

Integrated Science Assessment for Sulfur Oxides – Health Criteria

(Second External Review Draft)

Integrated Science Assessment for Sulfur Oxides – Health Criteria

National Center for Environmental Assessment-RTP Division
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Abbreviations and Acronyms

A	Alpha
ACS	American Cancer Society
ADS	annular denuder system
AHR	airways hyperreactiveness
AM	alveolar macrophages
APHEA	Air Pollution on Health: a European Approach (study)
APEX	Air Pollution Exposure (model)
APIMS	atmospheric pressure ionization mass spectrometer
ARIC	Atherosclerosis Risk in Communities (study)
ARP	Acid Rain Program
AQCD	Air Quality Criteria Document
asl	above sea level
atm	Atmosphere
β	beta; the calculated Health Effect Parameter
B[a]P	benzo[a]pyrene
BHR	bronchial hyperresponsiveness
BS	black smoke
CAMP	Childhood Asthma Management Program
CARB	California Air Resources Board
CASAC	Clean Air Scientific Advisory Committee
CASTNet	Clean Air Status and Trends Network
CDC	Centers for Disease Control and Prevention
CHAD	Consolidated Human Activities Database
CHF	congestive heart failure
CHS	Children's Health Study
CH ₃ -S-H	methyl mercaptan
CH ₃ -S-S-CH ₃	dimethyl disulfide
CI	confidence interval
CMSA	consolidated metropolitan statistical area
CO	carbon monoxide
CoH	coefficient of haze
CONUS	continental United States
COPD	chronic obstructive pulmonary disease
CS ₂	carbon disulfide
CVD	cardiovascular disease
DEN	diethylnitrosamine
DEP	diesel exhaust particle
DMS	dimethyl sulfide
ED	emergency department
ECG	electrocardiography; electrocardiogram
EIB	exercise-induced bronchial reactivity
ELF	epithelial lining fluid

EMECAM	Spanish Multicentre Study on Air Pollution and Mortality
EPA	U.S. Environmental Protection Agency
eNO	exhaled nitric oxide
ET	extrathoracic
Fe	iron
FEMs	Federal Equivalent Methods
FEV _{0.75}	forced expiratory volume in 0.75 second
FEV ₁	forced expiratory volume in 1 second
FPD	flame photometric detection
FPD-TA	flame photometric detection-thermal analysis
FRM	Federal Reference Method
FVC	forced vital capacity
GAM	Generalized Additive Model(s)
GIS	Geographic Information System
GLM	Generalized Linear Model(s)
GSH	glutathione; reduced glutathione
GST	glutathione S-transferase (e.g., GSTM1, GSTP1, GSTT1)
H ⁺	hydrogen ion
HEADS	Harvard-EPA Annular Denuder System
HEI	Health Effects Institute
HF	high frequency
HNO ₂	nitrous acid
HNO ₃	nitric acid
HO ₂	hydroperoxyl; hydroperoxy radical
H ₂ O	water
H ₂ O ₂	hydrogen peroxide
HR	heart rate
HRV	heart rate variability
H ₂ S	hydrogen sulfide
HSO ₃ ⁻	hydrogen sulfite, bisulfite
HSO ₄ ⁻	bisulfate ion
H ₂ SO ₄	sulfuric acid
hν	solar ultraviolet photon
IARC	International Agency for Research on Cancer
ICD9	International Classification of Diseases, Ninth Revision
ICDs	implanted cardioverter defibrillators
Ig	immunoglobulin (e.g., IgA, IgE, IgG)
IHD	ischemic heart disease
IIASA	International Institute for Applied Systems Analysis
IL	interleukin (e.g., IL-4, IL-6, IL-8)
IOM	Institute of Medicine
IQR	interquartile range
ISA	Integrated Science Assessment
ISAAC	International Study of Asthma and Allergies in Children

IUGR	intrauterine growth retardation
K	mass transfer coefficient
LF	low frequency
LOD	limit of detection
LRD	lower respiratory disease
MCh	methacholine
MENTOR	Modeling Environment for Total Risk for One-Atmosphere studies
MI	myocardial infarction
MEF _{50%}	maximal midexpiratory flow at 50% of forced vital capacity
MMEF	maximal midexpiratory flow
Mn	manganese
MONICA	Monitoring Trend and Determinants in Cardiovascular Disease (registry)
MOZART-2	Model for Ozone and Related Chemical Tracers, version 2
MSA	metropolitan statistical area
N, n	number of observations
NAAQS	National Ambient Air Quality Standards
NaCl	sodium chloride
NaCO ₃	sodium carbonate
NADP	National Atmospheric Deposition Program
NAMS	National Air Monitoring Stations
NAPAP	National Acid Precipitation Assessment Program
NAS	National Academy of Sciences
NCAR	National Center for Atmospheric Research
NCEP	National Center for Environmental Prediction
NCICAS	National Cooperative Inner-City Asthma Study
NCore	National Core Monitoring Network
NERL	National Exposure Research Laboratory
NH ₄ ⁺	ammonium ion
NHAPS	National Human Activity Pattern Survey
NHANES	National Health and Nutrition Examination Survey
NMMAPS	National Morbidity, Mortality, and Air Pollution Study
NO	nitric oxide
NO ₂	nitrogen dioxide
NO ₃ ⁻	nitrate radical
NO ₃ ⁻	nitrate ion
NOAA	National Oceanic and Atmospheric Administration
NO _x	oxides of nitrogen
NR	not reported
NRC	National Research Council
NTN	National Trends Network
NTP	National Toxicology Program
O ₂	molecular oxygen, diatomic oxygen
O ₃	ozone
OCS	carbonyl sulfide

OH	hydroxyl radical
OR	odds ratio
P, p	probability value
PAARC	Air Pollution and Chronic Respiratory Diseases (study)
PAH	polycyclic aromatic hydrocarbon
PC(SO ₂)	provocative concentration of SO ₂ that produces a 100% increase in specific airway resistance
PD20FEV ₁	20% decrease in forced expiratory volume in 1 second
PD20	provocative dose that produces a 20% decrease in FEV ₁
PD100	provocative dose that produces a 100% increase in sRAW
PEACE	Pollution Effects on Asthmatic Children in Europe (study)
PEC	pulmonary endocrine cell
PEF	peak expiratory flow
PEMs	personal exposure monitors
PF	pulsed fluorescence
PM	particulate matter
PM _{2.5}	particulate matter with 50% upper cut point aerodynamic diameter of 2.5 µm for sample collection; surrogate for fine PM
PM ₁₀	particulate matter with 50% upper cut point aerodynamic diameter of 10 µm for sample collection
PM _{10-2.5}	particulate matter with 10 µm as upper cut point aerodynamic diameter and 2.5 µm as lower cut point for sample collection; surrogate for thoracic coarse PM (does not include fine PM)
PM ₁₃	particulate matter with 50% upper cut point aerodynamic diameter of 13 µm for sample collection
PMT	photomultiplier tube
ppb	parts per billion
ppbv	parts per billion by volume
ppm	parts per million
pptv	parts per trillion by volume
PRB	policy relevant background
PS	passive sample
R, r	correlation coefficient
RAR	rapidly activating receptor
RAS	roll-around system
Raw	airway resistance
RH	relative humidity
r-MSSD	root mean square of successive differences in R-R intervals.
RR	rate ratio; relative risk
S ²⁻	sulfur radical
SAB	Science Advisory Board
SAPALDIA	Study of Air Pollution and Lung Diseases in Adults
SAVIAH	Small-Area Variation in Air Pollution and Health (study)
SD	standard deviation
SDNN	standard deviation of normal R-R intervals
SES	socioeconomic status

SHEDS	Simulation of Human Exposure and Dose System
SIDS	sudden infant death syndrome
SNP	single nucleotide polymorphism
³⁵ S	sulfur-35 radionuclide
SLAMS	State and Local Air Monitoring Stations
SO	sulfur monoxide
SO ₂	sulfur dioxide
SO ₃	sulfur trioxide
SO ₃ ²⁻	sulfite ion
SO ₄ ²⁻	sulfate ion
SO _x	sulfur oxides
S ₂ O	disulfur monoxide
SPM	suspended particulate matter
sRaw	specific airway resistance
STN	Speciation Trends Network
τ	tau; atmospheric lifetime
TBARS	thiobarbituric acid reactive substances
TEA	triethanolamine
TNF	tumor necrosis factor (e.g., TNF-α)
TSP	total suspended particles
URI	upper respiratory infections
UV	ultraviolet
\dot{V}_E	minute ventilation

Authors, Contributors, Reviewers

Authors

Dr. Jee Young Kim (SO_x Team Leader)—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Jeffrey Arnold—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. James S. Brown—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Barbara Buckley—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Ila Cote—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Douglas Johns—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Ellen Kirrane—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Thomas Long—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Thomas Luben—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Qingyu Meng—Oak Ridge Institute for Science and Education, Postdoctoral Research Fellow to National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Anu Mudipalli—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Joseph Pinto—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Mary Ross—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. David Svendsgaard—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Lori White—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. William Wilson—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Brett Grover—National Exposure Research Laboratory, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Douglas Bryant—Intrinsik Science, Mississauga, Ontario, Canada

Dr. Arlene Fiore—Geophysical Fluid Dynamics Laboratory/National Oceanographic & Atmospheric Administration, Princeton, NJ

Dr. Panos Georgopoulos—Computational Chemodynamics Laboratory, Environmental and Occupational Health Sciences Institute, Piscataway, NJ

Dr. Vic Hasselblad—Duke University Medical Center, Durham, NC

Dr. Larry Horowitz—Geophysical Fluid Dynamics Laboratory/National Oceanographic and Atmospheric Administration, Princeton University Forrestal Campus, Princeton, NJ

Ms. Annette Ianucci—Sciences International, Alexandria, VA

Dr. Kazuhiko Ito—Department of Environmental Medicine, New York University School of Medicine, Tuxedo, NY

Dr. Jane Koenig—Department of Environmental and Occupational Health Sciences, University of Washington, Seattle, WA

Dr. Therese Mar—Department of Environmental and Occupational Health Sciences, University of Washington, Seattle, WA

Dr. James Riddle—Sciences International, Alexandria, VA

Contributors

Ms. Rebecca Daniels—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Mr. Jason Sacks—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Dale Allen—Department of Atmospheric and Oceanic Sciences, University of Maryland, College Park, MD

Ms. Louise Camalier—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Russell Dickerson—Department of Atmospheric and Oceanic Science, University of Maryland, College Park, MD

Dr. Tina Fan—Environmental and Occupational Health Sciences Institute, Piscataway, NJ

Mr. William Keene—Department of Environmental Sciences, University of Virginia, Charlottesville, VA

Dr. Randall Martin—Department of Physics and Atmospheric Science, Dalhousie University, Halifax, Nova Scotia, Canada

Dr. Maria Morandi—Department of Environmental Sciences, School of Public Health, University of Texas – Houston Health Science Center, Houston, TX

Dr. William Munger—Division of Engineering and Applied Sciences, Harvard University, Cambridge, MA

Mr. Charles Piety—Department of Meteorology, University of Maryland, College Park, MD

Dr. Sandy Sillman—Department of Atmospheric, Ocean, and Space Sciences, University of Michigan, Ann Arbor, MI

Dr. Helen Suh—Department of Environmental Health, Harvard School of Public Health, Boston, MA

Dr. Charles Wechsler—Environmental and Occupational Health Sciences Institute, Piscataway, NJ

Dr. Clifford Weisel—Environmental and Occupational Health Sciences Institute, Piscataway, NJ

Dr. Jim Zhang—Environmental and Occupational Health Sciences Institute, Piscataway, NJ

Reviewers

Dr. Tina Bahadori—American Chemistry Council, Arlington, VA

Dr. Tim Benner—Office of Science Policy, U.S. Environmental Protection Agency, Washington, DC

Dr. Daniel Costa—National Program Director for Air, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Robert Devlin—National Health and Environmental Effects Research Laboratory, U.S. Environmental Protection Agency, Chapel Hill, NC

Dr. Judy Graham—American Chemistry Council, Arlington, VA

Dr. Stephen Graham—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Ms. Beth Hassett-Sipple—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Gary Hatch—National Health and Environmental Effects Research Laboratory, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Scott Jenkins—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. David Kryak—National Exposure Research Laboratory, U.S. Environmental Protection Agency, Research Triangle Park, NC

Mr. John Langstaff—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Morton Lippmann—Department of Environmental Medicine, New York University School of Medicine, Tuxedo, NY

Dr. Karen Martin—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. William McDonnell—William F. McDonnell Consulting, Chapel Hill, NC

Dr. Dave McKee—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Lucas Neas—National Health and Environmental Effects Research Laboratory, U.S. Environmental Protection Agency, Chapel Hill, NC

Dr. Russell Owen—National Health and Environmental Effects Research Laboratory, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Halûk Özkaynak—National Exposure Research Laboratory, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Jennifer Peel—Department of Environmental and Radiological Health Sciences, Colorado State University, Fort Collins, CO

Mr. Harvey Richmond—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Mr. Steven Silverman—Office of General Counsel, U.S. Environmental Protection Agency, Washington, DC

Dr. Michael Stewart—Office of Air Quality Planning and Standards, U.S. Environmental Protection Agency, Research Triangle Park, NC

Ms. Susan Stone—Office of Air Quality Planning and Standards, Office of Air and Radiation, U.S. Environmental Protection Agency, Research Triangle Park, NC

Ms. Chris Trent—Office of Air Quality Planning and Standards, Office of Air and Radiation, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. John Vandenberg—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Alan Vette—National Exposure Research Laboratory, U.S. Environmental Protection Agency, Research Triangle Park, NC

Ms. Debra Walsh—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Mr. Ron Williams—National Exposure Research Laboratory, U.S. Environmental Protection Agency, Research Triangle Park, NC

SO_x Project Team

Executive Direction

Dr. Ila Cote (Acting Director)—National Center for Environmental Assessment-RTP Division, U.S. Environmental Protection Agency, Research Triangle Park, NC

Ms. Debra Walsh (Deputy Director)—National Center for Environmental Assessment-RTP Division, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Mary Ross—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Scientific Staff

Dr. Jee Young Kim (SO_x Team Leader)—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Jeff Arnold—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. James S. Brown—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Barbara Buckley—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Douglas Johns—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Ellen Kirrane—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Dennis Kotchmar—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Thomas Long—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Thomas Luben—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Qingyu Meng—Oak Ridge Institute for Science and Education, Postdoctoral Research Fellow to National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Joseph Pinto—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Mr. Jason Sacks—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. David Svendsgaard—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. Lori White—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Dr. William Wilson—National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Technical Support Staff

Ms. Emily R. Lee—Management Analyst, National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Ms. Ellen Lorang—Information Manager, National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Ms. Christine Searles—Management Analyst, National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Mr. Richard Wilson—Clerk, National Center for Environmental Assessment, U.S. Environmental Protection Agency, Research Triangle Park, NC

Document Production Staff

Ms. Barbra H. Schwartz—Task Order Manager, Computer Sciences Corporation, Morrisville, NC

Mr. John A. Bennett—Technical Information Specialist, Library Associates of Maryland, Rockville, MD

Mr. David Casson—Publication/Graphics Specialist, TekSystems, Raleigh, NC

Ms. Melissa Cesar—Publication/Graphics Specialist, Computer Sciences Corporation, Morrisville, NC

Ms. Rebecca Early—Publication/Graphics Specialist, TekSystems, Raleigh, NC

Mr. Eric Ellis—Records Management Technician, InfoPro, Inc., McLean, VA

Ms. Kristin Hamilton—Publication/Graphics Specialist, TekSystems, Raleigh, NC

Ms. Stephanie Harper—Publication/Graphics Specialist, TekSystems, Raleigh, NC

Ms. Sandra L. Hughey—Technical Information Specialist, Library Associates of Maryland, Rockville, MD

Dr. Barbara Liljequist—Technical Editor, Computer Sciences Corporation, Morrisville, NC

Ms. Molly Windsor—Graphic Artist, Computer Sciences Corporation, Morrisville, NC

Clean Air Scientific Advisory Committee for NO_x and SO_x Primary NAAQS

Chairperson

Dr. Rogene Henderson*, Scientist Emeritus, Lovelace Respiratory Research Institute, Albuquerque, NM

Members

Mr. Ed Avol, Professor, Preventive Medicine, Keck School of Medicine, University of Southern California, Los Angeles, CA

Dr. John R. Balmes, Professor, Department of Medicine, Division of Occupational and Environmental Medicine, University of California, San Francisco, CA

Dr. Ellis Cowling*, University Distinguished Professor At-Large, North Carolina State University, Colleges of Natural Resources and Agriculture and Life Sciences, North Carolina State University, Raleigh, NC

Dr. James D. Crapo*, Professor, Department of Medicine, National Jewish Medical and Research Center, Denver, CO

Dr. Douglas Crawford-Brown*, Director, Carolina Environmental Program; Professor, Environmental Sciences and Engineering; and Professor, Public Policy, Department of Environmental Sciences and Engineering, University of North Carolina at Chapel Hill, Chapel Hill, NC

Dr. Terry Gordon, Professor, Environmental Medicine, NYU School of Medicine, Tuxedo, NY

Dr. Dale Hattis, Research Professor, Center for Technology, Environment, and Development, George Perkins Marsh Institute, Clark University, Worcester, MA

Dr. Patrick Kinney, Associate Professor, Department of Environmental Health Sciences, Mailman School of Public Health, Columbia University, New York, NY

Dr. Steven Kleeberger, Professor, Laboratory Chief, Laboratory of Respiratory Biology, NIH/NIEHS, Research Triangle Park, NC

Dr. Timothy Larson, Professor, Department of Civil and Environmental Engineering, University of Washington, Seattle, WA

* Members of the statutory Clean Air Scientific Advisory Committee (CASAC) appointed by the EPA Administrator

Preface

Legislative Requirements

1 Section 109 (42 U.S. Code, 2003b) directs the Administrator to propose and promulgate
2 “primary” and “secondary” National Ambient Air Quality Standards (NAAQS) for pollutants
3 listed under section 108. Section 109(b)(1) defines a primary standard as one “the attainment and
4 maintenance of which in the judgment of the Administrator, based on such criteria and allowing
5 an adequate margin of safety, are requisite to protect the public health.”¹ A secondary standard,
6 as defined in section 109(b)(2), must “specify a level of air quality the attainment and
7 maintenance of which, in the judgment of the Administrator, based on such criteria, is required to
8 protect the public welfare from any known or anticipated adverse effects associated with the
9 presence of [the] pollutant in the ambient air.”² The requirement that primary standards include
10 an adequate margin of safety was intended to address uncertainties associated with inconclusive
11 scientific and technical information available at the time of standard setting. It was also intended
12 to provide a reasonable degree of protection against hazards that research has not yet identified.
13 See *Lead Industries Association v. EPA*, 647 F.2d 1130, 1154 (D.C. Cir 1980), cert. denied, 449
14 U.S. 1042 (1980); *American Petroleum Institute v. Costle*, 665 F.2d 1176, 1186 (D.C. Cir. 1981)
15 cert. denied, 455 U.S. 1034 (1982). Both kinds of uncertainties are components of the risk
16 associated with pollution at levels below those at which human health effects can be said to
17 occur with reasonable scientific certainty. Thus, in selecting primary standards that include an
18 adequate margin of safety, the Administrator is seeking not only to prevent pollution levels that
19 have been demonstrated to be harmful but also to prevent lower pollutant levels that may pose an
20 unacceptable risk of harm, even if the risk is not precisely identified as to nature or degree.

21 In selecting a margin of safety, the EPA considers such factors as the nature and severity of
22 the health effects involved, the size of sensitive population(s) at risk, and the kind and degree of
23 the uncertainties that must be addressed. The selection of any particular approach to providing an
24 adequate margin of safety is a policy choice left specifically to the Administrator’s judgment. See
25 *Lead Industries Association v. EPA*, supra, 647 F.2d at 1161-62.

¹ The legislative history of section 109 indicates that a primary standard is to be set at “the maximum permissible ambient air level . . . which will protect the health of any [sensitive] group of the population,” and that for this purpose “reference should be made to a representative sample of persons comprising the sensitive group rather than to a single person in such a group” [S. Rep. No. 91-1196, 91st Cong., 2d Sess. 10 (1970)]. (Senate., 1970)

² Welfare effects as defined in section 302(h) [42 U.S.C. 7602(h)] include, but are not limited to, “effects on soils, water, crops, vegetation, man-made materials, animals, wildlife, weather, visibility and climate, damage to and deterioration of property, and hazards to transportation, as well as effects on economic values and on personal comfort and well-being.”

1 In setting standards that are “requisite” to protect public health and welfare, as provided in
2 section 109(b), EPA’s task is to establish standards that are neither more nor less stringent than
3 necessary for these purposes. In so doing, EPA may not consider the costs of implementing the
4 standards. See generally *Whitman v. American Trucking Associations*, 531 U.S. 457, 465-472,
5 475-76 (D.C. Cir. 2001).

6 Section 109(d)(1) requires that “not later than December 31, 1980, and at 5-year intervals
7 thereafter, the Administrator shall complete a thorough review of the criteria published under
8 section 108 and the national ambient air quality standards...and shall make such revisions in
9 such criteria and standards and promulgate such new standards as may be appropriate...” Section
10 109(d)(2) requires that an independent scientific review committee “shall complete a review of
11 the criteria...and the national primary and secondary ambient air quality standards...and shall
12 recommend to the Administrator any new...standards and revisions of existing criteria and
13 standards as may be appropriate...” Since the early 1980s, this independent review function has
14 been performed by the Clean Air Scientific Advisory Committee (CASAC) of EPA’s Science
15 Advisory Board.

History of Reviews of the Primary NAAQS for Sulfur Oxides

16 On April 30, 1971, the EPA promulgated primary NAAQS for sulfur oxides (SO_x). These
17 primary standards, which were based on the findings outlined in the original 1969 Air Quality
18 Criteria for Sulfur Oxides, were set at 0.14 parts per million (ppm) averaged over a 24-hour
19 period, not to be exceeded more than once per year, and 0.030 ppm annual arithmetic mean with
20 SO₂ as the indicator. In 1982, EPA published the *Air Quality Criteria for Particulate Matter and*
21 *Sulfur Oxides* (EPA, 1982) along with an addendum of newly published controlled human
22 exposure studies, which updated the scientific criteria upon which the initial standards were
23 based. In 1986, a second addendum was published presenting newly available evidence from
24 epidemiologic and controlled human exposure studies (EPA, 1986b). In 1988, EPA published a
25 proposed decision not to revise the existing standards (53 FR 14926). However, EPA specifically
26 requested public comment on the alternative of revising the current standards and adding a new
27 1-hour primary standard of 0.4 ppm.

28 As a result of public comments on the 1988 proposal and other post-proposal
29 developments, EPA published a second proposal on November 15, 1994 (59 FR 58958). The

1 1994 re-proposal was based in part on a supplement to the second addendum of the criteria
2 document, which evaluated new findings on short-term SO₂ exposures in asthmatics (EPA,
3 1994b). As in the 1988 proposal, EPA proposed to retain the existing 24-hour and annual
4 standards. The EPA also solicited comment on three regulatory alternatives to further reduce the
5 health risk posed by exposure to high 5-minute peaks of SO₂ if additional protection were judged
6 to be necessary. The three alternatives were: 1) revising the existing primary SO₂ NAAQS by
7 adding a new 5-minute standard of 0.60 ppm SO₂; 2) establishing a new regulatory program
8 under section 303 of the Act to supplement protection provided by the existing NAAQS, with a
9 trigger level of 0.60 ppm SO₂, one expected exceedance; and 3) augmenting implementation of
10 existing standards by focusing on those sources or source types likely to produce high 5-minute
11 peak concentrations of SO₂. On May 22, 1996, EPA's final decision, that revisions of the NAAQS
12 for SO_x were not appropriate at that time, was announced in the Federal Register (61 FR 25566).
13 In that decision, EPA announced an intention to propose guidance, under section 303 of the Act,
14 to assist states in responding to short-term peak levels of SO₂. The basis for the decision, and
15 subsequent litigation, is discussed in Annex A.

Chapter 1. Introduction

This second external review draft Integrated Science Assessment (ISA) presents a concise synthesis of the most policy-relevant science to form the scientific foundation for the review of the primary (health-based) NAAQS for SO_x. This document is intended to “accurately reflect the latest scientific knowledge useful in indicating the kind and extent of identifiable effects on public health which may be expected from the presence of [a] pollutant in ambient air” (Clean Air Act, Section 108, 2003a)¹. Contained herein are the key information and judgments formerly contained in the Air Quality Criteria Document (AQCD) for SO_x; additional details of the pertinent scientific literature published since the last review, as well as selected older studies of particular interest, are included in a series of annexes to the draft ISA. This second external review draft ISA thus serves to update and revise the information available at the time of the previous review of the NAAQS for SO_x in 1996.

SO₂ is the most important of the monomeric sulfur oxides (SO_x) for both atmospheric chemistry and health effects. SO_x is usually defined to include SO₃ and H₂SO₄ as well, but neither is present in the atmosphere in concentrations significant for human exposures. Descriptions of the atmospheric chemistry of SO_x include both gaseous and particulate species; a meaningful analysis would not be possible otherwise. Most studies on the health effects of gaseous SO_x focus on SO₂; effects of other gaseous species are considered as information is available. The health effects of particulate SO_x are included in the review of the NAAQS for particulate matter (PM). In evaluating the health evidence, this second external draft ISA considers possible influences of other atmospheric pollutants, including interactions of SO₂ with other co-occurring pollutants such as PM, nitrogen oxides (NO_x), carbon monoxide (CO), and ozone (O₃).

As discussed in the Integrated Plan for Review of the Primary NAAQS for SO_x (EPA, 2007), a series of policy-relevant questions frames this review to provide a scientific basis for a decision about whether the current primary NAAQS for SO_x should be retained or revised. The primary NAAQS for SO_x, with SO₂ serving as the indicator, is set at 0.14 parts per million (ppm), averaged over a 24-h period, not to be exceeded more than once per year, and 0.030 ppm annual arithmetic mean. This second external review draft ISA focuses on evaluation of the newly

¹A review of the secondary SO_x NAAQS, in conjunction with a review of the secondary NAAQS for NO_x, is underway independently, as is a review of the primary NAAQS for NO_x and a review of the primary and secondary effects of PM.

1 available scientific evidence to best inform consideration of these framing questions, including
2 the following:

- 3 ▪ How has new information altered/substantiated the scientific support for the occurrence
4 of health effects following short- and/or long-term exposure to levels of SO_x found in
5 the ambient air?
- 6 ▪ How does new information influence conclusions from the previous review regarding the
7 effects of SO_x on susceptible populations?
- 8 ▪ At what levels of SO_x exposure do health effects of concern occur?
- 9 ▪ How has new information altered conclusions from previous reviews regarding the
10 plausibility of adverse health effects caused by SO_x exposure?
- 11 ▪ To what extent have important uncertainties identified in the last review been reduced?
12 Have new uncertainties emerged?
- 13 ▪ What are the air quality relationships between short-term and long-term exposures
14 to SO_x?

1.1. Document Development

15 EPA initiated the current formal review of the NAAQS for SO_x on May 15, 2006 with a
16 call for information from the public (FR, 2006). In addition to the call for information,
17 publications are identified through an ongoing literature search process that includes extensive
18 computer database mining on specific topics. Additional publications were identified by EPA
19 scientists in a variety of disciplines by combing through relevant, peer-reviewed scientific
20 literature obtained through these ongoing literature searches, reviewing previous EPA reports,
21 and a review of reference lists from important publications. All relevant epidemiological, human
22 clinical, and animal toxicological studies, including those related to exposure-response
23 relationships, mechanism(s) of action, or susceptible subpopulations published since the last
24 review were considered. Added to the body of research were EPA's analyses of air quality and
25 emissions data, studies on atmospheric chemistry, transport, and fate of these emissions, as well
26 as issues related to exposure to SO_x. Further information was acquired from consultation with
27 content and area experts and the public. Annex A has more discussion of search strategies and
28 criteria for study selection.

1.2. Document Organization

This second external review draft ISA is composed of five chapters. This introductory chapter presents background information, discusses the purpose of the document, and characterizes the search, evaluation and retrieval process of policy-relevant scientific studies. Chapter 2 highlights key concepts or issues relevant to understanding the atmospheric chemistry, sources, exposure, and dosimetry of SO_x, following a “source-to-dose” paradigm. Chapter 3 evaluates and integrates epidemiological, human clinical, and animal toxicological information relevant to the review of the primary NAAQS for SO_x. Chapter 4 has information related to the public health impact of ambient SO_x exposure, with emphasis on potentially susceptible and vulnerable population groups. Finally, Chapter 5 summarizes key findings and conclusions from the atmospheric sciences, ambient air data analyses, exposure assessment, dosimetry, and health effects for consideration in the review of the NAAQS for SO_x.

A series of annexes supplement this second external review draft ISA. The annexes provide additional details of the pertinent literature published since the last review, as well as selected older studies of particular interest. These annexes contain information on:

- atmospheric chemistry of SO_x as well as the sampling and analytic methods for measurement of SO_x¹;
- environmental concentrations and human exposure to SO_x;
- toxicological studies of health effects in laboratory animals;
- human clinical studies of health effects related to peak (5-10 min) and short-term (1-h or longer) exposure to SO_x; and
- epidemiological studies of health effects from short- and long-term exposure to SO_x.

Detailed information about methods and results of health studies is summarized in tabular format, and generally includes information about: concentrations of SO_x and averaging times; study methods employed; results and comments; and quantitative results for relationships between effects and exposure to SO_x.

¹ This section also includes information on NO₂, in order to support the reviews of the primary and secondary NAAQS for both SO₂ and NO₂. The atmospheric chemistry of NO_x and SO_x are intricately linked; discussion of their combined chemistry is more effective and more efficient than a separate discussion of each pollutant.

1.3. EPA Framework for Causal Determinations

It is important to have a consistent and transparent basis to evaluate the causal nature of air pollution-induced health effects. The framework described below establishes uniform language concerning causality and brings more specificity to the findings. It draws standardized language from across the Federal government and wider scientific community, especially from the recent National Academy of Sciences (NAS) Institute of Medicine (IOM) document, *Improving the Presumptive Disability Decision-Making Process for Veterans* (IOM, 2007), the most recent comprehensive work on evaluating the causality of health effects. This section:

- describes the kinds of scientific evidence used in establishing a general causal relationship between exposure and health effects;
- defines cause, in contrast to statistical association;
- discusses the sources of evidence necessary to reach a conclusion about the existence of a causal relationship;
- highlights the issue of multifactorial causation;
- identifies issues and approaches related to uncertainty; and
- provides a framework for classifying and characterizing the weight of evidence in support of a general causal relationship.

Approaches to assessing the separate and combined lines of evidence (e.g., epidemiological, human clinical, animal toxicological, and in vitro studies) have been formulated by a number of regulatory and science agencies, including the IOM of the National Academies of Science (IOM, 2008), International Agency for Research on Cancer (IARC, 2006), EPA Guidelines for Carcinogen Risk Assessment (EPA, 2005), Centers for Disease Control and Prevention (CDC, 2004), and National Acid Precipitation Assessment Program (NAPAP, 1991). Highlights or excerpts from the various decision framework documents are included in Annex A.

These formalized approaches offer guidance for assessing causality. The frameworks are similar in nature, although adapted to different purposes, and have proved effective in providing a uniform structure and language for causal determinations. Moreover, these frameworks must support decision-making under conditions of uncertainty.

1.3.1. Scientific Evidence Used in Establishing Causality

1 The most compelling evidence of a causal relationship between pollutant exposures and
2 human health effects comes from human clinical studies. This type of study experimentally
3 evaluates the health effects of administered exposures in humans under highly-controlled
4 laboratory conditions.

5 In epidemiological or observational studies of humans, the investigator does not control
6 exposures or intervene with the study population. Broadly, observational studies can describe
7 associations between exposures and effects. These studies fall into several categories: cross-
8 sectional, prospective cohort, time-series, and panel studies. “Natural experiments” occur
9 occasionally in epidemiology; these include comparisons of health effects before and after a
10 change in population exposures, such as closure of a pollution source.

11 Experimental animal data complements the clinical and observational data; these studies
12 can help characterize effects of concern, exposure-response relationships, sensitive
13 subpopulations and modes of action. In the absence of clinical or epidemiological data, animal
14 data alone may be sufficient to support a likely causal determination, assuming that humans
15 respond similarly to the experimental species.

1.3.2. Association and Causation

16 Association and causation are not the same. “Cause” conveys the notion of a significant,
17 effectual relationship between an agent and an associated disorder or disease in the population.
18 “Association” is the statistical dependence among multiple (two or more) events, characteristics,
19 or other variables. An association is merely prima facie evidence for causation; alone, it is not
20 sufficient for proof of a causal relationship between exposure and disease. Unlike an association,
21 a causal claim supports the creation of counterfactual claims; that is, a claim about what the
22 world would have been like under different or changed circumstances (IOM, 2008). Currently,
23 much of the newly available health information evaluated in the draft ISA comes from
24 epidemiological studies that report a statistical association between exposure and health
25 outcome.

26 It is recognized that many of the health outcomes evaluated in ISAs have complex
27 etiologies. Most diseases, such as cancer or coronary heart disease, result from a complex web of
28 causation, whereby one or more agents can initiate a disease process. The outcome could depend

on many factors, including age, genetic susceptibility, nutritional status, immune competence, social factors, and others (Gee and Payne-Sturges, 2004; IOM, 2008). Figure 1-1 shows a diagram of a variety of etiologic factors that contribute to disease. Exposure to multiple agents together could result in synergistic or antagonistic effects that are different from what might result from exposure to each agent separately.¹ The results are the net effect of many actions and counteractions.

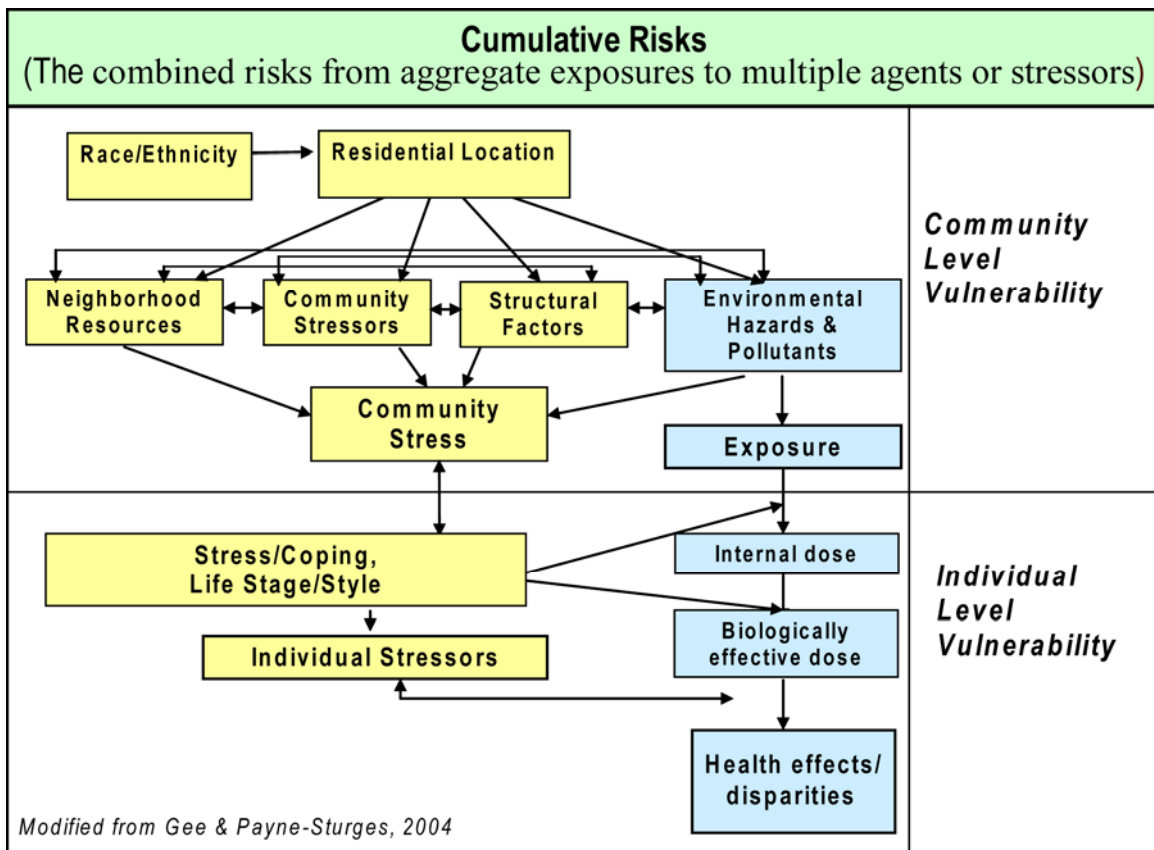


Figure 1-1. Exposure-disease-stress model for environmental health disparities.

1.3.3. Evidence for Going beyond Association to Causation

Moving from association to causation involves elimination of alternative explanations for the association. Human clinical studies are experiments in which subjects in a population are

¹ For example, a multiplicative interaction relative risk (RR) could be defined as $RR_{Int(mult)} = RR_{joint} / RR_E \times RR_S$. An additive interaction RR could be defined as $RR_{Int(add)} = RR_{joint} - RR_E - RR_S + 1$

1 randomly allocated into groups, usually called study and control groups, and exposed to a
2 pollutant or a sham. The results are assessed by rigorous comparison of rates of appropriate
3 outcomes between the study and control groups. Randomized human clinical studies are
4 generally regarded as the most scientifically rigorous method of hypothesis testing available. By
5 assigning exposure randomly, the study design attempts to remove the effect of any factor that
6 might influence exposure. Done properly, and setting aside randomness, only a causal
7 relationship between exposure and health outcome should produce observed associations in
8 randomized clinical trials. In another type of human clinical study, the same subject is exposed to
9 a pollutant and a sham at different time points, and the responses to the two types of exposures
10 are compared. This study design is also effective at controlling for any potential confounders,
11 since the subject is serving as his/her own control. A lack of observation of effects from human
12 clinical studies does not necessarily mean that a causal relationship does not occur. Human
13 clinical studies are often limited because the study population is generally small, which restricts
14 the ability to discern statistically significant findings. In addition, the most susceptible
15 individuals may be explicitly excluded (for ethical reasons), and other susceptible individuals or
16 groups, such as those with nutritional deficits, may not be included.

17 Inferring causation from epidemiological studies requires consideration of potential
18 confounders. When associations are found in epidemiological studies, one approach to remove
19 spurious association from possible confounders is statistical control on characteristics that may
20 differ between exposed and unexposed persons; this is frequently termed “adjustment.”
21 Multivariable regression models constitute one tool for estimating the association between
22 exposure and outcome after adjusting for characteristics of participants that might confound the
23 results. Another way to adjust for potential confounding is through stratified analysis, i.e.,
24 examining the association within homogeneous groups with the confounding variable. The use of
25 stratified analyses has an additional benefit: it allows examination of effect modification through
26 comparison of the effect estimates across different groups. If investigators successfully measured
27 characteristics that distort the results, adjustment of these factors help separate a spurious from a
28 true causal association. Appropriate statistical adjustment for confounders requires identifying
29 and measuring all reasonably expected confounders. Deciding which variables to control for in a
30 statistical analysis of the association between exposure and disease depends on knowledge about

possible mechanisms. Identifying these mechanisms makes it possible to control for potential sources that may result in a spurious association.

Measurement error is another problem encountered when adjusting for spurious associations. In multivariate analyses, the effects of a well-measured covariate may be overestimated, in contrast to a more poorly measured covariate. There are several components that contribute to exposure measurement error in these studies, including the difference between true and measured ambient concentrations, the difference between average personal exposure to ambient pollutants and ambient concentrations at central monitoring sites, and the use of average population exposure rather than individual exposure estimates.

It is difficult to identify and measure all potential confounders in epidemiological studies. Confidence that unmeasured confounders are not producing the findings is increased when multiple studies are conducted in various settings using different subjects or exposures; each of which might eliminate another source of confounding from consideration. Thus, multi-city studies which use a consistent method to analyze data from across locations with different levels of covariates can provide insight on potential confounding in associations. The number and degree of diversity of covariates, as well as their relevance to the potential confounders, remain matters of scientific judgment. Intervention studies, because of their experimental nature, can be particularly useful in characterizing causation.

In addition to clinical and epidemiological studies, the tools of experimental biology have been valuable for developing insights into human physiology and pathology. Laboratory tools have been extended to explore the effects of putative toxicants on human health, especially through the study of model systems in other species. Background knowledge of the biological mechanisms by which an exposure might or might not cause disease can prove crucial in establishing, or negating, a causal claim. At the same time, species can differ from each other in fundamental aspects of physiology and anatomy (e.g., metabolism, airway branching, hormonal regulation) that may limit extrapolation. Testable hypotheses about the causal nature of proposed mechanisms or modes of action are central to utilizing experimental data in causal determinations.

1.3.4. Multifactorial Causation

Scientific judgment is needed regarding likely sources and magnitude of confounding, together with judgment about how well the existing constellation of study designs, results, and analyses address this potential threat to inferential validity. One key consideration in this review is evaluation of the potential contribution of SO_x to health effects, when it is a component of a complex air pollutant mixture. There are multiple ways by which SO_x might cause or be associated with adverse health effects. First, the reported SO_x effect estimates in epidemiological studies may reflect independent SO_x effects on respiratory health. Second, ambient SO_x may be serving as an indicator of complex ambient air pollution mixtures that share the same source as SO_x (i.e., combustion of sulfur-containing fuels or metal smelting). Finally, copollutants may mediate the effects of SO_x or SO_x may influence the toxicity of copollutants. Epidemiologists use the term “interaction” or “effect modification” to denote the departure of the observed joint risk from what might be expected based on the separate effects of the factors. These possibilities are not necessarily exclusive. In addition, confounding can result in the production of an association between adverse health effects and SO₂ that is actually attributable to another factor that is associated with SO₂ in a particular study. Multivariate models are the most widely used strategy to address confounding in epidemiological studies, but such models are not readily interpreted when assessing effects of covarying pollutants such as PM, SO₂, and nitrogen dioxide (NO₂).

1.3.5. Uncertainty

The science of estimating the causal influence of an exposure on disease is an uncertain one. There are two distinct levels of uncertainty to be considered here:

- Model uncertainty—uncertainty regarding gaps in scientific theory required to make predictions on the basis of causal inferences.
- Parameter uncertainty—uncertainty as to the statistical estimates within each model.

Assessment of model uncertainty involves:

- whether exposure causes the health outcome;
- the set of confounders associated with exposure and health outcome;

- which parametric forms best describe the relations of exposure and confounders with outcome; and
- whether other forms of bias could be affecting the association.

Model uncertainty is not limited to the qualitative causal structure: it also involves factors such as uncertainty about the parametric form of the model specified, the variables included and whether or not measurement error is modeled. When mechanistic knowledge exists, this important source of uncertainty can be reduced. In contrast, uncertainty about the parameter estimates (regression coefficients) for a given model is a well-studied problem. The important point is that these reports of uncertainty are conditional on the model providing a sufficiently adequate approximation of reality so that inferences are valid. The overall scientific inference involves evaluation of model uncertainty and uncertainty about parameter estimates given to each model.

There are systematic, quantitative approaches for including uncertainty about the model in an assessment of overall uncertainty about a causal inference, such as sensitivity analysis and model averaging. Sensitivity analysis attempts to quantify the sensitivity of the parameter estimate in relation to assumptions about the model. Uncertainty ranges can be estimated using classical analysis (Robinson, 1989) or the Monte Carlo technique (Eggleston, 1993). By averaging over many different competing models, Bayesian Model Averaging incorporates model uncertainty into conclusions about parameters and prediction.

1.3.6. Application of Framework

EPA uses a two-step approach to evaluate the scientific evidence on health effects of exposure to criteria pollutants. These two steps address two policy relevant questions noted in the beginning of this chapter – what are (if any) the effects of SO_x on susceptible populations, given the total body of evidence, and at what levels of SO_x exposure do health effects of concern occur. The first step determines the weight of evidence in support of causation and characterizes the strength of any resulting causal classification. The second step includes further evaluation of the quantitative evidence regarding the concentration-response relationships and the levels, duration and pattern of exposures at which effects are observed.

To aid judgment, various “aspects”¹ of causality have been discussed by many philosophers and scientists. The most widely cited aspects of causality in epidemiology, and public health in general, were articulated by Sir Austin Bradford Hill in 1965 and have been widely used (EPA, 2005; IARC, 2006; Surgeon General, 2004; IOM, 2008). These nine aspects (Hill, 1965) have been modified (below) for use in causal determinations specific to health and environmental effects and pollutant exposures.²

Table 1-1 Aspects to aid in judging causality.

1. ***Consistency of the observed association.*** An inference of causality is strengthened when a pattern of elevated risks is observed across several independent studies. The reproducibility of findings constitutes one of the strongest arguments for causality. If there are discordant results among investigations, possible reasons such as differences in exposure, confounding factors, and the power of the study are considered.
2. ***Strength of the observed association.*** The finding of large, precise risks increases confidence that the association is not likely due to chance, bias, or other factors. A modest risk, however, does not preclude a causal association and may reflect a lower level of exposure, an agent of lower potency, or a common disease with a high background level.
3. ***Specificity of the observed association.*** As originally intended, this refers to increased inference of causality if one cause is associated with a single effect or disease (Hill, 1965). Based on our current understanding this is now considered one of the weaker guidelines for causality; for example, many agents cause respiratory disease and respiratory disease has multiple causes. The ability to demonstrate specificity under certain conditions remains, however, a powerful attribute of experimental studies. Thus, although the presence of specificity may support causality, its absence does not exclude it.
4. ***Temporal relationship of the observed association.*** A causal interpretation is strengthened when exposure is known to precede development of the disease.

¹The “aspects” describe by Hill (1965) have become, in the subsequent literature, more commonly described as “criteria.” The original term “aspects” is used here to avoid confusion with ‘criteria’ as it is used, with different meaning, in the Clean Air Act.

² The Hill aspects were developed for use with epidemiology data. They have been modified here for use with a broader array of data, i.e., epidemiological, controlled human exposure, and animal toxicological studies, as well as *in vitro* data, and to be more consistent with EPA’s Guidelines for Carcinogen Risk Assessment.

5. **Biological gradient (exposure-response relationship).** A clear exposure-response relationship (e.g., increasing effects associated with greater exposure) strongly suggests cause and effect, especially when such relationships are also observed for duration of exposure (e.g., increasing effects observed following longer exposure times). There are, however, many possible reasons that a study may fail to detect an exposure-response relationship. Thus, although the presence of a biologic gradient may support causality, the absence of an exposure-response relationship does not exclude a causal relationship.
6. **Biological plausibility.** An inference of causality tends to be strengthened by consistency with data from experimental studies or other sources demonstrating plausible biological mechanisms. A lack of biologic understanding, however, is not a reason to reject causality.
7. **Coherence.** An inference of causality may be strengthened by other lines of evidence (e.g., clinical and animal studies) that support a cause-and-effect interpretation of the association. The absence of other lines of evidence, however, is not a reason to reject causality.
8. **Experimental evidence (from human populations).** Experimental evidence is generally available from human populations for the criteria pollutants. The strongest evidence for causality can be provided when a change in exposure brings about a change in adverse health effect or disease frequency in either clinical or observational studies.
9. **Analogy.** Structure activity relationships and information on the agent's structural analogs can provide insight into whether an association is causal. Similarly, information on mode of action for a chemical, as one of many structural analogs, can inform decisions regarding likely causality.

While these aspects provide a framework for assessing the evidence, they do not lend themselves to being considered in terms of simple formulas or fixed rules of evidence leading to conclusions about causality (Hill, 1965). For example, one cannot simply count the number of studies reporting statistically significant results or statistically nonsignificant results for health effects and reach credible conclusions about the relative weight of the evidence and the likelihood of causality. Rather, these important considerations are taken into account with the goal of producing an objective appraisal of the evidence, informed by peer and public comment and advice, which includes weighing alternative views on controversial issues. Additionally, it is important to note that the principles in Table 1-1 cannot be used as a strict checklist, but rather to determine the weight of the evidence for inferring causality. In particular, the absence of one or more of the principles does not automatically exclude a study from consideration (e.g., see discussion in U.S. Surgeon General's Report, 2004).

1.3.7. First Step—Determination of Causality

In the ISA, EPA assesses results of recent publications, in light of evidence available during the previous NAAQS review, to draw conclusions on the causal relationships between relevant pollutant exposures and health outcomes. This second external review draft ISA uses a five-level hierarchy that classifies the weight of evidence for causation, not just association.¹; that is, whether the weight of scientific evidence makes causation at least as likely as not, in the judgment of the reviewing group. In developing this hierarchy, EPA has drawn on the work of previous evaluations, most prominently the IOM's Improving the Presumptive Disability Decision-Making Process for Veterans (IOM, 2008), EPA's Guidelines for Carcinogen Risk Assessment (EPA, 1986a), and the U.S. Surgeon General's smoking reports (U.S. Surgeon General's Report, 2004). These efforts are presented in more detail in Annex A. In the draft ISA, EPA uses a series of five descriptors to characterize the weight of evidence on whether associations are in fact causal. This weight of evidence evaluation is based on various lines of evidence from epidemiological studies, animal studies, or other mechanistic, toxicological, or biological sources. These separate judgments are integrated into a qualitative statement about the overall weight of the evidence and causality. The five descriptors for causal determination are described in Table 1-2.

Table 1-2. Weight of evidence for causal determination.

Sufficient to infer a causal relationship	Evidence is sufficient to conclude that there is a causal relationship between relevant pollutant exposure and the outcome. That is, a positive association has been observed between the pollutant and the outcome in studies in which chance, bias, and confounding could be ruled out with reasonable confidence. For example, controlled human exposures, epidemiologic “natural experiments,” or observational studies supported by other lines of evidence. Generally, determination is based on multiple studies by multiple investigators.
Sufficient to infer a likely causal relationship (i.e., more likely than not).	Evidence is sufficient to conclude a likely causal association between relevant pollutant exposures and the outcome. That is, a positive association has been observed between the pollutant and the outcome in studies in which chance and bias can be ruled out with reasonable confidence but potential confounding issues remain. For example, a) observational studies show positive associations but copollutant exposures are difficult to address and/or other lines of evidence (controlled human exposure, animal, or mechanism of action information) are limited or inconsistent; or b) animal evidence from multiple studies, sex, or species is positive but limited or no human data are available. Generally, determination is based on multiple studies by multiple investigators.

¹ It should be noted that the CDC and IOM frameworks use a four-category hierarchy for the strength of the evidence. A five-level hierarchy is used here to be consistent with the EPA Guidelines for Carcinogen Risk Assessment and to provide a more nuanced set of categories.

Suggestive, but not sufficient to infer a causal relationship	Evidence is suggestive of an association between relevant pollutant exposures and the outcome, but is limited because chance, bias and confounding cannot be ruled out. For example, at least one high-quality study shows a positive association but the results of other studies are inconsistent.
Inadequate to infer the presence or absence of a causal relationship	The available studies are of insufficient quality, consistency or statistical power to permit a conclusion regarding the presence or absence of an association between relevant pollutant exposure and the outcome. For example, studies fail to control for confounding, have inadequate exposure assessment, or fail to address latency.
Suggestive of no causal relationship	Several adequate studies, covering the full range of levels of exposure that human beings are known to encounter and considering sensitive subpopulations, are mutually consistent in not showing a positive association between exposure and the outcome at any level of exposures. In addition, the possibility of a very small elevation in risk at the levels of exposure studied can never be excluded.

1.3.8. Second Step—Evaluation of Population Response

Beyond judgments regarding causality are questions relevant to characterizing exposure and risk to populations (i.e., at what levels do health effects occur). Such questions include:

- Under what exposure conditions (dose or exposure, duration and pattern) are effects seen?
- What is the shape of the concentration-response or dose-response relationship?
- What population groups appear to be affected or more susceptible to effects?

On the population level, causal and likely causal claims typically characterize how risk—the probability of health effects—changes in response to exposure. Initially, the response is evaluated within the range of observation. Approaches to analysis of the range of observation of epidemiological and human clinical studies are determined by the type of study and methods of exposure/dose and response measurement. Extensive human data for concentration-response analyses exists for all criteria pollutants, unlike most other environmental pollutants. Animal data also can inform concentration-response, particularly relative to dosimetry, mechanisms of action, and characteristics of sensitive subpopulations.

An important consideration in characterizing the public health impacts associated with exposure to a pollutant is whether the concentration-response relationship is linear across the full concentration range encountered, or if nonlinear relationships exist along any part of this range. Of particular interest is the shape of the concentration-response curve at and below the level of the current standards. The complex molecular and cellular events that underlie cancer and noncancer toxicity are likely to be both linear and nonlinear, and vary depending on dose.

1 Additionally, many chemicals and agents may act by perturbing naturally occurring background
2 processes that lead to disease. At the human population level, however, various sources of
3 variability and uncertainty tend to smooth and “linearize” the concentration-response function
4 (such as the low data density in the lower concentration range, possible influence of
5 measurement error, and individual differences in susceptibility to air pollution health effects).
6 These attributes of population dose-response may explain why the available human data at
7 ambient concentrations for some environmental pollutants (e.g., ozone, lead [Pb], PM,
8 secondhand tobacco smoke, radiation) do not exhibit evident thresholds for cancer or noncancer
9 health effects, even though likely mechanisms of action include nonlinear processes for some
10 key events. These attributes of human population dose-response relationships have been
11 extensively discussed in the broader epidemiological literature (e.g., Rothman and Greenland,
12 1998).

1.4. Conclusions

13 This second external review draft ISA strives to present a concise review, synthesis, and
14 evaluation of the most policy-relevant science, and communicates critical science judgments
15 relevant to the NAAQS review. It reviews the most policy relevant evidence from
16 epidemiological, human clinical, and animal toxicological studies, including mechanistic
17 evidence from basic biological science. Annexes to the ISA provide additional details of the
18 literature published since the last review. A framework for making critical judgments concerning
19 causality is presented in this chapter. It relies on a widely accepted set of principles and
20 standardized language to express evaluation of the evidence. This approach can bring rigor and
21 clarity to the current and future assessments. This ISA should assist EPA and others, now and in
22 the future, to represent accurately what is presently known—and what remains unknown—
23 concerning the effects of sulfur oxides on human health.

Chapter 2. Source to Tissue Dose

This chapter contains basic information about concepts and findings in atmospheric sciences, human exposure assessment, and human dosimetry. It is meant to serve as a prologue for the detailed discussions of health effects data in Chapters 3 and 4. Section 2.1 gives an overview of the sources of SO₂. Atmospheric chemistry processes involved in the oxidation of SO₂ and those involved in the production of SO₂ from reduced sulfur gases in the atmosphere are discussed in Section 2.2. A description of SO₂ measurement methods and related issues are presented in Section 2.3. Data for ambient SO₂ concentrations are characterized in Section 2.4. Policy relevant background concentrations of SO₂, i.e., those concentrations defined to result from uncontrollable emissions, are also presented in Section 2.4. Factors related to personal exposure to SO₂ are discussed in Section 2.5. Finally, Section 2.6 covers the dosimetry of SO₂ in the respiratory tract. This organization generally follows that given in the National Research Council (NRC) paradigm for integrating air pollutant research (NRC, 1998).

2.1. Sources of Sulfur Oxides

Industrial emissions of SO₂ in the United States are mainly due to combustion of fossil fuels by electrical utilities (~66 %) and industry (~29%); transportation-related sources contribute minimally (~5%) (2002 statistics) (EPA, 2006d). Thus, most SO₂ emissions originate from point sources. Annex B has a detailed breakdown of emissions by source category. Almost all of the sulfur in fuel is released as volatile components (SO₂ or SO₃) during combustion. Hence, based on sulfur content in fuel stocks, sulfur emissions can be calculated to a higher degree of accuracy than other pollutants such as nitrogen oxides or primary PM. However, these estimates given above are national averages and may not accurately reflect the contribution of specific local sources for determining individual exposure to SO₂ at a particular location and time. For example, shipping and in-port activities may be a significant source of SO₂ in some coastal cities (Wang et al., 2007).¹

¹ Ships and commercial boats contribute approximately 25% of the SO₂ emissions in the South Coast Air Basin, and 50% of statewide SO₂ emissions (Dabdub and Vutukuru, 2008). Because of the importance of SO₂ emissions, the ports of Long Beach and Los Angeles are part of a Sulfur Emissions Control Area in which sulfur contents of fuels are not to exceed 1.5%. Modeling studies by Vutukuru (2008) also indicate that ships contribute just over 1 part per billion (ppb) SO₂ (for a 24-h avg) to Long Beach, and a few tenths of a ppb to locations further inland.

1 The largest natural sources of SO₂ are volcanoes and wildfires. Although SO₂ constitutes a
2 relatively minor fraction (0.005% by volume) of total volcanic emissions (Holland, 1978),
3 concentrations in volcanic plumes can be in the range of several to tens of ppm. Volcanic sources
4 of SO₂ in the U.S. are limited to the Pacific Northwest, Alaska, and Hawaii. Emissions of SO₂
5 from burning vegetation are generally in the range of 1 to 2% of the biomass burned (Levine and
6 Pinto, 1998). Sulfur is a component of amino acids in vegetation and is released during
7 combustion. Gaseous sulfur emissions from this source are mainly in the form of SO₂.

8 In addition to its role as an emitted primary pollutant, SO₂ is also produced by the
9 photochemical oxidation of reduced sulfur compounds such as dimethyl sulfide (CH₃-S-CH₃, or
10 DMS), hydrogen sulfide (H₂S), carbon disulfide (CS₂), carbonyl sulfide (OCS), methyl
11 mercaptan (CH₃-S-H), and dimethyl disulfide (CH₃-S-S-CH₃). The sources for these compounds
12 are mainly biogenic (see Annex Table B-6). Emissions of reduced sulfur species are associated
13 typically with marine organisms living either in pelagic or coastal zones, and with anaerobic
14 bacteria in marshes and estuaries. Emissions of DMS from marine plankton represent the largest
15 single atmospheric source of reduced sulfur species (Berresheim et al., 1995). Other than OCS,
16 which is lost mainly by photolysis (e-folding lifetime, [τ] ~6 months), species are lost mainly by
17 reaction with hydroxyl radical (OH) and NO₃ radicals, and are relatively short-lived; lifetimes
18 range from a few hours to a few days (see Annex Table B-2). Reaction with NO₃ radicals at night
19 most likely represents the major loss process for DMS and methyl mercaptan. Although the
20 mechanisms for the oxidation of DMS are not completely understood, excess sulfate in marine
21 aerosol appears related mainly to the production of SO₂ from the oxidation of DMS. Emissions of
22 sulfur from natural sources are small compared to industrial emissions within the U.S. However,
23 important exceptions occur locally as the result of volcanic activity, wildfires and in certain
24 coastal zones as described above.

25 Because OCS is relatively long-lived, it can survive oxidation in the troposphere and be
26 transported upward into the stratosphere. Crutzen (1976) proposed that its oxidation to sulfate in
27 the stratosphere serves as the major source of the stratospheric aerosol layer. However, Myhre
28 et al. (2004) proposed that SO₂ transported upward from the troposphere by deep convection is
29 the most likely source, since the flux of OCS is too small. In addition, in situ measurements of
30 the isotopic composition of sulfur in stratospheric sulfate do not match those of OCS (Leung,

2002). Thus, anthropogenic SO₂ emissions could be important precursors to the formation of the stratospheric aerosol layer.

2.2. Atmospheric Chemistry

The only forms of monomeric sulfur oxides of interest in tropospheric chemistry are SO₂ and SO₃. SO₃ can be emitted from the stacks of power plants and factories; however, it reacts extremely rapidly with H₂O in the stacks or immediately after release into the atmosphere to form H₂SO₄, which mainly condenses onto existing particles when particle loadings are high; it can nucleate to form new particles under lower concentration conditions. Thus, only SO₂ is present in the tropospheric boundary layer at concentrations of concern for human exposures. The gas phase oxidation of SO₂ is initiated by the reaction



where M is an atmospheric constituent such as N₂ and O₂ that helps stabilize the reaction product. Reaction 2-1 is followed by



Because the saturation vapor pressure of H₂SO₄ is extremely low, it will be removed rapidly by transfer to the aqueous phase of aerosol particles and cloud drops. Depending on atmospheric conditions and concentrations of ambient particles and gaseous species that can participate in new particle formation, it can also nucleate to form new particles. Rate coefficients for the reactions of SO₂ with either the hydroperoxyl radical (HO₂) or NO₃ are too low to be significant (Jet Propulsion Laboratory, 2003).

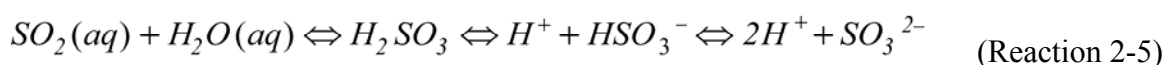
The major sulfur species in clouds are hydrogen sulfite (HSO₃⁻) and the sulfite ion (SO₃²⁻). Both are derived from the dissolution of SO₂ in water, and are referred to as S(IV); bisulfate ion (HSO₄⁻) and sulfate (sulfate) are referred to as S(VI). The chief species capable of oxidizing S(IV) to S(VI) in cloud water are O₃, peroxides (either hydrogen peroxide [H₂O₂] or organic peroxides), hydroxyl (OH) radicals, and ions of transition metals such as iron (Fe), manganese (Mn) and copper (Cu) that can catalyze the oxidation of S(IV) to S(VI) by O₂. The basic

mechanism of the aqueous phase oxidation of SO₂ has long been studied and can be found in numerous texts on atmospheric chemistry, e.g., Seinfeld and Pandis (1998), Finlayson-Pitts and Pitts (1999), Jacob (1999), and Jacobson (2002). Following Jacobson (2002), the steps involved in the aqueous phase oxidation of SO₂ can be summarized as

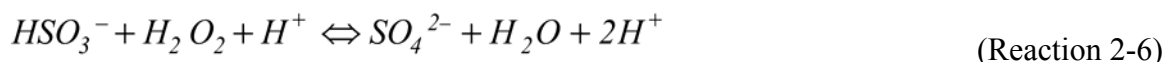
Dissolution of SO₂



The formation and dissociation of H₂SO₃



In the pH range commonly found in rainwater (pH 2 to 6), the most important reaction converting S(IV) to S(VI) is



as SO₃²⁻ is much less abundant than HSO₃⁻.

For pH up to about 5.3, H₂O₂ is the dominant oxidant, while at pH > 5.3, O₃ becomes dominant, followed by Fe(III), using characteristic values found in Seinfeld and Pandis (1998). However, differences in concentrations of oxidants result in differences in the pH at which this transition occurs. It should also be noted that the oxidation of SO₂ by O₃ and O₂ tends to be self-limiting: as sulfate is formed, the pH decreases and the rates of these reactions decrease. Higher pH levels are expected to be found mainly in marine aerosols. However, in marine aerosols, the chloride-catalyzed oxidation of S(IV) may be more important (Hoppel and Caffrey, 2005; Zhang and Millero, 1991). Because the ammonium ion (NH₄⁺) is so effective in neutralizing acidity, it affects the rate of oxidation of S(IV) to S(VI) and the rate of dissolution of SO₂ in particles and cloud drops.

A comparison of the relative rates of oxidation by gas and aqueous phase reactions by Warneck (1999) indicates that on average only about 20% of SO₂ is oxidized by gas phase reactions; the remainder is oxidized by aqueous phase reactions. In areas away from strong pollution sources, the SO₂ τ is ~7 days, based on measurements of the rate constant for Reaction 2-1 (Jet Propulsion Laboratory, 2003) and a nominal concentration for the OH radical of 10⁶/cm³.

1 However, the mechanism of SO₂ oxidation at a particular location depends on local
2 environmental conditions. For example, near stacks, oxidants such as OH radicals are depleted
3 and almost no SO₂ is oxidized in the gas phase. Further downwind, as the plume is diluted with
4 background air, the gas phase oxidation of SO₂ increases in importance. Finally, even further
5 downwind when conditions in the plume can become more oxidizing than in background air, the
6 SO₂ oxidation rate could exceed that in background air. SO₂ in the planetary boundary layer is
7 also removed from the atmosphere by dry deposition to moist surfaces, resulting in an
8 atmospheric τ with respect to dry deposition of approximately 1 day to 1 week. Wet deposition of
9 sulfur naturally depends on the variable nature of rainfall, but in general results in a τ of SO₂ ~7
10 days, too. These two processes, oxidation and deposition, lead to an overall lifetime of SO₂ in the
11 atmosphere of 3 to 4 days.

2.3. Measurement Methods and Associated Issues

12 Currently, ambient SO₂ is measured using instruments based on pulsed ultraviolet (UV)
13 fluorescence. The UV fluorescence monitoring method for atmospheric SO₂ was developed to
14 improve on the flame photometric detection (FPD) method, which in turn had replaced the
15 pararosaniline wet chemical method. This latter method is still the EPA's Federal Reference
16 Method (FRM) for atmospheric SO₂, but is rarely used due to its complexity and slow response,
17 even in its automated forms. Both the UV fluorescence and FPD methods are designated as
18 Federal Equivalent Methods (FEMs) by EPA, but UV fluorescence has largely supplanted the
19 FPD approach because of the UV method's inherent linearity and because the FPD method needs
20 consumable hydrogen gas.

21 In the UV fluorescence method, SO₂ molecules absorb UV light at one wavelength and
22 emit UV light at longer wavelengths in the process known as fluorescence, through excitation of
23 the SO₂ molecule to a higher energy (singlet) electronic state. Once excited, the molecule decays
24 nonradiatively to a lower-energy electronic state from which it then decays to the original or
25 electronic state by emitting a photon of light at a longer wavelength (i.e., a lower-energy photon)
26 than the original, incident photon. The intensity of the emitted light is thus proportional to the
27 number of SO₂ molecules in the sample gas.

28 In commercial analyzers, light from a high-intensity UV lamp passes through a bandwidth
29 filter, allowing only photons with wavelengths around the SO₂ absorption peak (near 214

nanometers [nm]) to enter the optical chamber. The light passing through the source bandwidth filter is collimated using a UV lens and passes through the optical chamber, where it is detected on the opposite side of the chamber by the reference detector. A photomultiplier tube (PMT) is offset from and placed perpendicular to the light path to detect the SO₂ fluorescence. Since the SO₂ fluorescence at 330 nm is different from its excitation wavelength, an optical bandwidth filter is placed in front of the PMT to filter out any stray light from the UV lamp. A lens is located between the filter and the PMT to focus the fluorescence onto the active area of the detector and optimize the fluorescence signal. The limit of detection (LOD) for a non-trace level SO₂ analyzer is required to be 10 ppb (FR, 2006). However, most commercial analyzers have detection limits of about 3 ppb; many monitors might have lower effective detection limits. The EPA, through its National Core (NCore) initiative (EPA, 2005) is in the process of supporting state, local, tribal, and federal networks in the implementation of newer trace-level SO₂ instrumentation. These new trace-level instruments have detection limits of 0.1 ppb or lower. More information related to SO₂ sampling and measurement is in Annex B.5.

2.3.1. Sources of Positive Interference

The most common source of interference to the UV fluorescence method for SO₂ is from other gases that fluoresce in a similar fashion when exposed to UV radiation. The most significant of these are polycyclic aromatic hydrocarbons (PAHs), of which naphthalene is a prominent example. Xylene is another common hydrocarbon that can cause fluorescent interference. Consequently, any such aromatic hydrocarbons in the optical chamber can act as positive interference. To remove this source of interference, high-sensitivity SO₂ analyzers, such as those to be used in the NCore network (EPA, 2005), have hydrocarbon scrubbers to remove these compounds from the sample stream before the sample air enters the optical chamber.

Luke (1997) reported positive artifacts of a modified pulsed fluorescence detector generated by the coexistence of nitric oxide (NO), carbon disulfide (CS₂), and a number of highly fluorescent aromatic hydrocarbons such as benzene, toluene, *o*-xylene, *m*-xylene, *p*-xylene, *m*-ethyltoluene, ethylbenzene, and 1,2,4-trimethylbenzene. The positive artifacts could be reduced by using a hydrocarbon “kicker” membrane. At a flow rate of 300 standard cc min⁻¹ and a pressure drop of 645 torr across the membrane, the interference from ppm levels of many aromatic hydrocarbons was eliminated. NO fluoresces in a spectral region close to that of SO₂.

1 However, in high-sensitivity SO₂ analyzers, the bandpass filter in front of the PMT is designed to
2 prevent NO fluorescence from being detected at the PMT. Care must be exercised when using
3 multicomponent calibration gases containing both NO and SO₂, so that the NO rejection ratio of
4 the SO₂ analyzer is sufficient to prevent NO interference.

5 The most common source of positive bias (as contrasted with positive spectral
6 interference) in high-sensitivity SO₂ monitoring is stray light in the optical chamber. Since SO₂
7 can be electronically excited by a broad range of UV wavelengths, any stray light with an
8 appropriate wavelength that enters the optical chamber can excite SO₂ in the sample and increase
9 the fluorescence signal. Furthermore, stray light at the wavelength of the SO₂ fluorescence that
10 enters the optical chamber may impinge on the PMT and increase the fluorescence signal.
11 Several design features minimize stray light, including the use of light filters, dark surfaces, and
12 opaque tubing.

13 Nicks and Benner (2001) reported a sensitive SO₂ chemiluminescence detector based on a
14 differential measurement: response from ambient SO₂ is determined by the difference between
15 air containing SO₂ and air scrubbed of SO₂ when both air samples contain other detectable sulfur
16 species. Assuming monotonic efficiency of the sulfur scrubber, all positive artifacts should also
17 be reduced with this technique.

2.3.2. Sources of Negative Interference

18 Nonradiative deactivation (quenching) of excited SO₂ molecules can occur from collisions
19 with common molecules in air, including nitrogen, oxygen, and water. During collisional
20 quenching, the excited SO₂ molecule transfers energy, kinetically allowing the SO₂ molecule to
21 return to the original lower energy state without emitting a photon. Collisional quenching results
22 in a decrease in the SO₂ fluorescence and, hence, an underestimation of SO₂ concentration in the
23 air sample. Of particular concern is the variable water vapor content of air. Luke (1997) reported
24 that the response of the detector could be reduced by an amount of ~7 to 15% at water vapor
25 mixing ratios of 1 to 1.5 mole percent (relative humidity [RH] = 35 to 50% at 20 to 25°C and 1
26 atmosphere [atm] for a modified pulsed fluorescence detector [Thermo Environmental
27 Instruments, Model 43s]). Condensation of water vapor in sampling lines must be avoided, as
28 water on the inlet surfaces can absorb SO₂ from the sample air. The simplest approach to avoid
29 condensation is to heat sampling lines to a temperature above the expected dew point and to

1 within a few degrees of the controlled optical bench temperature. At very high SO₂
2 concentrations, reactions between electronically excited SO₂ and ground state SO₂ might occur,
3 forming SO₃ and SO (Calvert et al., 1978). However, the possibility that this artifact might be
4 affecting measurements at very high SO₂ levels has not been examined.

2.3.3. Other Techniques for Measuring so₂

5 More sensitive techniques for measuring SO₂ are available, but most of these systems are
6 too complex and expensive for routine monitoring applications. However, techniques such as
7 those described by Luke (1997) can be used to improve the sensitivity of ambient SO₂ monitors
8 by eliminating sources of common interference. See descriptions in Annex section B.5.

2.4. Environmental Concentrations of SO_x

2.4.1. Design Criteria for the NAAQS so₂ Monitoring Networks¹

9 Trace level SO₂ monitoring is currently required at the approximately 75 proposed NCore
10 sites, as noted in CFR 40 Part 58 Appendices C and D. Continued operation of existing State and
11 Local Air Monitoring Sites (SLAMS) for SO₂ using Federal Reference Methods (FRM) or
12 Federal Equivalent Methods (FEM) is required until discontinuation is approved by the EPA
13 Regional Administrator. Where SLAMS SO₂ monitoring is required, at least one of the sites must
14 be a maximum concentration site for that specific area. In 2007, there were ~500 SO₂ monitors
15 reporting values to the EPA Air Quality System database (AQS).

16 The appropriate spatial scales for SO₂ SLAMS monitoring are the microscale, middle, and
17 possibly neighborhood scales.

18 ■ **Micro and middle scale**—Some data uses associated with microscale and middle scale
19 measurements for SO₂ include assessing the effects of control strategies to reduce
20 concentrations (especially for the 3-hour and 24-hour averaging times), and monitoring
21 air pollution episodes.

22 ■ **Neighborhood scale**—This scale applies where there is a need to collect air quality data
23 as part of an ongoing SO₂ stationary source impact investigation. Typical locations
24 might include suburban areas adjacent to SO₂ stationary sources, for example, or for

¹ This section is adapted from Code of the Federal Register 40 CFR Parts 53 and 58 and Appendix E to Part 58, as revised: Vol. 71, No. 200 / 17
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1 determining background concentrations as part of studies of population responses to
2 SO₂ exposure.

Horizontal and Vertical Placement

3 The probe, or at least 80 percent of the monitoring path, must be located between 2 and 15
4 meters above ground level for all SO₂ monitoring sites. The probe, or at least 90 percent of the
5 monitoring path, must be positioned at least 1 meter vertically or horizontally from any
6 supporting structure, walls, parapets, penthouses, etc., and away from dusty or dirty areas. If the
7 probe, or a significant portion of the monitoring path, is located near the side of a building, it
8 should be located on the windward side relative to the prevailing wind direction during the
9 season of highest concentration potential for the pollutant being measured.

Spacing from Minor Sources

10 Local minor sources of a primary pollutant such as SO₂ can affect concentrations of that
11 particular pollutant at a monitoring site. If the objective for that site is to investigate these local
12 primary pollutant emissions, then the site should be located where the spatial and temporal
13 variability in these emissions can be captured. This type of monitoring site would likely be the
14 microscale type. If a monitoring site is to be used to determine air quality over a much larger
15 area, such as a neighborhood or city, a monitoring agency should avoid placing a monitor probe,
16 path, or inlet near local, minor sources. The plume from the local minor sources should not be
17 allowed to inappropriately influence the air quality data collected.

18 To minimize these potential interferences, the probe, or at least 90 percent of the
19 monitoring path, must be placed away from furnace or incineration flues, or other minor sources
20 of SO₂. The separation distance should take into account the heights of the flues, type of waste or
21 fuel burned, and the sulfur content of the fuel.

Spacing from Obstructions

22 Buildings and other obstacles may possibly scavenge SO₂, and can act to restrict airflow
23 for any pollutant. To avoid this interference, the probe, inlet, or at least 90 percent of the
24 monitoring path must have unrestricted airflow and be located away from obstacles. The distance
25 from the obstacle to the probe, inlet, or monitoring path must be at least twice the height of the
26 obstruction's protrusion. An exception can be made for measurements taken in street canyons or

1 at source-oriented sites where buildings and other structures are unavoidable. Generally, a probe
2 or monitoring path located near or along a vertical wall is undesirable, because air moving along
3 the wall may be subject to possible removal mechanisms. A probe, inlet, or monitoring path must
4 have unrestricted airflow in an arc of at least 180 degrees. This arc must include the predominant
5 wind direction for the season of greatest pollutant concentration potential.

6 Special consideration must be devoted to the use of open path analyzers, due to their
7 inherent potential sensitivity to certain types of interferences, or optical obstructions. A
8 monitoring path must be clear of all trees, brush, buildings, plumes, dust, or other optical
9 obstructions, including potential obstructions that may move due to wind, human activity, growth
10 of vegetation, etc. Temporary optical obstructions, such as rain, particles, fog, or snow, should be
11 considered when locating an open path analyzer. Any temporary obstructions that are of
12 sufficient density to obscure the light beam will affect the ability of the open path analyzer to
13 measure pollutant concentrations continuously. Transient, but significant obscuration of
14 especially longer measurement paths could occur because certain meteorological conditions
15 (e.g., heavy fog, rain, snow) and/or aerosol levels are of sufficient density to prevent the
16 analyzer's light transmission. If certain compensating measures are not otherwise implemented at
17 the onset of monitoring (e.g., shorter path lengths, higher light source intensity), data recovery
18 during periods of greatest primary pollutant potential could be compromised. For instance, if
19 heavy fog or high particulate levels are coincident with periods of projected NAAQS-threatening
20 pollutant potential, the resulting data may not be representative for reflecting maximum pollutant
21 concentrations, despite the fact that the site may otherwise exhibit an acceptable, even
22 exceedingly high overall valid data capture rate.

Spacing from Trees

23 Trees can provide surfaces for SO₂ adsorption or reactions, and surfaces for particle
24 deposition. Trees can also act as obstructions in cases where they are located between the air
25 pollutant sources or source areas and the monitoring site, and where the trees are of sufficient
26 height and leaf canopy density to interfere with normal airflow around the probe, inlet, or
27 monitoring path. To reduce possible interference, the probe, inlet, or at least 90 percent of the
28 monitoring path must be at least 10 meters or further from the drip line of trees.

For microscale sites, no trees or shrubs should be located between the probe and the source under investigation, such as a roadway or a stationary source.

2.4.2. Monitor Locations in Selected Areas of the U.S.

Figures 2-1 through 2-6 illustrate the 2005 geospatial locations of monitors for SO₂, NO₂, CO, particulate matter ≤ 10 μm (PM₁₀), particulate matter ≤ 2.5 μm (PM_{2.5}), and O₃. These locations, sited in several cities in six states, were selected as relevant for SO₂ health effects studies; see the summaries and assessments of health effects in Chapter 4, and the discussion of intracity SO₂ correlations that follows. For each state, Figure A shows locations of each monitor for all six pollutants; Figure B shows only the SO₂ monitor locations. Totals for each monitor type are included. These figures demonstrate the important point that not all SO₂ monitors in any Consolidated Metropolitan Statistical Area (CMSA) are co-located with monitors for other pollutants. Two examples are given below.

Table 2-1. Monitor counts for California and San Diego County, 2005.

	SO ₂	NO ₂	O ₃	CO	PM ₁₀	PM _{2.5}
California (all)	35	105	176	86	177	97
San Diego County	4	9	10	6	7	7

Table 2-2. Monitor counts for Ohio and Cuyahoga County, 2005.

	SO ₂	NO ₂	O ₃	CO	PM ₁₀	PM _{2.5}
Ohio (all)	31	4	49	15	49	49
Cuyahoga County	4	2	3	4	6	7

Table 2-1 lists the totals for all criteria air pollutant monitors (except Pb) in California, as well as the subset of these monitors in San Diego County. At each of the four sites where SO₂ was measured, NO₂, CO, PM₁₀, PM_{2.5}, and O₃ were also measured, with the exception of PM_{2.5} at one site (AQS ID 060732007) in Otay Mesa, CA. Table 2-2 lists the totals for all criteria air pollutant monitors (except Pb) in Ohio, as well as the subset of in Cuyahoga County.

1 In Cuyahoga County, PM₁₀ and PM_{2.5} were measured at all four sites where SO₂ was also
2 measured in 2005, but O₃ and CO were not measured at any of those four sites; NO₂ was only
3 measured at one site (AQS ID 39050060) near Cleveland's city center and ~0.5 km from the
4 intersection of Interstate Highways 77 and 90.

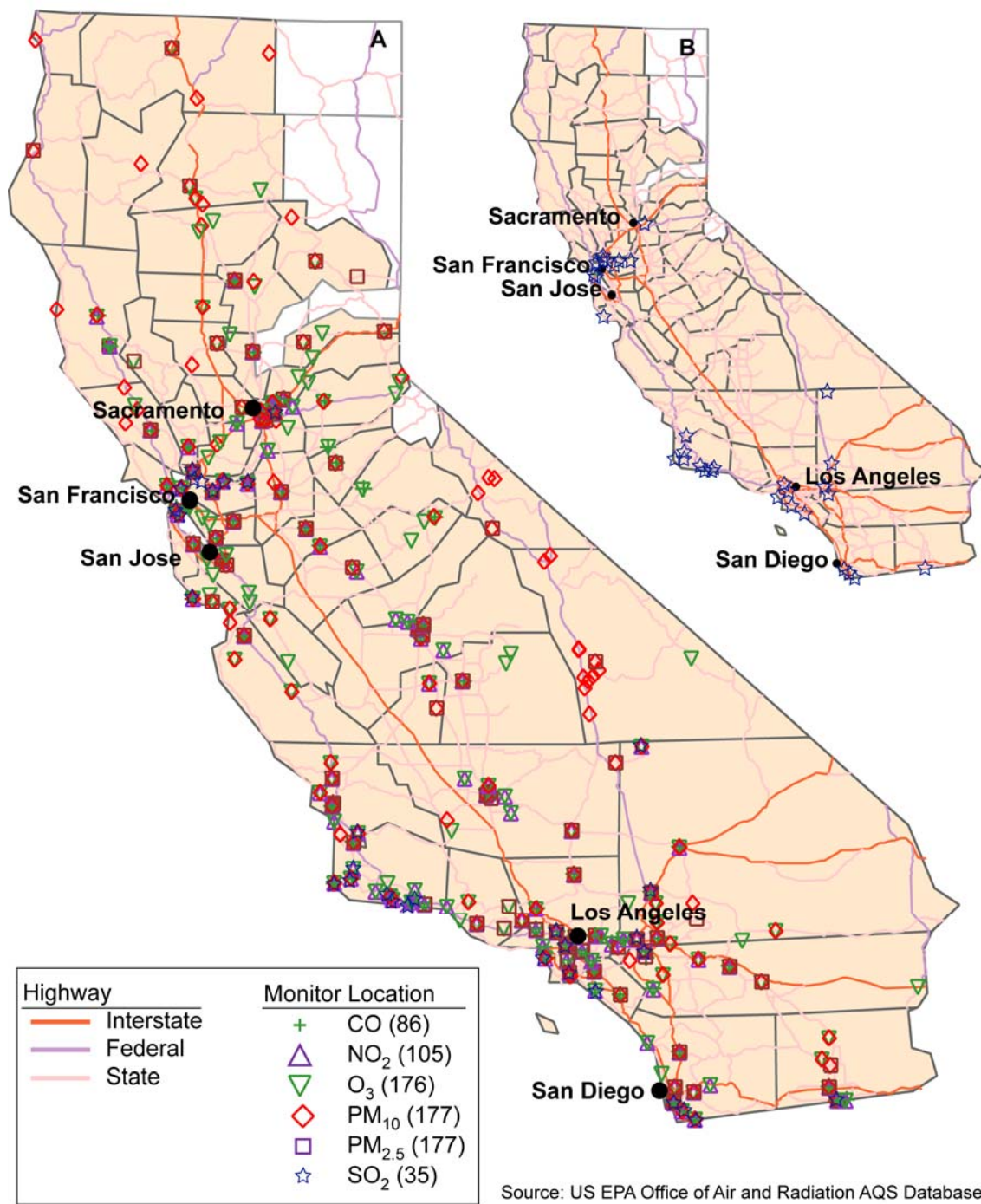


Figure 2-1. Criteria pollutant monitor locations (A) and SO₂ monitor locations (B), California, 2005. Shaded counties have at least one monitor.

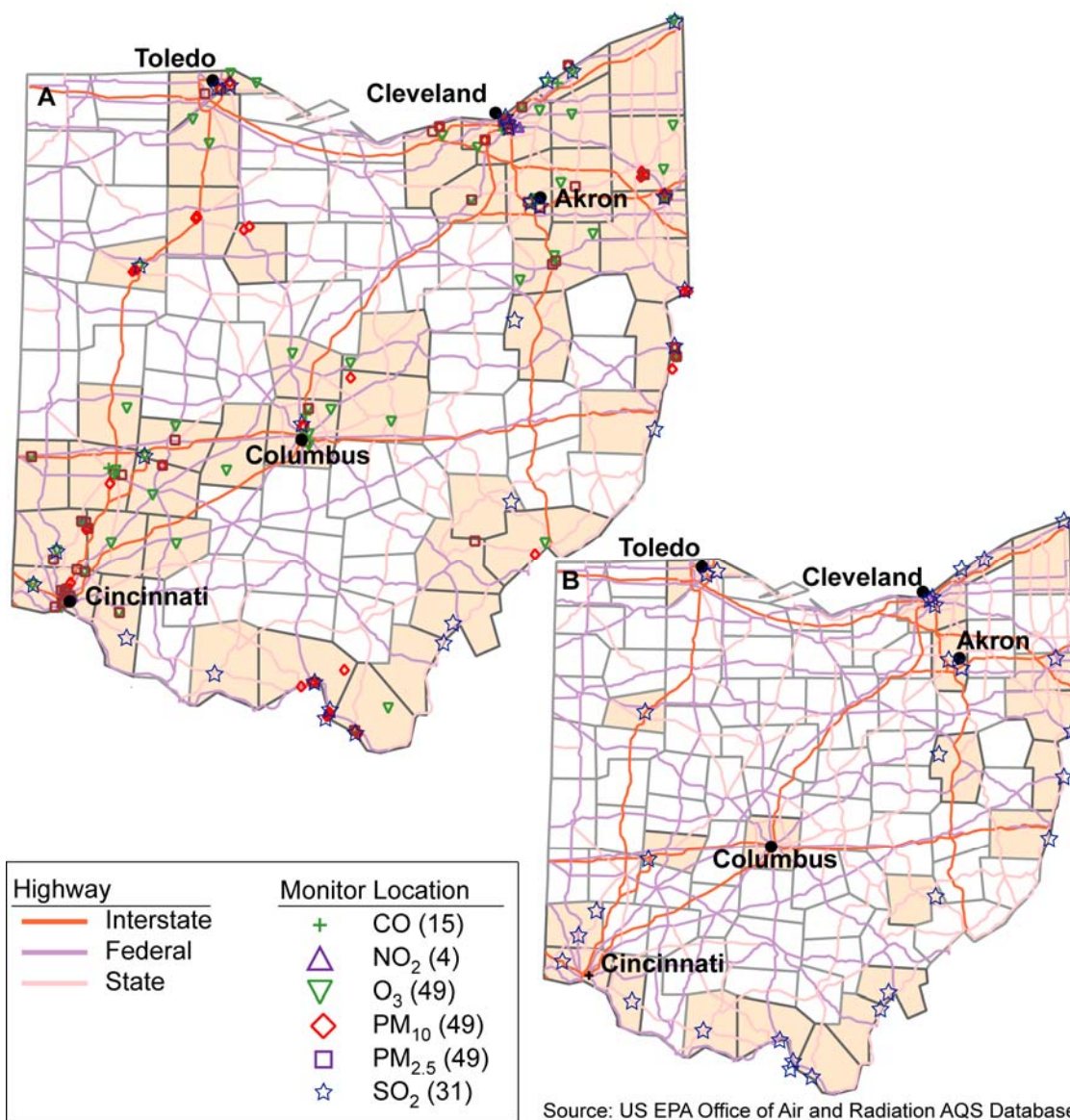


Figure 2-2. Criteria pollutant monitor locations (A) and SO₂ monitor locations (B), Ohio, 2005. Shaded counties have at least one monitor.

2.4.3. Ambient so₂ Concentrations in Relation to so₂ Sources

- 1 SO₂ data collected from the SLAMS and NAMS networks, like those illustrated in
- 2 Figures 2-1 through 2-6, show that the decline in SO₂ emissions from electric generating utilities
- 3 has substantially improved air quality. Not one monitored exceedance of the SO₂ annual ambient

1 air quality standard in the lower 48 States of the United States has been recorded since 2000,
 2 according to the EPA Acid Rain Program (ARP) 2005 Progress Report (EPA, 2006a). EPA's
 3 trends data (www.epa.gov/airtrends) reveal that the national composite average SO₂ annual mean
 4 ambient concentration decreased by 48% from 1990 to 2005; the largest single-year reduction
 5 was 1994-95, the ARP's first operating year (EPA, 2006a). Figure 2-7 depicts data for SO₂
 6 emissions in the contiguous United States (CONUS) during those years, with state-level totals.

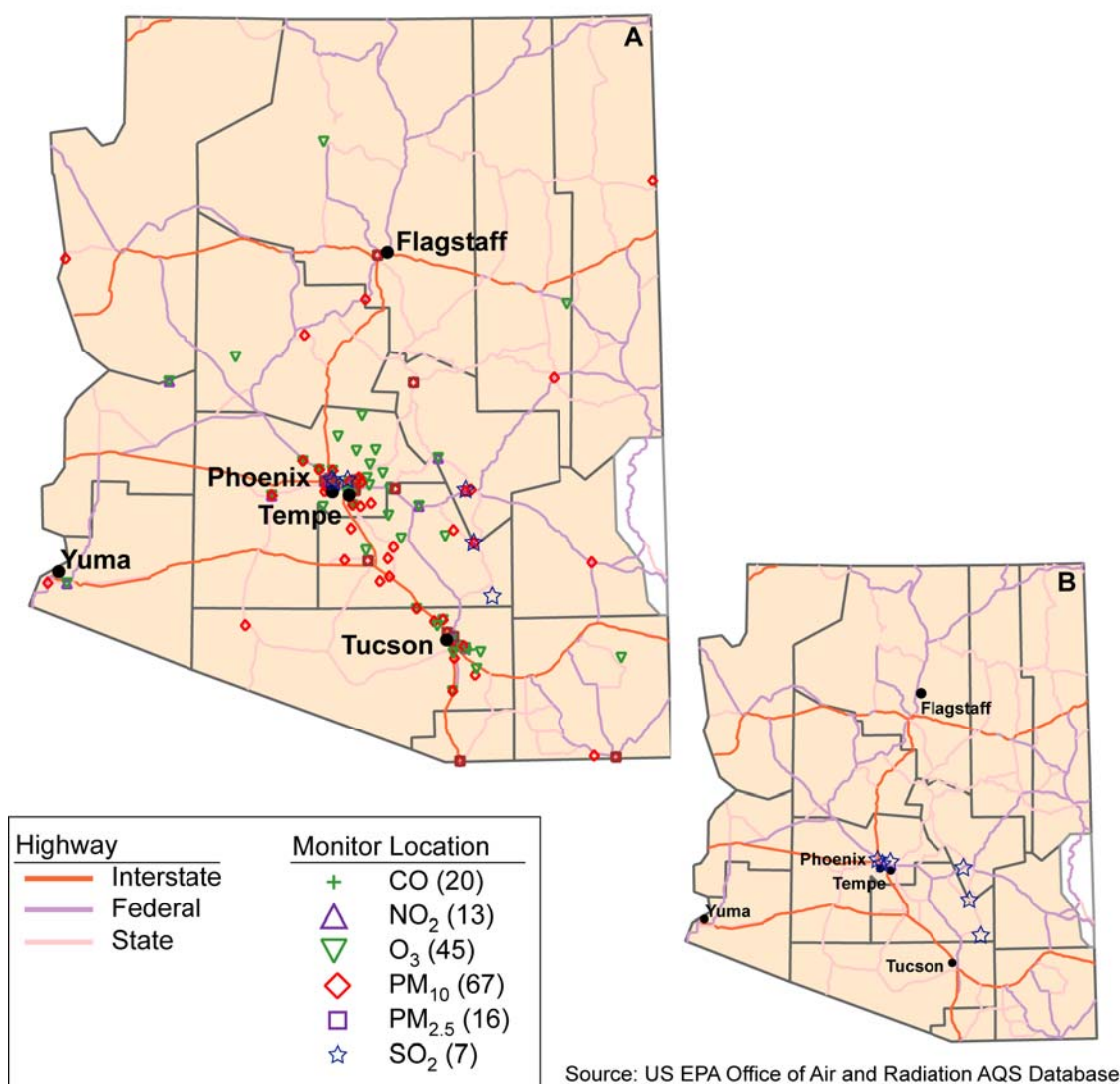


Figure 2-3. Criteria pollutant monitor locations (A) and SO₂ monitor locations (B), Arizona, 2005. Shaded counties have at least one monitor.

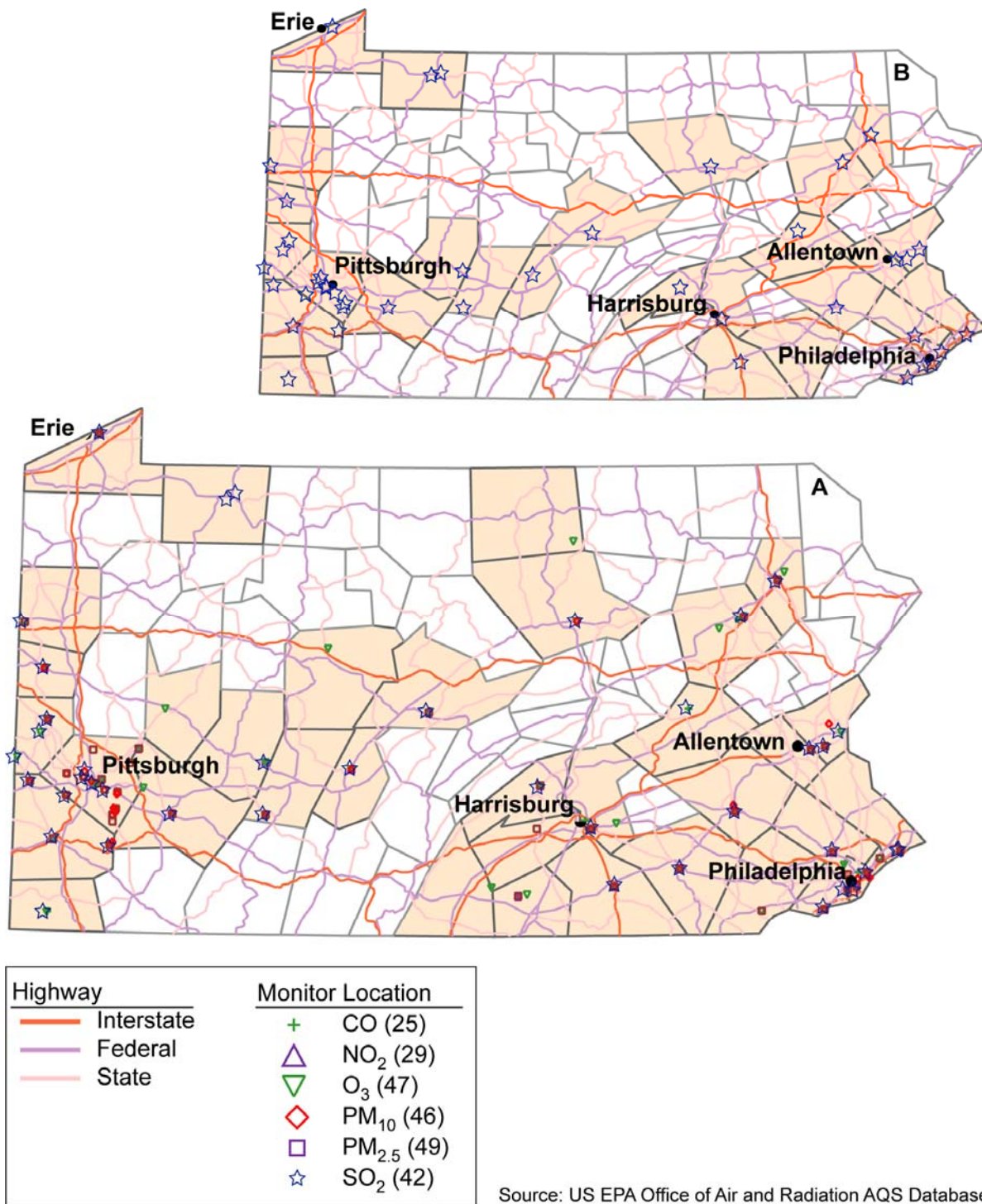


Figure 2-4. Criteria pollutant monitor locations (A) and SO₂ monitor locations (B), Pennsylvania, 2005. Shaded counties have at least one monitor.

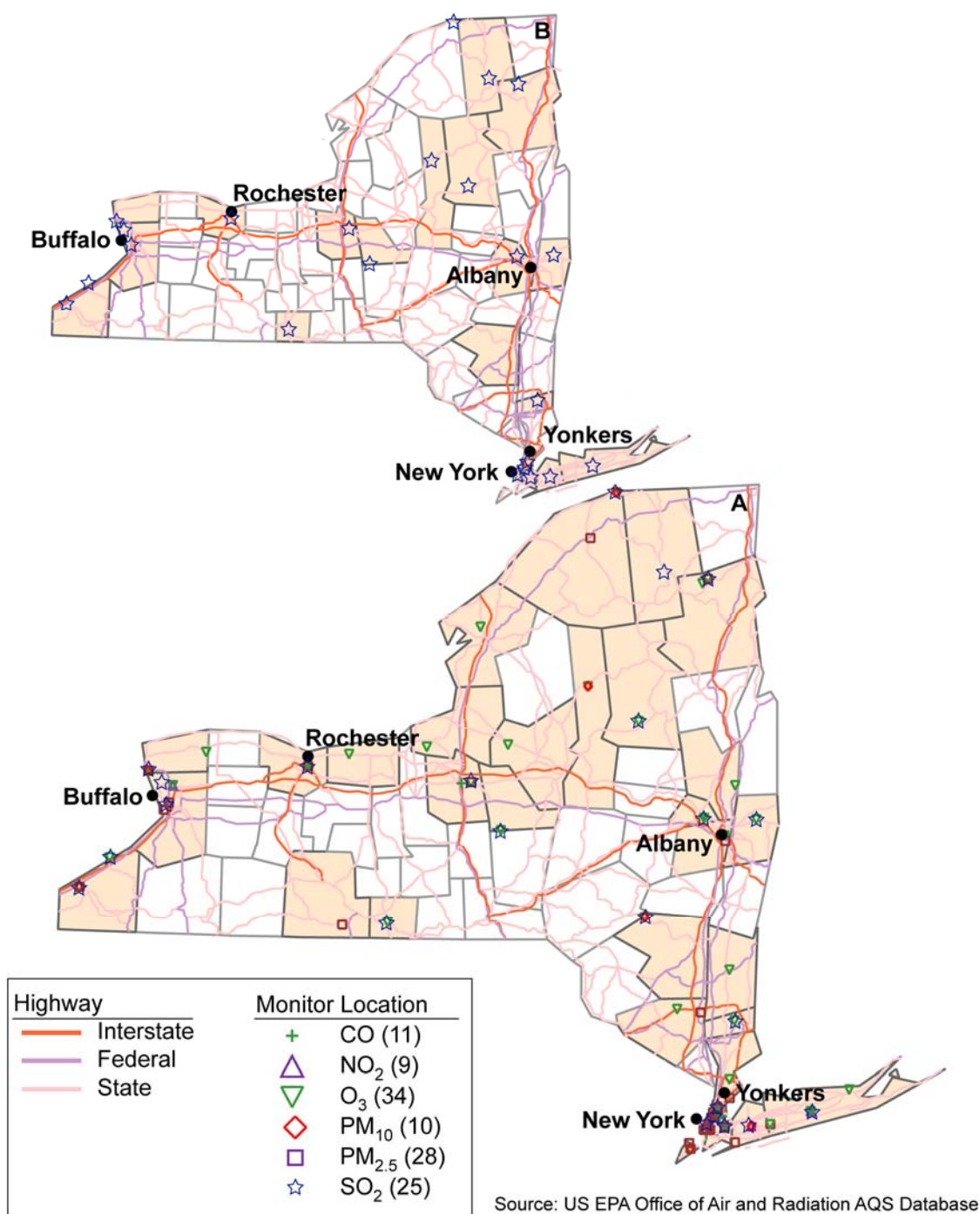


Figure 2-5. Criteria pollutant monitor locations (A) and SO₂ monitor locations (B), New York, 2005. Shaded counties have at least one monitor.

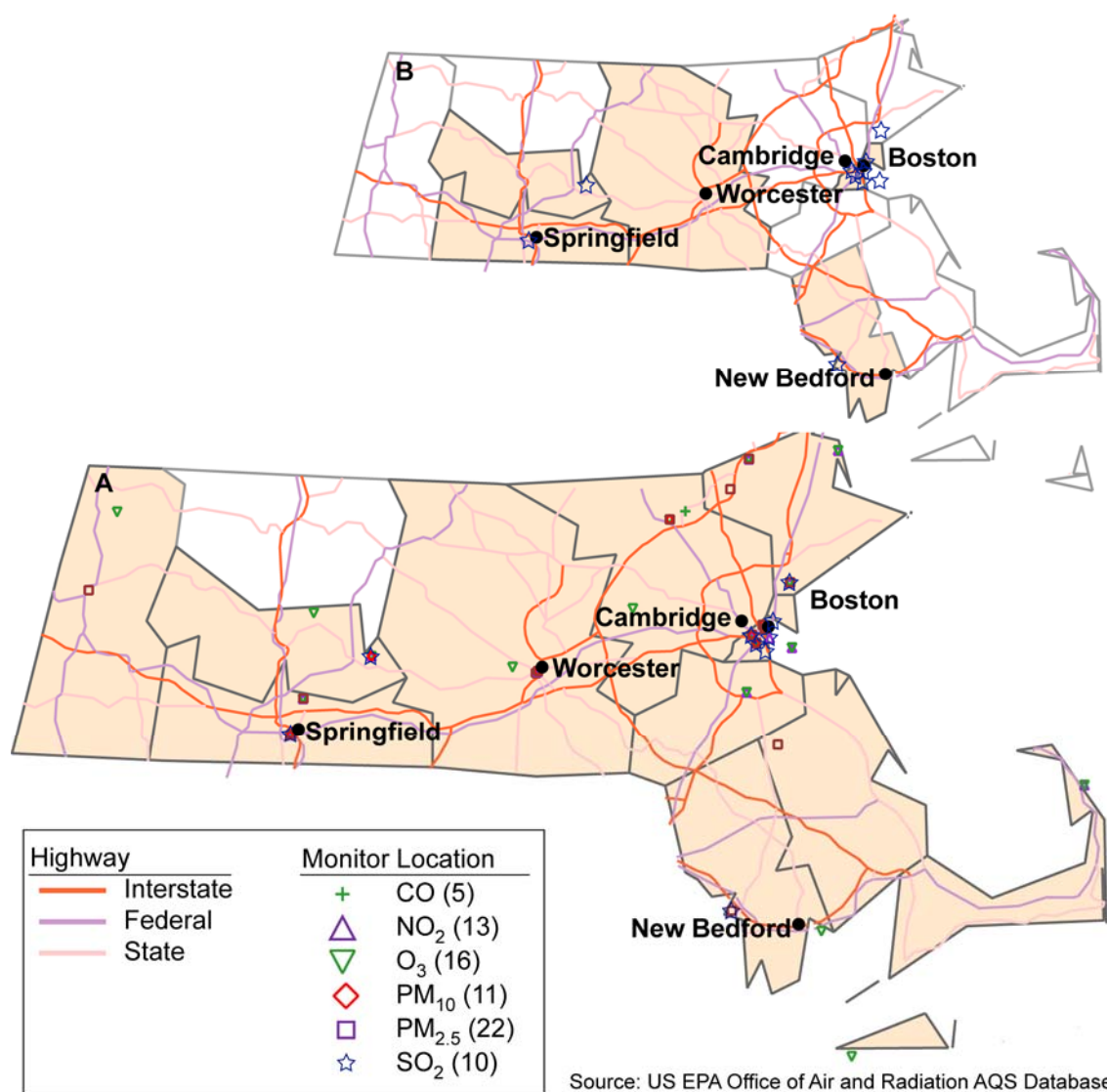


Figure 2-6. Criteria pollutant monitor locations (A) and SO₂ monitor locations (B), Massachusetts, 2005. Shaded counties have at least one monitor.

These emissions data trends are consistent with the trends in the observed ambient concentrations from the Clean Air Status and Trends Network (CASTNet). Following implementation of the Phase I controls on ARP sources between 1995 and 2000, significant reductions in SO₂ and ambient SO₄²⁻ concentrations were observed at CASTNet sites throughout the eastern United States. The mean annual concentrations of SO₂ and SO₄²⁻ from CASTNet's long-term monitoring sites can be compared using two 3-year periods, 1989–1991 and 2003–2005, shown in Figure 2-8 for SO₂ and Figure 2-9 for SO₄²⁻.

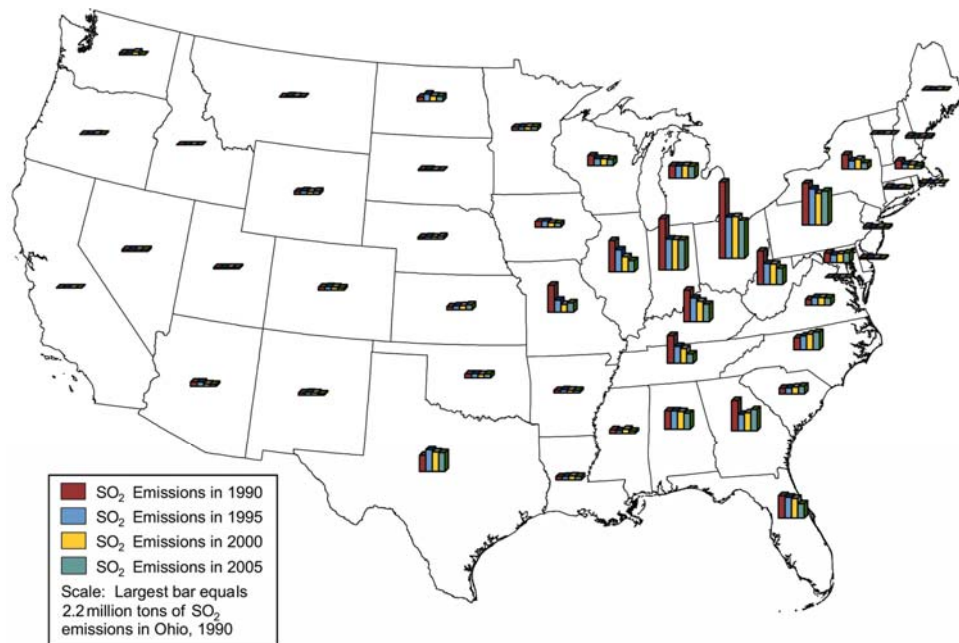


Figure 2-7. State-level SO₂ emissions, 1990-2005.

Source: Environmental Protection Agency Clean Air Markets Division (www.epa.gov/airmarkets/index.html).

From 1989 through 1991—that is, in the years prior to implementation of the ARP Phase I—the highest ambient mean concentrations of SO₂ and SO₄²⁻ were observed in western Pennsylvania and along the Ohio River Valley: > 20 µg/m³ (~8 ppb) SO₂ and > 15 µg/m³ SO₄²⁻. As with SO₂, in the years since the ARP controls were enacted, both the magnitude of SO₄²⁻ concentrations and their areal extent have been significantly reduced, with the largest decreases again along the Ohio River Valley.

Figure 2-10 depicts the magnitude and spatial distribution of SO₂ emissions in 2006 from sources in the ARP for the CONUS. This depiction clearly shows the continuing overrepresentation of SO₂ sources in the United States east of the Mississippi River, a trend even stronger in the central Ohio River Valley, as evident in the smoothed concentration plots in Figure 2-8. As shown in Table 2-3, regional distributions of SO₂ and SO₄²⁻ concentrations averaged for 2003–2005 reflect this geospatial emissions source difference as well.

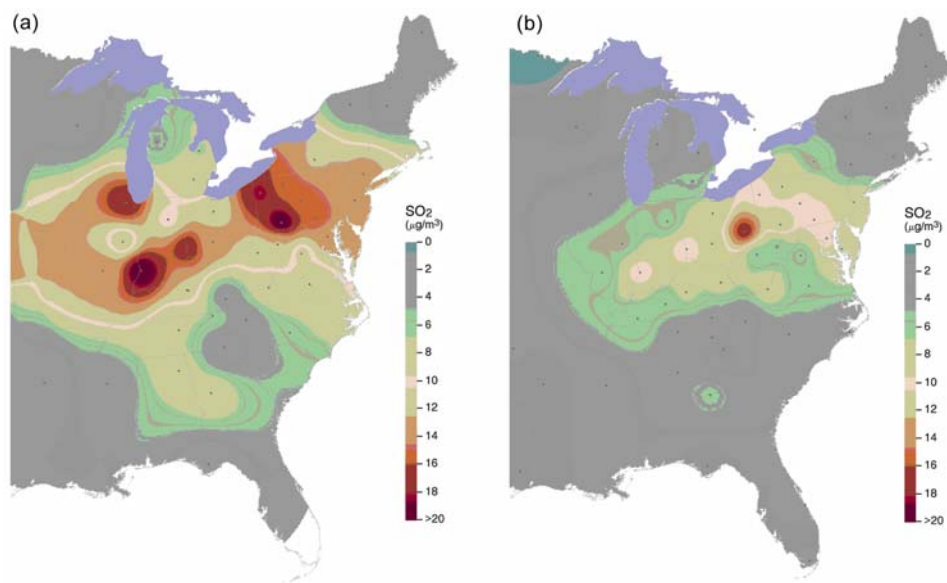


Figure 2-8. Annual mean ambient SO_2 concentration, 1989 through 1991 (a), and 2003 through 2005 (b).

Source: U.S. EPA CASTNet

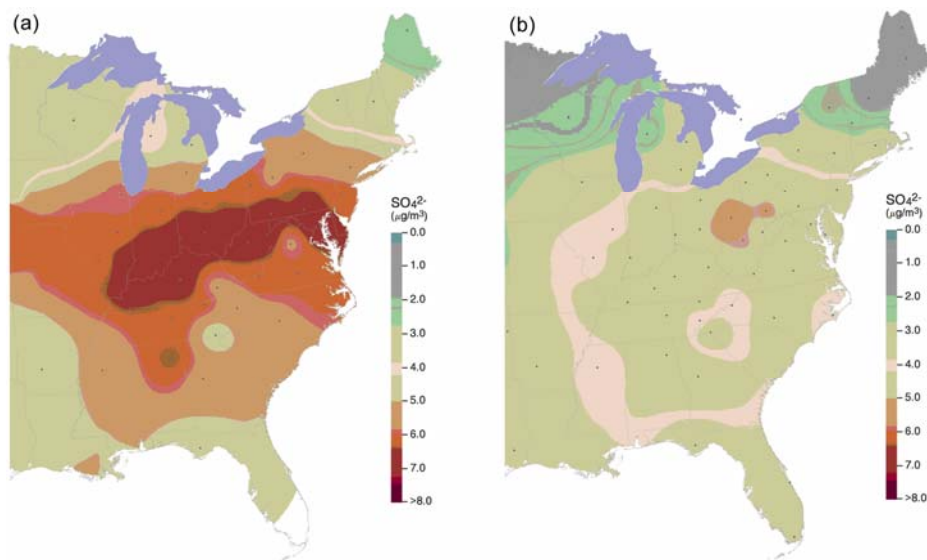


Figure 2-9. Annual mean ambient SO_4^{2-} concentration, 1989 through 1991 (a), and 2003 through 2005 (b).

Source: U.S. EPA CASTNet

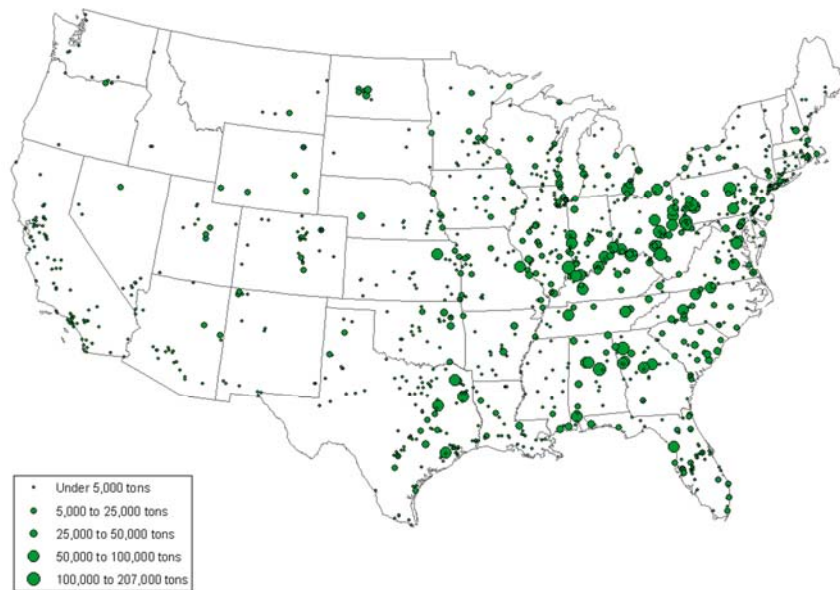


Figure 2-10. Annual SO₂ emissions for Acid Rain Program cooperating facilities, 2006.

Dots represent monitoring sites. Lack of shading for Southern Florida indicates lack of monitoring coverage.
Source: Environmental Protection Agency, Clean Air Markets Division (www.epa.gov/airmarkets/index.html).

2.4.4. Spatial and Temporal Variability of Ambient so₂ Concentrations

SO₂ concentrations have been falling throughout all regions of the CONUS, as demonstrated by the CASTNet data reviewed above. In and around most individual CMSAs, the trends are also toward lower SO₂ levels. Table 2-4 shows that many annual and even 1-h mean concentrations for the years 2003 through 2005 were consistently at or below the operating LOD of ~3 ppb for the standard sensitivity UV fluorescence SO₂ monitors deployed in the regulatory networks, while the aggregate mean value over all 3 years and all sites in and around the CMSAs was just above the LOD at ~4 ppb, and identical to the 1-h and 24-h means. Hence, it appears reasonable to aggregate up in time from available 1-h samples to daily and even annual exposure estimates.

Figure 2-11 shows the composite diel variation in hourly SO₂ concentrations in boxplot form from all monitors reporting SO₂ data into AQS. The AQS contains measurements of air pollutant concentrations in the 50 states, plus the District of Columbia, Puerto Rico, and the Virgin Islands, for the six criteria air pollutants (SO₂, NO₂, PM, CO, Pb, O₃), as well as for hazardous air pollutants.

Table 2-3. Regional distribution of SO₂ and SO₄²⁻ ambient concentrations, averaged for 2003-05.

REGION	CONCENTRATION	
	SO ₂ (ppb)	SO ₄ ²⁻ (µg/m ³)
Mid-Atlantic	3.3	4.5
Midwest	2.3	3.8
Northeast	1.2	2.5
Southeast	1.3	4.1

Table 2-4. Distributions of temporal averaging inside and outside CMSAs.

AVERAGING TIME MONITOR LOCATIONS	N	MEAN	PERCENTILES											MAX
			1	5	10	25	30	50	70	75	90	95	99	
1-h Max Concentration														
Inside CMSAs	332405	13	1	1	1	3	4	7	13	16	30	45	92	714
Outside CMSAs	53417	13	1	1	1	1	2	5	10	13	31	51	116	636
1-h Avg Concentration														
Inside CMSAs	7408145	4	1	1	1	1	1	2	4	5	10	15	34	714
Outside CMSAs	1197179	4	1	1	1	1	1	2	3	3	7	13	36	636
24-h Avg Concentration														
Inside CMSAs	327918	4	1	1	1	1	2	3	5	6	10	13	23	148
Outside CMSAs	52871	4	1	1	1	1	1	2	3	4	8	12	25	123
Annual Avg Concentration														
Inside CMSAs	898	4	1	1	1	1	2	4	5	6	8	10	12	15
Outside CMSAs	143	4	1	1	1	1	2	3	4	5	8	9	13	14
Aggregate 3-yr Avg Concentration, 2003-2005														
Inside CMSAs	283	4	1	1	1	2	3	3	5	5	8	10	12	14
Outside CMSAs	42	4	1	1	1	2	2	3	4	5	8	9	13	13

* Values are ppb

** CMSA = Consolidated Metropolitan Statistical Area

1 To be sure, the max 1-h concentration observed at some sites in and around some CMSAs
2 still exceeded the mean by a large margin, with max 1-h values of > 600 ppb. However, the 50th
3 percentile maximum value outside CMSAs, 5 ppb, was only slightly greater than the 1-h, 24-h,
4 and annual mean value, 4 ppb. The 50th percentile maximum value inside CMSAs, 7 ppb, was
5 75% greater than these longer-term averages, reflecting heterogeneity in source strength and

location. In addition, even with 1-h max values of > 600 ppb, the maximum annualized mean value for all CMSAs was still < 16 ppb, which is below the current annual primary SO₂ NAAQS.

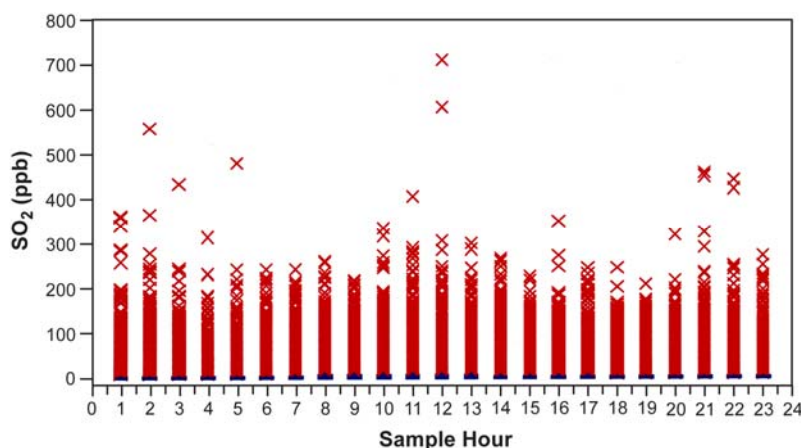


Figure 2-11. Boxplot of hourly SO₂ concentrations across all cities in focus.

The strong west-to-east increasing gradient in SO₂ emissions described above is well-replicated in the observed concentrations in individual CMSAs. For example, Table 2-5 shows the mean annual concentrations from 2003–2005 for the 12 CMSAs with four or more SO₂ regulatory monitors. Values ranged from a reported low of ~1 ppb in Riverside, CA and San Francisco, CA to a high of ~12 ppb in Pittsburgh, PA and Steubenville, OH, in the highest SO₂ source region.

The Pearson correlation coefficients (*r*) for multiple monitors in these CMSAs were generally very low for all cities, especially at the lower end of the observed concentration ranges, and even negative at the very lowest levels on the West Coast (see Table 2-5). This reflects strong heterogeneity in SO₂ ambient concentrations even within any one CMSA and, therefore, indicates possibly different exposures of spatially distinct subgroups of humans in these CMSAs to these very low concentrations of SO₂. At higher concentrations, the *r* values were also higher. In some CMSAs, this heterogeneity may result from meteorological effects, whereby a generally well-mixed subsiding air mass containing one or more SO₂ plumes with relatively high concentration would be more uniformly spread than faster-moving plumes with lower concentrations. However, instrument error may also play a role, because the highest *r* values, i.e., those > 0.7, correspond to the highest SO₂ concentrations, i.e., > 6 and > 10 ppb. Since the lowest SO₂ concentrations are at or below the operating LOD, and demonstrate the lowest correlation

across monitors that share at least some air mass characteristics most of the year, the unbiased instrument error in this range may be confounding interpretation of any possible correlation. This could be because the same actual ambient value would be reported by different monitors (with different error profiles) in the CMSA as different values in this lowest concentration range.

To better characterize the extent and spatiotemporal variance of SO₂ concentrations within each of the CMSAs having four or more SO₂ monitors, the means, minima, and maxima were computed from daily mean data across all available monitors for each month for the years 2003 through 2005. Because many of these CMSAs with SO₂ monitors also reported SO₄²⁻, it is possible to compute the degree of correlation between SO₂, the emitted species, and SO₄²⁻, the most prominent oxidized product from SO₂. SO₄²⁻ values, however, while averaged over all available data at each site are generally available at their monitoring sites on a schedule of only 1 in 3 days or 1 in 6 days. Furthermore, SO₂ and SO₄²⁻ monitors are not all co-located throughout the CMSAs. For each of the five example CMSAs in Figures 2-12 through 2-16, monthly aggregated values are depicted from daily means of: (a) the monthly mean, minimum, and maximum SO₂ concentrations; (b) the monthly mean, minimum and maximum SO₄²⁻ concentrations; and (c) a scatterplot of SO₂ versus SO₄²⁻ concentrations.

Table 2-5. Range of mean annual SO₂ concentrations and Pearson correlation coefficients in urban areas having at least four regulatory monitors, 2003–2005.

CMSA (# MONITORS)	MEAN SO ₂ CONCENTRATION (ppb)	PEARSON CORRELATION COEFFICIENT
Philadelphia, PA (10)	3.6 – 5.9	0.37 – 0.84
Washington, DC (5)	3.2 – 6.5	0.30 – 0.68
Jacksonville, FL (5)	1.7 – 3.4	–0.03 – 0.51
Tampa, FL (8)	2.0 – 4.6	–0.02 – 0.18
Pittsburgh, PA (10)	6.8 – 12	0.07 – 0.77
Steubenville, OH (13)	8.6 – 14	0.11 – 0.88
Chicago, IL (9)	2.4 – 6.7	0.04 – 0.45
Salt Lake City, UT (5)	2.2 – 4.1	0.01 – 0.25
Phoenix, AZ (4)	1.6 – 2.8	–0.01 – 0.48
San Francisco, CA (7)	1.4 – 2.8	–0.03 – 0.60
Riverside, CA (4)	1.3 – 3.2	–0.06 – 0.15
Los Angeles, CA (5)	1.4 – 4.9	–0.16 – 0.31

1 In Steubenville, OH (Figure 2-12), the area of highest SO₂ concentrations of all 12 CMSAs
2 with more than four monitors, all monthly mean SO₂ concentrations (a) were substantially < 30
3 ppb, though max daily means in some months were often > 60 ppb, or even > 90 ppb. Sulfate
4 data (b) at Steubenville were insufficient to make meaningful comparisons, though the 12
5 months of available SO₄²⁻ data suggest no correlation with SO₂ (c).

6 Next, consider Philadelphia, PA (Figure 2-13). SO₂ in Philadelphia, PA (a) is present at
7 roughly one-half the monthly mean concentrations in Steubenville, OH, and demonstrates a
8 strong seasonality with SO₂ concentrations peaking in winter. By contrast, SO₄²⁻ concentrations
9 in Philadelphia peak in the three summer seasons, with pronounced wintertime minima. This
10 seasonal anticorrelation still contains considerable monthly scatter, however.

11 Los Angeles, CA (Figure 2-14) presents a special case, since its size and power
12 requirements place a larger number of SO₂ emitters near it than would otherwise be expected on
13 the West Coast. Concentrations of SO₂ demonstrate weak seasonality in these 3 years, with
14 summertime means of ~3 to 4 ppb, and maxima generally higher than wintertime ones, though
15 the highest means and maxima occur during the winter of 2004–2005. SO₄²⁻ at Los Angeles
16 shows stronger seasonality, most likely because the longer summer days of sunny weather allow
17 for additional oxidation of SO₂ to SO₄²⁻ than would be available in winter. Weak seasonal effects
18 in SO₂ likely explain the complete lack of correlation between SO₂ and SO₄²⁻ here.

19 The Riverside, CA CMSA (Figure 2-15) presents the strongest example among the 12
20 examined for this study of correlation between SO₂ and SO₄²⁻, though even here the R² value is
21 merely 0.3. Seasonal peaks are obvious in summertime for SO₂ and SO₄²⁻, both at roughly one-
22 half the ambient concentrations seen in Los Angeles. This is very likely due to Riverside's
23 geographic location just downwind of the regionally large electric generating utility sources near
24 Los Angeles and the prevailing westerly winds in summer. Again, as with Los Angeles, the
25 summertime peaks in SO₄²⁻ are most likely due to the combination of peaking SO₂ and favorable
26 meteorological conditions allowing more complete oxidation.

27 Phoenix, AZ was the CMSA with the lowest monthly mean SO₂ and SO₄²⁻ concentrations
28 examined here (Figure 2-16). In Phoenix, nearly all monthly mean SO₂ values were at or below
29 the regulatory monitors' operating LOD of ~3 ppb. SO₄²⁻ concentrations were equivalently low,
30 roughly one-half the concentrations seen in Riverside, CA, for example. The monthly mean data
31 show strong summertime peaks for even these very low-level SO₄²⁻ observations, which, at ~1 to

1 $3 \mu\text{g}/\text{m}^3$, were generally one-half of those in Philadelphia. This suggests some seasonality in SO_2 ,
2 though anticorrelated with SO_4^{2-} ; however, the trend is very weak, as the correlation scatterplot
3 shows.

2.4.5. 5-Minute Sample Data in the Monitoring Network

4 Although the number of monitors across the CONUS varies somewhat each year, in 2006
5 there were ~500 SO_2 monitors in the NAAQS monitoring network (<http://www.epa.gov/air/data>).
6 The state and local agencies responsible for these monitors are required to report 1-h avg
7 concentrations to the EPA AQS. In addition, a very small number of sites—only 108 total from
8 1997 to 2006, and not the same sites in all years—voluntarily reported 5-min block avg to AQS.
9 Of these, 104 reported only the max 5-minute average, 15 reported all 12 5-minute avg in each
10 hour, and 11 of those 15 reported all 12 values each hour and maximum values for some fraction
11 of time between 1997 and 2006. See Table 2-6 and Table 2-7 for a breakdown of these monitor
12 locations and sampling periods, and Figure 2-17 for the geospatial distribution of these monitors
13 across the CONUS.

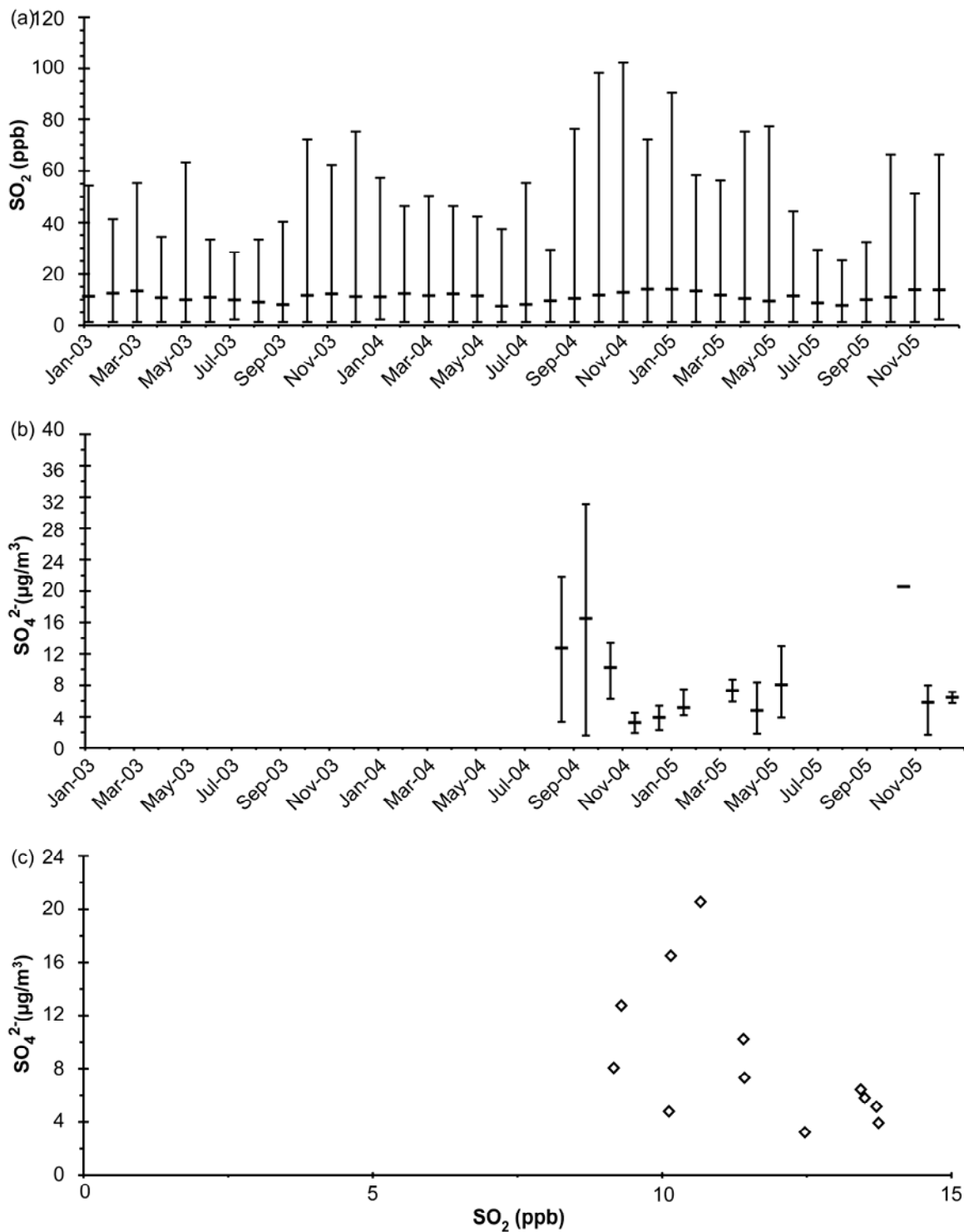


Figure 2-12. Steubenville, OH, 2003–2005. (a) Monthly mean, minimum, and maximum SO_2 concentrations. (b) Monthly mean, minimum, and maximum SO_4^{2-} concentrations. (c) Monthly mean SO_4^{2-} concentrations as a function of SO_2 concentrations.

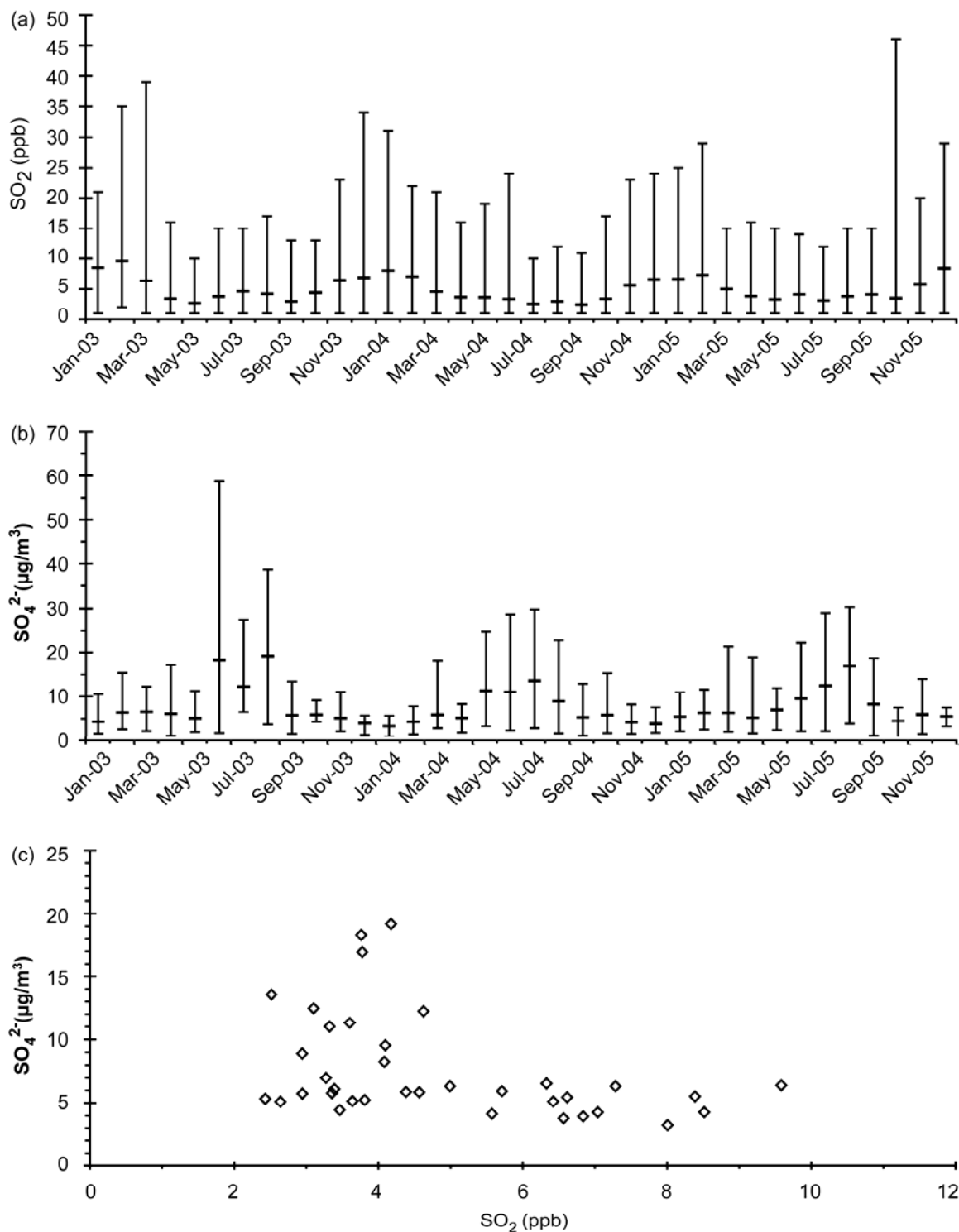


Figure 2-13. Philadelphia, 2003–2005. (a) Monthly mean, minimum, and maximum SO_2 concentrations. (b) Monthly mean, minimum, and maximum SO_4^{2-} concentrations. (c) Monthly mean SO_4^{2-} concentrations as a function of SO_2 concentrations.

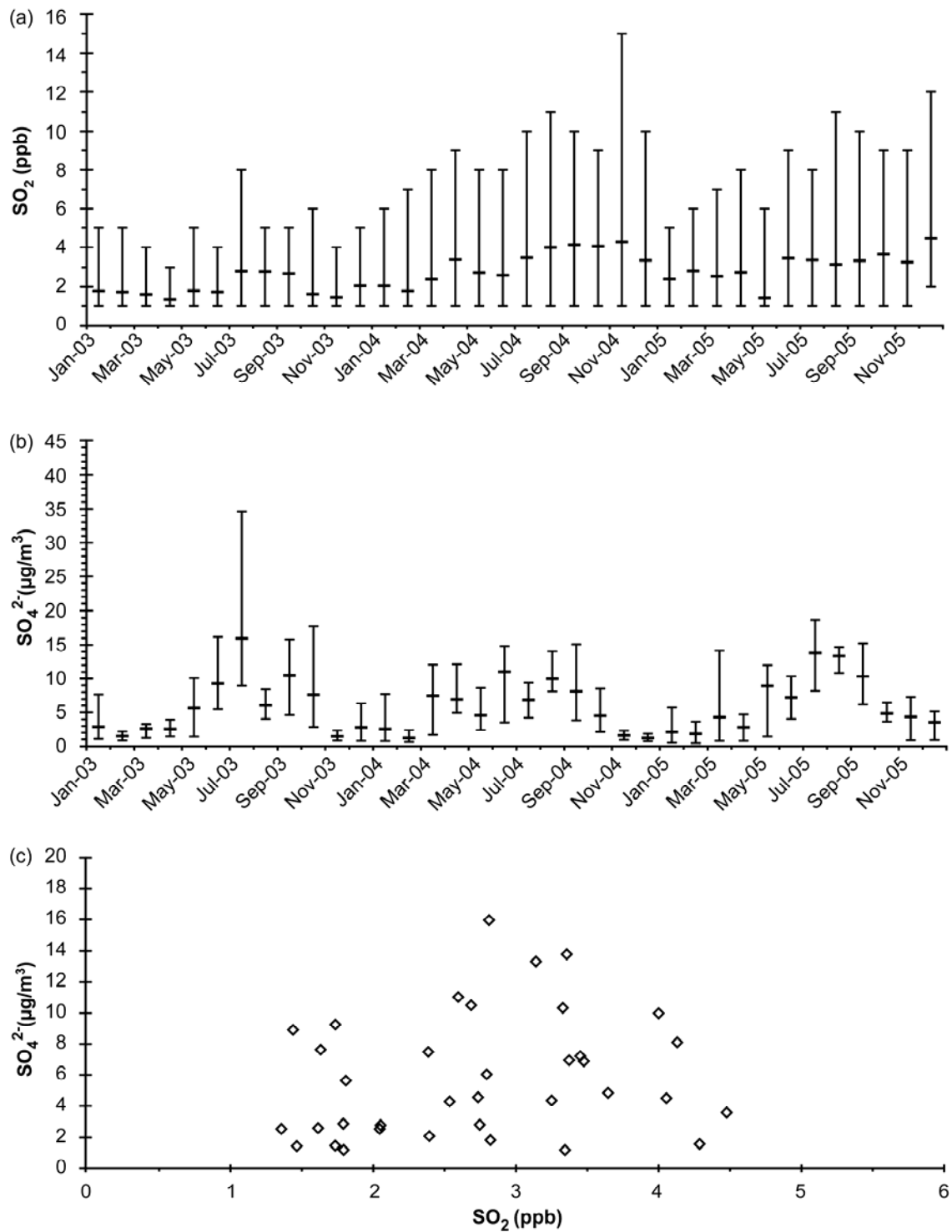


Figure 2-14. Los Angeles, 2003–2005. (a) Monthly mean, minimum, and maximum SO_2 concentrations. (b) Monthly mean, minimum, and maximum SO_4^{2-} concentrations. (c) Monthly mean SO_4^{2-} concentrations as a function of SO_2 concentrations.

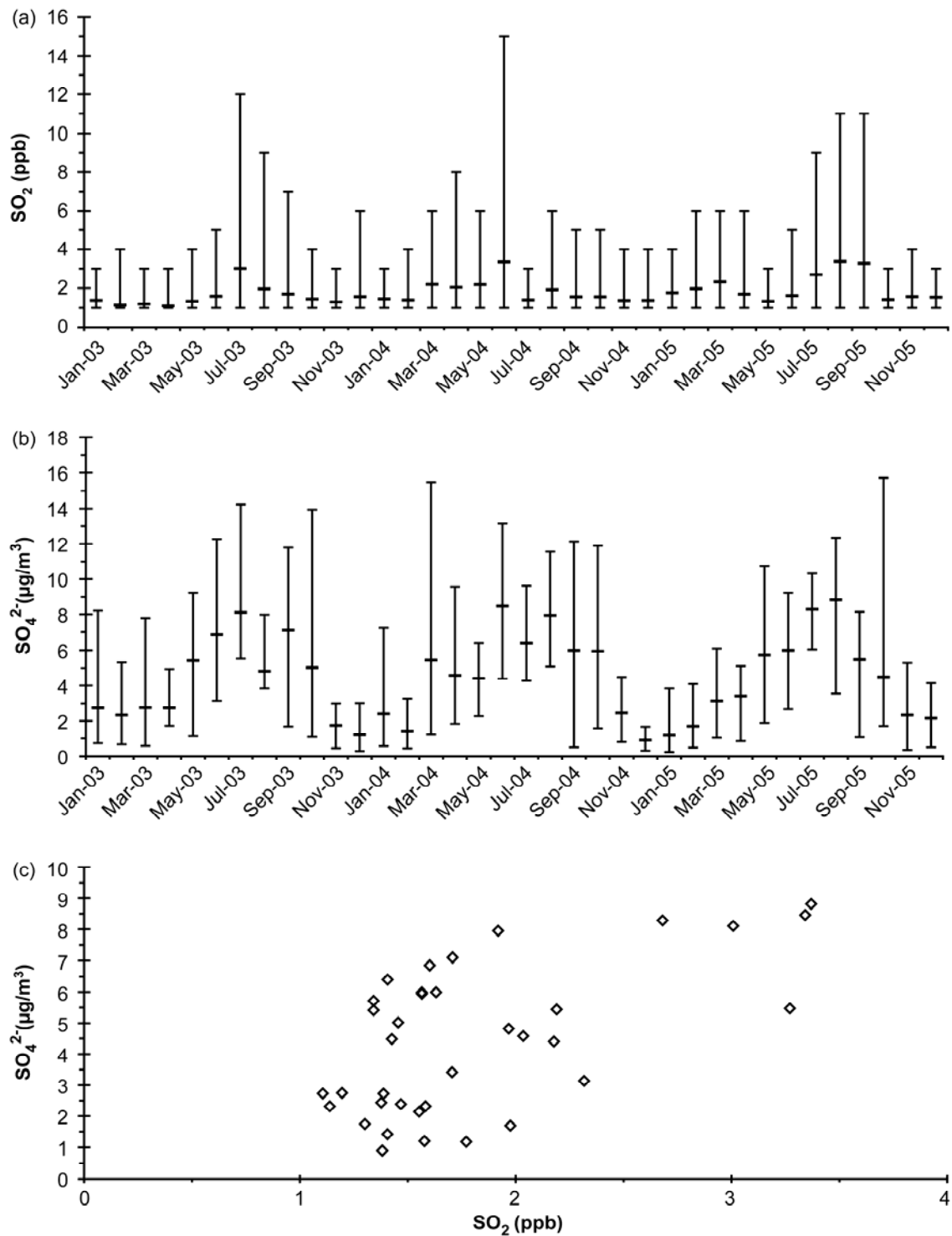


Figure 2-15. Riverside, CA, 2003–2005. (a) Monthly mean, minimum, and maximum SO_2 concentrations. (b) Monthly mean, minimum, and maximum SO_4^{2-} concentrations. (c) Monthly mean SO_4^{2-} concentrations as a function of SO_2 concentrations.

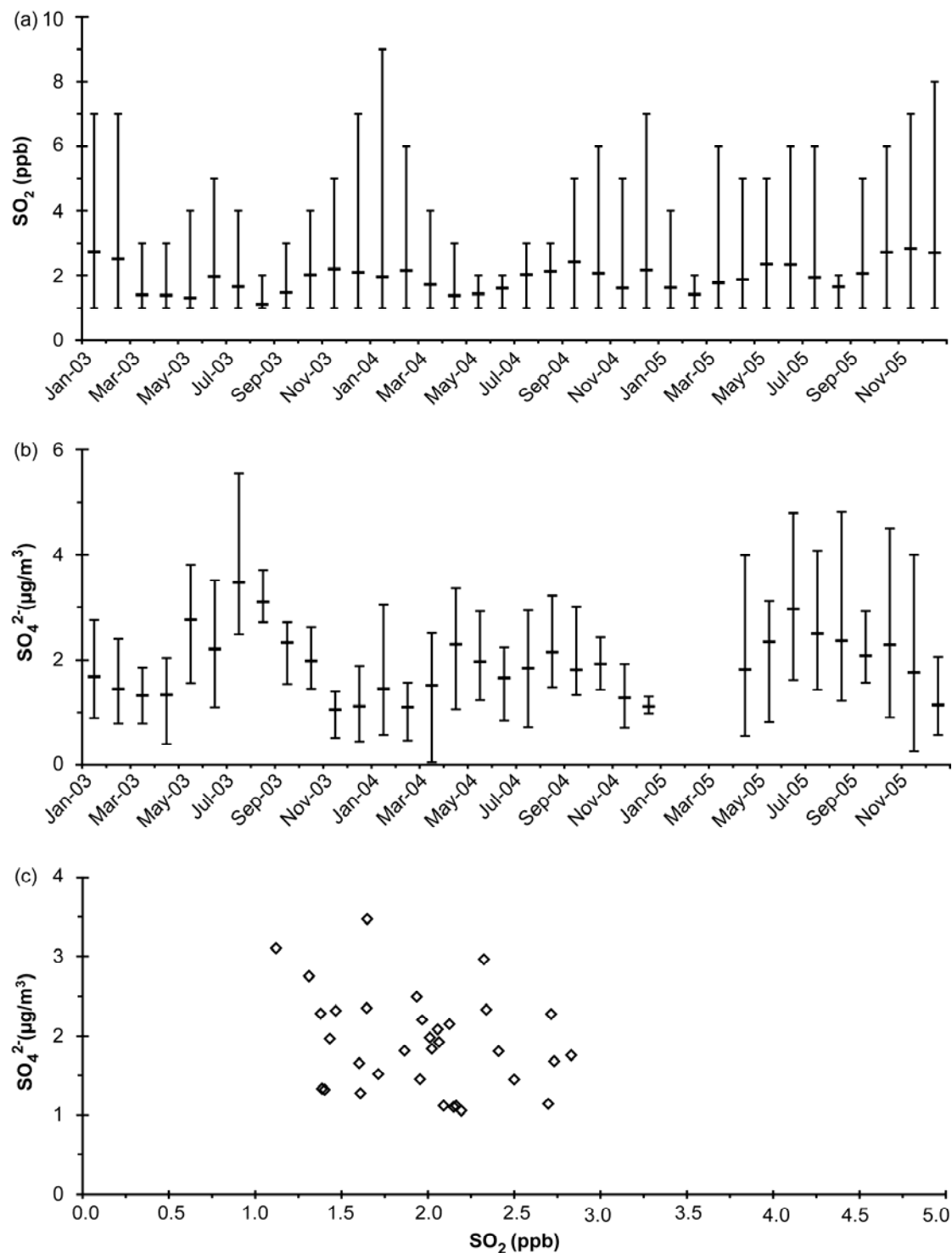


Figure 2-16. Phoenix, 2003–2005. (a) Monthly mean, minimum, and maximum SO_2 concentrations. (b) Monthly mean, minimum, and maximum SO_4^{2-} concentrations. (c) Monthly mean SO_4^{2-} concentrations as a function of SO_2 concentrations.

1 Although these 5-minute data meet AQS minimum quality assurance requirements, the
 2 voluntary nature of this reporting and the high variability across space and time make these data
 3 very difficult to use precisely.

Table 2-6. Locations, counts, and sampling periods of monitors reporting 5-minute maximum SO₂ values, 1997–2006.

STATE	NUMBER OF COUNTIES	NUMBER OF MONITORS	NUMBER OF YEARS	YEARS OPERATING
Arkansas	2	3	10	1997-2006
Colorado	1	1	10	1997-2006
Delaware	1	1	2	1997-1998
D.C.	1	1	5	2000-2004
Iowa	6	9	5	2001-2005
Louisiana	1	1	4	1997-2000
Missouri	7	14	10	1997-2006
Montana	1	7	10	1997-2006
North Carolina	1	1	8	1997-2004
North Dakota	11	19	10	1997-2006
Pennsylvania	8	23	7	1997-2003

Table 2-7. Locations, counts, and sampling periods of monitors reporting all 12 5-minute SO₂ values in each hour, 1997–2006.

STATE	NUMBER OF COUNTIES	NUMBER OF MONITORS	NUMBER OF YEARS	YEARS OPERATING
D.C.	1	1	1	2007
Florida	1	1	4	2002-2005
Missouri	1	2	4	2003-2006
Montana	1	4	1	2002
North Carolina	1	1	4	1999-2002
Pennsylvania	2	5	5	2002-2006
West Virginia	2	2	5	2001-2005

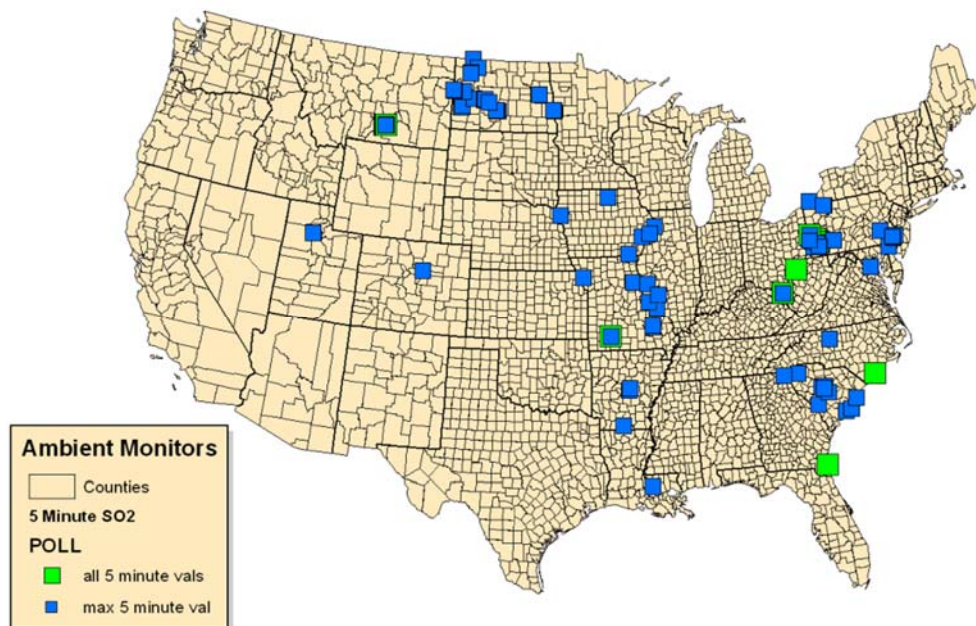


Figure 2-17. SO₂ monitors reporting maximum or continuous 5-minute average values for any period, 1997–2006.

2.4.6. Policy Relevant Background Contributions to SO₂ Concentrations

Background concentrations used for purposes of informing decisions about the NAAQS are referred to as Policy Relevant Background (PRB) concentrations. PRB concentrations are those concentrations that would occur in the United States in the absence of anthropogenic emissions in continental North America (defined here as the United States, Canada, and Mexico). PRB concentrations include contributions from natural sources everywhere in the world, and from anthropogenic sources outside these three countries. Background levels so defined facilitate separation of cases where pollution levels can be controlled by U.S. regulations (or through international agreements with neighboring countries), from cases where pollution is generally uncontrollable by the United States. EPA assesses risks to human health and environmental effects from SO₂ levels in excess of PRB concentrations.

Contributions to PRB concentrations include natural emissions of SO₂ and photochemical reactions involving reduced sulfur compounds of natural origin, as well as their long-range transport from outside of North America from any source. As an example, transport of SO₂ from Eurasia across the Pacific Ocean or the Arctic Ocean would carry PRB SO₂ into the U.S.

Annex B contains a schematic diagram showing the major photochemical processes involved in the sulfur cycle, including natural sources of reduced sulfur species from anaerobic microbial activity in wetlands and volcanic activity. Volcanoes and wildfires are the major natural source of SO₂. Biogenic emissions from agricultural activities are not considered in the formation of PRB concentrations. Discussions of the sources and estimates of emissions are given in Annex Section B.6.

The MOZART-2 global model of tropospheric chemistry (Horowitz et al., 2003) is used to estimate the PRB contribution to SO₂ concentrations. The model setup for the present-day simulation, i.e., including all sources in the U.S. Canada and Mexico, was published in a series of papers from a recent model intercomparison (Dentener et al., 2006; van Noije et al., 2006). MOZART-2 is driven by the National Oceanic and Atmospheric Administration's National Center for Environmental Prediction (NOAA/NCEP) meteorological fields and the International Institute for Applied Systems Analysis (IIASA) 2000 emissions at a resolution of 1.9° × 1.9° with 28 σ (sigma) levels in the vertical, and includes gas- and aerosol-phase chemistry. Results shown in Figure 2-18 are for the meteorological year 2001. An additional PRB simulation was conducted in which continental North American anthropogenic emissions were set to zero.

The role of PRB in contributing to SO₂ concentrations in surface air is examined first. Figure 2-18 shows the annual mean predicted SO₂ concentrations in surface air in the simulation including all sources, or the "base case" (top panel); the PRB simulation (middle panel); and the percentage contribution of the background to the total base case SO₂ (bottom panel). Maximum concentrations in the base case simulation, > 5 ppb, occur along the Ohio River Valley (upper panel). Background SO₂ concentrations are orders of magnitude smaller, below 10 parts per trillion (ppt) over much of the United States (middle panel). Maximum PRB concentrations of SO₂ are 30 ppt. In the Northwest where there are geothermal sources of SO₂, the contribution of PRB to total SO₂ is 70 to 80%; however absolute SO₂ concentrations are still of the order of a couple of ppb or less. With the exception of the West Coast where volcanic SO₂ emissions cause high PRB concentrations, PRB contributes < 1% to present-day SO₂ concentrations in surface air (bottom panel).

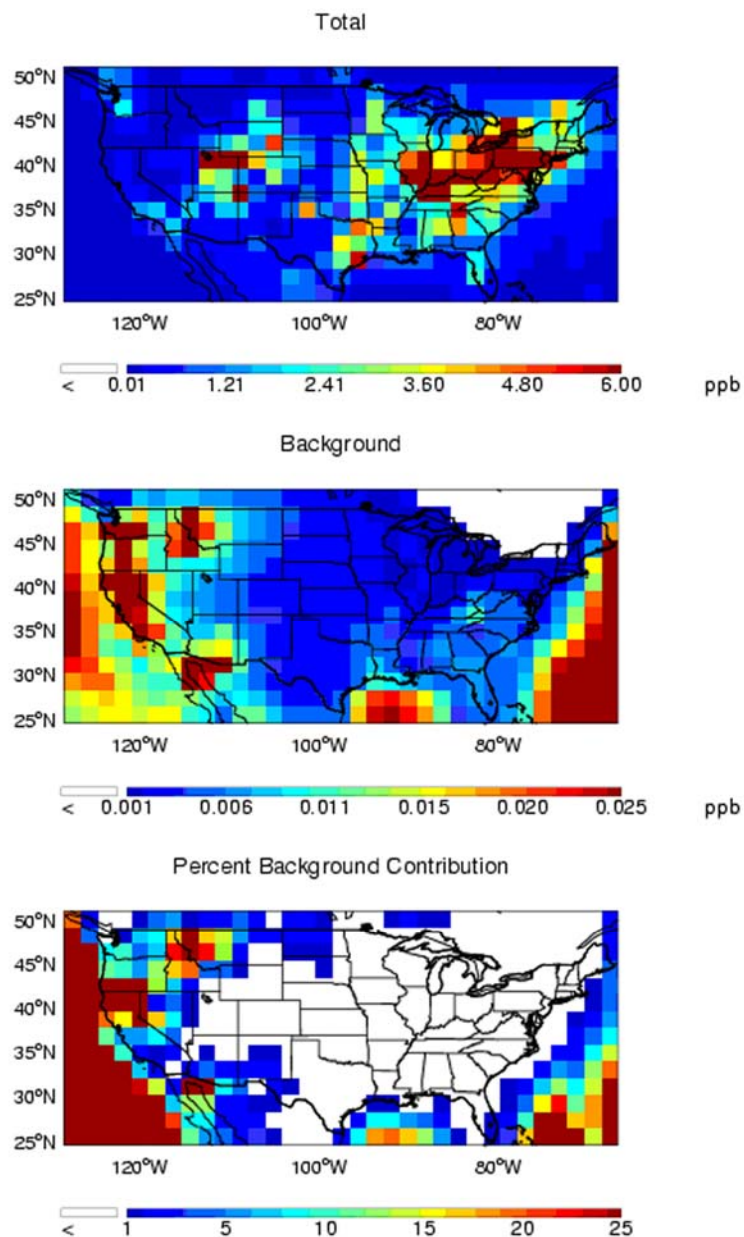


Figure 2-18. Annual mean model-predicted concentrations of SO₂ (ppb).

When estimating background concentrations it is instructive to consider measurements of SO₂ at relatively remote monitoring sites, i.e., sites located in sparsely populated areas not subject to obvious local sources of pollution. Berresheim et al. (1993) used a type of atmospheric pressure ionization mass spectrometer (APIMS) at Cheeka Peak, WA (48.30°N 124.62°W, 480 m asl), in April 1991 during a field study for DMS oxidation products. SO₂ concentrations ranged between 20 and 40 ppt. Thornton et al. (2002) have also used an APIMS with an

isotopically labeled internal standard to determine background SO₂ levels. SO₂ concentrations of 25 to 40 ppt were observed in northwestern Nebraska in October, 1999 at 150 m above ground using the National Center for Atmospheric Research (NCAR)'s C-130 research aircraft. These data are comparable to remote central South Pacific convective boundary layer SO₂ data (Thornton, 1999).

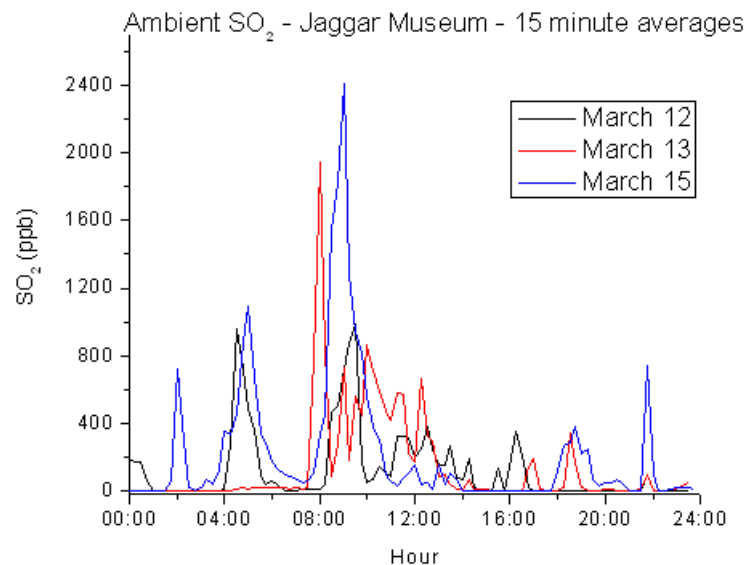


Figure 2-19. 15-minute average ambient SO₂ concentrations measured at Hawaii Volcanoes National Park monitoring sites, March 12, 13, and 15, 2007.

Source: National Park Service

As noted earlier, volcanic sources of SO₂ in the United States are found in the Pacific Northwest, Alaska, and Hawaii. The greatest potential domestic effects from volcanic SO₂ occurs on the island of Hawaii. Nearly continuous venting of SO₂ from Mauna Loa and Kilauea produces SO₂ in high concentrations (see Figure 2-19 and Figure 2-20) at two National Park sites near the Kilauea caldera and the nearby east rift zone. The latter emits several times as much SO₂ as the Kilauea caldera. The two measurement sites within the National Park are < 3 km from the summit emission source and ~10 km from the east rift source and are affected by the two sources during southerly and easterly winds. A number of communities and population centers are within the same distance from the east rift gas source that affects these two monitoring sites. When the

1 normal trade wind flows are disrupted, emissions from the sources can be brought directly to
2 these various communities. Since these communities are located at a similar distance from the
3 large east rift emission source as the Park monitoring stations, it is probable that these
4 communities experience SO₂ concentrations as high as those measured within Hawaii Volcanoes
5 National Park.

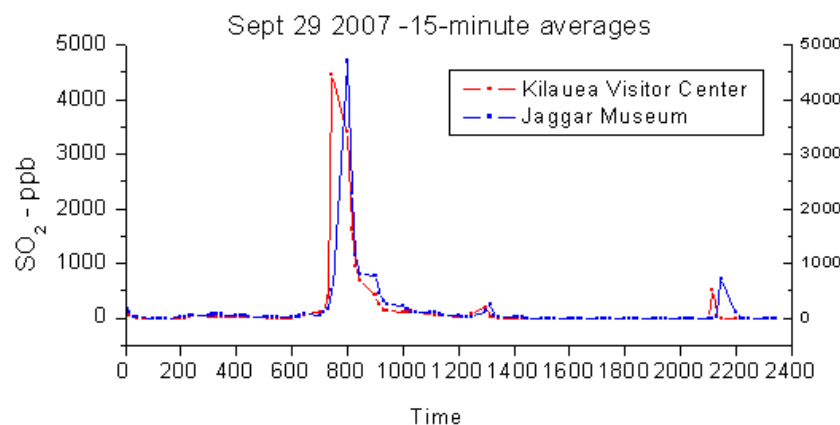


Figure 2-20. 15-minute average ambient SO₂ concentrations measured at the two National Park monitoring sites at Hawaii Volcanoes NP, Hawaii on September 29, 2007.

Source: National Park Service

6 Since 1980, the Mount St. Helens volcano (46.20°N, 122.18°W, summit 2549 m asl) in the
7 Washington Cascade range has been a variable source of SO₂. Its major effects came in the
8 explosive eruptions of 1980, which primarily affected the northwestern United States. The
9 Augustine volcano near the mouth of the Cook Inlet in southwestern Alaska (59.363°N,
10 153.43°W, summit 1252 m asl) has emitted variable quantities of SO₂ since its last major
11 eruptions in 1986. Volcanoes in the Kamchatka peninsula in far eastern Siberia do not
12 particularly affect the surface concentrations in northwestern North America.

13 Overall, the background contribution to SO₂ over the United States is relatively small, with
14 a max PRB of 0.030 ppb SO₂, except for areas with volcanic activity.

2.5. Issues Associated with Evaluating SO_2 Exposure

2.5.1. General Considerations for Personal Exposure

Human exposure to an airborne pollutant consists of contact between the human and the pollutant at a specific concentration for a specified period of time. People spend various amounts of time in different microenvironments characterized by different pollutant concentrations. The integrated exposure of a person to a given pollutant is the sum of the exposures over all time intervals for all microenvironments. Figure 2-21 represents a composite average of activity patterns across all age groups in the United States, based on data collected in the National Human Activity Pattern Survey (NHAPS) (Klepeis et al., 2001). The demographic distribution of the respondents was designed to be similar to that of overall U.S. Census data. Different cohorts, e.g., the elderly, young and middle-aged working adults, and children exhibit different activity patterns.

A person's exposure to a pollutant, such as SO_2 , can be represented by:

$$E_T = \sum_{i=1}^n C_i t_i \quad (2-7)$$

where E_T is an individual's total personal exposure for a specific time period, n is the total number of microenvironments encountered, C_i is the average concentration, and t_i is the time spent in the i th microenvironment. The exposure a person experiences can be characterized as: an instantaneous exposure; a peak exposure such as might occur during cooking; an average exposure; or an integrated exposure over all environments encountered. These distinctions are important because health effects caused by long-term low-level exposures may differ from those caused by short-term peak exposures.

An individual's total exposure (E_T) can also be represented by:

$$E_T = E_a + E_{na} = \{y_o + \sum_i y_i [P_i a_i / (a_i + k_i)]\} C_a + E_{na} = \{y_o + \sum_i y_i F_{inf_i}\} C_a + E_{na} \quad (2-8)$$

subject to the constraint

$$y_o + \sum_i y_i = 1 \quad (2-9)$$

where E_a is the ambient component of personal exposure, E_{na} is the nonambient component of personal exposure, y_o is the fraction of time spent outdoors, and y_i is the fraction of time spent in microenvironment i . F_{inf_i} , P_i , a_i , and k_i are the infiltration factor, penetration coefficient, air exchange rate, and decay rate, respectively for microenvironment i . In the case where an exposure occurs mainly in one microenvironment, Equation 2-8 may be approximated by Equation 2-10 where y is the fraction of time spent outdoors, and α is the ratio of personal exposure from a pollutant of ambient origin to the pollutant's ambient concentration (or the ambient exposure factor). Other symbols have the same definitions as in Equations 2-8 and 2-9.

$$E_T = E_a + E_{na} = \{y + (1-y)[Pa/(a+k)]\}C_a + E_{na} = \alpha C_a + E_{na} \quad (2-10)$$

If concentrations in a single microenvironment are considered, then Equation 2-10 can be recast as

$$C_{me} = C_a + C_{na} = [Pa/(a+k)]C_a + S/[V(a+k)] \quad (2-11)$$

where C_{me} is the concentration in a microenvironment, C_a and C_{na} are the contributions to C_{me} from ambient and nonambient sources, S is the microenvironmental source strength, and V is the volume of the microenvironment. (Bracketed symbols are same as Equation 2-8.) In this equation, it is assumed that microenvironments do not exchange air with each other, but only with the ambient environment.

Microenvironments in which people are exposed to air pollutants such as SO_2 typically include residential indoor environments, other indoor locations, near-traffic outdoor environments, other outdoor locations, and in vehicles, as shown in Figure 2-21. Indoor combustion sources such as gas stoves and space heaters need to be considered when evaluating exposures to SO_2 . Exposure misclassification may result when total human exposure is not disaggregated between various microenvironments, and this may obscure the true relationship between ambient air pollutant exposures and health outcomes.

NHAPS - Nation, Percentage Time Spent

Total n = 9,196

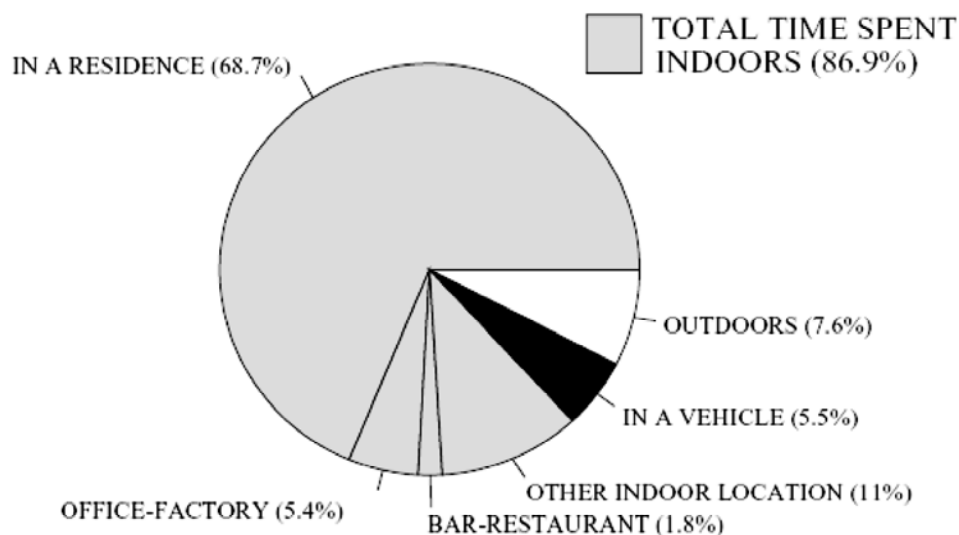


Figure 2-21. Percentage of time spent in various environments in the United States.¹

Source: Klepeis et al. (2001).

In a given microenvironment, the ambient component of a person's microenvironmental exposure to a pollutant is determined by the following physical factors:

- ambient concentration C_a
- air exchange rate a_i
- pollutant specific penetration coefficient P_i
- pollutant specific decay rate k_i
- fraction of time an individual spends in the microenvironment y_i

These factors are in turn affected by the following exposure factors:

- environmental conditions, such as weather and season

¹ For example, the cohort of working adults between the ages of 18 and 65 represents ~50% of the population. Of this total, about 60% work outside the home, spending ~24% (40 h/168 h) of their time in factory/office environments. Thus, this cohort is likely to spend considerably more time in offices and factories than shown in the figure (5.4%), which reflects the entire population, and is also likely to spend less time in a residence than small children or the elderly.

- dwelling conditions, such as the location of the house which determines proximity to sources and geographical features that can modify transport from sources, the amount of natural ventilation (e.g., open windows and doors, and the “draftiness” of the dwelling) and ventilation system (e.g., removal efficiency and operation cycle)
- personal activities (e.g., the time spent cooking or commuting)
- indoor sources and sinks of a pollutant

Microenvironmental exposures can also be influenced by the individual-specific factors such as age, gender, health or socioeconomic status.

Time-activity diaries, completed by study participants, are used to compile activity patterns for input to exposure models and assessments. The EPA’s National Exposure Research Laboratory (NERL) has consolidated the majority of the most significant human activity databases into one comprehensive database called the Consolidated Human Activity Database (CHAD). Eleven different human activity pattern studies were evaluated to obtain over 22,000 person-days of 24-h human activities in CHAD (McCurdy et al., 2000). These data can be useful in assembling population cohorts to be used in exposure modeling and analysis.

In general, the relationship between personal exposures and ambient concentrations can be modified by microenvironments. During infiltration, ambient pollutants can be lost through chemical and physical loss processes, and therefore, the ambient component of a pollutant’s concentration in a microenvironment is not the same as its ambient concentration but the product of the ambient concentration and the infiltration factor (F_{inf} or α if people spend 100% of their time indoors). In addition, exposure to nonambient, microenvironmental sources modifies the relationship between personal exposures and ambient concentrations.

In practice, it is extremely difficult to characterize community exposures by measurements of each individual’s personal exposures. Instead, the distribution of personal exposures in a community, or the population exposure, is simulated by extrapolating measurements of personal exposure using various techniques or by stochastic, deterministic or hybrid exposure modeling approaches such as APEX, SHEDS, and MENTOR (see Annex Section C.2 for a description of modeling methods). Variations in community-level personal exposures are determined by cross-community variations in ambient pollutant concentrations and the physical and exposure factors mentioned above. These factors also determine the strength of the association between population exposure to SO_2 of ambient origin and ambient SO_2 concentrations.

Of major concern is the ability of SO₂, measured by ambient monitors, to serve as a reliable indicator of personal exposure to SO₂ of ambient origin. The key question is what errors are associated with using SO₂ measured by ambient monitors as a surrogate for personal exposure to ambient SO₂ and/or its oxidation products in epidemiological studies. There are three aspects to this issue: (1) ambient and personal sampling issues; (2) the spatial variability of ambient SO₂ concentrations; (3) the associations between ambient concentrations and personal exposures as influenced by exposure factors, e.g., indoor sources and time spent indoors and outdoors. Items (1) and (3) are treated individually in the following sections; item (2) was treated previously in Section 2.4.2.

2.5.2. Methods Used for Monitoring Personal Exposure

Three basic methods of analysis have been used as personal exposure monitors (PEMs) to measure personal exposure to SO₂. The Harvard-EPA annular denuder system (HEADS) was initially developed to measure particles and acid gases simultaneously (Brauer et al., 1999; Koutrakis et al., 1988). The aerosol is initially sampled at 10 L/min through an impactor that is attached to an annular denuder to remove particles. Subsequently, the aerosol is sampled through an annular denuder coated with sodium carbonate (Na₂CO₃). This denuder is used to trap SO₂, nitric acid (HNO₃), and nitrous acid (HNO₂). Following sampling, the denuder is extracted with ultrapure water and analyzed by ion chromatography. Collection efficiencies of SO₂ in the denuder are typically around 0.993, which compares well with predicted values.

For a study conducted in Baltimore, MD, Chang et al. (2000) developed and employed a personal roll-around system (RAS, an active sampling system designed to measure short-term exposure) to measure personal exposure concentrations of several atmospherically relevant species, including SO₂. For the measurement of SO₂, the RAS employed an NO₂/SO₂ sorbent denuder worn on a vest by the study participant. The hollow glass denuder, encased in an aluminum jacket, is coated with triethanolamine (TEA) for the collection of SO₂ and NO₂, and aerosol is sampled through the denuder at 100 cc/min. Following sampling, the denuder can be extracted and analyzed for SO₂ concentrations by ion chromatography. The detection limit for 1-h sampling of SO₂ was reported to be 62 ppb, which resulted in many of the 1-h samples being below the LOD.

1 The most commonly employed SO₂ PEM method for personal exposure studies is the
2 passive badge sampler. A personal multipollutant sampler has been developed to measure
3 particulate and gaseous pollutants simultaneously (Demokritou et al., 2001). A single elutriator,
4 operating at 5.2 L/min, is employed to sample particulate pollutants. A passive SO₂ badge is
5 attached diametrically to the elutriator, which has been coated with Teflon to minimize reactive
6 gas losses. The passive badge sample is coated with TEA for the collection of SO₂ and NO₂.
7 Because wind speed can affect the collection rate of the passive badge sampler, this system
8 employs a constant face velocity across the passive badge sampler. For 24-h sampling times, the
9 estimated limit of detection (LOD) for SO₂ is 5 ppb.

10 Currently, limits exist for using PEM systems to measure personal exposure to SO₂.
11 Because SO₂ concentrations have been declining annually in the United States, little focus has
12 been placed on improving the methods of analysis. LODs for SO₂ PEMs (~5-10 ppb for 24 hr
13 sampling) are often greater than the concentrations of SO₂ that are typically observed in urban
14 ambient environments. However, much lower detection limits can be achieved by extending the
15 sampling time (Kasper-Giebl et al., 1999). Personal exposure monitoring studies often suffer
16 from having many of the daily SO₂ samples (e.g., 30 to 70%) collected below the sampler's LOD
17 (see Tables 2-10 and 2-11). Because of these issues, current methods can not characterize hourly
18 or shorter exposures unless these values are in the range of several tens to hundreds of ppb.

2.5.3. Relationship between Personal Exposure and Ambient Concentration

19 Because SO₂ concentrations have declined markedly over the past few decades, relatively
20 few studies have focused on SO₂. Another consideration is that current indoor and outdoor levels
21 in many areas are often beneath detection limits for passive personal SO₂ monitors.

2.5.3.1. Indoor Versus Outdoor SO₂ Concentrations

22 Several studies in the United States, Canada, Europe, and Asia have examined the
23 relationships of indoor, outdoor, and personal concentrations of SO₂ to ambient SO₂
24 concentrations. Perhaps the most comprehensive set of indoor-outdoor data was obtained by
25 Spengler et al. (1979) during the Harvard Six Cities Study. These data are shown in Figure 2-22.

1 Twenty-four-hour ambient and indoor SO₂ concentrations were measured every sixth day for
 2 1 year in a minimum of 10 homes or public facilities for each of the cities studied.

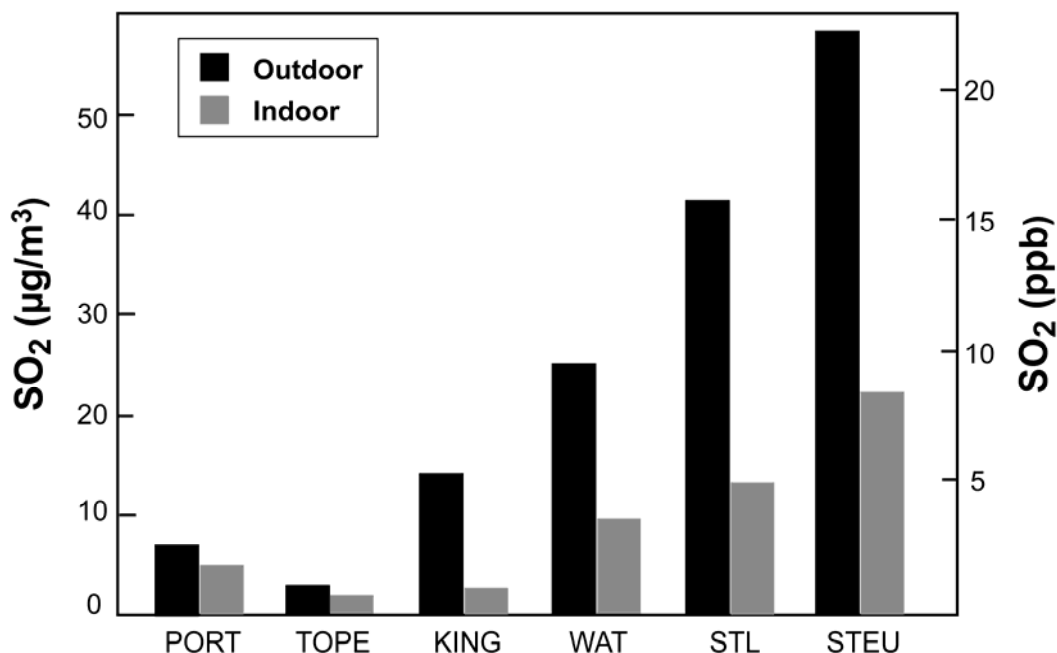


Figure 2-22. Average annual indoor and outdoor SO₂ concentrations for each of the six cities included in the analysis.

PORT = Portage, WI TOPE = Topeka, KS KING = Kingston, TN
 WAT = Watertown, MA STL = St. Louis, MO STEU = Steubenville, OH.

Source: Adapted from Spengler et al. (1979).

3 As can be seen from Table 2-8, a wide range is found in the ratio of indoor to outdoor
 4 concentrations among the different studies. These differences among studies could be due in part
 5 to differences in building characteristics (e.g., residences versus schools or other public
 6 buildings), in activities affecting air exchange rates, and in analytical capabilities. In several
 7 studies, high values for R² were found, suggesting that indoor levels were largely driven by
 8 outdoor levels. A few studies found higher levels of SO₂ indoors than outdoors in some samples.
 9 This situation could have arisen if there were indoor sources or because of analytical
 10 measurement issues. One would expect to find lower concentrations indoors than outdoors,
 11 because SO₂ is consumed by reactions on indoor surfaces, especially those that are moist. Chao
 12 acknowledged this point but could not account for the findings of this study. It was noted that

two samples had unusually high indoor to outdoor ratios and that the mean ratios would have been much lower otherwise. Winter-summer differences in the indoor:outdoor ratio are consistent with seasonal differences in air exchange rates, as noted by Brauer et al. (1991).

Table 2-8. Relationships of indoor to outdoor SO₂ concentrations.

REFERENCE	LOCATION	INDOOR TO OUTDOOR RATIO V(# SAMPLES)	NOTES
Spengler et al. (1979)	Portage, WI	0.67 (349)	One year during Harvard Six Cities Study. West-Gaeke method.
	Topeka, KS	0.50 (389)	
	Kingston, TN	0.08 (425)	
	Watertown, MA	0.33 (486)	
	St. Louis, MO	0.31 (543)	
	Steubenville, OH	0.39 (499)	
Stock et al. (1985)	Houston, TX	0.54 (2425)	May to October, continuous FRM for indoor and outdoor.
Meranger and Brule (1987)	Antigonish, NS, Canada	0.84 (8)	Early spring, 1 wk avg in 1 house with oil furnace, FPD-TA
Brauer et al. (1989)	Boston, MA	0.23 (24)	Summer, HEADS
Li and Harrison (1990)	Essex, UK	0.22	Summer
Brauer et al. (1991)	Boston, MA	0.39 (geom. mean) (29), R ² = 0.89	Summer, HEADS
		0.05 (geom. mean) (23), R ² = 0.73	Winter, HEADS
Chan et al. (1994)	Taipei, Taiwan	0.24 (15)	Summer, PS
		0.23 (37)	Winter, PS
Lee et al. (1999)	Hong Kong	0.92, R ₂ = 0.56	Winter, PF
Patterson and Eatough (2000)	Lindon, UT	0.027 ± 0.0023, R ² = 0.73	Winter, ADS, all samples
Kindzierski and Sembaluk (2001)	Boyle, Alberta, Canada	0.12 (12)	Late Fall, PS
	Sherwood Park, Alberta, Canada	0.14 (13)	
Chao (2001)	Hong Kong	1.01 ± 0.78 (10)	Summer. Windows mainly kept closed, PS
Kindzierski and Ranganathan (2006)	Fort McKay, Alberta, Canada	0.35 (30)	Fall. All indoor levels < LOD and set =1/2 LOD, PS

FPD-TA = Flame Photometric Detection-Thermal Analysis
HEADS = Harvard-EPA Annular Denuder System
PS = passive sampler

PF = pulsed fluorescence
FRM = Federal Reference Method
ADS = Annular Denuder System

Indoor, or nonambient, sources of SO₂ could complicate the interpretation of associations between personal exposure to ambient SO₂ in exposure studies. Possible sources of indoor SO₂ are associated with the use of sulfur-containing fuels, with higher levels expected when emissions are poorly vented. Brauer et al. (2002) noted that only one study (Biersteker et al., 1965) conducted inferential analyses of potential determinants of exposure to indoor SO₂ levels.

1 In the Biersteker et al. study, conducted in the Netherlands, indoor levels increased with oil, coal,
2 and gas heating, as well as smoking in homes and increased outdoor levels.

3 Triche et al. (2005) measured SO₂ levels in homes in which secondary heating sources
4 (fireplaces, kerosene heaters, gas space heaters, and wood stoves) were used. They found
5 elevated indoor levels of SO₂ when kerosene heaters were in use. Median levels of SO₂ when
6 kerosene heaters were used (6.4 ppb) were much higher than when they were not in use
7 (0.22 ppb). The maximum SO₂ level associated with kerosene heater use was 90.5 ppb. They did
8 not find elevated SO₂ levels when the other secondary heating sources were in use.

2.5.3.2. Relationship of Personal Exposure to Ambient Concentration

9 A few studies evaluated the association of personal exposure to SO₂ to ambient
10 concentrations (Brauer et al., 1989; Chang et al., 2000; Sarnat et al., 2000; 2001; 2005; 2006).
11 Some of these studies fall under the umbrella of the Health Effects Institute's Characterization of
12 Particulate and Gas Exposures of Sensitive Subpopulations Living in Baltimore and Boston
13 research plan (Koutrakis et al., 2005). However, the focus of many of these studies has been
14 exposure to particles, with acid gases included to evaluate confounder or surrogate issues.

15 Table 2-9 summarizes the longitudinal correlation coefficients between personal SO₂
16 exposures and ambient concentrations of SO₂, and Table 2-10 the pooled correlation coefficients.
17 Most of the studies examined lack the ability to quantify 24-h averaged personal SO₂ exposures
18 due to the low ambient SO₂ concentrations and the limitations of passive sampling, except two
19 studies conducted by Brauer et al. (1989) and Sarnat et al (2006).

20 Brauer et al. (1989) determined the slope of the regression line between personal and
21 ambient concentrations to be 0.13 ± 0.02 , $R^2 = 0.43$, based on 44 measurements made in Boston,
22 MA during the summer of 1988. Most if not all of the data points obtained using the HEADS
23 appeared to be above the working detection limits as defined by the authors in their publications
24 (Brauer et al., 1989; Koutrakis et al., 1988). Note that calculating detection limits in this way
25 could result in lower detection limits than if field blanks were used. The authors reported
26 significance at the $p < 0.001$ level, but the intercept was not significant at the $p < 0.001$ level.
27 Since the stationary monitoring site was located at an elevation of 250 m above street level, the
28 use of data from this ambient monitoring site will overestimate personal exposure, as the
29 concentration of SO₂ increases with height because it is emitted mainly by elevated point

sources. Indeed, the ambient concentrations are about a factor of two higher than the outdoor concentrations. Sarnat et al. (2006) reported that ambient SO₂ was observed to be significantly associated with personal SO₂ exposures during the fall (slope = 0.08 for overall population) in a study in Steubenville, OH. The authors also observed the effect of ventilation on the association between personal exposures and ambient concentrations (slope = 0.07 for subjects in buildings with low ventilation rates, and 0.13 for subjects in buildings with high ventilation rates).

Table 2-9. Association between personal exposure and ambient concentration (longitudinal correlations coefficients).

REFERENCE	STUDY DESIGN	SEASON	MEAN CONC. (ppb)	SLOPE	INTERCEPT	r, R ²	COMMENTS
Sarnat et al. (2000)	Longitudinal, Baltimore, 20 senior, healthy, nonsmoking people (average age 75), summer of 1998 and winter of 1999, 1 day averaged sample, for 12 consecutive days for each subject; four to six subjects were measured concurrently during each 12-day monitoring period.	Winter	Ambient: 6.6 – 10.2 Personal: -0.8 – 1.2	NR	NR	-0.75 to 0.65 (r) with a median of 0.02 (14 subjects)	The LOD for 24-h sampling was 6.5 ppb. All personal samples were below LOD.
Sarnat et al. (2001)	Longitudinal, Baltimore, 56 seniors, schoolchildren, and people with COPD, summer of 1998 and winter of 1999, 14 of 56 subjects participated in both sampling seasons; all subjects were monitored for 12 consecutive days (24-h avg samples) in each of the one or two seasons, with the exception of children who were measured for 8 consecutive days during the summer.	Winter	Ambient: 4 – 17 Personal: -2 – 3	-0.05* (N = 487 with 45 subjects)	0.54* (N = 487 with 45 subjects)	-0.75 to 0.6 (r) with a median of -0.1 (44 subjects)	1) Concentrations are estimated from Figure 1 in the paper. 2) Correlation coefficients are estimated from Figure 2 in the paper. 3) LOD was referred to Sarnat et al (2000), which was 6.5 ppb. Therefore, all personal samples were below LOD.
Sarnat et al. (2005)	Longitudinal, Boston, 43 seniors and schoolchildren, summer of 1999 and winter of 2000. Similar study design as Sarnat et al. (2001).	Summer	Ambient: 2.8 – 4.5 Personal: 0.3 – 0.5	0.00 (N = 335)	NR	-0.60 to 0.70 (r) with a median of 0.00 (Sample size NR)	1) Correlation coefficients are estimated from Figure 1 in the paper. 2) LOD was 2.3 ppb, and 96.5% of personal samples were below LOD.
		Winter	Ambient: 4.9 – 10.7 Personal: -0.3 – 1.9	-0.02 (N = 299)	NR	-0.55 to 0.60 (r) with a median of 0.10 (Sample size NR)	1) Correlation coefficients are estimated from Figure 1 in the paper. 2) LOD was 3.2 ppb, and 95.4% of personal samples were below LOD.

* significant at $\alpha = 0.05$ level

The associations between personal exposure and ambient concentration cannot be examined in the other studies because almost all the personal exposure concentrations were beneath

detection limits. For example, Chang et al. (2000) tested a new personal active sampling device (a RAS with a TEA-based denuder) on volunteer participants to measure hourly personal exposure to SO₂. However, the method detection limit was too high for SO₂ (62 ppb for 1-h sampling) to generate a robust SO₂ exposure dataset to perform further analysis, and so the authors did not use the SO₂ data.

Table 2-10. Association between personal exposure and ambient concentration (pooled correlations coefficients).

REFERENCE	STUDY DESIGN	SEASON	MEAN CONC. (ppb)	SLOPE	INTERCEPT	r, R ²	COMMENTS
Brauer et al. (1989)	Pooled, Boston, study population was NR, the number of participants was estimated to be 48, July and August of 1988 for 24 days, 1 day averaged sample, two subjects were monitored each day.	Summer	Ambient: 2.5 – 9.5 Personal: 0.4 – 1.8	0.13* (N = 44)	Not significant	0.43 (R ²)	1) Concentrations estimated from Figure 2 in the paper. 2) Central site monitor was 250 m above the ground level. 3) LOD for personal samples was ~0.19 ppb based on the way to determine the LOD for an active sampling system.
Sarnat et al. (2006)	Steubenville, 15 senior subjects, summer and fall of 2000, two consecutive 24-h samples were collected for each subject for each wk, 23 wks total. Correlation coefficients were calculated in the pooled data set.	Summer	Ambient: 2.7 ± 3.9 Personal: 1.5 ± 3.3	0.03 (N = 106)	NR	0.00 (R ²)	LOD was 5.5 ppb; 53.5% of personal samples were below LOD.
		Fall	Ambient: 5.4 ± 9.6 Personal: 0.7 ± 1.9	0.08* (N = 152)	NR	0.15 (R ²)	LOD was 3.8 ppb, and 31.6% of personal samples were below LOD.

* significant at $\alpha = 0.05$ level

In the context of determining the effects of ambient pollutants on human health, the association between the ambient component of personal exposures and ambient concentrations is more relevant than the association between personal total exposures (ambient component + nonambient component) and ambient concentrations. As described in Equations 2-8 and 2-10, personal total exposure can be decomposed into two parts; an ambient and a nonambient component. Usually, the ambient component of personal exposure is not directly measureable, but it can be estimated by exposure models, or the personal total exposure can be regarded as the personal exposure of ambient origin if there are no indoor or nonambient sources. It is expected that the association between ambient concentrations and the ambient component of personal exposures would be stronger than the association between ambient concentrations and personal total exposures as long as the ambient and nonambient component of personal total exposure are independent. None of the studies examined indoor sources, however, indoor sources are not

expected to be present. The correlation coefficients between personal ambient SO₂ exposures and ambient SO₂ concentrations in different types of exposure studies are relevant to different types of epidemiologic studies.

There are three types of correlations generated from different study designs and ways to analyze the data from exposure studies: longitudinal, “pooled,” and daily-average correlations (EPA, 2004). Longitudinal correlations¹ are calculated when data from a study includes measurements over multiple days for each subject (longitudinal study design). Longitudinal correlations describe the temporal relationship between daily personal SO₂ exposure or microenvironment concentration and daily ambient SO₂ concentration for the same subject. The longitudinal correlation coefficient can differ between subjects (i.e., each person may have a different correlation coefficient). The distribution of correlations for each subject across a population could be obtained with this type of data (e.g., Sarnat et al., 2000; 2001; 2005). A longitudinal correlation coefficient between the ambient component of personal exposures and ambient concentrations is relevant to the panel epidemiological study design. In Table 2-9, most longitudinal studies reported the association between personal total exposures and ambient concentrations for each subject; for some subjects the associations were strong and for some subjects the associations were weak. The weak personal and ambient associations do not necessarily mean that ambient concentrations are not a good surrogate for personal exposures, because the weak associations could have resulted from the day-to-day variation in the nonambient component of total personal exposure. The type of correlation analysis can have a substantial effect on the value of the resultant correlation coefficient.

Mage (1999) showed that very low correlations between personal exposure and ambient concentrations could be obtained when people with very different nonambient exposures are pooled, even though their individual longitudinal correlations are high.

$$r_{ax_i} = \frac{\sum_j (x_{ij} - \bar{x}_i)(a_j - \bar{a})}{(n-1)s_{x_i}s_a}$$

¹ where “r” is the longitudinal correlation coefficient between personal exposure and ambient concentration, “a” represents the ambient concentration, “x” represents exposure, “i” represents the ith subject, “j” represents the jth measurement (with the averaging time ranging from two days to two weeks for SO₂ measurement), “s” represents the standard deviation, and “n” in the longitudinal studies is the number of measurements for each subject. The ambient concentration a_j could be measured by one ambient monitor or the average of several ambient monitors.

Pooled correlations¹ are calculated when a study involves one or only a few measurements per subject and when different subjects are studied on subsequent days. Pooled correlations combine individual-subject/individual-day data for the calculation of correlations. Pooled correlations describe the relationship between daily personal NO₂ exposure and daily ambient SO₂ concentration across all subjects in the study (e.g., Brauer et al., 1989; Sarnat et al., 2006).

Daily-average correlations² are calculated by averaging exposure across subjects for each day. Daily-average correlations then describe the relationship between the daily average exposure and daily ambient pollutant concentration. This type of correlation (i.e., the association between community average exposures (ambient component) and ambient concentrations) is more directly relevant to community time-series and long-term cohort epidemiologic studies, in which ambient concentrations are used as a surrogate for community average exposure to pollutants of ambient origin. However, exposure of the population to SO₂ of ambient origin has not been reported in any of the studies examined.

Not only does the exposure study design determine the meaning of the correlation coefficients in the context of exposure assessment in epidemiologic studies, but it also affects the strength of the association between personal exposures and ambient concentrations. The strength of the association between personal exposures and ambient and/or outdoor concentrations for a population is determined by variations in several physical factors: indoor or other local sources, air exchange rate, penetration, and decay rate of the pollutant in different microenvironments and the time people spend in different microenvironments with different pollutant concentrations. For different types of correlation coefficients, the components of the variance of these physical factors are different, and therefore the strength of different types of correlation coefficients is different. Longitudinal correlation coefficients reflect the *inter-personal* variations of these physical factors; pooled correlation coefficient reflect both *inter-* and *intra-* personal variations of these physical factors; and for the association between community average exposures and ambient concentrations, *inter-personal* variations of these physical factors are reduced by

$$r_{ax} = \frac{\sum_{i,j} (x_{ij} - \bar{x})(a_j - \bar{a})}{(n-1)s_x s_a} \quad \text{where "n" is the number of paired measurements of exposure and ambient concentration, and all other symbols are defined the same way as those in the longitudinal correlation coefficient.}$$

$$r_{a\bar{x}} = \frac{\sum_j (\bar{x}_j - \bar{x})(a_j - \bar{a})}{(n-1)s_{\bar{x}} s_a} \quad \text{, where n is the number of measurement period, during each of which the exposure for all subjects are measured, and all other symbols are defined the same way as those in the longitudinal correlation coefficient.}$$

1 averaging personal exposures across a community. Therefore, the strength of the associations
2 between personal exposures and ambient concentrations may not be comparable directly,
3 although these associations are determined by the same set of physical factors (but affected in
4 different ways).

5 Since correlations are standardized quantities that depend on multiple features of the data,
6 in a correlation, not only is the linear “relatedness” (covariance) of the two quantities important,
7 but so is the variability of each, which can be affected by exposure factors in various ways. In the
8 following assessments, the effects of these physical factors on the strength of correlation
9 coefficients are primarily examined *within* a study, and the purpose of the inter-study comparison
10 is to examine the consistency of the effects across different types of studies.

11 The strength of the associations between personal exposures and ambient concentrations
12 could also be affected by the quality of the data collected during the exposure studies. There are
13 at least six aspects associated with the quality of the data: method precision, method accuracy
14 (compared with FRM), percent of data above method detection limits (based on field blanks),
15 completeness of the data collection, sample size, and soundness of the quality assurance/quality
16 control procedures. Unfortunately, not all studies reported the SIX aspects of the data quality
17 issue. The fraction of data below the detection limit might be a concern for some studies (see
18 Sarnat et al., 2000; 2001; 2005). Correlation coefficients would be biased low if data used in
19 their calculation are below detection limits. Sampling interferences associated with both ambient
20 (see Section 2.3) and personal sampling (see Section 2.5.2) could also affect data quality.
21 Therefore, caution must be exercised when interpreting the results in Table 2-9 and Table 2-10.
22 Sarnat et al. (2001; 2005; 2006) examined the associations between ambient SO₂ concentrations
23 and ambient or personal co-pollutant concentrations. Sarnat et al. (2001) reported that during the
24 winter of 1999, ambient SO₂ was significantly associated (at 5% significance level) with personal
25 exposure to fine particulate matter (PM_{2.5}) (slope = − 0.24), personal exposure to SO₄^{2−}
26 (slope = − 0.03), and personal exposure to PM_{2.5} of ambient origin (slope = − 0.16). However, it
27 should be noted that all the slopes are negative perhaps as the result of measurement error. Sarnat
28 et al. (2005) reported that significant associations between ambient SO₂ and either personal
29 exposures or ambient concentrations of other pollutants were found for personal SO₄^{2−} (winter,
30 slope = 0.06), personal SO₄^{2−} (summer, slope = 0.39), personal PM_{2.5} (summer, slope = 1.68),
31 ambient SO₄^{2−} (winter, slope = 0.19), and ambient PM_{2.5} (winter, slope = 0.80). In Sarnat et al.

(2006), ambient SO₂ was observed to be significantly associated with ambient PM_{2.5}, ambient SO₄²⁻ and ambient EC during the fall ($R^2 = 0.22, 0.33, \text{ and } 0.34$ respectively), and was significantly associated with personal PM_{2.5} during the summer, personal SO₄²⁻ and personal EC during the fall ($R^2 = 0.07, 0.06, \text{ and } 0.05$ respectively).

Of significant concern is the ability of currently available techniques for monitoring either personal exposures or ambient concentrations to measure SO₂ concentrations that are typically found in most urban environments. In some studies, most data, especially data for monitoring personal exposure and indoor concentrations, might be beneath detection limits. Indeed, in one study (Chang et al., 2000), the investigators had to discard data for SO₂, because the values were mostly beneath detection limits. In the study of Kindzierski and Ranganathan (2006), all indoor concentration data were beneath detection limits. In Sarnat et al. (2000), ~70% of personal measurements were beneath detection limits, and ~33% of personal measurements returned apparent negative concentration values. In such situations, associations between ambient concentrations and personal exposure are inadequately characterized. When personal exposure concentrations are above detection limits, a reasonably strong association is observed between personal exposures and ambient concentrations.

2.5.4. Exposure Measurement Errors in Epidemiological Studies

For the purposes of this draft, the effects of exposure error on epidemiological study results refers to changes in the health effects estimate expressed as the relative risk factor, β , and in the related standard error that results from using the ambient concentration of an air pollutant as an exposure indicator rather than using the actual personal exposure in the epidemiological statistical analysis. There are many assumptions made in going from the available measurement of a pollution indicator to an estimate of the personal exposure. The importance of these assumptions and their effect on β depend on the type of epidemiological study.

The considerations of exposure error for SO₂ are simplified compared to those for NO₂ and PM. The only experimental measure available is the ambient concentration of SO₂. In addition, indoor and other non-ambient sources of SO₂ are not thought to be important in population studies, lessening concerns about the possible influence of exposures other than to ambient SO₂.

2.5.4.1. Community Time-Series Studies

This section applies primarily to studies on the association of daily average SO₂ concentrations with daily measures of mortality or morbidity in a community. The following three exposure issues are of primary concern with respect to SO₂ time-series epidemiological analysis: (1) the relationship of the measured concentration of SO₂ to the true concentration; (2) the relationship of day-to-day variations in the concentrations of SO₂, as measured at a central monitoring site, with the corresponding variations in the average concentration of SO₂ over the geographic area from which the health measurements are drawn; and (3) the relationship of the community average concentration of SO₂ to the average personal exposure to ambient SO₂. These three issues are described below.

2.5.4.1.1. Relationship of Measured SO₂ to the True Concentration

Since there is always a random component to instrumental measurement error, the correlation of the measured SO₂ with the true SO₂, on either a 24-h or 1-h basis, will be less than 1. Sheppard et al. (2005) indicate that instrument error in the individual or daily average concentrations have “the effect of attenuating the estimate of α .” Zeger et al. (2000) suggest that instrument error has both Berkson and non-Berkson error components; however, the authors state that the “instrument error in the ambient levels is close to the Berkson type,” and in order for this error to cause substantial bias in β , the error term (the difference between the true concentrations and the measured concentrations) must be strongly correlated with the measured concentrations. Zeger et al. (2000) suggest that, “further investigations of this correlation in cities with many monitors are warranted.” Averaging across multiple unbiased ambient monitors in a region should reduce the instrument measurement error (Sheppard et al., 2005; Wilson and Brauer, 2006; Zeger et al., 2000). There are concerns about the precision and accuracy of the ambient concentration measurements, because SO₂ concentrations are much lower now than when the SO₂ standards were first promulgated. Typical ambient concentrations of SO₂ in the contiguous United States are nearly all at or beneath the detection limit of the monitors currently used in the regulatory network. Thus, greater relative error is most often observed at the lower ambient concentrations compared to the less frequent higher concentration exposures, as might occur because of plume downwash near local point sources or entrainment of plumes downwind from

1 large power plants or smelters. It is unclear how uncertainties in the true concentrations of SO₂,
2 i.e., instrument measurement error, will change β .

2.5.4.1.2. Relationship of Day-to-day Variations in the Ambient Concentration of SO₂ to Variations in the Community Average

3 There has been little analysis of the spatial variation of SO₂ across communities. SO₂
4 emissions arise mainly from coal fired power plants (see Annex Table B-4). Newer power plants
5 and smelters in the United States are no longer located within urban areas. However, some older
6 power plants and industrial facilities are located in many urban areas, especially in the Midwest
7 and Northeast. Downwash from the plumes emitted from these facilities can contribute to
8 elevated levels of SO₂ at the surface in these cities. However, it is anticipated that SO₂ will
9 behave largely as a regional pollutant in most areas. Site-to-site correlations of SO₂
10 concentrations, as shown for several cities in Table 2-3 vary from very low to very high values.
11 This suggests the concentration of SO₂, measured at any given monitoring site, may not be highly
12 correlated with the average community concentration in some areas. There are a number of
13 possible reasons for these findings: local sources that cause the SO₂ to be unevenly distributed
14 spatially; a monitoring site being chosen to represent a nearby source; terrain features that divide
15 the community into several sub-communities that differ in the temporal pattern of pollution; and
16 errors in the measurement of the low concentrations of SO₂ present at most sites. To the extent
17 that the correlation of the ambient concentration with the community average concentration is
18 < 1 , β will be reduced. Similarly, β will be reduced if there are subareas of the community where
19 the correlation of the subarea average concentrations with the concentrations measured at the
20 ambient monitoring site is < 1 . If concentrations in an area of a community impacted by plumes
21 from local SO₂ sources might be higher than, and not well-correlated with, the concentrations at
22 the ambient monitor, and if such high concentrations affected a sizable portion of the population
23 affected by a local source, that community might not be suitable for time-series epidemiological
24 analyses. On the other hand, if the plume impacts the ambient monitor, the high concentration of
25 SO₂ not accompanied by a corresponding high effect in the entire community will bias β toward
26 the null.

2.5.4.1.3. Relationship of Community Average Concentration of SO₂ to Average Personal Exposure to Ambient SO₂

People spend much of their time indoors and, in the absence of indoor sources, indoor concentrations are lower than outdoor concentrations. This is very likely the case with SO₂, since the only known significant indoor source of SO₂ in the United States is the use of kerosene heaters, not thought to be widespread enough to influence population studies. Differences in infiltration factors among homes can also result in differences among individuals' personal exposures. It is necessary to consider how this difference between the ambient concentration, which is used in epidemiological analyses, and the personal ambient source exposure concentration (which includes exposure to the full outdoor concentration while outdoors, and exposure of only a fraction of the outdoor concentrations while indoors) will affect the calculated β . The contribution of the ambient concentration of SO₂ to the personal exposure to ambient SO₂ is given by $Ea = \alpha \cdot Ca$ where Ea is exposure to ambient SO₂, α is the ambient exposure factor with values between 0 and 1, and Ca is the ambient SO₂ concentration as measured at a community monitoring site. Zeger et al. (2000) noted that for community time-series epidemiology, which analyzes the association between health effects and potential causal factors at the community scale rather than the individual scale, it is the correlation of the daily average ambient concentrations with the daily *community average* personal exposures that is important, not the correlation between the daily average ambient concentrations and *the individual* personal exposures. Thus, as mentioned in Section 2.5.3, the low correlation between daily average ambient concentrations and individual personal exposures, as frequently found in pooled panel exposure studies, is not relevant to community time-series epidemiological analysis. Unfortunately, no studies provide adequate information about the community average personal exposure to SO₂.

There has also been concern with the variation of α . Zeger et al. (2000) suggested (for PM) that variations in the individual daily values of α would be a Berkson error and would not change the point estimate of β . Sheppard et al. (2005) used simulations to confirm this for nonreactive pollutants. However, such variations increase the standard error. Day-to-day variations in the population average fraction of ambient exposure will not change the point estimate of β unless the population average fraction of ambient exposure is correlated with seasonal trends in ambient concentration, according to Sheppard et al.

Both Zeger et al. (2000) and Sheppard et al. (2005) show that if β_A is the health effect parameter that would be obtained with a time-series analysis using the ambient exposure and β_C is the health effect parameter that would be obtained with a time-series analysis using the ambient concentration, then $\beta_C = \alpha \cdot \beta_A$. Thus, time-series studies yield different parameters depending on whether they use concentration or exposure. However, the two parameters are related by α . Overestimation of exposure by substitution of the ambient concentration for the ambient exposure leads to underestimation of the effect estimate, or generally bias toward the null.

2.5.4.2. Short-Term Panel Studies

Panel epidemiology refers to time series studies that follow a relatively smaller number of subjects for a relatively short time. Each subject must be considered individually. Panel studies typically examine the association between symptoms or health outcomes and either ambient concentrations or personal exposures. Personal exposures to SO_2 are not measured; rather, ambient concentrations are used in panel studies. Similar types of exposure error as discussed for community time series apply to panel studies.

The ambient exposure factor (α) may differ for each person and each day leading to error in the exposure estimate. If a panel is composed of subjects who live in similar housing and have similar activity pattern, and the study is limited to a single season, the variation in α over time and individual subjects may be small. However, if the panels are composed of more diverse subjects or extend or more than one season, values of α may be quite variable. Such variability will cause error in the estimate of exposure for each subject.

2.5.4.3. Long-Term Cohort Studies

For long-term exposure epidemiological studies, concentrations are integrated over time periods of a year or more, and usually for spatial areas the size of a city, county, or metropolitan statistical area (MSA), although integration over smaller areas may be feasible. Health effects are then regressed, in a statistical model, against the average concentrations in the series of cities (or other areas). In time-series studies, a constant difference between the measured and the true concentration (instrument offset) will not affect β , nor will variations in the daily average α or

1 the daily average nonambient exposure, unless the variations are correlated with the daily
2 variations in concentrations. However, in long-term exposure epidemiological studies, if
3 instrument measurement errors, long-term average values of α , or long-term averages of
4 nonambient exposure differ for different cities (or other areas used in the analysis), the city-to-
5 city long-term ambient SO₂ concentrations will not be perfectly correlated with the long-term
6 average exposure to either ambient or total SO₂. This lack of correlation would be expected to
7 bias the point estimate of β .

2.5.4.4. Summary of Evaluation of Exposure Measurement Error in Epidemiological Studies

8 Exposure error caused by using ambient concentrations of SO₂ as a surrogate for exposure
9 to ambient SO₂ affect β in different ways, dependent upon the type of epidemiological study. In
10 community time-series and short-term panel epidemiological studies, in general, the nonambient
11 source component of personal exposure and the variation in the ambient exposure factor caused
12 by building ventilation practices and personal behaviors, will not change the estimate of β ; but
13 the spatial variation of SO₂ or the representativeness of the ambient monitor might bias the
14 estimate of β toward null. Therefore, β observed in SO₂ community time-series or panel
15 epidemiological studies would be stronger and less uncertain if exposure errors had been
16 adjusted and/or controlled for. In long-term cohort epidemiological studies, instrument
17 measurement errors, factors that influence exposure to ambient SO₂, or long-term averages of
18 nonambient exposure may differ for different cities, which may bias the estimate of β , but the
19 extent and direction of this bias is unclear.

2.6. Dosimetry of Inhaled Sulfur Oxides

20 This section is intended to present an overview of general concepts related to the dosimetry
21 of SO₂ in the respiratory tract. Dosimetry of SO₂ refers to the measurement or estimation of the
22 amount of SO₂ or its reaction products reaching and persisting at specific respiratory tract and
23 systemic sites after exposure. One of the principal effects of inhaled SO₂ is that it stimulates
24 bronchial epithelial irritant receptors and initiates a reflexive contraction of smooth muscles in
25 the bronchial airway. The compound most directly responsible for health effects may be the
26 inhaled SO₂, or perhaps its chemical reaction products. Complete identification of the causative

agents and their integration into SO₂ dosimetry is a complex issue that has not been thoroughly evaluated. Few studies have investigated SO₂ dosimetry in the interval since the 1982 AQCD and the 1986 Second Addendum.

2.6.1. Gas Deposition

The major factors affecting the transport and fate of gases and aerosols in the respiratory tract are: the morphology of the respiratory tract; the physiochemical properties of the mucous and surfactant layers; respiratory functional parameters such as tidal volume, flow rate, and route of breathing; physicochemical properties of the gas; and the physical processes that govern gas transport. Physicochemical properties of SO₂ relevant to respiratory tract uptake include its solubility and diffusivity in epithelial lining fluid (ELF), as well as its reaction-rate with ELF constituents. Henry's law relates the gas phase and liquid phase interfacial concentrations at equilibrium, and is a function of temperature and pressure. Henry's law shows that the amount of SO₂ in the aqueous phase is directly proportional to the partial pressure or concentration of SO₂ in the gas phase. Although the solubility of most gases in mucus and surfactant is not known, the Henry's law constant is known for many gases in water. The Henry's law constant for SO₂ is 0.048 (mole/liter) air / (mole/liter) water at 37° C and 1 atm; for comparison, the value for O₃ is 6.4 under the same conditions (Kimbell and Miller, 1999). In general, the more soluble a gas is in biological fluids, the more rapid, and proximal its absorption will be in the respiratory tract. When the partial pressure of SO₂ on mucosal surfaces exceeds that of the gas phase, such as during expiration, some desorption of SO₂ from the ELF may be expected.

Because SO₂ is highly soluble in water, it is expected to be almost completely absorbed in the nasal passages of both humans and laboratory animals under resting conditions. The dosimetry of SO₂ can be contrasted with the lower solubility gas, O₃, for which the predicted tissue doses (O₃ flux to liquid-tissue interface) are very low in the trachea and increase to a maximum in the terminal bronchioles or first airway generation in the pulmonary region (see Chapter 4, EPA, 2006c). Similar to O₃, the nasal passages remove SO₂ more efficiently than the oral pathway (Brain, 1970). With exercise, the pattern of SO₂ absorption shifts from the upper airway to the tracheobronchial airway in conjunction with a shift from nasal to oronasal breathing and increased ventilatory rates. Due to its effect on delivery and uptake, mode of breathing is also recognized as an important determinant of the severity of SO₂-induced

1 bronchoconstriction, with the greatest responses occurring during oral breathing followed by
2 oronasal breathing and the smallest responses observed during nasal breathing.

3 Melville (1970) measured the absorption of SO₂ (1.5 to 3.4 ppm) during nasal and oral
4 breathing in 12 healthy volunteers. Total respiratory tract absorption of SO₂ was significantly
5 greater ($p < 0.01$) during nasal than oral breathing (85 versus 70%, respectively) and was
6 independent of the inspired concentration. Respired flows were NR. Andersen et al. (1974)
7 measured the nasal absorption of SO₂ (25 ppm) in 7 volunteers at an average inspired flow of 23
8 L/min (i.e., eucapnic hyperpnea [presumably] to simulate light exertion). These investigators
9 reported that the oropharyngeal SO₂ concentration was below their limit of detection (0.25 ppm),
10 implying that at least 99% of SO₂ was absorbed in the nose of subjects during inspiration.
11 Speizer and Frank (1966) also measured the absorption of SO₂ (16.1 ppm) in 7 healthy subjects
12 at an average ventilation of 8.5 L/min (i.e., at rest). They reported that 14% of the inhaled SO₂
13 was absorbed within the first 2 cm into nose. The concentration of SO₂ reaching the pharynx was
14 below the limit of detection, suggesting that at least 99% was absorbed during inspiration.

15 Frank et al. (1969) and Brain (1970) investigated the oral and nasal absorption of SO₂ in
16 the surgically isolated upper respiratory tract of anesthetized dogs. Radiolabeled SO₂ (³⁵SO₂) at
17 the concentrations of 1, 10, and 50 ppm was passed separately through the nose and mouth at the
18 steady flows of 3.5 and 35 L/min for 5 min. The nasal absorption of SO₂ (1 ppm) was 99.9% at
19 3.5 L/min and 96.8% at 35 L/min. The oral absorption of SO₂ (1 ppm) was 99.56% at 3.5 L/min,
20 but only 34% at 35 L/min. The nasal absorption of SO₂ at 3.5 L/min increased with concentration
21 at 1, 10, and 50 ppm and was reported to be 99.9, 99.99, and 99.999%, respectively. This
22 increase in absorption with concentration was hypothesized to be due to increased mucous
23 secretion and increased nasal resistance at the higher SO₂ concentrations. The increased mucus
24 was thought to provide a larger reservoir for SO₂ uptake. The increased nasal resistance may
25 increase turbulence in the airflow and, thereby, decrease the boundary layer between the gas and
26 liquid phases. Dissimilar to the nose, SO₂ absorption in the mouth decreased from 99.56 to
27 96.3% when the concentration was increased from 1 to 10 ppm at 3.5 L/min. Frank et al. (1969)
28 noted that the aperture of the mouth may vary considerably, and that this variation may affect
29 SO₂ uptake in the mouth. Although SO₂ absorption was dependent on inhaled concentration, the
30 rate and route of flow had a greater effect on the magnitude of SO₂ absorption in the upper
31 airway.

1 Strandberg (1964) studied the uptake of SO₂ in the respiratory tract of rabbits. A tracheal
2 cannula with two outlets was utilized to allow sampling of inspired and expired air, and SO₂
3 absorption was observed to depend on inhaled concentration. The absorption during maximal
4 inspiration was 95% at high concentrations (100 to 700 ppm), reflecting an increased SO₂
5 removal in the extrathoracic (ET) airway, whereas it was only 40% at low concentrations (0.05 to
6 0.1 ppm). On expiration, the total SO₂ absorbed (i.e., inspiratory removal in the ET airway plus
7 removal in the lower airway) was 98% at high concentrations and only 80% at the lower
8 concentrations.

9 Amdur (1966) examined changes in airway resistance in guinea pigs due to SO₂ exposure.
10 Guinea pigs were exposed for 1-h to 0.1- to 800 ppm SO₂ during natural unencumbered breathing
11 or to 0.4 to 100 ppm while breathing through a tracheal cannula. At concentrations of 0.4- to 0.5
12 ppm SO₂, route of administration did not affect the airway resistance response, whereas at
13 concentrations of > 2 ppm, the responses were greater in animals exposed by tracheal cannula.
14 Based on the concentration-dependent absorption of SO₂ in the ET airway observed by
15 Strandberg (1964), Amdur (1966) concluded that the airway resistance responses at low-
16 exposure concentrations were independent of method of administration, because the lung
17 received nearly the same concentration with or without the cannula as evidenced by minimal ET
18 absorption.

19 More recently, Ben-Jebria et al. (1990) investigated the absorption of SO₂ in excised
20 porcine tracheae. Absorption was monitored over a 30-min period following the introduction of
21 SO₂ (0.1 to 0.6 ppm, inlet concentration) at a constant flow (2.7 to 11 L/min). The data were
22 analyzed using diffusion-reactor theory. An overall mass transfer coefficient (KSO₂) was
23 determined and separated into its contributions due to gas (convection and diffusion) and tissue
24 phase (diffusivity, solubility, and reaction rates) resistances. SO₂ in the liquid phase was assumed
25 to form HSO₃⁻ rapidly, in proportion with the gas phase SO₂ concentration, HSO₃⁻ then diffused
26 down the concentration gradient into the tissues where it reacted irreversibly with biochemical
27 substrates. Initially, KSO₂ was limited only by gas phase resistance, but decreased exponentially
28 over the first 5 to 10 min of SO₂ exposure to a smaller steady-state value because of tissue
29 resistance to SO₂ absorption. The initial and steady-state KSO₂ values were found to be
30 independent of inlet SO₂ concentration, i.e., for a given flow, the fractional absorption of SO₂ did
31 not depend on SO₂ concentration. An increased KSO₂ (initial and steady-state) was observed with

1 an increasing flow that was thought to be due to a decrease in the boundary layer near the walls
2 of the trachea for radial SO₂ transport. This is in agreement with Aharonson et al. (1974), who
3 also reported that the transfer rate coefficient for SO₂ increases with increasing flow. However,
4 the initial molar flux of SO₂ across the gas-tissue interface appears to increase purely as a
5 function of the increase in mass transport occurring with increasing flow (see Figure 5 in Ben-
6 Jebria et al., 1990). Given that the steady-state KSO₂ remained stable during the 10 to 30 min of
7 exposure and that no SO₂ leakage through the tissue was identified, the authors concluded that
8 there was an irreversible sink for SO₂ within the tissue.

9 Mathematical modeling specific to the regional respiratory uptake of SO₂ is unavailable for
10 humans and laboratory animals. More generally, the influence of age on gas dosimetry in humans
11 during light activity (on average) was examined by Ginsberg et al. (2005) using the U.S. EPA
12 reference concentration methodology (EPA, 1994a). For a highly soluble gas, such as SO₂, they
13 predicted that the majority of gas uptake would occur in the extrathoracic airway and that uptake
14 in these airways would be modestly greater in a 3-month-old infant than an adult. The rate of gas
15 uptake per surface, however, in the extrathoracic airway and large bronchial airway was not
16 markedly different between infants and adults. The smaller bronchial airway of adults were
17 predicted to receive a greater dose (i.e., uptake per unit time and surface area) relative to infants,
18 although the majority of the inhaled SO₂ would be removed proximal to these airways.

19 In summary, inhaled SO₂ is readily absorbed in the upper airway of both humans and
20 laboratory animals. During nasal breathing, the majority of available data suggests 95% or
21 greater SO₂ absorption occurs in the nasal passages, even under ventilation levels comparable to
22 exercise. Somewhat less SO₂ is absorbed in the oral passage than in the nasal passages. The
23 difference in SO₂ absorption between the mouth and the nose is highly dependent on respired
24 flow rates. With an increase in flow from 3.5 to 35 L/min, nasal absorption is relatively
25 unaffected, whereas, oral absorption is reduced from 100 to 34%. Thus, the rate and route of
26 breathing have a great effect on the magnitude of SO₂ absorption in the upper airway and so the
27 penetration of SO₂ to the lower airway. Overall, the available data clearly show that the pattern of
28 SO₂ absorption which shifts from the upper airway to the tracheobronchial airway in conjunction
29 a shift from nasal to oronasal breathing and associated increased ventilatory rates in exercising
30 humans. Mode of breathing is also recognized as an important determinant of the severity of

1 SO₂-induced bronchoconstriction, with the greatest responses occurring during oral breathing
2 followed by oronasal breathing and the smallest responses observed during nasal breathing.

2.6.2. Particles and Sulfur Oxide Mixtures

3 As already discussed, inhaled SO₂ is readily absorbed in the upper airway, particularly
4 during nasal breathing. It has been suggested that sulfur oxides may become absorbed to
5 particles and subsequently transported to more distal lung regions. Depending on atmospheric
6 conditions, SO₂ can be transformed to secondary sulfate particles and acid aerosols (H₂SO₄) and
7 can adsorb onto particulate matter. Jakab et al. (1996) observed that the conversion of SO₂ to
8 SO₄²⁻ on the surface of carbon black aerosols was dependent on high relative humidity (85%)
9 and SO₂ concentration. These investigators suggested that fine carbon black particles can be an
10 effective vector for delivery of SO₄²⁻ to the peripheral lung. Other studies investigating the
11 effects of SO₂ coated aerosols are briefly discussed in Section 3.1.5.

12 Sulfate aerosols are hygroscopic and grow in the respiratory tract. The implications of
13 hygroscopic growth on deposition have been reviewed extensively by Morrow (1986) and Hiller
14 (1991). In general, compared to nonhygroscopic particles of the same initial size, the deposition
15 of hygroscopic aerosols in different regions of the lung may be higher or lower, depending on the
16 initial size. For particles with initial sizes larger than 0.5 μm (aerodynamic diameter), the
17 influence of hygroscopicity would be to increase total deposition with a shift in regional
18 deposition from the distal to larger proximal airway; for smaller particles deposition would tend
19 to be decreased. A thorough review of respiratory deposition and clearance of particulate matter
20 is available elsewhere (EPA, 2004; 2006b). The intent herein was to briefly mention some issues
21 specific to sulfur oxides.

2.6.3. Distribution and Elimination of Sulfur Oxides

22 When SO₂ contacts the fluids lining the airway, it dissolves into the aqueous fluid and
23 forms hydrogen (H⁺) ions and bisulfite (HSO₃⁻) and sulfite (SO₃²⁻) anions (Bascom et al., 1996).
24 The majority of anions are expected to be present as HSO₃⁻ at a concentration proportional to the
25 gas phase concentration of SO₂ (Ben-Jebria et al., 1990). Because of the chemical reactivity of
26 these anions, various reactions are possible, leading to the oxidation of SO₃²⁻ to SO₄²⁻ (see
27 Section 12.2.1, EPA, 1982). Clearance of SO₃²⁻ from the respiratory tract may involve several

1 intermediate chemical reactions and transformations. Gunnison and Benton (1971) identified *S*-
2 sulfonate in blood as a reaction product of inhaled SO₂. Following inhalation of SO₂, the
3 clearance half-time of 4.1 days for *S*-sulfonate in rabbits has been reported (Gunnison and
4 Palmes, 1973).

5 Some SO₂ is also removed by desorption of from the respiratory tract. Desorption is
6 expected when the partial pressure of SO₂ in airway lining fluids exceeds that of the air. Speizer
7 and Frank (1966) found that on expiration, 12% of the SO₂ absorbed during inspiration was
8 desorbed into the expired air. During the first 15 min after the 25- to 30-min SO₂ exposure,
9 another 3% was desorbed. In total, 15% of the amount originally inspired and absorbed SO₂ was
10 desorbed from the nasal mucosa. Frank et al. (1969) reported that up to 18% of the SO₂ was
11 desorbed within ~10 min after exposure.

Chapter 3. Integrated Health Effects

This integrated discussion is structured to provide a coherent framework for the assessment of health risks associated with human exposure to ambient SO₂ in the United States. The main goals of this chapter are: (1) to integrate newly available epidemiological, human clinical, and animal toxicological evidence with consideration of key findings and conclusions from the 1982 AQCD for Sulfur Oxides and First Addendum (EPA, 1982), 1986 Second Addendum (EPA, 1986c), and 1994 Supplement to the Second Addendum, (EPA, 1994c); and (2) to draw conclusions about the causal nature of SO₂ in relation to a variety of health effects. These causal determinations utilize the framework outlined in Chapter 1.

This chapter is organized to present morbidity and mortality associated with short-term exposures to SO₂, followed by morbidity and mortality associated with long-term exposures. Human clinical studies examining the effect of peak exposures (less than 1-h, generally 5-10 min) of SO₂ on respiratory symptoms and lung function are discussed first. Later sections describe the findings of epidemiological studies that examine the association between short-term (generally 24-h avg) and long-term (generally months to years) ambient SO₂ exposure and health outcomes, such as respiratory symptoms in children and asthmatics, emergency department (ED) visits and hospital admissions for respiratory and cardiovascular diseases, and premature mortality. The human clinical and epidemiological evidence are presented with relevant animal toxicological data, when available.

Considerations in the Interpretation of Health Evidence

Human clinical studies are conducted in a controlled laboratory setting using fixed concentrations of air pollutants under carefully regulated environmental conditions and subject activity levels. Results of human clinical studies provide evidence of potential mechanisms for observed effects and a direct quantitative assessment of the SO₂ exposure-health response relationship among asthmatic individuals. Observed effects in these studies may underestimate the response in certain sensitive subpopulations for a number of reasons. First, study subjects must either be healthy, or have a level of illness which does not preclude them from participating in the study. Second, asthmatics who are unable to withhold the use of bronchodilators for at

1 least 6 hours prior to exposure and subjects with a recent history of upper respiratory tract
2 infections are typically excluded from clinical studies of exposure to SO₂. While human clinical
3 studies provide important information on the biological plausibility of associations observed
4 between SO₂ exposure and health outcomes in epidemiological studies, the concentration-
5 response relationships cannot necessarily be directly extrapolated to concentrations below those
6 administered in the laboratory. Further, human clinical studies are normally conducted on a
7 relatively small number of subjects, which reduces the power of the study to detect significant
8 differences in the health outcomes of interest between exposure to varying concentrations of SO₂
9 and clean air.

10 Epidemiological studies provide important information on the associations between health
11 effects and exposure of human populations to ambient levels of SO₂. These studies also help to
12 identify susceptible subgroups and associated risk factors. However, associations observed
13 between specific air pollutants and health outcomes in epidemiological studies may be
14 confounded by copollutants or meteorological conditions, and influenced by model
15 specifications in the analytical methods. Extensive discussion of these issues is provided in the
16 2004 AQCD for PM (EPA, 2004) and the 2006 AQCD for O₃ and Related Photochemical
17 Oxidants (EPA, 2006c), and therefore presented only briefly below.

18 The use of multipollutant regression models has been the prevailing approach for
19 controlling potential confounding by copollutants in air pollution health effects studies. Finding
20 the likely causal pollutant from multipollutant regression models is made difficult by the
21 possibility that one or more air pollutants may be acting as a surrogate for an unmeasured or
22 poorly-measured pollutant or for a particular mixture of pollutants. SO₂ presents an especially
23 interesting test of multipollutant effects models: correlations with sulfate, the principal
24 atmospheric oxidation product of SO₂, show temporal and spatial discrepancies that can
25 influence exposures and health effects. Short-term, mostly time-series epidemiological studies
26 generally use intracity ambient concentration data which show very little or no correlation
27 between emitted SO₂ and transformed sulfate. In contrast, long-term epidemiological studies
28 using intercity data can show correlations between SO₂ and sulfate on the order of 0.8 or higher.
29 In these studies the fine-scale spatiotemporal variations in the intracity data are significantly
30 reduced, since sulfate has sufficient time for production from SO₂, dispersed over a wide spatial
31 area, and mixed down to ground level. Layered over these spatial and fine-scale temporal

1 differences are seasonal and regional dissimilarities driven by cities' various SO₂ emissions
2 profiles and differing available time and sunlight conditions for oxidation. Thus, attempts to
3 distinguish gaseous and particle effects related to SO₂ using multipollutant epidemiological
4 models must be interpreted with caution. Despite these limitations, the use of multipollutant
5 models is still the prevailing approach employed in most studies of SO₂ and health effects, and
6 may provide some insight into the potential for confounding or interaction among pollutants.

7 Model specification and model selection also need to be considered in the interpretation of
8 the epidemiological evidence. The studies presented in this chapter investigated the association
9 between various measures of SO₂ (e.g., multiple lags and different exposure metrics) and various
10 health outcomes using different model specifications. The summary of health effects in this
11 chapter is vulnerable to the errors of publication bias and multiple testing. Efforts have been
12 made to reduce the impact of multiple testing errors. For example, although many studies
13 examined multiple single-day lag models, priority was given to effects observed at 0- or 1-day
14 lags, rather than at longer lags. Additional focus was placed on results from distributed and
15 moving average lags as they are able to take into consideration multiday effects. Both single- and
16 multiple-pollutant models were considered and examined for robustness of results. Additional
17 analyses of multiple model specifications for adjustment of temporal or meteorological trends are
18 considered to be sensitivity analyses.

19 In addition to issues related to confounding by copollutants and model selection, the
20 evaluation of the epidemiological evidence also considers study population and sample size, with
21 particular emphasis placed on multicity studies. Other factors considered are study location
22 (North America versus other regions), meaningfulness and reliability of the health endpoint
23 measurements, and appropriateness of the statistical analyses methods used. These
24 considerations in the interpretation of the epidemiological evidence lead to emphasis of certain
25 studies in the chapter text, tables, and figures.

26 Animal toxicological studies may provide further evidence for the potential mechanism of
27 an observed effect; however, most of these studies have been conducted at concentrations vastly
28 exceeding current ambient conditions. In discussing the mechanisms of SO_x toxicity, studies
29 conducted under atmospherically relevant conditions are emphasized whenever possible; studies
30 at higher levels are also considered, due to species-to-species differences and potential
31 differences in sensitivity between study subjects and especially susceptible human populations.

1 This chapter focuses on important new scientific studies, with emphasis on those
2 conducted at or near current ambient concentrations. Given their respective strengths and
3 limitations, evidence from human clinical, epidemiological and animal toxicological studies are
4 considered in order to evaluate the causality of SO_x–health effects associations. The annexes
5 supplement the information included here by presenting a more details of the literature.

3.1. Respiratory Morbidity Associated with Short-Term Exposure

3.1.1. Summary of Findings from the Previous Review

6 The majority of the SO₂ human clinical studies in the 1982 AQCD for Sulfur Oxides
7 evaluated respiratory effects of SO₂ exposure in healthy adults, with some limited data from
8 clinical studies of adults with asthma. SO₂-related respiratory effects such as increased airway
9 resistance and decreased forced expiratory volume in 1 s (FEV₁) were observed in healthy
10 individuals at concentrations > 1.0-5.0 ppm, and in asthmatics at concentrations < 1.0 ppm. The
11 1986 Second Addendum (EPA) and 1994 Supplement to the Second Addendum (EPA) reviewed
12 several additional controlled studies involving both healthy and asthmatic individuals. In general,
13 these studies found no pulmonary effects of SO₂ exposure in healthy subjects exposed to
14 concentrations < 1.0 ppm (Bedi et al., 1984; Folinsbee et al., 1985; Kulle et al., 1984; Stacy et
15 al., 1983). However, in exposures of asthmatic adults, respiratory effects were observed
16 following short-term exposures (5-10 min) to levels < 1.0 ppm (Balmes et al., 1987; Horstman et
17 al., 1988; Linn et al., 1987).

18 Only a few epidemiological studies reviewed in the 1982 AQCD were useful in
19 determining the concentration-response relationship of respiratory health effects from short-term
20 exposure to SO₂. The most notable study was by Lawther (1970), which examined the
21 association between air pollution and worsening health status in bronchitic patients residing in
22 London, England. It was concluded in the 1982 AQCD that worsening of health status among
23 chronic bronchitic patients was associated with daily black smoke (BS) levels of 250-500 µg/m³
24 in the presence of SO₂ levels in the range of 191-229 ppb. In the 1986 Second Addendum,
25 additional studies investigated morbidity associated with short-term exposure to SO₂. The most
26 relevant study was by Dockery (1982), which examined pulmonary function in school children in

1 Steubenville, OH, as part of the Harvard Six Cities Study. This study found that small but
2 statistically significant reversible decrements in forced vital capacity (FVC) and forced
3 expiratory volume in 0.75 s (FEV_{0.75}) were associated with increases in 24-h avg concentrations
4 of total suspended particles (TSP) at levels ranging up to 220-420 µg/m³ and SO₂ at levels
5 ranging up to 107-176 ppb. However, it was impossible to separate the relative contributions of
6 TSP and SO₂, and no threshold level for the observed effects could be discerned from the wide
7 range of exposure levels.

8 Epidemiological evidence for an association between SO₂ and respiratory morbidity, as
9 indicated by increased use of ED facilities or increased hospital admissions for respiratory
10 diseases, was also reported in the 1982 AQCD. Overall, these results suggested increased upper
11 respiratory tract morbidity during episodic marked elevations of PM or SO₂ (400-500 ppb),
12 especially among older adults. The 1982 AQCD further concluded that the studies reviewed
13 provided essentially no evidence for an association between asthma attacks and acute exposures
14 at typical ambient PM or SO₂ levels in the United States (the mean annual average SO₂
15 concentrations from 1972 to 1977 was approximately 6 ppb, with 90th percentile values ranging
16 from 15 to 20 ppb).

17 The 1982 AQCD for Sulfur Oxides (EPA, 1982) reported numerous effects on the
18 respiratory system in animals exposed to SO₂. Effects were generally observed at levels
19 exceeding those found in the ambient environment, and included morphological changes, altered
20 pulmonary function, lipid peroxidation, and changes in host lung defenses. The immediate effect
21 of acute SO₂ exposure in animals was increased pulmonary resistance to airflow, a measure of
22 bronchoconstriction. Bronchoconstriction was reported to be the most sensitive indicator of lung
23 function effects in acute SO₂ exposure.

24 Collectively, the human clinical, epidemiological and animal toxicological, studies
25 provided biological plausibility and coherent evidence of an adverse effect of ambient SO₂ on
26 respiratory health. Since the 1982 AQCD, 1986 Second Addendum, and 1994 Supplement to the
27 Second Addendum, additional studies have been conducted to determine the relationship
28 between short-term exposures to ambient SO₂ and adverse respiratory health effects, including
29 respiratory symptoms, lung function, airway inflammation, airway hyperresponsiveness, lung
30 host defenses, and ED visits and hospitalizations for respiratory causes. The epidemiological,
31 human clinical, and animal toxicological evidence on the effects of SO₂ on these various

endpoints are discussed below. The finding of the previous review are integrated below with the current literature.

3.1.2. Potential Mode of Action for Respiratory Health Effects

The 1982 AQCD (EPA, 1982) gave background information on the biochemistry of SO₂, chemical reactions of bisulfite (HSO₃⁻), metabolism of SO₂, and the activating or inhibiting effects of bisulfite on various enzymes. SO₂ readily dissolves in water, rapidly becoming hydrated to form sulfurous acid, which at physiological pH substantially dissociates to form bisulfite and sulfite (SO₃²⁻) ions. Studies in vitro have shown that SO₂ and/or bisulfite readily react with nucleic acids, proteins, lipids, and other classes of biomolecules. Bisulfite participates in three important types of reactions with biomolecules: sulfonation (sulfitolysis), autooxidation with generation of free radicals, and addition to cytosine. Products of sulfonation reactions have been shown to be long-lived in vivo and may be highly reactive. Products of autooxidation may be responsible for the initiation of lipid peroxidation, which, among other effects, could damage plasma membranes. In contrast, studies have shown that bisulfite can react with nucleic acids to convert cytosine to uracil, thus resulting in mutational events. A principal mechanism of detoxification of SO₂ (and sulfite/bisulfite) occurs through the enzymatic activity of sulfite oxidase, resulting in the production of sulfate. Sulfite oxidase is a molybdenum-containing enzyme, and the 1982 AQCD noted that depleting its activity through a low-molybdenum diet supplemented with the competitive inhibitor tungsten resulted in a significant lowering of the LD₅₀ for intraperitoneally injected bisulfite. It was also noted that while in vitro exposure to SO₂ or sulfite/bisulfite had been shown to either activate or inhibit a variety of enzymes, no such effects had yet been demonstrated for in vivo exposure.

As discussed in the 1982 AQCD, the immediate effect of acute SO₂ exposure in animals was bronchoconstriction. Reactions of SO₂ with respiratory tract fluids can result in the production of bisulfite, sulfite, and a lowering of the pH, which may be involved in the bronchoconstrictive response. It is now widely appreciated that bronchoconstriction following SO₂ exposure is mediated by chemosensitive receptors in the tracheobronchial tree. Rapidly activating receptors (RARs) and sensory C-fiber receptors found at all levels of the respiratory tract are sensitive to irritant gases such as SO₂ (Coleridge and Coleridge, 1994; Widdicombe, 2006). Activation of these vagal afferents causes central nervous system reflexes resulting in

1 bronchoconstriction, mucus secretion, mucosal vasodilation, cough, or apnea, followed by rapid
2 shallow breathing and effects on the cardiovascular system such as bradycardia and hypotension
3 or hypertension (Coleridge and Coleridge, 1994; Widdicombe and Lee, 2001; Widdicombe,
4 2003).

5 Early experiments demonstrated that SO₂-induced reflexes were mediated by cholinergic
6 parasympathetic pathways involving the vagus nerve and inhibited by atropine (Grunstein et al.,
7 1977; Nadel et al., 1965a; 1965b). Bronchoconstriction was found to involve smooth muscle
8 contraction since β -adrenergic agonists such as isoproterenol reversed the effects (Nadel et al.,
9 1965a; 1965b). Acetylcholine and histamine were also thought to be involved in SO₂-induced
10 bronchoconstriction (EPA, 1982).

11 More recent experiments in animal models conducted since 1982 have demonstrated that
12 both cholinergic and noncholinergic mechanisms may be involved in SO₂-induced effects. In two
13 studies utilizing bilateral vagotomy, vagal afferents were found to mediate the immediate
14 ventilatory responses to SO₂ (Wang et al., 1996), but not the prolonged bronchoconstrictor
15 response (Barthelemy et al., 1988). Other studies showed that atropine failed to block SO₂-
16 induced bronchoconstriction, and that a local axon reflex resulting in C-fiber secretion of
17 neuropeptides (i.e., neurogenic inflammation) was responsible for the effect (Atzori et al.,
18 1992a; Hajj et al., 1996). Neurogenic inflammation has been shown to play a key role in animal
19 models of airway inflammatory disease (Groneberg et al., 2004).

20 In humans, the mechanisms responsible for SO₂-induced bronchoconstriction are not fully
21 understood. In non-asthmatics, near complete attenuation of bronchoconstriction has been
22 demonstrated using the anticholinergic agents atropine and ipratropium bromide (Snashall and
23 Baldwin, 1982; Tan et al., 1982; Yildirim et al., 2005). However, in asthmatics, these same
24 anticholinergic agents (Field et al., 1996; Myers et al., 1986a), as well as short- and long-acting
25 β 2-adrenergic agonists (Gong et al., 1996; Linn et al., 1988), theophylline (Koenig et al., 1992),
26 cromolyn sodium (Myers et al., 1986), nedocromil sodium (Bigby and Boushey, 1993) and
27 leukotriene receptor antagonists (Gong et al., 2001; Lazarus et al., 1997) only partially blocked
28 SO₂-induced bronchoconstriction. That none of these therapies have been shown to completely
29 attenuate the effects of SO₂ implies the involvement of both parasympathetic pathways and
30 inflammatory mediators in asthmatics. Strong evidence of this is borne out in a study by Myers
31 et al. (1986), in which asthmatic adults were exposed to SO₂ following pretreatment with

1 cromolyn sodium (a mast cell stabilizer), atropine (a muscarinic receptor antagonist), and the two
2 medications together. While both treatments individually provided some protection against the
3 bronchoconstrictive effects of SO₂, there was a much stronger and statistically significant effect
4 following concurrent administration of the two medications.

5 It has been proposed that inflammation contributes to the enhanced sensitivity to SO₂ seen
6 in asthmatics by altering autonomic responses (Tunnicliffe et al., 2001), enhancing mediator
7 release (Tan et al., 1982) and/or sensitizing C-fibers and RARs (Widdicombe and Lee, 2001).
8 Whether local axon reflexes also play a role in SO₂-induced bronchoconstriction in asthmatics is
9 not known (Widdicombe and Lee, 2001; Widdicombe, 2003; Groneberg et al., 2004). However,
10 differences in respiratory tract innervation between rodents and humans suggest that C-fiber
11 mediated neurogenic inflammation may be unimportant in humans (Groneberg et al., 2004;
12 Widdicombe and Lee, 2001; Widdicombe, 2003).

3.1.3. Respiratory Effects Associated with Peak Exposure

13 SO₂-induced respiratory effects among exercising asthmatics are well-documented, and
14 have been consistently observed following peak exposures (defined here as 5-10 min exposures
15 to relatively higher concentrations, e.g., 0.4-1.0 ppm) (Balmes et al., 1987; Bethel et al., 1985;
16 Horstman et al., 1986; 1988; Linn et al., 1984b; 1987; 1990; Schachter et al., 1984; Sheppard et
17 al., 1981). Similar respiratory effects have been observed in some sensitive asthmatics at
18 concentrations as low as 0.2-0.3 ppm; however, these effects have not reached statistical
19 significance (Horstman et al., 1986; Linn et al., 1987; 1988; 1990). Since the publication of the
20 1994 Supplement, several additional human clinical studies have been published that provide
21 supportive evidence of SO₂-induced decrements in lung function and increases in respiratory
22 symptoms among exercising asthmatics (see Annex Table D-2). Descriptions of older studies are
23 presented in the 1994 Supplement, and will not be described in great detail in this document.
24 However, based on recent guidance from the American Thoracic Society (ATS) regarding what
25 constitutes an adverse health effect of air pollution (ATS, 2000a), some key older studies were
26 reviewed and analyzed along with studies published since 1994. In its official statement, the ATS
27 recommended that transient loss in lung function with accompanying respiratory symptoms
28 attributable to air pollution should be considered adverse. In addition, ATS concluded that a
29 decrease in health-related quality of life, which refers to an individual's perception of well being,

1 should also be considered to represent an adverse effect of air pollution. Therefore, whereas the
2 conclusions in the 1994 Supplement were based on SO₂ exposure concentrations which resulted
3 in large decrements in lung function along with moderate to severe respiratory symptoms, the
4 current review of data from human clinical studies focuses on moderate to large SO₂-induced
5 decrements in lung function combined with respiratory symptoms ranging from mild (perceptible
6 wheeze or chest tightness) to severe (breathing distress requiring the use of a bronchodilator).

3.1.3.1. Respiratory Symptoms

7 The 1994 Supplement to the Second Addendum described in detail several studies that
8 evaluated respiratory symptoms following controlled human exposures to SO₂. Briefly, following
9 5-min exposures to 0, 0.2, 0.4, and 0.6 ppm SO₂ during moderate to heavy levels of exercise (48
10 L/min), Linn et al. (1983) reported that the severity of respiratory symptoms (i.e., cough, chest
11 tightness, throat irritation) among asthmatics increased with increasing SO₂ concentration.
12 Relative to clean air exposures, exposures to SO₂ resulted in statistically significant increases in
13 respiratory symptoms at concentrations of 0.4 and 0.6 ppm. In a subsequent study, Linn et al.
14 (1987) observed a significant effect of SO₂ on respiratory symptoms in asthmatics who were
15 engaged in slightly lower levels of exercise (40 L/min) for a duration of 10 min. Clear increases
16 in respiratory symptoms were observed at concentrations of 0.6 ppm, with 43% of subjects
17 experiencing SO₂-induced symptoms. Some evidence of SO₂-induced increases in respiratory
18 symptoms was also demonstrated at concentrations as low as 0.4 ppm, with 15% of subjects
19 experiencing symptoms (Smith, 1994). It was also observed that these symptoms abated < 1 h
20 after exposure. Balmes et al. reported that 7 out of 8 asthmatic adults developed respiratory
21 symptoms, including wheezing and chest tightness, following 3-min exposures to 0.5 ppm SO₂
22 during eucapnic hyperpnea ($\dot{V}_E = 60$ L/min).

23 Additional human clinical studies published since the 1994 Supplement to the Second
24 Addendum have provided support for previous conclusions regarding the effect of peak
25 exposures to SO₂ on respiratory symptoms. In a human clinical study with SO₂-sensitive
26 asthmatics, Gong et al. (1995) reported that respiratory symptoms (i.e., shortness of breath,
27 wheeze, and chest tightness) increased with increasing SO₂ concentration (0, 0.5, and 1.0 ppm
28 SO₂) following exposures of 10 min with varying levels of exercise. It was also observed that
29 exposure to 0.5 ppm SO₂ during light exercise evoked a more severe symptomatic response than

heavy exercise in clean air. Trenga et al. (1999) observed a significant correlation between decreases in FEV₁ and increases in respiratory symptoms following 10 min exposures to 0.5 ppm SO₂.

3.1.3.2. Lung Function

In controlled exposures of healthy human subjects to SO₂, respiratory effects including increased respiration rates, decrements in peak flow, bronchoconstriction, and increased airway resistance have been observed at concentrations > 1 ppm (Abe, 1967; Amdur et al., 1953; Andersen et al., 1974; Frank et al., 1962; Lawther, 1955; Lawther et al., 1975; Sim and Pattle, 1957; Snell and Luchsinger, 1969). SO₂-induced decrements in lung function can be potentiated by increasing ventilation rate, either through eucapnic hyperpnea or by performing exercise during exposure. This effect is likely due to an increased uptake of SO₂ resulting from both the increase in \dot{V}_E as well as a shift from nasal breathing to oronasal breathing.

It has been clearly established that subjects with asthma are more sensitive to the respiratory effects of SO₂ exposure than healthy individuals without asthma. Asthmatic individuals exposed to SO₂ concentrations as low as 0.4-0.6 ppm for 5-10 min during exercise have been shown to experience moderate or greater bronchoconstriction, measured as an increase in sRaw ($\geq 100\%$) or decrease in FEV₁ ($\geq 15\%$) after correction for exercise-induced responses in clean air (Linn et al., 1983; 1984; 1987; 1988; 1990; Magnussen et al., 1990; Roger et al., 1985). Asthmatic subjects who are most sensitive to the respiratory effects of SO₂ have been observed to experience significant decrements in lung function following exposure to SO₂ at concentrations ≤ 0.3 ppm (Horstman et al., 1986; Sheppard et al., 1981). In some cases, bronchoconstrictive responses to SO₂ can occur in as little as 2 min after the start of exposure (Balmes et al., 1987; Horstman et al., 1988). Gong et al. (1995) demonstrated an exposure-response relationship between SO₂ and lung function by exposing 14 unmedicated, SO₂-sensitive asthmatics to 0, 0.5, and 1 ppm SO₂ under 3 different levels of exercise. It was shown that increasing SO₂ concentration had a greater effect on sRaw and FEV₁ than increasing exercise level. Trenga et al. (1999) observed that 25 out of 47 adult asthmatics experienced a drop in FEV₁ versus baseline of between 8 and 44% (mean = 17.2%) following a 10 min exposure to 0.5 ppm SO₂ during moderate exercise.

1 Since some of the studies involving asthmatic subjects have used change in sRaw as the
2 endpoint of interest while others measured changes in FEV₁ or both, a comparison of FEV₁ and
3 sRaw based on data from Linn et al. (1987; 1990) was provided in the 1994 Supplement to the
4 Second Addendum. Based on simple linear interpolation of the data from these two studies, a
5 100% increase in sRaw corresponded to a 12 to 15% decrease in FEV₁ and a 200% increase in
6 sRaw corresponded to a 25 to 30% decrease in FEV₁.

7 One of the aims of the Linn et al. (1987) study was to determine how the intensity of
8 response varied with asthma severity or status. In this study, 24 normal, 21 atopic (but not
9 asthmatic), 16 mild asthmatic, and 24 moderate/severe asthmatic subjects were exposed to SO₂
10 concentrations between 0 and 0.6 ppm. While the moderate/severe asthmatics were more
11 responsive than mild asthmatics following exposure to clean air during exercise, their increases
12 in response to increasing SO₂ concentrations were similar to those of the mild asthmatic group.
13 Thus, it was concluded that SO₂ response was not strongly dependent on the clinical severity of
14 asthma. However, the apparent lack of correlation between SO₂ response and asthma severity
15 should be interpreted with caution, since the SO₂ response may have been attenuated by
16 medication usage or its persistence. Three of the moderate/severe asthmatics were unable to
17 withhold medication usage during the exposure period. Conversely, a few of the asthmatics,
18 including some in the moderate/severe group, did not react to 0.6 ppm SO₂.

19 One of the key studies discussed in the 1994 Supplement to the Second Addendum was by
20 Horstman et al. (1986). In this study, 27 asthmatic subjects were exposed to concentrations of
21 SO₂ between 0- and 2 ppm SO₂ for 10 min on different days under exercising conditions (\dot{V}_E =
22 42 L/min). The authors reported that for 22% of the subjects, the concentration of SO₂ needed to
23 produce a doubling of sRaw compared to clean air exposure [PC(SO₂)] was < 0.5 ppm, with 2
24 subjects (7.4%) experiencing moderate decrements in lung function following exposure to
25 concentrations of SO₂ at or below 0.3 ppm (see Figure 3-1). For approximately 15% of the
26 subjects, the PC(SO₂) was > 2 ppm, with approximately 35% of asthmatic subjects experiencing
27 a doubling in sRaw versus clean air at ≤ 0.6-ppm SO₂.

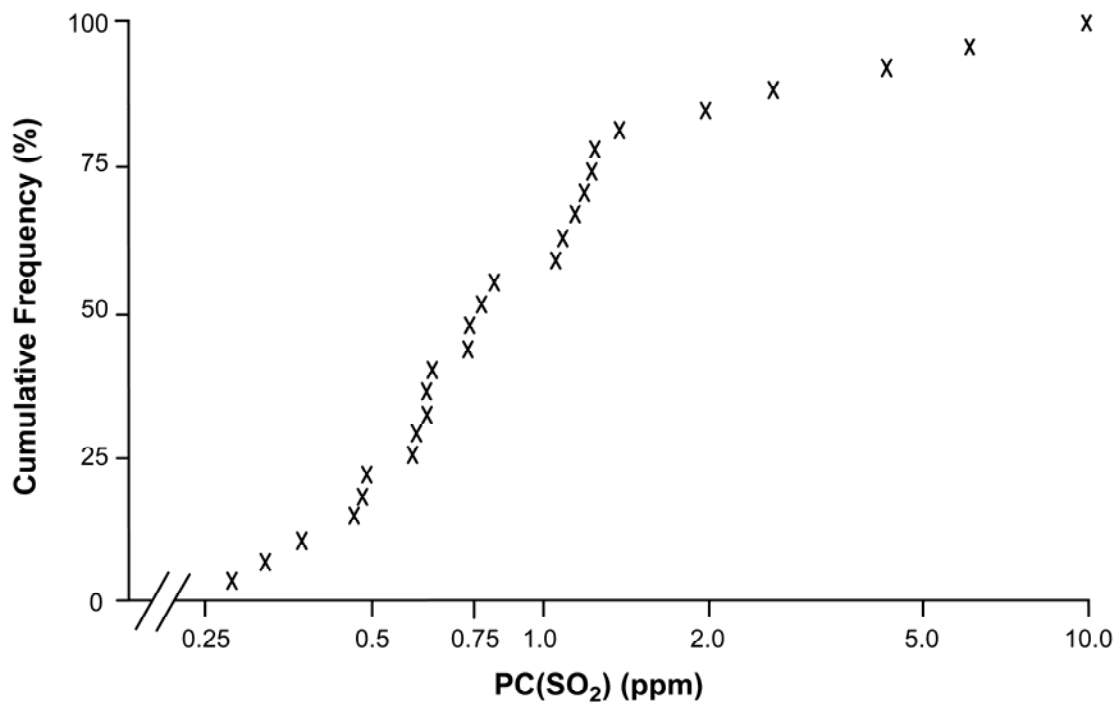


Figure 3-1. Distribution of individual airway sensitivity to SO₂. Each data point represents the value of PC(SO₂) for an individual subject. PC(SO₂) is defined as the concentration of SO₂ which resulted in a doubling of sRaw compared to clean air exposure.

Source: Horstman et al. (1986).

It is important to note that a transient decrement in lung function following exposure to an air pollutant is not automatically considered to represent an adverse effect. However, SO₂-induced decrements in lung function (increased sRaw and decreased FEV₁) have frequently been associated with increases in respiratory symptoms among asthmatics (Balmes et al., 1987; Gong et al., 1995; Linn et al., 1987; 1988; 1990; 1983; Roger et al., 1985), which together does constitute an adverse effect under the ATS guidelines. Linn et al. (1987) exposed 40 mild and moderate asthmatics during 10 min periods of exercise to 0, 0.2, 0.4, and 0.6 ppm SO₂. The effect of SO₂ on lung function and respiratory symptoms was assessed immediately following exposure, and the individual-specific results have been made available to the U.S. EPA by the study authors (Smith, 1994). Following exposure to 0.6 ppm SO₂ and after adjusting for effects of exercise in clean air, 21 of the 40 subjects demonstrated moderate or greater decrements in lung function, defined as a $\geq 15\%$ decrease in FEV₁, a $\geq 100\%$ increase in sRaw, or both. Of these

21 responders, 14 (67%) also experienced mild to severe respiratory symptoms (6 mild, 6 moderate, and 2 severe). In the same study, 14 asthmatics experienced moderate or greater decrements in lung function at 0.4 ppm SO₂, 5 of whom (36%) also experienced mild to moderate respiratory symptoms (2 mild, 3 moderate). Five asthmatics experienced moderate or greater decrements in lung function at the lowest SO₂ concentration tested (0.2 ppm), with 1 of the 5 (20%) also experiencing mild respiratory symptoms.

It has been proposed that, as in asthmatics, individuals with COPD may also be more susceptible to SO₂-induced respiratory health effects. However, this group has not been extensively studied in human clinical studies. Among a group of older adults with physician-diagnosed COPD, Linn et al. (1985) reported no significant effect on lung function following 15 min exposures to SO₂ at concentrations of 0.4 and 0.8 ppm. While it was concluded that older adults with COPD appear to be less sensitive to SO₂ when compared with younger adult asthmatics, the authors suggested that the lack of response may have been due in part to the very low levels of exercise used in the study ($\dot{V}_E = 18$ L/min), which would result in a lower dose of SO₂ reaching the lower airway. In contrast to studies with asthmatics, most of the subjects in this study regularly used bronchodilators and were permitted their use up to 4 h prior to the study.

In summary, SO₂-induced decrements in lung function have been observed following peak exposures in humans. These effects are particularly evident in exercising asthmatic individuals, with significant decreases in sRaw and increases in FEV₁ consistently demonstrated following 5-10 min exposures to 0.4-0.6 ppm SO₂. SO₂-induced decrements in lung function have frequently been associated with respiratory symptoms, and with increasing SO₂ exposure concentration from 0.2-1.0 ppm, both the magnitude of response among asthmatics and the percentage of asthmatics significantly affected have been shown to increase.

3.1.3.3. Airway Inflammation

A very limited number of human clinical studies have investigated the role of airway inflammation in the asthmatic response following peak exposure to SO₂. Gong et al. (2001) observed an SO₂-induced increase in sputum eosinophil counts in exercising asthmatics 2 h after a 10 min exposure to 0.75 ppm SO₂. The results of this study provide some evidence that SO₂ may elicit an allergic inflammatory response in the airways of asthmatics which extends beyond the short time period typically associated with SO₂ effects.

3.1.3.4. Evidence of the Effect of Peak Exposure from Animal Studies

1 In addition to the findings of human clinical studies involving asthmatics, SO₂-induced
2 decrements in lung function have been demonstrated following peak exposures to SO₂ in
3 laboratory animals. The 1982 AQCD reported bronchoconstriction, as indicated by increased
4 pulmonary resistance, as the most sensitive indicator of lung function effects of acute SO₂
5 exposure based on the observations of increased pulmonary resistance in guinea pigs that were
6 acutely exposed to 0.16 ppm SO₂. Since 1982, a few new animal toxicological studies have
7 demonstrated acute changes in lung function following SO₂ exposures of 45 min or less. These
8 studies are summarized below and in Annex Table E-1.

9 Lewis and Kirchner (1984) measured lung function in dogs exposed for 5 min to two doses
10 of SO₂ via an endotracheal tube. Increased pulmonary resistance and decreased compliance were
11 observed in conscious dogs exposed to 30 ppm SO₂, but not to 10 ppm SO₂.

12 All other studies focused on the role of local nervous system reflexes and/or C-fiber
13 receptors in mediating responses to SO₂. Barthelemy et al. (1988) measured lung function in
14 anesthetized rabbits exposed for 45 min by endotracheal tube to two doses of SO₂. Airway
15 resistance increased 16% and 50% following 0.5 and 5 ppm SO₂, respectively. Bivagal vagotomy
16 had little effect on the response to 5 ppm, indicating that the prolonged bronchoconstriction
17 response did not result from a vagal reflex. This study did not rule out the possibility that vagal
18 reflexes were involved in immediate bronchoconstriction following SO₂ exposure.

19 In another study, Atzori et al. (1992a) demonstrated bronchoconstriction, as measured by
20 changes in dynamic lung compliance and airway conductance, within the first 5 min following
21 exposure of isolated and perfused guinea pig lungs to 100 and 250 ppm SO₂ via an endotracheal
22 tube. This response was found to be due to a local nervous system reflex. However, this result
23 does not preclude involvement of central nervous system reflexes in SO₂-induced
24 bronchoconstriction under conditions of an intact vagus nerve. Furthermore, the formation of
25 sulfite was observed in perfusate following SO₂ exposure. Using the same model, Atzori et al.
26 (1992b) found that SO₂-induced bronchoconstriction was associated with the release of a sensory
27 neuropeptide and was inhibited when C-fiber receptors were blocked.

28 Other animal toxicological studies examined immediate respiratory effects from exposure
29 to very high SO₂ concentrations. Hajj et al. (1996) exposed anesthetized guinea pigs to six tidal
30 breaths of 500–2,000 ppm SO₂. Increased total pulmonary resistance, decreased dynamic

1 compliance, and systemic hypotension were observed within seconds. Tachykinin antagonists
2 blocked the changes in lung airway function, but not the changes in blood pressure in this model
3 system. Atropine failed to block the airway response. These results suggest that a local nervous
4 system reflex involving tachykinin release is an important mediator of bronchoconstriction
5 following high concentrations of SO₂. Wang et al. (1996) exposed anesthetized rats to two tidal
6 breaths of 0.5% SO₂ via an endotracheal tube. Immediate and transient bradypnea and
7 bradycardia were observed. Selective block of the C-fiber receptors and bilateral vagotomy
8 eliminated the SO₂-mediated effect on ventilation.

3.1.3.5. Summary of Evidence on the Effect of Peak Exposure on Respiratory Health

9 Collectively, evidence from earlier studies considered in the previous review, along with a
10 limited number of new human clinical studies, consistently indicates that with elevated
11 ventilation rates, asthmatic individuals experience moderate or greater decrements in lung
12 function, as well as increased respiratory symptoms, following peak exposures to SO₂ at
13 concentrations as low as 0.4-0.6 ppm (Balmes et al., 1987; Gong et al., 1995; Horstman et al.,
14 1986; Linn et al., 1987; Linn et al., 1983). These findings are consistent with our understanding
15 of the potential modes of action for respiratory health as described in Section 3.1.2. Some
16 sensitive asthmatics have been shown to experience moderate decrements in lung function at
17 concentrations below 0.3 ppm (Balmes et al., 1987; Linn et al., 1987; Sheppard et al., 1981),
18 although there is limited evidence of a significant increase in respiratory symptoms at these
19 exposure concentrations. Among asthmatics, both the magnitude of SO₂-induced decrements in
20 lung function and the percent of individuals affected have consistently been shown to increase
21 with increasing exposure to SO₂ concentrations between 0.2 and 1.0 ppm. This is summarized in
22 Table 3-1 along with supporting evidence of SO₂-induced increases in respiratory symptoms at
23 various exposure concentrations. The table includes data from all studies where individual data
24 are presented or have been made available by the authors (Smith, 1994). Although the vast
25 majority of human clinical studies involving controlled exposure to SO₂ have been conducted in
26 adult asthmatics, there is a relatively strong body of evidence to suggest that adolescents may
27 experience many of the same respiratory effects at similar SO₂ exposure concentrations (Koenig
28 et al., 1981; 1983; 1987; 1988; 1990; 1992). It should be noted, however, that in all of these

1 studies involving adolescents, SO₂ was administered via inhalation through a mouthpiece rather
2 than an exposure chamber. This exposure technique bypasses nasal absorption of SO₂, likely
3 resulting in a relative increase of pulmonary SO₂ uptake (see Section 2.6.1).

Table 3-1. Percentage of asthmatic individuals in controlled human exposures experiencing SO₂-induced decrements in lung function.

SO ₂ CONC (ppm)	EXPOSURE DURATION	NO. SUBJ	VENTILATION (L/MIN)	LUNG FUNCT	CUMULATIVE PERCENTAGE OF RESPONDERS (NUMBER OF SUBJECTS) ¹			REFERENCE	RESPIRATORY SYMPTOMS: SUPPORTING STUDIES
					≥ 100% ↑	sRaw ≥ 200% ↑	≥ 300% ↑		
					≥ 15% ↓	FEV ₁ ≥ 20% ↓	≥ 30% ↓		
0.2	10 min	40	~40	sRaw	5% (2)	0	0	Linn et al. (1987) ²	Some evidence of SO ₂ -induced increases in respiratory symptoms in the most sensitive individuals: Linn et al. (1987; 1988; 1990; 1984a; 1983), Schacter et al. (1984)
	10 min	40	~40	FEV ₁	13% (5)	5% (2)	3% (1)	Linn et al. (1987)	
0.25	5 min	19	~50-60	sRaw	32% (6)	16% (3)	0	Bethel et al. (1985)	
	5 min	9	~80-90	sRaw	22% (2)	0	0		
	10 min	28	~40	sRaw	4% (1)	0	0	Roger et al. (1985)	
0.3	10 min	20	~50	sRaw	10% (2)	5% (1)	5% (1)	Linn et al. (1988) ³	
	10 min	21	~50	sRaw	33% (7)	10% (2)	0	Linn et al. (1990) ³	
	10 min	20	~50	FEV ₁	15% (3)	0	0	Linn et al. (1988)	
	10 min	21	~50	FEV ₁	24% (5)	14% (3)	10% (2)	Linn et al. (1990)	
0.4	10 min	40	~40	sRaw	23% (9)	8% (3)	3% (1)	Linn et al. (1987)	Stronger evidence with some statistically significant increases in respiratory symptoms: Balme et al. (1987) ⁴ , Gong et al. (1995), Linn et al. (1987; 1983), Roger et al. (1985)
	10 min	40	~40	FEV ₁	30% (12)	23% (9)	13% (5)	Linn et al. (1987)	
0.5	5 min	10	~50-60	sRaw	60% (6)	40% (4)	20% (2)	Bethel et al. (1983)	
	10 min	28	~40	sRaw	21% (6)	4% (1)	4% (1)	Roger et al. (1985)	
	10 min	45	~30	sRaw	36% (16)	16% (7)	13% (6)	Magnussen et al. (1990) ⁴	
0.6	10 min	40	~40	sRaw	35% (14)	28% (11)	18% (7)	Linn et al. (1987)	Clear and consistent increases in SO ₂ -induced respiratory symptoms: Linn et al. (1987; 1988; 1984a; 1990), Gong et al. (1995), Horstman et al. (1988)
	10 min	20	~50	sRaw	60% (12)	35% (7)	10% (2)	Linn et al. (1988)	
	10 min	21	~50	sRaw	57% (12)	33% (7)	14% (3)	Linn et al. (1990)	
	10 min	40	~40	FEV ₁	53% (21)	45% (18)	20% (8)	Linn et al. (1987)	
	10 min	20	~50	FEV ₁	55% (11)	55% (11)	5% (1)	Linn et al. (1988)	
	10 min	21	~50	FEV ₁	45% (9)	35% (7)	19% (4)	Linn et al. (1990)	
1.0	10 min	28	~40	sRaw	54% (15)	25% (7)	14% (4)	Roger et al. (1985)	
	10 min	10	~40	sRaw	60% (6)	20% (2)	0	Kehrl et al. (1987)	

¹Data presented from all references from which individual data were available. Percentage of individuals who experienced greater than or equal to a 100, 200, or 300% increase in specific airway resistance (sRaw), or a 15, 20, or 30% decrease in FEV₁. Lung function decrements are adjusted for effects of exercise in clean air.

²Responses of mild and moderate asthmatics reported in Linn et al. (1987) have been combined.

³Analysis includes data from only mild (1988) and moderate (1990) asthmatics who were not receiving supplemental medication.

⁴Indicates studies in which exposures were conducted using a mouthpiece rather than a chamber.

1 In laboratory animals, SO₂-induced decrements in lung function were observed following
2 peak exposures in several studies conducted since the last review. Most of these experiments
3 were designed to evaluate the mode of action underlying SO₂-mediated bronchoconstriction.
4 They used high concentrations of administered SO₂, which in many cases were delivered using
5 an endotracheal tube. As a result, these studies are of limited usefulness in understanding the
6 effects of SO₂ at or near ambient levels or under conditions of nasal breathing.

3.1.4. Respiratory Effects Associated with Short-Term (≥ 1 h) Exposure

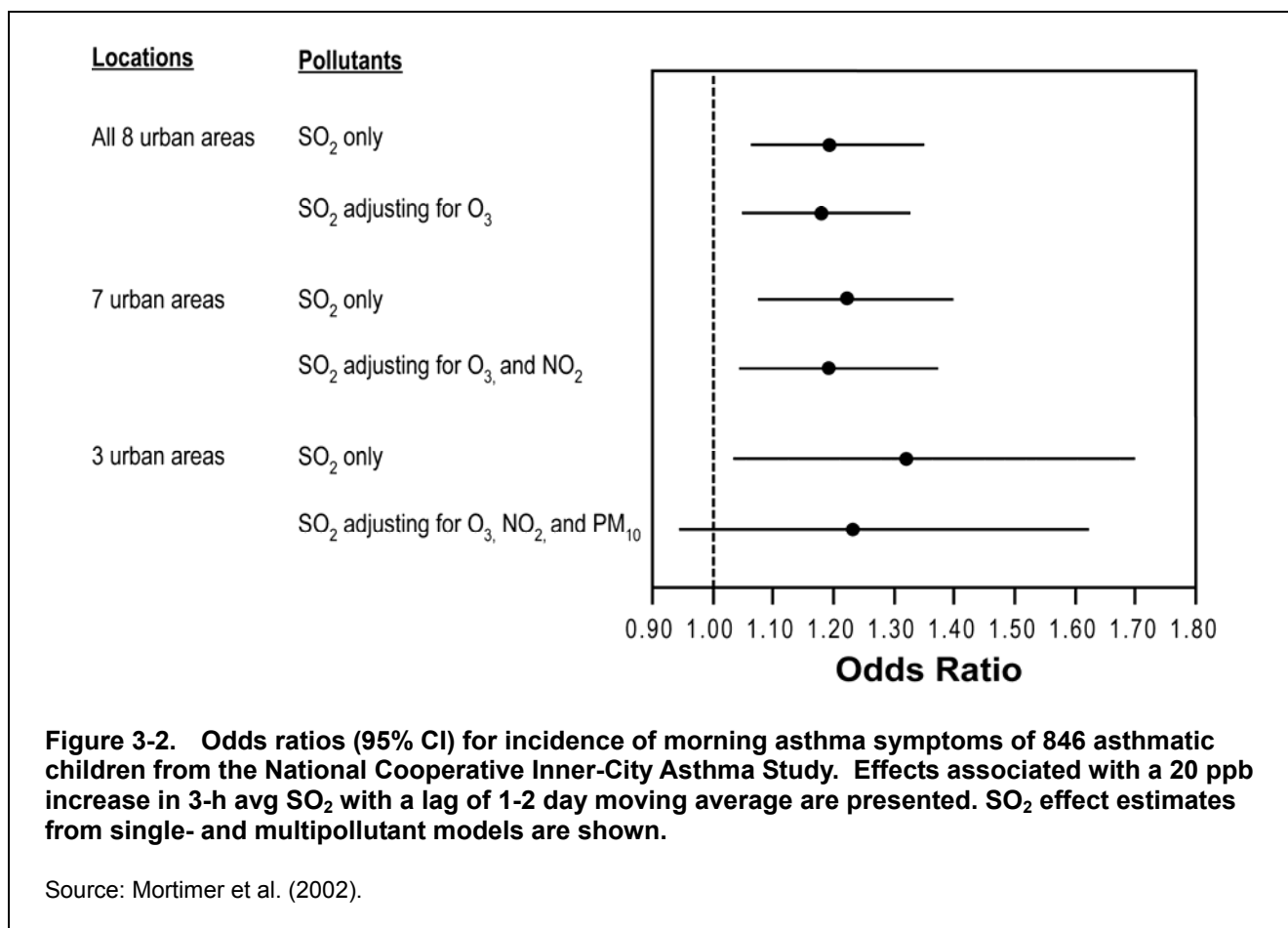
3.1.4.1. Respiratory Symptoms

7 Consideration of the mode of action suggests that SO₂ may contribute to respiratory
8 symptoms by stimulating mucus secretion and cough through activating central nervous system
9 reflexes. Recent studies in vitro have demonstrated increased expression of a gene encoding
10 mucin protein, MUC5AC, in human bronchial epithelial cells following exposure to the SO₂
11 derivatives sulfite and bisulfite at concentrations of 1-10 μM (Li and Meng, 2007). Increased
12 levels of MUC5AC protein were also reported. Sulfite and bisulfite were used, since SO₂
13 dissolves into the aqueous fluid and forms hydrogen ions and bisulfite and sulfite anions when it
14 contacts the fluids lining the airway. These same investigators conducted a related in vivo study
15 in which rats were exposed by inhalation to 2 ppm SO₂ for 1 h per day for 7 days. Rats which
16 were sensitized and challenged with ovalbumin, as well as exposed to SO₂, had increased
17 MUC5AC mRNA and protein levels compared with animals treated with ovalbumin or SO₂
18 alone (Li et al., 2007b). Further studies are required to determine the relevance of mucin gene
19 expression to mucous secretion and respiratory symptoms in allergic and non-allergic animals at
20 ambient levels of SO₂. However, evidence from toxicological studies such as these may provide
21 biological plausibility for the effects of SO₂ on respiratory symptoms in humans.

22 Epidemiological studies have examined the association between ambient SO₂
23 concentrations and respiratory symptoms in both adults and children. In air pollution field
24 studies, respiratory symptoms are usually assessed using questionnaire forms (or “daily diaries”)
25 completed by study subjects. Questions address the daily experience of coughing, wheezing,
26 shortness of breath (or difficulty breathing), production of phlegm, and others.

3.1.4.1.1. Children

1 Epidemiological studies on respiratory symptoms published since the last review are
2 summarized in Annex Table F-1; key studies are discussed in detail below.



3 The strongest epidemiological evidence for an association between respiratory symptoms
4 and exposure to ambient SO₂ comes from two large U.S. multicity studies (Mortimer et al., 2002;
5 Schildcrout et al., 2006). Mortimer et al. examined 846 asthmatic children from eight U.S. urban
6 areas in the National Cooperative Inner-City Asthma Study (NCICAS) for summertime air
7 pollution-related respiratory symptoms. Median 3-h avg SO₂ (8 to 11 a.m.) levels ranged from 17
8 ppb in Detroit, MI to 37 ppb in East Harlem, NY. Morning symptoms were found to be most
9 strongly associated with an average of a 1- to 2-day lag of SO₂ concentrations. In multipollutant
10 models with O₃ and NO₂ (measured in seven cities), the SO₂ association remained robust (see

Figure 3-2). When particulate matter with an aerodynamic diameter of $\leq 10 \mu\text{m}$ (PM_{10}) was also included in the multipollutant models, the SO_2 effect estimate decreased only slightly; however, it became nonsignificant, possibly due to reduced statistical power (only three of eight cities were included in this analysis) or collinearity resulting from adjustment of multiple pollutants. A similar decline was observed in the effect estimate for PM_{10} in the multipollutant model compared to the single-pollutant model.

In the Childhood Asthma Management Program (CAMP) study, the association between ambient air pollution and asthma exacerbations in children ($n = 990$) from eight North American cities was investigated (Schildcrout et al., 2006). SO_2 measurements were available in seven of the eight cities. The median 24-h avg SO_2 concentrations ranged from 2.2 ppb (interquartile range [IQR]: 1.7, 3.1) in San Diego, CA to 7.4 ppb (IQR: 5.3, 10.7) in St. Louis, MO. Results for the associations between asthma symptoms and all pollutants are shown in Figure 3-3. Analyses indicate that although SO_2 was positively related to increased risk of asthma symptoms at all lags, only the 3-day moving average was statistically significant. No associations were observed between SO_2 and rescue inhaler use. Stronger associations were observed for CO and NO_2 . The effect estimates appear to be slightly larger in joint-pollutant models with CO or NO_2 , particularly at a 2-d lag, but did not change much when PM_{10} was jointly considered.

A longitudinal study of 1,844 schoolchildren during the summer from the Harvard Six Cities Study suggested that the association between SO_2 and respiratory symptoms could be confounded by PM_{10} (Schwartz et al., 1994). The median 24-h avg SO_2 concentration during this period was 4.1 ppb (10th–90th percentile: 0.8, 17.9; max 81.9). SO_2 concentrations were found to be associated with cough incidence and lower respiratory tract symptoms. Of the pollutants examined, PM_{10} had the strongest associations with respiratory symptoms. In two-pollutant models, the effect of PM_{10} was found to be robust to adjustment for other copollutants, while the effect of SO_2 was substantially reduced after adjustment for PM_{10} . Because the PM_{10} concentrations were correlated strongly to SO_2 -derived sulfate particles ($r = 0.80$), the diminution of the SO_2 effect estimate may indicate that for PM_{10} dominated by fine sulfate particles, PM_{10} has a slightly stronger association than SO_2 . This study further investigated the concentration-response function and observed a nonlinear relationship between SO_2 concentrations and respiratory symptoms. Though an increasing trend was observed at concentrations as low as

- 1 10 ppb, no statistically significant increase in the incidence of lower respiratory tract symptoms
- 2 was seen until concentration exceeded a 24-h avg SO₂ of 22 ppb.

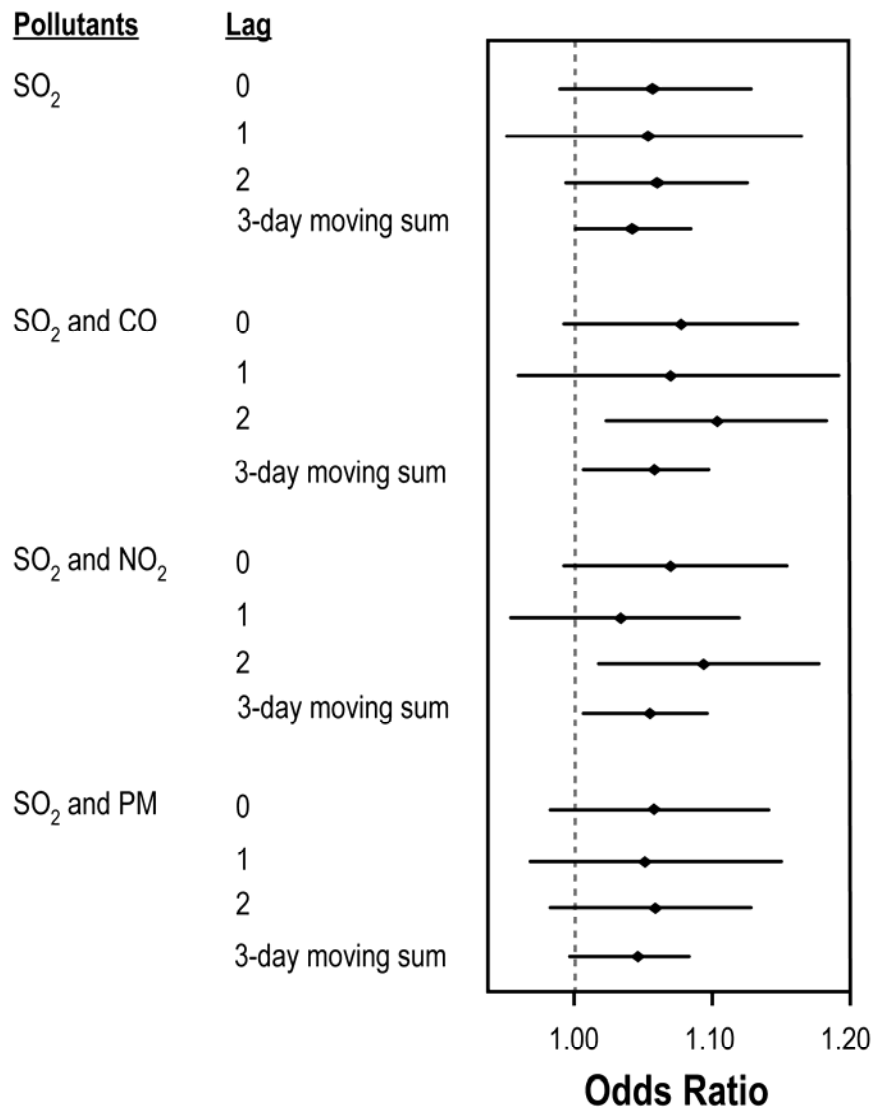


Figure 3-3. Odds ratios (95% CI) for daily asthma symptoms of 990 asthmatic children from the Childhood Asthma Management Program Study. Effects associated with a 10 ppb increase in within-subject concentrations of 24-h avg SO₂ are presented. Data collected from November 1993 to September 1995 were used. Results from single- and joint-pollutant models are shown.

Source: Schildcrout et al. (2006).

1 In the Pollution Effects on Asthmatic Children in Europe (PEACE) study, a multicenter
2 study of 14 cities across Europe, the effects of acute exposure to various pollutants including SO₂
3 on the respiratory health of children with chronic respiratory symptoms (n = 2,010) was
4 examined during the winter of 1993–1994 (Roemer et al., 1998). Mean 24-h avg SO₂
5 concentrations ranged from 1 ppb in the urban area of Umeå, Sweden, to 43 ppb in the urban
6 area of Prague, Czech Republic. No associations were observed between SO₂ and daily
7 prevalence of respiratory symptoms or bronchodilator use at any of the single- and multiday lags
8 considered. In addition, no associations were observed for any of the other pollutants examined.
9 It should be noted that during the study period, there were only two major air pollution episodes,
10 at the beginning and end of the study period. In the epidemiological model, the control for time
11 trend was accomplished through the use of linear and quadratic terms. Given the timing of the air
12 pollution episodes, the quadratic trend term would have removed most of the air pollution effect.
13 Other studies that participated in the PEACE study and analyzed results for longer periods of
14 time have observed statistically significant associations between SO₂ and respiratory symptoms
15 in children (van der Zee et al., 1999, presented below).

16 Additional studies have examined the relationship between respiratory symptoms and
17 ambient SO₂ concentrations and generally found positive associations, including two U.S. studies
18 (Delfino et al., 2003; Neas et al., 1995) and several European studies (Hoek and Brunekreef,
19 1994; Peters et al., 1996; Roemer et al., 1993; Segala et al., 1998; Timonen and Pekkanen, 1997;
20 van der Zee et al., 1999). However, some did not find a consistent association between
21 respiratory symptoms and SO₂ concentrations (e.g., Hoek and Brunekreef, 1993; 1995; Romieu
22 et al., 1996). Only one of these studies examined possible confounding of the SO₂ effect by
23 copollutants. Van der Zee et al. (1999) looked at the association between respiratory symptoms
24 and SO₂ in 7- to 11-year-old children (n = 633) with and without chronic respiratory symptoms
25 in the Netherlands. Significant associations with lower respiratory tract symptoms and increased
26 bronchodilator use were observed for SO₂, as well as PM₁₀, BS, and sulfate, in symptomatic
27 children living in urban areas (n = 142). In a two-pollutant model with PM₁₀, the results were
28 robust for bronchodilator use, but slightly reduced for lower respiratory tract symptoms. A
29 subgroup analysis of this cohort examining SO₂-related respiratory symptoms in children with
30 airway hyperresponsiveness and atopy (Boezen et al., 1999) is discussed in Section 3.1.4.4.

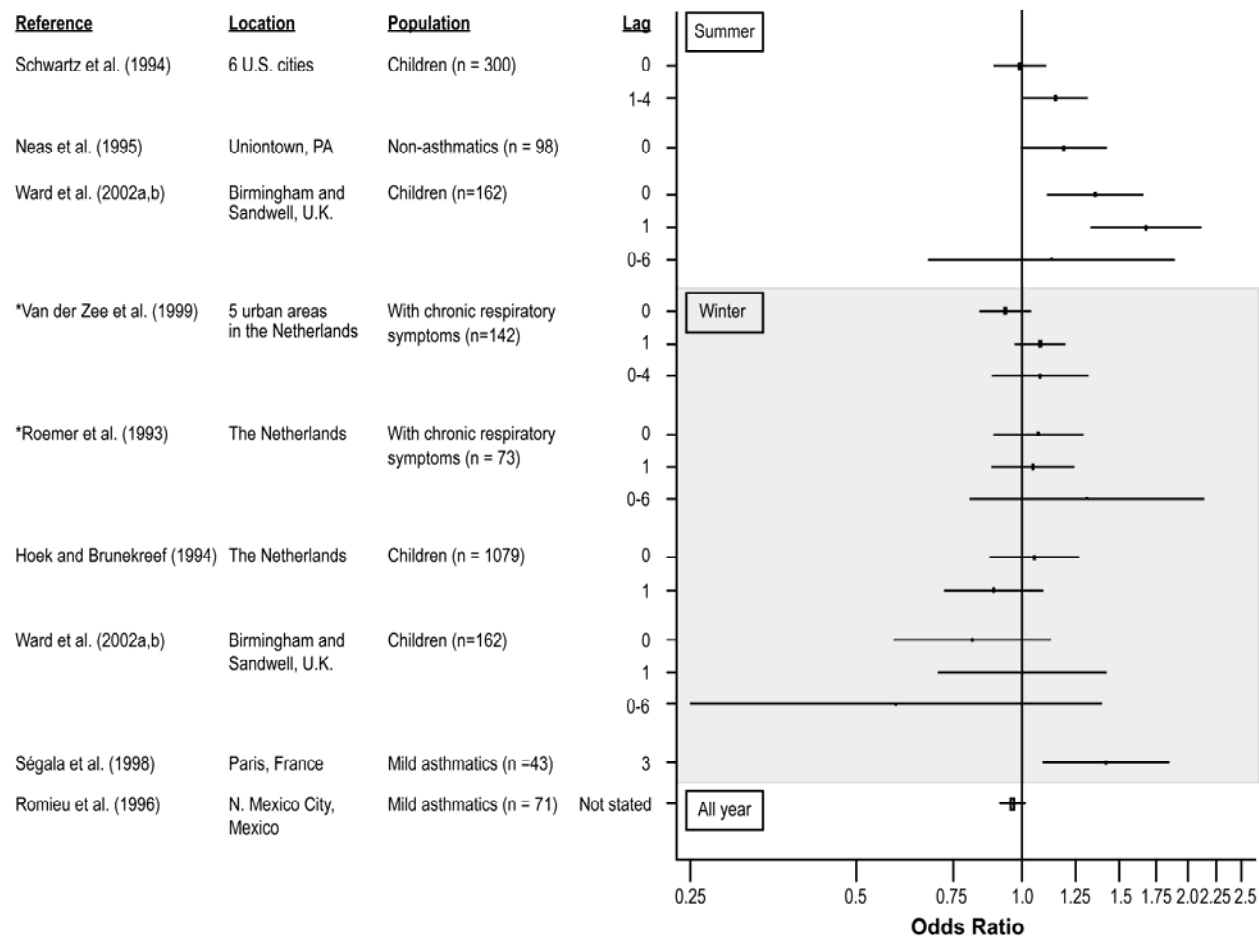
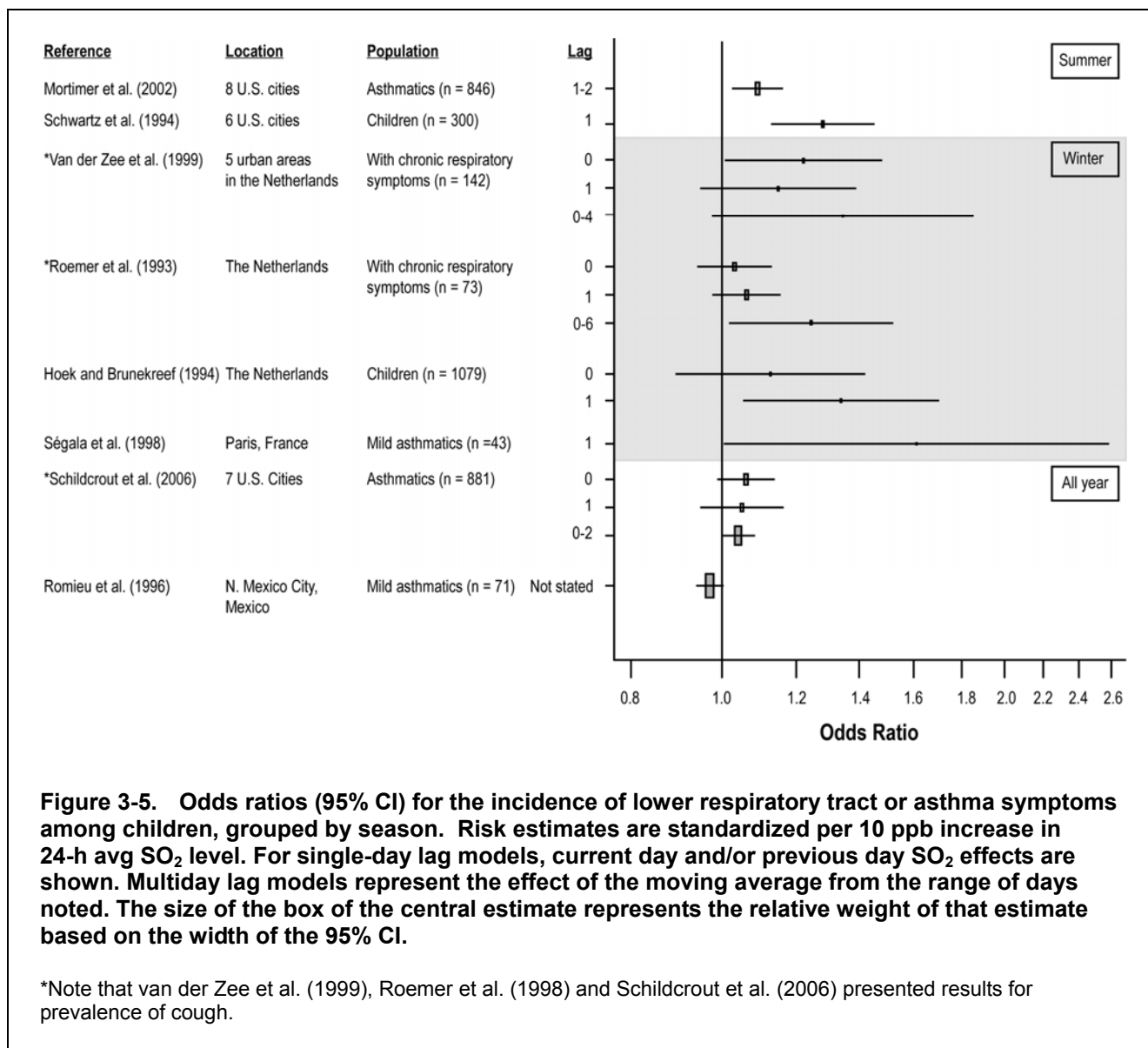


Figure 3-4. Odds ratios (95% CI) for incidence of cough among children, grouped by season. For single-day lag models, current day and/or previous day SO₂ effects are shown, except for Ségala et al. (1998), which only presented results for a 3-day lag. Multiday lag models represent the effect of the mean concentration from the range of days noted. Risk estimates are standardized per 10 ppb increase in 24-h avg SO₂ level. The size of the box of the central estimate represents the relative weight of that estimate based on the width of the 95% CI.

*Note that van der Zee et al. (1999) and Roemer et al.(1998) presented results for prevalence of cough.

Figure 3-4 and Figure 3-5 present the odds ratios for SO₂-related cough and lower respiratory tract or asthma symptoms, respectively, from several epidemiological studies with relevant data. The results for cough are somewhat variable with wide confidence intervals, as shown in Figure 3-4. The studies conducted in the summer generally indicate increased risk of cough from exposure to SO₂. A more consistent effect of SO₂ is observed on lower respiratory tract or asthma symptoms (Figure 3-5). Although there is some variability in the individual effect estimates, the majority of the odds ratios appear to be greater than one. As was the case with

cough, stronger associations with lower respiratory tract or asthma symptoms were observed in the summer, as opposed to the winter. There was some variability among the different lags of exposure; however, effects were generally observed with current day or previous day exposure and, in some cases, with a distributed lag of 2 to 3 days.



Overall, recent epidemiological studies provide evidence for an association between ambient SO₂ exposures and increased respiratory symptoms in children, particularly those with asthma or chronic respiratory symptoms. Recent U.S. multicity studies observed significant

1 associations between SO₂ and respiratory symptoms at a median range of 17 to 37 ppb (75th
2 percentile: ~25 to 50) across cities for 3-h avg SO₂ (NCICAS, Mortimer et al., 2002) and 2.2 to
3 7.4 ppb (90th percentile: 4.4 to 14.2) for 24-h avg SO₂ (CAMP, Schildcrout et al., 2006).
4 However, an earlier study that examined the concentration-response function found that a
5 statistically significant increase in the incidence of lower respiratory tract symptoms was not
6 observed until concentrations exceeded a 24-h avg SO₂ of 22 ppb, though an increasing trend
7 was observed at concentrations as low as 10 ppb (Harvard Six Cities Study, Schwartz et al.,
8 1994). In the limited number of studies that examined potential confounding by copollutants
9 through multipollutant models, the SO₂ effect was generally found to be robust after adjusting
10 for PM and other copollutants. More details of the literature published since the last review are
11 found in Annex Table F-1.

3.1.4.1.2. Adults

12 Compared to the number of studies conducted with children, fewer epidemiological studies
13 were performed that examined the effect of ambient SO₂ exposure on respiratory symptoms in
14 adults. Most of these studies focused on potentially susceptible populations, i.e., those with
15 asthma or COPD. One of the larger studies was conducted by van der Zee et al. (2000) in 50- to
16 70-year-old adults, with (n = 266) and without (n = 223) chronic respiratory symptoms in the
17 Netherlands. In adults both with and without chronic respiratory symptoms, no consistent
18 associations were observed between SO₂ levels and respiratory symptoms or medication use. A
19 subgroup analysis of this cohort examining SO₂-related respiratory symptoms in individuals with
20 airway hyperresponsiveness and atopy (Boezen et al., 2005) is discussed in Section 3.1.4.4.

21 Studies by Desqueyroux et al. (2002b; 2002a) examined the association between air
22 pollution and respiratory symptoms in other potentially susceptible populations, i.e., those with
23 severe asthma (n = 60, mean age 55 years) and COPD (n = 39, mean age 67 years), in Paris,
24 France. The mean 24-h avg SO₂ concentration was 3 ppb (range: 1, 10) in the summer and 7 ppb
25 (range: 1, 31) in the winter. No associations were observed between SO₂ concentrations and the
26 incidence of asthma attacks or episodes of symptom exacerbation in severe asthmatics or
27 individuals with COPD. O₃ was found to have the strongest effect in these studies.

28 Several other European studies did observe an association between ambient SO₂
29 concentrations and respiratory symptoms in adults with asthma or chronic bronchitis (Higgins et

al., 1995; Neukirch et al., 1998; Peters et al., 1996; Taggart et al., 1996). Only one of these studies examined possible confounding of the association by copollutants. Higgins et al. examined the effect of summertime air pollutant exposure on respiratory symptoms in 62 adults with either asthma, COPD, or both. The max 24-h avg SO₂ level was 45 ppb. An association was observed between SO₂ and symptoms of wheeze, and it remained robust after adjustment for O₃ and NO₂. The effects of PM were not examined in this study.

Results from the epidemiological studies examining the association between SO₂ and respiratory symptoms in adults are generally mixed, with some showing positive associations and others finding no relationship at current ambient levels. The overall epidemiological evidence that 24-h avg SO₂ exposures at or near ambient concentrations has an effect on adults is inconclusive. However, as discussed in Section 3.1.3.1, human clinical studies have observed an effect of peak exposures to SO₂ on respiratory symptoms, particularly among SO₂-sensitive asthmatics, with 10 min exposures to SO₂ concentrations as low as 0.4-0.6 ppm under exercise conditions. These effects in clinical studies are at levels that have sometimes been measured in ambient air for similarly short-time durations.

3.1.4.2. Lung Function

The 1982 AQCD reported bronchoconstriction, indicated by increased pulmonary resistance, as the most sensitive indicator of lung function effects of acute SO₂ exposure, based on the observations of increased pulmonary resistance in guinea pigs that were acutely exposed to 0.16 ppm SO₂. Since then, only a few new animal toxicological studies have measured lung function at or near ambient levels of SO₂. These studies, and those using higher concentrations of SO₂, are summarized in Annex Table E-4. Increased pulmonary resistance and decreased dynamic compliance were observed in conscious guinea pigs exposed to 1 ppm SO₂ for 1 h (Amdur et al., 1983). Effects were seen immediately after exposure and were not present 1 h post-exposure. No changes in tidal volume, minute volume or breathing frequency were found. These same investigators also exposed guinea pigs to 1 ppm SO₂ for 3 h/day for 6 days (Conner et al., 1985). No changes were observed in pulmonary function or respiratory parameters, i.e., diffusing capacity for carbon monoxide, functional reserve capacity, vital capacity, total lung capacity, respiratory frequency, tidal volume, pulmonary resistance or pulmonary compliance. In another study, Barthelemy et al. (1988) demonstrated a 16% increase in airway resistance

following a 45-min exposure of anesthetized rabbits to 0.5 ppm SO₂ via an endotracheal tube. This latter exposure is more relevant to oronasal than nasal breathing.

3.1.4.2.1. Children

Most epidemiological studies discussed in the previous section on respiratory symptoms also examined lung function. In these studies self-administered PEF meters were primarily used to assess lung function. PEF follows a circadian rhythm, with the highest values found during the afternoon and lowest values during the night and early morning (Borsboom et al., 1999). Therefore, these studies generally analyze PEF data stratified by time of day. The epidemiological studies on lung function are summarized in Annex Table F-1.

Mortimer et al. (2002) examined 846 asthmatic children from eight U.S. urban areas in the NCICAS for changes in PEF related to air pollution. The mean 3-h avg SO₂ was 22 ppb across the eight cities during the study period of June through August, 1993. No associations were observed between SO₂ concentrations and morning or evening PEF. Of all the pollutants examined, including PM₁₀, O₃, and NO₂, only O₃ was associated with changes in morning PEF.

In another U.S. study (Neas et al., 1995), 83 children from Uniontown, PA reported twice-daily PEF measurements during the summer of 1990. The mean daytime 12-h avg SO₂ concentration was 14.5 ppb (max 44.9). No associations were observed between daytime 12-h avg SO₂ concentrations and mean deviation in evening PEF, even after concentrations were weighted by the proportion of hours spent outdoors during the prior 12-h. Statistically significant associations were observed for O₃, total sulfate particles, and particle-strong acidity.

A study by van der Zee et al. (1999) observed associations between ambient SO₂ concentrations and daily PEF measurements in 7- to 11-year-old children (n = 142) with chronic respiratory symptoms living in urban areas of the Netherlands. The OR for a > 10% decrement in evening PEF per 10 ppb increase in 24-h avg SO₂ was 1.20 (95% CI: 0.97, 1.47) with same-day exposure. A greater effect was observed at a 2-day lag, OR = 1.40 (95% CI: 1.18, 1.67), and this effect remained robust in a two-pollutant model with PM₁₀, OR = 1.34 (95% CI: 1.08, 1.64).

Multipollutant analyses also were conducted in a study by Chen et al. (1999), which examined the effects of short-term exposure to air pollution on the pulmonary function of 895 children, ages 8 to 13 years, in three communities in Taiwan. The daytime 1-h max SO₂ the day before spirometry ranged from 0 to 72.4 ppb. In a single-pollutant model, 1-h max SO₂

1 concentration at a 2-day lag was significantly associated with FVC, -50.80 mL (95% CI: -97.06,
2 -4.54), or a 2.6% decline, per 40 ppb 1-h max SO₂. However, in multipollutant models, authors
3 noted that only O₃ remained significantly associated with FVC and FEV₁. Effect estimates for
4 SO₂ in multipollutant models were not provided.

5 While additional studies have observed associations between ambient SO₂ concentrations
6 and changes in lung function in children (Hoek and Brunekreef, 1993; Roemer et al., 1993;
7 Peters et al., 1996; Segala et al., 1998; Timonen and Pekkanen, 1997), several other studies did
8 not find a significant association between SO₂ and lung function parameters. In addition, in
9 studies that did observe an association, the correlations between SO₂ and other pollutants,
10 particularly PM indices, were high [for example, $r = 0.8-0.9$ in Peters et al. (1996), making it
11 difficult to separate the contributions of individual pollutants.

12 In conclusion, while some epidemiological studies observed a positive association between
13 short-term SO₂ exposure and lung function in children, several others, including a large U.S.
14 multicity study, did not observe such an association. The limited evaluation of potential
15 confounding by copollutants also indicated mixed results. Overall, the evidence is insufficient to
16 conclude that short-term exposure to ambient SO₂ has an independent effect on lung function in
17 children.

3.1.4.2.2. Adults

18 Only a limited number of epidemiological studies have been conducted examining the
19 association between ambient SO₂ concentrations and lung function in adults, as in the case of
20 respiratory symptoms. In a cross-sectional survey, Xu et al. (1991) investigated the effects of
21 indoor and outdoor air pollutants on the respiratory health of 1,140 adults (aged 40 to 69 years)
22 living in residential, industrial, and suburban areas of Beijing, China. The annual mean
23 concentrations of SO₂ in residential, industrial, and suburban areas from 1981 to 1985 were 49
24 ppb, 22 ppb, and 7 ppb, respectively. Log-transformed SO₂ and TSP were significantly
25 associated with reductions in FEV₁ and FVC. The authors cautioned that since SO₂ and TSP
26 concentrations were strongly correlated, the effect of SO₂ could not be separated.

27 Van der Zee et al. (2000) observed an association between SO₂ and morning PEF in 50- to
28 70-year-old adults ($n = 138$) with chronic respiratory symptoms living in urban areas of the
29 Netherlands. No associations were observed with evening PEF. The OR for a > 20% decrement

1 in PEF was 1.21 (95% CI: 0.76, 1.92) per 10 ppb increase in 24-h avg SO₂ with same-day
2 exposure and 1.56 (95% CI: 1.02, 2.39) at a 1-day lag. No associations were observed for a
3 > 10% decrement in PEF. The authors hypothesized that while SO₂ level did not have much
4 effect on PEF in most subjects, a small subgroup of individuals experienced fairly large PEF
5 decrements when SO₂ levels were high. No multipollutant analyses were conducted.

6 Higgins et al. (1995) examined the association between pulmonary function and air
7 pollution in 75 adults with either asthma, COPD, or both. Exposure to SO₂ was associated with
8 increased variation in PEF, but not with mean or minimum PEF. The SO₂ effects on PEF
9 variation were robust to adjustment for O₃ and NO₂. Effects of PM were not considered.
10 Neukirch et al. (1998) also observed associations between lung function and SO₂ concentrations
11 in a study of asthmatic adults in Paris, France; however, significant associations were found for
12 all pollutants examined, including BS, PM₁₃, and NO₂. Other epidemiological studies observed
13 only weak relationships between ambient SO₂ concentrations and lung function in adults (Peters
14 et al., 1996; Taggart et al., 1996).

15 Evidence from human clinical studies clearly indicates that asthmatic individuals
16 experience moderate or greater decrements in lung function, as well as increased respiratory
17 symptoms, following peak exposure (5-10 min) to SO₂ (Balmes et al., 1987; Gong et al., 1995;
18 Horstman et al., 1986; Linn et al., 1987; 1983) These effects were seen at peak concentrations as
19 low as 0.4-0.6 ppm. However, in a human clinical study by Tunnicliffe et al. (2003) that
20 evaluated the effect of 1-h exposures to 0.2 ppm SO₂ in resting healthy and asthmatic subjects,
21 no significant changes were observed in lung function as measured by FEV₁, FVC, and maximal
22 midexpiratory flow (MMEF).

23 In summary, the epidemiological studies examining adults do not provide strong evidence
24 for an association between short-term exposure to ambient SO₂ and lung function. While some
25 studies did observe associations between SO₂ exposure and decrements in lung function
26 parameters, the strong correlation between SO₂ and various copollutants in most studies, and the
27 lack of evidence evaluating potential confounding by copollutants, limit interpretation of
28 independent effects of SO₂ on lung function.

3.1.4.3. Airway Inflammation

The animal toxicological studies on airway inflammation are summarized in Annex Table E-1. In one study, guinea pigs were exposed to 1 ppm SO₂ for 3 h/day for 5 days and bronchoalveolar lavage was performed daily (Conner et al., 1985) No change in numbers of total cells or neutrophils was observed. However, in two models of allergic sensitization, SO₂ exposure increased airway inflammation. In one study (Park et al., 2001), guinea pigs were exposed to 0.1 ppm SO₂ for 5 h/day for 5 days and sensitized with 0.1% ovalbumin aerosols for 45 min on days 3-5. One week later, animals were subjected to bronchial challenge with 1.0% ovalbumin and bronchoalveolar lavage and histopathologic examination were performed 24 h later. Results demonstrated increased numbers of eosinophils in lavage fluid, and an infiltration of inflammatory cells, bronchiolar epithelial cell damage and plugging of the airway lumen with mucus and cells in the bronchial tissues of animals treated with both SO₂ and ovalbumin, but not in animals treated with ovalbumin or SO₂ alone.

In a second study, rats which were sensitized and challenged with ovalbumin and exposed to 2 ppm SO₂ for 1 h/day for 7 days had an increased number of inflammatory cells in bronchoalveolar lavage fluid and an enhanced histopathological response compared with those treated with ovalbumin or SO₂ alone (Li et al., 2007a). Similar responses were noted for ICAM-1, a protein involved in regulating inflammation. Further experiments are required to determine whether near ambient SO₂ also enhance inflammatory responses in non-allergic and allergic rats. Taken together, these animal experiments suggest that near-ambient levels of SO₂ may play a role in exacerbating allergic responses.

In a human clinical study, Tunnicliffe et al. (2003) measured levels of exhaled NO (eNO) in asthmatic and healthy adult subjects, before and after 1-h exposure to 0.2 ppm SO₂ under resting conditions. While eNO concentrations were higher in the asthmatic than in healthy subjects, no significant difference was observed between pre- and postexposure in either group.

One epidemiological study by Adamkiewicz et al. (2004) examined eNO as a biological marker for inflammation in 29 older adults (median age 70.7 years) in Steubenville, OH. The mean 24-h avg SO₂ concentration was 12.5 ppb (IQR 11.5). The authors reported that, while significant and robust associations were observed between increased daily levels of fine PM (PM_{2.5}) and increased eNO, no associations were observed with any of the other pollutants examined, including SO₂, NO₂, and O₃.

Overall, the very limited human clinical and epidemiological evidence does not indicate that exposure to SO₂ at current ambient concentrations is associated with inflammation in the airway. However, toxicological studies suggest that repeated exposures to SO₂ at concentrations as low as 0.1 ppm in guinea pigs, may exacerbate inflammatory responses in allergic animals.

3.1.4.4. Airway Hyperresponsiveness and Allergy

The toxicological studies describing SO₂-induced effects on airway obstruction, hypersensitivity and/or allergy in guinea pigs and sheep are summarized in Annex Table E-3. In one study, Amdur et al. (1988) exposed guinea pigs for 1 h to 1 ppm SO₂ and measured airway responsiveness to acetylcholine 2 h later. No airway hyperresponsiveness (AHR) was observed. In a second study, Douglas et al., (1994) found no AHR following a histamine challenge 24 h after exposure of rabbits to 5 ppm SO₂ for 2 h. In a third study, exposure of sheep for 4 h to 5 ppm SO₂ failed to result in AHR following carbachol (Gong et al., 2001). In a fourth study, a 5 min exposure to 30 ppm but not to 10 ppm SO₂ resulted in AHR in horses challenged with methacholine (Lewis and Kirchner, 1984). Collectively, these results show that a single exposure to SO₂ at a concentration of 10 ppm or less failed to induce AHR following challenge in 4 different animal models.

However, two other studies demonstrated increased airway responsiveness in guinea pigs exposed repeatedly to SO₂ and allergen. Riedel et al. (1988) studied the effect of SO₂ exposure on local bronchial sensitization to inhaled antigen. Guinea pigs were exposed by inhalation to 0.1, 4.3 and 16.6 ppm SO₂ for 8 h/d for 5 days. During the last 3 days, SO₂ exposure was followed by exposure to nebulized ovalbumin for 45 min. Following bronchial provocation with inhaled ovalbumin (0.1%) one week later, airway obstruction was measured by whole body plethysmography. In addition, specific antibodies against ovalbumin were measured in serum and bronchoalveolar fluids. Results show significantly higher bronchial obstruction in animals exposed to SO₂ (at all concentration levels) with ovalbumin compared with animals exposed only to ovalbumin. In addition, significant increases in anti-ovalbumin IgG antibodies were detected in bronchoalveolar lavage fluid of animals exposed to 0.1, 4.3 and 16.6 ppm SO₂ and in serum from animals exposed to 4.3 and 16.6 ppm SO₂ compared with controls exposed only to ovalbumin. These results demonstrate that repeated exposure to SO₂ can enhance allergic sensitization in the guinea pig at a concentration as low as 0.1 ppm. In a second study, guinea

pigs were exposed to 0.1 ppm SO₂ for 5 h/day for 5 days and sensitized with 0.1% ovalbumin aerosols for 45 min on days 3 to 5 (Park et al., 2001). One week later, animals were subjected to bronchial challenge with 1.0% ovalbumin and lung function was evaluated 24 h later by whole body plethysmography. Results demonstrated a significant increase in enhanced pause (P_{enh}), a measure of airway obstruction, in animals exposed to SO₂ with ovalbumin but not in animals treated with ovalbumin or SO₂ alone. These experiments also indicate that near ambient levels of SO₂ may play a role in exacerbating allergic responses in the guinea pig.

In a human clinical study evaluating SO₂-induced AHR to an inhaled allergen (house dust mite), Devalia et al. (1994) found that neither SO₂ (0.2 ppm) nor NO₂ (0.4 ppm) enhanced sensitization to the allergen in asthmatic individuals. However, following concurrent exposure (6 h) to SO₂ and NO₂ while at rest, subjects did exhibit increased sensitivity to the inhaled allergen. In a subsequent study, Rusznak et al. (1996) confirmed these findings and observed that the combination of SO₂ and NO₂ enhanced sensitization to house dust mite antigen up to 48 hours post-exposure.

A limited number of epidemiological studies also examined the association between SO₂ and AHR. Other studies considered individuals with AHR and atopy as a subgroup potentially susceptible to SO₂-related health effects. These studies are summarized in Annex Table F-1. Søyseth et al. (1995) investigated the effect of short-term exposure to SO₂ and fluoride on the number of capillary blood eosinophils, and the prevalence of AHR in schoolchildren, ages 7 to 13 years, (n = 620) from two regions in Norway, a valley containing an SO₂-emitting aluminum smelter and a similar but nonindustrialized valley. The median 24-h avg SO₂ concentration was 8 ppb (10th–90th percentile: 1, 33) in the exposed area and 1 ppb (10th–90th percentile: 0, 4) in the nonindustrialized valley. The mean number of eosinophils was significantly greater in children living near the aluminum smelter compared to the nonindustrialized area. However, within children in the exposed area, a negative concentration-response relationship was observed between mean eosinophils and previous-day 24-h avg SO₂. The observed association between SO₂ and eosinophils was limited to atopic children. In children living in the exposed area, a statistically significant positive association was observed between prevalence of AHR and previous-day 24-h avg SO₂ concentrations. Similar associations were observed for fluoride. The authors hypothesized that recent exposure to SO₂ may have induced changes in the airway

1 leading to AHR, in addition to recruitment of eosinophils to the airways in atopic subjects.
2 Exposure to PM was not assessed in this study.

3 A study by Taggart et al. (1996) examined the effect of summertime air pollution levels in
4 northwestern England on AHR in nonsmoking, asthmatic subjects ($n = 38$) aged 18 to 80 years
5 who were determined to be methacholine (MCh) reactors. Subjects were tested multiple times,
6 for a total of 109 evaluable challenge tests, with a range of two to four tests per subject. The max
7 24-h avg SO_2 concentration during the study period was 40 ppb. This study reported that
8 24-h avg SO_2 levels were marginally associated with a decreased dose of MCh required for a
9 20% drop in the postsaline FEV_1 ($\text{PD}_{20}\text{FEV}_1$).

10 Other epidemiological studies investigated the effect of exposure to SO_2 on children and
11 adults with AHR and atopy. Boezen et al. (1999) examined 7- to 11-year-old children ($n = 459$)
12 in the Netherlands and tested them for AHR and atopy. These children were a subset of a larger
13 cohort examined in van der Zee et al. (1999). It was hypothesized that children with AHR, as
14 measured using a MCh challenge, and atopy, indicated by raised serum total IgE ($> 60 \text{ kU/L}$, the
15 median value), may be susceptible to the effects of air pollution. One of the strengths of this
16 study was the use of AHR and serum IgE concentration to indicate susceptibility; these
17 measurements would be less prone to error than self-reported chronic respiratory symptoms. A
18 total of 121 children were found to have AHR and relatively high serum total IgE; 67 had AHR
19 and relatively low serum total IgE, 104 had no AHR but had a relatively high serum total IgE
20 concentration, and 167 were found to have neither AHR nor relatively high serum total IgE. For
21 the subset of children with relatively low serum total IgE with or without AHR, no associations
22 were observed between SO_2 and any respiratory symptoms. However, for children with relatively
23 high serum total IgE either with or without AHR, the prevalence of lower respiratory tract
24 symptoms increased with increasing SO_2 concentrations. For children with AHR and relatively
25 high serum total IgE, the OR for the prevalence of lower respiratory tract symptoms was 1.70
26 (95% CI: 1.26, 2.29) with a 5-day moving average for every 10 ppb increase in SO_2 . For children
27 without AHR but with relatively high serum total IgE, the OR was 1.82 (95% CI: 1.33, 2.50)
28 with a 5-day moving average.

29 Boezen et al. (2005) conducted a similar study in 50- to 70-year-old adults ($n = 327$) in the
30 Netherlands. The subgroup of individuals with elevated serum total IgE, both with ($n = 48$) and
31 without ($n = 112$) AHR, were found to be more susceptible to air pollutants when contrasted with

those who did not have elevated serum total IgE ($n = 167$). Significant associations were observed between previous-day 24-h avg SO_2 concentrations and the prevalence of upper respiratory tract symptoms in those with elevated serum total IgE. Stratified analyses by gender indicated that, among those with AHR and elevated IgE, only males ($n = 25$) were at a higher risk for respiratory symptoms. The OR for these males was 3.54 (95% CI: 1.79, 7.07) increase in 24-h avg SO_2 for a 5-day moving average, compared with 1.05 (95% CI: 0.59, 1.91) for the females.

In summary, the animal toxicological evidence suggests that repeated exposures to SO_2 at concentrations as low as 0.1 ppm in guinea pigs can exacerbate airway responsiveness following allergic sensitization. Two new human clinical studies have demonstrated an increase in sensitivity to an inhaled allergen in asthmatic subjects following exposures to a combination of 0.2-ppm SO_2 and 0.4-ppm NO_2 . These findings are consistent with the very limited epidemiological evidence that suggests that exposure to SO_2 may lead to AHR in atopic individuals.

3.1.4.5. Respiratory Illness-Related Absences

An additional concern has been the potential for SO_2 exposure to enhance susceptibility to, or the severity of illness resulting from respiratory infections, especially in children. School absenteeism is an indicator of morbidity in children caused by acute conditions. Respiratory conditions are the most frequent cause, particularly influenza and the common childhood infectious diseases. Park et al. (2002) examined the association between air pollution and school absenteeism in 1,264 first- to sixth-grade students attending school in Seoul, Korea. The study period extended from March 1996 to December 1999, with a mean 24-h avg SO_2 concentration of 9.19 ppb (SD 4.61). Note that analyses were performed using Poisson Generalized Additive Model (GAM) with default convergence criteria. Same-day SO_2 concentrations were positively associated with illness-related absences (16% excess risk [95% CI: 13, 22] per 10 ppb increase in 24-h avg SO_2), but inversely associated with non-illness-related absences (9% decrease [95% CI: 2, 15]). PM_{10} and O_3 concentrations also were positively associated with illness-related absences. In two-pollutant models containing SO_2 and either PM_{10} or O_3 , the SO_2 estimates were robust.

A study by Pönka (1990) observed results that were consistent with those from the Park et al. (2002) study. Pönka found that absenteeism due to febrile illnesses among children in day

care centers and schools and in adults was significantly higher on days of higher SO₂ concentrations (> 8.1 ppb weekly mean of 1-h avg), compared to days of lower SO₂ concentrations. In addition, on days of higher SO₂ concentrations, the mean weekly number of cases of upper respiratory tract infections and tonsillitis reported from health centers increased. Temperature, but not NO₂, was also found to be associated with febrile illnesses and respiratory tract infections. From these epidemiological studies, it is unknown whether SO₂ increases susceptibility to infection or whether its presence exacerbates preexisting morbidity following infection.

Pino et al. (2004) examined the association between air pollution and respiratory illnesses in a cohort of 504 infants recruited at 4 months of age from primary health care units in southeastern Santiago, Chile. The infants were followed through the first year of life. The mean 24-h avg SO₂ concentration was 11.6 ppb (5th–95th percentile: 3.0, 29.0). The most frequent diagnosis during follow-up was wheezing bronchitis. No associations were observed between current-day or previous-day SO₂ and wheezing bronchitis, but with a 7-day lag, a 21% (95% CI: 8, 39) excess risk in wheezing bronchitis was observed per 10 ppb increase in 24-h avg SO₂. However, it should be noted that stronger associations were observed with PM_{2.5}, which was well-correlated with SO₂ (r = 0.73). These epidemiologic studies are summarized in Annex Table F-1.

To summarize, very few studies have examined the association between ambient SO₂ concentrations and absences from school or work as a result of respiratory illnesses. The limited evidence suggests a possible association between exposure to SO₂ concentrations and increased respiratory illnesses, particularly among young children; however, this association was also seen with PM, which was correlated with SO₂.

3.1.4.6. Emergency Department Visits and Hospitalizations for Respiratory Diseases

Total respiratory causes for ED visits and hospital admissions typically include asthma, bronchitis and emphysema (collectively referred to as COPD), upper and lower respiratory tract infections, pneumonia, and other minor categories. Temporal associations between ED visits or hospital admissions for respiratory diseases and the ambient concentrations of SO₂ have been the subject of more than fifty peer-reviewed research publications since 1994. In addition to considerable statistical and analytical refinements, recent studies have examined responses of

morbidity in different age groups, the effect of seasons on ED and hospital usage, and multipollutant models to characterize the effects of copollutant mixtures. The epidemiological studies of ED visits and hospital admissions for respiratory causes are summarized in Annex Table F-2.

3.1.4.6.1. All Respiratory Diseases

There are relatively few studies of ED visits for all respiratory causes in contrast to the quantity of studies that examine hospital admissions for all respiratory causes. Collectively, studies of ED visits and hospitalizations provide suggestive evidence of an association between ambient SO₂ levels and ED visits and hospitalizations for all respiratory causes. When analyses were restricted by age, the results among children (0-14 years) and older adults (65+ years) were mainly positive, though not all statistically significant. The studies that examined the association of these outcomes and SO₂ levels among adults (15-64 years) reported a mix of positive and negative results. When all age groups were combined, the results of ED and hospitalization studies were mainly positive; however, the excess risk estimates were generally smaller compared to the children and older adults groups. It is possible that the effects observed in the combined age groups were driven by increases in the very young or older adult subpopulations. The results from the hospitalization and ED studies, separated by analyses among all ages and age-specific analyses, are shown in Figure 3-6. Overall, the effect estimates in this figure range from a -5% to 20% excess risk in ED visits or hospital admissions for respiratory causes per 10 ppb increase in 24-h avg SO₂, with the large majority of studies suggesting an increase in risk.

Wilson et al. (2005) examined ED visits for all respiratory causes in Portland, ME from 1996–2000 and in Manchester, NH from 1998–2000. The mean 1-h max SO₂ concentration in Portland was 11.1 ppb (SD 9.1), and was higher during the winter months (mean 17.1 ppb (SD 12.0)) and lower in the summer (mean 9.1 ppb [SD 8.0]). In Manchester, the mean 1-h max SO₂ concentration was 16.5 ppb (SD 14.7 ppb), and was higher in the winter months (mean 25.7 ppb [SD 15.8]) as opposed to the summer months (mean 10.6 ppb [SD 15.1]). Though the authors reported the 1-h max SO₂ concentrations, they used the 24-h avg SO₂ concentrations in their analyses. When all ages were included in analyses, Wilson et al. found positive associations between ED visits and SO₂, with an 8% (95% CI: 3.0, 11) and 11% (95% CI: 0.0, 20.0) excess

1 risk per 10 ppb increase in 24-h avg SO₂ at a 0-d lag in Portland, ME and Manchester, NH,
2 respectively.

3 Peel et al. (2005) investigated ED visits for all respiratory causes in Atlanta, GA from
4 1993–2000. This study included 484,830 ED visits. The mean 1-h max SO₂ concentration was
5 16.5 ppb (SD 17.1). The researchers found a weak positive relationship between ED visits and
6 SO₂, though the increased risk was not statistically significant (1.6% [95% CI: –0.6, 3.8] excess
7 risk per 40 ppb increase in 1-h max SO₂). Tolbert et al. (2007) recently reanalyzed these data
8 with four additional years of data and found similar results. An analysis by Dab et al. (1996)
9 examined the association between SO₂ and hospital admissions for all respiratory causes in Paris,
10 France, using both the 24-h avg and 1-h max. It should be noted that these researchers observed
11 similar effect estimates for both exposure metrics; however, only the estimate using 24-h avg
12 was statistically significant (1.1% [95% CI: 0.1, 2.0] excess risk per 10 ppb increase in 24-h avg
13 SO₂ versus 1.9% [95% CI: –1.3, 5.0]) per 40 ppb increase in 1-h max SO₂).

14 When analyses were stratified to include only children (0-14 years), evidence of a modest
15 association between SO₂ and ED visits or hospitalizations for all respiratory causes in children
16 was reported in several Australian (Barnett et al., 2005; Petroeschevsky et al., 2001) and
17 European (Anderson et al., 2001; Atkinson et al., 1999a; 1999b) studies. Excess risks ranging
18 from 3% to 22% per 10 ppb increase in 24-h avg SO₂ were reported by these studies. In a
19 multicity study spanning Australia and New Zealand, Barnett et al. (2005) compared hospital
20 admission data collected from 1998–2001 with ambient SO₂ concentrations, where the mean
21 24-h avg SO₂ concentration ranged from 0.9 to 4.8 ppb. The authors found a 5% (95% CI: 1, 9)
22 excess risk per 10 ppb increment in 24-h avg SO₂ among children (1-4 years) in these cities.
23 However, some additional U.S. (Wilson et al., 2005), European (Fusco et al., 2001; Ponce de
24 Leon et al., 1996), and Latin American (Braga et al., 1999; 2001) studies did not find statistically
25 significant associations between ambient SO₂ concentrations and hospitalizations for all
26 respiratory causes among children.

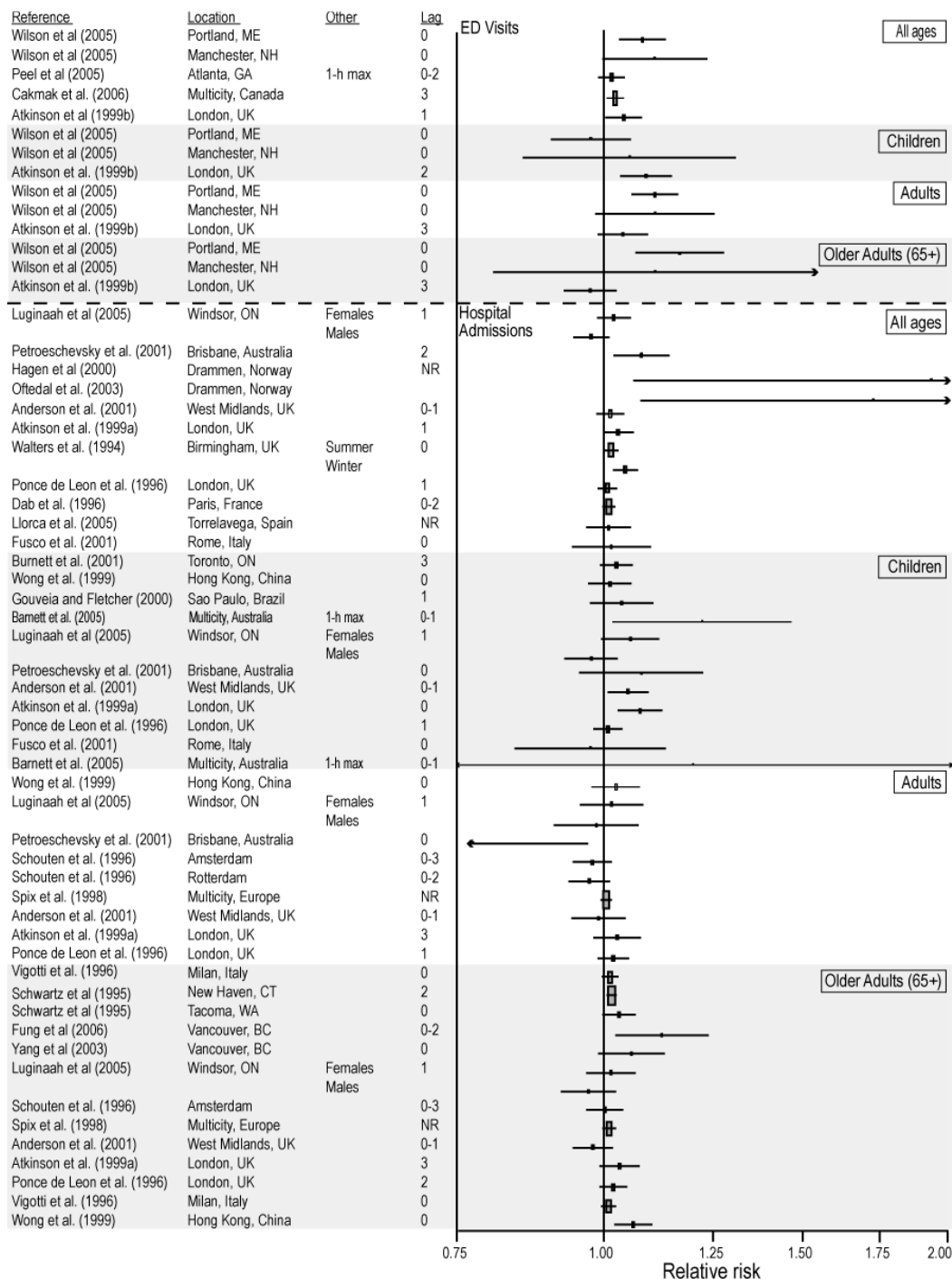


Figure 3-6. Relative risks (95% CI) of SO₂-associated emergency department visits and hospitalizations for all respiratory causes among all ages and separated by age group. Risk estimates are standardized per 10 ppb increase in 24-h avg SO₂ concentrations or 40 ppb increase in 1-h max SO₂. The size of the box of the central estimate represents the relative weight of that estimate based on the width of the 95% CI.

Wilson et al. (2005) found a positive association between ED visits and SO₂, with a 16% (95% CI: 8.0, 25.0) excess risk per 10 ppb increase in 24-h avg SO₂ at a 0-d lag, and no association in Manchester, NH when only older adults (65+ years) were considered. In another two-city study, Schwartz (1995) compared 13,740 hospital admission in New Haven, CT and Tacoma, WA from 1988–1990 with ambient SO₂ concentrations. The mean 24-h avg SO₂ concentration was 29.8 ppb (90th percentile: 159) in New Haven and 16.8 ppb (90th percentile: 74) in Tacoma. Schwartz found positive associations between hospitalizations and SO₂, with a 2% (95% CI: 1.0, 3.0) excess risk at a 2-d lag in New Haven and 3% (95% CI: 1.0, 6.0) excess risk at a 0-d lag in Tacoma per 10 ppb increase in 24-h avg SO₂. In two-pollutant models, the SO₂ effect estimate from New Haven, but not Tacoma, was found to be robust to adjustment for PM₁₀. Here, the term robust is used to indicate that there was little change in the magnitude of the central estimate, though statistical significance may have been lost. In Vancouver, BC, both Fung et al. (2006a) and Yang et al. (2003) also found positive associations between hospitalizations and SO₂. In a multipollutant model including coefficient of haze (CoH), NO₂, O₃, and CO, the SO₂ effect estimate diminished slightly (Jaffe et al., 2003).

Additional evidence of a positive association between ED visits or hospitalizations for all respiratory causes among older adults and SO₂ comes from several European (Spix et al., 1998; Sunyer et al., 2003; Vigotti et al., 1996) and Australian (Petroeschevsky et al., 2001) studies. Excess risks ranging from 1% to 12% per 10 ppb increase in 24-h avg SO₂ were reported by these studies. Petroeschevsky et al. (2001) examined 33,710 hospital admissions in Brisbane, Australia from 1987–1994. The mean 24-h avg SO₂ concentration was 4.1 ppb, and was highest in the winter months (4.8 ppb) and lowest in the spring (3.7 ppb). Petroeschevsky et al. found a 12% (95% CI: 2.0, 23.0) excess risk per 10 ppb increase in 24-h SO₂ at 0-d lag. Additional European studies did not find statistically significant associations between ambient SO₂ concentrations and hospitalizations for all respiratory causes among older adults (Anderson et al., 2001; Atkinson et al., 1999b; Ponce de Leon et al., 1996; Schouten et al., 1996).

In summary, many studies show a small, positive, though not statistically significant association between ambient SO₂ concentrations and ED visits and hospitalizations, particularly among children and older adults (65+ years). The positive evidence from these studies is supported by the results of panel, human clinical, and limited toxicological studies that also found a positive relationship between SO₂ levels and adverse respiratory outcomes.

3.1.4.6.2. Asthma

Studies of ED visits and hospitalizations provide suggestive evidence of an association between ambient SO₂ levels and ED visits and hospitalizations for asthma. The results from the hospitalization and ED studies, separated by analyses among all ages and age-specific analyses, are shown in Figure 3-7. Overall, central effect estimates in the figure range from a -10% to 40% excess risk in ED visits and hospitalizations for asthma per 10 ppb increase in 24-h avg SO₂. Most of the effect estimates are positive (suggesting an association with SO₂ and ED visits and hospitalizations for asthma), though few are statistically significant at the 95% confidence level. When all ages were included in the analyses, Wilson et al. (2005) found a positive association between ED visits and SO₂, with a 10% (95% CI: 2.0, 20.0) excess risk per 10 ppb increase in 24-h avg SO₂ at a 0-d lag in Portland, ME and a positive, though not statistically significant association in Manchester, NH. Ito et al. (2003) found a 36% (95% CI: 1.23, 1.51) excess risk in asthma ED visits per 10 ppb increase in 24-h avg SO₂, though this association was diminished once NO₂ was included in the model. A study conducted in (NY Dept of Health, 2006) found a 11% (95% CI: 6, 17) excess risk in asthma hospital admissions per 10 ppb increase in 24-h avg SO₂ for Bronx residents, but a null association for the residents of Manhattan. A study conducted in Atlanta (Peel et al., 2005) found a null relationship between asthma ED visits and 1-h max SO₂.

A study by Jaffe et al. (2003) examined the association between SO₂ and ED visits for asthma in three cities in Ohio – Cincinnati, Cleveland, and Columbus – in asthmatics aged 5 to 34 years. The mean 24-h avg SO₂ concentrations were 14 ppb (range: 1–50) in Cincinnati, 15 ppb (range: 1–64) in Cleveland, and 4 ppb (range: 0–22) in Columbus. A positive association was observed in the multicity analysis, with a 6.1% (95% CI: 0.5, 11.5) excess risk in asthma visits observed per 10 ppb increase in 24-h avg SO₂. In the city-stratified analyses, significant associations were only observed for Cincinnati (17.0% [95% CI: 4.6, 30.8]).

When analyses were stratified to include children (0-14 years) only, Wilson et al. (2005) found positive, but not statistically significant associations between ED visits and SO₂ in Portland, ME or Manchester, NH. Similarly, Lin et al. (2005) observed a weak positive association between hospitalizations for asthma and SO₂ among girls, and a null association for boys (Toronto, ON; mean 24-h avg SO₂ of 5.36 ppb [SD 5.90]). Stronger evidence comes from a study of childhood asthma hospitalizations conducted in Bronx County, New York (Lin et al.,

2003b). In this study, the authors conducted a case-control study of children aged 0-14 years and examined the association of daily ambient SO₂ concentrations (categorized into quartiles of both average and max levels) and cases admitted to the hospital for asthma or controls who were admitted for reasons other than asthma. The mean 24-h avg SO₂ was below 17 ppb for both cases and controls across all lag days examined. The authors found that cases were exposed to higher 24-h avg SO₂ than controls. When the highest exposure quartile was compared with the lowest, the ORs were strongest when a 3-day lag was employed (OR 2.16 [95% CI: 1.77, 2.65] for 24-h avg SO₂; OR 1.86 [95% CI: 1.52, 2.27] for 1-h max SO₂). The results were positive and statistically significant for all lag days examined. These results suggest a consistent positive association between SO₂ exposure and hospitalizations for childhood asthma.

Additional evidence of a positive association between ED visits or hospitalizations for asthma and SO₂ comes from several European (Anderson et al., 1998; Atkinson et al., 1999a; Hajat et al., 1999; Sunyer et al., 1997; 2003; Thompson et al., 2001) and Asian (Lee et al., 2002) studies. Excess risks ranging from 2% to 10% per 10 ppb increase in 24-h avg SO₂ were reported by these studies. Several of these studies observed that the SO₂ effect estimate was robust to adjustment for BS and NO₂ (Anderson et al., 1998; Sunyer et al., 1997), but one study observed that the SO₂ effect diminished considerably with adjustment for PM₁₀ and benzene (Thompson et al., 2001). Atkinson et al. (1999b) compared 165,032 hospital admissions in London from 1992–1994 with ambient SO₂ levels (mean 24-h avg of 7.2 ppb [SD 4.7]). They found a 10% (95% CI: 4.0, 16.0) excess risk per 10 ppb increase in 24-h avg SO₂ at 1-d lag. Additional European (Fusco et al., 2001), Australian (Barnett et al., 2005; Petroeschevsky et al., 2001), Asian (Ko et al., 2007; Lee et al., 2006) and Latin American (Gouveia and Fletcher, 2000) studies did not find statistically significant associations between ambient SO₂ concentrations and hospitalizations for all respiratory causes among children.

In summary, small, positive associations were observed between ambient SO₂ concentrations and ED visits and asthma hospitalizations. Evidence from these studies is further supported by the results of panel and human clinical studies that have also found SO₂-related respiratory effects in asthmatics.

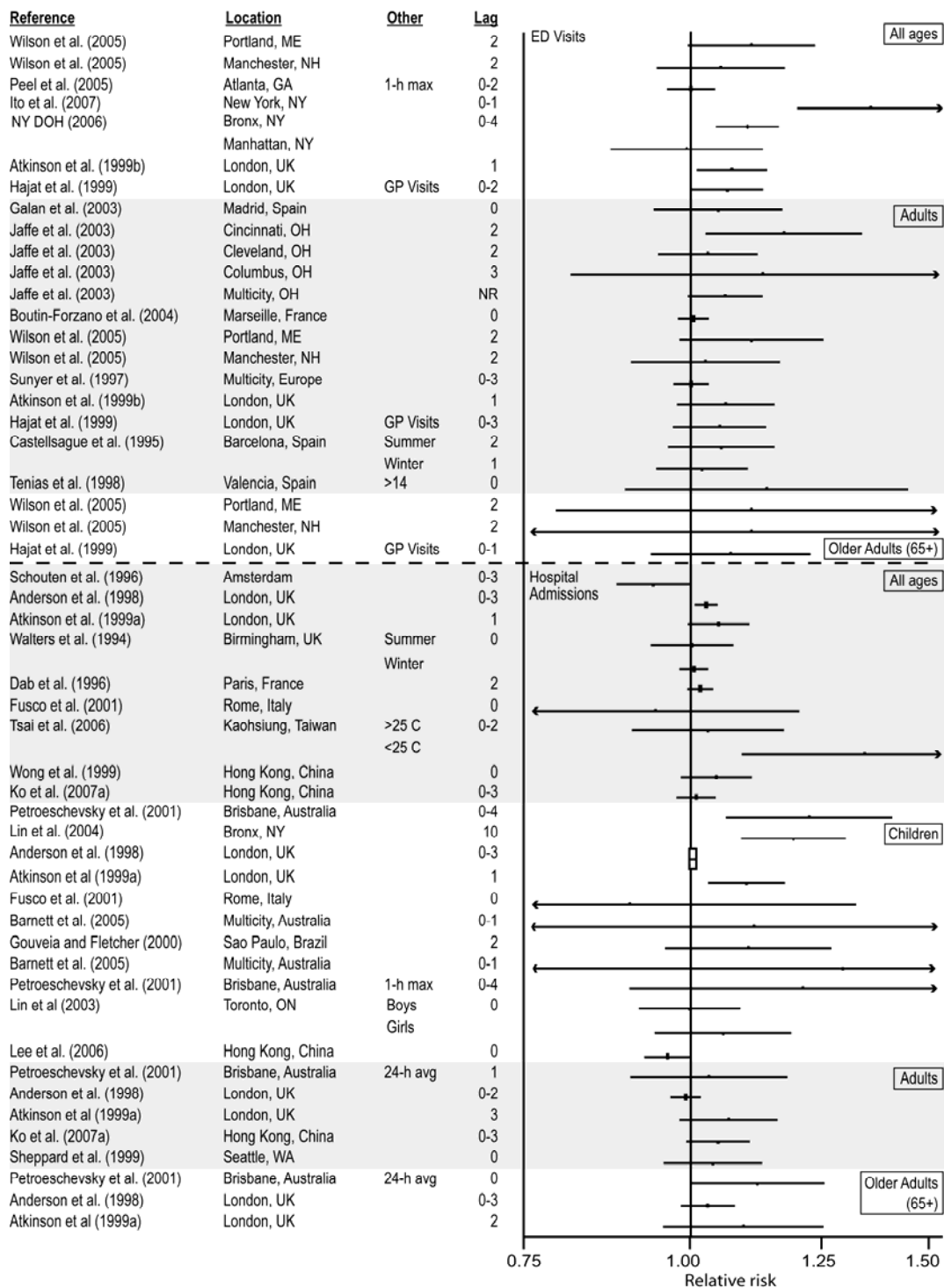


Figure 3-7. Relative risks (95% CI) of SO₂-associated emergency department visits and hospitalizations for asthma among all ages and age-specific groups. Risk estimates are standardized per 10 ppb increase in 24-h avg SO₂ concentrations or 40 ppb increase in 1-h max SO₂. The size of the box of the central estimate represents the relative weight of that estimate based on the width of the 95% CI.

3.1.4.6.3. Chronic Obstructive Pulmonary Disease

There are relatively few studies that have examined the association of ED visits and hospitalizations for COPD and ambient SO₂ levels, and very little evidence that an association exists. A recent study (Ko et al., 2007) found a significant association between hospital admissions for COPD (not including asthma) in Hong Kong (1.8% [95% CI: 0.3, 3.8]) excess risk per 10 ppb increase in 24-h avg SO₂ concentration). Three additional studies reported positive and statistically significant results for COPD and SO₂; all three studies included asthma in their diagnostic definition of COPD (Anderson et al., 2001; Moolgavkar, 2003; Sunyer et al., 2003). Anderson et al. (2001) reported a 12% (95% CI: 5.0, 20.0) excess risk per 10 ppb increase in 24-h avg SO₂ among children, while Moolgavkar (2003) and Sunyer et al. (2003) found 5% and 2% excess risks per 10 ppb increase in 24-h avg SO₂ among older adults populations, respectively. Other studies examining COPD did not find statistically significant results (Atkinson et al., 1999b; Burnett et al., 1999; Michaud et al., 2004).

Overall, this limited and inconsistent evidence does not support a relationship between ED visits and hospitalizations for COPD and ambient SO₂ levels.

3.1.4.6.4. Respiratory Diseases Other than Asthma or COPD

Studies of ED visits or hospital admissions for other respiratory diseases looked at several other specific outcomes. There are limited studies with mixed results for upper respiratory tract infections (Burnett et al., 1999; Hajat et al., 2002; Lin et al., 2005; Peel et al., 2005), pneumonia (Barnett et al., 2005; Moolgavkar et al., 1997; Peel et al., 2005), bronchitis (Barnett et al., 2005; Michaud et al., 2004), and allergic rhinitis (Hajat et al., 1999; Villeneuve et al., 2006). The limited evidence is suggestive of an association between SO₂ levels and ED visits for lower respiratory tract diseases (Atkinson et al., 1999b; Farhat et al., 2005; Hajat et al., 1999; Lin et al., 1999; Martins et al., 2002). All of the studies that characterized this relationship found a positive and statistically significant excess risk associated with increases in SO₂. Excess risks ranging from 3% to 33% per 10 ppb increase in 24-h avg SO₂ were reported by these studies.

In summary, few studies provide results with mixed respiratory health outcomes other than asthma and COPD. This makes it difficult to draw conclusions about the effects of SO₂ on these diseases. Limited evidence does exist to support a suggestive association between ambient SO₂ levels and ED visits for lower respiratory tract diseases.

3.1.4.6.5. Summary of Evidence on Emergency Department Visits and Hospitalizations for Respiratory Diseases

Small, positive associations exist between ambient SO₂ concentrations and ED visits and hospitalizations for all respiratory causes, particularly among children and older adults (65+ years), and for asthma, though not always statistically significant. The SO₂-related changes in ED visits or hospital admissions for respiratory causes ranged from -5% to 20% excess risk, with the large majority of studies suggesting an increase in risk. No association was observed between SO₂ levels and ED visits and hospitalizations for COPD. Given the limited number of studies with mixed results, it is difficult to draw conclusions about the effect of SO₂ on other respiratory diseases, though studies of lower respiratory tract diseases are somewhat suggestive of an association.

Multipollutant regression analyses indicate that SO₂ risk estimates, in general, are not sensitive to the inclusion of copollutants, including O₃ (Anderson et al., 1998; Hajat et al., 1999; Yang et al., 2003; 2005), PM (Hagen et al., 2000; Lin et al., 2003; 2005; Schwartz, 1995) CO (Farhat et al., 2005) and NO₂ (Anderson et al., 1998; Lin et al., 2004; Sunyer et al., 1997). Figure 3-8 presents SO₂ excess risk estimates with and without adjustment for various copollutants. PM and NO₂ are the main foci, since these pollutants have been found to be highly-correlated with SO₂ in epidemiological studies and have known respiratory health effects. Although the studies showed that copollutant adjustment had varying degrees of influence on the SO₂ effect estimates, the effect of SO₂ on respiratory health outcomes appears to be generally robust and independent of the effects of ambient particles or other gaseous copollutants.

The results of several studies (Anderson et al., 1998; Hajat et al., 1999; Schouten et al., 1996; Spix et al., 1998; Wong et al., 1999) have demonstrated a greater increase in ED visits and hospitalizations for respiratory illnesses during the summer months, despite the fact that the average concentrations for SO₂ in some of areas studied were greatest in winter. In contrast, some studies found the associations between ED visits and hospital admissions and respiratory disease with similar increases in SO₂ to be greater in winter than summer (Vigotti et al., 1996; Walters et al., 1994). Other studies were unable to discern a seasonal difference in ED visits and hospitalizations for respiratory causes (Castellsague et al., 1995; Tenias et al., 1998; Wong et al., 2002). These effects were not consistent across age groups. Warmer months were more likely to show evidence of an association with adverse respiratory outcomes in children, while older

- 1 adults appeared more likely to be affected during the cooler months. These seasonal associations
- 2 remain somewhat uncertain and require additional investigation.

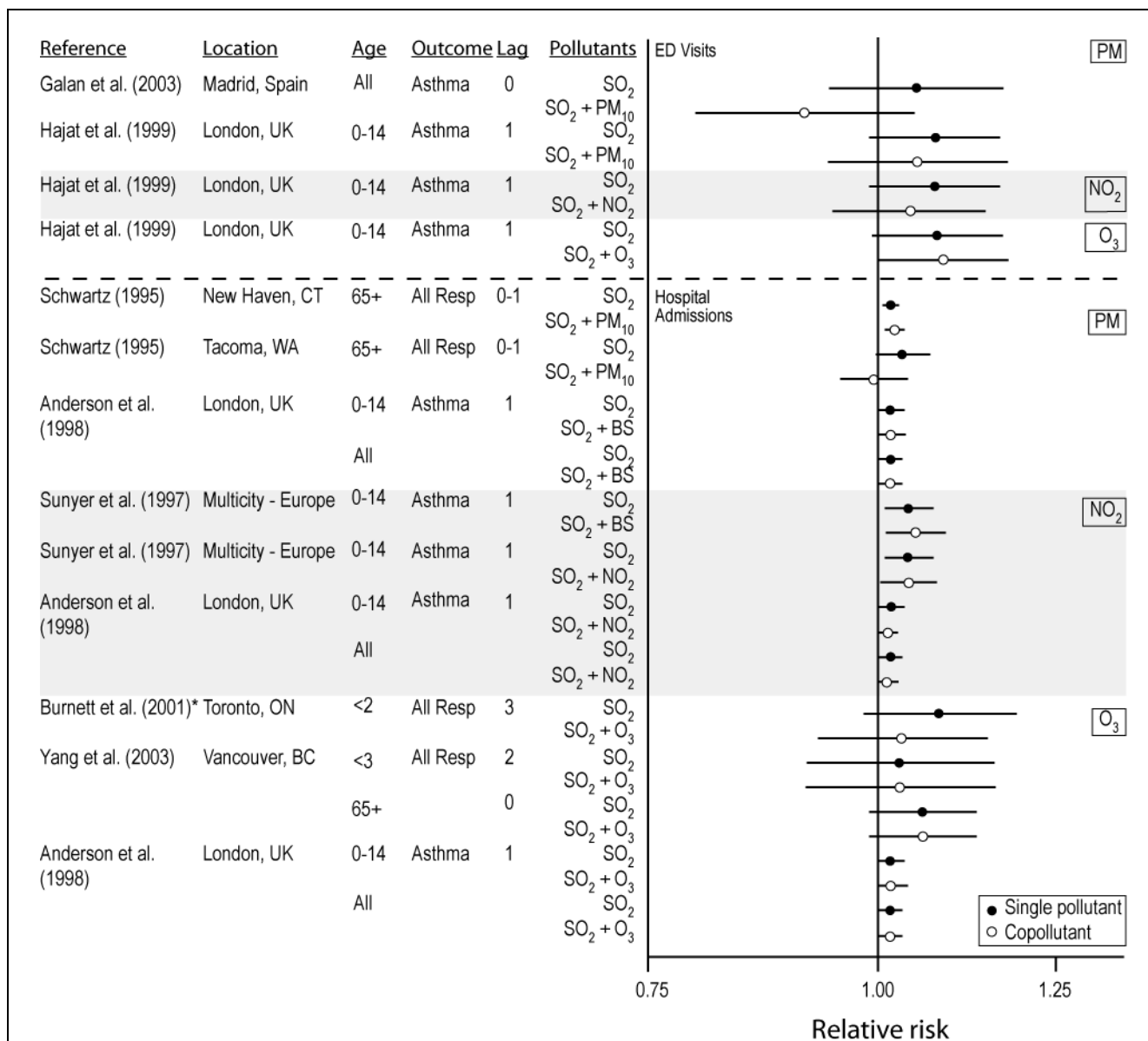


Figure 3-8. Relative risks (95% CI) of SO₂-associated emergency department visits and hospitalizations for all respiratory causes and asthma, with and without copollutant adjustment. Risk estimates are standardized per 10 ppb increase in 24-h avg SO₂ concentrations or 40 ppb increase in 1-h max SO₂. In Burnett et al. (2001), analyses were performed using default convergence criteria for Poisson GAM with a nonparametric LOESS prefilter applied to air pollution and hospitalization data.

1 In conclusion, a large number of epidemiologic studies provide evidence of an association
2 between ambient SO₂ concentrations and ED visits and hospitalizations for all respiratory causes,
3 in particular among children and older adults (65+ years), and for asthma. The findings are
4 generally robust when additional copollutants are included in the model. These associations are
5 supported by panel studies that observed SO₂-related increases in asthma and other respiratory
6 symptoms in children, and human clinical and animal toxicological studies that found a positive
7 relationship between SO₂ exposure and various respiratory outcomes.

3.1.4.7. Summary of Evidence on the Effect of Short-Term (≥ 1 h) Exposure on Respiratory Health

8 Numerous epidemiological studies have observed associations between short-term (≥ 1-h,
9 generally 24-h avg) exposure to SO₂ and respiratory health effects, ranging from respiratory
10 symptoms to ED visits and hospital admissions for respiratory causes. The associations between
11 ambient SO₂ concentrations and several respiratory outcomes were generally consistent, with the
12 large majority of studies showing positive associations, and multicity studies, as well as several
13 single-city studies, indicating statistically significant findings. The respiratory effects related to
14 short-term exposure to SO₂ found in animal toxicological studies, and to a more limited extent
15 the human clinical studies, provide coherence and biological plausibility for the observed
16 epidemiological associations. The causal effects of peak exposure to SO₂ on respiratory health
17 found in the human clinical studies (see Section 3.1.3.5) provide further evidence of biological
18 plausibility for the effects associated with short-term exposure to SO₂.

19 Two recent multicity studies (Mortimer et al., 2002; Schildcrout et al., 2006) and several
20 other studies (Delfino et al., 2003; Neas et al., 1995; van der Zee et al., 1999) have found an
21 association between short-term ambient SO₂ concentrations and respiratory symptoms in
22 children. In the limited number of studies that assessed potential confounding by copollutants
23 using multipollutant models, the SO₂ effect on respiratory symptoms was generally found to be
24 robust to adjustment for copollutants. These findings provide supportive evidence for an
25 association between short-term exposure to ambient SO₂ exposure and respiratory symptoms in
26 children, particularly those with asthma. Several recent studies (Desqueyroux et al., 2002a;
27 2002b; van der Zee et al., 2000) found no association between ambient SO₂ levels and
28 respiratory symptoms in adults, though there was limited epidemiological evidence which

suggested that atopic adults as well as children may be at increased risk for SO₂-induced respiratory symptoms (Boezen et al., 1999; 2005).

Animal toxicological studies in guinea pigs showed changes in lung function immediately following 1 ppm SO₂ exposure (Amdur et al., 1983). Guinea pigs, as a species, are typically more sensitive to air pollution than other laboratory animals and, thus, may provide a better model for characterizing the effects of air pollutants on lung function. Epidemiological studies do not provide strong evidence of an association between short-term ambient SO₂ exposure and lung function in either children (Mortimer et al., 2002; Roemer et al., 1998) or adults (e.g., Peters et al., 1996; Taggart et al., 1996). Several other studies reported positive results; however, the generally mixed findings, as well as the relative lack of evidence available to evaluate potential confounding by copollutants, limits the causal interpretation of ambient SO₂ on lung function.

Only one epidemiological study (Adamkiewicz et al., 2004) evaluated inflammation, as indexed by eNO, and found no association with SO₂ exposure. Animal toxicological studies found that repeated exposure to near ambient levels of SO₂ leads to increased airway inflammation in two models involving animals which were sensitized to an antigen (Park et al., 2001; Li et al., 2007). Studies of other ambient pollutants indicate that influx of macrophages and other inflammatory cells, with the related release of inflammatory cytokines, is a common response to — and may further contribute to — injury.

Effects of short-term exposure to SO₂ on AHR have been observed. In two animal toxicological studies, repeated exposure to 0.1 ppm SO₂ led to AHR in guinea pigs sensitized to an antigen (Riedel et al., 1988; Park et al., 2001). Human clinical studies by Devalia et al. (1994) and Rusznak et al. (1996) demonstrated increased sensitivity to an inhaled allergen in asthmatic subjects following exposure to a combination of SO₂ (0.2 ppm) and NO₂ (0.4 ppm). This effect was not observed following exposure to either SO₂ or NO₂ alone. These findings of increased pulmonary resistance are in concordance with the limited epidemiological findings of SO₂-induced AHR (Taggart et al., 1996).

Epidemiological studies provide suggestive evidence for an association between ambient SO₂ levels and ED visits and hospitalizations for all respiratory diseases in two susceptible populations: children (Dab et al., 1996; Petroschevsky et al., 2001; Walters et al., 1994) and older adults (65+ years) (Fung et al., 2006; Schwartz, 1995; Spix et al., 1998; Wong et al., 1999). Evidence for an association between ambient SO₂ levels and these outcomes in adults was less

1 consistent. A modest association between ambient SO₂ and ED visits and hospitalizations for
2 asthma was also suggested. SO₂ effect estimates were generally robust to the inclusion of
3 copollutants, including PM, O₃, CO and NO₂, indicating that the observed effects of SO₂ on
4 respiratory endpoints is independent of the effects of other ambient air pollutants.

3.1.5. Mixtures and Interactive Effects

3.1.5.1. Evidence from Human Clinical Studies

5 The interaction of SO₂ with other common air pollutants or the sequential exposure of SO₂
6 after prior exposure to another pollutant can potentially modify SO₂-induced respiratory effects.
7 However, only a few human clinical studies have looked at the interactive effects of coexisting
8 ambient air pollutants. In a human clinical study designed to simulate an ambient “acid summer
9 haze,” Linn et al. (1997) exposed healthy and asthmatic children (9–12 years of age) for 4 h with
10 intermittent exercise to a mixture of SO₂ (0.1 ppm), H₂SO₄ (100 µg/m³), and O₃ (0.1 ppm).
11 Compared with exposure to filtered air, exposure to the air pollution mixture did not result in
12 significant changes in lung function or respiratory symptoms. These findings are in agreement
13 with a series of similar studies conducted by Kleinman et al. (1981; 1984; 1985).

14 In a human clinical study of asthmatic adolescents (12- to 16-years-old), Koenig et al.
15 (1983) evaluated changes in FEV₁ following a 10-min exposure during moderate exercise to
16 1-mg/m³ NaCl alone and in combination with 0.5 and 1.0 ppm SO₂. Significant decreases of 15
17 and 23% were reported in FEV₁ following exposure to 1 mg/m³ NaCl in combination with 0.5-
18 and 1.0-ppm SO₂, respectively. No significant changes in FEV₁ were observed between pre- and
19 post-exposure to 1-mg/m³ NaCl without SO₂. The effect observed in this study may be the result
20 of the presence of hygroscopic particles that can carry SO₂ deeper into the lung.

21 Koenig et al. (1990) also examined the effect of 15-min exposures to 0.1 ppm SO₂ in
22 adolescent asthmatics engaged in moderate levels of exercise. Immediately preceding this
23 exposure, subjects were exposed for 45 min to 0.12 ppm O₃ during intermittent moderate
24 exercise. Subjects also underwent two additional exposure sequences with the same exercise
25 regimen: 15-min exposure to 0.1 ppm SO₂ following a 45-min exposure to clean air, and 15-min
26 exposure to 0.12 ppm O₃ following a 45-min exposure to 0.12 ppm O₃. The authors found that
27 the change in FEV₁ compared to baseline was significantly different following the O₃-SO₂

1 exposure (8% decrease) when compared to the change following the air-SO₂ or O₃-O₃ exposures
2 (decreases of 3 and 2%, respectively). In a more recent study, Trenga et al. (2001) reported that
3 among adult asthmatics, exposure to O₃ (0.12 ppm for 45 min) resulted in a slight increase in
4 bronchial responsiveness to SO₂ at a concentration of 0.25 ppm (6.5% decrease in FEV₁ with
5 pre-exposure to O₃, compared with a 3.4% decrease in FEV₁ with pre-exposure to filtered air).
6 Hazucha and Bates (1975) demonstrated a synergistic effect of concurrent exposure to SO₂
7 (0.37 ppm) and O₃ (0.37 ppm) on lung function in healthy asthmatics; however, no such effect
8 was observed in a similar study conducted by Bedi et al. (1979).

9 Jörres and Magnussen (1990) and Rubinstein et al. (1990) investigated the effects of a
10 prior NO₂ exposure on SO₂-induced bronchoconstriction in asthmatic adults. While Jörres and
11 Magnussen suggested that prior exposure to NO₂ increased the responsiveness to SO₂,
12 Rubinstein et al. did not find that NO₂ exacerbated the effects of SO₂. Linn et al. (1980) reported
13 no difference in lung function or respiratory symptoms among a group of exercising asthmatics
14 exposed to both clean air and a combination of NO₂ (0.5 ppm) and SO₂ (0.3 ppm).

15 In summary, although findings from some human clinical studies suggest that respiratory
16 effects of exposure to SO₂ may be enhanced when preceded by, or occurring concomitant with
17 exposure to other air pollutants, this evidence is quite limited and inconsistent.

3.1.5.2. Evidence from Animal Toxicological Studies

18 As discussed earlier, SO₂ is a component of complex air pollution mixtures that vary
19 geographically and temporally (e.g., by hour, week, and season). Depending on atmospheric
20 conditions, SO₂ can be transformed to secondary sulfate particles and acid aerosols (H₂SO₄) and
21 can adsorb onto particulate matter. Since SO₂, H₂SO₄ and PM share a common source—fossil
22 fuels—health effects of fossil-fuel derived air pollution mixtures may be determined by
23 interactions among individual components. Although epidemiological studies provide
24 information on real-world exposure, it is difficult to evaluate causative factors and quantitative
25 relationships from such studies. Animal studies are therefore useful in evaluating health effects
26 of mixtures. The studies discussed below demonstrate important interactions between SO₂ and
27 other air pollution components.

28 An informative study of complex air pollutants was conducted in dogs and addressed in the
29 1982 AQCD (EPA, 1982). In dogs that were exposed to SO₂ and H₂SO₄, with or without

1 irradiated or non-irradiated auto exhaust concentrations relevant to urban exposures, functional
2 lung changes were observed at 61 months of exposure and at 2 years after exposures ended.
3 Morphological and biochemical changes were observed at 2.5–3 years after exposure.

4 Since then, studies have demonstrated respiratory responses following inhalation of SO₂
5 which had been layered onto metal or carbon particles. The resulting particles were submicron in
6 size; they would be expected to deposit in the lower respiratory tract. As discussed in Section
7 3.2.2, chemosensitive receptors are present at all levels of the respiratory tract and are known to
8 activate reflexes involving the respiratory and cardiovascular systems. It has been postulated that
9 bronchial C-fiber receptors are more sensitive to chemical irritants than pulmonary C-fibers
10 receptors, but that more intense cardiovascular responses are triggered by the pulmonary
11 receptors (Coleridge and Coleridge, 1994; Widdicombe and Lee, 2001). Some of the work
12 involving SO₂ layered onto particles has been reviewed in the PM AQCD (EPA, 1996; 2004).
13 Important studies are described briefly to show biological plausibility for the health effects of
14 SO₂ which rarely, if at all, exists in nature in the absence of PM. This work is summarized in
15 Annex Tables F-15 through 5-1.

3.1.5.2.1. Effects of SO₂ Layered on Metallic Particles

16 Studies examining interactions between SO₂ and metallic or carbonaceous particles are
17 summarized in Annex Table E-14. Metal oxides may be released into the atmosphere with SO₂
18 during combustion of fossil fuels or by smelting operations (Lam et al., 1982). The 1982 AQCD
19 noted that sorption of SO₂ onto liquid or solid particles, which may act as carriers, tended to
20 increase its potency, but the mechanism for the effect was not known. The studies discussed
21 below strongly suggest that SO₂ adsorbed to particles penetrates to more distal regions of the
22 lung, compared with gaseous SO₂. In addition, SO₂ which is adsorbed to particles can transform
23 to sulfite, sulfate, H₂SO₄ and sulfur trioxide.

24 In an early study, guinea pigs were exposed to submicron zinc oxide aerosols alone or to
25 0.8-6 mg/m³ zinc oxide in combination with 1–2 ppm SO₂ for 3 h (Lam et al., 1982). The higher
26 concentrations of zinc oxide served to carry more sulfur in the aerosol (Amdur et al., 1988).
27 Animal exposure to zinc oxide alone resulted in a decrease in functional residual capacity;
28 animal exposure to the combination of zinc oxide and SO₂ resulted in dose-dependent decreases
29 in total lung capacity, vital capacity, functional residual capacity, residual volume, diffusion

1 capacity for carbon monoxide and alveolar volume (Amdur et al., 1988; Lam et al., 1982).
2 Another study exposed guinea pigs to ~1 ppm SO₂ and 1–2 mg/m³ zinc oxide for 1 h (Amdur et
3 al., 1983). Results demonstrated the formation of sulfite, sulfate, and sulfur trioxide under
4 conditions of high humidity and temperature, and greater than additive decrements in pulmonary
5 function, compared with exposures to SO₂ or zinc oxide alone or with a mixture of SO₂ and zinc
6 oxide where no transformation had taken place. Additional studies in guinea pigs involved 1 h
7 exposures to ~1 ppm SO₂ and 1-3 mg/m³ copper oxide (Chen et al., 1991) and ~1ppm SO₂ and
8 1-3 mg/m³ zinc oxide (Chen et al., 1992). In the copper oxide-SO₂ study, components were
9 mixed at either 37°C or 1,411°C prior to exposure (Chen et al., 1991). Results demonstrated
10 increased pulmonary resistance when the compounds were mixed at low temperature, leading to
11 the formation of sulfite; when components were mixed at high temperature, leading to the
12 formation of sulfate, no increase was found. This suggested that sulfite has greater biological
13 effect than sulfate. The zinc oxide-SO₂ study found a synergistic interaction between zinc oxide
14 and SO₂. Co-exposure, but not exposure to either component alone, led to airway
15 hyperresponsiveness following an acetylcholine challenge (Chen et al., 1991). A further study by
16 these same investigators determined that H₂SO₄ was the predominant species of sulfur associated
17 with the zinc oxide particles mixed with SO₂ under high temperature and humidity conditions
18 (Amdur et al., 1988). H₂SO₄ has a valence of S(VI), unlike that of SO₂ and sulfite, which have a
19 valence of S(IV). This study also correlated increased bronchial sensitivity to acetylcholine with
20 the formation of H₂SO₄ at 21 and 30 µg/m³ in animals exposed for 1 h to this mixture. The
21 authors concluded that the response might have been enhanced by H₂SO₄ being carried on the
22 metal particles. Furthermore, the H₂SO₄-coated zinc oxide particle was five- to tenfold more
23 potent in eliciting a response than H₂SO₄ alone.

24 Several subacute studies were also conducted. One involved the exposure of guinea pigs
25 for 3 h/day for 6 consecutive days to 6 mg/m³ submicron zinc oxide particles generated in a
26 humid furnace mixed with 1 ppm SO₂ or to SO₂ alone (Conner et al., 1985). Exposure to the zinc
27 oxide/SO₂ particles resulted in significant decreases in total lung capacity, vital capacity and
28 functional residual capacity compared with controls or exposure to SO₂ alone. These decreases
29 were maintained for at least 72 h following the last exposure. Decreases in diffusing capacity for
30 carbon monoxide and alveolar volume, as well as increased alveolar-duct inflammation, were
31 observed in response to zinc oxide-SO₂ but not to SO₂ alone. Because the changes noted in

response to zinc oxide-SO₂ were identical to those seen in a previous study using zinc oxide, the authors concluded that SO₂ played a minor role in these responses. Subsequent studies employed a lower concentration of zinc oxide (1 and 2.5 mg/m³) in order to unmask the effects of the SO₂-zinc oxide interaction (Amdur et al., 1988; Conner et al., 1989). In one of these, significant changes were observed in numbers of cells and other lavage fluid components in guinea pigs exposed for 3 h/day for 5 days to SO₂-zinc oxide compared with those exposed only to SO₂ or zinc oxide. As previously described, it was determined that the mixture of zinc oxide and SO₂ in a high temperature furnace with sufficient humidity resulted in the formation of a zinc oxide/H₂SO₄ aerosol. Exposure of guinea pigs for 3 h/day for 5 days to this H₂SO₄-coated ultrafine aerosol at a concentration of 7-11 µg/m³ S(VI) resulted in a significant decrease in diffusing capacity for carbon monoxide, an effect not seen with exposure to zinc oxide or SO₂ alone. The authors concluded that the response was due to the H₂SO₄ associated with the aerosol and delivered to the lower respiratory tract (Amdur et al., 1988).

3.1.5.2.2. Effects of SO₂ Layered on Carbon Particles

Effects of SO₂-containing mixtures on host defenses were examined in several studies. Host defense results will not be discussed since high concentrations (10 ppm) of SO₂ were used. However, these experiments were important since they demonstrated that mixing SO₂ and submicron particles of carbon black in the presence of 85% relative humidity led to significant adsorption of SO₂ onto the carbon black and oxidation of SO₂ to acid sulfate. No high temperature furnace was employed in these studies (Jakab et al., 1996; Clarke et al., 2000).

3.1.5.2.3. Effects of Sulfite Aerosols

Several studies used sulfite particles as a surrogate for SO₂ adsorbed onto carbonaceous or metallic particulates. These are discussed below and in Annex Table E-15. Sulfite and SO₂, both S(IV) species, were expected to have similar chemical reactivities. An acute 1-h exposure of guinea pigs to submicron aerosols of sodium sulfite (204-1,152 µg/m³ of sulfite) resulted in significant effects on pulmonary function (Chen et al., 1987). A 50% increase in resistance and 19% decrease in compliance were observed using 972 µg/m³ of sulfite, while dose-dependent decreases in total lung capacity, vital capacity, functional residual capacity, residual volume, diffusion capacity for carbon monoxide and increases in lung wet weight at concentrations of

204 $\mu\text{g}/\text{m}^3$ and above. The authors noted that aerosols of S(IV) were 6 times more potent than gaseous S(IV) in terms of bronchoconstriction and attributed these effects to greater pulmonary deposition of the S(IV) aerosol. An earlier study by the same group found that 1-h exposure of rabbits to greater than 1,200 $\mu\text{g}/\text{m}^3$ of sulfite led to accelerated clearance of a tracer aerosol from the bronchial tree (Chen and Schlesinger, 1983). Chronic studies involving sulfite aerosols are discussed in Section 3.4.2.5.

3.1.5.2.4. Other Mixtures

In addition to examining the interaction of SO_2 and particles, animal studies performed since the publication of the 1982 AQCD evaluated the effects of binary mixtures, laboratory-generated complex mixtures (e.g., simulation of regional air pollution), or actual ambient air mixtures. Annex Tables E-17 through E-20 summarize results from short-term studies on possible toxicity relationships between SO_2 and O_3 , and between SO_2 and sulfates as well as the effects of complex air pollution mixtures in healthy animals and disease models. Possible interactions between SO_2 and cold air were also examined (Annex Table E-20). Generally, most studies with ambient or laboratory-generated complex mixtures did not include a SO_2 -only exposure group, making it difficult to determine the contribution of SO_x . No definitive conclusions can be made from these studies.

3.1.5.2.5. Summary of Evidence on SO_2 Interactions with PM and Other Mixtures

The key findings by Amdur and others discussed above demonstrate that the effects of SO_2 may be enhanced when aerosol particles act as carriers and deliver SO_2 to the lower respiratory tract. Interaction of SO_2 and PM may also lead to transformation of SO_2 to other forms of SO_x which may have more potent biological effects than SO_2 . Studies discussed above reported transformation of SO_2 adsorbed onto metal oxide or carbon particles to sulfite, sulfate, sulfur trioxide, and H_2SO_4 depending on conditions of temperature and relative humidity.

3.1.6. Evidence of the Effect of SO_2 on Respiratory Morbidity from Intervention Studies

Many epidemiological studies have examined the association of short-term SO_2 concentrations and various respiratory morbidity outcomes. These studies collectively suggest

1 that increased ambient SO₂ concentrations are associated with increased risk of respiratory
2 outcomes, ranging from respiratory symptoms to ED visits and hospitalizations. Further
3 contributing to the evidence base are intervention studies that reported decreases in respiratory
4 morbidity following improvements in air quality, particularly reductions in SO₂ concentrations.

5 In Hong Kong, a sudden change in regulation in July 1990 required all power plants and
6 road vehicles to use fuel oil with a sulfur content of $\leq 0.5\%$ by weight. These regulations were
7 enforced quickly, and provided opportunities to observe changes in morbidity before and after
8 the intervention. Peters et al. (1996) followed 3,521 children (mean age 9.5 years) residing in
9 two districts with good and poor air quality before the intervention from 1989 to 1991. The
10 intervention resulted in large reductions in SO₂ (up to 80% in polluted district), along with a
11 modest reduction in sulfate (38% in polluted district). Only a small change in TSP levels was
12 observed after the intervention (15% decline in polluted district). In 1989 and 1990, an excess
13 risk of respiratory symptoms was observed in the polluted district. After the intervention, there
14 was a greater decline in reported symptoms of cough, sore throat, phlegm, and wheezing in the
15 polluted compared with the unpolluted district. For example, the OR for cough, comparing the
16 polluted to the unpolluted district, was 1.22 (95% CI: 1.05, 1.42) in 1989 and 1990, and
17 decreased to 0.92 (95% CI: 0.73, 1.15) in 1991.

18 A study by Keles et al. (1999) evaluated the prevalence of chronic rhinitis among high
19 school students before and after installation of a natural gas network for domestic heating and
20 industrial works, in a polluted area of Istanbul, Turkey. Concentrations of CO, NO₂, and
21 hydrocarbons were relatively low compared to SO₂ and TSP in this area. After the intervention,
22 the annual mean TSP concentration declined by 23% from 89.7 $\mu\text{g}/\text{m}^3$ to 68.8 $\mu\text{g}/\text{m}^3$. An even
23 greater decline (46%) was observed for SO₂, from an annual mean of 70.8 ppb to 38.2 ppb. The
24 prevalence of rhinitis decreased significantly from 62.5% to 51% of the student population ($p <$
25 0.05) following the installation of the natural gas network. Symptoms of rhinitis were associated
26 with air pollution levels, but not with any of the other factors considered, including sex,
27 household crowding, heating source, and smoking status. Although the effects from TSP could
28 not be separated from SO₂ effects, this study demonstrated that reductions in both pollutants
29 (with greater declines in SO₂) resulted in significant reductions in the prevalence of chronic
30 rhinitis in a highly polluted area.

Another study in Germany observed that reductions in air pollutant levels were associated with improvement in reported respiratory symptoms. Heinrich et al. (2002) examined the influence of reduced air pollution levels on respiratory symptoms in children aged 5 to 14 years ($n = 7,632$). Questionnaires were collected from the children during 1992–1993, 1995–1996, and 1998–1999 in three study areas. During the study period, SO_2 concentrations decreased by more than 90% and TSP concentrations decreased by approximately 60%. Concentrations of nucleation-mode particles (10–30 nm) increased during this time period. For most respiratory outcomes, the prevalence continued to decline in each of the three surveys. The temporal changes followed similar trends in all three study areas. Stronger effects between SO_2 and prevalence of respiratory symptoms were observed among children without indoor exposures. For those without indoor exposures, ORs of 1.21 (95% CI: 1.11, 1.32) were observed for prevalence of bronchitis and 1.11 (95% CI: 1.02, 1.22) for frequent colds per 5-ppb increase in the annual mean of SO_2 . Frye et al. (2003) reported changes in lung function parameters associated with declines in SO_2 concentrations in 2,493 children during this period as well. The researchers observed a 0.6% (95% CI: 0.1, 1.2) increase in FVC and a 0.4% (95% CI: -0.1, 0.9) increase in FEV_1 per 5-ppb decrease in the annual mean of SO_2 . They concluded that the decreasing prevalence of respiratory symptoms and the increase in lung function following decreases in air pollution levels might indicate the reversibility of adverse health effects in children.

In summary, these studies observed that improvements in air quality, in particular large decreases in SO_2 concentrations, were associated with improvements in respiratory symptoms and lung function. However, the decreased respiratory morbidity following large reductions in ambient SO_2 concentrations does not preclude the possibility that other constituents of the pollution mixture that share the same source as SO_2 are also responsible for adverse effects. In the German and Turkey studies, both SO_2 and TSP concentrations decreased dramatically. Although PM_{10} levels before and after the intervention were stable in Hong Kong, large reductions in ambient nickel and vanadium were observed concomitantly with reductions of sulfur after the intervention (Hedley et al., 2006). As discussed in Section 3.1.5, interactions of SO_2 and PM may lead to transformation of SO_2 to other forms of SO_x which have more potent biological effects; thus the improvements in respiratory health may also be attributable to both declines in SO_2 and PM. Nonetheless, considered collectively with the larger body of evidence

from epidemiological, human clinical, and animal toxicological studies, these studies are supportive of SO₂-related effects on respiratory morbidity.

3.1.7. Summary of Evidence of the Effect of Short-Term SO₂ Exposure on Respiratory Health

Evaluation of the health evidence led to the conclusion that it is *sufficient to infer a causal relationship between respiratory morbidity and short-term exposure to SO₂*. This conclusion is supported by the consistency, coherence, and plausibility of findings observed in human clinical studies with 5-10 min exposures, epidemiological studies using largely 24-h avg exposures and animal toxicological studies using exposures of minutes to hours.

The strongest evidence for this causal relationship comes from human clinical studies reporting respiratory symptoms and decreased lung function following peak exposures of 5-10 min duration to SO₂. These effects have been observed consistently across studies involving exercising mild to moderate asthmatics. Statistically significant decrements in lung function accompanied by respiratory symptoms including wheeze and chest tightness have been clearly demonstrated following exposure to 0.4-0.6 ppm SO₂. Although studies have not reported statistically significant respiratory effects following exposure to 0.2-0.3 ppm SO₂, some asthmatic subjects (5-20%) have been shown to experience moderate to large decrements in lung function at these exposure concentrations. A larger body of evidence supporting this determination of causality comes from numerous epidemiological studies reporting associations with respiratory symptoms, ED visits, and hospital admissions with short-term SO₂ exposures, generally of 24-h avg. Important new multicity studies and several other studies have found an association between 24-h avg ambient SO₂ concentrations and respiratory symptoms in children, particularly those with asthma. Furthermore, limited epidemiological evidence indicates that atopic children and adults may be at increased risk for SO₂-induced respiratory symptoms. Generally consistent and robust associations also were observed between ambient SO₂ concentrations and ED visits and hospitalizations for all respiratory causes, particularly among children and older adults (65+ years), and for asthma. Results of experiments in laboratory animals support these observations; studies in animals sensitized with antigen demonstrate that repeated exposure to near ambient SO₂ levels (as low as 0.1 ppm in guinea pigs) can exacerbate

1 allergic responses including mucin production, airway inflammation and airway
2 hyperresponsiveness.

3.2. Other Morbidity Associated with Short-Term SO₂ Exposure

3.2.1. Summary of Findings from the Previous Review

3 The studies reviewed in the 1982 AQCD primarily investigated respiratory health
4 outcomes. There were no key animal toxicological or human clinical studies available at the last
5 review to address effects of SO₂ exposure on the cardiovascular system. The only report was a
6 study in dogs exposed to air pollutant mixtures (SO₂ + H₂SO₄ with or without nonirradiated or
7 irradiated auto exhaust). No changes were observed in cardiovascular function at the end of 3
8 years of exposure and 3 years after exposure. No epidemiological studies linking exposure to
9 SO₂ with cardiovascular physiological endpoints or ED visits or hospital admissions for
10 cardiovascular causes were examined in the last review. Other organ systems in addition to
11 cardiovascular were not addressed in the 1982 AQCD.

3.2.2. Cardiovascular Effects Associated with Short-Term Exposure

12 The biological basis for SO₂-related cardiovascular health effects may lie in the stimulation
13 of chemosensitive receptors found in the respiratory tract which respond to irritants like SO₂.
14 Vagally-mediated responses may affect the cardiovascular system by inducing bradycardia and
15 either hypotension or hypertension, as discussed in Section 3.1.2. Alternatively oxidation
16 reactions mediated by the SO₂ metabolites sulfite and bisulfite which have been absorbed into
17 the systemic circulation may potentially alter cardiovascular function. In general, vagally-
18 mediated responses have been observed at lower concentrations of SO₂ than oxidative injury.

19 Since 1982, several animal toxicological studies have addressed the effects of SO₂ on
20 cardiovascular endpoints. These are summarized below and in Annex Table E-5. In addition,
21 there is one noteworthy study examining the hematological effects of short-term SO₂ exposure
22 (Annex Table E-8). Acute exposure of rats to 0.87 ppm SO₂ for 24 h resulted in increased
23 hematocrit, sulfhemoglobin and osmotic fragility as well as decreased whole blood and packed
24 cell viscosities (Baskurt, 1988). These results indicate a systemic effect of inhaled SO₂ at

1 concentrations near ambient levels and are consistent with an oxidative injury to red blood cells.
2 Only one study since 1982 measured systemic levels of sulfite or bisulfite following SO₂
3 inhalation (Gunnison et al., 1987; Annex Table E-22). Further studies are required to confirm that
4 inhalation exposures of SO₂ at or near ambient levels increase blood sulfite and bisulfite levels
5 sufficient for oxidative injury to blood cells or other tissues.

6 Recent epidemiological studies have examined the association between air pollution and
7 cardiovascular effects, including increased heart rate (HR), reduced heart rate variability (HRV),
8 incidence of ventricular arrhythmias, changes in blood pressure, incidence of myocardial
9 infarctions (MI), and ED visits and hospitalizations due to cardiovascular causes. The results of
10 these cardiovascular studies are summarized in Annex Tables F-3 and F-4.

3.2.2.1. Heart Rate and Heart Rate Variability

11 Heart rate variability (HRV) is generally determined by analyses of time (e.g., standard
12 deviation of normal R-R intervals [SDNN]) and frequency domains (e.g., low frequency [LF] /
13 high frequency [HF] ratio by power spectral analysis, reflecting autonomic balance) measured
14 during 24 h of electrocardiography (ECG). Brook et al. (2004) stated that HRV, resting heart rate,
15 and blood pressure are modulated by a balance between the two determinants of autonomic tone
16 (the sympathetic and parasympathetic nervous systems). An imbalance of cardiac autonomic
17 control may predispose susceptible people to greater risk of ventricular arrhythmias and
18 consequent cardiac deaths (Brook et al., 2004; Liao et al., 2004).

19 A limited number of human clinical studies examined the effect of SO₂ on HRV. During a
20 controlled exposure of 12 healthy subjects and 12 subjects with asthma to 0.2 ppm SO₂ for 1 h
21 under resting conditions, Tunnicliffe et al. (2001) reported that HF power, LF power, and total
22 power were higher with SO₂ exposures compared to air exposure in the healthy subjects, but that
23 these indices were reduced during SO₂ exposure in the subjects with asthma. The LF/HF ratios
24 were unchanged in both groups. The authors postulated two autonomic pathways for SO₂-
25 mediated bronchoconstriction. In healthy subjects, the dominant pathway was proposed to be the
26 rapidly adapting receptor/C-fiber route, which results in a central nervous system reflex with an
27 increase in vagal tone. In the asthmatic subjects, proximal airway narrowing was proposed as the
28 dominant response, possibly through neurogenic inflammation. This likely causes a
29 compensatory central nervous system-mediated reduction in vagal tone, resulting in

1 bronchodilation of the distal airway. While there were no detectable changes in symptoms or
2 lung function in either of the groups, this study provides some evidence that exposure to SO₂
3 may elicit systemic responses at these low levels (0.2 ppm).

4 In a similar study, Routledge et al. (2006) exposed patients with stable angina as well as
5 healthy subjects to 50 µg/m³ carbon particles, 0.2 ppm SO₂, alone and in combination, for 1 h
6 under resting conditions. HRV, C-reactive protein, and coagulation markers were measured. The
7 authors reported that for the healthy subjects, SO₂ exposure was associated with a decrease in
8 HRV markers of cardiac vagal control 4 h after exposure. However, it should be noted that there
9 was no apparent difference in the absolute value of the root mean square of successive RR
10 interval differences (r-MSSD) at 4 h postexposure between the control, SO₂, carbon, and
11 carbon/SO₂ groups. The significant difference reported in the change in r-MSSD from baseline to
12 4 h postexposure with SO₂ appears to be due to a higher baseline value of r-MSSD preceding the
13 SO₂ exposure compared to the baseline value of r-MSSD preceding the air exposure. There were
14 no changes in HRV among the patients with stable angina. The authors noted that this lack of
15 response in the heart patients may be due to a drug treatment effect rather than decreased
16 susceptibility; a large portion of the angina patients were taking beta blockers, which are known
17 to increase indices of cardiac vagal control.

18 In an epidemiological study, Liao et al. (2004) investigated short-term associations
19 between ambient pollutants and cardiac autonomic control from the fourth cohort examination
20 (1996 through 1998) of the population-based Atherosclerosis Risk in Communities (ARIC) study
21 using a cross-sectional study design. Men and women aged 45 to 64 years (n = 6,784) from three
22 U.S. study centers in North Carolina, Minnesota, and Mississippi were examined. Resting,
23 supine, and 5-min beat-to-beat R-R interval data were collected. The mean 24-h avg SO₂ level
24 measured 1 day prior to the HRV measurement was 4 ppb (SD 4). In addition to SO₂, the
25 potential effects of PM₁₀, O₃, CO, and NO₂ were evaluated. Previous-day SO₂ concentrations
26 were positively associated with HR and inversely associated with SDNN and LF power.
27 Consistently more pronounced associations were suggested between SO₂ and HRV among
28 persons with a history of coronary heart disease. Significant associations with HRV indices also
29 were observed for PM₁₀ and the other gaseous pollutants. When the regression coefficients for
30 each individual pollutant model were compared, the effects of PM₁₀ on HRV were considerably

larger than the effects for the gaseous pollutants, including SO₂. No multipollutant analyses were conducted.

Gold et al. (2000; reanalysis 2003) examined the effect of short-term changes in air pollution on HRV in a panel study of 21 older adults (aged 53 to 87 years) in Boston, MA. The study participants were observed up to 12 times from June to September 1997. The mean 24-h avg SO₂ concentration was 3.2 ppb (range: 0, 12.6). The 24-h avg SO₂ concentration was associated with decreased HR in the first 5-min rest period, but not in the overall 25-min study protocol. The effect estimate for SO₂ slightly diminished but remained marginally significant in a two-pollutant model with PM_{2.5}. The inverse association between SO₂ and HR observed in this study are in contrast to the SO₂-related increases in HR observed by Liao et al. (2004) and Peters et al. (1999). No associations were observed between HRV and SO₂. The strongest associations with HRV were observed for PM_{2.5} and O₃.

Another study of air pollutants and HRV was conducted in Boston, MA on 497 men from the Normative Aging Study (Park et al., 2005). The best 4-consecutive-min interval from a 7-min sample was used for the HRV calculations. For the exposure variable, 4-, 24-, and 48 h moving averages matched on the time of the ECG measurement for each subject were considered. The mean 24-h avg SO₂ concentration was 4.9 ppb (range: 0.95, 24.7). Associations with measures of HRV were reported for PM_{2.5} and O₃, but not with SO₂ for any of the averaging periods. In another study conducted in Boston, MA, Schwartz et al. (2005) found significant effects of increases in PM_{2.5} on measures of HRV, while no associations with SO₂ were observed. Other studies examined the relationship of SO₂ with HRV (Chan et al., 2005; de Paula Santos et al., 2005; Holguin et al., 2003; Luttmann-Gibson et al., 2006). Most of these studies, with the exception of de Paula Santos et al., did not observe associations with SO₂.

In the limited number of epidemiological studies that examined a possible effect of SO₂ on HRV, there were some suggestive findings; however, results reported from the human clinical studies were inconsistent. The overall evidence does not support the conclusion that SO₂ affects cardiac autonomic control.

3.2.2.2. Repolarization Changes

1 In addition to the role played by the autonomic nervous system in arrhythmogenic
2 conditions, myocardial vulnerability and repolarization abnormalities are believed to be key
3 factors contributing to the mechanism of such diseases.

4 Two in vitro studies (Nie and Meng, 2005) conducted with a 1:3 molar:molar mixture of
5 the SO₂ derivatives bisulfite and sulfite demonstrated effects of a 10-μm bisulfite:sulfite mixture
6 on sodium and L-type calcium currents (which included changes in inactivation and/or
7 activation, recovery from inactivation, and inactivation/activation time constants) in ventricular
8 myocytes. These in vitro observations suggested a potential role for L-type calcium current in
9 cardiac injury following SO₂ exposure. However, in vivo cardiovascular effects were observed
10 only at high SO₂ concentrations (10 ppm and higher). Additional toxicological studies are
11 necessary to evaluate repolarization changes at ambient levels of SO₂.

12 In an epidemiological study, Henneberger et al. (2005) examined the association of
13 repolarization parameters (QT duration, T-wave complexity, variability of T-wave complexity,
14 and T-wave amplitude) with air pollutants in patients with preexisting coronary heart disease (n =
15 56, all males) in East Germany. The patients were examined repeatedly once every 2 weeks for 6
16 months, for a total of 12 ECG recordings. The mean 24-h avg SO₂ concentration was 2 ppb
17 (range: 1, 4). Ambient SO₂ concentrations during the 24-h preceding the ECG were associated
18 with the QT interval duration, but not with any other repolarization parameters. Stronger
19 associations were observed between PM indices and QT interval duration, T-wave amplitude,
20 and T-wave complexity.

21 To summarize, the evidence, while suggestive, is too limited to draw conclusions on the
22 association of SO₂ exposure and repolarization changes at this time.

3.2.2.3. Cardiac Arrhythmias

23 One toxicological study examined the effects of PM, ultrafine carbon, and SO₂ on
24 spontaneous arrhythmia frequency in 18-month-old rats (Nadziejko et al., 2004). The rats were
25 exposed to 1 ppm SO₂ for 4 h. No significant change in the frequency of spontaneous
26 arrhythmias was found with SO₂ and ultrafine carbon exposure. However, rats exposed to

1 concentrated ambient PM had a significantly greater increase in the frequency of delayed beats
2 than rats exposed to air.

3 In a panel study of 100 patients with implanted cardioverter defibrillators (ICDs) in
4 Eastern Massachusetts, Peters et al (2000) tested the hypothesis that patients with ICDs would
5 experience life-threatening arrhythmias after an air pollution episode. The mean 24-h avg SO₂
6 concentration measured at two sites in Boston during the study period was 7 ppb (5th–95th
7 percentile: 1, 19). ICDs monitor ECG abnormalities, and treat ventricular fibrillation or
8 ventricular tachycardias by administering shock therapy to restore the normal cardiac rhythm.
9 The ICD device also stores information on each tachyarrhythmia and shock. There was no
10 association between SO₂ and defibrillator discharges in the 33 subjects who had any defibrillator
11 discharges during the follow-up period or in the 6 subjects who had at least 10 discharges. There
12 was some evidence that NO₂ was associated with increased defibrillatory interventions in the
13 subjects with any defibrillator discharges. Among the patients with at least 10 events, NO₂, CO,
14 and PM_{2.5} were found to be associated with defibrillator discharges.

15 In a follow-up study designed to confirm the findings of Peters et al. (2000), Dockery et al.
16 (2005) used a larger sample of ICD patients in Boston (n = 203) with a longer follow-up period.
17 The median concentration of 48-h avg SO₂ averaged across multiple sites in Boston was 4.9 ppb
18 (IQR 4.1). No significant associations were found between ventricular arrhythmic episode days
19 and any of the air pollutants. However, when the analysis was stratified by recent arrhythmias
20 (i.e., within 3 days), there was evidence of an excess risk of ventricular arrhythmia with SO₂,
21 PM_{2.5}, black carbon, NO₂, and CO. Since PM_{2.5}, black carbon, NO₂, and CO were correlated with
22 each other and with SO₂, the authors noted that differentiating the independent effects of the
23 pollutants would be difficult. A case-crossover analysis of the same data by Rich et al. (2005)
24 also observed associations with 48-h avg SO₂, but the SO₂ effect was not found to be robust to
25 adjustment by PM_{2.5}. In a similar study conducted in St. Louis, MO, an excess risk was
26 associated with SO₂ concentrations in the 24 h prior to an arrhythmia, but not with PM_{2.5} and O₃
27 (Rich et al., 2006). In this study, none of the other measured pollutants (PM, elemental carbon,
28 O₃, CO, NO₂) were correlated with SO₂. The authors suggested that the different effects observed
29 in St. Louis and Boston may be due to differences in the source or mix of air pollutants in these
30 cities. Finally, findings from a time series study of tachyarrhythmic events among 518 patients

over a 10 year period in Atlanta do not indicate an association with SO₂, nor with the other pollutants studied including PM_{2.5} and its components (Metzger et al., 2007).

Additional studies have examined the relationship of SO₂ with arrhythmias in Vancouver, and observed associations at very low ambient SO₂ concentrations (mean 24-h avg SO₂ of ~2.5 ppb with a max of 8.1 ppb). Vedal et al. (2004) stated that of all pollutants examined, the only one with somewhat consistent positive associations with arrhythmia events was SO₂. In season-stratified analyses, SO₂ was positively associated with arrhythmias in the winter, while in the summer the association was negative. On the other hand, in the Rich et al. (2004) study, positive associations were observed in the summer but not in the winter. The authors stated that it was difficult to interpret these findings.

Collectively, the epidemiological evidence for an association between short-term exposure to SO₂ and arrhythmias is inconsistent. The limited toxicological evidence does not provide biological plausibility for an effect.

3.2.2.4. Blood Pressure

Two animal toxicological studies examined the effect of SO₂ on blood pressure (Annex Table E-5). Hälinen et al. (2000) examined blood pressure changes in guinea pigs. The animals were hyperventilated to simulate exercise, and exposed to 1-, 2.5-, and 5 ppm SO₂ in cold, dry air. After 10-min exposures to each SO₂ concentration, separated by 15-min exposures to clean, warm, humid air, a transient increase in blood pressure was observed during 5 ppm SO₂ exposure in cold, dry air. In a second study (Halinen et al., 2000b), hyperventilated guinea pigs were exposed to cold, dry air alone or to 1 ppm SO₂ in cold, dry air for 60 min. The study reported similar increases in blood pressure and HR with exposure to cold, dry air or to SO₂ in cold, dry air. The increase in HR was gradual, while increases in blood pressure generally occurred during the first 10 to 20 min of exposure. Similar effects were observed with exposure to cold, dry air or to SO₂ in cold, dry air, suggesting that effects were associated with cold, dry air rather than with SO₂.

Ibald-Mulli et al. (2001) examined the association between blood pressure and SO₂ using survey data from the MONICA (Monitoring Trends and Determinants in Cardiovascular Disease) Project. Blood pressure measurements were taken from 2,607 men and women. The mean 24-h avg SO₂ concentration was 23 ppb (range: 5, 91). An increase in systolic blood pressure was

1 associated with 24-h avg SO₂ and TSP. However, in a two-pollutant model with TSP, the effect of
2 SO₂ on blood pressure was substantially reduced and became nonsignificant while the effect of
3 TSP was robust.

4 In a study by de Paula Santos et al. (2005), changes in blood pressure in association with
5 SO₂ were investigated in vehicular traffic controllers (n = 48) aged 31 to 55 years living in São
6 Paulo, Brazil, where vehicles are the primary source of air pollution. The mean 24-h avg SO₂
7 level, measured at six different stations around the city, was 7 ppb (SD 3). Blood pressure was
8 measured every 10 min when subjects were awake (6 a.m. to 11 p.m.) and every 20 min during
9 sleep (11 p.m. to 6 a.m.). Results indicated that SO₂, as well as CO, were associated with
10 increases in systolic and diastolic blood pressure. However, when a two-pollutant model was
11 used to test the robustness of the associations, only the CO effect remained statistically
12 significant.

13 Very few studies have examined the effects of short-term SO₂ exposure on blood pressure.
14 Collectively, the limited toxicological and epidemiological evidence does not suggest that
15 exposure to SO₂ has effects on blood pressure.

3.2.2.5. Blood Markers of Cardiovascular Risk

16 Folsom et al. (1997) demonstrated that elevated levels of fibrinogen, white blood cell
17 count, factor VIII coagulant activity (factor VIII-C), and von Willebrand factor were associated
18 with risk of cardiovascular disease. Schwartz (2001) investigated the association between various
19 blood markers of cardiovascular risk and air pollution among subjects in the Third National
20 Health and Nutrition Examination Survey (NHANES III) in the United States conducted between
21 1989 and 1994 across 44 counties. The NHANES III is a random sample of the U.S. population
22 with oversampling for minorities (30% of NHANES sample) and the elderly (20% of the
23 sample). The mean SO₂ concentration was 17.2 ppb (IQR 17) across the 25 counties where data
24 were available. This study looked at fibrinogen levels, platelet counts, and white blood cell
25 counts. After controlling for age, ethnicity, gender, body mass index, and smoking status and
26 number of cigarettes per day, SO₂ was found to be positively associated with white blood cell
27 counts. PM₁₀ was associated with all blood markers. In two-pollutant models, PM₁₀ remained a
28 significant predictor of white blood cell counts after controlling for SO₂, but not vice versa.

1 A recent cross-sectional study by (Liao et al., 2005) investigated the effects of air pollution
2 on plasma hemostatic and inflammatory markers in the ARIC study (n = 10,208). The authors
3 hypothesized that short-term exposure to air pollutants was associated with increased levels of
4 inflammatory markers and lower levels of albumin, as serum albumin is inversely associated
5 with inflammation. The mean 24-h avg SO₂ concentration was 5 ppb (SD 4). Significant
6 curvilinear relationships were observed between SO₂ and factor VIII-C, white blood cell counts,
7 and serum albumin. The authors noted that since no biological explanation could be offered for
8 the “U”-shaped curve between SO₂ and factor VIII-C and the “inverse U”-shape between SO₂
9 and albumin, generalization of the association should be exercised with caution. No associations
10 were observed between SO₂ and fibrinogen or von Willebrand factor.

11 In another large cross-sectional study of 7,205 office workers in London, Pekkanen et al
12 (2000) examined the association between plasma fibrinogen and ambient air pollutants. The
13 mean 24-h avg SO₂ was 9 ppb (10th–90th percentile: 5, 19). Associations with fibrinogen were
14 observed for all pollutants examined, either in all-year or summer-only analyses, except for SO₂
15 and O₃.

16 Taken together, results from the limited number of studies do not suggest that SO₂ is
17 associated with various blood markers of cardiovascular risk.

3.2.2.6. Acute Myocardial Infarction

18 The association between air pollution and the incidence of MI was examined in a small
19 number of studies. As part of the Determinants of Myocardial Infarction Onset Study, Peters
20 et al. (2001) examined 772 patients with MI living in greater Boston, MA. A case-crossover
21 design was used to assess changes in the risk of acute MI after exposure to potential triggers. The
22 mean 24-h avg SO₂ was 7 ppb (range: 1, 20) during the study period. Similarly, the mean 1-h avg
23 SO₂ was 7 ppb (range: 0, 23). In an analysis that considered both the 2-h avg (between 60 and
24 180 min before the onset of symptoms) and 24-h avg (between 24 and 48 h before the onset)
25 concentrations jointly, the study found no significant association between risk of MI and SO₂. Of
26 all the pollutants considered, only PM_{2.5} and PM₁₀ were found to be associated with an excess
27 risk of MI. In a study of 5793 confirmed cases of acute MI in King County Washington, Sullivan
28 et al. (2005) also used a case-crossover design to investigate the association of ambient levels of

several air pollutants 1, 2, 4 and 24 h before the MI onset. No association with SO₂ (or with PM_{2.5}) was observed. The mean SO₂ level was 9 ppb (range: 0-39 ppb) at the time of the study.

In the MONICA Project, the effect of air pollution on acute MI was studied in Toulouse, France, using a case-crossover study design (Ruidavets et al., 2005). The mean 24-h avg SO₂ level was 3 ppb (5th–95th percentile: 1, 5). A total of 399 cases of acute MI were recorded during the study period. O₃, but not SO₂ or NO₂, was found to be associated with the incidence of acute MI. Exposure to PM was not considered in this study.

Only a limited number of studies examined the association between ambient SO₂ concentrations and incidence of acute MI. These studies provide no evidence that exposure to SO₂ increases the risk of MI.

3.2.2.7. Emergency Department Visits and Hospitalizations for Cardiovascular Diseases

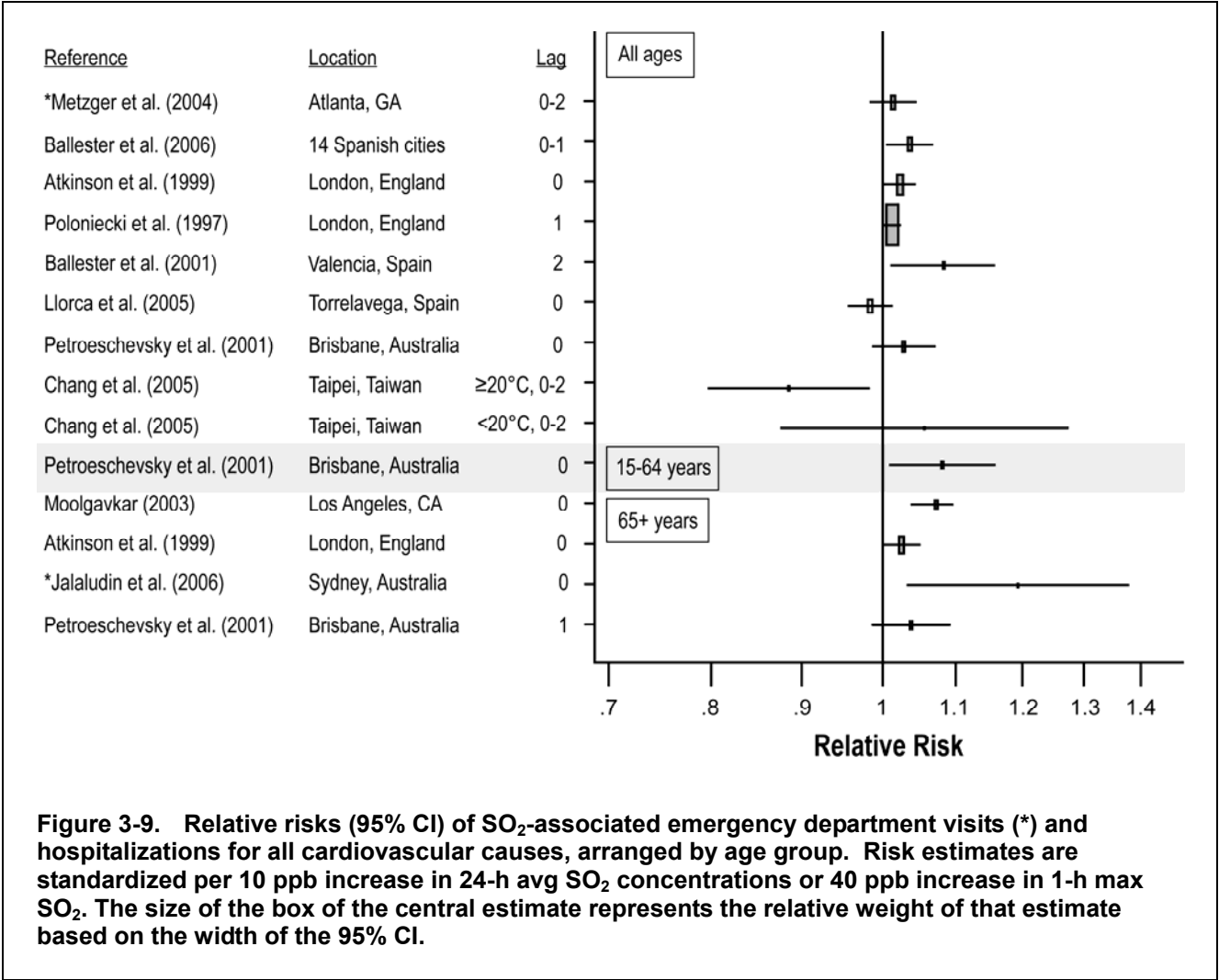
The current review includes more than 30 peer-reviewed studies that address the effect of SO₂ exposure on ED visits or hospitalizations for cardiovascular diseases. These studies are discussed briefly in this section and further summarized in Annex Table F-4.

3.2.2.7.1. All Cardiovascular Diseases

The disease grouping of all cardiovascular diseases typically includes all diseases of the circulatory system (e.g., heart diseases and cerebrovascular diseases, ICD9 Codes 390-459). A summary of the associations reported for ambient SO₂ concentrations with all cardiovascular diseases are presented in Figure 3-9.

In a study of 11 cities in Spain, an excess risk of 3.6% (95% CI: 0.6, 6.7) per 10 ppb increase in 24-h avg SO₂ at a 0-1 day lag was observed for all cardiovascular disease admissions (Ballester et al., 2006). The mean 24-h avg SO₂ level in the cities studied was 6.6 ppb. In addition, time-series data linking SO₂ with hospital admissions for cardiovascular diseases in three metropolitan areas in the United States (i.e., Cook, Maricopa, Los Angeles Counties) was conducted (Moolgavkar, 2000; reanalysis, 2003). Among older adults (65+ years) in Los Angeles County, a 13.7% (95% CI: 11.3, 16.1) excess risk in admissions was observed per 10 ppb increase in 24-h avg SO₂ at lag 0 day, using a Generalized Linear Model (GLM) and natural splines to adjust for temporal trends rather than GAM. The median 24-h avg SO₂ level for Los

1 Angeles County was 2 ppb during the study period. Results for Maricopa and Cook counties
 2 were not presented in the reanalysis.



3 In a large single city study Metzger et al. (2004) examined approximately 4.4 million
 4 hospital visits to 31 hospitals from 1993 to 2000 in Atlanta, GA and reported a 1.4% (95% CI: 1.5, 4.4) excess risk in ED visits for cardiovascular causes per 40-ppb increase in 1-h max SO₂.
 5 Peel et al. (2006) conducted analyses using the same dataset to compare results obtained with a
 6 case-crossover design to the Metzger et al. (2007) results, which were obtained using a time
 7 series approach. Peel et al. and Metzger et al. report similar findings. The median 1-h max SO₂
 8 level in Atlanta during the study period was 11 ppb (10th–90th percentile: 2–39). Results from
 9 several single-city studies in Europe, Australia, and Taiwan indicated positive associations with
 10

SO₂ (Atkinson et al., 1999a; Ballester et al., 2001; Jalaludin et al., 2006; Petroeschevsky et al., 2001; Poloniecki et al., 1997), though others observed negative associations (Chang et al., 2005; Llorca et al., 2005) (see Figure 3-9).

3.2.2.7.2. Specific Cardiovascular Diseases

Several studies examined the effect of ambient SO₂ on hospital admissions for cardiac disease (ICD9 Codes 390-429), ischemic heart disease (IHD, ICD9 Codes 410-414), dysrhythmia (ICD9 Code 427), congestive heart failure (CHF, ICD9 Code 428), MI (410) or cerebrovascular diseases (ICD9 Codes 430-438). A European multicity study reported statistically significant positive associations with cardiac disease admissions (Ballester et al., 2006). However, adjustment for PM₁₀ and CO in two-pollutant models diminished the association reported by Ballester et al. by approximately 50%. Findings for cardiac disease admissions reported in several additional single city studies conducted in the United States, Canada, Australia and Europe were inconsistent (Fung et al., 2005; Jalaludin et al., 2006; Llorca et al., 2005; Michaud et al., 2004).

Analyses restricted to diagnoses of IHD (Jalaludin et al., 2006; Lee et al., 2003; Lin et al., 2003a; Metzger et al., 2004; Peel et al., 2007), CHF (Koken et al., 2003; Metzger et al., 2004; Morris et al., 1995; Peel et al., 2007; Wellenius et al., 2005b), dysrhythmia (Koken et al., 2003; Metzger et al., 2004; Peel et al., 2007), MI (Koken et al., 2003; Lin et al., 2003a), and angina pectoris (Hosseinpoor et al., 2005) were also conducted. Metzger et al. (2004) observed weak nonsignificant or negative associations of 1-h max SO₂ with IHD, CHF, and dysrhythmia. Using the same dataset, Peel et al. (2007) investigated effect modification of cardiovascular disease outcomes across comorbid disease status categories, including hypertension, diabetes, COPD, dysrhythmia, and CHF. Authors observed only weak nonsignificant or negative associations for IHD, CHF, and dysrhythmia with ambient 1-h max SO₂ level in any comorbid disease category. Both increases in admissions or ED visits (Jalaludin et al., 2006; Koken et al., 2003; Wellenius et al., 2005a) and weak or negative associations (Hosseinpoor et al., 2005; Lee et al., 2003b; Lin et al., 2003a) were reported in other studies.

Studies of the effect of SO₂ on cerebrovascular admissions were also considered. Positive associations were reported for ischemic stroke (Villeneuve et al., 2006; Wellenius et al., 2005b; Wellenius et al., 2005a). However Wellenius et al. (2005b) reported stronger associations for

1 NO₂ and CO than SO₂, and the association reported by Villeneuve et al. (2006) was diminished
2 in two-pollutant models. No meaningful positive associations of ambient SO₂ with
3 cerebrovascular diseases were observed in several other studies (Henrotin et al., 2007; Jalaludin
4 et al., 2006; Metzger et al., 2004; Peel et al., 2007; Tsai et al., 2003).

3.2.2.7.3. Summary of Evidence on Emergency Department Visits and Hospitalizations from Cardiovascular Diseases

5 Several studies have observed positive associations between ambient SO₂ concentrations
6 and ED visits or hospital admissions for cardiovascular diseases (e.g., all cardiovascular diseases,
7 cardiac diseases, cerebrovascular diseases) particularly among individuals 65+ years of age, but
8 results are not consistent across studies. The strongest evidence comes from a large multicity
9 study conducted in Spain (Ballester et al., 2006) that observed statistically significant positive
10 associations between ambient SO₂ and cardiovascular disease admissions; however, the SO₂
11 effect was found to diminish by half with PM₁₀ and CO adjustment. Only a limited number of
12 studies assessed potential confounding by copollutants despite the moderate to strong correlation
13 between SO₂ and various copollutants in most studies. While some studies suggest that the
14 association of SO₂ with cardiovascular hospitalizations were generally robust to adjustment for
15 BS and PM₁₀ (Ballester et al., 2001; Fung et al., 2005), several other studies, including that by
16 Ballester et al. (2006), observed that the effect of SO₂ on cardiovascular ED visits and
17 hospitalizations may be confounded by copollutant exposures. Jalaludin et al. (2006) reported a
18 3% excess risk in cardiovascular disease hospital admissions per 0.75 ppb incremental change in
19 24-h avg SO₂ in single-pollutant models, which was reduced to null when CO was included.
20 Chang et al. (2005) noted that the observed negative association of SO₂ with all cardiovascular
21 disease hospitalizations they observed was strengthened after adjusting for NO₂, PM₁₀, and CO
22 in two-pollutant models. The authors attributed this finding to possible collinearity problems
23 between SO₂ and copollutants. None of the epidemiological studies examined effects of possible
24 interactions among copollutants.

3.2.2.8. Summary of Evidence on the Effect of Short-Term SO₂ Exposure on Cardiovascular Health

25 The overall evidence on the effect of short-term exposure to SO₂ on cardiovascular health
26 effects is *inadequate to infer the presence or absence of a causal relationship*. Epidemiological

1 studies of HRV, cardiac repolarization changes, and cardiac rhythm disorders provide limited
2 evidence of associations with short-term exposure to SO₂. There was some suggestive evidence
3 of an association between SO₂ exposure and HRV in the epidemiological studies, but the
4 evidence from two human clinical studies were weak and inconsistent. Similarly, several studies
5 observed positive associations between ambient SO₂ concentrations and ED visits or hospital
6 admissions for cardiovascular diseases, but results were not consistent across studies and specific
7 cardiovascular disease outcomes. In general, most epidemiological studies observed that these
8 cardiac outcomes were associated more strongly with PM compared to SO₂. In the limited
9 studies that examined potential confounding by copollutants using multipollutant models, the
10 SO₂ effect was generally found to diminish with adjustment for PM indices or CO.

11 Given the lack of coherence among the cardiovascular outcomes examined and the limited
12 evidence available to evaluate potential confounding and interaction by copollutants, the overall
13 evidence for cardiovascular health effects following exposure to ambient SO₂ is weak and
14 insufficient to make a causal determination.

3.2.3. Other Effects Associated with Short-Term SO₂ Exposure

15 The short-term effects of SO₂ on other organ systems were not examined in the previous
16 review. A review of animal toxicological studies published since the 1982 AQCD indicates a
17 limited number of research inquiries addressing the systemic effects of short-term SO₂ exposure
18 in various other organs. The most recent studies on these are summarized in Annex Tables E-6
19 through E-9 and E-22 through E-24.

20 Of note are three ex vivo acute exposure studies using SO₂ derivatives (sulfite and
21 bisulfite) on hippocampal or dorsal root ganglion neurons isolated from Wistar rats (Du and
22 Meng, 2004a; b; 2006). Perturbations were observed in potassium-, sodium-, and calcium-gated
23 channels at concentrations of 0.01-100 µM. These authors speculated that such effects might
24 correlate with the neurotoxicity that has been associated with SO₂ inhalation. However effects on
25 the nervous system have generally been studied using chronic exposures ≥ 5 ppm SO₂. Effects
26 observed at these levels are of questionable significance in evaluating the health effects at
27 ambient levels. These studies are summarized in Annex Table E-6.

3.3. Mortality Associated with Short-Term SO₂ Exposure

3.3.1. Summary of Findings from the Previous Review

The studies available to review in the 1982 AQCD were mostly from historical data including London, England, and New York City air pollution episodes. Effects of SO_x (mainly SO₂) were investigated along with PM indices because they shared a common source, coal burning, and separating their associations with mortality was a challenge that many of the earlier episodic studies could not resolve. The SO₂ levels observed in these air pollution episodes were several orders of magnitude higher than the current average levels observed in U.S. cities (e.g., in the 1962 New York City episode, SO₂ in Manhattan peaked at 400 to 500 ppb). Some of these London and New York City studies suggested that PM, not SO₂, was associated with observed mortality, but the 1982 AQCD could not resolve the relative roles of these two pollutants and suggested that the clearest mortality associations were seen when both pollutants were at high levels (24-h avg values of both BS and SO₂ exceeding 1000 µg/m³ [~400 ppb for SO₂]) and less so at lower ranges although the review of the studies and reanalyses found no clear evidence of a threshold for SO₂.

The 1986 Second Addendum to the 1982 AQCD reviewed more reanalyses of the London data and analyses of New York City, Pittsburgh, and Athens data. While these reanalyses and some new analyses confirmed earlier findings (and suggested stronger evidence of BS effects than of the SO₂ effects), given the remaining uncertainties with exposure error and statistical modeling, there was not sufficient information to quantitatively determine concentration-response relationships at lower concentrations of either PM or SO₂.

A series of short-term mortality effects studies in the late 1980s and early 1990s (Pope, 1989; Fairley, 1990; Dockery et al., 1992; Pope et al., 1992; Schwartz and Dockery, 1992) showed associations between mortality and PM indices at relatively low levels. Since then, a large number of epidemiological studies have investigated the adverse health effects of air pollution with hypotheses mainly focused on PM, and SO₂ was often analyzed as one of the potential confounders in these studies.

3.3.2. Associations of Mortality and Short-Term SO₂ Exposure in Multicity Studies and Meta-Analyses

In reviewing the range of SO₂ mortality effect estimates, multicity studies provide especially useful information because they analyze data from multiple cities using a consistent method, avoiding potential publication bias. There have been several multicity studies from the United States, Canada, and Europe, some of which will be discussed in the sections below. Meta-analysis studies also provide useful information on describing heterogeneity of effect estimates across studies; however, in contrast to multicity studies, the observed heterogeneity may reflect the variation in analytical approaches across studies. In addition, the effect estimate from a meta-analysis may be subject to publication bias, unless the analysis specifically examines such bias and adjusts for it. These studies, as well as many other single-city studies, are summarized in Annex Table F-5.

3.3.2.1.1. Multicity Studies

3.3.2.1.2. National Morbidity, Mortality, and Air Pollution Study

The time-series analysis of the largest 90 U.S. cities (Samet et al., 2000; reanalysis Dominici et al., 2003) in the National Morbidity, Mortality, and Air Pollution Study (NMMAPS) is by far the largest multicity study conducted to date to investigate the mortality effects of air pollution, but its primary focus was PM₁₀. It should also be noted that, according to the table of mean pollution levels in the original report (Samet et al., 2000), SO₂ data were missing in 28 of 90 cities. Annual 24-h avg mean SO₂ levels ranged from 0.4 ppb (Riverside, CA) to 14.2 ppb (Pittsburgh, PA), with a mean of 5.9 ppb during the study period of 1987 to 1994. The analysis in the original report used GAM models with default convergence criteria. Dominici et al. (2003) reanalyzed the data using GAM with stringent convergence criteria as well as using GLM. It should be noted that this model's adjustment for weather effects employs more terms than other time-series studies in the literature, suggesting that the model adjusts for potential confounders more aggressively than the models in other studies.

Figure 3-10 shows the all-cause mortality excess risk estimates for SO₂ from Dominici et al. (2003). The mortality excess risk estimate with a 1-day lag was 0.60% (95% CI: 0.26, 0.95) per 10 ppb increase in 24-h avg SO₂. PM₁₀ and O₃ (in summer) appeared to be more strongly associated with mortality compared to the other gaseous pollutants. The model with PM₁₀ and

1 NO₂ resulted in an appreciably reduced SO₂ excess risk estimate, 0.38% (95% CI: -0.62, 1.38)
 2 per 10 ppb increase in 24-h avg SO₂. These results suggest that the observed SO₂-mortality
 3 association could be confounded by PM₁₀ and NO₂. The authors stated that the results did not
 4 indicate associations of SO₂, NO₂, and CO with all-cause mortality.

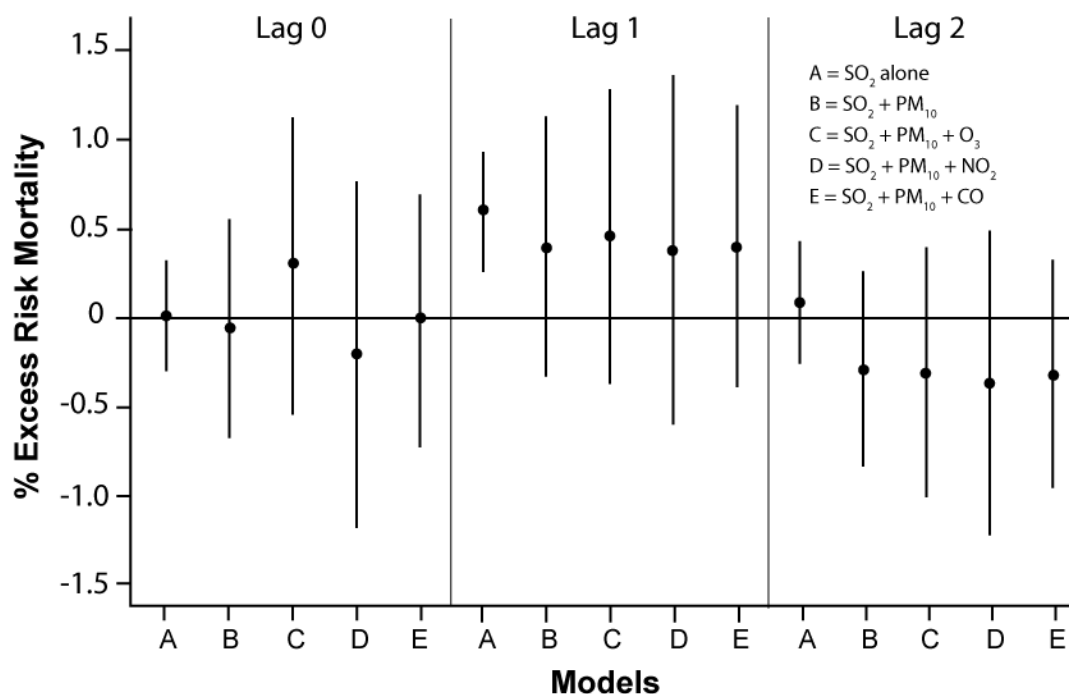


Figure 3-10. All cause mortality excess risk estimates for SO₂ from the National Morbidity, Mortality, and Air Pollution Study. Posterior means and 95% posterior intervals of national average estimates of SO₂ effects on all-cause mortality from non-external causes per 10 ppb increase in 24-h avg SO₂ at 0, 1, and 2-day lags within sets of the 62 cities with pollutant data available.

Source: Dominici et al. (2003).

3.3.2.1.3. Canadian Multicity Studies

5 There have been three Canadian multicity studies conducted by the same group of
 6 investigators examining the association between mortality and short-term exposure to air
 7 pollutants (Burnett et al., 1998; 2000b; 2004). This section focuses on Burnett et al. (2004) as
 8 this study is the most extensive Canadian multicity study, both in terms of the length and
 9 coverage of cities. The discussion in this study focused on NO₂, because NO₂ was the best

1 predictor of short-term mortality fluctuations among the pollutants. This was also the case in the
2 Burnett et al. (1998) study of the gaseous pollutants in 11 Canadian cities. The mean 24-h avg
3 SO₂ levels across the 12 cities was 5.8 ppb, with city means ranging from 1 ppb in Winnipeg to
4 10 ppb in Halifax. The population-weighted average was 5 ppb. The mean SO₂ levels in this
5 study were similar to those in the NMMAPS (mean 24-h avg SO₂ levels across the 62 NMMAPS
6 cities was 5.9 ppb).

7 All-cause (nonaccidental), cardiovascular, and respiratory mortality were analyzed in
8 Burnett et al. (2004). For SO₂, PM_{2.5}, PM_{10-2.5}, PM₁₀ (arithmetic addition of PM_{2.5} and PM_{10-2.5}),
9 CoH, and CO, the strongest mortality association was found at a 1-day lag, whereas for NO₂, it
10 was the 3-day moving average (i.e., average of 0, 1, and 2-day lags), and for O₃, it was the 2-day
11 moving average. The daily 24-h avg values showed stronger associations than the daily 1-h max
12 values for all the gaseous pollutants and CoH except for O₃. The SO₂ all-cause mortality excess
13 risk estimate was 0.74% (95% CI: 0.29, 1.19) per 10 ppb increase in the 24-h avg SO₂ with a 1-
14 day lag. After adjusting for NO₂, the SO₂ effect estimate was reduced to 0.42% (95% CI: 0.01,
15 0.84), while the NO₂ effect estimate was only slightly affected. In this analysis, no regression
16 analysis using both SO₂ and PM was conducted. The Burnett et al. (2000) analysis observed that
17 the simultaneous inclusion of SO₂ and PM_{2.5} in the model reduced the SO₂ effect estimate by
18 half, whereas the PM_{2.5} estimate was only slightly reduced. Overall, these results suggest that
19 SO₂ was not an important predictor of daily mortality in the Canadian cities and that its mortality
20 associations could be confounded with NO₂ or PM.

3.3.2.1.4. Air Pollution and Health: A European Approach

21 Several Air Pollution and Health: a European Approach (APHEA) analyses have reported
22 SO₂ mortality excess risk estimates. Katsouyanni et al. (1997) examined the association of PM₁₀,
23 BS, and SO₂ with all-cause mortality in 12 European cities using the standard APHEA (GLM)
24 approach. The same data set was reanalyzed to adjust for the seasonal cycles (Samoli et al., 2001;
25 2003). The reanalysis by Samoli et al. (2003) produced results that were similar to those in the
26 original analysis by Katsouyanni et al. (1997). Since the original analysis presented more results,
27 including multipollutant model results, discussion will focus on this analysis.

28 The study by Katsouyanni et al. includes seven western European cities (Athens,
29 Barcelona, Cologne, London, Lyon, Milan, and Paris) and five central eastern European cities

(Bratislava, Krakow, Lodz, Poznan, and Wroclaw). The data covered at least 5 consecutive years for each city within the years 1980 through 1992. The SO₂ levels in these cities were generally higher than in the United States or Canada, with the median 24-h avg SO₂ ranging from 5 ppb in Bratislava to 28 ppb in Krakow. Analysis was restricted to days when PM and SO₂ concentrations did not exceed 200 µg/m³ (76 ppb for SO₂) to evaluate the effects of moderate to low exposures. The data were analyzed by each center separately following a standardized method, but the lag for the “best” model was allowed to vary in these cities from 0 to 3 days. The city-specific effect estimates were then examined in the second stage for source of heterogeneity using city-specific variables such as mean pollution and weather variables, accuracy of the air pollution measurements, health of the population, smoking prevalence, and geographical differences.

The city-specific estimates were found to be heterogeneous and, among the explanatory variables, only the separation between western and central eastern European cities resulted in more homogeneous groups. The all-cause mortality excess risk estimates were 1.14% (95% CI: 0.88, 1.39), 1.99% (95% CI: 1.15, 2.83), and 0.46% (95% CI: -0.23, 1.15) for all the 12 cities combined, western cities, and central eastern cities, respectively, per 10 ppb increase in the 24-h avg SO₂ at variable single-day lags. Zmirou et al. (1998) analyzed cardiovascular and respiratory mortality in 10 of the 12 APHEA cities and observed that the cause-specific mortality excess risk estimates were higher than those for all-cause mortality. As in the analyses of all-cause mortality, SO₂ effect estimates for these cause specific deaths were higher in western European cities than in central eastern European cities.

Seasonal analyses indicated that the summer estimate was slightly higher than the winter estimate in the western cities, but the difference was not statistically significant. The results for the two-pollutant model with SO₂ and BS were presented for the western cities, with a similar extent (~30%) of reductions in the estimates of both pollutants (1.31% [95% CI: 0.40, 2.23] for SO₂). Furthermore, for western cities, they estimated effects for SO₂ for days with high or low BS levels and the corresponding BS effects for days with high or low SO₂ levels and found that the effects were similar in the stratified data. From these results, Katsouyanni et al. (1997) suggested that the effects of the two pollutants were independent.

Overall, the APHEA studies provide some suggestive evidence that the effect of short-term exposure to SO₂ on mortality is independent of PM. This is somewhat in contrast to the U.S. and

1 Canadian studies. The SO₂ levels were much higher in the European cities, but the type of PM
2 constituents also might be different.

3.3.2.1.5. The Netherlands Study

3 In the Netherlands studies by Hoek et al. (Hoek et al., 2000; 2001; reanalysis, Hoek, 2003),
4 the association between air pollutants and mortality were examined in a large population (14.8
5 million for the entire country) over the period of 1986 through 1994. The Netherlands were not
6 part of the APHEA analysis. The median 24-h avg SO₂ level in the Netherlands was 4 ppb (6 ppb
7 for the four major cities). All the pollutants examined, including PM₁₀, BS, O₃, NO₂, SO₂, CO,
8 sulfate, and nitrate, were associated with all-cause mortality, and for single-day models, a 1-day
9 lag showed the strongest associations for all the pollutants. The following effect estimates are all
10 from the GLM models with natural splines for smoothing functions. The SO₂ excess risk
11 estimate in a single-pollutant model was 1.31% (95% CI: 0.69, 1.93) per 10 ppb increase in
12 24-h avg SO₂ at a 1-day lag and 1.78% (95% CI: 0.86, 2.70) at an average of 0- to 6-day lag.
13 Seasonal analyses showed slightly greater effect estimates during the summer compared to the
14 winter. SO₂ was most highly correlated with BS (r = 0.70). The simultaneous inclusion of SO₂
15 and BS reduced the effect estimates for both pollutants (SO₂ effect estimate was 1.07% [95% CI:
16 -0.27, 2.42] per 10 ppb increase with an average of 0- to 6-day lag of 24-h avg SO₂). PM₁₀ was
17 less correlated with SO₂ (r = 0.65), and the simultaneous inclusion of these pollutants resulted in
18 an increase in the SO₂ effect estimate. These results from the analysis of the Netherlands data
19 suggested some indication of confounding between SO₂ and BS.

20 Cause specific analysis showed larger excess risk estimates for COPD (3.61% [95% CI: -
21 0.29, 7.67] per 10 ppb increase in the average of 0- through 6-day lags of 24-h avg SO₂) and
22 pneumonia (6.56% [95% CI: 1.16, 12.24]) deaths compared to that from all causes, but because
23 essentially all of the pollutants showed larger effect estimates for these sub-categories, it is
24 difficult to interpret these estimates as effects of SO₂ alone. Similarly, the effect estimates for
25 heart failure (7.1% [95% CI: 2.6, 11.7]) and thrombosis-related deaths (9.6% [95% CI: 3.1,
26 16.6]) were larger than that for total cardiovascular (2.7% [95% CI: 1.3, 4.1]) causes. Since the
27 same pattern was seen for other pollutants as well, it is difficult to interpret these cause-specific
28 effect estimates due to SO₂ alone or any one of the pollutants analyzed.

3.3.2.1.6. Other European Multicity Studies

Other European multicity studies were conducted in 8 Italian cities (Biggeri et al., 2005), 9 French cities (Le Tertre et al., 2002), and 13 Spanish cities (Ballester et al., 2002). The studies by Le Tertre et al. and Ballester et al. were conducted using GAM methods with the default convergence setting.

Biggeri et al. analyzed eight Italian cities (Turin, Milan, Verona, Ravenna, Bologna, Florence, Rome, and Palermo) for mortality and hospital admissions (mortality data were not available for Ravenna and Verona). The study period varied from city to city between 1990 and 1999. Only single-pollutant models were examined in this study. The SO₂ excess risk estimates were 4.14% (95% CI: 1.05, 7.33), 4.94% (95% CI: 0.41, 9.67), and 7.37% (95% CI: -3.58, 19.57) per 10 ppb increase with an average of 0- and 1-day lag of 24-h avg SO₂ for all-cause, cardiovascular, and respiratory deaths, respectively. Since all the pollutants showed positive associations with these mortality categories and the correlations among the pollutants were not presented, it is not clear how much of the observed associations are shared or confounded. The mortality excess risk estimates were not heterogeneous across cities for all the gaseous pollutants. It should be noted that in Turin, Milan, and Rome, the mean SO₂ values declined by 50% from the first half to the second half of the study period, while the levels of other pollutants declined by smaller fractions. This also complicates the interpretation of SO₂ effect estimates in this study, which are much higher than those from the APHEA studies.

The Le Tertre et al. study of nine French cities examined BS, SO₂, NO₂, and O₃ by generally following the APHEA protocol, but using GAM with default convergence criteria and using the average of lags 0 and 1 day for combined estimates. SO₂ data were not available in one of the nine cities (Toulouse). All four pollutants were positively associated with mortality outcomes. The study did not report descriptions of correlation among the pollutants or conduct multipollutant models, and therefore, it is difficult to assess the potential extent of confounding among these pollutants. The SO₂ effect estimates were homogeneous across cities, with the exception of Bordeaux, which was the only city that used strong acidity as a proxy for SO₂.

The Spanish Multicentre Study on Air Pollution and Mortality (EMECAM) examined the association of PM indices (i.e., PM₁₀, TSP, BS) and SO₂ with mortality in 13 cities (Ballester et al., 2002). These studies followed the APHEA protocol, but using the GAM approach. The daily mean 24-h avg SO₂ concentrations ranged from 3 to 17 ppb. In the seven cities where 1-h max

SO₂ data were also available, mean concentrations ranged from 21 to 43 ppb. The combined effect estimates for all-cause and respiratory mortality were statistically significant for both 24-h avg SO₂ and 1-h max SO₂. Controlling for PM indices substantially diminished the effect estimates for 24-h avg SO₂, but not for 1-h max SO₂. The authors reported that these results could indicate an independent impact of peak values of SO₂ more than an effect due to a longer exposure.

3.3.2.2. Meta-Analyses of Air Pollution-Related Mortality Studies

3.3.2.2.1. Meta-Analysis of All Criteria Pollutants

Stieb et al. (2002) reviewed time-series mortality studies published between 1985 and 2000, and conducted a meta-analysis to estimate combined effects for PM₁₀, CO, NO₂, O₃, and SO₂. Since many of the studies reviewed in that analysis used GAM with default convergence parameters, Stieb et al. (2003) updated the estimates by separating the GAM versus non-GAM studies. In addition, separate combined estimates were presented for single- and multipollutant models. There were more GAM estimates than non-GAM estimates for all the pollutants except for SO₂. For SO₂, there were 29 non-GAM estimates from single-pollutant models and 10 estimates from multipollutant models. The lags and multiday averaging used in these estimates varied. The combined estimate for all-cause mortality was 0.95% (95% CI: 0.64, 1.27) per 10 ppb increase in 24-h avg SO₂ from the single-pollutant models and 0.85% (95% CI: 0.32, 1.39) from the multipollutant models. Because these estimates are not from an identical set of studies, the difference (or lack of a difference, as in this case) between the two estimates may not necessarily be due to the effect of adding a copollutant in the model. Note that the data extraction procedure of this meta-analysis for the multipollutant models was to include from each study the multipollutant model that resulted in the greatest reduction in effect estimates compared with that observed in single-pollutant models. It should also be noted that all the multicity studies whose combined estimates have been discussed in the previous section were published after this meta-analysis.

3.3.2.2.2. Health Effects Institute Review of Air Pollution Studies in Asia

The Health Effects Institute (HEI) conducted a comprehensive review of air pollution health effects studies (HEI, 2004). They summarized the results from mortality and hospital

admission studies of the health effects of ambient air pollution in Asia (East, South, and Southeast) published in peer-reviewed scientific literature from 1980 through 2003. Of the 138 papers the report identified, most were studies conducted in East Asia (mainland China, Taipei, Hong Kong, South Korea, and Japan). The levels of SO₂ in these Asian cities were generally higher than in U.S. or Canadian cities, with more than half of these studies reporting mean 24-h avg SO₂ levels of > 10 ppb. Based on a comparison of the reported mean SO₂ levels from the same cities in different time periods, it is clear that the SO₂ levels declined significantly in the 1990s. The meta-analysis used the most recent estimate for each city to reflect recent pollution levels. Based on the criteria of having at least one year of data, model adjustment for major time-varying confounders, and reporting effect estimates per unit increase in air pollution, the meta-analysis included 28 time-series studies (11 from South Korea, 6 from mainland China, 6 from Hong Kong, and 1 each from Taipei, India, Singapore, Thailand, and Japan). The lags selected to compute combined estimates were inevitably variable; a systematic approach was used to favor the a priori lag stated in the study, followed by the most significant lag, and then the largest effect estimate.

Among the health outcomes examined in the meta-analysis, all-cause mortality was addressed in the largest number of studies (13 studies) and SO₂ was the most frequently studied pollutant (11 studies). The report generally focused on the results of single-pollutant models, as there were too few studies with results of comparable multipollutant models to allow meaningful analysis. The SO₂ mortality effect estimates showed evidence of heterogeneity. The combined estimate for all-cause mortality was 1.49% (95% CI: 0.86, 2.13) per 10 ppb increase in 24-h avg SO₂. One of the limitations noted in the report was that some degree of publication bias was present in these studies.

3.3.3. Evidence of the Effect of SO₂ on Mortality from an Intervention Study

Many time-series studies provide estimates of excess risk of mortality, but a question remains as to the likelihood of a reduction in deaths when SO₂ levels are actually reduced. A sudden change in regulation in Hong Kong in July 1990 required the conversion to fuel oil with low sulfur content. The reduction in respiratory symptoms among children living in the polluted district in Hong Kong after the intervention were previously discussed in Section 3.1.6. Hedley

et al. (2002) examined changes in mortality rates following the intervention. The SO₂ levels after the intervention declined about 50% (from about 17 ppb to 8 ppb), but the levels for PM₁₀, NO₂, and sulfate did not change and O₃ levels slightly increased. The seasonal mortality analysis results showed that the apparent reduction in seasonal death rate occurred only during the first winter, and this was followed by a rebound (i.e., higher than expected death rate) in the following winter, then returned to the expected pattern three to five years after the intervention. Using Poisson regression of the monthly deaths, the average annual trend in death rate significantly declined after the intervention for all causes (2.1%), respiratory causes (3.9%), and cardiovascular causes (2.0%), but not from other causes. These results seem to suggest that a reduction in SO₂ leads to an immediate reduction in deaths and a continuing decline in the annual trend in death rates. Hedley et al. estimated that the expected average gain in life expectancy per year due to the lower SO₂ levels was 20 days for females and 41 days for males.

Interpreting these results is somewhat complicated by an upward trend in mortality across the intervention point, which the authors noted was due to increased population size and aging. The results suggest that such an upward trend is less steep after the introduction of low sulfur fuel. While the Poisson regression model of monthly deaths does adjust for trend and seasonal cycles, the regression model does not specifically address the influence of influenza epidemics, which can vary from year to year. This issue also applies to the analysis of warm to cool season change in death rates. The most prominent feature of the time-series plot (or the fitted annual cycle of monthly deaths) presented in this study is the lack of a winter peak for respiratory and all-cause mortality during the year immediately following the intervention. Much could be made of this lack of a winter peak, but no discussion of the potential impact of (or a lack of) influenza epidemics is provided. These issues complicate the interpretation of the estimated decline in upward trend of mortality rate or the apparent lack of winter peak.

The decline in mortality following the intervention does not preclude the possibility that other constituents of the pollution mixture that share the same source as SO₂ are responsible for the adverse effects. Even though PM₁₀ levels before and after the intervention were stable in Hong Kong, it is possible that constituents that do not explain a major fraction of PM may have declined. As also noted previously in Section 3.1.6, Hedley et al. (2006) noted large reductions in ambient nickel and vanadium concomitantly with reductions of sulfur after the intervention. SO₂ also may be serving as a modifier of the effect of respirable particles. Thus, while the Hong

Kong data are supportive of SO₂-mortality effects, the possibility remains that mortality effects may be caused by constituents of SO₂-associated sources.

3.3.4. Summary of Evidence on the Effect of Short-Term SO₂ Exposure on Mortality

The epidemiological evidence on the effect of short-term exposure to SO₂ on all-cause (nonaccidental) and cardiopulmonary mortality is *suggestive but not sufficient to infer a causal relationship* at ambient concentrations. The epidemiological studies are generally consistent in reporting positive associations between SO₂ and mortality; however, there was a lack of robustness of the observed associations to adjustment for copollutants.

Figure 3-11 presents all-cause SO₂ mortality excess risk estimates from the multicity studies and meta-analyses. The mortality effect estimates from single-pollutant models range from 0.6% (NMMAPS) to 4.1% (Italian 8-cities study) per 10 ppb increase in 24-h avg SO₂ concentrations, but given the large confidence band in the Italian study, a more stable range may be 0.6 to 2%. It is noteworthy that the SO₂ effect estimates for the NMMAPS and Canadian 12-city studies are quite comparable (0.6 and 0.7%, respectively), considering the differences in the modeling approach. The heterogeneity of estimates in the multicity studies and meta-analyses may be due to several factors, including the differences in model specifications, averaging/lag time, SO₂ levels, and effect-modifying factors. Only the APHEA study examined possible sources of heterogeneity for SO₂-related mortality. They examined several potential effect modifiers such as the mean levels of pollution and weather variables, accuracy of the air pollution measurements, health of the population, smoking prevalence, and geographical differences. The only variable that could explain the heterogeneity of city-specific effect estimates was the geographic separation (western versus central eastern European cities) for both SO₂ and BS, but heterogeneity in the SO₂ effect estimates remained within the western cities.

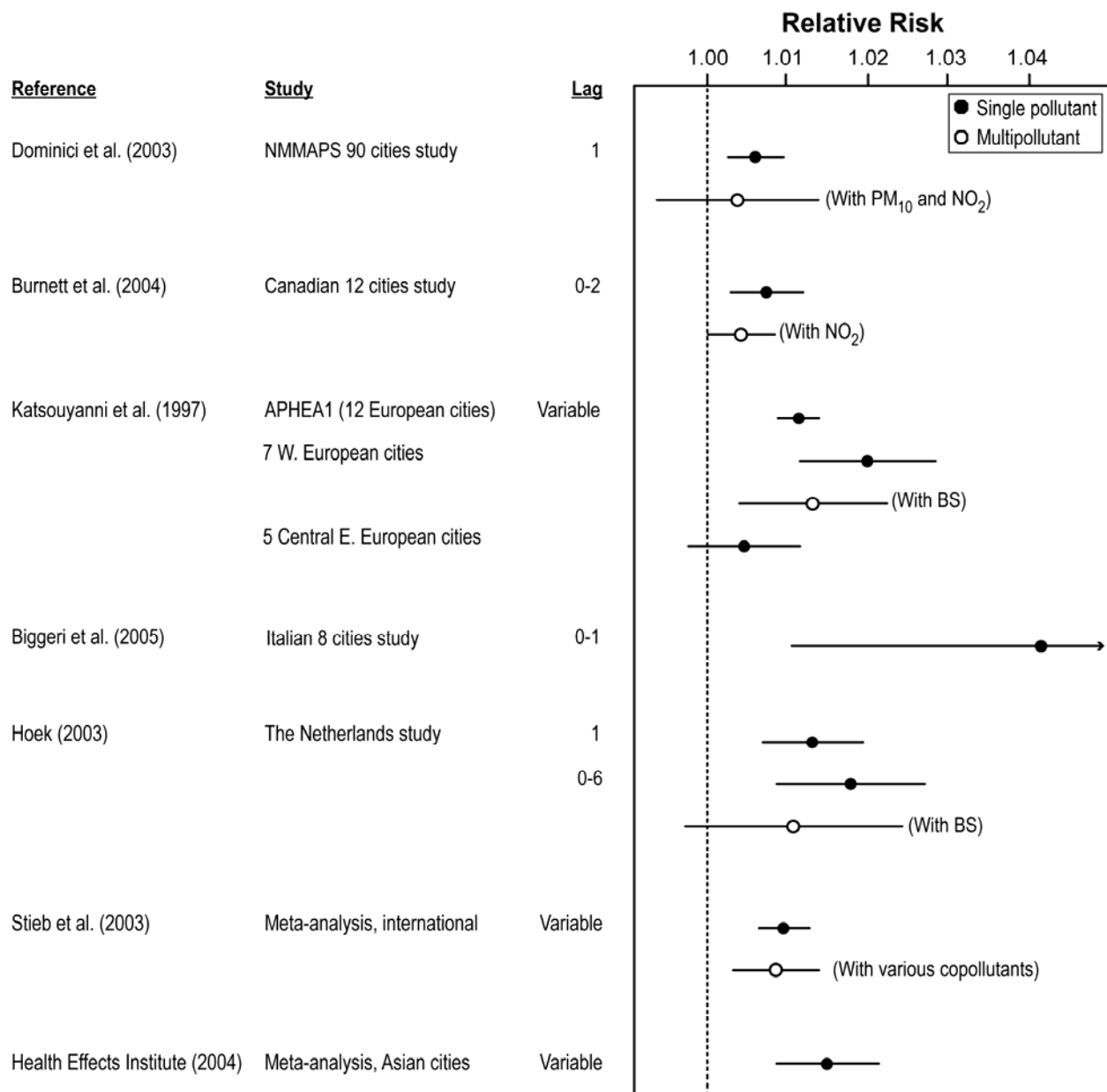


Figure 3-11. Relative risks (95% CI) of SO₂-associated all-cause (nonaccidental) mortality, with and without copollutant adjustment, from multicity and meta-analysis studies. Effect estimates are standardized per 10 ppb increase in 24-h avg SO₂ concentrations. For multipollutant models, results from the models that resulted in the greatest reduction in the SO₂ effects are shown. (NMMAPS = National Morbidity, Mortality, and Air Pollution Study; APHEA = Air Pollution and Health: a European Approach)

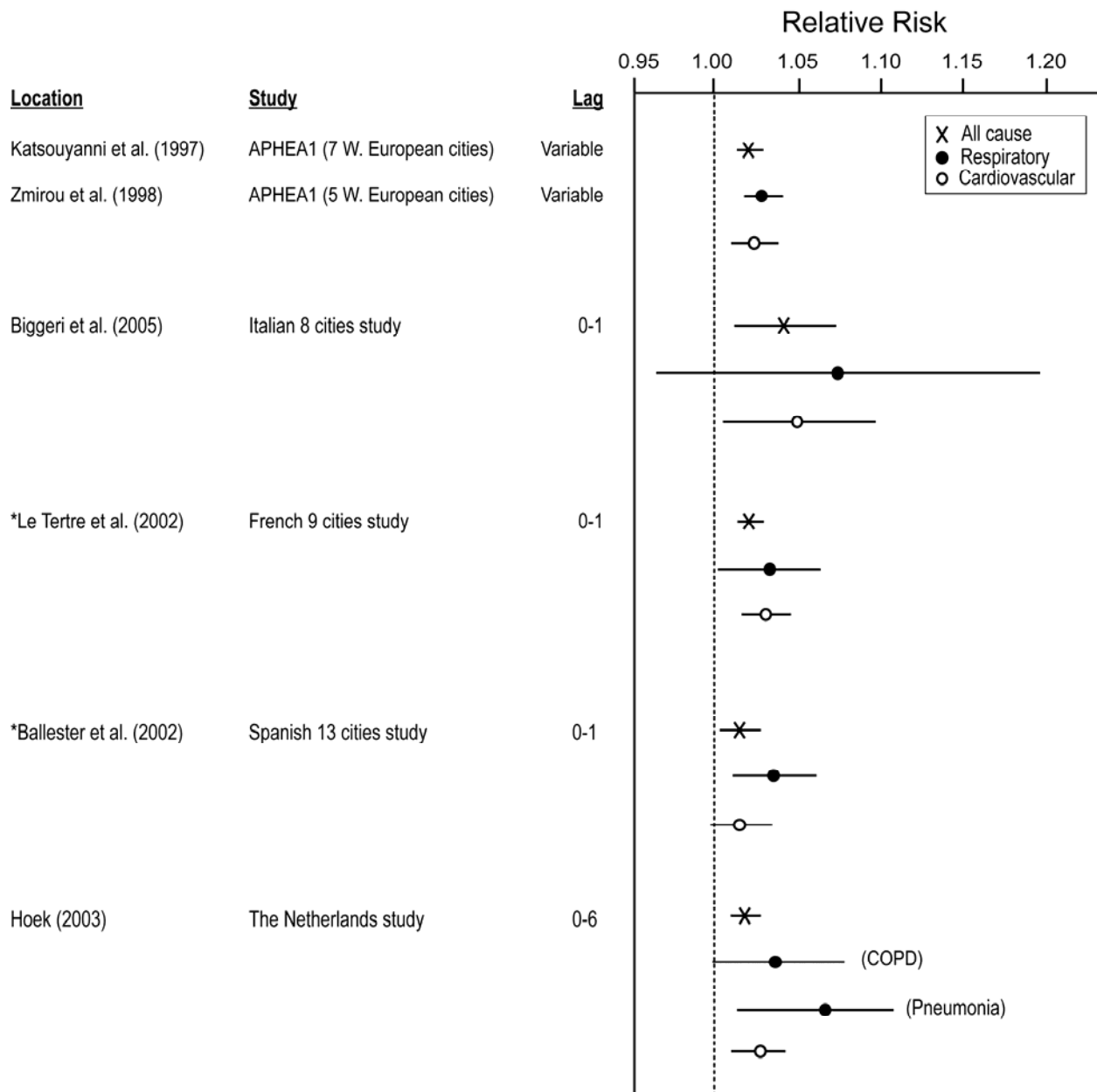


Figure 3-12. Relative risks (95% CI) of SO₂-associated mortality for all (nonaccidental), respiratory, and cardiovascular causes from multicity studies. Effect estimates are standardized per 10 ppb increase in 24-h avg SO₂ concentrations. (APHEA = Air Pollution and Health: a European Approach)

*Note: Le Tertre et al. (2002) and Ballester et al. (2002) performed analyses using Poisson GAM with default convergence criteria.

Several multicity studies provided effect estimates for broad cause-specific categories, typically respiratory and cardiovascular mortality. A summary of these effect estimates, along with the all-cause mortality estimates for comparison, are presented in Figure 3-12. These results from multicity studies suggest that the mortality effect estimates for cardiovascular and respiratory causes were generally larger than that for all-cause mortality, though in some cases the effects were not statistically significant, possibly because of reduced statistical power by which to examine cause-specific associations. In these studies, the effect estimates for respiratory mortality were also found to be larger than the cardiovascular mortality effect estimates, suggesting a stronger association of SO₂ with respiratory mortality compared to cardiovascular mortality. This suggestive finding is consistent with the observed greater effects of SO₂ on respiratory morbidity compared to cardiovascular morbidity.

As shown previously in Figure 3-11, the mortality effect estimates from the multipollutant models in the multicity studies suggest some extent of confounding between SO₂ and PM and/or results from multicity studies generally suggest some evidence of confounding, in the sense of instability of effect estimates in multipollutant models. It should be noted, however, that interpretation of the single- versus multipollutant model results are complicated by potential interaction among copollutants and differing degrees of measurement error for correlated pollutants.

Very few studies specifically examined possible interactions among the copollutants. Katsouyanni et al. (1997) examined the effect estimates for SO₂ and BS in seven western European cities for subsets stratified by high and low levels of the other pollutant and found that the estimates were similar for days with low or high levels of the other pollutant. From these results, Katsouyanni et al. suggested that the effects of SO₂ and BS were independent.

Other multi- or single-city studies did not consider examination of possible interaction effects between SO₂ and copollutants.

In summary, recent epidemiological studies have reported associations between mortality and SO₂, often at mean 24-h avg levels of < 10 ppb. The range of the excess risk estimates for SO₂ on all-cause mortality is 0.4 to 2% per 10 ppb increase in 24-h avg SO₂ in several multicity studies and meta-analyses. The effect estimates for more specific categories may be larger. The larger European study suggests that the observed heterogeneity in SO₂ effect estimates is at least in part regional. The intervention study from Hong Kong supports the idea that a reduction in

SO₂ levels results in a reduction in deaths, but this does not preclude the possibility that the causal agent is not SO₂ but rather something else that is associated with SO₂ sources. Results from the multicity studies suggest that SO₂-mortality excess risk estimates may be confounded to some extent by copollutants, making a definitive distribution of effects among the pollutants difficult. However, the interpretation of multipollutant model results also requires caution because of possible interaction among the copollutants and influence of varying measurement error. Very limited information was available to determine possible interaction effects between SO₂ and PM or other copollutants. Overall, the evidence that SO₂ is causally related to mortality at current ambient levels is suggestive, but limited by potential confounding and lack of understanding regarding the interaction of SO₂ with copollutants in the epidemiological data.

3.4. Morbidity Associated with Long-Term SO₂ Exposure

3.4.1. Summary of Findings from the Previous Review

The 1982 AQCD addressed some effects of long-term SO₂ exposure. It was reported that bronchoconstriction resulted from chronic exposure to 5.1 ppm SO₂ in dogs but not in monkeys. This increased pulmonary resistance was thought to occur as a result of morphological changes in the airway or hypersecretion of mucus leading to airway narrowing. However, there were no remarkable pulmonary pathological findings in monkeys and dogs in these studies. This could have been due to the conventional light microscopic examination applied, which could not detect alterations in surface membranes or subtle changes in cilia.

It was also noted that repeated exposures of rats ≥ 50 ppm SO₂ produced a chronic bronchitis similar to that seen in humans although there was no evidence to suggest that bronchitis developed in humans at ambient levels of SO₂. Furthermore, nasal mucosal alterations were observed in mice exposed to 10 ppm SO₂ for 72 h by inhalation. Lack of data on morphological effects of SO₂ at near ambient concentrations was noted. In addition, some alterations in lung host defenses were discussed with chronic exposure to SO₂ at doses exceeding ambient concentrations.

In the 1982 AQCD, only a few epidemiological studies provided sufficient quantitative evidence relating respiratory symptoms or pulmonary functions changes to long-term exposure

1 to SO₂. Briefly, a study by Lunn et al. (1967) in Sheffield, England, provided the strongest
2 evidence of an association between pulmonary function decrements and increased frequency of
3 lower respiratory tract symptoms in 5- to 6-year-old children chronically exposed to ambient BS
4 (annual level of 230 to 301 µg/m³) and SO₂ levels (69 to 105 ppb). A follow-up study in 1968 by
5 Lunn (1970) found no effect with much lower levels of BS (range: 48, 169 µg/m³) and SO₂
6 (range: 36, 97 ppb); it was suggested that this might be due to insufficient power to detect small
7 health effect changes.

8 The 1986 Second Addendum presented three additional studies that examined the effects of
9 long-term exposure on respiratory health. A study by Ware et al. (1986) reported that respiratory
10 symptoms were associated with annual average TSP in the range of ~30 to 150 µg/m³ in children
11 (n = 8,380) from six U.S. studies. Only cough was found to be significantly associated with SO₂.
12 Although the increase in symptoms did not appear concomitantly with any decrements in lung
13 function, this may indicate different mechanisms of effect. Other studies by Chapman et al.
14 (1985) and Dodge et al. (1985) also observed increased prevalence of cough among children and
15 young adults living in areas of higher SO₂ concentrations; however, it was noted that the
16 observed effects might have been due to intermittent high SO₂ peak concentrations.

17 In addition to respiratory effects from long-term exposure to SO₂, the potential
18 carcinogenicity of SO₂ or other SO_x was also examined in the previous review. The 1982 AQCD
19 concluded that little or no clear epidemiological evidence substantiated the hypothesized links
20 between SO₂ or other SO_x and cancer, though there was some animal toxicological evidence that
21 led to the conclusion that SO₂ may be considered a suspect carcinogen/cocarcinogen. There was
22 very limited consideration of the effects of long-term exposure to SO₂ on other organ systems.

23 Since the 1982 AQCD and the 1986 Second Addendum, a number of animal toxicological
24 and epidemiological studies have investigated the effect of long-term exposure to SO₂ on
25 respiratory morbidity, including asthma, bronchitis and respiratory symptoms, lung
26 function, morphological effects, and lung host defense. Additional studies have examined the
27 effect of long-term SO₂ exposure on genotoxic and carcinogenic effects, cardiovascular effects,
28 and prenatal and neonatal outcomes, which are also briefly discussed in this section.

3.4.2. Respiratory Effects Associated with Long-Term Exposure to SO₂

3.4.2.1. Asthma, Bronchitis, and Respiratory Symptoms

Several epidemiological studies have examined the association between long-term exposure to SO₂ and other air pollutants on asthma, bronchitis, and a variety of respiratory symptoms. These studies are summarized in Annex Table F-1. In the Six Cities Study of Air Pollution and Health, cross-sectional associations between air pollutants and respiratory symptoms were examined in 5,422 white children aged 10 to 12 years old from Watertown, MA, St. Louis, MO, Portage, WI, Kingston-Harriman, TN, Steubenville, OH, and Topeka, KS (Dockery et al., 1989). Annual means of 24 h avg SO₂ concentrations ranged from 3.5 ppb in Topeka to 27.8 ppb in Steubenville. Except for O₃, the correlations among pairs of pollution measures varied between 0.53 and 0.98. No associations were observed between SO₂ and a variety of respiratory symptoms, including bronchitis, chronic cough, chest illness, persistent wheeze, and asthma. Stronger associations were observed for PM indices.

Dockery et al. (1996) examined the respiratory health effects of acid aerosols in 13,369 white children aged 8 to 12 years old from 24 communities in the United States and Canada between 1988 and 1991. The city-specific annual mean SO₂ concentration was 4.8 ppb, with a range of 0.2 to 12.9 ppb. With the exception of the gaseous acids, nitrous and nitric acid, none of the particulate or gaseous pollutants, including SO₂, were associated with increased asthma or any asthmatic symptoms. Stronger associations with particulate pollutants were observed for bronchitis and bronchitic symptoms. For SO₂, the only significant association found was with chronic phlegm, with an OR of 1.19 (95% CI: 1.00, 1.40) per 5 ppb increase in SO₂.

Herbarth et al. (2001) performed a meta-analysis of three cross-sectional surveys conducted in East Germany investigating the relationship between lifetime exposure (from birth to completion of questionnaire survey) to SO₂ and TSP in children and the prevalence of chronic bronchitis. Using a logistic model that included variables on parental predisposition (mother or father with bronchitis) and environmental tobacco smoke exposure, the authors reported that the OR for bronchitis due to a lifetime exposure to SO₂ was 3.51 (95% CI: 2.56, 4.82) (the concentration change for which the OR was based was not presented). No associations were found between TSP and the prevalence of bronchitis in children.

As part of the international SAVIAH (Small-Area Variation in Air Pollution and Health) study, Pikhart et al. (2001) examined the respiratory health effects from long-term exposure to SO₂ in children (n = 6,959) from two central European cities with high pollution levels (Prague, Czech Republic, and Poznan, Poland). A novel technique was used to estimate the outdoor concentrations of SO₂ at a small-area level. Outdoor SO₂ was measured by passive samplers at 130 sites in the two cities during 2-week periods. Concentrations of SO₂ at each location in the study areas were estimated from these data by modeling using a geographic information system (GIS). The estimated mean exposure to outdoor SO₂ was 32 ppb, (range: 25, 37) in Prague and 31 ppb, (range: 17, 53) in Poznan. The prevalence of wheezing or whistling in the past 12 months was associated with SO₂ (OR of 1.08 [95% CI: 1.03, 1.13] per 5 ppb increase in SO₂). Moreover, the lifetime prevalence of wheezing or whistling (OR 1.03 [95% CI: 1.00, 1.07]) and lifetime prevalence of physician-diagnosed asthma (OR 1.09 [95% CI: 1.00, 1.19]) also were associated with SO₂ levels. In the SAVIAH study, the only other pollutant considered in relation to health outcomes was NO₂. An earlier publication by Pikhart et al. (2000) presented preliminary results of the Prague data and indicated that the observed associations between NO₂ and respiratory symptoms were generally similar to that of SO₂.

The International Study of Asthma and Allergies in Children (ISAAC) included thousands of children in several European countries and Taiwan (Hirsch et al., 1999; Hwang et al., 2005; Penard-Morand et al., 2005; Ramadour et al., 2000; Studnicka et al., 1997). Pénard-Morand et al. examined the effect of long-term exposures to air pollution and prevalence of exercise-induced bronchial reactivity (EIB), flexural dermatitis, asthma, allergic rhinitis, and atopic dermatitis in 9,615 children aged 9 to 11 years in six French communities. Using 3-year averaged concentrations of SO₂, the investigators reported that the prevalence of exercise-induced bronchial reactivity, lifetime asthma, and allergic rhinitis were significantly associated with increases in SO₂ exposure. The estimated 3-year averaged concentration of SO₂ was 2 ppb in the low-exposure schools and 4 ppb in the high-exposure schools. In a single-pollutant model, the ORs were 2.37 (95% CI: 1.44, 3.77) for EIB and 1.58 (95% CI: 1.00, 2.46) for lifetime asthma per 5 ppb increase in SO₂. In this study, SO₂ was correlated with PM₁₀ (r = 0.76) but not with O₃ (r = -0.02). Using a two-pollutant model that included PM₁₀, the associations of SO₂ with EIB and lifetime asthma were fairly robust (< 5% change).

1 In a German study of 5,421 children, the annual mean SO₂ concentration was associated
2 with morning cough reported in the last 12 months, but not bronchitis (Hirsch et al., 1999). This
3 study further observed that the association of SO₂ and other air pollutants with respiratory
4 symptoms were stronger in nonatopic than in atopic children. The authors noted that these
5 findings were in line with the hypothesis that these air pollutants induce nonspecific irritative
6 rather than allergic inflammatory changes in the airway mucosa, as irritative effects would affect
7 the clinical course in nonatopic children more strongly than in atopics whose symptoms are also
8 determined by allergen exposure.

9 In contrast to the studies noted above, other studies using the ISAAC protocol did not
10 observe an association between long-term exposure to SO₂ and respiratory symptoms. In France,
11 (Ramadour et al., 2000) performed an epidemiological survey of 2,445 children aged 13 to 14
12 years living in communities with contrasting levels of air pollution to determine the relationship
13 between long-term exposure to gaseous air pollutants and prevalence rate of rhinitis, asthma, and
14 asthma symptoms. The average SO₂ concentrations during the 2-month survey period ranged
15 from 7 ppb to 22 ppb across the seven communities. This study found no relationship between
16 the mean levels of SO₂, NO₂, or O₃ and the above-mentioned symptoms. Another study of 843
17 children from eight nonurban communities in Austria did not observe consistent associations
18 between SO₂ and prevalence of asthma and symptoms (Studnicka et al., 1997). Compared to the
19 lowest SO₂ concentration category, the ORs in the higher SO₂ concentration categories (third and
20 fourth quartiles) did not exceed one for any of the symptoms examined (wheeze, cough,
21 bronchitis, and asthma).

22 A cohort study was conducted by (Goss et al., 2004) to examine the effect of air pollutants
23 on a potentially susceptible population, patients with cystic fibrosis. Study participants included
24 11,484 patients (mean age 18.4 years) enrolled in the Cystic Fibrosis Foundation National Patient
25 Registry in 1999–2000. Exposure was assessed by linking air pollution values from ambient
26 monitors with the patient's home ZIP code. During the study period, the mean SO₂ concentration
27 was 4.9 ppb (SD 2.6, IQR: 2.7, 5.9). This study found no association between SO₂ and the odds
28 of having two or more pulmonary exacerbations. One of the limitations addressed by the authors
29 was the lack of information regarding tobacco use or environmental tobacco smoke, an important
30 risk factor for pulmonary exacerbations.

Several studies examined the effects of long-term exposure to SO₂ on asthma, bronchitis, and respiratory symptoms. The studies reported positive associations in children; the notable exception was the Harvard Six Cities Study. However, there were inconsistencies in the results observed: some found effects on bronchitic but not asthmatic symptoms; others found the converse. A major limitation was that some subjects were asked to recall prevalence of symptoms in the last 12 months or in a lifetime; such long recall periods may have caused significant recall bias. Another concern is the high correlation of long-term average SO₂ and copollutant concentrations, particularly PM, and the very limited evaluation of potential confounding in these studies. Overall, while the evidence is suggestive, the variety of outcomes examined and the inconsistencies in the observed results make it difficult to assess the direct impact of long-term exposure of SO₂ on asthma, bronchitis, or respiratory symptoms.

3.4.2.2. Lung Function

Only a few new animal toxicological studies involving longer-term inhalation exposures to SO₂ were conducted since the last review. These studies are summarized here and in Annex Table E-1. Rabbits that were neonatally immunized to *Alternaria tenuis* and exposed to 5 ppm SO₂ for 13 weeks beginning in the neonatal period (Douglas et al., 1994) did not demonstrate alterations in lung resistance, dynamic compliance, trans-pulmonary pressure, tidal volume, respiration rate or minute volume. Similarly, no changes in physiological function were noted in dogs exposed to 15 ppm SO₂ for 2 h/day and 4-5 days/week for 5 months (Scanlon et al., 1987), although changes were noted at 50 ppm. However, Smith et al. (1989) found decreased residual volume and quasistatic compliance in rats at 4 months of exposure to 1 ppm SO₂ for 5 h/day and 5 days/week.

Only a limited number of epidemiological studies examined the association between long-term exposure to SO₂ and changes in lung function. The Harvard Six Cities Study by Dockery et al. (1989) reported that no associations were observed between lung function and long-term exposure to air pollution, including SO₂, in a cohort of more than 5,000 children. An analysis of NHANES II data by Schwartz (1989), which included information on children and youths from 44 cities but was limited by a cross-sectional study design, also did not observe an association with SO₂, though inverse associations of FVC and FEV₁ with annual concentrations of TSP, NO₂ and O₃ were found. Additional studies conducted in Europe observed mixed results.

1 In a longitudinal cohort study of 1,150 children in nine communities in Austria, Frischer
2 et al. (1999) examined the effect of long-term exposure to air pollutants on lung function. Lung
3 function was measured in the spring and fall over a 3-year period from 1994 through 1996.
4 Annual mean SO₂ concentrations ranged from 2 to 6 ppb across the nine communities. The
5 authors reported no consistent associations between SO₂, PM₁₀, or NO₂ and lung function. For
6 SO₂, a negative parameter estimate was observed during the summer, but a positive estimate was
7 found during the winter period. Horak et al. (2002a; b) extended the study of Frischer et al.
8 (1999) with an additional year of data. The mean SO₂ concentration was 6 ppb in the winter and
9 3 ppb in the summer. This study found a positive association between wintertime SO₂
10 concentrations and changes in FVC, which became null with PM₁₀ in a two-pollutant model.

11 Jedrychowski et al. (1999) conducted a prospective cohort study of 1,001 preadolescent
12 children from two areas of Krakow, Poland, that differed in ambient air pollutants. In the city
13 center, which had higher pollution area, the mean annual level of SO₂ was 16.7 ppb (SD 12.5). In
14 comparison, the mean annual SO₂ level in the control area was 12.1 ppb (SD 8.4). A similar
15 difference in TSP levels was observed between the city center and control area. The adjusted
16 ORs comparing the city center to the control area for the occurrence of slower lung function
17 growth over a two-year period were 2.10 (95% CI: 1.27, 3.46) for FVC and 2.10 (95% CI: 1.27,
18 3.48) for FEV₁ in boys. The adjusted ORs for girls were 1.54 (95% CI: 0.89, 2.64) for FVC and
19 1.51 (95% CI: 0.90, 2.53) for FEV₁. However, as both TSP and SO₂ levels were higher in the city
20 center, the observed effects on lung function growth cannot be specifically attributable to SO₂.

21 One notable study examined the potential effect of long-term exposure to air pollution on
22 lung function in adults. The study by Ackermann-Lieblich et al. (1997) included 9,651 adults
23 aged 18 to 60 years old residing in eight different areas in Switzerland (Study on Air Pollution
24 and Lung Diseases in Adults [SAPALDIA]). They observed a 0.1% decrease in FEV₁ per 5 ppb
25 increase in SO₂ for adults. Significant associations also were observed for PM₁₀ and NO₂. The
26 limited number of study areas and high intercorrelation between the pollutants made it difficult
27 to assess the effect of an individual pollutant. The authors concluded that air pollution from fossil
28 fuel combustion, which was the main source of air pollution for SO₂, NO₂, and PM₁₀ in
29 Switzerland, was associated with decrements in lung function parameters in this study.

Collectively, the results from the limited number of animal toxicological and epidemiological studies do not give support to long-term exposure to ambient SO₂ having a detrimental effect on lung function.

3.4.2.3. Morphological Effects

Three animal toxicological studies of morphological effects resulting from subacute to chronic SO₂ exposures have been published since the 1982 AQCD. These studies are summarized in Annex Table E-11. No alveolar lesions (including electron microscopic evaluation) or changes in numbers of tracheal secretory cells were observed in guinea pigs exposed to 1 ppm SO₂ for 3 h/day for 6 days (Conner et al., 1985). No pulmonary or nasal lesions were observed in rats exposed to 5 ppm SO₂ for 5 days/week for 4 weeks (Wolff et al., 1989). A weakness of the latter study is that histopathological methods were not reported. However, a third study reported histopathological changes in the respiratory system involving lesions in the bronchioles. Smith et al. (1989) exposed rats for 4 to 8 months to 1 ppm SO₂ for 5 h/day and 5 days/week and observed increased incidence of bronchiolar epithelial hyperplasia and a small increase (12%) in numbers of nonciliated epithelial cells in terminal respiratory bronchioles at 4 but not 8 months of exposure. A limitation of the study was the examination of a single concentration, which does not allow for concentration-response assessment or identification of a no-effect-level.

In summary, results from these animal toxicological studies do not support an association between long-term exposure to ambient SO₂ and prolonged effects on lung morphology.

3.4.2.4. Lung Host Defense

The 1982 AQCD reported some detrimental effects of SO₂ on lung host defenses that generally occurred at concentrations exceeding ambient exposure concentrations. In rats exposed to 0.1 ppm SO₂ for ~2 to 3 weeks, clearance of labeled particles from the lung was accelerated at 10 and 23 days following exposure. In rats exposed to 1 ppm for ~2 to 3 weeks, clearance was accelerated at 10 days and slowed down at 25 days. Tracheal mucus flow was decreased with a 1-year exposure of dogs to 1 ppm SO₂, but was unaffected by a 30-minute exposure of donkeys to 25 ppm SO₂. Studies in mice suggested no effect on susceptibility to bacterial infection with

1 exposure to SO₂ concentrations of ≤ 5 ppm for 3 months. Antiviral defenses were impaired in
2 mice exposed to 7-10 ppm SO₂ for 7 days. No alterations in pulmonary immune system were
3 reported with chronic exposure of mice to 2 ppm SO₂.

4 Several studies on lung host defense have been conducted since the last review and are
5 summarized in Annex Table E-4. Only one study published after the last review evaluated
6 mucociliary clearance in rats after exposure to SO₂. In this subchronic study, no effect on
7 clearance of radiolabeled particles from the lung was observed in rats exposed to 5 ppm SO₂ for
8 2 h/day for 4 weeks (Wolff et al., 1989). These findings are in contrast to the altered clearance
9 reported in the 1982 AQCD. Three other recent studies were conducted evaluating the effects of
10 10 ppm SO₂ on immune responses.

11 In summary, animal toxicological studies do not provide much evidence for long-term
12 exposure to ambient SO₂ having detrimental effects on lung host defense.

3.4.2.5. SO₂ Interactions with PM and Other Mixtures

13 An elegant series of experiments was conducted in dogs exposed to 0.31 mg/m³ neutral
14 sulfite aerosol for 22.5 h/day for 290 days (Heyder et al., 1992). The aerosol particles were
15 submicron in size. These studies are summarized in Annex Table E-14. Although sulfite particles
16 are not usually found in nature, they were engineered and used in these studies for the purpose of
17 delivering SO₂-like reactivity to the lower respiratory tract. It should be noted that the reactivity
18 of SO₂ is due to the IV-valent sulfur, a feature shared by sulfite but not sulfate which has VI-
19 valent sulfur. The concentration of sulfite particles used in these studies was comparable to
20 ambient levels of SO₂ on smog-alert days in Germany (i.e., 0.25 ppm). Important findings from
21 these studies included a significant decrease in specific lung compliance and increase in alveolar-
22 capillary permeability in sulfite-exposed dogs compared with controls (Maier et al., 1992; Schulz
23 et al., 1992). In addition, macrophage respiratory burst activity and phagocytic capacity were
24 significantly decreased while intrapulmonary particle transport to the larynx was increased
25 (Maier et al., 1992; Kreyling et al., 1992). Morphological effects included hyperplastic changes
26 in the respiratory mucosa of the nasal cavity and a moderate mononuclear cell infiltration. Loss
27 of cilia in larynx and trachea was also noted. Some of the dogs also exhibited changes in the
28 larynx and trachea (Takenaka et al., 1992). The authors concluded that chronic exposure to a low
29 dose of sulfur (IV) aerosols can initiate a pathophysiological response.

1 A second set of studies was conducted by these same investigators in dogs exposed to
2 sulfite and sulfate aerosols for 13 months (Heyder et al., 1999). These are summarized in Annex
3 Table E-14. This protocol involved daily exposures of 16.5 h neutral sulfite aerosol at the same
4 concentration used in the previous study followed by 6 hrs of an acidic sulfate aerosol at a
5 concentration of 15.2 $\mu\text{mol}/\text{m}^3$ hydrogen ions. Both aerosols were about 1 μm MMAD in size.
6 The authors stated that the dose received by each dog in 13 months was equivalent to what a
7 person living for 70 years in an urban environment would receive. Results of these experiments
8 demonstrated no change in lung compliance or other measure of lung function in dogs exposed
9 consecutively to sulfite and sulfate each day (Schulz et al., 1999). Alveolar-capillary
10 permeability was no different than in controls (Maier et al., 1999). Intrapulmonary particle
11 transport to the larynx was decreased while transport to the tracheobronchial lymph nodes was
12 increased in dogs exposed to both sulfite and sulfate (Kreyling et al., 1999). No alteration in the
13 surfactant system was observed (Griese et al., 1999). Slight morphological effects were observed
14 in the proximal alveolar region but not in the nasal cavity, larynx or trachea (Takenaka et al.,
15 1999). The authors attributed this milder response to a modulating effect of the acidic sulfate
16 aerosol. They concluded that inhalation of low levels of sulfite and hydrogen ion is not likely to
17 constitute a health risk. These results are somewhat surprising given the pathophysiologic
18 response to sulfite alone found by these same authors in a similar model. Possibly they indicate
19 an antagonistic interaction between sulfate and sulfite.

20 In addition to studies examining the interaction of SO_2 and particles, other animal studies
21 performed since the 1982 AQCD involved binary mixtures, laboratory-generated complex
22 mixtures (e.g., simulation of regional air pollution), or actual ambient air mixtures (Annex Tables
23 E-18 through E-20). Generally, most studies with ambient or laboratory-generated complex
24 mixtures did not include an SO_2 -only exposure group, making it difficult to determine the
25 contribution of SO_x . No definitive conclusions can be made from these studies.

3.4.2.6. Summary of Evidence on the Effect of Long-Term Exposure on Respiratory Health

26 The overall epidemiological evidence on the respiratory effects of long-term exposure to
27 SO_2 is *inadequate to infer the presence or absence of a causal relationship*. Studies that
28 examined the effects of long-term exposure to SO_2 on asthma, bronchitis, and respiratory

1 symptoms observed positive associations in children. However, the variety of outcomes
2 examined and the inconsistencies in the observed results make it difficult to assess the impact of
3 long-term exposure of SO₂ on respiratory symptoms. In the limited number of studies examining
4 the SO₂ associations with lung function, results were generally mixed. A major consideration in
5 evaluating SO₂-related health effects in these epidemiological studies of long-term exposure is
6 the high correlation among the pollutant levels observed, particularly between long-term average
7 SO₂ and PM concentrations. The lack of evidence available to evaluate potential confounding by
8 copollutants limits the ability to make a causal determination based on these studies.

9 A limited number of animal toxicological have examined the effect of long-term exposure
10 to SO₂ on lung function. Results from these studies do not provide strong biological plausibility
11 for effects of long-term exposure to SO₂ on respiratory morbidity. These studies observed no
12 effects on physiological lung function at SO₂ concentrations ≤ 5 ppm in rabbits and dogs;
13 however, one study found decreased residual volume and quasistatic compliance at 1 ppm SO₂ in
14 rats. In addition, no morphological changes were found in guinea pigs exposed subacutely to
15 1 ppm SO₂, or in rats exposed subchronically to 5 ppm SO₂. While mild, bronchiolar epithelial
16 hyperplasia was observed in rats exposed to 1 ppm for 4 months, this change was not apparent at
17 8 months. Furthermore, animal toxicological studies provide no evidence for decrements in lung
18 host defense at or near ambient levels of SO₂.

19 Overall, results from the generally limited number of epidemiological and animal
20 toxicological studies do not give support to respiratory effects from long-term exposure to SO₂ at
21 ambient concentrations. However, chronic studies in dogs exposed to sulfite particles at
22 concentrations equivalent to near ambient levels of SO₂ demonstrated a mild pathophysiologic
23 response, suggesting that deposition of SO_x in the lower respiratory tract may lead to more
24 profound effects on the respiratory system than those observed with gaseous SO₂ alone. These
25 changes were modulated and in some cases reversed by sequential exposure to sulfate particles,
26 suggesting an antagonistic interaction among the different PM in the mixture.

3.4.3. Carcinogenic Effects Associated with Long-Term Exposure

27 The 1982 AQCD concluded that little or no clear epidemiological evidence substantiated
28 the hypothesized links between SO₂ or other SO_x and cancer. From the toxicological studies, it
29 was noted that while there were some indications of carcinogenicity for both SO₂ and SO₂ +

benzo[*a*]pyrene (B[*a*]P), complex exposure regimens, problematic dose determinations, and/or inadequately reported experimental details led to the conclusion that SO₂ could only be considered a suspect carcinogen/cocarcinogen.

Since the last review, numerous studies have examined the genotoxic effects of SO₂. These are summarized in Annex Table E-22. SO₂ and its metabolite sulfite were found not to be mutagenic or to induce DNA damage in vitro (Pool et al., 1988; Pool-Zobel et al., 1990). However, inhalation studies demonstrated increased mouse bone marrow micronucleated polychromatic erythrocytes and DNA damage in cells isolated from various organs when mice were exposed for 4-6 h/day for 7 days to 5-30 ppm SO₂ (Meng et al., 2002; 2005; Ruan et al., 2003). These in vivo studies suggest that inhaled SO₂ may have systemic effects at high concentrations, but they are of questionable significance in evaluating the effects of SO₂ at ambient levels.

The carcinogenic potential of SO₂ was examined in animal toxicological studies which are summarized in Annex Table E-11. Gunnison et al. (1988) conducted a two-part study in which rats were exposed either for 21 weeks (6 h/day, 5 days/week) by inhalation to 0, 10, or 30 ppm SO₂, or for 21 weeks to two tungsten-supplemented, molybdenum-deficient diets. This latter regimen induced a condition of sulfite oxidase deficiency, resulting in elevated systemic levels of sulfite:bisulfite relative to control values (e.g., in plasma, from 0 to 44 μM; and in tracheal tissue, from 33 to 69 or 550 nmol/g wet weight). Beginning with week 4, some groups from each regimen received weekly tracheal installations of 1-mg B[*a*]P for 15 weeks. Overall results indicated that squamous cell carcinoma was not induced, or in the B[*a*]P groups coinduced or promoted, by SO₂ inhalation or elevated systemic sulfite:bisulfite. Researchers found a very high incidence of animals with tumors in the groups exposed to only B[*a*]P (128/144). As a result, carcinogenicity or cocarcinogenicity of SO₂ or sulfite:bisulfite could only have been detected as a shortening of tumor induction time or an increase in rate of tumor appearance, and neither was observed. As noted by the authors, these findings do not support the conclusion that SO₂ exposure enhances the carcinogenicity of B[*a*]P. It was proposed that SO₂ exposure, by elevating systemic sulfite:bisulfite, would generate glutathione-*S*-sulfonates, which in turn could inhibit glutathione *S*-transferase (GST) and reduce intracellular GSH and, thus, interfere with a major detoxication pathway for B[*a*]P. See Annex Table E-21 for further discussion from the work of Menzel et al., (1986).

Two similar studies were published that investigated the ability of 10 to 11 months of exposure (16 h/day) to 4 ppm SO₂, 6 ppm NO₂, or their combination to affect the carcinogenicity of either urban suspended PM (SPM) (Ito et al., 1997) or diesel exhaust particle (DEP) (Ohyama et al., 1999) extract-coated carbon particles. The former study found that, while exposure to SPM extract-coated carbon particles significantly increased pulmonary endocrine cell (PEC) hyperplasia, coexposure to SO₂, NO₂, or their combination was without additional affect. Also, irrespective of gas coexposure, SPM extract-coated carbon particles demonstrated a few PEC papillomas versus control frequencies of zero.

Using Syrian golden hamsters, Heinrich et al. (1989) investigated whether coexposure to 10 ppm SO₂ and 5 ppm NO₂ for 6 to 8 months (5 days/week, 19 hours/day) could enhance tumorigenesis induced by a single subcutaneous injection of diethylnitrosamine (DEN) during week 2. The combined gas exposure did not affect body weight gain and only minimally shortened survival times. Compared to the DEN groups, serial sacrifices of gas-exposed animals demonstrated progressively increasing numbers of tracheal mucosal cells and aberrant tracheal cell cilia. In the lung, effects related to gas mixtures were largely limited to a progressive type of alveolar lesion that involved the lining of bronchiolar epithelium and the appearance of pigment-containing AM and to a mild, diffuse thickening of the alveolar septa. Exposure to the combined gases by itself did not induce tumors of the upper respiratory tract, nor did it enhance the induction of such tumors by DEN.

In addition to the animal toxicological studies that examined the genotoxic and carcinogenic potential of SO₂, a limited number of recent epidemiologic studies have investigated the relationship between long-term exposure to SO₂ and lung cancer incidence and mortality. These studies are summarized in Annex Table F-7. Nyberg et al. (2000) conducted a case-control study of men aged 40 to 75 years with (n = 1,042) and without (n = 2,364) lung cancer in Stockholm County, Sweden. They mapped residence addresses to a GIS database to assign individual exposures to SO₂ from defined emission sources (mainly local oil-fueled residential heating). Available SO₂ measurement data were used to calibrate the model. In this study, SO₂ was considered an indicator of air pollution from residential heating. Exposure to NO₂, considered to be a marker of traffic pollution, also was evaluated in this study. The 90th percentile 30-year average SO₂ level was 30 ppb. After adjusting for potential confounders (e.g., smoking, occupational exposures), long-term average heating-related SO₂ exposure was not

1 associated with an increase in risk of lung cancer incidence. A weak association for the 30-year
2 average traffic-related NO₂ exposure was observed.

3 Very similar results were reported in a Norwegian study by Nafstad et al. (2003). The study
4 population is a cohort of 16,209 men who enrolled in a study of cardiovascular diseases in 1972.
5 The Norwegian cancer registry identified 422 incident cases of lung cancer. SO₂ exposure data
6 were modeled based on residence using data for observed concentrations and emission from
7 point sources (e.g., industry and heating of buildings and private homes) and traffic. Once again,
8 no association was observed between long-term exposure to SO₂ and lung cancer incidence.

9 Three additional European cohort studies examined the associations between long-term
10 exposure to air pollution and lung cancer mortality (Beelen et al., 2008; Filleul et al., 2005;
11 Nafstad et al., 2004) in cohorts ranging in size from 14,284 to 120,852 subjects, who were
12 followed for 9 to > 20 years. Consistent with the results for lung cancer incidence, none of these
13 studies observed an association between long-term SO₂ exposure and lung cancer mortality.
14 These studies are discussed in further detail in Section 3.5.2.2.

15 Similar to the European cohort studies, studies conducted in the United States generally did
16 not observe an association between long-term exposure to SO₂ and lung cancer mortality. In the
17 reanalysis of the Harvard Six Cities study, Krewski et al. (2000) estimated a RR of 1.03 (95% CI:
18 0.91, 1.16) per 5 ppb increase in average SO₂ over the study period, while Pope et al. observed a
19 positive but not statistically significant (RR ~1.04 per 5 ppb increase in average SO₂ from 1982
20 to 1998) association in the extended analysis of the American Cancer Society (ACS) cohort. The
21 California Seventh-day Adventists study by Abbey et al. (1999) did observe a statistically
22 significant association between lung cancer mortality and SO₂ (and most of the pollutants
23 examined including PM₁₀, sulfate, O₃, and NO₂), but the number of lung cancer deaths in this
24 cohort was very small (12 for female, 18 for male) and, therefore, it is difficult to interpret these
25 estimates. More detailed discussions of these studies are provided in Section 3.5.2.2.

26 In conclusion, the toxicological studies indicate that SO₂ at high concentrations may cause
27 DNA damage but fails to induce carcinogenesis, cocarcinogenesis, or tumor promotion.
28 Furthermore, the epidemiological studies did not provide evidence that long-term exposure to
29 SO₂ is associated with an excess risk of lung cancer.

3.4.4. Cardiovascular Effects Associated with Long-Term Exposure

The effects of SO₂ on the cardiovascular system were not addressed in the 1982 AQCD. Since then, animal toxicological studies have reported oxidation (Meng et al., 2003) and glutathione (GSH) depletion (Langley-Evans et al., 1996; Meng et al., 2003; Wu and Meng, 2003) in the hearts of rodents which were exposed by inhalation to SO₂. However, as concentrations of SO₂ used in these studies were 5 ppm and above, the oxidative injury observed is probably not relevant to cardiovascular effects seen at ambient levels of SO₂. These and other animal toxicology studies measuring cardiovascular endpoints are summarized in Annex Table E-5.

A recent epidemiological study examined the association between long-term exposure to air pollution, including SO₂, and one or more fatal or nonfatal cardiovascular events. In the Women's Health Initiative cohort study, Miller et al. (2007) studied 65,893 postmenopausal women between the ages of 50 and 79 years without previous cardiovascular disease in 36 U.S. metropolitan areas from 1994 to 1998. Subjects' exposures to air pollution were estimated using residents' five-digit ZIP code, assigning the annual mean levels of air pollutants measured at the nearest monitor. A total of 1,816 women had one or more fatal or nonfatal cardiovascular events, including 261 deaths from cardiovascular causes. Hazard ratios for the first cardiovascular event were estimated. The results for models that only included subjects with non-missing exposure data for all pollutants (n = 28,402 subjects, resulting in 879 cardiovascular events) are described here. In the single-pollutant models, PM_{2.5} showed the strongest associations with cardiovascular events among the pollutants (Hazard Ratios = 1.24 [95% CI: 1.04, 1.48] per 10 µg/m³ increase in annual average), followed by SO₂ (1.07 [95% CI: 0.95, 1.20] per 5 ppb increase in the annual average). In the multipollutant model where all the pollutants (i.e., PM_{2.5}, PM_{10-2.5}, CO, SO₂, NO₂, O₃) were included in the model, the PM_{2.5} association with overall cardiovascular events was even stronger (1.53 [95% CI: 1.21, 1.94]). The association with SO₂ also became stronger (1.13 [95% CI: 0.98, 1.30]). Correlations among these pollutants were not described and, therefore, the extent of confounding among these pollutants in these associations could not be examined, but among all the air pollutants considered, PM_{2.5} was clearly the best predictor of cardiovascular events.

The available toxicological and epidemiological evidence to assess the effect of long-term exposure to SO₂ on cardiovascular health is too limited to make any conclusions at this time.

3.4.5. Prenatal and Neonatal Outcomes Associated with Long-Term Exposure

Several animal toxicological studies examined developmental effects of SO₂ and are summarized in Annex Table E-7. No changes in birth weight or neurobehavioral development were noted in mouse pups prenatally exposed to 5-30 ppm SO₂ (1996), while some behavioral modifications were seen in adults exposed prenatally to these same levels (Fiore et al.). However, effects observed at such high concentrations of SO₂ are of questionable relevance.

In recent years, the effects of prenatal and neonatal exposure to air pollution have been examined in epidemiologic studies by several investigators. These studies are summarized in Annex Table F-8. The most common endpoints studied are low birth weight, preterm delivery, and measures of intrauterine growth. Preterm birth and low birth weight may result in serious long-term health outcomes for the infant. Preterm birth is the leading cause of infant mortality and is a major determinant of a variety of adverse neurodevelopmental outcomes and chronic adverse respiratory effects (Berkowitz and Papiernik, 1993). Low birth weight has also been linked with increased risk of infant mortality and morbidity. Other studies have examined associations between maternal exposure to ambient air pollution and sudden infant death syndrome (SIDS) and neonatal hospitalizations.

These studies analyzed air pollution data and birth certificates from a given area. In evaluating the results of these studies, it is important to consider the limitations of these data. For example, the reliability and validity of birth certificate data have been reviewed (Buescher et al., 1993; Piper et al., 1993) and have been found to vary in degrees of reliability by specific variables. The variables considered the most reliable include birth weight, maternal age, race, and insurance status. Whereas gestational age, parity and delivery type (vaginal vs. cesarean) were reasonably reliable, obstetrical complications and maternal lifestyle factors such as smoking and alcohol consumption were not reliable. Another concern in these studies regards adequate control for potential confounders. While most of these studies adequately controlled for maternal education, parity, age, and sex of child, many did not adjust for socioeconomic status, occupational exposures, indoor pollution levels, maternal smoking, alcohol use, prenatal care, or concurrent temperature exposures as fetal growth is associated with all of these factors. This makes overall comparisons across studies a difficult task.

1 While most studies analyzed average SO₂ exposure for the whole pregnancy, many also
2 considered exposure during specific trimesters, or other time periods (e.g., first and last months
3 of gestation). Different exposure periods have been examined because the biological mechanisms
4 and timing of critical exposures that link air pollution to adverse birth outcomes are yet to be
5 determined. For example, fetal growth is much more variable during the third trimester;
6 therefore, exposure during the third trimester would have the greatest likelihood of an
7 association. However, insufficient placentation during the first trimester may be associated with
8 early environmental insult, whereby subsequent fetal growth is hindered. Similarly, it is possible
9 that preterm delivery is associated with insufficient placentation resulting from early exposure.
10 Furthermore, preterm delivery may be the result of acute exposures just prior to delivery.

11 Epidemiological studies examining the effects of air pollutants on low birth weight are
12 summarized in Figure 3-13. Maisonet et al. (2001) examined the association between air
13 pollution and low birth weight in six northeastern cities: Boston, MA; Hartford, CT;
14 Philadelphia, PA; Pittsburgh, PA; Springfield, MA; and Washington, DC. The study population
15 consisted of 89,557 singleton, full-term, live births (37-44 weeks of gestation) born between
16 January 1994 and December 1996. Low birth weight was classified as < 2,500 g (5.5 lbs.). This
17 study observed an association between low birth weight and SO₂ concentrations during each
18 trimester among Caucasians; however, the association was not consistent in other races and
19 ethnicities.

20 An excess risk for low birth weight associated with ambient SO₂ concentrations was
21 reported by Dugandzic et al. (2006) in a large cohort study of 74,284 women with full-term,
22 singleton births from 1988–2000 in Nova Scotia, Canada. The mean 24-h avg SO₂ concentration
23 over the study period was 10 ppb (IQR 7). These investigators found that exposure only during
24 the first trimester was associated with increased risk of low birth weight. The RR was 1.14 (95%
25 CI: 1.04, 1.26) per 5 ppb increase in SO₂ level.

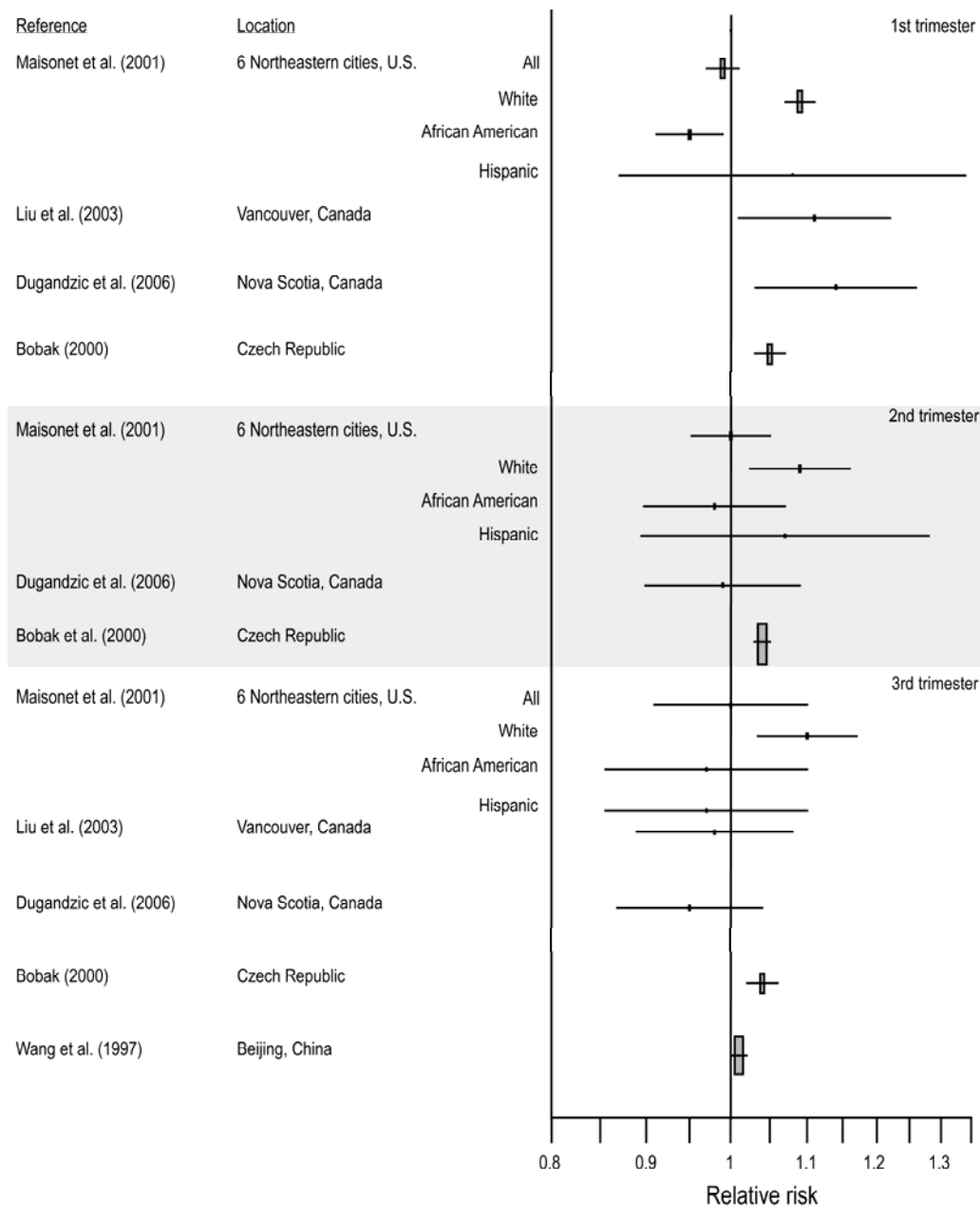


Figure 3-13. Relative risks (95% CI) for low birth weight, grouped by trimester of SO₂ exposure. Risk estimates are standardized per 5 ppb increase in SO₂ concentrations. The size of the box of the central estimate represents the relative weight of that estimate based on the width of the 95% CI.

1 Liu et al. (2003) found similar results in a study of pregnancy outcomes and air pollution in
2 Vancouver, Canada. The mean 24-h avg SO₂ concentration was 4.9 ppb (IQR 7.7) from 1985 to
3 1998. Maternal exposure during the first month was associated with an increased risk of low

1 birth weight (OR 1.11 [95% CI: 1.01, 1.22]). Additional studies from the United States, Europe,
2 Latin America and Asia have reported positive associations between low birth weight and
3 maternal exposure to SO₂ during the first (Bell et al., 2007; Bobak, 2000; Ha et al., 2001;
4 Mohorovic, 2004; Yang et al., 2003a), second (Bobak, 2000; Gouveia et al., 2003; Lee et al.,
5 2003a) and third (Bobak, 2000; Lin et al., 2004; Wang et al., 1997) trimesters.

6 Preterm delivery, intrauterine growth retardation (IUGR), and birth defects are additional
7 adverse birth outcomes that have been associated with ambient SO₂ levels. In a time-series
8 analysis using data from four Pennsylvania counties, (Sagiv et al., 2005) reported that the mean
9 6-week SO₂ exposure prior to birth was associated with increased risk of preterm birth, with a
10 RR of 1.05 (95% CI: 1.00, 1.10) per 5 ppb increase in SO₂. A 5 ppb increase in SO₂
11 concentrations three days before birth was associated with a RR of 1.02 (95% CI: 0.99, 1.05).
12 The authors discussed two plausible mechanisms for the effects of air pollution on preterm birth:
13 (1) changes in blood viscosity due to inflammation as a result of air pollution (citing Peters et al.,
14 1997); and (2) maternal infection during pregnancy as a consequence of impaired immunity from
15 air pollution exposure. Liu et al. (2003) reported that SO₂ exposure during the last month of
16 pregnancy was associated with preterm birth, with an OR of 1.09 (95% CI: 1.01, 1.19) for a
17 5 ppb increase in SO₂, in Vancouver, Canada. Similar results were found for studies conducted in
18 the Czech Republic (Bobak, 2000), Korea (Leem et al., 2006), and Beijing (Xu et al., 1995).

19 Liu et al. (2003) further reported that SO₂ exposure during the last month of pregnancy
20 was associated with IUGR (OR 1.07 [95% CI: 1.01, 1.13]). However, in a later study in the
21 Canadian cities of Calgary, Edmonton and Montreal, (Liu et al., 2007) did not observe
22 associations between maternal exposure to SO₂ and increased risk of IUGR.

23 Two Brazilian studies examined exposure to SO₂ and neonatal deaths. Pereira et al. (1998)
24 found a positive association between SO₂ and intrauterine mortality in São Paulo during a 2-year
25 period, though the effect was sensitive to model specifications and did not support a
26 concentration-response relationship. The most robust association was observed for an index of
27 three gaseous pollutants (NO₂, SO₂, CO) with mortality. Lin et al. (2004) found that a 5 ppb
28 increase in SO₂ was associated with an increase of 8.8% (95% CI: 5.8, 11.8). A similar
29 relationship was found for PM₁₀. The creation of an index containing both PM₁₀ and SO₂ allowed
30 the observation of their cumulative effects on daily death counts. The result of this analysis was

1 similar in magnitude to the effect of SO₂ alone. An ecologic cohort study of infant mortality in
2 the U.S. found no association with annual averages of SO₂ concentration (Lipfert et al., 2000a).

3 Gilboa et al. (2005) conducted a population-based case-control study to investigate the
4 association between maternal exposure and air pollutant exposure during weeks 3-8 of
5 pregnancy, the risk of selected cardiac birth defects and oral clefts in live births, and fetal deaths
6 between 1997 and 2000 in seven Texas counties. When the highest quartile of exposure was
7 compared to the lowest, the authors observed a positive association between SO₂ and isolated
8 ventricular septal defects (OR 2.16 [95% CI: 1.51, 3.09]). Although this is the only study to have
9 examined the effect of maternal exposure to SO₂ on birth defects, it supports the notion that the
10 developing embryo and growing fetus is susceptible to maternal air pollution exposure.

11 Several studies examined adverse health outcomes in relation to SO₂ concentrations during
12 the neonatal period. Dales et al. (2006) evaluated hospitalizations for respiratory disorders in
13 neonates < 4 weeks of age from hospitals in 11 large Canadian cities during a 15-year study
14 period (population-weighted average 24-h avg SO₂ of 4.3 ppb). The researchers observed a 5.5%
15 (95% CI: 2.8, 8.3) excess risk in respiratory hospitalizations associated with a 10 ppb increase in
16 24-h avg SO₂ concentrations with a 2-d lag. This effect was slightly attenuated after adjusting for
17 PM₁₀ and gaseous copollutants. To investigate the influence of ambient SO₂ concentrations on
18 SIDS, Dales et al. (2004) conducted a time-series analysis comparing daily rates of SIDS and
19 daily SO₂ concentrations from 12 large, Canadian cities during a 16-year period. The mean
20 24-h avg SO₂ level across the 12 cities was 5.51 ppb (IQR 4.92). There was an 18.0% (95% CI:
21 4.4, 33.4) excess risk in SIDS incidence for a 10 ppb increase in 24-h avg SO₂ levels. The
22 authors concluded that the effect of SO₂ was independent of sociodemographic factors, temporal
23 trends, and weather.

24 In summary, epidemiological studies on birth outcomes have found suggestive positive
25 associations between SO₂ exposure and low birth weight; however, toxicological studies provide
26 very little biological plausibility for reproductive outcomes related to SO₂ exposure. The
27 inconsistent results across trimesters of pregnancy and the lack of evidence regarding
28 confounding by copollutants further limit the interpretation of these studies. The limited number
29 of studies addressing preterm delivery, IUGR, birth defects, neonatal hospitalizations, and infant
30 mortality make it difficult to draw conclusions regarding the effect of SO₂ on these outcomes.

3.4.6. Other Organ System Effects Associated with Long-Term Exposure

The 1982 AQCD presented only one chronic exposure study which was relevant to nervous system effects. Dogs were exposed for 68 months to a mixture of SO₂ and H₂SO₄. No effects on visual evoked brain potentials during or immediately after exposure to the SO_x mixture were observed. Since then, numerous studies have examined brain lipid content, lipid peroxidation and glutathione content and antioxidant enzymes following inhalation exposure of rodents to SO₂ at concentrations of 10 ppm or higher. Concentrations of 5 ppm or higher SO₂ were used in studies examining neurobehavior and neurodevelopment in mice. These studies are summarized in Annex Table E-6.

In the past 25 years, numerous animal toxicological studies have evaluated the effects of long-term SO₂ exposure on other organ systems such as reproductive, hematological, gastrointestinal, renal, lymphatic, and endocrine systems. Most of these studies used concentrations of SO₂ of 5 ppm or higher. Many of these studies examined alteration profiles of lipid peroxidation and antioxidant levels (Langley-Evans et al., 1996; Meng and Bai, 2004; Meng et al., 2003b) and are summarized in Annex Table E-7 through E-9 and E-22 through E-24.

3.5. Mortality Associated with Long-Term SO₂ Exposure

3.5.1. Summary of Findings from the Previous Review

At the time of the 1982 AQCD, the available studies on the effects of long-term exposure to SO₂ on mortality were all ecological cross-sectional studies. This study design could not take into consideration such confounders as cigarette smoking, occupational exposures, and social status. In addition, there were questions regarding how representative the aerometric data used were for community exposure. Therefore, it was concluded that the epidemiological studies did not provide valid quantitative data relating respiratory disease or other types of mortality to long-term (annual average) exposures to SO₂ or PM.

The 1986 Secondary Addendum reviewed more studies of this type, with information on more detailed components of PM (inhalable and fine particles, and particulate sulfate). While some studies suggested importance of the size of PM, the fundamental problem of the study

1 design made it difficult to interpret the effect estimates. The 1986 Secondary Addendum also
2 reviewed a Japanese study in which the death rates from asthma and chronic bronchitis in a
3 highly polluted section of Yokkaichi, an industrial city with large SO₂ emissions from the largest
4 oil-fired power plant in Japan, were compared with those in a less polluted area of the same city.
5 SO_x levels (measured using the lead peroxide method) averaged across several monitoring sites
6 in the polluted harbor area ranged from around 1.0 to 2.0 mg/day (annual average) during 1964
7 through 1972 and then steadily declined to less than 0.5 mg/day in 1982. This is in contrast to
8 levels consistently < 0.3 mg/day in the low pollution areas throughout 1967 through 1982.
9 Annual average levels for other pollutants (i.e., NO₂, TSP, oxidants) monitored in the high
10 pollution area were consistently low from 1974 through 1982. The results indicated elevated
11 rates of chronic bronchitis mortality in the highly polluted area compared to the less polluted
12 area, but the 1986 Secondary Addendum could not conclude that this was due to SO₂ alone,
13 because sulfate or other particulate SO_x such as H₂SO₄ could have been responsible.

14 Several, more recent studies have examined long-term exposure effects of air pollution,
15 including SO₂, on mortality. These studies are summarized in Annex Table F-9. As with short-
16 term exposure studies, the focus of most of these studies was mainly on PM though some
17 focused on traffic-related air pollution. They all used Cox proportional hazards regression
18 models with adjustment for potential confounders. The designs of these studies were better than
19 earlier cross-sectional studies as the outcome and most of the potential confounders (e.g.,
20 smoking history, occupational exposure) were measured on an individual basis. However, the
21 geographic scale and method for exposure estimates varied across these studies.

3.5.2. Associations of Mortality and Long-Term Exposure in Key Studies

3.5.2.1. U.S. Cohort Studies

3.5.2.1.1. Harvard Six Cities Studies

22 Dockery et al. (1993) conducted a prospective cohort study to study the effects of air
23 pollution with the main focus on PM components in six U.S. cities. These cities were chosen
24 based on levels of air pollution, with Portage, WI and Topeka, KS representing the least polluted
25 cities and Steubenville, OH representing the most polluted city. Mean SO₂ levels ranged from

1 1.6 ppb in Topeka to 24.0 ppb in Steubenville from 1977 to 1985. Cox proportional hazards
2 regression was conducted with data from a 14- to 16-year follow-up study of 8,111 adults in the
3 six cities. Dockery et al. reported that lung cancer and cardiopulmonary mortality were more
4 strongly associated with the concentrations of inhalable and fine PM, and sulfate particles, than
5 with the levels of TSP, SO₂, NO₂, or acidity of the aerosol.

6 Krewski et al. (2000) conducted a sensitivity analysis of the Harvard Six Cities study and
7 examined associations between gaseous pollutants (i.e., O₃, NO₂, SO₂, and CO) and mortality.
8 SO₂ showed positive associations with total mortality (RR = 1.05 [95% CI: 1.02, 1.09] per 5 ppb
9 increase in average SO₂ over the study period) and cardiopulmonary deaths (1.05 [95% CI: 1.00,
10 1.10]), but in this dataset SO₂ was highly correlated with PM_{2.5} (r = 0.85), sulfate (r = 0.85), and
11 NO₂ (r = 0.84).

3.5.2.1.2. American Cancer Society Cohort Studies

12 Pope et al. (1995) investigated associations between long-term exposure to PM and the
13 mortality outcomes in the ACS cohort. Ambient air pollution data from 151 U.S. metropolitan
14 areas in 1981 were linked with individual risk factors in 552,138 adults who resided in these
15 areas when enrolled in the prospective study in 1982. Death outcomes were ascertained through
16 1989. PM_{2.5} and sulfate were associated with total, cardiopulmonary, and lung cancer mortality,
17 but not with mortality for all other causes. Gaseous pollutants were not analyzed in this study.

18 Krewski and co-investigators (Jerrett et al., 2003; Krewski et al., 2000) conducted an
19 extensive sensitivity analysis of the Pope et al. (1995) ACS data, augmented with additional
20 gaseous pollutants data. The mean SO₂ concentrations were 7.18 ppb in the warm season (April
21 to September) and 11.24 ppb in the cool season (October to March). Among the gaseous
22 pollutants examined, only SO₂ showed positive associations with mortality. The RR for total
23 mortality was 1.06 (95% CI: 1.05, 1.07) per 5 ppb increase in the annual average SO₂. Analysis
24 using SO₂ measured in different seasons produced a somewhat higher estimate for the warm
25 season than that for the cool season (7% compared to 5% excess risk per 5 ppb increase).

26 Although the subjects in the ACS cohort came from all regions of the United States, the majority
27 of the 151 cities were located in the eastern United States, where both SO₂ and sulfate tend to be
28 higher. PM_{2.5} levels are also higher in the east. To address the influence of these spatial patterns,
29 which may confound associations between mortality and these pollutants, Krewski et al. (2000)

1 conducted extensive two-stage regression modeling. In these models, the association between
2 SO₂ and mortality was diminished but persisted after adjusting for sulfate, PM_{2.5}, and other
3 variables. For example, in the spatial filtering model (which resulted in the largest reduction of
4 the SO₂ effect estimate when sulfate was included), the SO₂ total mortality RR estimate was 1.07
5 (95% CI: 1.03, 1.11) in the single-pollutant model and 1.04 (95% CI: 1.02, 1.06) with sulfate in
6 the two-pollutant model. The effect estimates for PM_{2.5} and sulfate also were diminished when
7 SO₂ was included in the models. The result further showed that SO₂ effect estimates were
8 generally insensitive to adjustment for spatial correlation. Thus, these results suggest that the
9 association between SO₂ and mortality may be confounded with PM, but the association cannot
10 be accounted for by PM_{2.5} or sulfate alone. Krewski et al. (2000) noted that their reanalysis of the
11 ACS and Harvard Six Cities studies suggested that mortality might be attributed to more than
12 one component of the complex mixture of ambient air pollutants in urban areas in the United
13 States.

14 The original Pope et al. (1995) study and the Krewski et al. (2000) reanalysis both used the
15 air pollution exposure estimates that are based on the average over the Metropolitan Statistical
16 Area (MSA), which consists of multiple counties. To investigate the effects of geographic scale
17 over which the air pollution exposures are averaged, Willis et al. (2003) reanalyzed the ACS
18 cohort data using the exposure estimates averaged over the county scale, and compared the
19 results with those based on the MSA-scale average exposure. Less than half of the cohort used in
20 the MSA-based study was used in the county-scale based analysis, because of the limited
21 availability of sulfate monitors and the reduced sample size due to the loss of subjects when
22 using the five-digit ZIP codes. The mean (9.3 ppb versus 10.7 ppb) and range (0.0 to 29.3 ppb
23 versus 0.0 to 27.2 ppb) of the MSA- and county-level SO₂ data sets were similar. In the analysis
24 comparing the two-pollutant model with sulfate and SO₂, they found that the inclusion of SO₂
25 reduced sulfate effect estimates substantially (> 25%) in the MSA-scale model but not
26 substantially (< 25%) in the county-scale model. In the MSA-level analysis (with 113 MSAs),
27 the SO₂ RR estimate was 1.04 (95% CI: 1.02, 1.06) per 5 ppb increase, with sulfate in the model.
28 In the county-level analysis (91 counties) with sulfate in the model, the corresponding estimate
29 was smaller (1.02 [95% CI: 1.00, 1.05]). It should also be noted that the correlation between
30 covariates were different between the MSA-level data and county-level data. The correlation
31 between SO₂ and sulfate was 0.48 in the MSA-level data, but it was 0.56 in the county-level

1 data. The correlation between poverty rate and SO₂ was -0.16 in the MSA-level data, but it was
2 0.15 in the county-level data. Thus, the extent of confounding between SO₂ and PM components
3 as well as among other covariates in the model can be affected by the geographic scale of
4 aggregation of exposure estimates. It is not clear, however, if the smaller geographic scale
5 increases or decreases exposure characterization error for SO₂, because a certain extent of
6 smoothing (averaging) over distance may reduce very local concentration peaks that are not
7 relevant to the city-wide population.

8 Pope et al. (2002) extended analysis of the ACS cohort with double the follow-up time (to
9 1998) and triple the number of deaths compared to the original Pope et al. (1995) study. In
10 addition to PM_{2.5}, all the gaseous pollutant data were retrieved for the extended period and
11 analyzed for their associations with death outcomes. As in the 1995 analysis, the air pollution
12 exposure estimates were based on the MSA-level averages. PM_{2.5} was associated with total,
13 cardiopulmonary, and lung cancer mortality but not with deaths for all other causes. SO₂ was
14 associated with all the mortality outcomes, including all other causes of deaths. The SO₂ RR
15 estimate for total mortality was 1.03 (95% CI: 1.02, 1.05) per 5 ppb increase (1982 to 1998
16 average). The association of SO₂ with mortality for all other causes (sulfate also showed this
17 pattern) makes it difficult to interpret the effect estimates. This lack of specificity for SO₂ (in
18 contrast to PM) is not consistent with causal inference.

3.5.2.1.3. The EPRI-Washington University Veterans' Cohort Mortality Studies

19 Lipfert et al. (2000b) conducted an analysis of a national cohort of ~70,000 male U.S.
20 military veterans who were diagnosed as hypertensive in the mid 1970s and were followed up for
21 about 21 years (up to 1996). This cohort was 35% black and 57% were current smokers (81% of
22 the cohort had been smokers at one time). PM_{2.5}, PM₁₀, PM_{10-2.5}, TSP, sulfate, CO, O₃, NO₂, SO₂,
23 and lead were examined in this analysis. No mean or median level of SO₂ was reported. The
24 county of residence at the time of entry to the study was used to estimate exposures. Four
25 exposure periods (from 1960 to 1996) were defined, and deaths during each of the three most
26 recent exposure periods were considered. The results for SO₂ were presented only qualitatively
27 as part of their preliminary screening regression results. Lipfert et al. (2000b) noted that lead and
28 SO₂ were not found to be associated with mortality, thus were not considered further. They also
29 noted that the pollution effect estimates were sensitive to the regression model specification,

1 exposure periods, and the inclusion of ecological and individual variables. The authors reported
2 that indications of concurrent mortality risks were found for NO₂ and peak O₃.

3 Lipfert et al. (2006b) examined associations between traffic density and mortality in the
4 same cohort, whose follow-up period was extended to 2001. As in their 2000 study, four
5 exposure periods were considered but included more recent years. The 95th percentiles of daily
6 average in each of the exposure periods were considered for SO₂. For the 1997–2001 data period,
7 the estimated mortality RR for SO₂ was 0.99 (95% CI: 0.97, 1.01) per 5 ppb increase in a single-
8 pollutant model. They reported that traffic density was a better predictor of mortality than
9 ambient air pollution variables with the possible exception of O₃. The log-transformed traffic
10 density variable was only weakly correlated with SO₂ (r = 0.32) and PM_{2.5} (r = 0.50) in this data
11 set.

12 Lipfert et al. (2006a) further extended analysis of the veterans' cohort data to include the
13 EPA's Speciation Trends Network (STN) data, which collected chemical components of PM_{2.5}.
14 They analyzed the STN data for year 2002, again using county-level averages. PM_{2.5} and gaseous
15 pollutants data for 1999 through 2001 were also analyzed. As in the previous Lipfert et al. (2006)
16 study, traffic density was the most important predictor of mortality, but associations were also
17 seen for elemental carbon, vanadium, nickel, and nitrate. O₃, NO₂, and PM₁₀ also showed
18 positive but weaker associations. Once again, no associations were observed between long-term
19 exposure to SO₂ and mortality.

3.5.2.1.4. Seventh-day Adventist Study

20 Abbey et al. (1999) investigated associations between long-term ambient concentrations of
21 PM₁₀, sulfate, SO₂, O₃, and NO₂ (1973 through 1992) and mortality (1977 through 1992) in a
22 cohort of 6,338 nonsmoking California Seventh-day Adventists. Monthly indices of ambient air
23 pollutant concentrations at 348 monitoring stations throughout California were interpolated to
24 ZIP codes according to home or work location of study participants, cumulated, and then
25 averaged over time. They reported associations between PM₁₀ and total mortality for males and
26 nonmalignant respiratory mortality for both sexes. SO₂ was not associated with total mortality
27 (RR 1.07 [95% CI: 0.92, 1.24] for males and 1.00 [95% CI: 0.88, 1.14] for females per 5 ppb
28 increase in multiyear average SO₂), cardiopulmonary deaths, or respiratory mortality for either
29 gender.

3.5.2.2. European Cohort Studies

A study by Beelen et al. (2008) examined associations between traffic-related air pollution and mortality. They analyzed data from the Netherlands Cohort Study on Diet and Cancer with 120,852 subjects who were followed from 1987 to 1996. BS, NO₂, SO₂, PM_{2.5}, and four types of traffic-exposure estimates were analyzed. While the local traffic component was estimated for BS, NO₂, and PM_{2.5}, no such attempt was made for SO₂, because there was “virtually no traffic contributions to this pollutant.” Thus, only “background” SO₂ levels were reflected in the exposure estimates. Traffic intensity on the nearest road was associated with all-cause mortality and a larger RR was observed for respiratory mortality. Results were similar for BS, NO₂ and PM_{2.5}, but no associations were found for SO₂ (RR = 0.98 [95% CI: 0.93, 1.03] per 5 ppb increase in multiyear average SO₂).

Nafstad et al. (2004) investigated the association between mortality and long-term exposure to air pollution exposure in a cohort of Norwegian men followed from 1972–1973 through 1998. Data from 16,209 males (aged 0 to 49 years) living in Oslo, Norway, in 1972–1973 were linked with data from the Norwegian Death Register and with estimates of the average annual air pollution levels at the participants’ home addresses. PM was not considered in this study because measurement methods changed during the study period. Exposure estimates for nitrogen oxides (NO_x) and SO₂ were constructed using models based on subject addresses, emission data for industry, heating, and traffic, and measured concentrations. While NO_x was associated with total, respiratory, lung cancer, and ischemic heart disease deaths, SO₂ did not show any associations with mortality. The authors noted that the SO₂ levels were reduced by a factor of 7 during the study period (from 5.6 ppb in 1974 to 0.8 ppb in 1995), whereas NO_x did not show any clear downward trend.

Filleul et al. (2005) investigated long-term effects of air pollution on mortality in 14,284 adults who resided in 24 areas from seven French cities when enrolled in the Air Pollution and Chronic Respiratory Diseases (PAARC) survey in 1974. Daily measurements of SO₂, TSP, BS, NO₂, and NO were made in the 24 areas for 3 years (1974 through 1976). Models were run before and after exclusion of six area monitors influenced by local traffic as determined by a NO:NO₂ ratio of > 3. Before exclusion of the six areas, none of the air pollutants was associated with mortality outcomes. After exclusion of these areas, analyses showed associations between total mortality and TSP, BS, NO₂, and NO but not SO₂ (RR = 1.01 [95% CI: 0.97, 1.06] per

5 ppb multiyear average) or acidimetric measurements. It should be noted that SO₂ levels in these French cities declined markedly between the 1974 through 1976 period and the 1990 through 1997 period by a factor of 2 to 3, depending on the city. The changes in air pollution levels over the study period complicate interpretation of reported effect estimates.

3.5.2.3. Cross-Sectional Analysis Using Small Geographic Scale

Elliott et al. (2007) examined associations of BS and SO₂ with mortality in Great Britain using a cross-sectional analysis. However, unlike the earlier ecological cross-sectional mortality analyses in the United States in which mortality rates and air pollution levels were compared using large geographic boundaries (i.e., MSAs or counties), in the Elliot et al. analysis, the mortality rates and air pollution were compared using a much smaller geographic unit, the electoral ward, with a mean area of 7.4 km² and a mean population of 5,301 per electoral ward. Death rates were computed for four successive 4-year periods from 1982 to 1994 and associated with 4-year exposure periods from 1966 to 1994. The number of deaths from all causes in the 10,520 wards was 420,776. Of note, SO₂ levels declined from 41.4 ppb in the 1966 to 1970 period to 12.2 ppb in 1990 to 1994. This type of analysis does not allow adjustments for individual risk factors, but the study did adjust for socioeconomic status data available for each ward from the 1991 census. Social deprivation and air pollution were more highly correlated in the earlier exposure windows. They observed associations for both BS and SO₂ and mortality outcomes. The estimated effects were stronger for respiratory illness than other causes of mortality for the most recent exposure period and most recent mortality period (when pollution levels were lower). The adjustment for social deprivation reduced the effect estimates for both pollutants. The adjusted mortality RRs for SO₂ for the pooled mortality periods using the most recent exposure windows were 1.021 (95% CI: 1.018, 1.024) for all causes, 1.015 (95% CI: 1.011, 1.019) for cardiovascular, and 1.064 (95% CI: 1.056, 1.072) for respiratory causes per 5 ppb increase in SO₂. The effect estimates for the most recent mortality period using the most recent exposure windows were larger. Simultaneous inclusion of BS and SO₂ reduced effect estimates for BS but not SO₂. Elliott et al. (2007) noted that the results were consistent with those reported in the Krewski et al. (2000) reanalysis of the ACS study. This analysis was ecological, but the exposure estimates in the smaller area compared to that in the U.S. cohort studies may have resulted in less exposure misclassification error, and the large underlying

1 population appears to be reflected in the narrow confidence bands of effect estimates. The results
2 from this study suggest an association between long-term exposures (especially in recent years)
3 to SO₂ and mortality.

3.5.3. Summary of Evidence on the Effect of Long-Term Exposure on Mortality

4 The available epidemiological evidence on the effect of long-term exposure to SO₂ on
5 mortality is *inadequate to infer the presence or absence of a causal relationship* at this time. The
6 ecological cross-sectional studies examined in the 1982 AQCD and 1986 Secondary Addendum
7 found suggestive relationships between long-term exposure to SO₂ and mortality. However, there
8 were concerns as to whether the observed association was due to SO₂ alone, because sulfate or
9 other particulate SO_x such as H₂SO₄ could have been responsible. In the more recent longitudinal
10 cohort studies, once again, positive associations have been observed between long-term exposure
11 to SO₂ and mortality; however, several issues affect the interpretation of these results.

12 Figure 3-14 presents all-cause mortality RR estimates associated with long-term exposure
13 to SO₂ from the U.S. and European cohort studies. The overall range of RRs spans 0.97 to 1.07
14 per 5 ppb increase in the annual (or longer period) average SO₂. The analyses of the Harvard Six
15 Cities and the ACS cohort data, which likely provide effect estimates that are most useful for
16 evaluating possible health effects in the United States, observed RRs of 1.02 to 1.07. Note that
17 each of the U.S. cohort data has its own advantages and limitations. The Harvard Six Cities data
18 have a small number of exposure estimates, but the location of the monitors were chosen
19 carefully for epidemiological purposes. The ACS cohort had far more subjects, but the
20 population was more highly educated than the representative U.S. population. Since educational
21 status appeared to be an important effect modifier of air pollution effects in both studies, the
22 overall effect estimate for the ACS cohort may underestimate that for the more general
23 population. However, it should also be noted that several other U.S. and European studies did not
24 observe an association between long-term exposure to SO₂ and mortality.

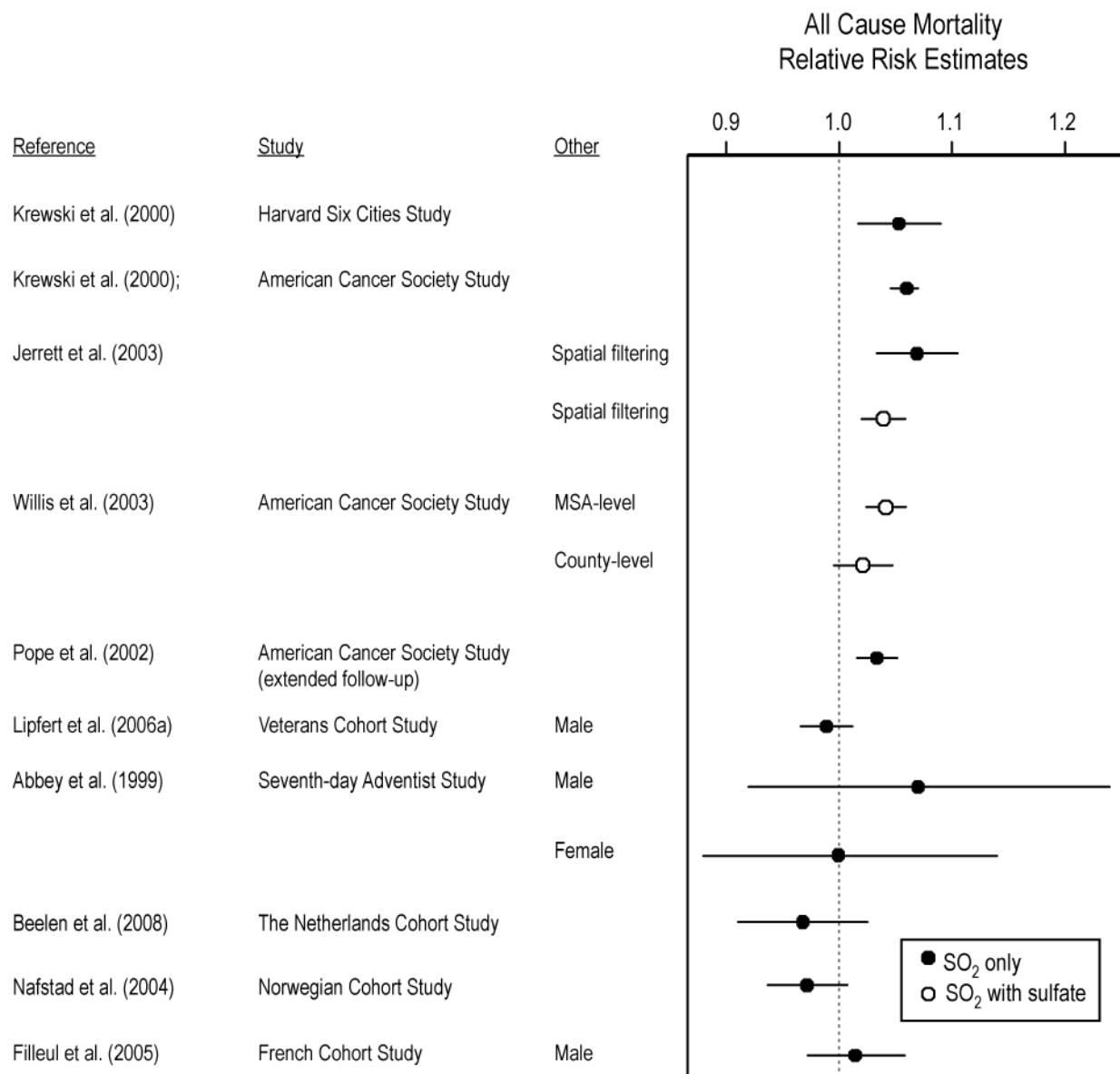


Figure 3-14. Relative risks (95% CI) of SO₂-associated all-cause (nonaccidental) mortality, with and without adjustment for sulfate, from longitudinal cohort studies. Effect estimates are standardized per 5 ppb increase in SO₂ concentrations. The exposure estimates for Krewski et al. (2000) and Pope et al. (2002) are based on MSA (Metropolitan Statistical Area)-level averaging; Lipfert et al. (2006a) used county-level averaging.

The geographic scale of analysis appears to influence SO₂ effect estimates and exposure error. In a reanalysis of the ACS data, the county-level analysis showed a smaller SO₂ effect estimate than MSA-level analysis. For sulfate, the opposite pattern was found. Thus, the impact of the geographic scale of analysis may also depend on the spatial distribution of air pollutants. The cross-sectional analysis in Great Britain using small-scale electoral wards observed an effect

1 estimate similar to the lower end of the range of effect estimates for all-cause mortality from
2 U.S. cohort studies, though it is not clear if the effect estimates from this cross-sectional study
3 are directly comparable to those from cohort studies.

4 Another important issue that these studies could not resolve was the possible confounding
5 and/or interaction among PM indices and SO₂. The possibility that the observed effects may not
6 be due to SO₂, but other constituents that come from the same source as SO₂, or that PM may be
7 more toxic in the presence of SO₂ or other components associated with SO₂, cannot be ruled out.
8 For example, the ACS cohort came from all regions of the United States, but a major fraction of
9 the ACS cities were located in the eastern United States, where both SO₂ and sulfate levels tend
10 to be higher. Therefore, even with sophisticated spatial modeling, separating possible
11 confounding of SO₂ effects by PM is challenging. Future and on-going studies that take into
12 consideration within- versus between-city variation of these pollutants may help elucidate this
13 issue.

14 Overall, the results from two major U.S. epidemiological studies observe an association
15 between long-term exposure to SO₂ or sulfur-containing particulate air pollution and mortality.
16 However, several other U.S. and European cohort studies did not observe an association. The
17 lack of consistency across studies, inability to distinguish potential confounding by copollutants,
18 and uncertainties regarding the geographic scale of analysis limit the interpretation of a causal
19 relationship.

Chapter 4. Public Health Impact

This chapter addresses several issues relating to the broader public health impact from exposure to ambient SO₂. First, the shape of the concentration-response relationship for SO₂ is discussed, with consideration of interindividual variability in responses and evaluation of the limited evidence available to assess a population threshold value for health effects. The next section identifies characteristics of subpopulations which may experience increased risks from SO₂ exposures, through either enhanced susceptibility (e.g., as a result of pre-existing disease, genetic factors, age) and/or differential vulnerability associated with increased exposure (e.g., close proximity to sources, activities). The final section discusses the potential public health impact from adverse health effects associated with SO₂ by examining the prevalence of susceptible individuals in the U.S. population.

4.1. Assessment of Concentration-Response Function and Potential Thresholds

An important consideration in characterizing the public health impacts associated with SO₂ exposure is whether the concentration-response relationship is linear across the full concentration range, or if there are concentration ranges where there are departures from linearity (i.e., nonlinearity). Of particular interest is the shape of the concentration-response curve at and below the level of the current SO₂ NAAQS level of a 24-h avg level of 0.14 ppm or the annual average of 0.03 ppm.

Some human clinical studies provide individual-level response data in relation to different levels of SO₂ exposure; this allows evaluation of both the percentage of individuals showing responses across the range of exposures as well as the concentration at which an individual begins to indicate a response. In epidemiological studies, rather than identifying interindividual differences in response, most studies evaluate whether there is a population-level threshold, which is the concentration of SO₂ that must be exceeded to elicit a health response in the study population. Low data density in the lower concentration range, measurement error in the response, and exposure measurement error are some of the factors that complicate the ability to determine the shape of the concentration-response curve, including the presence of any

threshold. Biological characteristics that tend to linearize concentration-response relationships include individual biological differences in susceptibility to air pollution health effects, additivity of SO₂-induced effects to a naturally occurring background level, and additivity of effects from other pollutant exposures. Epidemiological and human clinical studies that examined the shape of the concentration-response function for different averaging times or exposure durations are presented below. The discussion focuses on respiratory morbidity effects associated with short-term exposure to SO₂, for which the strongest causal evidence exists.

4.1.1. Evidence from Human Clinical Studies

In human clinical studies of exercising asthmatics, moderate SO₂-induced decrements in lung function have been observed at the lowest levels tested (i.e., 0.2 to 0.3 ppm, 5 to 10 min exposures) in the most sensitive individuals (approximately 5-20% of subjects). Statistically significant respiratory effects have been consistently observed at concentrations of 0.4-0.6 ppm, with 20-60% of asthmatics experiencing moderate to large decrements in lung function following 5-10 min exposures (see Table 3-1). Smaller, yet statistically significant decrements in lung function have also been demonstrated at SO₂ concentrations < 0.2 ppm when preceded by exposure to O₃ (see Section 3.1.5.1).

With increasing exposure concentration between 0.2 and 1.0 ppm, there is a clear increase both in magnitude of respiratory effect and percent of asthmatics affected (Table 3-1). A subset of the data presented in this table was taken from a series of studies conducted by Linn et al. (1987; 1988; 1990) and is presented graphically in Figure 4-1 through Figure 4-3. In these studies, mild and moderate asthmatics were exposed for 10 min to SO₂ concentrations between 0 and 0.6 ppm during moderate to heavy exercise. These particular studies were selected for inclusion in this meta-analysis owing to similarities between exposure protocols, with all subjects being exposed to multiple concentrations of SO₂. In the 1987 study, subjects were exposed to SO₂ concentrations of 0, 0.2, 0.4, and 0.6 ppm, while in the 1988 and 1990 studies, subjects were exposed to concentrations of 0, 0.3, and 0.6 ppm. The percent of asthmatics experiencing moderate or greater SO₂-induced decrements in lung function (increase in sRaw ≥ 100% or decrease in FEV₁ ≥ 15%) is shown in Figure 4-1. At 0.2 ppm, between 5 and 13% of subjects are affected, and this fraction increases with increasing concentration, with approximately 50% of subjects experiencing respiratory effects at a concentration of 0.6 ppm.

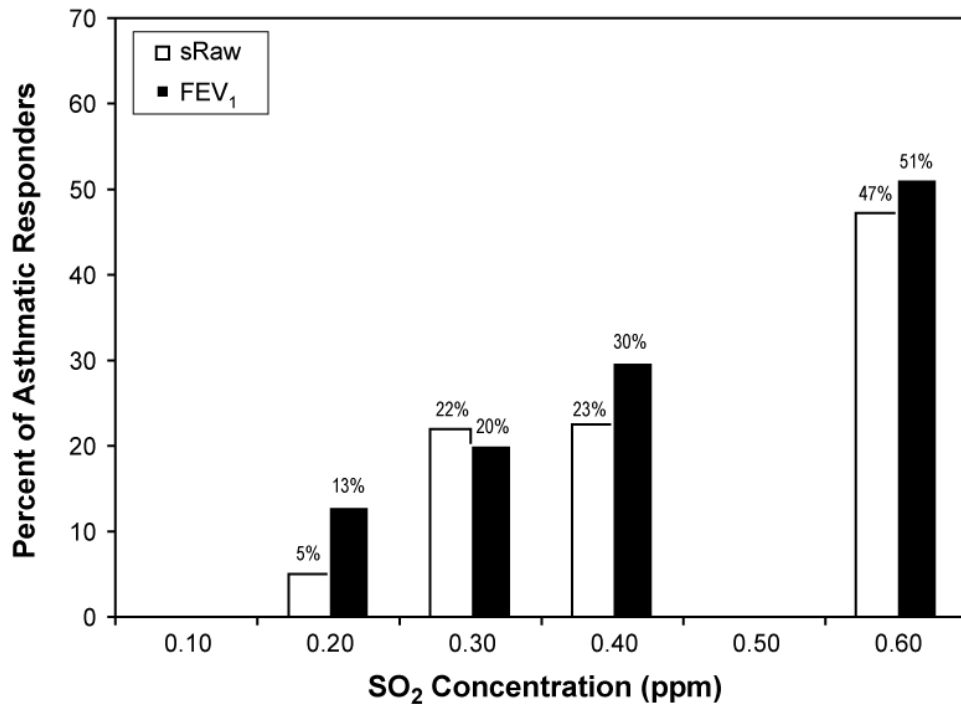


Figure 4-1. Percent of mild and moderate asthmatics ($\dot{V}_E = 40\text{-}50$ L/min) experiencing an SO_2 -induced increase in sRaw of $\geq 100\%$ or a decrease in FEV_1 of $\geq 15\%$, adjusted for effects of moderate to heavy exercise in clean air. The data represents lung function measurements from 40, 41, 40, and 81 subjects at concentrations of 0.2, 0.3, 0.4, and 0.6 ppm, respectively.

Source: Data taken from Linn et al. (1987; 1988; 1990)

Figure 4-2 and Figure 4-3 present the concentration-response relationship between SO_2 decrements in lung function in SO_2 -sensitive asthmatics, i.e., those asthmatics experiencing significant decrements in lung function at the highest exposure concentration (0.6 ppm) used in the Linn et al. studies (1987; 1988; 1990). This analysis demonstrates a clear increase in the magnitude of respiratory effects with increasing exposure concentration, with more marked effects observed at SO_2 concentrations greater than 0.3 ppm. The results of a study by Gong et al. (1995) support this conclusion: the authors observed a linear relationship between SO_2 concentration (0, 0.5, and 1.0 ppm) and both lung function (decrease in FEV_1 , and increase in sRaw) and respiratory symptoms.

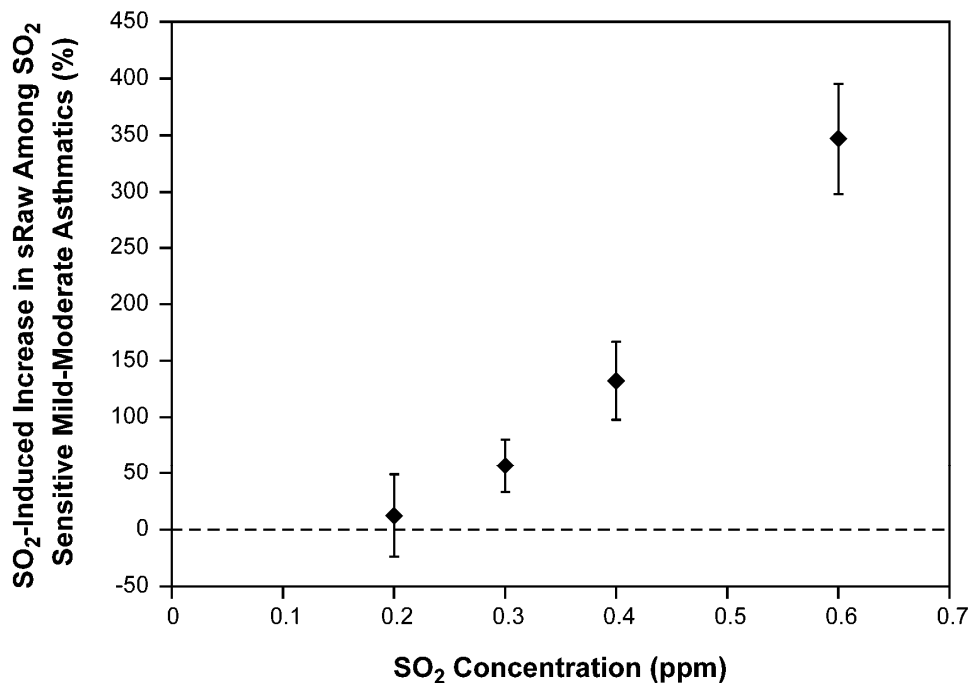


Figure 4-2. SO₂-induced increase in sRaw among SO₂-sensitive mild and moderate asthmatics (n=38) following 10-min exposures during moderate to heavy exercise ($V_E = 40\text{-}50$ L/min). Only SO₂-sensitive asthmatics, defined here as asthmatics experiencing $\geq 100\%$ SO₂-induced increase in sRaw at 0.6 ppm, were included in this analysis. The analysis includes data from 14 SO₂-sensitive subjects from Linn et al. (1987) exposed to concentrations of 0.2, 0.4, and 0.6 ppm, as well as 24 SO₂-sensitive subjects from Linn et al. 1988 and 1990 exposed to 0.3 and 0.6 ppm. Error bars = ± 1 SE.

4.1.2. Evidence from Epidemiological Studies

Although there are numerous epidemiological studies that examined the association between SO₂ and various health effects, only a few of these studies attempted to evaluate the concentration-response function. Most studies assumed a linear or log-linear relationship between ambient SO₂ concentrations and the health outcome in their evaluations.

Epidemiological studies have examined the concentration-response relationship for SO₂ using various statistical methods, including the comparison of effect estimates in increasing quartiles or quintiles, plotting the risk observed against increasing SO₂ concentrations, and using nonparametric smoothed curves to assess the nonlinearity of the SO₂-effect relationship. Most of the epidemiological studies that examined the concentration-response function between SO₂ exposure and respiratory morbidity observed that the relationship was linear across the entire concentration range.

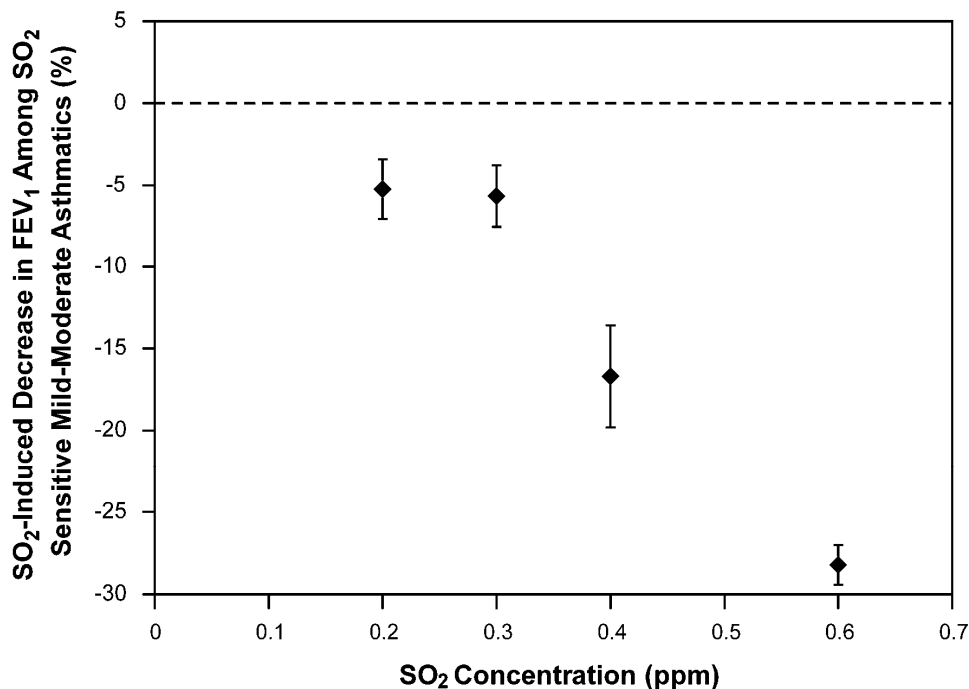


Figure 4-3. SO₂-induced decrease in FEV₁ among SO₂-sensitive mild and moderate asthmatics (n=41) following 10 min exposures during moderate to heavy exercise ($V_E = 40\text{-}50$ L/min). Only SO₂-sensitive asthmatics, defined here as asthmatics experiencing $\geq 15\%$ SO₂-induced decrease in FEV₁ at 0.6 ppm, were included in this analysis. The analysis includes data from 21 SO₂-sensitive subjects from Linn et al. (1987) exposed to concentrations of 0.2, 0.4, and 0.6 ppm, as well as 20 SO₂-sensitive subjects from Linn et al. 1988 and 1990 exposed to 0.3 and 0.6 ppm. Error bars = ± 1 SE.

1 Only one epidemiological study investigated the concentration-response function of peak
2 SO₂ exposures. The association between asthma hospitalizations and ambient 1-h max SO₂
3 concentrations was examined in a case-control study of children in Bronx County, NY (Lin et al.,
4 2004). The 1-h max concentration ranged from 2.9 to 66.4 ppb. The authors categorized 1-h max
5 SO₂ concentrations and estimated ORs for each category using the lowest exposure group as the
6 reference (2.9 to 9.2 ppb). They observed an increasing linear trend across the range of
7 concentrations, with more marked effects observed at 1-h max SO₂ concentrations greater than
8 40 ppb Figure 4-4. A 1-h max SO₂ concentration of 40 ppb falls between the 90th and 95th
9 percentiles of the ambient regulatory data for the years 2003-2005; during these years 24-h avg
10 SO₂ concentrations for the 90th and 95th percentiles were 10 and 13 ppb, respectively.

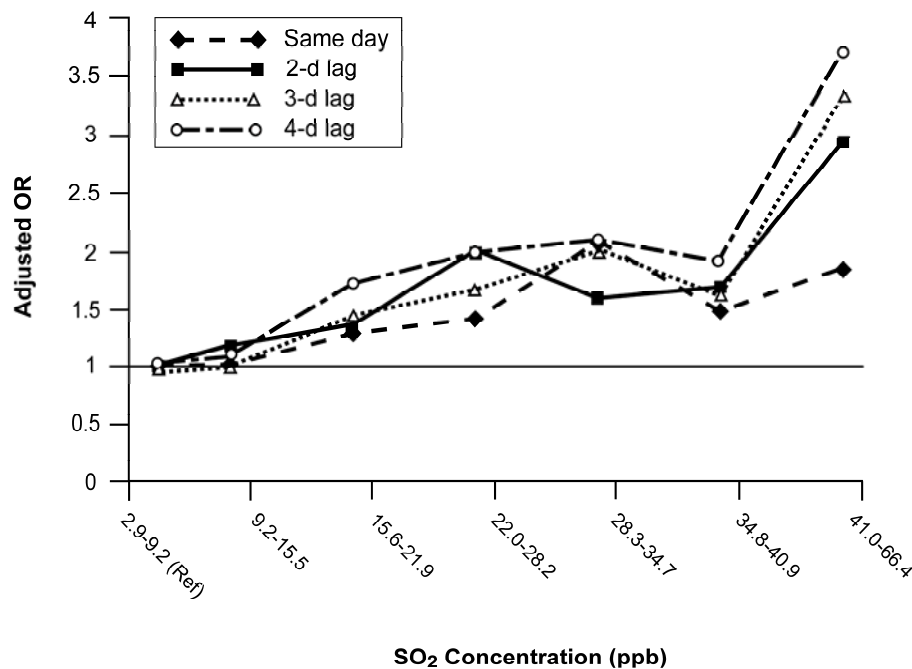


Figure 4-4. Adjusted odds ratios of asthma hospitalizations by groupings of 1-h max SO₂ concentrations in Bronx County, New York. All groups were compared with the lowest exposure group (2.9-9.2 ppb). ORs for 1-h max SO₂ concentrations on the same day, as well as from a 2-day, 3-day, and 4-day moving average lag are presented.

Source: Lin et al. (2004)

Most epidemiological studies investigating the concentration-response function examined the effects of short-term 24-h avg exposures to SO₂. The Harvard Six Cities study by Schwartz et al. (1994) investigated the concentration-response function and observed a nonlinear relationship between SO₂ concentrations and respiratory symptoms. A figure plotting the relative odds of incidence of lower respiratory tract symptoms against SO₂ concentrations lagged 1 day indicated that no statistically significant increase in the incidence of lower respiratory tract symptoms was seen until concentrations exceeded a 24-h avg SO₂ of 22 ppb though an increasing trend was observed at concentrations as low as 10 ppb (see Figure 4-5). In a study of respiratory hospitalizations, Ponce de Leon et al. (1996) found that a weak relationship with SO₂ was only observable at 24-h avg SO₂ concentrations above 23 ppb. In both this study and the study by Schwartz et al. (1994), a statistically significant increased risk was observable only at 24-h avg SO₂ concentrations that were above the 90th percentile. The nonlinearity observed in

1 these concentration-response functions is dependent on only a few influential observations; thus,
2 the results should be viewed with caution.

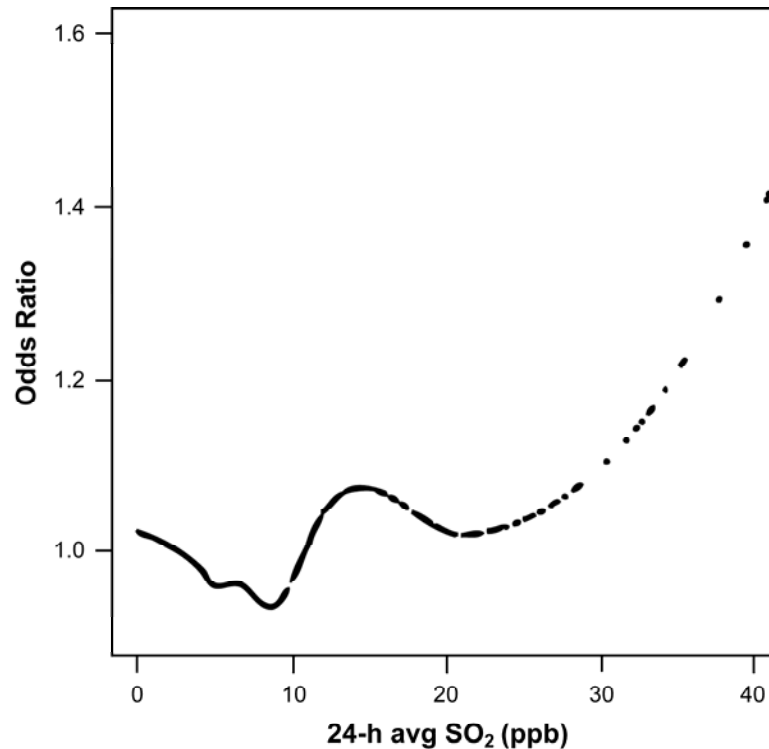


Figure 4-5. Relative odds ratio of incidence of lower respiratory tract symptoms smoothed against 24-h avg SO₂ concentrations on the previous day, controlling for temperature, city, and day of week.

Source: Schwartz et al. (1994).

3 A study by Jaffe et al. (2003) examined the association between SO₂ and emergency
4 department (ED) visits for asthma in three cities in Ohio, and found significant associations only
5 in Cincinnati using Poisson regression analysis. To examine the concentration-response function,
6 they also conducted quintile analyses. In Cincinnati, an increasing linear trend in risk was
7 observed across the range of concentrations. Wong et al. (2002; using GAM with default
8 convergence criteria) constructed a plot of risk against 24-h avg SO₂ concentrations to examine
9 the concentration-response relationship in Hong Kong and London. In general, a linear
10 relationship between risk of respiratory hospitalizations and SO₂ was observed across the range

of SO₂ concentrations in Hong Kong, but not in London. Several other studies that examined the concentration-response relationship found that the association between respiratory hospitalizations and SO₂ did not deviate from linearity (Atkinson et al., 1999a; Burnett et al., 1997a; 1997b; 1999; Hajat et al., 2002).

4.1.3. Summary of Evidence on Concentration-Response Functions and Thresholds

In the previous two sections, evidence from human clinical and epidemiological studies on the concentration-response function that may inform identification of any potential population threshold was presented. Evidence from human clinical studies indicates a clear increase in the magnitude of respiratory effects with increasing exposure concentration between 0.2 and 1.0 ppm with 5-10 min SO₂ exposures. Epidemiological studies also have observed generally increasing trends across the entire range of SO₂ exposures; however, in a limited number of studies a marked increase in effect was observed only at the higher concentrations (above 90th percentile values).

Discerning a possible population-level threshold for air pollution-related effects in epidemiological studies is quite challenging. Using PM_{2.5} as an example, Brauer et al. (2002) examined the relationship between ambient concentrations and mortality risk in a simulated population with specified common individual threshold levels. They found that no population threshold was detectable when a low threshold level was specified. Even at high-specified individual threshold levels, the apparent threshold at the population level was much lower than specified. Brauer et al. (2002) concluded that the use of surrogate measures of exposure (i.e., those from centrally located ambient monitors) that were not highly correlated with personal exposures made it difficult to discern population-level thresholds in epidemiological studies even if a common threshold exists for individuals within the population.

The wide interindividual variability in sensitivity to SO₂ exposure further hinders the ability to discern a potential threshold level in population studies. Human clinical studies have shown that asthmatics experience greater increases in sRaw following peak SO₂ exposures compared to healthy individuals (Linn et al., 1987). Among asthmatics, interindividual differences in response also have been noted, with some asthmatics experiencing SO₂-related effects at much lower levels than others (Horstman et al., 1986).

Another factor that complicates the identification of a possible threshold of effects is that currently deployed ambient monitors may be inadequate for accurate and precise measurements at lower 24-h avg SO₂ levels. Ambient concentrations of SO₂ have been declining since the 1980s and are now at or very near the limit of detection of the ambient monitors in the regulatory network. The mean 24-h avg SO₂ concentration across the metropolitan statistical areas (MSAs) from 2003 through 2005 was 4 ppb (5th–95th percentile: 1, 13). Thus, there is greater uncertainty at the lower concentration range compared to the higher concentrations, which likely limits the ability to detect any clear-cut potential threshold.

In conclusion, evidence from human clinical studies indicates wide interindividual variability in response to SO₂ exposures, with peak (5 to 10 min) exposures at levels as low as 0.2–0.3 ppm eliciting respiratory responses in some asthmatic individuals. Several epidemiological studies that examined the concentration-response function between short-term (24-h avg or 1-h max) exposure to SO₂ and respiratory morbidity observed that the relationship was linear across the entire concentration range, suggesting a lack of a threshold in effect. However, given the various limitations in observing a possible threshold in population studies, the lack of evidence does not necessarily indicate that there is indeed no threshold in SO₂ health effects. Some epidemiological studies did report that though there was generally an increasing trend at the lower SO₂ concentrations, a marked increase in SO₂-related respiratory health effects was observed at higher concentrations. However, as these observations were based on a few potentially influential data points (24-h avg SO₂ concentrations above the 90th percentile), the results should be interpreted with caution. The overall limited evidence from epidemiological studies examining the concentration-response function of SO₂ health effects is inconclusive regarding the presence of an effect threshold at current ambient levels.

4.2. Susceptible and Vulnerable Populations

Interindividual variation in human responses to air pollutants indicates that not all individuals exposed to pollutants respond similarly. That is, some subpopulations are at increased risk to the detrimental effects of pollutant exposure. The NAAQS are intended to provide an adequate margin of safety for both general populations and sensitive subpopulations, or those subgroups potentially at increased risk for ambient air pollution health effects. For the purposes of this document, a susceptible population is defined as one that might exhibit an adverse health

effect to a pollutant at concentrations lower than those needed to elicit the same response in the general population, and a vulnerable population as one that might be differentially exposed to higher concentrations of a pollutant than the general population, regardless of health outcome. The previous review of the SO₂ NAAQS identified certain groups within the population that may be more susceptible to the effects of SO₂ exposure, including asthmatics, individuals not diagnosed as asthmatic but with atopic disorders (e.g., allergies), and individuals with COPD or cardiovascular disease. Other subgroups considered to be somewhat sensitive in this document include children and older adults; people with other respiratory disease; genetic factors; socioeconomic status (SES); and populations experiencing heightened exposure levels (e.g., those living near roadways or other “hot spots” or engaged in outdoor work or exercise). Also of concern are individuals who generally may not be susceptible to SO₂-related health effects but may experience transient airways reactivity to respiratory irritants such as SO₂ following a recent viral respiratory infection (Stempel and Boucher, 1981). These groups comprise a large fraction of the U.S. population. Given the likely heterogeneity of individual responses to air pollution, the severity of health effects experienced by a susceptible subgroup may be much greater than that experienced by the population at large (Zanobetti et al., 2000).

4.2.1. Preexisting Disease as a Potential Risk Factor

Several researchers have investigated the effect of air pollution among potentially susceptible groups with preexisting medical conditions. A recent report of the National Research Council emphasized the need to evaluate the effect of air pollution on susceptible groups, including those with respiratory illnesses and cardiovascular diseases (National Research Council., 2004). Generally, asthma, COPD, conduction disorders, congestive heart failure (CHF), diabetes, and myocardial infarction (MI) are conditions believed to put persons at greater risk of adverse events associated with air pollution. Asthmatics are known to be one of the most SO₂-responsive subgroups in the population; the evidence related to respiratory illness, including asthma and other factors, is discussed in further detail below.

4.2.1.1. Individuals with Respiratory Diseases

The 1982 AQCD concluded that asthmatics are more susceptible to respiratory effects from SO₂ exposures than the general public. This conclusion was primarily drawn from the strong

1 human clinical evidence. Recent epidemiological studies have strengthened this conclusion,
2 reporting associations between a range of health outcomes with both short-term and long-term
3 SO₂ exposures in subjects with respiratory disease.

4 In human clinical studies, asthmatics have been shown to be more responsive to respiratory
5 effects of SO₂ exposures than healthy non-asthmatics. While SO₂-attributable decrements in lung
6 function generally have not been demonstrated at concentrations < 1.0 ppm in non-asthmatics
7 (Lawther et al., 1975; Linn et al., 1987; Schachter et al., 1984), statistically significant increases
8 in respiratory symptoms and decreases in lung function have been observed in exercising
9 asthmatics following peak (5 to 10 min) SO₂ exposures to concentrations as low as 0.4-0.6 ppm
10 (Gong et al., 1995; Horstman et al., 1986; Linn et al., 1983). Respiratory effects have been
11 observed in some sensitive asthmatics at concentrations as low as 0.2-0.3 ppm (Horstman et al.,
12 1986; Linn et al., 1987). There is no evidence that individuals with COPD have increased
13 susceptibility to SO₂-induced respiratory effects.

14 A number of epidemiological studies reported increased respiratory morbidity associated
15 with SO₂ exposures in asthmatics and atopic individuals. Notably, two U.S. multicity studies
16 observed associations between ambient SO₂ concentrations and respiratory symptoms in
17 asthmatic children (Mortimer et al., 2002; Schildcrout et al., 2006). Additional studies also have
18 generally indicated positive associations for asthma among children and included a U.S. study
19 (Delfino et al., 2003) and several European studies (Higgins et al., 1995; Neukirch et al., 1998;
20 Peters et al., 1996; Roemer et al., 1993; Segala et al., 1998; Taggart et al., 1996; Timonen and
21 Pekkanen, 1997; van der Zee et al., 1999). Studies of adults found no consistent association
22 between respiratory symptoms among asthmatics and SO₂ concentrations (Desqueyroux et al.,
23 2002a; 2002b; Romieu et al., 1996; van der Zee et al., 2000).

24 A suggestive association between ambient SO₂ concentrations and ED visits and
25 hospitalizations provides further evidence that asthmatics are susceptible to the effects of SO₂.
26 The associations between ambient concentrations of 24-h avg SO₂ and ED visits and
27 hospitalizations for asthma in the United States are generally positive (Jaffe et al., 2003; Lin et
28 al., 2004; Michaud et al., 2004; Wilson et al., 2005), though a large time-series study conducted
29 in Atlanta, GA did not find an association between ambient 1-h max SO₂ levels and ED visits
30 (Peel et al., 2005). Studies conducted outside the United States (Atkinson et al., 1999a; Hajat et
31 al., 1999; Sunyer et al., 1997; Thompson et al., 2001) also generally found positive results.

1 In summary, substantial evidence from epidemiological studies suggests that individuals
2 with preexisting respiratory diseases, particularly asthma, are more susceptible to respiratory
3 health effects, though not mortality, from SO₂ exposures than the general public. The
4 observations from human clinical studies indicating increased sensitivity to SO₂ exposures in
5 asthmatic subjects compared to healthy subjects provide coherence and biological plausibility for
6 these observations in epidemiological studies.

4.2.1.2. Individuals with Cardiovascular Diseases

7 The evidence available to evaluate the susceptibility of individuals with cardiovascular
8 disease for SO₂-related health effects is very limited. One human clinical study observed no
9 evidence to suggest that patients with stable angina were more susceptible to SO₂-related health
10 effects compared with healthy subjects (Routledge et al., 2006). The authors noted that this lack
11 of response in the heart patients may be due to a drug treatment effect rather than decreased
12 susceptibility, as a large portion of the angina patients were taking beta blockers, which are
13 known to increase indices of cardiac vagal control.

14 Liao et al. (2004) investigated short-term associations between ambient pollutants and
15 cardiac autonomic control and observed that consistently more pronounced associations were
16 found between SO₂ and heart rate variability among persons with a history of coronary heart
17 disease. In another epidemiological study, Henneberger et al. (2005) examined the association of
18 repolarization parameters with air pollutants in East German men with preexisting coronary heart
19 disease. Ambient SO₂ concentrations during the 24-h preceding the ECG were associated with
20 the QT interval duration, but not with any other repolarization parameters.

21 Evidence is inconsistent in studies analyzing the associations between ambient levels of air
22 pollutants and ED visits or hospitalizations for cardiovascular diseases. A recent epidemiological
23 study investigated the association of SO₂ with cardiac-related hospital admissions among persons
24 with preexisting cardiopulmonary conditions and observed no associations with ambient 1-h max
25 SO₂ level for any cardiac disease investigated (i.e., ischemic heart disease [IHD], CHF, and
26 dysrhythmia) across strata of comorbid disease status, including hypertension, diabetes, and
27 COPD (Peel et al., 2007).

28 Goldberg et al. (2003) compared the risk estimates for death with the underlying cause of
29 CHF and those deaths classified as having CHF 1 year before death and did not find associations

1 between air pollution and those with CHF as an underlying cause of death. The authors found
2 associations between some of the air pollutants examined (coefficient of haze [CoH], SO₂, and
3 NO₂) and the deaths that were classified as having CHF 1 year before death, but the association
4 with the specific cause of death was not unique to SO₂. This pattern of association, including but
5 not specific to SO₂, with specific causes of death also was observed in an additional cohort of
6 patients with CHF (Kwon et al., 2001).

7 In conclusion, the very limited evidence examining the susceptibility of individuals with
8 preexisting cardiovascular disease to adverse health effects from ambient SO₂ exposures is
9 inconclusive.

4.2.2. Genetic Factors for Oxidant and Inflammatory Damage from Air Pollutants

10 A consensus now exists among scientists that genetic factors related to health outcomes
11 and ambient pollutant exposures merit serious consideration (Gilliland et al., 1999; Kauffmann,
12 2004). Several criteria must be satisfied in selecting and establishing useful links between
13 polymorphisms in candidate genes and adverse respiratory effects. First, the product of the
14 candidate gene must be significantly involved in the pathogenesis of the effect of interest, which
15 is often a complex trait with many determinants. Second, polymorphisms in the gene must
16 produce a functional change in either the protein product or in the level of expression of the
17 protein. Third, in epidemiological studies, the issue of confounding by other genes or
18 environmental exposures must be carefully considered.

19 Several glutathione S-transferase (GST) families have common, functionally important
20 polymorphic alleles (e.g., homozygosity for the null allele at the GSTM1 and GSTT1 loci,
21 homozygosity for the A105G allele at the GSTP1 locus) that significantly reduce expression of
22 enzyme function in the lung. Exposure to radicals and oxidants from air pollution induces
23 decreases in GSH that increase GST transcription. Individuals with genotypes that result in
24 enzymes with reduced or absent glutathione peroxidase activity are likely to have reduced
25 oxidant defenses and increased susceptibility to inhaled oxidants and radicals.

26 Gilliland et al. (2002) examined effects of GSTM1, GSTT1, and GSTP1 genotypes and
27 acute respiratory illness, specifically respiratory illness-related absences from school. The goal
28 was to examine potential susceptibilities on this basis, but not specifically to air pollutants. They

1 concluded that fourth grade schoolchildren who inherited a GSTP1 Val-105 variant allele had a
2 decreased risk of respiratory illness-related school absences, indicating that GSTP1 genotype
3 influences the risk and/or severity of acute respiratory infections in school-aged children.

4 Lee et al. (2004) studied ninth grade schoolchildren with asthma in Taiwan for a gene-
5 environmental interaction between GSTP1-105 genotypes and outdoor pollution. They examined
6 general district air pollution levels of low (mean SO₂ level of 3.6 ppb from 1994 to 2001),
7 moderate (mean SO₂ of 6.2 ppb), and high (mean SO₂ of 8.6 ppb) and found that compared with
8 individuals with any Val-105 allele in the low air pollution district, Ile-105 homozygotes in the
9 high air pollution district had a significantly increased risk of asthma.

10 Gauderman et al. (2007) describe a study method that uses principal components analysis
11 computed on single nucleotide polymorphism (SNP) markers to test for an association between a
12 disease and a candidate gene. For example, they evaluated the association between respiratory
13 symptoms in children and four SNPs in the GSTP1 locus, using data from the Southern
14 California Children's Health Study (CHS). The authors observed stronger evidence of an
15 association using the principal components approach ($p = 0.044$) than using either a genotype-
16 based ($p = 0.13$) or haplotype-based ($p = 0.052$) approach. This method may be applied to
17 relationships in this and other databases to evaluate aspects of air pollutants such as SO₂.

18 Winterton et al. (2001) attempted to identify a genetic biomarker for susceptibility to SO₂.
19 They screened 62 asthmatic subjects for SO₂ responsiveness using an inhalation challenge and
20 collected genetic material via buccal swabs to test for associations between SO₂ sensitivity and
21 specific gene polymorphisms. Subjects inhaled 0.5 ppm SO₂ by mouthpiece for 10 min while
22 wearing noseclips during moderate exercise on a treadmill. Subjects were defined as SO₂-
23 sensitive if FEV₁ was decreased 12%. Genetic polymorphisms as biomarkers of susceptibility
24 were evaluated in five regions coding for the β 2-adrenergic receptor, the α subunit of the
25 interleukin-4 (IL-4) receptor, the Clara cell secretory protein (CC16), tumor necrosis factor- α
26 (TNF- α), and lymphotoxin- α (also known as TNF- β). The authors found a significant association
27 between response to SO₂ and the homozygous wild-type allele of TNF- α . All of the SO₂-
28 sensitive subjects had the homozygous wild-type allele for TNF- α , while 61% of the
29 nonresponders had this genotype. Homozygosity for the TNF-1 allele was associated with a 5-
30 fold increased risk of physician-diagnosed asthma relative to other genotypes. None of the other
31 polymorphisms showed significant trends.

1 In summary, the differential effects of air pollution among genetically diverse
2 subpopulations have been examined for a number of GST genes and other genotypes. The
3 limited number of studies may provide some insight into susceptible groups and a potential
4 genetic role in such. Only one of these studies specifically examined SO₂ as the exposure of
5 interest, and it found a significant association with the homozygous wild-type allele for TNF- α .
6 Khoury et al. (2005) states that while genomics is still in its infancy, opportunities exist for
7 developing, testing, and applying its tools to public health research of outcomes with possible
8 environmental causes. At this time, there are insufficient data on which to base a conclusion
9 regarding the effect of SO₂ exposure on genetically distinct subpopulations.

4.2.3. Age-Related Susceptibility

10 The American Academy of Pediatrics (2004) notes that children and infants are among the
11 most susceptible to many air pollutants, including SO₂. Eighty percent of alveoli are formed
12 postnatally and changes in the lung continue through adolescence; furthermore, the developing
13 lung is highly susceptible to damage from exposure to environmental toxicants (Dietert et al.,
14 2000). Children also have increased vulnerability as they spend more time outdoors, are highly
15 active, and have high minute ventilation, which collectively increase the dose they receive
16 (Plunkett et al., 1992; Wiley, 1991a; 1991b). In addition to children, the elderly are frequently
17 classified as being particularly susceptible to air pollution. The basis of the increased sensitivity
18 in the elderly is not known, but one possibility is that it may be related to changes in the
19 respiratory tract lining fluid antioxidant defense network (Kelly and Mudway, 2003) or a general
20 reduction in immune competence.

21 Adverse respiratory effects have been observed in adolescents following SO₂ exposure in a
22 laboratory setting (Koenig et al., 1981; 1983; 1987; 1988; 1990). However, there is no evidence
23 from human clinical studies to suggest that the respiratory effects in adolescents are more severe
24 than those observed in adults. Similarly, a number of epidemiological studies have observed
25 increased respiratory symptoms in children associated with increasing SO₂ exposures (Mortimer
26 et al., 2002; Schildcrout et al., 2006; Schwartz et al., 1994), though there is no evidence from a
27 limited number of studies suggesting this same effect in adults (Desqueyroux et al., 2002a;
28 Romieu et al., 1996; van der Zee et al., 2000).

1 A number of studies, investigating the association between ambient SO₂ levels and ED
2 visits or hospital admissions for all respiratory causes or asthma, stratified their analyses by age
3 group. Figure 4-6 summarizes the evidence of age-specific associations between SO₂ and acute
4 respiratory ED visits and hospitalizations. Several studies demonstrated that the excess risk of
5 ED visits or hospitalizations for all respiratory causes or asthma was higher for children
6 (Anderson et al., 2001; Atkinson et al., 1999a; Atkinson et al., 1999b; Petroeschevsky et al.,
7 2001) and older adults (Anderson et al., 1998; Petroeschevsky et al., 2001; Ponce de Leon et al.,
8 1996; Wilson et al., 2005) when compared to the risk for all ages together. Increased risks for
9 children and older adults were more prevalent in the studies of all respiratory disease than those
10 considering asthma as the outcome.

11 Cakmak et al. (2007) reported that among seven Chilean urban centers, the percent
12 increase in nonaccidental mortality associated with a 10 ppb increase in 24-h avg SO₂ was 3.4%
13 (95% CI: 0.7, 6.1) for those < 65 years of age and 5.6% (95% CI: 2.2, 9.1) for those > 85 years
14 of age. The authors concluded that the elderly are particularly susceptible to dying from air
15 pollution, and suggested that concentrations deemed acceptable for the general population may
16 not adequately protect the very elderly.

17 There is limited epidemiologic evidence to suggest that children and older adults (65+
18 years) are more susceptible to the adverse respiratory effects associated with ambient SO₂
19 concentrations when compared to the general population.

4.2.4. Other Potentially Susceptible Populations

20 There are a number of other potentially susceptible groups that, while not included here
21 due to a lack of data specific to SO₂ exposures, deserve mention in this document. These include
22 obese individuals, individuals in a chronic pro-inflammatory state (e.g., diabetics), and children
23 born prematurely or with low birth weight.

24 Enhanced susceptibility for air pollution-related cardiovascular events has been shown for
25 older individuals and persons with conditions associated with chronic inflammation, such as

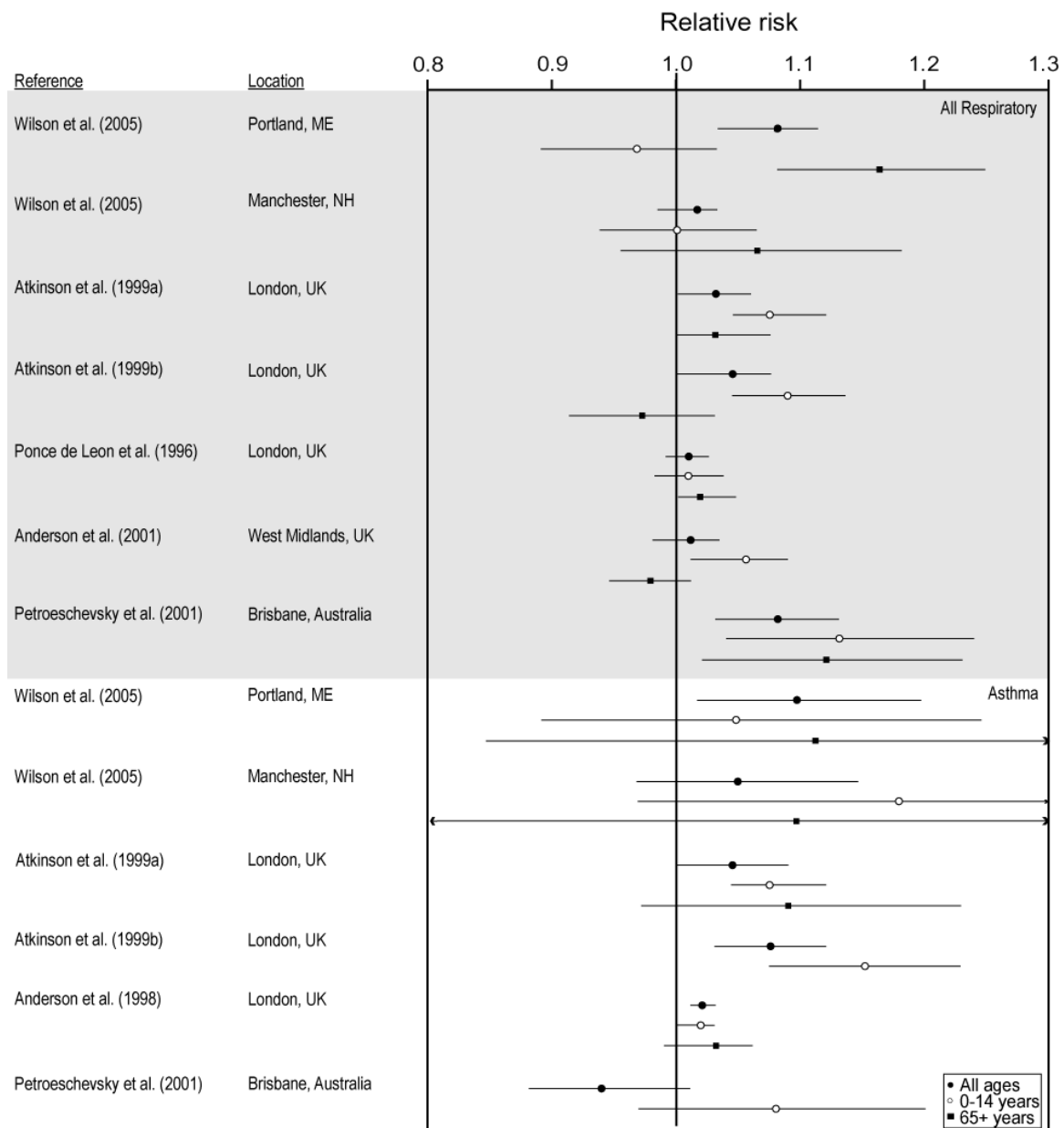


Figure 4-6. Relative risks (95% CI) of age-specific associations between short-term exposure to SO₂ and respiratory ED visits and hospitalizations. Risk estimates are standardized per 10 ppb increase in 24-h avg SO₂ concentrations or 40 ppb increase in 1-h max SO₂.

- 1 diabetes, coronary artery disease, and past myocardial infarctions (Bateson and Schwartz, 2004;
- 2 Goldberg et al., 2001; Zanobetti and Schwartz, 2002). Dubowsky et al. (2006) observed that
- 3 individuals with conditions associated with both chronic inflammation and increased cardiac risk
- 4 were more vulnerable to the short-term pro-inflammatory effects of air pollution. This included

1 individuals with diabetes; obesity; and concurrent diabetes, obesity and hypertension. Zanobetti
2 and Schwartz (2001) reported more than twice the risk for hospital admissions for heart disease
3 in persons with diabetes than in persons without diabetes associated with exposure to ambient air
4 pollution, indicating that persons with diabetes are an important at-risk group. Data from the
5 Third National Health and Nutrition Examination Survey indicate that 5.1% of the U.S.
6 population older than 20 years of age have diagnosed diabetes and an additional 2.7% have
7 undiagnosed diabetes (Harris et al., 1998). Moreover, another study found that subjects with
8 impaired glucose tolerance without type II diabetes also had reduced heart rate variability
9 (Schwartz, 2001). This suggests the at-risk population may be even larger.

10 Mortimer et al. (2000) reported that among asthmatic children, birth characteristics
11 continue to be associated with increased susceptibility to air pollution later in life, demonstrating
12 that air pollution-induced asthma symptoms are more severe in children born prematurely or of
13 low birth weight. Specifically, the authors revealed asthmatic children born more than three
14 weeks prematurely or weighing less than 2,500 grams (5.5 pounds) had a six-fold decrease in
15 breathing capacity associated with air pollution compared to full-weight, full-term children. The
16 low birth weight and premature children also reported a five-fold greater incidence of symptoms
17 like wheezing, coughing and tightness in the chest.

4.2.5. Factors that Potentially Increase Vulnerability to SO₂

18 A limited amount of information exists on exposures to SO₂ among vulnerable populations.
19 Because indoor and personal SO₂ concentrations are generally much lower than outdoor or
20 ambient measurements, individuals that spend most of their time indoors, such as older adults,
21 are not anticipated to be vulnerable to high SO₂ exposures, though in some cases they may be
22 more susceptible to the effects of these exposures than the general population due to preexisting
23 health factors.

24 Other individuals with increased vulnerability include those who spend a lot of time
25 outdoors at increased exertion levels, for example outdoor workers and individuals who exercise
26 or play outdoor sports. Exercise may cause an increase in uptake of SO₂ resulting from an
27 increase in ventilation rate and accompanying shift from nasal to oronasal breathing. Children,
28 who generally spend more time playing outdoors, may qualify as both a susceptible population

1 due to their developing physiology and as a vulnerable population since ambient SO₂
2 concentrations are several-fold higher than indoor concentrations.

3 Residential location is not as strong of a predictor of exposure vulnerability for SO₂ as for
4 traffic-related pollutants, because meteorological conditions have a greater impact on pollutant
5 plume direction from primary point sources such as coal-fired power plants.

6 Social economic status (SES) is a known determinant of health and there is evidence that
7 SES modifies the effects of air pollution (Makri and Stilianakis, 2007; O'Neill et al., 2003). Both
8 higher exposures to air pollution and greater susceptibility to its effects may contribute to a
9 complex pattern of risk among those with lower SES. Conceptual frameworks have been
10 proposed to explain the relationship between SES, susceptibility and exposure to air pollution.
11 Common to these frameworks is the consideration of the broader social context in which people
12 live and its effect on health in general (Gee and Payne-Sturges, 2004; O'Neill et al., 2003), as
13 well as on maternal and child health (Morello-Frosch and Shenassa, 2006), and asthma (Wright
14 et al., 2007) specifically. Multilevel modeling approaches that allow parameterization of
15 community level stressors such as increased life stress, as well as individual risk factors, are
16 considered by these authors. In addition, statistical methods that allow for temporal and spatial
17 variability in exposure and susceptibility, have been discussed in the recent literature (Jerrett and
18 Finkelstein, 2005; Kunzli, 2005).

19 Most studies to date have examined modification by SES indicators on the association
20 between mortality and PM (Finkelstein et al., 2003; Jerrett et al., 2004; Martins et al., 2004;
21 O'Neill et al., 2003; Romieu et al., 2004) or other indices such as traffic density, distance to
22 roadway or a general air pollution index (Finkelstein et al., 2005; Ponce et al., 2005; Woodruff et
23 al., 2003). However, modification of SO₂ associations has been examined in a few studies. For
24 example, in a study conducted in 10 large Canadian cities, living in communities in which
25 individuals have lower household education and income levels increased the individual's
26 vulnerability to air pollution (Cakmak et al., 2006). These effects were statistically significant for
27 several gaseous criteria pollutants, but not for SO₂. In addition, Finkelstein et al. (2003)
28 evaluated neighborhood levels of income and air pollution in southern Ontario, Canada. They
29 found that both income and SO₂ levels were associated with mortality differences. Specifically,
30 among people with below-median income, the relative risk for those with above-median
31 exposure to SO₂ was 1.18 (95% CI: 1.11, 1.26); the corresponding relative risk among subjects

1 with above-median income was 1.03 (95% CI: 0.83, 1.28). Overall, there is very limited
2 evidence available from which conclusions on the human health effects from the interaction
3 between SES and SO₂ can be drawn.

4.3. Potential Public Health Impacts

4 Exposure to ambient SO₂ is associated with a variety of outcomes including increases in
5 respiratory symptoms, particularly among asthmatic children, and ED visits and hospital
6 admissions for respiratory diseases among children and older adults (65+ years). In protecting
7 public health, a distinction must be made between health effects that are considered “adverse”
8 and those that are not. What constitutes an adverse health effect varies for different population
9 groups. Some changes in healthy individuals are not viewed as adverse while those of similar
10 type and magnitude in other susceptible individuals with preexisting disease are.

4.3.1. Concepts Related to Defining Adverse Health Effects

11 The American Thoracic Society (ATS) published an official statement titled “What
12 Constitutes an Adverse Health Effect of Air Pollution?” (ATS, 2000a). This statement updated
13 the guidance for defining adverse respiratory health effects that had been published 15 years
14 earlier (Society, 1985), taking into account new investigative approaches used to identify the
15 effects of air pollution and reflecting concern for impacts of air pollution on specific susceptible
16 groups. In the 2000 update, there was an increased focus on quality of life measures as indicators
17 of adversity and a more specific consideration of population risk. Exposure to air pollution that
18 increases the risk of an adverse effect to the entire population is viewed as adverse, even though
19 it may not increase the risk of any identifiable individual to an unacceptable level. For example,
20 a population of asthmatics could have a distribution of lung function such that no identifiable
21 single individual has a level associated with significant impairment. Exposure to air pollution
22 could shift the distribution to lower levels that still do not bring any identifiable individual to a
23 level that is associated with clinically relevant effects. However, this shift to a lower level of lung
24 function would be considered adverse because individuals within the population would have
25 diminished reserve function and, therefore, would be at increased risk if affected by another
26 agent.

Reflecting new investigative approaches, the ATS statement also describes the potential usefulness of research into the genetic basis for disease, including responses to environmental agents that provide insights into the mechanistic basis for susceptibility and provide markers of risk status. Likewise, biomarkers that are indicators of exposure, effect, or susceptibility may someday be useful in defining the point at which one or an array of responses should be considered an adverse effect.

The 2006 Ozone AQCD (EPA, 2006) provided information helpful in defining adverse health effects associated with ambient O₃ exposure by describing the gradation of severity and adversity of respiratory-related O₃ effects. The definitions that relate to responses in impaired individuals are reproduced and presented here in Table 4-1. The severity of effects described in the tables and the approaches taken to define their relative adversity are valid and reasonable in the context of the new ATS (2000b) statement.

Table 4-1. Gradation of individual responses to short-term SO₂ exposure in individuals with impaired respiratory systems.

FUNCTIONAL RESPONSE	NONE	SMALL	MODERATE	LARGE
FEV ₁ change	Decrements of < 3%	Decrements of 3 - 10%	Decrements of 10 - 20%	Decrements of > 20%
Nonspecific bronchial responsiveness ^a	Within normal range	Increases of < 100%	Increase of < 300%	Increases of > 300%
Airway resistance (sRaw)	Within normal range (±20%)	sRaw increased < 100%	sRaw increased up to 200% or up to 15 cm H ₂ O · s	sRaw increased > 200% or more than 15 cm H ₂ O · s
Duration of response	None	< 4 h	4 h - 24 h	> 24 h
SYMPTOMATIC RESPONSE	NORMAL	MILD	MODERATE	SEVERE
Wheeze	None	With otherwise normal breathing	With shortness of breath	Persistent with shortness of breath
Cough	Infrequent cough	Cough with deep breath	Frequent spontaneous cough	Persistent uncontrollable cough
Chest pain	None	Discomfort just noticeable on exercise or deep breath	Marked discomfort on exercise or deep breath	Severe discomfort on exercise or deep breath
Duration of response	None	< 4 h	4 h – 24 h	> 24 h
IMPACT OF RESPONSES	NORMAL	MILD	MODERATE	SEVERE
Interference with normal activity	None	Few individuals choose to limit activity	Many individuals choose to limit activity	Most individuals choose to limit activity
Medical treatment	No change	Normal medication as needed	Increased frequency of medication use or additional medication	Physician or emergency room visit

^aAn increase in nonspecific bronchial responsiveness of 100% is equivalent to a 50% decrease in PD₂₀ or PD₁₀₀. This table is adapted from the 2006 Ozone AQCD (Table 8-3, page 8-68) (EPA, 26).

1 As assessed in detail in earlier chapters of this document and briefly recapitulated in
2 preceding sections of this chapter, exposures to a range of SO₂ concentrations have been reported
3 to be associated with increasing severity of health effects, ranging from respiratory symptoms to
4 ED visits and hospital admission for respiratory causes. Respiratory effects associated with
5 short-term SO₂ exposures have been by far the most extensively studied and most clearly shown
6 to be causally related to SO₂ exposure. Such effects are observed among children, older adults,
7 and persons with respiratory impairments.

4.3.2. Estimation of Potential Numbers of Persons in At-Risk Susceptible Population Groups in the United States

8 Although SO₂-related health risk estimates may appear to be small, they may be significant
9 from an overall public health perspective due to the large numbers of individuals in the potential
10 risk groups. Several subpopulations have been identified as possibly having increased
11 susceptibility or vulnerability to adverse health effects from SO₂, including children, older adults,
12 and individuals with preexisting pulmonary diseases. One consideration in the assessment of
13 potential public health impacts is the size of various population groups that may be at increased
14 risk for health effects associated with SO₂-related air pollution exposure. Table 4-2 summarizes
15 information on the prevalence of chronic respiratory conditions in the U.S. population in 2004
16 and 2005 (NHIS, 2006a, 2006b).

17 Of most concern are those individuals with preexisting respiratory conditions, with
18 approximately 10% of adults and 13% of children having been diagnosed with asthma and 6% of
19 adults with COPD (chronic bronchitis and/or emphysema). As summarized in Section 3.1.3.5,
20 human clinical studies indicate that a significant fraction (20-60%) of asthmatic individuals
21 experience moderate or greater decrements in lung function as well as increased respiratory
22 symptoms following peak (5-10 min) SO₂ exposures to concentrations of as low as 0.4-0.6 ppm
23 (Table 3-1). Some sensitive asthmatics (5-20%) have been shown to experience moderate
24 decrements in lung function at concentrations between 0.2 and 0.3 ppm. Among asthmatics, both
25 the magnitude of SO₂-induced decrements in lung function as well as the percent of individuals
26 affected have been shown to increase with increasing exposure concentrations between 0.2 and
27 1.0 ppm.

In addition, subpopulations based on age group also comprise substantial segments of the population that may be potentially at risk for SO₂-related health impacts. Based on U.S. Census data from 2000, about 72.3 million (26%) of the U.S. population are under 18 years of age, 18.3 million (7.4%) are under 5 years of age, and 35 million (12%) are 65 years of age or older. Hence, large proportions of the U.S. population are included in age groups that are considered likely to have increased susceptibility and vulnerability for health effects from ambient SO₂ exposure. For example, Figure 4-6 demonstrates that the SO₂-related excess risk for asthma is, on average, 50% higher among children when compared to risk estimates that include all ages with a 10 ppb increase in 24-h avg SO₂ concentration.

Table 4-2. Prevalence of selected respiratory disorders by age group in the United States (2004 [U.S. adults] and 2005 [U.S. children] National Health Interview Survey).

CHRONIC CONDITION/DISEASE ADULTS (18+ YEARS)	AGE (YEARS)					
	ALL ADULTS		18-44	45-64	65-74	75+
	CASES (× 10 ⁶)	%	%	%	%	%
Respiratory Conditions: Asthma	14.4	6.7	6.4	7.0	7.5	6.6
COPD: Chronic Bronchitis	8.6	4.2	3.2	4.9	6.1	6.3
COPD: Emphysema	3.5	1.7	0.3	2	4.9	6.0
CHRONIC CONDITION/DISEASE CHILDREN (<18 YEARS)	ALL CHILDREN		0-4	5-11	12-17	
	CASES (× 10 ⁶)	%	%	%	%	
Respiratory Conditions	6.5	8.9	6.8	9.9	9.6	

Source: National Center for Health Statistics (2006a,b)

Evidence indicates that several groups are potentially at increased risks from SO₂ exposures compared to the average population. Susceptible subgroups include individuals with preexisting disease, especially asthma, and children and older adults. Other individuals with potentially increased vulnerability include those who spend a lot of time outdoors at increased exertion levels (e.g., outdoor workers, children, individuals who exercise or play sports) and those in proximity to large uncontrolled or poorly controlled sources. The considerable size of the population groups at risk indicates that exposure to ambient SO₂ could have a potentially significant impact on public health in the United States, with the greatest public health risks for the smaller subset of susceptible individuals exposed to relatively high peak SO₂ concentrations.

Chapter 5. Summary and Conclusions

Previous chapters present the most policy-relevant information related to the review of the NAAQS for SO_x, which are 0.14 ppm averaged over a 24-hour period not to be exceeded more than once per year, and 0.030 ppm annual arithmetic mean, with SO₂ as the indicator. This chapter summarizes and integrates key findings from atmospheric sciences, ambient air data analyses, exposure assessment, dosimetry, and health evidence. The EPA framework for causal determinations described in Chapter 1 has been applied to the body of evidence in order to judge the scientific data about exposure to SO_x and health effects in a two-step process. This framework draws from similar efforts across the Federal government and the wider scientific community, especially from the recent NAS Institute of Medicine document *Improving the Presumptive Disability Decision-Making Process for Veterans* (IOM, 2007). The first step is to determine the weight of evidence in support of causation at relevant pollutant exposures and characterize the strength of any resulting causal classification. The EPA framework applied here employs a five-level hierarchy for causal determination to be consistent with the *Guidelines for Carcinogen Risk Assessment* (EPA, 2005):

- Sufficient to infer a causal relationship
- Sufficient to infer a likely causal relationship (i.e., more likely than not)
- Suggestive but not sufficient to infer a causal relationship
- Inadequate to infer the presence or absence of a causal relationship
- Suggestive of no causal relationship

The second step evaluates the quantitative evidence regarding the concentration-response relationships including levels and exposure durations at which effects are observed. These two steps characterize the health effects of SO_x and levels at which effects may occur.

5.1. Emissions and Ambient Concentrations of SO₂

Anthropogenic SO₂ is emitted mainly by fossil fuel combustion (chiefly coal and oil) and metal smelting. The largest source of emissions is from elevated point sources such as the stacks of power plants and industrial facilities. Since 1990, in response to controls applied under the Acid Rain Program (EPA, 2006), SO₂ emissions from these sources have declined substantially.

1 Emissions demonstrate a strong gradient increasing from west to east, owing to the high
2 concentration of SO₂-emitting electric generating utilities in the Ohio River Valley and regions to
3 the south. PRB levels of SO₂ are estimated to be in the range of a few hundredths of a ppb (< 1%
4 of typical ambient levels) across most of the United States, though much higher values are found
5 in areas affected by volcanic or geothermal activity or in areas affected by episodic transport of
6 high concentration plumes from Asia and Europe.

7 The levels of the current primary NAAQS for SO_x are 0.14 ppm for 24-h avg SO₂
8 concentrations and 0.03 ppm for an annual avg SO₂ concentration. Exceedances in recent years
9 have become rare, as the mean 24-h and annual avg SO₂ concentrations in the United States for
10 the years 2003 to 2005 were ~4 ppb, with 99th percentile values of ~25 ppb for the 24-h avg, and
11 ~15 ppb for both the 99th percentile and max values of the annual avg. Mean 1-h max
12 concentrations in these years were ~13 ppb, with a 99th percentile value of ~120 ppb and
13 maximum value of ~700 ppb. The large differences between 99th percentile and maximum values
14 for the shorter term averages suggest that the maxima are strongly limited spatially and
15 temporally and are not a major determinant of the mean values. The nonuniform spatiotemporal
16 distribution of 5-min data, which are voluntarily supplied from a very few monitors without a
17 specific regulatory mandate makes them very difficult to use quantitatively for determining
18 concentrations and exposures at this very short time duration.

5.2. Health Effects of SO₂

19 Evaluation of the health evidence, with consideration of issues related to atmospheric
20 sciences, exposure assessment, and dosimetry, led to the conclusion that it is *sufficient to infer a*
21 *causal relationship between respiratory morbidity and short-term exposure to SO₂*. This
22 conclusion is supported by the consistency, coherence, and plausibility of findings observed in
23 human clinical studies with 5-10 min exposures, epidemiological studies mostly using 24-h avg
24 exposures,⁷ and animal toxicological studies using exposures of minutes to hours.

25 The respiratory health effects of SO₂ are consistent with the mode of action of SO₂ as it is
26 currently understood. The immediate effect of SO₂ on the respiratory system is
27 bronchoconstriction. This response is mediated by chemosensitive receptors in the
28 tracheobronchial tree. These receptors trigger reflexes at the central nervous system level
29 resulting in bronchoconstriction, mucus secretion, mucosal vasodilation, cough, and apnea

followed by rapid shallow breathing. In some cases, local nervous system reflexes also may be involved. Asthmatics are more sensitive to the effects of SO₂ likely resulting from preexisting inflammation associated with this disease. This inflammation may lead to enhanced release of mediators, alterations in the autonomic nervous system and/or sensitization of the chemosensitive receptors. These biological processes are likely to underlie the respiratory symptoms; exacerbations of airways inflammation, reactivity, and responsiveness; and decreased lung function observed in response to SO₂ exposure.

The strongest evidence for this causal relationship comes from human clinical studies reporting respiratory symptoms and decreased lung functions following peak exposures of 5-10 min duration to SO₂ at concentrations which have sometimes been measured in ambient air for similarly short-time durations. These effects are particularly evident among exercising asthmatics, with some sensitive asthmatics (5-20%) experiencing moderate or greater decrements in lung function at SO₂ concentrations as low as 0.2-0.3 ppm (see Table 5-1). At concentrations

Table 5-1. Key health effects of short-term exposure to SO₂ observed in human clinical studies.

CONCENTRATION	EXPOSURE	EFFECTS	STUDIES
0.2-0.3 ppm	5-10 min	Moderate to large reductions in FEV ₁ and increases in specific airway resistance (sRaw) observed among some asthmatic adults (5-20%) during moderate to heavy exercise. Bronchial responsiveness to SO ₂ may be enhanced when preceded by exposure to O ₃ . Limited evidence of SO ₂ -induced increases in respiratory symptoms.	Bethel et al. (1985); Horstman et al. (1986); Koenig et al. (1990); Linn et al. (1983, 1987; 1988; 1990); Schachter et al. (1984); Sheppard et al. (1981); Trenga et al. (2001)
	1-6 h	Enhanced sensitivity to an inhaled allergen following exposure to SO ₂ with NO ₂ in resting asthmatics. No evidence of respiratory symptoms or decrements in lung function in resting asthmatics or healthy adults. Some weak and inconsistent evidence to suggest that SO ₂ exposure may lead to changes in heart rate variability.	Devalia et al. (1994); Routledge et al. (2006); Rusznak et al. (1996); Tunnicliffe et al. (2001, 2003)
0.4-0.5 ppm	1-10 min	Decrements in lung function clearly demonstrated in asthmatics during exercise with significant interindividual variability in response (approximately 30% of asthmatics experienced moderate or greater decrements in lung function). Effects observed within 1-5 min of exposure generally not enhanced by increasing exposure duration. Respiratory symptoms (e.g., wheezing, chest tightness) observed at concentrations as low as 0.4 ppm and have been shown to increase with increasing exposure concentrations.	Balmes et al. (1987); Gong et al. (1995); Horstman et al. (1986); Koenig et al. (1983); Linn et al. (1983, 1987); Magnussen et al. (1990); Schachter et al. (1984); Sheppard et al. (1981); Trenga et al. (1999)
	~1-h	Decrements in lung function among asthmatics following 10 min of exercise at the end of a 60-75 min exposure are statistically significant, but less severe than effects observed following a 10 min period of exercise at the start of the exposure.	Linn et al. (1987); Roger et al. (1985)
0.6-1.0 ppm	1-10 min	Clear and consistent SO ₂ -induced increases in respiratory symptoms observed among exercising asthmatics. Moderate to large decrements in lung function demonstrated in 35-60% of asthmatics. Respiratory effects attributed to SO ₂ among asthmatics during exercise may be diminished after cessation of exercise, even with continued SO ₂ exposure. No respiratory effects reported in healthy, non-asthmatics.	Balmes et al. (1987); Gong et al. (1995); Hackney et al. (1984); Horstman et al. (1986, 1988); Koenig et al. (1983); Linn et al. (1987; 1988; 1990); Roger et al. (1985); Schachter et al. (1984)
	1-6 h	Decrements in lung function among asthmatics following 5-10 min of exercise at the end of a 1-6 h exposure are statistically significant, but less severe than effects observed following a 5-10 min period of exercise at the start of the exposure.	Linn et al. (1984; 1987); Hackney et al. (1984); Roger et al. (1985)

1 ≥ 0.4 ppm, a greater percentage (20-60%) of asthmatics experience SO₂-induced decrements in
2 lung function, which are frequently accompanied by respiratory symptoms. A clear
3 concentration-response relationship has been demonstrated following exposures to SO₂ at
4 concentrations between 0.2 and 1.0 ppm, both in terms of severity of effect and percentage of
5 asthmatics adversely affected. Animal toxicological studies have also reported
6 bronchoconstriction with short-term exposures of 0.5 to 1 ppm SO₂ (see Table 5-2).

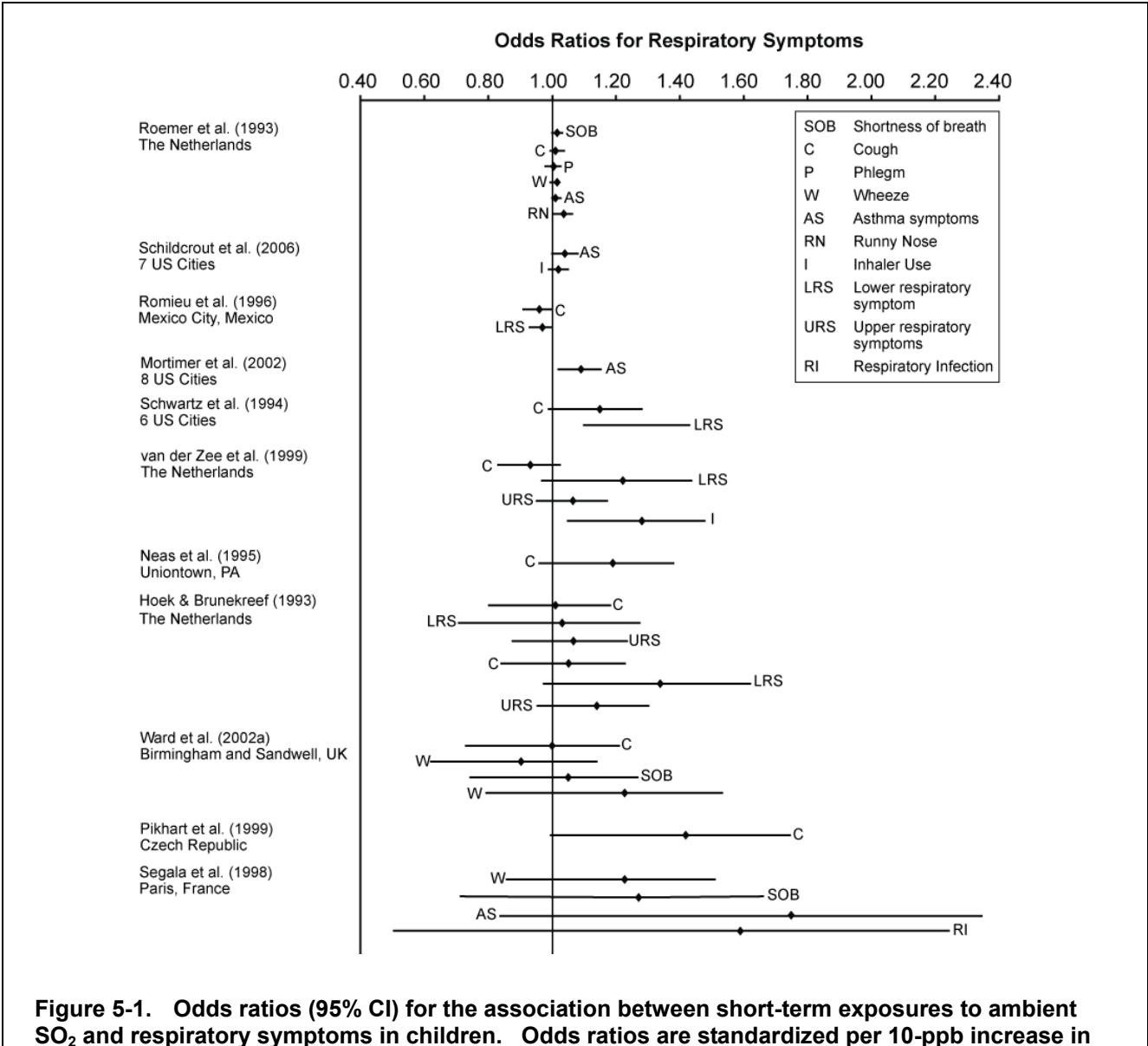
7 A larger body of evidence supporting this determination of causality comes from numerous
8 epidemiological studies reporting associations with respiratory symptoms, ED visits, and hospital
9 admissions with short-term SO₂ exposures, generally of 24-h avg. Almost all of these studies
10 were conducted in areas where the maximum ambient 24-h avg SO₂ concentration was consis-
11 tently below the current 24-h avg NAAQS level of 0.14 ppm. Important new multicity studies
12 and several other studies have found an association between 24-h avg ambient SO₂
13 concentrations and respiratory symptoms in children, particularly those with asthma.
14 Furthermore, limited epidemiological evidence indicates that atopic children and adults may be
15 at increased risk for SO₂-induced respiratory symptoms. Generally consistent associations also
16 were observed between ambient SO₂ concentrations and ED visits and hospitalizations for all
17 respiratory causes, particularly among children and older adults (≥ 65 years), and for asthma.
18 The SO₂-related changes in ED visits or hospital admissions for respiratory causes ranged from -
19 5% to 20% excess risk. Results of experiments in laboratory animals support these observations.
20 Studies in animals sensitized with antigen demonstrated that repeated exposure to SO₂ levels as
21 low as 0.1 ppm exacerbated allergic responses including mucin production, airway inflammation
22 and airway hyperresponsiveness. These responses are consistent with exacerbation of asthma in
23 humans.

24 The consistency and internal coherence of the epidemiological evidence for respiratory
25 effects associated with short-term exposure to SO₂ are illustrated in Figures 5-1 and 5-2, which
26 present effect estimates for respiratory symptoms, ED visits, and hospitalizations in children.
27 Associations between short-term ambient SO₂ concentrations and respiratory symptoms, ED
28 visits, and hospitalizations are largely positive, with several of the more precise effect estimates
29 (suggestive of greater study power) indicating statistical significance. The epidemiological
30 findings of asthma symptoms with 24-h avg SO₂ exposures are generally coherent with increases

Table 5-2. Key respiratory health effects of exposure to SO₂ in animal toxicological studies.

REFERENCE	EXPOSURE	SPECIES	EFFECTS
LUNG FUNCTION			
Amdur et al. (1983)	1 ppm SO ₂ for 1-h	Male Hartley guinea pigs	An 11% increase in pulmonary resistance and 12% decrease in dynamic compliance were observed. Neither effect persisted into the 1-h period following exposure. No effects were observed for breathing frequency, tidal volume, or min volume.
Conner et al. (1985)	1 ppm (2.62 mg/m ³); nose only; 3-h/day for 6 days; animals evaluated for up to 48-h following exposure	Hartley guinea pig, male, age not reported, 250-320 g, n = ≤18 group/time point	No effect was observed on residual volume, functional reserve capacity, vital capacity, total lung capacity, respiratory frequency, tidal volume, pulmonary resistance, pulmonary compliance, diffusing capacity for carbon monoxide or alveolar volume at 1- or 48-h after last exposure.
Barthélemy et al. (1988)	0.5 or 5 ppm (1.3 or 13.1 mg/m ³); intratracheal; 45 min	Rabbit, sex not reported, adult, mean 2.0 kg, n = 5-9/group; rabbits were mechanically ventilated	Lung resistance increased by 16% and 50% in response to 0.5 and 5 ppm SO ₂ , respectively. Bivagotomy had no effect on 5 ppm SO ₂ -induced increases in lung resistance (54% increase before vagotomy and 56% increase after vagotomy). Reflex bronchoconstrictive response to phenyldiguanide (intravenously administered) was eliminated by exposure to SO ₂ but SO ₂ had no effect on lung resistance induced by intravenously-administered histamine. The study authors concluded that (1) vagal reflex is not responsible for SO ₂ -induced increase in lung resistance at 45 min and (2) transient alteration in tracheobronchial wall following SO ₂ exposure may have reduced accessibility of airway nervous receptors to phenyldiguanide.
LUNG INJURY, INFLAMMATION AND MORPHOLOGY			
Conner et al. (1989)	1 ppm (2.62 mg/m ³); nose only 3-h/day for 5 days; bronchoalveolar lavage performed daily	Hartley guinea pig, male, age not reported, 250-320 g, n = 4	No change in numbers of total cells and neutrophils, protein levels or enzyme activity in lavage fluid following SO ₂ exposure.
Park et al. (2001)	0.1 ppm (0.26 mg/m ³); whole body; with and without exposure to ovalbumin, 5-h/day for 5 days	Dunkin-Hartley guinea pig, male, age not reported, 250-350 g, n = 7-12/group	After bronchial challenge, the ovalbumin/SO ₂ -exposed group had significantly increased eosinophil counts in BAL fluids compared with all other groups, including the SO ₂ -only group. The bronchial and lung tissue of the ovalbumin/SO ₂ -exposed group showed infiltration of inflammatory cells, bronchiolar epithelial damage, and mucus and cell plug in the lumen.
Li et al. (2007b)	2 ppm (5.24 mg/m ³), with and without exposure to ovalbumin, 1-h/day for 7 days	Wistar rats, male, age not reported	Increased number of inflammatory cells in BAL fluid, increased levels of MUC5AC and ICAM-1 and an enhanced histopathological response compared with those treated with ovalbumin or SO ₂ alone
Conner et al. (1985)	1 ppm, 3-h/day/6 day. Evaluated up to 72-h postexposure	Male Hartely guinea pigs	No alveolar lesions.
Smith et al. (1989)	1 ppm, 5-h/day, 5 day/wk up to 4 and 8 mos	Male Sprague-Dawley rats	Increased bronchial epithelial hyperplasia and number of nonciliated epithelial cells observed at 4 mos.
AIRWAY HYPERRESPONSIVENESS AND ALLERGY			
Riedel et al. (1988)	0.1, 4.3, or 16.6 ppm (0, 0.26, 11.3, or 43.5 mg/m ³); whole body; 8-h/day for 5 days; animals were sensitized to ovalbumin on the last 3 days of exposure	Perlbright-White Guinea pig, female, age not reported, 300-350 g, n = 5 or 6/group (14 controls)	Bronchial provocation with ovalbumin was conducted every other day for 2 wks, starting at 1 wk after the last exposure. Numbers of animals displaying symptoms of bronchial obstruction after ovalbumin provocation was increased in all SO ₂ groups compared to air-exposed groups. Anti-ovalbumin antibodies (IgG total and IgG1) were increased in BAL fluid and serum of SO ₂ -exposed compared to air-exposed controls, with statistical significance obtained for IgG total in BAL fluid at ≥4.3 ppm SO ₂ and in serum at all SO ₂ concentrations. Results indicate that in this model, subacute exposure to even low concentrations of SO ₂ can potentiate allergic sensitization of the airway.
Park et al. (2001)	0.1 ppm (0.26 mg/m ³); whole body; with and without exposure to ovalbumin; 5-h/day for 5 days	Dunkin-Hartley guinea pig, male, age not reported, 250-350 g, n = 7-12/group	After bronchial challenge, the ovalbumin/SO ₂ -exposed group had significantly increased enhanced pause (indicator of airway obstruction) compared with all other groups, including the SO ₂ group. Study authors concluded that low level SO ₂ may enhance the development of ovalbumin-induced asthmatic reactions in guinea pigs.

1 in symptoms reported in asthmatics in human clinical studies with 5-10 min exposures; it is
 2 possible that these epidemiological associations are determined in large part by peak exposures
 3 within a 24-h period. The effects of SO₂ on respiratory symptoms, lung function, and airway
 4 inflammation observed in the human clinical studies using peak exposures further provides a
 5 basis for a progression of respiratory morbidity resulting in increased ED visits and hospital
 6 admissions. Collectively, these findings provide biological plausibility for the observed
 7 associations between ambient SO₂ levels and ED visits and hospitalizations for all respiratory
 8 diseases and asthma, notably in children and older adults (≥ 65 years).



24-h avg SO₂ level. Studies are generally presented in the order of increasing width of the 95% CI.

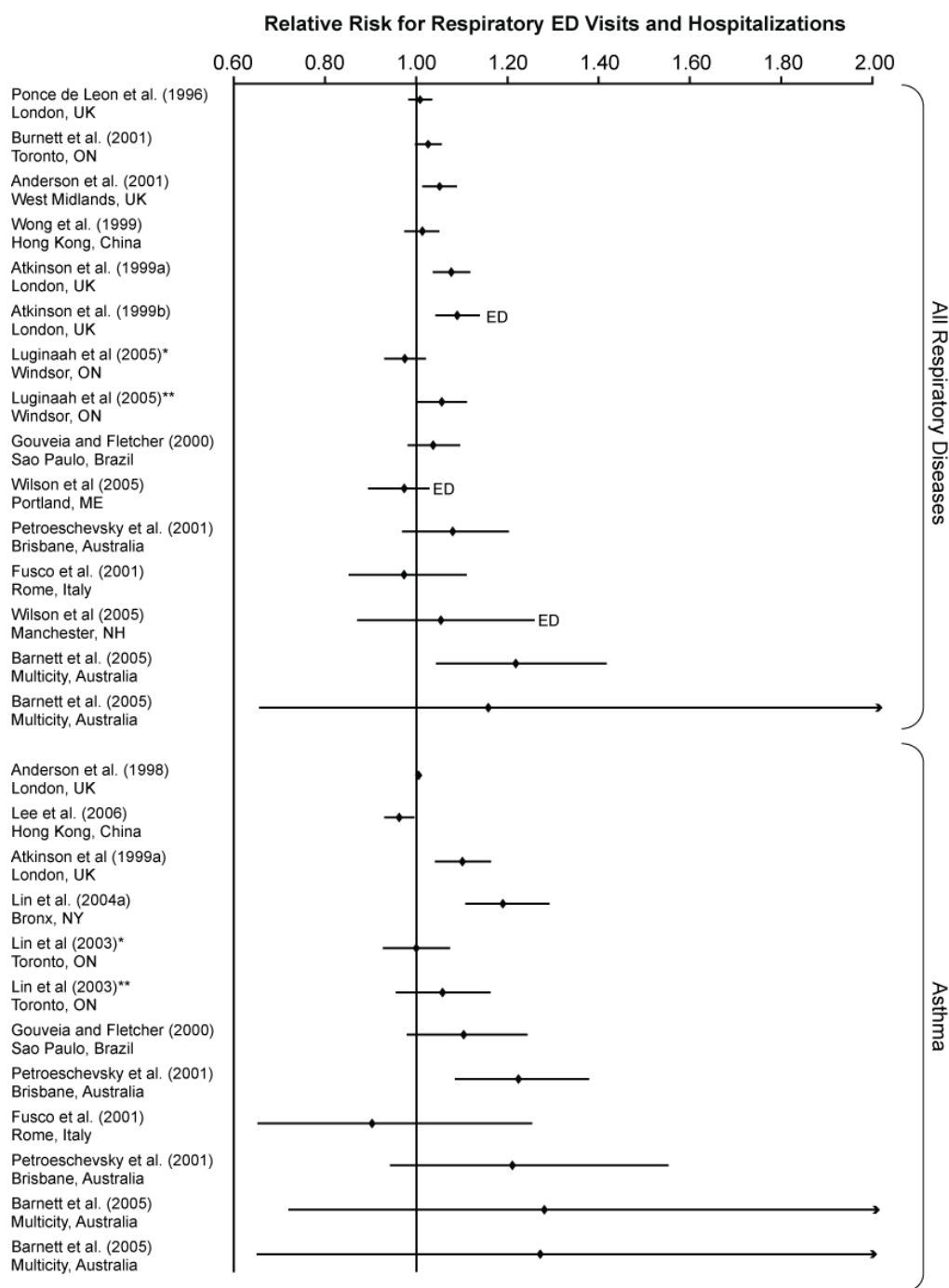


Figure 5-2. Relative risks (95% CI) for the association between short-term exposures to ambient SO₂ and emergency department (ED) visits/hospitalizations for all respiratory diseases and asthma in children. Relative risks are standardized per 10-ppb increase in 24-h avg SO₂ level. The

studies are generally presented in the order of increasing width of the 95% CI. For Luginaah et al. (2005) and Lin et al. (2003), risk estimates for males (*) and females () are shown separately.**

Overall, the epidemiological evidence for respiratory morbidity is consistent, with associations reported in studies conducted in numerous locations using a variety of methodological approaches. In the epidemiological studies that assessed potential confounding by copollutants using multipollutant models, SO₂ effect estimates were generally robust to the inclusion of copollutants, including PM, O₃, CO, and NO₂, suggesting that the observed effects of SO₂ on respiratory endpoints occur independent of the effects of other ambient air pollutants.

Intervention studies provide additional evidence that supports a causal relationship between SO₂ exposure and respiratory health effects. The proposition that intervention studies can provide strong support for causal inferences was emphasized by Hill (1965). Two notable studies conducted in several cities in Germany and in Hong Kong reported that decreases in SO₂ concentrations were associated with improvements in respiratory symptoms. In eastern Germany, a decrease in the prevalence of respiratory symptoms was correlated with a steep decline in ambient SO₂ concentrations of more than 90% from 1992-1993 to 1998-1999. During this study period, decreases in other ambient air pollutants, including ~60% lower TSP concentrations, also occurred in these cities. In Hong Kong, respiratory health improved with similarly large reductions in SO₂ of up to 80% in the polluted district but with much smaller reductions in TSP (less than 20%) compared with those in the cities in eastern Germany. The possibility remains that these health improvements may be partially attributable to declining concentrations of air pollutants other than SO₂, most notably PM or constituents of PM. Animal toxicological studies have reported that interactions of SO₂ and PM may lead to transformation of SO₂ to other sulfur-containing compounds which may have more potent biological effects; thus the improvements in respiratory health may be jointly attributable to declines in both SO₂ and PM.

The draft ISA also evaluates the evidence of other health outcomes and exposure durations. For short-term exposure to SO₂ and mortality, the evidence was found to be suggestive but not sufficient to infer a casual relationship. Recent epidemiological studies have consistently reported positive associations between mortality and SO₂, with slightly larger effect estimates observed for respiratory mortality compared to cardiovascular mortality. However, the SO₂ effect estimates were generally reduced after adjusting for copollutants in the regression models, indicating some extent of confounding among these pollutants. The evidence between short-term

SO₂ exposure and cardiovascular effects, and morbidity and mortality with long-term SO₂ exposures is inadequate to infer a causal relationship. The key conclusions on the health effects of SO₂ exposure are briefly summarized in Section 5.5.

5.3. Interpretation of the Epidemiological Evidence

This section highlights some key considerations for the evaluation of epidemiological evidence in this draft ISA. As discussed above, clinical studies provide the strongest evidence that short-term SO₂ exposure is associated with respiratory morbidity. Numerous epidemiological studies report associations for a broader range of respiratory health outcomes, at lower concentrations than the clinical studies, at levels below the current standard. There is, however, uncertainty about the magnitude of the epidemiological effects estimates. Several sources of uncertainty and the implications for risk assessment are discussed below.

Although the numerous epidemiological studies provide supportive evidence in making a causal determination for the effect of SO₂ on respiratory health, much uncertainty remains in the magnitude of the effect estimates related to ambient SO₂ exposures. Exposure measurement error is a key source of this uncertainty as there are questions about the extent to which concentrations measured by the regulatory ambient monitoring network typically used in epidemiological studies can accurately represent an individual's exposure to SO₂ of ambient origin. Factors contributing to exposure measurement error include the spatial variation in ambient SO₂, variation in time-activity patterns and the infiltration characteristics of microenvironments, as well as instrument error in the ambient and personal monitors.

SO₂ monitors currently deployed in the regulatory monitoring networks are adequate to determine compliance with current standards, since both the 24-h avg and annual standards are substantially above the operating limit of detection of these monitors. However, these monitors are inadequate for accurate and precise measurements at or near the current ambient mean 24-h avg SO₂ levels of ~4 ppb. Also, typical 24-h avg personal SO₂ exposures are often below the detection limit of commonly deployed passive SO₂ monitors. Therefore, the association between ambient concentrations and personal exposure may be inadequately characterized in recent studies at lower ambient concentrations.

For community time-series and short-term panel epidemiological studies using daily SO₂ concentrations from ambient monitors, these exposure and analytical measurement errors would

1 tend to bias the effect estimate towards the null, leading to uncertainty in accurately quantifying
2 the magnitude of the effect. In long-term exposure studies, the variable ambient measurement
3 and exposure error could also result in bias, but the extent and direction of this bias is unclear.

4 Another factor that contributes to uncertainty in estimating the SO₂-related effect from
5 epidemiological studies is that SO₂ is one component of a complex air pollution mixture
6 including various other components, e.g., PM and NO₂, known to affect respiratory health.

7 As a consequence of these uncertainties, the epidemiological observations of SO₂ health
8 effects can be interpreted in several ways which are not mutually exclusive. First, the reported
9 SO₂ effect estimates in epidemiological studies may reflect independent SO₂ effects on
10 respiratory health. This is supported by evidence from human clinical studies which indicate that
11 peak exposures (5-10 min) to SO₂ at levels as low as 0.2-0.3 ppm are capable of eliciting
12 respiratory responses in some sensitive asthmatics. Because pure SO₂ does not appear alone in
13 real-world ambient conditions but rather is part of a pollutant mixture, it is difficult to relate
14 these peak exposures in the human clinical studies unequivocally to the 24-h avg SO₂
15 concentrations typically assessed in epidemiological studies. It is possible that higher, shorter-
16 term concentrations of SO₂ may be driving the observed associations in epidemiological studies.
17 Among the limited number of epidemiological studies evaluating the concentration-response
18 function, several reported a linear relationship across the entire range of concentrations,
19 suggesting the lack of a population effects threshold. However, other studies found that a marked
20 increase in SO₂-related respiratory health effects was only observed at higher concentrations
21 (above 90th percentile values).

22 A second interpretation is that ambient SO₂ may be serving as an indicator of complex
23 ambient air pollution mixtures sharing the same source as SO₂ (i.e., combustion of sulfur-
24 containing fuels or metal smelting). Other components of mixed emissions from these sources
25 include trace elements such as vanadium, nickel, selenium, and arsenic. Distinguishing effects of
26 individual pollutants in multipollutant regression models is made difficult by the possibility that
27 a given air pollutant may be acting as a surrogate for a less-well-measured or unmeasured
28 pollutant, or that several pollutants may all be acting as surrogates for the same mixtures of
29 pollutants. Therefore, reported SO₂-related effects may represent those of the overall mixture or
30 other chemical components within the mixture. However, analysis of ambient data compiled
31 monthly for the years 2003 to 2005 showed that SO₂ concentrations were uncorrelated with SO₄²⁻

1 in 12 CMSAs having multiple monitors. Moreover, in multipollutant models adjusted for PM
2 indices, SO₂ effect estimates were generally found to be robust.

3 A third interpretation is that in situations of complex pollution mixtures, copollutants may
4 enhance the toxic capability of SO₂ or that SO₂ may influence the toxicity of copollutants.
5 Findings from animal toxicological studies demonstrate that the effects of SO₂ may be
6 exacerbated when aerosol particles act as carriers and deliver sulfur-containing compounds more
7 effectively to the lower respiratory tract. The synergism observed with combined exposure to
8 SO₂ and PM in the animal toxicological studies provides supportive evidence for the SO₂-related
9 respiratory effects observed under ambient conditions in the epidemiological studies.

5.4. Susceptible and Vulnerable Populations

10 Evidence from epidemiological and human clinical studies has indicated that certain
11 subgroups within the population are more susceptible and/or vulnerable to the effects of SO₂
12 exposure. There is substantial evidence from epidemiological and human clinical studies
13 indicating that asthmatics are more susceptible to respiratory health effects from SO₂ exposures
14 than the general public. Limited epidemiological evidence further suggests that children and
15 older adults (≥ 65 years) are more susceptible to the adverse respiratory effects associated with
16 ambient SO₂ concentrations when compared to the general population. A number of potentially
17 susceptible groups, including obese individuals, individuals in a chronic pro-inflammatory state
18 like diabetics, and children born prematurely or with low birth weight ($< 2,500$ grams), may
19 experience increased adverse effects associated with exposure to air pollution, but these
20 relationships have not been examined specifically in relation to SO₂. The differential effects of
21 air pollution among genetically diverse subpopulations have been examined for a number of GST
22 genes and other genotypes. While limited in number, these studies provide some insight into a
23 potential genetic role in the determination of susceptibility to air pollution.

24 Human clinical studies have clearly shown that exercising asthmatics are at greatest risk of
25 experiencing adverse respiratory effects related to SO₂ exposure. Oronasal breathing during
26 exercise increases vulnerability as it allows a larger fraction of inhaled SO₂ to reach the lower
27 airways. Therefore, individuals with increased vulnerability for SO₂-related respiratory health
28 effects include those who spend time outdoors at increased exertion levels, for example children,
29 outdoor workers, and individuals who exercise or play sports.

5.5. Conclusions

1 The important findings of this draft ISA on the health effects of SO₂ exposure, including
2 the levels at which effects are observed, are briefly summarized in Table 5-3. Also summarized
3 are conclusions drawn in the previous review for comparison.

4 Collectively, the epidemiological, human clinical, and animal toxicological data support
5 the finding of a causal relationship between short-term exposure to SO₂ and respiratory
6 morbidity. Observed associations between SO₂ exposure and an array of respiratory outcomes,
7 including respiratory symptoms, lung function, airway inflammation, airway
8 hyperresponsiveness, and ED visits and hospitalizations from the human clinical, animal
9 toxicological, and epidemiological studies, in combination, provide clear and convincing
10 evidence of consistency, specificity, temporal and biologic gradients, biological plausibility, and
11 coherence.

12 Human clinical studies provide strong evidence of respiratory morbidity among asthmatics
13 following peak exposures (5-10 min) to SO₂ concentrations ≥ 0.4 ppm, with some evidence of
14 respiratory effects at concentrations as low as 0.2 ppm in the most sensitive asthmatics. In the
15 epidemiological studies, the SO₂-related respiratory effects were consistently observed in areas
16 where the maximum ambient 24-h avg SO₂ concentration was below the current 24-h avg
17 NAAQS level of 0.14 ppm (Tables 5-4 and 5-5). Potentially susceptible and vulnerable
18 subgroups include asthmatics, children, older adults, and individuals who spend a lot of time
19 outdoors at increased exertion levels.

20 In addition to respiratory morbidity related to short-term exposure to SO₂, studies of other
21 health outcomes and exposure durations were also evaluated in this draft ISA. The evidence is
22 suggestive but not sufficient to infer a causal relationship between short-term exposure to SO₂
23 and mortality. The evidence linking short-term SO₂ exposure and cardiovascular effects, and
24 morbidity and mortality with long-term exposures to SO₂ is inadequate to infer a causal
25 relationship.

Table 5-3. Key findings on the health effects of SO₂ exposure

Short-Term Exposure: RESPIRATORY MORBIDITY

Sufficient to infer a causal relationship

RESPIRATORY SYMPTOMS

Previous Conclusion: Among exercising asthmatics, there is a clear, statistically significant increase in respiratory symptoms following peak exposures (5-10 min) to 0.6-1.0 ppm SO₂. Significant, but less severe symptoms are associated with peak SO₂ exposures at concentrations of 0.4-0.5 ppm in human clinical studies.

In the epidemiological studies, an association with aggravation of bronchitis is consistently observed at 24-h avg SO₂ levels of 0.19 to 0.23 ppm and in some cases at levels below 0.19 ppm.

Current Conclusion: Recent human clinical studies provide additional evidence of respiratory symptoms in asthmatics following peak exposures (5-10 min) with exercise to 0.5 ppm SO₂. Statistically significant increases in respiratory symptoms are observed at SO₂ concentrations of as low as 0.4 ppm, with the severity of symptoms shown to increase with increasing concentration between 0.4 and 0.6 ppm.

Epidemiological studies provide consistent evidence of an association between ambient SO₂ exposure and increased respiratory symptoms in children, particularly those with asthma or chronic respiratory symptoms. Multicity studies have observed these associations at a median range of 17 to 37 ppb (75th percentile: ~25 to 50) across cities for 3-h avg SO₂ and 2.2 to 7.4 ppb (90th percentile: 4.4 to 14.2) for 24-h avg SO₂.

In contrast, the epidemiological evidence on the association between SO₂ and respiratory symptoms in adults are generally mixed, with some showing positive associations and others finding no relationship at current ambient levels.

LUNG FUNCTION

Previous Conclusion: Bronchoconstriction has been found to be the most sensitive indicator of lung function effects following acute exposure to SO₂. Guinea pigs were found to be the most sensitive species, with bronchoconstriction observed using 0.16 ppm SO₂. In human clinical studies, ≤ 10-20% of exercising asthmatic individuals experience large decrements in lung function (i.e., sRaw increases ≥ 200% or FEV₁ decreases ≥ 20%) following 5-10 min exposures to SO₂ concentrations of 0.2-0.5 ppm. At 0.6-1.0 ppm SO₂, ≥ 20-25% of exercising asthmatics are similarly affected.

Small, reversible declines in lung function in children are observed in epidemiological studies at levels of 0.10 to 0.18 ppm but not at levels of 0.04 to 0.08 ppm.

Current Conclusion: Evaluation of the human clinical evidence focused on moderate or greater decrements in lung function in exercising asthmatics. SO₂-induced increases in sRaw (≥ 100%) or decreases in FEV₁ (≥ 15%) following 5-10 min exposures are observed in 5-10% of individuals at 0.2 ppm, 10-20% of individuals at 0.3 ppm, and 20-60% of individuals at 0.4-1.0 ppm.

The results are inconsistent for the association between 24-h avg SO₂ and lung function in children and adults in the epidemiological studies.

AIRWAYS INFLAMMATION

Previous Conclusion: No overall conclusion.

Current Conclusion: A limited number of health studies have evaluated the effect of SO₂ on airway inflammation. One human clinical study observed an SO₂-induced increase in sputum eosinophil counts in exercising asthmatics 2 h after a 10 min exposure to 0.75 ppm SO₂. The results of this study provide some evidence that SO₂ may elicit an allergic inflammatory response in the airways of asthmatics which extends beyond the short time period typically associated with SO₂ effects.

Animal toxicological studies suggest that repeated exposures to SO₂, at concentrations as low as 0.1 ppm in guinea pigs, may exacerbate inflammatory responses in allergic animals.

AIRWAYS RESPONSIVENESS

Previous Conclusion: No conclusions in the previous review.

Current Conclusion: Animal toxicological evidence suggests that repeated exposures to SO₂, at concentrations as low as 0.1 ppm in guinea pigs, can exacerbate airway responsiveness following allergic sensitization. In a human clinical study, concurrent exposure (6 h) to 0.2 ppm SO₂ and 0.4 ppm NO₂ has been observed to enhance airway responsiveness to an inhaled antigen among resting asthmatics. These findings are consistent with the very limited epidemiological evidence that suggests that exposure to SO₂ may lead to airway hyperresponsiveness in atopic individuals.

ED VISITS/HOSPITALIZATIONS

Previous Conclusion: No conclusions in the previous review.

Current Conclusion: Epidemiological studies provide evidence of an association between ambient SO₂ concentrations and ED visits and hospitalizations for all respiratory causes, particularly among children and older adults (age 65+ years), and for asthma. The SO₂ effect estimates ranged from a 5% decreased risk to a 20% excess risk per 10-ppb increase in 24-h avg SO₂, with the large majority of studies suggesting an increase in risk. These effects were observed in studies with mean 24-h avg concentrations as low as 4 ppb, but two studies evaluating the concentration-response function observed that a marked increase in SO₂-related effects was only observed higher concentrations (above 90th percentile values).

Short-Term Exposure: CARDIOVASCULAR MORBIDITY

Inadequate to infer the presence or absence of a causal relationship

CARDIOVASCULAR EFFECTS; ED VISITS/ HOSPITALIZATIONS

Previous Conclusion: No conclusions in the previous review.

Current Conclusion: There was some suggestive evidence of an association between 24-h avg SO₂ exposure and heart rate variability in the epidemiological studies, but the evidence from two human clinical studies with were weak and inconsistent. Some epidemiological studies have observed positive associations between ambient SO₂ concentrations and ED visits or hospital admissions for cardiovascular diseases, but results are not consistent across studies and the SO₂ effect estimate was generally not robust to copollutant adjustment.

Short-Term Exposure: MORTALITY

Suggestive but not sufficient to infer a causal relationship

NONACCIDENTAL AND CARDIOPULMONARY MORTALITY

Previous Conclusion: Epidemiological studies based on historical air pollution episodes observed the clearest mortality associations when both black smoke (BS) and SO₂ concentrations were at high levels (24-h avg values exceeding 1,000 µg/m³ [~400 ppb for SO₂]). Later studies observed that an increased risk of mortality was associated with exposure to BS and SO₂ levels in the range 0.19 to 0.38 ppm. Because of the high colinearity between BS and SO₂ levels, it is difficult to readily separate the effects of these pollutants on mortality.

Current Conclusion: Recent epidemiological studies have consistently reported positive associations between mortality and SO₂, often at mean 24-h avg levels < 10 ppb. The range of SO₂ excess risk estimates for nonaccidental mortality is 0.4 to 2% per 10 ppb increase in 24-h avg SO₂ in several multicity studies and meta-analyses. SO₂ effect estimates for respiratory mortality were generally larger than the cardiovascular mortality estimates, suggesting a stronger association of SO₂ with respiratory mortality compared to cardiovascular mortality. The SO₂ effect estimates were generally reduced when copollutants were added in the model, indicating some extent of confounding among these pollutants.

Long-Term Exposure: RESPIRATORY MORBIDITY

Inadequate to infer the presence or absence of a causal relationship

RESPIRATORY SYMPTOMS AND LUNG FUNCTION

Previous Conclusion: The limited available epidemiological data indicated associations between respiratory illnesses and symptoms and persistent exposures to SO₂ in areas with long-term averages exceeding 0.04 ppm.

Current Conclusion: Several epidemiological studies that examined the effects of long-term exposure to SO₂ on asthma, bronchitis, and respiratory symptoms observed positive associations in children. While the evidence is suggestive, the variety of outcomes examined and the inconsistencies in the observed results make it difficult to assess the direct impact of long-term exposure of SO₂ on respiratory symptoms. The epidemiological and animal toxicological evidence generally do not indicate that long-term exposure to SO₂ has a detrimental effect on lung function.

Long-Term Exposure: OTHER MORBIDITY

Inadequate to infer the presence or absence of a causal relationship

CARCINOGENIC EFFECTS

Previous Conclusion: Epidemiological evidence did not substantiate the hypothesized links between SO₂ or other SO_x and cancer, though there was some animal toxicological evidence that led to the conclusion that SO₂ may be considered a suspect carcinogen/cocarcinogen.

Current Conclusion: Animal toxicological studies indicate that SO₂ at high concentrations may cause DNA damage but fails to induce carcinogenesis, cocarcinogenesis, or tumor promotion. Epidemiological studies did not provide evidence that long-term exposure to SO₂ is associated with an increased incidence of or mortality from lung cancer.

PRENATAL AND NEONATAL OUTCOMES

Previous Conclusion: No conclusions in the previous review.

Current Conclusion: Epidemiological studies on birth outcomes have found suggestive positive associations between SO₂ exposure and low birth weight. However, the inconsistent results across trimesters of pregnancy and the lack of evidence to evaluate confounding by copollutants limit the interpretation of these studies.

Long-Term Exposure: MORTALITY

Inadequate to infer the presence or absence of a causal relationship

NONACCIDENTAL AND CARDIOPULMONARY MORTALITY

Previous Conclusion: The available studies on the effects of long-term exposure to SO₂ on mortality were all ecological cross-sectional studies which did not take into consideration potential confounders. In addition, it was concluded that effects from PM and SO₂ could not be distinguished in these studies.

Current Conclusion: Two major U.S. epidemiological studies observed associations between long-term exposure to SO₂ and mortality, but several other U.S. and European cohort studies did not observe an association. The relative risks ranged from 0.97 to 1.07 per 5-ppb increase in the long-term average SO₂. Evaluation of these studies is further limited by the inability to distinguish potential confounding by copollutants and uncertainties regarding the geographic scale of analysis.

Table 5-4. Effects of short-term exposure to SO₂ on respiratory symptoms among children.

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
United States							
Schildcrout et al. (2006) Multicity, North America Seattle, WA; Baltimore, MD; St. Louis, MS (Nov 1993– Aug 1995); Denver, CO; San Diego, CA (Nov 1993– Jul 1995); Toronto, ON (Dec 1993–Jul 1995); Boston, MA (Jan 1994–Sep 1995) No SO ₂ data available in Albuquerque, NM	Asthmatic children (n = 990)	24-h avg: 2.2-7.4 (range of city-specific medians)	NR	NR	NR	75th: 3.1, 10.7 90th: 4.4, 14.2 (range in city specific estimates)	Asthma symptoms: SO ₂ alone: 1.04 (1.00, 1.08), 3-day sum SO ₂ & NO ₂ : 1.04 (1.00, 1.09), 3-day sum SO ₂ & PM ₁₀ : 1.04 (0.99, 1.08), 3-day sum
Schwartz et al. (1994) Multicity, United States Watertown, MA (Apr-Aug 1985); Kingston-Harriman, TN; St. Louis, MO (Apr-Aug 1986); Steubenville, OH; Portage, WI (Apr-Aug 1987); Topeka, KS (Apr- Aug 1988)	Children in grades 2-5 (n = 1,844)	24-h avg: 4.1 (median)	NR	NR	NR	75th: 8.2 90th: 17.9 Max: 81.9	Cough incidence: SO ₂ alone: 1.15 (1.02- 1.31), 4-day avg SO ₂ , adjusting for PM ₁₀ : 1.08 (0.93, 1.25), 4-day avg SO ₂ , adjusting for NO ₂ : 1.09 (0.94, 1.30), 4-day avg
Neas et al. (1995) Uniontown, PA Summer 1990	Children in grades 4-5 (n = 83)	12-h avg: 0.2 5.9 overnight 14.5 daytime	NR	NR	IQR: 11.1	Max: 44.9	Evening cough: 1.19 (1.00, 1.42), lag 12-h
Mortimer et al. (2002) Multicity, United States Bronx, NY; East Harlem, NY; Baltimore, MD; Washington, DC; Detroit, MI; Cleveland, OH; Chicago, IL; St. Louis, MO (Jun-Aug 1993)	Asthmatic children, aged 4-9 (n = 846)	3-h avg: 22 (shown in figure)	NR	NR	0-75 ppb (shown in graph)	NR	Asthma symptoms: SO ₂ alone (8 cities): 1.19 (1.06, 1.35), lag 1-2 SO ₂ , adjusting for O ₃ & NO ₂ (7 cities): 1.19 (1.04, 1.37), lag 1-2 SO ₂ , adjusting for O ₃ , NO ₂ & PM ₁₀ (3 cities): 1.23 (0.94, 1.62), lag 1-2
Europe							
Timonen and Pekkanen (1997) Kuopio (urban and suburban) Finland Winter 1994	Children 7-12 yrs with asthma or cough symptoms (n = 169)	24-h avg: 2.3	NR	NR	NR	75th: 2.7 Max: 12.3	Upper respiratory symptoms: 2.71 (1.19, 6.17), lag 0 3.17 (1.21, 8.78), lag 1
Ward et al. (2002) Birmingham and Sandwell, UK Jan-Mar 1997 May-Jul 1997	Children, age at enrollment 9 yrs (n = 162)	24-h avg: Median 5.4, Winter 4.7, Summer	NR	NR	2, 18 Winter 2, 10 Summer	NR	Cough: Winter: 0.59 (0.25, 1.40), Summer: 0.90 (0.49, 1.66) Shortness of breath: Winter: 0.59 (0.32, 1.09), Summer: 0.81 (0.30, 2.17) Wheeze: Winter: 0.79 (0.38, 1.63), Summer: 0.77 (0.28, 2.08) (7-day avg lag for above)

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
Segala et al. (1998) Paris, France Nov 1992-May 1993	Children 7-15 yrs with physician- diagnosed asthma (n = 84)	24-h avg: 8.3 (5.2)	NR	NR	1.7-32.2	NR	Prevalent asthma: 1.32 (1.08, 1.62), lag 0 1.26 (0.93, 1.71), lag 1 Prevalent shortness of breath: 1.17 (0.53, 2.62), lag 0 1.21 (0.99, 1.49) lag 1 Incident asthma 1.73 (1.15, 2.60), lag 0 1.60 (1.01, 2.53), lag 1 Incident wheeze 1.22 (0.95, 1.58), lag 0 1.13 (0.68, 1.88), lag 1
Boezen et al. (1998) Amsterdam and Meppel (urban and rural), the Netherlands Winter 1993-1994	Children 7-11 yrs, with and w/o BHR and high serum concentrations of total IgE (n = 632)	24-h avg: Means: 1.7, 8.7; Medians: 1.4, 8.3 (range in city- specific estimates)	NR	NR	1.9, 23.6	NR	Among children with BHR and relatively high serum total IgE - lower respiratory symptoms: 1.27 (1.09, 1.49), lag 0 1.25 (1.06, 1.48), lag 1 1.69 (1.26, 2.28), 5-day avg
Roemer et al. (1993) Wageningen, the Netherlands Winter 1990-1991	Children 6-12 yrs with chronic respiratory conditions (n = 73)	24-h avg 1-h max	NR	NR	0, 40.4 (24-h avg)	Max: 56.5 (1-h max)	Asthma attack: 1.79 (1.35, 2.38). 7-day avg Wheeze: 1.97 (1.42, 2.72), 7-day avg Waken with symptoms: 1.79 (1.12, 2.87), 7-day avg Shortness of breath: 1.48 (1.06, 2.07), 7-day avg Cough: 1.97 (1.03, 3.77), 7-day avg
Hoek and Brunekreff (1993) Wageningen, the Netherlands Winter 1991	Children 7-11 yrs, nonurban area (n = 112)	24-h avg	NR	NR	NR	Max: 40.4	Cough: 1.22 (0.20, 7.39), lag 0 0.25 (0.04, 1.65), lag 1 3.67 (0.002, 7.331.974), 7-day avg Lower resp symptoms: 1.82 (0.14, 24.3), lag 0 0.33 (0.02, 6.05), lag 1 0.005 (0.0, 44.7), 5-day avg

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
Van der Zee et al. (1999) Urban and nonurban areas The Netherlands 3 winters, 1992-1995	Children 7-11 yrs, with and without chronic respiratory symptoms (n = 633)	24-h avg: 1.4, 8.8 (range in city-specific medians)	NR	NR	NR	Max: 6.5, 58.5 (range in city- specific maxi- mums)	Lower respiratory symp- toms, urban, SO ₂ alone: 1.22 (1.01, 1.46), lag 0 1.14 (0.95, 1.38), lag 1 1.34 (0.98, 1.82), 5-day avg Lower respiratory symptoms, urban, SO ₂ , adjusting for PM ₁₀ : 1.18 (0.96, 1.45), lag 0 1.03 (0.83, 1.27), lag 1 1.08 (0.72, 1.63), 5-day avg Lower respiratory symptoms, nonurban: 0.94 (0.79, 1.12), lag 0 0.94 (0.78, 1.13), lag 1 1.10 (0.75, 1.63), 5-day avg Cough, urban: 0.93 (0.84, 1.03), lag 0 1.08 (0.98, 1.19), lag 1 1.08 (0.89, 1.30) 5-day avg Cough, nonurban: 1.05 (0.96, 1.15), lag 0 0.98 (0.90, 1.08), lag 1 1.04 (0.83, 1.30), 5-day avg

^a24-h avg SO₂ and 12-h avg SO₂ standardized to 10- ppb incremental change; 3-h avg SO₂ standardized to 20-ppb incremental change; and 1-h max SO₂ standardized to 40-ppb incremental change.
NR = Not Reported
BHR = Bronchial Hyperresponsiveness

Table 5-5. Effects of short-term SO₂ exposure on emergency department visits and hospital admissions for respiratory outcomes.

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
EMERGENCY DEPARTMENT VISITS – ALL RESPIRATORY							
UNITED STATES							
Wilson et al. (2005) Portland, ME Jan 1998-Dec 2000 Manchester, NH Jan 1996-Dec 2000	≈ 84,000 ED visits	1-h max: Portland: 11.1 (9.1) Manchester: 16.5 (14.7)	NR	NR	NR	NR	Portland: All ages: 8% (3, 11) 0-14: -2.6% (-10.3, 2.7) 15-64: 11% (5.4, 13.9) 65+: 16.8% (8.2, 25.8) Manchester: All ages: 0% (-3, 5) 0-14: 0% (8, 8) 15-64: 0% (-3, 5) 65+: 8% (6, 23)
Tolbert et al. (2007) Atlanta, GA Jan 1993-Dec 2004	> 1,000,000 ED visits for all respiratory causes	1-h max: 14.9	NR	NR	1.0 - 149.0	75th: 20.0 90th: 35.0	0.75% (-0.75, 2.3)

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
Peel et al. (2005) Atlanta, GA Jan 1993-Aug 2000	484,830 ED visits, all ages from 31 hospitals	1-h max: 16.5 (17.1)	NR	NR	NR	90th: 39.0	1.6% (-0.6, 3.8)
EUROPE							
Atkinson et al. (1999a) London, UK Jan 1992-Dec 1994	98,685 ED visits from 12 hospitals	24-h avg: 8.0 (2.9)	NR	NR	2.8, 30.9	50th: 7.4 90th: 11.7	All Ages: 4.2% (1.1, 7.4) 0-14: 9.0% (4.4, 13.8) 15-64: 4.0% (-0.3, 8.5) 65+: -2.7% (-5.4, 3.3)
EMERGENCY DEPARTMENT VISITS – ASTHMA							
UNITED STATES							
Ito et al. (2007) New York, NY Jan 1999-Dec 2002	Asthma ED visits, all ages from 11 hospitals	24-h avg: 7.8 (4.6)	NR	NR	NR	75th: 10 95th: 17	35% (23%, 51%)
NY Department of Health (2006) Bronx & Manhattan, NY Jan 1999-Dec 2000	Asthma ED visits among children from 22 hospitals	24-h avg : 11 (7.2)	NR	NR	NR	NR	5-day moving average: Manhattan: -1% (-12, 12) Bronx: 11% (6, 17)
Jaffe et al. (2003) Cincinnati, OH Cleveland, OH Columbus, OH Jul 1991-Jun 1996	4,416 ED visits for asthma, age 5-34	24-h avg: Cincinnati: 13.5 (9.4) Cleveland: 14.7 (9.5) Columbus: 4.2 (3.2)	NR	NR	Cincinnati: 0.6, 49.6 Cleveland: 0.98, 62.8 Columbus 0, 21.4	NR	Cincinnati: 17.3% (4.7, 30.8) Cleveland: 3.1% (- 3.8, 10.7) Columbus: 13.1% (- 14.2, 48.6) All Cities: 6.2% (0.5, 11.6)
Wilson et al. (2005) Portland, ME Jan 1998-Dec 2000 Manchester, NH Jan 1996-Dec 2000	≈ 84,000 ED visits	1-h max: Portland: 11.1 (9.1) Manchester: 16.5 (14.7)	NR	NR	NR	NR	Portland: All ages: 11.0% (0.0, 19.7) 0-14: 5.4% (-12.8, 25.8) 15-64: 11% (0, 22.7) 65+: 11.0% (-15.2, 48.4) Manchester: All ages: 5.4% (-2.6, 16.8) 0-14: 19.7% (-2.6, 51.8) 15-64: 2.7% (-7.8, 13.9) 65+: 11.0% (-28.8, 77.2)
Peel et al. (2005) Atlanta, GA Jan 1993-Aug 2000	Asthma ED visits, all ages and 2-18 yrs from 31 hospitals	1-h max: 16.5 (17.1)	NR	NR	NR	90th: 39.0	0.2% (-3.2, 3.4)

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
EUROPE							
Atkinson et al. (1999a) London, UK Jan 1992-Dec 1994	98,685 ED visits from 12 hospitals	24-h avg: 8.0 (2.9)	NR	NR	2.8, 30.9	50th: 7.4 90th: 11.6	All ages: 7.4% (2.3, 12.8) 0-14: 15.0% (7.1, 23.5) 15-64: 6.3% (-0.8, 13.8)
Hajat et al. (1999) London, UK Jan 1992-Dec 1994	General practitioner visits for asthma	24-h avg: All yr: 8.0 (2.9) Warm: 7.7 (2.4) Cool: 8.3 (3.4)	NR	NR	NR	All yr: 90th: 11.6 Warm: 90th: 10.7 Cool: 90th: 12.4	All ages: 6.6% (1.3, 11.9) 0-14: 6.6% (-1.0, 14.7) 15-64: 5.2% (-1.5, 12.3) 65+: 7.2% (-4.3, 20.1)
Boutin-Forzano et al. (2004) Marseille, France Apr 1997-Mar 1998	549 ED visits for asthma	24-h avg: 8.5	NR	NR	0.0, 35.3	NR	3-49 yrs: 0.6% (-1.4, 2.7)
Galan et al. (2003) Madrid, Spain Jan 1995-Dec 1998	4,827 ED visits for asthma	24-h avg: 8.9 (5.8)	NR	NR	1.9, 45.6	50th: 7.0 75th: 11.8 90th: 16.5	All ages: 4.9% (-4.2, 15.0)
Tenias et al. (1998) Valencia, Spain Jan 1993-Dec 1995	734 ED visits for asthma	24-h avg: 10.0 Cold: 11.9 Warm: 8.2 1-h max: 21.2 Cold: 24.3 Warm: 18.1	NR	NR	NR	24-h avg: 50th: 9.8 75th: 12.9 95th: 16.0 1-h max: 50th: 19.6 75th: 27.1 95th: 35.8	> 14 yrs: 13.9% (-7.0, 39.4)
Sunyer et al. (1997) Multicity, Europe Barcelona, Spain; Helsinki, Finland; Paris, France; London, UK Jan 1986-Dec 1992	All ED visits for asthma	24-h avg: Barcelona: 15.4 Helsinki: 6.0 London: 11.6 Paris: 8.6	NR	NR	Barcelona: 0.8, 60.2 Helsinki: 1.1, 35.7 London: 3.4, 37.6 Paris: 0.4, 82.3	NR	0-14 yrs: 3.2% (-0.2, 6.8) 15-64: 0.2% (-2.2, 2.6)
Castellsague et al. (1995) Barcelona, Spain Jan 1986-Dec 1989	ED visits for asthma from 4 hospitals	24-h avg: Summer: 15.3 Winter: 19.5	NR	NR	NR	Summer: 50th: 13.5 75th: 20.3 95th: 30.8 Winter: 50th: 18.4 75th: 25.2 95th: 35.3	15-64 yrs, summer: 5.5% (-2.1, 13.8) 15-64 yrs, winter: 2.1% (-4.2, 9.0)
HOSPITAL ADMISSIONS – ALL RESPIRATORY							
UNITED STATES							

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
Schwartz (1995) New Haven, CT Tacoma, WA Jan 1988-Dec 1990	≈ 13,470 Hospital admissions, ages 65+	24-h avg: New Haven: 29.8 Tacoma: 16.8	NR	NR	NR	New Haven: 75th: 37.6 90th: 59.8 Tacoma: 75th: 21.1 90th: 27.8	New Haven: 1.6% (1.1, 2.6) Tacoma: 3.2% (0.5, 6.2)
CANADA							
Fung et al. (2006b) Vancouver, BC Jun 1995-Mar 1999	≈ 41,000 respiratory admissions for elderly (65+ yrs)	24-h avg: 3.46 (1.82)	NR	NR	0.0, 12.5	NR	12.6% (4.1, 22.0)
Cakmak et al. (2006) Multicity, Canada Calgary, Edmonton, Halifax, London, Ottawa, Saint John, Toronto, Vancouver, Windsor, Winnipeg Jan 1993-Dec 2000	> 200,000 hospital admissions for all respiratory causes	24-h avg: 4.6	NR	NR	2.8, 10.2	NR	2.4% (1.1, 3.9)
Yang et al. (2003b) Vancouver, BC Jan 1986-Dec 1998	Respiratory hospital admissions among young children (< 3 yrs) and elderly (≥ 65 yrs)	24-h avg: 4.84 (2.84)	NR	NR	NR	75th: 6.25 100th: 24.00	< 3 yrs: 3% (-6, 15) 65+ yrs: 5.8% (0.0, 11.9)
*Burnett et al. (2001) Toronto, ON Jan 1980-Dec 1994	All respiratory admissions for young children (< 2 yrs)	1-h max: 11.8	NR	55	NR	75th: 15 95th: 32 100th: 110	11% (-0.3, 23.6)
Luginaah et al. (2005) Windsor, ON Apr 1995-Dec 2000	All respiratory admissions ages 0-14, 15-64, and 65+ from 4 hospitals	1-h max: 27.5 (16.5)	NR	NR	0, 129	NR	All ages, female: 2.1% (-0.7, 5.0) All ages, male: - 2.5% (-5.3, 0.5) 0-14, female: 5.6% (0.6, 10.9) 0-14, male: -2.5% (-6.8, 1.9) 15-64, female: 1.6% (-3.7, 7.2) 15-64, male: - 4.5% (-8.4, 5.8) 65+, female: 1.5% (-2.6, 5.8) 65+, male: -3.1% (-7.5, 1.5)

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
AUSTRALIA							
Barnett et al. (2005) Multicity, Australia/New Zealand Auckland, Brisbane, Canberra, Christchurch, Melbourne, Perth, Sydney Jan 1998-Dec 2001	All respiratory hospital admissions	24-h avg: Auckland: 4.3 Brisbane: 1.8 Christchurch: 2.8 Sydney: 0.9 1-h max: Brisbane: 7.6 Christchurch: 10.1 Sydney: 3.7 NA in Auckland, Canberra, Melbourne, and Perth	NR	NR	24-h avg: Auckland: 0, 24.3 Brisbane: 0, 8.2 Christchurch: 0, 11.9 Sydney: 0, 3.9 1-h max Brisbane: 0, 46.5 Christchurch: 0.1, 42.1 Sydney: 0.1, 20.2		1-4 yrs: 5.1% (0.0, 9.1) 5-14: 3.7% (-9.9, 19.5)
Petroeschevsky et al. (2001) Brisbane, Australia Jan 1987-Dec 1994	33,710 hospital admissions	24-h avg: 4.1 1-h max: 9.2	NR	NR	NR	NR	All ages: -5.9% (-12.4, 1.1) 0-14: 8.0% (-2.9, 20.1) 15-64: -21.6% (-34.4, -6.2)
EUROPE							
Ofstedal et al. (2003) Drammen, Norway Jan 1994-Dec 2000	All respiratory hospital admissions	24-h avg: 1.1 (0.8)	NR	NR	NR	NR	All ages: 71.8% (15.5, 152.7)
Fusco et al. (2001) Rome, Italy Jan 1995-Oct 1997	All respiratory hospital admissions	24-h avg: 3.4 (2.2)	NR	NR	NR	50th: 3.0 75th: 4.5	All age: 1.6% (-4.9, 8.8) 0-14: -2.7% (-4.6, 10.8)
Llorca et al. (2005) Torrelavega, Spain Jan 1992-Dec 1995	Hospital admissions from one hospital	24-h avg: 5.0 (6.3)	NR	NR	NR	NR	All ages: 1.0% (-2.8, 4.7)
Anderson et al. (2001) West Midlands conurbation, UK Oct 1994-Dec 1996	Hospital admissions stratified by age	24-h avg: 7.2 (4.7)	NR	NR	1.9, 59.8	90th: 12.3	All ages: 1.4% (-0.8, 3.8) 0-14: 5.1% (1.6, 8.7) 15-64: -1.0% (-5.3, 3.7) 65+: -2.2% (-5.4, 1.2)
Atkinson et al. (1999a) London, UK Jan 1992-Dec 1994	165,032 hospital admissions	24-h avg: 8.0 (2.9)	NR	NR	2.8, 30.9	50th: 7.4 90th: 11.7	All ages: 3.0% (0.4, 5.6) 0-14: 7.7% (3.8, 11.7) 15-64: 2.8% (-1.2, 7.0) 65+: 3.3% (-0.1, 6.9)
Schouten et al. (1996) Multicity, The Netherlands Amsterdam, Rotterdam Apr 1977-Sep 1989	All respiratory hospital admissions	24-h avg: Amsterdam: 10.5 Rotterdam: 15.0 1-h max: Amsterdam: 24.4 Rotterdam: 37.2	NR	NR	NR	NR	Amsterdam: 15-64: -2.3% (-5.5, 0.9) 65+: 0.2% (-2.8, 3.3) Rotterdam: 15-64: -2.9% (-6.2, 0.5)

STUDY	POPULATION	MEAN CONCENTRATION	SO ₂ (ppb) 98TH%	SO ₂ (ppb) 99TH%	SO ₂ (ppb) RANGE	SO ₂ (ppb) UPPER %TILE	STANDARDIZED ODDS RATIO (95% CI) ^a
Spix et al. (1998) Multicity, Europe London, UK; Amsterdam & Rotterdam, the Netherlands; Paris, France; Milan, Italy Jan 1977-Dec 1991	All respiratory hospital admissions	24-h avg: London: 10.9 Amsterdam: 7.9 Rotterdam: 9.4 Paris: 8.6 Milan: 24.8	NR	NR	NR	NR	15-64 yrs: 0.5% (-0.4, 1.3) 65+: 1.1% (0.3, 2.4)
Dab et al. (1996) Paris, France Jan 1987-Sep 1992	Hospital admissions from 27 hospitals	All yr: 24-h avg: 11.2 1-h max: 22.5 Warm season 24-h avg: 7.6 1-h max: 16.1 Cold season 24-h avg: 15.1 1-h max: 29.4	NR	NR	NR	All yr: 24-h avg: 50.0 99th: 50.0 1-h max: 99th: 87.5 Warm season 99th: 18.5 1-h max: 99th: 50.3 Cold season 24-h avg: 99th: 56.0 1-h max: 99th: 100.9	All ages: 1.1% (0.1, 2.1)

^a24-h avg SO₂ and 12-h avg SO₂ standardized to 10-ppb incremental change; 3-h avg SO₂ standardized to 20-ppb incremental change; and 1-h max SO₂ standardized to 40-ppb incremental change.

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