothelioma will have that disease recorded (in some field) on their death certificate. Their results are based on histopathological analysis ([Ribak et al., 1991](#)) of a very large cohort of insulators, with more than 450 mesothelioma cases. Despite medical advances, diagnosis of mesothelioma is still very challenging, and histopathology is still a standard diagnostic tool today (e.g., [Mossman et al., 2013](#)). Using their results on the most common misdiagnoses of mesothelioma (mesothelioma diagnosed as lung, colon, and pancreatic cancers; and conversely, other diseases misdiagnosed as mesothelioma) and likelihoods of corresponding misdiagnoses, [Kopylev et al. \(2011\)](#) conducted a simulation study randomly

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creating a new data set with the number of mesothelioma deaths simulated in the full Libby cohort to match the [Selikoff and Seidman \(1992\)](#) results. That simulated data set included 24 mesothelioma cases that were obtained using the underascertainment estimate of 37% derived from [Selikoff and Seidman \(1992\)](#). The full cohort was used for the simulation study because of the larger number of mesothelioma cases and because limitations of exposure information is not relevant for that analysis. Using the Poisson model and MCMC simulation similarly to that described in Section 5.1.3.1, [Kopylev et al. \(2011\)](#) calculated the mean of underascertainment of risk and its 90% confidence interval to be 1.39 and (0.80; 2.17).

This method to adjust for underascertainment was applied to the Libby workers subcohort; mesothelioma mortality-adjusted unit risks are listed in Table 5-50.

**Table 5-50. Mesothelioma unit risks for the subcohort hired after 1959 adjusted for underascertainment**

Exposure metric	Lag yr	DIC	Adjusted central risk estimate	Adjusted unit risk
CE—5-yr half-life	15	70.6	0.044	0.074
CE—5-yr half-life	10	72.8	0.075	0.122
CE—10-yr half-life	10	73.9	0.039	0.065
CE—10-yr half-life	15	74.0	0.028	0.044
CE—15-yr half-life	10	75.7	0.031	0.050
CE—15-yr half-life	15	76.1	0.024	0.038
CE—20-yr half-life	10	76.7	0.028	0.044
CE—20-yr half-life	15	77.2	0.022	0.035

Similarly, for the subcohort data, the Peto model and the Peto model with clearance unit risks are presented in Table 5-51.

**Table 5-51. Mesothelioma unit risks for the subcohort hired after 1959 based on the Peto model and the Peto model with clearance adjusted for mesothelioma underascertainment**

Model	Power	Decay %	DIC	Adjusted central risk estimate	Adjusted unit risk
Peto with clearance	5.4	0.15	95.3	0.021	0.035
	3.9	0.068	95.4	0.049	0.086
Peto	3	No	98.4	0.163	0.265

#### 5.4.5.2. Unit Risk Estimates for Lung Cancer Mortality

Computational details of the methodology and tables for deriving the unit risk for lung cancer mortality are presented in Appendix G. For lung cancer, the life-table procedure involves application of the hazard rates of lung cancer mortality estimated in the Libby workers to the age-specific background rates of lung cancer in the general population (accounting for the age-specific survival distribution of the general population) to compute the age-specific risks of lung cancer mortality expected at specific LAA exposure concentrations. Although there is uncertainty in applying these models for occupational mortality to the estimation of risks for different exposure levels and time patterns (see Section 5.4.6), following the recommendations of the *Guidelines for Carcinogen Risk Assessment* (U.S. EPA, 2005a), a linear low-dose extrapolation below the POD was used because the mode of action for LAA for lung cancer is undetermined. The nine exposure-response models shown in Table 5-45 all had reasonably similar goodness of fits. No single model stands out as statistically superior; however, there is a range of quality of fit within the set that could be considered adequate. The lung cancer mortality unit risks are shown in Table 5-52.

**Table 5-52. Unit risks for subset of lung cancer models with lagged exposures that yielded statistically significant model fit ( $p < 0.05$ ) and exposure metric fit ( $p < 0.05$ ) to the epidemiologic data**

Exposure metric	Lag	AIC	Exposure $p$ -value	Central risk estimate (based on EC <sub>01</sub> )	Unit risk (based on LEC <sub>01</sub> )
C—10-yr half-life	10	358.400	0.0009	0.0260	0.0389
CE—5-yr half-life	10	358.502	0.0010	0.0195	0.0293
CE—15-yr half-life	10	358.777	0.0015	0.0300	0.0455
CE—20-yr half-life	10	359.122	0.0022	0.0326	0.0501
CE—5-yr half-life	15	359.910	0.0032	0.0167	0.0260
CE—10-yr half-life	15	360.543	0.0079	0.0231	0.0375
CE	10	361.073	0.0188	0.0399	0.0679
CE—15-yr half-life	15	361.129	0.0162	0.0258	0.0434
CE—20-yr half-life	15	361.533	0.0254	0.0280	0.0486

LEC<sub>01</sub> = 95% lower confidence limit of the exposure concentration associated with a 1% increased risk.

Using the results of the exposure model with the lowest AIC value (i.e., cumulative exposure with a 10-year half-life for decay and a 10-year lag for cancer latency) alone, the 95% lower confidence limit of the exposure concentration associated with a 1% increased risk (LEC<sub>01</sub>) yielded a lifetime unit risk of 0.0389 per fiber/cc. The value of the risk that would

correspond to the measure of central tendency involves EC<sub>01</sub> rather than LEC<sub>01</sub>. The EC<sub>01</sub> yielded a lifetime central estimate of 0.0260 per fiber/cc.

Using the results of the exposure model based on cumulative exposure with a 10-year lag for cancer latency, the LEC<sub>01</sub> for the adult-only exposures was determined to be 0.191 fiber/cc. This LEC<sub>01</sub> yielded a lifetime unit risk of 0.0679 per fiber/cc. The EC<sub>01</sub> for the adult-only exposures was determined to be 0.325 per fiber/cc. When divided into a POD of 1%, this EC<sub>01</sub> yielded a lifetime central estimate of 0.0399 per fiber/cc.

The resulting unit risks in Table 5-52 ranged from 0.0260 to 0.0679 per fiber/cc. This shows that the unit risk (i.e., 0.0389 per fiber/cc) based on the exposure metric with the lowest AIC value (i.e., cumulative exposure with a 10-year half-life for decay and a 10-year lag for cancer latency) is in the center of this range, and is thus statistically robust. However, because this estimate is in the middle of the range, it does not capture the uncertainty across metrics with similar goodness of fit. As noted (see Section 5.4.3.6., an argument can be made that the CE metric with a 10-year lag and no half-life is implicitly more parsimonious (simpler) because it was not explicitly adjusted to include decay, although this metric is mathematically equivalent to CE metric with a 10-year lag and an infinitely long decay half-life. Conceptually, the AIC values are penalized for increased model complexity (thereby increasing the AIC). The CE metric with a 10-year lag does fit these data—both statistically and by examination of the residuals. Further, the CE metric is a simpler and more straightforward metric, and has an extensive tradition of use in the epidemiologic literature and in the practice of risk assessment.

#### **5.4.5.3. Inhalation Unit Risk (IUR) Derivation for Combined Mesothelioma and Lung Cancer Mortality**

For mesothelioma, the exposure-response models developed by EPA using personal exposure data on the subcohort (see Table 5-50) provided better fit to the subcohort data than the Peto model and the Peto model with clearance that have been proposed in the asbestos literature (see Table 5-51). These variations of the Peto model have been shown to predict mesothelioma mortality more precisely than the original Peto model in a large crocidolite-exposed cohort of 6,908 workers with 329 mesothelioma deaths from Wittenoom, Australia ([Berry et al., 2012](#))—specifically, the best fitting models for that cohort was the Peto model with a power  $k = 3.9$  and decay rate of  $\lambda = 0.068$  (approximately 10-year half-life) and Peto model with a power  $k = 5.4$  and decay rate of  $\lambda = 0.15$  (approximately 5-year half-life).

The exposure-response models developed by EPA using personal exposure data on the subcohort (see Table 5-50) show estimated lifetime unit risks of 0.074 and 0.122 per fiber/cc. The two Peto models with different clearances and powers of  $k$  show lifetime unit risks of 0.086 and 0.035 per fiber/cc, respectively (see Table 5-51). The results of the two different approaches to modeling, the first based only on the LAA data and the second based on a much larger population of workers exposed to another but different amphibole asbestos, reveal a relatively

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small degree of uncertainty in the derivation of the mesothelioma lifetime unit risk. Therefore, EPA selected the model derived directly from the Libby data with the strong support of the model that best fit the mesothelioma risk in a much larger cohort of amphibole-exposed workers. EPA's selected model based on cumulative exposure with a 10-year lag and 5-year half-life yielded a lifetime unit risk for mesothelioma of 0.122 per fiber/cc which encompasses other three risk estimates. Table 5-53 shows the combined IUR for mesothelioma and lung cancer based on the selected mesothelioma model, the two best-fitting mesothelioma models from the epidemiologic literature (the Peto model with clearance), as well as the combined IUR based on Peto model.

**Table 5-53. Estimates of the combined central estimate of the unit risk for mesothelioma and lung cancer and the combined upper-bound lifetime unit risks for mesothelioma and lung cancer risks (the Inhalation Unit Risk) for different combination of mesothelioma and lung cancer models<sup>a</sup>**

Lung cancer	Mesothelioma	Combined central estimate (per fiber/cc)	Combined upper bound (per fiber/cc)
Selected IUR based directly on the Libby data			
CE10 Subcohort	CE10 5-yr half-life	0.115	0.169
Best models from the epidemiologic literature (Peto model with clearance)			
CE10 Subcohort	Peto with clearance Decay rate of 6.8%/yr Power of time = 3.9 Subcohort	0.089	0.135
CE10 Subcohort	Peto with clearance Decay rate of 15%/yr Power of time = 5.4 Subcohort	0.061	0.092
Alternative model from the epidemiologic literature (Peto model)			
CE10 Subcohort	Peto No decay Power of time = 3 Subcohort	0.203	0.308

<sup>a</sup>Note that for the IUR values shown in this table, the fiber concentration are presented here as continuous lifetime exposure in fiber/cc where exposure measurements are based on analysis of air filters by PCM. Current analytical instruments used for PCM analysis have resulted in a standardization of minimum fiber width considered visible by PCM between 0.2 and 0.25  $\mu\text{m}$ . Historical PCM analysis (1960s and early 1970s) generally had less resolution, and fibers with minimum widths of 0.4 or 0.44  $\mu\text{m}$  were considered visible by PCM ([Amandus et al., 1987b](#); [Rendall and Skikne, 1980](#)). Methods are available to translate exposure concentrations measured in other units into PCM units for comparison.



1 For lung cancer, this assessment selected the upper bound among the lung cancer lifetime  
2 unit risks from the plausible exposure metrics (regardless of the small residual differences in  
3 quality of fit). Because there were few metrics with unit risks higher than the best fitting  
4 metric's unit risk for lung cancer mortality endpoint, this method effectively selects the highest  
5 lifetime unit risk among those considered for the lung cancer mortality endpoint. Based on the  
6 selected model for lung cancer mortality using the cumulative exposure with a 10-year lag in the  
7 Libby subcohort data yields a central unit risk estimate of 0.040 per fiber/cc and upper-bound  
8 lifetime unit risk of 0.0680 per fiber/cc.

9 Once the cancer-specific lifetime unit risks are selected, the two are then combined. It is  
10 important to note that this estimate of overall potency describes the risk of mortality from cancer  
11 at either of the considered sites and is not just the risk of both cancers simultaneously. Because  
12 each of the unit risks is itself an upper-bound estimate, summing such upper-bound estimates  
13 across mesothelioma and lung cancer mortality is likely to overpredict the overall risk.  
14 Therefore, following the recommendations of the *Guidelines for Carcinogen Risk Assessment*  
15 ([U.S. EPA, 2005a](#)), a statistically appropriate upper bound on combined risk was derived in order  
16 to gain an understanding of the overall risk of mortality resulting from mesothelioma and from  
17 lung cancer.

18 Because the estimated risks for mesothelioma and lung cancer mortality were derived  
19 using Poisson and Cox proportional hazards models, respectively, it follows from statistical  
20 theory that each of these estimates of risk is approximately normally distributed. For  
21 independent normal random variables, a standard deviation for a sum is easily derived from  
22 individual standard deviations, which are estimated from confidence intervals: standard  
23 deviation = (unit risk – central risk) ÷  $Z_{0.95}$ , where  $Z_{0.95}$  is a standard normal quantile equal to  
24 1.645. For normal random variables, the standard deviation of a sum is the square root of the  
25 sum of the squares of individual standard deviations.

26 As shown in Table 5-50, the upper bound of the selected mesothelioma mortality unit  
27 risks was 0.122 per fiber/cc (highest adjusted unit risk value). The associated central estimate of  
28 risk was 0.075 per fiber/cc for mesothelioma mortality. Table 5-52 shows the upper bound of the  
29 selected lung cancer mortality unit risk was 0.068 per fiber/cc (highest unit risk value based on  
30 LEC<sub>01</sub>). The associated central estimate of risk was 0.040 per fiber/cc for lung cancer mortality.

31 It is important to mention here that the assumption of independence of the estimated risks  
32 for mesothelioma and lung cancer mortality (note above) is a theoretical assumption, as there is  
33 insufficient data on independence of mesothelioma and cancer risks for LAA. However, in a  
34 somewhat similar context of different tumors in animals, [NRC \(1994\)](#) stated: "...a general  
35 assumption of statistical independence of tumor-type occurrences within animals is not likely to  
36 introduce substantial error in assessing carcinogenic potency." To provide numerical bounding  
37 analysis of impact of this assumption, EPA used results of [Chiu and Crump \(2012\)](#) on the upper  
38 and lower limits on the ratio of the true probability of a tumor of any type and the corresponding

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probability assuming independence of tumors. The lower limit is calculated by  $[1 - \min(p_1, p_2)] / (1 - p_1 \times p_2)$  and the upper limit is  $\min(1, 2 - p_1 - p_2) / (1 - p_1 \times p_2)$ . Substituting the risk of lung cancer ( $p_1$ ) is 0.040 and the risk of mesothelioma ( $p_2$ ) is 0.075, the lower limit is 0.963 and the upper limit is 1.003. A value of 1.0 indicates independence. Because lower and upper values are both very close to the value of 1.0, this demonstrates that the assumption of independence in this case does not introduce substantial error, consistent with what [NRC \(1994\)](#) has stated.

In order to combine the unit risks, first obtain an estimate of the standard deviation of the sum of the individual unit risks as:

$$\sqrt{[[(0.122 - 0.075) \div 1.645]^2 + (0.068 - 0.040) \div 1.645]^2} = 0.033 \text{ per fiber/cc} \quad (5-15)$$

Then, the combined central estimate of risk of mortality from either mesothelioma or lung cancer is  $0.040 + 0.075 = 0.115$  per fiber/cc, and the combined IUR is  $0.115 + 0.033 \times 1.645 = 0.169$  per fibers/cc.

To illustrate the uncertainty in the selected IUR, Table 5-53 shows central risks and upper bounds for the combined IUR for selected metrics for each cancer (CE10 for lung cancer and CE10 with 5-year half-life for mesothelioma) and for selected lung cancer model (CE10) with other mesothelioma models suggested in the literature from Tables 5-51. The selected IUR does address issues of model uncertainty because a higher risk is only given by the Peto model, but the Peto model tends to overestimate mortality from mesothelioma in asbestos cohorts with long follow-up (e.g., [Berry et al., 2012](#); [Barone-Adesi et al., 2008](#)).

#### *Age-dependent adjustment factor*

As discussed in Section 4.7.1.1, there is no chemical-specific information for LAA, or general asbestos, that would allow for the computation of a chemical-specific age-dependent adjustment factor for assessing the risk of exposure that include early-life exposures.

The review of mode-of-action information in this assessment (see Section 4.6.2.2) concluded that the available information on the mode of action by which LAA causes lung cancer or mesothelioma is complex and a mode of action is not established at this time. Thus, in accordance with EPA's *Supplemental Guidance for Assessing Susceptibility from Early-Life Exposure to Carcinogens* ([U.S. EPA, 2005b](#)), the application of the age-dependent adjustment factors for substances that act through a mutagenic mode of action is not recommended.

**5.4.5.3.1. Comparison with other published studies of Libby workers cohort.** For lung cancer, two alternative analytic approaches to the use of EPA's extended Cox proportional hazards models are considered here for the calculation of a unit risk of lung cancer mortality. All of the

choices are based on different analyses of the Libby worker cohort; however, inclusion criteria differ among the analyses as does the length of mortality follow-up. Each of the two alternative approaches has two options to estimate the slope of the exposure-response relationship in place of the regression slope estimated from the Cox proportional hazards model.

The first approach would be to use the published categorical results based on [Sullivan \(2007\)](#), which offers two options: (1) estimate a slope to those categorical data or (2) use the slope estimated in a published reanalysis of categorical data of the [Sullivan \(2007\)](#) cohort by [Berman and Crump \(2008\)](#). The second approach would be to use the published regression results of other researchers who modeled the underlying continuous data. There are two options under this approach: (1) use the slope estimated by [Larson et al. \(2010b\)](#) or (2) use the slope estimated by [Moolgavkar et al. \(2010\)](#).

For comparison purposes, the lung cancer unit risk from these alternatives is computed; however, as all analyses are based upon different subsets of the Libby workers cohort and used different analytic methods, the results are not necessarily interchangeable. Table 5-54 summarizes lung cancer risks derived from these studies.

**Table 5-54. Lung cancer regression results from different analyses of cumulative exposure in the cohort of workers in Libby, MT.** All analyses used NIOSH-collected exposure data but used different cohort definitions, lengths of follow-up, and lengths of exposure lags to account for cancer latency.

Lung cancer analysis	Cohort definition	Follow-up	Lung cancer cases/N	Slope per fiber/cc-yr $\times 10^{-3}$ (calendar yr)	Risk based on upper confidence limit (UCL) on the slope (per fiber/cc)
This current assessment	Hired post-1959 Exposures 1960–1982	2006	32/880	5.8	0.068
<a href="#">Sullivan (2007)</a>	Still alive post-1959 White males Exposures 1960–1982	2001	99/1,672	4.2	0.037
<a href="#">Moolgavkar et al. (2010)<sup>a</sup></a>	Still alive post-1959 White males Exposures 1960–1982	2001	95/1,662	1.69	0.011
<a href="#">Berman and Crump (2008)<sup>b</sup></a>	Still alive post-1959 White males Exposures 1960–1982	2001	93/1,672	3.96	0.079
<a href="#">Larson et al. (2010b)</a>	Full cohort Exposures 1935–1993	2006	98/1,862	1.61	0.010

<sup>a</sup>Reanalysis of [Sullivan \(2007\)](#).

<sup>b</sup>[Sullivan \(2007\)](#) and reanalysis of [Sullivan \(2007\)](#) state slightly different number of lung cancers. It is impossible to reconcile these numbers from published information.

1 The first alternative analytical approach to estimating the extra risk from a linear  
2 regression of individual mortality data was to use a technique that is standard in EPA cancer risk  
3 assessments ([U.S. EPA, 2005a](#)) when individual-level data are not available. This approach used  
4 a weighted linear regression of standardized risk ratio (SRR) estimators for lung cancer mortality  
5 in white males, as calculated in the NIOSH cohort analysis ([Sullivan, 2007](#)), with categorical  
6 cumulative exposure and a 15-year lag. The [Sullivan \(2007\)](#) analysis was based only on those  
7 workers who had not died or been lost to follow-up before January 1, 1960 (in contrast to  
8 employment beginning after January 1, 1960), because the NIOSH software program (Life-Table  
9 Analysis System) used for this analysis only has statistics on external comparison rates for  
10 asbestosis (one of the primary outcomes of interest in the [Sullivan \(2007\)](#) analysis) beginning in  
11 1960. The SRR analysis involves internal comparisons of lung cancer mortality rates in the  
12 higher exposure categories to the lung cancer mortality rates in the lowest exposure category.  
13 The weights used for the SRRs were the inverses of the variances. Midpoints of the exposure  
14 intervals were used, and for the unbounded interval, the midpoint was assumed to be twice the  
15 starting point of that interval.

16 Using this approach, a regression coefficient of  $4.2 \times 10^{-3}$  per fibers/cc-yr (standard error  
17 [SE] =  $7.7 \times 10^{-4}$  per fibers/cc-yr,  $p = 0.03$ ) was obtained from the weighted linear regression of  
18 the categorical SRR results. Because the data from [Sullivan \(2007\)](#) were already adjusted for the  
19 length of an occupational year (240 days) to the length of a calendar year (365 days), only the  
20 standard adjustment for inhaled air volume was performed. The concentration estimate obtained  
21 using this regression modeling and the life-table analysis procedure was  $LEC_{01} = 0.272$  fiber/cc,  
22 resulting in the lung cancer unit risk of 0.0368 per fiber/cc.

23 The [Berman and Crump \(2008\)](#) reanalysis was based on the [Sullivan \(2007\)](#) summary  
24 results except the authors used a lag of 10 years ([personal communication with Sullivan in 2008](#)  
25 [as cited by Berman and Crump \[2008\]](#)). They fit the IRIS IUR ([U.S. EPA, 1988a](#)) lung cancer  
26 model to aggregate data using an extra multiplicative parameter  $\alpha$ . In this model, the relative  
27 risk at zero exposure is  $\alpha$  rather than 1 (unity). With  $\alpha = 1$ , their model did not fit, and with  $\alpha$   
28 estimated, the fit was satisfactory. [Berman and Crump \(2008\)](#) chose the central estimate of the  
29 slope from the fit with  $\alpha$  estimated, but constructed an “informal” 90% confidence interval by the  
30 union of two confidence intervals (this upper bound is shown in Table 5-54). This was done to  
31 address uncertainty in the estimated parameter  $\alpha$ , similar to what is done in this current  
32 assessment with estimated lag and decay. Note also that [Berman and Crump \(2008\)](#) provided a  
33 UF to adjust for several sources of uncertainty in exposures, resulting in an upper-bound risk of  
34 0.3162.

35 The second alternative analytic approach to estimating the extra risk of lung cancer from  
36 a Cox regression with time-dependent covariates of individual mortality data was to use the  
37 results published by [Larson et al. \(2010b\)](#), with cumulative exposure and a 20-year lag. This  
38 analysis of lung cancer mortality was based on the full cohort of 1,862 workers, updated until

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2006 and using the same model form as the current EPA analysis (the extended Cox proportional hazards model). [Larson et al. \(2010b\)](#) reported a regression coefficient of  $1.06 \times 10^{-3}$  per fibers/cc-yr (SE =  $3.1 \times 10^{-4}$  per fibers/cc-yr,  $p = 0.0006$ ).<sup>30</sup> EPA assumed that the cumulative exposures reported by [Larson et al. \(2010b\)](#) were based on years of occupational exposure (240 days per year) during a 365-day calendar year. In order to account for exposure on every day of the year for a calculation of unit risk, an adjustment for exposures during the length of an occupational year (240 days) to the length of an calendar year (365 days) and an adjustment for the volume of inhaled air were performed to match EPA's analyses. The concentration estimate obtained using the [Larson et al. \(2010b\)](#) regression modeling and the life-table analysis procedure was  $LEC_{01} = 1.26$  fibers/cc, resulting in a lung cancer unit risk of 0.0103 per fiber/cc.

[Moolgavkar et al. \(2010\)](#) also used the Cox proportional hazards model with time-dependent covariates for analysis of the [Sullivan \(2007\)](#) cohort with a 15-year lag. The parameter in this study estimates  $1.11 \times 10^{-3}$  per fibers/cc-yr (SE =  $2.5 \times 10^{-4}$  per fibers/cc-yr), which is very close to the [Larson et al. \(2010b\)](#) value, and therefore, the lung cancer unit risk based on their analysis would be very close to the [Larson et al. \(2010b\)](#) value. Comparison with [McDonald et al. \(2004\)](#) is difficult because in that article outcome is defined as respiratory cancer (ICD-9 160–165), which is more expansive than other researchers' definitions of the outcome as lung cancer, and their subcohort of 406 white men employed before 1963—a time period when exposure assessment was less reliable and more likely to include significant exposure-measurement error. Nonetheless, the parameter estimate resulting from the Poisson analysis by [McDonald et al. \(2004\)](#) was  $3.6 \times 10^{-3}$  per fibers/cc-yr.

The differences in the results in Table 5-54 appear to be mostly attributable to the time periods of analysis and various degrees of exposure measurement error corresponding to these time periods rather than the analytic approach. EPA based their analyses on the exposures that occurred after 1959, while the [Sullivan \(2007\)](#), [Larson et al. \(2010b\)](#), and [Moolgavkar et al. \(2010\)](#) analyses were based on the cohort including those hired before 1960, and [McDonald et al. \(2004\)](#) included only workers hired before 1964. The small discrepancy between observed lung cancer deaths between this current assessment and [Larson et al. \(2010b\)](#), described in Section 4.1.1.1, is unlikely to play a role in the difference among risk estimates. Moreover, for the subcohort hired after 1959, all deaths are included in the [Larson et al. \(2010b\)](#) lung cancer counting rules.

As explained in detail in the discussion on uncertainty in the exposure assessment (see Section 5.4.6), there were only several measurements from the 1950s and one from 1942, and most of the exposure estimation for the early years of the cohort's experience was based on estimates of the ratio of dust to fibers estimated in the late 1960s and extrapolated backwards in

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<sup>30</sup>Note that EPA results based on the subcohort hired after 1959 were from the same model form but based on the cumulative exposure with a 10-year lag and had a slope of  $5.81 \times 10^{-3}$  per fiber/cc-yr (SE =  $2.48 \times 10^{-3}$  per fiber/cc-yr,  $p = 0.018$ ).

1 time for several decades. Moreover, 706 of the workers hired before 1960 (not necessarily short-  
2 term workers) did not have an exposure measurement assigned to them at all, leading to a much  
3 larger measurement error. These limitations in the underlying exposure assessment for the years  
4 before 1968 likely resulted in exposure measurement error that could have attenuated the  
5 analytic regression results (see the discussion in Lenters et al. ([2012](#); [2011](#)) on the impact the  
6 quality of the exposure information has on estimates of dose-response; also Bateson and Kopylev  
7 ([2014](#))), thereby yielding a smaller effect estimate for the whole cohort compared to the  
8 subcohort hired after 1959.

9 None of the approaches used by [McDonald et al. \(2004\)](#), [Sullivan \(2007\)](#), or [Larson et al.](#)  
10 [\(2010b\)](#) could have been appropriately used for the unit risk of mesothelioma because these  
11 approaches are not based on absolute risk metrics of association, which the current assessment  
12 considered to be the relevant metric of association. [Berman and Crump \(2008\)](#) did not evaluate  
13 the risk of mesothelioma. [Moolgavkar et al. \(2010\)](#) used an absolute risk model for  
14 mesothelioma. These results are summarized in Table 5-55. The upper-bound results for the full  
15 cohort presented by [Moolgavkar et al. \(2010\)](#) are about 80% of the [U.S. EPA \(1988a\)](#) estimate of  
16 the mesothelioma slope factor, leading to an approximately 80% estimate of the mesothelioma  
17 unit risk as dependence is linear in the mesothelioma slope factor (see eq 5-11). This is very  
18 close to the current assessment's estimate based on the subcohort, which is also about 80% of the  
19 [U.S. EPA \(1988a\)](#) estimate of mesothelioma risk. Duration of exposure, but neither department  
20 code nor job category, was known for 706 of 991 (71%) workers hired from 1935 to 1959.  
21 Because of that limitation, duration of employment is the best metric for the full cohort, and it  
22 does not support exposure-response estimation.

**Table 5-55. Mesothelioma analysis results from different analyses of cumulative exposure in the Libby workers cohort.** All analyses used NIOSH-collected exposure data but different cohort definitions, lengths of follow-up, and lengths of exposure lags to account for cancer latency.

Mesothelioma analysis	Cohort definition	Follow-up	Mesothelioma cases/N	Mesothelioma risk (absolute risk model) (per fiber/cc)
This current assessment	Hired post-1959 Exposures 1960–1982	2006	7/880	Upper Bound = 0.122 Central = 0.075
<a href="#">Sullivan (2007)</a>	Still employed post-1959 White males Exposures 1960–1982	2001	15/1,672	No estimates of absolute risk
<a href="#">Moolgavkar et al. (2010)<sup>a</sup></a>	Still employed post-1959 White males Exposures 1960–1982	2001	15/1,662	Upper Bound $\approx$ 0.13 Central $\approx$ 0.08
<a href="#">Larson et al. (2010b)</a>	Full cohort Exposures 1935–1993	2006	19/1,862	No estimates of absolute risk
<a href="#">Berman and Crump (2008)<sup>a</sup></a>	Still employed post-1959 White males Exposures 1960–1982	2001	15/1,672	No estimates provided

<sup>a</sup>Reanalysis of [Sullivan \(2007\)](#).

#### 5.4.6. Uncertainties in the Cancer Risk Values

Uncertainties in the derivation of the IUR are important to consider. This assessment does not involve extrapolation from high doses in animals to low doses in humans. It is based on a well-documented and well-studied cohort of workers with adequate years of follow-up to evaluate mesothelioma and lung cancer mortality risks with PODs within the range of the data. The discussions below explore uncertainty in the derivation of the IUR to provide a comprehensive and transparent context for the resulting cancer mortality risk estimates.

##### 5.4.6.1. Sources of Uncertainty

Sources of uncertainty in this assessment include:

- 1) *Uncertainty in low-dose extrapolation,*
- 2) *Uncertainty in exposure assessment, including analytical measurements uncertainty,*
- 3) *Uncertainty in model form,*
- 4) *Uncertainty in selection of exposure metric,*
- 5) *Uncertainty in assessing mortality corresponding to other cancer endpoints,*

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- 6) *Uncertainty in control of potential confounding in modeling lung cancer mortality,*
- 7) *Uncertainty due to potential effect modification,*
- 8) *Uncertainty due to length of follow-up,*
- 9) *Uncertainty in use of life tables to calculate cancer mortality inhalation unit risks,*
- 10) *Uncertainty in combining of risks to derive a composite cancer inhalation unit risk (IUR), and*
- 11) *Uncertainty in extrapolation of findings in adults to children.*

**5.4.6.1.1. Uncertainty in low-dose extrapolation.** A common source of uncertainty in quantitative cancer risk assessments generally derives from extrapolating from high doses in animals to low doses in humans. Compared to assessments based on animal data, the uncertainty from low-dose extrapolation in this assessment, which uses occupational epidemiology data, is considered to be lower for the following reasons. The NIOSH worker cohort developed by [Sullivan \(2007\)](#) includes 410 workers employed less than 1 year among the 880 workers hired on or after January 1, 1960. Although short-term workers on average experience a mean exposure intensity per day worked greater than workers employed more than a year ([Sullivan, 2007](#)), the cohort nevertheless includes many short-term workers with relatively low cumulative occupational exposures. Further, inclusion of salaried workers in the NIOSH cohort ([Sullivan, 2007](#)) adds many workers with lower workplace exposure. Thus, while occupational exposure concentrations may be generally higher than typical ongoing environmental concentrations, the low-dose exposures in this occupational database may be more representative of nonoccupational exposures.

While many occupational epidemiology studies are based on relatively high exposure levels that are beyond the range of common environmental exposures, many in the Libby worker cohort experienced exposures that were near or below the PODs derived from the life-table analysis (i.e., the estimated PODs are in the range of the observed data). The POD for the selected lung cancer mortality exposure metric was 0.191 fiber/cc. The POD for the selected mesothelioma mortality exposure metric was 0.106 fiber/cc. Among the workers hired after 1959 who had at least 1 year of occupational exposure ( $n = 470$ ; 20 lung cancer deaths), there were 19 (4%) with average occupational exposure concentrations of less than 0.3 fiber/cc, including one lung cancer death (5%).

Although data might have been modeled down to a very low cumulative exposure level, the *Guidelines for Carcinogen Risk Assessment* ([U.S. EPA, 2005a](#)) recommends defining a POD for low-dose extrapolation to increase the stability of the IUR estimate at lower exposures where fewer cancers might be expected. Thus, the uncertainty associated with low-dose extrapolation is somewhat mitigated because the linear extrapolation from the dose associated with the POD

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from the life-table analyses of each cancer endpoint was encompassed within the observed data range. Nonetheless, some uncertainty remains in the extrapolation from occupational exposures to lower environmental exposures when using a POD.

**5.4.6.1.2. *Uncertainty in exposure assessment.*** Accurate exposure assessment is generally considered to be a major challenge for occupational epidemiologic studies and is a challenge well recognized by the NIOSH investigators ([Amandus et al., 1987b](#)). As stated previously in Section 5.4.3.3, while it is generally true that the use of more data is an advantage in statistical analyses because it allows for the computation of more statistically precise effect estimates, this advantage in precision may be offset by a negative impact on the accuracy of the effect estimate if an increase in sample size is accompanied by greater exposure misclassification or other biases. In this case, EPA decided to base this LAA-specific human health risk assessment upon the mortality experience of workers hired on or after January 1, 1960. EPA's use of the subcohort analysis is based on the belief that it is important to accurately estimate the true underlying exposure-response relationships by relying on the most accurate exposure data. The use of this subcohort greatly reduces the uncertainty in exposure error compared to evaluations based on the full cohort. More specifically:

- a) Job category and department codes were completely unknown for 706 of the 991 workers' jobs from 1935 to 1959 (71% of the cohort for this time period). These workers were assigned by [Sullivan \(2007\)](#) the same exposure concentration (66.5 fibers/cc) for all years without this information. Examination of the post-1959 cohort removes this significant source of exposure misclassification (only 9 of 880 subcohort workers did not have department code and job category information).
- b) Using the more recently hired cohort minimizes the uncertainty in estimated worker exposures based on the JEM, which was informed by air sampling data available in 1956 and later years. Although uncertainties still exist in the task-specific exposure estimates from 1960–1967, uncertainty in the assessment of earlier exposure levels is considerably greater.
- c) Exposure measurements were collected from the area samples and represented exposures for all the workers with the same job code. Statistically, this causes the Berkson measurement error effect, which is described later in this section.

As EPA exposure-response modeling for mesothelioma and lung cancer mortality is based on the post-1959 subcohort, the remaining discussion of uncertainty in exposure measurement will address these data.

**5.4.6.1.2.1. *Sources of uncertainty in job history information.*** Worker exposures for EPA exposure-response modeling were calculated based on job histories and the JEM from 1960 through 1982 (see Figure 5-5). Overall, there is little uncertainty in the job history information.

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Regarding exposure estimation for the occupational cohort, the NIOSH investigators ([Amandus et al., 1987b](#)) conducted a detailed retrospective exposure assessment to estimate the individual worker exposures. NIOSH used extensive occupational exposure data to construct the time-specific JEM, spanning decades ([Amandus et al., 1987b](#)). These data were reabstracted from the workers' employment records for quality assurance ([Sullivan, 2007](#)). NIOSH records on work histories and job-specific exposure extended from the 1930s through May 1982. However, the vermiculite mining and milling operation continued on for several years, and some workers were retained through 1993 for plant close-out activities. Only 148 members of the post-1959 cohort ( $n = 880$ ) were employed as of the May 1982 employment records when the cohort was enumerated by NIOSH ([Sullivan, 2007](#)). Because exposure concentrations in 1982 (see Table 5-21) were generally below 1 fiber/cc with only two locations having concentrations of 1.2 fibers/cc, it is unlikely that these workers' exposures were significantly underestimated.

#### *Sources of uncertainty in exposure intensity for the identified location operations*

The available exposure data that inform the JEM include over 4,000 air samples, the majority of which were collected after 1967 (see Table 4-1). All of the job location exposure estimates (see Table 5-21) from 1968–1982 were directly informed from air samples collected on membrane filters and analyzed for fibers by PCM. The availability of site-and task-specific air samples for these years provides a good basis for the exposure estimates. However, some uncertainties exist in estimating asbestos exposures using air samples analyzed by PCM.

- 1) PCM analysis does not determine the mineral or chemical make-up of the fiber: The PCM method defines and counts fibers based on the size (length, width, and aspect ratio) of the particle without regard for the material that makes up the particle being viewed. The PCM method was developed for use in occupational environments where asbestos was present, and the nature of the fibers should be further evaluated to confirm the fibers viewed under PCM are asbestos. McGill University researchers evaluated the fibers collected on membrane filters in the early 1980s and confirmed the presence of asbestos fibers in the tremolite-actinolite solution series consistent with the LAA ([McDonald et al., 1986a](#)). NIOSH researchers confirmed the presence of tremolite asbestos in bulk dust samples but not in air samples from the facility ([Amandus et al., 1987b](#)). Although less specific to fibers, 60–80% of the airborne dust in the mills in 1968 was tremolite, further supporting the presence of asbestos in the air (based on State of Montana air sampling, and x-ray diffraction analysis by the Public Health Service [correspondence, October 17, 1968]). However, although the presence of mineral fibers in the actinolite-tremolite series was confirmed in the work environment, it is possible that fibers were also counted by PCM from other materials (such as textiles from clothes and packaging materials). Therefore, it is unknown from these data what proportion of the counted PCM fibers were mineralogically asbestos and what proportion were other materials present in the air workplace.

- 2) PCM defines fibers as particles with an aspect ratio of 3:1 or greater. There is an ongoing debate in the literature on asbestos toxicity regarding the influence of aspect

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ratio on relative toxicity. Specifically, in mining environments, it has been speculated that a larger proportion of low aspect ratio fibers from mineral dusts may significantly impact the apparent cancer potency of the measured PCM fibers in those environments ([Berman, 2010](#); [U.S. EPA, 1988a](#)). Few data are available to understand fiber morphology and fiber aspect ratios in the Libby cohort working environment. Considering the post-1959 cohort, PCM fiber size distribution and aspect ratio data only exist for a set of eight air samples (599 fibers) collected from the wet mill and screening operations and analyzed by the NIOSH researchers ([Amandus et al., 1987b](#)). For these air samples, over 96% of the fibers viewed by PCM had an aspect ratio greater than 10:1 (see Table 4-2) ([Amandus et al., 1987b](#)).<sup>31</sup> However, because these samples were provided by the company in the early 1980s, they do not represent conditions in the old wet mill or dry mill operations, which were significantly dustier environments ([Amandus et al., 1987b](#)). It is possible that prior to IH modifications in 1974, the dry and old wet mills generated proportionally more mineral dusts than the screening plant and new wet mill operations after IH modifications. No data are available for the mining environment, which would also be expected to generate a range of mineral dusts. Therefore, there is a significant uncertainty about the size and aspect ratio of fibers included in PCM fiber counts for the majority of the post-1960 workers cohort, but it is not possible to judge the direction or magnitude of such uncertainty.

- 3) The resolution of visible PCM fibers: Current analytical instruments used for PCM analysis have resulted in a standardization of minimum fiber width considered visible by PCM between 0.2 and 0.25  $\mu\text{m}$ . Historical PCM analysis (1960s and early 1970s) generally had less resolution, and fibers with minimum widths of 0.4 or 0.44  $\mu\text{m}$  were considered visible by PCM ([Amandus et al., 1987b](#); [Rendall and Skikne, 1980](#)). [McDonald et al. \(1986a\)](#) compared fibers viewed by PCM and TEM and estimated that approximately one-third of the total fibers could be viewed by PCM. Because 38% of the fibers were  $<5 \mu\text{m}$  in length, this implies approximately 30% were not viewable by optical microscopy for other reasons, such as width. However, it is unknown what proportion of that 30% would be viewed with the minimum width resolution of 0.25  $\mu\text{m}$  for later optical microscopy. It is likely that early PCM counts were underestimated relative to the later data for the cohort but by less than a factor of 2.

Before 1968, no air sampling data were available for 23 of the 25 job location operations (see Table 4-2), and the exposure estimates were extrapolated from later air sampling data. [Amandus et al. \(1987b\)](#) recognized there was significant uncertainty in the extrapolation of available air sampling data to previous time periods. The researchers considered major changes in operations and interviewed employees in the early 1980s regarding previous years of operation. The assumptions used to make these extrapolations are clearly stated for each of the plant operations. For four operations, high and low estimates of pre-1968 exposures were

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<sup>31</sup>Although [Amandus et al. \(1987b\)](#) report the sizing of PCM fibers, the details of the methodology are not given regarding how these fibers were identified. No method is cited, and it is unclear if the sizing was done by PCM or TEM for fibers in the reported size categories.

provided based on different sets of exposure assumptions (see Table 5-21). For ore loading, there were negligible differences in the exposure estimates for the period from 1960–1967 (10.7 versus 9 fibers/cc). For drilling, the river dock, and the bagging plant, there were 3.4-, 2.6-, and 2.8-fold differences, respectively, between the high and low estimates of exposure between 1960 and 1968.

Dry mill exposures between 1960 and 1968 were informed by air sampling for total dust collected in the dry mill facility from 1956–1969 (where total dust was collected by midge impingers). [Amandus et al. \(1987b\)](#) derived a conversion factor of 4.0 fibers/cc per mppcf to apply to the two location operations in the dry mill during these years. A range of conversion factors was considered for the dry mill depending on how the dust and fiber air samples (PCM) were grouped and averaged (1.2 to 11.5 fibers/cc per mppcf). A subset of dust and fiber samples available over the same time period (1967–1968) resulted in a ratio of 8.0 fibers/cc per mppcf. In contrast, a ratio of 1.9 fibers/cc resulted when total dust samples from 1969 were compared with fiber samples from 1970. However, both of these subsets had limited numbers of samples available. Therefore, the conversion factor of 4.0 fibers/cc per mppcf was selected based on using the maximum samples available over a time period when the dry mill exposures were considered similar: dust samples (1965–1969) and fiber samples (1967–1971).

**5.4.6.1.2.2. Sources of uncertainty in the calculation of the job-exposure matrix (JEM).** The exposures in the JEM (see Figure 5-5) were calculated from the exposure intensities of the various task-specific exposure intensities shown by job location operation (see Table 5-21). The uncertainties in the exposure intensity for the job location operations will impact the JEM. Additionally, for each of the job categories in the JEM, NIOSH researchers defined which tasks (job location operations) were conducted and for what proportion of the work day. A TWA exposure for each job category across time was calculated based upon these assumptions and the task-specific exposure estimates. There is a measure of uncertainty in these assumptions for each job category. Additionally, there is interindividual variation within the job categories. These uncertainties are common to exposure reconstruction for epidemiological cohorts.

**5.4.6.1.2.3. Uncertainty in the exposure metric.** The PCM measurement is the available exposure metric for analysis of the Libby worker cohort at this time. Currently, there is no optimal choice of the best dose metric for asbestos, in general, and for LAA, in particular. Uncertainties related to PCM analytical method are discussed in Section 2. Briefly, PCM cannot distinguish between asbestos and nonasbestos material or differentiate among specific types of asbestos. Further, due to limitations of this methodology, PCM does not take into account fibers shorter than 5  $\mu\text{m}$  in length.

**5.4.6.1.2.4. Evaluation of the effects of uncertainties in exposure measurement.** An understanding of the effects of exposure measurement error on the risks estimated from epidemiologic analyses is important to place these possible exposure measurement errors in context. The effect of exposure measurement error on estimates of the risk of mesothelioma or lung cancer mortality attributable to exposure depends upon the degree to which that error may be related to the likelihood of the outcome of interest (mesothelioma or lung cancer mortality). Exposure measurement error that is similar among workers who died of lung cancer, and those who did not die of lung cancer, is termed nondifferential exposure measurement error. Exposure measurement error that is associated with the outcome (error that is differential with respect to disease status) can cause bias in an effect estimate towards or away from the null, while nondifferential exposure error typically results in bias towards the null ([Rothman and Greenland, 1998](#)). From the above evaluation of uncertainties, there is no indication that the uncertainties in job history information, exposure estimates for specific tasks, or calculation of the JEM would be differential based on the cancer health outcome data. Therefore, these uncertainties are considered unrelated to disease status and the general result is likely to be an attenuation in risk estimates towards the null (i.e., the addition of random noise to a clear signal tends to reduce the clarity of the observed signal, and the avoidance of random noise results in a stronger observed signal).

Generally speaking, if the exposure concentrations estimated by NIOSH were systematically too high, then the associated risks of exposure estimated in the regression analysis would be low because the same actual risk would be spread across a larger magnitude of exposure. Similarly, if the exposure concentrations estimated by NIOSH were systematically too low, then the associated risks of exposure estimated in the regression analysis would be too high. From the above evaluation, the majority of the sources of uncertainty are not systematic. There are a few areas of uncertainty that may be classified as biased:

- 1) High- and low-exposure estimates for four job location operations were provided between 1960 and 1967. [Amandus et al. \(1987b\)](#) chose the high estimates of exposure for these job location operations when calculating the JEM. Therefore, there will be a bias towards the high end for the job categories informed by these data. There was a 1.1- to 3.4-fold difference between the high and low estimates. This difference will be less pronounced where these exposure concentrations are averaged with other job location operations in the JEM, and across multiple jobs (as was the case for the majority of the workers; see Figure 5-5).
- 2) Current PCM analysis would count more fibers relative to early PCM methods based on minimum fiber width resolution. For example, [Amandus et al. \(1987b\)](#) used a minimum width cutoff of 0.44  $\mu\text{m}$  in their review of PCM fibers in the 1980s, which may have resulted in as much as a twofold underestimate compared to current PCM methods with a width resolution of 0.25  $\mu\text{m}$  or less. Additionally, as PCM methodology has developed over time, it is unknown when PCM results from

company records would be considered relatively standard to a minimum width resolution between 0.2 and 0.25  $\mu\text{m}$ . Also, prior to standardization of PCM to 0.25- $\mu\text{m}$  minimum width, there was interlaboratory variability as well. Therefore, the size distribution of PCM fibers (e.g., minimum width) reported in the JEM may have changed over time. Although theoretically a systematic bias, given the years for which PCM data are available, this is likely an insignificant effect.

- 3) Asbestos was a contaminant of vermiculite that was the primary object of production. Mine, old dry mill, and wet mill ambient air may have contained material other than asbestos that could have contributed to PCM fiber count. The exposures in the old dry and wet mills and mine location may have included a greater proportion of dust to fibers than tasks using the ore and refined vermiculite after the new wet mill became operational. It is possible there is a systematic overcount of fibers in the dusty environment due to interference from nonmineral fragments. This likely affects the exposure intensity for 23 of 25 job location operations within the mine and old dry mill. Estimated exposures from job categories that include these operations may be biased upwards.

Nondifferential measurement error in a continuous exposure can be of the classical or Berkson type and typically arises in environmental and occupational settings as a mixture of the two forms ([Zeger et al., 2000](#)). Classical measurement error occurs when true exposures are measured with additive error ([Carroll et al., 2006](#)) and the average of many replicate measurements, conditional on the true value, equals the true exposure ([Armstrong, 1998](#)). This error is statistically independent of the true exposure being measured and attenuates true linear effects of exposure, resulting in effect estimates in epidemiologic studies that are biased towards the null ([Heid et al., 2004](#); [Zeger et al., 2000](#); [Armstrong, 1998](#)). Such errors occur, for example, when the mean values of multiple local air samples are used.

Berkson measurement error is independent of the surrogate measure of exposure ([Heid et al., 2004](#); [Berkson, 1950](#)) and is present when the average of individuals' true exposures, conditional on the assigned measurement, equals the assigned measurement. Berkson measurement error can arise from the use of local area mean sampled exposures to represent the individual exposures of people in that area—even when the estimated area mean is equal to the true underlying mean (i.e., no classical measurement error). Examples of random variability in personal behavior that may produce Berkson measurement error in personal exposure estimates include the volume of air breathed per day among the workers and the effectiveness of an individual's nasal filtration at removing contaminants. In general, Berkson measurement error is not thought to bias effect estimates but rather increases their standard errors ([Zeger et al., 2000](#)). However, some epidemiologic studies have suggested that Berkson measurement error can produce a quantitatively small bias towards the null in some analyses ([Bateson and Wright, 2010](#); [Kim et al., 2006](#); [Reeves et al., 1998](#); [Burr, 1988](#)). Uncertainties in the levels and time course of asbestos exposure for the Libby workers also adds uncertainty in evaluating the relative fit of different exposure metrics.

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1 **5.4.6.1.2.5. Exposure to other kinds of asbestos and residential exposure.** Another source of  
2 uncertainty in the estimation of exposures in the Libby workers cohort is the potential  
3 contribution of nonoccupational or residential exposures as well as exposures to other kinds of  
4 asbestos in employment before or after working in Libby.

5 Many of the workers resided in Libby, MT before and/or after their employment at the  
6 mining and milling facilities ended. The vermiculite from the mine had been used at numerous  
7 sites around the town including baseball fields around the expansion plant, high- and  
8 middle-school tracks, attic and wall insulation in homes, and as a soil amendment in gardens  
9 ([U.S. EPA, 2010a, 2001](#)). Exposure to asbestos could have occurred among individuals outside  
10 of the workplace, particularly through activities with the potential of stirring up soil or other  
11 materials that had been mixed with the vermiculite ([Weis, 2001](#)). The results of community  
12 sampling indicated that even 10 years after mill operations ceased, asbestos fiber concentrations  
13 in the air could exceed OSHA standards established for the protection of workers during some  
14 activities ([Weis, 2001](#)).

15 Therefore, the workers' actual personal exposures as the sum of occupational and  
16 nonoccupational exposures are likely to have been underestimated by the use of estimated  
17 Libby-related occupational exposure alone. The difficulty stems from the lack of data on  
18 residential exposures and lack of information on pre- and postemployment residence of the  
19 Libby workers. Nonoccupational exposures were likely to have been smaller in magnitude than  
20 the occupational exposures, but workers may have lived in and around Libby, MT for many  
21 more years than they were exposed occupationally. The effect of residential exposure could be  
22 more prominent for workers with lower occupational exposure who resided in Libby for a long  
23 time. [Whitehouse et al. \(2008\)](#) has reported several cases of mesothelioma among residents of  
24 the Libby, MT region who were not occupationally exposed. However, because the report by  
25 [Whitehouse et al. \(2008\)](#) details only the cases and does not define or enumerate the population  
26 from which those cases were derived, computed relative risks from nonoccupational exposures  
27 were not available. [ATSDR \(2000\)](#) reported higher relative risks of mesothelioma among the  
28 population of Libby, MT, including former workers residing in Libby, but did not provide  
29 relative risk for nonoccupational exposure. Instead, the ATSDR report on mortality [ATSDR](#)  
30 [\(2000\)](#) grouped cases among the former workers with nonoccupationally exposed cases.  
31 Therefore, it is not clear what the magnitude of the contribution of workers' nonoccupational  
32 exposures was to their overall risk.

33 Some of the occupational workers with lower exposures, such as short-term workers, may  
34 have either been high school or college students working during the summer or may have been  
35 transient workers who may not have stayed for a long time in Libby. It is interesting to note that  
36 the lung cancer rates by age at first exposure show very low rates for those first exposed before  
37 age 25 (see Table 5-42). [Sullivan \(2007\)](#) analyzed differences between short- and long-term  
38 workers and reported little difference among the groups except for age at hire. As the short-term

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workers were younger on average, this supported the suggestions that some of the short-term workers may have been college students working during the summer. This population of short-term workers is not well defined; however, while it is possible that short-term transient workers could potentially have been exposed to other kinds of asbestos or other lung carcinogens in their non-Libby occupational career, which might have affected their pre- and post-Libby risk profile for asbestos exposure, lung cancer rates for those with less than 1 year of exposure are in line with those with less than 5 years of exposure (see Table 5-27). While their occupational histories other than working in Libby are unknown, it is very unlikely that they include exposures of the magnitude that were encountered in the Libby mine and mill. The impact of these uncertainties on regression slopes is difficult to evaluate. However, the slope may be somewhat underestimated, as an observed increase in risk would be attributed to a larger exposure differential than might have been present due to the addition of nonoccupational exposures. There will also be a downward bias from random exposure measurement error with lower occupational exposure affected disproportionately; however, the magnitude of this bias would be expected to be small.

**5.4.6.1.2.6. Conclusion regarding uncertainty in exposure assessment.** Overall, there are likely to be multiple sources of uncertainty attributable to exposure measurement error. It is possible that systematic error may have been introduced into the exposure intensities assigned to several of the job location operations discussed above. In each case, these errors in estimating exposures were overestimates, which in general, might lead to underestimations of risk for lung cancer, but the results are unclear for the risk of mesothelioma. The magnitude of the potential overestimates of drilling and dry and old wet mill exposures is uncertain. The dust-to-fiber conversion ratio applied to the dry mill during 1960–1967 could be an over- or underestimate by as much as twofold, as [Amandus et al. \(1987b\)](#) derived a conversion factor of 4.0 fibers/cc per mppcf, but subsequent samples available during 1967–1968 resulted in a ratio of 8.0 fibers/cc per mppcf, while samples from 1970 yielded a ratio of 1.9 fibers/cc. Random error in the measurement of dust or fibers would likely have produced an underestimation of risk. There is no known bias in the assumptions to extrapolate exposure to pre-1968 location operations outside of the dry mill, and random bias would also likely have produced an underestimation of risk.

1 **5.4.6.1.3. Uncertainty in model form.** For mesothelioma mortality, the Poisson model is  
2 commonly used for rare outcomes and has been applied by [McDonald et al. \(2004\)](#) to model  
3 mesothelioma risk in the Libby worker cohort. For lung cancer mortality, the Cox proportional  
4 hazards model is a well-established method that is commonly used in cohort studies, including  
5 by [Larson et al. \(2010b\)](#) and [Moolgavkar et al. \(2010\)](#) for the Libby worker cohort, because this  
6 type of survival analysis takes into account differences in follow-up time among the cohort.  
7 [Larson et al. \(2010b\)](#) conducted Poisson analyses and reported that their lung cancer results  
8 using this different model form were similar to those from their extended Cox proportional  
9 hazards models, but those results were not shown.

10 Both of these model forms allow for the evaluation and control of important potential  
11 confounding factors such as age, gender, and race, and for the modeling of exposure as a  
12 continuous variable. Both model forms yielded exposure-response results with good fit to the  
13 occupational exposure data. The default assumption of the extended Cox proportional hazards  
14 model as well as the Poisson model is that all censoring (due to death or loss to follow-up) is  
15 assumed to be independent of exposure to the LAA. However, exposure to LAA may cause  
16 death from other causes such as asbestosis or nonmalignant respiratory disease ([Larson et al.,](#)  
17 [2010b](#)), which is referred to as dependent censoring. The concern is that the observation of lung  
18 cancer mortality may be precluded by mortality from other causes.

19 In the cohort of 880 workers hired after 1959, 32 died of lung cancer, while 10 died of  
20 asbestosis, and 21 died of nonmalignant respiratory disease. The mean length of follow-up from  
21 the date of hire until death for the workers who died of lung cancer was 24.9 years. However,  
22 the mean length of follow-up for the workers who died of asbestosis or nonmalignant respiratory  
23 disease was 30.4 years, so it does not appear that early deaths from other causes associated with  
24 exposure to the LAA ([Larson et al., 2010b](#)) would have precluded many cases of lung cancer.  
25 This implies that any potential bias in the lung cancer risk estimates due to dependent competing  
26 risks is small.

27 With respect to mesothelioma mortality, it should be noted that the exposure-response  
28 modeling is limited by the number of deaths. However, dependent censoring, as described  
29 above, is not accounted for in the Poisson model and likely causes a downward bias in the  
30 estimation of risk. The mean length of follow-up for the workers who died of mesothelioma was  
31 30.1 years, and there is some evidence that early deaths from other exposure-related causes  
32 precluded an individual's risk of death from mesothelioma; only lung cancer exhibited a shorter  
33 average follow-up time compared to mesothelioma, and in 419 cases of mesothelioma,  
34 mesothelioma and lung cancer were never coidentified ([Roggli and Vollmer, 2008](#)).  
35



1 **5.4.6.1.4. Uncertainty in selection of exposure metric.** There is uncertainty about what metric  
2 should be used for modeling exposure to LAA. This current assessment evaluated models  
3 proposed in the asbestos literature for modeling mesothelioma and models that include unlagged  
4 and lagged cumulative exposure with and without a half-life of various lengths, and RTW  
5 exposure with and without a half-life. In the analysis of comparative model fit based on the  
6 empirical data, lagged cumulative exposure with a half-life provided the best fits for both  
7 mesothelioma and lung cancer mortality associated with LAA. However, evaluation of 20-year  
8 lag and longer lag times for mesothelioma was not possible, as the earliest mesothelioma death  
9 happened less than 20 years from the start of the exposure; hence, exposure was zeroed out, and  
10 the fit of any model with 20-year lag was very poor. Latency time for mesothelioma may be as  
11 long as 60–70 years ([e.g., Bianchi and Bianchi, 2009](#)), so the precise lag time is uncertain.

12 In evaluating the data on lung fiber burden, [Berry et al. \(2009\)](#) estimated the range of the  
13 half-life for crocidolite to be between 5 and 10 years. That range is consistent with the finding of  
14 a 5- to 10-year half-life with 10–15 years lag that provided the best fit to the Libby workers  
15 cohort mesothelioma mortality data. Similarly, recent publications indicate that the relative risk  
16 of lung cancer due to asbestos exposure declines 15–20 years after the cessation of exposure to  
17 asbestos ([Magnani et al., 2008](#); [Hauptmann et al., 2002](#)). The marginally best fit for the Libby  
18 workers cohort lung cancer mortality data was for CE models with a 5- to 20-year half-life and  
19 10-year lag. However, the precise lag and half-life times are somewhat uncertain. Sensitivity  
20 analysis that excluded people with high exposure during 1960–1963 (see Section 5.4.3.6.4)  
21 provides further evidence that distinguishing between various lags and decays may be difficult  
22 with these data. A limitation of this sensitivity analysis is the decrease in the number of cases,  
23 especially for mesothelioma. Resolving this uncertainty would require longer follow-up time,  
24 which would allow for a subcohort analysis of workers hired in 1967 or afterwards (when  
25 exposure estimates began to be based on PCM measurements) until a sufficient number of cases  
26 would be available for additional analysis.

27 These simulated decay models were derived mathematically to approximate underlying  
28 biological processes that are not well understood, and are based on maximizing the likelihood for  
29 the workers cohort and may not necessarily apply to the environmental exposure patterns.  
30 Nonetheless, while the mode of action for carcinogenicity is unknown, the models incorporating  
31 a half-life in the exposure metric were clearly preferable for mesothelioma mortality, and the  
32 goal of the regression modeling effort was to identify the best fitting exposure model for the  
33 Libby worker cohort.

34 Table 5-53 illustrates uncertainty in the IUR due to exposure metric selection. The  
35 quantitative uncertainty is about threefold.

1 **5.4.6.1.5. *Uncertainty in assessing of mortality corresponding to other cancer endpoints.***

2 There is evidence that other cancer endpoints may also be associated with exposure to the  
3 commercial forms of asbestos. IARC concluded that there was sufficient evidence in humans  
4 that commercial asbestos (chrysotile, crocidolite, amosite, tremolite, actinolite, and  
5 anthophyllite) was causally associated with lung cancer and mesothelioma, as well as cancer of  
6 the larynx and the ovary ([Straif et al., 2009](#)). Among the entire Libby workers cohort, only  
7 two deaths were found to be due to laryngeal cancer, and there were no deaths from ovarian  
8 cancer among the 24 deaths of 84 female workers. The lack of sufficient number of workers to  
9 estimate risk of ovarian cancer is an uncertainty in an overall cancer health assessment.

10 The remaining uncertainties attributed to assessing mortality corresponding to the cancer  
11 endpoints are considered to be low.

12  
13 **5.4.6.1.6. *Uncertainty in control of potential confounding in modeling lung cancer mortality.***

14 It is well known that smoking is a strong independent risk factor for lung cancer and may have a  
15 synergistic effect with asbestos exposure ([Wraith and Mengersen, 2007](#)). In contrast, smoking is  
16 not considered a risk factor for mesothelioma ([Selikoff and Lee, 1978](#); [Anderson et al., 1976](#)).

17 As an important potential confounder of the lung cancer mortality analysis, the possible  
18 effect of smoking on the estimated risk of lung cancer mortality associated with exposure to  
19 LAA needs to be evaluated to the fullest extent possible. This consideration was discussed in  
20 [Amandus and Wheeler \(1987\)](#) and in Section 4.1.1.3.

21 Additionally, W.R. Grace and Co. instituted a smoking ban on the property in 1979  
22 ([Peacock, 2003](#)). Information is not available as to the effect of this smoking ban at work on  
23 smoking patterns outside of the work environment. About 30% of the subcohort was still  
24 employed in 1979 and all of the post-1959 cohort had been terminated by May 1982, so the  
25 effect of a workplace smoking ban on cohort smoking history may explain the higher proportion  
26 of former smokers in the [Amandus and Wheeler \(1987\)](#) data. Lung cancer risks in ex-smokers  
27 decrease over time compared to lung cancer risks in continued smokers. A reduction of smoking  
28 in the Libby worker population may lead to fewer observations of lung cancer deaths in later  
29 years of the cohort study than would have occurred in the absence of the smoking restrictions.  
30 Changes in smoking behavior during the course of the epidemiological observation period would  
31 lead to changes in the observed time course of lung cancer death rates. This issue is related to  
32 potential effect modification of lung cancer mortality described in Section 5.4.6.1.7.

33 Without high-quality individual-level data on smoking that could be used to control for  
34 potential confounding, it is still possible to comment upon the likelihood and potential magnitude  
35 of confounding and the impact any confounding would be expected to have on the lung cancer  
36 mortality risk estimates. Confounding can be controlled for in a number of ways including by  
37 modeling and by restriction. Restriction of the study population can reduce any potential  
38 confounding by making the resulting population more similar. For instance, there can be no

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1 confounding by gender when a study population is restricted to only men. This assessment  
2 restricted the study population to those workers hired after 1959. Smoking habits have changed  
3 over time, and it can reasonably be assumed that the range of smoking habits among those hired  
4 after 1959 is less variable than that among the whole cohort, particularly because of the narrower  
5 range of birth cohorts represented in this subcohort. This should have the effect of reducing  
6 some of the potential for confounding. Analytic examinations of potential confounding are  
7 discussed below.

8 [Richardson \(2010\)](#) describes a method to determine if an identified exposure relationship  
9 with lung cancer is confounded by unmeasured smoking in an occupational cohort study. EPA  
10 implemented this methodology to model the potential effects of LAA on the risk of COPD  
11 mortality on the subcohort of workers hired after 1959 (see Section 5.4.3.8). Summarizing these  
12 findings, EPA used the method described by [Richardson \(2010\)](#) to evaluate whether exposures to  
13 LAA predicted mortality from COPD as an indication of potential confounding by smoking and  
14 found a statistically nonsignificant negative relationship, which was inconsistent with  
15 confounding by smoking.

16  
17 **5.4.6.1.7. *Uncertainty due to potential effect modification.*** Among the 32 deaths from lung  
18 cancer in workers hired after 1959 that were used to estimate the unit risk of lung cancer  
19 mortality (see Section 5.4.5.2), data on smoking listed 16 as smokers, 4 as former smokers, and  
20 12 of the 32 had missing data. Thus, data to support an estimate of the risk of LAA among  
21 known nonsmokers were not available.

22 It is theoretically possible that the risk of lung cancer mortality estimated in this current  
23 assessment is a reflection of a positive synergy between smoking and asbestos, and that the  
24 adverse effect of LAA among the potentially nonsmoking workers has been overestimated. The  
25 unit risk of the lung cancer estimate herein and the combined mesothelioma and lung cancer  
26 mortality IUR would then be health protective for any population that had a lower prevalence of  
27 smoking than that of the Libby worker cohort. However, if the smoking ban did diminish the  
28 effect of smoking, then any overestimation would be somewhat mitigated.

29  
30 **5.4.6.1.8. *Uncertainty due to length of follow-up.*** There is some potential uncertainty regarding  
31 the length of follow-up for cancer mortality, even more so with the restriction of the cohort to  
32 those workers hired after 1959. The hire dates among this subset of the cohort ranged from  
33 January 1960 to November 1981 (the mean date of hire was May 1971). Follow-up continued  
34 until the date of death or December 31, 2006, whichever occurred first. Therefore, the range of  
35 follow-up was from 25 to 46 years, with a mean of more than 35 years.

36 However, for mesothelioma mortality, the length of the latency period is considerably  
37 longer. [Suzuki and Yuen \(2001\)](#) reviewed 1,517 mesothelioma cases from 1975 through 2000  
38 and was able to estimate the latency for 800. [Suzuki and Yuen \(2001\)](#) reported 17% of cases had

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1 a latency of less than 30 years with 52% of cases with a latency of less than 40 years. [Bianchi](#)  
2 [and Bianchi \(2009\)](#) estimated the mesothelioma latency in 552 cases and reported mean latency  
3 periods of 35 years among insulators, 46 years among various industries, and 49 years among  
4 shipyard workers.

5 According to the results of [Suzuki and Yuen \(2001\)](#) and of [Bianchi and Bianchi \(2009\)](#) a  
6 mean length of follow-up of 35 years may only have captured half of all eventual mesothelioma  
7 mortality cases among the Libby workers hired after 1959. If this were so, then the unit risk of  
8 mesothelioma mortality could be larger than was estimated from existing data, suggesting  
9 continued examination.

10  
11 **5.4.6.1.9. Uncertainty in use of life tables to calculate cancer mortality inhalation unit risk**  
12 **(IUR).** The life-table procedure computes the extra risk of death from birth up to 85 years of  
13 age, in part, because this is how national cancer incidence and mortality rate data that are one  
14 basis of the life tables are made available ([see SEER, 2010, Table 15.10, age-specific U.S. death](#)  
15 [rates](#)). Because the prevalence of cancer mortality is a function of increasing age, this cut-off at  
16 age 85 ignores a small additional risk of lung cancer mortality among a small percentage of  
17 people who have the higher background risk. This has the effect of slightly underestimating the  
18 IUR that would be derived if the life table were extended for an additional period of time,  
19 accounting for longer life spans. Extension of the life-table analysis to people over the age of  
20 85 requires an additional assumption. Assuming that having attained the age of 85 years, the  
21 additional life expectancy is 5 years, then the lung cancer mortality unit risk based on the LEC<sub>01</sub>  
22 would be somewhat larger—on the order of 5–10%—slightly more than the additional  
23 mesothelioma mortality risk if the life tables were extended.

24  
25 **5.4.6.1.10. Uncertainty in combining of risk for composite cancer inhalation unit risk (IUR).**  
26 For the purpose of combining risks, it is assumed that the unit risks of mesothelioma and lung  
27 cancer mortality are normally distributed. Because risks were derived from a large  
28 epidemiological cohort, this is a reasonable assumption supported by the statistical theory. EPA  
29 conducted a bounding analysis and showed that the related uncertainty is very low.

30  
31 **5.4.6.1.11. Uncertainty in extrapolation of findings in adults to children.** The analysis of lung  
32 cancer mortality specifically tested the assumption that the relative risk of exposure is  
33 independent of age within the observed age range of the occupational subcohort hired after 1959  
34 and did not find evidence of age dependence, although such a dependence among a working-age  
35 study population has been reported in another asbestos-exposed cohort ([Richardson, 2009](#)).  
36 However, it should be noted that no comparable data are available to estimate the lifetime risk  
37 from early life exposures. Note that default age-dependent adjustment factors (ADAFs) are not  
38 recommended because a mutagenic MOA was not identified.

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## 5.4.6.2. Summary

Section 5.4.6.1 details the several sources of uncertainty in the assessment of the cancer exposure-response relationships and the use of those data to derive the inhalation unit risk. The text that follows summarizes the primary sources of uncertainty and, where possible, the expected direction of effect on the exposure-response risk estimates and the inhalation units risk.

### 1) *Uncertainty in low-dose extrapolation (see Section 5.4.6.1.1)*

- There remains some uncertainty in the extrapolation of risks based on occupational exposure to environmental exposure levels but this uncertainty is considered to be low as the lower range of occupational exposure overlaps with expected environmental exposure levels.

### 2) *Uncertainty in exposure assessment, including analytical measurements uncertainty (see Section 5.4.6.1.2)*

- The JEM was based on the “high” exposure estimate for each job according to [Amandus et al. \(1987b\)](#) and to some extent this could be an overestimate of exposure. The associated cancer risk would be somewhat underestimated resulting in a somewhat underestimated IUR.
- The JEM was largely based on estimated fiber concentration using PCM measurement (with some extrapolations in time), and because PCM may count all long and thin objects as fibers, these measurement could overestimate the true LAA fiber concentrations leading to an overestimate of exposure and a somewhat underestimated cancer risk resulting in a somewhat underestimated IUR.
- PCM measurements in the era of NIOSH measurements in Libby used a lower resolution, and therefore, included only somewhat thicker fibers thereby counting fewer fibers than would have been counted by later PCM standards. These earlier measurement could underestimate the true LAA fiber concentrations leading to an underestimate of exposure and an overestimate of cancer risk resulting in a somewhat overestimated IUR.
- The PCM measurement is the available exposure metric for analysis of the Libby worker cohort at the time of this assessment. Currently, there is no optimal choice of the best dose metric for asbestos, in general and in particular, for LAA. Uncertainties related to PCM analytical method are discussed in Section 2 and such uncertainties cannot be related to the IUR at the time of this assessment.
- Random measurement error in the assignment of exposures could have the effect of underestimating the risk of lung cancer mortality as that measure of risk is based on a relative measure. The effect would be to somewhat underestimate the risk of lung cancer resulting in a somewhat underestimated unit risk for lung cancer. It is unclear what the impact of such measurement error would be on the absolute risk of mesothelioma.

- Exposure to other kinds of asbestos and residential exposure to LAA may have caused workers' actual personal exposures (as the sum of occupational and nonoccupational exposures) to have been underestimated by the use of estimated Libby occupational exposure information alone. This could underestimate the true LAA fiber exposures leading to an overestimate of the associated cancer risk resulting in a somewhat overestimated IUR.

3) *Uncertainty in model form (see Section 5.6.4.1.3)*

- For mesothelioma, the Poisson model is the standard epidemiologic form and it is considered to be the most appropriate model form for rare health outcomes; therefore, uncertainty is considered to be low. For lung cancer mortality, the Cox proportional hazards model is the standard epidemiologic form. It is considered to be the most appropriate model form for health outcomes with time-varying exposure data, and thus, uncertainty is considered to be low.

4) *Uncertainty in selection of exposure metric (see Section 5.6.4.1.4)*

- There is uncertainty about what metric should be used for modeling exposures to LAA. Table 5-53 illustrates the uncertainty in the IUR due to exposure metric selection. The quantitative uncertainty is about threefold.

5) *Uncertainty in assessing mortality corresponding to other cancer endpoints (see Section 5.6.4.1.5)*

- The lack of sufficient numbers of workers to estimate the risk of other cancers potentially related to LAA exposure is an uncertainty of unclear direction but is considered to be low due to the rarity of those cancers.

6) *Uncertainty in control of potential confounding in modeling lung cancer mortality (see Section 5.4.6.1.6)*

- The uncertainty in control of potential confounding by smoking is considered to be low, as the described sensitivity analysis did not show evidence of potential confounding.

7) *Uncertainty due to potential effect modification (see Section 5.4.6.1.7)*

- Smoking was not considered to be related to LAA exposure, and therefore, smoking is not considered to be a likely effect modifier of cancer risk. Age has been shown to be a potential effect modifier of lung cancer risk but there was no evidence of this relationship in the subcohort.

8) *Uncertainty due to length of follow-up (see Section 5.4.6.1.8)*

- There is uncertainty related to the limited follow-up for cancer mortality, and it is possible that with subsequent mortality follow-up the IUR could change in a direction that is unknown.

1 9) *Uncertainty in the use of life tables to calculate cancer mortality IUR (see*  
2 *Section 5.4.6.1.9)*

- 3 • The life-table procedure computes the extra risk of death from birth to 85 years of  
4 age. If the life tables were extended from 85 to 90 years to account for longer life  
5 spans, the selected lung cancer mortality unit risk (Table 5-52 shows this as  
6 0.068) would be somewhat larger, about 5–10%, and the selected mesothelioma  
7 unit risk (Table 5-50 shows this as 0.122) would be slightly less (about 3%).  
8 Taking both effects into consideration, the uncertainty in the IUR is considered to  
9 be low.

10 10) *Uncertainty in combining of mortality risks to derive a composite cancer mortality*  
11 *inhalation unit risk (IUR) (see Section 5.4.6.1.10)*

- 12 • EPA assumed that the cancer risks were independent, conducted a bounding  
13 analysis and showed the related uncertainty to be very low.

14 11) *Uncertainty due to extrapolation of findings in adults to children (see*  
15 *Section 5.4.6.1.11)*

- 16 • There is uncertainty in the assumption that risks are independent of age and that  
17 children are at the same exposure-related risk as adults. The lack of published  
18 information on cancer risks associated with exposures during childhood remains an  
19 uncertainty of unclear magnitude.  
20



## 6. MAJOR CONCLUSIONS IN THE CHARACTERIZATION OF HAZARD AND EXPOSURE RESPONSE

Libby Amphibole asbestos (LAA),<sup>32</sup> present in vermiculite ore from the mine near Libby, MT, is a complex mixture of amphibole fibers—both mineralogically and morphologically (see Section 2.2). The mixture primarily includes winchite, richterite, tremolite, magnesio-riebeckite, magnesio-arfvedsonite, and edenite (84:11:6:1:1:1) amphibole minerals that exhibit a range of fiber morphologies (e.g., asbestiform, acicular, prismatic) ([Meeker et al., 2003](#)). Given the exposure potential to LAA—and its characteristic mineral composition—a hazard characterization and cancer exposure-response assessment are presented.

As discussed in Section 1, no RfC for asbestos currently exists, and the EPA IRIS IUR for asbestos is based on a synthesis of 14 epidemiologic studies that included occupational exposure to chrysotile, amosite, or mixed mineral fibers (chrysotile, amosite, and crocidolite) ([U.S. EPA, 1988a](#)). Some uncertainty exists in applying the resulting IUR to environments and minerals not included in the studies considered for the asbestos IUR derivation ([U.S. EPA, 1988a](#)). Published mortality studies on the Libby, MT worker cohort have become available since the derivation of the IRIS asbestos IUR (i.e., [Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). This assessment documents noncancer and cancer health effects from inhalation exposure to LAA. Data are not available to support derivation of either a RfD or a cancer oral slope factor (OSF) following oral exposures to LAA.

### 6.1. HUMAN HAZARD POTENTIAL

#### 6.1.1. Exposure

Several different groups of humans have the potential for exposure to fibers from vermiculite ore from the mine in Libby, MT, and hence the potential for exposure to the LAA associated with this material. These groups include not only the former workers at the mine and mill site, but also residents in the community of Libby, MT, as well as workers at other locations who processed the vermiculite product. When the mine in Libby, MT, was active, miners, mill workers, and those working in the processing plants were exposed to vermiculite ore, silica dust, and amphibole structures released to air from the ore during the mining and processing operations ([Meeker et al., 2003](#); [Amandus et al., 1987b](#); [McDonald et al., 1986a](#)). In some cases, workers may have inadvertently transported contaminated materials from the workplace to vehicles, homes, and other establishments, typically on the clothing, shoes, and hair. This

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<sup>32</sup>The term “Libby Amphibole asbestos” is used in this document to identify the mixture of amphibole mineral fibers of varying elemental composition (e.g., winchite, richterite, tremolite, etc.) that have been identified in the Rainy Creek complex near Libby, MT. It is further described in Section 2.2.

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1 transported material may have resulted in “take-home exposure” for the workers, their families,  
2 and other coresidents. The magnitude of these historic take-home exposures was not measured,  
3 so the levels at which individuals in the home might have been exposed are unknown.

4 While some vermiculite concentrate was exfoliated and used in Libby, MT, most of the  
5 concentrate was transported to expansion plants at other locations across the country where it  
6 was exfoliated and distributed. A review of company records from 1964–1990 indicates that  
7 more than 6 million tons of vermiculite concentrate was shipped to over 200 facilities outside of  
8 Libby, MT ([ATSDR, 2008](#)). Because expanded vermiculite from Libby was widely used in  
9 numerous consumer and construction products in the United States, even people not associated  
10 with Libby or other communities with expansion plants may also have the potential for exposure  
11 to LAA (see Table 2-2). Vermiculite was most notably used as attic insulation ([VAL, Versar, 2003](#)),  
12 as a soil amendment for gardening, fireproofing agent, and in the manufacturing of  
13 gypsum wallboard. Other residents living in communities near the expansion plants may also  
14 have been subjected to some of the same exposure pathways as was the Libby community. The  
15 2008 ATSDR Summary Report observed that individuals in a community with a vermiculite  
16 expansion and processing plant could have been exposed by breathing airborne emissions from  
17 the facility or by inhalation exposure to contaminants brought into the home on workers’  
18 clothing or from outdoor sources ([ATSDR, 2008](#)).

### 19 20 **6.1.2. Fiber Toxicokinetics**

21 Although oral and dermal exposure to fibers does occur, inhalation is considered the main  
22 route of human exposure to mineral fibers, and therefore, has been the focus of more fiber  
23 toxicokinetic analyses in the literature. As with other forms of asbestos, exposure to LAA is  
24 presumed to be through all three routes of exposure; however, this assessment specifically  
25 focuses on the inhalation pathway of exposure. Generally, fiber deposition in the respiratory  
26 tract is fairly well defined based on fiber dimensions and density, although the same cannot be  
27 said for fiber translocation to extrapulmonary sites (e.g., pleura). The deposition location within  
28 the pulmonary and extrapulmonary tissues plays a role in the clearance of the fibers from the  
29 organism.

30 Fiber clearance from the respiratory tract can occur through physical and biological  
31 mechanisms. Limited mechanistic information is available on fiber clearance mechanisms in  
32 general, and no information specific to clearance of LAA fibers is available. Fibers have been  
33 observed in various pulmonary and extrapulmonary tissues following exposure, suggesting  
34 translocation occurs to a variety of tissues. Studies have also demonstrated that fibers may be  
35 cleared through physical mechanisms (coughing, sneezing) or through dissolution of fibers.

36 Multiple fiber characteristics (e.g., dimensions, density, and durability) play a role in the  
37 toxicokinetics and toxicity of fibers. The literature examining a variety of fiber determinants and  
38 their role in disease is extensive, with a focus on fiber length, width, and durability; however,

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these studies are often contradictory, making conclusions difficult for fibers in general. This is in part due to the variety of fibers analyzed, inadequate study design, and/or lack of information on fiber dimensions in earlier studies. However, due to the importance in understanding the role of these fiber determinants in the biological response, careful attention has been paid to these fiber characteristics when analyzing research studies on LAA and asbestiform tremolite, an amphibole fiber that comprises part of LAA (see Appendix D). No toxicokinetic data are currently available specific to LAA or its components (e.g., winchite, richterite, tremolite, magnesio-riebeckite, magnesio-arfvedsonite, and edenite). When available, fiber characteristic data are presented in the discussion of each study in relation to the toxic endpoints described.

### **6.1.3. Noncancer Health Effects in Humans and Laboratory Animals**

The predominant noncancer health effects observed following inhalation exposure to LAA are on the lungs and pleural lining surrounding the lungs. These effects have been observed primarily in studies of exposed workers and community members, and are supported by laboratory animal studies. Recent studies have also examined other noncancer health effects following exposure to Libby Amphibole, including autoimmune effects and cardiovascular disease; this research base is currently not as well developed as that of respiratory noncancer effects. Adequate data are not available to differentiate the health effects of the predominant mineralogical forms composing LAA. Although the adverse effects of tremolite are reported in the literature, the contribution of winchite and richterite to the aggregate effects of LAA has not been determined.

Noncancer health effects identified in humans following inhalation exposure to LAA include pleural abnormalities, asbestosis, and reduced lung function, as well as increased mortality from noncancer causes. Two cohorts of workers exposed to LAA have been studied: workers at the mine and related operations in Libby, MT and employees in the O.M. Scott plant in Marysville, OH, where the vermiculite ore was exfoliated and used as an inert carrier in lawn care products. Radiographic assessments of study participants in both cohorts indicate radiographic abnormalities consistent with asbestos-related disease, specifically pleural effects and small interstitial opacities (indicative of interstitial fibrosis) ([Rohs et al., 2008](#); [Amandus et al., 1987a](#); [McDonald et al., 1986b](#); [Lockey et al., 1984](#)). These studies provided quantitative exposure estimates and were considered suitable for exposure-response analysis to support an RfC derivation. Additionally, five cohort mortality studies of Libby, MT workers identified an increased risk of mortality from noncancer causes, including nonmalignant respiratory disease (e.g., asbestosis) ([Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)) and cardiovascular disease ([Larson et al., 2010b](#)).

ATSDR conducted health screening of community members in and around Libby, MT (including past workers), and identified an increase in radiographic abnormalities with an increased number of exposure pathways ([Peipins et al., 2004a](#); [Peipins et al., 2003](#); [ATSDR,](#)

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2001b). Other researchers have also used these data to identify the increased prevalence of respiratory symptoms in children (Vinikoor et al., 2010) and to evaluate the prevalence of radiographic abnormalities and reduced lung function in nonworker participants (Weill et al., 2011). Radiographic abnormalities were more prevalent in mine/mill workers versus other exposure categories (i.e., household contacts, ‘dusty trades’, and community-only exposures) (Weill et al., 2011). Prevalence of pleural effects increased with age and within each exposure group. Decreased pulmonary function (as percentage of the predicted forced vital capacity) is reported for participants with radiographic abnormalities (Weill et al., 2011). A nested case-control study based on ATSDR community health screening also identified a potential for increased prevalence of autoimmune disease (Noonan et al., 2006), and other experimental work has examined mechanistic steps relating to autoimmunity (Marchand et al., 2012; Pfau et al., 2005). Further development of this area of research could provide additional insights into the range of health effects possibly linked to LAA.

Laboratory animal and mechanistic studies of LAA are consistent with the noncancer health effects observed in workers exposed to LAA in Libby, MT and Marysville, OH, as well as exposed community members. Pleural fibrosis was increased in hamsters after intrapleural injections of LAA (Smith, 1978). More recent studies have demonstrated increased collagen deposition and inflammation consistent with fibrosis following intratracheal instillation of LAA fibers in mice and rats (Cyphert et al., 2012b; Cyphert et al., 2012a; Padilla-Carlin et al., 2011; Shannahan et al., 2011a; Shannahan et al., 2011b; Smartt et al., 2010; Putnam et al., 2008). Pulmonary fibrosis, inflammation, and granulomas were observed after tremolite inhalation exposure in Wistar rats (Bernstein et al., 2005; Bernstein et al., 2003) and intratracheal instillation in albino Swiss mice (Sahu et al., 1975). Davis et al. (1985) also reported pulmonary effects after inhalation exposure in Wistar rats, including increases in peribronchiolar fibrosis, alveolar wall thickening, and interstitial fibrosis.

Limited research is available on noncancer health effects occurring outside the respiratory system and pleura. Larson et al. (2010b) examined cardiovascular disease-related mortality in the cohort of exposed workers from Libby (see Section 4.1). Mechanistic studies have examined the potential role of iron and the associated inflammation for both respiratory and cardiovascular disease (Shannahan et al., 2012a; Shannahan et al., 2012c; Shannahan et al., 2012b; Shannahan et al., 2012d; Shannahan et al., 2011b). Recent studies in the Libby, MT community examined the association between asbestos exposure and autoimmune disease (Noonan et al., 2006) or autoantibodies and other immune markers (Pfau et al., 2005; see Table 4-16). Mechanistic studies examining the role of LAA exposure in autoimmune disease have shown limited effects but did observe an increase in autoantibodies in the serum of exposed animals (Salazar et al., 2013; Salazar et al., 2012). These results are supported by recent in vitro studies demonstrating increased autoantibodies to mesothelial cells, leading to collagen deposition (Serve et al., 2013). These recent studies have examined the association between

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asbestos exposure and autoimmune disease; additional research in this area could enhance understanding of this potential mode of action for noncancer effects ([Salazar et al., 2013](#); [Serve et al., 2013](#); [Rasmussen and Pfau, 2012](#); [Salazar et al., 2012](#); [Blake et al., 2008](#); [Pfau et al., 2008](#); [Hamilton et al., 2004](#)). However, limitations in the number, scope, and design of these studies make it difficult to reach conclusions as to the role of asbestos exposure in either cardiovascular disease or autoimmune disease.

Limited in vitro studies have demonstrated oxidative stress following LAA exposures in various cell types ([Duncan et al., 2014](#); [Duncan et al., 2010](#); [Hillegass et al., 2010](#); [Pietruska et al., 2010](#); [Blake et al., 2007](#)). LAA fibers increased intracellular ROS in both murine macrophages and human epithelial cells ([Duncan et al., 2010](#); [Blake et al., 2007](#)). The role of surface iron on inflammatory marker gene expression and inflammasome activation was shown to be increased following exposure to LAA in human epithelial cells ([Duncan et al., 2014](#); [Shannahan et al., 2012a](#); [Shannahan et al., 2012c](#); [Shannahan et al., 2012b](#); [Shannahan et al., 2012d](#); [Shannahan et al., 2011b](#); [Duncan et al., 2010](#); [Pietruska et al., 2010](#); see Table 4-18). Tremolite studies also demonstrate cytotoxicity in various cell culture systems (see Table 4-22). However, evidence is currently insufficient to establish the noncancer mode of action for LAA.

#### **6.1.4. Carcinogenicity in Humans and Laboratory Animals**

There is convincing evidence of a causal association between exposure to LAA mesothelioma and lung cancer in workers from the Libby, MT vermiculite mining and milling operations as well as workers from the Marysville, OH plant ([Larson et al., 2010b](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus et al., 1988](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). [Whitehouse et al. \(2008\)](#) documented 11 mesothelioma cases in nonworkers exposed to LAA in Libby, MT. Increased lung cancer and mesothelioma deaths are also reported for worker cohorts exposed to other forms of amphibole fibers (amosite and crocidolite) ([de Klerk et al., 1989](#); [Seidman et al., 1986](#); [Henderson and Enterline, 1979](#)). These findings are consistent with the increased cancers reported for communities exposed to various rocks and soils containing tremolite fibers ([Hasanoglu et al., 2006](#); [Sichletidis et al., 1992b](#); [Baris et al., 1987](#); [Langer et al., 1987](#); [Baris et al., 1979](#); [Yazicioglu, 1976](#)). Although potency, fiber dimension, and mineralogy differ among amphiboles, these studies are supportive of the hazard identification of LAA fibers described in this assessment.

Although experimental data in animals and data on toxicity mechanisms are limited for LAA, tumors were observed in tissues similar to those seen in humans (e.g., mesotheliomas, lung cancer) indicating that the existing data are consistent with the cancer effects observed in humans exposed to LAA. [Smith \(1978\)](#) reported increased incidence of mesotheliomas in hamsters after intrapleural injections of LAA. Additionally, studies in laboratory animals (rats and hamsters) exposed to tremolite via inhalation ([Bernstein et al., 2005](#); [Bernstein et al., 2003](#); [Davis et al., 1985](#)), intrapleural injection ([Roller et al., 1997, 1996](#); [Davis et al., 1991](#); [Wagner et al., 1982](#);

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1 [Smith et al., 1979](#)), or implantation ([Stanton et al., 1981](#)) have shown increases in mesotheliomas  
2 and lung cancers. The tremolite used in these studies was from various sources and varied in  
3 fiber content and potency (see Section 4.2, Appendix D). Although [McConnell et al. \(1983a\)](#)  
4 observed no increase in carcinogenicity following oral exposure to nonfibrous tremolite, the  
5 ability of this study to inform the carcinogenic potential of fibrous tremolite through inhalation is  
6 unclear, and the study results contribute little weight to the evaluation of the carcinogenicity of  
7 fibrous LAA.

8 The available mechanistic information suggests LAA induces effects that may play a role  
9 in carcinogenicity (see Section 4.2, Appendix D). Several in vitro studies have demonstrated  
10 oxidative stress and genotoxicity following LAA exposures in various cell types ([Duncan et al.,](#)  
11 [2010](#); [Hillegass et al., 2010](#); [Pietruska et al., 2010](#); [Blake et al., 2007](#)). LAA increased  
12 intracellular ROS in both murine macrophages and human epithelial cells ([Duncan et al., 2010](#);  
13 [Blake et al., 2007](#)). Additionally, surface iron, inflammatory marker gene expression,  
14 inflammasome, and aneugenic micronuclei were increased following exposure to LAA in human  
15 epithelial cells ([Duncan et al., 2010](#); [Pietruska et al., 2010](#)). Tremolite studies demonstrate  
16 cytotoxic and clastogenic effects (e.g., micronucleus induction and chromosomal aberrations) of  
17 the fibers in various cell culture systems.

#### 19 **6.1.5. Susceptible Populations**

20 Certain populations could be more susceptible than the general population to adverse  
21 health effects from exposure to LAA. In general, factors that may contribute to increased  
22 susceptibility from environmental exposures include lifestage, gender, race/ethnicity, genetic  
23 polymorphisms, health status, and lifestyle. However, little data exist to address the potential of  
24 increased susceptibility to cancer or noncancer effects from exposure to the LAA.

25 Most occupational studies of workers exposed to LAA have examined the effects only in  
26 men because this group represents the vast majority of workers in these settings ([Moolgavkar et](#)  
27 [al., 2010](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus et al., 1988](#); [Amandus et al., 1987b](#);  
28 [Amandus and Wheeler, 1987](#); [Amandus et al., 1987a](#); [McDonald et al., 1986a](#); [McDonald et al.,](#)  
29 [1986b](#)). The analysis presented here includes all workers; however, there were few women in  
30 the cohort, and therefore, no determination can be made regarding increased susceptibility to  
31 lung cancer or mesothelioma by gender. Gender-related differences in exposure patterns,  
32 physiology, and dose-response are some of the factors that may contribute to gender-related  
33 differences in risk from asbestos exposure ([Smith, 2002](#)). The limited data available from  
34 community-based studies ([ATSDR, 2000](#)) do not provide a basis for drawing conclusions  
35 regarding gender-related differences in carcinogenic effects from LAA. Racial diversity among  
36 workers exposed to LAA is also limited, and data on ethnic groups are absent, precluding the  
37 ability to examine racial and ethnicity-related differences in the mortality risks within the Libby,  
38 MT worker cohort. Finally, the potential modifying effects of genetic polymorphisms,

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preexisting health conditions, nutritional status, and other lifestyle factors have not been studied sufficiently to determine their potential contribution to variation in risk in the population.

#### 6.1.6. Mode-of-Action Information

Research on multiple types of elongate mineral fibers supports the role of multiple modes of action following exposure to LAA. Of the MOAs described in Section 4.4, the evidence that chronic inflammation, genotoxicity and cytotoxicity, and cellular proliferation may all play a role in the carcinogenic response to LAA is only suggestive (see Table 4-23). In vitro studies provide evidence that amphibole asbestos is capable of eliciting genotoxic and mutagenic effects in mammalian respiratory cells; however, direct evidence linking mutagenicity to respiratory cells following inhalation exposure is lacking. Results of the in vivo studies described here are consistent with the hypothesis that some forms of amphibole asbestos act through a MOA dependent on cellular toxicity, based on the observations that cytotoxicity and reparative proliferation occur following subchronic exposure and that bronchiolar tumors are produced at exposure levels that produce cytotoxicity and reparative proliferation. However, dose-response data in laboratory animal studies for damage/repair and tumor development are limited due to the limited number of inhalation studies that used multiple doses of fibers. Although evidence is generally supportive of a MOA involving chronic inflammation or cellular toxicity and repair, there is insufficient evidence to establish a MOA; thus, a linear approach is used to calculate the inhalation cancer unit risk in accordance with the default recommendation of the 2005 *Guidelines for Carcinogen Risk Assessment* ([U.S. EPA, 2005a](#)). It is possible that multiple MOA discussed above, or an alternative MOA, may be responsible for tumor induction.

#### 6.1.7. Weight-of-Evidence Descriptor for Cancer Hazard

Under the EPA *Guidelines for Carcinogen Risk Assessment* ([U.S. EPA, 2005a](#)), LAA is *carcinogenic to humans* following inhalation exposure based on epidemiologic evidence that shows a convincing association between exposure to LAA fibers and increased lung cancer and mesothelioma mortality ([Larson et al., 2010b](#); [Moolgavkar et al., 2010](#); [Sullivan, 2007](#); [McDonald et al., 2004](#); [Amandus and Wheeler, 1987](#); [McDonald et al., 1986a](#)). These results are further supported by animal studies that demonstrate the carcinogenic potential of LAA fibers and tremolite fibers in rodent bioassays (see Section 4.1, 4.2, Appendix D). As LAA is a durable mineral fiber of respirable size, this weight-of-evidence descriptor is consistent with the extensive published literature that documents the carcinogenicity of amphibole fibers (as reviewed in [Aust et al., 2011](#); [Broaddus et al., 2011](#); [Bunderson-Schelvan et al., 2011](#); [Huang et al., 2011](#); [Mossman et al., 2011](#)).

EPA's *Guidelines for Carcinogenic Risk Assessment* ([U.S. EPA, 2005a](#)) indicate that for tumors occurring at a site other than the initial point of contact, the weight of evidence for

1 carcinogenic potential may apply to all routes of exposure that have not been adequately tested at  
2 sufficient doses. An exception occurs when there is convincing information (e.g., toxicokinetic  
3 data) that absorption does not occur by other routes. Information on the carcinogenic effects of  
4 LAA via the oral and dermal routes in humans or animals is absent. The increased risk of lung  
5 cancer and mesothelioma following inhalation exposure to LAA has been established by studies  
6 in humans, but these studies do not provide a basis for determining the risk from other routes of  
7 exposure. Mesothelioma occurs in the pleural and peritoneal cavities, and therefore, is not  
8 considered a portal-of-entry effect. However, the role of indirect or direct interaction of asbestos  
9 fibers in disease at these extrapulmonary sites is still unknown. There is no information on the  
10 translocation of LAA to extrapulmonary tissues following either oral or dermal exposure, and  
11 limited studies have examined the role of these routes of exposure in cancer. Therefore, LAA is  
12 considered *carcinogenic to humans* by the inhalation route of exposure.

## 14 **6.2. EXPOSURE-RESPONSE**

### 15 **6.2.1. Noncancer/Inhalation**

16 There were three potential candidate studies for the derivation of the RfC—two of these  
17 were occupationally exposed cohorts, that is, the Libby worker cohort ([Larson et al., 2012a](#)) and  
18 the Marysville worker cohort ([Rohs et al., 2008](#)) and the third was of community residents in  
19 Minneapolis, MN ([Minneapolis cohort, Alexander et al., 2012](#)). Each of these studies provided  
20 individual exposure estimates and documented increased hazard of pleural effects. As detailed in  
21 Section 5.2.1, each of the available studies has strengths and weaknesses. The cohort of  
22 Marysville, OH workers (Lockey et al. (1984) and the follow-up by Rohs et al. (2008)) was  
23 selected as the principal cohort over the Libby worker cohort for several reasons: (1) lack of  
24 confounding by residential and community exposure; (2) availability of information on important  
25 covariates (e.g., BMI); (3) an exposure-response relationship defined for lower cumulative  
26 exposure levels (particularly the workers hired in 1972 or later and evaluated in 2002–2005);  
27 (4) adequate length of follow-up; (5) use of more recent criteria for evaluating radiographs ([ILO,](#)  
28 [2002](#)); (6) availability of high-quality exposure estimates based on numerous industrial hygiene  
29 samples and work records (see Section 5.2.1 for details); and (7) availability of data on TSFE  
30 matched to the exposure data. The study of Libby workers ([Larson et al., 2012a](#)) had many of  
31 these same attributes (e.g., adequate follow-up and high-quality exposure estimates), but  
32 exposure levels were generally higher in this group compared to the Marysville workers, and the  
33 Libby workers may have experienced greater levels of undocumented “take home” and other  
34 nonoccupational exposure for which TSFE data were more uncertain. The main limitation in the  
35 study of Minneapolis community residents ([Alexander et al., 2012](#)) was relatively lower quality  
36 exposure information; exposure estimates were based on a small number of total dust  
37 measurements from stack emissions combined with air dispersion modeling, and the authors

1 estimate that the individual exposure estimates are likely to have an order of magnitude of  
2 uncertainty. Thus, the study of Marysville workers ([Rohs et al., 2008](#)) was selected as the  
3 principal study for RfC derivation.

4 The MOA for LPT and the results of other asbestos epidemiology studies could  
5 potentially inform noncancer modeling decisions and suggest exposure metrics to use in  
6 modeling. However, the conclusion of Section 3 of this assessment is that the data are not  
7 sufficient to establish a MOA for the pleural and/or pulmonary effects of exposure to LAA. A  
8 general understanding of the biology and epidemiology of pleural health endpoints suggests that  
9 the timing of exposure, the exposure intensity, and the duration of exposure may be important  
10 explanatory variables and these variables were carried forward for modeling using three  
11 exposure metrics called “mean exposure” or “mean concentration” (C), “cumulative exposure”  
12 (CE), and “residence time-weighted exposure” (RTW).

13 LPT was selected as the critical effect for derivation of the RfC, with a BMR of 10%  
14 extra risk. LPT was selected because, among the noncancer radiographic endpoints evaluated in  
15 the principal study, it is the endpoint that generally appears soonest after exposure and at the  
16 lowest levels of exposure, and it was deemed the most sensitive endpoint. LPT is a pathological  
17 change associated with decreased pulmonary function, and thus is considered an appropriate  
18 adverse effect for deriving the RfC (see Section 5.2.2.3 and Appendix I).

19 The RfC is derived based on data from the Marysville workers who were evaluated in  
20 2002–2005 and hired in 1972 or later. These workers were selected due to the greater certainty  
21 in their exposure assessment. BMC modeling was used to derive the POD. Statistical models  
22 were evaluated based on biological and epidemiological considerations (see Section 5.2.2.6.1)  
23 and EPA’s Benchmark Dose Technical Guidance ([U.S. EPA, 2012](#)). Considerations included  
24 (1) the nature of the data set (i.e., cross-sectional, dichotomous health outcome data), (2) ability  
25 to estimate the effect of exposure and of covariates, (3) appropriate inclusion of a plateau term  
26 representing theoretical maximal prevalence of the outcome, and (4) appropriate estimation of  
27 the background rate of the outcome. A number of models were evaluated, and the Dichotomous  
28 Hill model with the plateau parameter fixed at a literature-derived value of 85% was selected for  
29 the derivation of a POD and sensitivity analyses. This model had very similar fit to others  
30 evaluated and was thought to provide the greatest flexibility and ability to determine sensitivity  
31 of model results to various assumptions. EPA considered several exposure metrics informed by  
32 general biology and the epidemiologic literature, including mean exposure intensity, cumulative  
33 exposure (which incorporates duration of exposure), and RTW exposure (which incorporates  
34 TSFE by weighting more heavily exposures occurring in the more distant past).

35 Another important feature of the exposure-response analysis is the ability to include  
36 effects of TSFE in the modeling. TSFE has been shown in the literature to be important in  
37 evaluating risk of LPT, and studies have shown that prevalence of LPT can increase with  
38 increasing TSFE even after cessation of exposure. EPA evaluated TSFE as a predictor in the

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primary analytic group of workers hired after 1972 and evaluated in 2002–2005, but found that TSFE was not significantly associated with LPT in this group—likely due to the very low variability in TSFE for this particular population. Thus, EPA used a hybrid modeling approach to “borrow” information on the effect of TSFE from a larger subset of the Marysville workers with greater variability in TSFE. The model was fit in the group of all workers evaluated in 2002–2005 (regardless of hire date), including both LAA exposure and TSFE as predictors. The regression coefficient corresponding to TSFE was then set as a fixed parameter in the model for the primary analytic group of workers hired in 1972 or later. In this hybrid modeling, mean exposure was used due to its superior model fit compared to cumulative exposure. RTW exposure was not used since TSFE was included as a separate covariate (to avoid collinearity of predictors). Using this modeling approach (details in Section 5.2.2.6.2), the resulting BMC<sub>10</sub> under these modeling assumptions is 0.00923 fiber/cc; the corresponding lower 95% confidence limit of the BMC<sub>10</sub> (BMCL<sub>10</sub>) is 0.026 fiber/cc.

The RfC is obtained by applying uncertainty factors as needed. Three UFs have been applied for a composite UF of 300 (intraspecies uncertainty factor, UF<sub>H</sub> = 10; database uncertainty factor, UF<sub>D</sub> = 3; subchronic-to-chronic uncertainty factor, UF<sub>S</sub> = 10 ) (see Section 5.2.5). As shown below, the chronic RfC is  $9 \times 10^{-5}$  fiber/cc for LAA, calculated by dividing the POD by a composite UF of 300:

$$\begin{aligned}\text{Chronic RfC} &= \text{POD} \div \text{UF} \\ &= 0.026 \text{ fiber/cc} \div 300 \\ &= 8.67 \times 10^{-5} \text{ fiber/cc, rounded to } 9 \times 10^{-5} \text{ fiber/cc}\end{aligned}\tag{6-1}$$

Note that for the primary RfC as well as for all the alternative RfCs, the fiber concentration are presented here as continuous lifetime exposure in fiber/cc, where exposure measurements are based on analysis of air filters by PCM. Current analytical instruments used for PCM analysis have resulted in a standardization of minimum fiber width considered visible by PCM between 0.2 and 0.25  $\mu\text{m}$ . Historical PCM analysis (1960s and early 1970s) generally had less resolution, and fibers with minimum widths of 0.4 or 0.44  $\mu\text{m}$  were considered visible by PCM ([Amandus et al., 1987b](#); [Rendall and Skikne, 1980](#)). Methods are available to translate exposure concentrations measured in other units into PCM units for comparison.

EPA conducted RfC derivation from the same subcohort with alternative definitions of the health endpoint (see Section 5.2.3.1 and 5.2.3.2). The chronic RfC value for “any pleural thickening” (APT) was also  $9 \times 10^{-5}$  fiber/cc and the same value was derived for “any radiographic change” (ARC). Although EPA based the primary value of the RfC on the model based on mean exposure, Section 5.2.4 illustrates an alternative derivation of an RfC of  $1 \times 10^{-4}$  fiber/cc from the same cohort with an alternative exposure metric of cumulative exposure. EPA also conducted alternative modeling of the Marysville cohort, including all individuals who

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1 participated in the health examination in 1980 ([Lockey et al., 1984](#)) and 2002–2005 ([Rohs et al.,](#)  
2 [2008](#)) and who were not exposed to asbestos from a source outside of the Marysville facility (see  
3 Section 5.2.4 and Appendix E). The modeling of this full cohort ( $n = 434$  individuals) was  
4 performed using an alternative critical effect of “any pleural thickening” (APT), a slightly  
5 different definition of LPT (diagnostic criteria changed slightly over time) and with “any  
6 radiographic change” (ARC). These analyses yielded five other RfC values. A summary table  
7 of the primary and alternative derivation of the RfC is provided in Table 5-11 in Section 5.2.5;  
8 all eight alternative derivations of the RfC were within threefold of the primary RfC, ranging  
9 from  $3 \times 10^{-5}$  fiber/cc up to  $2 \times 10^{-4}$  fiber/cc. This series of derivations further substantiates the  
10 primary RfC derived from the Marysville workers evaluated in 2002–2005, and hired in 1972 or  
11 later.

12 Confidence in the principal study is considered medium. The data used are human  
13 epidemiological data which are preferred to animal bioassays, and the principal study ([Rohs et](#)  
14 [al., 2008](#)) is conducted in a population of occupationally exposed workers with long term,  
15 relatively low intensity, exposures. While deriving the primary analysis from the group of  
16 workers evaluated in 2002–2005 and hired after 1972 resulted in a smaller data set with fewer  
17 cases to model, alternative RfC derivations based on the larger group of workers without  
18 restriction on the date of hire (and many more cases) yielded similar values of the RfC. The  
19 exposure assessment in the principal study is based on measured data. The main source of  
20 uncertainty in the exposure estimates is incomplete exposure measurements for some of the  
21 occupations/tasks before industrial hygiene improvements that started about 1973 or 1974 and  
22 continued throughout the 1970s (see Appendix F, Figure F-1). The principal study assessed the  
23 health outcome cross sectionally and this may underrepresent the true health burden as  
24 individuals with more severe disease could have left employment or may have died and not been  
25 included in the follow up study, resulting in an underestimation of overall toxicity. However, for  
26 health outcomes not considered to be frank effects, such as LPT, this underestimation should be  
27 minimal. Further, [Rohs et al. \(2008\)](#) compared the study participants with the complete study  
28 population and there was no evidence of major differences in the two group’s exposure  
29 distributions. Thus, the potential for selection bias is considered to be low. In terms of the  
30 sensitivity of the principal study to detect the critical effect (LPT) by radiograph, it is known that  
31 HRCT can identify asbestos related lesions in the respiratory tract, which cannot be identified by  
32 standard radiographs (e.g., [Lebedova et al., 2003](#); [Janković et al., 2002](#); [Šimundić et al., 2002](#)).  
33 Thus, the technology employed for determining the prevalence of radiographic changes in the  
34 Marysville cohort will likely underestimate the prevalence of pleural lesions that could be  
35 detected using HRCT.

36 Confidence in the completeness of the overall database is medium. The database consists  
37 of long term mortality and morbidity studies in humans exposed via inhalation to LAA. The  
38 mortality studies do not provide appropriate data for RfC derivation for pleural abnormalities,

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although the two other morbidity studies ([Alexander et al., 2012](#); [Larson et al., 2012a](#)) do support the conclusion that low levels of exposure to LAA is associated with increased prevalence of LPT. It is known that inhaled asbestos fibers migrate out of the lung and into other tissues (see Section 3.1), which leads to uncertainty regarding the assumption that other health effects would not be expected. While a potential for autoimmune effects and cardiovascular disease is noted in exposed individuals, there are insufficient data to provide a quantitative exposure response relationship for these endpoints. It is unknown whether an RfC based on these other health would result in a higher or lower estimate for the RfC. Nor is there evidence as to whether any of these other effects would occur earlier following exposure to LAA than LPT occurs. There are no data in laboratory animals or humans on general systemic effects. Therefore, overall confidence in the RfC is medium, reflecting medium confidence in the principal study and medium confidence in the completeness of the overall database.

#### **6.2.1.1. Uncertainty and Sensitivity Analyses for Reference Concentration (RfC) Derivation**

It is important to consider the sources of uncertainties in the derivation of the RfC for LAA. These include the following:

*Measurement error in exposure assessment and assignment.* The estimated exposure for each individual relied on self-reported employment history, which may be subject to recall error. Only data from 1972 and later were used for an RfC derivation based on lack of fiber measurements prior to this date, but some uncertainty remains due to the limited amount of industrial hygiene data collected in 1972–1973. There is also uncertainty in the post-1972 data regarding asbestos content and potency of fibers originating from other ore sources (Virginia, South Carolina, and South Africa). Although LAA was not used in the facility after 1980, industrial hygiene measurements collected after 1980 showed low levels of fibers. Regarding nonoccupational exposure, any exposure to LAA outside of the workplace is not likely to contribute significantly to cumulative exposure—~10% of workers reported bringing raw vermiculite home, and the majority showered and changed clothes before leaving the workplace. As a sensitivity analysis, EPA evaluated the change in the POD when setting all exposure measurements after 1980 to zero, and when using the geometric mean rather than the arithmetic mean to summarize multiple fiber measurements for a given task/location/time period. These analyses showed a difference of –65 to +50% in the POD.

*Radiographic assessment of localized pleural thickening.* Conventional radiographs—rather than the more sensitive high-resolution computed tomography—were used to determine the health outcome. Localized pleural thickening may be difficult to detect on these radiographs, leading to the potential for outcome misclassification. However, uncertainty in the detection of LPT in each individual is considered minimal due to the use of a team of highly qualified chest radiologists evaluating the radiographic films and the use of consensus diagnosis.

*Use of an alternative critical effect.* In addition to the primary analysis using a critical effect of LPT, EPA also derived an RfC based on the alternative endpoint of any pleural

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thickening (APT), and based on any radiographic change (ARC). In the primary analytic group of workers evaluated in 2002–2005 and hired in 1972 or later, the two alternative endpoints are identical to the critical effect of LPT, because the one individual with DPT also had LPT, and none had interstitial changes. In the larger group of workers used to estimate the effect of TSFE (all evaluated in 2002–2005), there were 69 cases of APT and 71 with ARC, of which the majority ( $n = 66$ ) were LPT. Consequently, the RfC derived using these two alternative endpoints of APT or ARC were identical (to one significant digit) to that derived for LPT,  $9 \times 10^{-5}$  fibers/cc.

*Length of follow-up.* Time from first exposure to x-ray was 23.2–32.7 years in the primary analytic group of workers evaluated in 2002–2005, and hired in 1972 or later (mean of 28.2 years). The literature shows that the prevalence of LPT may increase with time, beyond this observed range of time from first exposure. The lack of observed data beyond ~30 years after first exposure (on average) is a source of uncertainty when characterizing the exposure-response relationship for a full lifetime of exposure (e.g., 70 years).

*Model Form.* A number of model forms were explored in the initial stages of analysis, and generally showed reasonably close fits as measured by the AIC. The Dichotomous Hill model with a plateau fixed at 85% was selected for RfC derivation due to its greater flexibility and ability to evaluate sensitivity to model assumptions. EPA also evaluated the sensitivity of the fixed plateau parameter and found that the POD changed very little (<16%) when fixing the plateau at different values (70, 100%) or when estimating the plateau from the Marysville data.

*Effect of covariates.* Information on a number of covariates was available for the Marysville workers, including demographic characteristics (gender, smoking status, BMI) as well as potentially exposure-related factors (hire year, job tenure, exposure duration, and age at x-ray). The potential for these factors to confound the association between LAA exposure and LPT was investigated in two ways. First, each was evaluated for association with the exposure, association with the outcome, and whether it was an intermediate in the pathway between exposure and outcome (i.e., did they meet the theoretical definition for a confounder). By these standards, none of the covariates was a confounder. Second, each covariate was included in the final model to evaluate the impact on the estimated effects of LAA exposure and TSFE; the differences were quite small, and none of the covariates were significantly associated with risk of LPT in these models.

## **6.2.2. Cancer/Inhalation**

### **6.2.2.1. Background and Methods**

The most appropriate data set for deriving quantitative cancer risk estimates based on LAA exposure in humans is the cohort of workers employed at the vermiculite mining and milling operation near Libby, MT (see Section 4.1.1.1). The Libby, MT worker cohort has been the focus of two epidemiologic investigations by the NIOSH scientists. A database created by NIOSH in the 1980s contains demographic data, work history, and vital status at the end of May of 1982 for 1,881 workers at the vermiculite mine, mill, and processing plant in Libby, MT (see

1 Section 4.1.1.1). Vital status follow-up was completed by NIOSH through 2006 using the  
2 National Death Index ([Bilgrad, 1997](#)). Nearly 54% of workers in the cohort ( $n = 1,009$ ) had died  
3 by December 31, 2006. The data from this update (provided by NIOSH) is the basis of EPA  
4 exposure-response modeling.

5 EPA does not have sufficient information to select models for the epidemiology data on  
6 the basis of the biological mechanism of action for lung cancer or mesothelioma (see Section 3).  
7 In this situation, EPA's practice is to investigate several modeling options to determine how to  
8 best empirically model the exposure-response relationship in the range of the observed data, as  
9 well as to consider exposure-response models suggested in the epidemiologic literature. For  
10 LAA, possible exposure metrics were explored for model fit to the chosen models forms. The  
11 exposure metric options were selected to provide a range of shapes that was sufficiently flexible  
12 to allow for a variety of ways that time and duration might relate to cancer risk in the data being  
13 modeled. EPA then evaluated how well the models and exposure metric combinations fit the  
14 data being modeled. Then EPA calculated a reasonable upper bound on risk using selected  
15 exposure metrics. This is explained in more detail below and in Section 5.4.5. However, there  
16 are uncertainties in the modeling of the epidemiological data that may impact the IUR and these  
17 are described in Section 6.2.8 below and in greater detail in Section 5.4.6.

18 In the Libby, MT worker cohort data developed by NIOSH and used by EPA in this  
19 assessment, detailed work histories, together with job-specific exposure estimates, allowed for  
20 the reconstruction of each individual's occupational exposure experience over time to define  
21 multiple exposure metrics. From this information-rich, individual-level data set from NIOSH,  
22 EPA constructed a suite of the different metrics of occupational exposure which had been  
23 proposed in the asbestos literature or used in EPA health assessment on general asbestos  
24 exposures ([U.S. EPA, 1988a](#)) as well as modifications proposed ([Berry et al., 2012](#)). This suite  
25 of models was defined a priori to encompass a reasonable set of proposed exposure metrics to  
26 allow sufficient flexibility in model fit to these data. These exposure metrics were evaluated in  
27 analytic-regression models to test which exposure metrics were the best empirical predictors of  
28 observed cancer mortality, and the better fitting models were advanced for consideration as the  
29 basis of the exposure-response relationship for the IUR. The types of exposure metrics evaluated  
30 were intended to allow for variations of the classic metric of cumulative exposure, allowing for  
31 more or less weight to be placed on earlier or later exposures. These simulated exposure metrics  
32 were derived mathematically to approximate underlying processes that are not well understood.  
33 Thus, the empirical fit of various exposure metrics to the observed epidemiologic data is  
34 evaluated statistically, and the exposure metrics have epidemiological interpretation but do not  
35 necessarily have direct biological interpretations.

36 Exposure estimates for all exposure metrics were adjusted to account for the time period  
37 between the onset of cancer and mortality. The lag period defines an interval before death, or  
38 end of follow-up, during which any exposure is excluded from the calculation of the exposure

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metric. There was one important limitation of the NIOSH JEM. Of the 991 workers hired before 1960, 706 workers with unknown department code and unknown job assignments hired between 1935 and 1959 were assigned the same average estimated exposure intensity. The lack of information on specific job assignments for such a large portion of these early workers when exposures were higher resulted in the misclassification of the exposure and effectively yielded exposure metrics that were differentiated only by the duration of each worker's employment. For this reason and because there was little measured fiber exposure data during the earlier period, identifying an adequate exposure-response model fit was unsuccessful. The two biggest problems were that the duration of employment was the best-fitting metric for modeling mesothelioma and that the Cox model assumptions were violated in modeling lung cancer mortality (see Section 5.4.3.4). As a result, this assessment developed a subcohort analysis by dividing the whole cohort into two groups: those hired before 1960 and those hired after 1959. This removed all but nine cohort members with missing department code and job category information and lessened the effect of estimates of early exposures where no air sampling data were available. For the subcohort of those hired after 1959, those two biggest problems were resolved: the assumptions of the Cox model were satisfied, and a lagged cumulative exposure with a decay (rather than duration of exposure, as for the full cohort) was the best-fitting metric for mesothelioma.

Of the 880 workers hired after 1959, 230 (26%) had died by December 31, 2006. The number of mesothelioma deaths in the subcohort is relatively small ( $n = 7$ , two deaths coded in ICD-10 and five deaths coded in ICD-9), but the rate of mesothelioma mortality was very similar in the subcohort (24.7 per 100,000 person-years versus 26.8 per 100,000 person-years for the full cohort [18 mesothelioma deaths], a difference of less than 10%).

### 6.2.3. Modeling of Mesothelioma Exposure Response

A Poisson model is employed for estimating the absolute risk of mesothelioma following exposure to LAA, as the Poisson distribution is an appropriate model to use with data that are counts of a relatively rare outcome, such as observed mesothelioma deaths in the Libby, MT worker cohort. Estimation of the exposure-response relationship for mesothelioma using the Poisson model was performed in WinBUGS software by a MCMC Bayesian approach with an uninformative or diffuse (almost flat) prior. The model was run to fit the mortality data to exposure data for various exposure metrics described above. To comparatively evaluate how much better one model fits than another, the DIC was used. DIC is used in Bayesian analysis and is an analogue of AIC ([Burnham and Anderson, 2002](#)). Use of the DIC and AIC is standard practice in comparing the fit of nonnested models to the same data set with the same dependent outcome variable but different independent covariates.

Modeling of mesothelioma mortality included an exposure metric with a cubic function of time (see eq 5-9), originally proposed in [Peto et al. \(1982\)](#) and employed in derivation of the

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IUR for asbestos ([U.S. EPA, 1988a, 1986a](#)), as well as modifications of Peto model (Peto model with clearance) proposed in the asbestos literature ([Berry et al., 2012](#)). See Section 5.4.3 for further details.

For the subcohort hired after 1959, two cumulative exposure metrics with decay provided the best model fits. Both metrics had a common 5-year half-life, with lag times of either 10 or 15 years. The Peto model and Peto model with clearance did not fit as well. As it is less likely that exposure during the last few years before death were contributory to the development of the cancer and cancer mortality, the zero-lag metrics were dropped from further consideration. The selected metric as well as the Peto model and Peto model with clearance were retained for derivation of the IUR (see Section 6.2.7 below and for additional detail see Section 5.4.5). The selected exposure metric was cumulative exposure with a 5-year half-life and a 10-year lag time with a central estimate for the slope (KM) of  $3.11 \times 10^{-4}$  per fiber/cc-year with a 95% upper confidence limit (UCL) of  $5.08 \times 10^{-4}$  per fiber/cc-year.

#### **6.2.4. Unit Risk Estimates for Mesothelioma Mortality**

The increased risk of mesothelioma mortality attributable to continuous fiber exposure was estimated using a life-table procedure based on the general U.S. population. The life-table procedure involved the application of the estimated LAA toxicity to a structured representation of the general U.S. population in such a manner as to yield age-specific risk estimates for cancer mortality in the presence or absence of exposure to LAA (see Section 5.4.5; Appendix G).

A default linear low-dose extrapolation was used because the mode of action by which LAA causes mesothelioma was not established. The lower limit on the effective concentration ( $LEC_{01}$ ) yielded a unit risk for mesothelioma mortality of 0.053 per fiber/cc (POD of 1% divided by the  $LEC_{01}$ ).

The value of the effective concentration (EC) that would correspond to the measure of central tendency is the  $EC_{01}$ . This value is used in the derivation of a combined risk of mesothelioma and of lung cancer. The  $EC_{01}$  yielded a lifetime central estimate value of 0.032 per fiber/cc.

For mesothelioma, the undercounting of cases (underascertainment) is a particular concern given the limitations of the ICD classification systems used before 1999. In practical terms, this means that some true occurrences of mortality due to mesothelioma are missed on death certificates and in almost all administrative databases such as the National Death Index. Even after introduction of special ICD code for mesothelioma with introduction of ICD-10 in 1999, detection rates are still imperfect ([Camidge et al., 2006](#); [Pinheiro et al., 2004](#)), and the reported numbers of cases typically reflect an undercount of the true number. [Kopylev et al. \(2011\)](#) reviewed the literature on this underascertainment and developed methods to account for the likely numbers of undocumented mesothelioma deaths.

To compensate for mesothelioma underascertainment attributable to ICD coding, the mesothelioma mortality unit risk was further adjusted following the analysis of [Kopylev et al. \(2011\)](#). The adjusted mesothelioma central (i.e., maximum likelihood estimate) risk, corresponding to the best-fit metric, was 0.044 per fiber/cc, and the adjusted mesothelioma mortality unit risk was 0.074 per fiber/cc.

The adjusted mesothelioma risks for the Peto model and Peto model with clearance ranged from 2-fold lower (0.035 per fiber/cc) to 3.6-fold higher (0.265 per fiber/cc). Thus, there is uncertainty in mesothelioma risks generated from similar-fitting models from different exposure metrics (see details in Section 5.4.5.3).

## **6.2.5. Modeling of Lung Cancer Exposure Response**

All multivariate extended Cox models were fit to the subcohort hired after 1959 with covariates for gender, race, date of birth, and exposure. Exposure for each of the 40 exposure parameterizations was calculated independently, and the fit of these exposure metrics was evaluated one at a time. Of the 40 exposure-response metrics, 14 demonstrated an adequate fit to the data as measured by the overall model fit with the likelihood ratio test ( $p < 0.05$ ) as well as having statistically significant exposure metrics ( $p < 0.05$ ). However, only the nine models that demonstrated adequate model and exposure metric fit and incorporated a lag period to account for cancer latency were considered further in the development of the IUR (see Tables 5-43 and 5-50).

Lagging exposure by 10 years was a better predictor of lung cancer mortality compared to other lags. As it is less likely that exposure during the last few years before death were contributory to the development of the cancer and cancer mortality, the zero lag metrics were dropped from further consideration. The residence time-weighted cumulative exposure, both with and without decay of the exposure metric, did not fit these lung cancer mortality data well compared to the other models (see Table 5-44); this form of exposure metric does not demonstrate evidence of an empirical fit to these epidemiologic data.

The model with the smallest AIC was for cumulative exposure with a 10-year half-life for decay and a 10-year lag for cancer latency. The extended Cox model estimated a beta (the lung cancer slope factor: KL) of  $1.26 \times 10^{-2}$  per fibers/cc-yr based on a 365-day year, and the 95<sup>th</sup> percentile upper bound was  $1.88 \times 10^{-2}$  per fibers/cc-yr. The  $p$ -value for the LAA regression coefficient beta (slope) was  $<0.001$ . The slopes and confidence interval for the other exposure metrics, which had similar fits to these data are reported in Table 5-45. Uncertainty in the choice of the exposure metric is considered in the derivation of the unit risk (see details in Section 5.4.5.2), representing the range of unit risks that are derived from these similarly fitting metrics. The model results that were ultimately selected to reflect the upper bound among the range of results were based on the cumulative exposure with a 10-year lag exposure metric



(CE10). The extended Cox model estimated a beta (slope) of  $5.28 \times 10^{-3}$  per fibers/cc-yr based on a 365-day year, and the 95<sup>th</sup> percentile upper bound was  $1.00 \times 10^{-2}$  per fibers/cc-yr.

#### **6.2.5.1. Analysis of Potential Confounding of Lung Cancer Results by Smoking in the Subcohort**

EPA used two approaches to address the confounding issue, including restriction of the cohort and an analytic evaluation of the potential for confounding by smoking including the method described by [Richardson \(2010\)](#). [Richardson \(2010\)](#) describes a method to determine whether an identified exposure relationship with lung cancer is confounded by unmeasured smoking in an occupational cohort study. EPA implemented this methodology to model the potential effects of LAA on the risk of COPD mortality on the subcohort of workers hired after 1959 (see Section 5.4.3.8). Summarizing these findings, EPA used the method described by [Richardson \(2010\)](#) to evaluate whether exposures to LAA predicted mortality from COPD as an indication of potential confounding by smoking and found a nonsignificant negative relationship, which was inconsistent with confounding by smoking.

#### **6.2.6. Unit Risk Estimates for Lung Cancer Mortality**

The increased risk of lung cancer mortality attributable to continuous fiber exposure was estimated using a life-table procedure based on the general U.S. population. The life-table procedure involved applying the estimated LAA-specific toxicity to a structured representation of the general U.S. population in such a manner as to yield age-specific risk estimated for cancer mortality in the presence or absence of exposure to LAA (see Section 5.4.5; Appendix G). A default linear low-dose extrapolation was used because the mode of action by which LAA causes lung cancer was not established.

The nine exposure-response models retained in Table 5-45 all had reasonably similar goodnesses of fit. No single model stands out as clearly statistically superior; however, there is a range of quality of fit within the set that could be considered to have adequate fit. The lung cancer mortality unit risks are shown in Table 5-52.

Using the results of the exposure model based on cumulative exposure with a 10-year lag for cancer latency, the LEC<sub>01</sub> yielded a lifetime unit risk of 0.0679 per fiber/cc. The value of the risk that would correspond to the measure of central tendency involves the EC<sub>01</sub> rather than the LEC<sub>01</sub>. The EC<sub>01</sub> yielded a lifetime central estimate of 0.0399 per fiber/cc.

The resulting unit risks in Table 5-52 ranged from 0.0260 to 0.0679 per fiber/cc, for a lifetime continuous exposure. This shows that the unit risk based on the exposure metric with the lowest AIC value (i.e., cumulative exposure with a 10-year half-life for decay and a 10-year lag for cancer latency) is in the center of this range (i.e., 0.0389 per fiber/cc). This estimate is in the middle of the range of possible unit risks and does not capture the uncertainty across metrics with similar goodness of fit (see details in Section 5.4.6).

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1 The model results selected to represent the upper-bound risk among the range of  
2 reasonable results are based on a CE10 metric with a 10-year lag. The model results selected to  
3 reflect the upper bound among the range of results are based on the CE10 exposure metric with a  
4 10-year lag, providing a unit risk of 0.0679 per fiber/cc.

#### 6 6.2.7. Inhalation Unit Risk (IUR) Derivation Based on Combined Mesothelioma and Lung 7 Cancer Mortality from Exposure to Libby Amphibole Asbestos

8 Once the cancer-specific lifetime unit risks are selected, the two are then combined. It is  
9 important to note that this estimate of overall potency describes the risk of mortality from cancer  
10 at either of the considered sites and is not just the risk of both cancers simultaneously. Because  
11 each of the unit risks is itself an upper-bound estimate, summing such upper-bound estimates  
12 across mesothelioma and lung cancer mortality is likely to overpredict the overall risk.  
13 Therefore, following the recommendations of the *Guidelines for Carcinogen Risk Assessment*  
14 ([U.S. EPA, 2005a](#)), a statistically appropriate upper bound on combined risk was derived to gain  
15 an understanding of the overall risk of mortality resulting from mesothelioma and from lung  
16 cancer. For mesothelioma, the exposure-response models developed by EPA using personal  
17 exposure data on the subcohort (see Table 5-50) provided better fit to the subcohort data than the  
18 Peto model and the Peto model with clearance that have been proposed in the asbestos literature.  
19 For lung cancer, this assessment selected the upper bound among the lung cancer lifetime unit  
20 risks from the plausible exposure metrics (regardless of the small residual differences in quality  
21 of fit). Because there were few metrics with unit risks higher than the best fitting metric's unit  
22 risk for lung cancer mortality endpoint, this method effectively selects the highest lifetime unit  
23 risk among those considered for the lung cancer mortality endpoint.

24 Table 6-1 shows cancer-specific unit risks as well as combined risk of mesothelioma and  
25 lung cancer. The IUR value of 0.17 per fiber/cc, continuous lifetime exposure, accounts for  
26 important quantitative uncertainties in the selection of the specific exposure metric that may have  
27 remained in an IUR that might have been based on the best-fitting exposure models alone.  
28 Additional uncertainties are discussed in detail in Section 5.4.6.

**Table 6-1. Estimates of the combined central estimate of the unit risk for mesothelioma and lung cancer and the combined upper-bound lifetime unit risks for mesothelioma and lung cancer risks (the Inhalation Unit Risk) for different combination of mesothelioma and lung cancer models<sup>a</sup>**

Lung cancer	Mesothelioma	Combined central estimate per fiber/cc	Combined upper bound per fiber/cc
Selected IUR based directly on the Libby data			
<b>CE10 subcohort</b>	<b>CE10 5-yr half-life</b>	<b>0.115</b>	<b>0.169</b>
Best models from the epidemiologic literature (Peto model with clearance)			
CE10 subcohort	Peto with clearance Decay rate of 6.8%/yr Power of time = 3.9 Subcohort	0.089	0.135
CE10 subcohort	Peto with clearance Decay rate of 15%/yr Power of time = 5.4 Subcohort	0.061	0.092
Alternative model from the epidemiologic literature (Peto model)			
CE10 subcohort	Peto No decay Power of time = 3 Subcohort	0.203	0.308

<sup>a</sup>Note that for all the IUR values presented in this table, the fiber concentration is presented here as continuous lifetime exposure in fiber/cc, where exposure measurements are based on analysis of air filters by PCM. Current analytical instruments used for PCM analysis have resulted in a standardization of minimum fiber width considered visible by PCM between 0.2 and 0.25  $\mu\text{m}$ . Historical PCM analysis (1960s and early 1970s) generally had less resolution, and fibers with minimum widths of 0.4 or 0.44  $\mu\text{m}$  were considered visible by PCM ([Amandus et al., 1987b](#); [Rendall and Skikne, 1980](#)). Methods are available to translate exposure concentrations measured in other units into PCM units for comparison.

#### 1 *Age-dependent adjustment factor*

2 As discussed in Section 4.7.1.1, there is no chemical-specific information for LAA, or  
3 general asbestos, that would allow for the computation of a chemical-specific age-dependent  
4 adjustment factor for assessing the risk of exposure that includes early-life exposures.

5 The review of mode-of-action information in this assessment (see Section 4.6.2.2)  
6 concluded that the available information on the mode of action by which LAA causes lung  
7 cancer or mesothelioma is complex and a mode of action is not established at this time. Thus, in  
8 accordance with EPA's *Supplemental Guidance for Assessing Susceptibility from Early-Life*  
9 *Exposure to Carcinogens* ([U.S. EPA, 2005b](#)), the application of the age-dependent adjustment  
10 factors for substances that act through a mutagenic mode of action is not recommended.

#### 6.2.7.1. Comparison with Other Published Studies of Libby, MT Workers Cohort

Several published studies have previously evaluated risk of mesothelioma and lung cancer (i.e., [Larson et al., 2010b](#); [Moolgavkar et al., 2010](#); [Berman and Crump, 2008](#); [Sullivan, 2007](#)) in Libby, MT workers cohort. For mesothelioma, only [Moolgavkar et al. \(2010\)](#) provided an exposure-response relationship for absolute risk of mesothelioma mortality that would be comparable with this current assessment. Based on the full cohort, with mortality data through 2001 and a modification of the Peto/Nicholson exposure metric, life-table analysis would provide an upper-bound unit risk of approximately 0.13 per fiber/cc continuous lifetime exposure. Therefore, utilization of the exposure response modeling of [Moolgavkar et al. \(2010\)](#), would provide an IUR for excess mesothelioma mortality in close agreement with the IUR derived in this assessment (see Section 5.4.5.3.1 for more details).

For lung cancer, all of the studies provide exposure-response relationships in terms of relative risk of lung cancer mortality, and thus, may provide risk estimates comparable to this assessment. However, inclusion criteria, length of mortality follow-up, and analytic methods differ among the analyses—thus, the results are not necessarily interchangeable. For comparison purposes, the lung cancer unit risks from these studies are computed from life-table analyses (see Table 5-54). The lung cancer unit risks calculated based on the published literature, ranged from 0.010 to 0.079 per fiber/cc (based on the upper confidence limit). This is in close agreement with this current assessment where an upper-bound estimate of 0.068 per fiber/cc, continuous lifetime exposure is derived (see Section 5.4.5.3.1 for more details).

#### 6.2.8. Uncertainty in the Cancer Risk Values

It is important to consider the uncertainties in the derivation of the mesothelioma and lung cancer mortality risks in this assessment in the context of uncertainties in animal-based health assessments. This assessment does not involve extrapolation from high dose in animals to low dose in humans. The current assessment is based on a well-documented and well-studied cohort of workers with adequate years of follow-up to evaluate mesothelioma and lung cancer mortality risks with PODs within the range of the data. The discussions in Section 5.4.6 explore uncertainty in the derivation of the IUR in order to provide a comprehensive and transparent context for the resulting cancer mortality risk estimates.

Section 5.4.6.1 details the several sources of uncertainty in the assessment of the cancer exposure-response relationships and the use of those data to derive the inhalation unit risk. The text that follows summarizes the primary sources of uncertainty and, where possible, the expected direction of effect on the exposure-response risk estimates and the inhalation units risk.

1           1) *Uncertainty in low-dose extrapolation (see Section 5.4.6.1.1)*

- 2           • Some uncertainty remains in the extrapolation of risks based on occupational  
3           exposure to environmental exposure levels, but this uncertainty is considered to  
4           be low as the lower range of occupational exposure overlaps with expected  
5           environmental exposure levels.

6           2) *Uncertainty in exposure assessment, including analytical measurements uncertainty*  
7           *(see Section 5.4.6.1.2)*

- 8           • The JEM was based on the “high” exposure estimate for each job according to  
9           [Amandus et al. \(1987b\)](#), and to some extent this could be an overestimate of  
10          exposure. The associated cancer risk would be somewhat underestimated  
11          resulting in a somewhat underestimated IUR.
- 12          • The JEM was largely based on estimated fiber concentration using PCM  
13          measurement (with some extrapolations in time), and because PCM may count all  
14          long and thin objects as fibers, these measurement could overestimate the true  
15          LAA fiber concentrations, leading to an overestimate of exposure and a somewhat  
16          underestimated cancer risk resulting in a somewhat underestimated IUR.
- 17          • PCM measurements in the era of NIOSH measurements in Libby used a lower  
18          resolution, and therefore, included only somewhat thicker fibers, thereby counting  
19          fewer fibers than would have been counted by later PCM standards. These earlier  
20          measurements could underestimate the true LAA fiber concentrations, leading to  
21          an underestimate of exposure and an overestimate of cancer risk resulting in a  
22          somewhat overestimated IUR.
- 23          • The PCM measurement is the available exposure metric for analysis of the Libby  
24          worker cohort at the time of this assessment. Currently, there is no optimal choice  
25          of the best dose metric for asbestos, in general and in particular, for LAA.  
26          Uncertainties related to PCM analytical method are discussed in Section 2, and  
27          such uncertainties cannot be related to the IUR at the time of this assessment.
- 28          • Random measurement error in the assignment of exposures could have the effect  
29          of underestimating the risk of lung cancer mortality because that measure of risk  
30          is based on a relative measure. The effect would be to somewhat underestimate  
31          the risk of lung cancer, resulting in a somewhat underestimated unit risk for lung  
32          cancer. It is unclear what the impact of such measurement error would be on the  
33          absolute risk of mesothelioma.
- 34          • Exposure to other kinds of asbestos and residential exposure to LAA may have  
35          caused workers’ actual personal exposures (as the sum of occupational and  
36          nonoccupational exposures) to have been underestimated by the use of estimated  
37          Libby occupational exposure information alone. This could underestimate the  
38          true LAA fiber exposures leading to an overestimate of the associated cancer risk  
39          resulting in a somewhat overestimated IUR.

1           3) *Uncertainty in model form (see Section 5.6.4.1.3)*

- 2           • For mesothelioma, the Poisson model is the standard epidemiologic form and  
3           considered to be the most appropriate model form for rare health outcomes;  
4           therefore, uncertainty is considered to be low. For lung cancer mortality, the Cox  
5           proportional hazards model is the standard epidemiologic form. It is considered  
6           to be the most appropriate model form for health outcomes with time-varying  
7           exposure data and thus uncertainty is considered to be low.

8           4) *Uncertainty in selection of exposure metric (see Section 5.6.4.1.4)*

- 9           • There is uncertainty about what metric should be used for modeling exposures to  
10          LAA. Table 5-53 illustrates the uncertainty in the IUR due to exposure metric  
11          selection. The quantitative uncertainty is about threefold.

12          5) *Uncertainty in assessing mortality corresponding to other cancer endpoints (see  
13          Section 5.6.4.1.5)*

- 14          • The lack of sufficient numbers of workers to estimate the risk of other cancers  
15          potentially related to LAA exposure is an uncertainty of unclear direction but is  
16          considered to be low due to the rarity of those cancers.

17          6) *Uncertainty in control of potential confounding in modeling lung cancer mortality  
18          (see Section 5.4.6.1.6)*

- 19          • The uncertainty in control of potential confounding by smoking is considered to  
20          be low as the described sensitivity analysis did not show evidence of potential  
21          confounding.

22          7) *Uncertainty due to potential effect modification (see Section 5.4.6.1.7)*

- 23          • Smoking was not considered to be related to LAA exposure, and therefore,  
24          smoking is not considered to be a likely effect modifier of cancer risk. Age has  
25          been shown to be a potential effect modifier of lung cancer risk but there was no  
26          evidence of this relationship in the subcohort.

27          8) *Uncertainty due to length of follow-up (see Section 5.4.6.1.8)*

- 28          • There is uncertainty related to the limited follow-up for cancer mortality, and it is  
29          possible that with subsequent mortality follow-up, the IUR could change in a  
30          direction that is unknown.

31          9) *Uncertainty in the use of life-tables to calculate cancer mortality IUR (see  
32          Section 5.4.6.1.9)*

- 33          • The life-table procedure computes the extra risk of death from birth to 85 years of  
34          age. If the life tables were extended from 85 to 90 years to account for longer life  
35          spans, the selected lung cancer mortality unit risk (Table 5-53 shows this as  
36          0.068) would be somewhat larger, about 5–10%, and the selected mesothelioma  
37          unit risk (Table 5-53 shows this as 0.122) would be slightly less (about 3%).

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1 Taking both effects into consideration, the uncertainty in the IUR is considered to  
2 be low.

3 *10) Uncertainty in combining mortality risks to derive a composite cancer mortality IUR*  
4 *(see Section 5.4.6.1.10)*

- 5 • EPA assumed that the cancer risks were independent, conducted a bounding  
6 analysis, and showed the related uncertainty to be very low.

7 *11) Uncertainty due to extrapolation of findings in adults to children (see*  
8 *Section 5.4.6.1.11)*

- 9 • There is uncertainty in the assumption that risks are independent of age and that  
10 children are at the same exposure-related risk as adults. The lack of published  
11 information on cancer risks associated with exposures during childhood remains  
12 an uncertainty of unclear magnitude.



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